# **REGRESSION ANALYSIS PROJECT**

Variable Model Selection



# **Submitted by – Group 8**

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# 1 ABSTRACT

"Essentially, All models are wrong, but some are useful"

— George Box

The main objective of this project is to establish the statistical relationship between policyholder charges and several explanatory variables using Multiple Linear Regression. To accomplish this, we've selected a dataset from Kaggle, which we'll discuss in detail later on.

At first, we sort predictor variables into categories: categorical and non-categorical. Then, we split the dataset into two parts: 70% for training and 30% for testing. We focus our analysis on the training data. We study each numerical predictor variable against the response to see if there's any relationship between them. we have also used added variable plots to pick the variables for the multiple regression model.

We choose the best subset model based on measures like Adjusted  $R^2$ , AIC, and BIC. Then, we conduct outlier and residual analysis to verify the validity of the linear regression assumptions for all of three models. Additionally, we assess whether multicollinearity impacts our model by examining the Variance Inflation Factor(VIF). Afterward, we apply each model to the test dataset and choose the most accurate one based on their MSPR values and other considerations.

# 2 Description of the Dataset

#### 2.1 Source

The data is openly available in multiple online sources. We have collected this specific dataset from Kaggle. The link of the dataset is provided below.

https://www.kaggle.com/datasets/simranjain17/insurance

#### 2.2 About Variables

A brief description of the variables is given below.

Variables	Type	Description
age	Continuous	The Age of the policyholder
sex Character The Gender of the policyholder		The Gender of the policyholder
bmi	Continuous	The Body Mass Index of the Policyholder
children	Integer	Number of Children of the Policyholder
smoker	Character	Indicates whether the Policyholder is Smoker or No Smoker
region	Character	The Region where the Policyholder belongs to
charges	Continuous	The Premium Charged to the Policyholder

Table 1: Description of the variables

# 3 Relationship among the variables

In this section, we will explore the association between the predictors and the response. We have taken **charges** as the response.

## 3.1 Univariate Analysis

#### 1. 'region'

	Region	Count	Proportion
0	southeast	364	27.23 %
1	southwest	325	24.31 %
2	northwest	324	24.23 %
3	northeast	324	24.23 %

#### Distribution of Region

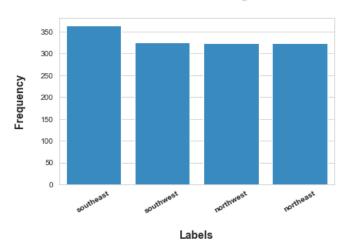


Figure 1: Distribution of Region

**Comment:** From the diagram, we observe that the proportion of policyholders is more or less the same for every region.

## 2. 'children'

	Children	Count	Proportion
0	0	573	42.86 %
1	1	324	24.23 %
2	2	240	17.95 %
3	3	157	11.74 %
4	4	25	1.87 %
5	5	18	1.35 %

# Distribution of Children 500 400 200 100 Labels

Distribution of Smoker

Labels

Figure 2: Distribution of Children

**Comment:** The distribution of children in the diagram is positively skewed, indicating that there are more families with fewer children, while fewer families have a higher number of children.

#### 3. 'smoker'

#### 1000 800 Smoker Count Proportion Frequency 600 1063 79.51 % no 20.49 % yes 274 400 200 ĸρ

Figure 3: Distribution of Smoker

**Comment:** From the diagram, we observe that most of the policyholders are non-smokers.

## 4. 'sex'

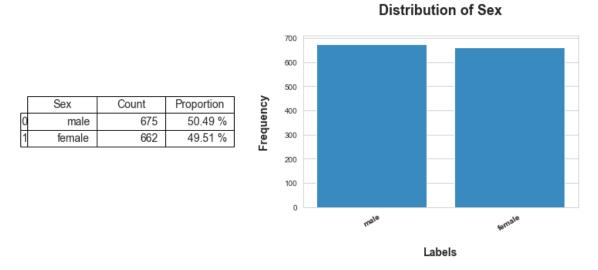


Figure 4: **Distribution of Sex** 

**Comment:** From the diagram, we observe that the proportions of males and females are almost same in this case.

#### 5. 'bmi'

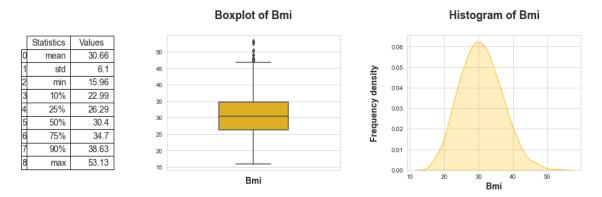


Figure 5: **Distribution of BMI** 

**Comment:** From the boxplot, we see that the distribution is almost symmetrical and the histogram supports that. The presence of some outliers is also evident.

## 6. 'age'

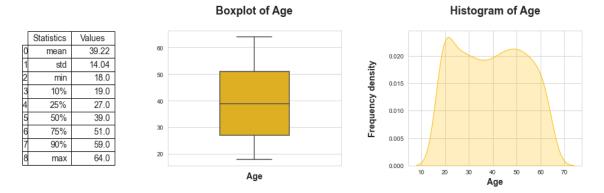


Figure 6: **Distribution of BMI** 

**Comment:** From the above figure, we observe that the distribution is almost symmetric and there are no outliers.

## 7. 'charges'

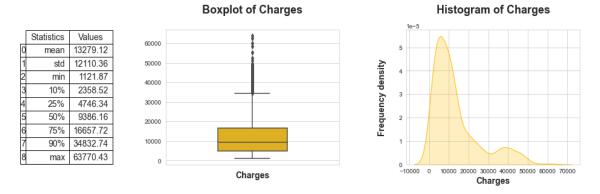


Figure 7: Distribution of Charges

**Comment:** From the above diagram, we observe that the distribution of charges is highly positively skewed with the evidence of the presence of large number of outliers.

## 3.2 Paired plots

Here we are interested in examining the association between the predictors and the response variable.

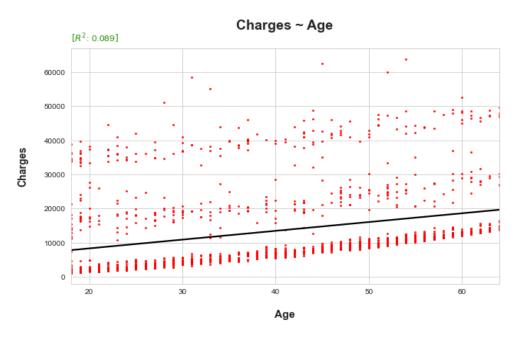


Figure 8: Relationship between Age and Charges

1. **Comment:** From the above plot, we can see that the relationship between Age and Charges is almost linear.

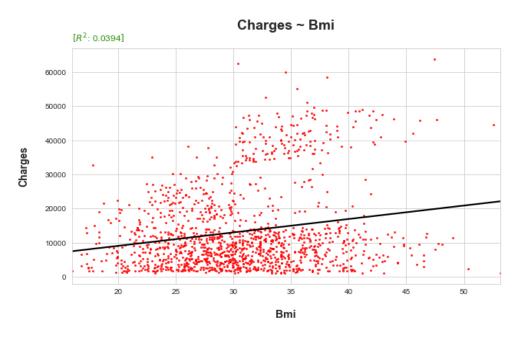


Figure 9: Relationship between BMI and Charges

2. **Comment:** From the above plot, we can see that there is a slight positive correlation between BMI and the response Charges.

# 4 Train-Test Split

Here we add dummy variables corresponding to each categorical column.

Now, we have split the whole dataset into two parts - one consisting of randomly chosen 70% of the rows (called the train data) and the other consisting of the remaining 30% of the rows (called the test data).

The main analysis will be based on the training dataset and models will be fitted using this data.

Later, the validity and accuracy of the models will be verified using the test data.

#### 5 Box-Cox Transformation

Previously we have observed that the distribution of Charges is highly positively skewed. So, fitting a regression line with the original data would be inappropriate as the response will violate the normality assumption. Hence, the Box-Cox transformation is used to make it more symmetric.

**Transformation Formula:** 

$$y^{(\lambda)} = \begin{cases} \frac{y^{\lambda} - 1}{\lambda}, & \text{if } \lambda \neq 0\\ \log_e(y), & \text{if } \lambda = 0 \end{cases}$$

The distribution of Charges before and after the transformation is shown below.

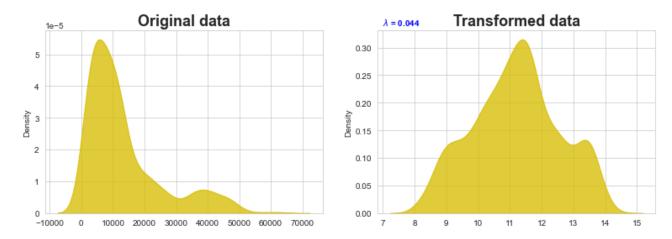


Figure 10: The effect of transformation on the response

**Comment:** Observe that the distribution of the response has become almost symmetrical.

# 6 Box-Tidwell Transformation

After implementing the Box-Cox transformation, from the scatter plot of continuous variables with the response, we observe that there is no significant linear relationship. To get linearity among them, we use the classical Box-Tidwell transformation.

