Machine Learning for econometrics

Event studies: Causal methods for pannel data

Authors

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Motivation

Estimation of the effect of a treatment when data is:

Aggregated: country-level data such as employment rate, GDP...



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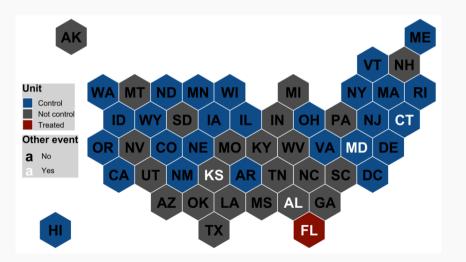


Figure from (Degli Esposti et al., 2020)

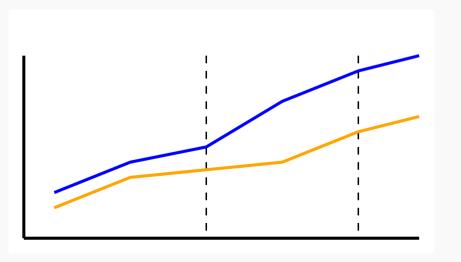
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Aggregated: country-level data such as employment rate, GDP...

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Staggered adoption of the treatment: units adopt the policy/treatment at different times...

This setup is known as

Panel data, event studies, longitudinal data, time-series data.

Examples of event studies

Archetypal questions

- Did the new marketing campaign had an effect on the sales of a product?
- Did the new tax policy had an effect on the consumption of a specific product?
- Did the guidelines on the prescription of a specific drug had an effect on the practices?

Examples of event studies

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Modern examples

- What is the effect of the extension of Medicaid on mortality? (Miller et al., 2019)
- What is the effect of Europe's protected area policies (*Natura 2000*) on vegetation cover and on economic activity? (Grupp et al., 2023)
- Which policies achieved major carbon emission reductions? (Stechemesser et al., 2024)

Setup: event studies are quasi-experiment

Quasi-experiment

A situation where the treatment is not randomly assigned by the researcher but by nature or society.

It should introduce *some* randomness in the treatment assignment: enforcing treatment exogeneity, ie. ignorability (ie. unconfoundedness).

Setup: event studies are quasi-experiment

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A situation where the treatment is not randomly assigned by the researcher but by nature or society.

It should introduce *some* randomness in the treatment assignment: enforcing treatment exogeneity, ie. ignorability (ie. unconfoundedness).

Other quasi-experiment designs

- Instrumental variables: a variable that is correlated with the treatment but not with the outcome.
- Regression discontinuity design: the treatment is assigned based on a threshold of a continuous variable.

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Reminder on difference-in-differences

Difference-in-differences

History

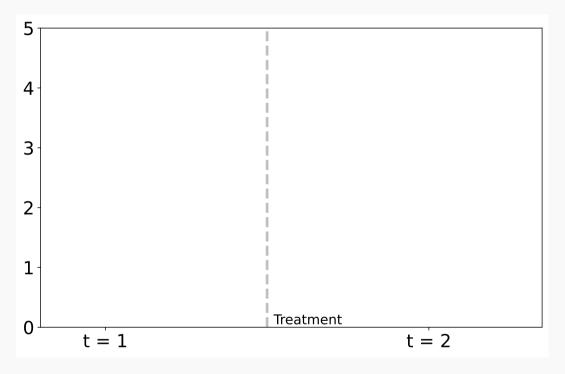
- First documented example (though not formalized): John Snow showing how cholera spread through the water in London (Snow, 1855)¹
- Modern usage introduced formally by (Ashenfelter, 1978), applied to labor economics

Idea

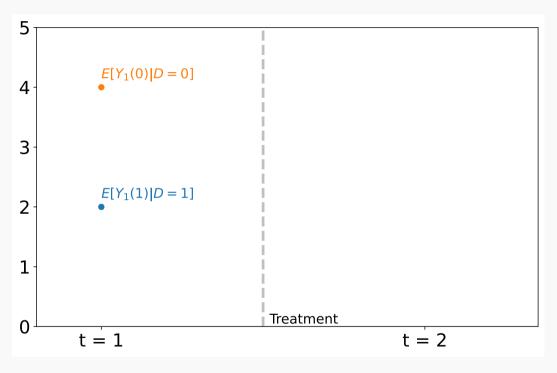
- Contrast the temporal effect of the treated unit with the control unit temporal effect:
- The difference between the two differences is the treatment effect

¹Good description: https://mixtape.scunning.com/09-difference_in_differences#john-snows-cholera-hypothesis

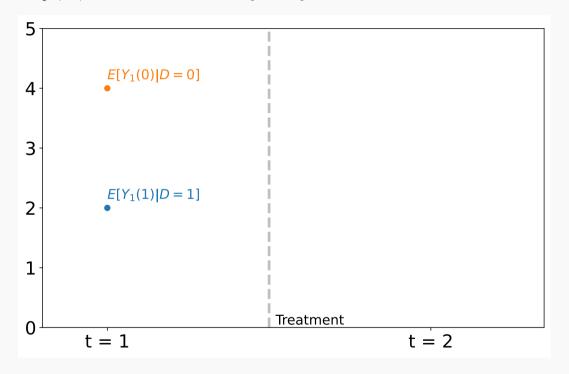
Two period of times: t=1, t=2



Potential outcomes: $Y_t(d)$ where $d=\{0,1\}$ is the treatment at period 2

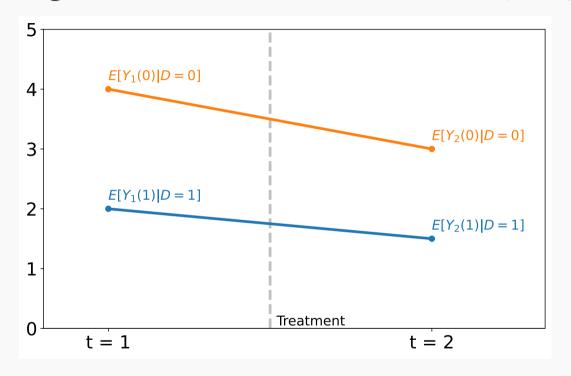


Potential outcomes: $Y_t(d)$ where $d = \{0, 1\}$ is the treatment at period 2



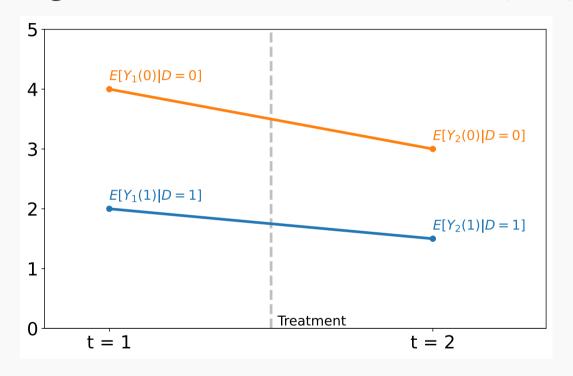
$$\mathbb{E}[Y_1(1)] = \mathbb{E}[Y_1(1) \mid D = 1] \mathbb{P}(D = 1) + \mathbb{E}[Y_1(1) \mid D = 0] \mathbb{P}(D = 0)$$
 but we only observe $\mathbb{E}[Y_1(1) \mid D = 1]$

