

Transformations of Random Vectors

Let $\underline{f} : \mathbb{R}^n \rightarrow \mathbb{R}^m$ be a measurable fn.

$$\underline{X} = (X_1, \dots, X_n)$$

$$Y_1 = f_1(X_1, \dots, X_n)$$

\vdots

$$Y_m = f_m(X_1, \dots, X_n)$$

$$\underline{Y} = (Y_1, \dots, Y_m)$$

\underline{Y} is m -dimensional
random vector

For specific types of functions sometimes

mgf is useful. For example, in finding distributions of sums of independent r.v.'s.

Example: let $x_1, \dots, x_n \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2)$

$$Y = \sum_{i=1}^n x_i$$

$$\begin{aligned} M_Y(t) &= \prod_{i=1}^n M_{x_i}(t) = \prod_{i=1}^n \left[e^{\mu t + \frac{1}{2}\sigma^2 t^2} \right] \\ &= e^{n\mu t + \frac{1}{2}n\sigma^2 t^2} \end{aligned}$$

ie $Y \sim N(n\mu, n\sigma^2)$

$$\bar{X} = \frac{Y}{n} = \frac{\sum X_i}{n} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

If r. v.'s are discrete, we may have to use pmf directly to derive pmf of the

transformed variables.

Ex. Let $X, Y \stackrel{i.i.d}{\sim} B(n, p)$

$$U = X + Y \sim \text{Bin}(2n, p).$$

$$V = X - Y, \quad v \rightarrow -n, -(n-1), \dots, -1, 0, 1, \dots, n$$

$$P(V=v) = P(X-Y=v) = P(X=v+Y)$$

$$= \sum_{y=0}^n P(X=v+y, Y=y)$$

$$= \sum_{y=0}^n P(X=v+y) P(Y=y) \quad 0 \leq v+y \leq n$$

$$= \sum \binom{n}{u+y} p^{u+y} (1-p)^{n-u-y}$$

$$\binom{n}{y} p^y (1-p)^{n-y}$$

$$= \sum_{y=0}^n \binom{n}{u+y} \binom{n}{y} p^{u+2y} (1-p)^{2n-u-2y}$$

$u+y = 0, 1, \dots, n$

2. $U = \frac{X}{Y+1}, \quad V = Y+1$

$(U, V) \quad ?? \quad (\text{dist}^n)$

$$V \rightarrow 1, 2, \dots, (n+1)$$

$$U \rightarrow \begin{matrix} 0, 1, \frac{1}{2}, \frac{1}{3}, \dots, \frac{1}{n+1} \\ 2, \frac{2}{2}=1, \frac{2}{3}, \dots, \frac{2}{n+1} \\ n, \frac{n}{2}, \dots, \frac{n}{n+1} \end{matrix} \left. \vphantom{\begin{matrix} 0, 1, \frac{1}{2}, \frac{1}{3}, \dots, \frac{1}{n+1} \\ 2, \frac{2}{2}=1, \frac{2}{3}, \dots, \frac{2}{n+1} \\ n, \frac{n}{2}, \dots, \frac{n}{n+1} \end{matrix}} \right\}$$

The joint pmf of (U, V)

$$\begin{aligned} P(U=u, V=v) &= P(X=uv, Y=v-1) \\ &= P(X=uv) P(Y=v-1) \end{aligned}$$

$$= \binom{n}{uv} p^{uv} (1-p)^{n-uv} \cdot \binom{n}{v-1} p^{v-1} (1-p)^{n-v+1}$$

3. Let (X, Y) have joint pmf

$Y \backslash X$	-1	0	1
-2	$1/6$	$1/12$	$1/6$
1	$1/6$	$1/12$	$1/6$
2	$1/12$	0	$1/12$

$$U = |X|, V = Y^2$$

$(U, V) \rightarrow ???$
joint pmf

$$U \rightarrow 0, 1, \quad V \rightarrow 1, 4$$

$V \backslash U$	0	1
1	$\frac{1}{12}$	$\frac{1}{3}$
4	$\frac{1}{12}$	$\frac{1}{2}$

CDF Approach

Ex Let (X, Y) have joint pdf

$$f_{X,Y}(x,y) = \begin{cases} \frac{1+xy}{4}, & |x| < 1, |y| < 1 \\ 0, & \text{ew} \end{cases}$$

$$U = X^2, \quad V = Y^2$$

The joint cdf of (U, V)

$$F_{U,V}(u, v) = P(U \leq u, V \leq v), \quad \begin{matrix} 0 \leq u < 1 \\ 0 \leq v < 1 \end{matrix}$$

$$= P(-\sqrt{u} \leq X \leq \sqrt{u}, -v < Y \leq \sqrt{v})$$

$$= \int_{-\sqrt{v}}^{\sqrt{v}} \int_{-\sqrt{u}}^{\sqrt{u}} \left(\frac{1+xy}{4} \right) dx dy = \boxed{\sqrt{u}\sqrt{v}}$$