#### **Transformation Formula:**

$$x^* = x^{\alpha}, \quad \alpha \in \mathbb{R}$$

To see how the transformation has helped to attain linearity between the response and the continuous explanatory variables, we look at the scatter plots before and after the transformation.

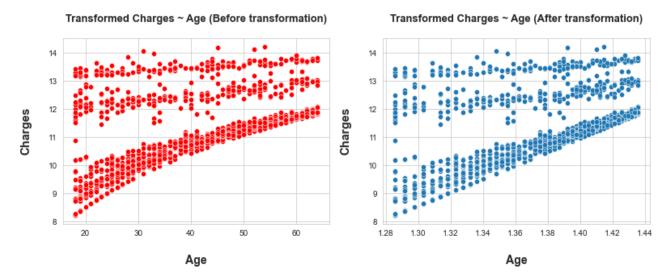


Figure 11: Effect of Box-Tidwell Transformation (Age)

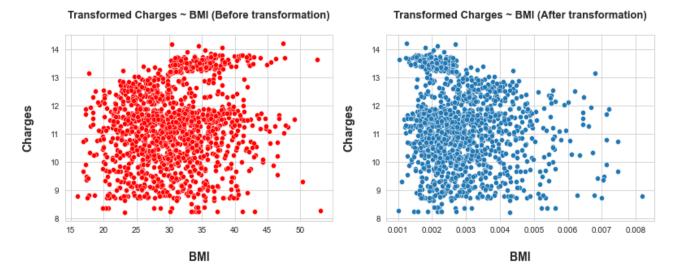


Figure 12: Effect of Box-Tidwell Transformation (BMI)

**Comment:** The above diagrams suggest that the degree of linear relationship has increased for both cases. Further, the nature of the linear relationship between transformed charges and BMI has changed after the transformation.

# 7 Added Variable Plot

# 7.1 Why?

Added variable plots (partial regression plots) are refined residual plots that provide graphic information about the marginal importance of a predictor variable  $X_k$  given the other predictor variables already in the model.

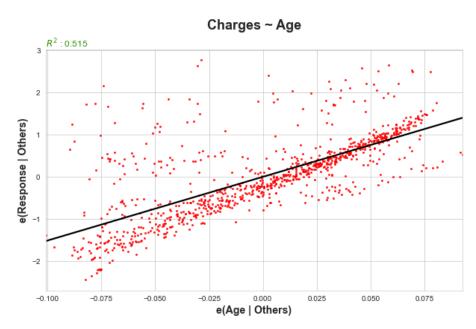


Figure 13: Effect of Box-Tidwell Transformation (BMI)

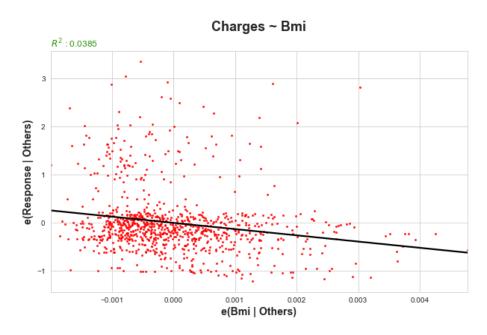


Figure 14: Effect of Box-Tidwell Transformation (BMI)

**Comment:** The Added variable plots of Transformed charges with respect to age and BMI show a linear relationship. Therefore, the marginal influences of these continuous predictors on the response are moderately strong.

#### **8 Model Construction**

In this section, we are aiming to build a robust model with maximum interpretability. To achieve our goal, we will use the Best Subset Selection method to find the most important predictors and consider only those for our final model.

# 8.1 Criteria for Selecting Best Model

#### 8.1.1 Terminology

- n: Total number of rows in the Train data.
- p : Total number of parameters in the model.
- SSE<sub>p</sub>: Sum of Squares due Errors of a p-subset model. (Number of parameters is (p-1))
- SSTO: Total Sum of Squares.
- $\bullet$   $MSE_{full}$  : Mean Square Error of the Full Model.

#### 8.1.2 Chosen Criterion

## 1. Adjusted R<sup>2</sup>

The formula for adjusted  $R^2$  is given by:

$$R_{a,p}^2 = 1 - \left(\frac{n-1}{n-p}\right) \frac{SSE_p}{SSTO}$$

## 2. Akaike Information Criterion (AIC)

The formula of **AIC** is given by:

$$AIC_p = n \log \left( \frac{SSE_p}{n} \right) + 2p$$

## 3. Bayesian Information Criterion (BIC)

The formula of **BIC** is given by:

$$\mathrm{BIC}_p = n\log(\mathrm{SSE}_p) - n\log(n) + [\log(n)]p$$

# 8.2 Selected Models

- $\bullet$  The best subset of predictors obtained by using the  $\mathbf{R_{a,p}^2}$  criterion are listed below:
  - 1. Age
  - 2. Bmi
  - 3. Children
  - 4. Sex\_male
  - 5. Smoker\_yes
  - 6. Region\_northwest
  - 7. Region\_southeast
  - 8. Region\_southwest
- The best subset of predictors obtained by using the **AIC** criterion are listed below:
  - 1. Age
  - 2. Bmi
  - 3. Children
  - 4. Sex\_male
  - 5. Smoker\_yes
  - 6. Region\_southeast
  - 7. Region\_southwest
- The best subset of predictors obtained by using the **BIC** criterion are listed below:
  - 1. Age
  - 2. Bmi
  - 3. Children
  - 4. Smoker\_yes

**Comment:** Notice that,  $R_{a,p}^2$  selection criteria is returning the full set of predictors. Whereas AIC and BIC selection criteria return a proper subset of the predictors. Further, the number of predictors returned by BIC is the lowest.

# 9 Residual Analysis

Here, we will be checking whether the following assumption holds:

$$\epsilon_i \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2), \ \forall \ i \ (\sigma^2 \ \text{is a constant})$$

To verify the assumption, the following plots and tests are done.

## 9.1 Necessary Plots

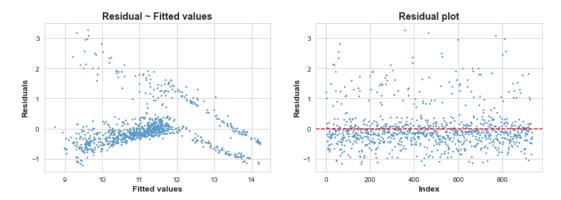


Figure 15: Residual vs. Fitted plot and Residual Plot (For  ${\bf R^2_{a,p}}$ )

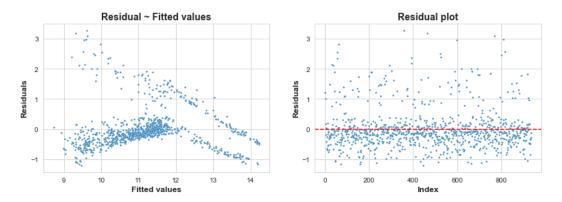


Figure 16: Residual vs. Fitted plot and Residual Plot (For AIC)

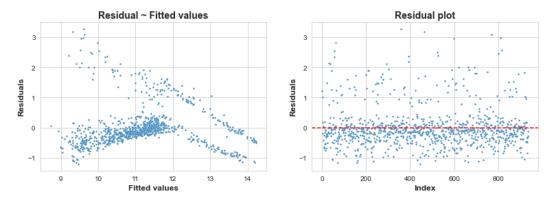


Figure 17: Residual vs. Fitted plot and Residual Plot (For BIC)

**Comment:** As we can observe the plots are quite similar for each of the selection criteria. In each of these cases, the residual versus fitted plot shows a funnel-like shape with some evident clusters. Also, the residual plot indicates the absence of auto-correlation.

#### 9.2 Brown-Forsythe Test

This test is implemented to check the homogeneity of error variance, i.e. to test,

$$H_0:\sigma_i^2=\sigma^2({
m Constant})\ orall\ i=1 (1)n$$
 against  $H_1:{
m At\ least\ one\ inequality}$ 

**Test Statistic** 

$$F = \frac{(N-k)}{(k-1)} \frac{\sum_{j=1}^{k} n_j (\bar{z}_j - \bar{z})^2}{\sum_{j=1}^{k} \sum_{j=1}^{n_j} (z_{ij} - \bar{z}_j)^2}$$

Where,

- $z_{ij} = |y_{ij} \tilde{y}_j|$  ( $\tilde{y}_j$  denotes median of the group  $j, \ j = 1, 2, \dots, k$ )
- k: Number of groups
- ullet  $n_j$ : Number of observations under group j
- $\bullet$  N: Total number of instances
- $\bar{z}_i$ : Mean of group j
- $\bar{z}$  : Overall mean

Here.

- $F \stackrel{H_0}{\sim} F_{k-1,N-k}$
- N = 935
- k = 3
- Groups are of the same size.
- Here we consider 5% level of significance.

