

gibbs

July 11, 2019

```
In [1]: import numpy as np
import pandas as pd
from scipy import stats
import matplotlib.pyplot as plt

from scipy.special import digamma
```

0.1 Make the function

Make the required functions, which calculate XX' , *expit* and draw the autocorelation plot

```
In [2]: def bernoulliSample(p=0.5,n=1):
    lst = []
    for i in range(n):
        if p > np.random.uniform(low=0,high=1):
            lst.append(1)
        else:
            lst.append(0)
    if n==1:
        return(lst[0])
    else:
        return(np.array(lst))
```

```
In [3]: def product(a):
    n = len(a)
    out = np.zeros([n,n])
    for i in range(n):
        for j in range(n):
            out[i,j] = a[i]*a[j]
    return(out)
```

```
In [4]: def expit(x):
    #if x < 100:
    return(np.exp(x)/(1+np.exp(x)))
    #else:
    #return(1)
```

```

In [5]: def acf(sampl,lag =40):
        sampl= np.array(sampl)
        base = sampl.dot(sampl)/len(sampl)
        acr = [1]
        for t in range(1,lag):
            acr.append((sampl[t:].dot(sampl[:-t]))/(len(sampl)-t))/base)
        x = list(range(lag))
        plt.bar(x,acr,color='gray')
        plt.title('Autocorrelation plot')
        plt.hlines(0.05,xmin=-0.5 ,xmax=lag,colors='r',linestyles='dashed')
        plt.hlines(-0.05,xmin=-0.5 ,xmax=lag,colors='r',linestyles='dashed')
        #plt.show()

In [90]: def hist(sampled,tlr,nbin =19,gr=4):
        grid =np.linspace(-gr,gr,1000)
        plt.hist(sampled, bins=nbin,color='gray',density=True)
        #plt.plot(grid , stats.norm.pdf(grid), 'black')
        plt.title(tlr)
        plt.xlabel('x')
        #plt.xlim(-4,4)
        plt.ylabel('density')
        #plt.show()

```

1 Gibbs Sampling

1.1 Prior

$$\begin{aligned}
 Y|\beta, \sigma^2, \Gamma &\sim N(Z\Gamma\beta, \sigma^2 \cdot I) \\
 \sigma^2 &\sim \text{Inverse} - \text{Gamma}(a, b) \\
 \beta_j|\sigma_{\beta_j}^2 &\sim^{ind} N(0, \sigma_{\beta_j}^2) \\
 \sigma_{\beta_j}^2 &\sim^{iid} \text{Inverse} - \text{Gamma}(c, d) \\
 \gamma_j &\sim^{iid} \text{Bernoulli}(\rho) \\
 \rho &\sim \text{Beta}(u, v)
 \end{aligned}$$

where

Z is $n \times p$ design matrix

$\Gamma = \text{diag}(\gamma_j)$ for $j = 1, \dots, p$

a, b, c, d, u, v is flat prior

1.2 Posterior

$$\begin{aligned}
p(\beta, \sigma^2, \sigma_{\beta_j}^2, \Gamma | Y) &\propto p(\beta, \sigma^2, \sigma_{\beta_j}^2, \Gamma, Y) \\
&\propto p(Y | \beta, \sigma^2, \Gamma) p(\beta | \sigma_{\beta_j}^2) p(\Gamma) p(\sigma^2) p(\sigma_{\beta_j}^2) p(\rho) \\
&\propto (\sigma^2)^{-n/2} \exp \left(-\frac{1}{2\sigma^2} (Y - Z\Gamma\beta)' (Y - Z\Gamma\beta) \right) \\
&\quad \times \prod_{j=1}^p (\sigma_{\beta_j}^2)^{-1/2} \exp \left(-\frac{1}{2} \sum_{j=1}^p \frac{\beta_j^2}{\sigma_{\beta_j}^2} \right) \\
&\quad \times \prod_{j=1}^p \rho^{\gamma_j} (1 - \rho)^{1-\gamma_j} \\
&\quad \times (\sigma^2)^{-a-1} \exp \left(-\frac{b}{\sigma^2} \right) \\
&\quad \times \prod_{j=1}^p (\sigma_{\beta_j}^2)^{-c-1} \exp \left(-\sum_{j=1}^p \frac{d}{\sigma_{\beta_j}^2} \right) \\
&\quad \times \rho^{u-1} (1 - \rho)^{v-1}
\end{aligned}$$

