

Student Name: Subham Kumar

Roll Number: 160707

Date: August 19, 2019

Since $\mathbf{z}^{(1)}, \mathbf{z}^{(2)}, \dots, \mathbf{z}^{(S)}$ are drawn i.i.d from $p(\mathbf{z})$,

$$E[f(\mathbf{z}^{(i)})] = E[f(\mathbf{z}^{(j)})] = E[f(\mathbf{z})] \quad \forall i \neq j$$

$$\text{var}[f(\mathbf{z}^{(i)})] = \text{var}[f(\mathbf{z}^{(j)})] = \text{var}[f(\mathbf{z})] \quad \forall i \neq j$$

Now,

$$\begin{aligned} E[\hat{f}] &= E\left[\frac{1}{S} \sum_{s=1}^S f(\mathbf{z}^s)\right] \\ &= \frac{1}{S} \sum_{s=1}^S E[f(\mathbf{z}^s)] \quad (\text{using linearity of expectation}) \\ &= \frac{S \times E[f(\mathbf{z})]}{S} = E[f(\mathbf{z})] \end{aligned}$$

Hence monte-carlo approximation is unbiased.

Also,

$$\text{var}[\hat{f}] = \text{var}\left[\frac{1}{S} \sum_{s=1}^S f(\mathbf{z}^s)\right]$$

Using the fact that $\text{var}[a_1 X_1 + a_2 X_2 + \dots + a_s X_s] = a_1^2 \text{var}[X_1] + a_2^2 \text{var}[X_2] + \dots + a_s^2 \text{var}[X_s]$ where X_1, X_2, \dots, X_s are i.i.d., we have,

$$\begin{aligned} \text{var}\left[\frac{1}{S} \sum_{s=1}^S f(\mathbf{z}^{(s)})\right] &= \frac{1}{S^2} \sum_{s=1}^S \text{var}[f(\mathbf{z}^{(s)})] \\ &= \frac{S \times \text{var}[f(\mathbf{z})]}{S^2} = \frac{1}{S} E[(f - E[f])^2] \end{aligned}$$

Student Name: Subham Kumar

Roll Number: 160707

Date: August 19, 2019

For the augmented model $p(y_n, z_n | \mathbf{w}, \mathbf{x}_n, \sigma^2, \nu) = \mathcal{N}(y_n | \mathbf{w}^T \mathbf{x}_n, \frac{\sigma^2}{z_n}) \text{Gamma}(z_n | \frac{\nu}{2}, \frac{\nu}{2})$

$$p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{z}, \rho^2, \sigma^2) \propto p(\mathbf{y} | \mathbf{w}, \mathbf{X}, \mathbf{z}, \sigma^2) p(\mathbf{w} | \rho^2)$$

$$\begin{aligned} & \propto \prod_{n=1}^N p(y_n | \mathbf{w}, \mathbf{x}_n, z_n, \sigma^2) p(\mathbf{w} | \rho^2) \\ & \propto \prod_{n=1}^N \mathcal{N}(y_n | \mathbf{w}^T \mathbf{x}_n, \frac{\sigma^2}{z_n}) \mathcal{N}(\mathbf{w} | 0, \rho^2 \mathbf{I}_D) \\ & \propto \mathcal{N}\left(\mathbf{y} | \mathbf{X} \mathbf{w}, \text{Diag}\left\{\frac{\sigma^2}{z_1}, \frac{\sigma^2}{z_2}, \dots, \frac{\sigma^2}{z_N}\right\}\right) \mathcal{N}(\mathbf{w} | 0, \rho^2 \mathbf{I}_D) \end{aligned}$$

