0725_code

July 25, 2019

First import required modules

```
In [1]: import numpy as np
    import pandas as pd
    import matplotlib.pyplot as plt
    import scipy.stats as stats
    from scipy.special import digamma
    from scipy.linalg import sqrtm
```

1 Regression Spline

Assume that the range of x is [a, b]. Let the point

$$a < \xi_1 < \cdots < \xi_K < b$$

be a partion of the interval [a, b] $\{\xi_1, \dots, \xi_K\}$ are called knots.

Then make the function which return the knot points

2 Radial Basis Function

A RBF φ is a real valued function whose value depends only on the distance from origin. A real function $\varphi: [0,\infty) \to \mathbb{R}$ with a metric on space $\|\cdot\|: V \to [0,\infty)$ a function $\varphi_c = \varphi(\|\mathbf{x} - \mathbf{c}\|)$ is said to be a radial kernel centered at c. A radial function and the associated radial kernels are said to be radial basis function

we use radial basis functions defined by

$$\mathbf{b}(u) = \left\{ u, \left| \frac{u - \tau_1}{c} \right|^3, \cdots, \left| \frac{u - \tau_K}{c} \right|^3 \right\}$$

where c is sample standard deviation

Then we can make the function which retrun the basis

Nonparametric linear model can be represented as

$$Y = \mathbf{b}(X)\boldsymbol{\beta} + \varepsilon$$

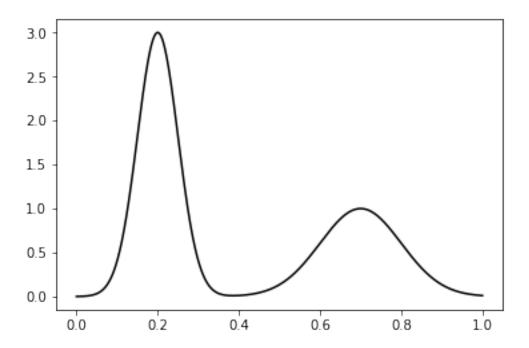
where $Y \in \mathbb{R}^{n \times 1}$, $X \in \mathbb{R}^{n \times 1}$ and $\varepsilon \sim N(0, \tau^{-1})$

3 Make toy data

Let

$$y = 3 \exp(-200(x - 0.2)^2) + \exp(-50(x - 0.7)^2)$$

Plotting true distribution of *Y* is

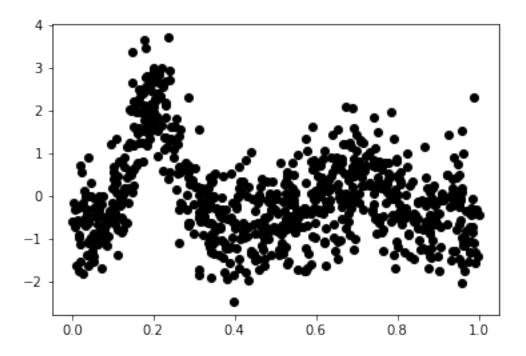


make the simulation function which make the obs with error N(0, 0.5)

Plotting the distribution of simulated data

$$y = 3 \exp(-200(x - 0.2)^2) + \exp(-50(x - 0.7)^2) + \varepsilon$$

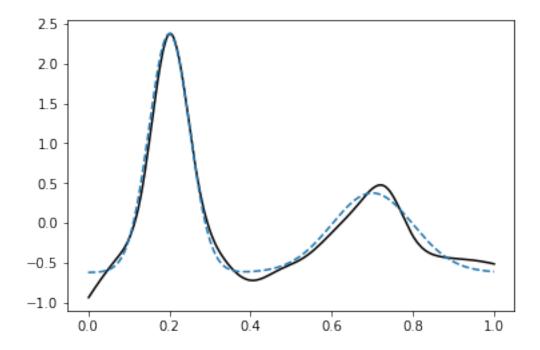
where $\varepsilon \sim N(0, 0.5)$



Calculate the standard deviation of observed data and define the knot and make design matrix

4 LSE method

plotting the fitted value



Blue dashed line is true function and solid line is LSE estimated function

5 MFVB method

setting prior as

$$p(Y|\tau, \beta) \sim N(X\beta, \tau^{-1} \cdot I_N)$$

 $p(\beta_i|\gamma_i) \sim^{ind} N(0, \gamma^{-1}) \text{ for } i = 1, \dots p$
 $p(\gamma) \sim Gamma(a, b)$
 $p(\tau) \sim Gamma(c, d)$

