Cognitive Interference Management in Retransmission-Based Wireless Networks

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Abstract—Cognitive radio methodologies have the potential to dramatically increase the throughput of wireless systems. Herein, control strategies which enable the superposition in time and frequency of primary and secondary user transmissions are explored in contrast to more traditional sensing approaches which only allow the secondary user to transmit when the primary user is idle. In this paper, the optimal transmission policy for the secondary user when the primary user adopts a retransmission-based error control scheme is investigated. The policy aims to maximize the secondary users' throughput, with a constraint on the throughput loss and failure probability of the primary user. Due to the constraint, the optimal policy is randomized, and determines how often the secondary user transmits according to the retransmission state of the packet being served by the primary user. The resulting optimal strategy of the secondary user is proven to have a unique structure. In particular, the optimal throughput is achieved by the secondary user by concentrating its transmission, and thus its interference to the primary user, in the first transmissions of a primary user packet. The rather simple framework considered in this paper highlights two fundamental aspects of cognitive networks that have not been covered so far: 1) the networking mechanisms implemented by the primary users (error control by means of retransmissions in the considered model) react to secondary users' activity; 2) if networking mechanisms are considered, then their state must be taken into account when optimizing secondary users' strategy, i.e., a strategy based on a binary active/idle perception of the primary users' state is suboptimal.

Index Terms—Automatic retransmission request (ARQ), cognitive radios, Markov processes, reactive primary users, wireless networks.

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I. Introduction

▼ OGNITIVE radio has been the subject of intense research of late, e.g., [1]-[9], due to its potential to increase the efficiency of wireless networks. Unlicensed secondary users adapt their operations around those of the primary users and the surrounding network environment to opportunistically exploit available resources while limiting their interference with licensed primary users. Most prior work [1]-[4] focuses on a white space approach, where the secondary users sense the channel in order to detect time/frequency slots left unused by the primary users and exploit them for transmission. Pure white space approaches are based on a zero-interference rationale, i.e., the objective of the secondary user is to not interfere at all with the primary user. However, sensing errors may lead to unwanted collisions, thus degrading the throughput achieved by the latter. Typically, primary users are modeled via a fixed Markov chain tracking the idle-busy channel state, irrespective of the operations of the secondary users, according to the general assumption that primary users are dumb and nonadaptive devices.

However, such a model may not always be accurate. For instance, a collision may force a primary user to schedule a retransmission and enter a backoff period. As a consequence, a collision may modify the arrival rate of the packets at the primary destination, while changing the characterization of the generated traffic (burstiness of idle/busy slots). As a consequence, the interaction between the primary users and the secondary users must be considered when analyzing the network. Additionally, the use of signal processing methods (multiuser detection and multiple-input multiple-output systems) enables the superposition of secondary transmissions over a primary transmission while achieving accurate decoding of the primary user packet.

There exists some prior literature investigating the coexistence in the same time/frequency band of primary and secondary users with a focus on physical layer methods for static scenarios [5]–[8], [10], [11]. A thorough discussion of spectrum sharing under performance constraints from an information theoretic perspective can be found in [12]. Those approaches, though valuable in some broadcasting network scenarios, do not characterize the dynamic interaction between the two classes of users. In contrast, our prior work, which inspires this paper, studies concurrent transmission by secondary and primary users in a highly dynamic environment [13]. We explicitly consider an interference mitigation scenario, where the secondary user is allowed to transmit concurrently to the primary user, with a constraint on the performance loss suffered by the latter, in terms of either a reduced throughput or an increased failure probability.

In this paper, we study access control policies for secondary users in wireless networks where nodes implement a retransmission-based error control scheme. Some prior literature has investigated networks of primary users implementing automatic retransmission request (ARQ). In [14], the secondary user exploits the retransmissions of primary user packets in order to achieve a higher transmission rate. In fact, the secondary receiver can potentially decode the primary user's packet in the first transmission and then opportunistically cancel interference [11] in the following retransmissions. However, the framework in [14] does not consider the dynamics of the network and the bias in the channel availability generated by interference. In [15], Eswaran et al. propose a framework where the secondary user exploits ARQ feedback to estimate the throughput loss of the primary user and tune the transmission policy accordingly, by using information theoretic results. In [16], Zhang proposed a learning algorithm for a scenario in which the primary user adapts the transmitted power in response to interference.

The contribution of this paper is to introduce the reactive primary user scenario, where the activity of the secondary users biases the temporal evolution of the stochastic process tracking the state of the primary users. A Markov model is proposed and the optimization problem is formulated as a constrained Markov decision process (MDP). The structure of the optimal policy is derived analytically for a specific case. We focus on a network with two mutually interfering links, one primary and one secondary. In the framework considered, a packet may be retransmitted a finite number of times, due to transmission failure, before being discarded by the transmitter. Packet arrivals at the primary user are modeled with a fixed probability that an empty slot, i.e., a slot in which a retransmission is not scheduled, is accessed for transmission.

We study the interference that the secondary user causes to the primary user and how this interference impacts the retransmission process of the latter. We explicitly consider an interference mitigation scenario, where the secondary user is allowed to transmit concurrently to the primary user, with a constraint on the performance loss suffered by the latter, in terms of either a reduced throughput or an increased failure probability. Our analysis is based on a detailed Markov model of the network, accounting for the distortion of the retransmission process caused by secondary user transmissions. Remarkably, this simple model captures a fundamental aspect of cognitive networks: the control mechanisms implemented by the primary users react to the activity of the secondary users. An accurate stochastic model for the activity of the primary users should include this effect. In the model considered herein, interference from the secondary source increases the probability that primary source's transmissions fail. As a consequence, retransmissions are triggered more often, and the stochastic characterization of primary source's channel occupation changes. Moreover, as the motion law of the state of the primary user depends on both the action of the secondary user and the current state of the primary user (the retransmission index of the packet being served in the considered model), then the secondary users' strategy should be based on the state of the primary user. This means that a binary active/idle representation of the state of the primary users leads to suboptimal policies.

In this framework, interference due to the activity of the secondary user not only reduces the instantaneous average revenue collected by the primary user in each state of the Markov chain modeling the network, but also changes the transition probabilities, and thus the steady-state distribution, of the chain. The optimization problem can be formalized through a linear program (LP). Due to the constraint on the maximum performance loss of the primary user, the solution is a randomized policy, i.e., the optimal policy assigns a probability to each action in the action set given the state of the underlying Markov process. As we focus on a binary transmission/idleness action set, the randomized policy simply determines how often the secondary user transmits given the state of the network.

This problem, though conceptually simple, unveils important issues and general behaviors. As the primary user implements a retransmission-based error control mechanism, the activity of the secondary user biases the retransmission process via interference. Interference at the primary receiver increases the failure probability of primary user's transmissions. Therefore, due to the activity of the secondary user, the average number of transmissions of a primary user's packet gets larger, together with the average time required to return to primary user's idle state. Interestingly, the increase of the average number of transmissions of primary user's packets depends on the index of the interfered transmission. For instance, while interference from the secondary user in the first transmission of the primary user's packets potentially leads to a significant increase of the number of transmissions per packet, transmission by the secondary user in the last allowed transmission of primary user's packets does not increase the average number of transmissions at all. Thus, as observed earlier, the impact of secondary users' transmission in the various states critically depends on the state of the primary network. On the other hand, first transmissions occur more frequently than last transmissions, and, thus, the overall throughput collected by the secondary user as a function of the strategy greatly depends on the states in which it concentrates its transmissions. The interplay between the cost of the primary user and the reward of the secondary one due to the modifications of the steady-state distribution determines the optimal strategy.

An important observation concerns the availability of time slots in which the primary user is idle, i.e., the white spaces. As the primary user implements a retransmission-based error control mechanism, failed decoding at the primary receiver triggers a further transmission of the packet, until the maximum number of transmissions per packet is reached. Therefore, the interference generated by the activity of the secondary user increases the fraction of channel resource occupied by the primary user. This means that the availability of white spaces decreases as the activity of the secondary user increases. This is an additional reason for carefully designing the strategy of the secondary user.

If transmission by the primary user does not affect reception at the secondary receiver, and either throughput or packet failure probability is considered as the metric for the primary user performance, the optimal transmission strategy of the secondary user is shown to have a unique structure. The throughput-optimal strategy concentrates transmissions by the secondary user in the region of the state space corresponding to the first transmissions of a primary user's packet. According to such optimal policy, the secondary user transmits with probability 1 up to the N_1 th transmission of a primary user's packet, with probability in [0,1] in time slots in which the primary user is performing the N_1 th transmission of a packet and with probability 0 otherwise. The boundary state and the associated transmission probability are determined to result into a bounded reduction of the time-average performance of the primary user. This result also

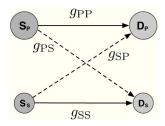


Fig. 1. Considered network. Direct links and interfering links are represented by solid and dashed arrows, respectively.

provides a simple algorithm to solve the LP resulting from the constrained optimization problem.

