

Unit-5  
Test of Significance (Small Samples)

Test of Significance for Small Samples:

Definition: If the Sample size less than 30, it is called a Small Sample.

The following are some important tests for  
small samples.

- (i) Student's 't' Test
- (ii) F - Test
- (iii)  $\chi^2$  - Test.

Test of Hypothesis for Single Mean ( $\mu$ ):

$$\text{Test Statistic } t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}} \text{ where}$$

if 's' is not given directly

$\bar{x}$  = Sample mean

$\mu$  = population mean

s = Sample S.D

n = Sample Size

$$t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n-1}}} \text{ if 's' is given directly.}$$

Here Degrees of freedom =  $(n-1)$

Note:-

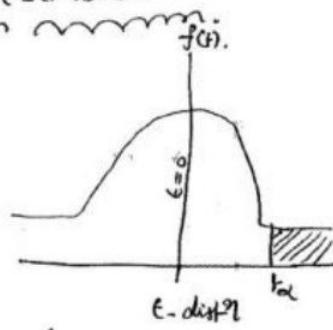
① For one tailed Test  $|t_{tab}| = \text{See } t\text{-Table } (n-1) \text{ at } \alpha$

② For two tailed Test  $|t_{tab}| = \text{See } t\text{-Table } (n-1) \text{ at } \frac{\alpha}{2}$

Confidence interval for  $\mu$ :

Confidence interval for  $\mu$  is  $(\bar{x} - t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}}, \bar{x} + t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}})$

Properties of  $t$ -distribution: (3) Characteristics of  $t$ -distribution



1. The shape of  $t$ -distribution is bell-shaped, which is similar to that of a normal distribution and is symmetrical about the mean.
2. The  $t$ -distribution curve is also asymptotic to  $t$ -axis. i.e. The two tails of the curve on both sides of  $t=0$  extends to infinity.
3. It is symmetrical about the line  $t=0$ .
4. The form of probability Curve varies with degrees of freedom.
5. It is unimodal with Mean = Median = Mode.
6. The shape of  $t$ -distribution depends on Sample size  $n$ .

Applications of  $t$ -distribution (3) Uses of  $t$ -distribution

- (1) To test the significance of the sample mean, when population Variance is not given
- (2) To test the significance of the mean of the sample  
i.e. To test if the sample mean differs significantly from the population mean
- (3) To test the significance of the difference between two sample means.
- (4) To test the significance of an observed sample correlation Co-efficient and sample regression Co-efficient.

(2)

- ① A Random Sample of Six steel beams has mean Compressive strength of 58,392 P.S.I (pounds per square inch) with standard deviation of 648 P.S.I. Use this information and the level of significance  $\alpha = 0.05$  to test whether the true average Compressive strength of steel from which this sample came is 58,000 P.S.I.  
Assume normality.

Sol Given Sample size  $n = 6$  (Small Sample), Degrees of freedom  $= n-1 = 5$   
Sample mean  $\bar{x} = 58,392$   
sample standard deviation  $s = 648$ .  
population mean  $\mu = 5800$

- (i) Null Hypothesis  $H_0 : \mu = 58000$
- (ii) Alternative Hypothesis  $H_1 : \mu \neq 58000$  <two-tailed Test>
- (iii) Level of Significance:  $\alpha = 0.05$
- (iv) Test statistic:  $t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n-1}}} \quad \langle s.d. \text{ of sample given directly} \rangle$   
 $= \frac{58,392 - 58000}{\frac{648}{\sqrt{6-1}}} = \frac{392 \times \sqrt{6}}{648} = 1.353$   
 $\therefore |t_{\text{cal}}| = 1.353$

(v) Conclusion:  $|t_{\text{tab}}| = 2.571$ .

$\begin{cases} \text{see t-table } (n-1) \text{ at } \frac{\alpha}{2}. \\ \text{see t-Table } 5 \text{ at } \frac{0.05}{2} = 0.025 \end{cases}$

 $\therefore |t_{\text{cal}}| < |t_{\text{tab}}|$

$\therefore H_0$  is accepted

Hence the average Compressive strength of steel beam is 58000 P.S.I /

- ② A Sample of 26 bulbs gives a mean life of 990 hours S.D of 20 hours. The manufacturer claims that the mean life of bulbs is 1000 hours. Is the sample not up to the standard.

Sol Given  $n = 26$  (Small sample),  $\bar{x} = 990$ ,  $s = 20$ ,  $\mu = 1000$   
 $d.o.f = 26-1 = 25$ .

(i)  $H_0 : \mu = 1000$ , (ii)  $H_1 : \mu < 1000$  (iii)  $\alpha = 5\% = 0.05$  (iv)  $t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n-1}}} = -2.5$

$\therefore |t_{\text{cal}}| = 2.5$

(v) Conclusion:  $|t_{\text{tab}}| = 1.708$   $\begin{cases} \text{See t-Table } (n-1) \text{ at } \alpha. \\ \text{See t-Table } 25 \text{ at } 0.05 \end{cases}$   $\begin{cases} \text{d.o.f Sample} \\ \text{given directly} \end{cases}$

$\therefore H_0$  is rejected (i.e. sample is not up to standard).

- (3) A new process of producing synthetic diamonds can be operated at a profitable level only if the average weight of the diamonds is greater than 0.5 carat. To test the profitability of the process, 6 diamonds are produced with weights 0.45, 0.60, 0.52, 0.49, 0.58 and 0.54 carat respectively. Do the 6 measurements present sufficient evidence to indicate that the average weight of the diamonds produced by the process is in excess of 0.5 carat.

*< Dec. 2014 >*

Sol

Given  $n = 6$  (small sample),  $\mu = 0.5$

$$\bar{x} = \frac{\sum x_i}{n} = \frac{1}{6} (0.45 + 0.60 + 0.52 + 0.49 + 0.58 + 0.54) = 0.53.$$

$$\begin{aligned}\text{Sample Variance } s^2 &= \frac{\sum (x_i - \bar{x})^2}{(n-1)} \\ &= \left[ (0.45 - 0.53)^2 + (0.60 - 0.53)^2 + (0.52 - 0.53)^2 + \right. \\ &\quad \left. (0.49 - 0.53)^2 + (0.58 - 0.53)^2 + (0.54 - 0.53)^2 \right] / 5\end{aligned}$$

$$s^2 = 0.0156/5$$

$$s = \sqrt{\frac{0.0156}{5}}$$

$$s = 0.0558, \text{ Degrees of freedom } = n-1 = 6-1 = 5$$

(i) Null Hypothesis  $H_0: \mu = 0.5$  (ii) Alternative Hypothesis  $H_1: \mu > 0.5$  (one-tailed)

(iii) Level of Significance:  $\alpha = 5\% = 0.05$

(iv) Test Statistic:  $t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}}$  *( $s$  of sample not given directly)*

$$t_{\text{cal}} = \frac{0.53 - 0.5}{\frac{0.0558}{\sqrt{6}}} = 1.3155$$

$$\therefore |t_{\text{cal}}| = 1.3155.$$

(v) Conclusion:  $|t_{\text{Tabl}}| = 2.015$

$$\therefore |t_{\text{cal}}| < |t_{\text{Tabl}}|$$

$\therefore H_0$  is accepted.

*See t-Table  $(n-1)$  at  $\alpha$*   
*See t-Table 5 at 0.05*

$\therefore$  The mean of the population of the diamonds produced by the new process is  $\mu = 0.5$  //

(3)

- ④ Labels on one-gallon cans of paint indicate the drying time and the area that can be covered in one coat.

A manufacturer of paints claims that the brand of paint they manufacture will cover 420 sq.ft per gallon. To test this,

a random sample of 10 one-gallon cans of paints was tested.

The actual areas painted in Sq.ft are 362, 356, 413, 422, 372, 416, 376, 434, 388 and 421. Do the data present sufficient evidence to indicate that the average differs from 420 sq.ft?

Sol

Given  $\mu = 420$ ,  $n = 10$  < small sample >

Mean and S.D are not given, first we have to find them.

$$\bar{x} = \frac{\sum x_i}{n} = \frac{1}{10} (362 + 356 + 413 + 422 + 372 + 416 + 376 + 434 + 388 + 421)$$

$$\boxed{\bar{x} = 396}$$

$$s^2 = \frac{\sum (x_i - \bar{x})^2}{n-1} = \frac{1}{9} \left[ (362 - 396)^2 + (356 - 396)^2 + (413 - 396)^2 + (422 - 396)^2 + (372 - 396)^2 + (416 - 396)^2 + (376 - 396)^2 + (434 - 396)^2 + (388 - 396)^2 + (421 - 396)^2 \right]$$

$$s^2 = \frac{7230}{9} = 803.3333$$

$$s = \sqrt{803.3333} = 28.343$$

$$\text{Degrees of freedom} = n-1 = 10-1 = 9.$$

(i) Null Hypothesis  $H_0: \mu = 420$  (ii) Alternative Hypothesis  $H_1: \mu \neq 420$

(iii) Level of Significance:  $\alpha = 5\% = 0.05$  (iv) Test Statistic:

$$t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}} \quad <\because \text{S.D of Sample not given directly}>$$

$$= \frac{396 - 420}{\frac{28.3333}{\sqrt{10}}} = -2.678$$

$$\therefore |t_{-call}| = 2.678$$

(v) Conclusion:  $|t_{-tab}| = 2.262$

$\leftarrow$  See t-Table 9 at  $\frac{0.05}{2} = 0.025$

$$\therefore |t_{-call}| > |t_{-tab}|$$

$\therefore H_0$  is rejected.

⑤ Find 95% Confidence limits for the mean of a normally distributed population from which the following sample was taken 15, 17, 10, 18, 16, 9, 7, 11, 13, 14.

Sol Given  $n = 10$ ,  $\bar{x} = \frac{1}{10}(15+17+10+18+16+9+7+11+13+14) = 13$ .

$$s^2 = \frac{\sum (x_i - \bar{x})^2}{(n-1)} = \frac{40}{3}$$

$$s = \sqrt{\frac{40}{3}}, \text{ degrees of freedom} = n-1 = 10-1 = 9$$

Given Confidence limit = 95%.

i.e. Level of Significance  $\alpha = 5\% = 0.05$

$$t_{\frac{\alpha}{2}} = 2.262$$

$$\begin{cases} \text{See t-Table } (n-1) \text{ at } \frac{\alpha}{2} \\ \text{See t-Table } 9 \text{ at } \frac{0.05}{2} = 0.025 \end{cases}$$

$\therefore$  95% Confidence limits for  $H$  is  $(\bar{x} - t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}}, \bar{x} + t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}})$

$$(13 - 2.262 \frac{\sqrt{40}}{\sqrt{3}}, 13 + 2.262 \frac{\sqrt{40}}{\sqrt{3}})$$

$$(13 - 2.6, 13 + 2.6)$$

$$(10.4, 15.6) //$$

⑥ A sample of size 10 and standard deviation 0.03 is taken from a population. Find Maximum error with 99% Confidence. (May 2015)

Sol Given  $n = 10$ ,  $s = 0.03$ . Confidence limit = 99%.  $\therefore \alpha = 1\% = 0.01$

degrees of freedom =  $10-1 = 9$ .

$$t_{\frac{\alpha}{2}} = 3.25$$

$$\begin{cases} \text{See t-Table } (n-1) \text{ at } \frac{\alpha}{2} \\ \text{See t-Table } 9 \text{ at } \frac{0.01}{2} = 0.005 \end{cases}$$

$$\therefore \text{Maximum Error } E = t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}} = \frac{3.25 (0.03)}{\sqrt{9}} = 0.0325 //$$

⑦ A Random Sample of size 16 values from a normal population showed a mean of 53 and a sum of squares of deviations from the mean equals to 150. Can this sample be regarded as taken from the population having 56 as mean? Obtain 95% confidence limits of the mean of the population.

Sol Given  $n = 16$ ,  $\bar{x} = 53$ ,  $\mu = 56$  and  $\sum (x_i - \bar{x})^2 = 150$ .  $\therefore s^2 = \frac{\sum (x_i - \bar{x})^2}{n-1} = 10$   
 $d.o.f = 16-1 = 15$

$$(i) H_0: \mu = 56 \quad (ii) \mu \neq 56 \quad (iii) \alpha = 5\% = 0.05 \quad (iv) t = \frac{\bar{x} - \mu}{s/\sqrt{n}} = \frac{53 - 56}{\sqrt{10}} = -3.79 \quad |z_{tab}| = 2.131$$

$$(v) |t| > |z_{tab}| \text{ and 95% confidence limits } (51.31, 54.68)$$

(4)

Testing of Two means:

$$\text{Test Statistic } t = \frac{\bar{x} - \bar{y}}{\sqrt{s^2(\frac{1}{n_1} + \frac{1}{n_2})}}$$

$$\text{where } \bar{x} = \frac{\sum x_i}{n_1}, \quad \bar{y} = \frac{\sum y_i}{n_2}$$

$$s^2 = \frac{n_1 s_1^2 + n_2 s_2^2}{n_1 + n_2 - 2} \quad (8)$$

$$\text{Degrees of freedom} = n_1 + n_2 - 2$$

$$s^2 = \frac{1}{n_1 + n_2 - 2} \left[ \sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2 \right]$$

Note: Confidence limits for the difference of two population means are  $(\bar{x} - \bar{y}) - t_{\alpha/2} \cdot \sqrt{s^2(\frac{1}{n_1} + \frac{1}{n_2})}, (\bar{x} - \bar{y}) + t_{\alpha/2} \cdot \sqrt{s^2(\frac{1}{n_1} + \frac{1}{n_2})}$

- Q. ① The IQs of 16 students from one area of a city showed a mean of 107 with standard deviation of 10, while the IQs of 14 students from another area of city showed a mean of 112 with standard deviation of 8. Is there a significant difference between the IQs of the two groups at a 0.05 level of significance.