Our target is the average treatment effect on the treated (ATT)



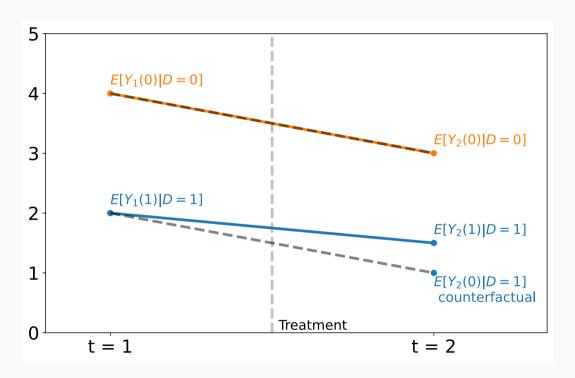
$$\tau_{\mathrm{ATT}} = \mathbb{E}[Y_2(1)|\ D=1] - \mathbb{E}[Y_2(0)|\ D=1]$$

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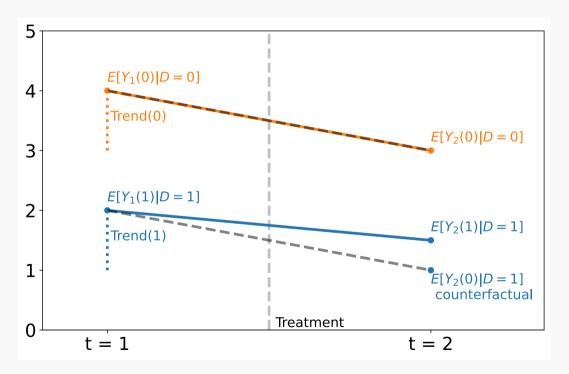


$$\tau_{\text{ATT}} = \underbrace{[Y_2(1)|\ D=1]}_{\text{treated outcome for t=2}} - \underbrace{\mathbb{E}[Y_2(0)|\ D=1]}_{\text{unobserved counterfactual}}$$

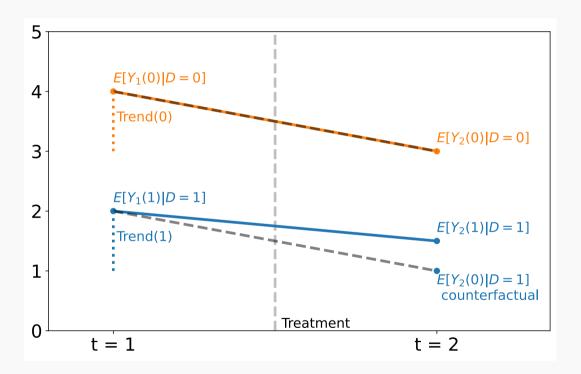
$$\mathbb{E}[Y_2(0) - Y_1(0) \mid D = 1] = \mathbb{E}[Y_2(0) - Y_1(0) \mid D = 0]$$



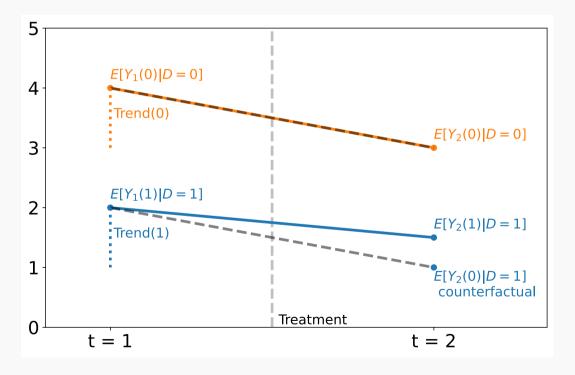
$$\underbrace{ \begin{bmatrix} Y_2(0) - Y_1(0) \mid D = 1 \end{bmatrix}}_{\mathbf{Trend}(1)} = \underbrace{ \mathbb{E}[Y_2(0) - Y_1(0) \mid D = 0]}_{\mathbf{Trend}(0)}$$



$$\mathbb{E}[Y_2(0) \mid D=1] = \mathbb{E}[Y_1(0) \mid D=1] + \mathbb{E}[Y_2(0) - Y_1(0) \mid D=0]$$

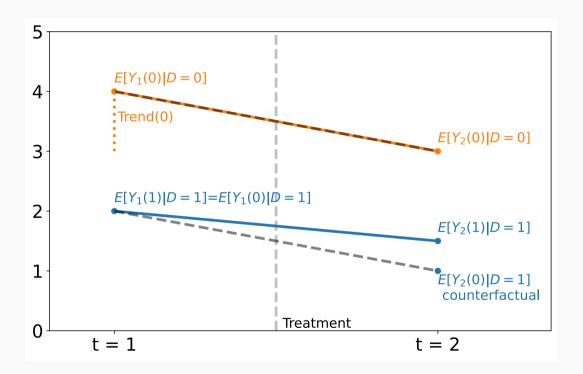


$$\mathbb{E}[Y_2(0) \mid D=1] = \underbrace{[Y_1(0) \mid D=1]}_{\text{unobserved counterfactual}} + \mathbb{E}[Y_2(0) - Y_1(0) \mid D=0]$$



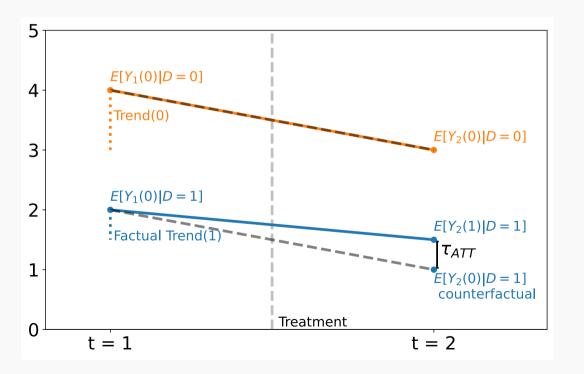
Second assumption, no anticipation of the treatment

$$E[Y_1(1)|D=1] = E[Y_1(0)|D=1]$$



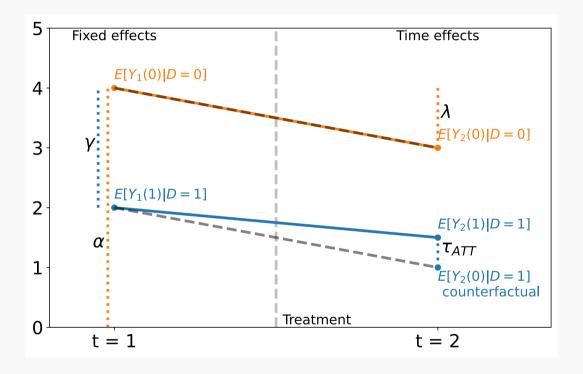
Difference-in-differences framework: identification of ATT

$$\begin{split} \tau_{\text{ATT}} &= \mathbb{E}[Y_2(1)|\ D=1] - \mathbb{E}[Y_2(0)|\ D=1] \\ &= \underbrace{\mathbb{E}[Y_2(1)|\ D=1] - \mathbb{E}[Y_1(0)|D=1]}_{\text{Factual Trend}(1)} - \underbrace{\mathbb{E}[Y_2(0)|D=0] - \mathbb{E}[Y_1(0)|D=0]}_{\text{Trend}(0)} \end{split}$$



