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Associations between diabetes medication use and risk of second breast cancer events and mortality

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Abstract

Purpose Diabetes and certain diabetes medications have been shown to influence breast cancer (BC) risk. Less is known about their relation to BC outcomes. Our objective was to evaluate the effects of diabetes and diabetes medications on risk of second breast cancer events (SBCE) and mortality.

Methods This population-based cohort study was conducted among women diagnosed with early-stage (I–II) BC and enrolled in an integrated health plan. Exposures of interest were diabetes and medication classes including insulin, metformin, and sulfonylureas. Outcomes of interest were SBCE defined as recurrence or second primary BC,

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BC-specific mortality, and all-cause mortality. We used multivariable Cox proportional hazards models to estimate hazard ratios (HR) and 95 % confidence intervals (CI) for diabetes and medication use while accounting for potential confounders and competing risks.

Results Among 4,216 women, 13 % developed SBCE during a median follow-up of 6.3 years. 610 women had diabetes of which 76 % used oral diabetes medication and/or insulin. Findings suggested that diabetes increased the risk of recurrence (HR = 1.57; 95 % CI 1.09–2.25) but not overall SBCE (HR = 1.29; 95 % CI 0.94–1.76) or second primary BC (HR = 0.74; 95 % CI 0.39–1.41). Among women with diabetes, insulin use was associated with increased risks of recurrence (HR = 1.94; 95 % CI 1.08–3.48) and all-cause mortality (HR = 2.33; 95 % CI 1.70–3.20). Metformin use was associated with lower all-cause mortality (HR = 0.55; 95 % CI 0.38–0.79).

Conclusions Our findings show an association between diabetes and increased recurrence risk, and risk may be greater among insulin users. Metformin may reduce all-cause mortality among BC survivors. Given the growing breast cancer survivor population, further research in larger, more diverse populations is warranted.

Keywords Breast cancer recurrence · Mortality · Diabetes · Insulin · Metformin · Sulfonylureas

Introduction

There are an estimated 2.8 million breast cancer (BC) survivors in the USA [1]. As this growing population ages, more women are burdened with chronic comorbid conditions such as diabetes and require treatment with multiple medications that could have further implications for health outcomes following BC [2, 3].



Diabetes is associated with an increased risk of incident BC [4] possibly through the direct effect of high levels of insulin on breast tissue and/or indirect effects on increases in sex steroids due to the inhibition of sex hormone-binding globulin, disruption of adipokines, and increased insulin-like growth factor (IGF)-I production [5]. Diabetes may also be a risk factor for second breast cancer events (SBCE) including recurrence [6] and second primary BC [7]. Also, metaanalyses have described a 49 % increased risk of all-cause mortality in BC survivors with diabetes [8]. Some commonly used diabetes medications, such as metformin, may influence the relation between diabetes and BC by decreasing primary BC risk [9], while other diabetes medications, such as insulin glargine, are found to be associated with an increased risk of incident BC in some [10-12] but not all epidemiologic studies [13–15]. Less is known about the relation between diabetes and diabetes medications with risk of SBCE and survival among women with BC.

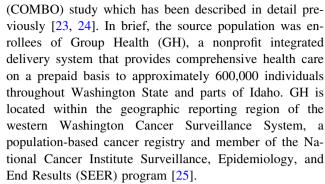
There is growing evidence to support a possible role for diabetes medications to influence BC outcomes. Metformin, for instance, decreases circulating insulin levels and activates the adenosine monophosphate-activated protein kinase (AMPK) pathway [16]. These effects are relevant to risk of SBCE in that insulin can promote cancer growth through modulators such as insulin-like growth factors, and AMPK activation can suppress several tumorigenic metabolic pathways [16]. As such, metformin is associated with increased rates of pathologic complete response after chemotherapy [17] and reduced tumor proliferation markers in BC patients without diabetes [18]. On the other hand, it is possible that medications such as sulfonylureas that stimulate insulin secretion from betacells or insulin analogues have increased mitogenicity and carcinogenic potential to influence risk of SBCE [19]. Insulin glargine, in particular, is a long-acting structurally altered human insulin analogue that has altered binding kinetics for insulin receptors [20] and increased affinity for IGF-I receptors [21]. Recent data suggest that among patients with diabetes, serum of those treated with insulin glargine is more mitogenic to BC cells compared to those treated with regular human insulin or insulin detemir [22].

The objective of this study was to evaluate the relation between diabetes and diabetes medications including insulin, metformin, and sulfonylureas with risks of SBCE, BC-specific mortality, and all-cause mortality.

Materials and methods

Study population and data collection

We conducted a retrospective cohort study within the Commonly Used Medications and Breast Cancer Outcomes



The COMBO cohort included women aged ≥18 years, residing within the 13 Washington State counties covered by SEER, diagnosed with an incident, histologically confirmed stage I or II asynchronous (non-bilateral) BC between 1 January 1990 and 31 December 2008, and enrolled in GH's integrated group practice for 1+ year before and after the incident BC diagnosis unless they died [23, 24]. Chart abstraction began in 2009 and continued through August 2011.

