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1. Given that

$$p(f_n | \mathbf{x}_n, \mathbf{Z}, \mathbf{t}) = \mathcal{N}(\mathbf{x}_n | \mathbf{k}_n^T \mathbf{K}_M^{-1} \mathbf{t}, \kappa(x_n, x_n) - \mathbf{k}_n^T \mathbf{K}_M^{-1} \mathbf{k}_n)$$

$$p(\mathbf{f} | \mathbf{X}, \mathbf{Z}, \mathbf{t}) = \prod_{n=1}^N p(f_n | \mathbf{x}_n, \mathbf{Z}, \mathbf{t})$$

$$= \mathcal{N}(\mathbf{f} | \mathbf{K}_{NM} \mathbf{K}_M^{-1} \mathbf{t}, \mathbf{\Lambda})$$

Here  $\mathbf{K}_{NM}$  is  $N \times M$  matrix with  $[\mathbf{K}_{NM}]_{nm} = \kappa(\mathbf{x}_n, \mathbf{z}_m)$  and  $\mathbf{K}_M$  is  $M \times M$  matrix with  $[\mathbf{K}_M]_{mm} = \kappa(\mathbf{z}_m, \mathbf{z}_m)$ . Also  $\mathbf{\Lambda}$  is a diagonal matrix with  $[\mathbf{\Lambda}]_{ii} = \kappa(\mathbf{x}_i, \mathbf{x}_i) - \mathbf{k}_i^T \mathbf{K}_M^{-1} \mathbf{k}_i$ . Also

$$p(y_* | \mathbf{x}_*, \mathbf{X}, \mathbf{f}, \mathbf{Z}) = \int p(y_* | \mathbf{x}_*, \mathbf{X}, \mathbf{f}, \mathbf{Z}, \mathbf{t}) p(\mathbf{t} | \mathbf{X}, \mathbf{f}, \mathbf{Z}) d\mathbf{t}$$

Using Bayes's rule to get posterior over  $\mathbf{t}$ , we have:

$$p(\mathbf{t} | \mathbf{X}, \mathbf{f}, \mathbf{Z}) \propto p(\mathbf{f} | \mathbf{X}, \mathbf{t}, \mathbf{Z}) p(\mathbf{t} | \mathbf{Z})$$

Since the pseudo sample points are modelled by same G.P., we have  $p(\mathbf{t} | \mathbf{Z}) = \mathcal{N}(\mathbf{t} | 0, \mathbf{K}_M)$ . Writing the terms in exponent in the R.H.S. of the above proportionality in information form of Gaussian, we get:

$$p(\mathbf{t} | \mathbf{X}, \mathbf{f}, \mathbf{Z}) = \mathcal{N}(\mathbf{t} | \boldsymbol{\mu}_{\mathbf{t}|\mathbf{f}}, \boldsymbol{\Sigma}_{\mathbf{t}|\mathbf{f}})$$

Where  $\boldsymbol{\Sigma}_{\mathbf{t}|\mathbf{f}} = (\mathbf{K}_M^{-1} \mathbf{K}_{NM}^T \mathbf{\Lambda}^{-1} \mathbf{K}_{NM} \mathbf{K}_M^{-1} + \mathbf{K}_M^{-1})^{-1}$  and  $\boldsymbol{\mu}_{\mathbf{t}|\mathbf{f}} = \boldsymbol{\Sigma}_{\mathbf{t}|\mathbf{f}} \mathbf{K}_M^{-1} \mathbf{K}_{NM}^T \mathbf{\Lambda}^{-1} \mathbf{f}$ .

Since  $y_* = f_*$ . We can write  $f_* = \mathbf{k}_*^T \mathbf{K}_M^{-1} \mathbf{t} + \epsilon$ , where  $\mathbf{k}_*$  is  $M \times 1$  vector with  $[\mathbf{k}_*]_i = \kappa(\mathbf{x}_*, \mathbf{z}_i)$ , and  $\epsilon = \mathcal{N}(0, \kappa(\mathbf{x}_*, \mathbf{x}_*) - \mathbf{k}_*^T \mathbf{K}_M^{-1} \mathbf{k}_*)$ . Now using the property of linear gaussian model we get,

$$p(f_* | \mathbf{x}_*, \mathbf{X}, \mathbf{f}, \mathbf{Z}) = \mathcal{N}(f_* | \boldsymbol{\mu}_*, \boldsymbol{\Sigma}_*)$$

$$\text{where } \boldsymbol{\mu}_* = \mathbf{k}_*^T \mathbf{K}_M^{-1} \boldsymbol{\Sigma}_{\mathbf{t}|\mathbf{f}} \mathbf{K}_M^{-1} \mathbf{K}_{NM}^T \mathbf{\Lambda}^{-1} \mathbf{f}$$

$$\text{and } \boldsymbol{\Sigma}_* = \mathbf{k}_*^T \mathbf{K}_M^{-1} \boldsymbol{\Sigma}_{\mathbf{t}|\mathbf{f}} \mathbf{K}_M^{-1} \mathbf{k}_* + \kappa(\mathbf{x}_*, \mathbf{x}_*) - \mathbf{k}_*^T \mathbf{K}_M^{-1} \mathbf{k}_*$$

Note that the computation of posterior predictive is mainly dominated by the term  $\boldsymbol{\Sigma}_{\mathbf{t}|\mathbf{f}}$  whose computation cost is now  $\mathcal{O}(M^2 N)$  which is much less than the earlier version of GP (provided  $M \ll N$ ) where the computation cost for inversion of  $\mathbf{K}_N$  in posterior predictive was  $\mathcal{O}(N^3)$ .

2.

$$p(\mathbf{f} | \mathbf{X}, \mathbf{Z}) = \int p(\mathbf{f} | \mathbf{X}, \mathbf{Z}, \mathbf{t}) p(\mathbf{t} | \mathbf{Z}) d\mathbf{t}$$

Note that we can also write  $\mathbf{f} = \mathbf{K}_{NM} \mathbf{K}_M^{-1} \mathbf{t} + \boldsymbol{\epsilon}$  where  $\boldsymbol{\epsilon} = \mathcal{N}(0, \mathbf{\Lambda})$ .

