



# Università degli Studi di Trieste

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Dipartimento di Scienze Economiche, Aziendali, Matematiche e Statistiche  
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TESI DI LAUREA

## Application of GLM Advancements to Non Life Insurance Pricing

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*The data scientist is a person who is better at statistics than any  
software engineer  
and better at software engineering than any statistician.*

Josh Wills

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# Abstract

This is my abstract ...

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# Introduction

La mia introduzione ...

## **Thesis aim**

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## **Actuary and datascientist figure**

Lorem ipsum ...

## **Thesis structure**

Lorem ipsum ...

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# 1

## Non-Life Insurance Pricing

In this chapter I am going to provide an overview on how non-life insurance works from an actuarial point of view with a specific focus on the pricing process.

### 1.1 What a Non-Life Insurance is

The Italian Civil Code provides the following definition of insurance contract:

**Definition 1.1** (Insurance Contract, Art. 1882, Italian Civil Code). The insurance is the contract with which an insurer, in exchange of the payment of a certain premium, obliged himself, within the agreed limits:

1. to pay an indemnity to the insured equivalent to the damage caused by an accident;
2. or to pay an income or a capital if a life-related event occurs.

This definition identifies two parties: the *Insurer* and the *Policyholder*. The policyholder pays to the Insurer a certain *Premium* at the beginning of the insurance coverage and the insurer will pay a benefit if a certain event (*Claim*) occurs. This event could happen zero, one or more than one times, so it is possible to have more than one claim.

Usually, in non-life insurance, the benefit is the payment of a sum. This sum could be predetermined (e.g. in motor theft insurance, where the benefit is usually the value of the insured vehicle) or defined by the entity of the claim (e.g. in motor third party liability insurance, it depends on the damage the policyholder has provided to a third party). Regarding the “agreed limits”, another peculiarity of non-life insurances is that the coverage period is defined as a fixed amount of time, usually corresponding to 1 year.

Starting from this legal definition, we can formalize a non-life insurance contract as follows.

Let’s:

- $]t_1, t_2]$ , with  $t_1 < t_2$ , be the coverage period;
- $P > 0$  be the premium payed by the policyholder to the insurer;
- $N \in \mathbb{N}$  be the number of claims occurred during the coverage period (*claims count*);
- $\tau_1, \tau_2, \dots, \tau_N$ , with  $t_1 < \tau_1 < \tau_2 < \dots < \tau_N < t_2$ , be the timing of each claim;
- $Z_1, Z_2, \dots, Z_N > 0$  be the amount of each claim (*claims severities* or *claims sizes*).

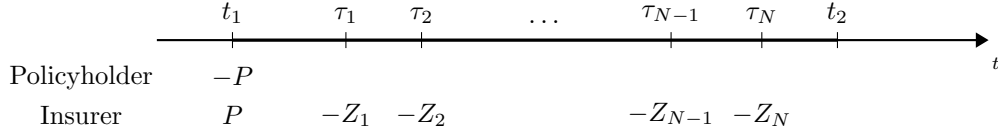
The total cost of claims for the insurance is:

$$S = \begin{cases} 0 & \text{if } N = 0 \\ \sum_{i=1}^N Z_i & \text{if } N > 0 \end{cases}$$

For semplicity, in the following we are going to just use the notation  $S = \sum_{i=1}^N Z_i$  with the meaning of 0 if  $N = 0$ .

Figure 1.1 shows the cash flows corresponding to the insurance contract. From this representation we can interpret the entering into an insurance contract by the policyholder as a way to exchange the negative cash flows  $-Z_1, -Z_2, \dots, -Z_N$  with one single negative cash flow  $-P$ . On the other hand, the insurer undertakes the negative cash flows  $-Z_1, -Z_2, \dots, -Z_N$  in exchange for a positive cash flow  $+P$ .

The major difference between these cash flows is that  $P$  is a certain amount, while  $Z_1, Z_2, \dots, Z_N$ , at the time  $t_1$ , are uncertain in the amount, in the count ( $N$ )



**Figure 1.1:** Insurance Contract cash flows.

and in the timing  $(\tau_1, \tau_2, \dots, \tau_N)$ . So, the policyholder, paying a premium  $P$ , is giving his risk to the insurer.

This representation points out the inversion of the production cycle typical of the insurance activity. From the insurer point of view, the revenue emerges at the beginning of the economic activity, in  $t_1$ , while the costs will emerge later. In most of other economic activities, the costs emerge before the selling of the product, so the agent can choose the selling price taking into account how much that product costed him. In insurance activity, the insurer, when is selling his product (the insurance coverage), doesn't know the amount of costs he is going to pay for that product. Thus, it is crucial to properly estimate the future costs and determine an adequate premium.

From a statistical point of view, we can translate this uncertainty saying that  $N$  and  $Z_1, Z_2, \dots, Z_N$  are random variables. Therefore, we can say that  $\{N, Z_1, Z_2, \dots\}$  is a stochastic process. Usually, in non-file insurance pricing, the variables  $\tau_1, \tau_2, \dots, \tau_N$  are not taken into account because the coverage span is short and from a financial point of view the timing of the claims occurrences is negligible.

Previously we said that  $Z_1, Z_2, \dots, Z_N$  are all  $> 0$ . This assumption corresponds to the fact that we are excluding the null claims, i.e. the claims that have been opened, but result in no payment due by the insurer. For the values of  $Z_i$  with  $N < i$  we can use the rule that  $\{N < i\} \Rightarrow \{Z_i = 0\}$ , so  $Z_{N+1} = 0, Z_{N+2} = 0, \dots$ . Therefore, we can say that:

$$\{N < i\} \iff \{Z_i = 0\}$$

## 1.2 Non-Life insurance pricing

In insurances, the premium that the insurer offers to the policyholder in exchange for the insurance coverage is not the same for every policyholder. The insurer evaluates the risk related to that policy and determine a “proper” premium taking into account risk related factors and commercial related factors. The process of *pricing* corresponds in defining the set of rules for determining this “proper” premium  $P_i$  for a specific policyholder  $i$ , given the known information on him. In the next sections I am going to better explain what “proper” means.

### 1.2.1 Compound distribution hypotheses

The first step for evaluating the stochastic process  $\{N, Z_1, Z_2, \dots\}$  is to introduce some probabilistic hypotheses. The usual hypotheses assumed are the following:

**Definition 1.2** (Compound distribution). Let’s assume that:

1. for each  $n > 0$ , the variables  $Z_1|N = n, Z_2|N = n, \dots, Z_n|N = n$  are stochastically independent and identically distributed;
2. the probability distribution of  $Z_i|N = n, i \leq n$  does not depend on  $n$ .

Under these hypotheses we say that:

$$S = \sum_{i=1}^N Z_i$$

has a compound distribution.

The variable  $Z_i|N = n$  used in this definition can be interpreted as the *claim severity for the  $i^{th}$  claim under the hypothesis that  $n$  claims occurred*. The two hypotheses provided in definition 1.2 imply that the distribution of  $Z_i|N = n, i < n$  does not depend from  $i$  nor from  $n$ . For this reason, in the following, we are going to use the notation  $Z$  to represent a random variable with the  $Z_i|N = n, i < n$  distribution and  $F_Z(\cdot)$  for its cumulative distribution function (i.e.  $F_Z(z) = P(Z \leq z)$ ).