We also observe that the maximum *aggressiveness* of the secondary user depends on the arrival rate at the primary user. In fact, when the primary source spends most of its time idle, a longer retransmission process has a less deleterious effect on throughput.

The rest of this paper is organized as follows. Section II describes the network scenario considered throughout this paper. Section III defines the optimization problem and derives the Markov model of the network. In Section IV, the structure of the optimal strategy for the case in which primary users' transmission does not affect packet reception at the secondary receiver is derived. Section V discusses the optimal transmission policy for the general case. In Section VII, numerical results highlighting the fundamental issues and behaviors described in the previous sections are shown. Section VIII concludes the paper.

II. NETWORK DESCRIPTION

Consider the network in Fig. 1, with a primary and a secondary source, namely $S_{\rm P}$ and $S_{\rm S}$. The primary source $S_{\rm P}$ and the secondary source $S_{\rm S}$ transmit packets to their respective destinations, namely $D_{\rm P}$ and $D_{\rm S}$.

The reception of a packet at a particular destination is interfered with by the transmission of the other source. Our model subsumes the white space approach which typically assumes that a collision results in a decoding failure. An alternative view is that the collision approach implies a secondary access policy that will result in no throughput loss for the primary user. In contrast, we assign decoding error probabilities to the primary and secondary destinations as an abstraction of various interference mitigation methods. This in turn will result in some throughput loss and increase of the packet failure probability as a function of the access strategy of the secondary user. It is trivial to show that the white space approach is optimal for the constraint of no collisions.

We assume a quasi-static channel model, where time is divided into slots of fixed duration and the channel gain of a certain link remains constant within a slot, and is independent of the channel gains in the other slots. We denote by $g_{\rm PP}, g_{\rm PS}, g_{SS}$, and $g_{\rm SP}$, the random variables corresponding to the channel coefficients, respectively, between $S_{\rm P}$ and $D_{\rm P}$, $S_{\rm P}$ and $D_{\rm S}$, $S_{\rm S}$ and $D_{\rm S}$, and $S_{\rm S}$ and $D_{\rm P}$, and with $\zeta_{\rm PP}(g), \zeta_{\rm PS}(g), \zeta_{\rm SS}(g)$, and $\zeta_{\rm SP}(g)$ their respective probability density function.

Assuming that the transmission of a packet fits a slot, the performance of the receiver can be modeled via the average decoding failure probability that depends on the packet encoding, transmission rates, structure of the receiver, and average channel gains, as well as the activity of the concurrent source. The average decoding failure probability at the primary destination $D_{\rm P}$ associated with a silent secondary source is denoted by $\rho{>}0$, while the same probability when the secondary source transmits is $\rho^*{\geq}\rho$. Analogously, the average decoding failure probability at the secondary destination $D_{\rm S}$ when the primary source is silent and transmitting is denoted with $\nu{>}0$ and $\nu^*{\geq}\nu$, respectively.

The construction fits many models and assumptions on the architecture of the physical layer and the transmission protocols. For instance, one may assume that the primary destination performs signal decoding unaware of the presence of the secondary source, and thus treats its signal as noise, whereas the secondary receiver adopts a smarter decoding strategy, by either treating as noise or decoding and canceling the signal from the primary source according to the transmission rates, powers, and channel coefficients [17].

Denoting the transmission rate and power of the primary and secondary sources with R_P , P_P , R_S , P_S , respectively, we obtain the following failure probabilities for the primary link:

$$\rho = \mathcal{P}\left\{R_{P} > \mathcal{C}\left(g_{PP}P_{P}\right)\right\} \tag{1}$$

$$\rho^* = \mathcal{P}\left\{R_{\rm P} > \mathcal{C}\left(\frac{g_{\rm PP}P_{\rm P}}{1 + g_{\rm SP}P_{\rm S}}\right)\right\}$$
 (2)

where $C(x) = \log(1+x)$.

For the secondary link, we obtain

$$\nu = \mathcal{P}\left\{R_{S} > \mathcal{C}\left(g_{SS}P_{S}\right)\right\} \tag{3}$$

$$\nu^* = \mathcal{P}\left\{\left\{R_{\mathcal{P}}, R_{\mathcal{S}}\right\} \notin \Psi\right\} \tag{4}$$

where Ψ is the set of all the rate pairs $\{R_{\rm P}, R_{\rm S}\}$ such that

$$R_{\rm S} \le \mathcal{C} \left(g_{\rm SS} P_{\rm S} \right)$$
 (5)

$$R_{\rm P} + R_{\rm S} \le \mathcal{C} \left(g_{\rm PP} P_{\rm P} + g_{\rm SS} P_{\rm S} \right) \tag{6}$$

or

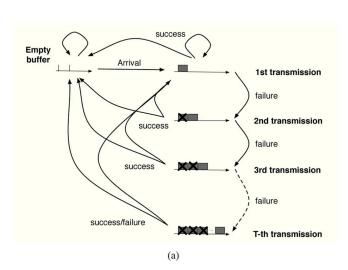
$$R_{\rm S} \le \mathcal{C} \left(\frac{g_{\rm SS} P_{\rm S}}{1 + q_{\rm PS} P_{\rm P}} \right)$$
 (7)

where (5) and (6) refer to the achievable rate region corresponding to the secondary receiver performing interference cancellation, while (7) refers to the case in which the signal from the primary source is treated as noise by the secondary receiver. The failure probabilities listed previously admit a simple integral form and can be easily computed.

We remark that we do not consider a specific physical layer architecture or transmission technique, but rather, we refer to the simple construction based on the average decoding probabilities described earlier.

In order to improve reliability, the primary source implements a retransmission-based error control scheme, by which a failed packet is retransmitted in the subsequent slot. We consider a finite-retransmission process, where each packet can be transmitted at most T times [see Fig. 2(a)]. Delayed retransmissions do not alter the following discussion. If the packet has been transmitted T times, it is discarded by the primary source. It is assumed that the destination sends an acknowledgment packet after each received packet, in order to make the source aware

 1 In the following example, rates $R_{\rm P}$ and $R_{\rm S}$ are expressed in [bit/s/Hz] and the transmission powers $P_{\rm P}$ and $P_{\rm S}$ are normalized to the noise power.



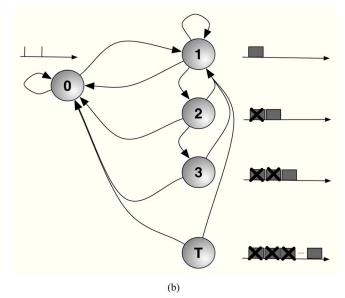


Fig. 2. (a) Primary source's state scheme. (b) Graphical representation of the Markov chain of the system.

of the outcome of the transmission. Note that this scheme can be classified either as an ARQ or a type-I hybrid ARQ scheme depending on whether or not the packets are encoded before transmission. For the sake of simplicity, the secondary source is assumed to transmit each packet only once. This assumption is consistent with the common characterization of secondary users as opportunistic sources without strict quality of service guarantees ("best effort").

Unless a retransmission is scheduled, the primary source $S_{\rm P}$ accesses the channel in each slot to transmit a fresh packet with fixed probability α , with $0<\alpha<1$. The secondary source is assumed to be backlogged, i.e., it always has a packet to transmit. However, a packet arrival process at the secondary source can be included in the model with some straightforward modifications to the analysis in Section III. Nevertheless, its inclusion does not add any insight to the discussion presented in this paper, while it complicates the formulae.

The channel access strategy of the secondary source follows a policy μ , whose action set is U=0,1, where 0 and 1 correspond to a silent and a transmitting source, respectively. We remark that transmission by the secondary source increases the probability of decoding failure at the primary receiver. Thus, the transition probabilities of the Markov chain strongly depend on the secondary user's activity and there is an explicit dependence between the stochastic characterization of the primary source activity and the activity of the secondary source.

The following discussion is specialized to a constraint defined on the throughput loss of the primary source. A constraint posed on the increase of the failure probability only results in a different definition of the average primary source's cost, as reported in Section IV-B.

