

# Astrostatistics: 13 Mar 2019

<https://github.com/CambridgeAstroStat/PartIII-Astrostatistics-2019>

- Make-Up Lecture, Thu 14 Mar, 11am-12: MR13
- Example Class, Fri 15 Mar 1pm: MR 12
- Today:
- Hierarchical Bayes & Shrinkage Estimators

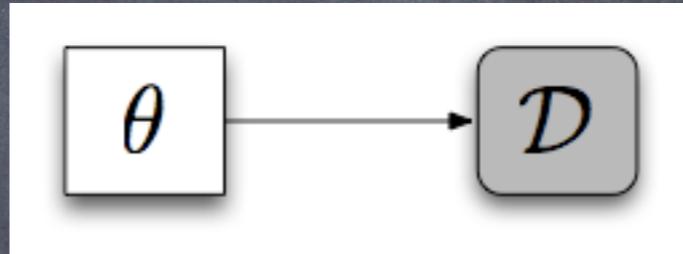
# Aside on Product of Gaussian densities

Review:

Hierarchical Bayes and Probabilistic Graphical Models:  
a visual way to understand complex statistical models

Simple Bayes:

$$\mathcal{D} | \theta \sim \text{Model}(\theta)$$

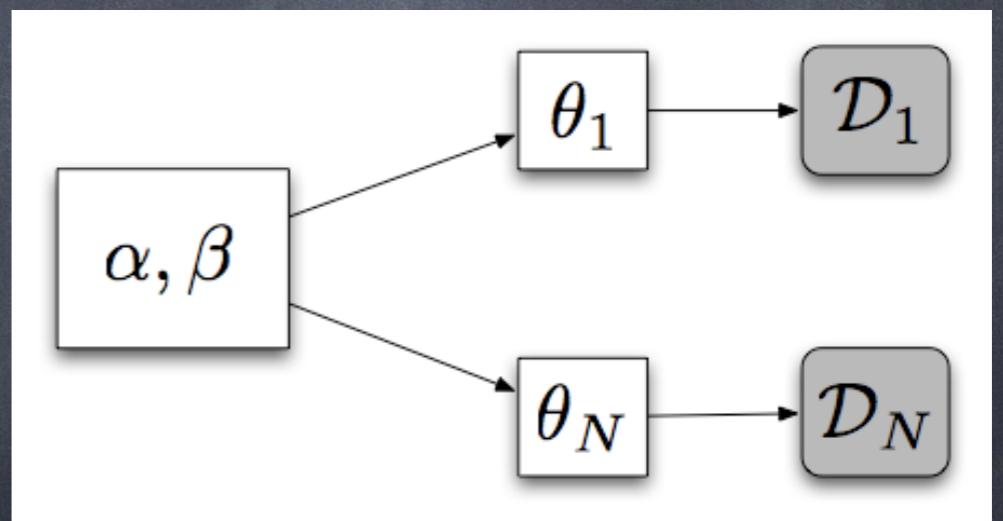


$$P(\theta | \mathcal{D}) \propto P(\mathcal{D} | \theta)P(\theta)$$

Hierarchical Bayes:

$$\mathcal{D}_i | \theta_i \sim \text{Model}(\theta_i)$$

$$\theta_i | \alpha, \beta \sim \text{PopModel}(\alpha, \beta)$$



$$P(\{\theta_i\}, \alpha, \beta | \{\mathcal{D}_i\}) \propto \left[ \prod_{i=1}^N P(\mathcal{D}_i | \theta_i) P(\theta_i | \alpha, \beta) \right] P(\alpha, \beta)$$

Build up complexity by layering conditional probabilities

# Review: Hierarchical Bayesian “Normal-Normal” Model

**Level 1:** Population Distribution of Latent Variables (Absolute Mags)

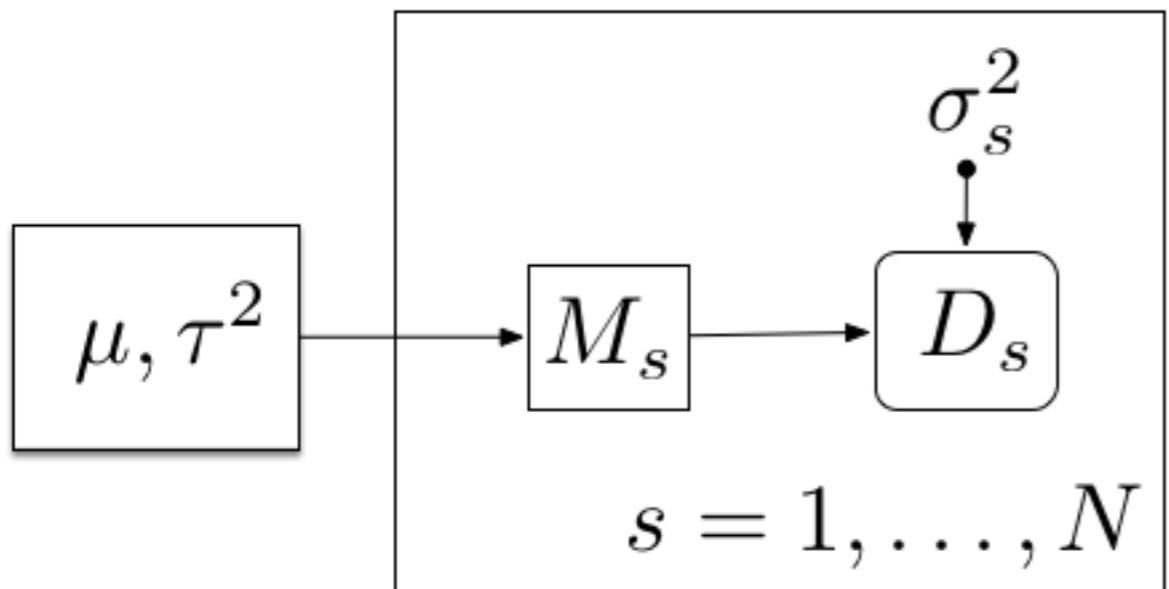
$$M_s \sim N(M_0, \tau^2) \quad s = 1 \dots N$$

**Level 2 :** Measurement Error Process

$$D_s | M_s \sim N(M_s, \sigma_s^2) \quad s = 1 \dots N$$

Joint Probability Density of Data, Latent Variables, Hyperparameters

$$P(\{D_s\}, \{M_s\}, H = M_0, \tau^2)$$



Joint factors into Conditional and Marginal densities based on Model Assumptions

# Hierarchical vs Regular Bayes

- Could regard as just a general Bayesian inference problem in a very high dimensional parameter space, e.g.

$$\boldsymbol{\theta} = \{M_1, \dots, M_N; M_0, \tau^2\} = \{\mathbf{M}; M_0, \tau^2\}$$

$$P(\boldsymbol{\theta}|\mathbf{D}) \propto P(\mathbf{D}|\boldsymbol{\theta})P(\boldsymbol{\theta})$$

$$P(\boldsymbol{\theta}|\mathbf{D}) \propto P(\mathbf{D}|\mathbf{M})P(\mathbf{M}|M_0, \tau^2)P(M_0, \tau^2)$$

$$P(\boldsymbol{\theta}|\mathbf{D}) \propto \left[ \prod_{s=1}^N P(D_s|M_s)P(M_s|M_0, \tau^2) \right] P(M_0, \tau^2)$$

- However, special hierarchical structure is useful for modelling, estimation, and computation
- For large N, wouldn't want to do an N+2 dimensional Metropolis MCMC!

# Hierarchical Bayesian “Normal-Normal” Model

**Level 1:** Population Distribution of Latent Variables (Absolute Mags)

$$M_s \sim N(M_0, \tau^2)$$

Population Dist'n / Prior

**Level 2 :** Measurement Error Process

$$D_s | M_s \sim N(M_s, \sigma_s^2)$$

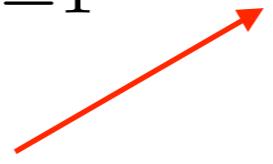
Measurement Likelihood

Joint Probability Density of ALL THE THINGS:  
Data, Latent Variables, Hyperparameters

$$P(\{D_s\}, \{M_s\}, H) = \left[ \prod_{s=1}^N P(D_s | M_s) P(M_s | M_0, \tau^2) \right] \times P(H)$$

(Derive on board)

