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## 1 Regression Spline

Assume that the range of  $x$  is  $[a, b]$ . Let the point

$$a < \xi_1 < \cdots < \xi_K < b$$

be a partition of the interval  $[a, b]$   $\{\xi_1, \dots, \xi_K\}$  are called knots.

### 1.1 Radial Basis Function

A RBF  $\varphi$  is a real valued function whose value depends only on the distance from origin. A real function  $\varphi : [0, \infty) \rightarrow \mathbb{R}$  with a metric on space  $\|\cdot\| : V \rightarrow [0, \infty)$  a function  $\varphi_c = \varphi(\|\mathbf{x} - \mathbf{c}\|)$  is said to be a radial kernel centered at  $c$ . A radial function and the associated radial kernels are said to be radial basis function

We use radial basis functions defined by

$$\mathbf{b}(u) = \left\{ u, \left| \frac{u - \tau_1}{c} \right|^3, \dots, \left| \frac{u - \tau_K}{c} \right|^3 \right\}$$

where  $c$  is sample standard deviation

## 2 Simulation

Let

$$y = \sum_{l=1}^4 f_l(X_l) + \sum_{k=1}^4 Z_k \theta_k + e$$

$$f_1(x) = 3\exp(-30(x - 0.3)^2) + \exp(-50(x - 0.7)^2)$$

$$f_2(x) = \sin(2\pi x)$$

$$f_3(x) = x$$

$$f_4(x) = 0$$

$$\theta_1 = 0.6$$

$$\theta_2 = -1$$

$$\theta_3 = \theta_4 = 0$$

Make spline and centerize the data we can get  $\tilde{y}$

$$\tilde{y} = y - \bar{y} = b_1(X_1)\beta_1 + b_2(X_2)\beta_2 + b_3(X_3)\beta_3 + b_4(X_4)\beta_4 + \sum_{k=1}^4 Z_k\theta_k + e$$

## 2.1 MFVB method

Setting prior as

$$\begin{aligned} Y|\tau, \beta &\sim N(X\beta, \sigma^2 \cdot I_N) \\ \beta_i|\gamma_i &\sim^{ind} N(0, \sigma_\beta^2) \text{ for } i = 1, \dots, p \\ \sigma_\beta^2 &\sim \text{Inverse} - \text{Gamma}(a, b) \\ \sigma^2 &\sim \text{Gamma}(c, d) \end{aligned}$$

By Baye's rule

$$p(\tau, \gamma, \beta|Y) \propto p(Y|\tau, \beta)p(\beta|\gamma)p(\tau)p(\gamma)$$

Then variational distribution is

$$p(\tau, \gamma, \mu|Y) \approx q(\tau, \gamma, \mu) = q_1(\tau)q_2(\gamma)q_3(\mu)$$

we can maximize ELBO by coordinate descent algorithm

$$\begin{aligned} q_1^*(\sigma^2) &= E_{q_2, q_3}[p(\sigma^2, \gamma, \beta|Y)] \propto E_{q_2, q_3}[p(Y|\sigma^2, \beta)p(\tau)] \\ q_2^*(\sigma_\beta^2) &= E_{q_1, q_3}[p(\tau, \sigma_\beta^2, \beta|Y)] \propto E_{q_1, q_3}[p(\beta|\sigma_\beta^2)p(\sigma_\beta^2)] \\ q_3^*(\beta) &= E_{q_1, q_2}[p(\tau, \gamma, \beta|Y)] \propto E_{q_1, q_2}[p(Y|\tau, \beta)p(\beta|\gamma)] \end{aligned}$$

Then

$$\begin{aligned} q_1^* &\sim \text{Gamma}\left(c + \frac{N+1}{2}, d + \frac{1}{2}\{Y'Y - E_{q_3}[\beta'](X'Y)\} + \text{tr}[X(\text{var}_{q_3}[\beta] + E_{q_3}[\beta]E_{q_3}[\beta'])X']\right) \\ q_2^* &\sim \prod_{i=1}^p \text{Inverse} - \text{Gamma}\left(a + \frac{1}{2}, b + \frac{1}{2}\{\text{var}_{q_3}[\beta]_{i,i} + E_{q_3}[\beta_i]^2\}\right) \\ q_3^* &\sim N\left(E_{q_1}[\sigma^2]\Sigma X'Y, (\text{diag}(E_{q_2}[\sigma_\beta^2]) + E_{q_1}[\sigma^2]X'X)^{-1} = \Sigma\right) \end{aligned}$$