Data were collected through health plan administrative databases, SEER, and medical record review from one year before incident BC through the end of study follow-up defined as the earliest of death, disenrollment from GH (>90 day lapse in enrollment), or end of study (chart abstraction date). GH's administrative data files include demographics, start and end dates of enrollment, inpatient and outpatient diagnoses and procedures, breast screening services and results, pharmacy dispensings, laboratory results, vital signs, and death [26]. The pharmacy database includes all medications dispensed at GH's outpatient pharmacies as well as claims from contracting pharmacies and is estimated to be 97 % complete [26–28]. SEER was the primary source of information for incident BC characteristics including diagnosis year, American Joint Committee on Cancer (AJCC) stage [29], lymph node status, hormone receptor status, and tumor size. Data on BC risk factors and surveillance were obtained from GH's Breast Cancer Screening Recruitment and Reminder (BSRR) survey taken at the time of each mammography [30]. The source of death data was GH data files that contain date of death for enrollees who die in a hospital while covered by GH or whose survivors inform GH (for instance, to terminate insurance coverage), required reporting of death notice from CMS for Medicare beneficiaries, death information from SEER, and an ongoing link to Washington State death tapes [31] which includes records of those who die out of state. Medical record review was done to confirm the incident BC diagnosis and to collect its treatment information, some comorbidities, and details on recurrence and second primary BC that occurred during the follow-up period.

A total of 4,426 potentially eligible subjects underwent medical record review. Women were further excluded on



the basis of no medical record (n=72), bilateral disease (n=6), recurrences or second primary BC that were incorrectly identified as incident BC (n=79), and no definitive surgery (n=44). A final cohort of 4,216 women were included in the study after excluding five patients who died and four patients with metastases that occurred prior to 120 days after completing surgery which was defined as the first date becoming at risk of SBCE.

Exposures

Our exposures of interest included diabetes and use of diabetes medications classes. A woman was defined as having diabetes if she had ≥ 2 outpatient (within a 6-month period) or one inpatient encounter that was assigned an ICD-9 code of 250.00-250.93 during the year prior to the incident BC through the end of study follow-up [32–34]. Date of diabetes diagnosis was the first date when either of these criteria were met. Women were defined as users of a particular diabetes medication class including insulin, metformin, and sulfonylureas during the follow-up period if they had >1 dispensing of a medication in the class of interest after the incident BC diagnosis. The presence of diabetes and use of diabetes medication classes were modeled as time-varying covariates that once a woman had the diagnosis or started the medication, she would remain as having diabetes or a user during the remaining of the follow-up. In addition, the three diabetes medication classes were not mutually exclusive. Therefore, medication use among women with diabetes accounted for the progression of monotherapy to combination therapy by allowing medication users to have progressive, time-varying exposure to multiple diabetes medication classes [35]. Insulin was further classified as long-acting, short-acting, and rapid-acting.

Outcomes

The primary end point of interest was SBCE, defined as the first of recurrence in any regional or distant sites or second primary ductal carcinoma in situ (DCIS) or invasive cancer of the ipsilateral or contralateral breast [36]. Secondary end points were BC-specific mortality and all-cause mortality defined based on the cause-of-death information recorded in the State death tapes.

Statistical analysis

Cox proportional hazards regression models were used to evaluate the association between diabetes and diabetes medication use with risks of SBCE and mortality. We estimated hazard ratios (HR) and their 95 % confidence intervals (CI) while adjusting for potential confounders. We modeled time from the incident BC diagnosis with a

delayed entry at 120 days post-surgery when a woman was defined to become at risk of a SBCE to the first of SBCE, death, disenrollment from the health plan, or end of study date. Individual events (i.e., recurrence and second primary BC) that made up the composite outcome of SBCE were also modeled to obtain a comprehensive assessment of the medication effects [37]. Specifically, in the analyses of recurrence as an outcome, women were censored at the earliest of disenrollment, end of follow-up, and competing events including death and second primary BC. Analyses for BC-specific mortality and all-cause mortality were modeled in a similar manner but without competing events in the models.

Our unadjusted model included adjustment for age at diagnosis and AJCC stage for the incident BC. We considered potential confounders *a priori* and similar to other studies [38–41] further adjusted for the following covariates in the multivariable models: calendar year, hormone receptor status, and primary treatment for the incident BC diagnosis; body mass index (BMI), smoking status, and menopausal status defined at the time of diagnosis; timevarying covariates including endocrine therapy for the incident BC, Charlson comorbidity score [42], statin use, prescription nonsteroidal anti-inflammatory (NSAID) medication use, Cox-2 inhibitors, and aspirin; and receipt of a screening mammogram in the prior 12 months.

We included all 4,216 women in the analyses of diabetes on risks of SBCE, BC-specific, and all-cause mortality. In separate models, we further evaluated the effects of medication exposures by including the three classes of diabetes medications. Diabetes medications were also evaluated among 610 women with diabetes any time during the year prior to incident BC through end of follow-up. In this subgroup analysis, if the date of diabetes diagnosis occurred after the at risk date for SBCE (i.e., 120 days post-surgery for incident BC), then the date of diabetes became the new at risk date (i.e., the new delayed entry date in the Cox models).

We tested for interaction between the exposures of interest and the logarithm of follow-up time to evaluate proportional hazards assumptions. There was no evidence to suggest violation of the assumptions.

All analyses were performed using SAS statistical software version 9.3 (SAS Institute Inc., Cary, NC). The Institutional Review Board at GH approved this study.

Results

Analyses among all women

Characteristics of the 4,216 women with incident stage I or II BC included in the COMBO study are described in detail



elsewhere [23, 24] (Table 1). The majority of women in the cohort were postmenopausal and had a Charlson comorbidity score of zero at diagnosis. The majority of incident BC were AJCC stage I, estrogen receptor (ER)-positive/progesterone receptor (PR)-positive, treated with breast-conserving surgery with or without radiation, not treated with chemotherapy, and treated with endocrine therapy.

