

Student Name: Subham Kumar

Roll Number: 160707

Date: February 8, 2019

The MLE objective can be written as:

$$\hat{\theta} = \arg \max_{\theta} \sum_{\mathbf{x} \in \mathbf{X}} \log p(\mathbf{x}|\theta)$$

Here $\mathbf{X} = (\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N)$

$$\mathcal{KL}(p||q) = - \sum_{\mathbf{x} \in \mathbf{X}} p(\mathbf{x}) \log \frac{q(\mathbf{x})}{p(\mathbf{x})}$$

putting $p = p_{data}(\mathbf{x})$ and $q = p(\mathbf{x}|\theta)$, the KL-divergence looks like:

$$\mathcal{KL}(p_{data}||p) = - \sum_{\mathbf{x} \in \mathbf{X}} p_{data}(\mathbf{x}) \log \frac{p(\mathbf{x}|\theta)}{p_{data}(\mathbf{x})}$$

The objective function minimizing the KL-divergence w.r.t. θ can be written as:

$$\begin{aligned} \hat{\theta} &= \arg \min_{\theta} \mathcal{KL}(p||q) \\ &= - \arg \min_{\theta} \sum_{\mathbf{x} \in \mathbf{X}} \log p(\mathbf{x}|\theta) \\ &= \arg \max_{\theta} \sum_{\mathbf{x} \in \mathbf{X}} \log p(\mathbf{x}|\theta) \end{aligned}$$

Observe that both the objectives are the same and hence doing MLE is equivalent to finding θ that minimizes the $\mathcal{KL}(p_{data}||p)$. Note that doing the other way round will give an extra factor of $p(\mathbf{x}|\theta)$ in the numerator minimizing this objective function won't give MLE.

Student Name: Subham Kumar

Roll Number: 160707

Date: February 8, 2019

Moment generating function(M) of univariate gaussian with mean μ and variance σ is $\exp(\mu t + \frac{1}{2}\sigma^2 t^2)$.

Denote $\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i$.

$$\begin{aligned} M_{\bar{x}} &= E[\exp(t\bar{x})] \\ &= E\left[\exp\left(\frac{t}{N} \sum_{i=1}^N x_i\right)\right] \\ &= \prod_{i=1}^N E\left[\exp\left(\frac{t}{N} x_i\right)\right] \quad \text{since data is i.i.d.} \\ &= \prod_{i=1}^N \exp\left(\mu \frac{t}{N} + \frac{1}{2} t^2 \frac{\sigma^2}{N^2}\right) \\ &= \exp\left(\mu t + \frac{1}{2} t^2 \frac{\sigma^2}{N}\right) \end{aligned}$$

which is simply the Moment generating function of a gaussian with mean μ and variance $\frac{\sigma^2}{N}$.

Student Name: Subham Kumar

Roll Number: 160707

Date: February 8, 2019

Let $\bar{x}^{(m)} = \frac{1}{N_m} \sum_{i=1}^{N_m} x_i^{(m)}$. Then using the result from Problem 1 we have:

$$\begin{aligned} p(\bar{x}^{(m)}|\mu_m) &= \mathcal{N}\left(\mu_m, \frac{\sigma^2}{N_m}\right) \\ p(\mu|\mathbf{x}, \mu_0, \sigma_0) &\propto \prod_{m=1}^M p(\bar{x}^{(m)}|\mu_m) p(\mu_m|\mu_0, \sigma_0^2) \\ &= \prod_{m=1}^M \mathcal{N}\left(\bar{x}^{(m)}|\mu_m, \frac{\sigma^2}{N_m}\right) \mathcal{N}(\mu_m|\mu_0, \sigma_0^2) \end{aligned}$$

Using completing the square trick:

$$\begin{aligned} p(\mu|\mathbf{x}, \mu_0, \sigma_0) &= \prod_{m=1}^M \mathcal{N}\left(\mu_m \middle| \frac{\bar{x}^{(m)}\sigma_0^2 + \mu_0\frac{\sigma^2}{N_m}}{\sigma_0^2 + \frac{\sigma^2}{N_m}}, \frac{\sigma_0^2\frac{\sigma^2}{N_m}}{\sigma_0^2 + \frac{\sigma^2}{N_m}}\right) \\ &= \prod_{m=1}^M \mathcal{N}\left(\mu_m \middle| \frac{\sum_{i=1}^{N_m} x_i^{(m)}\sigma_0^2 + \mu_0\sigma^2}{N_m\sigma_0^2 + \sigma^2}, \frac{\sigma_0^2\sigma^2}{N_m\sigma_0^2 + \sigma^2}\right) \end{aligned}$$

