



Nonparametric Statistics

J. Lozano

University of Goettingen

Department of Genetic Epidemiology

Interdisciplinary PhD Program in
Applied Statistics & Empirical Methods

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Binomial test

Friedman's ANOVA

Wilcoxon-Mann-Whitney

McNemar



Median test

Sign Test

Fisher's Exact

Spearman

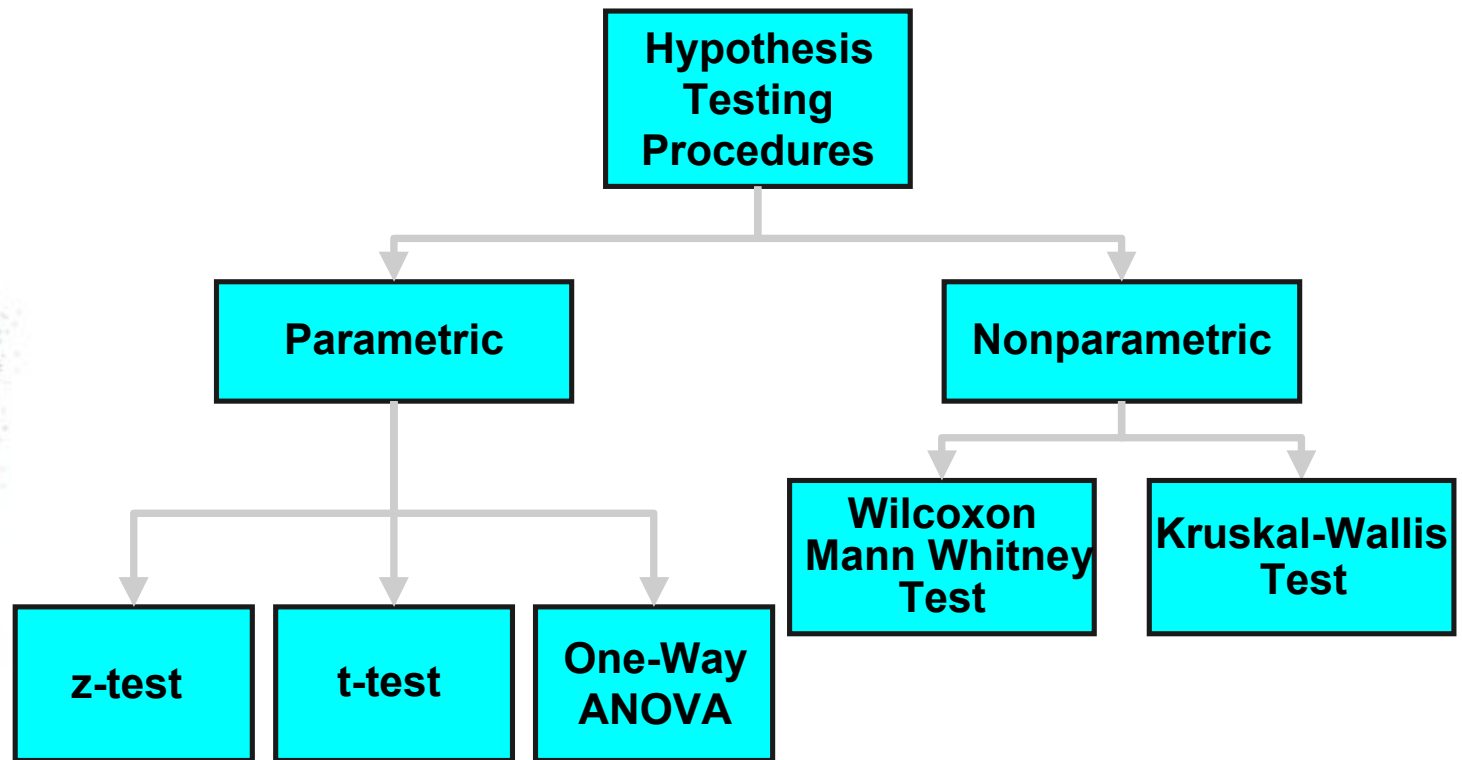
Wilcoxon Signed Ranks Test

Kruskal-Wallis

Kappa

Run's Test
Run's Test

Parametric vs. Nonparametric



Many More Tests Exist!