Measurement Likelihood



Population Dist'n / Prior

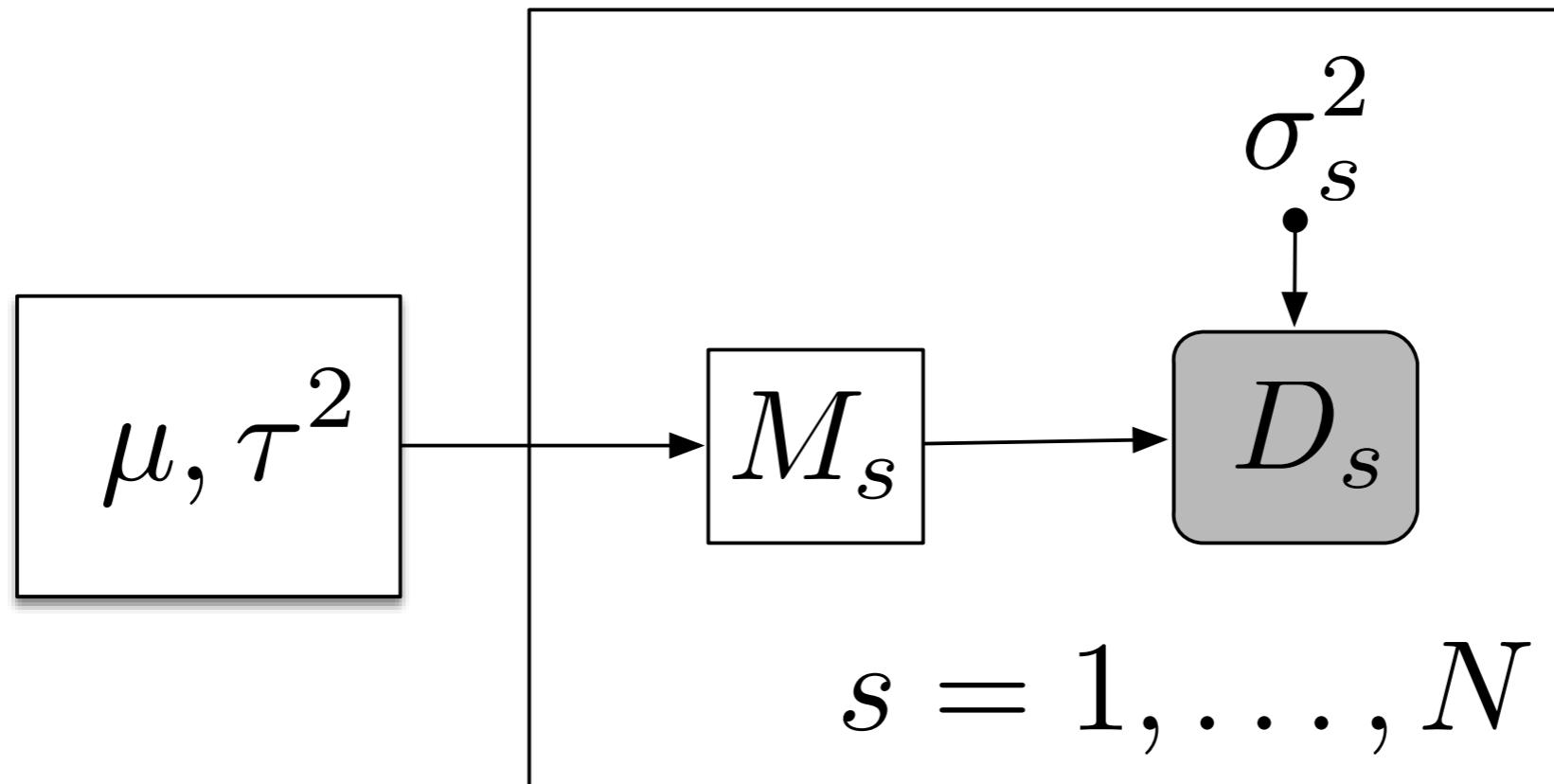


Hyperprior



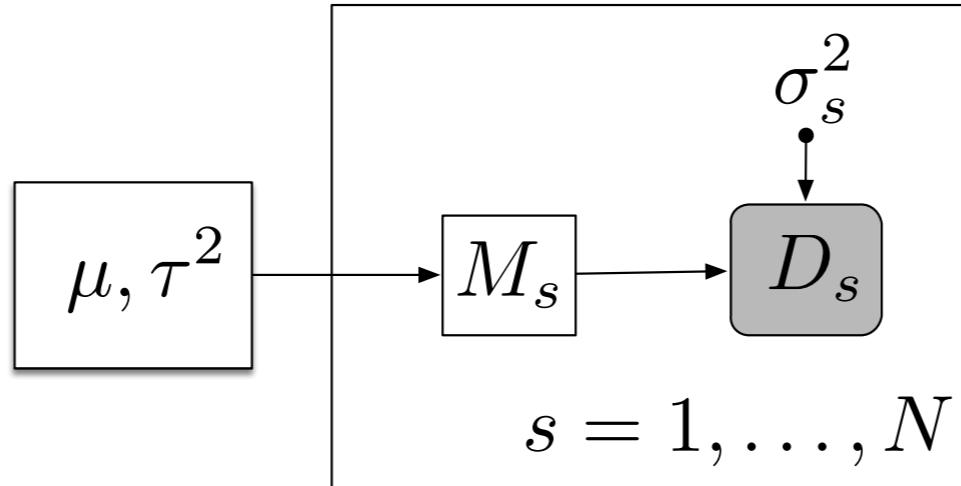
# Sampling the Hierarchical Bayesian posterior

$$P(\{M_s\}, H | \{D_s\}) \propto \left[ \prod_{s=1}^N P(D_s | M_s) P(M_s | M_0, \tau^2) \right] \times P(H)$$

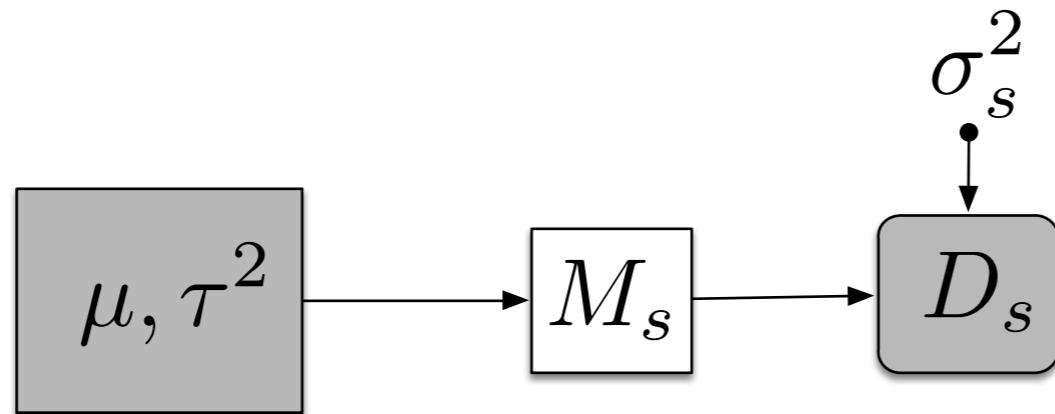


Utilises Conditional Independence structure of PGM to derive conditional posterior densities

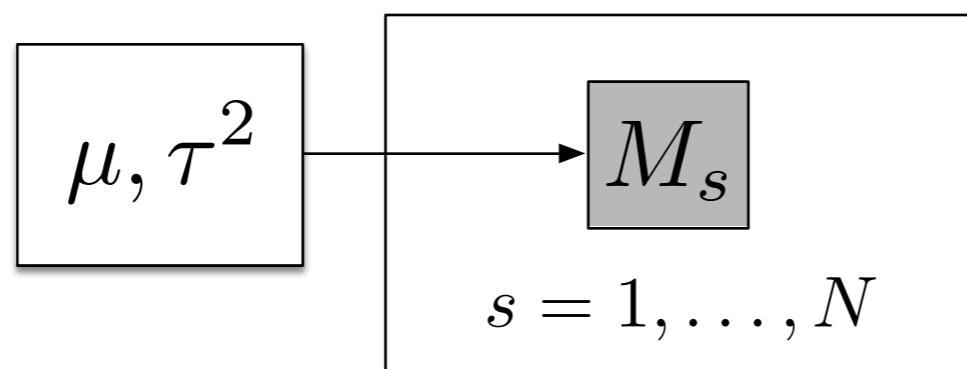
# Gibbs sampling & Hierarchical Bayes



1. For  $s = 1 \dots N$ : Sample Latent Variables Conditional on Data and **Hyperparameters**



2. Sample Hyperparameters from Conditional on Data and Latent Variables:

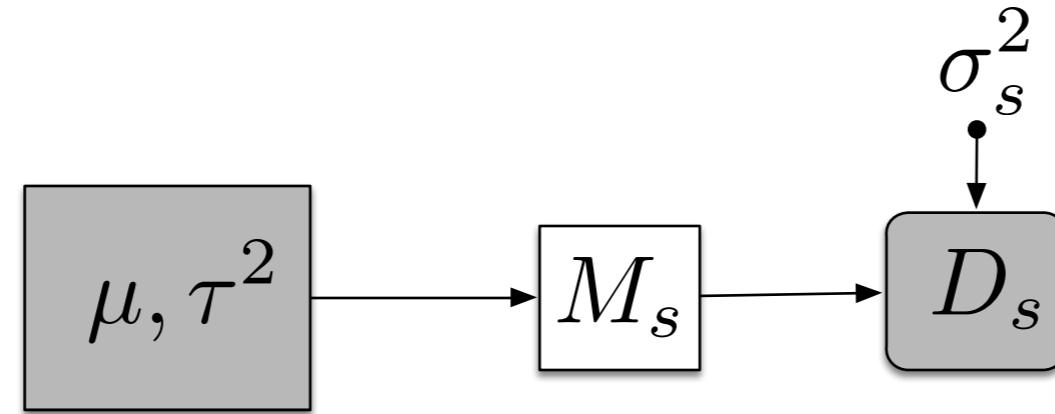


# Gibbs sampling & Hierarchical Bayes

Utilises Conditional Independence structure of PGM to derive conditional posterior densities

$$P(\{M_s\}, \mathbf{H} | \{D_s\}) \propto \left[ \prod_{s=1}^N P(D_s | M_s) P(M_s | M_0, \tau^2) \right] \times P(\mathbf{H})$$

