

第3-2章 点估计 EM算法

《统计推断》 第7章

感谢清华大学自动化系江瑞教授提供PPT

内容

- Motivation
- EM算法理论
- EM算法应用

Bernoulli MLE

- 在抛硬币的Bernoulli试验中，前面我们计算了其MLE就是用频率来估计。
- 如果在Bernoulli试验中有两个硬币，但不知道每次试验抛出的是哪一个，存在缺失数据。
- 这是一个典型的混合模型

$$\pi \text{Ber}(p_1) + (1 - \pi) \text{Ber}(p_2)$$

Basic Setting in EM

- X is a set of data points: **observed** data
- Θ is a parameter vector.
- EM is a method to find θ_{ML} where

$$\begin{aligned}\theta_{ML} &= \arg \max_{\theta \in \Omega} l(\Theta) \\ &= \arg \max_{\theta \in \Omega} \log P(X | \Theta)\end{aligned}$$

- Calculating $P(X | \theta)$ directly is hard.
- Calculating $P(X, Z | \theta)$ is much simpler, where Z is “hidden” data (or “missing” data).

The Basic Setting in EM

- $Y = (X, Z)$
 - Y : complete data (“augmented data”)
 - X : observed data (“incomplete” data)
 - Z : hidden data (“missing” data)
- Given a fixed x , there could be many possible z ’s.
 - Ex: given a sentence x , there could be many state sequences in an HMM that generates x .

The Iterative Approach for MLE

- When missing data is available, it's hard to find the MLE directly

$$\theta_{ML} = \underset{\theta}{\operatorname{Argmax}} \log \left(\sum_Z P(X, Z | \theta) \right)$$

- An alternative is to find a sequence

$$\theta^{(0)}, \theta^{(1)}, \dots, \theta^{(t)}, \dots,$$

$$\text{s.t. } l(\theta^{(0)}) < l(\theta^{(1)}) < \dots < l(\theta^{(t)}) < \dots$$

$$\begin{aligned}
l(\theta) - l(\theta^{(t)}) &= \log P(X|\theta) - \log P(X|\theta^{(t)}) \\
&= \log \left(\frac{\sum_Z P(X, Z|\theta)}{\sum_Z P(X, Z|\theta^{(t)})} \right) \\
&= \log \left(\sum_Z \frac{P(X, Z|\theta)}{\sum_{Z'} P(X, Z'|\theta^{(t)})} \right) \\
&= \log \left(\sum_Z \frac{P(X, Z|\theta)}{\sum_{Z'} P(X, Z'|\theta^{(t)})} \times \frac{P(X, Z|\theta^{(t)})}{P(X, Z|\theta^{(t)})} \right) \\
&= \log \left(\sum_Z \frac{P(X, Z|\theta^{(t)})}{\sum_{Z'} P(X, Z'|\theta^{(t)})} \times \frac{P(X, Z|\theta)}{P(X, Z|\theta^{(t)})} \right)
\end{aligned}$$

$$\begin{aligned}
l(\theta) - l(\theta^{(t)}) &= \log \left(\sum_Z \frac{P(X, Z|\theta^{(t)})}{\sum_{Z'} P(X, Z'|\theta^{(t)})} \times \frac{P(X, Z|\theta)}{P(X, Z|\theta^{(t)})} \right) \\
&= \log \left(\sum_Z P(Z|X, \theta^{(t)}) \times \frac{P(X, Z|\theta)}{P(X, Z|\theta^{(t)})} \right) \\
&\geq \sum_Z P(Z|X, \theta^{(t)}) \times \log \left(\frac{P(X, Z|\theta)}{P(X, Z|\theta^{(t)})} \right) \\
&= E_{P(Z|X, \theta^{(t)})} \left[\log \left(\frac{P(X, Z|\theta)}{P(X, Z|\theta^{(t)})} \right) \right] \\
&= E_{P(Z|X, \theta^{(t)})} [\log P(X, Z|\theta)] \\
&\quad - E_{P(Z|X, \theta^{(t)})} [\log P(X, Z|\theta^{(t)})]
\end{aligned}$$