Let’s consider the variable  $Z_i|N > i$ . We can interpret it as the *claim severity for the  $i^{th}$  claim under the hypothesis that the  $i^{th}$  claim occurred*. From the hypotheses

provided in definition 1.2 we can obtain that also  $Z_i|N > i$  has the same distribution of  $Z_i|N = n$ ,  $i < n$ . This can be easily obtained as follows:

$$P(Z_i|N \geq i) = P\left(Z_i \left| \bigvee_{n=i}^{+\infty} (N = n) \right.\right) = \quad (1.1)$$

$$= \sum_{n=i}^{+\infty} \underbrace{P(Z_i \leq z|N = n)}_{=F_Z(z)} P(N = n|N \geq i) = \quad (1.2)$$

$$= \sum_{n=i}^{+\infty} F_Z(z) P(N = n|N \geq i) = \quad (1.3)$$

$$= F_Z(z) \underbrace{\sum_{n=i}^{+\infty} P(N = n|N \geq i)}_{=1} = \quad (1.4)$$

$$= F_Z(z)$$

Where:

- the step (1.1) and the step (1.2) are given by the fact that the event  $\{N \geq i\}$  can be decomposed as  $\{N \geq i\} = \{\bigvee_{n=i}^{+\infty} (N = n)\}$  and that the events  $\{N = n\}, n \in \{i, i+1, i+2, \dots\}$  are two-by-two disjoint, so they constitute a partition of  $\{N \geq i\}$ , that allows us to use the disintegrability property of the probability;
- the step (1.3) is due to the fact that the distribution of  $Z_i \leq z|N = n$  depends neither on  $i$  nor on  $n$ ;
- the equivalence  $\sum_{n=i}^{+\infty} P(N = n|N \geq i) = 1$  at step (1.4) is due to the fact that the events  $\{N = n\}, n \in \{i, i+1, i+2, \dots\}$  are a partition of  $\{N \geq i\}$ .

Therefore,  $Z$  can be considered as the *claim severity for a claim under the hypothesis that that claim occurred*.

### 1.2.2 Distribution of the total cost of claims

Under the hypotheses defined in definition 1.2, it is possible to obtain the full distribution of  $S$  given the distribution of  $N$  and  $Z$ . In this chapter we are going to



provide only the formula of the expected value  $E(S)$ , but, with the same approach one can obtain all the moments.

The expected value of the total cost of claims  $E(S)$  can be obtained from the expected value of the claims count  $E(N)$  and the expected value of the claim severity  $E(Z)$  as follows:

$$E(S) = \sum_{n=0}^{+\infty} P(N = n) E(S|N = n) = \quad (1.5)$$

$$= \sum_{n=0}^{+\infty} P(N = n) E\left(\sum_{i=1}^n Z_i \middle| N = n\right) = \quad (1.6)$$

$$= \sum_{n=0}^{+\infty} P(N = n) \sum_{i=1}^n \underbrace{E(Z_i|N = n)}_{=E(Z)} = \quad (1.7)$$

$$= \sum_{n=0}^{+\infty} P(N = n) n E(Z) = \quad (1.8)$$

$$= E(Z) \underbrace{\sum_{n=0}^{+\infty} n P(N = n)}_{=E(N)} = \quad (1.9)$$

$$= E(N)E(Z) \quad (1.10)$$

Where:

- the step (1.5) is given by the fact that the events  $\{N = 0\}, \{N = 1\}, \{N = 2\}, \dots$  constitute a partition of the certain event  $\Omega$ , that allows us to use the disintegrability property of the expected value;
- the step (1.6) is due to the definition of  $S$ ;
- the step (1.7) is due to the linearity of the expected value;
- the steps (1.8) and (1.9) are due to the fact that, as assumed by the compound distribution hypotheses,  $E(Z_i|N = n)$  does not depends on  $i$  and  $n$ ;
- the step (1.10) is due to the definition of the expected value  $E(N) = \sum_{n=0}^{+\infty} n P(N = n)$ .

This result tells us that, under the hypotheses of the compound distribution, it is possible to easily obtain  $E(S)$  from  $E(N)$  and  $E(Z)$ . That means that we can model

separately  $E(N)$  and  $E(Z)$  and, from them, obtain  $E(S)$ . That result is particularly useful in personalization (paragraph 1.3), because, for each individual  $i$ , given the information we have on him  $x_i = (x_{i1}, x_{i2}, \dots, x_{ip})$ , we can estimate his expected claim size  $E(N_i)$  and his expected claim severity  $E(Z_i)$  and obtain his expected total cost of claims as  $E(S_i) = E(N_i)E(Z_i)$ .

### 1.2.3 Risk premium and Technical Price

The expected cost of claims  $E(S)$  is important because it gives us a first interpretation of what “proper” premium means.

**Definition 1.3** (Risk Premium). Said  $S$  the total cost of claims of a policyholder, his *Risk Premium* is given by:

$$P^{(risk)} = E(S)$$

The *Risk Premium* is the premium that on average covers the total cost of claims. As mentioned above, as the coverage spans are usually short, we are not taking into account the timing of the claims and we don’t discount the fact that the claims occur later than the premium payment.

It is clear that this premium, that only covers the cost of claims, is not “proper” in the practice.

First of all, the insurer has to cover also the expenses related to the policy (commission on sales and expenses related to the claim settlement) and the general expenses of the company. Adding the expenses, we obtain the *Technical Price*.

**Definition 1.4** (Technical Price). Said  $S$  the total cost of claims of a policyholder and  $E$  the expenses related to his policy, his *Technical Price* is given by:

$$P^{(tech)} = E(S) + E = P^{(risk)} + E$$

Secondly, even if the policyholder would pay a premium that on average covers claims and expenses, undertaking that risk with nothing in return would not make

sense for the insurer. So, to the Technical Price, a loading must be added. This loading can be justified as a loading for the risk or as a profit loading.

The result of the Technical Price with these loadings can be further modified based on business logic, as I am going to discuss later.

### **1.3 Modeling and Personalization**

In this section We are going to better explain how pricing based on policyholder information works.

#### **1.3.1 Pricing variables**

Usually for every policyholder we have a certain amount of information on him that is considered relevant for his risk evaluation. This information must be reliable and observable at the moment of the underwriting of the policy.

In motor insurances, this information could be:

- Information on the insured vehicle: make, model, engine power, vehicle mass, age of the vehicle;
- General information of the policyholder: age, sex, address (region, city, postcode), ownership of a private box where he parks the car;
- Insurance specific information of the policyholder: number of claims caused in the previous years, how long he has been covered, bonus-malus class;
- Policy options: amount of the maximum coverage, presence and amount of a deductible, presence of other insurance guarantees, how many drivers will drive the vehicle;
- Customer information on the policyholder: how many years he has been a customer of the insurer, how many other policies he owns.
- Telematic data: how many kilometers per year the policyholder travelled in the previous years, how many sharp accelerations and decelerations per kilometer the policyholder performed in the previous years.

These pieces of information are usually called *pricing variables*.

We must observe that some of these variables are available for every potential customer (such as his age and address), while others are only available for policyholder that are already customers (such as telematic data that is available only if the policyholder agreed on installing on their car the device that collects this data).

Moreover, even considering the variables that are available for every customer, it is important to be aware on how reliable they are. Some of them comes from official documents (as customer age and address or bonus-malus class), but others could be declared by the customer and his statements are not easily verifiable by the insurer (as the ownership of a private box or how many drivers will drive the vehicle).

This topic of variables reliability fits in the wider framework of fraud detection. Insurance companies put a lot of effort in preventing frauds. This is done with active actions, such as documents checks and inspections, and with predictive fraud detections models. The two most common categories of frauds are underwriting frauds (such as false declaration on insurance related data) and settlement frauds (such as faking an accident). The customer information on the policyholder is usually important to predict both underwriting frauds and settlement frauds. Usually customers that have a longer relationship with the company and own many policies are less likely to commit frauds.

Regarding the topic of variables reliability, the Italian Insurance Associations ([ANIA](#)) in the last years made some big steps forward by collecting in its databases a lot of information about policyholders and vehicles and making it available to insurance companies. For example, by logging in these databases it is possible, at the moment of the quote request, to retrieve useful insurance specific information such as the number of claims caused by the customer in the previous years or how long he has been covered and useful information on his vehicle such as when it has been registered or how many changes of ownership did it experienced.