Again using property of linear gaussian model we have,

$$p(\mathbf{f} | \mathbf{X}, \mathbf{Z}) = \mathcal{N}(\mathbf{f} | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$\text{where } \boldsymbol{\mu} = \mathbf{K}_{NM} \mathbf{K}_M^{-1} \times 0 = 0$$

$$\text{and } \boldsymbol{\Sigma} = \mathbf{K}_{NM} \mathbf{K}_M^{-1} \mathbf{K}_{NM}^T + \boldsymbol{\Lambda}$$

Hence to solve for  $\mathbf{Z}$  via **MLE-II**, we have the following objective:

$$\begin{aligned} \hat{\mathbf{Z}} &= \arg \max_{\mathbf{Z}} p(\mathbf{f} | \mathbf{X}, \mathbf{Z}) \\ &= \arg \max_{\mathbf{Z}} \left( -\frac{1}{2} \log |\boldsymbol{\Sigma}| - \frac{1}{2} \mathbf{f}^T \boldsymbol{\Sigma}^{-1} \mathbf{f} \right) \end{aligned}$$

Note that this objective can be solved using gradient ascent.

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**Flavor 1.**

$$p(\mathbf{x}_n | c_n = m, \mathbf{z}_n, \Theta) = \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_m + \mathbf{W}_m \mathbf{z}_n, \sigma_m^2 \mathbf{I}_D)$$

$$p(\mathbf{z}_n | c_n = m, \Theta) = \mathcal{N}(\mathbf{z}_n | 0, \mathbf{I}_K)$$

Here  $\Theta = \{\pi_m, \boldsymbol{\mu}_m, \mathbf{W}_m, \sigma_m^2\}_{m=1}^M$

Given  $c_n = m$ ,  $\mathbf{x}_n$  can be written as  $\mathbf{x}_n = \boldsymbol{\mu}_m + \mathbf{W}_m \mathbf{z}_n + \boldsymbol{\epsilon}_n$  where  $\boldsymbol{\epsilon}_n = \mathcal{N}(0, \sigma_m^2 \mathbf{I}_D)$ .

Using property of linear gaussian model, we have

$$p(\mathbf{x}_n | c_n = m, \Theta) = \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_m, \mathbf{K}_m)$$

where  $\mathbf{K}_m = \mathbf{W}_m \mathbf{W}_m^T + \sigma_m^2 \mathbf{I}_D$ .

The CLL can be written as:

$$\begin{aligned} \log p(\mathbf{X}, \mathbf{c} | \Theta) &= \log \left[ \prod_{n=1}^N \prod_{m=1}^M (p(\mathbf{x}_n | c_n = m, \Theta) p(c_n = m | \Theta))^{c_{nm}} \right] \\ &= \sum_{n=1}^N \sum_{m=1}^M c_{nm} (\log p(\mathbf{x}_n | c_n = m, \Theta) + \log p(c_n = m | \Theta)) \\ &= \sum_{n=1}^N \sum_{m=1}^M c_{nm} \left[ \log \pi_m - \frac{1}{2} (\mathbf{x}_n - \boldsymbol{\mu}_m)^T \mathbf{K}_m^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_m) - \frac{1}{2} \log |\mathbf{K}_m| + \text{const} \right] \end{aligned}$$

Here  $c_{nm} = \mathbb{I}[c_n = m]$ .

Expected CLL:

$$E[\log p(\mathbf{X}, \mathbf{c} | \Theta)] = \sum_{n=1}^N \sum_{m=1}^M E[c_{nm}] \left[ \log \pi_m - \frac{1}{2} (\mathbf{x}_n - \boldsymbol{\mu}_m)^T \mathbf{K}_m^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_m) - \frac{1}{2} \log |\mathbf{K}_m| + \text{const} \right]$$

The only required expectation i.e.  $E[c_{nm}]$  can be calculated as:

$$E[c_{nm}] = 0 \times p(c_n \neq m | \mathbf{x}_n, \Theta) + 1 \times p(c_n = m | \mathbf{x}_n, \Theta)$$

$$\propto p(\mathbf{x}_n | c_n = m, \Theta) p(c_n = m | \Theta)$$

$$\text{Hence } E[c_{nm}] = \gamma_{nm} = \frac{\pi_m \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_m, \mathbf{K}_m)}{\sum_{l=1}^M \pi_l \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_l, \mathbf{K}_l)}$$

Note that the conditional posterior  $p(c_n = m | \mathbf{x}_n, \Theta)$  is same as  $\gamma_{nm}$

The M-step parameter updates are as follows:

$$\begin{aligned} \hat{\boldsymbol{\mu}}_m &= \frac{\sum_{n=1}^N \gamma_{nm} \mathbf{x}_n}{\sum_{n=1}^N \gamma_{nm}} \\ \hat{\pi}_m &= \frac{\sum_{n=1}^N \gamma_{nm}}{N} \end{aligned}$$

$$\hat{\sigma}_m^2 = \frac{1}{D-K} \sum_{k=K+1}^D \lambda_k$$

$$\hat{\mathbf{W}}_m = \mathbf{U}_K (\mathbf{L}_K - \hat{\sigma}_m^2 \mathbf{I}_K)^{\frac{1}{2}} \mathbf{R}$$

Here  $\mathbf{U}_K$  is  $D \times K$  matrix of top  $K$  eigenvectors of  $\mathbf{S}_m = \frac{\sum_{n=1}^N \gamma_{nm} (\mathbf{x}_n - \hat{\boldsymbol{\mu}}_m)(\mathbf{x}_n - \hat{\boldsymbol{\mu}}_m)^T}{\sum_{n=1}^N \gamma_{nm}}$ ,  $\mathbf{L}_K$ :  $K \times K$  diagonal matrix of top  $K$  eigvalues  $\lambda_1, \dots, \lambda_K$ ,  $\mathbf{R}$  is a  $K \times K$  arbitrary rotation matrix.