Jensen's inequality

Maximizing the Lower Bound

- The Jensen's inequality gives a lower bound to maximize,

$$\theta^{(t+1)} = \underset{\theta}{\operatorname{Argmax}} E_{P(Z|X, \theta^{(t)})} [\log P(X, Z|\theta)]$$

- Q-function

$$Q(\theta|\theta^{(t)}) = E_{P(Z|X, \theta^{(t)})} [\log P(X, Z|\theta)]$$

Increasing the Likelihood

- Increasing the likelihood by maximizing the lower bound

$$l(\theta) - l(\theta^{(t)}) \geq Q(\theta|\theta^{(t)}) - Q(\theta^{(t)}|\theta^{(t)})$$

$$Q(\theta^{(t+1)}|\theta^{(t)}) > Q(\theta^{(t)}|\theta^{(t)}) \Rightarrow l(\theta^{(t+1)}) > l(\theta^{(t)})$$

- Which means that a better estimation of the parameter.

Summary: EM Algorithm

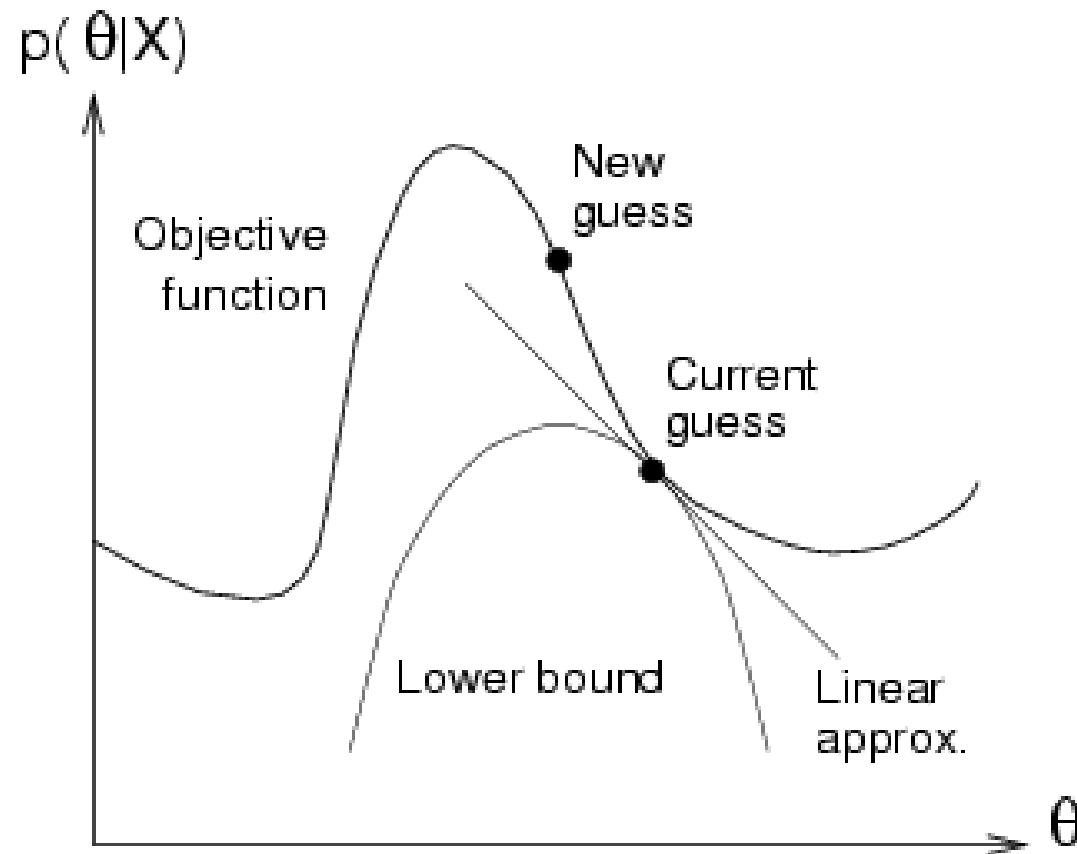
- Define a auxiliary function

$$\begin{aligned} Q(\theta|\theta') &= \sum_Z P(Z|X, \theta') \log P(X, Z|\theta) \\ &= E_{P(Z|X, \theta')} [\log P(X, Z|\theta)] \end{aligned}$$

