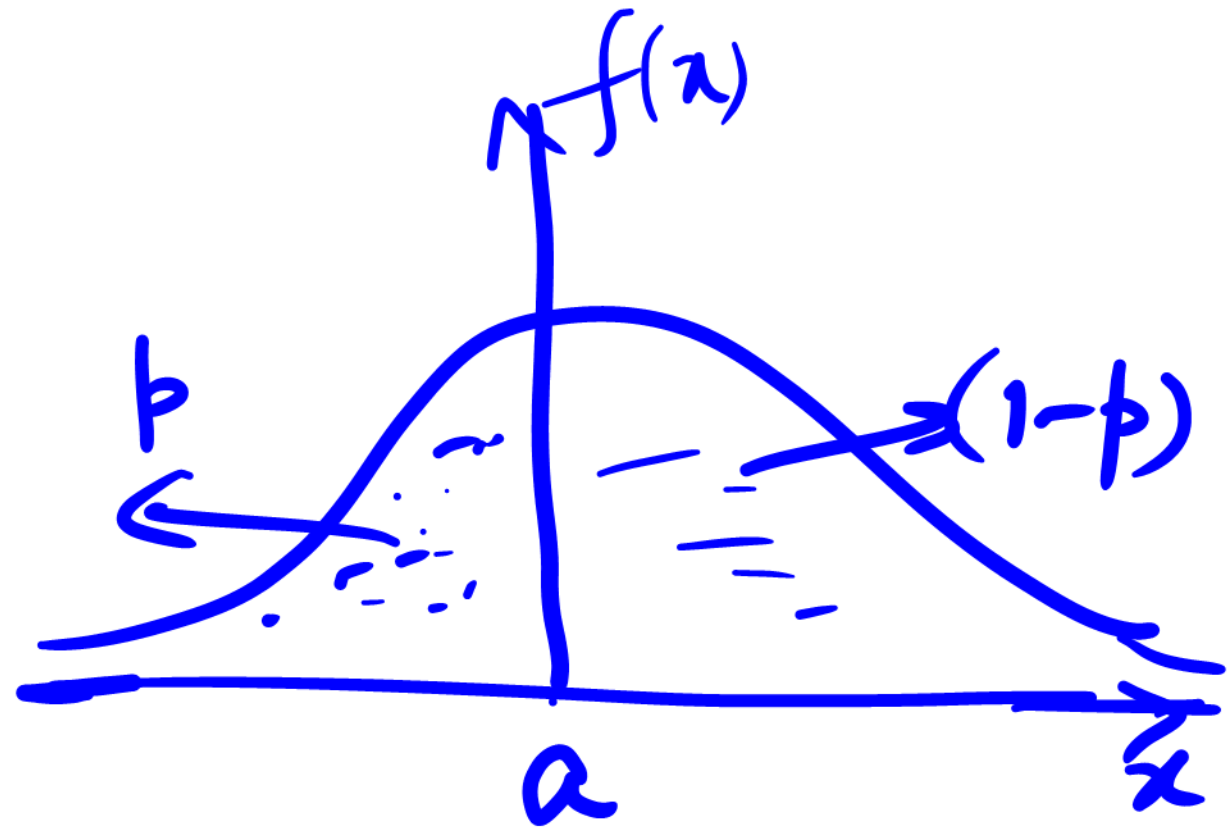


# Quantiles :

$$P(X \leq a) = p$$



A real number  $Q_p$  satisfying

$$P(X \leq Q_p) \geq p \quad \text{and}$$

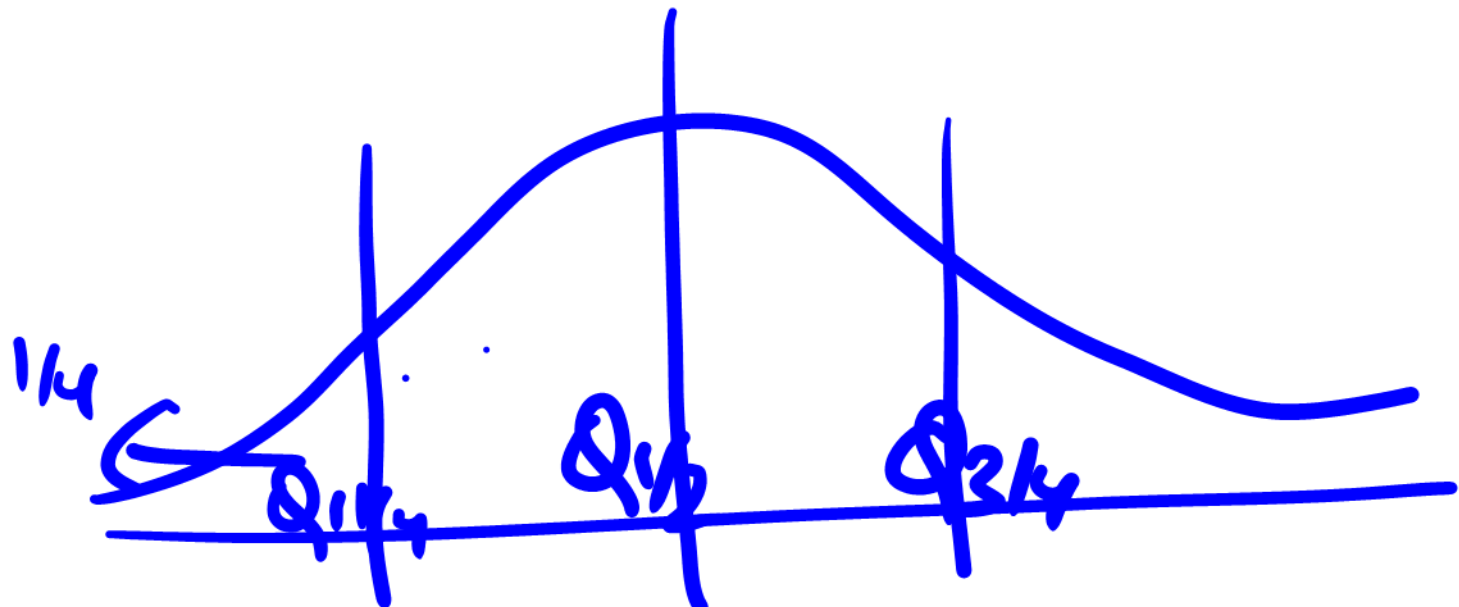
$$P(X \geq Q_p) \geq 1-p, \quad 0 < p < 1$$

is called  $p^{\text{th}}$  quantile (or quantile of order  $p$ ) of distribution of  $X$ .