1.3 Sampling the β from

$$N(\mu, \Sigma)$$

where

$$\Sigma = \left(\text{diag}(\sigma_{\beta_j}^2) + \frac{1}{\sigma^2} \Gamma' Z' Z \Gamma \right)^{-1}, \quad \mu = \frac{1}{\sigma^2} \Sigma \Gamma' Z' y$$

```

In [7]: def sampleBeta(gamma,s2,sb2):
        Gamma = np.diag(gamma)
        D = np.diag(1/sb2)
        sinv = D + (1/s2)*Gamma.T.dot(Z.T.dot(Z.dot(Gamma)))
        Sigma = np.linalg.inv(sinv)
        mu = (1/s2)*Sigma.dot(Gamma.T).dot(Z.T).dot(y)
        out = np.random.multivariate_normal(mu,Sigma)
        return(out)

```

1.4 Sampling σ^2 from

$$Inverse - Gamma \left(a + \frac{N}{2}, b + \frac{1}{2} (y - \Gamma\beta)' (y - \Gamma\beta) \right)$$

```

In [8]: def sampleS2(beta,gamma):
        Gamma = np.diag(gamma)
        alpha = a + N/2
        igbeta = b + 0.5*(y-Z.dot(Gamma).dot(beta)).T.dot((y-Z.dot(Gamma).dot(beta)))
        out = stats.invgamma.rvs(a=alpha,scale=igbeta)
        return(out)

```

1.5 Sampling $\sigma_{\beta_j}^2$ from

$$Inverse - Gamma \left(c + \frac{1}{2}, d + \frac{1}{2} \beta_j^2 \right)$$

```
In [9]: def sampleSb2(beta):
    lst = []
    for j in range(p):
        alpha = c + 1/2
        igbeta = d+0.5*(beta[j]**2)
        lst.append(stats.invgamma.rvs(a=alpha,scale=igbeta))
    out = np.array(lst)
    return(out)
```

1.6 Sampling γ_j from

$$Bernoulli \left(\expit \left(\text{logit}(\rho) + \frac{\beta_j}{\sigma^2} Z_j' (y - Z_{-j} \Gamma_{-j} \beta_{-j}) - \frac{\beta_j^2}{2\sigma^2} Z_j' Z_j \right) \right)$$

expit is inverse function of *logit*

```
In [10]: def samplegamma(rho,s2,beta,gamma):
    gam = gamma.copy()
    wlst = []
    for j in range(p):
        Gamma= np.diag(gam)
        eta = np.log(rho/(1-rho))-((beta[j]**2)/s2)*Z[:,j].T.dot(Z[:,j])\
        +(beta[j]/s2)*Z[:,j].T.dot(y-np.delete(Z,j,1).dot(np.diag(np.delete(gamma,j)).d
        w = expit(eta)
        wlst.append(w)
        gam[j] = w
    outlst = []
    for k in gam:
        outlst.append(stats.bernoulli.rvs(k))
    out = np.array(outlst)
    return(out)
```

1.7 Sampling ρ from

$$Beta \left(\sum_{j=1}^p \gamma_j + u, p - \sum_{j=1}^p \gamma_j + v \right)$$

```
In [11]: def sampleRho(gamma):
    alpha = sum(gamma) +u
    bbeta = p-sum(gamma) + v
    out = np.random.beta(alpha,bbeta)
    return(out)
```