There were 432 (10 %) women who experienced a recurrence and 153 (4 %) second primary BC during study follow-up, yielding a total of 558 (13 %) women having a SBCE (with 415 recurrences and 143 s primary BC as the first SBCE events). Median time from incident BC diagnosis to the SBCE was 3.3 years. Among recurrences, 67 % were distant, 32 % local or regional, and 1 % DCIS. Among second primary cancers, 21 % were DCIS, 49 % stage I, 21 % stage II, 4 % stage III/IV, and 5 % unknown stage. At the end of study follow-up, 22 % of women died, and 6 % were BC-specific.

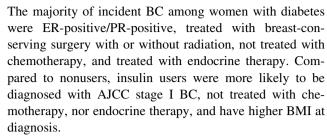
Women who experienced a SBCE were more likely to be peri- or premenopausal, diagnosed with AJCC stage II, lymph node-positive, ER- and/or PR-negative, tumor size >2 cm, HER2-positive, treated by mastectomy, treated with chemotherapy, not treated with endocrine therapy, and detected by a diagnostic versus screening mammography compared to women without a SBCE during the follow-up period (Table 1).

Unadjusted and adjusted models did not differ substantially, so we present only results from adjusted models. We observed a statistically significant association between diabetes diagnosis and an increased risk of recurrence (HR = 1.57; 95 % CI 1.09–2.25), but not of SBCE (HR = 1.29; 95 % CI 0.94–1.76) or second primary BC (HR = 0.74; 95 % CI 0.39–1.41) (Fig. 1). Results suggested a slightly elevated risk of both BC-specific and all-cause mortality among women with diabetes compared to those without, though the confidence intervals included 1.0.

Associations between oral diabetes medication classes and SBCE were not statistically significant, but point estimates suggested an increased risk of BC-specific mortality with use of sulfonylureas (HR = 1.67; 95 % CI 0.98-2.86). Metformin users had a nonsignificant decrease in risk of BC-specific mortality (HR = 0.67; 95 % CI 0.35-1.27) and a significant decrease in risk of all-cause mortality (HR = 0.58; 95 % CI 0.42-0.79; Fig. 1). Compared to nonusers, insulin users had an increased risk of BC-specific mortality (HR = 2.22; 95 % CI 1.27-3.91) and all-cause mortality (HR = 2.02; 95 % CI 1.56-2.61).

Analyses among women with diabetes

Among the 610 women with diabetes, most women were postmenopausal, had BMI \geq 30 kg/m², and a Charlson comorbidity score of one or greater at diagnosis (Table 2).



During follow-up, 76 % had ever used oral or injectable diabetes medications during the study follow-up, and 30 % used both insulin and oral medications. Individually or in combination, 40 % used insulin, 45 % used metformin, and 50 % used sulfonylureas (Table 2). Among the 246 insulin users, 19 % used long-acting insulin, 80 % used short-acting insulin, and 25 % used rapid-acting insulin analogues alone or in combination. Compared to these women, those with diabetes who did not use oral or injectable diabetes medications during follow-up had higher median age, lower proportions of obese women and lower Charlson scores (Online Resource 1).

At the end of study follow-up, 58 (9 %) of the 610 women with diabetes had a recurrence and 16 (3 %) had a second primary BC, giving a total of 71 (12 %) women experienced a SBCE (with 57 recurrences and 14 s primary BC as the first SBCE). There were 187 (31 %) deaths of all causes of which 37 were due to BC. Since there were few second primary BC and BC-specific deaths observed in this subgroup, our analyses were limited to evaluating the composite SBCE outcome, recurrences, and all-cause mortality.

Results from the multivariable analyses of women with diabetes are shown in Fig. 2. There were no significant associations between insulin use and risk of SBCE (HR = 1.49; 95 % CI 0.88-2.53). However, insulin use was associated with an increased risk of BC recurrence (HR = 1.94; 95 % CI 1.08-3.48) and all-cause mortality (HR = 2.33; 95 % CI 1.70-3.20). These associations with recurrence and all-cause mortality were seen in all types of insulin use: use of long-acting insulin only, rapid-/short-acting insulin only, or both (Table 3).

We did not observe any associations between oral diabetes medications and risk of SBCE (Fig. 2). Use of metformin but not sulfonylureas was associated with a decreased risk of all-cause mortality (HR = 0.55; 95 % CI 0.38-0.79).

Discussion

In a large cohort of women diagnosed with an early-stage BC, we found an increased risk of BC recurrence among women with diabetes compared to those without. We did not find an association between diabetes diagnosis and BC-



Table 1 Descriptive characteristics of 4,216 women included in the COMBO study by SBCE status