If  $F$  is absolutely continuous then  $F(Q_p) = p$  i.e. there is a unique quantile.

For  $p = \frac{1}{2}$ ,  $Q_{\frac{1}{2}} = M$  is called the median.

$Q_{\frac{1}{4}}$ ,  $Q_{\frac{1}{2}}$ ,  $Q_{\frac{3}{4}}$  are called quantiles

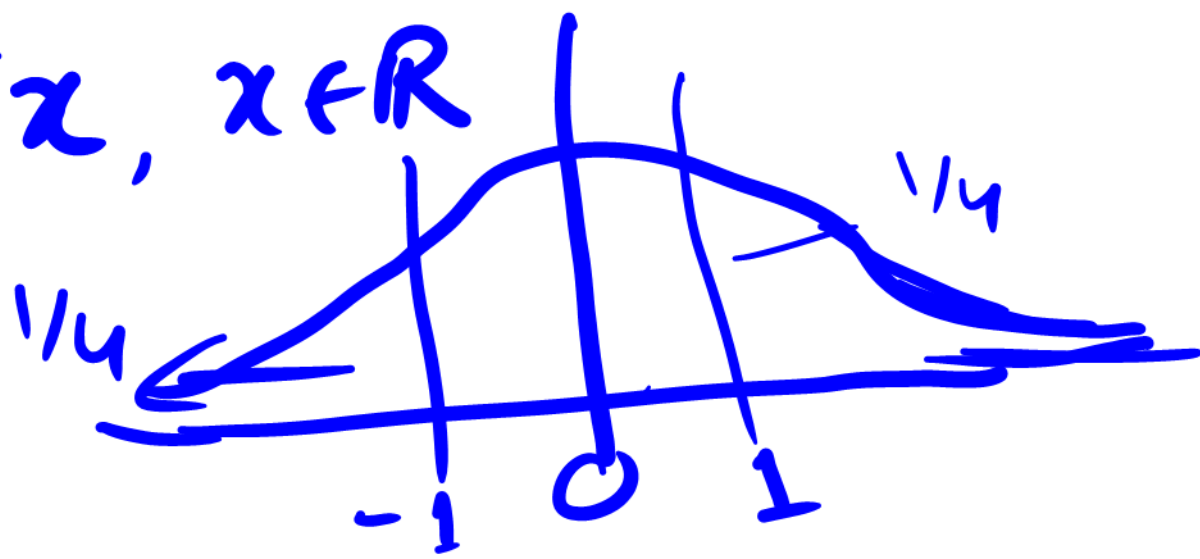


$Q_{1/10}, Q_{2/10}, \dots, Q_{9/10}$  Deciles

$Q_{1/100}, Q_{2/100}, \dots, Q_{99/100}$   
 $\rightarrow$  percentiles

Examples:  $f_X(x) = \frac{1}{\pi(1+x^2)}$ ,  
Cauchy dist<sup>n</sup>  
 $-\infty < x < \infty$

$$F(x) = \frac{1}{2} + \frac{1}{\pi} \tan^{-1} x, \quad x \in \mathbb{R}$$



$$F(0) = \frac{1}{2}$$

So  $M=0$  is the median

$$E(X) = \int_{-\infty}^{\infty} \frac{1}{\pi} \frac{x}{(1+x^2)} dx$$

does not exist.