1.8 Gibbs sampling function

```
In [12]: def gibbs():
    lst = []
    for i in range(1000):
        beta = sampleBeta(gamma,s2,sb2)
        s2 = sampleS2(beta,gamma)
        sb2 = sampleSb2(beta)
        gamma = samplegamma(rho,s2,beta,gamma)
        rho = sampleRho(gamma)

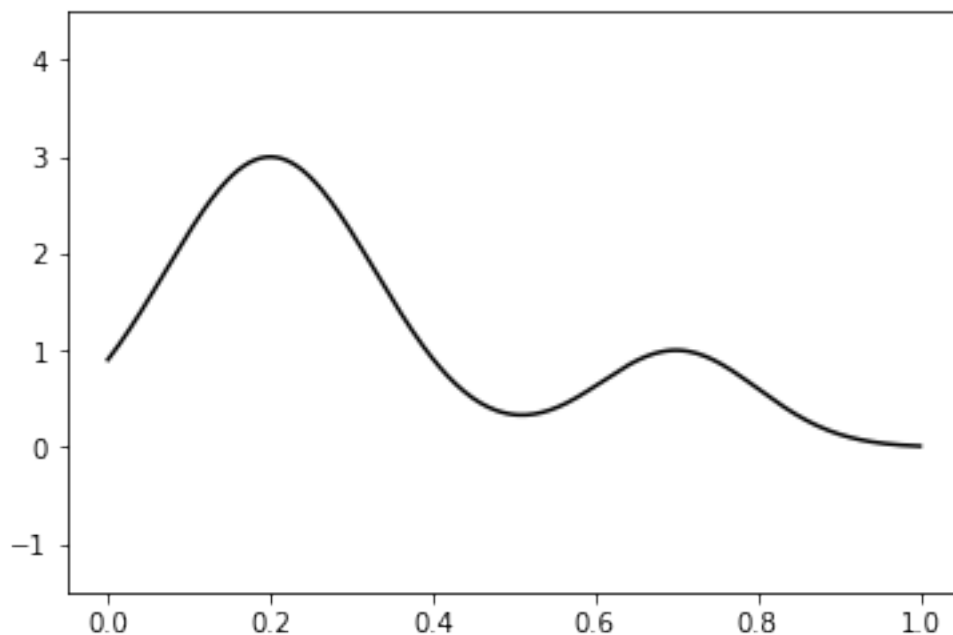
        params = np.array([beta,s2,sb2,gamma,rho])
        lst.append(params)
    return(pd.DataFrame(lst))
```

1.9 Simulation Data

$$f(x) = 3\exp(-30(x - 0.2)^2) + \exp(-50(x - 0.7)^2)$$

```
In [13]: def f(x):
    out = 3*np.exp(-30*((x-0.2)**2))+np.exp(-50*((x-0.7)**2))
    return(out)
```

```
In [14]: x = np.linspace(0,1,300)
    y = f(x)
    plt.plot(x, y, 'k')
    plt.ylim(-1.5, 4.5)
    plt.show()
```