1. For  $s = 1 \dots N$ : Sample Latent Variables Conditional on Data and **Hyperparameters**



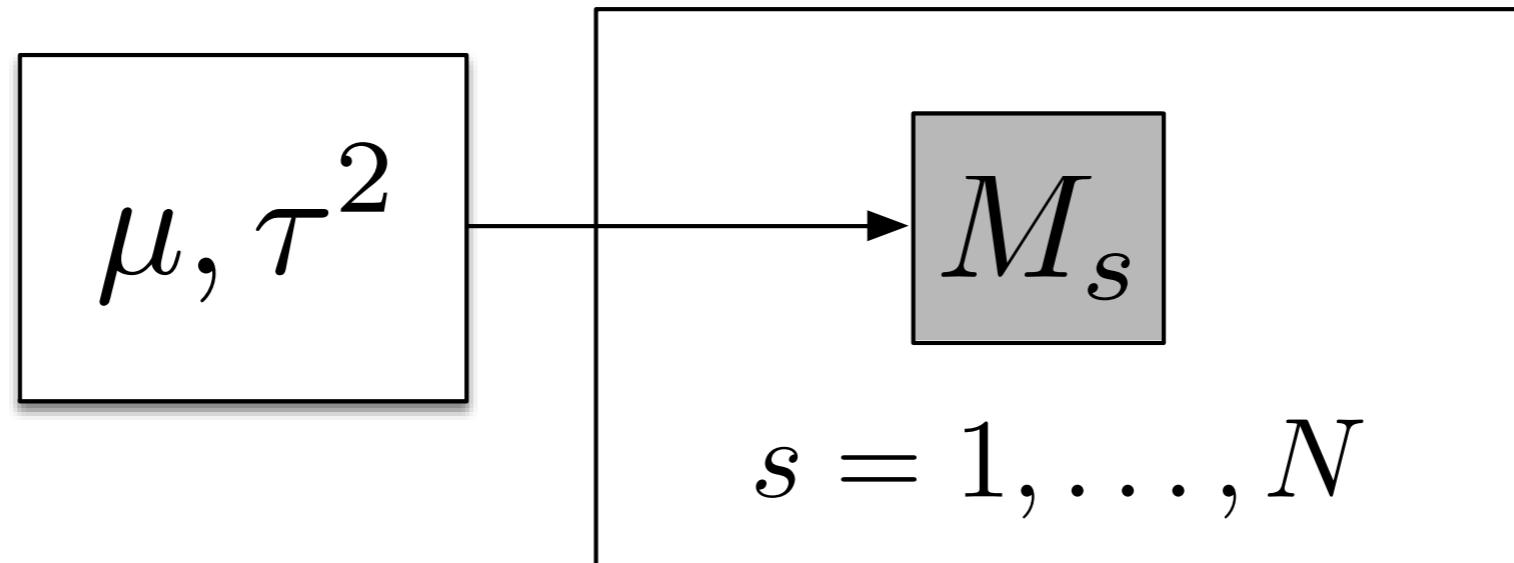
$$\begin{aligned} P(M_s | \mu, \tau^2, D_S) &\propto P(D_s | M_s) \times P(M_s | M_0, \tau^2) \\ &\propto N(D_s | M_s, \sigma^2) \times N(M_s | M_0, \tau^2) \end{aligned}$$

# Gibbs sampling & Hierarchical Bayes

Utilises Conditional Independence structure of PGM to derive conditional posterior densities

$$P(\{M_s\}, \mathbf{H} | \{D_s\}) \propto \left[ \prod_{s=1}^N P(D_s | M_s) P(M_s | M_0, \tau^2) \right] \times P(\mathbf{H})$$

2. Sample Hyperparameters from Conditional on Data and Latent Variables:



Hyperprior:  $P(\mu, \tau^2) \propto (\tau^2)^{-1/2}$

$$P(M_0, \tau^2 | \{M_s\}; \{D_s\}) = P(M_0, \tau^2 | \{M_s\}) = P(M_0 | \tau^2, \{M_s\}) P(\tau^2 | \{M_s\})$$

(See Example Sheet 1, Problem 4)

(Gaussian)

(Inv- $\chi^2$ )

# Gibbs sampling & Hierarchical Bayes

Utilises Conditional Independence structure of PGM to derive conditional posterior densities

$$P(\{M_s\}, \mathbf{H} | \{D_s\}) \propto \left[ \prod_{s=1}^N P(D_s | M_s) P(M_s | M_0, \tau^2) \right] \times P(\mathbf{H})$$

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Hyperprior:  $P(\mu, \tau^2) \propto (\tau^2)^{-1/2}$  (Gaussian) (Inv- $\chi^2$ )

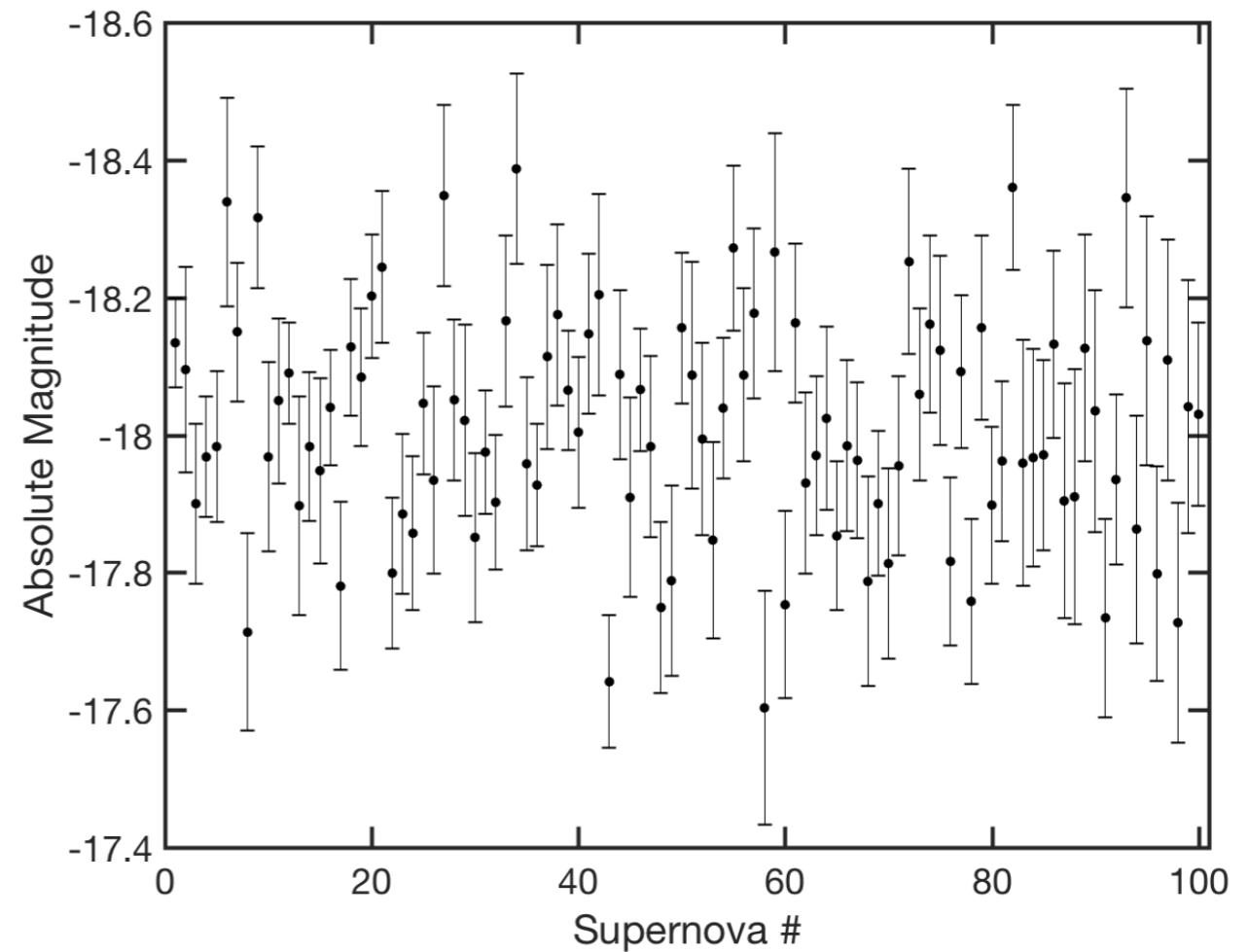
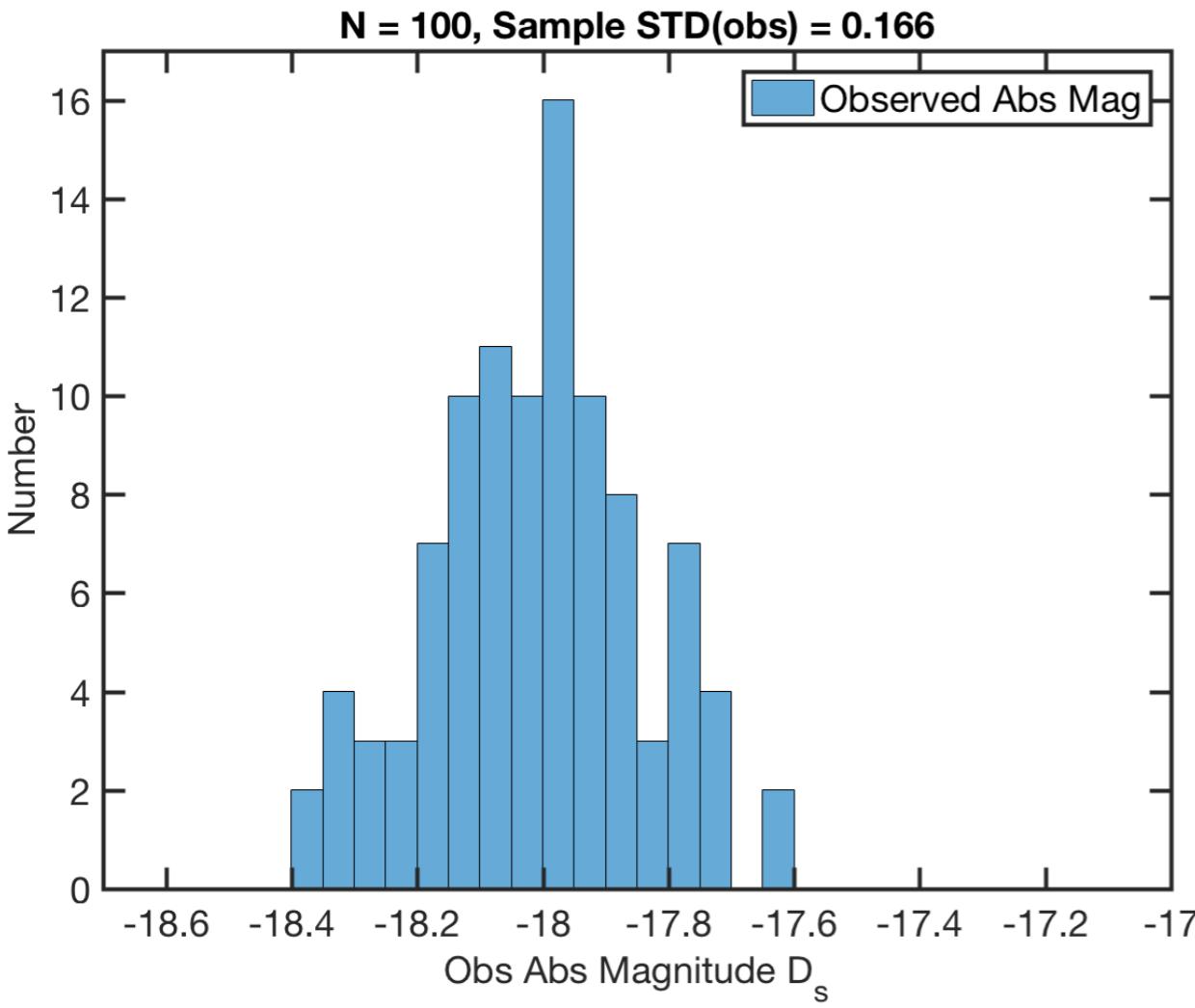
Conditional Posteriors:  $\tau^2 | \{M_s\} \sim \text{Inv-}\chi^2 \left( N - 2, \frac{(N - 1)}{(N - 2)} s^2 \right)$