## 2.2 MFVB method with variable selection

Variable selection model is

$$Y = X\Gamma\beta + \epsilon, \quad \epsilon \sim N(0, \sigma^2 I)$$

where

$$Y|\beta, \sigma^2, \gamma \sim N(X\Gamma\beta, \sigma^2 I)$$

$$\sigma^2 \sim \text{Inverse} - \text{Gamma}(A, B) \quad A = 0, B = 0$$

$$\beta_j \sim N(0, \sigma_\beta^2)$$

$$\gamma_j \sim \text{Bernoulli}(\rho) \quad \rho = 0.5$$

## 2.3 MFVB method with variable selection hierarchical model

Variable selection hierarchical model is

$$Y = X\Gamma\beta + \epsilon, \quad \epsilon \sim N(0, \sigma^2 I)$$

where

$$Y|\beta, \sigma^2, \gamma \sim N(Z\Gamma\beta, \sigma^2 I)$$

$$\sigma^2 \sim \text{Inverse} - \text{Gamma}(A, B) \quad A = 0, B = 0$$

$$\beta_j \sim N(0, \sigma_\beta^2)$$

$$\gamma_j \sim \text{Bernoulli}(\rho)$$

$$\rho \sim \text{Beta}(C, D) \quad C = 1, D = 1.4$$

$$p(\beta, \sigma^2, \sigma_\beta^2, \Gamma|Y) \approx q(\beta, \sigma^2, \sigma_\beta^2, \Gamma) = \prod_{j=1}^p q_1^*(\beta_j) q_2^*(\sigma^2) q_3^*(\rho) \prod_{j=1}^p q_4^*(\gamma_j)$$

Use coordinate ascent algorithm,  $q$  density of  $\beta$  is

$$\begin{aligned}
q_1^*(\beta) &\propto E_{-q_1} [p(\beta, \sigma^2, \sigma_\beta^2, \Gamma, \rho, Y)] \\
&\propto E_{-q_1} \left[ \exp \left( -\frac{1}{2\sigma^2} (Y - Z\Gamma\beta)' (Y - Z\Gamma\beta) - \frac{1}{2} \sum_{j=1}^p \frac{\beta_j^2}{\sigma_\beta^2} \right) \right] \\
&\propto \exp \left( -\frac{1}{2} \beta' \langle D \rangle - \frac{1}{2} \left\langle \frac{1}{\sigma^2} \right\rangle (\beta' \langle \Gamma \rangle' Z' Z \langle \Gamma \rangle \beta - 2\beta' \langle \Gamma \rangle' Z' Y) \right) \\
&\propto \exp \left( -\frac{1}{2} \left[ \beta' \left( \langle D \rangle + \left\langle \frac{1}{\sigma^2} \right\rangle \langle \Gamma \rangle' Z' Z \langle \Gamma \rangle \right) \beta - 2 \left\langle \frac{1}{\sigma^2} \right\rangle \beta' \langle \Gamma \rangle' Z' Y \right] \right) \\
&\sim N(\mu, \Sigma)
\end{aligned}$$

Where  $D = \text{diag}(\frac{1}{\sigma_\beta^2})$ ,  $\langle \cdot \rangle$  means expectation under  $q$  functions and

$$\Sigma = \left( \langle D \rangle + \beta' \left\langle \frac{1}{\sigma^2} \right\rangle \langle \Gamma \rangle' Z' Z \langle \Gamma \rangle \right)^{-1}, \quad \mu = \left\langle \frac{1}{\sigma^2} \right\rangle \Sigma \langle \Gamma \rangle' Z' Y$$

