COMP 680 Statistics for Computing and Data Science Week 2: Probability and Random Variable II

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Outline

- Expectations, Variance and Covariance
- Multivariate Distributions
- Probability Inequalities
- Transformation of Random Variables

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Expectation (mean)

- The expectation of a random variable X, denoted $\mathbb{E}[X]$
 - $\mathbb{E}[X] = \sum_{i} x_i \cdot \mathbb{P}(X = x_i)$ if X is discrete
 - $\mathbb{E}[X] = \int_{X \in \mathbb{R}} x \cdot f_X(x) dx$ if X is continuous

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 - $\mathbb{E}[X] = \int_{X \in \mathbb{R}} x \cdot f_X(x) dx$ if X is continuous
- For any function $g: \mathbb{R} \to \mathbb{R}$, the expectation of g(X) is:
 - $\mathbb{E}[g(X)] = \sum_{i} g(x_i) \cdot \mathbb{P}(X = x_i)$ if X is discrete
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- Examples:

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Properties of Expectation

• Linearity: $\mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]$ for constants a and b and random variables X and Y.

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- Independence: If two random variables X and Y are independent, then $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$.
 - can you show this?

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Properties of Expectation

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- Independence: If two random variables X and Y are independent, then $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$.
 - can you show this?
- Intuition: the "typical" value of the random variable.

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Variance and Standard Deviation

• The variance of a random variable *X* is defined to be the mean square deviation from the mean:

$$Var(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

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• The standard deviation σ_X of X is:

$$\sigma_X = \sqrt{\mathsf{Var}(X)}$$

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• The standard deviation σ_X of X is:

$$\sigma_X = \sqrt{\mathsf{Var}(X)}$$

• $Var(aX + b) = a^2Var(X)$ for any constants a and b

Covariance

• The covariance of two random variables X and Y is defined to be:

$$\mathsf{Cov}(X,Y) = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

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Covariance

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• One can show: Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)

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Covariance

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- One can show: Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)
- If X and Y are independent:
 - $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$
 - Cov(X, Y) = 0
 - Var(X + Y) = Var(X) + Var(Y)

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Correlation Coefficient

Pearson's correlation:

$$\rho_{X,Y} = \frac{\mathsf{Cov}(X,Y)}{\sigma_X \sigma_Y}$$

• $-1 \le \rho_{X,Y} \le 1$: a standardized measure

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Correlation Coefficient

Pearson's correlation:

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- $-1 \le \rho_{X,Y} \le 1$: a standardized measure
- Describe the linear association between X and Y:
 - $\rho_{X,Y} = 1$: Y = aX + b where $a = \sigma_Y / \sigma_X$
 - $\rho_{X,Y} = -1$: Y = aX + b where $a = -\sigma_Y/\sigma_X$
 - $\rho_{X,Y} = 0$: no linear association or "uncorrelated"

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Outline

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- Multivariate Distributions
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Multinomial Distributions

• Multinomial distributions with parameters n and $\sum_{i=1}^k p_i = 1$

$$\mathbb{P}(X_1 = x_1, \dots X_k = x_k) = \binom{n}{x_1 \dots x_k} p_1^{x_1} \dots p_k^{x_k} = \frac{n!}{x_1! \dots x_k!} p_1^{x_1} \dots p_k^{x_k}$$

where x_i 's are integers such that $\sum_{i=1}^k x_i = n$

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Multinomial Distributions

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where x_i 's are integers such that $\sum_{i=1}^k x_i = n$

• Binomial distribution is multinomial with k=2

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Gaussian Distributions

ullet Gaussian distribution with parameters μ and Σ

$$f_{oldsymbol{X}}(oldsymbol{x}) = rac{1}{(2\pi)^{k/2} |oldsymbol{\Sigma}|^{1/2}} \exp\left[-rac{1}{2}(oldsymbol{x}-oldsymbol{\mu})^T oldsymbol{\Sigma}^{-1}(oldsymbol{x}-oldsymbol{\mu})
ight]$$

- mean vector μ : a vector of length k
- covariance matrix Σ : a $k \times k$ symmetric and positive definite matrix

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Bivariate Normal Distribution

Example of 2-dimensional Gaussian $N_2(\mu, \Sigma)$

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- mean vector $\boldsymbol{\mu} = (\mu_1, \mu_2)$
- covariance matrix $\Sigma = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}$

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Bivariate Normal Distribution

Example of 2-dimensional Gaussian $N_2(\mu, \Sigma)$

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- mean vector $\boldsymbol{\mu} = (\mu_1, \mu_2)$
- covariance matrix $\Sigma = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}$
- ullet 2-d random vector $oldsymbol{X} = (X_1, X_2) \sim \mathcal{N}_2(oldsymbol{\mu}, oldsymbol{\Sigma})$
 - $\mathbb{E}[X_1] = \mu_1$, $\mathbb{E}[X_2] = \mu_2$
 - $Var[X_1] = \sigma_1^2$, $Var[X_2] = \sigma_2^2$
 - $Cov[X_1, X_2] = \rho \sigma_1 \sigma_2$

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Properties of Gaussian Distribution

- Let X be a multivariate Gaussian random vector, then the following also follow Gaussian distribution:
 - The marginal distribution of all subsets of the components
 - The conditional distribution of the components
 - Linear combination of the components

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Properties of Gaussian Distribution