The throughput achieved by the primary source under policy μ can be written as

$$\mathcal{W}_{\mathbf{P}}(\mu) = \lim_{N \to +\infty} \sup \frac{1}{\tau N} \sum_{n=1}^{N} \mathbb{E}[I(\Xi_{\mathbf{P}}^{n}(\mu))] L_{\mathbf{P}}$$
 (8)

where $L_{\rm P}$ is the size, in bits, of the packets sent by the primary source, au is the duration of a slot, $\Xi^n_{\rm P}(\mu)$ is the event cor-

responding to a successfully delivered packet by the primary source in slot n, I is the indicator function, and $\mathbb E$ denotes average. The throughput of the secondary user admits an analogous expression.

The goal of the secondary source is to maximize its own achieved throughput while limiting throughput loss to the primary source. In particular, let us denote as μ_0 the policy by which the secondary source never transmits. The optimization problem can be written as the following infinite horizon constrained MDP [18]:

$$\widehat{\mu} = \arg\min_{\mu} \mathcal{J}_{S}(\mu) \text{ s.t. } \mathcal{W}_{P}(\mu_{0}) - \mathcal{W}_{P}(\mu) \leq \sigma$$
 (9)

where $\mathcal{J}_S(\mu)$ is the average cost incurred by the secondary source² and can be computed as $\mathcal{J}_S(\mu) = L_S/\tau - \mathcal{W}_S(\mu)$. For two arbitrary policies μ_1 and μ_2 , we refer to $\mathcal{W}_P(\mu_1) - \mathcal{W}_P(\mu_2)$ as $\Delta(\mu_1, \mu_2)$. Note that the throughput loss can also be defined as the difference of average costs $\Delta(\mu_1, \mu_2) = \mathcal{J}_P(\mu_2) - \mathcal{J}_P(\mu_1)$.

According to [19], the solution of the optimization problem (9) is a past-independent randomized policy. Moreover, as the number of independent constraints is equal to one, then, in the optimal stationary policy, randomization occurs in at most one state, i.e., the map is either deterministic in all states or deterministic in all states except one in which the decision is randomized.

III. MARKOV CHAIN AND OPTIMIZATION OF THE NETWORK

The state of the network can be modeled as a homogeneous Markov process $\Theta = \{\Theta_1, \Theta_2, \ldots\}$ taking values in the state space $\mathcal{X} = \{0, 1, \ldots, T\}$, where $\Theta_n = 0$ and $\Theta_n = \theta, 1 \leq \theta \leq T$, correspond to S_P not accessing the channel and performing the θ th transmission of a packet in slot n, respectively. Since the secondary source is backlogged and transmits each packet only once, its status is the same in each slot and we do not need to account for it in the model. A graphical representation of the Markov chain is depicted in Fig. 2(b).

 $^2 In$ the following, we will denote by $\mathcal{J}_{\rm P}$ the analogous cost defined for the primary source.

It can be shown that the solution of the problem in (9) is a randomized past-independent stationary policy [19]. Thus, the policy μ maps the state of the network $\theta \in \mathcal{X}$ to the probability that the secondary source takes the actions in \mathcal{U} . The action selected in the time slot n is referred to as $u_n \in \mathcal{U}$ in the following. We define $\mu(\theta, u)$ as the probability that the secondary source takes action u when the network is in state θ . As \mathcal{U} is binary, the policy can be defined as the vector $\underline{\kappa} = \{\kappa_0, \kappa_1, \ldots, \kappa_T\}$, where $0 \le \kappa_\theta = \mu(\theta, 1) \le 1$. κ_θ is the probability that the secondary source accesses the channel when the network is in state θ . Policy μ_0 corresponds to the all-zero vector $\underline{0}$. For the sake of simplicity, in the following, $L_P/\tau = L_S/\tau$ is set to one.

The transition probability from state Θ to state Θ' , Θ , $\Theta' \in \mathcal{X}$, conditioned on the action u is defined as

$$\zeta_u(\Theta, \Theta') = \mathcal{P} \left\{ \Theta_{n+1} = \Theta' \middle| \Theta_n = \Theta, u_n = u \right\}$$
 (10)

and does not depend on n. Note that since the policy μ is past-independent, the stochastic process which models the temporal evolution of the network is a Markov process. We remark that due to the mutual interference, the probability that the Markov process transitions from one state to another in the state space depends on the action taken by the secondary user.

The transition matrix of the chain, which collects the transition probabilities $\zeta^{\mu}(\Theta, \Theta')$, is

$$\begin{bmatrix} 1-\alpha & \alpha & 0 & \dots & 0 \\ (1-\alpha)(1-\rho_1) & \alpha(1-\rho_1) & \rho_1 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ (1-\alpha)(1-\rho_{T-1}) & \alpha(1-\rho_{T-1}) & 0 & \dots & \rho_{T-1} \\ (1-\alpha) & \alpha & 0 & \dots & 0 \end{bmatrix}$$
(11)

where ρ_{θ} depends on the transmission probability of the secondary user κ_{θ} , and represents the failure probability of the primary source in state t conditioned on the transmission strategy. Thus, from state 0, the network moves to state 1 (new packet in the buffer of $S_{\rm P}$) with probability α , and remains in 0 otherwise. In each state θ , $1 \le \theta < T$, the network moves to state $\theta + 1$ if a failure occurs, while it returns to 0 or 1 according to the arrival probability α if the primary packet is successfully delivered. From state T, the transmission of the current packet is terminated regardless of failure or success, and thus the network returns to state 0 and 1 with probability $1-\alpha$ and α , respectively.

The steady-state distribution π_{μ} of the Markov chain is the solution of the following system of equations:

$$\pi_{\mu}(0) = (1 - \alpha)\pi_{\mu}(0) + (1 - \alpha)\sum_{t=1}^{T-1} (1 - \rho_t)\pi_{\mu}(t) + (1 - \alpha)\pi_{\mu}(T)$$

$$\pi_{\mu}(\theta) = \rho_{\theta}\pi_{\mu}(\theta - 1) \quad \text{for } 2 \le \theta \le T$$
(12)

with the normalization condition $\pi_{\mu}(1)=1 - \pi_{\mu}(0) - \sum_{t=2}^{T} \pi_{\mu}(t)$.

As intuition suggests, states corresponding to a larger number of transmissions are hit by the process a smaller number of times with respect to those associated with a smaller number of transmissions of the same packet, i.e., $\pi_{\mu}(\theta+1) \leq \pi_{\mu}(\theta)$ for any $\theta > 0$. In fact, the process enters state $\theta + 1$, $\theta > 0$, only by passing through θ . This can be observed in (12), by which we get

 $\pi_{\mu}(\theta) = \pi_{\mu}(1) \prod_{i=1}^{\theta-1} \rho_i$, for $2 \leq \theta \leq T$, where $\prod_{i=1}^{\theta-1} \rho_i \leq 1$. The steady-state distribution is

$$\pi_{\mu}(0) = \frac{1 - \alpha}{1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} \rho_{i}}$$

$$\pi_{\mu}(1) = \frac{\alpha}{1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} \rho_{i}}$$

$$\pi_{\mu}(\theta) = \frac{\alpha \prod_{i=1}^{\theta-1} \rho_{i}}{1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} \rho_{i}} \quad \text{for } 2 \le t \le T \quad (13)$$

The average cost of the primary source can be rewritten as $\mathcal{J}_{P}(\mu) = \sum_{\theta \in \mathcal{X}} \pi_{\mu}(\theta) \tilde{\gamma}_{P}(\mu, \theta)$, where $\tilde{\gamma}_{P}(\mu, \theta)$ is the average cost collected by the primary source in state θ under policy μ .

The average cost difference $\Delta(\mu_1, \mu_2)$ is equal to

$$\Delta(\mu_1, \mu_2) = \sum_{\theta \in \mathcal{X}} (\pi_{\mu_2}(\theta) \tilde{\gamma}_{P}(\mu_2, \theta) - \pi_{\mu_1}(\theta) \tilde{\gamma}_{P}(\mu_1, \theta)) \quad (14)$$

Different policies result in different average costs collected in each state, but correspond to different steady-state distributions as well.