$q$  density of  $\sigma^2$  is

$$\begin{aligned}
q_2^*(\sigma^2) &\propto E_{-q_2} [p(\beta, \sigma^2, \sigma_\beta^2, \Gamma, \rho, Y)] \\
&\propto E_{-q_2} \left[ (\sigma^2)^{-(\frac{n}{2}+a)-1} \exp \left( -\frac{1}{\sigma^2} (b + (Y - Z\Gamma\beta)' (Y - Z\Gamma\beta)) \right) \right] \\
&\propto \text{Inverse-Gamma} \left( a + \frac{n}{2}, b + \frac{1}{2} (Y'Y - 2\langle \beta \rangle' \langle \Gamma \rangle' Z' Y + \text{tr}((Z'Z \odot \Omega)(\mu\mu' + \Sigma))) \right)
\end{aligned}$$

Where  $\odot$  is hadamard product and

- $\gamma = (\gamma_1, \dots, \gamma_p)$
- $\Omega = \langle \gamma \rangle \langle \gamma \rangle' + \langle \Gamma \rangle \odot (I - \langle \Gamma \rangle)$

$q$  density of  $\rho$  is

$$\begin{aligned}
q_3^*(\rho) &\propto E_{-q_3} [p(\beta, \sigma^2, \sigma_\beta^2, \Gamma, \rho, Y)] \\
&\propto \rho^{C-1} (1-\rho)^{D-1} \prod_{j=1}^p \rho^{\gamma_j} (1-\rho)^{1-\gamma_j} \\
&\propto \rho^{C+\sum_{j=1}^p \gamma_j - 1} (1-\rho)^{D+p-\sum_{j=1}^p \gamma_j - 1} \\
&\sim \text{Beta}(C + \sum_{j=1}^p \gamma_j, D + p - \sum_{j=1}^p \gamma_j)
\end{aligned}$$

$q$  density of  $\gamma$  is

$$q_4^*(\gamma)E_{-q_4} \left[ \prod_{j=1}^p \rho^{\gamma_j} (1-\rho)^{-\gamma_j} \exp\left(-\frac{1}{2} \frac{1}{\sigma^2} (\beta' \Gamma' Z' Z \Gamma \beta - 2\beta' \Gamma' Z' y)\right) \right] \\ \propto \exp \left[ \gamma_j \left( \langle \log(\rho/(1-\rho)) \rangle - \frac{1}{2} \left\langle \frac{1}{\sigma^2} \right\rangle \langle \beta_j^2 \rangle Z_j' Z_j + \left\langle \frac{1}{\sigma^2} \right\rangle Z_j' [Y \mu_j - X_{-j} \langle \Gamma_{-j} \rangle (\mu_{-j} \mu_j + \Sigma_{-j,j})] \right) \right]$$

Where

- $X_j$  means  $j$ th coloumn of  $X$
- $X_{-j}$  means without  $j$ th column
- $X_{-i,j}$  means  $j$ th column without  $i$ th component
- $\mu_j$  is  $j$ th component of vector and  $\mu_{-j}$  means without  $j$ th component

### 2.3.1 Posterior of $\rho$

$$f_1(x) = 3\exp(-30(x-0.3)^2) + \exp(-50(x-0.7)^2)$$

$$f_2(x) = \sin(2\pi x)$$

$$f_3(x) = x$$

$$f_4(x) = 0$$

$\rho \sim$	$Beta(1, 1.4)$	$Beta(1, 1)$	$Beta(1, 2)$	$Beta(2, 1)$	$Beta(2, 2)$
f1 $q^*(\rho) \sim$	beta(30.0 2.4)	beta(31.0 1.0)	beta(4.0, 29.0)	beta(32.0 1.0)	beta(22.8,10.2)
f2 $q^*(\rho) \sim$	beta(5.0 27.4)	beta(31.0 1.0)	beta(3.0, 30.0)	beta(32.0 1.0)	beta(26.9 6.1)
f3 $q^*(\rho) \sim$	beta( 4.0, 28.4)	beta(31.0 1.0)	beta(4.0, 30.0)	beta(32.0 1.0)	beta(26.9 6.1)
f4 $q^*(\rho) \sim$	beta(1.0 31.4)	beta(1.0 31.0)	beta(1.0 32.0)	beta(32.0 1.0)	beta(30.0 3.0)