#### **Observations:**

- 1. For R<sub>a,p</sub>
  - Value of F-statistic = 15.24
  - p-value =  $3.12 \times 10^{-7}$
- 2. For AIC
  - Value of F-statistic = 15.53
  - p-value =  $2.33 \times 10^{-7}$

## 3. For BIC

- Value of F-statistic = 13.68
- p-value =  $1.41 \times 10^{-6}$

**Comment:** As we can see, in all of the above cases, the assumption of homoscedasticity is getting violated.

# 9.3 Inspecting Normality of Residuals

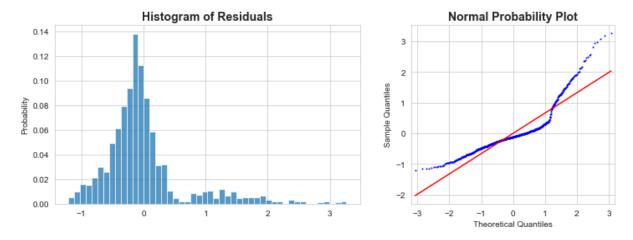


Figure 18: Distribution and the QQ Plot of the Residuals (For  $\mathbf{R}^{2}_{a,p}$ )

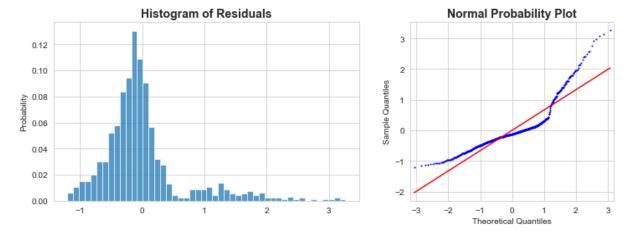


Figure 19: Distribution and the QQ Plot of the Residuals (For AIC)

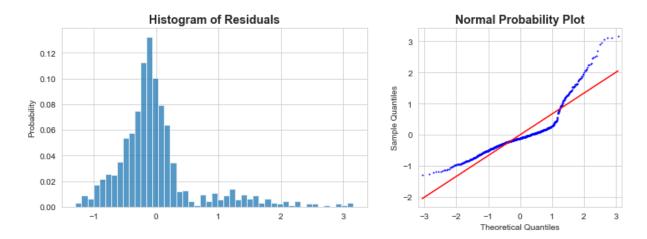


Figure 20: Distribution and the QQ Plot of the Residuals (For BIC)

**Comment:** The above diagrams suggest that the distribution of residuals is positively skewed in each of the cases. So the assumption of normality is violated, which is also evident from the QQ plot.

# 10 Outlier Analysis

#### 10.1 Outlier w.r.t. the Response

Here, our aim is to find out the outliers with respect to the response. By outliers, we mean those observations for which  $|r_i| > 2$ , where  $|r_i|$  is the studentized residual value for  $i^{th}$  observation.

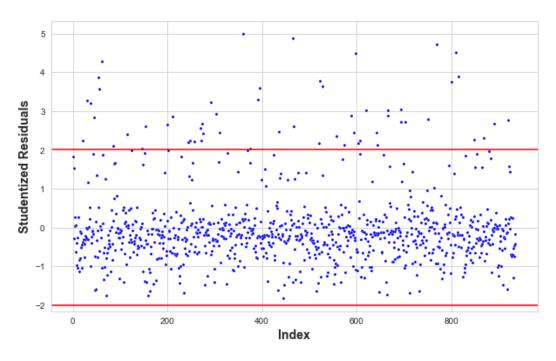


Figure 21: Studentized Residual Plot (For  $\mathbb{R}^2_{a,p}$ )

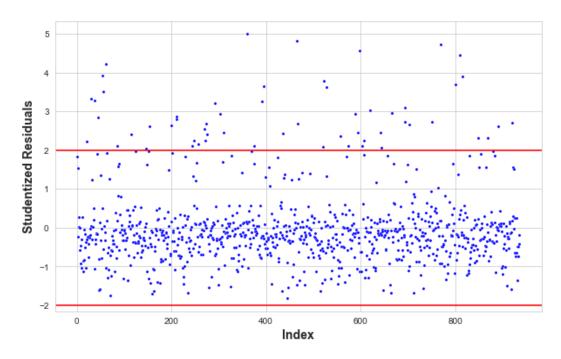


Figure 22: Studentized Residual Plot (For AIC)

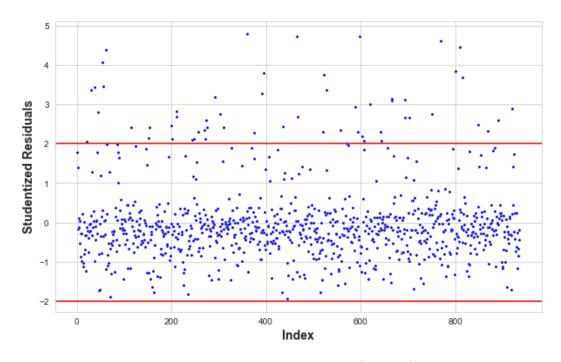


Figure 23: Studentized Residual Plot (For BIC)

**Comment:** The presence of outliers is evident in all of the three models.

#### 10.2 Outlier w.r.t. the Predictors

Here, we compute leverage values for each of the observations for all three models and mark those observations as outliers for which the leverage value is more than  $\frac{2p}{n}$ . (Where p is the number of parameters in the model and n is the number of observations in the data)

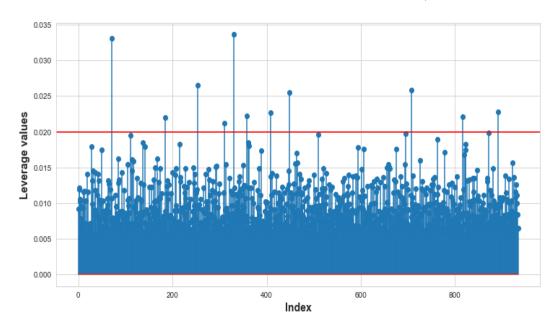


Figure 24: Plot of the Leverages (For  $\mathbb{R}^2_{a,p}$ )

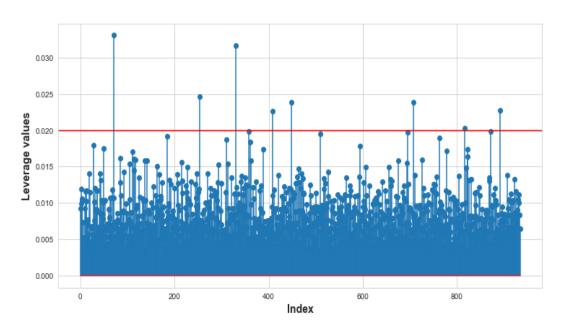


Figure 25: Plot of the Leverages (For AIC)

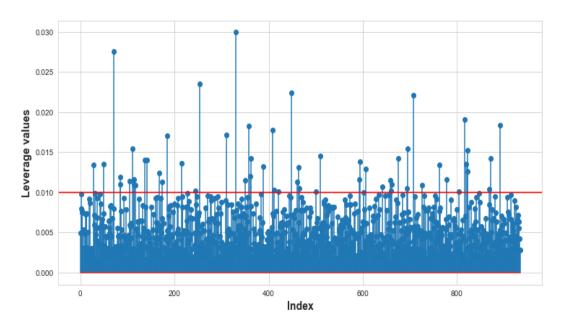


Figure 26: Plot of the Leverages (For BIC)

**Comment:** For the best model w.r.t adjusted  $R^2$  and AIC, the threshold,  $\frac{2p}{n}$  turns out to be 0.02 and that for BIC is 0.01 since, number of parameters is least for BIC. This leads to the occurrence of a greater number of outliers in the case of BIC compared to other models.

#### 10.3 Cook's Distance

Cook's distance is a summary of how much the regression model gets changed when  $i^{th}$  observation is removed. It takes into account both the leverage and residual of each of the observations.

In this section, we will try to find out if the outliers are influential points using Cook's Distance.

#### Formula:

$$D_i = \frac{\sum_{j=1}^{n} (\hat{Y}_j - \hat{Y}_{j(i)})^2}{p \cdot MSE} \sim F_{p,n-p}$$

Where the notations carry their usual meaning.

If the percentile value of F-distribution is less than 10% or 20% for  $i^{th}$  case, then we conclude that  $i^{th}$  case is not much influential. However, if the percentile is 50 or more then it will be influential.

Model	<b>10</b> %	20%
AIC	0.44	0.57
BIC	0.32	0.47
$R_{a,p}^2$	0.46	0.61

Table 2: Table showing the 10<sup>th</sup> and 20<sup>th</sup> percentile of respective F-distribution

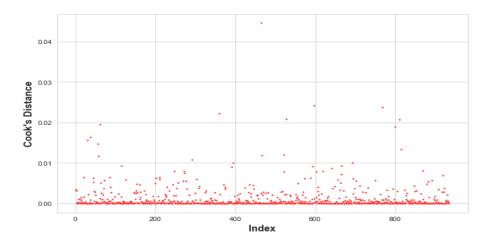


Figure 27: Plot of Cooks Distance (For  $\mathbf{R_{a,p}^2}$ )

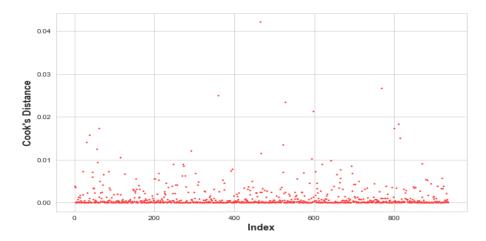


Figure 28: Plot of Cooks Distance (For AIC)

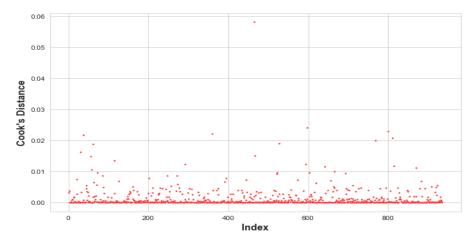


Figure 29: Plot of Cooks Distance (For BIC)

**Comment:** We can observe from the above plots that no influential point exists in the best models obtained by  $R_{a,p}^2$  and AIC criteria. However, there is one influential point present in the case of the best model selected by the BIC criterion.

