
Statistical view of Spline

Spline

- In computational field, it was populous to interpolate points.
- The property of convex function $x = \lambda x_0 + (1 - \lambda) x_1$ is fundamental of it.
- Spline is the method of such an interpolation that it connect the points through piecewise polynomial curves.
- (Thm) Let $(\mathbf{c}_i)_{i=1}^n$ be a set of control points for a spline curve f of degree d , with nondecreasing knots $(t_i)_{i=1}^{n+d+1}$,

$$f(t) = \sum_{i=d+1}^n p_{i,d}(t) B_{i,0}(t)$$

where $p_{i,d}$ is given recursively by

$$p_{i,d-r+1}(t) = \frac{t_{i+1} - t}{t_{i+r} - t_i} p_{i-1,d-r}(t) + \frac{t - t_i}{t_{i+r} - t_i} p_{i,d-r}(t)$$

for $i = d - r + 1, \dots, n$, and $r = d, d - 1, \dots, 1$, while $p_{i,0}(t) = \mathbf{c}_i$ for $i = 1, \dots, n$. This representation is equivalent to

$$f(t) = \sum_{i=1}^n \mathbf{c}_i B_{i,d}(t)$$

where $B_{i,d}$ is given recursively by

$$B_{i,d}(t) = \frac{t - t_i}{t_{i+1} - t_i} B_{i,d-1}(t) + \frac{t_{i+1+d} - t}{t_{i+1+d} - t_{i+1}} B_{i+1,d-1}(t)$$

where $B_{i,0}$ is

$$B_{i,0}(t) = \begin{cases} 1 & t \in [t_i, t_{i+1}) \\ 0 & \text{otherwise.} \end{cases}$$

- Our interest is to apply a spline method to approximate nonlinear function.
- Note that situation we face is a little different with the interpolation in that there are much data than control points.

Cubic spline

- Basis expansions : the input space can be expanded through a function $h_m : \mathbb{R}^p \rightarrow \mathbb{R}$. We call it m th transformation $m = 1, \dots, M$.

$$f(X) = \sum_{m=1}^M \beta_m h_m(X),$$

which is called as “a linear basis expansion in X ”.

- Let ξ_1 and ξ_2 be knots in X , then cubic spline is constructed with following transformation functions.

$$h_1(X) = 1$$

$$h_2(X) = X$$

$$h_3(X) = X^2$$

$$h_4(X) = X^3$$

$$h_5(X) = (X - \xi_1)_+^3$$

$$h_6(X) = (X - \xi_2)_+^3,$$

Total degree of freedom is 6 in that 3 separated spaces that have cubic polynomial functions respectively, and each knot has three restriction, continuous on function, first derivative and second derivative, respectively. Thus,

$$4 \times 3 - 2 \times 3 = 6.$$

- This truncated formual is equal to original cubic spline with knots ξ_1 and ξ_2 .

Natural cubic spline

- The behaviour of polynomials fit to data tends to be erratic near the boundaries, and extrapolation can be dangerous.
- By adding constraints that the function is linear beyond the boundary knots, the erratic behaviour would be alleviated.
- Two constraints each in both boundary regions.
- Basis functions with K knots are follows.

$$N_1(X) = 1$$

$$N_2(X) = X$$

$$N_{k+2}(X) = d_k(X) - d_{K-1}(X)$$

$$d_k(X) = \frac{(X - \xi_k)_+^3 - (X - \xi_K)_+^3}{\xi_K - \xi_k}.$$

B-spline

- (Natural) cubic spline is quite reasonable, but could make numerical unstability due to overflow, $100^3 = 1,000,000$.

- In this sense, B -spline can be a replacement in order to obtain efficient and stable calculation.
- It has recursive formula as degree d increases.

$$f(X) = \sum_{k=1}^K \beta_k B_{k,d}(X)$$

where $B_{k,d}(X)$ is a basis function of B -spline that has recursive formula,

$$B_{k,d}(X) = \frac{t - t_k}{t_{k+1} - t_k} B_{k,d-1}(X) + \frac{t_{k+1+d} - t}{t_{k+1+d} - t_{k+1}} B_{k+1,d-1}(X)$$

with $B_{k,0}$

$$B_{k,0}(t) = \begin{cases} 1 & t \in [t_k, t_{k+1}) \\ 0 & \text{otherwise.} \end{cases}$$

Varying coefficients model

Consider linear model $Y = X\beta + \epsilon$. If there is an another variable t so that it makes an interaction effect with X , then the linear model can be represented as

$$Y = X\beta(t) + \epsilon.$$

$$y_i = x_i\beta(t) + \epsilon_i$$

We call this model varying coefficient model. Let $Y \in \mathbb{R}^{n \times 1}$, $X \in \mathbb{R}^{n \times 1}$, $t \in \mathbb{R}^{n \times 1}$, and $\epsilon \sim N(0, \tau^{-1})$.