- EM algorithm iterates with two step
 - E-Step, compute $Q(\theta|\theta^{(t)})$
 - M-Step:

$$\theta^{(t+1)} = \underset{\theta}{\text{Argmax}} Q(\theta|\theta^{(t)})$$

Illustration of EM Algorithm



Jensen's Inequality

- Convex function

$$\forall x_1, x_2 \in (a, b), \lambda \in [0, 1]$$

$$f(\lambda x_1 + (1 - \lambda)x_2) \leq \lambda f(x_1) + (1 - \lambda)f(x_2)$$



Jensen's Inequality

- For convex function $f(x)$

$$E[f(X)] \geq f(E[X])$$

- For discrete random variable with two mass points

$$E[X] = p_1x_1 + p_2x_2$$

$$E[f(X)] = p_1f(x_1) + p_2f(x_2)$$

$$\geq f(p_1x_1 + p_2x_2) = f(E[x])$$

- It's easy to induce to random variable with more points

Jensen's Inequality Corollary

- $\log(x)$ is a concave function, for any positive function $g(x)$

$$\log(E[g]) \geq E[\log(g)]$$

$$\log\left(\sum_j q_j g(j)\right) \geq \sum_j q_j \log(g(j))$$

where

$$q_j \in [0, 1], \quad \sum_j q_j = 1$$

Example

- Rao (1965, pp.368-369), *Genetic Linkage Model*
- Suppose 197 animals are distributed multinomially into four categories,

$$X = (125, 18, 20, 34) = (x_1, x_2, x_3, x_4)$$

- A genetic model for the population specifies cell probabilities

$$\left(\frac{1}{2} + \frac{\theta}{4}, \frac{1}{4} - \frac{\theta}{4}, \frac{1}{4} - \frac{\theta}{4}, \frac{\theta}{4} \right)$$

Multinomial Distribution

- Likelihood function

$$L(\theta) = \frac{197!}{x_1!x_2!x_3!x_4!} \left(\frac{1}{2} + \frac{\theta}{4}\right)^{x_1} \left(\frac{1}{4} - \frac{\theta}{4}\right)^{x_2+x_3} \left(\frac{\theta}{4}\right)^{x_4}$$

- log-likelihood function

$$l(\theta) = \log \frac{197!}{x_1!x_2!x_3!x_4!} \\ + x_1 \log\left(\frac{1}{2} + \frac{\theta}{4}\right) + (x_2 + x_3) \log\left(\frac{1}{4} - \frac{\theta}{4}\right) + x_4 \log\left(\frac{\theta}{4}\right)$$

MLE

- Take derivative, solve equation

$$\frac{\partial l(\theta)}{\partial \theta} = \frac{1}{4} \times \frac{x_1}{\frac{1}{2} + \frac{\theta}{4}} - \frac{1}{4} \times \frac{x_2 + x_3}{\frac{1}{4} - \frac{\theta}{4}} + \frac{1}{4} \times \frac{x_4}{\frac{\theta}{4}} = 0$$

- It's not easy to solve this equation!

$$\frac{x_1}{2 + \theta} - \frac{x_2 + x_3}{1 - \theta} + \frac{x_4}{\theta} = 0$$

Missing Data Problem

- Split the first category into two group

$$x_1 = z_1 + z_2, \quad z_1, z_2 \text{ missing}$$

With Probability

$$p(z_1) = \frac{1}{2}, p(z_2) = \frac{\theta}{4}$$

- Log-likelihood function of complete data

$$l(\theta) = \log \frac{197!}{z_1!z_2!x_2!x_3!x_4!} \\ + z_1 \log\left(\frac{1}{2}\right) + (z_2 + x_4) \log\left(\frac{\theta}{4}\right) + (x_2 + x_3) \log\left(\frac{1}{4} - \frac{\theta}{4}\right)$$