Learning Objectives:

1. Differentiate nonparametric from parametric statistics
2. Discuss the advantages and disadvantages of nonparametric statistics
3. Enumerate and differentiate the different nonparametric tests
4. Apply some commonly used nonparametric test to hypothesis testing problems

Parametric Test Procedures

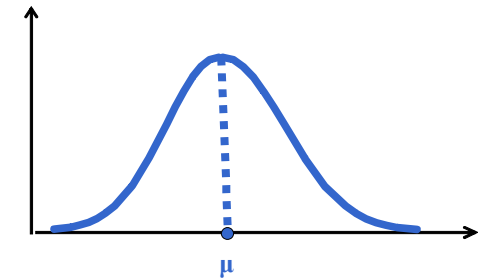
1. Involve Population Parameters

Example: Population Mean

$$t = \frac{\bar{X} - \mu_0}{(s_x / \sqrt{n})}$$

2. Have Stringent Assumptions

Example: Normal Distribution



3. Require Interval Scale or Ratio Scale

Whole Numbers or Fractions

Example: Height in Inches (72, 60.5, 54.7)

4. Examples: z-Test, t-Test, ANOVA

Nonparametric Test Procedures



1. Do Not Involve Population Parameters
2. No Stringent Distribution Assumptions
“Distribution-free”
3. Data Measured on Any Scale
 - Ratio or Interval
 - Ordinal
 - Example: Good-Better-Best
 - Nominal
 - Example: Male-Female
4. Example: Wilcoxon-Mann-Whitney Test



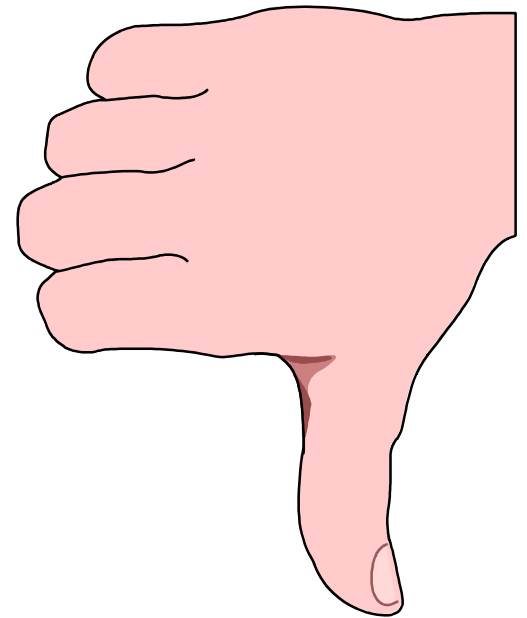
Advantages of Nonparametric Tests

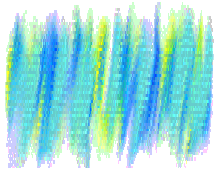
1. Used With All Scales
2. Easier to Compute
Developed Originally Before
Wide Computer Use
3. Make Fewer Assumptions
4. Need Not Involve Population
Parameters
5. Results May Be as Exact as
Parametric Procedures



Disadvantages of Nonparametric Tests

1. May Waste Information
If Data Permit Using Parametric Procedures
Example: Converting Data From Ratio to Ordinal Scale
2. Require a larger sample size than the corresponding parametric test in order to achieve the same power
3. Difficult to Compute by Hand for Large Samples
4. Stat tables are not readily available