Writing $\mathbf{y} = \mathbf{X} \mathbf{w} + \boldsymbol{\epsilon}$ where $p(\mathbf{w}) = \mathcal{N}(\mathbf{w} | 0, \rho^2 \mathbf{I}_D)$ and $p(\boldsymbol{\epsilon}) = \mathcal{N}(\boldsymbol{\epsilon} | 0, \boldsymbol{\Sigma})$ where $\boldsymbol{\Sigma} = \text{Diag}\left\{\frac{\sigma^2}{z_1}, \frac{\sigma^2}{z_2}, \dots, \frac{\sigma^2}{z_N}\right\}$.
 Now using formula for gaussian conditional posterior, we have $p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{z}, \sigma^2, \rho^2) = \mathcal{N}(\mathbf{w} | \boldsymbol{\mu}_{\mathbf{w} | \mathbf{y}}, \boldsymbol{\Sigma}_{\mathbf{w} | \mathbf{y}})$
 where,

$$\begin{aligned} \boldsymbol{\Sigma}_{\mathbf{w} | \mathbf{y}} &= \left(\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{X} + \frac{\mathbf{I}_D}{\rho^2} \right)^{-1} \\ \boldsymbol{\mu}_{\mathbf{w} | \mathbf{y}} &= \boldsymbol{\Sigma}_{\mathbf{w} | \mathbf{y}} \mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{y} \end{aligned}$$

Note that $\boldsymbol{\Sigma}^{-1} = \text{Diag}\left\{\frac{z_1}{\sigma^2}, \frac{z_2}{\sigma^2}, \dots, \frac{z_N}{\sigma^2}\right\}$

Similarly writing conditional posterior for z_n ,

$$\begin{aligned} p(z_n | \mathbf{y}, \mathbf{w}, \mathbf{X}, \mathbf{z}_{-n}, \sigma^2, \nu) & \propto p(y_n | z_n, \mathbf{w}, \mathbf{x}_n, \sigma^2) p(z_n | \nu) \\ & \propto \mathcal{N}\left(y_n | \mathbf{w}^T \mathbf{x}_n, \frac{\sigma^2}{z_n}\right) \text{Gamma}(z_n | \frac{\nu}{2}, \frac{\nu}{2}) \\ & \propto z_n^{\frac{\nu+1}{2}-1} \exp\left(-z_n \left\{ \frac{(y_n - \mathbf{w}^T \mathbf{x}_n)^2}{2\sigma^2} + \frac{\nu}{2} \right\}\right) \end{aligned}$$

which is similar to Gamma distribution with shape parameter $= \frac{\nu+1}{2}$ and rate parameter $= \left\{ \frac{(y_n - \mathbf{w}^T \mathbf{x}_n)^2}{2\sigma^2} + \frac{\nu}{2} \right\}$.

Hence $p(z_n | \mathbf{y}, \mathbf{w}, \mathbf{X}, \mathbf{z}_{-n}, \sigma^2, \nu) = \text{Gamma}\left(z_n | \frac{\nu+1}{2}, \frac{(y_n - \mathbf{w}^T \mathbf{x}_n)^2}{2\sigma^2} + \frac{\nu}{2}\right)$

The Gibbs Sampling Algorithm:

1. Initialize $\mathbf{w}^{(0)}$ randomly. Set $t=1$

2. While $t \leq T$,

Sample $z_n^{(t)}$ from $\text{Gamma}\left(z_n | \frac{\nu+1}{2}, \frac{(y_n - \mathbf{w}^{(t-1)T} \mathbf{x}_n)^2}{2\sigma^2} + \frac{\nu}{2}\right) \quad \forall n$

Sample $\mathbf{w}^{(t)}$ from $\mathcal{N}\left(\mathbf{w} | \boldsymbol{\mu}_{\mathbf{w} | \mathbf{y}}, \boldsymbol{\Sigma}_{\mathbf{w} | \mathbf{y}}\right)$

Set $t=t+1$

Student Name: Subham Kumar

Roll Number: 160707

Date: August 19, 2019

$$p(z_{d,n} = k | \mathbf{z}_{-d,n}, \mathbf{w}) \propto p(w_{d,n} | z_{d,n} = k, \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n}) p(z_{d,n} = k | \mathbf{z}_{-d,n})$$

The first term in R.H.S. can be written as:

$$\begin{aligned} p(w_{d,n} | z_{d,n} = k, \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n}) &= \int p(w_{d,n} | \phi_k) p(\phi_k | \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n}) d\phi_k \\ &= \int \phi_{k,w_{d,n}} p(\phi_k | \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n}) d\phi_k \\ &= E_{p(\phi_k | \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n})} [\phi_{k,w_{d,n}}] \end{aligned}$$

