

SDSS Stripe 82 : quasar variability from forced photometry

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Accepted XXX. Received YYY; in original form ZZZ

ABSTRACT

1 VARIABILITY

Many lightcurves display intrinsic variability, in addition to the error-induced noise. A lightcurve consists of a set of N measurements of the object brightness x_i and associated errors e_i . In this analysis we assume that x_i are drawn from a Gaussian distribution $\mathcal{N}(\mu, \sigma)$, and that errors e_i are homoscedastic, so that the distribution of measurements is a Gaussian. In this framework μ describes the median brightness, which for non-variable objects is the true brightness. Using the Bayesian approach, we find μ by maximizing the posterior probability distribution function (pdf) of μ given x_i and e_i : $p(\mu|x_i, \sigma_i)$. We can proceed analogously to find the width of the distribution, σ , which describes the departure from the mean.

To find μ and σ , we follow Ivezić+2014, with the two-step approach. First, we find approximate values of μ_0 and σ_0 , and then we evaluate the full logarithm of the posterior pdf in the vicinity of the approximate solution. The maximum of the 2D likelihood becomes our full solution - σ_{full} and μ_{full} , as shown on Fig. 1 (see Appendix B for the detailed calculation).

For each lightcurve, we also calculate mean-based χ^2_{DOF} and median-based χ^2_R (the latter is more robust against any outliers in the distribution) :

$$\chi^2_{dof} = \frac{1}{N-1} \sum \left(\frac{x_i - \mu}{e_i} \right)^2 \quad (1)$$

and

$$\chi^2_R = 0.7414(Z_{75\%} - Z_{25\%}) \quad (2)$$

with $Z = (x_i - \text{median}(x_i))/e_i$.

Initially, we evaluate μ_{full} , σ_{full} , χ^2_{dof} , and χ^2_R for the entire lightcurve. Then, only if either $\sigma_{full} > 0$ or $\chi^2 > 1$, which hints some intrinsic variability, we also calculate μ_{full} , σ_{full} , and χ^2 for the seasonally-binned portions of the lightcurve.

On Fig. 1 we plot the stages of calculating μ and σ . Left and middle panels compare two methods of calculating our variability parameters. The initial approximation (dashed) is based on bootstrapped resampling of x_i , e_i points from the lightcurve. With bootstrapping we resample the lightcurve M times, so that instead of a single sample with $N \approx 10 - 70$ points we have M samples. Random resampling means that points may be chosen multiple times in a random fashion.

This allows us to calculate M approximate values of μ, σ . Dashed lines trace the histogram for $M = 1000$ resamples of our variability parameters. The approximate values are in turn used to provide bounds for the 200×70 grid of μ, σ on which we evaluate the full posterior likelihood density function (right panel). This ensures that we are sampling the peak of the underlying distribution.

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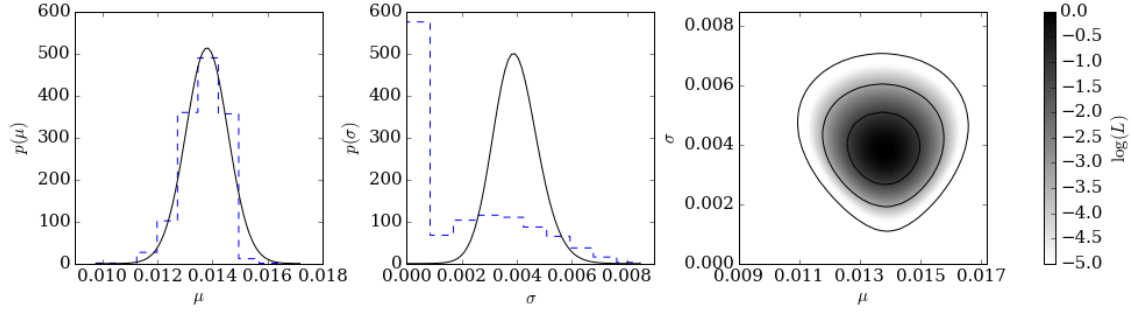


Figure 1. Two-step approach to finding μ and σ via μ_0 and σ_0 for an object 217720894888346446. In this calculation we use raw psf flux, before employing the faint source treatment outlined in Section ???. On the left and middle panels, solid lines trace marginalized posterior pdfs for μ and σ , while dashed lines depict histogram distributions of 10,000 bootstrap resamples for μ_0 and σ_0 . The right panel shows the logarithm of the posterior probability density function for μ and σ .