For further analysis of the model selected by the BIC criterion, we will be removing the influential instance and re-build the model.

The plot of Cook's Distance after removing the influential point is given below.

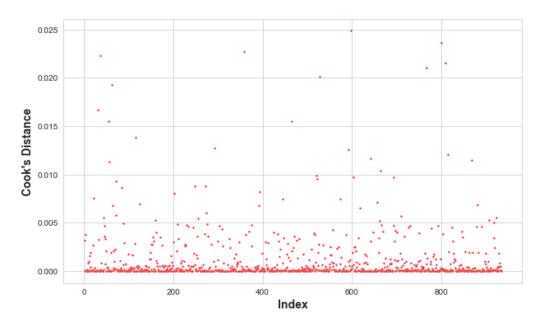


Figure 30: Plot of Cooks Distance (For BIC after removing influential point)

**Comment:** Notice that there is no influential point present here.

# 11 Multicollinearity

By multicollinearity, we mean moderate or high correlation among two or more of the predictors in a regression model.

Here, to check for multicollinearity, we use the VIF technique. Formula:

$$(\text{VIF})_k = \frac{1}{1 - R_k^2}$$

Where  $R_k^2$  is the multiple coefficient of determination when  $X_k$  is regressed on other p-2 predictors,  $(k=1,2,\ldots,p-1)$ 

If  $\max(\text{VIF})_k$  is greater than 15, then we can assume that there is multicollinearity present in the dataset.

#### **Observation**

Features	VIF
Age	9.15
Bmi	8.23
Children	1.83
Smoker_yes	1.26

Table 3: VIF values for different predictors (BIC)

Features	VIF
Age	13.12
Bmi	8.65
Children	1.83
Sex_male	2.11
Smoker_yes	1.28
Region_southeast	1.65
Region_southwest	1.52

Table 4: VIF values for different predictors (AIC)

Features	VIF
Age	15.54
Bmi	8.65
Children	1.83
Sex_male	2.11
Smoker_yes	1.28
Region_northwest	2.07
Region_southeast	2.25
Region_southwest	2.07

Table 5: VIF values for different predictors  $\mathbf{R_{a,p}^2}$ 

**Comment:** From the above tables, it is clear that there is no multicollinearity present among the predictors in the best models selected by AIC and BIC. However, the same is not the case for the model obtained by  $R_{a,p}^2$  selection criterion. In this case, VIF value for the predictor 'Age' is more than 15. So, before proceeding further with model validation, we remove 'Age' from the set of predictors.

#### 12 Model Validation

In this section, we will validate the models by checking their performance on the Test Data. We have taken MSPR as the metric for judging the performance of our models on unseen data.

#### Formula:

MSPR = 
$$\frac{1}{n^*} \sum_{i=1}^{n^*} (Y_i - \hat{Y}_i)^2$$

Where,  $n^*$  is the number of instances in the Test Data.

Model	MSPR
AIC	0.3966
BIC	0.4110
$R_{a,p}^2$	0.9404

Table 6: MSPR values for different models

**Comment:** We can observe that the value of MSPR is significantly higher for  $R_{a,p}^2$  model compared to the other models. This is because of the omission of a significant predictor (viz. 'Age').

## 13 Final Model

In the previous section, we have seen that the MSPR values for the models selected by AIC and BIC are almost similar. Moreover, the best subset of predictors is lesser in the case of BIC than that of AIC which suggests greater interpretability (i.e. lesser complexity).

Considering all the evidence and outcomes discussed above, we choose the best model obtained by BIC as our final model.

