布朗运动情形的破产模型

第9讲 2019.10.31

主要内容

市朗运动风险过程

市朗运动情形的破产概率

▶ 布朗运动与复合泊松盈余过程的近似

一、布朗运动风险过程

- ▶ 布朗运动的定义(三个特征)
 - > 初值
 - 独立增量、平稳
 - > 正态性:均值和方差为时间的线性函数
- > 标准布朗运动
 - > 均值为零
 - > 方差为时间的线性函数
- 市朗运动的性质
 - > 轨道连续
 - 处处不可微

布朗运动为盈余过程的极限

▶ 长期复合Poisson风险过程的跳跃次数增多(泊松参数增大)、跳跃的幅度减小(每次损失金额变量的期望减小)的极限情形为布朗运动。

$$Z(t) = U(t) - u = ct - S(t)$$
为漂移项大于零的BM

考虑如下的参数关系

$$\mu = \frac{E[Z(t)]}{t} = c - \lambda E[X] = \theta \lambda E[X]$$

$$\sigma^{2} = \frac{Var[Z(t)]}{t} = \lambda E[X^{2}], \quad \lambda = \frac{\sigma^{2}}{E[X^{2}]}$$

定义极限参数

- 通过一个公共的参数将跳跃频率和跳跃程度的极限过程综合表现
- ▶ 并保证对应的布朗运动有相同的均值和方差
- 定义如下的比例变换
- > 定义一个不变元Y

$$X = \alpha Y$$

$$\lambda = \frac{\sigma^2}{E[Y^2]} \frac{1}{\alpha^2}$$

极限

$$\lambda = \frac{\sigma^2}{E[Y^2]} \frac{1}{\alpha^2}$$

$$\alpha \rightarrow 0 \Rightarrow$$

$$\lambda \to \infty$$
, $E[X] \to 0$, $\lambda E[X] \to 0$

$$\lim_{\alpha \to 0} M_{Z(t)}(r) = \exp(r\mu t + \frac{r^2}{2}\sigma^2 t)$$

 $N(\mu t, \sigma^2 t)$ 的矩母函数

实际的解释:

业务量加大时,索赔频率加快,单位时间内的跳跃次数增加

相对于初始盈余,每次索赔的量充分的小

极限的证明

$$\frac{\log M_{z(t)}(r)}{t} = rc + \lambda [M_X(-r) - 1]$$

$$= r[\mu + \lambda E(X)] + \lambda [1 - rE(X) + \frac{r^2}{2!} E(X^2) + o(\alpha^2) - 1]$$

$$= \mu r + \frac{r^2}{2!} \lambda E(X^2) + \lambda \cdot o(\alpha^2)$$

$$= \mu r + \frac{\sigma^2}{2} r^2 + \lambda \cdot o(\alpha^2)$$

二、布朗运动的破产概率

- ▶考虑如下的盈余过程 U(t) = u + W(t) u > 0
- 》净累积收益模型 W(t) 是初值为零的布朗运动。 $\mu>0$

$$\psi(u,\tau) = \Pr(T < \tau) = \Pr\{\min_{0 < t < \tau} U(t) < 0\}$$

$$= \Pr\{\min_{0 < t < \tau} W(t) < -U(0) = -u\}$$

定理2-10

▶给定时间,上述盈余过程的有限时间的破产概率 可表示为:

$$\psi(u,\tau) = \Phi(-\frac{u+\mu\tau}{\sqrt{\sigma^2\tau}}) + \exp(-\frac{2\mu}{\sigma^2}u)\Phi(-\frac{u-\mu\tau}{\sqrt{\sigma^2\tau}})$$

$$= \Pr(U(\tau) \le 0) + \exp(-\frac{2\mu}{\sigma^2}u) \Pr(U(\tau) \le \mu\tau)$$

$$\psi(u,\tau) = \Pr(T \le \tau)$$

$$= \Pr(T \le \tau, U(\tau) \le 0) + \Pr(T \le \tau, U(\tau) > 0)$$

$$= \Pr(U(\tau) \le 0) + \Pr(T < \tau, U(\tau) > 0)$$

$$U(\tau) \sim N(u + \mu \tau, \sigma^2 \tau)$$

$$\Pr(U(\tau) \le 0 \mid U(0) = u) = \Phi(-\frac{u + \mu\tau}{\sqrt{\sigma^2\tau}})$$

证明

$$\bigcup (\tau -) = \{U(t), 0 < t < \tau; U(\tau) < 0\}
\{T \le \tau\} = \{T < \tau, U(\tau) < 0\} \cup \{T < \tau, U(\tau) \ge 0\}
= \bigcup (\tau -) \cup \{T < \tau, U(\tau) \ge 0\}$$

$$\Pr(T \le \tau) = \int_{-\infty}^{0} \Pr(U(\tau) = x) \left[1 + \frac{\Pr(B_x)}{\Pr(A_x)}\right] dx$$

$$\frac{\Pr(B_{x})}{\Pr(A_{x})} = \frac{\int_{0}^{\tau} \Pr(B_{x} \mid T = t) \Pr(T = t) dt}{\int_{0}^{\tau} \Pr(A_{x} \mid T = t) \Pr(T = t) dt} = \frac{\int_{0}^{\tau} \Pr(U(\tau) = -x \mid T = t) \Pr(T = t) dt}{\int_{0}^{\tau} \Pr(U(\tau) = x \mid T = t) \Pr(T = t) dt}$$