The average cost in state θ can be computed as

$$\tilde{\gamma}_{P}(\mu, \theta) = \sum_{\theta_{1} \in \mathcal{X}} \sum_{u \in \mathcal{U}} \mu(\theta, u) \gamma_{P}(\theta, \theta_{1}) \zeta_{u}(\theta, \theta_{1})$$
 (15)

where $\gamma_{\rm P}(\theta,\theta_1)$ is the cost incurred by $S_{\rm P}$ during the transition from θ to θ_1 . The cost $\gamma_{\rm P}(\theta,\theta_1)$ is equal to zero for the transitions in which a packet is successfully delivered and to $L_{\rm P}/\tau\!=\!1$ when a packet incurs failure.³ Note that when $\theta\!=\!T$, the cost of any transition is $\rho_T L_{\rm P}/\tau\!=\!\rho_T$. The throughput can be similarly defined as the sum of the steady-state distribution weighed by the average rewards $\tilde{\omega}(\mu,\theta)\!=\!L_{\rm P}/\tau\!-\!\tilde{\gamma}_{\rm P}(\mu,\theta)\!=\!1\!-\!\tilde{\gamma}_{\rm P}(\mu,\theta)$. Analogous definitions can be stated for the secondary link. In the following, with a slight abuse of notation, we denote the average cost and reward from state θ when the action u is selected as $\tilde{\gamma}(u,\theta)$, $\tilde{\omega}(u,\theta)$, respectively.

The average costs in the various states are trivially $\tilde{\gamma}_{\rm P}(\mu,0) = L_{\rm P}/\tau = 1$ and $\tilde{\gamma}_{\rm P}(\mu,\theta) = \rho_{\theta}L_{\rm P}/\tau = \rho_{\theta}$, $1 \le \theta \le T$. The average cost of the primary source can be, thus, written as

$$\mathcal{J}_{P}\left(\underline{\rho}\right) = \frac{(1-\alpha) + \alpha \sum_{t=1}^{T} \prod_{i=1}^{t} \rho_{i}}{1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} \rho_{i}} \le 1 \tag{16}$$

with $\rho = \{\rho_1, \ldots, \rho_T\}$.

The interference by the secondary source in a certain state has two effects on the performance of the primary source.

- 1) if θ >0, interference increases the instantaneous cost collected in that state by the primary source;
- 2) if $0 < \theta < T$, interference increases the probability that the process moves to $\theta+1$.

Clearly, transmission by $S_{\rm S}$ in state 0 does not have any effect on the primary source, while in T it only increases the cost associated with that state, as the packet being served by $S_{\rm P}$ is discarded after this transmission.

As observed in Section I, if the secondary source transmits in a state θ , with $0 < \theta < T$, the average number of transmissions of the packets of the primary source increases. This means that the fraction of time spent by the primary source in the idle state

 $^{^3}$ We recall that, without any loss of generality, $L_{\rm P}/\tau$ is set to unity.

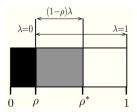


Fig. 3. Graphical representation of the failure probability increasing factor associated with the interference generated by the secondary source to the primary source's transmission.

decreases. By interfering with the primary source, the secondary source is then decreasing the number of idle slots, that is, the white spaces reduce.

The average failure probability of the primary source in state θ >0, conditioned on the policy, is ρ_{θ} = $(1-\kappa_{\theta})\rho+\kappa_{\theta}\rho^*$. In fact, when in state θ , the primary source incurs a failure probability equal to ρ^* if the secondary source transmits and equal to ρ if the secondary source does not transmit.

In order to provide a more intuitive explanation of the dependence between the decoding performance degradation at $D_{\rm P}$ and transmission by $S_{\rm S}$, we define the failure probability increasing factor λ , such that $\rho^* = \rho + (1-\rho)\lambda$. Thus, λ determines the impact of transmission by the secondary source on the decoding probability at the primary receiver: the larger λ , the closer to one the probability of failure. In particular, for $\lambda = 0$ and $\lambda = 1$, the failure probability at the primary source is ρ and 1, respectively (see Fig. 3 for a graphical representation). The resulting average failure rate in θ is $\rho_{\theta} = \rho + (1-\rho)\lambda\kappa_{\theta}$. Note that λ parameterizes the difference between the failure probability with and without interference from the secondary user and does not presume the use of a linear model for the failure probability as a function of the interference power.

Consider κ_0 , the probability that the secondary source transmits when the primary source is idle. As increasing κ_0 does not affect the cost to the primary user (which is idle), the optimal value for κ_0 is one.⁴ Thus, in the sequel, we set κ_0 =1.

The average cost collected by the primary source is then

$$\mathcal{J}_{P}(\underline{\kappa}) = \frac{(1-\alpha) + \alpha \sum_{t=1}^{T} \prod_{i=1}^{t} (\rho + (1-\rho)\lambda \kappa_{i})}{1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} (\rho + (1-\rho)\lambda \kappa_{i})}$$
(17)

The average throughput achieved by the secondary source, also referred to as the *reward* in the following, can be computed as in (18), shown at the bottom of the page.

⁴This may not hold if we consider more complex networks or energy consumption metrics. This is left for future research.

The optimization problem (9) is equivalent to the following LP [19]:

$$\widehat{Z} = \arg \max_{Z} \sum_{\theta \in \mathcal{X}} \sum_{u \in \mathcal{U}} \widetilde{\omega}(u, \theta) z_{u}(\theta)$$
s.t.
$$\sum_{\theta \in \mathcal{X}} \sum_{u \in \mathcal{U}} \widetilde{\gamma}(u, \theta) z_{u}(\theta) \leq \sigma + \mathcal{J}_{P}(\mu_{0})$$

$$\sum_{\theta \in \mathcal{X}} \sum_{u \in \mathcal{U}} z_{u}(\theta) = 1$$

$$\sum_{u \in \mathcal{U}} z_{u}(\theta_{1}) = \sum_{\theta \in \mathcal{X}} \sum_{u \in \mathcal{U}} z_{u}(\theta) \zeta_{u}(\theta, \theta_{1}), \ \forall \theta_{1}$$

$$z_{u}(\theta) \geq 0, \ \forall u, \ \theta$$

$$(19)$$

where $Z=\{z_u(\theta)\}_{\theta\in\mathcal{X},u\in\mathcal{U}}$, and $z_u(\theta)$ represents the joint probability that the Markov chain is in state θ and action u is selected. The first constraint bounds the maximum performance loss of the primary user, while the others force the solution to be a valid stationary distribution for the Markov chain.

The LP defined previously thus optimizes the steady-state distribution of state—action pairs. The involved expression of the average reward and cost functions defining the objective and constraint of the original problem are thus translated into *linear combinations* of the optimization variables.⁵ The condition for optimality is that the Markov chain under all the policies is unichain [20], i.e., it has a single recurrent class and an arbitrary number of transient classes. This property holds, in our case, for any policy and any set of parameters as defined throughout this paper.

The optimal policy is then $\widehat{\mu}(\theta,a)=\widehat{z}_a(\theta)/(\sum_u\widehat{z}_u(\theta))$ if $\sum_{u\in\mathcal{U}}z_u(\theta)=1$, i.e., θ is recurrent. If $\sum_{u\in\mathcal{U}}z_u(\theta)=0$, i.e., θ is transient, then the map in θ is $\widehat{\mu}(\theta,a)=1$ for a randomly chosen $a\in\mathcal{U}$ and $\widehat{\mu}(\theta,a)=0$ otherwise. In the model at hand, which considers a binary action whose randomization corresponds to the probability that the secondary source transmits given that the transmission probability κ_θ simply corresponds to $\mu(\theta,1)$. Note that $\sum_{\theta=0}^T z_1(\theta)$ is the total fraction of time in which the secondary source transmits.

It is also shown in [19] that the number of *randomizations*, i.e., the number of states in which the policy is nondeterministic, is equal to or smaller than the number of independent constraints in (19). Thus, in the model at hand, the optimal policy found via the aforementioned LP is nondeterministic in at most one state and the optimal vector $\hat{\underline{\kappa}}$ is a vector with N_1 ones, N_0 zeros, and N_r elements in (0,1), with $N_1+N_0+N_r=T+1$, $0\leq N_1\leq T$, $0\leq N_0\leq T$, and $N_r=1$ or 0. The space of the vectors described

⁵A constraint on the failure probability can be formalized as a linear constraint as well through straightforward manipulation.

$$W_{S}(\underline{\kappa}) = \pi_{\underline{\kappa}}(0)\kappa_{0}(1-\nu) + \sum_{\theta=0}^{T} \pi_{\underline{\kappa}}(\theta)\kappa_{\theta}(1-\nu^{*}) = \frac{(1-\alpha)(1-\nu)+(1-\nu^{*})\lambda_{S}\alpha\sum_{t=1}^{T} \kappa_{t}\prod_{i=1}^{t-1}(\rho+(1-\rho)\lambda\kappa_{i}))}{1+\alpha\sum_{t=1}^{T-1}\prod_{i=1}^{t}(\rho+(1-\rho)\lambda\kappa_{i})}$$
(18)

by the aforementioned conditions is denoted in the following with \mathcal{M}_r .

