

SLN
(2) 证 Kolmogorov 强大数律.

$$\text{对 } \forall \varepsilon > 0. \text{ 令 } Y_n = \frac{\tilde{S}_n}{n}.$$

$$P(\max_{2^m \leq n < 2^{m+1}} |Y_n| > \varepsilon) \leq P(\max_{2^m \leq n < 2^{m+1}} |\tilde{S}_n| > 2^m \cdot \varepsilon) \leq P(\max_{1 \leq n < 2^{m+1}} |\tilde{S}_n| > 2^m \cdot \varepsilon)$$

$$\leq \frac{1}{(2^m \varepsilon)^2} \sum_{i=1}^{2^{m+1}-1} \text{Var}(X_i)$$

$$\begin{aligned} \sum_{m=1}^{+\infty} P(\max_{2^m \leq n < 2^{m+1}} |Y_n| > \varepsilon) &\leq \sum_{m=1}^{+\infty} \frac{1}{(2^m \varepsilon)^2} \sum_{i=1}^{2^{m+1}-1} \text{Var}(X_i) = \sum_{i=1}^{\infty} \frac{1}{\varepsilon^2} \text{Var}(X_i) \sum_{m=m(i)}^{\infty} \frac{1}{2^{2m}}, m(i) = \min\{m: i < 2^{m+1}\} \\ &= \frac{1}{\varepsilon^2} \sum_{i=1}^{\infty} \text{Var}(X_i) \cdot \frac{1}{2^{2m(i)}} = \frac{4}{\varepsilon^2} \sum_{i=1}^{\infty} \frac{\text{Var}(X_i)}{2^{2m(i)}} = \frac{16}{\varepsilon^2} \sum_{i=1}^{\infty} \frac{\text{Var}(X_i)}{2^{2m(i)+2}} < \infty \end{aligned}$$

hw: 7.4.1, 7.11.17, 7.11.20

$$\text{又 } \sum_{i=1}^{+\infty} \frac{\text{Var}(X_i)}{i^2} < +\infty, 2^{m(i)} \leq i < 2^{m(i)+1} \Rightarrow \sum_{i=1}^{+\infty} \frac{\text{Var}(X_i)}{2^{2m(i)+2}} < \sum_{i=1}^{+\infty} \frac{\text{Var}(X_i)}{i^2} < +\infty$$

$$P(\max_{2^m \leq n < 2^{m+1}} |Y_n| \geq \varepsilon \text{ i.o.}) = 0 \quad P(|Y_n| \geq \varepsilon \text{ i.o.}) = 0 \quad Y_n \xrightarrow{\text{a.s.}} 0$$

$$\text{定理 } \{X_n\} \text{ i.i.d. } S_n = \sum_{k=1}^n X_k \quad \frac{S_n}{n} - \mu \xrightarrow{\text{a.s.}} 0 \Leftrightarrow E[X_n] = \mu$$

$$\text{引理 } X \geq 0. \text{ 则 } E(X^k) = \int_0^{+\infty} k x^{k-1} P(X > x) dx$$

$$\begin{aligned} \text{证: } \int_0^{+\infty} k x^{k-1} P(X > x) dx &= \int_0^{+\infty} k x^{k-1} \int_n^{\infty} I_{\{x > \alpha\}} dp dx = \int_n^{\infty} \int_0^{+\infty} k x^{k-1} I_{\{x > \alpha\}} dx dp \\ &= \int_n^{\infty} \int_0^x k x^{k-1} dx dp = \int_n^{\infty} x^k dp. \end{aligned}$$

$$E|X_n| = \int_0^{+\infty} P(|X_n| > x) dx = \sum_{n=1}^{+\infty} \int_{n-1}^n P(|X_n| > x) dx \geq \sum_{n=1}^{\infty} P(|X_n| > n)$$

$$E|X_n| \leq \sum_{n=1}^{\infty} P(|X_n| > n-1) = \sum_{n=1}^{\infty} P(|X_n| > n) + P(|X_n| \in (n-1, n]) \leq \sum_{n=1}^{\infty} P(|X_n| > n) + 1$$

证明 i.i.d. 情形下 SLN

$$\Rightarrow \frac{X_n}{n} = \frac{S_n - S_{n-1}}{n} \xrightarrow{\text{a.s.}} 0 \quad \text{对 } \forall \varepsilon > 0, \sum_{n=1}^{\infty} P(|\frac{X_n}{n}| > \varepsilon) < \infty$$

$$\text{取 } \varepsilon = 1, \sum_{n=1}^{\infty} P(|X_n| > n) < \infty \quad E[X_n] \text{ 存在 } E[\frac{S_n}{n}] = E[X_1] = \mu.$$

$$\Leftarrow \text{令 } Y_n = \begin{cases} X_n, & |X_n| \leq n \\ 0, & |X_n| > n \end{cases}$$

$$P(X_n \neq Y_n) = P(|X_n| > n), \sum_{n=1}^{\infty} P(X_n \neq Y_n) = \sum_{n=1}^{\infty} P(|X_n| > n) < \infty$$

