Homework 3

2 Days, 23 Hours Late

Student

Subhajyoti Saha

Total Points

96 / 120 pts

Question 1

- + 15 pts Fully correct answer
- + 5 pts Part 1 (unbiased expectation) correct
- + 4 pts Part 1 mostly correct with some minor errors
- + 3 pts Part 1 somewhat correct with some major errors but the basic approach is correct
- + 10 pts Part 2 (variance of the approximation) correct
- → + 8 pts Part 2 mostly correct with some minor errors
 - + 6 pts Part 2 somewhat correct with some major errors but the basic approach is correct
 - + 0 pts Incorrect or not attempted
- In part 2, you have not stated why you can write the variance in LHS of line 1 as RHS. It is only when the r.v.'s are independent but this needed to be stated explicitly. You have directly written it without any explanation.

Question 2

Problem 2 26 / 30 pts

- - +8 pts Basic steps correct
 - + 6 pts Correct mean-field posterior for \boldsymbol{w}
 - + 6 pts Correct mean-field posterior for β
 - + 6 pts Correct mean-field posterior for α
 - **+ 4 pts** Overall approach (the VI algorithm steps) correctly shown with appropriate expectations correctly specified
- ✓ 4 pts Didn't have this: Overall approach (the VI algorithm steps) correctly shown with appropriate expectations correctly specified. At least the expectations required in the mean-field distributions should be specified clearly.
 - + 0 pts Incorrect or not attempted

Problem 3 12 / 20 pts

- + 20 pts Fully correct answer for all parts
- + 7 pts Correct CP for each λ
- **✓ +4 pts** Somewhat correct CP for λ but some mistakes
 - + 7 pts Correct CP for α
- \checkmark +4 pts Somewhat correct CP for α but some mistakes
 - **+ 6 pts** Correct CP for β
- **+ 4 pts** Somewhat correct CP for β but some mistakes
 - + 0 pts Incorrect or not attempted

Question 4

Problem 4 5 / 15 pts

- + 15 pts Fully correct answer for all parts
- ✓ + 5 pts Correct answer for the mean
 - + 4 pts Mostly correct answer for the mean but with some small errors
 - + 10 pts Correct answer for the variance
 - + 8 pts Mostly correct answer for the variance but with some small errors
 - + 0 pts Incorrect or not attempted
- Incorrect expression for the variance

Question 5

Problem 5 40 / 40 pts

- → + 20 pts Correct implementation of the rejection sampler
 - + 15 pts Mostly correct implementation of the rejection sampler with some minor mistakes
- \checkmark + 20 pts Correct implementation of the MH sampler, correct plots, and correct answer to the questions about σ^2 and rejection rate
 - + 15 pts Mostly correct implementation of the MH sampler with some minor mistakes in implementation or not answering the questions about σ^2 and rejection rate correctly
 - + 0 pts Did not attempt

Question assigned to the following page: 1	

QUESTION

1

Student Name: Subhajyoti Saha Roll Number: 21111269 Date: November 18, 2022

1 Monte-Carlo Approximations (15 marks)

$$\hat{f} = \frac{1}{S} \sum_{s=1}^{S} f(z^{(s)})$$

$$\Rightarrow E[\hat{f}] = \frac{1}{S} \sum_{s=1}^{S} E[f(z^{(s)})]$$

$$= \frac{1}{S} S.E[f]$$

$$= E[f] \text{ [Prooved]}$$

It is prooved that the MCMC preserves the mean.

$$Var[\hat{f}] = \frac{1}{S^2} \sum_{i=1}^{S} Var[f(z^{(s)})]$$
$$= \frac{1}{S^2} S.Var[f]$$
$$= \frac{1}{S} Var[f]$$
$$= \frac{1}{S} E[(f - E[f])^2]$$

From the above equation, it is proven that the variation of MCMC samplers got reduced as the no of samples got incressed.