	All women $n = 4,216$	SBCE ^a		
		No $(n = 3,658)$	Yes $(n = 558)$	
Characteristics at incident breast cancer diagnosis	1			
Year of diagnosis				
1990–1994	950 (22.5)	755 (20.6)	195 (34.9)	
1995–1999	1,191 (28.2)	1,020 (27.9)	171 (30.6)	
2000–2004	1,201 (28.5)	1,073 (29.3)	128 (22.9)	
2005–2008	874 (20.7)	810 (22.1)	64 (11.5)	
Age at diagnosis, years				
Median (interquartile range)	63 (52–73)	63 (52–73)	62 (50–72)	
18–39	139 (3.3)	112 (3.1)	27 (4.8)	
40–49	646 (15.3)	544 (14.9)	102 (18.3)	
50–59	995 (23.6)	866 (23.7)	129 (23.1)	
60–69	1,018 (24.1)	889 (24.3)	129 (23.1)	
70–79	940 (22.3)	824 (22.5)	116 (20.8)	
80+	478 (11.3)	423 (11.6)	55 (9.9)	
Menopausal status at diagnosis				
Peri- or premenopausal	1,145 (27.2)	956 (26.1)	189 (33.9)	
Postmenopausal	3,071 (72.8)	2,702 (73.9)	369 (66.1)	
Body mass index (kg/m ²)				
Mean (SD)	28 (6.2)	28 (6.2)	28.3 (6.6)	
Median	26.8	26.7	27.5	
<18.5	69 (1.6)	55 (1.5)	14 (2.5)	
18.5–24.9	1,453 (34.6)	1,269 (34.8)	184 (33.3)	
25.0–29.9	1,362 (32.5)	1,186 (32.6)	176 (31.8)	
30.0–34.9	766 (18.3)	666 (18.3)	100 (18.1)	
35+	546 (13.0)	467 (12.8)	79 (14.3)	
Unknown	20	15	5	
Smoking status at diagnosis				
Current	253 (6.0)	230 (6.3)	23 (4.1)	
Past	352 (8.3)	318 (8.7)	34 (6.1)	
Never/unknown	3,611 (85.6)	3,110 (85.0)	501 (89.8)	
AJCC stage				
I	2,648 (62.8)	2,384 (65.2)	264 (47.3)	
IIA	1,078 (25.6)	906 (24.8)	172 (30.8)	
IIB	490 (11.6)	368 (10.1)	122 (21.9)	
ER/PR status				
ER(-)/PR(-)	667 (15.8)	531 (14.5)	136 (24.4)	
ER(+)/PR(-)	383 (9.1)	319 (8.7)	64 (11.5)	
ER(-)/PR(+)	61 (1.4)	47 (1.3)	14 (2.5)	
ER(+)/PR(+)	2,888 (68.5)	2,572 (70.3)	316 (56.6)	
ER and/or PR unknown	217 (5.1)	189 (5.2)	28 (5.0)	
HER2 status				
Test performed	2,074 (79.7)	1,874 (80.5)	200 (73.0)	
Positive/borderline	353 (17.0)	311 (16.6)	42 (21.0)	
Negative	1,714 (82.6)	1,556 (83.0)	158 (79.0)	
No result	7 (0.3)	7 (0.4)	0 (0.0)	
Surgical procedure				
Mastectomy including radical \pm radiation	1,521 (36.1)	1,289 (35.2)	232 (41.6)	



Table 1 continued

	All women $n = 4,216$	SBCE ^a		
		No $(n = 3,658)$	Yes $(n = 558)$	
Breast-conserving, + radiation	2,172 (51.5)	1,927 (52.7)	245 (43.9)	
Breast-conserving, no radiation	523 (12.4)	442 (12.1)	81 (14.5)	
Other treatment				
Any chemotherapy	1,376 (32.6)	1,142 (31.2)	234 (41.9)	
Completed course	1,212 (88.1)	1,003 (87.8)	209 (89.3)	
Any endocrine therapy	2,363 (56.0)	2,101 (57.4)	262 (47.0)	
Charlson score at diagnosis				
0	3,229 (76.6)	2,784 (76.1)	445 (79.7)	
1	704 (16.7)	625 (17.1)	79 (14.2)	
2+	283 (6.7)	249 (6.8)	34 (6.1)	
Diabetes				
Diabetes (ever) throughout follow-up	610 (14.5)	539 (14.7)	71 (12.7)	
At BC diagnosis (year prior to BC)	329 (53.9)	294 (54.5)	35 (49.3)	
After BC diagnosis through follow-up	281 (46.1)	245 (45.5)	36 (50.7)	
Diabetes medication use in year prior to breast ca	ancer diagnosis ^b			
Any oral or injectable medications	277 (6.6)	241 (6.6)	36 (6.5)	
Any oral medications	225 (5.3)	200 (5.5)	25 (4.5)	
Metformin	106 (47.1)	97 (48.5)	9 (36.0)	
Sulfonylurea	175 (77.8)	152 (76.0)	23 (92.0)	
Insulin	103 (2.4)	88 (2.4)	15 (2.7)	
Long-acting	12 (11.7)	11 (12.5)	1 (6.7)	
Short-acting	88 (85.4)	74 (84.1)	14 (93.3)	
Rapid-acting	16 (15.5)	13 (14.8)	3 (20.0)	
Characteristics throughout follow-up				
Diabetes medication use throughout follow-up ^b				
Any oral or injectable medications	464 (11.0)	408 (11.2)	56 (10.0)	
Any oral medications	403 (9.6)	358 (9.8)	45 (8.1)	
Metformin	275 (68.2)	248 (69.3)	27 (60.0)	
Sulfonylurea	307 (76.2)	270 (75.4)	37 (82.2)	
Insulin	246 (5.8)	218 (6.0)	28 (5.0)	
Long-acting	47 (19.1)	43 (19.7)	4 (14.3)	
Short-acting	198 (80.5)	173 (79.4)	25 (89.3)	
Rapid-acting	62 (25.2)	56 (25.7)	6 (21.4)	
% Follow-up years with yearly screening mamm		,		
<50 %	939 (22.3)	793 (21.7)	146 (26.2)	
50–80 %	1,439 (34.1)	1,284 (35.1)	155 (27.8)	
>80 %	1,838 (43.6)	1,581 (43.2)	257 (46.1)	
Years of follow-up	1,000 (1010)	1,001 (10.2)	207 (1011)	
Mean (SD)	7.1 (4.3)	7.5 (4.3)	4.4 (3.5)	
Median	6.3	6.7	3.3	

