

1. A fair six-sided die is rolled independently and repeatedly. Recall that the outcome of a roll of a fair six-sided die is an element of  $\{1, 2, 3, 4, 5, 6\}$ , each occurring with probability  $1/6$ .

- (a) What is the probability that none of the first  $n$  rolls results in a 6? (3 points)

**Solution:** Let  $X_i$  be the outcome of the  $i$ -th roll. Then,

$$\mathbb{P}[X_i \neq 6] = \frac{5}{6}.$$

Since the rolls are independent, the probability that all of the first  $n$  rolls are not 6 is

$$\left(\frac{5}{6}\right)^n.$$

- (b) What is the probability that a 6 appears for the first time on the  $n$ -th roll? (3 points)

**Solution:**  $\left(\frac{5}{6}\right)^{n-1} \frac{1}{6}$

- (c) What is the expected number of rolls required to obtain the first 6? (4 points)

**Solution:** This follows a geometric distribution with success probability

$$p = \frac{1}{6}.$$

Hence, the expected number of rolls is  $\frac{1}{p} = 6$ .

2. A fair six-sided die is rolled  $n$  times.

- (a) Let  $Y$  denote the number of times a 6 appears. Compute  $\mathbb{E}Y^2$ . (2 points)

*Hint:* For  $Z \sim \text{Ber}(p)$ , we have  $\text{Var}(Z) = p(1 - p)$ .

**Solution:** First, we compute  $\text{Var}(Y)$  by considering  $Y$  as a sum of Bernoulli random variables each with bias  $p = 1/6$ . Therefore,

$$\text{Var}(Y) = np(1 - p) = \frac{5n}{36}.$$

Thus,

$$\mathbb{E}Y^2 = \text{Var}(Y) + (\mathbb{E}Y)^2 = \frac{5n + n^2}{36}.$$

- (b) Show that the probability of obtaining more than  $1/\sqrt{6}$  fraction of 6s is at most  $1/\sqrt{6}$ . (3 points)

**Solution:** We have  $\mathbb{E}Y = n/6$ . Then, by Markov's inequality,

$$\begin{aligned} \mathbb{P}\left[Y \geq \frac{n}{\sqrt{6}}\right] &\leq \frac{\mathbb{E}Y}{n/\sqrt{6}} \\ &= \frac{n/6}{n/\sqrt{6}} \\ &= \frac{1}{\sqrt{6}}. \end{aligned}$$

- (c) Show that the probability of obtaining more than  $1/7$  fraction of 6s is at least  $1/49$  as  $n \rightarrow \infty$ .  
(5 points)

*Hint:* Use the Paley-Zygmund inequality: for  $Y \geq 0$  and  $0 < \alpha < 1$ ,

$$\mathbb{P}[Y \geq \alpha \mathbb{E}Y] \geq (1 - \alpha)^2 \frac{(\mathbb{E}Y)^2}{\mathbb{E}Y^2}.$$

**Solution:** We have  $\mathbb{E}Y = n/6$  and  $\mathbb{E}Y^2 = \frac{n^2 + 5n}{36}$ .

$$\begin{aligned} \mathbb{P}\left[Y \geq \frac{n}{7}\right] &= \mathbb{P}\left[Y \geq \frac{6}{7}\mathbb{E}Y\right] \\ &\geq \left(\frac{1}{7}\right)^2 \frac{(n/6)^2}{\frac{n^2}{36} + \frac{5n}{36}} \\ &= \frac{1}{49} \cdot \frac{n}{n+5} \end{aligned}$$

which goes to  $1/49$  as  $n \rightarrow \infty$ .

3. Let  $X_1, \dots, X_n$  be independent random variables with

$$\mathbb{P}[X_i = 1] = p_i = 1 - \mathbb{P}[X_i = -1]$$

where  $p_i \in (0, 1), i = 1, \dots, n$ . Show that, for any  $t \geq 0$ , we have

$$\mathbb{P}\left[\sum_{i=1}^n (X_i - \mathbb{E}X_i) \geq t\right] \leq \exp\left\{-\frac{t^2}{2n}\right\}.$$

(10 points)

*Hint:* Use Chernoff bound. You may assume the inequality: for all  $\lambda > 0$  and  $0 < p < 1$ ,

$$pe^{2(1-p)\lambda} + (1-p)e^{-2p\lambda} \leq e^{\lambda^2/2}.$$

**Solution:** We have

$$\mathbb{E}X_i = p_i - (1 - p_i) = 2p_i - 1$$

for  $i \in [n]$ . The Chernoff bound is given by

$$\mathbb{P}\left[\sum_{i=1}^n (X_i - \mathbb{E}X_i) \geq t\right] \leq e^{-\lambda t} \prod_{i=1}^n M_{(X_i - \mathbb{E}X_i)}(\lambda)$$

where

$$\begin{aligned} M_{(X_i - \mathbb{E}X_i)}(\lambda) &= \mathbb{E}\left[e^{\lambda(X_i - (2p_i - 1))}\right] \\ &= p_i e^{2(1-p_i)\lambda} + (1-p_i)e^{-2p_i\lambda} \\ &\leq e^{\lambda^2/2}. \end{aligned}$$

Therefore,

$$\begin{aligned} \mathbb{P}\left[\sum_{i=1}^n (X_i - \mathbb{E}X_i) \geq t\right] &\leq \min_{\lambda > 0} e^{-\lambda t} e^{n\lambda^2/2} \\ &= \exp\left\{-\frac{t^2}{2n}\right\}. \end{aligned}$$