$$M_0 | \tau^2; \{M_s\} \sim N(\bar{M}, \tau^2/N)$$

$$\bar{M} \equiv \frac{1}{N} \sum_{s=1}^N M_s \quad s^2 \equiv \frac{1}{N-1} \sum_{s=1}^N (M_s - \bar{M})^2$$

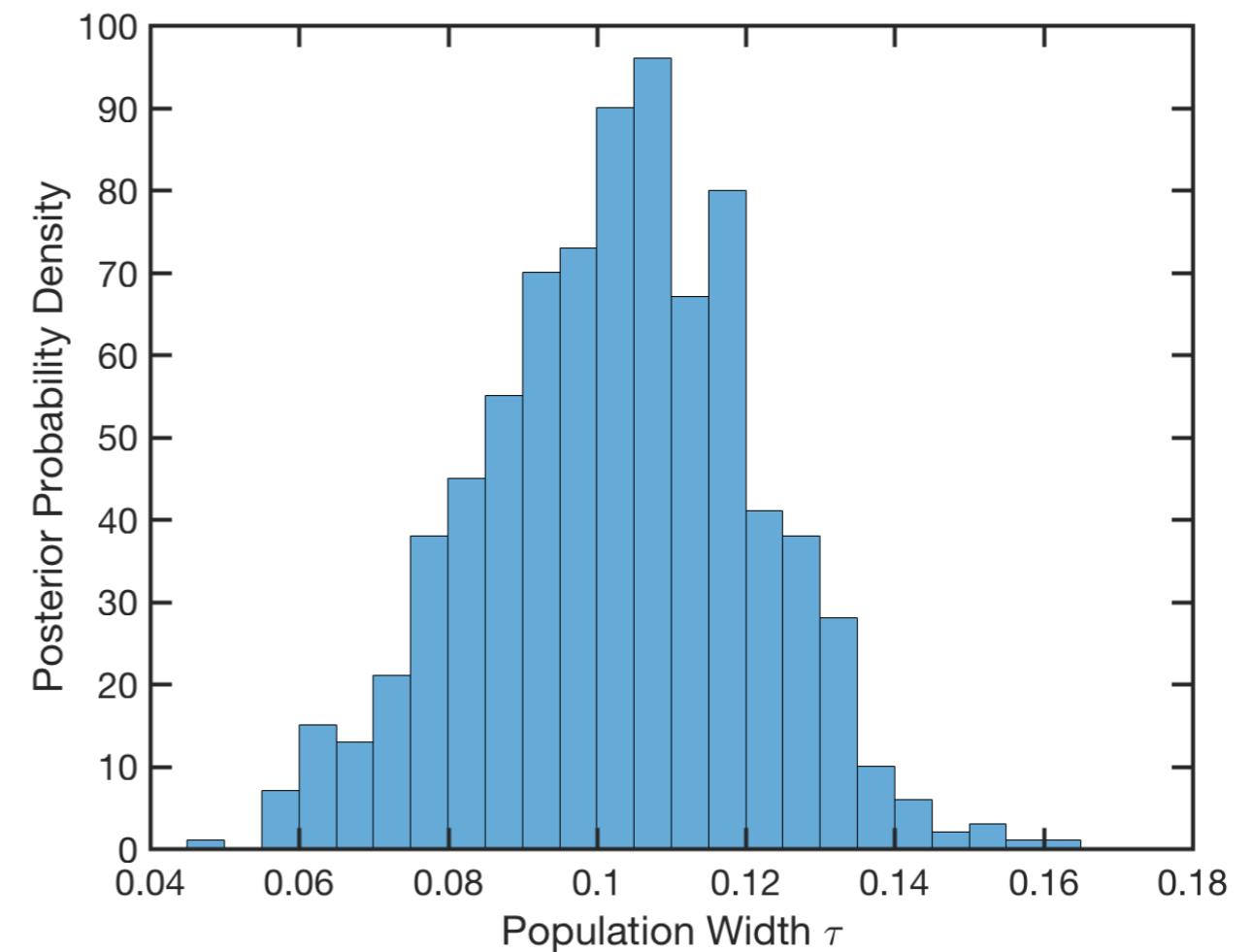
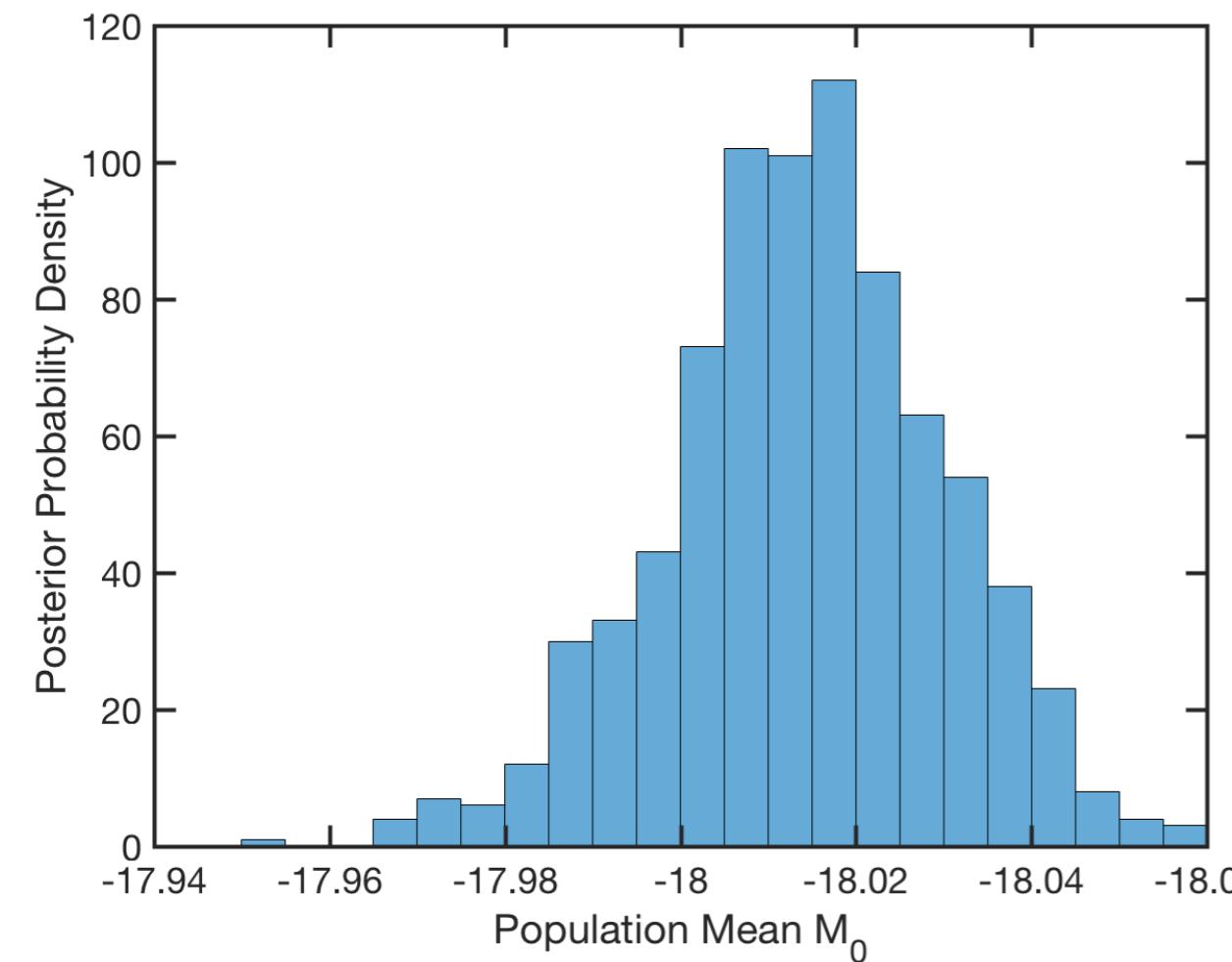
# Code/Data Demo