One of the roles of the actuary is to understand how reliable the information on the policyholder is and to decide how to use that information.

### 1.3.2 Pricing variables encoding

Formally the pricing variables can be encoded as a vector of real numbers.  $x_i = (x_{i1}, x_{i2}, \dots, x_{ip}) \in \mathcal{X} \subseteq \mathbb{R}^p$ . In the modeling framework they can be also called explanatory variables, covariates, predictors or features.

The pricing variables can be of two types:

1. *Quantitative variables*: variables, like policyholder age or vehicle mass, that can be easily represented as a number;
2. *Qualitative variables*: variables, like policyholder sex or vehicle make, that represent a category and are usually represented with strings.

The quantitative variables, with eventual transformations, are already suitable to be used.

To facilitate the use of the qualitative variables, they are usually encoded as sets of binary variables.

If a variable  $x$  has only 2 possible modalities, it can be easily encoded in a binary variable  $z$  that assigns 0 to one modality and 1 to the other. For example, if  $x = \text{sex}$ , it can be encoded this way:

$$z = \begin{cases} 1 & \text{if sex} = \text{'Male'} \\ 0 & \text{if sex} = \text{'Female'} \end{cases}$$

In general, if a variable  $x$  has  $K$  modalities, it can be encoded in  $K - 1$  binary variables  $z_1, z_2, \dots, z_{K-1}$ . For example, if  $x = \text{make}$  and it can have 4 possible modalities ('Fiat', 'Alfa-Romeo', 'Lancia', 'Ferrari') it can be encoded this way:

$$z_1 = \begin{cases} 1 & \text{if make} = \text{'Fiat'} \\ 0 & \text{otherwise} \end{cases}$$

$$z_2 = \begin{cases} 1 & \text{if make} = \text{'Alfa-Romeo'} \\ 0 & \text{otherwise} \end{cases}$$

$$z_3 = \begin{cases} 1 & \text{if make} = \text{'Lancia'} \\ 0 & \text{otherwise} \end{cases}$$

The variables  $z_1, z_2, z_3$  are called dummy variables. We can observe that all the information about the make is embedded in just these 3 variables, so a fourth dummy variable that indicate the modality 'Ferrari' is not needed. Indeed:

$$\text{make} = \text{'Ferrari'} \iff z_1 = z_2 = z_3 = 0$$

In table 1.1 the dummy variable encoding is illustrated.

**Table 1.1:** Dummy variables encoding.

make	z1	z2	z3
Fiat	1	0	0
Alfa-Romeo	0	1	0
Lancia	0	0	1
Ferrari	0	0	0

For some models it is suggested to use also the dummy variable that indicates the  $K^{\text{th}}$  modality. This encoding is called one-hot encoding and it is mainly used in Neural Networks. For the models considered in this paper it is preferred the  $K - 1$  dummy variables encoding, so we will always consider it.

In the following, when I use the notation  $x_i = (x_{i1}, x_{i2}, \dots, x_{ip})$ , I'll always consider that the qualitative variables have been already encoded as dummy variables, so  $(x_{i1}, x_{i2}, \dots, x_{ip}) \in \mathcal{X} \subseteq \mathbb{R}^p$

### 1.3.3 Pricing Rule and Modeling

The pricing variables are used as input of a *Pricing Rule*.

**Definition 1.5** (Pricing Rule). A *Pricing Rule* is a function  $f(\cdot)$  that from an instance of a set of pricing variables  $x_i \in \mathcal{X}$  returns a price:

$$\begin{aligned} f: \mathcal{X} &\longrightarrow R_+ \\ x_i &\longmapsto P_i \end{aligned}$$

The process of pricing consists in defining a Pricing Rule based on observed data from the past and assumptions on the future.

The first step for defining a Pricing Rule is to model the total cost of claims  $S$  and obtain a pricing rule for the risk premium  $P^{(risk)}$ .

**Definition 1.6** (Modeling). Modeling a *response variable*  $Y$  means finding a function

$$r: \mathcal{X} \rightarrow \mathcal{C}$$

that, given a set of explanatory variables  $x_i = (x_{i1}, x_{i2}, \dots, x_{ip}) \in \mathcal{X} \subseteq \mathbb{R}^p$ , returns the expected value of the response variable  $E(Y)$  and eventually other moments of  $Y$  or even the full distribution of  $Y$ .

In definition 1.6 I used a generic  $\mathcal{C}$  as codomain of the function  $r(\cdot)$  to not specify whether the model describes just  $E(Y)$  (and so  $\mathcal{C} = \mathbb{R}$ ) or something more, such as the couple  $(E(Y), Var(Y))$  or the full distribution of  $Y$ .

As we observed in section 1.2.2, under the compound distribution hypotheses, it is not needed to model directly the total cost of claims  $S$ , but we can separately model  $N$  and  $Z$ .

### 1.3.4 Response variables and distributions

Usually in statistical modeling, the response variables are seen as random variables with a distribution belonging to a specified family.

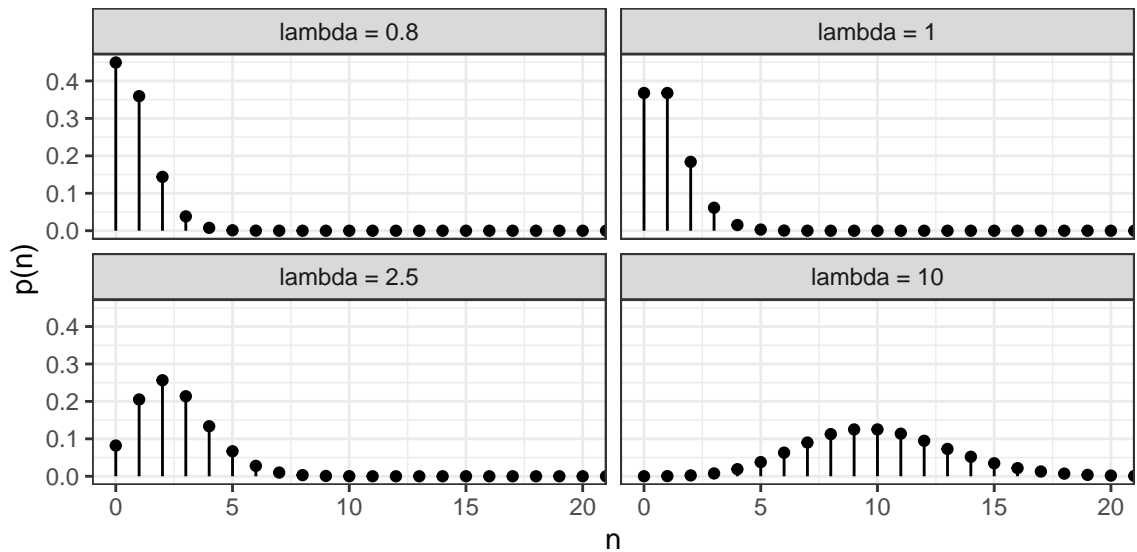
### 1.3.4.1 Distribution for the claims count $N$

The claim count  $N$  is a discrete variable with determination in  $\{0, 1, 2, 3, \dots\}$ . Even if in practice the number of claims can't be arbitrarily high,  $N$  is usually modeled with distributions that give a positive probability to all the numbers in its support. One of the most common distribution used for  $N$  is the Poisson distribution.

**Definition 1.7** (Poisson Distribution). A random variable  $N$  with support  $\{0, 1, 2, 3, \dots\}$  has a Poisson distribution, if its probability function is:

$$p_N(n) = P(N = n) = e^{-\lambda} \frac{\lambda^n}{n!}, \quad \lambda > 0$$

We will indicate it with the notation  $N \sim \text{Poisson}(\lambda)$ .