a SBCE second breast cancer event includes recurrence or second primary breast cancer, in situ and invasive

specific mortality, although estimates were suggestive of an increased risk of all-cause mortality. Among women with diabetes, insulin use was associated with increased risks of both recurrence and all-cause mortality. We did not

observe a statistically significant association between use of metformin or sulfonylureas with risk of SBCE, although metformin users had a reduced risk of all-cause mortality compared to nonusers.



b Not mutually exclusive

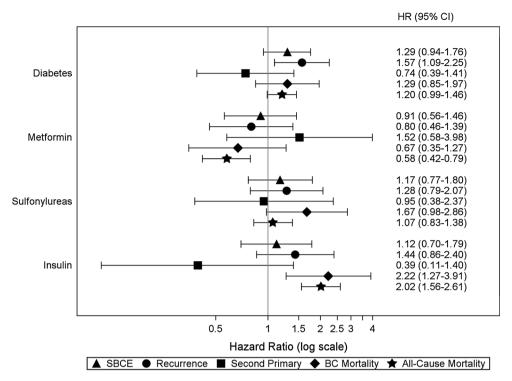


Fig. 1 Risk of SBCE, breast cancer-specific mortality, and all-cause mortality in relation to diabetes and diabetes medication classes *asterisk* among all women in the COMBO cohort. *SBCE* second breast cancer event (recurrence and second primary BC), *BC* breast cancer, *HR* hazard ratio, *CI* confidence interval. *Note* All hazard ratios are adjusted for age at diagnosis (18–49, 50–59, 60–69, 70–79, 80+ years); diagnosis year (1990–1994, 1995–1999, 2000–2004, 2005–2008); AJCC stage (I, IIA, IIB); hormone receptor status (estrogen receptor [ER] –/progesterone receptor [PR] –, ER +/PR –, ER –/PR +, ER +/PR +, and ER and/or PR unknown); primary treatment for initial BC (mastectomy, breast-conserving surgery with radiation, breast-conserving surgery without radiation); endocrine therapy for the incident BC (yes/no, time-varying); body mass index

(BMI) at diagnosis (<18.5, 18.5–24.9, 25.0–29.9, 30.0–34.9, 35+ kg/m²); smoking status at diagnosis (current, past, never/unknown); menopausal status at diagnosis (peri- or premenopausal, post-menopausal); Charlson comorbidity score (0, 1, 2+, time-varying); statin use (yes/no, time-varying); prescription nonsteroidal anti-inflammatory medication use, Cox-2 inhibitors, and aspirin (yes/no, time-varying); and receipt of screening mammogram in the 12 months prior to events (yes/no, time-varying). *Asterisk* hazard ratio for diabetes is from a model that did not adjust for diabetes medication use; hazard ratios for individual diabetes medication classes are from a separate model adjusting for all diabetes medication classes of interest (ever/never use, time-varying) and diabetes diagnosis (yes/no, time-varying)

Our finding of significantly increased risks of BC recurrence and all-cause mortality among insulin users warrants further studies in larger populations powered to investigate possible differences in incident and prevalent insulin users, duration of insulin use, and length of time with diagnosed diabetes. Multiple other explanations could underlie worse prognosis among BC survivors with diabetes and using insulin. For example, insulin use could be a surrogate for more advanced diabetes or poorer control of diabetes [43] which may lead to higher mortality. So while we found a significantly increased risk of breast cancer recurrence associated with insulin use, confounding by indication (i.e., diabetes severity) makes this observed association less clear [44].

If insulin users do have more severe diabetes and greater complications (e.g., neuropathy, end-organ damage), otherwise appropriate adjuvant chemotherapy treatment may be used less [45]. In our cohort, a slightly smaller

proportion of women who were ever insulin users were treated with chemotherapy compared to nonusers (21.5 vs. 28.6 %). Women with severe diabetes and the greatest degree of hyperinsulinemia and insulin resistance would also likely have very reduced IGF binding proteins and high bioavailability of IGF-I, a growth factor and gonadotrophic factor in BC [46]. He et al. [41] evaluated different classes of diabetes medications in relation to BCspecific mortality and all-cause mortality in a cohort of 154 women with diabetes and stages II-IV HER2+ BC. When comparing insulin users to nonusers, no association was found with risk of BC-specific mortality (HR = 1.31; 95 % CI 0.64-2.69) and all-cause mortality (HR = 1.38; 95 % CI 0.72–2.67) in multivariable models. Some considerable differences between He et al.'s study and our study include their selection of higher stage, HER2+ BC cases only, classification of medication use at BC presentation only, not differentiating by type of insulin analogues, and a



Table 2 Descriptive characteristics of 610 women with diabetes in the COMBO study by insulin use