E Step: Multinomial

$$E \left(\log f(x, \theta) | \theta^{(k)} \right) = E \left(\log \frac{197!}{z_1! z_2! x_2! x_3! x_4!} \right) \\ + z_1^{(k)} \log\left(\frac{1}{2}\right) + (z_2^{(k)} + x_4) \log\left(\frac{\theta}{4}\right) + (x_2 + x_3) \log\left(\frac{1}{4} - \frac{\theta}{4}\right)$$

- Where

$$\begin{cases} E(z_1) = 125 \frac{\frac{1}{2}}{\frac{1}{2} + \frac{\theta^{(k)}}{4}} = z_1^{(k)} \\ E(z_2) = 125 \frac{\frac{\theta^{(k)}}{4}}{\frac{1}{2} + \frac{\theta^{(k)}}{4}} = z_2^{(k)} \end{cases}$$

M Step: Multinomial

- Take derivative

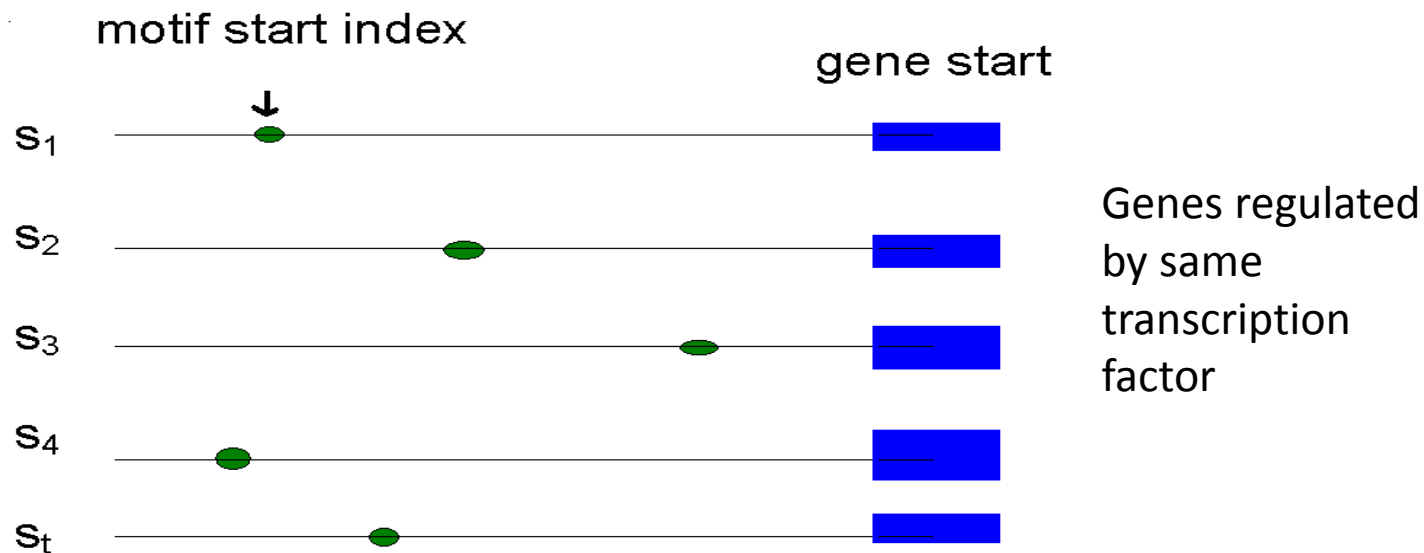
$$E \left(\log f(x, \theta) | \theta^{(k)} \right) = E \left(\log \frac{197!}{z_1! z_2! x_2! x_3! x_4!} \right) \\ + z_1^{(k)} \log\left(\frac{1}{2}\right) + (z_2^{(k)} + x_4) \log\left(\frac{\theta}{4}\right) + (x_2 + x_3) \log\left(\frac{1}{4} - \frac{\theta}{4}\right)$$