Now,

$$\begin{aligned} p(\phi_k | \mathbf{z}, \mathbf{w}, \eta) &\propto p(\mathbf{w} | \mathbf{z}, \phi_k) p(\phi_k | \eta) \\ &\propto \prod_{v=1}^V \phi_{k,v}^{\sum_{d=1}^D \sum_{n=1}^{N_d} \mathbb{I}[w_{d,n}=v] \mathbb{I}[z_{d,n}=k] + \eta_v - 1} \end{aligned}$$

which is $Dir(\eta_1, \eta_2, \dots, \eta_V)$ where $\eta_v = \sum_{d=1}^D \sum_{n=1}^{N_d} \mathbb{I}[w_{d,n} = v] \mathbb{I}[z_{d,n} = k] + \eta_v$.

Hence,

$$p(\phi_k | \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n}) = Dir(\eta'_1, \eta'_2, \dots, \eta'_V)$$

where $\eta'_v = \sum_{t=1}^D \sum_{m=1}^{N_d} \mathbb{I}[w_{t,m} = v] \mathbb{I}[z_{t,m} = k] + \eta_v$ not including $(t, m) = (d, n)$.

Hence, $E_{p(\phi_k | \mathbf{z}_{-d,n}, \mathbf{w}_{-d,n})} [\phi_{k,w_{d,n}}] = \frac{\eta'_{w_{d,n}}}{V \eta'_{k,-dn}}$ where $N_{k,-dn} = \sum_{t=1}^D \sum_{m=1}^{N_d} \mathbb{I}[z_{t,m} = k]$ not including $(t, m) = (d, n)$.

Now solving for the second term i.e.

$$\begin{aligned} p(z_{d,n} = k | \mathbf{z}_{-d,n}) &= \int p(z_{d,n} = k | \mathbf{z}_{-d,n}, \boldsymbol{\theta}_d) p(\boldsymbol{\theta}_d | \mathbf{z}_{-d,n}) d\boldsymbol{\theta}_d \\ &= \int \theta_{d,k} p(\boldsymbol{\theta}_d | \mathbf{z}_{-d,n}) d\boldsymbol{\theta}_d \\ &= E_{p(\boldsymbol{\theta}_d | \mathbf{z}_{-d,n})} [\theta_{d,k}] \end{aligned}$$

Now,

$$\begin{aligned} p(\boldsymbol{\theta}_d | \mathbf{z}_{d,1:N_d}) &\propto p(\mathbf{z}_{d,1:N_d} | \boldsymbol{\theta}_d) p(\boldsymbol{\theta}_d | \alpha) \\ &\propto \prod_{k=1}^K \theta_{d,k}^{\sum_{n=1}^{N_d} \mathbb{I}[z_{d,n}=k] + \alpha_k - 1} \end{aligned}$$

which is $Dir(\alpha_1, \alpha_2, \dots, \alpha_K)$ where $\alpha_k = \sum_{n=1}^{N_d} \mathbb{I}[z_{d,n} = k] + \alpha_k$. Hence

$$p(\boldsymbol{\theta}_d | \mathbf{z}_{-d,n}) = Dir(\alpha'_1, \alpha'_2, \dots, \alpha'_K)$$

where $\alpha'_k = \sum_{m=1, m \neq n}^{N_d} \mathbb{I}[z_{d,m} = k] + \alpha$.

Hence, $E_{p(\theta_d | \mathbf{z}_{-d,n})}[\theta_{d,k}] = \frac{\alpha'_k}{K\alpha + N_d - 1}$

Hence

$$p(z_{d,n} = k | \mathbf{z}_{-d,n}, \mathbf{w}) \propto \frac{\eta'_{w_{d,n}}}{V\eta + N_{k,-dn}} \frac{\alpha'_k}{K\alpha + N_d - 1}$$

$$p(z_{d,n} = k | \mathbf{z}_{-d,n}, \mathbf{w}) = \pi_k = \frac{\frac{\eta'_{w_{d,n}}}{V\eta + N_{k,-dn}} \frac{\alpha'_k}{K\alpha + N_d - 1}}{\sum_{l=1}^K \frac{\eta'_{w_{d,n}}}{V\eta + N_{l,-dn}} \frac{\alpha'_l}{K\alpha + N_d - 1}}$$

So, $p(z_{d,n} | \mathbf{z}_{-d,n}, \mathbf{w})$ is a *multinoulli*($\pi_1, \pi_2, \dots, \pi_K$).