In Section IV, we will show that, if $\nu^* = \nu$, the optimal policy $\underline{\widehat{\kappa}}$ has a precise structure that enables its calculation through a simple algorithm, thereby avoiding the need to solve the linear problem stated earlier. In particular, the optimal policy concentrates transmissions by the secondary source in the first transmissions of the primary source packets. Therefore, the N_1 unit elements and the N_0 zero elements are the first N_1 and the last N_0 elements of the vector $\underline{\widehat{\kappa}}$, respectively. If $N_1 + N_0 = T - 1$, then randomization occurs at the $N_1 + 1$ th state, otherwise the policy is deterministic.

As a side comment, we observe that in a pure collision scenario, where the failure a policy such that $\kappa_0{=}1$ and $\kappa_\theta{=}0$, $0{<}\theta{\leq}T$, is optimal. This is the white spaces approach. In fact, if $\rho_{\rm P}^*$ and $\rho_{\rm S}^*$ are both set to one, the secondary source gains nothing when transmitting concurrently with the primary source, while increasing the cost of the latter. In general, if the secondary source bases its strategy on channel sensing only it can distinguish between an idle slot $(\theta{=}0)$ and a non-idle slot $(0{<}\theta{\leq}T)$. The resulting strategy assigns the transmission probabilities $\kappa_0{=}1$ and $\kappa_\theta{=}\kappa$, $\forall 0{<}\theta{\leq}T$. We will show through numerical results that this policy is suboptimal.

IV. OPTIMAL TRANSMISSION STRATEGY FOR THE Z-INTERFERENCE CHANNEL

In this section, we address the structure of the optimal transmission strategy in the particular case in which $\nu^* = \nu$, that is, the transmission by the primary source does not affect the successful decoding probability of the packet of the secondary source by the secondary receiver.

This assumption can be referred to the well-known Z-interference channel framework, where the interference link between the primary source and the secondary destination is removed. We observe that this does not mean that the interference channel between the primary source and the secondary destination is simply removed. For instance, this model also fits the case in which $g_{PS} \gg g_{SS}$ with high probability, or the primary source transmits with a rate R_P sufficiently low to allow the secondary destination to decode and cancel the interference from the primary source with high probability.

In this case $\nu^*=\nu$, and thus, the failure probability at the secondary destination does not influence the solution of the optimization problem. In fact, the success probability $1-\nu$ only represents a scaling factor for the reward achieved by the secondary source. Thus, in the following, with $\mathcal{W}_S(\underline{\kappa})$ we refer to the normalized reward $\mathcal{W}_S(\underline{\kappa})/(1-\nu)$. We remark that the optimal policy $\underline{\hat{\kappa}}$ when maximizing the reward or the normalized reward of the secondary source is the same, and that the throughput is simply the normalized reward multiplied by the success probability.

A. Structure of the Optimal Policy

In the following, we show that the optimal transmission policy for the secondary source when $\nu=\nu^*$ has a specific structure. The transmission strategy maximizing the throughput

of the secondary source, given the constraint on the primary source's throughput loss, concentrates interference in the first transmissions of each of the packets sent by the primary source. The policy has the structure described in Theorem 5, where the secondary user transmits with probability 1 in states $\theta < N_1$, probability $\kappa_{N_1} \in [0,1]$ in state N_1 , and probability equal to zero in states $\theta > N_1$. The values of N_1 and κ_{N_1} are functions of the parameters of the system and of the throughput constraint. It can be shown that the same structure applies if the constraint is on the failure probability of primary source's packets. The definitions and proof for this last case are provided in Section IV-B.

As discussed earlier, interference from the secondary source in different states has a different effect. In fact, if the secondary source increases its transmission probability in state j, with 0 < j < T, it also increases the average failure probability ρ_j . This means that the primary source fails more often in the jth transmission of a packet. Therefore, the Markov process hits more frequently the states with indices larger than j, and less frequently all the other states, that is, the steady-state probability $\pi_{\mu}(t)$ of the states t > j grows, while the same probability associated with the states $t \le j$ decreases.

Moreover, as observed earlier, if j < r, then the steady-state probability associated with state j is larger than that of state r. Thus, if the secondary source increases its transmission probability in state j, it increases the overall level of interference more than if the same increase is applied to state r. Thus, the state in which the interference is increased influences the bias on the stochastic process of the primary source, due to the activity of the secondary source as well as the overall cost incurred by the former. The normalized reward of the secondary source counts the fraction of slots in which the secondary source transmits. Since $\pi_{\mu}(i) > \pi_{\mu}(r)$, the overall reward grows more if the transmission probability is increased in state j than if the same increase is applied to state r > j. On the other hand, note that if κ_j is increased, then $\pi_{\mu}(j)$ decreases. Nevertheless, we will shown in the following that, if $\nu^* = \nu$, an increased transmission probability in any of the states results in a larger secondary source's throughput.

The main intuition behind the structure of the optimal transmission policy is that, when considering the same increase of the transmission probability, the reward of the secondary source grows faster than the cost of the primary source. Moreover, the difference between the increase of the reward of the secondary source and the increase of the cost of the primary source grows much faster if the transmission probability is increased in state j, with respect to the same quantity measured if the transmission probability is increased in state r > j. Based on these observations, it is possible to show that among the set of the transmission policies resulting in the same cost for the primary source, the one that most concentrates the interfering transmissions in the first transmissions of the primary source's packets achieves the optimal throughput.

We first state the following theorems:

Theorem 1: $\mathcal{J}_{P}(\underline{\kappa})$ is a strictly increasing function of κ_{θ} , with $\theta > 0.6$

⁶We remark that the cost is independent of κ_0 , whose value has been set to one by assumption.

Theorem 2: $W_S(\underline{\kappa})$ is a strictly increasing function of κ_{θ} , with $\theta > 0$.

Formal proofs of these theorems are provided in Appendixes A and B.

Theorem 1 states that the cost of the primary source increases as the fraction of slots in which the secondary user accesses the channel gets larger. This is rather intuitive, as a larger amount of interference cannot result in a larger throughput for the interfered link, at least in the framework considered herein.

Theorem 2 states that the average throughput of the secondary source increases as the fraction of slots in which it accesses the channel gets larger. Although this result also agrees with intuition, it must be observed that transmission by the secondary source in a certain slot also modifies the steady-state distribution of the Markov chain for the primary source. For instance, the steady-state distribution of state 0, in which the secondary source can always transmit, decreases as κ_{θ} gets larger, with $0 < \theta < T.^7$ However, the theorem states that the gain outweighs the potential loss under the assumptions on the decoding failure at the secondary receiver stated earlier.

The previously stated theorems guarantee that the optimal policy lies in the space of policies where the constraint on the primary throughput loss of (9) is active, i.e., $\Delta(\underline{0},\underline{\hat{\kappa}}) = \sigma$, unless $\Delta(\underline{0},\underline{1}) \leq \sigma$, where $\underline{1}$ is a T+1—long vector whose elements are all ones. In fact, in this latter case, the secondary source transmits in all the slots with probability one, and, if this policy results in a cost for the primary source smaller than the maximum admitted, then a policy that activates the constraint does not exist. Moreover, under the policy $\underline{1}$, the secondary user achieves the maximum possible throughput, i.e., $\mathcal{W}_S(\underline{1}) = L_S/\tau = 1.8$ Thus, if $\underline{1}$ is admissible, then it is also optimal.

Let us consider now the case $\Delta(\underline{0},\underline{1}) > \sigma$ and define \underline{u}_i as a T+1-long vector of all zeros except for the ith element that is equal to one, $0 \le i \le T$.

Consider a policy $\underline{\kappa}'\neq\underline{1}$ such that $\Delta(\underline{0},\underline{\kappa}')<\sigma$. Since \mathcal{J}_P is continuous, there exist $\delta>0$ and $0\leq j\leq T$ such that $\Delta(\underline{0},\underline{\kappa}'')\leq\sigma$, where $\underline{\kappa}''=\underline{\kappa}'+\underline{u}_j\delta$. Due to Theorem 2, the reward achieved by policy $\underline{\kappa}''$ is larger than that achieved by policy $\underline{\kappa}'$, i.e., $\mathcal{W}_S(\underline{\kappa}'')>\mathcal{W}_S(\underline{\kappa}')$. Note that for any $\delta>0$, we also have $\mathcal{J}_P(\underline{\kappa}'')>\mathcal{J}_P(\underline{\kappa}')$, and thus $\sigma-\Delta(\underline{0},\underline{\kappa}'')<\sigma-\Delta(\underline{0},\underline{\kappa}')$.