Let one explanatory variable exist. The varying coefficient can be estimated via B -spline model.

$$\begin{aligned} \beta(t) &= \sum_{k=1}^K \phi_k B_{k,d}(t) \\ &= \mathbf{B}\phi, \end{aligned}$$

where $B_{k,d}(t)$ is basis function of B -spline with $\xi_1, \xi_2, \dots, \xi_K$ as knots, and \mathbf{B} is

$$\mathbf{B}\phi = \begin{bmatrix} B_{1,d}(t_1) & B_{2,d}(t_1) & \cdots & B_{K-1,d}(t_1) & B_{K,d}(t_1) \\ B_{1,d}(t_2) & B_{2,d}(t_2) & \cdots & B_{K-1,d}(t_2) & B_{K,d}(t_2) \\ \vdots & \vdots & \cdots & \vdots & \vdots \\ B_{1,d}(t_{n-1}) & B_{2,d}(t_{n-1}) & \cdots & B_{K-1,d}(t_{n-1}) & B_{K,d}(t_{n-1}) \\ B_{1,d}(t_n) & B_{2,d}(t_n) & \cdots & B_{K-1,d}(t_n) & B_{K,d}(t_n) \end{bmatrix} \begin{bmatrix} \phi_1 \\ \phi_2 \\ \vdots \\ \phi_{K-1} \\ \phi_K \end{bmatrix} = \begin{bmatrix} B_{\cdot,d}(t_1)^\top \phi \\ B_{\cdot,d}(t_2)^\top \phi \\ \vdots \\ B_{\cdot,d}(t_{n-1})^\top \phi \\ B_{\cdot,d}(t_n)^\top \phi \end{bmatrix}$$

Then, the model $E[Y | X] = X\beta(t) = X\mathbf{B}\phi$,

$$\begin{aligned}
X\mathbf{B}\phi &= \begin{bmatrix} x_1 B_{\cdot,d}(t_1)^\top \phi \\ x_2 B_{\cdot,d}(t_2)^\top \phi \\ \vdots \\ x_{n-1} B_{\cdot,d}(t_{n-1})^\top \phi \\ x_n B_{\cdot,d}(t_n)^\top \phi \end{bmatrix} \\
&= \begin{bmatrix} x_1 B_{1,d}(t_1) & x_1 B_{2,d}(t_1) & \cdots & x_1 B_{K-1,d}(t_1) & x_1 B_{K,d}(t_1) \\ x_2 B_{1,d}(t_2) & x_2 B_{2,d}(t_2) & \cdots & x_2 B_{K-1,d}(t_2) & x_2 B_{K,d}(t_2) \\ \vdots & \vdots & \cdots & \vdots & \vdots \\ x_{n-1} B_{1,d}(t_{n-1}) & x_{n-1} B_{2,d}(t_{n-1}) & \cdots & x_{n-1} B_{K-1,d}(t_{n-1}) & x_{n-1} B_{K,d}(t_{n-1}) \\ x_n B_{1,d}(t_n) & x_n B_{2,d}(t_n) & \cdots & x_n B_{K-1,d}(t_n) & x_n B_{K,d}(t_n) \end{bmatrix} \begin{bmatrix} \phi_1 \\ \phi_2 \\ \vdots \\ \phi_{K-1} \\ \phi_K \end{bmatrix}.
\end{aligned}$$

Let $X\mathbf{B} = W$.

Let's study simulation!

Specifying prior distributions

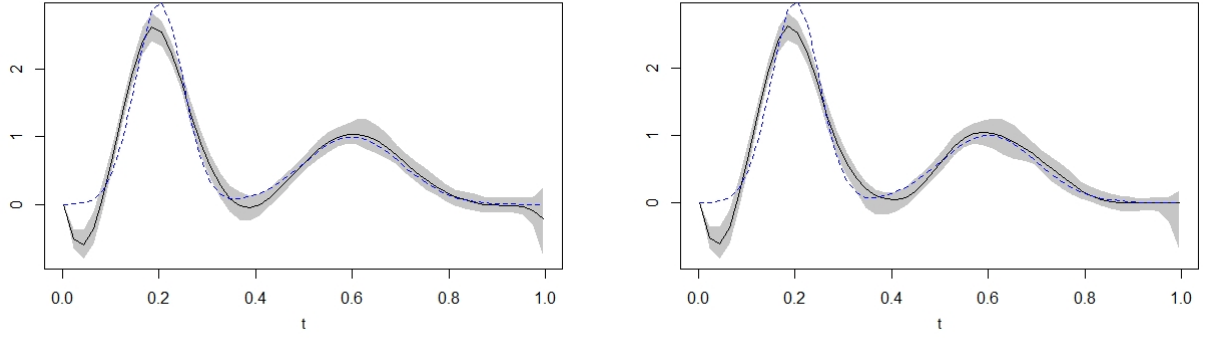
Following prior distributions are considered for each parameter.

$$\begin{aligned}
\phi_k &\sim \text{N}(0, \alpha_k^{-1}) \quad \forall k = 1, \dots, K \\
\alpha_k &\sim \text{Gamma}(a, b) \quad \forall k = 1, \dots, K \\
\tau &\sim \text{Gamma}(c, d).
\end{aligned}$$

Setting

Varying coefficient model $\beta(t) = 3 \exp(-200(t - 0.2)^2) + \exp(-50(t - 0.6)^2)$.