**Model:** 
$$Y = -9.87 + 15.21X_1 - 122.16X_2 + 0.11X_3 + 2.32X_4$$

Here,

• 
$$Y = \frac{(\text{charges})^{0.044} - 1}{0.044}$$

• 
$$X_1 = (age)^{0.087}$$

• 
$$X_2 = (bmi)^{-1.735}$$

• 
$$X_3 = \text{Children}$$

• 
$$X_4 = \text{smoker\_yes}$$

# 14 Python Codes

```
# Importing necessary Libraries:
import numpy as np
3 import pandas as pd
4 import matplotlib.pyplot as plt
5 import seaborn as sns
6 from scipy import stats as sts
7 from scipy.stats import f
8 from pandas.plotting import table
9 import statsmodels.api as sm
from statsmodels.formula.api import ols
in from itertools import permutations as pmr
12 from itertools import chain, combinations
from sklearn.model_selection import train_test_split
14 from statsmodels.stats.outliers_influence import variance_inflation_factor
pd.set_option('display.max_rows', None)
sns.set_style('whitegrid')
19 # Calling the data:
20 df = pd.read_csv('D:\Datasets\Insurance\insurance.csv')
22 # Visualizing the first and last 5 rows of the data:
24 print (df.head(5))
25 print('_' *65)
26 print(df.tail(5))
28 # Checking the data:
30 df.shape
31 df.info()
32 df.duplicated()
33 df = df.drop_duplicates()
34 df.shape
df.isnull().sum()
36 df.columns = df.columns.str.title()
37 #-----
38 # Function to plot the univariate distributions:
39 #===
40
     ## CASE I (Categorical)
41
42 def uni_func(x):
     d = df[x].value_counts().reset_index().rename(columns = {'index':x,
43
                                                              x:'Count'})
     d['Proportion'] = (100*d['Count']/sum(d['Count'])).round(2).astype(str) + ' %'
     fig = plt.figure(figsize = (12, 4))
     gs = fig.add_gridspec(1, 2, width_ratios = [1, 2])
48
      # creating the table:
50
     ax1 = fig.add_subplot(gs[0])
51
     plt.axis('off')
52
     t = table(ax1, d, loc = 'center')
53
     t.auto_set_font_size(False)
     t.scale(1.5, 1.8)
     t.set fontsize(14)
  # creating the barplot:
```

```
ax2 = fig.add_subplot(gs[1])
59
       s = sns.barplot(x = x, y = 'Count',
60
              data = d, color = '#218FD6', ax = ax2)
61
       s.set\_title('Distribution of ' + x + ' \n',
62
             fontdict = {'weight':'bold', 'size':18})
63
      s.set_xlabel('\nLabels', fontdict = {'weight':'bold', 'size':14})
64
       s.set\_ylabel('Frequency\n', fontdict = {'weight':'bold', 'size':14})
65
       s.set_xticklabels(s.get_xticklabels(), rotation = 30,
66
                        fontdict = {'weight':'bold'})
67
68
69
      plt.subplots_adjust(wspace = 0.5)
      plt.show()
71
       # Viewing the distributions
72
73 uni_func('Region')
74 uni_func('Children')
uni_func('Sex')
76 uni_func('Smoker')
79
      ## CASE II (Continuous)
80 def cont func(x):
      d = df[x].describe(percentiles = [0.1, 0.25, 0.5, 0.75, 0.9]).drop('count').reset_index(
81
       ).rename(columns = {'index':'Statistics', x:'Values'}).round(2)
82
83
      fig = plt.figure(figsize = (16, 4))
84
      gs = fig.add_gridspec(1, 3, width_ratios = [1, 3, 3])
85
87
       # creating the table:
      ax1 = fig.add_subplot(gs[0])
      plt.axis('off')
89
      t = table(ax1, d, loc = 'center')
       t.auto_set_font_size(False)
      t.scale(1.5, 1.8)
      t.set_fontsize(14)
       # creating the boxplot:
       ax2 = fig.add_subplot(gs[1])
       s = sns.boxplot(y = df[x], ax = ax2, color = '#FDBD04',
97
                      width = 0.4)
       s.set_title('Boxplot of '+ x + ' n',
99
                  fontdict = {'weight':'bold', 'size':18})
100
      s.set_xlabel(x, fontdict = {'weight':'bold', 'size':14})
101
      s.set_vlabel(' ')
102
103
104
       # creating the histogram:
       ax3 = fig.add_subplot(gs[2])
105
       s = sns.kdeplot(x = df[x], fill = True, ax = ax3, color = '#FDBD04')
106
107
       s.set\_title('Histogram of '+ x + ' \n',
108
                  fontdict = {'weight':'bold', 'size':18})
109
       s.set_xlabel(x, fontdict = {'weight':'bold', 'size':14})
       s.set_ylabel('Frequency density\n', fontdict = {'weight':'bold', 'size':14})
110
      plt.subplots_adjust(wspace = 0.5)
      plt.show()
       # Viewing the distributions
cont_func('Age')
117 cont_func('Bmi')
cont_func('Charges')
```

```
120 # Visualizing the relationship between the Predictors and the Response
121
       ## CASE I (Continuous Predictors vs. Continuous Response)
124 def cont_cont(x):
      r_sq = (df['Charges'].corr(df[x])**2).round(4)
125
126
       plt.figure(figsize = (10,6))
       s = sns.regplot(x = x, y = 'Charges',
128
                  data = df, color = '#DD1400', ci = False,
129
                      scatter_kws = {'s':4, "color": "red"}, line_kws = {"color": "black"})
       s.set_title('Charges ~ ' + x + '\n',
131
                   fontdict = {'weight':'bold', 'size':18})
       s.set_xlabel('\n' + x, fontdict = {'weight':'bold',
134
       s.set_ylabel('Charges\n', fontdict = {'weight':'bold',
                                             'size':14})
136
       plt.text(np.min(df[x]), 69000, r'$[R^2$: ' + str(r_sq) + '$]$',
137
                fontsize = 12, color = '#2C9203')
138
       plt.show()
139
140
       # fiting model:
141
       X = sm.add\_constant(df[x])
142
       Y = df['Charges']
143
144
       # Fit the linear regression model
145
       model = sm.OLS(Y, X).fit()
146
147
148
       # View the summary table
149
       print (model.summary())
       ## Visualization:
152 cont_cont('Bmi')
cont_cont('Age')
       ## CASE II (Categorical Predictors vs. Continuous Response)
155
156
       plt.figure(figsize = (10,5))
157
       s = sns.boxplot(x = x, y = 'Charges', data = df, width = 0.5,
158
               flierprops = dict(marker = 'o', markerfacecolor = '#5207C4', markersize = 4),
159
              color = '#18ADE5')
160
       s.set_title('Charges ~ ' + x + '\n',
161
                   fontdict = {'weight':'bold', 'size':18})
162
       s.set_xlabel(' \n' + x, fontdict = {'weight':'bold',}
163
164
                                             'size':14})
       s.set_ylabel('Charges\n', fontdict = {'weight':'bold',
165
166
                                             'size':14})
167
       plt.show()
168
169
      print('-'*100,'\nANOVA Result:')
170
       model = ols('Charges ~ ' + x, data = df).fit()
171
       print(sm.stats.anova_lm(model, type = 2))
       # Visualization:
175 cont_cat('Region')
176 cont_cat('Smoker')
cont_cat('Sex')
178 cont_cat('Children')
```

```
180 # Box-Cox Transformation
181 #==
182 Y = df['Charges']
183 Y_new, lambda_hat = sts.boxcox(Y)
184
fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (13,4))
186 s1 = sns.kdeplot(x = Y, fill = True, alpha = 0.8, ax = ax[0], color = '#D9C008')
sl.set_title('Original data', fontdict = {'weight':'bold', 'size':18})
s1.set_xlabel(' ')
189 s2 = sns.kdeplot(x = Y_new, fill = True, alpha = 0.8, ax = ax[1], color = '#D9C008')
190 s2.set_title('Transformed data', fontdict = {'weight':'bold', 'size':18})
191 plt.text(7, 0.34, r'$\lambda$ = ' + str(round(lambda_hat, 3)),
         fontdict = {'weight':'bold', 'color':'blue'})
194 plt.show()
       # saving the transformed data:
197 df['Y_new'] = Y_new
198 df_copy = df.drop('Charges', axis = 1)
200 # Box-Tidwell Transformation
201 #==
for i in df_copy[['Age','Bmi']].columns:
      d1 = pd.DataFrame({"1":np.ones_like(df_copy[i]),"2":df_copy[i]})
203
       \texttt{d2} = \texttt{pd.DataFrame} ( \texttt{"1":np.ones\_like(df\_copy[i]),"2":df\_copy[i],"3":df\_copy[i]*np.log(df\_copy[i],"3":df\_copy[i] ) } ) 
204
      il)})
      beta1 = np.linalg.lstsq(d1.values, df_copy['Y_new'].values, rcond=None)[0]
205
      beta2 = np.linalg.lstsq(d2.values, df_copy['Y_new'].values, rcond=None)[0]
206
      alpha = (beta2[-1]/beta1[-1])+1
207
      print(round(alpha,3))
208
      df_copy[i] = df_copy[i]**alpha
       # Scatterplot of transformed predictors and transformed Response
211
fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (13,4))
214 \text{ sl} = \text{sns.scatterplot}(x = 'Age', y = 'Y_new', data = df, ax = ax[0],
                       color = 'red')
216 sl.set_title('Transformed Charges ~ Age (Before transformation)' + '\n',
                  fontdict = {'weight':'bold', 'size':13})
s1.set_xlabel('\nAge', fontdict = {'weight':'bold',
219
220 s1.set_ylabel('Charges\n', fontdict = {'weight':'bold',
                                           'size':14})
222 s2 = sns.scatterplot(x = 'Age', y = 'Y_new', data = df_copy, ax = ax[1])
223 s2.set\_title('Transformed Charges ~ Age (After transformation)' + '\n',
                  fontdict = {'weight':'bold', 'size':13})
224
s2.set_xlabel('\nAge', fontdict = {'weight':'bold',
226
                                           'size':14})
s2.set_ylabel('Charges\n', fontdict = {'weight':'bold',
228
                                           'size':14})
229 plt.show()
230
231
233 fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (13,4))
235 \text{ sl} = \text{sns.scatterplot}(x = 'Bmi', y = 'Y_new', data = df, ax = ax[0],
                      color = 'red')
237 s1.set\_title('Transformed Charges ~ BMI (Before transformation)' + '\n',
```

```
fontdict = {'weight':'bold', 'size':13})
239 s1.set_xlabel('\nBMI', fontdict = {'weight':'bold',
                                        'size':14})
240
s1.set_ylabel('Charges\n', fontdict = {'weight':'bold',
                                        'size':14})
242
243 s2 = sns.scatterplot(x = 'Bmi', y = 'Y_new', data = df_copy, ax = ax[1])
s2.set_title('Transformed Charges \tilde{\ } BMI (After transformation)' + '\n',
                 fontdict = {'weight':'bold', 'size':13})
246 s2.set_xlabel('\nBMI', fontdict = {'weight':'bold',
248 s2.set_ylabel('Charges\n', fontdict = {'weight':'bold',
250 plt.show()
251 #=====
252 # Train-Test Split
254 Y = df_copy['Y_new']
255 X = df_copy.drop('Y_new', axis = 1)
257 X_train, X_test, Y_train, Y_test = train_test_split(X, Y, test_size = 0.30, random_state = 42)
259 X train.head(5)
260 X train.shape
261 #=
262 # Added Variable Plot.
264 def add var(x):
      def resid(x,y):
265
         x = sm.add_constant(pd.get_dummies(x, drop_first = True))
266
267
         M = sm.OLS(y,x).fit()
         return M.resid
268
      # Storing the set of other variables:
      X_{new} = X_{train.drop}(x, axis = 1)
      # y ~ x_others
      r1 = resid(X_new, Y_train)
      # x ~ x_others
276
      r2 = resid(X_new, X_train[x])
277
278
      r_sq = (r1.corr(r2)**2).round(4)
279
280
      plt.figure(figsize = (10,6))
281
      s = sns.regplot(x = r2, y = r1, color = '\#DD1400', ci = False,
282
                    scatter_kws = {'s':4, "color": "red"}, line_kws = {"color": "black"})
283
      s.set_title('Charges ~ ' + x + '\n',
284
                 fontdict = {'weight':'bold', 'size':18})
285
286
      s.set_xlabel('e(' + x + ' | Others)', fontdict = {'weight':'bold',
                                        'size':14})
287
288
      s.set_ylabel('e(Response | Others)', fontdict = {'weight':'bold',
                                        'size':14})
289
      plt.text(np.min(r2), np.max(r1) + 0.3, r'$R^2$ : ' + str(r_sq),
290
291
              fontsize = 12, color = '#2C9203')
      plt.show()
      # Visualization:
296 add_var('Age')
297 add_var('Bmi')
```

```
299 # Best Subset Selection
300 #==
301 def powerset(x):
     S = chain.from\_iterable(combinations(x, r) for r in range(len(x)+1))
302
303
     return [list(k) for k in list(S)][1:]
304
305 X_train_dumm = pd.get_dummies(X_train, drop_first = True)
306 C = powerset(X_train_dumm.columns)
308 AIC = []
309 BIC = []
Adj_Rsq = []
Rsq = []
313 for var in C:
     X = sm.add_constant(X_train_dumm[var])
      Y = Y_train
315
316
      M = sm.OLS(Y, X).fit()
317
318
      Adj_Rsq.append(round(M.rsquared_adj,3))
319
      Rsq.append(round(M.rsquared, 3))
320
      AIC.append(round(M.aic,3))
321
      BIC.append(round(M.bic, 3))
322
323
324 best_df = pd.DataFrame({
      'Variable set': C,
325
326
      'Adjusted Rsq': Adj_Rsq,
327
      'R_Square': Rsq,
      'AIC': AIC, 'BIC': BIC
328
329 })
330
      # Best variables concerning different criteria:
331
332 best_subset_vars_adj = best_df.at[best_df['Adjusted Rsq'].idxmax(),'Variable set']
333 best_subset_vars_adj
335 best_subset_vars_BIC = best_df.at[best_df['BIC'].idxmin(),'Variable set']
336 best_subset_vars_BIC
338 best_subset_vars_AIC = best_df.at[best_df['AIC'].idxmin(),'Variable set']
339 best subset vars AIC
340
      # Creating dummy variables:
341
342 X_train_dumm_adj = X_train_dumm[best_subset_vars_adj]
343 X_train_dumm_BIC = X_train_dumm[best_subset_vars_BIC]
344 X_train_dumm_AIC = X_train_dumm[best_subset_vars_AIC]