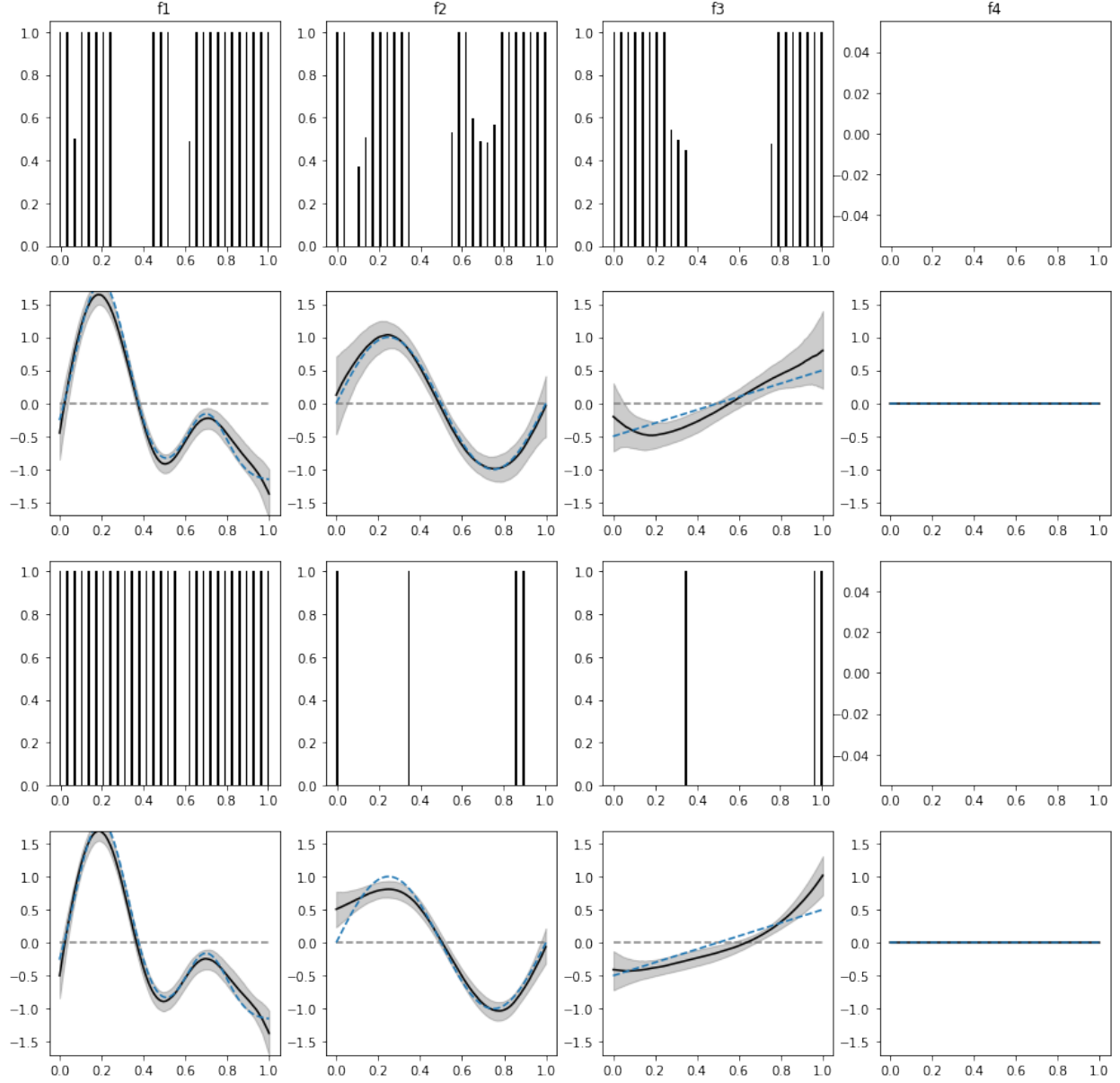


Figure 1: Compare the variable selection model without Beta prior and without Beta prior. First and Second row has plat prior  $\rho = 0.5$ . Third and Forth row has beta prior  $\rho \sim \text{Beta}(1, 1.4)$