**Figure 1.2:** Poisson distribution for some values of  $\lambda$ .

The Poisson distribution is a parametric distribution that depends on only the parameter  $\lambda$ . In figure 1.2, for different levels of  $\lambda$  the distribution is represented. These plots show how for larger values of  $\lambda$ , the distribution is shifted to larger values and it is wider.

Indeed, the first two moments are:



$$E(N) = \lambda$$

$$Var(N) = \lambda$$

Thus, increasing  $\lambda$ , both  $E(N)$  and  $Var(N)$  increase.

Looking to the distribution shape, we can see that:

- if  $\lambda < 1$ , the mode is in  $n = 0$ ;
- if  $\lambda = 1$ ,  $p(0) = p(1) = \frac{1}{e}$ ;
- if  $\lambda > 1$ , the mode is in a value greater than 0 and, as  $\lambda$  increases, the distribution assumes a bell shape similar to the Normal distribution one. The convergence to the Normal distribution can be obtained with the *Central Limit Theorem*.

In non-life insurance we usually are in the case with  $\lambda < 1$ . E.g. the average number of claims for motor third party liability insurances in Italy, in 2018 has been 5.68%<sup>1</sup>.

The property  $Var(N) = E(N)$  is an important constraint when the distribution is used in practice. It is possible that the observed data shows a different pattern. Often the observed data shows a situation where  $Var(N) > E(N)$ . This phenomenon is called *overdispersion*.

To address this issue it is possible to use more flexible distributions, such as Negative-Binomial distribution, or to adopt less assumptions on the response variable distribution. One common technique is the assumption of Quasi-Poisson distribution, that we will describe in chapter 2.

#### 1.3.4.2 Exposure

In section 1.1 we said that non-life insurances usually have a fixed coverage period that usually spans for one year. Often we work with portfolios of insurances with different coverage periods. For example, this could be due to the presence of

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<sup>1</sup>ANIA yearly statistical report for motor third party liability

insurances born with shorter coverage periods or to the presence of insurances that has been closed earlier. Moreover, in companies data, often insurance data are collected for accounting years. This means that, if an insurance coverage  $c$  spans in two consecutive years  $a$  and  $a + 1$ , it is collected as two records: the couple  $(c, a)$  and the couple  $(c, a + 1)$ . This situation is quite common, as usually coverages start during the year and not all at the first of the year.

The coverage span for an insurance coverage is called *exposure* and it is usually measured in years-at-risk. For instance, if an insurance coverage spans for 3 months, it corresponds to a quarter of year, so the exposure, measured in years-at-risk, is  $v = \frac{1}{4}$ . The term year-at-risk comes from the fact that the policyholder exposure is a risk for the insurer, so the exposure is the period in which the insurer is exposed to the risk of paying claims.

It is natural to assume that, if a policyholder has a longer exposure, it is expected for him to experience more claims. Considering that we have to work with policies with different exposures, in order to take this aspect into account, the usual assumption taken in the following. Said  $M$  the number of claims the policyholder will experience during his period of exposure  $v$  and  $N$  the number of claims the policyholder would experience during one year, we assume  $E(M) = vE(N)$ .

This assumption can be further extended if we assume that the claims come from a *Poisson process*.

**Definition 1.8** (Counting Process). A stochastic process  $\{N(t), t \geq 0\}$  is called *counting process* if:

1. The determination of  $N(t)$  are natural numbers  

$$N(t) \in \{0, 1, 2, \dots\} \quad t \geq 0$$
2. The process is not decreasing  

$$s < t \Rightarrow N(s) \leq N(t)$$

In a counting process  $\{N(t), t \geq 0\}$ :

- $N(t)$  can be interpreted as the number of events or arrivals that occur in the period  $[0, t]$ ;

- $N(t) - N(s)$ ,  $s \leq t$  can be interpreted as the number of events or arrivals that occur in the period  $]s, t]$ .  $N(t) - N(s)$  is also called *increment* of the process.

The counting process can be used to model the number of claims that occur to a specific policy.

**Definition 1.9** (Poisson Process). A counting process  $\{N(t), t \geq 0\}$  is a *Poisson process* with intensity  $\lambda$  if:

1. The increments of the process are stochastically independent  
 $\forall n \geq 0, \forall s_1 < t_1 \leq \dots \leq s_n < t_n$   
 $\Rightarrow N(t_1) - N(s_1), \dots, N(t_n) - N(s_n)$  are stochastically independent;
2. The probability of arrival in an interval is proportional to the size of the interval  
 $\forall t \geq 0, \forall \Delta t > 0 \Rightarrow P(N(t + \Delta t) - N(t) = 1) = \lambda \Delta t + o(\Delta t)$   
 where  $\lim_{\Delta t \rightarrow 0} \frac{o(\Delta t)}{\Delta t} = 0$
3. Multiple arrivals are excluded  
 $\forall t \geq 0, \forall \Delta t > 0 \Rightarrow P(N(t + \Delta t) - N(t) \geq 2) = o(\Delta t)$
4. Arrivals at time 0 are almost impossible  
 $P(N(0) = 0) = 1$

Under these hypotheses we obtain the following result:

**Theorem 1.1** (Poisson Process). *If  $\{N(t), t \geq 0\}$  is a Poisson process with intensity  $\lambda$ , then:*

$$\forall t \geq 0, \forall \Delta t > 0, \Rightarrow N(t + \Delta t) - N(t) \sim \text{Poisson}(\lambda \Delta t)$$

This result tells us that the distribution of the number of events in any interval  $]t, t + \Delta t]$  only depends on the size of the interval  $\Delta t$ . Moreover, for the Poisson property we saw in section 1.3.4.1, we get:

$$E(N(t + \Delta t) - N(t)) = \lambda \Delta t$$

So, the expected number of arrivals is proportional to the size of the interval  $\Delta t$ . The intensity of the process  $\lambda$  can be also interpreted as the expected number of claims in a unitary period.

If we assume that the claims that occur to a policy comes from a Poisson process with intensity  $\lambda$ , if we observe that policy for the period  $]t, t + v]$ , the claims count in that exposure period  $M$  are distributed as:

$$M \sim \text{Poisson}(\lambda \Delta t)$$

In particular, if the observed period spans 1 year, we get:

$$M = N \sim \text{Poisson}(\lambda)$$

#### 1.3.4.3 Distribution for the claims size $Z$

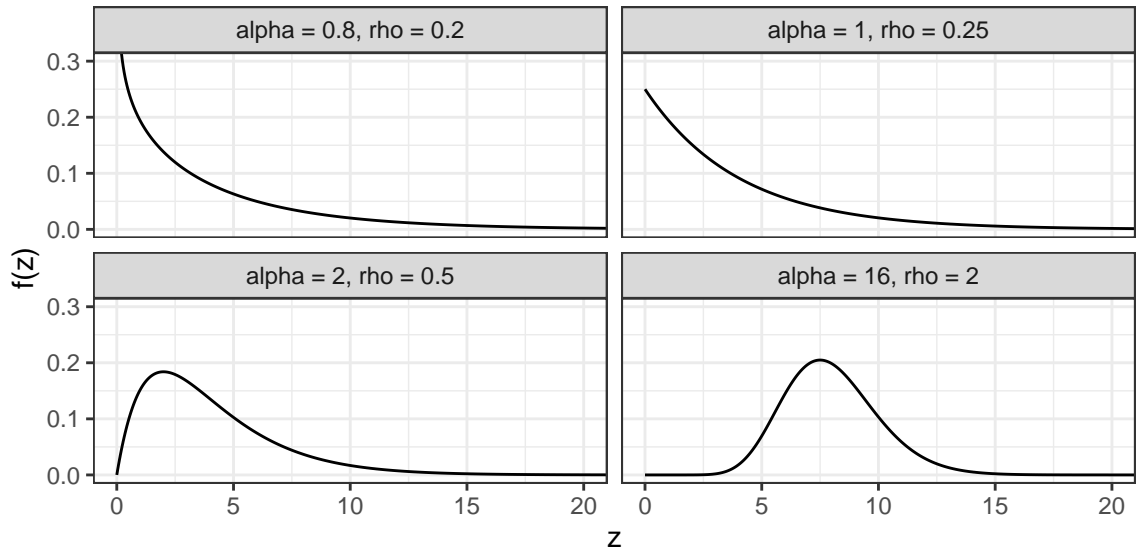
The claim size  $Z$  is a continuous variable with determination in  $[0, +\infty[$ . As for the claims count  $N$ , even if in practice it can't be arbitrarily high, it is usually modeled with distributions that give a positive probability to all the numbers in  $]0, +\infty[$ . As the null claims are excluded, it is natural to assume  $P(Z = 0) = 0$ . One of the most common distribution used for  $Z$  is the Gamma distribution.