**The EM algorithm:**

1. Initialize  $\Theta = \{\pi_m, \boldsymbol{\mu}_m, \mathbf{W}_m, \sigma_m^2\}_{m=1}^M$  as  $\Theta^{(0)}$ , set  $t=1$ ;

2. **E-step:**  $\forall n, m$ :

$$E[c_{nm}^{(t)}] = \gamma_{nm}^{(t)} = \frac{\pi_m^{(t-1)} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_m^{(t-1)}, \mathbf{K}_m^{(t-1)})}{\sum_{l=1}^M \pi_l^{(t-1)} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_l^{(t-1)}, \mathbf{K}_l^{(t-1)})}$$

3. **M-step:**  $\forall m$ :

$$\boldsymbol{\mu}_m^{(t)} = \frac{\sum_{n=1}^N \gamma_{nm}^{(t)} \mathbf{x}_n}{\sum_{n=1}^N \gamma_{nm}^{(t)}}$$

$$\pi_m^{(t)} = \frac{\sum_{n=1}^N \gamma_{nm}^{(t)}}{N}$$

$$\mathbf{S}_m^{(t)} = \frac{\sum_{n=1}^N \gamma_{nm}^{(t)} (\mathbf{x}_n - \boldsymbol{\mu}_m^{(t)}) (\mathbf{x}_n - \boldsymbol{\mu}_m^{(t-1)})^T}{\sum_{n=1}^N \gamma_{nm}^{(t)}}$$

$$\sigma_m^{2(t)} = \frac{1}{D-K} \sum_{k=K+1}^D \lambda_k^{(t)}$$

$$\mathbf{W}_m^{(t)} = \mathbf{U}_K^{(t)} (\mathbf{L}_K^{(t)} - \sigma_m^{2(t)} \mathbf{I}_K)^{\frac{1}{2}} \mathbf{R}^{(t)}$$

4. If not converged, then set  $t=t+1$ , Go to step 2;

**The Stepwise EM algorithm:**

1. Initialize  $\Theta = \{\pi_m, \boldsymbol{\mu}_m, \mathbf{W}_m, \sigma_m^2\}_{m=1}^M$  as  $\Theta^{(0)}$ , set  $t=1$ ;

2. While not converged:

Set learning rate  $\gamma_t$ , pick a random sample  $\mathbf{x}_n$  and compute  $\gamma_{nm}^{(t)} \forall m$ .

Compute  $\hat{\Theta}$  using only this sample

Update  $\Theta^{(t)} = (1 - \gamma_t) \Theta^{(t-1)} + \gamma_t \hat{\Theta}$

Set  $t=t+1$ ;

**Flavor 2.**

$$p(\mathbf{z}_n | \mathbf{x}_n, c_n = m, \Theta) \propto p(\mathbf{x}_n | \mathbf{z}_n, c_n = m, \Theta) p(\mathbf{z}_n | c_n = m)$$

Writing  $\mathbf{x}_n = \boldsymbol{\mu}_m + \mathbf{W}_m \mathbf{z}_n + \boldsymbol{\epsilon}_n$  where  $\boldsymbol{\epsilon}_n = \mathcal{N}(0, \sigma_m^2 \mathbf{I}_D)$  and using properties of Linear gaussian model to get gaussian posterior, we have:

$$p(\mathbf{z}_n | \mathbf{x}_n, c_n = m, \Theta) = \mathcal{N}(\mathbf{z}_n | \boldsymbol{\mu}_{nm}, \boldsymbol{\Sigma}_{nm})$$

$$\text{where } \boldsymbol{\Sigma}_{nm} = \left( \mathbf{I}_K + \frac{\mathbf{W}_m^T \mathbf{W}_m}{\sigma_m^2} \right)^{-1}$$

$$\text{and } \boldsymbol{\mu}_{nm} = \boldsymbol{\Sigma}_{nm} \left( \frac{\mathbf{W}_m^T (\mathbf{x}_n - \boldsymbol{\mu}_m)}{\sigma_m^2} \right)$$

Conditional posterior for  $\mathbf{z}_n$  can be found by summing over possible values of  $c_n$  i.e.:

$$\begin{aligned} p(\mathbf{z}_n|\mathbf{x}_n, \Theta) &= \sum_{m=1}^M p(\mathbf{z}_n|\mathbf{x}_n, c_n = m, \Theta)p(c_n = m|\Theta) \\ &= \sum_{m=1}^M \pi_m \mathcal{N}(\mathbf{z}_n|\boldsymbol{\mu}_{nm}, \boldsymbol{\Sigma}_{nm}) \end{aligned}$$

The conditional posterior for  $c_n$  i.e.  $p(c_n = m|\mathbf{x}_n, \Theta)$  will remain same as in flavor 1. The CLL can be written as :

$$\begin{aligned} \log p(\mathbf{X}, \mathbf{Z}, \mathbf{c}|\Theta) &= \log \left[ \prod_{n=1}^N \prod_{m=1}^M (p(\mathbf{x}_n|\mathbf{z}_n, c_n = m, \Theta)p(\mathbf{z}_n|c_n = m, \Theta)p(c_n = m|\Theta))^{c_{nm}} \right] \\ &= - \sum_{n=1}^N \sum_{m=1}^M c_{nm} \left[ \frac{D}{2} \log \sigma_m^2 + \frac{1}{2\sigma_m^2} \|\mathbf{x}_n - \boldsymbol{\mu}_m\|^2 - \frac{\mathbf{z}_n^T \mathbf{W}_m^T (\mathbf{x}_n - \boldsymbol{\mu}_m)}{\sigma_m^2} + \frac{1}{2\sigma_m^2} \text{Tr}(\mathbf{z}_n \mathbf{z}_n^T \mathbf{W}_m^T \mathbf{W}_m) \right. \\ &\quad \left. + \frac{1}{2} \text{Tr}(\mathbf{z}_n \mathbf{z}_n^T) - \log \pi_m + \text{const} \right] \end{aligned}$$