Hence posterior distribution of μ_m is $\mathcal{N}\left(\mu_m \middle| \frac{\sum_{i=1}^{N_m} x_i^{(m)}\sigma_0^2 + \mu_0\sigma^2}{N_m\sigma_0^2 + \sigma^2}, \frac{\sigma_0^2\sigma^2}{N_m\sigma_0^2 + \sigma^2}\right)$

$$p(\mathbf{x}|\mu_0, \sigma_0^2, \sigma^2) = \int_{\mu} p(\mathbf{x}|\mu, \mu_0, \sigma_0^2, \sigma^2) p(\mu) d\mu$$

Note that here μ is M dimensional and since each μ_m is i.i.d the above integral can be written as product of M independent integrals of the form:

$$\begin{aligned} &\int_{\mu_m} p(\bar{x}^{(m)}|\mu_m, \mu_0, \sigma_0^2, \sigma^2) p(\mu_m) d\mu_m \\ &= \mathcal{N}\left(\bar{x}^{(m)}|\mu_0, \sigma_0^2 + \frac{\sigma^2}{N_m}\right) \end{aligned}$$

Hence

$$\begin{aligned} p(\mathbf{x}|\mu_0, \sigma_0^2, \sigma^2) &= \prod_{m=1}^M \mathcal{N}\left(\bar{x}^{(m)}|\mu_0, \sigma_0^2 + \frac{\sigma^2}{N_m}\right) \\ \log p(\mathbf{x}|\mu_0, \sigma_0^2, \sigma^2) &= - \sum_{m=1}^M \frac{(\bar{x}^{(m)} - \mu_0)^2}{\sigma_0^2 + \frac{\sigma^2}{N_m}} + \text{constant} \end{aligned}$$

For MLE-II estimate of μ_0 :

$$\begin{aligned} \nabla_{\mu_0} \log p(\mathbf{x}|\mu_0, \sigma_0^2, \sigma^2) &= 0 \\ \iff 2 \sum_{m=1}^M \frac{(\bar{x}^{(m)} - \mu_0)}{\sigma_0^2 + \frac{\sigma^2}{N_m}} &= 0 \end{aligned}$$

$$\Longleftrightarrow \mu_0 = \frac{\sum_{m=1}^M \frac{N_m \bar{x}^{(m)}}{N_m \sigma_0^2 + \sigma^2}}{\sum_{m=1}^M \frac{N_m}{\sigma_0^2 N_m + \sigma^2}}$$

where $\bar{x}^{(m)} = \frac{1}{N_m} \sum_{i=1}^{N_m} x_i^{(m)}$.

Substituting this value of μ_0 in posterior derived in part(2), $\mathcal{N} \left(\mu_m \left| \frac{\sum_{i=1}^{N_m} x_i^{(m)} \sigma_0^2 + \frac{\sum_{m=1}^M \frac{N_m \bar{x}^{(m)}}{N_m \sigma_0^2 + \sigma^2} \sigma^2}{\sum_{m=1}^M \frac{N_m}{\sigma_0^2 N_m + \sigma^2}}}{N_m \sigma_0^2 + \sigma^2}, \frac{\sigma_0^2 \sigma^2}{N_m \sigma_0^2 + \sigma^2} \right. \right)$

There is no change in the form of the solution i.e the posterior is still a normal distribution with a different mean .Moreover, it increases the probability of marginal likelihood of X conditioned on the hyperparameters.