Sol Given  $n_1 = 16, \bar{x} = 107, s_1 = 10$  and  $n_2 = 14, \bar{y} = 112, s_2 = 8$  (June 2015)

- (i) Null Hypothesis:  $H_0: \mu_1 = \mu_2$  (ii) Alternative Hypothesis  $H_1: \mu_1 \neq \mu_2$   
 (iii) Level of Significance:  $\alpha = 0.05$

(iv) Test Statistic:  $t = \frac{\bar{x} - \bar{y}}{\sqrt{s^2(\frac{1}{n_1} + \frac{1}{n_2})}}$  where  $s^2 = \frac{n_1 s_1^2 + n_2 s_2^2}{n_1 + n_2 - 2}$

$$= \frac{16 \times 100 + 14 \times 16}{16 + 14 - 2}$$

$$t_{\text{cal}} = \frac{107 - 112}{\sqrt{89.14 \left( \frac{1}{16} + \frac{1}{14} \right)}} = -1.447$$

$$s^2 = 89.14$$

$$s = 9.44$$

$$\therefore |t_{\text{cal}}| = 1.447$$

$$(v) \text{ Conclusion: } |t_{\text{cal}}| = 2.048$$

See t-Table  $(n_1+n_2-2)$  at  $\frac{\alpha}{2}$   
 See t-Table 28 at  $\frac{0.05}{2} = 0.025$

$$\therefore |t_{\text{cal}}| < |t_{\text{tab}}|$$

$H_0$  is accepted.

Hence the difference between the IQs of two groups is not significant.

- ② In a manufacturing Company, a one-month training programme is conducted to each employee. A new training programme has been developed. To test the new programme, 2 groups of 9 employees each were assigned a job, the first group trained under the old system and second trained under the new programme. Their performance timings in minutes are recorded as follows.

Old programme :	31	38	36	30	42	34	32	33	39
New programme :	34	33	26	36	31	28	27	29	35

Do the job execution times present sufficient evidence to indicate that the mean time is less for the new programme?

$$\text{Sol} \quad \text{Given } n_1 = 9, \bar{x} = \frac{\sum x_i}{n} = \frac{31+38+36+30+42+34+32+33+39}{9}$$

$$\therefore \boxed{\bar{x} = 35}$$

$$n_2 = 9, \bar{y} = \frac{\sum y_i}{n} = \frac{34+33+26+36+31+28+27+29+35}{9}$$

$$\therefore \boxed{\bar{y} = 31}$$

$$\sum (x_i - \bar{x})^2 = (-4)^2 + 3^2 + 1^2 + (-5)^2 + 7^2 + (-1)^2 + (-3)^2 + (-2)^2 + 5^2 = 130$$

$$\sum (y_i - \bar{y})^2 = 3^2 + 2^2 + (-5)^2 + 5^2 + 0^2 + (-3)^2 + (-4)^2 + (-2)^2 + 4^2 = 108$$

$$\text{now } s^2 = \frac{1}{n_1+n_2-2} \left[ \sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2 \right]$$

$$= \frac{1}{16} [130 + 108] = \frac{238}{16} = 14.875$$

$$\boxed{s^2 = 14.875}$$

(5)

- (i) Null Hypothesis  $H_0: \mu_1 = \mu_2$   
(ii) Alternative Hypothesis  $H_1: \mu_1 > \mu_2$   $\leftarrow$  one tailed Test  
(iii) Level of Significance:  $\alpha = 0.05$ .  
(iv) Test Statistic:  $t = \frac{\bar{x} - \bar{y}}{\sqrt{s^2(\frac{1}{n_1} + \frac{1}{n_2})}} = \frac{35 - 31}{\sqrt{14.875(\frac{1}{9} + \frac{1}{9})}}$   
 $= \frac{4 \times 3}{\sqrt{14.875 \times 2}} = 2.2$   
 $\therefore |t_{cal}| = 2.2$

- (v) Conclusion:  $|t_{tab}| = 1.746$  See t-Table  $n_1+n_2-2$  at  $\alpha$   
 $\therefore |t_{cal}| > |t_{tab}|$  See t-Table 16 at 0.05  
 $\therefore H_0$  is rejected.

- Q. Two independent samples of 8 and 7 items respectively had the following values.

Sample I	11	11	13	11	15	9	12	14
Sample II	9	11	10	13	9	8	10	-

Is the difference between the means of samples significant? (Dec 2017)

Sol Given  $n_1 = 8, n_2 = 7$  and  $\bar{x} = \frac{1}{8}(11+11+13+11+15+9+12+14) = 12$   
 $\bar{y} = \frac{1}{7}(9+11+10+13+9+8+10) = 10$   
 $\sum (x_i - \bar{x})^2 = (-1)^2 + (-1)^2 + 1^2 + (-1)^2 + 3^2 + (-3)^2 + 0^2 + 2^2 = 26$   
 $\sum (y_i - \bar{y})^2 = (-1)^2 + 1^2 + 0^2 + 3^2 + (-1)^2 + (-2)^2 + 0^2 = 16$   
 $s^2 = \frac{1}{n_1+n_2-2} [\sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2] = \frac{1}{13}(42) = 3.23$ .

- (i) Null Hypothesis  $H_0: \mu_1 = \mu_2$  (ii) Alternative Hypothesis  $H_1: \mu_1 \neq \mu_2$   
(iii) L.O.S:  $\alpha = 0.05$  (iv) Test statistic:  $t = \frac{\bar{x} - \bar{y}}{\sqrt{s^2(\frac{1}{n_1} + \frac{1}{n_2})}} = 2.15$   
 $\therefore |t_{cal}| = 2.15$

- (v) Conclusion:  $|t_{tab}| = 2.16$  See t-Table 13 at  $\frac{0.05}{2} = 0.025$ .  
 $\therefore |t_{cal}| < |t_{tab}| \therefore H_0$  is accepted.

### Paired Sample t-Test:

mm mm mm

Paired observations arise in many practical situations where each homogeneous experimental unit receives both population conditions. As result, each experimental unit has a pair of observations, one for each population.

For instance, to test the effectiveness of "drug" some 11 persons blood pressure is measured "before" and "after" the intake of certain drug. Here the individual person is the experimental unit and two populations are blood pressure "before" and "after" the drug is given.

$$\text{Test Statistic } t = \frac{\bar{d} - \mu}{\frac{s}{\sqrt{n}}}, \text{ Here } \mu = 0$$

$$t_{\text{cal}} = \frac{\bar{d}}{\frac{s}{\sqrt{n}}}$$

$$\text{where } \bar{d} = \frac{1}{n} \sum d_i \text{ and } s^2 = \frac{1}{n-1} \sum (d_i - \bar{d})^2$$

$$s^2 = \frac{1}{n-1} \left[ \sum d^2 - n(\bar{d})^2 \right] \quad \textcircled{a}$$

$$\text{Degrees of freedom} = (n-1).$$

Note: ① By using paired sample t-Test, we test two means.

② confidence interval for difference of two populations

$$\text{means } \left( \bar{d} - t_{\alpha/2} \cdot \frac{s}{\sqrt{n}}, \bar{d} + t_{\alpha/2} \cdot \frac{s}{\sqrt{n}} \right)$$

① The following are the average weekly losses of work-hours due to accidents in 10 industrial units before and after introduction of certain safety measures.

use 0.05 LOS to test whether the safety measures are effective.

Before	43	66	109	35	56	85	36	38	49	42
After	36	50	96	37	53	78	28	48	41	42

Sol Given  $n = 10$

differences  $d_i = (43-36), (66-50), (109-96), (35-37), (56-53)$   
 $(85-78), (36-28), (38-48), (49-41), (42-42)$

$$d_i = 7, 16, 13, -2, 3, 7, 8, -10, 8, 0$$

$$\bar{d} = \frac{\sum d_i}{n} = \frac{1}{10} (7+16+13-2+3+7+8-10+8+0)$$

$$\therefore \bar{d} = 5.$$

$$\text{now } s^2 = \frac{1}{(n-1)} \sum (d_i - \bar{d})^2$$

$$= \frac{1}{9} [7^2 + 16^2 + 13^2 + (-2)^2 + 3^2 + 7^2 + 8^2 + (-10)^2 + 8^2 + (-5)^2]$$

$$s^2 = 57.1111$$

$$s = 7.5571$$

$$\text{Degrees of freedom} = n-1 = 10-1 = 9.$$

(i) Null Hypothesis  $H_0: \mu_1 = \mu_2$

(ii) Alternative Hypothesis  $H_1: \mu_1 > \mu_2$   $\langle \text{one-tailed} \rangle$

(iii) Level of Significance:  $\alpha = 0.05$

(iv) Test statistic:  $t = \frac{\bar{d} - \mu}{\frac{s}{\sqrt{n}}} \quad \langle \mu = 0 \rangle$

$$= \frac{\bar{d}}{\frac{s}{\sqrt{n}}} = \frac{5}{\frac{7.5571}{\sqrt{10}}} = 2.0922$$

$$\therefore |t_{\text{cal}}| = 2.0922$$

(v) Conclusion:  $|t_{\text{cal}}| = 1.833$

See t-Table  $n-1$  at  $\alpha$   
 See t-Table 9 at 0.05.  
 $|t_{\text{cal}}| > |t_{\text{Tabl}}$   
 $\therefore H_0$  is rejected.

Hence safety measures are effective.

- K ② Find 90% Confidence interval for the mean improvement in lost worker-hours for above problem.

Sol Given  $n = 10$ , From ① problem  $\bar{d} = 5$ ,  $s = 7.56$

Given Confidence limit = 90%.

level of significance = 10% = 0.1

$$\therefore t_{\frac{\alpha}{2}} = 1.833$$

See t-Table  $(n-1)$  at  $\frac{\alpha}{2}$   
 See t-Table 9 at  $\frac{0.1}{2} = 0.05$

90% Confidence interval is

$$\left( \bar{d} - t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}}, \bar{d} + t_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}} \right)$$

$$\left( 5 - 1.833 \cdot \frac{7.56}{\sqrt{10}}, 5 + 1.833 \cdot \frac{7.56}{\sqrt{10}} \right)$$

$$(0.62, 9.38)$$

- (3) Memory capacity of 10 students were tested before and after training. State whether the training was effective or not from the following scores.

Before Training	12	14	11	8	7	10	3	0	5	6
After Training	15	16	10	7	5	12	10	2	3	8

Sol Same as ①, Given  $n = 10$ ,  $\bar{d} = 1.2$ ,  $s = 2.78$

(i)  $H_0: \mu_1 = \mu_2$  (ii)  $H_1: \mu_1 < \mu_2$  (iii) LOS:  $\alpha = 0.05$  (iv)  $t = \frac{\bar{d}}{s/\sqrt{n}} = -1.365$

$$\therefore |t_{\text{cal}}| = 1.365$$

(v) Conclusion:  $|t_{\text{Tabl}}| = 1.833 \therefore |t_{\text{cal}}| < |t_{\text{Tabl}}$

$\therefore H_0$  is accepted.

### F-Test

This is very useful in testing the equality of population means by comparing sample variances.

To test whether there is any significant difference between two estimates of population variance (8) To test if the two samples have come from same population, we use F-Test.

In this case, we set up Null Hypothesis  $H_0: \sigma_1^2 = \sigma_2^2$  (i.e. populations variances are same).

The Test statistic is  $F = \frac{s_1^2}{s_2^2}$  where  $s_1^2 = \frac{\sum (x_i - \bar{x})^2}{n_1 - 1}$

$$\text{and } s_2^2 = \frac{\sum (y_i - \bar{y})^2}{n_2 - 1}$$

Degrees of freedom are  $v_1 = n_1 - 1$ ,  $v_2 = n_2 - 1$ .  $n_1 = \text{first sample size}$   
 $n_2 = \text{second sample size}$

Note:- (1) we will take greater of the variances  $s_1^2$  or  $s_2^2$  in the numerator and adjust for the degrees of freedom.

$$\text{i.e. } F = \frac{\text{Greater Variance}}{\text{Smaller Variance}}$$

(2) If sample variance  $s^2$  is given, we can obtain population variance  $\sigma^2$  by using the relation  $n\sigma^2 = (n-1)s^2$  and vice versa.

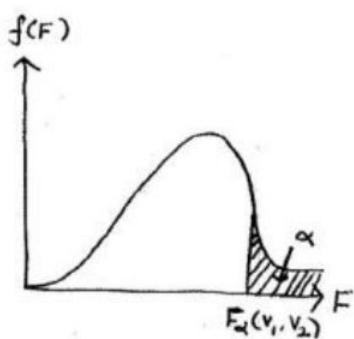
(3)  $F_\alpha(v_1, v_2)$  is the value of F with  $v_1$  and  $v_2$  degrees of freedom such that the area under the F-distribution to the right of  $F_\alpha$  is  $\alpha$ . In Tables  $F_\alpha$  tabulated for  $\alpha = 0.05$  and  $\alpha = 0.01$  for various combinations of d.o.f  $v_1$  and  $v_2$ . Clearly, the value of F at 5% significance is lower than 1%.

properties of F-distribution:

- ① F-distribution curve lies entirely in first quadrant.
- ② The F-curve depends not only on the two parameters but also on the order in which they are stated.

③  $F_{1-\alpha}(v_1, v_2) = \frac{1}{F_\alpha(v_2, v_1)}$  where  $F_\alpha(v_1, v_2)$  is the value of F with  $v_1$  and  $v_2$  d.o.f such that the area under the F-distribution curve to the right of  $F_\alpha$  is  $\alpha$ .