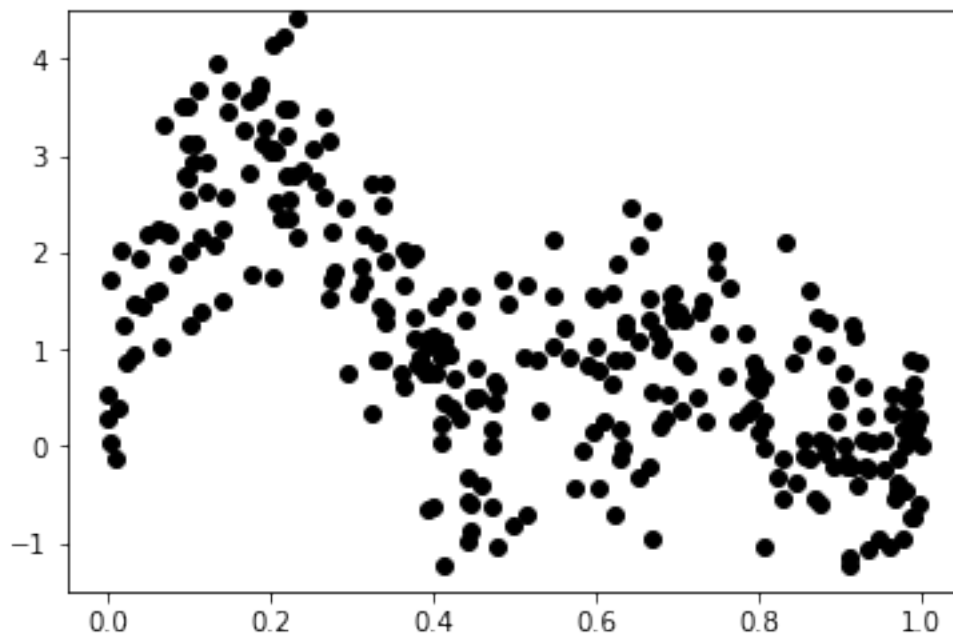


1.10 Make data wiht N(0,0.5) error

```
In [15]: def mkToy(n=300,tau = 0.5):  
        x = np.random.uniform(size = n)  
        e = np.random.normal(0,np.sqrt(0.5), size= n)  
        y = f(x)+e  
        #out = np.column_stack([x,y])  
        return(x,y)
```

```
In [16]: x,y = mkToy()
```

```
In [17]: plt.plot(x,y,'ko')  
        plt.ylim(-1.5, 4.5)  
        plt.show()
```



1.11 Radial basis

we use radial basis functions defined by

$$\mathbf{b}(u) = \left\{ u, \left| \frac{u - \tau_1}{c} \right|^3, \dots, \left| \frac{u - \tau_K}{c} \right|^3 \right\}$$

where c is sample standard deviation

```
In [18]: def defineKnot(X,K=10):  
        upper = max(X)
```

```

        lower = min(X)
        out = np.linspace(start=lower, stop=upper, num=K+2)[1:K+1]
        return(out)
def radialbasis(u, tau, sd):
    lst = []
    lst.append(u)
    for i in tau:
        lst.append(abs((u-i)/sd)**3)
    out = np.array(lst)
    return(out)

```

```
In [19]: sd = np.std(x)
```

```
In [20]: knot = defineKnot(x)
        d_x = radialbasis(x, knot, sd).T
```

```
In [86]: #initial value
        Z = d_x
        N, p = Z.shape
        a,b,c,d = [10**-7]*4
        e,f = [1,1]
        gamma = bernoulliSample(0.5,p)
        sb2 = np.repeat(0.5,p)
        s2 = 0.5
        rho=0.5
        u,v = 1,1
        print(gamma)

```

```
[0 1 1 1 0 0 0 0 1 0 0]
```

```
In [87]: lst = []
        blst = []
        rholst = []
        s2lst = []
        for i in range(2000):
            beta = sampleBeta(gamma,s2,sb2)
            s2 = sampleS2(beta,gamma);s2lst.append(s2)
            #sb2 = sampleSb2(beta)
            gamma = samplegamma(rho,s2,beta,gamma)
            rho = sampleRho(gamma);rholst.append(rho)

            #params = np.array([beta,s2,sb2,gamma,rho])
            #lst.append(params)
            blst.append(beta)

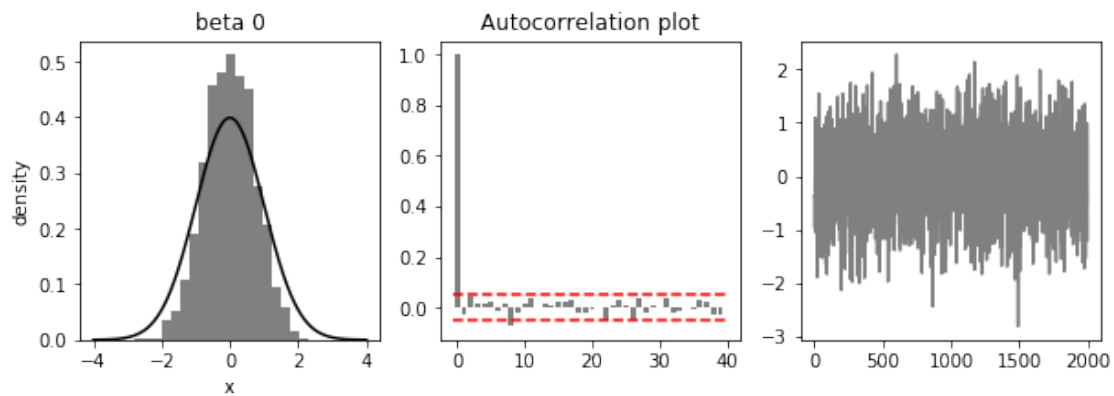
```

```
In [80]: df = pd.DataFrame(lst)
```