Thus, for any policy $\underline{\kappa}'$ resulting in a maximum performance loss below σ , there exists an admissible policy $\underline{\kappa}''$ such that the secondary user achieves an improved throughput, while the cost of the primary user increases. Any policy $\underline{\kappa}'$ such that $\Delta(\underline{0},\underline{\kappa}')$ is strictly smaller than σ is thus nonoptimal.

As a consequence of the previous statements, if the problem in (9) is feasible, then the optimal policy $\underline{\hat{\kappa}}$ lies in the space $\mathcal{M}_{\sigma} = \{\underline{\kappa} : \Delta(\underline{0}, \underline{\kappa}) = \sigma\} \cup \{\underline{1}\}.$

We now formalize the intuition discussed earlier by stating the following theorem:

Theorem 3: Consider a policy $\underline{\kappa}$ such that $\kappa_j = \kappa_r$ and $\kappa_{\theta} = 0$, $\forall \theta > r$ and with $0 < j < r \le T$.

 7 In the case θ =T, transmission by the secondary source does not modify the steady-state distribution.

⁸We recall that, throughout this section, we normalize the throughput of the secondary source normalized to the success probability $1-\nu=1-\nu^*$.

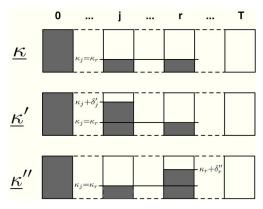


Fig. 4. Policies κ' and κ'' as defined in Theorems 3 and 4.

Define the two policies $\underline{\kappa}'$ and $\underline{\kappa}''$ as $\underline{\kappa}' = \underline{\kappa} + \underline{u}_j \delta_j'$ and $\underline{\kappa}'' = \underline{\kappa} + \underline{u}_r \delta_r''$, with $0 < \delta_j' \le 1 - \kappa_j$ and $0 < \delta_r'' \le 1 - \kappa_r$ (see Fig. 4 for a graphical representation). If $\mathcal{J}_P(\underline{\kappa}') = \mathcal{J}_P(\underline{\kappa}'')$, then $\mathcal{W}_S(\underline{\kappa}') > \mathcal{W}_S(\underline{\kappa}'')$.

The proof of the theorem is provided in Appendix C.

Theorem 3 states that, starting from a policy $\underline{\kappa}$ respecting the hypothesis, if the policy obtained by increasing κ_j and the policy obtained by increasing κ_r incur the same average primary source's cost, then, if j < r, the reward associated with the former is larger than the reward associated with the latter. As discussed earlier, this result is due to the difference between the reward and cost increase corresponding to an increased transmission probability in a certain state. This quantity grows faster if the transmission probability is increased in state j with respect to state r, with j < r.

Similarly, it can be shown that if the policy obtained by decreasing κ_j and the policy obtained by decreasing κ_r result in the same average primary source's cost, then the reward achieved by the latter is larger than the reward achieved by the former. Formally:

Theorem 4: Consider a policy $\underline{\kappa}$ such that $\kappa_j = \kappa_r$ and $\kappa_{\theta} = 0$, $\forall \theta > r$ and with $0 < j < r \le T$.

Define the two policies $\underline{\kappa}'$ and $\underline{\kappa}''$ as $\underline{\kappa}' = \underline{\kappa} - \underline{u}_j \delta_j'$ and $\underline{\kappa}'' = \underline{\kappa} - \underline{u}_r \delta_r''$, with $0 < \delta_j' \le \kappa_j$ and $0 < \delta_r'' \le \kappa_r$ (see Fig. 4 for a graphical representation). If $\mathcal{J}_{\mathrm{P}}(\underline{\kappa}') = \mathcal{J}_{\mathrm{P}}(\underline{\kappa}'')$ then $\mathcal{W}_{S}(\kappa') < \mathcal{W}_{S}(\kappa'')$.

The proof of the theorem is provided in Appendix D.

Theorems 3 and 4 are the basis for the derivation of the structure of the optimal policy $\hat{\kappa}$, defined by the following theorem:

Theorem 5: The optimal policy $\widehat{\underline{\kappa}}$ has the following structure:

$$\widehat{\underline{\kappa}} = [\underline{1}_{N_1}, \ \kappa_{N_1}, \ \underline{0}_{N_0}] \tag{20}$$

where $\underline{1}_{N_1}$ and $\underline{0}_{N_0}$ are vectors of N_1 ones and N_0 zeros, respectively, and $0 \le \kappa_{N_1} \le 1$.

Thus, the optimal policy concentrates transmission by the secondary source in those states associated with the first transmissions of primary source's packets. Intuitively, if the interference generated by the primary user's transmission has a small impact on the reception of secondary user's packets, then the difference between the secondary user reward increase and the primary user cost increase corresponding to an increase of the transmission probability in the early retransmissions is positive

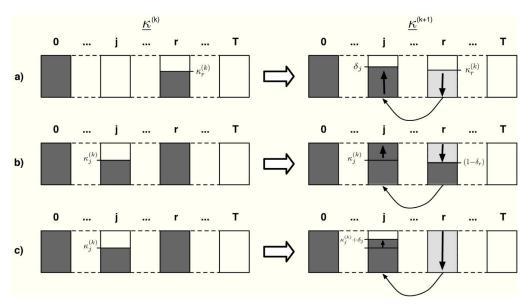


Fig. 5. Policies $\underline{\kappa}^{(k)}$ and $\underline{\kappa}^{(k+1)}$.

and larger than that corresponding to the same increase in the late retransmissions. In fact, the throughput achieved by the secondary user is not affected by the access rate of the primary user, and thus, a transmission probability increase corresponds to a positive reward in all the states. Moreover, the cost increase of the primary user accounts for the fact that additional primary user's retransmission due to secondary user interference takes place in otherwise idle slots with a positive probability, that is, there is a positive probability that retransmissions do not affect the throughput of the primary user. This reduces the cost increase speed in the early retransmissions of primary user packets and results into the unique structure of the optimal transmission policy discussed earlier.

We remark that the optimal transmission strategy of the secondary user is defined under the constraint on the maximum performance loss of the primary user. Therefore, the transmission probability in all the states $\Theta > 0$ is bounded by the constraint. We also observe that the transmission strategies proposed in prior literature addressing cognitive networks do not consider the long-term impact of interference. Therefore, these strategies may fail to guarantee the minimum performance to the primary user in those scenarios in which the primary user implements protocols and mechanisms which react to interference and packet failure.

Theorem 5 has a very intuitive proof, sketched in the following. As observed earlier, if the problem in (9) is feasible, then the policy lies in the space of policies $\mathcal{M}_{\sigma} = \{\underline{\kappa}: \Delta(\underline{0},\kappa) = \sigma\} \cup \{\underline{1}\}$. Moreover, according to [19], the optimal policy is a randomized policy with randomization in at most one state. We recall that the space of transmission probability vectors associated with those policies, i.e., the space of the vectors with N_1 ones, N_0 zeros, and N_r elements in (0,1), with $N_1 + N_0 + N_r = T + 1$, $0 \le N_1 \le T$, $0 \le N_0 \le T$, and $N_r = 1$ or 0, is denoted with \mathcal{M}_r . Therefore, the optimal transmission probability vector lies in the space $\mathcal{M}_r \cap \mathcal{M}_{\sigma}$.

If $\underline{1}$ is admissible, then it is the optimal policy and Theorem 5 holds with N_1 =T+1, N_0 =0, and κ_{N_1} =1.

Assume now that $\underline{1}$ is not admissible, i.e., $\Delta(\underline{0},\underline{1}) > \sigma$. If the optimization problem is feasible, then there exists a policy $\underline{\kappa}^{(1)} \in \mathcal{M}_r \cap \mathcal{M}_\sigma$. Starting from $\underline{\kappa}^{(1)}$, it is possible to construct a sequence of policies $\underline{\kappa}^{(1)},\underline{\kappa}^{(2)},\ldots$ in $\mathcal{M}_r \cap \mathcal{M}_\sigma$ such that $\mathcal{W}_S(\underline{\kappa}^{(k+1)}) > \mathcal{W}_S(\underline{\kappa}^{(k)})$ and converging to the optimal policy $\kappa^{(K)}$, where $\underline{\kappa}^{(k+1)}$ has the structure described in Theorem 5.

Consider a policy $\underline{\kappa}^{(k)} \in \mathcal{M}_r \cap \mathcal{M}_\sigma$ and fix $r = \max\{\theta : \kappa_{\theta}^{(k)} > 0\}$, i.e., r is the largest state with a nonzero transmission probability.