(4) The mode of F-distribution is less than unity.



Examples

① For an F-distribution, find (a)  $F_{0.05}$  with  $v_1=7$  and  $v_2=15$   
(b)  $F_{0.99}$  with  $v_1=28$  and  $v_2=12$ .

Sol @ From the Table,  $F_{0.05}$  with  $v_1=7$  and  $v_2=15$  is 2.71

$$⑥ F_{0.99}(28, 12) = \frac{1}{F_{0.01}(12, 28)} = \frac{1}{2.90} = 0.3448211$$

① The measurements of the output of two units have given the following results. Assuming the both samples have been obtained from the populations at 10% significant level, Test the whether the two populations have the Same Variance.

Unit - A	14.1	10.1	14.7	13.7	14.0
Unit - B	14.0	14.5	13.7	12.7	14.1

< June 2017 >

Sol Given  $n_1 = 5$ ,  $n_2 = 5$ ,  $\bar{x} = \frac{\sum x_i}{n_1} = \frac{1}{5}(14.1 + 10.1 + 14.7 + 13.7 + 14.0) = 13.32$

$$\bar{y} = \frac{\sum y_i}{n_2} = \frac{1}{5}(14 + 14.5 + 13.7 + 12.7 + 14.1) = 13.8$$

calculation for mean & and S.D of Samples :-

$x_i$	$(x_i - \bar{x})$	$(x_i - \bar{x})^2$	$y_i$	$(y_i - \bar{y})$	$(y_i - \bar{y})^2$
14.1	0.78	0.6084	14.0	0.2	0.04
10.1	-3.22	10.3684	14.5	0.7	0.49
14.7	1.38	1.9044	13.7	-0.1	0.01
13.7	0.38	0.1444	12.7	-1.1	1.21
14.0	0.68	0.4624	14.1	0.3	0.09
		$\sum (x_i - \bar{x})^2 = 13.488$			$\sum (y_i - \bar{y})^2 = 1.84$

$$S_1^2 = \frac{\sum (x_i - \bar{x})^2}{n_1 - 1} = \frac{13.488}{4} = 3.372$$

$$S_2^2 = \frac{\sum (y_i - \bar{y})^2}{n_2 - 1} = \frac{1.84}{4} = 0.46$$

(i) Null Hypothesis  $H_0: \sigma_1^2 = \sigma_2^2$

(ii) Alternative Hypothesis  $H_1: \sigma_1^2 \neq \sigma_2^2$

(iii) Level of Significance:  $\alpha = 10\% = 0.01$ .

(iv) Test Statistic:  $F = \frac{s_1^2}{s_2^2} < s_1^2 > s_2^2$

$$F_{\text{cal}} = \frac{3.372}{0.46} = 7.33.$$

$$\therefore |F_{\text{cal}}| = 7.33.$$

(v) Conclusion: Degrees of freedom are  $(n_1-1, n_2-1)$ .  
i.e.  $(4, 4)$ .

$\therefore$  Tabulated F at 10% level for  $(4, 4)$  degrees of freedom is 15.98

$$\text{i.e. } F_{0.01}(4, 4) = 15.98$$

$$\text{i.e. } |F_{\text{cal}}| = 15.98$$

$$\therefore |F_{\text{cal}}| < |F_{\text{tab}}|$$

$\therefore H_0$  is Accepted

i.e. There is no significant difference between the Variances.

② In one sample of 10 observations, the sum of the squares of the deviations of the sample values from sample mean was 120 and in the other sample of 12 observations, it was 314. Test whether the difference is significant at 5% level.

Qd Given  $n_1 = 10, n_2 = 12, \sum (x_i - \bar{x})^2 = 120, \sum (y_i - \bar{y})^2 = 314$ .

$$\therefore s_1^2 = \frac{\sum (x_i - \bar{x})^2}{n_1 - 1} = \frac{120}{9} = 13.33$$

$$s_2^2 = \frac{\sum (y_i - \bar{y})^2}{n_2 - 1} = \frac{314}{11} = 28.54$$

(i) Null Hypothesis  $H_0: \sigma_1^2 = \sigma_2^2$  (ii) Alternative Hypothesis  $H_1: \sigma_1^2 \neq \sigma_2^2$

(iii) Level of Significance:  $\alpha = 5\% = 0.05$  (iv) Test Statistic  $F = \frac{s_2^2}{s_1^2} < s_2^2 > s_1^2$

$$\therefore F_{\text{cal}} = \frac{28.54}{13.33} = 2.14. \quad \therefore |F_{\text{cal}}| = 2.14$$

(v) Conclusion: Degrees of freedom are  $(n_2-1, n_1-1) = (11, 9)$

$\therefore$  Tabulated F at 5% level for  $(11, 9)$  is 2.90

$$\text{i.e. } |F_{\text{tab}}| = 2.9$$

$$\therefore |F_{\text{cal}}| < |F_{\text{tab}}|$$

$\therefore H_0$  is accepted.

5

- ③ Two random samples gave the following results

Sample	size	Sample mean	Sum of squares of deviation from mean
1	10	15	90
2	12	14	108

check whether the samples came from the same normal population. (June 2016)

Sd

To test whether the two samples came from same normal population, i.e. ( $\mu_1 = \mu_2$  and  $\sigma_1^2 = \sigma_2^2$ ), we have to test (I) The Equality of Variances by F-test (II) The Equality of means by using t-Test.

(I) F-test (To test Equality of Variances)

$$\text{Given } n_1 = 10, \bar{x} = 15, \sum (x_i - \bar{x})^2 = 90 \\ n_2 = 12, \bar{y} = 14, \sum (y_i - \bar{y})^2 = 108$$

$$S_1^2 = \frac{\sum (x_i - \bar{x})^2}{n_1 - 1} = \frac{90}{9} = 10$$

$$S_2^2 = \frac{\sum (y_i - \bar{y})^2}{n_2 - 1} = \frac{108}{11} = 9.82$$

Null Hypothesis  $H_0: \sigma_1^2 = \sigma_2^2$

Alternative Hypothesis  $H_1: \sigma_1^2 \neq \sigma_2^2$

Level of significance:  $\alpha = 5\% = 0.05$

Test statistic :  $F = \frac{S_1^2}{S_2^2}$   $\because S_1^2 > S_2^2$

$$F_{\text{cal}} = \frac{10}{9.82} = 1.018$$

$$\therefore |F_{\text{cal}}| = 1.018$$

Conclusion: Degree of freedom  $(n_1 - 1, n_2 - 1) = (9, 11)$

$\therefore$  Tabulated F at 5% level for (9, 11) d.o.f is 2.90

$$\text{i.e. } |F_{\text{tab}}| = 2.90$$

$$\therefore |F_{\text{cal}}| < |F_{\text{Tab}}|$$

$\therefore H_0$  is accepted.

$$\text{i.e. } \boxed{\sigma_1^2 = \sigma_2^2}$$

(II)  $t$ -Test (To test Equality of means)

Given  $\bar{x} = 15$ ,  $\bar{y} = 14$ ,  $n_1 = 10$ ,  $n_2 = 12$

$$\sum (x_i - \bar{x})^2 = 90, \sum (y_i - \bar{y})^2 = 108$$

$$\therefore s^2 = \frac{1}{n_1 + n_2 - 2} \left[ \sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2 \right]$$

$$= \frac{1}{20} [90 + 108] = 9.9$$

$$\boxed{s^2 = 9.9}$$

Null Hypothesis  $H_0: \mu_1 = \mu_2$

Alternative Hypothesis  $H_1: \mu_1 \neq \mu_2$

Level of Significance:  $\alpha = 5\% = 0.05$ .

$$\text{Test Statistic: } t = \frac{\bar{x} - \bar{y}}{\sqrt{s^2 \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}} = \frac{15 - 14}{\sqrt{9.9 \left( \frac{1}{10} + \frac{1}{12} \right)}} = 0.74$$

$$\therefore |t_{\text{cal}}| = 0.74$$

Conclusion: Degrees of freedom  $= n_1 + n_2 - 2 = 20$

$$|t_{\text{Tab}}| = 2.086 \quad \leftarrow \text{See t-Table 20 at } \frac{0.05}{2} = 0.025$$

$$\therefore |t_{\text{cal}}| < |t_{\text{Tab}}|$$

$\therefore H_0$  is accepted.

$$\text{i.e. } \boxed{\mu_1 = \mu_2}$$

Hence from I and II, The given samples have been

drawn from the same normal population

Hence we accept the Null Hypothesis  $\sigma_1^2 = \sigma_2^2$  and  
 $\mu_1 = \mu_2$

(A) The nicotine contents in milligrams in two samples of tobacco were found to be as follows.

Sample A	24	27	26	21	25	-
Sample B	27	30	28	31	22	36

Can it be said that the two samples have come from the same normal population? (DEC 2016)

Sol To test whether the two samples came from same population (i.e.  $\sigma_1^2 = \sigma_2^2$  and  $\mu_1 = \mu_2$ ), we have to test (I) the equality of variances by using F-Test (II) The equality of means by using t-Test.

Given  $n_1 = 5$ ,  $n_2 = 6$

Calculation for mean & S.D of samples.

$x_i$	$x_i - \bar{x}$	$(x_i - \bar{x})^2$	$y_i$	$(y_i - \bar{y})$	$(y_i - \bar{y})^2$
24	0.6	0.36	27	-2	4
27	2.4	5.76	30	1	1
26	1.4	1.96	28	-1	1
21	3.6	12.96	31	2	4
25	0.4	0.16	22	-7	49
			36	7	49
123		21.2	174		108

$$\therefore \bar{x} = \frac{\sum x_i}{n_1} = \frac{123}{5} = 24.6, \quad \bar{y} = \frac{\sum y_i}{n_2} = \frac{174}{6} = 29$$

$$\sum (x_i - \bar{x})^2 = 21.2, \quad \sum (y_i - \bar{y})^2 = 108.$$

$$\therefore s_1^2 = \frac{\sum (x_i - \bar{x})^2}{n_1 - 1} = \frac{21.2}{4} = 5.3$$

$$s_2^2 = \frac{\sum (y_i - \bar{y})^2}{n_2 - 1} = \frac{108}{5} = 21.6.$$

(I) F-Test (To test equality of variances)

Null Hypothesis  $H_0: \sigma_1^2 = \sigma_2^2$

Alternative Hypothesis  $H_1: \sigma_1^2 \neq \sigma_2^2$

Level of Significance  $\therefore \alpha = 5\% = 0.05$

Test Statistic:  $F = \frac{s_2^2}{s_1^2} \quad (\because s_2^2 > s_1^2)$

$$F_{\text{cal}} = \frac{21.6}{5.3} = 4.075$$

$$\therefore |F_{\text{cal}}| = 4.075.$$

Conclusion: Degrees of freedom  $(n_2-1, n_1-1) = (5, 4)$ .

Tabulated Value of F at 0.05 with (5, 4) d.o.f is 6.26

i.e.  $F_{0.05}(5, 4) = 6.26$

i.e.  $|F_{\text{tab}}| = 6.26$

$\therefore |F_{\text{cal}}| < |F_{\text{tab}}|$

$\therefore H_0$  is accepted.

i.e.  $s_1^2 = s_2^2$

(II) t-Test (To test equality of means)

From the given data  $\bar{x} = 24.6, \bar{y} = 29$

$$\sum (x_i - \bar{x})^2 = 21.2, \sum (y_i - \bar{y})^2 = 108$$

$$S^2 = \frac{1}{n_1+n_2-2} \left[ \sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2 \right]$$

$$S^2 = \frac{1}{5+6-2} [21.2 + 108] = 14.35$$

Null Hypothesis  $H_0: \mu_1 = \mu_2$

Alternative Hypothesis  $H_1: \mu_1 \neq \mu_2$

Level of Significance:  $\alpha = 5\% = 0.05$

Test Statistic:  $t_{\text{cal}} = \frac{\bar{x} - \bar{y}}{\sqrt{S^2(\frac{1}{n_1} + \frac{1}{n_2})}} = \frac{24.6 - 29}{\sqrt{14.35(\frac{1}{5} + \frac{1}{6})}} = -1.92$

$$|t_{\text{cal}}| = 1.92$$

Conclusion: Degree of freedom  $= n_1 + n_2 - 2 = 9$

$$|t_{\text{Tab}}| = 2.26 \quad \leftarrow \text{See t-Table 9 at } \frac{0.05}{2} = 0.025.$$

$\therefore |t_{\text{cal}}| < |t_{\text{Tab}}|$

$\therefore H_0$  is accepted

i.e.  $\boxed{\mu_1 = \mu_2}$

$\therefore$  From (I) and (II) we conclude that two samples come from the same normal population.

## $\chi^2$ - Test for goodness of fit

Suppose we are given a set of "observed frequencies" obtained under some experiment and we want to test if the experimental results support a particular hypothesis or theory. Karl Pearson developed a test for testing the significance of discrepancy between experimental values and theoretical values obtained under some theory or hypothesis.

This Test known as  $\chi^2$ - test of goodness of fit. Karl Pearson proved the test statistic

$$\chi^2 = \sum \frac{(O - E)^2}{E}, \text{ where } O = \begin{matrix} \text{Observed frequency} \\ (\text{Experimental}) \end{matrix}$$

$E = \begin{matrix} \text{Expected frequency} \\ (\text{Theoretical}) \end{matrix}$

$\chi^2$ - Test is used to "test whether the difference between observed and Expected frequencies are significant"

Note:- ① If the data is given in series of "n" numbers  
Then D.O.F = n-1.

② In Case of Binomial distribution D.O.F = n-1

In Case of Poisson distribution D.O.F = n-2

In Case of Normal distribution D.O.F = n-3.