^{346} # Model fitting with the best subset of variables
347 #===========
348
    # For Adjusted R^2:
350 X = sm.add_constant(X_train_dumm_adj)
Y = Y_train
model_adj = sm.OLS(Y,X).fit()
354 resid_adj = model_adj.resid
yhat_adj = model_adj.fittedvalues
index = np.arange(1,len(resid_adj)+1)
```

```
_{358} fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (13,4))
359 	ext{ s1} = 	ext{sns.scatterplot}(x = 	ext{yhat\_adj}, y = 	ext{resid\_adj}, s = 6, ax = ax[0])
360 sl.set_title('Residual ~ Fitted values', fontdict = {'weight':'bold',
                                                         'size':14})
361
s1.set_xlabel('Fitted values', fontdict = {'weight':'bold',
                                                        'size':12})
363
s1.set_ylabel('Residuals', fontdict = {'weight':'bold', 'size':12})
365 	ext{ s2} = 	ext{sns.scatterplot}(x = 	ext{index}, y = 	ext{resid\_adj}, s = 6, ax = ax[1])
s2.set_title('Residual plot', fontdict = {'weight':'bold',
                                                        'size':14})
s2.set_xlabel('Index', fontdict = {'weight':'bold', 'size':12})
369 s2.set_ylabel('Residuals', fontdict = {'weight':'bold', 'size':12})
plt.axhline(y = 0, color = 'r', linestyle = '--')
      # For BIC:
374 X = sm.add_constant(X_train_dumm_BIC)
Y = Y_train
376 model_BIC = sm.OLS(Y,X).fit()
378 resid_BIC = model_BIC.resid
yhat_BIC = model_BIC.fittedvalues
380
index = np.arange(1, len(resid_adj)+1)
382 fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (13,4))
383 s1 = sns.scatterplot(x = yhat_adj, y = resid_adj, s = 6, ax = ax[0])
s1.set_title('Residual ~ Fitted values', fontdict = {'weight':'bold',
                                                        'size':14})
385
s1.set_xlabel('Fitted values', fontdict = {'weight':'bold',
387
                                                         'size':12})
sl.set_ylabel('Residuals', fontdict = {'weight':'bold', 'size':12})
389 s2 = sns.scatterplot(x = index, y = resid_adj, s = 6, ax = ax[1])
390 s2.set_title('Residual plot', fontdict = {'weight':'bold',
s2.set_xlabel('Index', fontdict = {'weight':'bold', 'size':12})
s2.set_ylabel('Residuals', fontdict = {'weight':'bold', 'size':12})
plt.axhline(y = 0, color = 'r', linestyle = '--')
395 plt.show()
      # For AIC:
397
398 X = sm.add_constant(X_train_dumm_AIC)
399 Y = Y train
400 model_AIC = sm.OLS(Y,X).fit()
401
402 resid_AIC = model_AIC.resid
403 yhat_AIC = model_AIC.fittedvalues
404
index = np.arange(1, len(resid_adj)+1)
fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (13,4))
_{407} s1 = sns.scatterplot(x = yhat_adj, y = resid_adj, s = 6, ax = ax[0])
408 sl.set_title('Residual ~ Fitted values', fontdict = {'weight':'bold',
                                                        'size':14})
s1.set_xlabel('Fitted values', fontdict = {'weight':'bold',
411
412 s1.set_ylabel('Residuals', fontdict = {'weight':'bold', 'size':12})
413 s2 = sns.scatterplot(x = index, y = resid_adj, s = 6, ax = ax[1])
s2.set_title('Residual plot', fontdict = {'weight':'bold',
416 s2.set_xlabel('Index', fontdict = {'weight':'bold', 'size':12})
417 s2.set_ylabel('Residuals', fontdict = {'weight':'bold', 'size':12})
```

```
plt.axhline(y = 0, color = 'r', linestyle = '--')
419 plt.show()
420 #===
# Brown-Forsythe test
k = len(Y_train) // 3
424
             # For BIC:
425
426 d = pd.DataFrame({'Y_fit':yhat_BIC, 'res':resid_BIC})
427 y = d.sort_values(by = 'Y_fit')['res']
_{428} print(sts.levene(y[:k], y[k:2*k], y[2*k:3*k], center = 'median'))
            # For AIC:
d = pd.DataFrame({'Y_fit':yhat_AIC, 'res':resid_AIC})
432 y = d.sort_values(by = 'Y_fit')['res']
433 print(sts.levene(y[:k], y[k:2*k], y[2*k:3*k], center = 'median'))
              # For Adj R^2:
d = pd.DataFrame({'Y_fit':yhat_adj, 'res':resid_adj})
437 y = d.sort_values(by = 'Y_fit')['res']
y = \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} \left( \frac{1}{2} \right) + \frac{1}{2} \left( \frac
439 #=
440 # Check for Normality
441 #=
442 # For BIC:
fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (12,4),
                                                              gridspec_kw = {'width_ratios': [1.5, 1]})
444
45 sns.histplot(resid_BIC, stat = 'probability', ax = ax[0])
sm.qqplot(resid_BIC, line = 's', ax = ax[1])
plt.setp(fig.axes[1].get_lines(), markersize = 1)
ax[0].set_title('Histogram of Residuals', fontdict = {'weight':'bold',
450 ax[1].set_title('Normal Probability Plot', fontdict = {'weight':'bold',
                                                                                                                                                'size':14})
452 plt.show()
454
            # For AIC
455 fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (12,4),
                                                                  gridspec_kw = {'width_ratios': [1.5, 1]})
456
sns.histplot(resid_AIC, stat = 'probability', ax = ax[0])
458 sm.qqplot(resid_AIC, line = 's', ax = ax[1])
459 plt.setp(fig.axes[1].get_lines(), markersize = 1)
460 ax[0].set_title('Histogram of Residuals', fontdict = {'weight':'bold',
461
ax[1].set_title('Normal Probability Plot', fontdict = {'weight':'bold',
463
                                                                                                                                               'size':14})
464 plt.show()
465
466
              # For Adjusted R^2
fig, ax = plt.subplots(nrows = 1, ncols = 2, figsize = (12,4),
468
                                                                 gridspec_kw = {'width_ratios': [1.5, 1]})
sns.histplot(resid_adj, stat = 'probability', ax = ax[0])
470 sm.qqplot(resid_adj, line = 's', ax = ax[1])
471 plt.setp(fig.axes[1].get_lines(), markersize = 1)
472 ax[0].set_title('Histogram of Residuals', fontdict = {'weight':'bold',
474 ax[1].set_title('Normal Probability Plot', fontdict = {'weight':'bold',
                                                                                                                                               'size':14})
476 plt.show()
```

```
479 # Outlier Analysis
480 #===
481
      # For AIC:
482
resid_std_AIC = model_AIC.outlier_test()['student_resid']
index = np.arange(1, len(resid_std_AIC)+1)
486
487 plt.figure(figsize = (10,6))
488 s = sns.scatterplot(x = index, y = resid_std_AIC, s = 10, color = 'blue')
489 plt.axhline(y = -2, color = 'red')
490 plt.axhline(y = 2, color = 'red')
491 s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
492 s.set_ylabel('Studentized Residuals', fontdict = {'weight':'bold', 'size':14})
     # For BIC:
496 resid_std_BIC = model_BIC.outlier_test()['student_resid']
498 index = np.arange(1, len(resid_std_BIC)+1)
499
500 plt.figure(figsize = (10,6))
501 s = sns.scatterplot(x = index, y = resid_std_BIC, s = 10, color = 'blue')
plt.axhline(y = -2, color = 'red')
plt.axhline(y = 2, color = 'red')
source s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
sos s.set_ylabel('Studentized Residuals', fontdict = {'weight':'bold', 'size':14})
506 plt.show()
508
     # For Adjusted R^2
resid_std_adj = model_adj.outlier_test()['student_resid']
510
index = np.arange(1, len(resid_std_adj)+1)
513 plt.figure(figsize = (10,6))
514 s = sns.scatterplot(x = index, y = resid_std_adj, s = 10, color = 'blue')
plt.axhline(y = -2, color = 'red')
plt.axhline(y = 2, color = 'red')
s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
518 s.set_ylabel('Studentized Residuals', fontdict = {'weight':'bold', 'size':14})
519 plt.show()
520
521 ### Checking for Leverage values:
522
523
      # For ATC:
10 lev = model_AIC.get_influence().hat_matrix_diag
thrs = round(2*(len(best_subset_vars_AIC)+1)/935, 2)
526
plt.figure(figsize = (12,6))
528 plt.stem(lev)
529 plt.axhline(y = thrs, color = 'r')
530 plt.xlabel('Index', fontdict = {'weight':'bold', 'size':14})
ssi plt.ylabel('Leverage values', fontdict = {'weight':'bold', 'size':14})
532 plt.show()
533
     # For Adjusted R^2
10 lev = model_adj.get_influence().hat_matrix_diag
thrs = round(2*(len(best_subset_vars_adj)+1)/935, 2)
```

```
538 plt.figure(figsize = (12,6))
539 plt.stem(lev)
540 plt.axhline(y = thrs, color = 'r')
plt.xlabel('Index', fontdict = {'weight':'bold', 'size':14})
542 plt.ylabel('Leverage values', fontdict = {'weight':'bold', 'size':14})
543 plt.show()
      # For BIC:
545
1546 lev = model_BIC.get_influence().hat_matrix_diag
thrs = round(2*(len(best_subset_vars_BIC)+1)/935, 2)
549 plt.figure(figsize = (12,6))
550 plt.stem(lev)
plt.axhline(y = thrs, color = 'r')
ss2 plt.xlabel('Index', fontdict = {'weight':'bold', 'size':14})
553 plt.ylabel('Leverage values', fontdict = {'weight':'bold', 'size':14})
554 plt.show()
557 ### Cook's Distance
     ## For AIC:
559
560 cooks_AIC = model_AIC.get_influence().cooks_distance[0]
561 p = len(X_train_dumm_AIC.columns)
562
index = np.arange(1, len(resid_AIC)+1)
564
565 plt.figure(figsize = (10,6))
s = sns.scatterplot(x = index, y = cooks_AIC, s = 6, color = 'red')
s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
s.set_ylabel("Cook's Distance", fontdict = {'weight':'bold', 'size':14})
569 plt.show()
570
       # For BIC:
571
cooks_BIC = model_BIC.get_influence().cooks_distance[0]
p = len(X_train_dumm_BIC.columns)
574
index = np.arange(1, len(resid_BIC)+1)
576
plt.figure(figsize = (10,6))
s = sns.scatterplot(x = index, y = cooks_BIC, s = 6, color = 'red')
s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
580 s.set_ylabel("Cook's Distance", fontdict = {'weight':'bold', 'size':14})
581 plt.show()
582
c = pd.DataFrame(cooks_BIC).rename(columns = {0:'cooks'})
584 d = pd.concat([X_train_dumm_BIC.reset_index().drop('index', axis = 1),
585
            pd.DataFrame(Y_train).reset_index().drop('index', axis = 1), c],
586
             axis = 1)
d = d[d['cooks'] < 0.05]
588
589 # After update:
590 X_train_dumm_BIC = d.drop(['Y_new','cooks'], axis = 1)
591 Y_train_BIC = d['Y_new']
593 X = sm.add_constant(X_train_dumm_BIC)
594 Y = Y_train_BIC
section model_BIC_new = sm.OLS(Y,X).fit()
597 cooks_BIC = model_BIC_new.get_influence().cooks_distance[0]
```

```
index = np.arange(1, len(resid_BIC))
599
600 plt.figure(figsize = (10,6))
s = sns.scatterplot(x = index, y = cooks_BIC, s = 6, color = 'red')
s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
s.set_ylabel("Cook's Distance", fontdict = {'weight':'bold', 'size':14})
604 plt.show()
605
      # For Adjusted R^2:
607 cooks_adj = model_adj.get_influence().cooks_distance[0]
608 p = len(X_train_dumm_adj.columns)
index = np.arange(1, len(resid_adj)+1)
612 plt.figure(figsize = (10,6))
613 s = sns.scatterplot(x = index, y = cooks_adj, s = 6, color = 'red')
s.set_xlabel('Index', fontdict = {'weight':'bold', 'size':14})
615 s.set_ylabel("Cook's Distance", fontdict = {'weight':'bold', 'size':14})
616 plt.show()
617 #======
# Variance Inflation Factor (VIF)
620 ## For BIC:
621 vif = pd.DataFrame()
622 vif["Variable"] = X_train_dumm_BIC.columns
623 vif["VIF"] = [variance_inflation_factor(X_train_dumm_BIC.values, i)
               for i in range(X_train_dumm_BIC.shape[1])]
624
625
626 Vif
627
    # For AIC:
628
630 vif = pd.DataFrame()
vif["Variable"] = X_train_dumm_AIC.columns
632 vif["VIF"] = [variance_inflation_factor(X_train_dumm_AIC.values, i)
                for i in range(X_train_dumm_AIC.shape[1])]
634
635 Vif
637
     ## For Adjusted R^2:
638 vif = pd.DataFrame()
639 vif["Variable"] = X_train_dumm_adj.columns
640 vif["VIF"] = [variance_inflation_factor(X_train_dumm_adj.values, i)
                for i in range(X_train_dumm_adj.shape[1])]
641
642
643 Vif
644
645 # Update:
646 X_train_dumm_adj = X_train_dumm_adj.drop('Age', axis = 1)
647
648 vif = pd.DataFrame()
vif["Variable"] = X_train_dumm_adj.columns
650 vif["VIF"] = [variance_inflation_factor(X_train_dumm_adj.values, i)
               for i in range(X_train_dumm_adj.shape[1])]
652
653 Vif
655 X = sm.add_constant(X_train_dumm_adj)
656 Y = Y_train
657 model_adj = sm.OLS(Y,X).fit()
```

```
659 # Model Validation
660 #========
661 X_test_dumm = pd.get_dummies(X_test, drop_first = True)
662 X_test_dumm.head(5)
663
    ## MSPR
664
n = len(Y_test)
667 def mspr_calc(var):
df_new = sm.add_constant(X_test_dumm[M.model.exog_names[1:]])
    yhat = M.predict(df_new)
     return round(np.sum((Y_test - yhat)**2)/n, 4)
print('For AIC: ' + str(mspr_calc('best_subset_vars_AIC')))
674 print('For BIC: ' + str(mspr_calc('best_subset_vars_BIC')))
print('For adj: ' + str(mspr_calc('best_subset_vars_adj')))
      # Printing the model parameters:
677
678 print (model_BIC.params)
679 print('_' *30)
680 print (model_BIC_new.params)
```