$$Q_{1/4} = -1$$

$$\text{and } Q_{3/4} = 1$$

as

$$F(-1) = 1/4,$$

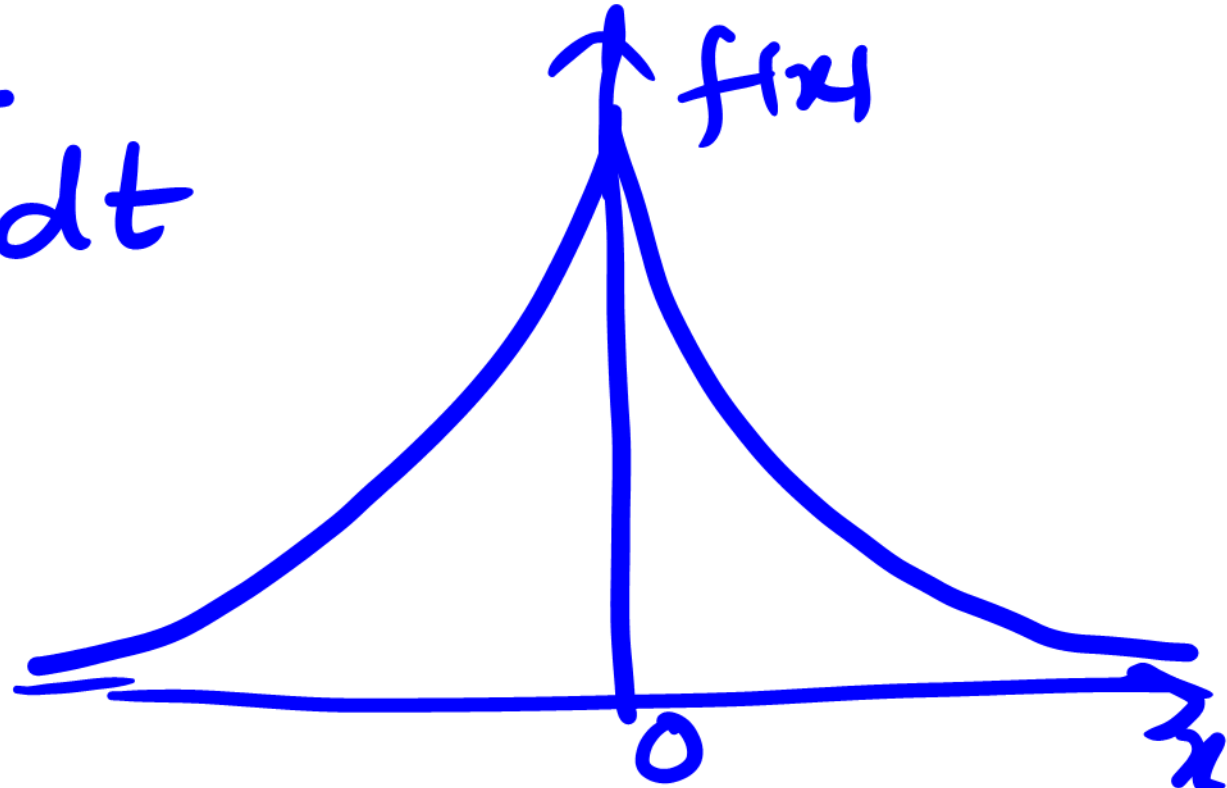
$$F(1) = 3/4.$$

$$2: f_x(x) = \frac{1}{2} e^{-|x|}, \quad -\infty < x < \infty$$

double exponential or Laplace dist<sup>n</sup>

$$F_X(x) = \int_{-\infty}^x \frac{1}{2} e^t dt$$

$$= \frac{1}{2} e^x, \quad x < 0$$



$$= \frac{1}{2} + \int_0^x \frac{1}{2} e^{-t} dt = \frac{1}{2} - \frac{1}{2} e^{-t} \Big|_0^x$$

$$= 1 - \frac{1}{2} e^{-x}, \quad x > 0$$



$$F_X(x) = \begin{cases} \frac{1}{2}e^x, & x < 0 \\ 1 - \frac{1}{2}e^{-x}, & x \geq 0 \end{cases}$$

$$F_X(0) = \frac{1}{2} \quad \text{So Median} = 0$$

$$F_X(x) = \frac{1}{4} \Rightarrow \frac{1}{2}e^x = \frac{1}{4}$$

$$x = -\ln_e 2 \Rightarrow Q_{1/4}$$

$$F_X(x) = \frac{3}{4}, 1 - \frac{1}{2} e^{-x} = \frac{3}{4}$$

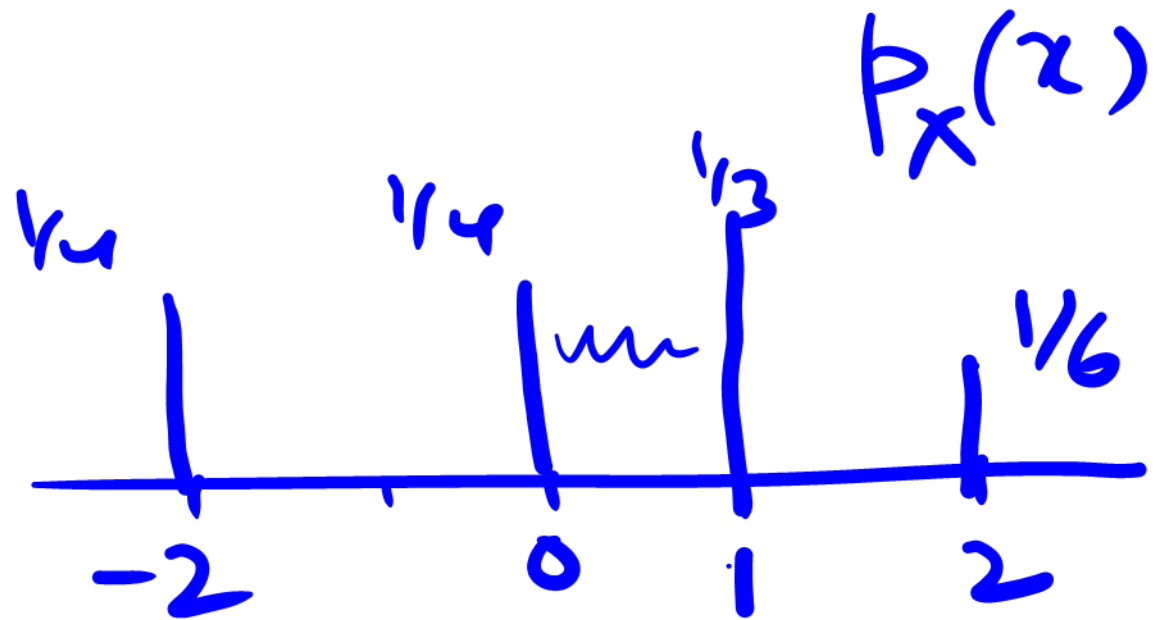
$$\frac{1}{4} = \frac{1}{2} e^{-x} \Rightarrow x = \ln 2 \rightarrow Q_{3/4}$$