	All women with diabetes $n = 610$	Insulin use		
		Nonuser $(n = 364)$	User Yes $(n = 246)$	
Characteristics at incident breast cancer diagn	osis			
Year of diagnosis				
1990–1994	97 (15.9)	54 (14.8)	43 (17.5)	
1995–1999	200 (32.8)	119 (32.7)	81 (32.9)	
2000–2004	181 (29.7)	109 (29.9)	72 (29.3)	
2005–2008	132 (21.6)	82 (22.5)	50 (20.3)	
Age at diagnosis, years				
Median (interquartile range)	67 (58–74)	67 (56–74)	67 (59–74)	
18–39	5 (0.8)	3 (0.8)	2 (0.8)	
40–49	35 (5.7)	16 (4.4)	19 (7.7)	
50–59	122 (20.0)	83 (22.8)	39 (15.9)	
60–69	197 (32.3)	114 (31.3)	83 (33.7)	
70–79	179 (29.3)	103 (28.3)	76 (30.9)	
80+	72 (11.8)	45 (12.4)	27 (11.0)	
Menopausal status at diagnosis				
Peri- or premenopausal	81 (13.3)	45 (12.4)	36 (14.6)	
Postmenopausal	529 (86.7)	319 (87.6)	210 (85.4)	
Body mass index (kg/m ²)				
Mean (SD)	32.1 (6.9)	31.5 (6.1)	33.1 (7.8)	
Median	31.2	30.7	32.3	
<18.5	5 (0.8)	2 (0.5)	3 (1.2)	
18.5–24.9	78 (12.8)	46 (12.6)	32 (13.1)	
25.0–29.9	173 (28.4)	119 (32.7)	54 (22.0)	
30.0–34.9	175 (28.7)	103 (28.3)	72 (29.4)	
35+	178 (29.2)	94 (25.8)	84 (34.3)	
Unknown	1	0	1	
Smoking status at diagnosis				
Current	37 (6.1)	16 (4.4)	21 (8.5)	
Past	59 (9.7)	42 (11.5)	17 (6.9)	
Never/unknown	514 (84.3)	306 (84.1)	208 (84.6)	
AJCC stage	,	, ,	, ,	
I	383 (62.8)	212 (58.2)	171 (69.5)	
IIA	150 (24.6)	103 (28.3)	47 (19.1)	
IIB	77 (12.6)	49 (13.5)	28 (11.4)	
ER/PR status		,	, ,	
ER(-)/PR(-)	93 (15.2)	53 (14.6)	40 (16.3)	
ER(+)/PR(-)	45 (7.4)	31 (8.5)	14 (5.7)	
ER(-)/PR(+)	8 (1.3)	3 (0.8)	5 (2.0)	
ER(+)/PR(+)	441 (72.3)	264 (72.5)	177 (72.0)	
ER and/or PR unknown	23 (3.8) 13 (3.6)		10 (4.1)	
HER2 status		,	, ,	
Test performed	312 (77.6)	189 (78.1)	123 (76.9)	
Positive/borderline	56 (17.9)	37 (19.6)	19 (15.4)	
Negative	256 (82.1)	152 (80.4) 104 (84.6)		
Surgical procedure		, ,	,	
Mastectomy including radical \pm radiation	228 (37.4)	132 (36.3)	96 (39.0)	
Breast-conserving, +radiation	311 (51.0)	183 (50.3)	128 (52.0)	



Table 2 continued

	All women with diabetes $n = 610$	Insulin use		
		Nonuser $(n = 364)$	User Yes $(n = 246)$	
Breast-conserving, no radiation	71 (11.6)	49 (13.5)	22 (8.9)	
Other treatment				
Any chemotherapy	157 (25.7)	104 (28.6)	53 (21.5)	
Completed course	133 (84.7)	92 (88.5)	41 (77.4)	
Any endocrine therapy	358 (58.7)	224 (61.5)	134 (54.5)	
Charlson score at diagnosis				
0	209 (34.3)	163 (44.8)	46 (18.7)	
1	223 (36.6)	116 (31.9)	107 (43.5)	
2+	178 (29.2)	85 (23.4)	93 (37.8)	
Diabetes				
Diabetes (ever) throughout follow-up	610 (14.5)	364 (100)	246 (100)	
At BC diagnosis (year prior to BC)	329 (53.9)	159 (43.7)	170 (69.1)	
After BC diagnosis through follow-up	281 (46.1)	205 (56.3)	76 (30.9)	
Diabetes medication use in year prior to breas	et cancer diagnosis ^a			
Any oral or injectable medications	277 (45.4)	96 (26.4)	181 (73.6)	
Any oral medications	225 (36.9)	94 (25.8)	131 (53.3)	
Metformin	106 (47.1)	38 (40.4)	68 (51.9)	
Sulfonylurea	175 (77.8)	71 (75.5)	104 (79.4)	
Insulin	103 (16.9)		103 (41.9)	
Long-acting	12 (11.7)		12 (11.7)	
Short-acting	88 (85.4)		88 (85.4)	
Rapid-acting	16 (15.5)		16 (15.5)	
Characteristics throughout follow-up				
Oral diabetes medication use during follow-up	o ^a			
Any oral or injectable medications	464 (76.1)	218 (59.9)	246 (100)	
Any oral medications	403 (66.1)	218 (59.9)	185 (75.2)	
Metformin	275 (45.1)	141 (38.7)	134 (54.5)	
Sulfonylureas	307 (50.3)	155 (42.6)	152 (61.8)	
Insulin	246 (40.3)		246 (100)	
Long-acting	47 (7.7)		47 (19.1)	
Short-acting	198 (32.5)		198 (80.5)	
Rapid-acting	62 (10.2)		62 (25.2)	
% Follow-up years with yearly screening ma				
<50 %	161 (26.4)	93 (25.5)	68 (27.6)	
50–80 %	207 (33.9)	117 (32.1)	90 (36.6)	
>80 %	242 (39.7)	154 (42.3)	88 (35.8)	
Years of follow-up				
Mean (SD)	6.8 (3.8)	6.8 (3.8)	6.8 (3.7)	
Median	6.5	6.5	6.4	