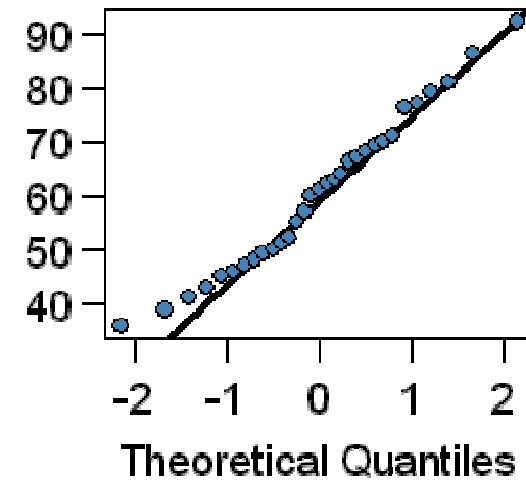
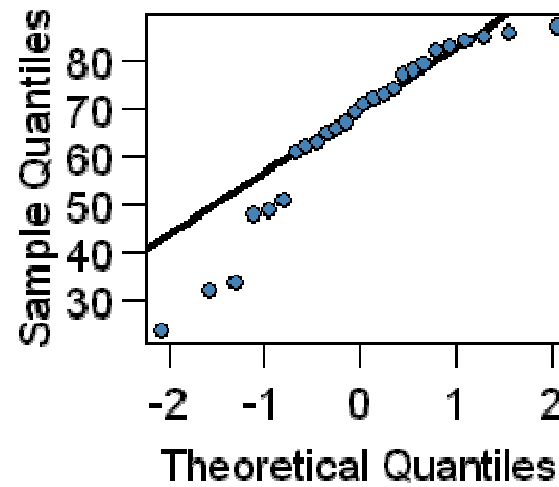
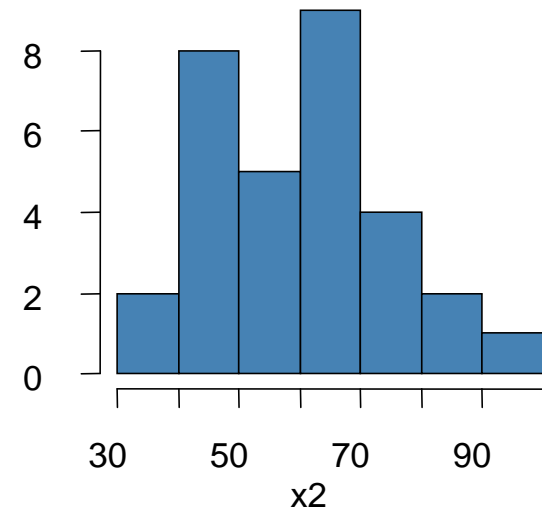
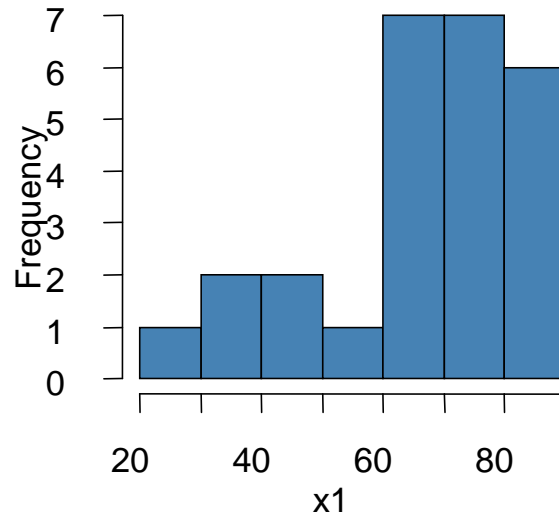


Summary Table of Statistical Tests

Level of Measurement	Sample Characteristics					Correlation
	1 Sample	2 Sample		K Sample (i.e., >2)		
		Independent	Dependent	Independent	Dependent	
Categorical or Nominal	Bi-nomial χ^2	χ^2 Fisher Exact	Mc Nemar's	χ^2	Cochran's Q	Kappa Agreement Test
Rank or Ordinal	Run's test Kolmogorov Smirnov	Median Test Wilcoxon-Mann-Whitney	Sign Test Wilcoxon Signed Ranks	Kruskal-Wallis	Friendman's ANOVA	Spearman's rho Kendall Rank
Parametric (Interval & Ratio)	z-test or t-test	t- test between groups	Paired t-test	1 way ANOVA between groups	1 way ANOVA (within or repeated measure)	Pearson's r
		Factorial ANOVA				



Application of Commonly Used Nonparametric Statistics



=> nonparametric statistics...

Wilcoxon-Mann-Whitney Test

The background of the slide features a solid green upper section. The bottom of the slide is decorated with two overlapping wavy shapes: a pink one on the left and an orange one on the right, which together form a horizontal band across the bottom.