Assume $\kappa_r^{(k)} < 1$, i.e., randomization occurs in state r. Then, the policy in any state $\theta \neq r$ is either $\kappa_\theta^{(k)} = 1$ or $\kappa_\theta^{(k)} = 0$. If $\exists \theta : \theta < r$, $\kappa_\theta^{(k)} = 0$, i.e., the transmission probability in θ is zero, then define $\underline{\kappa}^{(k+1)} = \underline{\kappa}^{(k)} + \underline{u}_j \delta_j - \underline{u}_r \kappa_r^{(k)}$ [see Fig. 5(a)], where $j = \min\{\theta : \kappa_\theta^{(k)} = 0\}$ and with $\delta_j > 0$ such that

$$\mathcal{J}_{P}(\underline{\kappa}^{(k)} + \underline{u}_{i}\delta_{i} - \underline{u}_{r}\kappa_{r}^{(k)}) = \mathcal{J}_{P}(\underline{\kappa}^{(k)}). \tag{21}$$

We observe that such a δ_j always exists, due to the continuity of the cost function, Theorem 1 and the fact that $\partial \mathcal{J}_{\mathrm{P}}(\underline{\kappa})/\partial \kappa_j > \partial \mathcal{J}_{\mathrm{P}}(\underline{\kappa})/\partial \kappa_r$. Note that $\underline{\kappa}^{(k+1)} \in \mathcal{M}_r \cap \mathcal{M}_\sigma$. Moreover, $\mathcal{W}_{\mathrm{S}}(\underline{\kappa}^{(k+1)}) > \mathcal{W}_{\mathrm{S}}(\underline{\kappa}^{(k)})$. In fact, define the policy $\underline{\kappa}^* = \underline{\kappa}^{(k)} - \underline{u}_r \kappa_r^{(k)}$. Thus, $\underline{\kappa}^{(k)} = \underline{\kappa}^* + \underline{u}_r \kappa_r^{(k)}$ and $\underline{\kappa}^{(k+1)} = \underline{\kappa}^* + \underline{u}_j \delta_j$ [see Fig. 5(b)]. Since $\kappa_j^* = \kappa_r^* = 0$, $\kappa_\theta^* = 0$, $\forall \theta > r$, and $\mathcal{J}_{\mathrm{P}}(\underline{\kappa}^{(k)}) = \mathcal{J}_{\mathrm{P}}(\underline{\kappa}^{(k+1)})$, then, according to Theorem 3, $\mathcal{W}_{\mathrm{S}}(\underline{\kappa}^{(k+1)}) > \mathcal{W}_{\mathrm{S}}(\underline{\kappa}^{(k)})$. Note that $\underline{\kappa}^{(k+1)}$ is obtained from $\underline{\kappa}^{(k)}$ by draining transmission probability in state r, and pumping it into state j < r. In fact, $\kappa_j^{(k+1)} = \delta_j > \kappa_j^{(k)} = 0$ and $\kappa_r^{(k)} > \kappa_r^{(k+1)} = 0$.

If $\kappa_r^{(k)} = 1$, then there may exist a state $\theta < r$ such that

If $\kappa_r^{(k)} = 1$, then there may exist a state $\theta < r$ such that $0 < \kappa_\theta^{(k)} < 1$, i.e., the randomization occurs in state θ . If such a state does not exist, i.e., the map in all states $\theta < r$ is deterministic, and there exists instead at least one state $\theta : \theta < r$, $\kappa_\theta^{(k)} = 0$, then fix $j = \min\{\theta : \kappa_\theta^{(k)} = 0\}$. $\underline{\kappa}^{(k+1)}$ is then constructed from $\underline{\kappa}^{(k)}$ as described earlier, and via the same considerations it can be shown that it achieves an improved throughput.

⁹This intuitive inequality, not proved herein, can be derived by using the expression for the partial derivatives reported in Appendix A.

If $\kappa_r^{(k)} = 1$ and $\exists \theta < r : 0 < \kappa_\theta^{(k)} < 1$, then we fix $j = \theta$. If there exists $0 < \delta_r < \kappa_r^{(k)} = 1$ such that

$$\mathcal{J}_{P}(\underline{\kappa}^{(k)} + \underline{u}_{j}(1 - \kappa_{j}^{(k)}) - \underline{u}_{r}\delta_{r}) = \mathcal{J}_{P}(\underline{\kappa}^{(k)})$$
 (22)

i.e., there exists a policy obtained by decreasing the transmission probability in r and setting the transmission probability in j to unity which incurs the same average cost of policy $\underline{\kappa}^{(k)}$; then, $\underline{\kappa}^{(k+1)}$ is defined as $\underline{\kappa}^{(k+1)} = \underline{\kappa}^{(k)} + \underline{u}_j (1 - \kappa_j^{(k)}) - \underline{u}_r \delta_r$ [see Fig. 5(b)]. We then define the policy $\underline{\kappa}^* = \underline{\kappa}^{(k)} + \underline{u}_j (1 - \kappa_j^{(k)})$. Thus, since $\kappa_j^* = \kappa_r^* = 1$, $\underline{\kappa}^{(k)} = \underline{\kappa}^* - \underline{u}_j (1 - \kappa_j^{(k)})$, $\underline{\kappa}^{(k+1)} = \underline{\kappa}^* - \underline{u}_r \delta_r$, $\kappa_\theta^* = 0$, $\forall \theta > r$ and

$$\mathcal{J}_{P}(\underline{\kappa}^{*} - \underline{u}_{j}(1 - \kappa_{j}^{(k)})) = \mathcal{J}_{P}(\underline{\kappa}^{*} - \underline{u}_{r}\delta_{r})$$
 (23)

then, due to Theorem 4

$$W_{S}(\underline{\kappa}^{(k+1)}) > W_{S}(\underline{\kappa}^{(k)}).$$
 (24)

Assume now $\sharp \delta_r : 0 < \delta_r < \kappa_r^{(k)} = 1$ such that

$$\mathcal{J}_{P}(\underline{\kappa}^{(k)} + \underline{u}_{i}(1 - \kappa_{i}^{(k)}) - \underline{u}_{r}\delta_{r}) = \mathcal{J}_{P}(\underline{\kappa}^{(k)})$$
 (25)

i.e., the cost obtained by setting $\kappa_j^{(k)} = 1$ and nulling the transmission probability in state r is larger than $\mathcal{J}_{\mathrm{P}}(\underline{\kappa}^{(k)})$. In this case, the policy $\underline{\kappa}^{(k+1)}$ is defined as $\underline{\kappa}^{(k+1)} = \underline{\kappa}^{(k)} + \underline{u}_j \delta_j - \underline{u}_r$ [see Fig. 5(c)], with $0 < \delta_j < 1 - \kappa_j^{(k)}$ and

$$\mathcal{J}_{P}(\underline{\kappa}^{(k)} + \underline{u}_{i}\delta_{i} - \underline{u}_{r}) = \mathcal{J}_{P}(\underline{\kappa}^{(k)})$$
 (26)

Therefore, the policy $\underline{\kappa}^{(k+1)}$ is obtained from $\underline{\kappa}^{(k)}$ by setting to zero the rth element of the vector and increasing accordingly the jth element. We show in the following that the reward achieved by policy $\underline{\kappa}^{(k+1)}$ is larger than that achieved by policy $\underline{\kappa}^{(k)}$ also in this case. The general problem of optimizing κ_i and κ_r given the transmission probabilities in all the other states (set according to $\underline{\kappa}^{(k)}$) can be seen as a reduced version of the LP (19). The optimal solution of this reduced problem is again a randomized policy with randomization in at most one state, i.e., at least one between κ_i and κ_r is set to either unity or zero. In the case we are considering, the only two reduced policies, corresponding to pairs (κ_i, κ_r) , with at most one randomization activating the constraint on the maximum performance loss are $(\kappa_j^{(k)}, \kappa_r^{(k)})$ and $(\kappa_j^{(k+1)}, \kappa_r^{(k+1)})$ as defined earlier. The solution of the reduced LP is then either $(\kappa_i^{(k)}, \kappa_r^{(k)})$ or $(\kappa_i^{(k+1)}, \kappa_r^{(k+1)})$. Fortunately, Theorem 3 ensures that there exists at least one policy achieving a reward larger than $\underline{\kappa}^{(k)}$ with the same cost. Therefore, $\underline{\kappa}^{(k)}$ is suboptimal, and $\underline{\kappa}^{(k+1)}$ is optimal. Define the policy $\underline{\kappa}^* = \underline{\kappa}^{(k)} - \underline{u}_r (1 - \kappa_i^{(k)})$. Thus, $\kappa_i^* = \kappa_r^* = \kappa_i^{(k)}$. It can be shown that there exists δ_j' , with $0 < \delta_i' < 1 - \kappa_i^{(k)}$ such that

$$\mathcal{J}_{P}(\kappa') = \mathcal{J}_{P}(\kappa^{(k)}) \tag{27}$$

where $\underline{\kappa}' = \underline{\kappa}^* + \underline{u}_j \delta_j'$. Since $\kappa_j^* = \kappa_r^*$, $\kappa_\theta^* = 0$, $\forall \theta > r$, and the aforementioned equalities, according to Theorem 3 we have

$$W_{S}(\underline{\kappa}') > W_{S}(\underline{\kappa}^{(k)}).$$
 (28)

As a consequence, $(\kappa_j^{(k+1)}, \kappa_r^{(k+1)})$ is the optimal solution of the reduced LP introduced earlier, and policy $\underline{\kappa}^{(k+1)}$ achieves the maximum reward given the constraint and once fixed the other transmission probabilities.