Theorem: Let $\underline{X} = (X_1, \dots, X_n)$ be a continuous random vector with joint pdf

$$f_{\underline{X}}(\underline{x}), \quad \underline{x} = (x_1, \dots, x_n).$$

(a) Let $U_i = g_i(\underline{x}), \quad i=1, \dots, n$

$\underline{U} = (U_1, \dots, U_n) : \mathbb{R}^n \rightarrow \mathbb{R}^n$ be one-to-one

Let $x_i = h_i(\underline{u}), \quad i=1, \dots, n$ be inverses.

(b) Let the function \underline{g} inverse both be continuous

(c) Assume partial derivatives $\frac{\partial x_i}{\partial u_j}$, $i, j = 1, \dots, n$ exist and are continuous

(d) Assume that the Jacobian of transformation

$$J = \begin{vmatrix} \frac{\partial x_1}{\partial u_1} & \frac{\partial x_1}{\partial u_2} & \dots & \frac{\partial x_1}{\partial u_n} \\ \frac{\partial x_2}{\partial u_1} & \frac{\partial x_2}{\partial u_2} & \dots & \frac{\partial x_2}{\partial u_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial u_1} & \frac{\partial x_n}{\partial u_2} & \dots & \frac{\partial x_n}{\partial u_n} \end{vmatrix} \neq 0$$

in the range of the transformation.

Then the r. vector $\underline{U} = (U_1, \dots, U_n)$ is continuous and has joint pdf given by

$$f_{\underline{U}}(\underline{u}) = f_{\underline{X}}(h_1(\underline{u}), \dots, h_n(\underline{u})) |J|.$$

Example. 1. Let $X_1, X_2, X_3 \stackrel{i.i.d.}{\sim} \text{Exp}(1)$

So the joint pdf of $\underline{X} = (X_1, X_2, X_3)$

$$f_{\underline{X}}(\underline{x}) = \prod_{i=1}^3 f_{X_i}(x_i) = e^{-(x_1 + x_2 + x_3)}, \quad x_i > 0, \quad i=1,2,3$$

Let $\underline{Y} = (Y_1, Y_2, Y_3)$, where

$$Y_1 = X_1 + X_2 + X_3, \quad Y_2 = \frac{X_1 + X_2}{X_1 + X_2 + X_3}, \quad Y_3 = \frac{X_1}{X_1 + X_2}$$

We want joint pdf of \underline{Y} and marginal pdf's of Y_1, Y_2, Y_3 .

$$x_1 = y_1 y_2 y_3$$

$$x_2 = y_1 y_2 (1 - y_3)$$

$$x_3 = y_1 (1 - y_2)$$

$$y_1 > 0$$

$$0 < y_2 < 1$$

$$0 < y_3 < 1$$

$$|J| = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \frac{\partial x_1}{\partial y_3} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} & \frac{\partial x_2}{\partial y_3} \\ \frac{\partial x_3}{\partial y_1} & \frac{\partial x_3}{\partial y_2} & \frac{\partial x_3}{\partial y_3} \end{vmatrix} = \begin{vmatrix} y_2 y_3 & y_1 y_3 & y_1 y_2 \\ y_2(1-y_3) & y_1(1-y_2) & -y_1 y_2 \\ 1-y_2 & -y_1 & 0 \end{vmatrix} = -y_1^2 y_2$$

So the joint pdf of $\underline{y} = (y_1, y_2, y_3)$ is

$$f_{\underline{y}}(\underline{y}) = \begin{cases} y_1^2 y_2 e^{-y_1}, & y_1 > 0, 0 < y_2 < 1, \\ & 0 < y_3 < 1 \\ 0, & \text{ew} \end{cases}$$

The marginal (joint) pdf of Y_1, Y_2 is

$$f_{Y_1, Y_2}(y_1, y_2) = \begin{cases} y_1^2 y_2 e^{-y_1}, & y_1 > 0, 0 < y_2 < 1 \\ 0, & \text{else} \end{cases}$$

The marginal pdf of Y_1 is

$$f_{Y_1}(y_1) = \begin{cases} \frac{1}{2} y_1^2 e^{-y_1}, & y_1 > 0 \\ 0, & \text{else} \end{cases}$$

Gamma(3, 1)

The marginal pdf of Y_2 is

$$f_{Y_2}(y_2) = \begin{cases} 2y_2, & 0 < y_2 < 1 \\ 0, & \text{else} \end{cases}$$

Beta(2, 1)

The marginal pdf of Y_3 is

$$f_{Y_3}(y_3) = \begin{cases} 1, & 0 < y_3 < 1 \\ 0, & \text{else} \end{cases} \quad \boxed{U(0,1)}$$

Since $f_{\underline{Y}}(\underline{y}) = \prod_{i=1}^3 f_{Y_i}(y_i) \quad \forall \underline{y} \in \mathbb{R}^3$

we conclude that Y_1, Y_2, Y_3 are also independent.