The Expected CLL is :

$$\begin{aligned} &= - \sum_{n=1}^N \sum_{m=1}^M E[c_{nm}] \left[ \frac{D}{2} \log \sigma_m^2 + \frac{1}{2\sigma_m^2} \|\mathbf{x}_n - \boldsymbol{\mu}_m\|^2 - \frac{E[\mathbf{z}_n]^T \mathbf{W}_m^T (\mathbf{x}_n - \boldsymbol{\mu}_m)}{\sigma_m^2} \right. \\ &\quad \left. + \frac{1}{2\sigma_m^2} \text{Tr}(E[\mathbf{z}_n \mathbf{z}_n^T] \mathbf{W}_m^T \mathbf{W}_m) + \frac{1}{2} \text{Tr}(E[\mathbf{z}_n \mathbf{z}_n^T]) - \log \pi_m + \text{const} \right] \end{aligned}$$

Required expectations are  $E[c_{nm}], E[\mathbf{z}_n|c_n = m], E[\mathbf{z}_n \mathbf{z}_n^T|c_n = m]$ .

$$E[c_{nm}] = \gamma_{nm} = \frac{\pi_m \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_m, \mathbf{K}_m)}{\sum_{l=1}^M \pi_l \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_l, \mathbf{K}_l)}$$

$$E[\mathbf{z}_n|c_n = m] = \boldsymbol{\mu}_{nm}$$

$$E[\mathbf{z}_n \mathbf{z}_n^T|c_n = m] = \boldsymbol{\Sigma}_{nm} + \boldsymbol{\mu}_{nm} \boldsymbol{\mu}_{nm}^T$$

The M-step parameter updates are as follows:

$$\begin{aligned} \hat{\boldsymbol{\mu}}_m &= \frac{\sum_{n=1}^N \gamma_{nm} (\mathbf{x}_n - \mathbf{W}_m E[\mathbf{z}_n|c_n = m])}{\sum_{n=1}^N \gamma_{nm}} \\ \hat{\pi}_m &= \frac{\sum_{n=1}^N \gamma_{nm}}{N} \end{aligned}$$

$$\begin{aligned} \hat{\mathbf{W}}_m &= \left( \sum_{n=1}^N \gamma_{nm} (\mathbf{x}_n - \boldsymbol{\mu}_m) E[\mathbf{z}_n|c_n = m]^T \right) \left( \sum_{n=1}^N \gamma_{nm} E[\mathbf{z}_n \mathbf{z}_n^T|c_n = m] \right)^{-1} \\ \hat{\sigma}_m^2 &= \frac{\sum_{n=1}^N \gamma_{nm} \|\mathbf{x}_n - \boldsymbol{\mu}_m\|^2 + \text{Tr}(E[\mathbf{z}_n \mathbf{z}_n^T] \mathbf{W}_m^T \mathbf{W}_m) - 2E[\mathbf{z}_n^T] \mathbf{W}_m^T (\mathbf{x}_n - \boldsymbol{\mu}_m)}{D \sum_{n=1}^N \gamma_{nm}} \end{aligned}$$

**The EM algorithm:**

1. Initialize  $\Theta = \{\pi_m, \boldsymbol{\mu}_m, \mathbf{W}_m, \sigma_m^2\}_{m=1}^M$  as  $\Theta^{(0)}$ , set  $t=1$ ;

2. **E-step:**  $\forall n, m$ :

$$E[c_{nm}^{(t)}] = \gamma_{nm}^{(t)} = \frac{\pi_m^{(t-1)} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_m^{(t-1)}, \mathbf{K}_m^{(t-1)})}{\sum_{l=1}^M \pi_l^{(t-1)} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_l^{(t-1)}, \mathbf{K}_l^{(t-1)})}$$

$$\boldsymbol{\Sigma}_{nm}^{(t)} = \left( \mathbf{I}_K + \frac{\mathbf{W}_m^{T(t-1)} \mathbf{W}_m^{(t-1)}}{\sigma_m^{2(t-1)}} \right)^{-1}$$

$$E[\mathbf{z}_n | c_n = m]^{(t)} = \boldsymbol{\mu}_{nm}^{(t)} = \boldsymbol{\Sigma}_{nm}^{(t)} \left( \frac{\mathbf{W}_m^{T(t-1)} (\mathbf{x}_n - \boldsymbol{\mu}_m^{(t-1)})}{\sigma_m^{2(t-1)}} \right)$$

$$E[\mathbf{z}_n \mathbf{z}_n^T | c_n = m]^{(t)} = \boldsymbol{\Sigma}_{nm}^{(t)} + \boldsymbol{\mu}_{nm}^{(t)} \boldsymbol{\mu}_{nm}^{T(t)}$$

3. **M-step:**  $\forall m$ :

$$\boldsymbol{\mu}_m^{(t)} = \frac{\sum_{n=1}^N \gamma_{nm}^{(t)} (\mathbf{x}_n - \mathbf{W}_m^{(t-1)} E[\mathbf{z}_n | c_n = m]^{(t)})}{\sum_{n=1}^N \gamma_{nm}^{(t)}}$$

$$\pi_m^{(t)} = \frac{\sum_{n=1}^N \gamma_{nm}^{(t)}}{N}$$

$$\mathbf{W}_m^{(t)} = \left( \sum_{n=1}^N \gamma_{nm}^{(t)} (\mathbf{x}_n - \boldsymbol{\mu}_m^{(t)}) E[\mathbf{z}_n | c_n = m]^{T(t)} \right) \left( \sum_{n=1}^N \gamma_{nm}^{(t)} E[\mathbf{z}_n \mathbf{z}_n^T | c_n = m]^{(t)} \right)^{-1}$$

$$\sigma_m^{2(t)} = \frac{\sum_{n=1}^N \gamma_{nm}^{(t)} \left( \left\| \mathbf{x}_n - \hat{\boldsymbol{\mu}}_m^{(t)} \right\|^2 + \text{Tr} (E[\mathbf{z}_n \mathbf{z}_n^T | c_n = m]^{(t)} \mathbf{W}_m^{T(t)} \mathbf{W}_m^{(t)}) - 2 E[\mathbf{z}_n | c_n = m]^{T(t)} \mathbf{W}_m^{T(t)} (\mathbf{x}_n - \boldsymbol{\mu}_m^{(t)}) \right)}{D \sum_{n=1}^N \gamma_{nm}^{(t)}}$$