Wilcoxon-Mann-Whitney Test

- Also known as Wilcoxon-test, Wilcoxon rank sum test, U-test, Mann-Whitney-U-test
- Tests Two Independent Population
 - compare medians
- Corresponds to t-Test for 2 Independent Means
- Assumptions
 - Independent, Random Samples
 - Populations Are Continuous
- Can Use Normal Approximation If $n_i \geq 20$



Wilcoxon-Mann-Whitney Test Procedure

1. Assign Ranks, r_i , to the $n_1 + n_2$ Sample Observations
If unequal sample sizes, let n_1 refer to smaller-sized sample
Smallest Value = 1
Average Ties
2. Sum the Ranks, R_i , for Each Sample
3. Test Statistic
 - Null hypothesis: both samples come from the same underlying distribution

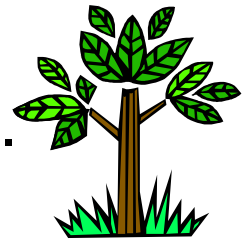
$$U = n_1 n_2 + \frac{n_1(n_1 + 1)}{2} - R_1 \quad \text{where } R_1 \text{ is the sum of ranks in sample 1}$$

Compute also $(n_1 n_2 - U)$. If this expression is larger than the above U , then this becomes the final U statistic.

** At $N \geq 20$, U begins to approximate t , so the test stat changes to a t -value.

Wilcoxon-Mann-Whitney Test Example

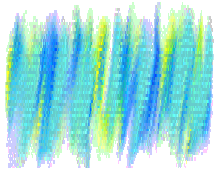
- You're an agriculturist. You want to determine if there was a difference in the biomass of male and female Juniper trees.



- Randomly select 6 trees of each gender from the field. Dry them to constant moisture and weigh in kg.



- Male trees Data: 71, 73, 78, 75, 72, 74
- Female trees Data: 80, 73, 83, 84, 82, 79



Wilcoxon-Mann-Whitney Test Solution

• H_0 :

• H_a :

☐ $\alpha =$

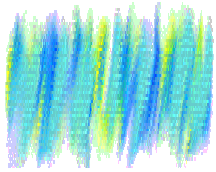
• $n_1 =$ $n_2 =$

• Critical Value:

Test Statistic:

Decision:

Conclusion:



Wilcoxon-Mann-Whitney Test Computation Table

Male Trees		Female Trees	
Weight	Rank	Weight	Rank
71		80	
73		73	
78		83	
75		84	
72		82	
74		79	

Rank Sum

Wilcoxon-Mann-Whitney Test Solution

Median weights =

• H_a : Medians not =

□ $\alpha = 0.05$

• $n_1 = 6$ $n_2 = 6$

• Critical Value:

U table = 31

* Depends on the U table you are using

Test Statistic:

$$U = (6)(6) + (6)(7)/2 - 24.5 = 32.5$$

$$n_1 n_2 - U = (6)(6) - 32.5 = 3.5$$

$$*!!! \text{ final } U_{\text{Calc}} = 32.5$$

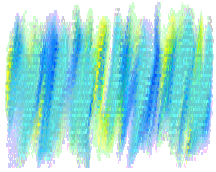
Decision:

Reject H_0 at $\alpha = .05$ since

$$U_{\text{calc}} > U_{\text{table}}$$

Conclusion:

There is evidence that the medians are not equal.



Wilcoxon-Mann-Whitney Test Solution (SPSS Output)

NPar Tests

Mann-Whitney Test

Ranks

Gender of Tree		N	Mean Rank	Sum of Ranks
Weight in kg	Male	6	4.08	24.50
	Female	6	8.92	53.50
	Total	12		

Test Statistics^b

	Weight in kg
Mann-Whitney U	3.500
Wilcoxon W	24.500
Z	-2.326
Asymp. Sig. (2-tailed)	.020
Exact Sig. [2*(1-tailed Sig.)]	.015 ^a

$p < \alpha$; Reject H_0

a. Not corrected for ties.

b. Grouping Variable: Gender of Tree

Sign Test

The background of the slide is a solid green color. At the bottom, there is a decorative wavy line that separates the green area from a base composed of two colors: a bright pink on the left and a bright orange on the right.



Sign Test

1. Tests One Population Median, η (eta)
2. Corresponds to t-Test for 1 Mean
3. Assumes Population Is Continuous
4. Small Sample Test Statistic: # Sample Values Above (or Below) Median
5. Can Use Normal Approximation If $n \geq 10$

Sign Test Example

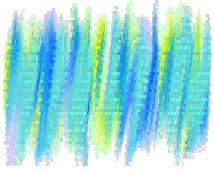
- You're a marketing analyst for Chefs-R-Us. You've asked 8 people to rate a new ravioli on a 5-point Likert scale
1 = terrible to
5 = excellent

The ratings are:

2 4 1 2 1 1 2 1

At the **.05** level, is there evidence that the **median** rating is **at least 3**?