100 Supernovae with  
app mag and  
distance measurements



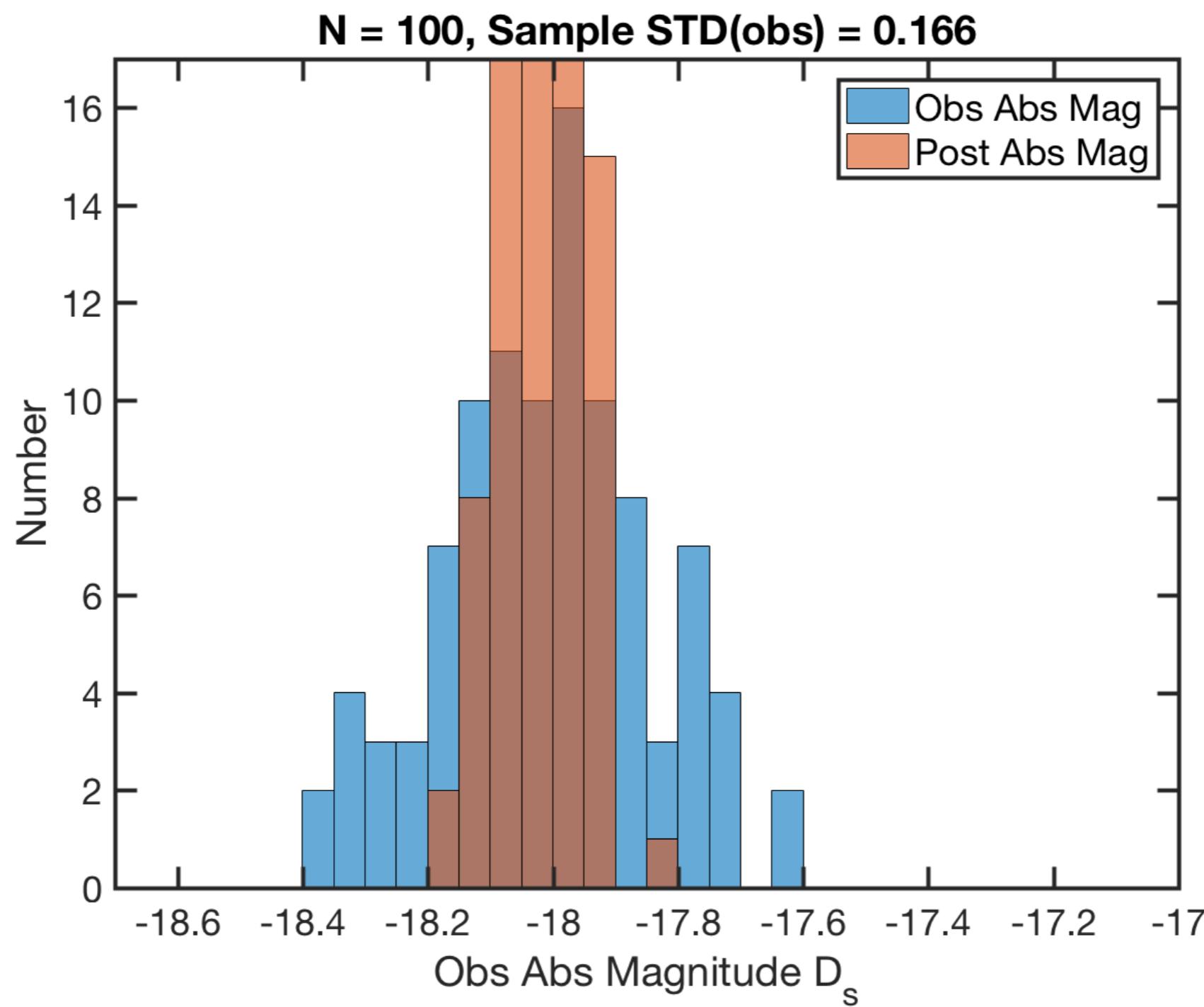
# Hyperparameter Inference

Marginal Posterior Histogram Estimates  
from Gibbs Sampling in 102 dimensions



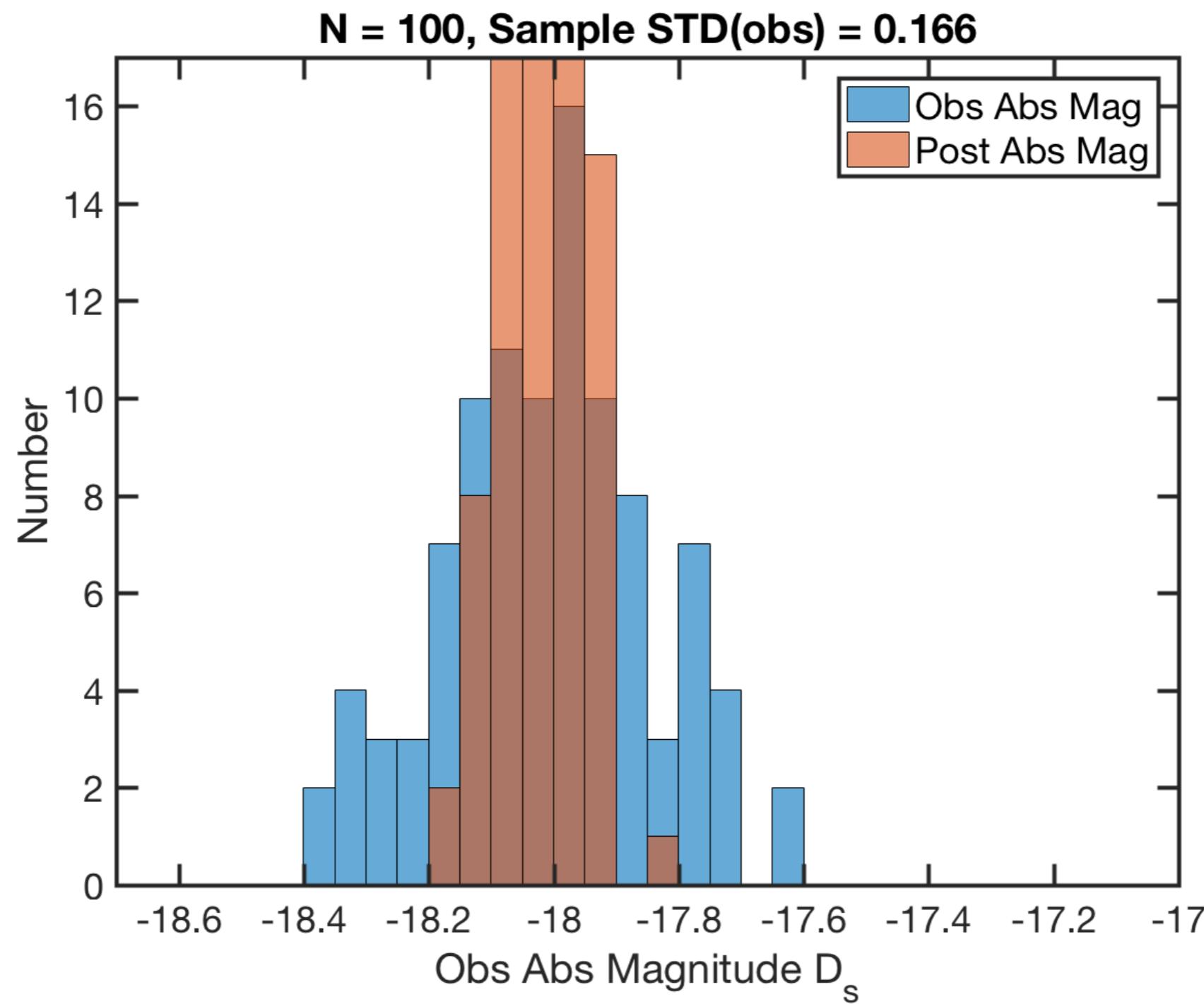
# Latent Variable Inference

Histogram of  $\mathbb{E}[M_s | \mathcal{D}]$  Estimates from MCMC



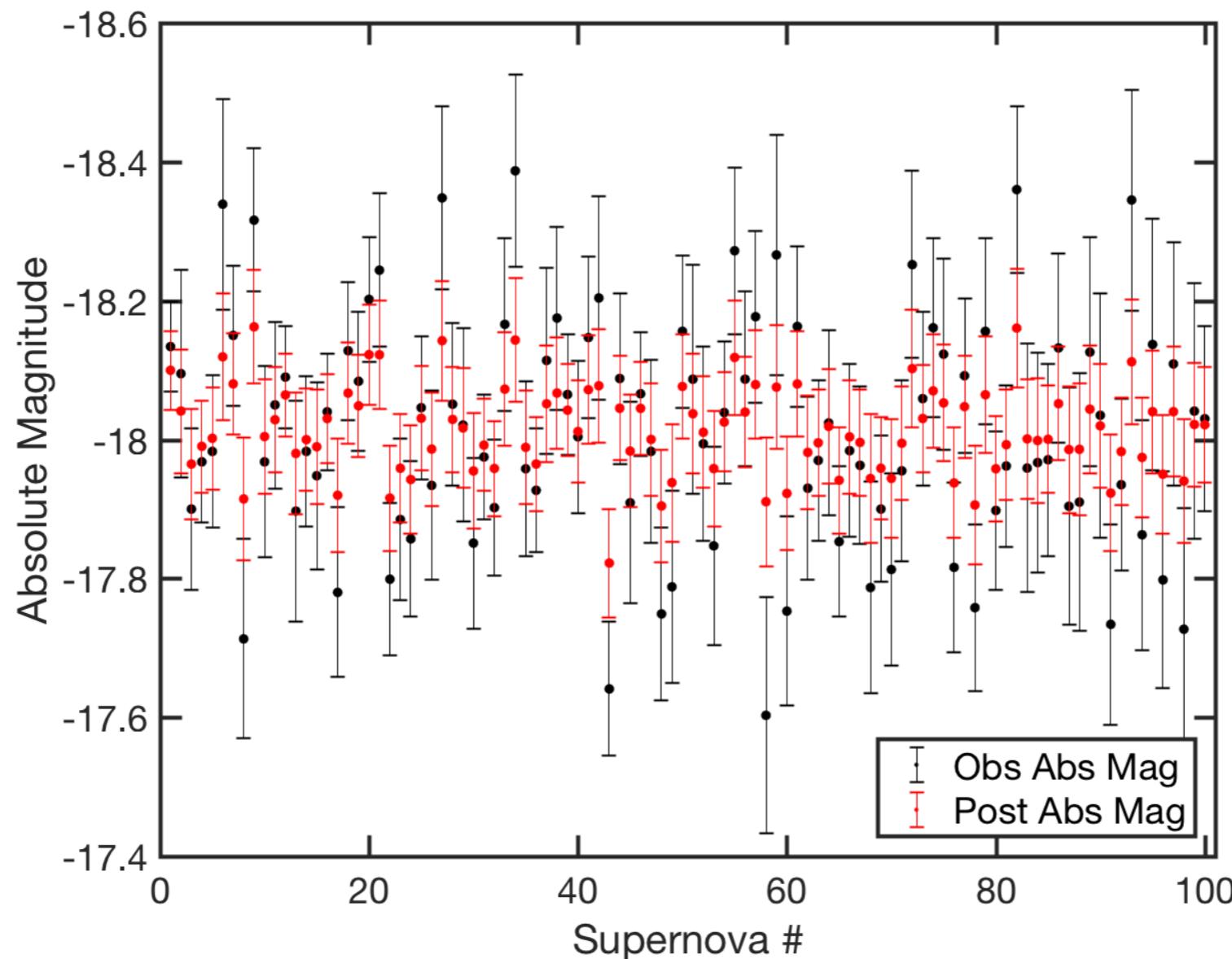
# Latent Variable(s) Inference

Histogram of  $\mathbb{E}[M_s | D]$  posterior estimates from MCMC



# Latent Variable(s) Inference

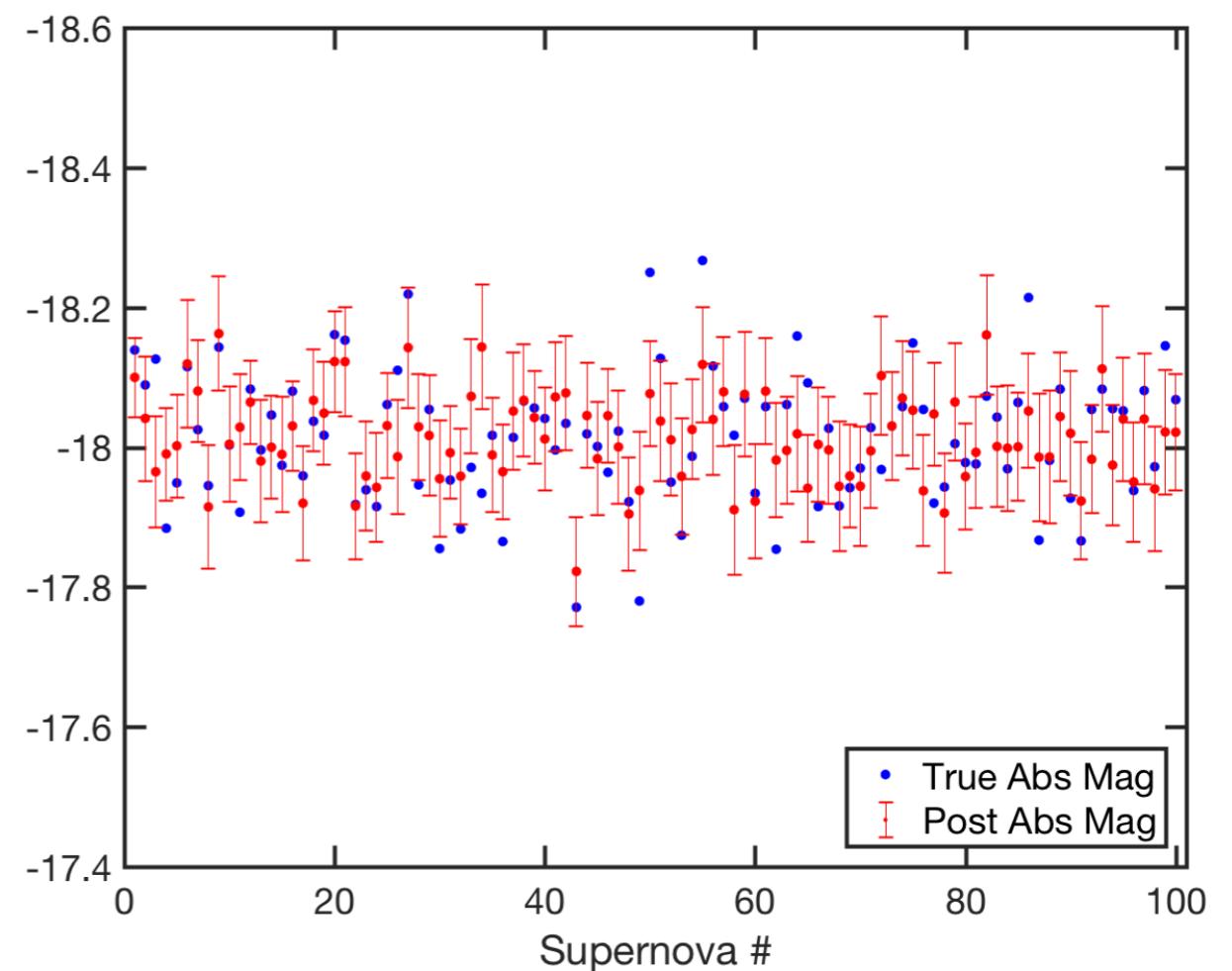
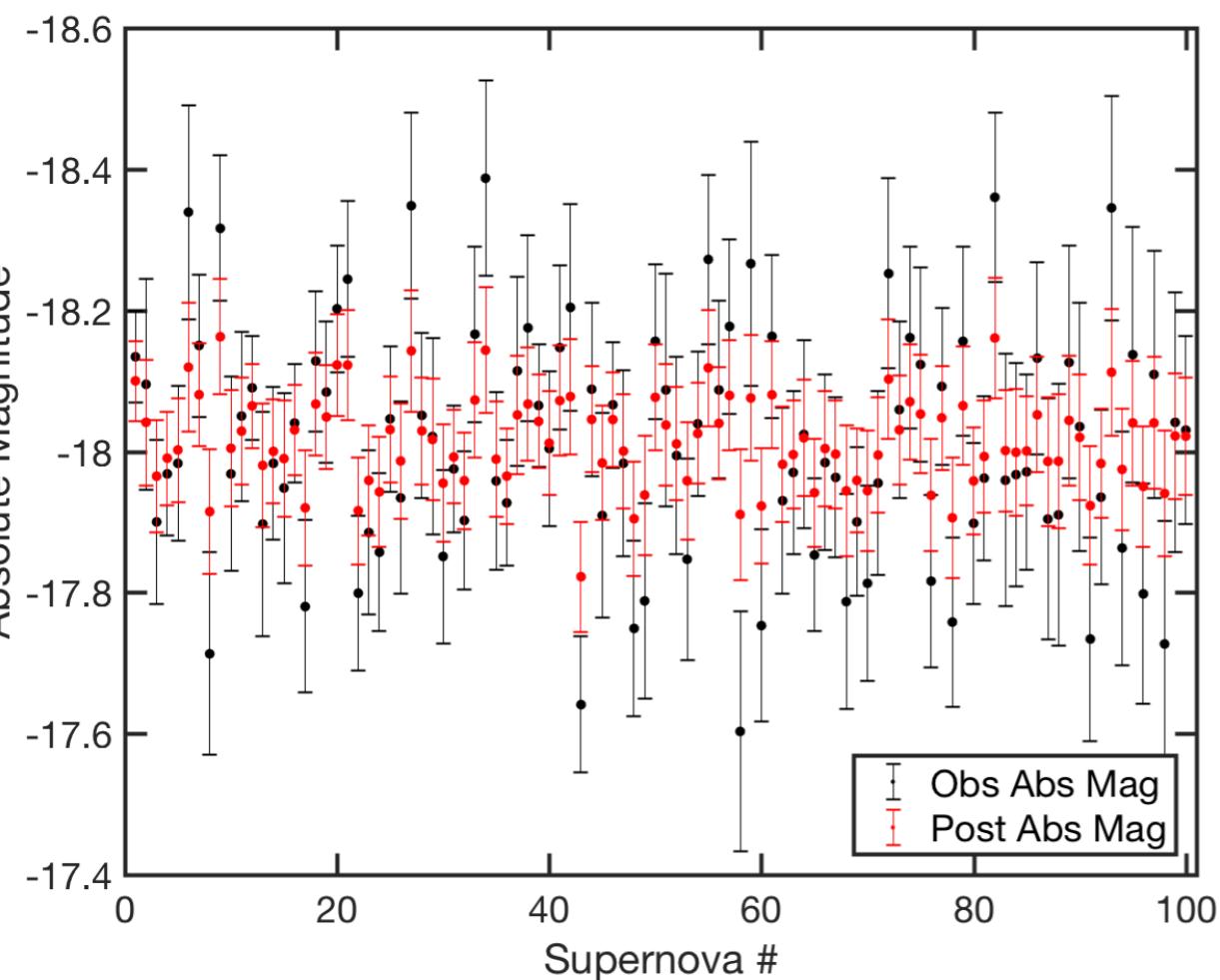
Histogram of  $\mathbb{E}[M_s | D]$  posterior estimates from MCMC



What's going on here?

# Latent Variable(s) Inference

Comparing the Data, Posterior Latent Mags vs. true Mags



What's going on here?

# HB models: Partial Pooling, Shrinkage, and “Borrowing of Strength”

- Common Sense Procedure:

- Analyze each individual object's data  $D_i$  separately and get each individual  $\text{MLE}_i$  estimate (with error)
- “Plug-in” all  $\{\text{MLE}_i\}$  to estimate population hyperparameters



## SHRINKAGE

Sometime it hides like a frightened turtle

- **Problem:** Each individual  $\theta_i$  estimate may be unbiased but collectively give a biased estimate of population (e.g. variance).