Estimation: link with two way fixed effect (TWFE)

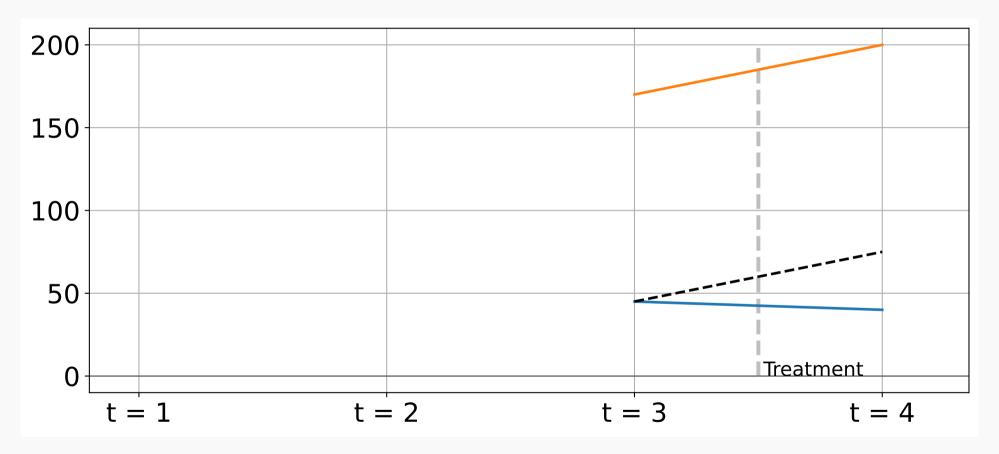
$$Y = \alpha + \gamma D + \lambda \mathbb{1}(t=2) + \tau_{\text{ATT}} D\mathbb{1}(t=2)$$



Mechanic link: works only under parallel trends and no anticipation assumptions.

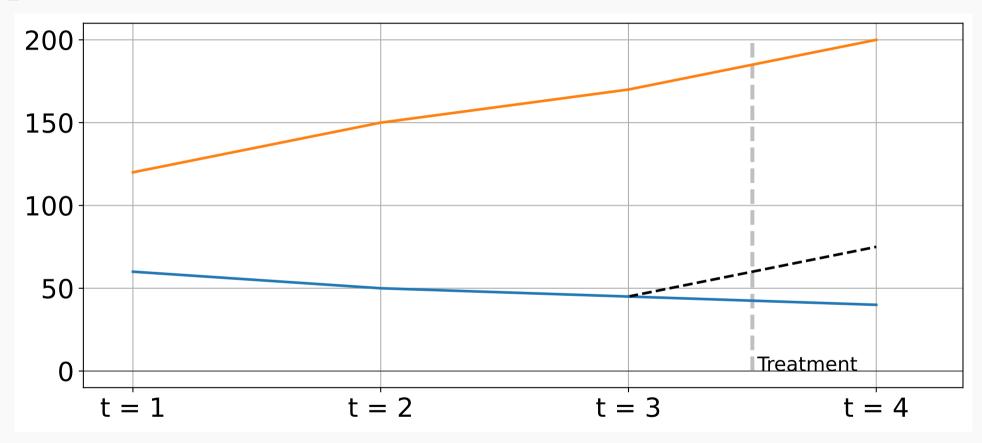
Failure of the parallel trend assumption

Seems like the treatment decreases the outcome!



Failure of the parallel trend assumption

Oups...



DID estimator for more than two time units

Target estimand: sample average treatment effect on the treated (SATT)

$$\tau_{\text{SATT}} = \frac{1}{|\{i:D_i=1\}|} \sum_{i:D_i=1}^{T} \frac{1}{T-H} \sum_{t=H+1}^{T} Y_{it}(1) - Y_{it}(0)$$

DID estimator

$$\begin{split} \widehat{\tau_{\text{DID}}} &= \frac{1}{|\{i:D_i=1\}|} \sum_{i:D_i=1} \left[\frac{1}{T-H} \sum_{t=H+1}^T Y_{it} - \frac{1}{H} \sum_{t=1}^H Y_{it} \right] - \\ &\frac{1}{|\{i:D_i=0\}|} \sum_{i:D_i=0} \left[\frac{1}{T-H} \sum_{t=H+1}^T Y_{it} - \frac{1}{H} \sum_{t=1}^H Y_{it} \right] \end{split}$$

Assumption

No anticipation of the treatment: $Y_{it}(0) = Y_{it}(1) \forall t = 1, ..., H$.

Parallel trend: $\mathbb{E}[Y_{it}(0,\infty)-Y_{i1}(0,\infty)]=\beta_t, t=2,...,T.$

See (Wager, 2024) for a clear proof of consistancy.

DID: Take-away

Pros

- Extremely common in economics and quite simple to implement.
- Can be extended to (Wager, 2024)
 - more than two time periods: exact same formulation
 - staggered adoption of the treatment: a bit more complex

Cons

- Very strong assumptions: parallel trends and no anticipation.
- Does not account for heterogeneity of treatment effect over time (De Chaisemartin & d'Haultfoeuille, 2020).

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Can we do better: ie. robust to the parallel trend assumption?

References

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Idea

Find a weighted average of controls that predicts well the treated unit outcome before treatment.

Example

What is the effect of tobacco tax on cigarettes sales? (Abadie et al., 2010)

Examples of application of synthetic controls to epidemiology

• What is the effect of taxes on sugar-based product consumption (Puig-Codina et al., 2021)

Context

1988: 25-cent tax per pack of cigarettes, ban of on cigarette vending machines in public areas accessible by juveniles, and a ban on the individual sale of single cigarettes.