Student Name: Subham Kumar

Roll Number: 160707

Date: February 8, 2019

$$\begin{aligned}
 p(Z_{nk}|\pi_k) &= (\pi_k)^{Z_{nk}}(1 - \pi_k)^{(1-Z_{nk})} \\
 p(\mathbf{Z}|\pi) &= \prod_{k=1}^K \prod_{n=1}^N (\pi_k)^{Z_{nk}}(1 - \pi_k)^{(1-Z_{nk})} \\
 &= \prod_{k=1}^K (\pi_k)^{\sum_{n=1}^N Z_{nk}}(1 - \pi_k)^{(N - \sum_{n=1}^N Z_{nk})} \\
 p(\mathbf{Z}|\pi, \alpha) &= p(\mathbf{Z}|\pi)p(\pi|\alpha) \\
 &= \frac{1}{\text{Beta}(\frac{\alpha}{K}, 1)^K} \prod_{k=1}^K (\pi_k)^{\sum_{n=1}^N Z_{nk} + \frac{\alpha}{K} - 1} (1 - \pi_k)^{(N - \sum_{n=1}^N Z_{nk})} \\
 p(\mathbf{Z}|\alpha) &= \frac{1}{\text{Beta}(\frac{\alpha}{K}, 1)^K} \int_{[0,0,\dots,0]}^{[1,1,\dots,1]} \prod_{k=1}^K (\pi_k)^{\sum_{n=1}^N Z_{nk} + \frac{\alpha}{K} - 1} (1 - \pi_k)^{(N - \sum_{n=1}^N Z_{nk})} d\pi_1 d\pi_2 \dots d\pi_K
 \end{aligned}$$

Note that there will be K-independent integral(each π_k is i.i.d.) and each integral will be the p.d.f. of a beta distribution multiplied by some constant. So final expression will be:

$$\begin{aligned}
 &= \prod_{k=1}^K \frac{\text{Beta}\left(\sum_{n=1}^N Z_{nk} + \frac{\alpha}{K}, N + 1 - \sum_{n=1}^N Z_{nk}\right)}{\text{Beta}(\frac{\alpha}{K}, 1)} \\
 p(Z_{nk} = 1|Z_{-nk}) &= \int p(Z_{nk} = 1|\pi_k)p(\pi_k|Z_{-nk})d\pi_k \\
 &= \int \pi_k p(\pi_k|Z_{-nk})d\pi_k
 \end{aligned}$$

This is expectation of π_k w.r.t. $p(\pi_k|Z_{-nk})$ which is simply its mean. Using Bayes Rule:

$$\begin{aligned}
 p(\pi_k|Z_{-nk}) &\propto p(Z_{-nk}|\pi_k)p(\pi_k) \\
 &\propto (\pi_k)^{\sum_{m=1, m \neq n}^N Z_{mk} + \frac{\alpha}{K} - 1} (1 - \pi_k)^{N - \sum_{m=1, m \neq n}^N Z_{mk} - 1}
 \end{aligned}$$

Hence

$$p(\pi_k|Z_{-nk}) = \text{Beta}\left(\sum_{m=1, m \neq n}^N Z_{mk} + \frac{\alpha}{K}, N - \sum_{m=1, m \neq n}^N Z_{mk}\right)$$

and $p(Z_{nk} = 1|Z_{-nk}) = E[\pi_k] = \frac{\sum_{m=1, m \neq n}^N Z_{mk} + \frac{\alpha}{K}}{\frac{\alpha}{K} + N}$ Hence

$$p(Z_{nk}|Z_{-nk}) = \text{Ber}(p(Z_{nk} = 1|Z_{-nk}))$$

Rewriting the $E[\pi_k]$ as :

$$E[\pi_k] = \frac{(N-1) \frac{\sum_{m=1, m \neq n}^N Z_{mk}}{(N-1)} + \frac{\frac{\alpha}{K}}{(\frac{\alpha}{K}+1)} (\frac{\alpha}{K} + 1)}{(\frac{\alpha}{K} + 1) + (N-1)}$$

This result actually makes sense. Note that here before observing Z_{nk} , $E[Z_{nk}] = p(Z_{nk} = 1) = \frac{\frac{\alpha}{K}}{\frac{\alpha}{K}+1}$. This is our prior belief with $\frac{\alpha}{K} + 1$ samples (unobserved). Taking into account only Z_{nk} the value $\frac{\sum_{m=1, m \neq n}^N Z_{mk}}{(N-1)}$ is accounted by $N-1$ observed samples. Hence the expectation $E[\pi_k]$ is weighted average of our prior and posterior belief.