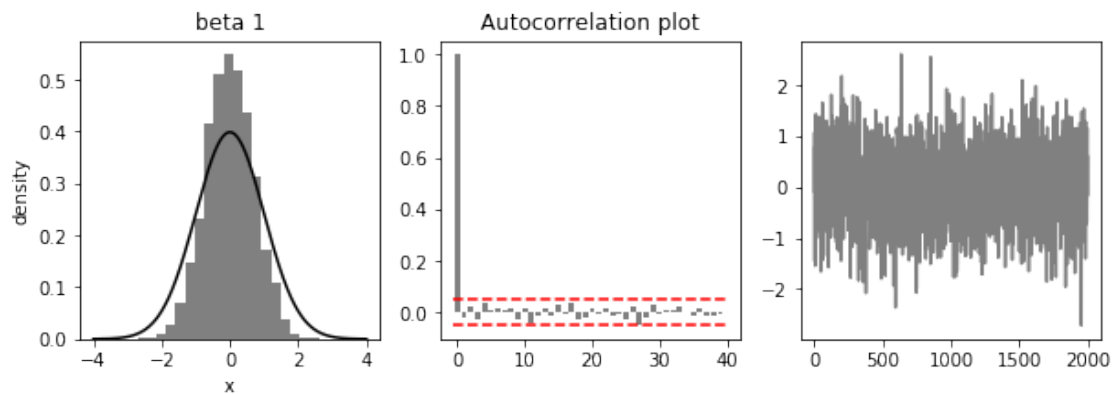
```
In [81]: bdf = pd.DataFrame(blst)
```

```
In [82]: def betaplt(j):
plt.figure(figsize=(10, 3))
plt.subplot(1,3,1)
hist(bdf[j],tlt='beta %d'%j)
plt.subplot(1,3,2)
acf(bdf[j])
plt.subplot(1,3,3)
plt.plot(bdf[j],color='gray')
plt.show()
```