The left is the result of variational inference, and the right is of ordinary Gibbs sampler.



Variational inference

Steps of variational inference are introduced at [https://jinwonsohn.github.io/statistics/bayesian/2019/03/22/Variational-Inference-\(3\).html](https://jinwonsohn.github.io/statistics/bayesian/2019/03/22/Variational-Inference-(3).html)

Gibbs sampler

Construction of Gibbs sampler is as follows.

$$Y = W\phi + \epsilon$$

$$\phi \sim N(0, \alpha^{-1}I) \in \mathbb{R}^{p \times 1}$$

where $\alpha = (\alpha_1, \dots, \alpha_p)$ has prior as

$$\alpha_k \sim \text{Gamma}(a, b), \forall k = 1, \dots, p$$

$$\epsilon \sim N(0, \tau^{-1}) \in \mathbb{R}^{N \times 1}$$

where τ has prior as

$$\tau \sim \text{Gamma}(c, d).$$

The joint likelihood is

$$\begin{aligned} p(\phi, \alpha, \tau | Y) &\propto p(Y | \phi, \alpha, \tau) p(\tau) p(\phi | \alpha) p(\alpha) \\ &= N(W\phi, \tau^{-1}I) \text{Gamma}(\tau | c, d) N(\phi | 0, \alpha^{-1}I) \prod_{k=1}^p \text{Gamma}(\alpha_k | a, b) \end{aligned}$$

Then, Gibbs sampler is

1. For ϕ ,

$$\begin{aligned}
p(\phi \mid -) &\propto \mathcal{N}(W\phi, \tau^{-1}I) \times \mathcal{N}(\phi \mid 0, \alpha^{-1}I) \\
&= \det(2\pi\tau^{-1}I)^{-1/2} \exp\left(-\frac{\tau}{2}(Y - W\phi)^\top (Y - W\phi)\right) \times \det(2\pi\alpha^{-1}I) \exp\left(-\frac{\alpha}{2}\phi^\top \phi\right) \\
&\propto \exp\left(-\frac{\tau}{2}[\phi^\top W^\top W\phi - 2\phi^\top W^\top Y] - \frac{\alpha}{2}\phi^\top \phi\right) \\
&= \exp\left(-\frac{1}{2}[\tau\phi^\top W^\top W\phi - 2\tau\phi^\top W^\top Y + \alpha\phi^\top \phi]\right) \\
&= \exp\left(-\frac{1}{2}[\phi^\top (\tau W^\top W + \alpha I)\phi - 2\phi^\top \tau W^\top Y]\right) \\
&\sim \mathcal{N}\left(\tau(\tau W^\top W + \alpha I)^{-1} W^\top Y, (\tau W^\top W + \alpha I)^{-1}\right).
\end{aligned}$$

2. For τ ,

$$\begin{aligned}
p(\tau \mid -) &\propto \mathcal{N}(W\phi, \tau^{-1}I) \text{Gamma}(\tau \mid c, d) \\
&\propto \det(2\pi\tau^{-1}I)^{-1/2} \exp\left(-\frac{\tau}{2}(Y - W\phi)^\top (Y - W\phi)\right) \times \tau^{c-1} \exp(-d\tau) \\
&\propto \tau^{N/2+c-1} \exp\left(-\tau\left(\frac{1}{2}(Y - W\phi)^\top (Y - W\phi) + d\right)\right) \\
&\sim \text{Gamma}\left(c + \frac{N}{2}, d + \frac{1}{2}(Y - W\phi)^\top (Y - W\phi)\right).
\end{aligned}$$

3. For α_k ,

$$\begin{aligned}
p(\alpha_k \mid -) &\propto \mathcal{N}(\phi_k \mid 0, \alpha_k^{-1}) \text{Gamma}(\alpha_k \mid a, b) \\
&\propto \det(2\pi\alpha_k^{-1})^{-1/2} \exp\left(-\frac{\alpha_k}{2}\phi_k^2\right) \times \alpha_k^{a-1} \exp(-b\alpha_k) \\
&\propto \alpha_k^{1/2+a-1} \exp\left(-\alpha_k\left(\frac{\phi_k^2}{2} + b\right)\right) \\
&\sim \text{Gamma}\left(a + \frac{1}{2}, \frac{\phi_k^2}{2} + b\right), \quad \forall k = 1, \dots, p
\end{aligned}$$