Sign Test Solution

- H_0 :

P-Value:

- H_a :

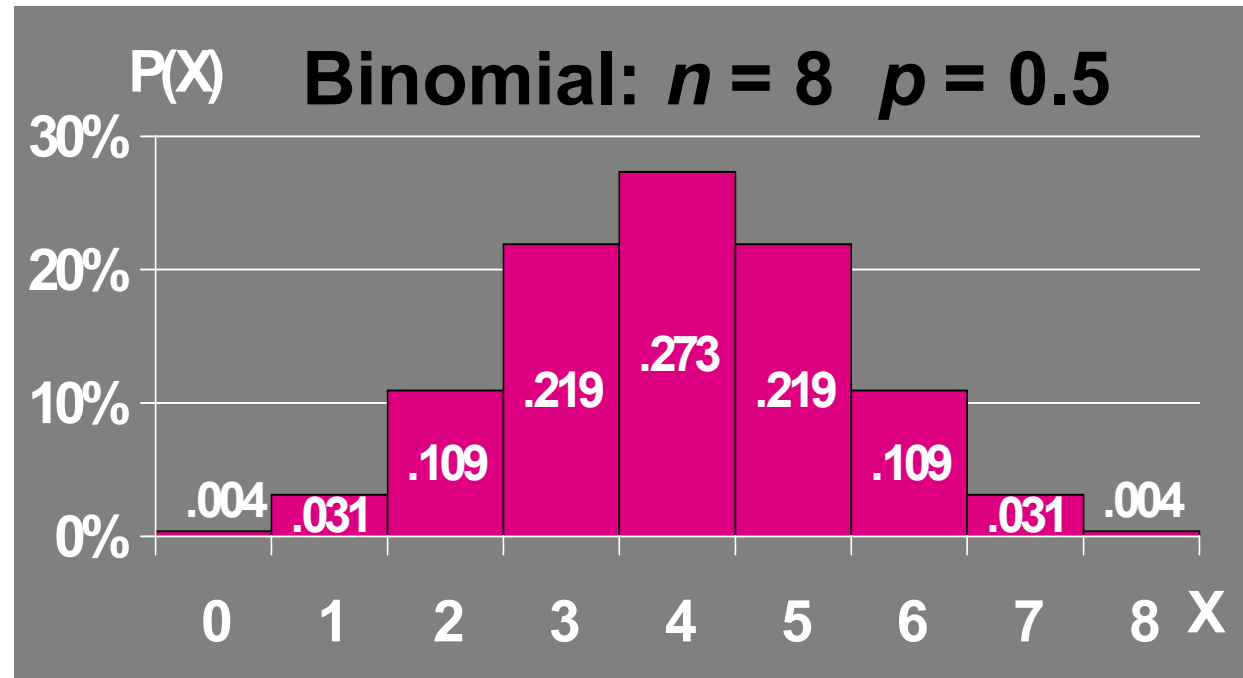
☐ $\alpha =$

- Test Statistic:

Decision:

Conclusion:

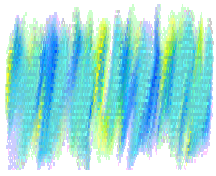
Sign Test Uses P-Value to Make Decision



P-Value is the probability of getting an observation at Least as extreme as we got.

If 7 of 8 Observations 'Favor' H_a , Then P-Value = $P(x \geq 7)$ = .031 + .004 = .035.

If $\alpha = .05$, Then Reject H_0 Since P-Value $\leq \alpha$.



Sign Test Solution

- $H_0: \eta = 3$
- $H_a: \eta < 3$
- $\alpha = .05$
- **Test Statistic:**
 $S = 7$
(Ratings 1 & 2 are less than $\eta = 3$:
2 4 1 2 1 1 2 1)

P-Value:

$P(x \geq 7) = .031 + .004 = .035$
(Binomial Table, $n = 8$, $p = 0.50$)

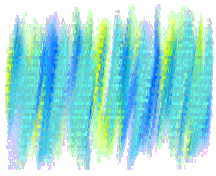
$$P(X) = \frac{n!}{X!(n-X)!} \pi^x (1-\pi)^{n-X}$$

Decision:

Reject H_0 at $\alpha = .05$

Conclusion:

There is evidence that the median is less than 3



Sign Test (R Output)