- Let X be a multivariate Gaussian random vector, then the following also follow Gaussian distribution:
 - The marginal distribution of all subsets of the components
 - The conditional distribution of the components
 - Linear combination of the components
- Example $\boldsymbol{X} = (X_1, X_2) \sim N_2(\boldsymbol{\mu}, \boldsymbol{\Sigma})$:
 - ullet $\Sigma^{-1/2}(oldsymbol{X}-\mu)\sim \mathcal{N}_2(oldsymbol{0},oldsymbol{I})$
 - $X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$
 - $X_1|X_2 = a \sim N(\mu_1 + \frac{\sigma_1}{\sigma_2}\rho(a \mu_1), (1 \rho^2)\sigma_1^2)$
 - For a constant vector a, $a^T \mathbf{X} \sim N(a^T \mu, a^T \Sigma a)$

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Markov's Inequality

• Let X be a non-negative random variable and suppose that $\mathbb{E}[X]$ exists. For any t > 0,

$$\mathbb{P}(X \ge t) \le \frac{\mathbb{E}[X]}{t}$$

Proof:

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Chebyshev's Inequality

• Let X be a random variable with $\mathbb{E}[X] = \mu$ and $\mathsf{Var}(X) = \sigma^2$, then

$$\mathbb{P}(|X - \mu| \ge t) \le \frac{\sigma^2}{t^2}$$

in particular taking $t = k\sigma$, then

$$\mathbb{P}(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

Proof:

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Jensen's Inequality

• If g is a convex function then $\mathbb{E}[g(X)] \ge g(\mathbb{E}[X])$; and if g is a concave function then $\mathbb{E}[g(X)] \le g(\mathbb{E}[X])$

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Jensen's Inequality

- If g is a convex function then $\mathbb{E}[g(X)] \geq g(\mathbb{E}[X])$; and if g is a concave function then $\mathbb{E}[g(X)] \leq g(\mathbb{E}[X])$
- If a function g is a convex then it lies above any line that touches g at any point; g is concave if -g is convex.

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Jensen's Inequality

- If g is a convex function then $\mathbb{E}[g(X)] \geq g(\mathbb{E}[X])$; and if g is a concave function then $\mathbb{E}[g(X)] \leq g(\mathbb{E}[X])$
- If a function g is a convex then it lies above any line that touches g at any point; g is concave if -g is convex.
- Examples:
 - $g(x) = x^2$ is convex, $\mathbb{E}[X^2] \ge \mathbb{E}[X]^2$
 - g(x) = 1/x is convex for x > 0, $\mathbb{E}[1/X] \ge 1/\mathbb{E}[X]$
 - $g(x) = \log x$ is concave, $\mathbb{E}[\log X] \leq \log \mathbb{E}[X]$

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 You flip a fair coin 100 time, give an upper bound of probability of getting 70 heads or more.

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- You flip a fair coin 100 time, give an upper bound of probability of getting 70 heads or more.
- How accurate is this estimate?
 - $X \sim \text{Binomial}(n = 100, p = 0.5)$
 - estimate $\mathbb{P}(X \geq 70)$
 - apply Markov's Inequality:

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 You would like to know the percentage p in general population that support certain legislation. You start a poll online to randomly survey 100 people for a Yes/No question and you estimate the proportion of Yes in the survey response.

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- You would like to know the percentage p in general population that support certain legislation. You start a poll online to randomly survey 100 people for a Yes/No question and you estimate the proportion of Yes in the survey response.
- How accurate is this estimate?
 - $X_1, X_2, \cdots X_n \stackrel{\text{i.i.d.}}{\sim} \text{Bernoulli}(p)$
 - estimate p by $\bar{X}_n = \sum_{i=1}^n X_i/n$
 - apply Chebyshev's Inequality:

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Functions of Random Variables

- Let X be a random variable and Y = g(X) be another random variable defined by a function $g : \mathbb{R} \to \mathbb{R}$.
- The cdf of Y is given by:

$$F_Y(y) = \mathbb{P}(Y \le y) = \mathbb{P}(g(X) \le y) = \int_{\{x \mid g(x) \le y\}} f_X(x) dx$$

If g is a strictly increasing function, then

$$F_Y(y) = \int_{-\infty}^{g^{-1}(y)} f_X(x) dx = F_X(g^{-1}(y))$$
$$f_Y(y) = f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

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General Steps of Transformation

- Let $\mathbf{X} = (X_1, X_2, \dots X_n)$ be a random vector and $Y = g(X_1, X_2, \dots X_n)$ be another random variable defined by a function $g : \mathbb{R}^n \to \mathbb{R}$.
- To find the pdf of Y :

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General Steps of Transformation

- Let $\boldsymbol{X} = (X_1, X_2, \dots X_n)$ be a random vector and $Y = g(X_1, X_2, \dots X_n)$ be another random variable defined by a function $g : \mathbb{R}^n \to \mathbb{R}$.
- To find the pdf of Y :
 - For each value of y, find the set

$$A_{v} = \{x_{1}, x_{2}, \cdots x_{n} : g(x_{1}, x_{2}, \cdots x_{n}) \leq y\}$$

Find the cdf

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(g(x_1, x_2, \dots x_n) \leq y) = \int_{A_y} f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}$$

• The pdf is $f_Y(y) = F'_Y(y)$

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• Let $X_1, X_2, \cdots X_n \overset{\text{i.i.d.}}{\sim}$ Uniform[0, 1], find the probability density function for the following random variables:

•
$$U = e^{X_1}$$

•
$$X^{(n)} = \max(X_1, X_2, \cdots X_n)$$

•
$$Y = X_1 + X_2$$

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