Expected number of ones in k-th column will be:

$$\begin{aligned} E \left[\sum_{n=1}^N Z_{nk} \right] \\ &= \sum_{n=1}^N E[Z_{nk}] \\ &= \sum_{n=1}^N p(Z_{nk} = 1) \\ &= \frac{N\alpha}{K + \alpha} \end{aligned}$$

Expected number of ones in Z will be:

$$\begin{aligned} E \left[\sum_{k=1}^K \sum_{n=1}^N Z_{nk} \right] \\ &= \sum_{k=1}^K \sum_{n=1}^N E[Z_{nk}] \\ &= \sum_{k=1}^K \sum_{n=1}^N p(Z_{nk} = 1) \\ &= \frac{NK\alpha}{K + \alpha} \end{aligned}$$

Student Name: Subham Kumar

Roll Number: 160707

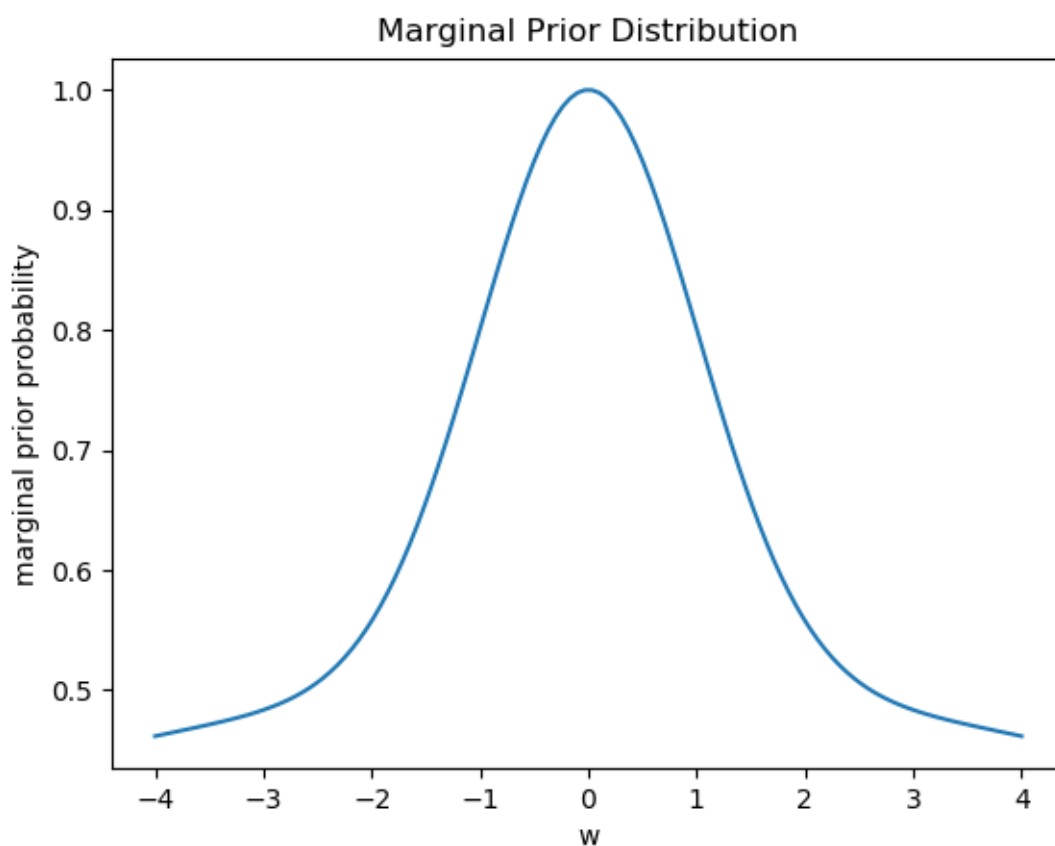
Date: February 8, 2019

a. Since b is discrete:

$$\begin{aligned}
 p(w|\sigma_{spike}^2, \sigma_{slab}^2) &= p(w|b=0, \sigma_{spike}^2, \sigma_{slab}^2)p(b=0) + p(w|b=1, \sigma_{spike}^2, \sigma_{slab}^2)p(b=1) \\
 &= \frac{1}{2} (\mathcal{N}(w|0, \sigma_{spike}^2) + \mathcal{N}(w|0, \sigma_{slab}^2))
 \end{aligned}$$