$$\Rightarrow P(X_n \neq Y_n \text{ i.o.}) = 0 \quad \text{验证 } \sum_{k=1}^n Y_k \xrightarrow{\text{a.s.}} \mu \quad \{Y_n\} \text{ 独立.}$$

$$\begin{aligned} \sum_{n=1}^{\infty} \frac{\text{Var}(Y_n)}{n^2} &\leq \sum_{n=1}^{\infty} \frac{E(Y_n^2)}{n^2} \leq \sum_{n=1}^{\infty} \frac{1}{n^2} \sum_{k=1}^n k^2 P(k-1 \leq |X_n| < k) = \sum_{k=1}^{\infty} k^2 P(k-1 \leq |X_n| < k) \sum_{n=k}^{\infty} \frac{1}{n^2} \\ &\leq \sum_{k=1}^{\infty} k^2 \left(\frac{1}{k^2} + \frac{1}{k} \right) P(k-1 \leq |X_n| < k) \leq 2 + \sum_{k=1}^{\infty} (k-1) P(k-1 \leq |X_n| < k) \\ &\leq 2 + E|X_n| < \infty \end{aligned}$$

$$\Rightarrow \frac{Y_1 + \dots + Y_n}{n} \xrightarrow{\text{a.s.}} \mu \quad \frac{X_1 + \dots + X_n}{n} \xrightarrow{\text{a.s.}} \mu$$

§ 7.4 中心极限定理

$$\{X_n\} \text{ i.i.d. } S_n = X_1 + \dots + X_n \quad \frac{S_n - E(S_n)}{\sqrt{\text{Var}(S_n)}} \xrightarrow{D} N(0, 1)$$

$$\text{Feller CLT} + \lim_{n \rightarrow \infty} \max_{1 \leq k \leq n} \frac{\sigma_k^2}{\sigma_1^2 + \dots + \sigma_n^2} = 0$$

Lindeberg 条件: $\{X_n\}$ 独立, $a_k = E(X_k)$, $b_k^2 = \text{Var}(X_k)$ $B_n^2 = \sum_{k=1}^n b_k^2$ $F_k(x)$ 为 X_k 的分布函数.

$$\begin{aligned} \text{对 } \forall \varepsilon > 0, \lim_{n \rightarrow \infty} \frac{1}{B_n^2} \sum_{k=1}^n \int_{|x-a_k| > \varepsilon B_n} (x-a_k)^2 dF_k(x) &= 0 \\ \Leftrightarrow \lim_{n \rightarrow \infty} \frac{1}{B_n^2} \sum_{k=1}^n E((X_k - a_k)^2 I_{\{|X_k - a_k| > \varepsilon B_n\}}) &= 0 \end{aligned}$$

注记: (1) $\{X_n\}$ 满足 Lindeberg 条件, 则 $\max_{1 \leq k \leq n} \left| \frac{X_k - a_k}{B_n} \right| \xrightarrow{P} 0$
 (2) L 条件成立, 则 Feller 条件成立. $\lim_{n \rightarrow \infty} \max_{1 \leq k \leq n} \frac{b_k^2}{B_n^2} = 0$

$$\begin{aligned} \text{证: (1) } \forall \varepsilon > 0, P\left(\max_{1 \leq k \leq n} \left| \frac{X_k - a_k}{B_n} \right| > \varepsilon\right) &= P\left(\max_{1 \leq k \leq n} |X_k - a_k| > \varepsilon B_n\right) \\ &= P\left(\bigcup_{1 \leq k \leq n} |X_k - a_k| > \varepsilon B_n\right) \leq \sum_{k=1}^n P(|X_k - a_k| > \varepsilon B_n) \\ &\leq \sum_{k=1}^n E\left[\frac{(X_k - a_k)^2}{\varepsilon^2 B_n^2} I_{\{|X_k - a_k| > \varepsilon B_n\}}\right] \rightarrow 0 \end{aligned}$$

$$\begin{aligned} \text{(2) } \max_{1 \leq k \leq n} \frac{b_k^2}{B_n^2} &= \frac{1}{B_n^2} \max_{1 \leq k \leq n} E((X_k - a_k)^2) \\ &= \frac{1}{B_n^2} \max_{1 \leq k \leq n} (E((X_k - a_k)^2 \cdot I_{\{|X_k - a_k| > \varepsilon B_n\}}) \\ &\quad + E((X_k - a_k)^2 I_{\{|X_k - a_k| \leq \varepsilon B_n\}})) \end{aligned}$$

$$\leq \frac{1}{B_n^2} \sum_{k=1}^n E((X_k - a_k)^2 I_{\{|X_k - a_k| > \varepsilon B_n\}}) + \varepsilon^2 \rightarrow 0 \quad (\text{令 } \varepsilon \rightarrow 0)$$

定理 Lindeberg-Feller CLT

$\{X_k\}$ 独立随机变量列 $E[X_k] = a_k$, $\text{Var}(X_k) = b_k^2$, $B_n^2 = \sum_{k=1}^n b_k^2$, $S_n = X_1 + \dots + X_n$.