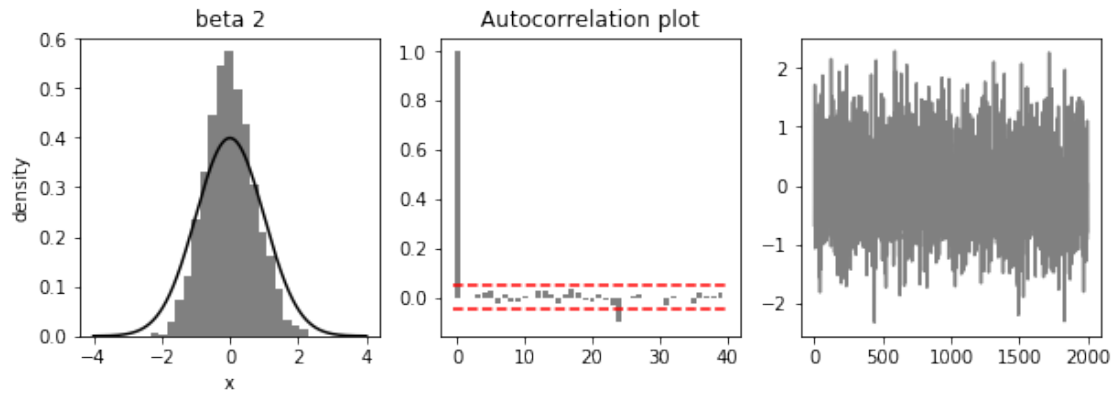
```
In [66]: betaplt(0)
```



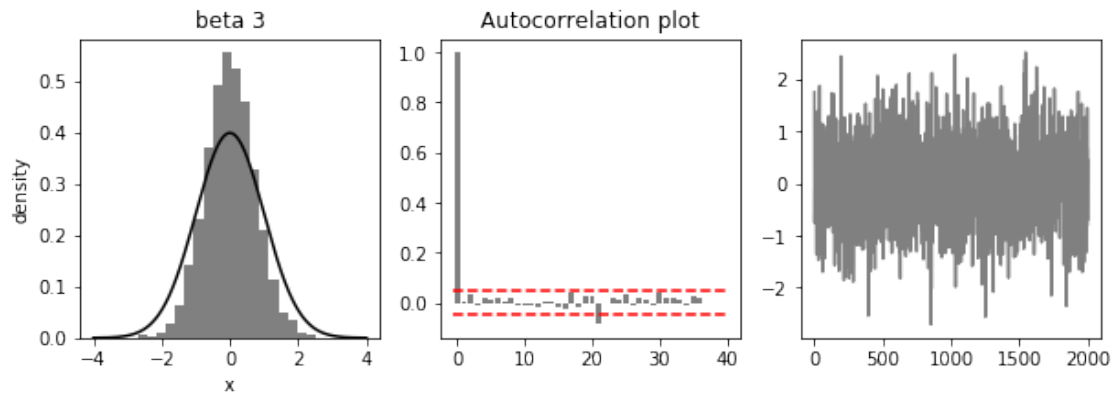
```
In [67]: betaplt(1)
```



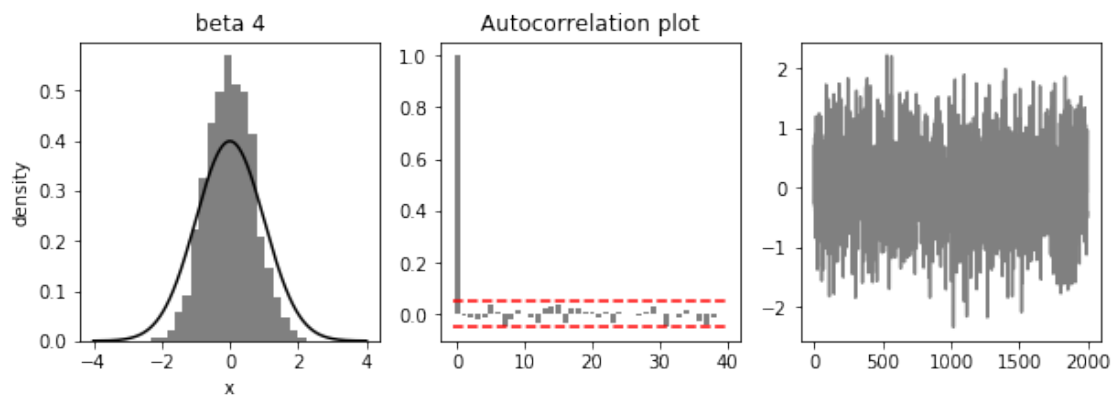
```
In [68]: betaplt(2)
```

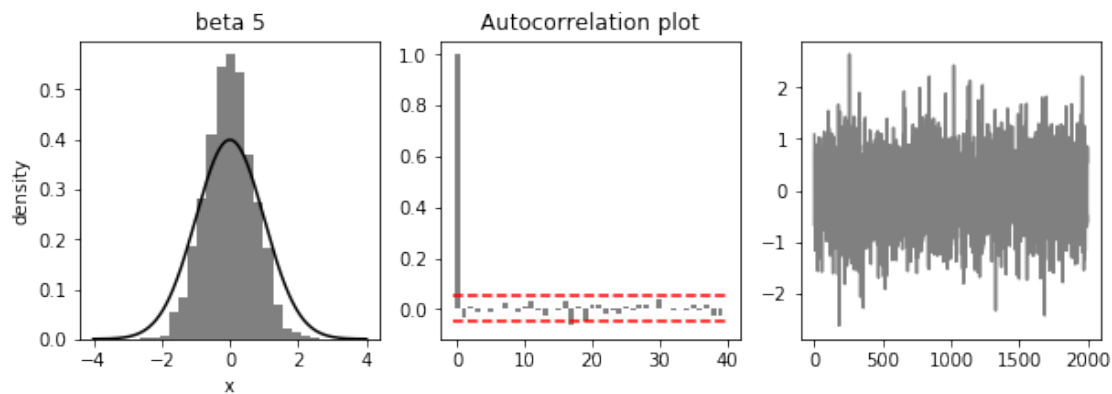