Example :  $P(X=-2) = P(X=0) = \frac{1}{4}$   
 $P(X=1) = \frac{1}{3}, P(X=2) = \frac{1}{6}$

$$0 \leq M \leq 1$$

$\downarrow$   
 median  
 first quartile

$$-2 \leq Q_{1/4} \leq 0$$



Moment Generating Function

$$M_X(t) = E(e^{tx}), \quad t \in \mathbb{R}$$

is called mgf of  $X$  if it exists

Example:  $f_X(x) = \begin{cases} 2x, & 0 < x < 1 \\ 0, & \text{ew} \end{cases}$

$$M_X(t) = E(e^{tX}) = \int_0^1 e^{tx} f_X(x) dx$$

$$= \int_0^1 2x e^{tx} dx$$

$$= 2 \left[ \left. \frac{x e^{tx}}{t} \right|_0^1 - \left. \frac{e^{tx}}{t^2} \right|_0^1 \right]$$

$$= 2 \left[ \frac{e^t}{t} - \frac{e^t}{t^2} + \frac{1}{t^2} \right]$$

2.  $p_x(0) = \frac{1}{2}, p_x(1) = \frac{1}{2}$

$$M_x(t) = \frac{1}{2} e^{0 \cdot t} + \frac{1}{2} e^{1 \cdot t} = \left( \frac{1 + e^t}{2} \right)$$

$$M_X(t) = E(e^{tX})$$

$$= E \left[ 1 + \frac{tX}{1!} + \frac{t^2 X^2}{2!} + \dots \right]$$

$$= 1 + \frac{t}{1!} \mu_1' + \frac{t^2}{2!} \mu_2' + \dots$$

Coefficient of  $\frac{t^k}{k!}$  in the

expansion of mgf is  $\mu'_k$ .

$$\frac{d}{dt} M_X(t) \Big|_{t=0} = \mu'_1$$

$$\frac{d^2}{dt^2} M_X(t) \Big|_{t=0} = \mu'_2$$

$$\left. \frac{d^k}{dt^k} M_X(t) \right|_{t=0} = \mu'_k$$

Theorem: If the moment of order  $t$  ( $>0$ ) exists then the moment of order  $s$  ( $0 < s < t$ ) exists for a given r.v.  $X$ . If the moment of order  $s$  ( $>0$ ) does not



exist, then the moment of order  $t$  ( $t > s$ ) does not exist.

Theorem: The mgf uniquely determines a cdf.

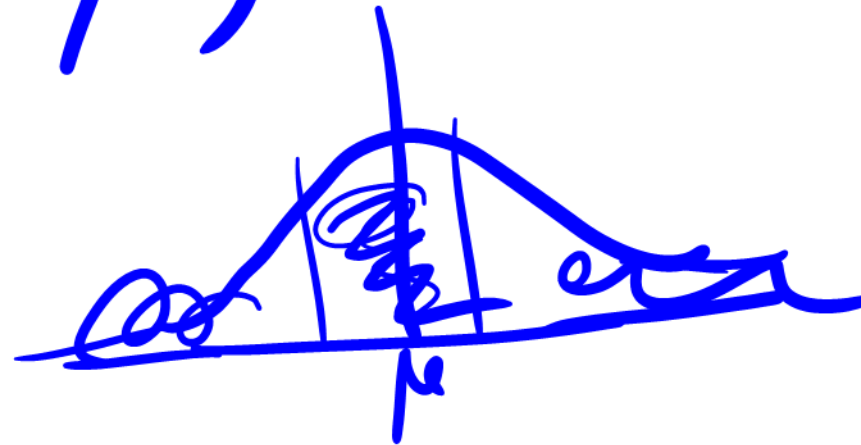
Chebyshev's Inequality: Let  $X$  be a r.v. with mean  $\mu$  and variance  $\sigma^2$ . Then for any  $k > 0$ ,

$$P(|X - \mu| \geq k) \leq \frac{\sigma^2}{k^2}.$$

Pf. Let  $X$  be continuous r.v.  
with pdf  $f_X(x)$ .

$$\sigma^2 = \text{Var}(X) = E(X - \mu)^2$$

$$= \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) dx$$



$$|x - \mu| \leq k \rightarrow \begin{aligned} & -k \leq x - \mu \leq k \\ \Rightarrow & \mu - k \leq x \leq \mu + k \end{aligned}$$

$$\sigma^2 \geq \int (x - \mu)^2 f_x(x) dx$$

$$|x - \mu| \geq k$$

$$\geq k^2 \int_{|x - \mu| \geq k} f_x(x) dx$$

$$= k^2 P(|X - \mu| \geq k)$$

$$\text{or } P(|X - \mu| \geq k) \leq \frac{\sigma^2}{k^2}$$

$$P(|X - \mu| < k) \geq 1 - \frac{\sigma^2}{k^2}$$

Other forms

$$P(|X - \mu| < k\sigma) \geq 1 - \frac{1}{k^2}$$

$$P(|x - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

let us take  $k=1$

$$P(|x - \mu| < \sigma) \geq 0$$

trivial statement

$k=2$

$$P(|x - \mu| < 2\sigma) \geq \frac{3}{4}$$

$$P(\mu - 2\sigma < x < \mu + 2\sigma) \geq 0.75$$

Example: 1. Let  $X$  be no. of costly purchases from a jewellery store in a day. Suppose  $\mu = 18$ ,  $\sigma = 2.5$ .

With what prob. can we assert that there will be between 8 to 28 costly purchases.

$$\begin{aligned} P(8 \leq X \leq 28) &= P(-10 \leq X - 18 \leq 10) \\ &= P(|X - 18| \leq 10) \geq 1 - \frac{\sigma^2}{100} = \frac{15}{16} \end{aligned}$$

2. Independent observations are available from a pop<sup>n</sup> with mean  $\mu$  and variance 1. How many observations are needed in order that prob is at least 0.9 that the mean of observations differs from  $\mu$  by not more than 1?