**Definition 1.10** (Gamma Distribution). A random variable  $Z$  with support  $[0, +\infty[$  has a Gamma distribution, if its probability density function is:

$$f_Z(z) = \frac{\rho^\alpha}{\Gamma(\alpha)} z^{\alpha-1} e^{-\rho z}, \quad \alpha > 0, \rho > 0$$

where  $\Gamma(\alpha) = \int_0^{+\infty} z^{\alpha-1} e^{-z} dz$

We will indicate it with the notation  $Z \sim \text{Gamma}(\alpha, \rho)$ .



**Figure 1.3:** Gamma distribution for some values of  $\alpha$  and  $\rho$ .

The Gamma distribution is a parametric distribution that depends on two parameters:

- $\alpha > 0$ , called shape parameter
- $\rho > 0$ , called scale parameter

The first two moments of the Gamma distribution are:

$$E(Z) = \frac{\alpha}{\rho}$$

$$Var(Z) = \frac{\alpha}{\rho^2}$$

In figure 1.3, for different levels of  $\alpha$  and  $\gamma$  the distribution is represented. These plots show how changing the values of  $\alpha$  and  $\gamma$ , the shape changes. We can see that:

- if  $\alpha < 1$ ,  $f_z(\cdot)$  is not defined in 0 and it has a vertical asymptote in  $z = 0$ . In  $[0, +\infty]$  it is monotonically decreasing.

- if  $\alpha = 1$ ,  $f_z(\cdot)$  starts from  $f(0) = \rho$  and then decreases monotonically. In this case, the density function becomes  $f_z(z) = \rho e^{-\rho z}$  and the distribution is also called exponential distribution.
- if  $\alpha > 0$ ,  $f_z(\cdot)$  starts from  $f(0) = 0$ , increases until the mode and then decreases.

In figure 1.3 the first three distributions represented have the same expected value  $E(Z) = \frac{\alpha}{\rho} = 4$ , but different shapes. The third and the fourth have the same variance  $Var(Z) = \frac{\alpha}{\rho^2} = 8$ , but different expected values. As the shape parameter  $\alpha$  increases, the distribution assumes a bell shape similar to the Normal distribution one. The convergence to the Normal distribution can be obtained with the *Central Limit Theorem*.

From the parametrical form of the moments, we get that  $Var(Z)$  and  $E(Z)$  are linked as follows:

$$Var(Z) = \frac{\alpha}{\rho^2} = \frac{1}{\alpha} (E(Z))^2$$

From this relation we find that the coefficient of variation is:

$$CV(Z) = \frac{\sqrt{Var(Z)}}{E(Z)} = \frac{1}{\sqrt{\alpha}}$$

This result means that, given the shape parameter  $\alpha$ , the coefficient of variation is constant. As we saw for Poisson distribution, it is possible that observed data shows a different pattern.

Another characteristic of Gamma distribution that could be problematic in modeling claims size is that it has a light tail. This means that, as  $z$  goes to  $+\infty$ ,  $f_z(z)$  approaches 0 quite fast. This could lead to a poor fitting for *large claims*. Other distributions with heavier tails are for example the *log-Normal* and the *Pareto*.

#### 1.3.4.4 Large claims

Modeling large claims is quite difficult in practice because usually there is not a lot of observed data on them, so it is hard to understand if they are related to some risk factors (identifiable by the pricing variables) or they happen just by chance.

First of all, to model large claims, we must define what a large claim is. What is usually done in practice is just choosing a threshold  $\bar{z}$  and considering large all the claims with a size that exceeds that threshold. The value  $\bar{z}$  must be chosen sufficiently big to consider large the claims above  $\bar{z}$ , but not so big that there are not enough observed claims that exceeds  $\bar{z}$ . One common choice for Motor Third Party Liability in European markets could be  $\bar{z} = 100'000\text{€}$ .

**Definition 1.11** (Large and Attritional Claims). Given a predetermined threshold  $\bar{z}$ , we say that:

- a claim  $Z$  is a *large claim* if  $Z > \bar{z}$
- a claim  $Z$  is an *attritional claim* if  $Z \leq \bar{z}$

For each claim  $Z$  we call:

- *Capped Claim Size*  
 $Z' = \min(Z, \bar{z});$
- *Excess Over the Threshold*  
 $Z'' = \max(Z - \bar{z}, 0).$

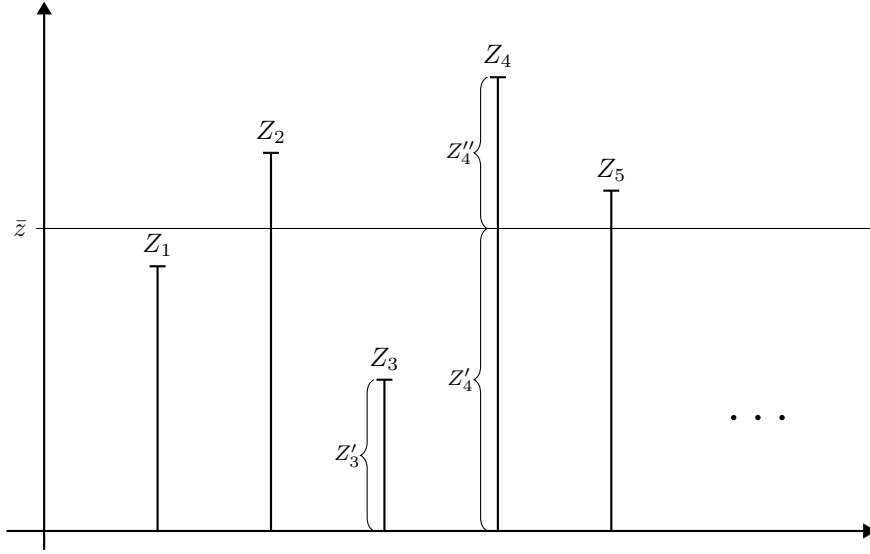
In figure 1.4 the *Capped Claim Size* and the *Excess Over the Threshold* are shown. It is easy to show that  $Z$  can be decomposed as:

$$Z = Z' + Z''$$

Given the total number of claims  $N$ , it can be decomposed as:

$$N = N^{(a)} + N^{(l)}$$

where



**Figure 1.4:** Large claims.

- $N^{(a)}$  is the attritional claims count, i.e. the number of claims with size  $Z \leq \bar{z}$ ;
- $N^{(l)}$  is the large claims count, i.e. the number of claims with size  $Z > \bar{z}$ ;