The expression implies that the probability of the word $w_{d,n}$ belonging to topic k depends on the proportion of the number of times the words across the document belonged to topic k (excluding current occurrence) and proportion of the number of times the word $w_{d,n}$ across the corpus belonged to topic k (excluding current occurrence). $z_{d,n}$ which is drawn from θ_d depends on the document d , so we look across the document d , whereas for word $w_{d,n}$ we look across entire corpus because it depends on topic vectors which are kind of support for the entire corpus.

The Gibbs Sampling Algorithm:

1. Initialize $\mathbf{z}^{(0)}$, set $t=1$.

2. While $t \leq T$,

Sample $z_{d,n}^{(t)}$ from *multinoulli*($\pi_1, \pi_2, \dots, \pi_K$) $\forall n, d$

Note: Most recent values of $z_{l,m} ((l, m) \neq (d, n))$ are used to calculate $\pi_1, \pi_2, \dots, \pi_K$

Set $t=t+1$

We will use monte-carlo approximation using the samples \mathbf{z} for $E[\theta_d] = (E[\theta_{d,1}], \dots, E[\theta_{d,K}])$ as well as $E[\phi_k] = (E[\phi_{k,1}], \dots, E[\phi_{k,V}])$, where

$$E[\theta_{d,k}] = \frac{1}{S} \sum_{s=1}^S \frac{\alpha + N_{d,k}^{(s)}}{K\alpha + N_d}$$

and $N_{d,k}^{(s)} = \sum_{n=1}^{N_d} \mathbb{I}[z_{d,n}^{(s)} = k]$ which is number of words in document d assigned to topic k given $\mathbf{z}_{d,1:N_d}^{(s)}$. Also

$$E[\phi_{k,v}] = \frac{1}{S} \sum_{s=1}^S \frac{\eta + N_{kv}^{(s)}}{V\eta + N_k^{(s)}}$$

where $N_k^{(s)} = \sum_{d=1}^D \sum_{n=1}^{N_d} \mathbb{I}[z_{d,n}^{(s)} = k]$ (i.e. number of words belonging to topic k over the corpus given $\mathbf{z}^{(s)}$) and $N_{kv}^{(s)} = \sum_{d=1}^D \sum_{n=1}^{N_d} \mathbb{I}[w_{d,n} = v] \mathbb{I}[z_{d,n}^{(s)} = k]$ (i.e. number of times word v belonged to topic k in complete corpus given $\mathbf{z}^{(s)}$).

Student Name: Subham Kumar

Roll Number: 160707

Date: August 19, 2019

Given $p(X_{nm}|\mathbf{u}_n, \mathbf{v}_m) = \text{Poisson}(X_{nm}|\sum_{k=1}^K u_{nk}v_{mk})$. We can augment the representation of X_{nm} as $X_{nm} = \sum_{k=1}^K X_{nmk}$ so that,

$$p(X_{nmk}|u_{nk}, v_{mk}) = \text{Poisson}(X_{nmk}|u_{nk}v_{mk})$$

Note that now we also need to infer these latent variables now.

$$\begin{aligned} p(u_{nk}|X_{n\{1:M\}k}, \mathbf{u}_{-nk}, \Theta) &\propto \prod_{m=1}^M p(X_{nmk}|u_{nk}, v_{mk}, \Theta)p(u_{nk}|a, b) \\ &\propto \prod_{m=1}^M \text{Poisson}(X_{nmk}|u_{nk}, v_{mk})\text{Gamma}(u_{nk}|a, b) \\ &\propto (u_{nk})^{\sum_{m=1}^M X_{nmk}+a-1} \exp\left(-u_{nk}\left(b + \sum_{m=1}^M v_{mk}\right)\right) \end{aligned}$$

which resembles a Gamma distribution with shape parameter $= \sum_{m=1}^M X_{nmk} + a$ and rate parameter $= b + \sum_{m=1}^M v_{mk}$.