$$\forall x, 0 < T \le \tau$$

$$\Pr\{U(\tau) = x \mid T = t\}$$
= \Pr\{U(\tau) - U(t) = x \cong T = t\}
= \Pr\{U(\tau - t) = x \cong T = t\}

$$U(\tau - t) \mid T = t \sim N((\tau - t)\mu, (\tau - t)\sigma^{2})$$

$$\frac{\Pr(B_x)}{\Pr(A_x)} = \frac{\int\limits_0^\tau \Pr(U(\tau) = -x \mid T = t) \Pr(T = t) dt}{\int\limits_0^\tau \Pr(U(\tau) = x \mid T = t) \Pr(T = t) dt}$$

$$= \frac{\int\limits_0^\tau \frac{1}{\sqrt{2\pi(\tau - t)\sigma^2}} \exp\left(-\frac{\left(-x - (\tau - t)\mu\right)^2}{2(\tau - t)\sigma^2}\right) \Pr(T = t) dt}{\int\limits_0^\tau \frac{1}{\sqrt{2\pi(\tau - t)\sigma^2}} \exp\left(-\frac{\left(x - (\tau - t)\mu\right)^2}{2(\tau - t)\sigma^2}\right) \Pr(T = t) dt}$$

$$= \exp\left(-\frac{\left(2x\mu\right)}{\sigma^2}\right) \int\limits_0^\tau \frac{1}{\sqrt{2\pi(\tau - t)\sigma^2}} \exp\left(-\frac{\left(x - (\tau - t)\mu\right)^2}{2(\tau - t)\sigma^2}\right) \Pr(T = t) dt}{\int\limits_0^\tau \frac{1}{\sqrt{2\pi(\tau - t)\sigma^2}} \exp\left(-\frac{\left(x - (\tau - t)\mu\right)^2}{2(\tau - t)\sigma^2}\right) \Pr(T = t) dt}$$

$$\Pr(T \le \tau) = \int_{-\infty}^{0} \Pr(U(\tau) = x) [1 + \frac{\Pr(B_x)}{\Pr(A_x)}] dx$$

$$= \Phi(-\frac{u + \mu\tau}{\sqrt{\sigma^2\tau}}) + \int_{-\infty}^{0} \frac{1}{\sqrt{2\pi\tau\sigma^2}} \exp\left(-\frac{\left(x - (u + \mu\tau)\right)^2}{2\tau\sigma^2}\right) \exp\left(-\frac{2x\mu}{\sigma^2}\right) dx$$

$$= \Phi(-\frac{u + \mu\tau}{\sqrt{\sigma^2\tau}}) + \exp\left(-\frac{2\mu}{\sigma^2}u\right) \int_{-\infty}^{0} \frac{1}{\sqrt{2\pi\tau\sigma^2}} \exp\left(-\frac{\left(x - (u - \mu\tau)\right)^2}{2\tau\sigma^2}\right) dx$$

$$= \Phi(-\frac{u + \mu\tau}{\sqrt{\tau\sigma^2}}) + \exp\left(-\frac{2\mu}{\sigma^2}u\right) \Phi(-\frac{u - \mu\tau}{\sqrt{\tau\sigma^2}})$$

推论2-4 最终破产概率

$$\psi(u) = \exp(-\frac{2\mu}{\sigma^2}u)$$

$$\psi(u) = \exp(-Ru)$$

推论2-5 (有限)破产时刻

$$T^* = T \mid T < \infty$$

$$\Pr(T^* \le \tau) = \frac{\Pr(T \le \tau)}{\Pr(T < \infty)} \qquad \tau > 0$$

三、布朗运动与复合泊松盈余过程的近似

用复合Poisson模型的参数近似表示的有限时间破产概率

$$\psi(u,\tau) \approx \Phi(-\frac{u + \theta \lambda E[X]\tau}{\sqrt{\lambda E[X^2]\tau}}) + \exp(-\frac{2\theta E[X]}{E[X^2]}u)\Phi(-\frac{u - \theta \lambda E[X]\tau}{\sqrt{\lambda E[X^2]\tau}})$$

$$\psi(u) \approx \exp(-\theta \frac{2E[X]}{E[X^2]}u)$$

进而有, (有限) 破产时刻的密度函数:

$$f_{T^*}(\tau) \approx \frac{u}{\sqrt{2\pi\lambda E[X^2]}} \tau^{-\frac{2}{3}} \exp\left(-\frac{(u - \theta\tau\lambda E[X])^2}{2\tau\lambda E[X^2]}\right)$$

$$E(T^*) \approx \frac{u}{\mu} = \frac{u}{\theta \lambda E(X)}$$

思考

- ▶若盈余过程的随机项为Poisson与BM的独立和,破产问题有什么结论?
- ▶ Poisson跳过程与BM的连续过程在破产问题上的本质 差异?
- 用本节的方法讨论对应的期权定价问题。
- ▶ 信用风险模型