By Baye's rule

$$p(\tau,\gamma,\beta|Y) \propto p(Y|\tau,\beta)p(\beta|\gamma)p(\tau)p(\gamma)$$

Then variational distribution is

$$p(\tau,\gamma,\mu|Y)\approx q(\tau,\gamma,\mu)=q_1(\tau)q_2(\gamma)q_3(\mu)$$

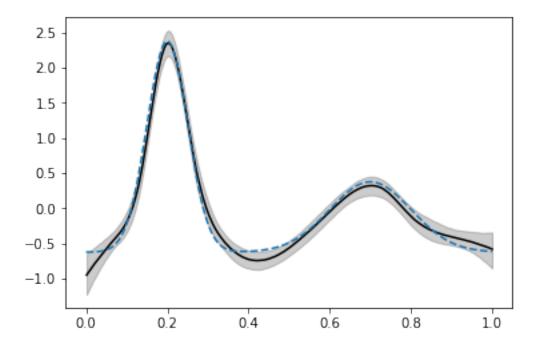
we can maximize ELBO by coordinate descent algorithm

$$\begin{split} q_1^*(\tau) &= E_{q_2,q_3}[p(\tau,\gamma,\beta|Y)] \propto E_{q_2,q_3}[p(Y|\tau,\beta)p(\tau)] \\ q_2^*(\gamma) &= E_{q_1,q_3}[p(\tau,\gamma,\beta|Y)] \propto E_{q_1,q_3}[p(\beta|\gamma)p(\gamma)] \\ q_3^*(\beta) &= E_{q_1,q_2}[p(\tau,\gamma,\beta|Y)] \propto E_{q_1,q_2}[p(Y|\tau,\beta)p(\beta|\gamma)] \end{split}$$