^a Not mutually exclusive

smaller sample size for comparison (48 insulin users vs. 106 nonusers). While the smaller proportion of certain types of insulin analogues (i.e., long-acting) is reflected in our wider confidence intervals, our associations for increased risks of BC recurrence and all-cause mortality

remained statistically significant when evaluating the type of insulin use among women with diabetes.

Our observed reduction in risk of all-cause mortality among BC survivors with diabetes who used metformin is in agreement with findings from some epidemiologic



Table 3 Risk of SBCE, recurrence, and all-cause mortality in relation to diabetes medication classes among women with diabetes in the COMBO cohort (insulin analogues stratified by type)

	SBCE		Recurrence	Recurrence		All-cause mortality	
	HR	95 % CI	HR	95 % CI	HR	95 % CI	
Diabetes medication classes							
Long-acting insulin only	7.10	(0.85-59.2)	13.9	(1.58–122)	4.95	(1.10-22.2)	
Short-/rapid-acting insulin only	1.58	(0.91-2.75)	2.07	(1.12-3.81)	1.93	(1.38-2.71)	
Both long- and short-/rapid-acting insulin	2.23	(0.64–7.75)	3.62	(1.02-12.9)	3.24	(1.83-5.73)	
Metformin	1.00	(0.57-1.77)	1.08	(0.56-2.06)	0.59	(0.40-0.85)	
Sulfonylureas	1.00	(0.59-1.70)	0.99	(0.54–1.81)	1.10	(0.79-1.53)	

All hazard ratios for individual diabetes medication classes are adjusted for other medication classes of interest (ever/never use, time-varying); age at diagnosis (18–49, 50–59, 60–69, 70–79, 80+ years); diagnosis year (1990–1994, 1995–1999, 2000–2004, 2005–2008); AJCC stage (I, IIA, IIB); hormone receptor status (estrogen receptor [ER]—/progesterone receptor [PR]—, ER+/PR—, ER-/PR+, ER+/PR+, and ER and/or PR unknown); primary treatment for initial BC (mastectomy, breast-conserving surgery with radiation, breast-conserving surgery without radiation); endocrine therapy for the incident BC (yes/no, time-varying); body mass index (BMI) at diagnosis (<25.0, 25.0–29.9, 30.0–34.9, 35+ kg/m²); smoking status at diagnosis (current, past, never/unknown); menopausal status at diagnosis (peri- or premenopausal, postmenopausal); Charlson comorbidity score (<2, 2+, time-varying); statin use (yes/no, time-varying); prescription nonsteroidal anti-inflammatory medication use, Cox-2 inhibitors, and aspirin (yes/no, time-varying); and receipt of screening mammogram in the 12 months prior to events (yes/no, time-varying) SBCE second breast cancer event (recurrence and second primary BC), HR hazard ratio, CI confidence interval

studies [40, 41] but not others [38, 39]. In a Danish registry-based cohort of 1,058 women with BC and diabetes who were treated with diabetes medications [40], current metformin users compared to nonusers had a significant reduction in all-cause mortality (HR = 0.74; 95 % CI 0.58-0.96). Other studies reported a beneficial effect of metformin in BC survivors including reduced BC-specific mortality in HER2+ women [41] and reduced risk of distant metastases in triple receptor-negative BC [38]. Another population-based cohort study of 2,361 BC survivors ages >66 years with diabetes [39] found no association between cumulative duration of past metformin use with risks of all-cause mortality (HR = 0.97; 95 % CI 0.92-1.02) and BC-specific mortality (HR = 0.91; 95 % CI 0.81-1.03). However, this study lacked information on important clinical data and did not account for some important potential confounders, including BMI, family history, tumor stage, and recurrence.

Our study has several strengths that include use of a well-characterized and population-based cohort of BC survivors; ascertainment of incident BC cases through a validated registry; detailed information on incident BC characteristics and treatment with extensive follow-up; review of medical charts to determine recurrences and obtain characteristics on recurrent cancer; health care utilization including medications dispensed and breast services; comprehensive information on comorbidities to examine confounding; and application of analytic methods to address potential competing risks and informative censoring in a robust manner. Also, we were able to collect date of diabetes diagnosis and dispensing of all diabetes medications during the year prior to the incident BC

through the end of study follow-up, allowing us to update and evaluate the effects of diabetes and diabetes medication classes during the entire follow-up period.

Findings from our study should be interpreted in the context of the following limitations. First, COMBO uses data from a single health plan and includes an insured, educated, and primarily Caucasian population. This may limit generalizability. Loss to follow-up is a possible source of bias with 18 % censored due to disenrollment from the health plan.