Let's indicate with  $Z_{(i)}$  the  $i^{\text{th}}$  in order from the smallest to the bigger. Sorting the claims we can separate the attritional claims from the large claims as follows:

$$\underbrace{Z_{(1)}, Z_{(2)}, \dots, Z_{(N^{(a)})}}_{\text{Attritional Claims}}, \underbrace{Z_{(N^{(a)}+1)}, Z_{(N^{(a)}+2)}, \dots, Z_{(N^{(a)}+N^{(l)})}}_{\text{Large Claims}}$$

In order to model the large claims it is possible to use the following three decompositions of the total cost of claims  $S$ :



$$\begin{aligned}
 S &= \underbrace{Z_{(1)} + Z_{(2)} + \cdots + Z_{(N^{(a)})}}_{\text{Attritional Claims}} + \underbrace{Z_{(N^{(a)}+1)} + Z_{(N^{(a)}+2)} + \cdots + Z_{(N^{(a)}+N^{(l)})}}_{\text{Large Claims}} \\
 &= \underbrace{\sum_{i=1}^{N^{(a)}} Z_{(i)}}_{=S^{(a)}} + \underbrace{\sum_{i=N^{(a)}+1}^{N^{(a)}+N^{(l)}} Z_{(i)}}_{=S^{(l)}} = S^{(a)} + S^{(l)} \tag{1.11}
 \end{aligned}$$

$$S = \sum_{i=1}^N Z_i = \sum_{i=1}^N (\{Z_i | Z_i > \bar{z}\} I_{Z_i > \bar{z}} + \{Z_i | Z_i \leq \bar{z}\} I_{Z_i \leq \bar{z}}) \tag{1.12}$$

$$S = \sum_{i=1}^N Z_i = \sum_{i=1}^N (Z'_i + Z''_i) = \sum_{i=1}^N (Z'_i + Z''_i I_{Z_i > \bar{z}}) \tag{1.13}$$

These three decompositions of  $S$  are useful because they provide three decompositions of  $E(S)$ :

$$\begin{aligned}
 E(S) &= E(S^{(a)}) + E(S^{(l)}) \\
 &= E(N^{(a)})E(Z|Z \leq \bar{z}) + E(N^{(l)})E(Z|Z > \bar{z}) \tag{1.14}
 \end{aligned}$$

$$\begin{aligned}
 E(S) &= E(N)E(Z) \\
 &= E(N) [P(Z \leq \bar{z})E(Z|Z \leq \bar{z}) + P(Z > \bar{z})E(Z|Z > \bar{z})] \\
 &= E(N) [(1 - P(Z > \bar{z}))E(Z|Z \leq \bar{z}) + P(Z > \bar{z})E(Z|Z > \bar{z})] \tag{1.15}
 \end{aligned}$$

$$\begin{aligned}
 E(S) &= E(N)E(Z) \\
 &= E(N) [E(Z') + P(Z > \bar{z})E(Z'')] \tag{1.16}
 \end{aligned}$$

1.14, 1.15 and 1.16 provide three approaches to model attritional and large claims.

1. Looking to 1.14 we can model separately attritional claims and large claims. Modeling  $N^{(a)}$  and  $Z|Z \leq \bar{z}$  we estimate the total cost of claims for the

attritional part  $S^{(a)}$ ; modeling  $N^{(l)}$  and  $Z|Z > \bar{z}$  we estimate the total cost of claims for the large part  $S^{(l)}$ .

2. Looking to 1.15 we can model together the claim count  $N$ , and then we can model the cost of the attritional claims  $Z|Z \leq \bar{z}$ , the cost of the large claims  $Z|Z > \bar{z}$  and the probability to exceed the threshold  $P(Z > \bar{z})$ .
3. Looking to 1.16 we can model together the claim count  $N$ , and then we can model the capped claims size  $Z'$ , the excess over the threshold  $Z''$  and the probability to exceed the threshold  $P(Z > \bar{z})$ .

If the large claims component weights a lot on the total cost of claims, these approaches could lead to quite different estimations of  $E(S)$ . In particular, if in the observed data the number of large claims is small, it will be hard to model both  $N^{(l)}$  and  $P(Z > \bar{z})$ , so the modeling process could lead to a flat or almost flat model for these components. However, with the first approach, a flat model for  $N^{(l)}$  leads to distribute the observed total cost of large claims proportionally to all the policies, while with the second and the third, a flat model for  $P(Z > \bar{z})$  leads to distribute the observed total cost of large claims proportionally to the expected number of claims  $E(N)$ . So, with the first approach, a flat model brings to more solidarity between policies, while, with the second approach, a flat model could bring to an exacerbation of the differences identified by modeling  $N$ .

For the second approach we must also introduce a distribution suitable for modeling  $P(Z > \bar{z})$ .

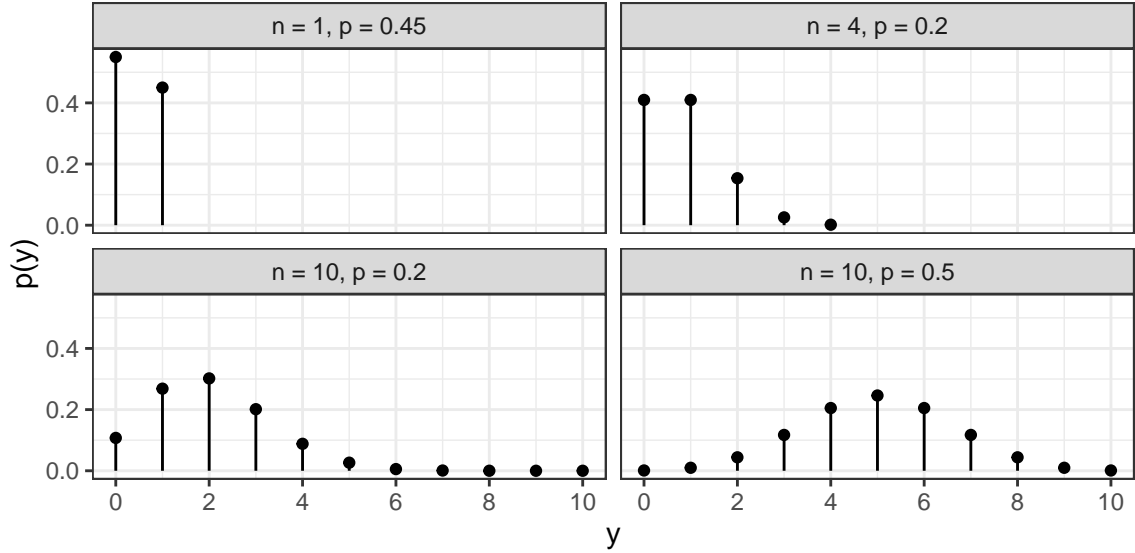
#### 1.3.4.5 Binomial distribution

The *binomial distribution* is used to model the counting on events that occurs (successes) in a fixed amount of trials  $n$ . For example we can use it to model the number of large claims withing a fixed number of  $n$  claims.

**Definition 1.12** (Binomial Distribution). A random variable  $Y$  with support  $\{0, 1, 2, \dots, n\}$  has a Binomial distribution, if its probability function is:

$$p_Y(y) = P(Y = y) = \binom{n}{y} p^y (1 - p)^{n-y}, \quad p \in [0, 1]$$

We will indicate it with the notation  $Y \sim \text{Binomial}(n, p)$ .



**Figure 1.5:** Binomial distribution for some values of  $n$  and  $p$ .

The binomial distribution is a parametric distribution that depends on the parameters  $n$  and  $p$ .  $n$  represents the number of trials, while  $p$  represents the probability for a trial to success. The assumption is that the  $n$  trials are identical, so they have all the same probability  $p$  to success. In figure 1.5, for different levels of  $n$  and  $p$ , the distribution is represented.

The first two moments of the binomial distribution are:

$$E(N) = np$$

$$\text{Var}(N) = np(1 - p)$$

If  $n = 1$ , the binomial distribution assumes only the values 1 (with probability  $p$ ) and 0 (with probability  $1 - p$ ). In this case it is also called *Bernoullian distribution* and it can be used to model the indicator of an event  $I_E$ .