Then

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q_{1}^{*} \sim Gamma\left(c + \frac{N+1}{2}, d + \frac{1}{2}\left\{Y'Y - E_{q3}[\beta'](X'Y)\right\} + tr\left[X(var_{q3}[\beta] + E_{q3}[\beta]E_{q3}[\beta'])X'\right]\right)
  q_2^* \sim \prod_{i=1}^p Gamma(a + \frac{1}{2}, b + \frac{1}{2} \{var_{q3}[\beta]_{i,i} + E_{q3}[\beta_i]^2\})
  q_3^* \sim N\left(E_{q1}[\tau]\Sigma X'Y, \left(diag(E_{q2}[\gamma]) + E_{q1}[\tau]X'X\right)^{-1} = \Sigma\right)
In [12]: def product(a):
               n = len(a)
               out = np.zeros([n,n])
               for i in range(n):
                    for j in range(n):
                        out[i,j] = a[i]*a[j]
               return(out)
In [13]: def mfvb(X,y,max_iter=1000):
               N,p = X.shape
               a ,b, c, d = [10**(-7)]*4
               a_tilde = np.repeat(a + 0.5, p)
               b_tilde = np.repeat(b,p)
               c_{tilde} = c + (N+1)/2
               d_{tilde} = d
               mu_coeffs = np.repeat(0,p)
               sigma_coeffs = np.diag(np.repeat(1,p))
               for i in range(max_iter):
                    expected_coeffs = mu_coeffs
                    double_expected_coeffs = sigma_coeffs + product(mu_coeffs)
                    diagonal_sigma = np.diag(sigma_coeffs)
                    expected_alpha = np.array(list(map(lambda x : a_tilde[x]/b_tilde[x] , np.arange
                    log_expected_alpha = np.array(list(map(lambda x : digamma(a_tilde[x])-np.log(b_
                    expected_tau = c_tilde / d_tilde
                    log_expected_tau = digamma(c_tilde)-np.log(d_tilde)
                    sigma_coeffs = np.linalg.inv(np.diag(expected_alpha)+expected_tau*(X.T.dot(X)))
                    mu_coeffs = expected_tau*sigma_coeffs.dot(X.T.dot(y))
                    b_tilde = np.array(list(map(lambda x : (diagonal_sigma[x]+mu_coeffs[x]**2)/2 +
                    d_{tilde} = d+0.5*(y.T.dot(y)) - expected_coeffs.T.dot((X.T.dot(y))) + 0.5*sum(np.detilde)
               return mu_coeffs, sigma_coeffs
In [14]: m,c = mfvb(d_x,y)
In [15]: m.round(1)
Out[15]: array([ 21.9, -88.7, 129. , -76. , 12.8, 0. , 0. , -0. , -1. ,
                    -0., 8., -7.5, -0., 1.2])
```

In [17]: ci95(m,c,n=1000)



Dashed line is True function, solid line is median estimator and gray filled area is 95% confidence interval

6 MFVB method with variable selection

Variable selection model is

$$Y = X\Gamma\beta + \epsilon, \quad \epsilon \sim N(0, \sigma^2 I)$$

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where
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Y|\beta,\sigma^2,\gamma\sim N(X\Gamma\beta,\sigma^2I)
                                   \sigma^2 \sim Inverse - Gamma(A, B)
                                   \beta_j \sim N(0, \sigma_\beta^2)
                                   \gamma_i \sim Bernoulli(\rho)
In [18]: '''
         n = 100
         x1 = np.random.uniform(size = n)
          x2 = np.random.uniform(size = n)
          x3 = np.random.uniform(size = n)
          x4 = np.random.uniform(size = n)
          x5 = np.random.uniform(size = n)
          x6 = np.random.uniform(size = n)
          X = np.array([x1, x2, x3, x4, x5, x6]).T
          Beta_true = np.array([0.02,0.03,0.4,1,0,0])
          y = X.dot(Beta_true) + np.random.normal(size=n)
          N, p = X.shape
          111
Out[18]: '\nn = 100\nx1 = np.random.uniform(size = n)\nx2 = np.random.uniform(size = n)\nx3 = np
In [19]: def expit(x):
              if x < 100:
                  return(np.exp(x)/(1+np.exp(x)))
              else:
                  return(1)
In [20]: def vselect(X,y,maxiter=100):
              N,p = X.shape
              sigmab = 1
              A = 10**(-7)
              B = 10**(-7)
              tau = 1
              rho = 0.5
              w = np.repeat(0.5,p)
              lamb= np.log(rho/(1-rho))
              t = 0
              for iteration in range(maxiter):
                  test= False
                  W = np.diag(w)
                  omega = product(w) + W.dot(np.eye(p)-W)
                  sigma = np.linalg.inv(tau*np.multiply(X.T.dot(X),omega)+ (1/sigmab) * np.eye(p)
                  mu = tau*sigma.dot(W.dot(X.T.dot(y)))
                  s = B + 0.5*(np.linalg.norm(y)**2 -2*y.T.dot(X).dot(W).dot(mu) + np.trace(np.mu)
                  tau = (A+N/2)/s
```

```
eta =np.zeros(len(w))
                   for j in range(p):
                       \mathtt{eta[j]} = \mathtt{lamb} - 0.5 * \mathtt{tau} * (\mathtt{mu[j]} * * 2 + \mathtt{sigma[j,j]}) * \mathtt{np.linalg.norm}(\mathtt{X[:,j]}) * * 2
                       w[j] = expit(eta[j])
                   t=t+1
              return(mu,sigma,w)
In [21]: m2,c2,w2 = vselect(d_x,y)
In [34]: def cil95(m,c,w,n=100):
              np.random.seed(4428)
              \#smaple\_gam = np.random.r
              sampled_coef = np.random.multivariate_normal(m*w,c,size=n)
              y_grid = np.array([d_x.dot(b) for b in sampled_coef])
              quantile = np.array([np.sort(x)[[int(n*0.05),int(n*0.5),int(n*0.95)]]  for x in y_gr
              xq = np.array(sorted(np.array([x,quantile[:,0],quantile[:,1],quantile[:,2]]).T,key=
              plt.fill_between(xq[:,0], xq[:,1],xq[:,3], color =(0,0,0,0.2))
              y = f(grid_x) - f(grid_x).mean()
              plt.plot(xq[:,0],xq[:,2],'k',grid_x, y, '--')
              #plt.plot(x_grid,y_grid[10],'k',x_grid, f(x_grid), '--')
              plt.ylim(-1.5,3.5)
              plt.show()
```