Suppose observations are  $X_1, X_2 \dots X_n$

$$E(X_i) = \mu, \quad V(X_i) = 1$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i, \quad E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i) \\ = \frac{n\mu}{n} = \mu$$

$$V(\bar{X}) = \frac{1}{n^2} \sum V(X_i) = \frac{n}{n^2} = \frac{1}{n}$$

$$P(|\bar{X} - \mu| < 1) \geq 1 - \frac{1}{n} > 0.9 \\ \Rightarrow n > 10$$

Exercise: Let  $X$  be a contr. v



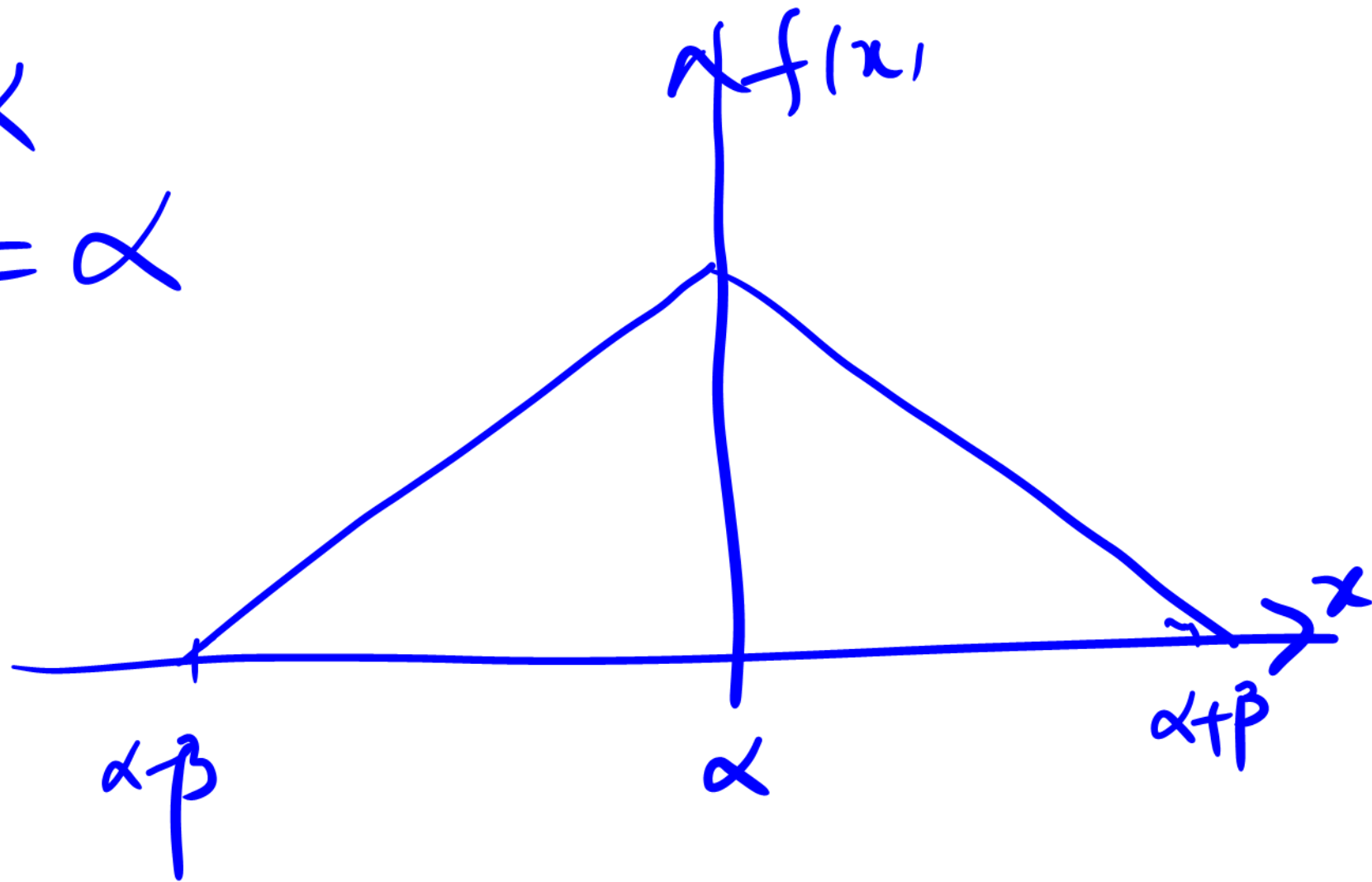
$$f_X(x) = \frac{1}{\beta} \left\{ 1 - \frac{|x - \alpha|}{\beta} \right\},$$

$$\alpha - \beta < x < \alpha + \beta$$

Find  $E(X)$ ,  $V(X)$ ,  $\text{Med}(X)$ ,  
 $Q_{1/4}$ ,  $Q_{3/4}$ , Measures of skewness  
 & kurtosis

$$E(X) = \alpha$$

$$\text{Med}(X) = \alpha$$



$$V(X) = E(X - \mu)^2$$

$$= \int_{\alpha-\beta}^{\alpha+\beta} (x-\alpha)^2 \frac{1}{\beta} \left(1 - \left|\frac{x-\alpha}{\beta}\right|\right) dx$$

$$= \beta^2 \int_{-1}^1 y^2 (1 - |y|) dy \quad \begin{array}{l} \frac{x-\alpha}{\beta} = y \\ \frac{1}{\beta} dx = dy \end{array}$$

