

DS5616: Foundations of Statistical Learning

Test 1

Jan-May 2025

Date: 17 Feb 2025

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- The test comprises 3 questions for a total of 30 points and is for a duration of 50 minutes.
 - Write your fair answer *after* you have tried it out in the rough space. **Only work written in the space below the corresponding questions shall be graded.** If you do not wish something to be graded, strike it out neatly.
 - This is a closed-book test. You may not use the textbook, class notes, mobile phone, tablet, laptop, or such devices.
 - This is also an opportunity for you to evaluate yourself – in your best interests, be honest.

Question	Points	Score
1.	10	10
2.	10	5
3.	10	1
Total	30	16

1. A coin is tossed independently and repeatedly with the probability of heads equal to p .

(a) What is the probability of getting only heads in the first n tosses?

(3 points)

probability of getting heads = p .

$$P(\text{heads in } n \text{ tosses}) = \left(\frac{p}{1-p}\right)^n.$$

Since heads occurred in ' n ' tosses.

(b) What is the probability of obtaining the first tail in the n -th toss?

(3 points)

Sol:- Probability of getting heads = p (success)

Probability of getting tail = $1-p$ (failure)

$$P(\text{first tail in } n^{\text{th}} \text{ toss}) = (p)^{n-1} (1-p).$$

since $(n-1)$ tosses heads occurred and in $(n-1)^{\text{th}}$ toss tail occurred (only 1 time).

(c) What is the expected number of tosses required to obtain the first tail?

(4 points)

$E[X] \Rightarrow$ no. of tosses required to obtain first tail in $(n-1)^{\text{th}}$ toss

We know, $E[X] = \sum_{x=1}^{\infty} x \cdot P(X=x)$

We know $P(X=x) = (p)^{n-1} (1-p)$

$$E[X] = \sum_{n=1}^{\infty} x \cdot (p)^{n-1} (1-p) \quad \text{From eq ① and ②}$$

$$E[X] = (1-p) \sum_{n=1}^{\infty} x \cdot (p)^{n-1} \rightarrow ①$$

We have this inequality,

$$\sum_{n=0}^{\infty} x^n = \frac{1}{1-x} \text{ for } |x| < 1$$

differentiating both sides:-

$$\sum_{n=1}^{\infty} n \cdot x^{n-1} = \frac{-1+1}{(1-x)^2} \quad \begin{array}{l} \text{Here from eq} \\ \text{①} \end{array}$$

$$\Rightarrow \sum_{n=1}^{\infty} n x^{n-1} = \frac{1}{(1-x)^2} \rightarrow ②$$

$$\text{Q.E.D.}$$

$$E[X] = \frac{1-p}{(1-p)^2}$$

$$E[X] = \frac{1}{1-p}$$

So, the expected number of tosses to obtain the first tail is $\frac{1}{1-p}$.

$$= \frac{1}{1-p}$$

failure

2. A fair coin is tossed n times.

(a) Show that the probability of obtaining more than $\frac{1}{\sqrt{2}}$ fraction of heads is at most $\frac{1}{\sqrt{2}}$.

Using Hoeffding's inequality, (5 points)

$$\text{The probability} = \frac{1}{2} \cdot (\text{fair coin}) \leq e^{-\frac{(0.707 - 0.5)^2}{2n}} \cdot (0.273)^2$$

$$P(X \geq nt + p) \leq e^{-\frac{t^2}{2n}} \leq e^{-\frac{(0.1207)^2}{2n}}$$

$$\text{Given: } P\left(\frac{X}{n} \geq \sqrt{\frac{1}{2}}\right) = P(X \geq n\sqrt{\frac{1}{2}}) \leq e^{-0.025628}$$

We know from the eqn,

$$P(X \geq n(\frac{1}{2} - \frac{1}{2}))$$

$$t = \frac{1}{\sqrt{2}} - \frac{1}{2}$$

$$P(X \geq \frac{1}{\sqrt{2}}) \leq \frac{1}{\sqrt{2}}$$

So, the probability of obtaining more than
and is atmost $\frac{1}{6}$. Hence proved.