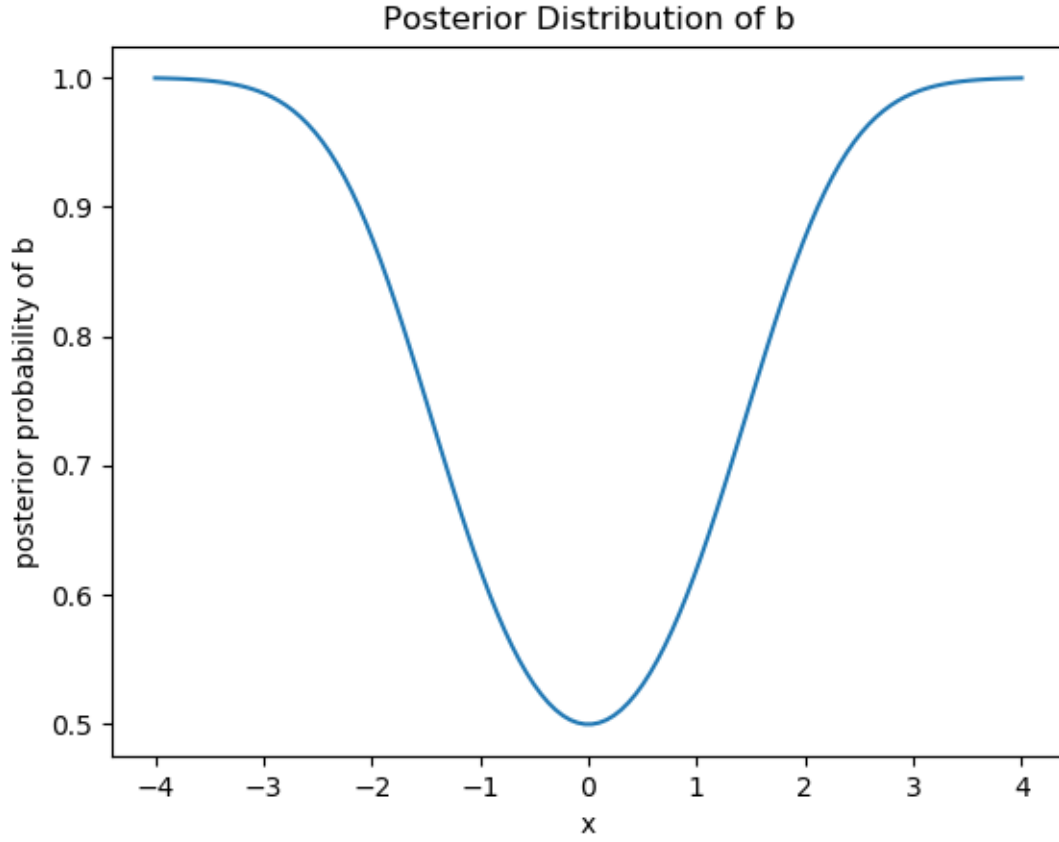
which is mixture of gaussians.

b.



The shape is not as aggressive as $\mathcal{N}(0, 1)$ around the mean in the sense that not all the values of w will be forced to be close to zero. This is a bit fat tailed and less peaky as compared to $\mathcal{N}(0, 1)$.

c.



$$p(b = 1|x, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2) = \frac{p(x|b = 1, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2)p(b = 1)}{p(x|\sigma_{spike}^2, \sigma_{slab}^2, \rho^2)}$$

$$p(x|b = 1, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2) = \mathcal{N}(x|0, \sigma_{spike}^2 + \rho^2)$$

$$p(x|b = 0, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2) = \mathcal{N}(x|0, \sigma_{spike}^2 + \rho^2) \text{ and } p(b = 1) = 0.5$$