- One can obtain

$$\theta^{(k+1)} = \frac{z_2^{(k)} + x_4}{z_2^{(k)} + x_4 + x_2 + x_3} = \frac{z_2^{(k)} + 34}{z_2^{(k)} + 18 + 20 + 34}$$

Motif Finding Problem

- Find promoter motifs associated with **co-regulated** or **functionally related** genes



Probabilistic Model

- Positional weighted matrix (PWM) for motif
 - $L \times 4$ matrix, where L is the length of the motif
 - Each position is a probability distribution ($p(A)$, $p(C)$, $p(G)$, $p(T)$)
 - Independence between different position
- Random background with distribution ($p_0(A)$, $p_0(C)$, $p_0(G)$, $p_0(T)$)

PWM

HEM13 CCCATT

HEM13 TTTCTG

HEM13 TCAATT

ANB1 CTCATT

ANB1 TCCATT

ANB1 CCTATT

ANB1 TCCATT

ROX1 CCAATT

	1	2	3	4	5	6
A	0	0	0.25	0.875	0	0
C	0.5	0.75	0.5	0.125	0	0.125
G	0	0	0	0	0	0.875
T	0.5	0.25	0.25	0	1.0	0

Motif Finding

- Given the missing data, it's a multinomial distribution

$$Pr(X_i, Z_{ij} = 1 | P) = \underbrace{\prod_{k=1}^{j-1} p_{x_{ik},0}}_{\text{before motif}} \underbrace{\prod_{k=j}^{j+w-1} p_{x_{ik},k-j+1}}_{\text{motif}} \underbrace{\prod_{k=j+w}^L p_{x_{ik},0}}_{\text{after motif}}$$

X_i is the i th sequence

Z_{ij} is 1 if motif starts at position j in sequence i

Log-likelihood

$$l(p) = \sum_{k=1}^{j-1} \log p_{x_{ik},0} + \sum_{k=j}^{j+W-1} \log p_{x_{ik},k-j+1} + \sum_{k=j+W}^L \log p_{x_{ik},0}$$

- Q function

$$\begin{aligned} Q(p|p^{(t)}) &= E_{P(Z|X,p^{(t)})} [\log P(X, Z|p)] \\ &= \sum_Z P(Z|X, p^{(t)}) \log P(X, Z|p) \end{aligned}$$

Q-function

$$\begin{aligned} Q(p|p^{(t)}) &= \sum_Z P(Z|X, p^{(t)}) \log P(X, Z|p) \\ &= \sum_Z P(Z|X, p^{(t)}) \sum_{k=1}^{j-1} \log p_{x_{ik},0} \\ &\quad + \sum_Z P(Z|X, p^{(t)}) \sum_{k=j}^{j+W-1} \log p_{x_{ik},k-j+1} \\ &\quad + \sum_Z P(Z|X, p^{(t)}) \sum_{k=j+W}^L \log p_{x_{ik},0} \end{aligned}$$

Q-function

- For each sequence i , the missing value Z_{ij} can take value

$$Z_{i1} = 1, Z_{i2} = 1, \dots, Z_{i,L-W+1} = 1$$

- So the coefficient of $\log P_{c,k}$ is

$$\sum_i \sum_{m=1}^{L-W+1} P(Z_{im} = 1 | X_i, p^t) \delta(X_{i,m+k}, c)$$

Q-function

- The coefficient of $\log P_{c,0}$ is

$$\sum_i \sum_{m=1}^{L-W+1} P(Z_{im} = 1 | X_i, p^t) \left(\sum_{k=1}^{m-1} \delta(X_{i,k}, c) + \sum_{k=m+W}^L \delta(X_{i,k}, c) \right)$$

M Step: Optimization

- For multinomial distribution, the optimization is of form

$$\begin{cases} \text{Max: } \sum_k c_k \log x_k \\ \text{subject to: } \sum_k x_k = 1 \end{cases}$$

$$\text{Estimation: } x_i = \frac{c_i}{\sum_k c_k}, i = 1, \dots, N.$$

M Step: Optimization

- So the estimation of $p_{c,k}$ is

$$\frac{\sum_i \sum_{m=1}^{L-W+1} P(Z_{im} = 1|X_i, p^t) \delta(X_{i,m+k}, c)}{\sum_b \sum_i \sum_{m=1}^{L-W+1} P(Z_{im} = 1|X_i, p^t) \delta(X_{i,m+k}, b)}$$