4. If not converged, then set  $t=t+1$ , Go to step 2;

**The Stepwise EM algorithm:**

1. Initialize  $\Theta = \{\pi_m, \boldsymbol{\mu}_m, \mathbf{W}_m, \sigma_m^2\}_{m=1}^M$  as  $\Theta^{(0)}$ , set  $t=1$ ;

2. While not converged:

Set learning rate  $\gamma_t$ , pick a random sample  $\mathbf{x}_n$

Compute  $\gamma_{nm}^{(t)}, \boldsymbol{\Sigma}_{nm}^{(t)}, E[\mathbf{z}_n | c_n = m]^{(t)}, E[\mathbf{z}_n \mathbf{z}_n^T | c_n = m]^{(t)} \quad \forall m$ .

Compute  $\hat{\Theta}$  using only this sample

Update  $\Theta^{(t)} = (1 - \gamma_t) \Theta^{(t-1)} + \gamma_t \hat{\Theta}$

Set  $t=t+1$ ;

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The joint distribution can be written as:

$$p(\mathbf{y}, \mathbf{w}, \beta, \boldsymbol{\alpha} | \mathbf{X}, \theta) = \left[ \prod_{n=1}^N \mathcal{N}(y_n | \mathbf{w}^T \mathbf{x}_n, \beta^{-1}) \mathcal{N}(\mathbf{w} | 0, \text{diag}(\alpha_1^{-1}, \dots, \alpha_D^{-1})) \text{Gamma}(\beta | a_0, b_0) \prod_{d=1}^D \text{Gamma}(\alpha_d | e_0, f_0) \right]$$

Here  $\boldsymbol{\alpha} = [\alpha_1, \dots, \alpha_D]^T$  and  $\theta = \{a_0, b_0, e_0, f_0\}$  The log-joint can be now written as:

$$\begin{aligned} \log p(\mathbf{y}, \mathbf{w}, \beta, \boldsymbol{\alpha} | \mathbf{X}, \theta) = & \left( \left( a_0 - 1 + \frac{N}{2} \right) \log \beta - \beta \left[ b_0 + \frac{\sum_{n=1}^N (y_n - \mathbf{w}^T \mathbf{x}_n)^2}{2} \right] \right. \\ & \left. + \left( e_0 + \frac{1}{2} - 1 \right) \sum_{d=1}^D \log \alpha_d - \sum_{d=1}^D \alpha_d \left( \frac{w_d^2}{2} \right) - f_0 \sum_{d=1}^D \alpha_d + \text{constant} \right) \end{aligned}$$

Now consider the mean field V.I. approximation of posterior as :

$$p(\mathbf{w}, \beta, \boldsymbol{\alpha} | \mathbf{y}, \mathbf{X}) = q(\mathbf{w} | \lambda_1) q(\beta | \lambda_2) \prod_{d=1}^D q(\alpha_d | \phi_d)$$

VI approximation for  $\beta$  can be found as follows:

$$\begin{aligned} \log q^*(\beta) &= E_{q(\boldsymbol{\alpha})q(\mathbf{w})} [\log p(\mathbf{y}, \mathbf{w}, \beta, \boldsymbol{\alpha} | \mathbf{X}, \theta)] \\ &= \left( a_0 - 1 + \frac{N}{2} \right) \log \beta - \beta \left[ b_0 + \frac{E_{q(\mathbf{w})} \left[ \sum_{n=1}^N (y_n - \mathbf{w}^T \mathbf{x}_n)^2 \right]}{2} \right] + \text{const} \\ &= \left( a_0 - 1 + \frac{N}{2} \right) \log \beta - \beta \left[ b_0 + \frac{E_{q(\mathbf{w})} \left[ (\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w}) \right]}{2} \right] + \text{const} \\ &= \left( a_0 - 1 + \frac{N}{2} \right) \log \beta - \beta \left[ b_0 + \frac{\mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{X} E_{q(\mathbf{w})} [\mathbf{w}] + \text{Tr} (\mathbf{X}^T \mathbf{X} E_{q(\mathbf{w})} [\mathbf{w}\mathbf{w}^T])}{2} \right] + \text{const} \end{aligned}$$

The above form is similar to log of a Gamma Distribution whose shape and rate parameters are as follows:

$$q^*(\beta) = \text{Gamma}(\beta | \tau_1, \tau_2)$$

$$\text{where } \tau_1 = a_0 + \frac{N}{2}$$

$$\text{and } \tau_2 = b_0 + \frac{\mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{X} E_{q(\mathbf{w})} [\mathbf{w}] + \text{Tr} (\mathbf{X}^T \mathbf{X} E_{q(\mathbf{w})} [\mathbf{w}\mathbf{w}^T])}{2}$$

Similarly, solving for  $\boldsymbol{\alpha}$  component-wise

$$\log q^*(\alpha_d) = \left( e_0 + \frac{1}{2} - 1 \right) \log \alpha_d - \alpha_d \left( f_0 + \frac{E_{q(\mathbf{w})} [w_d^2]}{2} \right) + \text{constant} \quad \forall d$$

This is again similar to log of a Gamma Distribution whose shape and rate parameters are as follows:

$$q^*(\alpha_d) = \text{Gamma}(\alpha_d | e_d, f_d)$$

$$\text{where } e_d = e_0 + \frac{1}{2}$$

$$\text{and } f_d = \left( f_0 + \frac{E_{q(\mathbf{w})}[w_d^2]}{2} \right)$$

Now solving for  $\mathbf{w}$ :

$$\log q^*(\mathbf{w}) = E_{q(\beta)} \left[ -\beta \left[ b_0 + \frac{(\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w})}{2} \right] - \frac{1}{2} \mathbf{w}^T \text{diag}(\alpha_1, \dots, \alpha_D) \mathbf{w} \right] + \text{const}$$

$$= \frac{-1}{2} (\mathbf{w}^T [E_{q(\beta)}[\beta] \mathbf{X}^T \mathbf{X} + \text{diag}(E_{q(\alpha_1)}[\alpha_1], \dots, E_{q(\alpha_D)}[\alpha_D])]) \mathbf{w} - 2E_{q(\beta)}[\beta] \mathbf{w}^T \mathbf{X}^T \mathbf{y} + \text{const}$$