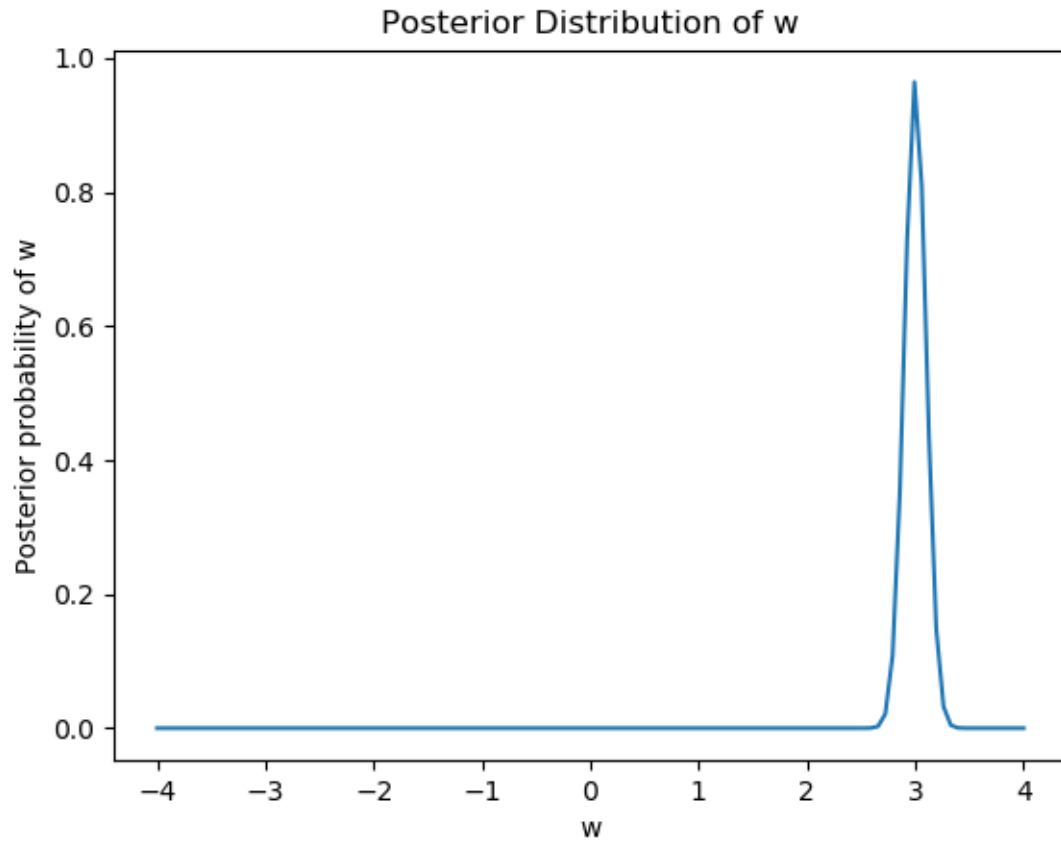
So,

$$p(x|\sigma_{spike}^2, \sigma_{slab}^2, \rho^2) = p(x|b = 0, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2)p(b = 0) + p(x|b = 1, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2)p(b = 1)$$

Hence,

$$p(b = 1|x, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2) = \frac{\mathcal{N}(x|0, \sigma_{slab}^2 + \rho^2)}{\mathcal{N}(x|0, \sigma_{slab}^2 + \rho^2) + \mathcal{N}(x|0, \sigma_{spike}^2 + \rho^2)}$$

d.



$$p(w|x, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2) = \frac{p(x|w, \sigma_{spike}^2, \sigma_{slab}^2, \rho^2)p(w|\sigma_{spike}^2, \sigma_{slab}^2, \rho^2)}{p(x|\sigma_{spike}^2, \sigma_{slab}^2, \rho^2)}$$

$$= \frac{\mathcal{N}(x|w, \rho^2) \left(\mathcal{N}(w|0, \sigma_{spike}^2) + \mathcal{N}(w|0, \sigma_{slab}^2) \right)}{\mathcal{N}(x|0, \sigma_{slab}^2 + \rho^2) + \mathcal{N}(x|0, \sigma_{spike}^2 + \rho^2)}$$

Student Name: Subham Kumar

Roll Number: 160707

Date: February 8, 2019

c. log marginal likelihood for $k = 1$ is -32.35201528

log marginal likelihood for $k = 2$ is -22.77215318

log marginal likelihood for $k = 3$ is -22.07907064

log marginal likelihood for $k = 4$ is -22.38677618

The mapping/model with $k=3$ seems to explain the data best as it has highest log marginal likelihood.

d. log likelihood for $k = 1$ is -28.09400438

log likelihood for $k = 2$ is -15.36066366

log likelihood for $k = 3$ is -10.93584688

log likelihood for $k = 4$ is -7.22529126

The model/mapping with $k=4$ seems to explain the data best as it has highest log likelihood.

The highest log marginal likelihood is more reasonable as it doesn't simply rely on point estimate of w but rather does posterior averaging.

e. Note that region between $[-4, -2.5]$ (roughly) has much higher variance as compared to other places. So getting an x' in this region would be quite helpful in reducing the uncertainty.