$$= \beta^2 / 6$$

$$\mu_4 = E(x-\alpha)^4 = \beta^4 \int_{-1}^1 y^4 (1 - |y|) dy$$

$$Q_1 = \alpha + \beta \left( -\frac{1}{\sqrt{2}} \right) ?$$

Complete these calculations

## Some Special Discrete Distributions

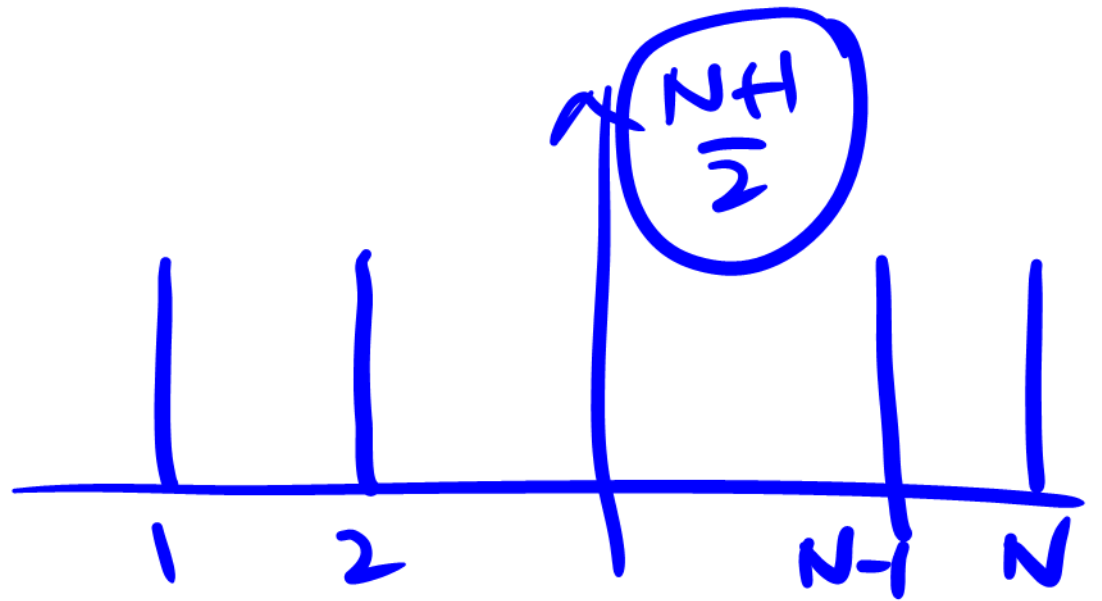
### 1. Discrete Uniform Distribution

$$X \rightarrow 1, 2, \dots, N$$

$$P(X=i) = \frac{1}{N}, \quad i=1, \dots, N$$

$$\mu_1' = E(X) = \sum_{i=1}^N \frac{i}{N} = \frac{N+1}{2}$$

$$\begin{aligned} \mu_2' = E(X^2) &= \sum_{i=1}^N \frac{i^2}{N} \\ &= \frac{(N+1)(2N+1)}{6} \end{aligned}$$



$$V(X) = \mu_2' - \mu_1'^2 = \frac{N^2 - 1}{12}$$

Moments of all orders exist.

MGF.  $M_X(t) = E(e^{tx})$

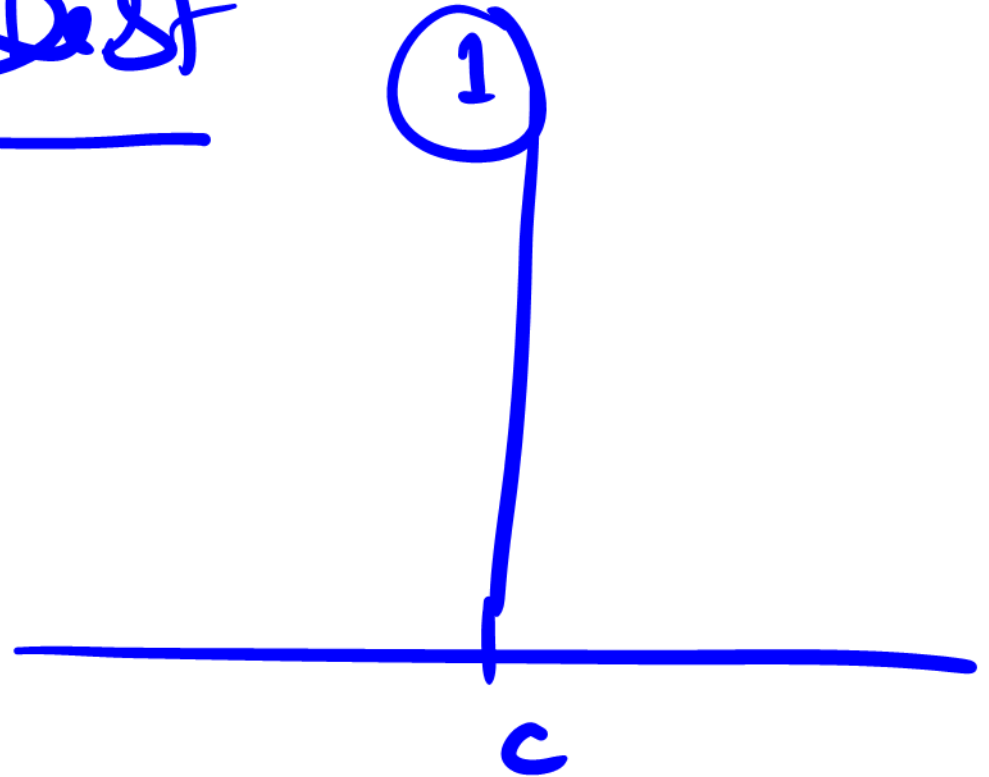
$$= \sum_{i=1}^N \frac{e^{it}}{N} = \begin{cases} \frac{e^t(e^{Nt}-1)}{N(e^t-1)}, & t \neq 0 \\ 1, & t = 0 \end{cases}$$