Comparing it with log of gaussian distribution in its information form, we get:

$$q^*(\mathbf{w}) = \mathcal{N}(\mathbf{w} | \boldsymbol{\mu}_N, \boldsymbol{\Sigma}_N)$$

$$\text{where } \boldsymbol{\Sigma}_N = (E_{q(\beta)}[\beta] \mathbf{X}^T \mathbf{X} + \text{diag}(E_{q(\alpha_1)}[\alpha_1], \dots, E_{q(\alpha_D)}[\alpha_D]))^{-1}$$

$$\text{and } \boldsymbol{\mu}_N = E_{q(\beta)}[\beta] \boldsymbol{\Sigma}_N \mathbf{X}^T \mathbf{y}$$

Note that all the expectations can be easily found using corresponding distributions. Hence,

$$E_{q(\beta)}[\beta] = \frac{\tau_1}{\tau_2}$$

$$E_{q(\alpha_d)}[\alpha_d] = \frac{e_d}{f_d} \quad \forall d$$

$$E_{q(\mathbf{w})}[\mathbf{w}] = \boldsymbol{\mu}_N$$

$$E_{q(\mathbf{w})}[\mathbf{w}\mathbf{w}^T] = \boldsymbol{\Sigma}_N + \boldsymbol{\mu}_N \boldsymbol{\mu}_N^T$$

$$E_{q(\mathbf{w})}[w_d^2] = [\boldsymbol{\Sigma}_N]_{dd} + \boldsymbol{\mu}_{Nd}^2 \quad \forall d$$

Note that updates of these distributions depend on each-other thus require cyclic updates:

### The Mean Field VI Algorithm

1. Given  $\theta, \mathbf{X}, \mathbf{y}$ , compute the following as:

$$\tau_1 = a_0 + \frac{1}{2}$$

$$e_d = e_0 + \frac{1}{2}$$

2. Set  $t=0$ , Initialize  $\boldsymbol{\Sigma}_N^{(0)} = \mathbf{I}_D$  and  $\boldsymbol{\mu}_N^{(0)} = 0$  so that

$$f_d^{(0)} = f_0 + \frac{1}{2}$$

$$\tau_2^{(0)} = b_0 + \frac{\mathbf{y}^T \mathbf{y} + \text{Tr}(\mathbf{X}^T \mathbf{X})}{2}$$

$$E_{q(\beta)}[\beta]^{(0)} = \frac{\tau_1}{\tau_2^{(0)}}$$



$$E_{q(\alpha_d)}[\alpha_d]^{(0)} = \frac{e_d}{f_d^{(0)}} \quad \forall d$$

**3.**Set  $t=t+1$ ;

$$\Sigma_N^{(t)} = \left( E_{q(\beta)}[\beta]^{(t-1)} \mathbf{X}^T \mathbf{X} + \text{diag}\left( E_{q(\alpha_1)}[\alpha_1]^{(t-1)}, \dots, E_{q(\alpha_D)}[\alpha_D]^{(t-1)} \right) \right)^{-1}$$

$$\boldsymbol{\mu}_N^{(t)} = E_{q(\beta)}[\beta]^{(t-1)} \Sigma_N^{(t)} \mathbf{X}^T \mathbf{y}$$

$$\tau_2^{(t)} = b_0 + \frac{\mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{X} E_{q(\mathbf{w})}[\mathbf{w}]^{(t)} + \text{Tr}\left(\mathbf{X}^T \mathbf{X} E_{q(\mathbf{w})}[\mathbf{w}\mathbf{w}^T]^{(t)}\right)}{2}$$

$$f_d^{(t)} = \left( f_0 + \frac{E_{q(\mathbf{w})}[w_d^2]^{(t)}}{2} \right) \quad \forall d$$

$$E_{q(\beta)}[\beta]^{(t)} = \frac{\tau_1}{\tau_2^{(t)}}$$

$$E_{q(\alpha_d)}[\alpha_d]^{(t)} = \frac{e_d}{f_d^{(t)}} \quad \forall d$$

**4.**If not converged,Go to step 3.

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**1.Score Function Gradient Method:**

$$\nabla_{\phi} \mathcal{L}(q) \approx \frac{1}{S} \sum_{s=1}^S \nabla_{\phi} \log q(\mathbf{w}_s | \phi) (\log p(\mathbf{y}, \mathbf{w}_s | \mathbf{X}) - \log q(\mathbf{w}_s | \phi))$$

Given that  $q(\mathbf{w} | \phi) = \mathcal{N}(\mathbf{w} | \boldsymbol{\mu}, \boldsymbol{\Sigma})$ , we can rewrite  $\mathbf{w} = \boldsymbol{\mu} + \mathbf{L}\mathbf{v}$  where  $\boldsymbol{\Sigma} = \mathbf{L}\mathbf{L}^T$  and  $\mathbf{v} = \mathcal{N}(\mathbf{v} | 0, \mathbf{I}_D)$

$$\log q(\mathbf{w} | \phi) = -\frac{1}{2}(\mathbf{w} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{w} - \boldsymbol{\mu}) - \frac{1}{2} \log |\boldsymbol{\Sigma}| - \frac{D}{2} \log 2\pi$$

Using chain rule:

$$\begin{aligned} \nabla_{\boldsymbol{\mu}} \log q(\mathbf{w} | \phi) &= (\mathbf{L}\mathbf{L}^T)^{-1}(\mathbf{w} - \boldsymbol{\mu}) \\ \nabla_{\mathbf{L}} \log q(\mathbf{w} | \phi) &= (\mathbf{L}\mathbf{L}^T)^{-1}(\mathbf{w} - \boldsymbol{\mu})(\mathbf{w} - \boldsymbol{\mu})^T (\mathbf{L}\mathbf{L}^T)^{-1} \mathbf{L} - \mathbf{L}^{-T} \end{aligned}$$