Question assigned to the following pa	age: <u>2</u>	

QUESTION

2

Student Name: Subhajyoti Saha

Roll Number: 21111269 Date: November 18, 2022

$$\begin{split} & \ln q^*(w) = E_{\beta,\alpha_1,\dots\alpha_D}[\ln P(y,X,w,\beta,\alpha_1,\dots\alpha_D)] + const \\ & = E_{\beta,\alpha_1,\dots\alpha_D}[\ln P(y|X,w)P(w)P(\beta)P(\alpha_1)\dots P(\alpha_D)] + const \\ & = \int q(\beta) \ln P(y|X,w)d\beta + \int q(\alpha_1)\dots q(\alpha_D) \ln P(w)d\alpha_1\dots d\alpha_D + const \\ & = \sum_{i=1}^n \int q(\beta) \ln P(y_i|x_i,w)d\beta + \int q(\alpha_1)\dots q(\alpha_D) \sum_{i=1}^D \ln P(w_i|\alpha_i)d\alpha_1\dots d\alpha_D + const \\ & = \sum_{i=1}^n \int q(\beta) \ln P(y_i|x_i,w)d\beta + \sum_{i=1}^D \int q(\alpha_i) \ln P(w_i|\alpha_i)d\alpha_i + const \\ & = \sum_{i=1}^n \int q(\beta) \ln N(y_i|w_i^Tx_i,\beta^{-1})d\beta + \sum_{i=1}^D \int q(\alpha_i) \ln N(w_i|0,\alpha_i^{-1})d\alpha_i + const \\ & = \sum_{i=1}^n -E_\beta[\frac{\beta}{2}(y_i-w^Tx_i)^2] - \sum_{i=1}^D E_{\alpha_i}[\frac{\alpha_i}{2}w_i^2] + const \\ & = -\sum_{i=1}^n \frac{E_\beta[\beta]}{2}(y_i-w^Tx_i)^2 - \sum_{i=1}^D \frac{E_{\alpha_i}[\alpha_i]}{2}w_i^2 + const \\ & = -\sum_{i=1}^n \frac{E_\beta[\beta]}{2}(w^Tx_ix_i^Tw - 2y_iw^Tx_i) - \sum_{i=1}^D \frac{E_{\alpha_i}[\alpha_i]}{2}w_i^2 + const \\ & = -\frac{1}{2}w^T[E_\beta[\beta]\sum_{i=1}^n x_ix_i^T + diag(E_{\alpha_i}[\alpha_i],\dots E_{\alpha_i}[\alpha_i])]w + w^T(\sum_{i=1}^n y_ix_i)E_\beta[\beta] + const \\ \Rightarrow \ln q^*(w) = \ln N(w|\mu,\Sigma) \\ \Rightarrow (\Sigma^{-1}) = \left[E_\beta[\beta]\sum_{i=1}^n x_ix_i^T + diag(E_{\alpha_i}[\alpha_i],\dots E_{\alpha_i}[\alpha_i])\right] \\ \mu = \Sigma(\sum_{i=1}^n y_ix_i)E_\beta[\beta] \end{split}$$

Question assigned to the following pa	age: <u>2</u>	

$$\begin{split} \ln q^*(\beta) &= E_{w,\alpha_1,\dots\alpha_D}[\ln P(y,X,w,\beta,\alpha_1,\dots\alpha_D)] + const \\ &= E_w[\ln P(y|X,w) + \ln P(\beta)] + const \\ &= \sum_{i=1}^n E_w[\ln N(y_i|w^Tx_i,\beta^{-1})] + \ln Gamma(\beta|a_o,b_o) + const \\ &= \frac{n}{2}\ln\beta - \frac{\beta}{2}\sum_{i=1}^n E_w[(y_i-w^Tx_i)^2] + (a_o-1)\ln\beta - b_o\ln\beta \\ &= \frac{n}{2}\ln\beta - \frac{\beta}{2}\sum_{i=1}^n (y_i-E_w[w]^Tx_i)^2 + (a_o-1)\beta - b_o\ln\beta + const \\ &= (\frac{n}{2}+a_o-1)\ln\beta - (b_o+\frac{1}{2}\sum_{i=1}^n (y_i-E_w[w]^Tx_i)^2)\beta + const \\ &\Rightarrow \ln q^*(\beta) = \ln Gamma(\beta|\hat{a},\hat{b}) \\ &\Rightarrow \hat{a} = (\frac{n}{2}+a_o) \\ &\Rightarrow \hat{b} = (b_o+\frac{1}{2}\sum_{i=1}^n (y_i-E_w[w]^Tx_i)^2) \\ &\ln q^*(\alpha_i) = E_{w,\beta,\alpha_{j;j\neq i}}[\ln P(y,X,w,\beta,\alpha_{j;j\neq i})] + const \\ &= E_w[\ln P(w) + \ln P(\alpha_i)] + const \\ &= E_w[\sum_{i=1}^D \ln P(w_i) + \ln P(\alpha_i)] + const \\ &= E_w[\sum_{i=1}^D \ln N(w_i|0,\alpha_i) + \ln P(\alpha_i)] + const \\ &= E_w[\frac{1}{2}\ln\alpha_i - \frac{\alpha_i}{2}w_i^2 + (e_o-1)\ln\alpha_i - f_o\alpha_i] + const \\ &= (\frac{1}{2}+e_o-1)\ln\alpha_i - (\frac{E_w[w_i^2]}{2} + f_o)\alpha_i + const \\ &\Rightarrow \ln q^*(\alpha_i) = \ln Gamma(\alpha_i|\hat{c},\hat{f}) \\ &\Rightarrow \hat{e} = \frac{1}{2} + e_o \\ &\Rightarrow \hat{f} = \frac{E_w[w_i^2]}{2} + f_o \end{split}$$