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Time period: $t \in \{1, ...T\} = \{1970, ...2000\}$ and treatment time $T_0 = 1988$

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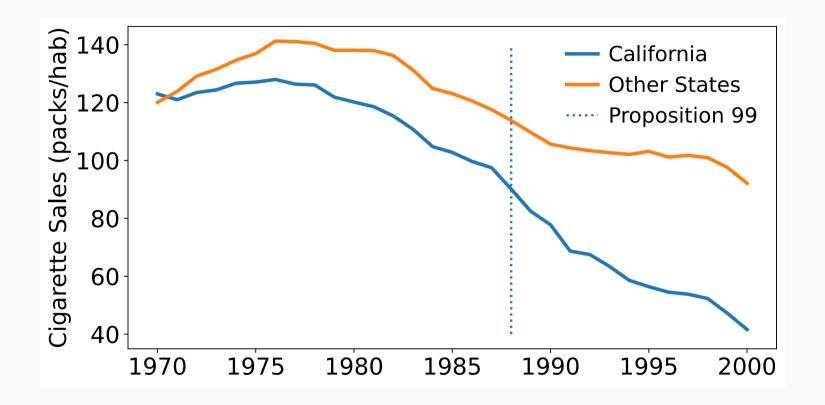
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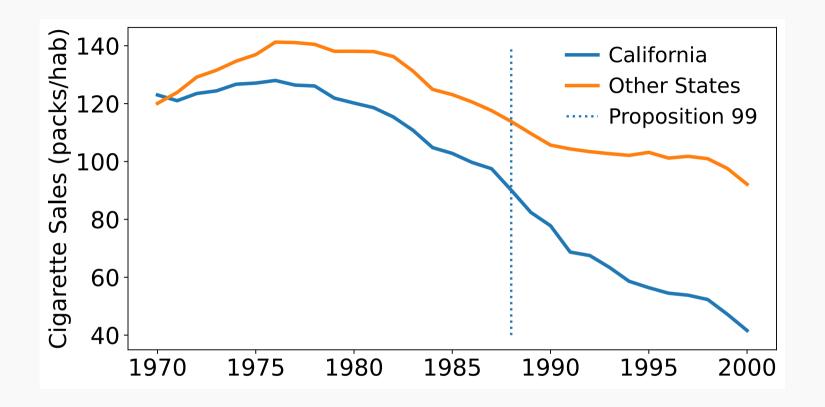
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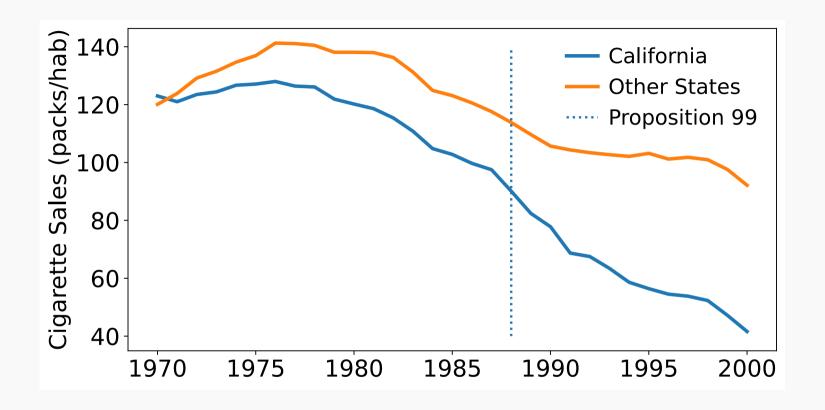
Time period: $t \in \{1, ...T\} = \{1970, ...2000\}$ and treatment time $T_0 = 1988$

Covariates $X_{j,t}$: cigarette price, previous cigarette sales.

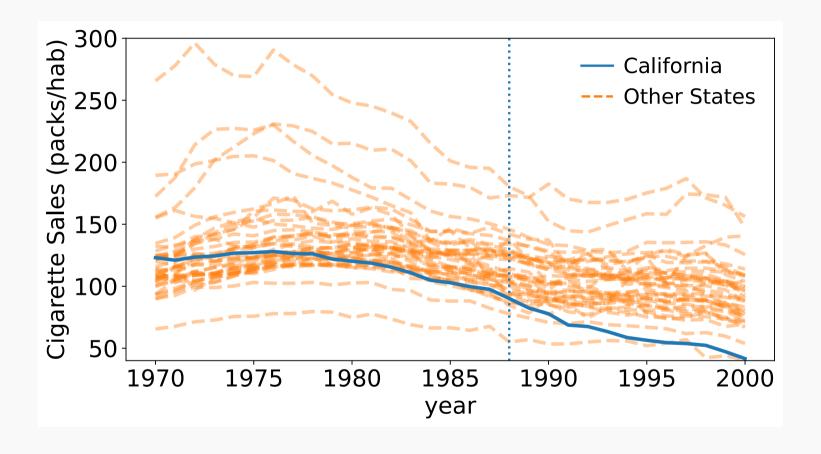


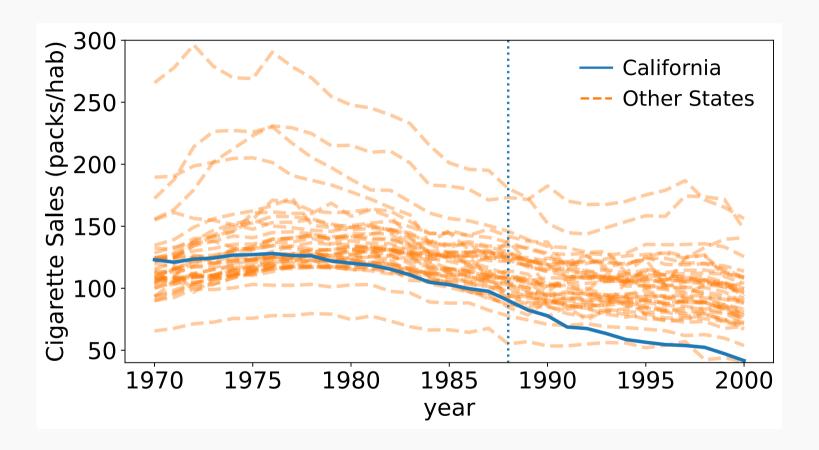


Pecrease in cigarette sales in California.



- Decrease in cigarette sales in California.
- Decrease began before the treatment and occured also for other states.

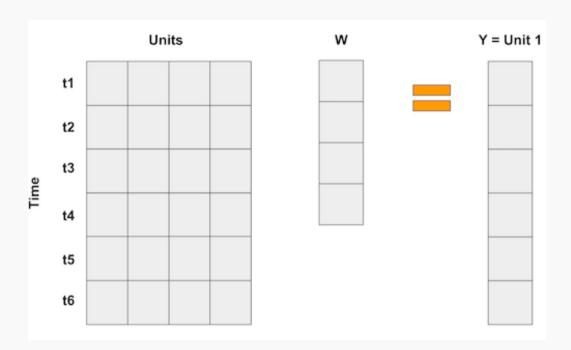




Force parallel trends: Find a weighted average of other states that predicts well the pre-treatment trend of California (before $T_0 = 1988$).