Second, women with diabetes in our cohort were classified as medication users based on time-varying exposure to diabetes medications from pharmacy dispensing records following incident BC diagnosis. However, data from COMBO only went back one year prior to incident BC diagnosis, so we could not distinguish prevalent and incident medication users nor did we have complete information on total length of time with diabetes prior to the study period. We also could not distinguish women with type I versus type II diabetes. Another limitation to our classification of "ever/never" use of diabetes medications was the inability to examine subcategories of duration or cumulative use.

Third, despite our rigorous study design and statistical approach, residual confounding is possible. We ascertained and considered the majority of potential confounders, but lacked information on certain modifiable lifestyle factors such as diet, physical activity, over-the-counter (OTC) aspirin/NSAID use, and alcohol intake. People who fill a prescription but do not ingest the medication and who obtain medications at non-GH pharmacies with no claim submitted to GH may be misclassified, but GH enrollees



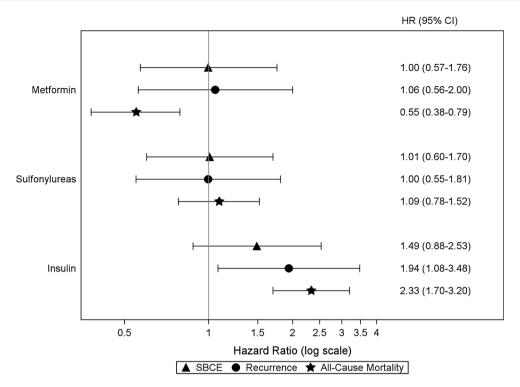


Fig. 2 Risk of SBCE, recurrence, and all-cause mortality in relation to diabetes medication classes among women with diabetes in the COMBO cohort. *SBCE* second breast cancer event (recurrence and second primary BC), *BC* breast cancer, *HR* hazard ratio, *CI* confidence interval. *Note* All hazard ratios for individual diabetes medication classes are adjusted for other medication classes of interest (ever/never use, time-varying); age at diagnosis (18–49, 50–59, 60–69, 70–79, 80+ years); diagnosis year (1990–1994, 1995–1999, 2000–2004, 2005–2008); AJCC stage (I, IIA, IIB); hormone receptor status (estrogen receptor [ER]–/progesterone receptor [PR]–, ER +/PR–, ER-/PR+, ER+/PR+, and ER and/or

PR unknown); primary treatment for initial BC (mastectomy, breast-conserving surgery with radiation, breast-conserving surgery without radiation); endocrine therapy for the incident BC (yes/no, time-varying); body mass index (BMI) at diagnosis (<25.0, 25.0–29.9, 30.0–34.9, 35+ kg/m²); smoking status at diagnosis (current, past, never/unknown); menopausal status at diagnosis (peri- or premenopausal, postmenopausal); Charlson comorbidity score (<2, 2+, time-varying); statin use (yes/no, time-varying); prescription non-steroidal anti-inflammatory medication use, Cox-2 inhibitors, and aspirin (yes/no, time-varying); and receipt of screening mammogram in the 12 months prior to events (yes/no, time-varying)

obtain almost all of their prescription medications at GH pharmacies or contracting pharmacies and any misclassification is likely non-differential [26–28].

Finally, due to limited statistical power, we did not evaluate other hypotheses related to risk of SBCE such as the interaction between medication classes or differences in risk within subgroups according to medication adherence, clinical control of diabetes, BC stage, and BMI. This also limited the ability to test our hypotheses within restricted groups, such as only women with diabetes prior to BC. Our median follow-up was 6.3 years overall and 3.3 years in women who developed SBCE. Given the small number of second primary BC and BC-specific deaths, additional studies in larger and more diverse populations are warranted to better understand the effects of diabetes medications on these BC outcomes.

Given the biological plausibility and results of our and other studies, it appears likely that women with diabetes have worse BC outcomes than similar women without diabetes. This is of public health importance. The relationship between diabetes and BC is complex, with one affecting the other in potentially dramatic ways; yet, optimal tailoring of treatment for these patients is unclear. Our results should be used to spur discussions about the importance of BC screening and surveillance but also prompt questions on whether there may be differences in treatment management that help explain differences in health outcomes. For instance, recent studies report that greater diabetes severity is associated with lower concordance with National Comprehensive Cancer Network guidelines for locoregional BC treatment among middle-aged women and adjuvant chemotherapy [45]. Studies such as ours that provide additional information on differences in outcomes by diabetes treatments can aide in the decision-making process of appropriately managing diabetes while also optimizing cancer outcomes. Additional studies are necessary to confirm the possible adverse influences of insulin therapies on risk of BC recurrence and all-cause mortality, as well as the potential benefit of metformin therapy to reduce all-cause mortality in women with BC and diabetes.



If confirmed, then future interventions to improve patientcentered outcomes such as diabetes management and medication adherence as part of ongoing cancer prevention efforts would be well motivated.

Conclusion

Future studies are warranted in larger and more diverse populations who are powered to examine differences by length of time with diabetes or disease severity and the influence of diabetes medications on BC-specific mortality in women with diabetes. Questions related to the potential benefit of diabetes medications such as metformin in the adjuvant BC therapy setting and among women without diabetes will be further clarified in ongoing randomized trials of metformin versus placebo on risk of recurrence and survival [47]. Nevertheless, we believe that our study contributes to the limited evidence available to understand diabetes medication safety in relation to SBCE and the potential benefits of some diabetes medications in early-stage BC survivors with diabetes.

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Conflict of interest The authors declare that they have no conflict of interest.

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