The above distribution is same for all $q^*(\alpha_i)$ for i = 1, ...D.

Question assigned to the following page: 3	

2

QUESTION

Student Name: Subhajyoti Saha

Roll Number: 21111269 Date: November 18, 2022

The CPs of all the variable are shown below:

$$\begin{split} P(\lambda_i|\lambda_{jj\neq i},\alpha,\beta,x_1,..x_n) &= P(\lambda_i|x_1,\alpha,\beta) \\ &= P(x_i|\lambda_i)P(\lambda_i|\alpha,\beta) \\ &= Poisson(x_i|\lambda_i)Gamma(\lambda_i|\alpha,\beta) \\ &= \frac{\lambda_i^{x_i}e^{-\lambda_i}}{x_i!}\frac{1}{\beta^{\alpha}\Gamma(\alpha)}\lambda_i^{\alpha-1}e^{-\frac{\lambda_i}{\beta}} \\ &\propto \lambda_i^{x_i+\alpha-1}e^{-\frac{i(1+\beta)}{\beta}} \\ &\Rightarrow P(\lambda_i|x_1,\alpha,\beta) = Gamma(_i|\hat{\alpha},\hat{\beta}) \\ &\Rightarrow \hat{\alpha} = x_i + \alpha \\ &\Rightarrow \hat{\beta} = \frac{\beta}{1+\beta} \\ P(\alpha|\lambda_1,\lambda_2,..\lambda_N,x_1,...x_N,\beta) &= P(\alpha|\lambda_1,...\lambda_N,\beta) \\ &= \prod_{i=1}^N P(\lambda_i|\alpha,\beta)P(\alpha)P(\beta) \\ &= \prod_{i=1}^N Gamma(\lambda_i|\alpha,\beta)Gamma(\alpha|a,b)Gamma(\beta|c,d) \end{split}$$

The above CP is not in closed form.

We may use MH to sample from this.

$$P(\beta|\lambda_{1},..\lambda_{N},x_{1},...x_{N},\alpha) = P(\beta|\lambda_{1},..\lambda_{N},\alpha)$$

$$= \prod_{i=1}^{N} P(\lambda_{i}|\alpha,\beta)P(\beta)P(\alpha)$$

$$= \prod_{i=1}^{N} Gamma(i|\alpha,\beta)Gamma(\alpha|a,b)Gamma(\beta|c,d)$$

The above CP is also not in closed form.

question assigned to the following page: $ frac{4}{}$	

QUESTION

4

Student Name: Subhajyoti Saha

Roll Number: 21111269 *Date:* November 18, 2022

$$\begin{split} P(r_{ij}|R) &\approx P(r_{ij}|R) \\ &= \frac{1}{S} \sum_{s=1}^{S} P(r_{ij}|u_{i}^{(s)}, v_{j}^{(s)}) \\ &= \frac{1}{S} \sum_{s=1}^{S} N(r_{ij}|u_{i}^{(s)T}v_{j}^{(s)}, \beta^{-1}) \\ &\Rightarrow E[r_{ij}|R)] = \int r_{ij}p(r_{ij}|R)dr_{ij} \\ &= \frac{1}{S} \int r_{ij} \sum_{s=1}^{S} N(r_{ij}|u_{i}^{(s)T}v_{j}^{(s)}, \beta^{-1})dr_{ij} \\ &= \frac{1}{S} \sum_{s=1}^{S} \int r_{ij}N(r_{ij}|u_{i}^{(s)T}v_{j}^{(s)}, \beta^{-1})dr_{ij} \\ &= \frac{1}{S} \sum_{s=1}^{S} (u_{i}^{(s)T}v_{j}^{(s)}) \\ &\Rightarrow Var_{P(r_{ij}|R}[r_{ij}|R)] = Var[u_{i}^{T}v_{j} + \epsilon] \\ &= Var[\epsilon] \text{ [The other one is const.]} \\ &= \frac{1}{\beta} \end{split}$$