Build a predictor for $Y_{1,t}$ (California):

$$\hat{Y}_{1,t} = \sum_{j=2}^{n_0+1} \hat{w}_j Y_{j,t}$$

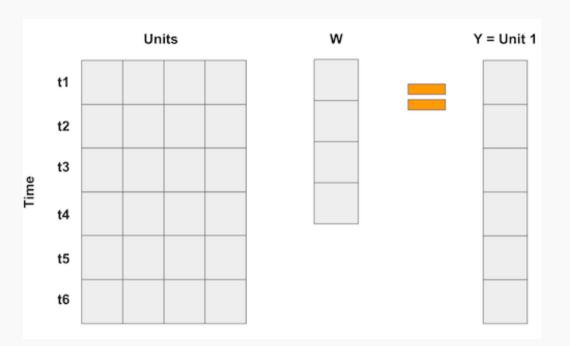


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Begin How to choose the weights?

Minimize some distance between the treated and the controls.

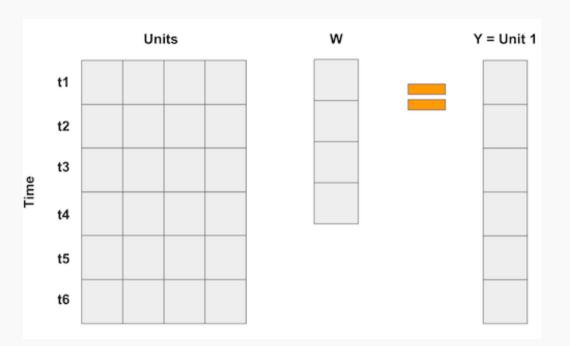


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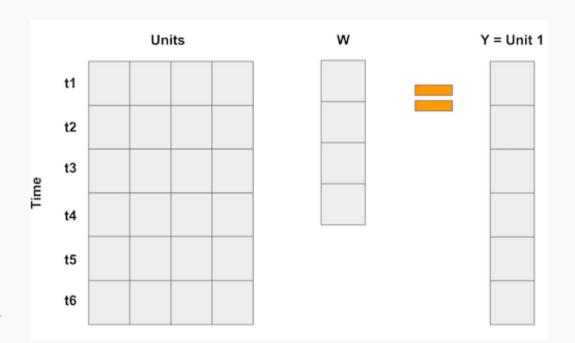
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How to choose the weights?

Minimize some distance between the treated and the controls.

This is called a balancing estimator: kind of Inverse Probability Weighting (Wager, 2024, chapter 7)



Characteristics

Pre-treatment characteristics concatenate pre-treatment outcomes and other pre-treatment predictors Z_1 eg. cigarette prices:

$$X_{ ext{treat}} = X_1 = \begin{pmatrix} Y_{1,1} \\ Y_{1,2} \\ & \ddots \\ & Y_{1,T_0} \\ & Z_1 \end{pmatrix} \in R^{p imes 1}$$

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Let the control pre-treatment characteristics be: $X_{\text{control}} = (X_2, ..., X_{n_0+1}) \in \mathbb{R}^{p \times n_0}$

Minimization problem

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$$w^* = \operatorname{argmin}_w \|X_{\operatorname{treat}} - X_{\operatorname{control}} w\|_V^2$$

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 where $\|X\|_V = \sqrt{X^T V X}$ with $V \in \operatorname{diag}(R^p)$

This gives more importance to some features than others.

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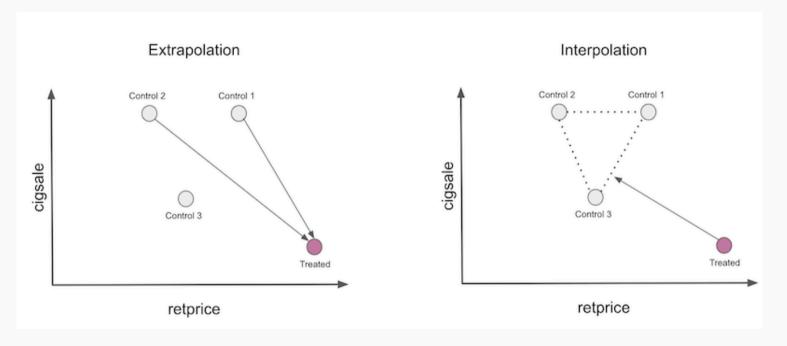
Minimization problem with constraints

$$\begin{split} w^* &= \operatorname{argmin}_w \ \|X_{\operatorname{treat}} - X_{\operatorname{control}} w\|_V^2 \\ s.t. \ w_j &\geq 0, \\ \sum_{j=2}^{n_0+1} w_j &= 1 \end{split}$$

Synthetic controls: Why choose positive weights summing to one?

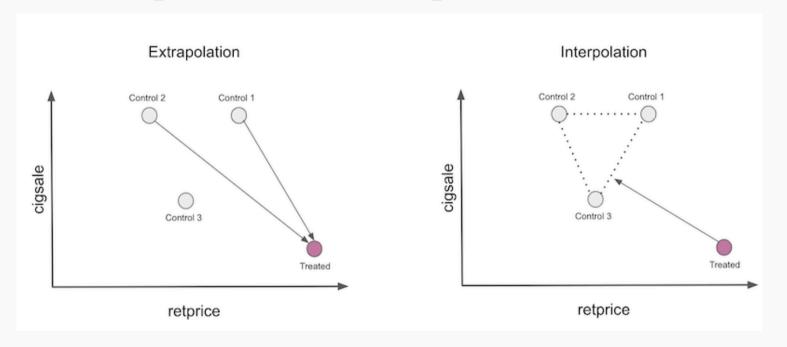
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Interpolation enforces regularization, thus limits overfitting

Same kind of regularization than L1 norm in Lasso: forces some coefficient to be zero.

 $p = 2T_0$ covariates:

$$X_{j} = \begin{pmatrix} Y_{j,1} \\ .. \\ Y_{j,T_{0}} \\ Z_{j,1} \\ .. \\ Z_{j,T_{0}} \end{pmatrix}^{T} \in R^{2T_{0}}$$

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Y cigarette sales, Z cigarette prices.