Properties of  $\chi^2$ - distribution;

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①  $\chi^2$ - distribution curve is not symmetrical, lies entirely in the first quadrant and hence not a normal curve, since  $\chi^2$  values from 0 to  $\infty$ .

② It depends on the degrees of freedom v.

③ If  $v_1, v_2$  are the degree of freedom of  $\chi^2_1$  and  $\chi^2_2$  then  $(v_1 + v_2)$  is the degree of freedom of  $\chi^2_1 + \chi^2_2$

(4) Mean =  $\nu$  and Variance =  $2\nu$ .

Applications of  $\chi^2$ -distribution:

- (i) To test the goodness of fit
- (ii) To test the independence of attributes.
- (iii) To test if the population has a specified value of the variance  $\sigma^2$ .

① The number of automobile accidents per week in a certain community are as follows: 12, 8, 20, 2, 14, 10, 15, 6, 9, 4. Are these frequencies in agreement with the belief that accident conditions were the same during this 10 week period.

Sol: Expected frequency of accidents each week =  $\frac{100}{10} = 10$

Null Hypothesis  $H_0$ : The accident conditions were the same during the 10 weeks.

Alternative Hypothesis  $H_1$ : The accident conditions are different during the 10 weeks.

| Observed Frequency ( $O$ ) | Expected Frequency ( $E$ ) | $(O - E)$ | $\frac{(O - E)^2}{E}$ |
|----------------------------|----------------------------|-----------|-----------------------|
| 12                         | 10                         | 2         | 0.4                   |
| 8                          | 10                         | -2        | 0.4                   |
| 20                         | 10                         | 10        | 10.0                  |
| 2                          | 10                         | -8        | 6.4                   |
| 14                         | 10                         | 4         | 1.6                   |
| 10                         | 10                         | 0         | 0.0                   |
| 15                         | 10                         | 5         | 2.5                   |
| 6                          | 10                         | -4        | 1.6                   |
| 9                          | 10                         | -1        | 0.1                   |
| 4                          | 10                         | -6        | 3.6                   |
| 100                        | 100                        |           | 26.6                  |

Test statistic:

$$\text{Now } \chi^2 = \sum \frac{(O-E)^2}{E}$$

$$\chi^2 = 26.6$$

$$\therefore |\chi^2_{\text{cal}}| = 26.6$$

Conclusion:

Degrees of freedom =  $n-1 = 10-1 = 9$

Since no. of observations =  $n=10$ .

The Tabulated value of  $\chi^2$  for 9 d.o.f at 5% level of Significance  
is 16.9

$$\therefore |\chi^2_{\text{Tabl}}| = 16.9$$

$$\therefore |\chi^2_{\text{cal}}| = |\chi^2_{\text{Tabl}}|$$

$\angle \text{See } \chi^2 \text{ Table 9 at 0.05} \\ = 16.919$

$\therefore H_0$  is rejected.

i.e The accidents conditions are different during  
the 10 week period

- ③ A sample analysis of Examination results of 500 students was made. It was found that 220 students had failed, 170 has secured a third class, 90 were placed in Second class and 20 got a first class. Do these figures Commensurate with the general examination result which is in the ratio of 4:3:2:1 for the various categories respectively.

Sol Total frequency =  $220 + 170 + 90 + 2 = 500$

Expected frequencies are in the ratio of 4:3:2:1

$$\text{Sum of Ratio} = 4+3+2+1 = 10$$

If we divide the total frequency 500 in the ratio 4:3:2:1,

we get Expected frequencies as  $\frac{4}{10} \times 500 = 200$ ,  $\frac{3}{10} \times 500 = 150$ ,  $\frac{2}{10} \times 500 = 100$ ,

$$\text{and } \frac{1}{10} \times 500 = 50.$$

i.e Expected frequencies are 200, 150, 100, 50.

| Class  | Observed Frequency ( $O$ ) | Expected Frequency ( $E$ ) | $(O-E)$ | $\frac{(O-E)^2}{E}$ |
|--------|----------------------------|----------------------------|---------|---------------------|
| Failed | 220                        | 200                        | 20      | 2.00                |
| Third  | 170                        | 150                        | 20      | 2.667               |
| Second | 90                         | 100                        | -10     | 1.000               |
| First  | 20                         | 50                         | -30     | 18.00               |
|        | 500                        | 500                        |         | 23.667              |

Null Hypotheses  $H_0$ : The observed results Commensurate with the general Examination results.

Alternative Hypotheses  $H_1$ : The observed results don't Commensurate with the general Examination results.

Level of Significance :  $\alpha = 5\%$ .

Test statistic:  $\chi^2 = \sum \frac{(O-E)^2}{E} = 23.667$   
 $\therefore |\chi^2_{\text{cal}}| = 23.667$

Conclusion: Degrees of freedom =  $(n-1) = 4-1 = 3$ .

Tabulated Value of  $\chi^2$  at 5% level for 3 d.o.f = 7.81

i.e  $|\chi^2_{\text{Tab}}| = 7.81$ .

$\therefore |\chi^2_{\text{cal}}| > |\chi^2_{\text{Tab}}|$

$\therefore H_0$  is rejected.

$\begin{cases} \text{see } \chi^2 \text{ Table 3 at } 0.05 \\ = 7.815 \end{cases}$

i.e The observed results are not Commensurate with the general Examination results.

② A die is thrown 264 times with the following results. Show that the die is biased ( Given  $\chi^2_{0.05} = 11.07$  for 5 d.o.f ).

|                         |    |    |    |    |    |    |
|-------------------------|----|----|----|----|----|----|
| No. appeared on the die | 1  | 2  | 3  | 4  | 5  | 6  |
| Frequency               | 40 | 32 | 28 | 58 | 54 | 52 |

Sol Expected frequency of each of the numbers 1, 2, 3, 4, 5, 6 is  $\frac{264}{6} = 44$   
 Null Hypothesis  $H_0$ : The die is unbiased.  
 Alternative Hypothesis  $H_1$ : The die is biased

Level of Significance :  $\alpha = 5\% = 0.05$

Test Statistic:  $\chi^2 = \sum \frac{(O-E)^2}{E}$

| No. appeared on the die | observed frequency (O) | Expected frequency (E) | $(O-E)^2$ | $\frac{(O-E)^2}{E}$ |
|-------------------------|------------------------|------------------------|-----------|---------------------|
| 1                       | 40                     | 44                     | 16        | 0.3636              |
| 2                       | 32                     | 44                     | 144       | 3.2727              |
| 3                       | 28                     | 44                     | 256       | 5.8181              |
| 4                       | 58                     | 44                     | 196       | 4.4545              |
| 5                       | 54                     | 44                     | 100       | 2.2727              |
| 6                       | 52                     | 44                     | 64        | 1.4545              |
|                         | 264                    | 264                    |           | 17.6362             |

$$\therefore \chi^2 = \sum \frac{(O-E)^2}{E} = 17.6362$$

$$\text{i.e. } |\chi^2_{\text{cal}}| = 17.6362$$

Conclusion: Degrees of freedom =  $n-1 = 5$

Tabulated Value of  $\chi^2$  for 5 d.o.f at 5% level = 11.07

$$\text{i.e. } |\chi^2_{\text{Tab}}| = 11.07$$

$$|\chi^2_{\text{cal}}| > |\chi^2_{\text{Tab}}|$$

$\leftarrow$  See  $\chi^2$  Table 5 at 0.05  
 $= 11.07$

$\therefore H_0$  is rejected

i.e. The die is biased !!

(4) A pair of dice are thrown 360 times and frequency of each sum is indicated below.

| Sum       | 2 | 3  | 4  | 5  | 6  | 7  | 8  | 9  | 10 | 11 | 12 |
|-----------|---|----|----|----|----|----|----|----|----|----|----|
| Frequency | 8 | 24 | 35 | 37 | 44 | 65 | 51 | 42 | 26 | 14 | 14 |

Would you say that the dice are fair on the basis of the  $\chi^2$ -test at 0.05 level of significance.

5d  
Null Hypothesis  $H_0$ : The dice are fair

Sum of frequency = 360

Alternative Hypothesis  $H_1$ : The dice are not fair

Level of Significance:  $\alpha = 5\% = 0.05$ .

The probabilities of getting Sum 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12 are

| $x$    | 2              | 3              | 4              | 5              | 6              | 7              | 8              | 9              | 10             | 11             | 12             |
|--------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|
| $P(x)$ | $\frac{1}{36}$ | $\frac{2}{36}$ | $\frac{3}{36}$ | $\frac{4}{36}$ | $\frac{5}{36}$ | $\frac{6}{36}$ | $\frac{5}{36}$ | $\frac{4}{36}$ | $\frac{3}{36}$ | $\frac{2}{36}$ | $\frac{1}{36}$ |

| Sum | Observed Frequency ( $O$ ) | Expected Frequency<br>$E = 360 \times P(x)$ | $(O - E)^2$ | $\frac{(O - E)^2}{E}$ |
|-----|----------------------------|---------------------------------------------|-------------|-----------------------|
| 2   | 8                          | 10                                          | 4           | 0.4                   |
| 3   | 24                         | 20                                          | 16          | 0.8                   |
| 4   | 35                         | 30                                          | 25          | 0.833                 |
| 5   | 37                         | 40                                          | 9           | 0.225                 |
| 6   | 44                         | 50                                          | 36          | 0.72                  |
| 7   | 65                         | 60                                          | 25          | 0.417                 |
| 8   | 51                         | 50                                          | 1           | 0.02                  |
| 9   | 42                         | 40                                          | 4           | 0.1                   |
| 10  | 26                         | 30                                          | 16          | 0.53                  |
| 11  | 14                         | 20                                          | 86          | 1.8                   |
| 12  | 14                         | 10                                          | 16          | 1.6                   |
|     | 360                        | 360                                         |             | 7.445                 |

$$\chi^2 = \sum \frac{(O - E)^2}{E} = 7.445 \text{ i.e. } |\chi^2_{\text{cal}}| = 7.445$$

Conclusion: Degrees of freedom =  $n-1 = 11-1 = 10$

$$|\chi^2_{\text{Tab}}| = 18.3 \Rightarrow |\chi^2_{\text{cal}}| < |\chi^2_{\text{Tab}}|$$

$\therefore H_0$  is accepted  
i.e. The dice are fair.

See  $\chi^2$  Table 10 at 0.05  
= 18.307

⑤ Fit a poisson distribution to the following data and test the goodness of fit at 0.05 LOS:

| $x$    | 0   | 1   | 2   | 3  | 4  | 5 | 6 | 7 |
|--------|-----|-----|-----|----|----|---|---|---|
| $f(x)$ | 305 | 366 | 210 | 80 | 28 | 9 | 2 | 1 |

5d Null Hypothesis  $H_0$ : poisson distribution is a good fit  
Alternative Hypothesis  $H_1$ : poisson distribution is not a good fit.  
Level of Significance :  $\alpha = 0.05$

$$\text{Mean} = \lambda = \frac{\sum x f(x)}{\sum f(x)} = \frac{0 \times 305 + 1 \times 366 + 2 \times 210 + 3 \times 80 + 4 \times 28 + 5 \times 9 + 6 \times 2 + 7 \times 1}{1001}$$

$$= \frac{12.02}{1001}$$

$$\therefore \boxed{\lambda = 1.2}$$

$$\text{By poisson distn } P(x) = \frac{e^{-\lambda} \lambda^x}{x!}, x=0,1,2, \dots, 7.$$

$$P(x) = \frac{e^{1.2} (1.2)^x}{x!}$$

$$\therefore \text{Expected frequency } E = N \times P(x)$$

$$E = \frac{1001 \times e^{-1.2} (1.2)^x}{x!} = \frac{1001 \times 0.3 (1.2)^x}{x!}$$

$$\boxed{E = (300 \cdot 3) \left( \frac{(1.2)^x}{x!} \right)}$$

| $x$ | Observed Frequency ( $O$ ) | Expected Frequency ( $E$ )                           | $(O-E)^2$ | $\frac{(O-E)^2}{E}$ |
|-----|----------------------------|------------------------------------------------------|-----------|---------------------|
| 0   | 305                        | $300 \cdot 3 \times \frac{1.2^0}{0!} = 300 \cdot 3$  | 22.09     | 0.0736              |
| 1   | 366                        | $300 \cdot 3 \times \frac{1.2^1}{1!} = 360 \cdot 36$ | 31.810    | 0.0883              |
| 2   | 210                        | $300 \cdot 3 \times \frac{1.2^2}{2!} = 216.216$      | 38.639    | 0.1781              |
| 3   | 80                         | $300 \cdot 3 \times \frac{1.2^3}{3!} = 86.48$        | 41.990    | 0.4855              |
| 4   | 28                         | $300 \cdot 3 \times \frac{1.2^4}{4!} = 25.946$       | 4.219     | 0.1626              |
| 5   | 9                          | $300 \cdot 3 \times \frac{1.2^5}{5!} = 6.227$        | 7.689     | 1.2348              |
| 6   | 2                          | $300 \cdot 3 \times \frac{1.2^6}{6!} = 1.245$        | 0.570     | 0.4578              |
| 7   | 1                          | $300 \cdot 3 \times \frac{1.2^7}{7!} = 0.213$        | 0.619     | 2.9061              |
|     | 1001                       |                                                      |           | 5.5874              |

$$\chi^2 = \sum \frac{(O-E)^2}{E} = 5.5874$$

$$\therefore |\chi^2_{\text{cal}}| = 5.5874$$

Conclusion: Degrees of freedom =  $n-2 = 8-2 = 6$  See  $\chi^2$  Table 6 at 0.05

Tabulated Value of  $\chi^2$  at 5% level for 6 d.o.f = 12.592

$$\therefore |\chi^2_{\text{Tab}}| = 12.592$$

$$\therefore |\chi^2_{\text{cal}}| < |\chi^2_{\text{Tab}}|$$

$\therefore H_0$  is accepted  
i.e. poisson distribution is a good fit.