If  $n > 1$ , the binomial distribution assumes a shape centered on its expected value  $E(Y) = np$  and fading for values of  $y$  that moves away from  $E(Y)$ . As  $n$

increases, the distribution assumes a bell shape similar to the Normal distribution one. The convergence to the Normal distribution can be obtained with the *Central Limit Theorem*.

In non-life insurance pricing, the binomial distribution can be used to model the probability for a claim to have specific characteristics. For example we can use it to model the probability that a certain claim is a large one  $P(Z > \bar{z})$  in order to model separately attritional claims severity  $\{Z|Z \leq \bar{z}\}$  and large claims severity  $\{Z|Z > \bar{z}\}$ , as we have seen in section 1.3.4.4.

Another example is the decomposition between claims with only damage to things and claims with also injuries to people. Modeling separately these two components is useful because they usually have a different distribution for the claim size.

As for large claims we can decompose  $S$  in the following two ways:

$$\begin{aligned} E(S) &= E(S^{(things)}) + E(S^{(inj)}) \\ &= E(N^{(things)})E(Z|\bar{J}) + E(N^{(inj)})E(Z|J) \end{aligned} \quad (1.17)$$

$$\begin{aligned} E(S) &= E(N)E(Z) \\ &= E(N) [P(\bar{J})E(Z|\bar{J}) + P(J)E(Z|J)] \\ &= E(N) [(1 - P(J))E(Z|\bar{J}) + P(J)E(Z|J)] \end{aligned} \quad (1.18)$$

where:

- $N^{(things)}$  is the number of claims with only damage to things;
- $N^{(inj)}$  is the number of claims with injuries;
- $J$  is the event that represent that a specific claim presents injuries; such as  $Z$  is a representative for  $Z_1, Z_2, \dots, Z_N$ ,  $J$  is a representative for  $J_1, J_2, \dots, J_N$ .

Combining this decomposition with what we have seen in large claims decomposition, we can further develop our decomposition taking into account both the presence or absence of injuries and the occurrence or not of a large claim. One example could be:

$$\begin{aligned}
 E(S) &= E(N) [(1 - P(J)) E(Z|\bar{J}) + P(J) E(Z|J)] \\
 &= E(N) \{ \\
 &\quad (1 - P(J)) E(Z|\bar{J}) \\
 &\quad + P(J) [ \\
 &\quad \quad P(Z \leq \bar{z}|J) E(Z | Z \leq \bar{z} \wedge J) \\
 &\quad \quad + P(Z < \bar{z}|J) E(Z | Z < \bar{z} \wedge J) \\
 &\quad ] \\
 &\quad \}
 \end{aligned}$$

This way, we are decomposing only the claims with injuries between attritional and large. That makes sense because claims that don't produce injuries usually have small severities.

### 1.3.5 Model fitting and data available

Once we have chosen how to decompose  $S$ , we have to model the response variables needed for that decomposition  $(N, Z, I_J, \dots)$  with the explanatory variables. Thus we have to estimate a function  $r : \mathcal{X} \rightarrow \mathcal{C}$  as defined in 1.6.

In order to estimate  $r(\cdot)$  we have also to take some assumptions on the distribution of the response variable and on the shape of  $r(\cdot)$ . We will call *model* a set of assumptions on the response variable and on the shape of  $r(\cdot)$ . We will discuss some of the most widespread models for claims count and claims severity in chapter 2.

Defined the model, we have to estimate it using observed data. In general, to model a response variable  $Y_i$  with the explanatory variables  $x_i = (x_{i1}, x_{i2}, \dots, x_{ip}) \in \mathcal{X} \subseteq \mathbb{R}^p$ , the observed data is in the form:

$$(x_1, w_1, y_1), (x_2, w_2, y_2), \dots, (x_i, w_i, y_i), \dots, (x_n, w_n, y_n)$$

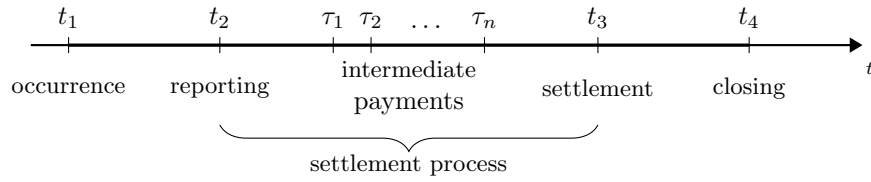
where:

- $n$  is the number of observations in the dataset;
- $x_i \in \mathcal{X} \subseteq \mathbb{R}^p$  is the set of explanatory variable for the observation  $i$ ;
- $w_i$  is the weight for the observation  $i$ ;
- $y_i \in \mathcal{Y} \subseteq \mathbb{R}$  is the realization of the response variable  $Y_i$  for the observation  $i$ .

What an observation is, depends on the variable we are modeling. For instance:

- If we are modeling the yearly claim count  $N_i$ , each observation could be a policy (or a couple (policy, accounting year)), the weights could be the exposures  $v_i$  and the realizations of response variables could be the number of observed claims for that policy (or couple (policy, accounting year)).
- If we are modeling the claim severity  $Z_j$ , each observation could be a claim  $j$ , the weights could all be 1 and the realizations of response variables could be the observed cost for the claim  $j$ . It is also possible to model the claim severity taking into account the total cost of claims for the policy  $S_i = \sum_{j=1}^{N_i} Z_j$ . In this case, each observation would be a policy  $i$ , the weights would be the number of claims for each policy  $n_i$  and the realizations of response variables would be the total observed cost for the claims of the policy  $i$ .
- If we are modeling the occurrence of injuries in a claim  $I_{Jj}$ , each observation could be a claim  $j$ , the weights could be all 1 and the realizations of response variables could be an indicator that assume the value 1 if the claim  $j$  caused injuries and 0 otherwise. As for the claim severity, we can also aggregate data for policy, so each observation would be a policy  $i$ , the weights would be the number of claims  $n_i$  for the policy  $i$  and the realizations of response variables would be the number of claims that caused injuries among the claims of the policy  $i$ .

In each of these cases,  $y_i$  is seen as a realization of the random variable  $Y_i$ . With an inferential process we obtain estimations on  $Y_i$  based on observations of their realizations  $y_i$ .



**Figure 1.6:** Claim timeline.

### 1.3.5.1 Settlement process and IBNR

One of the challenges in non-life insurance pricing is that obtaining the observed data is not so straightforward. In many insurance coverages, the settlement process could last many years, so, if we want to develop models using data from recent years, not all the information is available. To better understand this aspect we have to discuss how the settlement process works.

In figure 1.6 the settlement process for a claim is represented. At time  $t_1$  the insured event (e.g. an accident) occurs. From this moment a liability for the insurer emerges, even if the insurer has not been notified yet. This liability is called *Outstanding Loss Liability*. In  $t_2$  the claim is reported and the insurance is notified about the occurrence of the event. From this moment the settlement process starts. This process consists in evaluating the event and understanding the responsibilities of the parts and the entity of the damage. During this process, controversies between the parts can emerge and, in particular if injuries occurred, the damage evaluation can take a lot of time. When the situation is clear and everything is defined, the claim is settled and the liabilities are paid. In  $t_3$  we have the settlement and in  $t_4$  the claim is closed. It is possible that  $t_4 = t_3$ , but in general it is not always the case. If the settlement process takes a long time and the insurer already knows he will have to pay something, he can pay some partial payments during the period  $[t_2, t_3]$ . These intermediate payments are paid in the times  $\tau_1, \tau_2, \dots, \tau_n \in [t_2, t_3]$ . It is also possible that a claim is opened and then gets closed without any payment. After the closing ( $t_4$ ) it is also possible that a claim is reopened and that more payments emerge.