R version 2.3.1 (BSDA and eV10 package must be loaded)

```
> x <- c(2,4,1,2,1,1,2,1)
> sign.test(x,md=3, alternative="less")
```

One-sample Sign-Test

data: x

s = 1, **p-value = 0.03516**

alternative hypothesis: true median is less than 3

95 percent confidence interval:

-Inf 2

sample estimates:

median of x

1.5

\$Confidence.Intervals

	Conf.Level	L.E.pt	U.E.pt
Lower Achieved CI	0.8555	-Inf	2
Interpolated CI	0.9500	-Inf	2
Upper Achieved CI	0.9648	-Inf	2

Warning message:

multi-argument returns are deprecated in: return(rval, Confidence.Intervals)

Mc Nemar Change Test





McNemar Change Test (MCT)

- Typical examples:
 - Testing the shift in the proportion of abnormal responses from before and after treatment in the same group of patients
 - Comparing two ocular treatments when both are given to each patient, one in each eye

Data Layout: MCT

		Condition 2		
		No. of 'Responders'	No. of 'Non- Responders'	TOTAL
Condition 1	No. of 'Responders'	A	B	A + B
	No. of 'Non- Responders'	C	D	C + D
	TOTAL	A + C	B + D	$N = A + B + C + D$



McNemar Change Test (MCT)

- The hypothesis of interest is the equality of response proportions, p_1 and p_2 , under conditions 1 and 2, respectively.
- The test statistic is based on the difference in the discordant cell frequencies (B, C).



Test Summary: MCT

- Null hypothesis: $H_0 : p_1 = p_2$
- Alternative hypothesis: $H_a : p_1 \neq p_2$
- Test Statistic:

$$\chi^2 = \frac{(B - C)^2}{B + C}$$

- Decision Rule: Reject H_0 if $\chi^2 > \chi_1^2(\alpha)$
 - where $\chi_1^2(\alpha)$ is the critical value from the chi-square table with significance level α and 1 degree of freedom

Example: MCT

- Bilirubin Abnormalities Following Drug Treatment

86 patients were treated with an experimental drug for 3 months.

Pre-post study clinical laboratory results showed abnormally high total bilirubin values.

Is there evidence of a change in the pre-to post-treatment rates of abnormalities?





Solution: MCT

Let p_1 and p_2 represent the proportions of patients with abnormally high bilirubin values ('Y') before and after treatment, respectively.

		POST-Treatment		TOTAL
		'N'	'Y'	
PRE-Treatment	'N'	60	14	74
	'Y'	6	6	12
TOTAL		66	20	86

'Y' = T.Bilirubin above upper limit of normal range



Test Summary: MCT

- Null hypothesis: $H_0 : p_1 = p_2$
- Alternative hypothesis: $H_a : p_1 \neq p_2$

- Test Statistic:
$$\chi^2 = \frac{(B-C)^2}{B+C} = \frac{(14-6)^2}{14+6} = \frac{64}{20} = 3.20$$

- Decision: Do not reject H_0 since

$$\chi^2 < \chi_1^2 (\alpha_{0.05}) = 3.841$$

- Conclusion:

There is no sufficient evidence to conclude at 0.05 level of significance that a shift in abnormality rates occurs with treatment.

SAS Computer Output for MCT

McNemar's Test

Example: Bilirubin Abnormalities Following Drug Treatment

TABLE OF PRE BY PST

PRE(PRE)	PST(PST)		
Frequency,			
Percent ,			
Row Pct ,			
Col Pct ,N		Y	Total
fffffffff^fffffffff^fffffffff^			
N	60	14	74
	69.77	16.28	86.05
	81.08	18.92	
	90.91	70.00	
fffffffff^fffffffff^fffffffff^			
Y	6	6	12
	6.98	6.98	13.95
	50.00	50.00	
	9.09	30.00	
fffffffff^fffffffff^fffffffff^			
Total	66	20	86
	76.74	23.26	100.00

STATISTICS FOR TABLE OF PRE BY PST

McNemar's Test

Statistic = 3.200

DF = 1

Prob = 0.074



References

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Siegel S. and Castellan N.J. Nonparametric Statistics for the Behavioral Sciences (2nd edition). New York: McGraw Hill (1988).

Walker G.A. Common Statistical Methods for Clinical Research with SAS Examples, SAS Institute, Inc, Cary NC (1997).