$\chi^2$  Test for Independence of Attributes:

Attribute: An Attribute means a quality or characteristic.

Examples of attributes are drinking, smoking, honesty, beauty etc.

An Attribute may be marked by its presence (position) or absence in a number of a given population.

Let us consider two attributes A and B.

A is divided into two classes and B is divided into two classes. The various cell frequencies can be expressed in the following table known as  $2 \times 2$  Contingency table.

|   |   |   |
|---|---|---|
| A | a | b |
| B | c | d |

| a     | b     | $a+b$       |
|-------|-------|-------------|
| c     | d     | $c+d$       |
| $a+c$ | $b+d$ | $N=a+b+c+d$ |

A-1

The Expected frequencies are given by

|                               |                               |             |
|-------------------------------|-------------------------------|-------------|
| $E(a) = \frac{(a+c)(a+b)}{N}$ | $E(b) = \frac{(b+d)(a+b)}{N}$ | $a+b$       |
| $E(c) = \frac{(a+c)(c+d)}{N}$ | $E(d) = \frac{(b+d)(c+d)}{N}$ | $c+d$       |
| $a+c$                         | $b+d$                         | $N=a+b+c+d$ |

Note:- In this  $\chi^2$  Test, The Test if two Attributes A and B under consideration are independent or not.

Null Hypothesis  $H_0$ : Attributes are independent.

Degrees of freedom =  $(r-1)(s-1)$  where  $r = \text{no. of rows}$   
 $s = \text{no. of columns}$ .

- ① On the basis of information given below about the treatment of 200 patients suffering from a disease, state whether the new treatment is comparatively superior to the conventional treatment.

|              | Favourable | Not Favourable | Total |
|--------------|------------|----------------|-------|
| New          | 60         | 30             | 90    |
| Conventional | 40         | 70             | 110   |

- Sol
- (i) Null Hypothesis  $H_0$ : No difference between new and conventional treatment
  - (ii) New and conventional treatments are independent.
  - (iii) Alternative Hypothesis  $H_1$ : New and conventional treatments are not independent.
  - (iv) Level of Significance :  $\alpha = 5\% = 0.05$ .

(v) Computations

|              | Favourable | Not favourable | Total |
|--------------|------------|----------------|-------|
| New          | 60         | 30             | 90    |
| Conventional | 40         | 70             | 110   |
| Total        | 100        | 100            | 200   |

Table of Expected frequencies

|                                   |                                   |     |
|-----------------------------------|-----------------------------------|-----|
| $\frac{100 \times 90}{200} = 45$  | $\frac{100 \times 90}{200} = 45$  | 90  |
| $\frac{100 \times 110}{200} = 55$ | $\frac{100 \times 110}{200} = 55$ | 110 |
| 100                               | 100                               | 200 |

| Observed Frequency ( $O$ ) | Expected Frequency ( $E$ ) | $(O-E)^2$ | $\frac{(O-E)^2}{E}$                               |
|----------------------------|----------------------------|-----------|---------------------------------------------------|
| 60                         | 45                         | 225       | 5                                                 |
| 30                         | 45                         | 225       | 5                                                 |
| 40                         | 55                         | 225       | 4.09                                              |
| 70                         | 55                         | 225       | 4.09                                              |
|                            |                            |           | $\Sigma \left[ \frac{(O-E)^2}{E} \right] = 18.18$ |

$$\therefore \chi^2 = \sum \left[ \frac{(O-E)^2}{E} \right]$$

$$= 18.18$$

$$\therefore |\chi^2_{\text{cal}}| = 18.18$$

Conclusion: Degrees of freedom =  $(3-1)(5-1) = (2-1)(2-1) = 1$

$$|\chi^2_{\text{Tab}}| = 3.841. \quad \begin{array}{l} \text{See } \chi^2 \text{ Table at } 0.05 \\ = 3.841 \end{array}$$

$$\therefore |\chi^2_{\text{cal}}| > |\chi^2_{\text{Tab}}|$$

$\therefore H_0$  is rejected.

i.e New and Conventional treatment are not independent.

- ② From the following data, find whether there is any significant liking in the habit of taking soft drinks among the categories of employees.

| Soft drinks | employees |          |          |
|-------------|-----------|----------|----------|
|             | Clerks    | Teachers | Officers |
| Pepsi       | 10        | 25       | 65       |
| Thums up    | 15        | 30       | 65       |
| Fanta       | 50        | 60       | 30       |

*< Dec 2017, 15 >*

Sol

- (i) Null Hypothesis  $H_0$ : There is no significant liking in the habit of taking soft drinks
- (ii) Alternative Hypothesis  $H_1$ : There is significant liking.
- (iii) Level of Significance:  $\alpha = 0.05$ .
- (iv) Computations

I A-2

| Soft drinks | employees |          |          | Total |
|-------------|-----------|----------|----------|-------|
|             | Clerks    | Teachers | Officers |       |
| Pepsi       | 10        | 25       | 65       | 100   |
| Thums up    | 15        | 30       | 65       | 110   |
| Fanta       | 50        | 60       | 30       | 140   |
| Total       | 75        | 115      | 160      | 350   |

Table of Expected frequencies

|                                    |                                     |                                     |     |
|------------------------------------|-------------------------------------|-------------------------------------|-----|
| $\frac{75 \times 100}{350} = 21.4$ | $\frac{115 \times 100}{350} = 32.9$ | $\frac{160 \times 100}{350} = 45.7$ | 100 |
| $\frac{75 \times 110}{350} = 23.6$ | $\frac{115 \times 110}{350} = 36.1$ | $\frac{160 \times 110}{350} = 50.3$ | 110 |
| $\frac{75 \times 140}{350} = 30$   | $\frac{115 \times 140}{350} = 46$   | $\frac{160 \times 140}{350} = 64$   | 140 |
| 75                                 | 115                                 | 160                                 | 350 |

| Observed Frequency ( $O$ ) | Expected Frequency ( $E$ ) | $(O-E)^2$ | $\frac{(O-E)^2}{E}$ |
|----------------------------|----------------------------|-----------|---------------------|
| 10                         | 21.4                       | 129.96    | 6.073               |
| 25                         | 32.9                       | 62.41     | 1.897               |
| 65                         | 45.7                       | 372.49    | 8.151               |
| 15                         | 23.6                       | 73.96     | 3.134               |
| 30                         | 36.1                       | 37.21     | 1.031               |
| 65                         | 50.3                       | 216.09    | 4.3                 |
| 50                         | 30                         | 400       | 13.333              |
| 60                         | 46                         | 196       | 4.261               |
| 30                         | 64                         | 1156      | 18.062              |
|                            |                            |           | 60.2425             |

$$\text{Test Statistic } \chi^2 = \sum \frac{(O-E)^2}{E} = 60.2425$$

$$|\chi^2_{\text{cal}}| = 60.2425.$$

Conclusion: Degrees of freedom =  $(2-1)(5-1) = (3-1)(3-1) = 4$ .

$$|\chi^2_{\text{Tab}}| = 9.488$$

$\leftarrow$  See  $\chi^2$  Table 4 at 0.05  
= 9.488

$$\therefore |\chi^2_{\text{cal}}| > |\chi^2_{\text{Tab}}|$$

$\therefore H_0$  is rejected.

i.e The Habit of taking soft drink depends on categories of employees.

H.W

- ③ Given the following Contingency table for hair colour & eye colour.  
Find the value of  $\chi^2$ . Is there good Association between the two?

|            |       | Hair colour |       |       | Total |
|------------|-------|-------------|-------|-------|-------|
|            |       | Fair        | Brown | Black |       |
| Eye colour | Blue  | 15          | 5     | 20    | 40    |
|            | Grey  | 20          | 10    | 20    | 50    |
|            | Brown | 25          | 15    | 20    | 60    |
| Total      |       | 60          | 30    | 60    | 150   |

Sol  
Hunks

$H_0$ : Hair and eye colour are independent

$H_1$ : Hair and eye colour are dependent

$\alpha = 0.05$

$$|\chi^2_{\text{cal}}| = 3.6458$$

$$\angle D.O.F = (3-1)(3-1) = 4.$$

$$|\chi^2_{\text{Tab}}| = 9.488$$

$H_0$  is accepted

i.e Hair and eye colour are independent.

i.e Hair and eye colour are not associated.

A-4

$\chi^2$ - Test for population Variance:

Suppose that a random sample is drawn from a normal population with mean  $\mu$  and Variance  $\sigma^2$

under the Null Hypothesis  $H_0$  that the population Variance  $\sigma^2$  has a Specified Value  $\sigma_0^2$ . Test statistic is

$$\chi^2 = \sum \frac{(x_i - \bar{x})^2}{\sigma_0^2} = \frac{n s^2}{\sigma_0^2}$$

It follows  $\chi^2$ - distribution with  $(n-1)$  d.o.f.

- ① A firm manufacturing rivets wants to limit Variations in their lengths as much as possible. The lengths (in cms) of rivets manufactured by new process are.

|      |      |      |      |      |
|------|------|------|------|------|
| 2.15 | 1.99 | 2.05 | 2.12 | 2.17 |
| 2.01 | 1.98 | 2.03 | 2.25 | 1.93 |

Examine whether the new process can be considered superior to the old if the old population has standard deviation  $0.145$  cm?

Sol Given  $n = 10$ ,  $\bar{x} = \frac{\sum x_i}{n} = \frac{20.68}{10} = 2.068$

$$s^2 = \frac{\sum (x_i - \bar{x})^2}{n-1} = \frac{0.09096}{9} = 0.0101$$

and  $\sigma_0 = 0.145$ .

(i) Null Hypothesis  $H_0: \sigma^2 = \sigma_0^2$

(ii) Alternative Hypothesis  $H_1: \sigma^2 \neq \sigma_0^2$

(iii) Level of Significance :  $\alpha = 0.05$

(iv) Test Statistic :  $\chi^2 = \frac{n s^2}{\sigma_0^2} = \frac{10 \times 0.0101}{(0.145)^2} = 4.8$

$$\therefore |\chi^2_{\text{cal}}| = 4.8$$

(v) Conclusion: d.o.f =  $n-1=9$

$$|\chi^2_{\text{Tab}}| = 16.919$$

$$\therefore |\chi^2_{\text{cal}}| < |\chi^2_{\text{Tab}}|$$

$\therefore H_0$  is accepted.

see  $\chi^2$  Table 9 at 0.05  
 $= 16.919$

② A random sample of size 20 from a normal population gives a mean of 42 and variance 25. Test the hypothesis that the population standard deviation is 8 at 5% level.

Sol. Given  $n=20$ ,  $\bar{x}=42$ ,  $s^2=25$ ,  $\sigma_0=8$

(i) Null Hypothesis  $H_0: \sigma^2 = \sigma_0^2$

(ii) Alternative Hypothesis  $H_1: \sigma^2 \neq \sigma_0^2$

(iii) Level of Significance:  $\alpha=0.05$

(iv) Test Statistic:  $\chi^2 = \frac{ns^2}{\sigma_0^2} = \frac{20 \times 25}{8^2} = 7.8125$   
 $\therefore |\chi^2_{\text{cal}}| = 7.8125$

(v) Conclusion:  $d.o.f = n-1 = 19$ . *see  $\chi^2$  Table 19 at 0.05*  
 $\therefore |\chi^2_{\text{Tab}}| = 30.14$   $= 30.14$

$\therefore |\chi^2_{\text{cal}}| < |\chi^2_{\text{Tab}}|$   
 $\therefore H_0 \text{ is accepted.}$