## 2. Degenerate Dist<sup>n</sup>

$$P(X=c) = 1$$

$$\mu'_k = c^k$$

$$\mu'_1 = c, \mu'_2 = c^2, \dots$$



3. Bernoulli Trials: In a random expt if we have two possible

outcomes we associate them with  
success (s) and failure (f)

$$\Omega = \{s, f\}$$

$X \rightarrow$  no of  
success  
Bernoulli  
Dist

$$X(s) = 1, \quad X(f) = 0$$

$$P(X=1) = p,$$

$$P(X=0) = 1-p$$

(=q)

$$0 < p < 1$$

$$\mu_1' = p, \mu_2' = p, \dots$$

$$\mu_k' = p$$



$$\mu_2 = \mu_2' - \mu_1'^2 = p - p^2 = p(1-p) = pq$$

$$M_X(t) = (1-p)e^{0 \cdot t} + pe^{1 \cdot t} = (q + pet)$$

4. Binomial Dist<sup>n</sup>: Suppose independent

Bernoulli trials are conducted under identical conditions with prob of

Success  $p$ . Let  $X \rightarrow$  no of successes  
in  $n$  trials

$$X \rightarrow 0, 1, 2, \dots, n$$

$$p_X(k) = P(X=k) = \binom{n}{k} p^k q^{n-k}$$

$k=0, 1, \dots, n$   
 $0 < p < 1$

$$\sum_{k=0}^n p_X(k) = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k}$$

$$= (q+p)^n = 1.$$

$$\mu_1' = E(X) = \sum_{k=0}^n k \binom{n}{k} p^k q^{n-k}$$

$$= \sum_{k=1}^n \frac{n!}{(k-1)!(n-k)!} p^k q^{n-k}$$

$$= np \sum_{l=0}^{n-1} \binom{n-1}{l} p^l q^{n-1-l} \quad l = k-1$$

$$= np (q+p)^{n-1} = np.$$

$$E(X(X-1)) = \sum_{k=0}^n k(k-1) \binom{n}{k} p^k q^{n-k}$$

$$= \sum_{k=2}^n \frac{n!}{(k-2)!(n-k)!} p^k q^{n-k} \quad k-2=l$$

$$n(n-1)p^2 \sum_{l=0}^{n-2} \binom{n-2}{l} p^{l-2} q^{n-2-l}$$

$$= n(n-1)p^2$$

$$\begin{aligned} E(X^2) &= E[X(X-1)] + E(X) \\ &= n(n-1)p^2 + np \end{aligned}$$

$$\begin{aligned}
 \mu_2 = V(X) &= \mu_2' - \mu_1'^2 \\
 &= n(n-1)p^2 + np - n^2p^2 \\
 &= np(1-p) = npq.
 \end{aligned}$$