Here $\Theta = \{a, b, c, d\}$ which are assumed to be known.

Similarly $p(v_{mk}|X_{\{1:N\}mk}, \mathbf{v}_{-mk}, \Theta) = \text{Gamma}\left(v_{mk}|\sum_{n=1}^N X_{nmk} + c, d + \sum_{n=1}^N u_{nk}\right)$

Lemma: If $\mathbf{X} = (X_1, X_2, \dots, X_K)$ are poisson distributed with parameters $\lambda_1, \lambda_2, \dots, \lambda_K$ respectively, then the joint distribution of \mathbf{X} conditioned on $Y = \sum_{i=1}^K X_i$ will be a multinouli distribution.

Proof:

$$p(\mathbf{X}) = \prod_{i=1}^K \text{Poisson}(X_i|\lambda_i)$$

Also note that $p(Y) = \text{Poisson}(Y|\lambda)$ where $\lambda = \sum_{i=1}^K \lambda_i$ Now

$$p(\mathbf{X}|Y) = \frac{p(\{X_1 = x_1, X_2 = x_2, \dots, X_K = x_k\} \cap \{Y = y\})}{p(Y = y)}$$

$$p(\{X_1 = x_1, X_2 = x_2, \dots, X_K = x_k\} \cap \{Y = y\}) = \begin{cases} 0 & \text{if } y \neq \sum_{i=1}^K x_i \\ p(\mathbf{X}) & \text{otherwise} \end{cases}$$

Hence,

$$p(\mathbf{X}|Y) = \frac{\prod_{i=1}^K \lambda_i^{x_i} \times y!}{\lambda^{\sum_{i=1}^K x_i} x_1! x_2! \dots x_k!} = \frac{y!}{x_1! x_2! \dots x_k!} \prod_{i=1}^K \left(\frac{\lambda_i}{\lambda}\right)^{x_i}$$

which is $\text{Multinoulli}(\mathbf{X}|\frac{\lambda_1}{\lambda}, \frac{\lambda_2}{\lambda}, \dots, \frac{\lambda_K}{\lambda})$

Now to infer latent variables we will use above lemma,

$$p(X_{nm1}, X_{nm2}, \dots, X_{nmK}|X_{nm}, \mathbf{u}, \mathbf{v}, \Theta) = \text{Multinoulli}\left((X_{nm1}, X_{nm2}, \dots, X_{nmK}) \middle| \frac{u_{n1}v_{m1}}{\mathbf{u}_n^T \mathbf{v}_m}, \dots, \frac{u_{nK}v_{mK}}{\mathbf{u}_n^T \mathbf{v}_m}\right)$$

The Gibbs Sampling Algorithm:

1. $\forall n, m, k$ Sample $u_{nk}^{(0)}, v_{mk}^{(0)}$. Set $t=1$;

2. while $t \leq T$,

$\forall n, m$ Sample $\{X_{nmk}\}_{k=1}^K$ from $Multinoulli \left((X_{nm1}, X_{nm2}, \dots, X_{nmK}) \mid \left\{ \frac{u_{n1}v_{m1}}{\mathbf{u}_n^T \mathbf{v}_m}, \dots, \frac{u_{nK}v_{mK}}{\mathbf{u}_n^T \mathbf{v}_m} \right\}^{(t-1)} \right)$

$\forall n, k$ Sample $u_{nk}^{(t)}$ from $Gamma(u_{nk} \mid \sum_{m=1}^M X_{nmk}^{(t)} + a), b + \sum_{m=1}^M v_{mk}^{(t-1)})$

$\forall m, k$ Sample $v_{mk}^{(t)}$ from $Gamma(v_{mk} \mid \sum_{n=1}^N X_{nmk}^{(t)} + c, d + \sum_{n=1}^N u_{nk}^{(t)})$

Set $t=t+1$