2.  $t_{\alpha}$ - Critical values of  $t$ -distribution

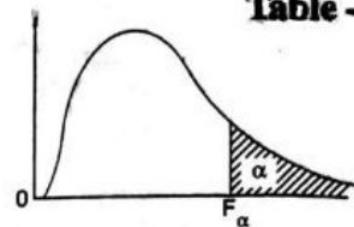
| v        | $\alpha$ |       |       |       |       |       |        |
|----------|----------|-------|-------|-------|-------|-------|--------|
|          | 0.40     | 0.30  | 0.20  | 0.15  | 0.10  | 0.05  | 0.025  |
| 1        | 0.325    | 0.727 | 1.376 | 1.963 | 3.078 | 6.314 | 12.706 |
| 2        | 0.289    | 0.617 | 1.061 | 1.386 | 1.886 | 20920 | 4.303  |
| 3        | 0.277    | 0.584 | 0.978 | 1.250 | 1.638 | 2.353 | 3.182  |
| 4        | 0.271    | 0.569 | 0.941 | 1.190 | 1.533 | 2.132 | 2.776  |
| 5        | 0.267    | 0.559 | 0.920 | 1.156 | 1.476 | 2.015 | 2.571  |
| 6        | 0.265    | 0.553 | 0.906 | 1.134 | 1.440 | 1.943 | 2.447  |
| 7        | 0.263    | 0.549 | 0.896 | 1.119 | 1.415 | 1.895 | 2.365  |
| 8        | 0.262    | 0.546 | 0.889 | 1.108 | 1.397 | 1.860 | 2.306  |
| 9        | 0.261    | 0.543 | 0.883 | 1.100 | 1.383 | 1.833 | 2.262  |
| 10       | 0.260    | 0.542 | 0.879 | 1.093 | 1.372 | 1.812 | 2.228  |
| 11       | 0.260    | 0.540 | 0.876 | 1.088 | 1.363 | 1.796 | 2.201  |
| 12       | 0.259    | 0.539 | 0.873 | 1.083 | 1.356 | 1.782 | 2.179  |
| 13       | 0.259    | 0.537 | 0.870 | 1.079 | 1.350 | 1.771 | 2.160  |
| 14       | 0.258    | 0.537 | 0.868 | 1.076 | 1.345 | 1.761 | 2.145  |
| 15       | 0.258    | 0.536 | 0.866 | 1.074 | 1.341 | 1.753 | 2.131  |
| 16       | 0.258    | 0.535 | 0.865 | 1.071 | 1.337 | 1.746 | 2.120  |
| 17       | 0.257    | 0.534 | 0.863 | 1.069 | 1.333 | 1.740 | 2.110  |
| 18       | 0.257    | 0.534 | 0.862 | 1.067 | 1.330 | 1.734 | 2.101  |
| 19       | 0.257    | 0.533 | 0.861 | 1.066 | 1.328 | 1.729 | 2.093  |
| 20       | 0.257    | 0.533 | 0.860 | 1.064 | 1.325 | 1.725 | 2.086  |
| 21       | 0.257    | 0.532 | 0.859 | 1.063 | 1.323 | 1.721 | 2.080  |
| 22       | 0.256    | 0.532 | 0.858 | 1.061 | 1.321 | 1.717 | 2.074  |
| 23       | 0.256    | 0.532 | 0.858 | 1.060 | 1.319 | 1.714 | 2.069  |
| 24       | 0.256    | 0.531 | 0.857 | 1.059 | 1.318 | 1.711 | 2.064  |
| 25       | 0.256    | 0.531 | 0.856 | 1.058 | 1.316 | 1.708 | 2.060  |
| 26       | 0.256    | 0.531 | 0.856 | 1.058 | 1.315 | 1.706 | 2.056  |
| 27       | 0.256    | 0.531 | 0.855 | 1.057 | 1.314 | 1.703 | 2.052  |
| 28       | 0.256    | 0.530 | 0.855 | 1.056 | 1.313 | 1.701 | 2.048  |
| 29       | 0.256    | 0.530 | 0.854 | 1.055 | 1.311 | 1.699 | 2.045  |
| 30       | 0.256    | 0.530 | 0.854 | 1.055 | 1.310 | 1.697 | 2.042  |
| 40       | 0.255    | 0.529 | 0.851 | 1.050 | 1.303 | 1.684 | 2.021  |
| 60       | 0.254    | 0.527 | 0.848 | 1.045 | 1.296 | 1.671 | 2.000  |
| 120      | 0.254    | 0.526 | 0.845 | 1.041 | 1.289 | 1.658 | 1.980  |
| $\infty$ | 0.253    | 0.524 | 0.842 | 1.036 | 1.282 | 1.645 | 1.960  |

3.  $t_{\alpha}$ -Critical values of  $t$ -distribution

| v        | $\alpha$ |        |        |        |        |         |         |
|----------|----------|--------|--------|--------|--------|---------|---------|
|          | 0.02     | 0.015  | 0.01   | 0.0075 | 0.005  | 0.0025  | 0.0005  |
| 1        | 15.895   | 21.205 | 31.821 | 42.434 | 63.657 | 127.322 | 636.590 |
| 2        | 4.849    | 5.643  | 6.965  | 8.073  | 9.925  | 14.089  | 31.598  |
| 3        | 3.482    | 3.896  | 4.541  | 5.047  | 5.841  | 7.453   | 12.924  |
| 4        | 2.999    | 3.298  | 3.747  | 4.088  | 4.604  | 5.598   | 8.610   |
| 5        | 2.757    | 3.003  | 3.365  | 3.634  | 4.032  | 4.773   | 6.869   |
| 6        | 2.612    | 2.829  | 3.143  | 3.372  | 3.707  | 4.317   | 5.959   |
| 7        | 2.517    | 2.715  | 2.998  | 3.203  | 3.499  | 4.029   | 5.408   |
| 8        | 2.449    | 2.634  | 2.896  | 3.085  | 3.355  | 3.833   | 5.041   |
| 9        | 2.398    | 2.574  | 2.821  | 2.998  | 3.250  | 3.690   | 4.781   |
| 10       | 2.359    | 2.527  | 2.764  | 2.932  | 3.169  | 3.581   | 4.587   |
| 11       | 2.328    | 2.491  | 2.718  | 2.879  | 3.106  | 3.497   | 4.437   |
| 12       | 2.303    | 2.461  | 2.681  | 2.836  | 3.055  | 3.428   | 4.318   |
| 13       | 2.282    | 2.436  | 2.650  | 2.801  | 3.012  | 3.372   | 4.221   |
| 14       | 2.264    | 2.415  | 2.624  | 2.771  | 2.977  | 3.326   | 4.140   |
| 15       | 2.249    | 2.397  | 2.602  | 2.746  | 2.947  | 3.286   | 4.073   |
| 16       | 2.235    | 2.382  | 2.583  | 2.724  | 2.921  | 3.252   | 4.051   |
| 17       | 2.224    | 2.368  | 2.567  | 2.706  | 2.898  | 3.222   | 3.965   |
| 18       | 2.214    | 2.356  | 2.552  | 2.689  | 2.878  | 3.197   | 3.922   |
| 19       | 2.205    | 2.346  | 2.539  | 2.674  | 2.861  | 3.174   | 3.883   |
| 20       | 2.197    | 2.336  | 2.528  | 2.661  | 2.845  | 3.153   | 3.849   |
| 21       | 2.189    | 2.328  | 2.518  | 2.649  | 2.831  | 3.135   | 3.819   |
| 22       | 2.183    | 2.320  | 2.508  | 2.639  | 2.819  | 3.119   | 3.792   |
| 23       | 2.177    | 2.313  | 2.500  | 2.629  | 2.807  | 3.104   | 3.768   |
| 24       | 2.172    | 2.307  | 2.492  | 2.620  | 2.797  | 3.091   | 3.745   |
| 25       | 2.167    | 2.301  | 2.485  | 2.612  | 2.787  | 3.078   | 3.725   |
| 26       | 2.162    | 2.296  | 2.479  | 2.605  | 2.779  | 3.067   | 3.707   |
| 27       | 2.158    | 2.291  | 2.473  | 2.598  | 2.771  | 3.057   | 3.690   |
| 28       | 2.154    | 2.286  | 2.467  | 2.592  | 2.763  | 3.047   | 3.674   |
| 29       | 2.150    | 2.282  | 2.462  | 2.586  | 2.756  | 3.038   | 3.659   |
| 30       | 2.147    | 2.278  | 2.457  | 2.581  | 2.750  | 3.030   | 3.646   |
| 40       | 2.125    | 2.250  | 2.423  | 2.542  | 2.704  | 2.971   | 3.551   |
| 60       | 2.099    | 2.223  | 2.390  | 2.504  | 2.660  | 2.915   | 3.460   |
| 120      | 2.076    | 2.196  | 2.358  | 2.468  | 2.617  | 2.860   | 3.373   |
| $\infty$ | 2.054    | 2.170  | 2.326  | 2.432  | 2.576  | 2.807   | 3.291   |

## Critical Values of the F-Distribution

**Table - 5**



| Values of $F_{0.05}(v_1, v_2)$ |       |       |       |       |       |       |       |       |       |  |
|--------------------------------|-------|-------|-------|-------|-------|-------|-------|-------|-------|--|
| $v_2$                          | $v_1$ |       |       |       |       |       |       |       |       |  |
|                                | 1     | 2     | 3     | 4     | 5     | 6     | 7     | 8     | 9     |  |
| 1                              | 161.4 | 199.5 | 215.7 | 224.6 | 230.2 | 234.0 | 236.8 | 238.9 | 240.5 |  |
| 2                              | 18.51 | 19.00 | 19.16 | 19.25 | 19.30 | 19.33 | 19.35 | 19.37 | 19.38 |  |
| 3                              | 10.13 | 9.55  | 9.28  | 9.12  | 9.01  | 8.94  | 8.89  | 8.85  | 8.81  |  |
| 4                              | 7.71  | 6.94  | 6.59  | 6.39  | 6.26  | 6.16  | 6.09  | 6.04  | 6.00  |  |
| 5                              | 6.61  | 5.79  | 5.41  | 5.19  | 5.05  | 4.95  | 4.88  | 4.82  | 4.77  |  |
| 6                              | 5.99  | 5.14  | 4.76  | 4.53  | 4.39  | 4.28  | 4.21  | 4.15  | 4.10  |  |
| 7                              | 5.59  | 4.74  | 4.35  | 4.12  | 3.97  | 3.87  | 3.79  | 3.73  | 3.68  |  |
| 8                              | 5.32  | 4.46  | 4.07  | 3.84  | 3.69  | 3.58  | 3.50  | 3.44  | 3.39  |  |
| 9                              | 5.12  | 4.26  | 3.86  | 3.63  | 3.48  | 3.37  | 3.29  | 3.23  | 3.18  |  |
| 10                             | 4.96  | 4.10  | 3.71  | 3.48  | 3.33  | 3.22  | 3.14  | 3.07  | 3.02  |  |
| 11                             | 4.84  | 3.98  | 3.59  | 3.36  | 3.20  | 3.09  | 3.01  | 2.95  | 2.90  |  |
| 12                             | 4.75  | 3.89  | 3.49  | 3.26  | 3.11  | 3.00  | 2.91  | 2.85  | 2.80  |  |
| 13                             | 4.67  | 3.81  | 3.41  | 3.18  | 3.03  | 2.92  | 2.83  | 2.77  | 2.71  |  |
| 14                             | 4.60  | 3.74  | 3.34  | 3.11  | 2.96  | 2.85  | 2.76  | 2.70  | 2.65  |  |
| 15                             | 4.54  | 3.68  | 3.29  | 3.06  | 2.90  | 2.79  | 2.71  | 2.64  | 2.59  |  |
| 16                             | 4.49  | 3.63  | 3.24  | 3.01  | 2.85  | 2.74  | 2.66  | 2.59  | 2.54  |  |
| 17                             | 4.45  | 3.59  | 3.20  | 2.96  | 2.81  | 2.70  | 2.61  | 2.55  | 2.49  |  |
| 18                             | 4.41  | 3.55  | 3.16  | 2.93  | 2.77  | 2.66  | 2.58  | 2.51  | 2.46  |  |
| 19                             | 4.38  | 3.52  | 3.13  | 2.90  | 2.74  | 2.63  | 2.54  | 2.48  | 2.42  |  |
| 20                             | 4.35  | 3.49  | 3.10  | 2.87  | 2.71  | 2.60  | 2.51  | 2.45  | 2.39  |  |
| 21                             | 4.32  | 3.47  | 3.07  | 2.84  | 2.68  | 2.57  | 2.49  | 2.42  | 2.37  |  |
| 22                             | 4.30  | 3.44  | 3.05  | 2.82  | 2.66  | 2.55  | 2.46  | 2.40  | 2.34  |  |
| 23                             | 4.28  | 3.42  | 3.03  | 2.80  | 2.64  | 2.53  | 2.44  | 2.37  | 2.32  |  |
| 24                             | 4.26  | 3.40  | 3.01  | 2.78  | 2.62  | 2.51  | 2.42  | 2.36  | 2.30  |  |
| 25                             | 4.24  | 3.39  | 2.99  | 2.76  | 2.60  | 2.49  | 2.40  | 2.34  | 2.28  |  |
| 26                             | 4.23  | 3.37  | 2.98  | 2.74  | 2.59  | 2.47  | 2.39  | 2.32  | 2.27  |  |
| 27                             | 4.21  | 3.35  | 2.96  | 2.73  | 2.57  | 2.46  | 2.37  | 2.31  | 2.25  |  |
| 28                             | 4.20  | 3.34  | 2.95  | 2.71  | 2.56  | 2.45  | 2.36  | 2.29  | 2.24  |  |
| 29                             | 4.18  | 3.33  | 2.93  | 2.70  | 2.55  | 2.43  | 2.35  | 2.28  | 2.22  |  |
| 30                             | 4.17  | 3.32  | 2.92  | 2.69  | 2.53  | 2.42  | 2.33  | 2.27  | 2.21  |  |
| 40                             | 4.08  | 3.23  | 2.84  | 2.61  | 2.45  | 2.34  | 2.25  | 2.18  | 2.12  |  |
| 60                             | 4.00  | 3.15  | 2.76  | 2.53  | 2.37  | 2.25  | 2.17  | 2.10  | 2.04  |  |
| 120                            | 3.92  | 3.07  | 2.68  | 2.45  | 2.29  | 2.17  | 2.09  | 2.02  | 1.96  |  |
| $\infty$                       | 3.84  | 3.00  | 2.60  | 2.37  | 2.21  | 2.10  | 2.01  | 1.94  | 1.88  |  |