From the moment the claim is reported ( $t_2$ ), the insurer estimates how much he is going to pay for that claim and he allocates that sum in a reserve. As new information emerges and some payments are settled, the reserve is updated. The

aim for this reserve is to have a best estimate for the future payments for the claims already emerged. As the claim gets settled, the sum between the payed and the reserved converges to the final cost of the claim.

From this description emerges that:

- In the period  $]t_1, t_2[$  the insurer has an outstanding loss liability for an event that has not been reported yet; in this case we will talk about *Incurred But Not Yet Reported* claim (IBNyR).
- In the period  $[t_2, t_3[$  the insurer has an outstanding loss liability for an event that has been reported, but has not been totally settled yet, so that liability is just an estimate; in this case we will talk about *Incurred But Not Enough Reported* claim (IBNeR).

### 1.3.5.2 Model fitting with available data

The IBNyR and IBNeR issue is particularly challenging when we have to perform a risk evaluation at a specific time  $t$ . In general  $t_1, t_2, \dots$  are not known a priori, so we don't know if in the future more claims for accidents occurred in the past will be reported and we don't know if the ones that are already reported will experience a revaluation. That means that, in general, when we model  $N$  and  $Z$  at a specific time  $t$ , we can't observe the total number of claims occurred to each policy  $n_i$  and the payments for each claim  $z_j$ . What we can use is:

- $n_i^{(t)} = n_i^{(\text{reported in } t)}$

where:

–  $n_i^{(\text{reported in } t)}$  is the number of reported claims in  $t$  for the policy  $i$ ;

- $z_j^{(t)} = z_j^{(\text{payed in } t)} + z_j^{(\text{reserved in } t)}$

where:

–  $z_j^{(\text{payed in } t)}$  is the amount already payed in  $t$  for the claim  $j$ ;

–  $z_j^{(\text{reserved in } t)}$  is the amount reserved in  $t$  for the claim  $j$ .



When we use this data for modeling the total cost of claims we must be particularly aware on what we are using. In general:

$$\begin{aligned} n_i^{(t)} &\neq n_i \\ z_j^{(t)} &\neq z_j \end{aligned}$$

The common case is that  $n_i^{(t)} < n_i$  and  $z_j^{(t)} < z_j$ . If we used  $n_i^{(t)}$  and  $z_j^{(t)}$  without any correction, we would underestimate both  $E(N)$  and  $E(Z)$ , obtaining a distorted estimate for  $E(S)$ .

To tackle these problems what is usually done is fitting the models for  $S_i$  with  $n_i^{(t)}$  and  $z_j^{(t)}$  and then apply a flat corrective coefficients  $\alpha$  to  $\widehat{E(S_i)}$  based on an aggregated estimate of  $E(S)$  that takes into account the long settlement process.

An estimate for the expected total cost of claims for a generic policy in the portfolio  $E(S)$  can be obtained with techniques based on runoff triangles, such as the *Chain Ladder*. We are not going to discuss these techniques in this thesis. We just have to know that these techniques provide us with an estimate for  $E(S)$ . Let's call it  $\widehat{E(S)}^{CL}$ . This estimate does not depend on explanatory variables; it is a sort of average total cost of claims for the policies in the portfolio.

Meanwhile, with the available data  $n_i^{(t)}$  and  $z_j^{(t)}$ , the fitting for all the models needed in the decomposition of  $S$  is performed and, for each policy  $i \in \{1, 2, \dots, n\}$  for  $E(S_i)$  is obtained. Let's call it  $\widehat{E(S_i)}'$ . As we used the data available in  $t$  that comes from claims not totally settled,  $\widehat{E(S_i)}'$  is a distorted estimate for  $E(S_i)$ .

We can then balance the estimates  $\widehat{E(S_i)}'$  with  $\widehat{E(S)}^{CL}$  by computing:

$$\alpha = \frac{n}{\sum_{i=1}^n \widehat{E(S_i)}'} \widehat{E(S)}^{CL}$$

and applying to the estimates as follows:

$$\widehat{E(S_i)} = \alpha \widehat{E(S_i)}'$$

We will call  $\widehat{E(S_i)}$  rebalanced estimates.

The property of these balanced estimates is that on average they point to  $\widehat{E(S)}^{CL}$ :

$$\begin{aligned} \frac{\sum_{i=1}^n \widehat{E(S_i)}}{n} &= \frac{\sum_{i=1}^n \alpha \widehat{E(S_i)}'}{n} \\ &= \alpha \frac{\sum_{i=1}^n \widehat{E(S_i)}'}{n} \\ &= \frac{\frac{\sum_{i=1}^n \widehat{E(S_i)}'}{\sum_{i=1}^n \widehat{E(S_i)}} \widehat{E(S)}^{CL} \frac{\sum_{i=1}^n \widehat{E(S_i)}'}{\sum_{i=1}^n \widehat{E(S_i)}}}{\frac{\sum_{i=1}^n \widehat{E(S_i)}}{\sum_{i=1}^n \widehat{E(S_i)}}} \\ &= \widehat{E(S)}^{CL} \end{aligned}$$

So, if  $\widehat{E(S)}^{CL}$  is a unbiased estimator for  $E(S)$ , we obtain:

$$E\left(\frac{\sum_{i=1}^n \widehat{E(S_i)}}{n}\right) = E\left(\widehat{E(S)}^{CL}\right) = E(S)$$

This procedure can be further developed by balancing not directly the total cost of claims  $E(S)$ , but its components. For example, we could separately balance the total cost of claims that only caused damage to things and the total cost of claims that caused injuries. This separation in components can lead to a more precise estimate because usually claims that caused injuries have a slower settlement process so they will have a higher corrective coefficient  $\alpha$ .

The fact that the final estimates  $\widehat{E(S_i)}$  are rebalanced on  $\widehat{E(S)}^{CL}$  means that the explanatory variables effects estimated with  $n_i^{(t)}$  and  $z_j^{(t)}$  are used just as relative effects and not absolute ones. For instance, if the model says that young people have an expected total cost of claims  $\widehat{E(S_i)}'$  that is two times the old people one, that relative coefficient 2 will be kept also in the balanced estimate  $\widehat{E(S_i)}$ .

For this reason, in practice, often the modeling is considered composed in 2 parts:

1. *Tariff Requirement* (or *Fabbisogno Tariffario*): the estimate of  $\widehat{E(S)}^{CL}$  by aggregated data;

2. *Personalization*: the estimate of  $\widehat{E(S_i)}$  and the relative coefficients.

The techniques used for *Tariff Requirement* are employed also to estimating the general liability position for the company. This information is particularly important and it is reported in the company financial statement.

## 1.4 Commercial Tariff and Offer Price

In section 1.2.3 we defined the *Risk Premium*  $P^{(risk)} = E(S)$  and the *Technical Price*  $P^{(tech)} = E(S) + E$  and in section 1.3 we described how the risk premium is estimated. In this thesis we are not going to deal with the estimate of the expenses. In this section we are going to discuss what the tariff and the offer price are and which are the further needs that the offer should satisfy.

## 1.5 Non-Life Insurance in Italy

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## 1.6 The actuary role

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Questa è una citazione (Shea et al., [2014](#); Lottridge et al., [2012](#))

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# 2

## Statistical models for Non Life Insurance Pricing

### 2.1 Statistical Models

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### 2.1.1 GLM

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### 2.1.2 Elastic Net

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### 2.1.3 GAM

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### 2.1.4 GBM

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## 2.2 Model comparison

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### 2.3 The actuary importance

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## 2.4 Implementation

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# 3

## Practical application

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### 3.1 Data description

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### 3.2 Model used

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### 3.3 Model assessment

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### 3.4 Results

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