In [69]: `betaplt(3)`



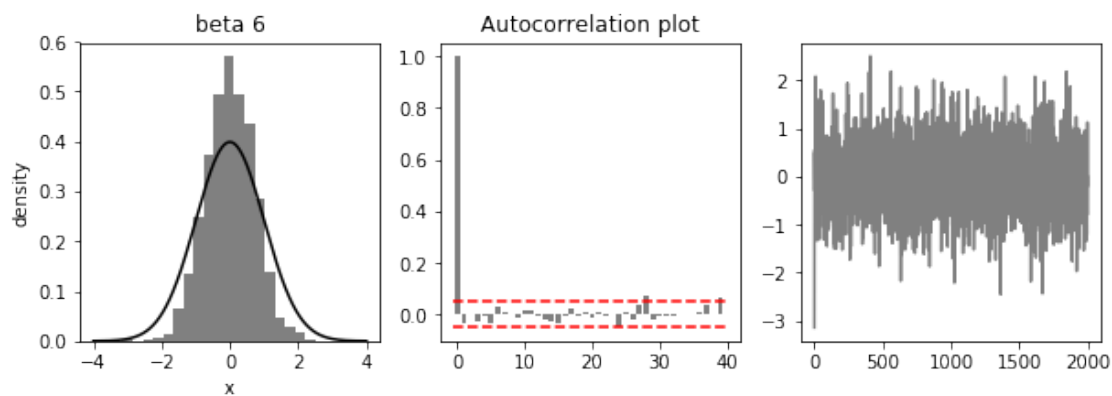
In [70]: `betaplt(4)`



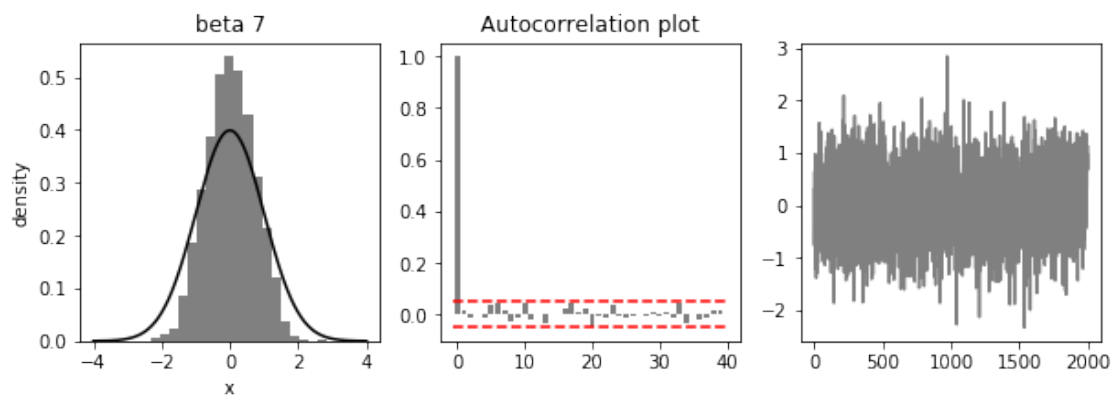
In [71]: betaplt(5)



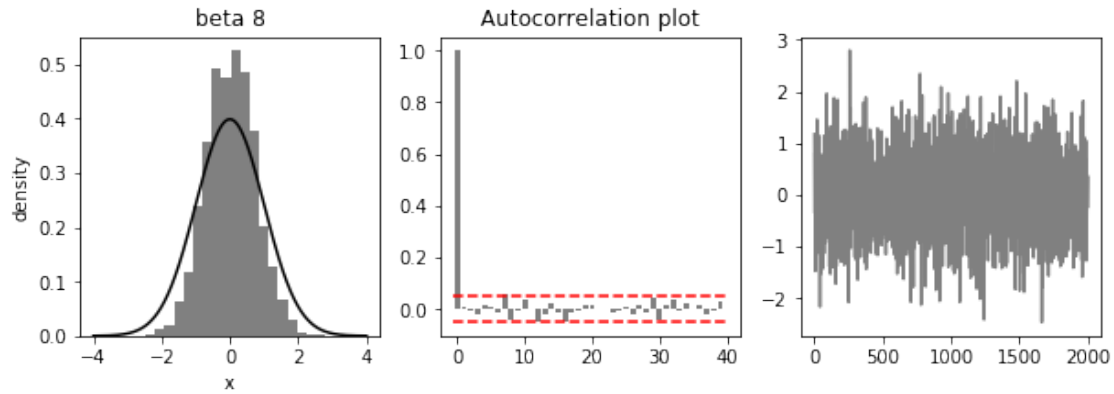
In [72]: betaplt(6)