- **Solution:** Use HB to model and infer individuals & population simultaneously and get better estimates of both

# Aside on Shrinkage (derivation on board)

# HB models: Partial Pooling, Shrinkage, and “Borrowing of Strength”

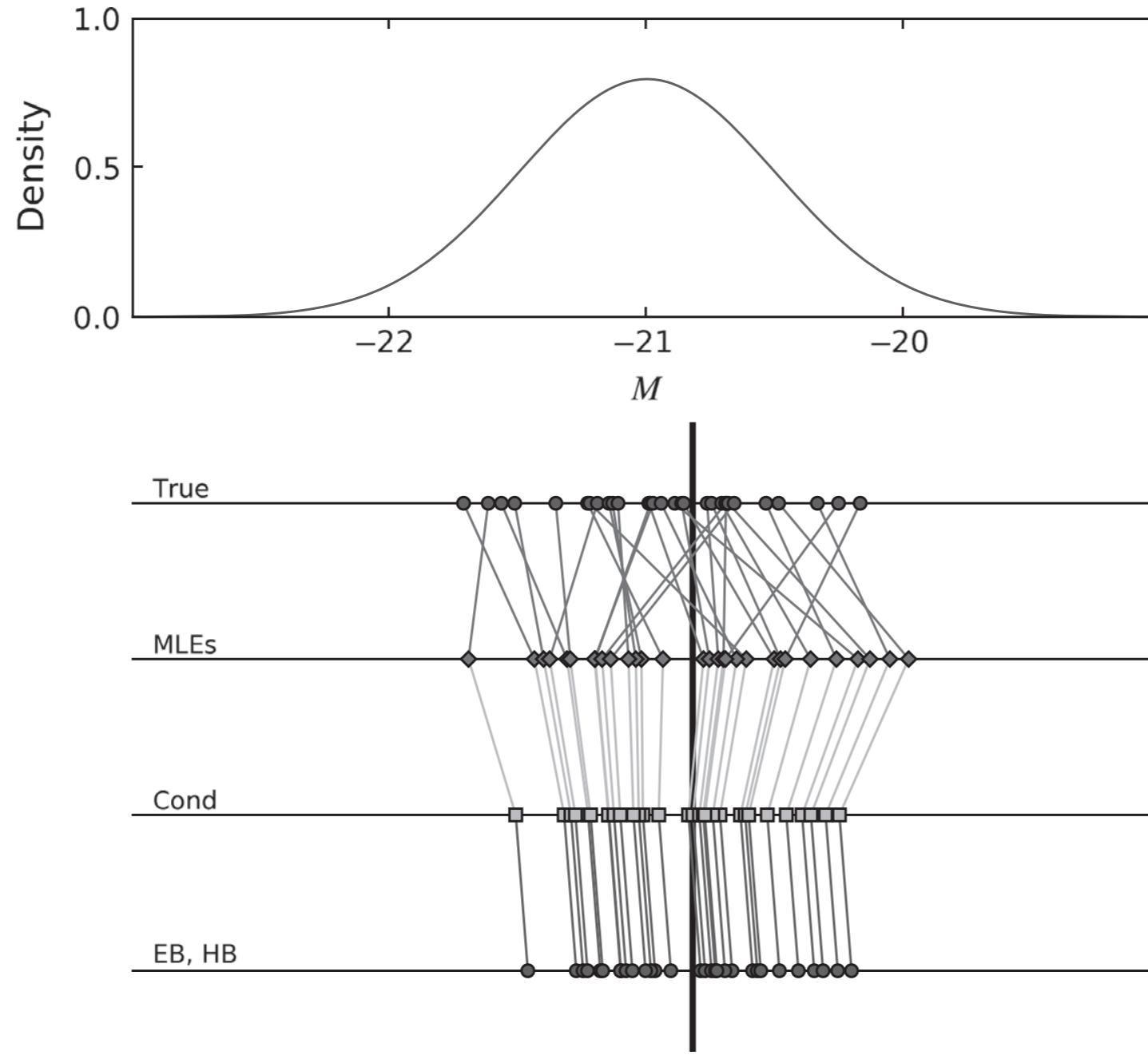
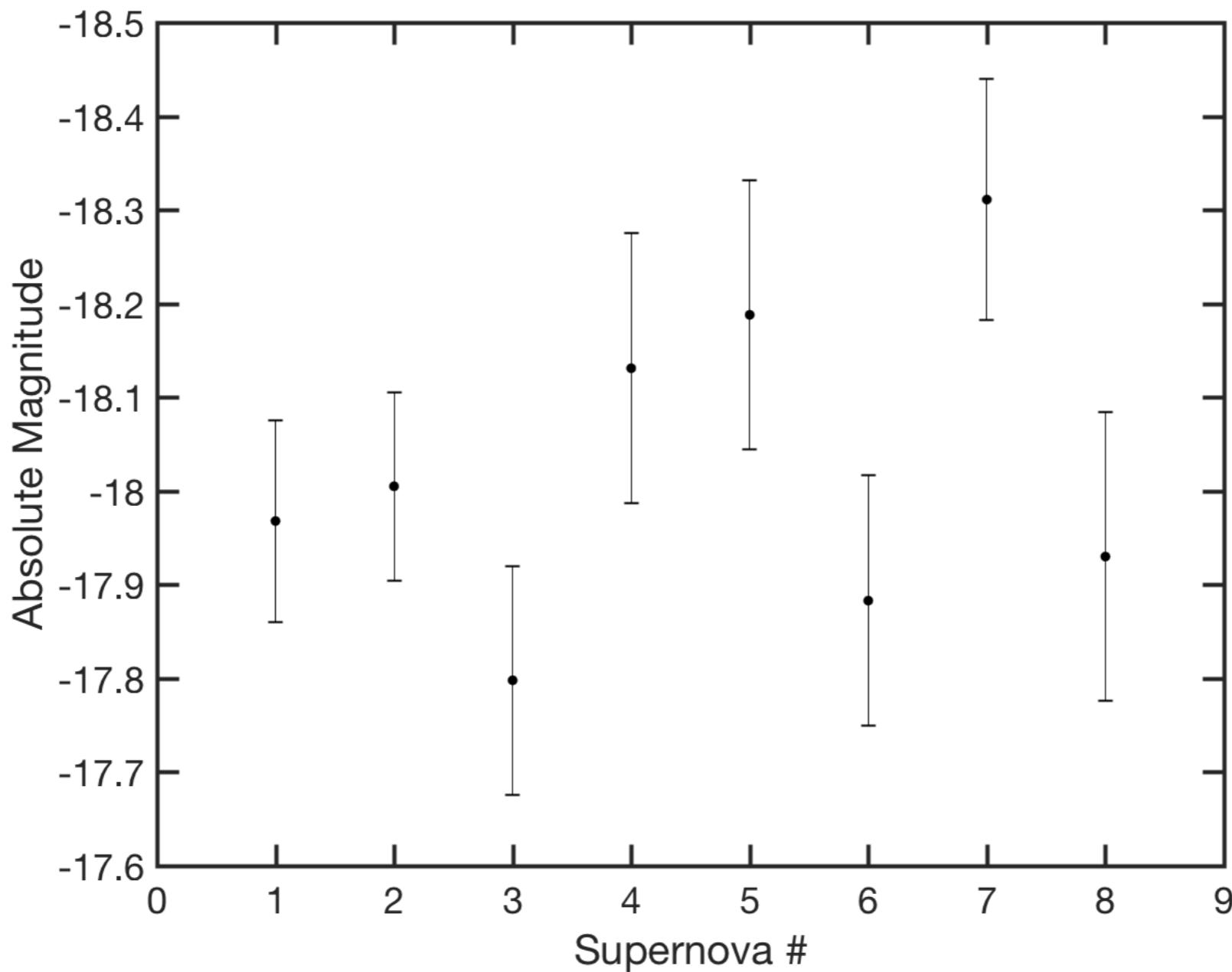