| $v_2$    | Values of $F_{0.05}(v_1, v_2)$ |       |       |       |       |       |       |       |          |       |
|----------|--------------------------------|-------|-------|-------|-------|-------|-------|-------|----------|-------|
|          | $v_1$                          |       |       |       |       |       |       |       |          |       |
| 10       | 12                             | 15    | 20    | 24    | 30    | 40    | 60    | 120   | $\infty$ |       |
| 1        | 241.9                          | 243.9 | 245.9 | 248.0 | 249.1 | 250.1 | 251.1 | 252.2 | 253.3    | 254.3 |
| 2        | 19.40                          | 19.41 | 19.43 | 19.45 | 19.45 | 19.46 | 19.47 | 19.48 | 19.49    | 19.50 |
| 3        | 8.79                           | 8.74  | 8.70  | 8.66  | 8.64  | 8.62  | 8.59  | 8.57  | 8.55     | 8.53  |
| 4        | 5.96                           | 5.91  | 5.86  | 5.80  | 5.77  | 5.75  | 5.72  | 5.69  | 5.66     | 5.63  |
| 5        | 4.74                           | 4.68  | 4.62  | 4.56  | 4.53  | 4.50  | 4.46  | 4.43  | 4.40     | 4.36  |
| 6        | 4.06                           | 4.00  | 3.94  | 3.87  | 3.84  | 3.81  | 3.77  | 3.74  | 3.70     | 3.67  |
| 7        | 3.64                           | 3.57  | 3.51  | 3.44  | 3.41  | 3.38  | 3.34  | 3.30  | 3.27     | 3.23  |
| 8        | 3.35                           | 3.28  | 3.22  | 3.15  | 3.12  | 3.08  | 3.04  | 3.01  | 2.97     | 2.93  |
| 9        | 3.14                           | 3.07  | 3.01  | 2.94  | 2.90  | 2.86  | 2.83  | 2.79  | 2.75     | 2.71  |
| 10       | 2.98                           | 2.91  | 2.85  | 2.77  | 2.74  | 2.70  | 2.66  | 2.62  | 2.58     | 2.54  |
| 11       | 2.85                           | 2.79  | 2.72  | 2.65  | 2.61  | 2.57  | 2.53  | 2.49  | 2.45     | 2.40  |
| 12       | 2.75                           | 2.69  | 2.62  | 2.54  | 2.51  | 2.47  | 2.43  | 2.38  | 2.34     | 2.30  |
| 13       | 2.67                           | 2.60  | 2.53  | 2.46  | 2.42  | 2.38  | 2.34  | 2.30  | 2.25     | 2.21  |
| 14       | 2.60                           | 2.53  | 2.46  | 2.39  | 2.35  | 2.31  | 2.27  | 2.22  | 2.18     | 2.13  |
| 15       | 2.54                           | 2.48  | 2.40  | 2.33  | 2.29  | 2.25  | 2.20  | 2.16  | 2.11     | 2.07  |
| 16       | 2.49                           | 2.42  | 2.35  | 2.28  | 2.24  | 2.19  | 2.15  | 2.11  | 2.06     | 2.01  |
| 17       | 2.45                           | 2.38  | 2.31  | 2.23  | 2.19  | 2.15  | 2.10  | 2.06  | 2.01     | 1.96  |
| 18       | 2.41                           | 2.34  | 2.27  | 2.19  | 2.15  | 2.11  | 2.06  | 2.02  | 1.97     | 1.92  |
| 19       | 2.38                           | 2.31  | 2.23  | 2.16  | 2.11  | 2.07  | 2.03  | 1.98  | 1.93     | 1.88  |
| 20       | 2.35                           | 2.28  | 2.20  | 2.12  | 2.08  | 2.04  | 1.99  | 1.95  | 1.90     | 1.84  |
| 21       | 2.32                           | 2.25  | 2.18  | 2.10  | 2.05  | 2.01  | 1.96  | 1.92  | 1.87     | 1.81  |
| 22       | 2.30                           | 2.23  | 2.15  | 2.07  | 2.03  | 1.98  | 1.94  | 1.89  | 1.84     | 1.78  |
| 23       | 2.27                           | 2.20  | 2.13  | 2.05  | 2.01  | 1.96  | 1.91  | 1.86  | 1.81     | 1.76  |
| 24       | 2.25                           | 2.18  | 2.11  | 2.03  | 1.98  | 1.94  | 1.89  | 1.84  | 1.79     | 1.73  |
| 25       | 2.24                           | 2.16  | 2.09  | 2.01  | 1.96  | 1.92  | 1.87  | 1.82  | 1.77     | 1.71  |
| 26       | 2.22                           | 2.15  | 2.07  | 1.99  | 1.95  | 1.90  | 1.85  | 1.80  | 1.75     | 1.69  |
| 27       | 2.20                           | 2.13  | 2.06  | 1.97  | 1.93  | 1.88  | 1.84  | 1.79  | 1.73     | 1.67  |
| 28       | 2.19                           | 2.12  | 2.04  | 1.96  | 1.91  | 1.87  | 1.82  | 1.77  | 1.71     | 1.65  |
| 29       | 2.18                           | 2.10  | 2.03  | 1.94  | 1.90  | 1.85  | 1.81  | 1.75  | 1.70     | 1.64  |
| 30       | 2.16                           | 2.09  | 2.01  | 1.93  | 1.89  | 1.84  | 1.79  | 1.75  | 1.68     | 1.62  |
| 40       | 2.08                           | 2.00  | 1.92  | 1.84  | 1.79  | 1.74  | 1.69  | 1.64  | 1.58     | 1.51  |
| 60       | 1.99                           | 1.92  | 1.84  | 1.75  | 1.70  | 1.65  | 1.59  | 1.53  | 1.47     | 1.39  |
| 120      | 1.91                           | 1.83  | 1.75  | 1.66  | 1.61  | 1.55  | 1.50  | 1.43  | 1.35     | 1.25  |
| $\infty$ | 1.83                           | 1.75  | 1.67  | 1.57  | 1.52  | 1.46  | 1.39  | 1.32  | 1.22     | 1.00  |

|                |       | Values of $F_{0.01}(v_1, v_2)$ |       |       |       |       |       |       |       |   |
|----------------|-------|--------------------------------|-------|-------|-------|-------|-------|-------|-------|---|
|                |       | v <sub>1</sub>                 |       |       |       |       |       |       |       |   |
| v <sub>2</sub> |       | 1                              | 2     | 3     | 4     | 5     | 6     | 7     | 8     | 9 |
| 1              | 405.2 | 4999.5                         | 5403  | 5625  | 5764  | 5859  | 5928  | 5981  | 6022  |   |
| 2              | 98.50 | 99.00                          | 99.17 | 99.25 | 99.30 | 99.33 | 99.36 | 99.37 | 99.39 |   |
| 3              | 34.12 | 30.82                          | 29.46 | 28.71 | 28.24 | 27.91 | 27.67 | 27.67 | 27.35 |   |
| 4              | 21.20 | 18.00                          | 16.69 | 15.98 | 15.52 | 15.21 | 14.98 | 14.80 | 14.66 |   |
| 5              | 16.26 | 13.27                          | 12.06 | 11.39 | 10.97 | 10.67 | 10.46 | 10.29 | 10.16 |   |
| 6              | 13.75 | 10.92                          | 9.78  | 9.15  | 8.75  | 8.47  | 8.26  | 8.10  | 7.98  |   |
| 7              | 12.25 | 9.55                           | 8.45  | 7.85  | 7.46  | 7.19  | 6.99  | 6.84  | 6.72  |   |
| 8              | 11.26 | 8.65                           | 7.59  | 7.01  | 6.63  | 6.37  | 6.18  | 6.03  | 5.91  |   |
| 9              | 10.56 | 8.02                           | 6.99  | 6.42  | 6.06  | 5.80  | 5.61  | 5.47  | 5.35  |   |
| 10             | 10.04 | 7.56                           | 6.55  | 5.99  | 5.64  | 5.39  | 5.20  | 5.06  | 4.94  |   |
| 11             | 9.65  | 7.21                           | 6.22  | 5.67  | 5.32  | 5.07  | 4.89  | 4.74  | 4.63  |   |
| 12             | 9.33  | 6.93                           | 5.95  | 5.41  | 5.06  | 4.82  | 4.64  | 4.50  | 4.39  |   |
| 13             | 9.07  | 6.70                           | 5.74  | 5.21  | 4.86  | 4.62  | 4.44  | 4.30  | 4.19  |   |
| 14             | 8.86  | 6.51                           | 5.56  | 5.04  | 4.69  | 4.46  | 4.28  | 4.14  | 4.03  |   |
| 15             | 8.68  | 6.36                           | 5.42  | 4.89  | 4.56  | 4.32  | 4.14  | 4.00  | 3.89  |   |
| 16             | 8.53  | 6.23                           | 5.29  | 4.77  | 4.44  | 4.20  | 4.03  | 3.89  | 3.78  |   |
| 17             | 8.40  | 6.11                           | 5.18  | 4.67  | 4.34  | 4.10  | 3.93  | 3.79  | 3.68  |   |
| 18             | 8.29  | 6.01                           | 5.09  | 4.58  | 4.25  | 4.01  | 3.84  | 3.71  | 3.60  |   |
| 19             | 8.18  | 5.93                           | 5.01  | 4.50  | 4.17  | 3.94  | 3.77  | 3.63  | 3.52  |   |
| 20             | 8.10  | 5.85                           | 4.94  | 4.43  | 4.10  | 3.87  | 3.70  | 3.56  | 3.46  |   |
| 21             | 8.02  | 5.78                           | 4.87  | 4.37  | 4.04  | 3.81  | 3.64  | 3.51  | 3.40  |   |
| 22             | 7.95  | 5.72                           | 4.82  | 4.31  | 3.99  | 3.76  | 3.59  | 3.45  | 3.35  |   |
| 23             | 7.88  | 5.66                           | 4.76  | 4.26  | 3.94  | 3.71  | 3.54  | 3.41  | 3.30  |   |
| 24             | 7.82  | 5.61                           | 4.72  | 4.22  | 3.90  | 3.67  | 3.50  | 3.36  | 3.26  |   |
| 25             | 7.77  | 5.57                           | 4.68  | 4.18  | 3.85  | 3.63  | 3.46  | 3.32  | 3.22  |   |
| 26             | 7.72  | 5.53                           | 4.64  | 4.14  | 3.82  | 3.59  | 3.42  | 3.29  | 3.18  |   |
| 27             | 7.68  | 5.49                           | 4.60  | 4.11  | 3.78  | 3.56  | 3.39  | 3.26  | 3.15  |   |
| 28             | 7.64  | 5.45                           | 4.57  | 4.07  | 3.75  | 3.53  | 3.36  | 3.23  | 3.12  |   |
| 29             | 7.60  | 5.42                           | 4.54  | 4.04  | 3.73  | 3.50  | 3.33  | 3.20  | 3.09  |   |
| 30             | 7.56  | 5.39                           | 4.51  | 4.02  | 3.70  | 3.47  | 3.30  | 3.17  | 3.07  |   |
| 40             | 7.31  | 5.18                           | 4.31  | 3.83  | 3.51  | 3.29  | 3.12  | 2.99  | 2.89  |   |
| 60             | 7.08  | 4.98                           | 4.13  | 3.65  | 3.34  | 3.12  | 2.95  | 2.82  | 2.72  |   |
| 120            | 6.85  | 4.79                           | 3.95  | 3.48  | 3.17  | 2.96  | 2.79  | 2.66  | 2.56  |   |
| $\infty$       | 6.63  | 4.61                           | 3.78  | 3.32  | 3.02  | 2.80  | 2.80  | 2.51  | 2.41  |   |

