

3. Multivariate linear regression is widely used to model the relationships between multiple related responses and a set of predictors. Suppose we have  $N$  observations of  $m$ -dimensional responses  $Y_i = (y_{i,1}, \dots, y_{i,m})^\top$  and  $p$ -dimensional predictors  $X_i = (x_{i,1}, \dots, x_{i,p})^\top$ , for  $i = 1, \dots, N$ . Let  $\mathbf{Y} = (Y_1, \dots, Y_N)^\top$  be the  $N \times m$  response matrix and  $\mathbf{X} = (X_1, \dots, X_N)^\top$  be the  $N \times p$  covariates matrix. The multivariate linear regression model assumes

$$\mathbf{Y} = \mathbf{X}\beta + E,$$

where  $B$  is a  $p \times m$  matrix of unknown regression parameters and  $E = (\epsilon_1, \dots, \epsilon_N)^\top$  is an  $N \times m$  matrix of regression errors, with  $\epsilon_i$ 's independently sampled from an  $m$ -dimensional Gaussian distribution  $N(\mathbf{0}, \Sigma)$ .

- (a) For the multivariate linear regression model, derive the Maximum Likelihood Estimator (MLE) of  $B$  and  $\Sigma$ .
- (b) Similarly to the unidimensional linear regression model, the Ordinary Least Squares (OLS) estimator of  $B$  is defined as the minimizer of the Residual Sum of Squares (RSS) defined as

$$RSS(B) = \sum_{i=1}^N (Y_i^\top - X_i^\top B)(Y_i^\top - X_i^\top B)^\top.$$

Derive the OLS estimator of  $B$ . Is it the same as the MLE?

$$(a) L(B, \Sigma) = \prod_{i=1}^N \frac{1}{(2\pi)^{\frac{m}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(Y_i - B^\top X_i)^\top \Sigma^{-1} (Y_i - B^\top X_i)\right) \quad \text{—— Likelihood function}$$

$$L(B, \Sigma) = \ln(L(B, \Sigma)) = -\frac{N}{2} \ln(2\pi) - \frac{N}{2} \ln|\Sigma| - \frac{1}{2} \sum_{i=1}^N (Y_i - B^\top X_i)^\top \Sigma^{-1} (Y_i - B^\top X_i)$$

$$\text{here } \sum_{i=1}^N (Y_i - B^\top X_i)^\top \Sigma^{-1} (Y_i - B^\top X_i) = \text{tr}(\Sigma^{-1} (Y - XB)^\top (Y - XB)) \Rightarrow L(B, \Sigma) = (-\frac{N}{2} \ln|\Sigma| - \frac{1}{2} \text{tr}(\Sigma^{-1} (Y - XB)^\top (Y - XB)))$$

$$\text{according to } \frac{\partial}{\partial B} \text{tr}(\Sigma^{-1} (Y - XB)^\top (Y - XB)) = -2X^\top (Y - XB) \Sigma^{-1}, \text{ let } -2X^\top (Y - XB) \Sigma^{-1} = 0 \Rightarrow X^\top Y = X^\top XB$$

$$\Rightarrow \hat{B}_{MLE} = (X^\top X)^{-1} X^\top Y$$

$$\text{define } S := (Y - XB)^\top (Y - XB) \Rightarrow L(\Sigma) = -\frac{N}{2} \ln|\Sigma| - \frac{1}{2} \text{tr}(\Sigma^{-1} S)$$

$$\text{Here } \frac{\partial L}{\partial \Sigma^{-1}} = \frac{N}{2} \Sigma - \frac{1}{2} S = 0 \Rightarrow \hat{\Sigma}_{MLE} = \frac{1}{N} S = \frac{1}{N} (Y - X\hat{B}_{MLE})^\top (Y - X\hat{B}_{MLE})$$

$$(b) RSS(B) = \text{tr}((Y - XB)^\top (Y - XB)) = \text{tr}(Y^\top Y - Y^\top XB - B^\top X^\top Y + B^\top X^\top XB)$$

$$\frac{\partial RSS}{\partial B} = \frac{\partial}{\partial B} \text{tr}(-2Y^\top XB + B^\top X^\top XB) = -2X^\top Y + 2X^\top XB, \text{ let } \frac{\partial RSS}{\partial B} = 0 \Rightarrow X^\top XB = X^\top Y \Rightarrow \hat{B}_{OLS} = (X^\top X)^{-1} X^\top Y$$

$$\hat{B}_{OLS} = \hat{B}_{MLE}$$