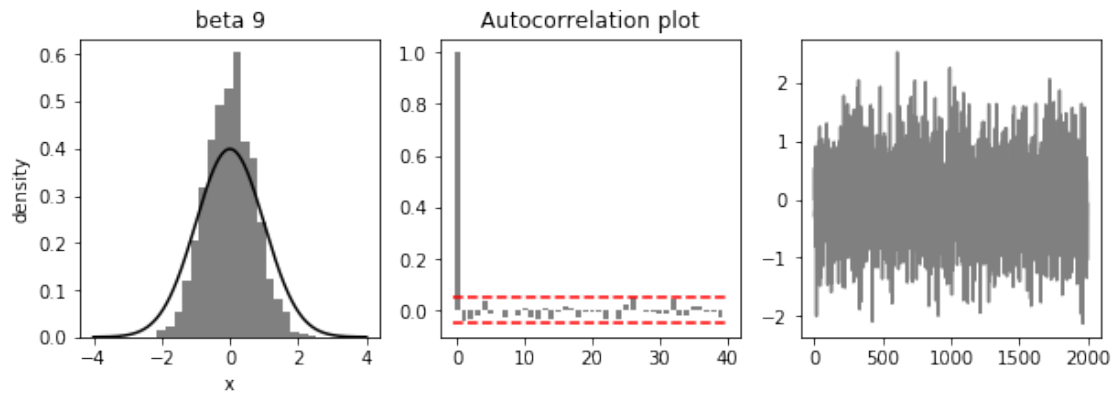
In [73]: betaplt(7)



```
In [74]: betaplt(8)
```



```
In [75]: betaplt(9)
```



```
In [56]: bdfarray = np.array(bdf)
```

```
In [51]: def mfvb(X,y,max_iter=100):
```

```
    N,p = X.shape
    a ,b, c, d = [10**(-7)]*4
    a_tilde = np.repeat(a + 0.5, p)
    b_tilde = np.repeat(b,p)
    c_tilde = c + (N+1)/2
    d_tilde = d
```

```

mu_coeffs = np.repeat(0,p)
sigma_coeffs = np.diag(np.repeat(1,p))

for i in range(max_iter):
    expected_coeffs = mu_coeffs
    double_expected_coeffs = sigma_coeffs + product(mu_coeffs)
    diagonal_sigma = np.diag(sigma_coeffs)
    expected_alpha = np.array(list(map(lambda x : a_tilde[x]/b_tilde[x] , np.arange
    log_expected_alpha = np.array(list(map(lambda x : digamma(a_tilde[x])-np.log(b_
    expected_tau = c_tilde / d_tilde
    log_expected_tau = digamma(c_tilde)-np.log(d_tilde)
    sigma_coeffs = np.linalg.inv(np.diag(expected_alpha)+expected_tau*(X.T.dot(X)))
    mu_coeffs = expected_tau*sigma_coeffs.dot(X.T.dot(y))
    b_tilde = np.array(list(map(lambda x : (diagonal_sigma[x]+mu_coeffs[x]**2)/2 +
    d_tilde = d+0.5*(y.T.dot(y)) - expected_coeffs.T.dot((X.T.dot(y)))+ 0.5*sum(np.
    return mu_coeffs,sigma_coeffs

```

In [52]: m,c = mfvb(Z,y)

In [53]: m

Out[53]: array([-0.03524236, -2.92799365, 5.4149432 , -0.01378641,
-0.07064352, -5.46560637, 0.14415652, 0.07018293,
9.37113655, -11.89600525, 4.49279968])

```

In [91]: plt.figure(figsize=(10, 3))
plt.subplot(1,3,1)
hist(s2lst,plt='s2')
plt.subplot(1,3,2)
acf(s2lst)
plt.subplot(1,3,3)
plt.plot(s2lst,color='gray')
plt.show()

```