2. Let $X, Y \stackrel{i.i.d.}{\sim} U(0,1)$

$$U = X + Y, \quad V = X - Y.$$

Find joint & marginal pdf's of U & V .

The joint pdf of (X, Y) is

$$f_{X,Y}(x,y) = \begin{cases} 1, & 0 < x < 1, 0 < y < 1 \\ 0, & \text{ew} \end{cases}$$

$$x = \frac{u+v}{2}$$

$$y = \frac{u-v}{2}$$

$$J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} 1/2 & 1/2 \\ 1/2 & -1/2 \end{vmatrix} = -\frac{1}{2}$$

The joint pdf of (U, V) is

$$f_{U,V}(u,v) = \begin{cases} \frac{1}{2}, & 0 < u+v < 2, \quad 0 < u-v < 2 \\ & \text{where } 0 < u < 2 \\ & \quad -1 < v < 1 \\ 0, & \text{else} \end{cases}$$

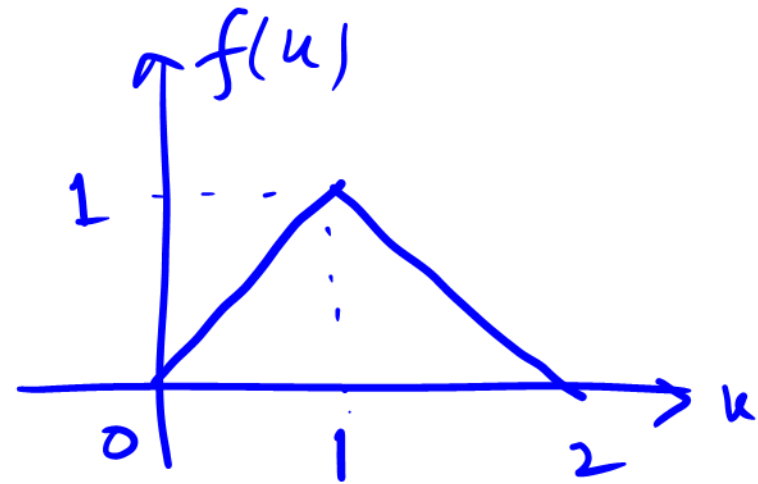
The marginal pdf of U is

$$f_U(u) = \begin{cases} \frac{1}{2} \int_{-u}^u dv, & 0 < u < 1 \end{cases}$$

$$\left(\frac{1}{2} \int_{u-2}^{2-u} du \right),$$

$$1 \leq u < 2$$

$$= \begin{cases} u, & 0 < u < 1 \\ 2-u, & 1 < u < 2 \\ 0, & \text{else} \end{cases}$$

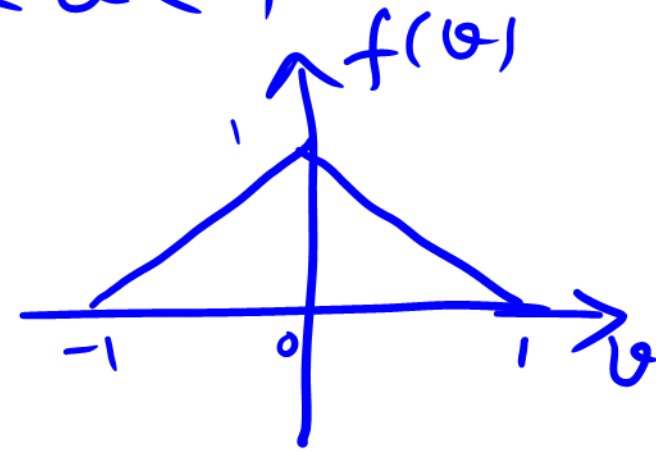


The marginal pdf of V is

$$f_V(v) = \begin{cases} \frac{1}{2} \int_{-v}^{v+2} du \end{cases}$$

$$-1 < v \leq 0$$

$$\begin{aligned}
 & \left(\frac{1}{2} \int_0^{2-u} du, \quad 0 < u < 1 \right. \\
 & = \begin{cases} 1+u, & -1 < u \leq 0 \\ 1-u, & 0 < u < 1 \\ 0, & \text{else} \end{cases}
 \end{aligned}$$



The sum and difference of two independent $U(0,1)$ r.v.'s have triangular distributions.