满足 Lindeberg 条件, 则 $\frac{S_n - E[S_n]}{B_n} \xrightarrow{D} N(0, 1)$

证: 令 $X_{nk} = \frac{X_k - a_k}{B_n}$, $k=1, 2, \dots, n$

$E(X_{nk}) = 0$, $E(X_{nk}^2) = \frac{b_k^2}{B_n^2}$, X_{nk} 的特征函数为 $\varphi_{nk}(t)$.

$\frac{S_n - E[S_n]}{B_n}$ 特征函数记为 $\varphi_n(t) = \prod_{k=1}^n \varphi_{nk}(t)$. 下面证明 $\lim_{n \rightarrow \infty} \prod_{k=1}^n \varphi_{nk}(t) = e^{-\frac{t^2}{2}}$

$$\varphi_{nk}(t) = E[e^{itX_{nk}}] = 1 + \frac{i^2 t^2}{2} E[X_{nk}^2] + r_{nk}(t)$$

tips: X 的特征函数 $\varphi(t) = E[e^{itx}]$

$$(1) |e^{it} - 1 - \frac{it}{1} - \frac{(it)^2}{2!}| \leq \frac{|t|^3}{3!} \quad (2) |e^{it} - 1 - it - \frac{(it)^2}{2!}| \leq |e^{it} - 1 - it| + \frac{t^2}{2} \leq \frac{t^2}{2} + \frac{t^2}{2} = t^2$$

$$|r_{nk}(t)| = |E(e^{itX_{nk}} - (1 + itX_{nk} + \frac{i^2 t^2}{2} X_{nk}^2))|$$

$$\leq E(|tX_{nk}|^2 \wedge |t^3 X_{nk}^3|)$$

$$\leq E(|t^3 X_{nk}^3| \cdot I_{\{|X_{nk}| \leq \varepsilon\}} + |t^2 X_{nk}^2| \cdot I_{\{|X_{nk}| > \varepsilon\}})$$

$$\leq |t^3| \cdot \varepsilon E(X_{nk}^2) + t^2 E[X_{nk}^2 \cdot I_{\{|X_{nk}| > \varepsilon\}}]$$

tips: $|a_k| \leq 1$, $|b_k| \leq 1$, $|a_1 \dots a_n - b_1 \dots b_n| \leq \sum_{k=1}^n |a_k - b_k|$

$$|\prod_{k=1}^n \varphi_{nk}(t) - \prod_{k=1}^n (1 - \frac{t^2 b_k^2}{2 B_n^2})| \leq \sum_{k=1}^n |\varphi_{nk}(t) - (1 + \frac{t^2 b_k^2}{2 B_n^2})| = \sum_{k=1}^n |r_{nk}| \leq \sum_{k=1}^n \frac{b_k^2}{B_n^2} + t^2 \sum_{k=1}^n E[X_{nk}^2 I_{\{|X_{nk}| > \varepsilon\}}]$$

$$= \varepsilon \cdot |t|^3 + t^2 \cdot \frac{1}{B_n^2} \cdot E[(X_k - a_k)^2 I_{\{|X_k - a_k| > \varepsilon B_n\}}] \rightarrow 0 \quad \text{令 } \varepsilon \rightarrow 0$$

$$\lim_{n \rightarrow \infty} \prod_{k=1}^n \varphi_{nk}(t) = \lim_{n \rightarrow \infty} \prod_{k=1}^n (1 - \frac{t^2 b_k^2}{2 B_n^2})$$

tips: $e^{x(1-x)} \leq 1+x \leq e^x$

$$\prod_{k=1}^n (1 - \frac{t^2 b_k^2}{2 B_n^2}) \leq \prod_{k=1}^n e^{-\frac{t^2 b_k^2}{2 B_n^2}} = e^{-\frac{t^2}{2}} \cdot \prod_{k=1}^n (1 - \frac{t^2 b_k^2}{2 B_n^2}) \geq \prod_{k=1}^n (e^{-\frac{t^2 b_k^2}{2 B_n^2} + \frac{t^4 b_k^4}{4 B_n^4}})$$

$$= e^{-\frac{t^2}{2}} \cdot \exp(-\frac{t^4}{4} \sum_{k=1}^n (\frac{b_k^2}{B_n^2})^2)$$

$$= e^{-\frac{t^2}{2}} \exp(-\max_{1 \leq k \leq n} \frac{b_k^2}{B_n^2} \cdot \frac{t^4}{4} \cdot \sum_{k=1}^n \frac{b_k^2}{B_n^2})$$

$$\rightarrow e^{-\frac{t^2}{2}}$$

定理 Lyapunov CLT $\{X_n\}$ 独立. 若 $\exists \delta > 0$, s.t. $\frac{1}{B_n^{2+\delta}} \sum_{k=1}^n E(|X_k - a_k|^{2+\delta}) \rightarrow 0$

则 $\frac{S_n - E[S_n]}{B_n} \xrightarrow{D} N(0, 1)$