$$\mu_3 = np(1-p)(1-2p),$$

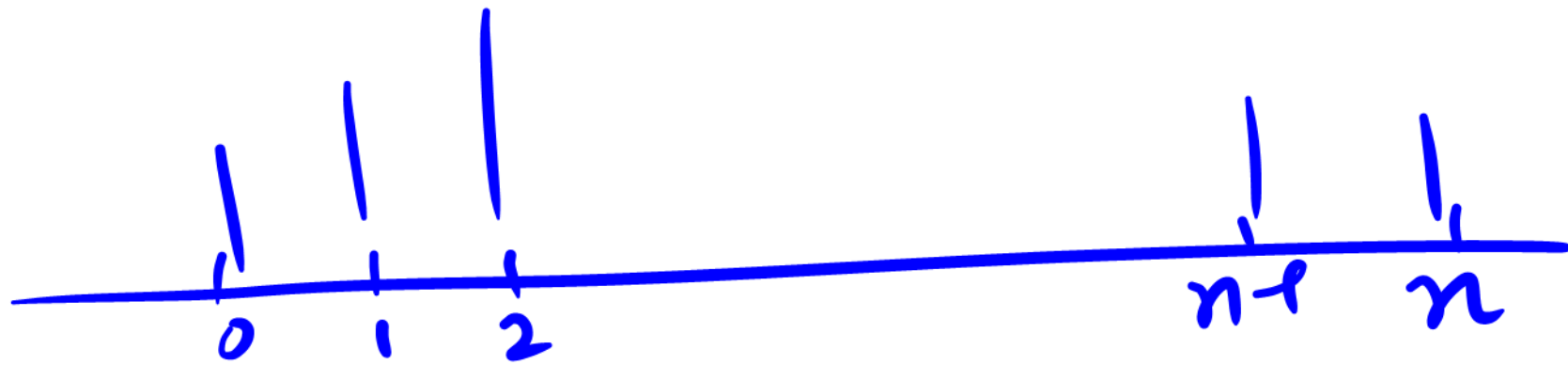
$$\mu_4 = 3(npq)^2 + npq(1-6pq)$$

⊕  
Ex

Measures of skewness & kurtosis

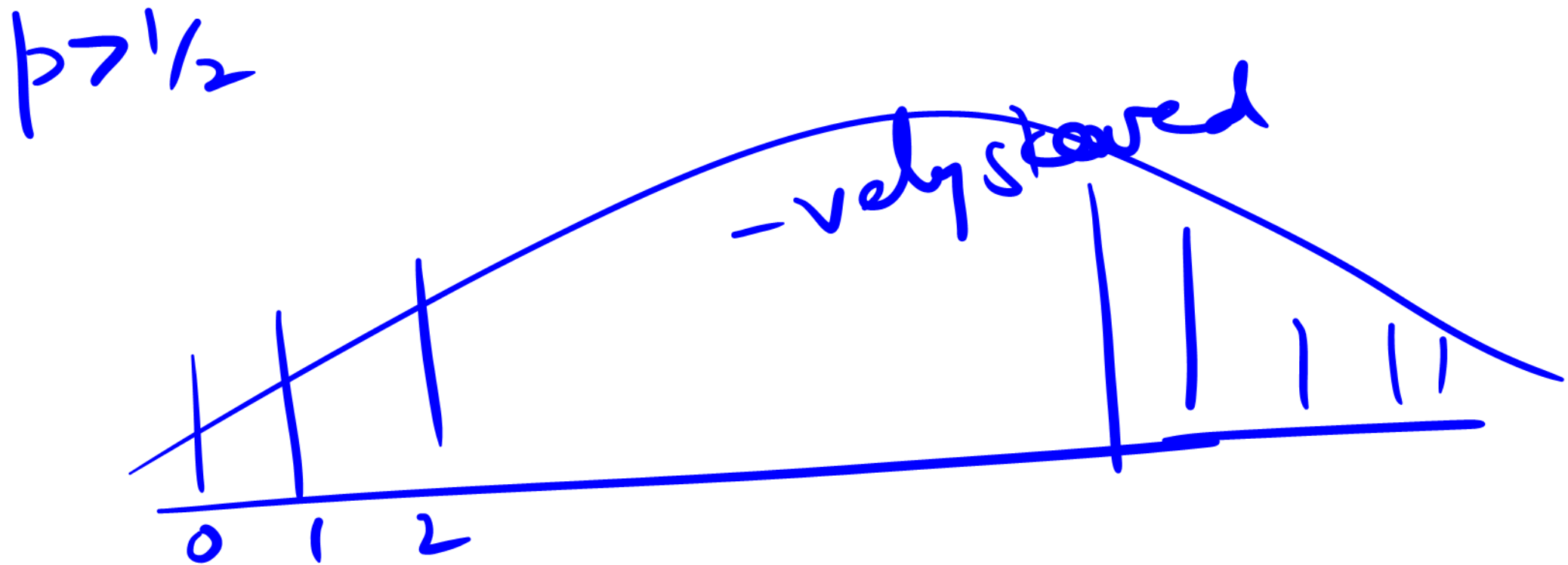
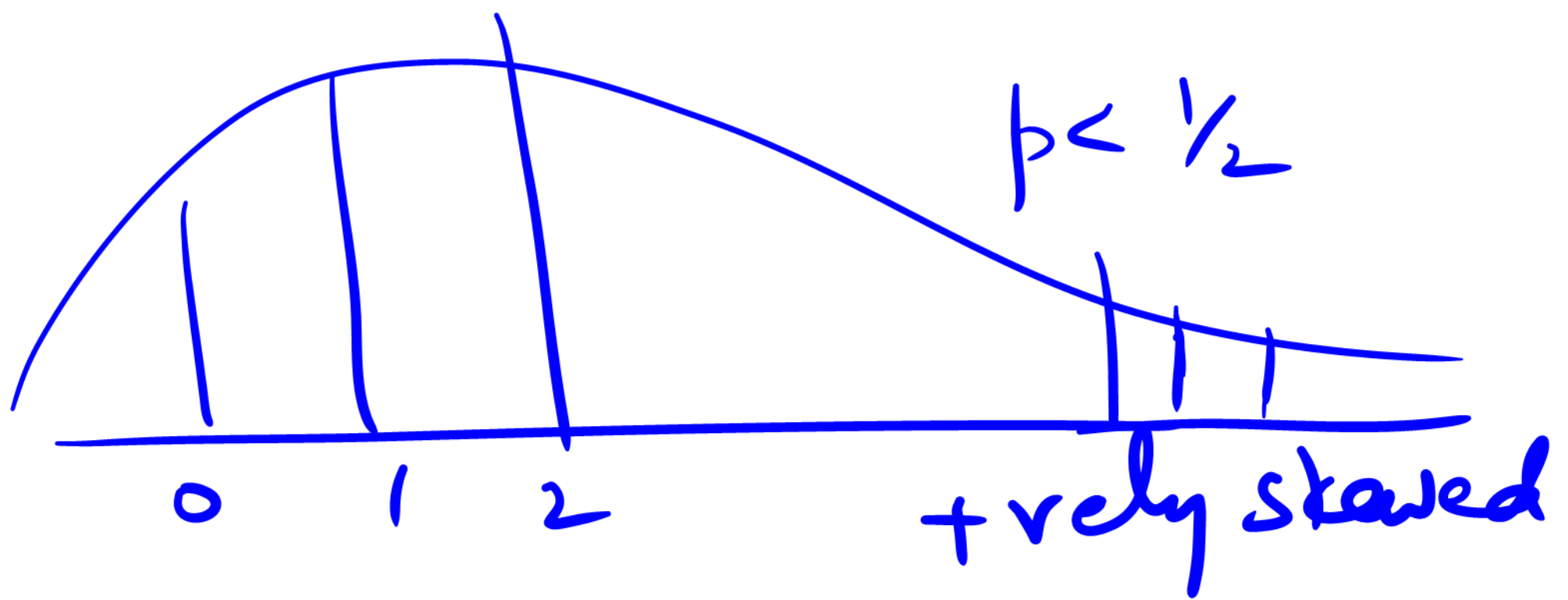
$$\beta_1 = \frac{\mu_3}{\sigma_3} = \frac{1-2p}{(npq)^{1/2}} = 0, \quad p = 1/2$$

perfectly symmetric



$$p = 1/2$$

$\beta_1 > 0$  for  $p < 1/2$



$$\beta_2 = \frac{\mu_4}{\mu_2^2} - 3 = \frac{1 - 6pq}{npq} = 0, \quad pq = \frac{1}{6}$$

$$npq > 0 \quad pq < \frac{1}{6}$$

$$< 0 \quad pq > \frac{1}{6}$$

(\*) Interpret it by taking specific values of  $p$