Suffices to use
Markov

(b) Show that the probability of obtaining more than $\frac{1}{4}$ fraction of heads is at least $\frac{1}{4}$ as $n \rightarrow \infty$.
(5 points)

Hint: Use the Paley-Zygmund inequality: for $X \geq 0$ and $0 < \alpha < 1$,

$$\mathbb{P}[X \geq \alpha \mathbb{E}X] \geq (1 - \alpha)^2 \frac{(\mathbb{E}X)^2}{\mathbb{E}X^2}.$$

$$\text{由 } P(X \geq \frac{1}{y}) = P\left(\frac{X}{n} \geq \frac{1}{y}\right) \Rightarrow P\left(X \geq \frac{n}{y}\right)$$

3. Let X_1, \dots, X_n be independent random variables with

$$\mathbb{P}[X_1 = 1] = p_i = 1 - \mathbb{P}[X = -1]$$

where $p_i \in (0, 1), i = 1, \dots, n$. Show that, for any $t \geq 0$, we have

$$\mathbb{P} \left[\sum_{i=1}^n (X_i - \mathbb{E}X_i) \geq t \right] \leq \exp \left\{ -\frac{t^2}{2n} \right\}.$$

(10 points)

Hint: Use Chernoff bound. You may assume the inequality: for all $\lambda > 0$ and $0 < p < 1$,

$$pe^{2(1-p)\lambda} + (1-p)e^{-2p\lambda} \leq e^{\lambda^2/2}.$$

Chernoff bound:-

$$\mathbb{P}[X \geq (t+\delta)\mathbb{E}[X]] \leq e^{-\frac{\delta^2 \mathbb{E}[X]}{2+\delta}}$$

$X = x_1 + x_2 + \dots + x_n$

$$\Rightarrow \mathbb{P}\left[\sum_{i=1}^n (x_i - \mathbb{E}[x_i]) \geq t\right]$$

$$\Rightarrow \mathbb{P}[(X - \mathbb{E}[X]) \geq t] \quad \text{Multiplying with } \mathbb{E}[X]$$

$$\Rightarrow \mathbb{P}[X \geq t + \mathbb{E}[X]] \Rightarrow \mathbb{P}\left[\frac{X - \mathbb{E}[X]}{\mathbb{E}[X]} \geq \frac{t + \mathbb{E}[X] - \mathbb{E}[X]}{\mathbb{E}[X]}\right]$$

let $t = 1 + \delta$

① then $\mathbb{P}\left[X \geq \frac{t + \delta \mathbb{E}[X]}{\mathbb{E}[X]}\right]$

$$\mathbb{P}\left(X \geq \frac{(t+1)\mathbb{E}[X]}{\mathbb{E}[X]}\right) \leq e^{-\frac{\delta^2 \mathbb{E}[X]}{2+\delta}}$$

$$\leq e^{-\frac{t^2 \mathbb{E}[X]}{2+\delta}}$$

1. A box has r red and b blue balls. A ball is chosen at random from the box (so that each ball is equally likely to be chosen), and then a second ball is drawn at random from the remaining balls in the box. Find the probability that

- (a) Both balls are of the same colour.
- (b) Both balls are of different colour.

Do they add up to 1?

Sol: Total no. of balls = $r+b$

(a) Both balls are of same colour.
i, let us think the colour is red.

$$P(\text{red}) = \frac{r}{r+b} \times \frac{r-1}{r+b-1}$$

ii, The colour is blue

$$P(\text{blue}) = \frac{b}{r+b} \times \frac{b-1}{r+b-1}$$

$$\text{Total probability} = \frac{r(r-1) + b(b-1)}{(r+b)(r+b-1)}$$

Yes they add up to 1. $\frac{r(r-1) + b(b-1) + 2rb}{(r+b)(r+b-1)} = \frac{r^2 + r^2 - r + b^2 - b + 2rb}{(r+b)(r+b-1)} = \frac{r^2 + b^2 + 2rb}{(r+b)(r+b-1)} = 1$

2. Let $X = 1$ if the second ball drawn is red; 0 otherwise. What is $E[X]$?

Sol: $E[X] = \sum_{n=0}^{\infty} n \cdot P(X=n)$

✓ MARKED

$$E[X] = \sum_{n=0}^{\infty} 1 \cdot P(X=1) \quad [\text{second ball drawn is red}]$$

[Because it may be different & equal]

$$E[X] = \frac{rb}{(r+b)(r+b-1)} + \frac{r(r-1)}{(r+b)(r+b-1)}$$

$$= \frac{rb + r(r-1)}{(r+b)(r+b-1)} = \frac{rb + r^2 - r}{(r+b)(r+b-1)}$$

$$= \frac{r(b+r-1)^1}{(r+b)(b+r-1)} = \frac{r}{r+b}$$