- So the estimation of $p_{c,0}$ is

$$\frac{\sum_i \sum_{m=1}^{L-W+1} P(Z_{im} = 1|X_i, p^t) \left(\sum_{k=1}^{m-1} \delta(X_{i,k}, c) + \sum_{k=m+W}^L \delta(X_{i,k}, c) \right)}{\sum_b \sum_i \sum_{m=1}^{L-W+1} P(Z_{im} = 1|X_i, p^t) \left(\sum_{k=1}^{m-1} \delta(X_{i,k}, b) + \sum_{k=m+W}^L \delta(X_{i,k}, b) \right)}$$

Example

- Finding motif (length 3) in following sequences

A C A G C A

A G G C A G

T C A G T C

EM Updating

- Let

$$z_{ij}(c) = Pr(Z_{ij} = 1 | X_i, p^{(t)}) \delta(x_{i,m+k}, c)$$

1	2	3	1	2	3	1	2	3
z11(A)	z11(C)	z11(A)	z21(A)	z21(G)	z21(G)	z31(T)	z31(C)	z31(A)
z12(C)	z12(A)	z12(G)	z22(G)	z22(G)	z22(C)	z32(C)	z32(A)	z32(G)
z13(A)	z13(G)	z13(C)	z23(G)	z23(C)	z23(A)	z33(A)	z33(G)	z33(T)
z14(G)	z14(C)	z14(A)	z24(C)	z24(A)	z24(G)	z34(G)	z34(T)	z34(C)

EM Updating

$$p_{A,1} = \frac{z_{11} + z_{13} + z_{21} + z_{33}}{z_{11} + z_{12} + z_{13} + z_{14} + \cdots + z_{31} + z_{32} + z_{33} + z_{34}}$$

$$p_{C,1} = \frac{z_{12} + z_{24} + z_{32}}{z_{11} + z_{12} + z_{13} + z_{14} + \cdots + z_{31} + z_{32} + z_{33} + z_{34}}$$

$$p_{G,1} = \frac{z_{14} + z_{22} + z_{23} + z_{32}}{z_{11} + z_{12} + z_{13} + z_{14} + \cdots + z_{31} + z_{32} + z_{33} + z_{34}}$$

$$p_{T,1} = \frac{z_{31}}{z_{11} + z_{12} + z_{13} + z_{14} + \cdots + z_{31} + z_{32} + z_{33} + z_{34}}$$

Background

- z11: A,C,G
- z12: 2A,C
- z13:2A,C
- z14: 2A, C
- z21:A,C,G
- z22:2A,G
- z23:A,2G
- z24:A,2G
- z31:C,G,T
- z32:C,2T
- z33:2C,T
- z34:A,C,G

Background Updating

- **A** $z_{11} + 2z_{12} + 2z_{13} + 2z_{14} + z_{21} + 2z_{22} + z_{23} + z_{24} + z_{34}$
- **C** $z_{11} + z_{12} + z_{13} + z_{14} + z_{21} + z_{31} + z_{32} + 2z_{33} + z_{34}$
- **G** $z_{11} + z_{21} + z_{22} + 2z_{23} + 2z_{24} + z_{31} + z_{34}$
- **T** $z_{31} + 2z_{32} + z_{33}$

- Normalization factor

$$3(z_{11} + z_{12} + z_{13} + z_{14} + z_{21} + z_{22} + z_{23} + z_{24} + z_{31} + z_{32} + z_{33} + z_{34})$$

Reference

- Dempster, A.P., Laird, N.M., Rubin, D.B. (1977). Maximum Likelihood from Incomplete Data via the EM Algorithm. *Journal of the Royal Statistical Society. Series B (Methodological)*, Vol. 39, No. 1, , pp. 1-38