Model:
$$\underbrace{X_{\text{treat}}}_{p \times 1} \sim \underbrace{X_{\text{control}}}_{p \times n_0} \underbrace{w}_{n_0}$$

-> simple linear regression estimated by OLS

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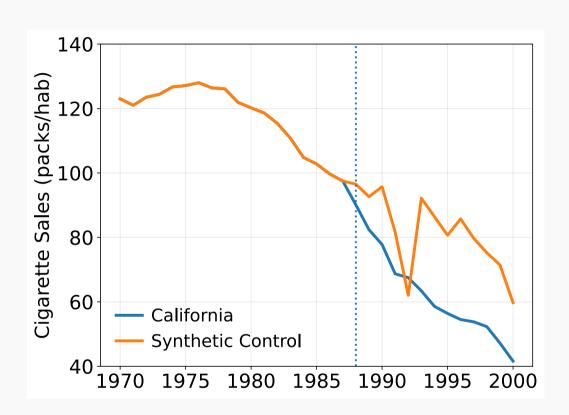
Prediction:
$$\hat{Y}_{\text{synth}} = (Y_{t,j})_{\substack{t=1..T \ j=2..n_0+1}} w$$

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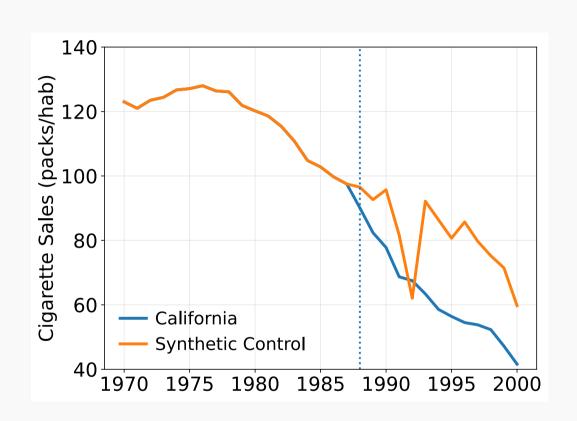


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Synthetic controls: How to choose the predictor weights V?

- 1. Don't choose: set $V = I_p$, ie. $||X||_V = ||X||_2$.
- 2. Rescale by the variance of the predictors:

$$V = \operatorname{diag}\left(\operatorname{var}(Y_{j,1})^{-1}, ..., \operatorname{var}(Y_{j,T_0})^{-1}, \operatorname{var}(Z_{j,1})^{-1}, ..., \operatorname{var}(Z_{j,T_0})^{-1}\right).$$

3. Minimize the pre-treatment mean squared prediction error (MSPE) of the treated unit:

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3. Minimize the pre-treatment mean squared prediction error (MSPE) of the treated unit:

$$\begin{split} \text{MSPE}(V) &= \sum_{t=1}^{T_0} \left[Y_{1,t} - \sum_{j=2}^{n_0+1} w_j^*(V) Y_{j,t} \right]^2 \\ &= \left\| \ \left(Y_{1,t} \right)_{t=1..T_0} - \left(Y_{j,t} \right)_{\substack{j=2..n_0+1 \\ t=1..T_0}}^T \hat{w} \ \right\|_2^2 \end{split}$$

This solution is solved by running two optimization problems:

- Inner loop solving $w^*(V) = \operatorname{argmin}_w \|X_{\operatorname{treat}} X_{\operatorname{control}} w\|_V^2$
- Outer loop solving $V^* = \operatorname{argmin}_V \operatorname{MSPE}(V)$

Synthetic controls: estimation without the outer optimization problem

Same coviarates:
$$X_j = \begin{pmatrix} Y_{j,1} \\ .. \\ Y_{j,T_0} \\ Z_{j,1} \\ .. \\ Z_{j,T_0} \end{pmatrix}^T$$

SCM minization with
$$V=I_p$$
, hence,
$$\|X\|_V=\|X\|_2.$$

$$\begin{split} w^* &= \operatorname{argmin}_w \ \|X_{\operatorname{treat}} - X_{\operatorname{control}} w\|_2^2 \\ s.t. \ w_j &\geq 0, \\ \sum_{j=2}^{n_0+1} w_j &= 1 \end{split}$$

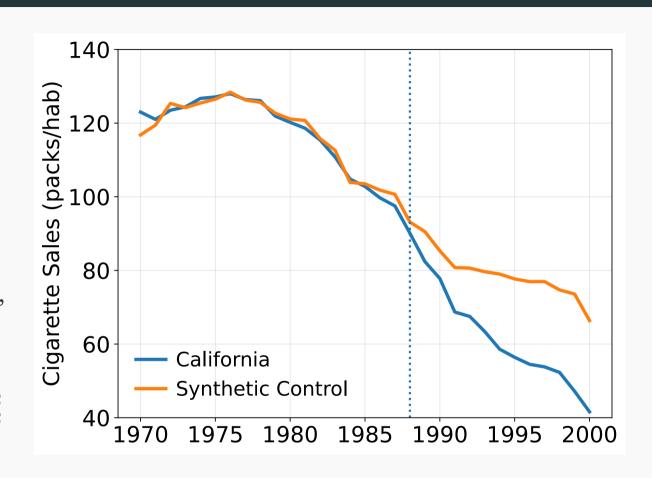
Synthetic controls: estimation without the outer optimization problem

Synthetic controls: estimation with a same coviarates:
$$X_j = \begin{pmatrix} Y_{j,1} \\ ... \\ Y_{j,T_0} \\ Z_{j,1} \\ ... \\ Z_{j,T_0} \end{pmatrix}^T$$

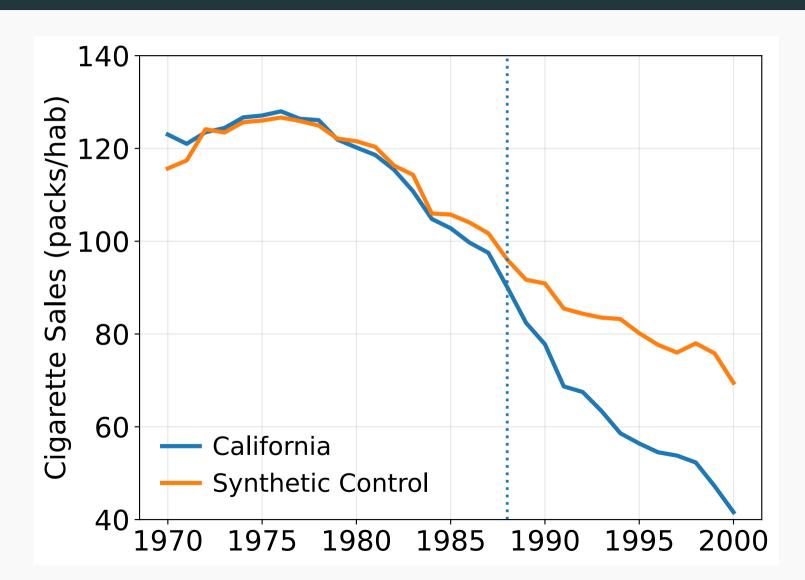
Y cigarette sales, Z cigarette prices.

SCM minization with $V=I_p$, hence, $\|X\|_V=\|X\|_2.$

$$\begin{split} w^* &= \operatorname{argmin}_w \ \|X_{\operatorname{treat}} - X_{\operatorname{control}} w\|_2^2 \\ s.t. \ w_j &\geq 0, \\ \sum_{j=2}^{n_0+1} w_j &= 1 \end{split}$$



Synthetic controls: estimation with the outer optimization problem



Synthetic controls: inference

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... but from the variability of the chosen control units

Treatment assignment introduces more noise than outcome variability.