| $v_2$    | Values of $F_{0.01}(v_1, v_2)$ |       |       |       |       |       |       |       |       |          |
|----------|--------------------------------|-------|-------|-------|-------|-------|-------|-------|-------|----------|
|          | $v_1$                          |       |       |       |       |       |       |       |       |          |
|          | 10                             | 12    | 15    | 20    | 24    | 30    | 40    | 60    | 120   | $\infty$ |
| 1        | 6056                           | 6106  | 6157  | 6209  | 6235  | 6261  | 6287  | 6313  | 6339  | 6366     |
| 2        | 99.40                          | 99.42 | 99.43 | 99.46 | 99.46 | 99.47 | 99.47 | 99.48 | 99.49 | 99.50    |
| 3        | 27.23                          | 27.05 | 26.87 | 26.60 | 26.60 | 26.50 | 26.41 | 26.32 | 26.22 | 26.13    |
| 4        | 14.55                          | 14.37 | 14.20 | 13.93 | 13.93 | 13.84 | 13.75 | 13.65 | 13.56 | 13.46    |
| 5        | 10.05                          | 9.89  | 9.72  | 9.47  | 9.47  | 9.38  | 9.29  | 9.20  | 9.11  | 9.02     |
| 6        | 7.87                           | 7.72  | 7.56  | 7.31  | 7.31  | 7.23  | 7.14  | 7.06  | 6.97  | 6.88     |
| 7        | 6.62                           | 6.47  | 6.31  | 6.07  | 6.07  | 5.99  | 5.91  | 5.82  | 5.74  | 5.65     |
| 8        | 5.81                           | 5.67  | 5.52  | 5.28  | 5.28  | 5.20  | 5.12  | 5.03  | 4.95  | 4.86     |
| 9        | 5.26                           | 5.11  | 4.96  | 4.73  | 4.73  | 4.65  | 4.57  | 4.48  | 4.40  | 4.31     |
| 10       | 4.85                           | 4.71  | 4.56  | 4.33  | 4.33  | 4.25  | 4.17  | 4.08  | 4.00  | 3.91     |
| 11       | 4.54                           | 4.40  | 4.25  | 4.02  | 4.02  | 3.94  | 3.86  | 3.78  | 3.69  | 3.60     |
| 12       | 4.30                           | 4.16  | 4.01  | 3.78  | 3.78  | 3.70  | 3.62  | 3.54  | 3.45  | 3.36     |
| 13       | 4.10                           | 3.96  | 3.82  | 3.59  | 3.59  | 3.51  | 3.43  | 3.34  | 3.25  | 3.17     |
| 14       | 3.94                           | 3.80  | 3.66  | 3.43  | 3.43  | 3.35  | 3.27  | 3.18  | 3.09  | 3.00     |
| 15       | 3.80                           | 3.67  | 3.52  | 3.29  | 3.29  | 3.21  | 3.13  | 3.05  | 2.96  | 2.87     |
| 16       | 3.69                           | 3.55  | 3.41  | 3.18  | 3.18  | 3.10  | 3.02  | 2.93  | 2.84  | 2.75     |
| 17       | 3.59                           | 3.46  | 3.31  | 3.08  | 3.08  | 3.00  | 2.92  | 2.83  | 2.75  | 2.65     |
| 18       | 3.51                           | 3.37  | 3.23  | 3.00  | 3.00  | 2.92  | 2.84  | 2.75  | 2.66  | 2.57     |
| 19       | 3.43                           | 3.30  | 3.15  | 2.92  | 2.92  | 2.84  | 2.76  | 2.67  | 2.58  | 2.49     |
| 20       | 3.37                           | 3.23  | 3.09  | 2.86  | 2.86  | 2.78  | 2.69  | 2.61  | 2.52  | 2.42     |
| 21       | 3.31                           | 3.17  | 3.03  | 2.80  | 2.80  | 2.72  | 2.64  | 2.55  | 2.46  | 2.36     |
| 22       | 3.26                           | 3.12  | 2.98  | 2.75  | 2.75  | 2.67  | 2.58  | 2.50  | 2.40  | 2.31     |
| 23       | 3.21                           | 3.07  | 2.93  | 2.70  | 2.70  | 2.62  | 2.54  | 2.45  | 2.35  | 2.26     |
| 24       | 3.17                           | 3.03  | 2.89  | 2.66  | 2.66  | 2.58  | 2.49  | 2.40  | 2.31  | 2.21     |
| 25       | 3.13                           | 2.99  | 2.85  | 2.62  | 2.62  | 2.54  | 2.45  | 2.36  | 2.27  | 2.17     |
| 26       | 3.09                           | 2.96  | 2.81  | 2.58  | 2.58  | 2.50  | 2.42  | 2.33  | 2.23  | 2.13     |
| 27       | 3.06                           | 2.93  | 2.78  | 2.55  | 2.55  | 2.47  | 2.38  | 2.29  | 2.20  | 2.10     |
| 28       | 3.03                           | 2.90  | 2.75  | 2.52  | 2.52  | 2.44  | 2.35  | 2.26  | 2.17  | 2.06     |
| 29       | 3.00                           | 2.87  | 2.73  | 2.49  | 2.49  | 2.41  | 2.33  | 2.23  | 2.14  | 2.03     |
| 30       | 2.98                           | 2.84  | 2.70  | 2.47  | 2.47  | 2.39  | 2.30  | 2.21  | 2.11  | 2.01     |
| 40       | 2.80                           | 2.66  | 2.52  | 2.29  | 2.29  | 2.20  | 2.11  | 2.02  | 1.92  | 1.80     |
| 60       | 2.63                           | 2.50  | 2.35  | 2.12  | 2.12  | 2.03  | 1.94  | 1.84  | 1.73  | 1.60     |
| 120      | 2.47                           | 2.34  | 2.19  | 1.95  | 1.95  | 1.86  | 1.76  | 1.66  | 1.53  | 1.38     |
| $\infty$ | 2.32                           | 2.18  | 2.04  | 1.79  | 1.79  | 1.70  | 1.59  | 1.47  | 1.32  | 1.00     |

#### 4. $\chi^2$ -Critical values of the chi-squared distribution

| v  | $\alpha$ |         |         |         |         |        |        |        |        |        |
|----|----------|---------|---------|---------|---------|--------|--------|--------|--------|--------|
|    | 0.995    | 0.99    | 0.98    | 0.975   | 0.95    | 0.90   | 0.80   | 0.75   | 0.75   | 0.50   |
| 1  | 0.04393  | 0.03157 | 0.03628 | 0.03982 | 0.00393 | 0.0158 | 0.0642 | 0.102  | 0.148  | 0.455  |
| 2  | 0.0100   | 0.0201  | 0.0404  | 0.103   | 0.103   | 0.211  | 0.446  | 0.575  | 0.713  | 1.386  |
| 3  | 0.0717   | 0.115   | 0.185   | 0.352   | 0.352   | 0.584  | 1.005  | 1.213  | 1.424  | 2.366  |
| 4  | 0.207    | 0.297   | 0.429   | 0.711   | 0.711   | 1.064  | 1.649  | 1.923  | 2.195  | 3.357  |
| 5  | 0.412    | 0.554   | 0.752   | 1.145   | 1.145   | 1.610  | 2.343  | 2.675  | 3.000  | 4.351  |
| 6  | 0.676    | 0.872   | 1.134   | 1.635   | 1.635   | 2.204  | 3.070  | 3.455  | 3.828  | 5.348  |
| 7  | 0.989    | 1.239   | 1.564   | 2.167   | 2.167   | 2.833  | 3.822  | 4.255  | 4.671  | 6.346  |
| 8  | 1.344    | 1.646   | 2.032   | 2.733   | 2.733   | 3.490  | 4.594  | 5.071  | 5.527  | 7.344  |
| 9  | 1.735    | 2.088   | 2.532   | 3.325   | 3.325   | 4.168  | 5.380  | 5.899  | 6.393  | 8.343  |
| 10 | 2.156    | 2.558   | 3.059   | 3.940   | 3.940   | 4.865  | 6.179  | 6.737  | 7.267  | 9.342  |
| 11 | 2.603    | 3.053   | 3.609   | 4.575   | 4.575   | 5.578  | 6.989  | 7.584  | 8.148  | 10.341 |
| 12 | 3.074    | 3.571   | 4.178   | 5.226   | 5.226   | 6.304  | 7.807  | 8.438  | 9.034  | 11.340 |
| 13 | 3.565    | 4.107   | 4.765   | 5.892   | 5.892   | 7.042  | 8.634  | 9.299  | 9.926  | 12.340 |
| 14 | 4.075    | 4.660   | 5.368   | 6.571   | 6.571   | 7.790  | 9.467  | 10.165 | 10.821 | 13.339 |
| 15 | 4.601    | 5.229   | 5.985   | 7.261   | 7.261   | 8.547  | 10.307 | 11.036 | 11.721 | 14.339 |
| 16 | 5.142    | 5.812   | 6.614   | 7.962   | 7.962   | 9.312  | 11.152 | 11.912 | 12.624 | 15.338 |
| 17 | 5.697    | 6.408   | 7.255   | 8.672   | 8.672   | 10.085 | 12.002 | 12.792 | 13.531 | 16.338 |
| 18 | 6.265    | 7.015   | 7.906   | 9.390   | 9.390   | 10.865 | 12.857 | 13.675 | 14.440 | 17.338 |
| 19 | 6.844    | 7.633   | 8.567   | 10.117  | 10.117  | 11.651 | 13.716 | 14.562 | 15.352 | 18.338 |
| 20 | 7.434    | 8.260   | 9.237   | 10.851  | 10.851  | 12.443 | 14.578 | 14.452 | 16.266 | 19.337 |
| 21 | 8.034    | 8.897   | 9.915   | 11.591  | 11.591  | 13.240 | 15.445 | 16.344 | 17.182 | 20.337 |
| 22 | 8.643    | 9.542   | 10.600  | 12.338  | 12.338  | 14.041 | 16.314 | 17.240 | 18.101 | 21.337 |
| 23 | 9.260    | 10.196  | 11.293  | 13.091  | 13.091  | 14.848 | 17.187 | 18.137 | 19.021 | 22.337 |
| 24 | 9.886    | 10.856  | 11.992  | 13.848  | 13.848  | 15.659 | 18.062 | 19.037 | 19.943 | 23.337 |
| 25 | 10.520   | 11.524  | 12.697  | 14.611  | 14.611  | 16.473 | 18.940 | 19.939 | 20.867 | 24.337 |
| 26 | 11.160   | 12.198  | 13.409  | 15.379  | 15.379  | 17.292 | 19.820 | 20.843 | 21.792 | 25.336 |
| 27 | 11.808   | 12.879  | 14.125  | 16.151  | 16.151  | 18.114 | 20.703 | 21.749 | 22.719 | 26.336 |
| 28 | 12.461   | 13.565  | 14.847  | 16.928  | 16.928  | 18.939 | 21.588 | 22.657 | 23.647 | 27.336 |
| 29 | 13.121   | 14.256  | 15.574  | 17.708  | 17.708  | 19.768 | 22.475 | 23.567 | 24.577 | 28.336 |
| 30 | 13.787   | 14.953  | 16.306  | 18.493  | 18.493  | 20.599 | 23.364 | 24.478 | 25.508 | 29.336 |

$\chi^2_{\alpha}$ - Critical values of the chi-squared distribution

| v  | $\alpha$ |        |        |        |        |        |        |        |        |        |
|----|----------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
|    | 0.30     | 0.25   | 0.20   | 0.10   | 0.05   | 0.025  | 0.02   | 0.01   | 0.005  | 0.001  |
| 1  | 1.074    | 1.323  | 1.642  | 2.706  | 3.841  | 5.024  | 5.412  | 6.635  | 7.879  | 10.827 |
| 2  | 2.408    | 2.773  | 3.219  | 4.605  | 5.991  | 7.378  | 7.824  | 9.210  | 10.597 | 13.815 |
| 3  | 3.665    | 4.108  | 4.642  | 6.251  | 7.815  | 9.348  | 9.837  | 11.345 | 12.838 | 16.268 |
| 4  | 4.878    | 5.385  | 5.989  | 7.779  | 9.488  | 11.143 | 11.668 | 13.277 | 14.860 | 18.465 |
| 5  | 6.064    | 6.626  | 7.289  | 9.236  | 11.070 | 12.832 | 13.388 | 15.086 | 16.750 | 20.517 |
| 6  | 7.231    | 7.841  | 8.558  | 10.645 | 12.592 | 14.449 | 15.033 | 16.812 | 18.548 | 22.457 |
| 7  | 8.383    | 9.037  | 9.803  | 12.017 | 14.067 | 16.013 | 16.622 | 18.475 | 20.278 | 24.322 |
| 8  | 9.524    | 10.219 | 11.030 | 13.362 | 15.507 | 17.535 | 18.168 | 20.090 | 21.955 | 26.125 |
| 9  | 10.656   | 11.389 | 12.242 | 14.684 | 16.919 | 19.023 | 19.679 | 21.666 | 23.589 | 27.877 |
| 10 | 11.781   | 12.549 | 13.442 | 15.987 | 18.307 | 20.483 | 21.161 | 23.209 | 25.188 | 29.588 |
| 11 | 12.899   | 13.701 | 14.631 | 17.275 | 19.675 | 21.920 | 22.618 | 24.725 | 26.757 | 31.264 |
| 12 | 14.011   | 14.845 | 15.812 | 18.545 | 21.026 | 23.337 | 24.054 | 26.217 | 28.300 | 32.909 |
| 13 | 15.119   | 15.984 | 16.985 | 19.812 | 22.362 | 24.736 | 25.472 | 27.688 | 29.819 | 34.528 |
| 14 | 16.222   | 17.117 | 18.151 | 21.064 | 23.685 | 26.119 | 26.873 | 29.141 | 31.319 | 36.123 |
| 15 | 17.322   | 18.245 | 19.311 | 22.307 | 24.996 | 27.488 | 28.259 | 30.578 | 32.801 | 37.697 |
| 16 | 18.418   | 19.369 | 20.465 | 23.542 | 26.296 | 28.845 | 29.633 | 32.000 | 34.267 | 39.252 |
| 17 | 19.511   | 20.489 | 21.615 | 24.769 | 27.587 | 30.191 | 30.995 | 33.409 | 35.718 | 40.790 |
| 18 | 20.601   | 21.605 | 22.760 | 25.989 | 28.869 | 31.526 | 32.346 | 34.805 | 37.156 | 42.312 |
| 19 | 21.689   | 22.718 | 23.900 | 27.204 | 30.144 | 32.852 | 33.687 | 36.191 | 38.582 | 43.820 |
| 20 | 22.775   | 23.828 | 25.038 | 28.412 | 31.410 | 34.170 | 35.020 | 37.566 | 39.997 | 45.315 |
| 21 | 23.858   | 24.935 | 26.171 | 29.615 | 32.671 | 35.479 | 36.343 | 38.932 | 41.401 | 46.797 |
| 22 | 24.939   | 26.039 | 27.301 | 30.813 | 33.924 | 36.781 | 37.659 | 40.289 | 42.796 | 48.268 |
| 23 | 26.018   | 27.141 | 28.429 | 32.007 | 35.172 | 38.076 | 38.968 | 41.638 | 44.181 | 49.728 |
| 24 | 27.096   | 28.241 | 29.553 | 33.196 | 36.415 | 39.364 | 40.270 | 42.980 | 45.558 | 51.179 |
| 25 | 28.172   | 29.339 | 30.675 | 34.382 | 37.652 | 40.646 | 41.566 | 44.314 | 46.928 | 52.620 |
| 26 | 29.246   | 30.434 | 31.795 | 35.563 | 38.885 | 41.923 | 42.856 | 45.642 | 48.290 | 54.052 |
| 27 | 30.319   | 31.528 | 32.912 | 36.741 | 40.113 | 43.194 | 44.140 | 46.963 | 49.645 | 55.476 |
| 28 | 31.391   | 32.620 | 34.027 | 37.916 | 41.337 | 44.461 | 45.419 | 48.278 | 50.993 | 56.893 |
| 29 | 32.461   | 33.711 | 35.139 | 39.087 | 42.557 | 45.722 | 46.693 | 49.588 | 52.336 | 58.302 |
| 30 | 33.530   | 34.800 | 36.250 | 40.256 | 43.773 | 46.979 | 47.962 | 50.892 | 53.672 | 59.703 |