# 15 Summary Table for Final Models

# Adjusted $\mathbb{R}^2$

		LS Regress	ion Results			
Dep. Variable:		Y_new	R-squared:		0.759	
Model:		OLS	Adj. R-squar	ed:	0.	.757
Method:	Least	Squares	F-statistic:		36	53.8
Date:	Thu, 02	May 2024	Prob (F-stat	istic):	1.10e-279	
Time:		14:48:37	Log-Likeliho	od:	-942	2.34
No. Observations:		935	AIC:		19	903.
Df Residuals:		926	BIC:		1946.	
Df Model:		8				
Covariance Type:	r	onrobust				
	coef	std err	t	P> t	[0.025	0.975
const	-9.4702	0.675	-14.038	0.000	-10.794	-8.14
Age	15.0557	0.482	31.212	0.000	14.109	16.00
Bmi	-129.5855	21.297	-6.085	0.000	-171.381	-87.79
Children	0.1101	0.018	6.012	0.000	0.074	0.14
Sex_male	-0.1013	0.044	-2.313	0.021	-0.187	-0.01
Smoker_yes	2.3225	0.054	42.910	0.000	2.216	2.42
Region_northwest	-0.0826	0.063	-1.314	0.189	-0.206	0.04
Region_southeast	-0.1855	0.064	-2.918	0.004	-0.310	-0.06
Region_southwest	-0.1642	0.063	-2.593	0.010	-0.289	-0.04
Omnibus: 35		353.982	======================================		1.927	
Prob(Omnibus):		0.000	Jarque-Bera (JB):		1275.302	
Skew:		1.826	Prob(JB):		1.18e-	-277
Kurtosis:	<pre>(urtosis: 7.405 Cond. No. 2.20e+03</pre>				e+03	

Figure 31: For Training Data (Before removing 'Age')

OLS REGRESSION RESULTS						
Dep. Variable:	Y_new	Y_new R-squared:		0.505		
Model: OLS		Adj. R-squar	red:	0.501		
Method:	Least	Squares	F-statistic:	:	1	35.0
Date:	Thu, 02	May 2024	Prob (F-stat	tistic):	9.32e-137	
Time:		16:06:59	Log-Likeliho	ood:	-12	78.4
No. Observations:		935	AIC:		2	573.
Df Residuals:		927	BIC:		2	611.
Df Model:		7				
Covariance Type:	r	nonrobust				
=======================================	=======================================	=======	=========			========
	coef	std err	t	P> t	[0.025	0.975]
					44.460	
const	11.4071	0.126				
Bmi			-7.372			
Children	0.1722					
Sex_male	-0.1100	0.063				
			29.063		2.098	
Region_northwest		0.090	-1.568		-0.318	
Region_southeast		0.091	-3.674	0.000	-0.512	-0.155
Region_southwest	-0.2131	0.091	-2.351	0.019	-0.391	-0.035
Omnibus:	========	.======= 0.873	Dunbin Watse	:======= >n :	======================================	
Prob(Omnibus):		0.646			1.956	
Skew:		0.013	1 /		0.928	
Kurtosis:		2.848	Prob(JB): 0.629 Cond. No. 1.82e+03			
vai (0212.	=======	Z.040	=========		1.02	====