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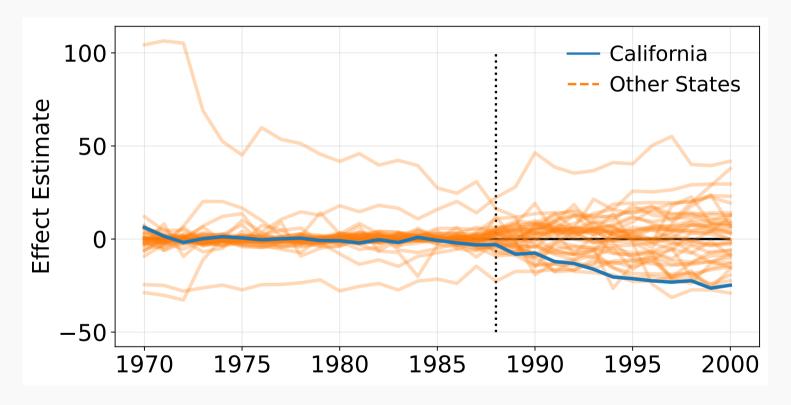
Treatment assignment introduces more noise than outcome variability.

(Abadie et al., 2010) introduced the placebo test to assess the variability of the synthetic control.

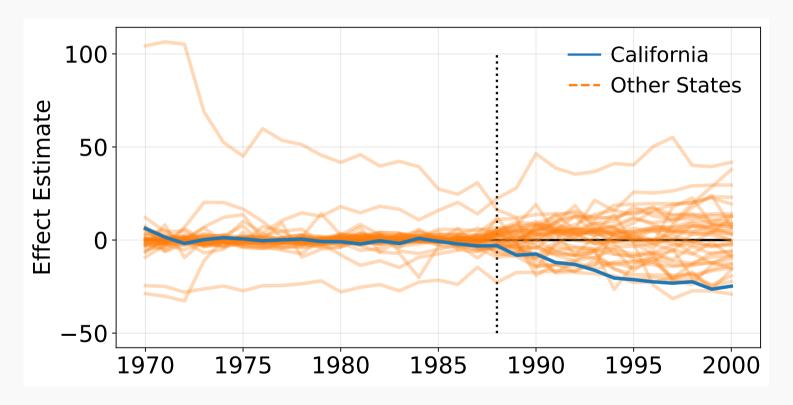
Idea of Fisher's Exact tests

- Permute the treated and control exhaustively.
- For each unit, we pretend it is the treated while the others are the control: we call it a placebo
- Compute the synthetic control for each placebo: it should be close to zero.

Placebo estimation for all 38 control states

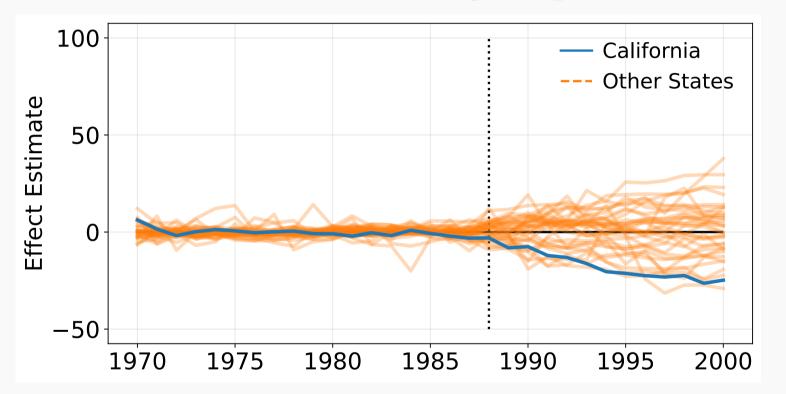


Placebo estimation for all 38 control states



- More variance after the treatment for California than before.
- Some states have pre-treatment trends which are hard to predict.

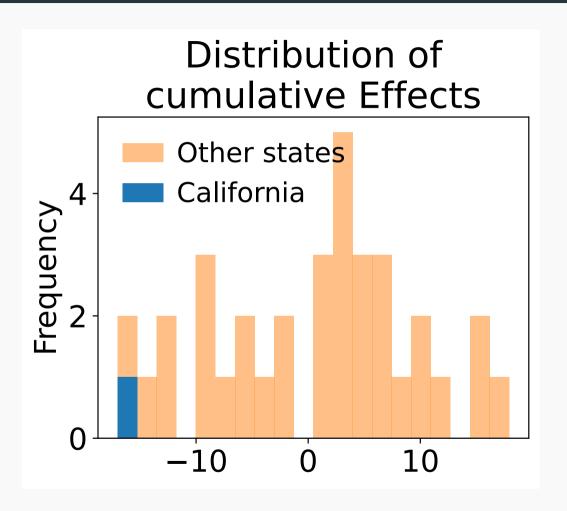
Placebo estimation for 34 control states with "good" pre-treatment fit



I removed the states above the 90 percentiles of the distribution of the pre-treatment fit.

California absolute cumulative effect

$$\hat{\tau}_{\text{scm, california}} = -17.00$$

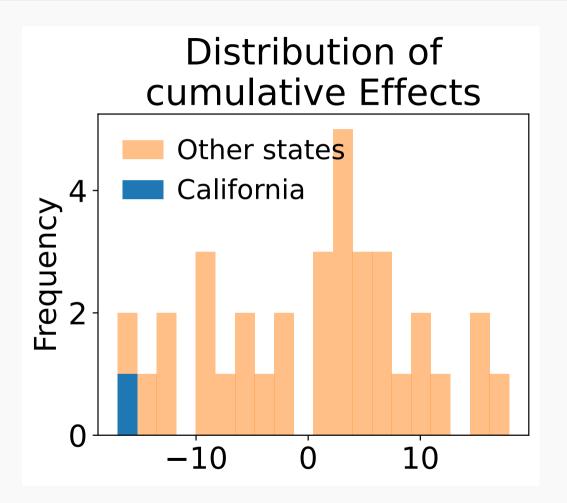


California absolute cumulative effect

$$\hat{\tau}_{\text{scm, california}} = -17.00$$

Get a p-value

$$PV = \frac{1}{n_0} \sum_{j=2}^{n_0} \mathbb{1}(|\hat{\tau}_{\text{scm, california}}| > |\hat{\tau}_{\text{scm,}j}|)$$
$$= 0.029$$



Synthetic controls: inference with conformal prediction

Synthetic controls: Take-away

Pros

- More convincing for parallel trends assumption.
- Simple for multiple time periods.
- Gives confidence intervals.

Cons

- Requires many control units to yield good pre-treatment fits.
- Might be prone to overfitting during the pre-treatment period.
- Still requires a strong assumption: the weights should also balance the post-treatment unexposed outcomes. See (Arkhangelsky et al., 2021) for discussions.
- Still requires the no-anticipation assumption.

