

DS5616: Foundations of Statistical Learning

Test 1

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- The test comprises 3 questions for a total of 30 points and is for a duration of 50 minutes.
- Write your fair answer *after* you have tried it out in the rough space. **Only work written in the space below the corresponding questions shall be graded.** If you do not wish something to be graded, strike it out neatly.
- This is a closed-book test. You may not use the textbook, class notes, mobile phone, tablet, laptop, or such devices.
- This is also an opportunity for you to evaluate yourself – in your best interests, be honest.

Question	Points	Score
1.	10	10
2.	10	5
3.	10	1
Total	30	16

1. A coin is tossed independently and repeatedly with the probability of heads equal to p .

(a) What is the probability of getting only heads in the first n tosses?

(3 points)

probability of getting heads = p .
 $P(\text{heads in 'n' tosses}) = (p)^n$.

Since heads occurred in 'n' tosses.

(b) What is the probability of obtaining the first tail in the n -th toss?

(3 points)

Sol:- Probability of getting heads = p (success)
 Probability of getting tail = $1-p$ (failure)

(10) $P(\text{first tail in } n^{\text{th}} \text{ toss}) = (p)^{n-1} (1-p)$

Since $(n-1)$ tosses heads occurred and in $(n-1)^{\text{th}}$ toss tail occurred (only 1 time).

(c) What is the expected number of tosses required to obtain the first tail?

(4 points)

$E[X] \Rightarrow$ no. of tosses required to obtain first tail in $(n-1)^{\text{th}}$ toss

We know, $E[X] = \sum_{x=1}^{\infty} x \cdot P(X=x)$

We know $P(X=x) = (p)^{x-1} (1-p)$

$E[X] = \sum_{x=1}^{\infty} x \cdot (p)^{x-1} (1-p)$ From eq ① and ②

$E[X] = (1-p) \sum_{x=1}^{\infty} x \cdot (p)^{x-1}$ ①

We have this inequality,

$$\sum_{n=0}^{\infty} x^n = \frac{1}{1-x} \text{ for } |x| < 1$$

differentiating both sides:-

$$\sum_{n=0}^{\infty} n \cdot x^{n-1} = \frac{-1}{(1-x)^2}$$

Here from ① $x = p$ ②

$$\Rightarrow \sum_{n=0}^{\infty} n x^{n-1} = \frac{1}{(1-x)^2} \rightarrow ②$$

$$E[X] = 1-p \cdot \frac{1}{[1-(p)]^2}$$

$$E[X] = \frac{1-p}{(1-p)^2}$$

$$E[X] = \frac{1}{1-p}$$

So, the expected number of tosses to obtain the first tail is $\frac{1}{1-p}$

= $\frac{1}{\text{failure}}$

2. A fair coin is tossed n times.

(a) Show that the probability of obtaining more than $\frac{1}{\sqrt{2}}$ fraction of heads is at most $\frac{1}{\sqrt{2}}$.

Using Hoeffding's inequality,

The probability = $\frac{1}{2}$ (fair coin)

$$P(X \geq nt + p) \leq e^{-\frac{t^2}{2n}}$$

Given:- $P\left(\frac{X}{n} \geq \sqrt{\frac{1}{2}}\right) = P\left(X \geq n\sqrt{\frac{1}{2}}\right)$

We know from the eqn,

$$P\left(X \geq n\left(\sqrt{\frac{1}{2}} - \frac{1}{2}\right)\right)$$

$$\begin{aligned} t &= \sqrt{\frac{1}{2}} - \frac{1}{2} \\ &\leq e^{-\frac{\left(\sqrt{\frac{1}{2}} - \frac{1}{2}\right)^2}{2n}} \end{aligned}$$

$$\leq e^{-\frac{(0.707 - 0.5)^2}{2n}}$$

$$\leq e^{-\frac{(0.207)^2}{2n}}$$

$$\leq e^{-\frac{0.021428}{n}}$$

$$\leq e^{-\frac{0.02142}{n}} \text{ as } n \rightarrow \text{large}$$

$$\approx \frac{1}{\sqrt{2}}$$

$$P\left(X \geq \frac{1}{\sqrt{2}}\right) \leq \frac{1}{\sqrt{2}}$$

Suffices to use Markov

So, the probability of obtaining more than

$\frac{1}{\sqrt{2}}$ fraction of heads is at most $\frac{1}{\sqrt{2}}$. Hence proved.