Figure 32: For Training Data (After removing 'Age')

			:ion kesuits			
Dep. Variable:		Y_new	R-squared:		0.525	
Model: OLS		Adj. R-squar	ed:	0.516		
Method:	Least	Squares	F-statistic:		62	2.12
Date:	Thu, 02	May 2024	Prob (F-stat	istic):	8.02e-60	
Time:		16:07:00			-554.67	
No. Observations:		402	AIC:		11	125.
Df Residuals:		394	BIC:		11	L57.
Df Model:		7				
Covariance Type:	n	onrobust				
=======================================	coef	std err	:======= t	P> t	[0.025	0.9751
const	11.2760	0.180	62.552	0.000	10.922	11.630
Bmi	-173.1094	43.607	-3.970	0.000	-258.841	-87.378
Children	0.1651	0.040	4.118	0.000	0.086	0.244
Sex_male	-0.1964	0.099	-1.976	0.049	-0.392	-0.001
Smoker_yes	2.4431	0.123	19.884	0.000	2.202	2.685
Region_northwest	-0.0549	0.141	-0.390	0.697	-0.332	0.222
Region_southeast	-0.2234	0.136	-1.648	0.100	-0.490	0.043
Region_southwest	-0.2191	0.139	-1.577	0.116	-0.492	0.054
Omnibus: 5.259 Durbin-V				:====== on:	1.	==== .738
Prob(Omnibus):		0.072	Jarque-Bera (JB):		3.796	
Skew:		-0.095			0.150	
Kurtosis:		2.564	Cond. No.		1.716	e+03
==========	========	=======	========	=======	=========	====

Figure 33: For Testing Data

# **AIC**

=======================================	========	=======	========	=======	========	:===	
Dep. Variable:		Y_new	R-squared:		0.758		
Model:		OLS	Adj. R-squar	Adj. R-squared:		0.756	
Method:	Least	Squares	F-statistic:	:	41	15.2	
Date:	Thu, 02	May 2024	Prob (F-stat	tistic):	1.23e-280		
Time:		14:44:59	Log-Likeliho	ood:	-943.21		
No. Observations:		935	AIC:		19	902.	
Df Residuals:		927	BIC:		19	941.	
Df Model:		7					
Covariance Type:	n	onrobust					
=======================================	coef	std err	t	P> t	======== [0.025	0.975]	
const	-9.5426	0.673	-14.187	0.000	-10.863	-8.222	
Age	15.0746	0.482	31.252	0.000	14.128	16.021	
Bmi	-128.2363	21.280	-6.026	0.000	-169.999	-86.473	
Children	0.1100	0.018	6.002	0.000	0.074	0.146	
Sex_male	-0.1030	0.044	-2.352	0.019	-0.189	-0.017	
Smoker_yes	2.3253	0.054	42.979	0.000	2.219	2.431	
Region_southeast	-0.1420	0.054	-2.615	0.009	-0.249	-0.035	
Region_southwest	-0.1212	0.054	-2.235	0.026	-0.228	-0.015	
Omnibus:	========	351.154	======= Durbin-Watso	:======: on :	 . 1	928	
Prob(Omnibus):		0.000	Jarque-Bera		1250.		
Skew:		1.815	Prob(JB):	(55).	2.49e-		
Kurtosis:		7.351	Cond. No.		2.186		
===========	========	=======	========	=======	========	:===	

Figure 34: For Training Data

OLS Regression Results							
Dep. Variable: Model: Method: Date: Time: No. Observations: Df Residuals: Df Model:	 Least	TES Regress Y_new OLS Squares May 2024 15:10:04 402 394 7	R-squared: Adj. R-squar F-statistic:	:istic):	0 2:07e -380 7:	==== .800 .796 25.0 -133 0.72 77.4	
Covariance Type:	r 	nonrobust 	.=======		.=======		
	coef	std err	t	P> t	[0.025	0.975]	
const Age Bmi Children Sex_male Smoker_yes Region_southeast Region_southwest	-10.7250 15.9221 -111.3008 0.1426 -0.0902 2.3833 -0.2612 -0.2148	0.950 0.684 28.408 0.026 0.064 0.080 0.078 0.080	-11.290 23.292 -3.918 5.507 -1.405 29.893 -3.349 -2.686	0.000 0.000 0.000 0.161 0.000 0.001 0.008	-12.593 14.578 -167.151 0.092 -0.216 2.227 -0.415 -0.372	-8.857 17.266 -55.450 0.193 0.036 2.540 -0.108 -0.058	
Omnibus: Prob(Omnibus): Skew: Kurtosis:		184.168 0.000 2.032 8.549	Durbin-Watso Jarque-Bera Prob(JB): Cond. No.		1 792 8.64e 2.04e	-173 e+03	

Figure 35: For Testing Data

# **BIC**

OLS Regression Results							
Dep. Variab	======== le:	Y	_==== _new	R-sq	======== uared:		0.755
Model:			OLS	Adj.	R-squared:		0.754
Method:		Least Squ	ares	F-st	atistic:		715.0
Date:	•	Thu, 02 May			(F-statistic	):	6.51e-282
Time:		15:5	0:38	Log-	Likelihood:		-950.04
No. Observa			935	AIC:			1910.
Df Residual	s:		930	BIC:			1934.
Df Model:			4				
Covariance	Type:	nonro	bust				
========	coef	======= std err	=====	===== t	======== P> t	 [0.025	 0.975]
const	-9.8213	0.671	-14	.638	0.000	-11.138	-8.505
Age	15.1569	0.484	31	.324	0.000	14.207	16.107
Bmi	-112.5968	20.699	-5	.440	0.000	-153.220	-71.974
Children	0.1113	0.018	6	.039	0.000	0.075	0.147
Smoker_yes	2.3122	0.054	42	.676	0.000	2.206	2.418
Omnibus:		343	.342	Durb:	======== in-Watson:		1.927
Prob(Omnibu	s):	0	.000	Jarq	ue-Bera (JB):		1213.530
Skew:		1	.772	Prob	(JB):		3.06e-264
Kurtosis:		7	.311	Cond	. No.		2.03e+03
	=======	========	=====	:			========

Figure 36: For Training Data (Before removing Influential point)

========	========			=====	=====	=========		========
Dep. Variab	le:		Υ	_new	R-sq	uared:		0.760
Model:				OLS	Adj.	R-squared:		0.759
Method:		Least	t Squ	iares	F-st	atistic:		736.1
Date:		Thu, 02	May	2024	Prob	(F-statistic	):	3.41e-286
Time:			15:5	1:36	Log-	Likelihood:		-938.42
No. Observa	tions:			934	AIC:			1887.
Df Residual	s:			929	BIC:			1911.
Df Model:				4				
Covariance	Type:	1	nonro	bust				
========	=======		====	=====	=====	=========	=======	========
	coef	std				P> t	-	-
const	-9.8741	0				0.000		
						0.000		
•						0.000		
Children	0.1136	9	.018	6	.231	0.000	0.078	0.149
Smoker_yes	2.317	0	. 054	43	.254	0.000	2.212	2.423
======= Omnibus:			225	===== .806	====	======== in-Watson:	=======	1.914
Prob(Omnibu	(e) ·			.000		ue-Bera (JB):		1157.134
Skew:	13).			742		(JB):		5.39e-252
Kurtosis:				1.194		. No.		2.04e+03
				• 1 2 4		. 140.		2.046+03

Figure 37: For Training Data (After removing Influential point)

After removing the influential point, we get the final model which is provided above for the training dataset. Now, the summary table of the corresponding model for test data is given below:

OLS F	Regress	ion R	Result	S
-------	---------	-------	--------	---

=======================================	==========		=========	========
Dep. Variable:	Y_new	R-squared:		0.791
Model:	_ OLS	Adj. R-square	d:	0.789
Method:	Least Squares	F-statistic:		376.1
Date:	Thu, 02 May 2024	Prob (F-stati	stic):	1.46e-133
Time:	15:11:24	Log-Likelihoo	d:	-389.30
No. Observations:	402	AIC:		788.6
Df Residuals:	397	BIC:		808.6
Df Model:	4			
Covariance Type:	nonrobust			
=======================================				
coe	f std err	t P> t	[0.025	0.9/5]
const -10.964	1 0.962 -	11.398 0.00	0 -12.855	-9.073
Age 15.927			0 14.564	
Bmi -84.168	1 27.975	-3.009 0.00	3 -139.165	-29.171
Children 0.136	8 0.026	5.215 0.00	0.085	0.188
Smoker_yes 2.354	7 0.080	29.266 0.00	0 2.196	2.513
	474 500	=========		4 050
Omnibus:		Durbin-Watson		1.852
Prob(Omnibus):	0.000		lR):	701.659
Skew:	1.937	` '		4.33e-153
Kurtosis:	8.184	Cond. No.		1.91e+03

Figure 38: For Testing Data