Conditional difference-in-differences

Time-series modelisation: methods without a control group

Interrupted Time Series: intuition

Setup

- One treated unit, no control unit.
- Multiple time periods.
- Sometimes, predictors are availables: there are called exogeneous covariates.

Interrupted Time Series: intuition

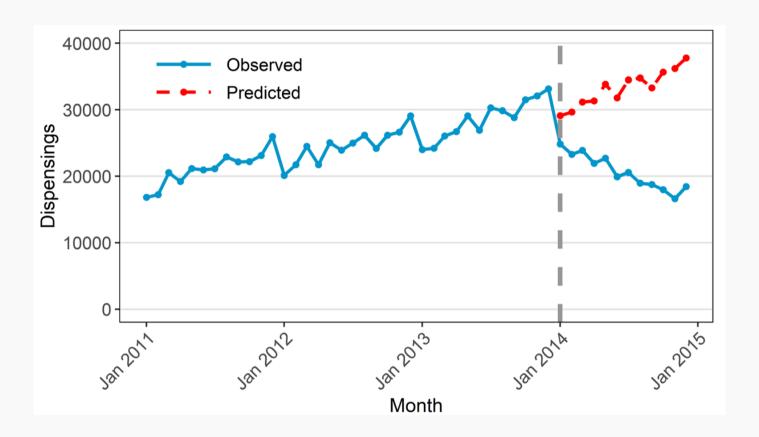
Setup

- One treated unit, no control unit.
- Multiple time periods.
- Sometimes, predictors are availables: there are called exogeneous covariates.

Intuition

- Model the pre-treatment trend: $Y_{t(1)}$ for $t < T_0$
- Predict post-treatment trend as the control: $\hat{Y}_t(0)$ for $t > T_0$
- Obtain treatment effect by taking the difference between observed and predicted post-treatment observations: $Y_t(1) \hat{Y}_t(0)$

Interrupted Time Series: illustration from (Schaffer et al., 2021)



 Y_t : Dispensations of quetiapine, an anti-psychotic medicine.

Treatment: Restriction of the conditions under which quetiapine could be subsidised.

Modelization of a time-series

Tools

• ARIMA models: AutoRegressive Integrated Moving Average

Motivation of ARIMA

Take into account:

- the structure of autodependance between observation,
- linear trends,
- seasonality.

A good reference for ARIMA: Forecasting: Principles and Practice, chapter 8

ARIMA are State Space Models (SSM) says the machine learning community

Why showing this formulation?

- I better understand ARIMA formulated as state space models.
- SSM are more general than ARIMA models.
- ARIMA are (almost always) fitted with SSM optimization algorithms.

What is a state space model?

We saw ARIMA models and the more general class of state space models.

However, we could any model that we want to fit the pre-treatment trend!

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- Facebook prophet model (Taylor & Letham, 2018) uses Generalized Additive Models (GAM).
- Any sklearn estimator could do the trick: Linear regression, Random Forest, Gradient Boosting...

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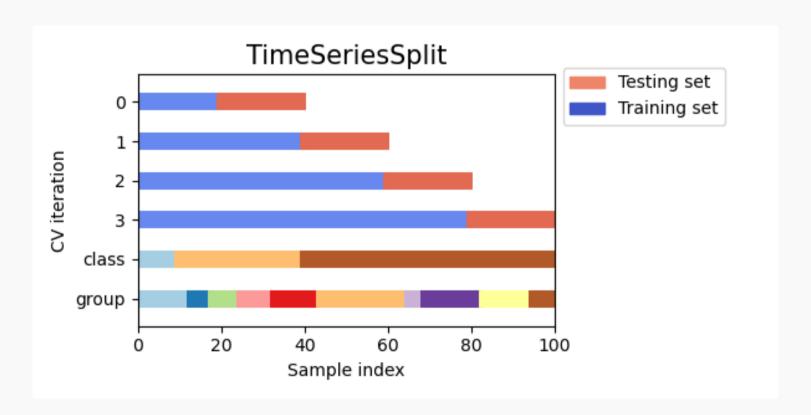
You should pay attention to appropriate train/test split when cross-validating a time-series model not to use the future to predict the past.

Relevant remark for all time series models (even ARIMA or state space models).

Cross-validation for time-series models

1 from sklearn.model_selection import TimeSeriesSplit

python



This avoids to use the future to predict the past.

Take-away on ITS

Pros

- Suitable when no control unit is available. The pre-treatment trend is the control.
- Handles multiple time periods.
- A lot of software available (eg. ARIMA models).
- Simple: few parameters to tune.

Cons

- Prone to bias by other events happening around the treatment time and impacting the outcome trend.
- Prone to overfitting of the pre-treatment trend.

An attempt to map event study methods

Methods	Characteristics	Hypotheses	Community	Introduction	Good reference
DID/TWFE	Treated/control	Parallel trends, no	Economics	Causal Inference for	(Arkhangelsky &
	units, few time peri-	anticipation, prone		the Brave and True,	Imbens, 2024)
	ods, no predictors	to overfitting		chapter 13	
ARIMA, ITS	No controls, no/few	Stationnarity , no	Epidemiology, Eco-	Forecasting:	(Schaffer et al., 2021)
	predictors, seasonal-	anticipation, prone	nomics	Principles and	
	ity	to overfitting		Practice	
State space models	Multiple time peri-	Contional ignorabil-	Machine learning,	(Brockwell & Davis,	(Murphy, 2022,
	ods, control units or	ity on predictors,	bayesian methods	2016, chapter 9)	chapter 18)
	predictors, general-	goodness of fit pre-			
	ization of ARIMA	treatment			
Synthetic control	Treated/control	Conditional parallel	Economics	Causal Inference for	(Abadie, 2021)
	units, multiple time	trend on controls,		the Brave and True	
	periods	goodness of fit pre-			
		treatment			

A summary on R packages for event studies

Package name	Methods	Predictors	Control units	Multiple time periods
did	Difference-in-differ-	X	X	X
	ences			
forecast	ARIMA, ITS	V	X	V
Synth	Synthetic control	X	V	
Causal impact	Bayesian state space	V	X	V
	models			

A summary on Python packages for event studies

Package name	Methods	Predictors	Control units	Multiple time periods
statsmodels.OLS	Difference-in-differences,	X	×	×
	TWFE			
statsmodels	ARIMA(X), ITS, bayesian		×	
	state space models			
pmdarima	ARIMA(X), ITS	V	X	V
SyntheticControlMethods	Synthetic control	X	V	~
pysyncon	Synthetic control	X	V	
causalimpact (pymc	Bayesian state space models	V	X	
implementation)				
causal-impact (statsmodels	Bayesian state space models	V	X	
implementation)				

Python hands-on

To your notebooks 🎑!

• url: https://github.com/strayMat/causal-ml-course/tree/main/notebooks

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