Also,

$$\begin{aligned} \log p(\mathbf{y}, \mathbf{w} | \mathbf{X}) &= \sum_{n=1}^N \log p(y_n | \mathbf{x}_n, \mathbf{w}) + \log p(\mathbf{w}) \\ &= \sum_{n=1}^N (y_n \mathbf{w}^T \mathbf{x}_n - \log(1 + \exp(y_n \mathbf{w}^T \mathbf{x}_n))) - \frac{\lambda \mathbf{w}^T \mathbf{w}}{2} + \frac{D}{2} \log \frac{\lambda}{2\pi} \end{aligned}$$

Substitute  $\mathbf{w} = \boldsymbol{\mu} + \mathbf{L}\mathbf{v}$  in above expression to get  $\log p(\mathbf{y}, \boldsymbol{\mu} + \mathbf{L}\mathbf{v} | \mathbf{X})$

Hence

$$\begin{aligned} \nabla_{\mathbf{L}} \mathcal{L}(q) &\approx \frac{1}{S} \sum_{s=1}^S ((\mathbf{L}\mathbf{L}^T)^{-1}(\mathbf{w}_s - \boldsymbol{\mu})(\mathbf{w}_s - \boldsymbol{\mu})^T (\mathbf{L}\mathbf{L}^T)^{-1} \mathbf{L} \\ &\quad - \mathbf{L}^{-T}) \left( \sum_{n=1}^N (y_n (\mathbf{w}_s)^T \mathbf{x}_n - \log(1 + \exp(y_n (\mathbf{w}_s)^T \mathbf{x}_n))) - \frac{\lambda (\mathbf{w}_s)^T (\mathbf{w}_s)}{2} + \frac{D}{2} \log \lambda \right. \\ &\quad \left. + \frac{1}{2} (\mathbf{w}_s - \boldsymbol{\mu})^T (\mathbf{L}\mathbf{L}^T)^{-1} (\mathbf{w}_s - \boldsymbol{\mu}) + \frac{1}{2} \log |\mathbf{L}\mathbf{L}^T| \right) \\ \nabla_{\boldsymbol{\mu}} \mathcal{L}(q) &\approx \frac{1}{S} \sum_{s=1}^S ((\mathbf{L}\mathbf{L}^T)^{-1}(\mathbf{w}_s - \boldsymbol{\mu})) \left( \sum_{n=1}^N (y_n (\mathbf{w}_s)^T \mathbf{x}_n - \log(1 + \exp(y_n (\mathbf{w}_s)^T \mathbf{x}_n))) \right. \\ &\quad \left. - \frac{\lambda (\mathbf{w}_s)^T (\mathbf{w}_s)}{2} + \frac{D}{2} \log \lambda + \frac{1}{2} (\mathbf{w}_s - \boldsymbol{\mu})^T (\mathbf{L}\mathbf{L}^T)^{-1} (\mathbf{w}_s - \boldsymbol{\mu}) + \frac{1}{2} \log |\mathbf{L}\mathbf{L}^T| \right) \end{aligned}$$

**The VI algorithm using B=1:**

1. Initialize  $\phi^{(0)} = \{\boldsymbol{\mu}, \mathbf{L}\}^{(0)}$ , choose learning rate  $\eta$ , set  $t=1$ ;
2. Draw  $S$  samples  $\{\mathbf{w}_1, \dots, \mathbf{w}_S\}^{(t)}$  from  $q(\mathbf{w} | \phi^{(t-1)})$ .

3. Pick a random example  $(\mathbf{x}_n, y_n)$ ;  
 4. Using only this chosen example, update

$$\begin{aligned}\nabla_{\boldsymbol{\mu}} \mathcal{L}(q)^{(t)} &\approx \frac{1}{S} \sum_{s=1}^S (\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T})^{-1} (\mathbf{w}_s^{(t)} \\ &\quad - \boldsymbol{\mu}^{(t-1)}) \left( \left( y_n (\mathbf{w}_s^{(t)})^T \mathbf{x}_n - \log(1 + \exp(y_n (\mathbf{w}_s^{(t)})^T \mathbf{x}_n)) \right) - \frac{\lambda (\mathbf{w}_s^{(t)})^T (\mathbf{w}_s^{(t)})}{2} + \frac{D}{2} \log \lambda \right. \\ &\quad \left. + \frac{1}{2} (\mathbf{w}_s^{(t)} - \boldsymbol{\mu}^{(t-1)})^T (\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T})^{-1} (\mathbf{w}_s^{(t)} - \boldsymbol{\mu}^{(t-1)}) + \frac{1}{2} \log |\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T}| \right) \\ \boldsymbol{\mu}^{(t)} &= \boldsymbol{\mu}^{(t-1)} + \eta \nabla_{\boldsymbol{\mu}} \mathcal{L}(q)^{(t)}\end{aligned}$$

$$\begin{aligned}\nabla_{\mathbf{L}} \mathcal{L}(q)^{(t)} &\approx \frac{1}{S} \sum_{s=1}^S \left( (\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T})^{-1} (\mathbf{w}_s^{(t)} - \boldsymbol{\mu}^{(t)}) (\mathbf{w}_s^{(t)} - \boldsymbol{\mu}^{(t)})^T (\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T})^{-1} \mathbf{L}^{(t-1)} \right. \\ &\quad \left. - \mathbf{L}^{(t-1)-T} \right) \left( \left( y_n (\mathbf{w}_s^{(t)})^T \mathbf{x}_n - \log(1 + \exp(y_n (\mathbf{w}_s^{(t)})^T \mathbf{x}_n)) \right) - \frac{\lambda (\mathbf{w}_s^{(t)})^T (\mathbf{w}_s^{(t)})}{2} \right. \\ &\quad \left. + \frac{D}{2} \log \lambda + \frac{1}{2} (\mathbf{w}_s^{(t)} - \boldsymbol{\mu}^{(t)})^T (\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T})^{-1} (\mathbf{w}_s^{(t)} - \boldsymbol{\mu}^{(t)}) + \frac{1}{2} \log |\mathbf{L}^{(t-1)} \mathbf{L}^{(t-1)T}| \right) \\ \mathbf{L}^{(t)} &= \mathbf{L}^{(t-1)} + \eta \nabla_{\mathbf{L}} \mathcal{L}(q)^{(t)}\end{aligned}$$