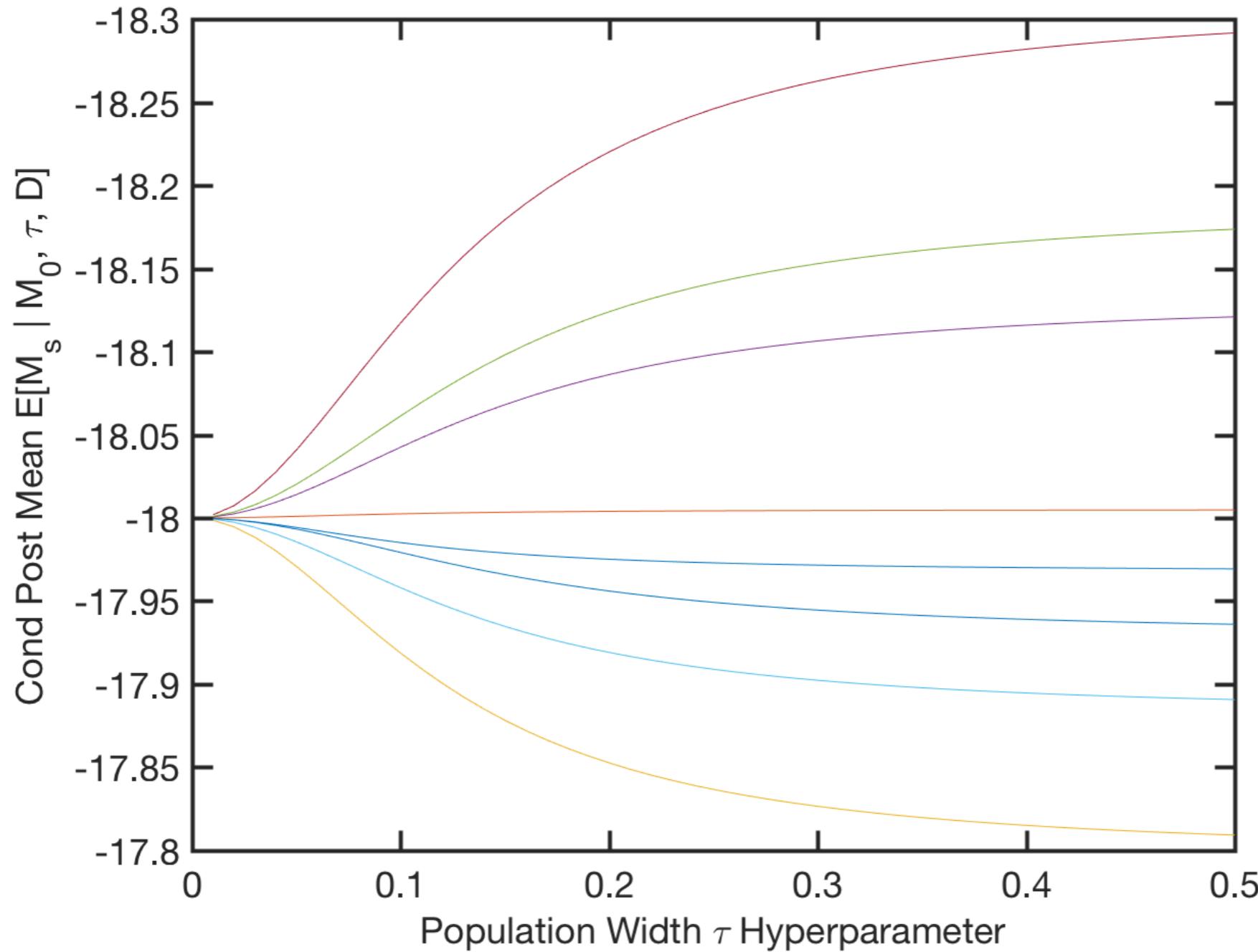
Fig. 11.2. Shrinkage in a simple normal–normal model. Top panel shows population distribution. ‘True’ axis shows  $M_i$  values of 30 samples. Remaining axes show estimates from measurements with  $\sigma = 0.3$  normal error: MLEs, conditional (on the true mean), and empirical/hierarchical Bayes estimates.

# How does Hierarchical Bayes implement shrinkage?

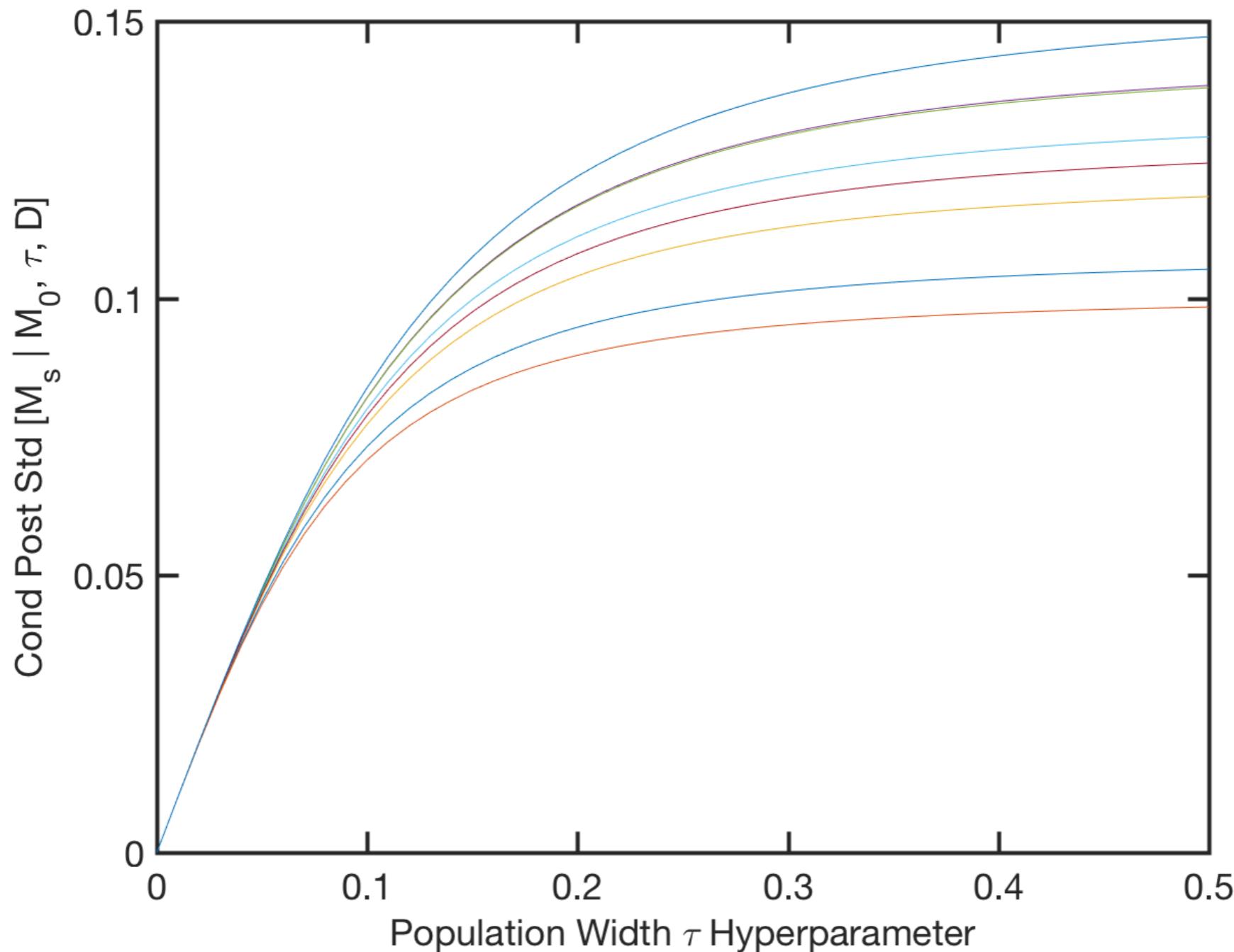
First, let's shrink the dataset



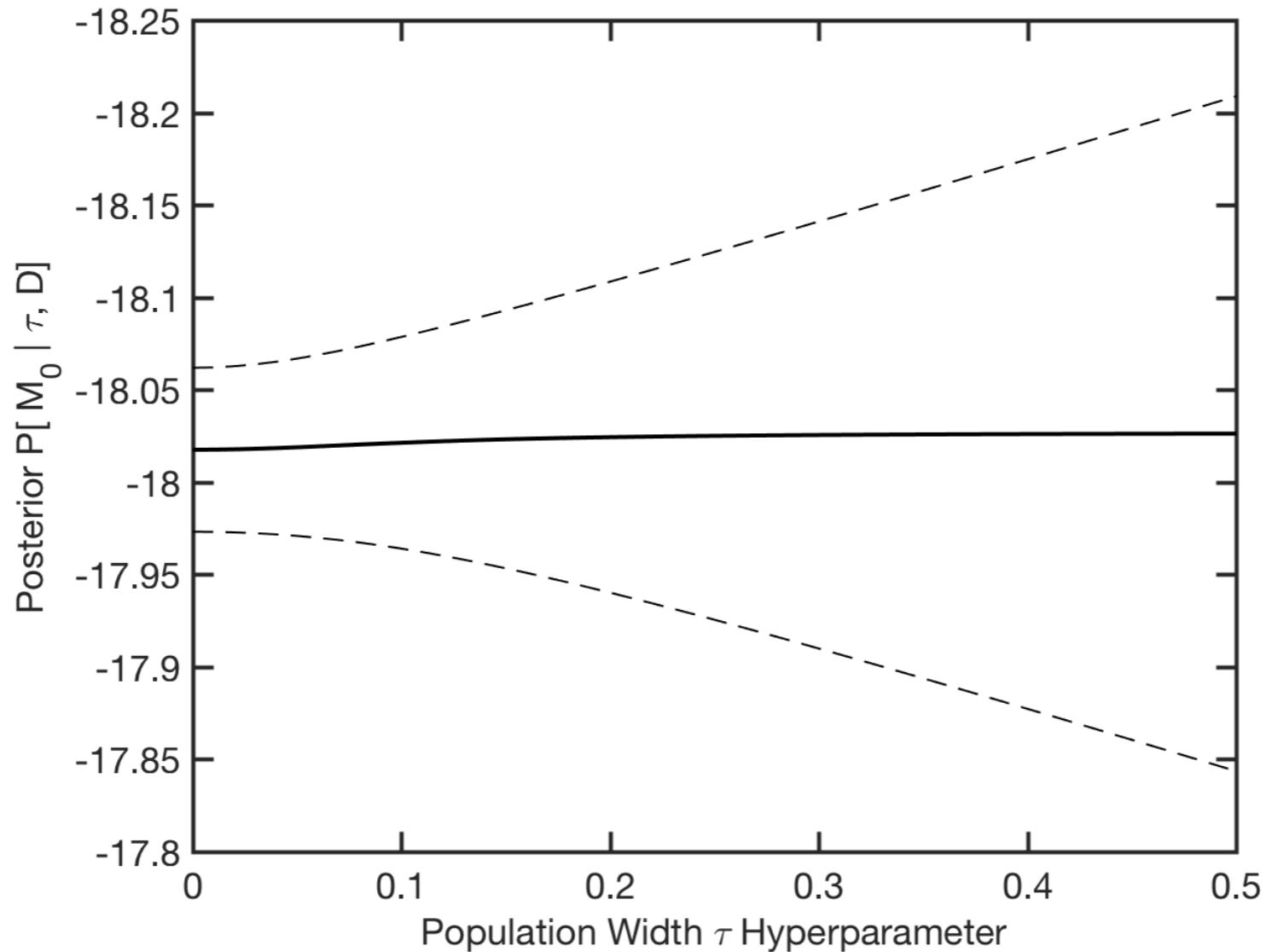
# How does Hierarchical Bayes implement shrinkage?



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