(b) Show that the probability of obtaining more than $\frac{1}{4}$ fraction of heads is at least $\frac{1}{4}$ as $n \rightarrow \infty$. (5 points)

Hint: Use the Paley-Zygmund inequality: for $X \geq 0$ and $0 < \alpha < 1$,

$$P[X \geq \alpha EX] \geq (1 - \alpha)^2 \frac{(EX)^2}{EX^2}$$

Sol:- $P\left(X \geq \frac{1}{4}\right) = P\left(\frac{X}{n} \geq \frac{1}{4}\right) \Rightarrow P\left(X \geq \frac{n}{4}\right)$

3. Let X_1, \dots, X_n be independent random variables with

$$\mathbb{P}[X_1 = 1] = p_i = 1 - \mathbb{P}[X = -1]$$

where $p_i \in (0, 1)$, $i = 1, \dots, n$. Show that, for any $t \geq 0$, we have

$$\mathbb{P} \left[\sum_{i=1}^n (X_i - \mathbb{E}X_i) \geq t \right] \leq \exp \left\{ -\frac{t^2}{2n} \right\}.$$

(10 points)

Hint: Use Chernoff bound. You may assume the inequality: for all $\lambda > 0$ and $0 < p < 1$,

$$pe^{2(1-p)\lambda} + (1-p)e^{-2p\lambda} \leq e^{\lambda^2/2}.$$

Chernoff bound:-

$$\mathbb{P}[X \geq (1+\delta)\mathbb{E}[X]] \leq e^{-\frac{\delta^2 \mathbb{E}[X]}{2+\delta}} \quad X = X_1 + X_2 + \dots + X_n$$

$$\Rightarrow \mathbb{P} \left[\sum_{i=1}^n (X_i) - \sum_{i=1}^n \mathbb{E}[X_i] \geq t \right]$$

$$\Rightarrow \mathbb{P}[(X - \mathbb{E}[X]) \geq t] \quad \text{Multiplying with } \mathbb{E}[X]$$

$$\Rightarrow \mathbb{P}[(X \geq t + \mathbb{E}[X]) \Rightarrow \mathbb{P} \left[(X \geq \frac{t\mathbb{E}[X] + \mathbb{E}[X]}{\mathbb{E}[X]}) \right]$$

let $t = 1 + \delta$

① then $\mathbb{P} \left[(X \geq \frac{t\mathbb{E}[X]}{\mathbb{E}[X]}) \right] \leq e^{-\frac{\delta^2 \mathbb{E}[X]}{2+\delta}}$

$$\mathbb{P} \left((X \geq \frac{(t+1)\mathbb{E}[X]}{\mathbb{E}[X]}) \right) \leq e^{-\frac{\delta^2 \mathbb{E}[X]}{2+\delta}}$$

$$\leq e^{-\frac{t^2 \mathbb{E}[X]}{2+\delta}}$$

1. A box has r red and b blue balls. A ball is chosen at random from the box (so that each ball is equally likely to be chosen), and then a second ball is drawn at random from the remaining balls in the box. Find the probability that

- (a) Both balls are of the same colour.
(b) Both balls are of different colour.

Do they add up to 1?

(6 points)

sol:- Total no. of balls = $r+b$

(a) Both balls are of same colour.

i, let us think the colour is red.

$$P(\text{red}) = \frac{r}{r+b} \times \frac{r-1}{r+b-1}$$

ii, The colour is blue

$$P(\text{blue}) = \frac{b}{r+b} \times \frac{b-1}{r+b-1}$$

$$\text{Total probability} = \frac{r(r-1) + b(b-1)}{(r+b)(r+b-1)}$$

(b) Both balls are different color
i, ~~red~~ and then next ball is blue

$$P(\text{1st ball}) =$$

$$P(\text{red, blue}) = \frac{r}{r+b} \times \frac{b}{r+b-1}$$

ii, blue and then red

$$P(\text{blue, red}) = \frac{b}{r+b} \times \frac{r}{r+b-1}$$

$$\text{Total probability} = \frac{2rb}{(r+b)(r+b-1)}$$

Yes they add up to 1.
$$\frac{r(r-1) + b(b-1) + 2rb}{(r+b)(r+b-1)} = \frac{r^2 - r + b^2 - b + 2rb}{(r+b)(r+b-1)} = \frac{r^2 + b^2 + 2rb - r - b}{(r+b)(r+b-1)} = \frac{(r+b)^2 - (r+b)}{(r+b)(r+b-1)} = 1$$

2. Let $X = 1$ if the second ball drawn is red; 0 otherwise. What is $E[X]$?

sol:-
$$E[X] = \sum_{n=0}^{\infty} n \cdot P(X=n)$$

$$E[X] = \sum_{i=0}^{\infty} 1 \cdot P(X=1) \text{ (second ball drawn is red)}$$

[Because it may be different @ equal]

$$E[X] = \frac{rb}{(r+b)(r+b-1)} + \frac{r(r-1)}{(r+b)(r+b-1)}$$

$$= \frac{rb + r(r-1)}{(r+b)(r+b-1)} = \frac{rb + r^2 - r}{(r+b)(r+b-1)}$$

$$= \frac{r(b+r-1)}{(r+b)(r+b-1)} = \frac{r}{r+b}$$