5. If ELBO not converged, then set  $t=t+1$ , Go to step 2

## 2. Pathwise Gradient Method:

$$\nabla_{\phi} \mathcal{L}(q) \approx \frac{1}{S} \sum_{s=1}^S \nabla_{\phi} \log p(\mathbf{y}, \mathbf{w}_s | \mathbf{X}) - \nabla_{\phi} E_{q(\mathbf{w}|\phi)} [\log q(\mathbf{w}|\phi)]$$

Note that the entropy form is integrable given that  $q(\mathbf{w}|\phi)$  is Gaussian and the value is:

$$-E_{q(\mathbf{w}|\phi)} [\log q(\mathbf{w}|\phi)] = \frac{1}{2} \log (2\pi |\boldsymbol{\Sigma}| e^D)$$

Reparametrizing  $\mathbf{w} = \boldsymbol{\mu} + \mathbf{L}\mathbf{v}$  and finding the required gradients, we have:

$$\begin{aligned}\nabla_{\boldsymbol{\mu}} \log p(\mathbf{y}, \boldsymbol{\mu} + \mathbf{L}\mathbf{v} | \mathbf{X}) &= \sum_{n=1}^N \frac{y_n \mathbf{x}_n}{1 + \exp(y_n (\boldsymbol{\mu} + \mathbf{L}\mathbf{v})^T \mathbf{x}_n)} - \lambda (\boldsymbol{\mu} + \mathbf{L}\mathbf{v}) \\ \nabla_{\mathbf{L}} \log p(\mathbf{y}, \boldsymbol{\mu} + \mathbf{L}\mathbf{v} | \mathbf{X}) &= \left( \sum_{n=1}^N \frac{y_n \mathbf{x}_n}{1 + \exp(y_n (\boldsymbol{\mu} + \mathbf{L}\mathbf{v})^T \mathbf{x}_n)} - \lambda (\boldsymbol{\mu} + \mathbf{L}\mathbf{v}) \right) \mathbf{v}^T \mathbf{1}_D \mathbf{1}_D\end{aligned}$$

Also Gradients of the entropy are:

$$\begin{aligned}\frac{1}{2} \nabla_{\boldsymbol{\mu}} \log (2\pi |\mathbf{L}\mathbf{L}^T| e^D) &= 0 \\ \frac{1}{2} \nabla_{\mathbf{L}} \log (2\pi |\mathbf{L}\mathbf{L}^T| e^D) &= \mathbf{L}^{-T}\end{aligned}$$

Summing these up the gradient of ELBO w.r.t.  $\phi'$  comes out to be:

$$\nabla_{\boldsymbol{\mu}} \mathcal{L}(q) \approx \frac{1}{S} \sum_{s=1}^S \left( \sum_{n=1}^N \frac{y_n \mathbf{x}_n}{1 + \exp(y_n(\boldsymbol{\mu} + \mathbf{L}\mathbf{v}_s)^T \mathbf{x}_n)} - \lambda(\boldsymbol{\mu} + \mathbf{L}\mathbf{v}_s) \right)$$

$$\nabla_{\mathbf{L}} \mathcal{L}(q) \approx \left( \frac{1}{S} \sum_{s=1}^S \left( \sum_{n=1}^N \frac{y_n \mathbf{x}_n}{1 + \exp(y_n(\boldsymbol{\mu} + \mathbf{L}\mathbf{v}_s)^T \mathbf{x}_n)} - \lambda(\boldsymbol{\mu} + \mathbf{L}\mathbf{v}_s) \right) \mathbf{v}_s^T \right) + \mathbf{L}^{-T}$$

**The VI algorithm using B=1:**

1. Generate S samples from  $\mathcal{N}(\mathbf{v}|0, \mathbf{I}_D)$ .
2. Initialize  $\phi' = \{\boldsymbol{\mu}, \mathbf{L}\}$  as  $\phi'^{(0)}$ , choose learning rate  $\eta$ , set  $t=1$ ;
3. pick a random example  $(\mathbf{x}_n, y_n)$ ;
4. Using only this chosen example, update

$$\nabla_{\boldsymbol{\mu}} \mathcal{L}(q)^{(t)} \approx \frac{1}{S} \sum_{s=1}^S \left( \frac{y_n \mathbf{x}_n}{1 + \exp(y_n(\boldsymbol{\mu}^{(t-1)} + \mathbf{L}^{(t-1)}\mathbf{v}_s)^T \mathbf{x}_n)} - \lambda(\boldsymbol{\mu}^{(t-1)} + \mathbf{L}^{(t-1)}\mathbf{v}_s) \right)$$

$$\boldsymbol{\mu}^{(t)} = \boldsymbol{\mu}^{(t-1)} + \eta \nabla_{\boldsymbol{\mu}} \mathcal{L}(q)^{(t)}$$

$$\nabla_{\mathbf{L}} \mathcal{L}(q)^{(t-1)} \approx \left( \frac{1}{S} \sum_{s=1}^S \left( \frac{y_n \mathbf{x}_n}{1 + \exp(y_n(\boldsymbol{\mu}^{(t)} + \mathbf{L}^{(t-1)}\mathbf{v}_s)^T \mathbf{x}_n)} - \lambda(\boldsymbol{\mu}^{(t)} + \mathbf{L}^{(t-1)}\mathbf{v}_s) \right) \mathbf{v}_s^T \right) + \mathbf{L}^{(t-1)-T}$$

$$\mathbf{L}^{(t)} = \mathbf{L}^{(t-1)} + \eta \nabla_{\mathbf{L}} \mathcal{L}(q)^{(t)}$$

5. If ELBO not converged, then set  $t=t+1$ , Go to step 3;