In all the cases presented, the transmission probability is drained from state r and pumped into state j < r. Note that it is possible to continue the iterations as long as there exists a pair (j,r): j < r, $\kappa_j^{(k)} < \kappa_r^{(k)}$. If such indices cannot be found, the iterations terminate with the policy $\underline{\kappa}^{(K)} \in \mathcal{M}_r \cap \mathcal{M}_\sigma$. It can be easily seen that the iterations terminate with the unique policy in $\mathcal{M}_r \cap \mathcal{M}_\sigma$ characterized by the structure indicated in Theorem 5, i.e.

$$\underline{\kappa}^{(K)} = [\underline{1}_{N_1}, \, \kappa_{N_1}, \, \underline{0}_{N_0}] \tag{29}$$

where $\underline{1}_{N_1}$ and $\underline{0}_{N_0}$ are vectors of N_1 ones and N_0 zeros, respectively, and $0 \le \kappa_{N_1} \le 1$.

Since from any policy $\underline{\kappa}^{(1)} \in \mathcal{M}_r \cap \mathcal{M}_\sigma$, the iterations produce a policy $\underline{\kappa}^{(K)} \in \mathcal{M}_r \cap \mathcal{M}_\sigma$ such that $\mathcal{W}_{\mathbf{S}}(\underline{\kappa}^{(K)}) \geq \mathcal{W}_{\mathbf{S}}(\underline{\kappa}^{(1)})$, then $\underline{\kappa}^{(K)}$ is the optimal policy, i.e., $\underline{\kappa}^{(K)} = \underline{\hat{\kappa}}$.

Theorem 5, besides unveiling an important feature of the optimal interference control strategy in retransmission-based systems, also has an immediate practical meaning. In fact, the optimal policy can be computed through a simple algorithm that generates a sequence $\underline{\kappa}^{(1)},\underline{\kappa}^{(2)},\ldots$ of at most T policies terminating with $\widehat{\underline{\kappa}}$.

Let us fix $\underline{\kappa}^{(1)} = \underline{1}$. If $\Delta(\underline{0},\underline{1}) \leq \sigma$, then the optimal policy is $\underline{\widehat{\kappa}} = \underline{1}$. Otherwise, if $\Delta(\underline{0},\underline{\kappa}^{(1)} - \underline{u}_T) \leq \sigma$, the algorithm terminates with the optimal policy $\underline{\widehat{\kappa}} = \underline{\kappa}^{(1)} - \underline{u}_T \delta_T$, where δ_T is the unique solution of $\Delta(\underline{0},\underline{\kappa}^{(1)} - \underline{u}_T \delta) = \sigma$. If instead $\Delta(\underline{0},\underline{\kappa}^{(1)} - \underline{u}_T) > \sigma$, the algorithm sets $\underline{\kappa}^{(2)} = \underline{\kappa}^{(1)} - \underline{u}_T$, and continues with the next iteration.

Similarly to the previous step, if $\Delta(\underline{0},\underline{\kappa}^{(2)}-\underline{u}_{T-1})\leq\sigma$, the algorithm terminates with the optimal policy $\underline{\widehat{\kappa}}=\underline{\kappa}^{(2)}-\underline{u}_{T-1}\delta_{T-1}$, where δ_{T-1} is the solution of $\Delta(\underline{0},\underline{\kappa}^{(2)}-\underline{u}_{T-1}\delta)=\sigma$. Otherwise, the algorithm sets $\underline{\kappa}^{(3)}=\underline{\kappa}^{(2)}-\underline{u}_{T-1}$ and so on.

Thus, the algorithm sequentially evaluates the variables κ_j in decreasing order from T and terminates with the optimal policy as soon as it finds the first nonzero element.

The structure of the optimal policy leads to another important observation. Consider a secondary source adopting a sensing approach, such that it always transmits when the channel is sensed idle, and transmits with fixed probability κ when the channel is sensed busy [see Fig. 6(a)]. Thus, the secondary source transmits with probability equal to κ in all the states θ , with $0 < \theta \le T$. We call this strategy *horizontal flooding*, meaning that the secondary source equalizes the transmission probability such that it reaches the same level in all the states in which the primary source transmits.

The optimal transmission strategy defined previously determines the transmission probabilities in the various states of the Markov chain under the constraint on the maximum throughput loss of the primary user. As the constraint becomes tighter, the water level of the secondary user is *drained* from the upper states in Fig. 6(b), corresponding to later retransmissions of primary user's packets in order to reduce the impact of the activity of the secondary user on the primary user's throughput. Note that if

 ϵ =0, the secondary user is always silent unless the primary user is idle. The optimal policy corresponds to a *vertical flooding*, where states with a smaller index are flooded with water, i.e., transmission probability, first [see Fig. 6(b)]. The horizontal approach, while sometimes simpler to implement, is suboptimal due to Theorem 3.

Finally, we observe that the arrival rate at the primary source influences the *aggressiveness* of the secondary source. Clearly, as α decreases, also the average throughput of the primary source decreases, as the fraction of time spent sending packets decreases. As the elements of κ get larger, if α is small, the impact of the increased transmission probability is small, as the secondary source is increasing its access rate in states with low probability. Interestingly, the *fraction* of throughput lost by the primary source decreases as the arrival rate α gets smaller.

B. Constraint on the Average Failure Probability

As shown in the previous section, if the constraint on the average performance loss at the primary transmitter is defined for the throughput loss, then the optimal policy concentrates transmission and interference of the secondary source in the first transmissions of the primary source's packets.

Remarkably, the same structure applies to an analogous optimization problem in which the constraint is defined for the increase of the failure probability of the primary source's packets.

The average cost is trivially the probability that all the transmissions of a packet fail, i.e.

$$\mathcal{J}_{\mathbf{P}}^{\mathbf{fp}}(\underline{\kappa}) = \prod_{t=1}^{T} (\rho + (1 - \rho)\lambda \kappa_{t}). \tag{30}$$

The aforementioned expression for $\mathcal{J}_{P}^{fp}(\underline{\kappa})$ is also obtained by assigning the following average cost to the various states:

$$\tilde{\gamma}_{P}(\underline{\kappa}, \theta) = 0 \ \forall \theta = 0, 1, \dots, T-1$$
 (31)

$$\tilde{\gamma}_{\rm P}(\kappa, T) = (\rho + (1 - \rho)\lambda\kappa_T)/\pi_{\kappa}(1). \tag{32}$$

In fact, recalling the steady-state probabilities provided in (13), the resulting average cost is

$$\mathcal{J}_{P}^{fp}(\underline{\kappa}) = \sum_{\theta=0}^{T} \pi_{\underline{\kappa}}(\theta) \tilde{\gamma}_{P}(\underline{\kappa}, \theta) = \frac{\pi_{\underline{\kappa}}(T)}{\pi_{\underline{\kappa}}(1)} (\rho + (1 - \rho) \lambda \kappa_{T})$$
(33)
$$= \prod_{t=1}^{T} (\rho + (1 - \rho) \lambda \kappa_{t})$$
(34)

Intuitively, the failure probability is the ratio between the fraction of slots in which the process is in state T and a packet fails, i.e., $\pi_{\underline{\kappa}}(T)\rho_T$, 10 and the fraction of slots in which the process starts the transmissions of a new packet, i.e., $\pi_{\underline{\kappa}}(1)$. While the average throughput can be expressed as time average of a sampling function [see (8)], the packet failure probability is then the ratio of the time averages of sampling functions associated with state F and state 1 multiplied by the failure probability in state F.

The cost in state T is, thus, a function of the steady-state distribution. The optimization problem can be reduced to an LP

¹⁰We recall that $\rho_T = (\rho + (1 - \rho)\lambda \kappa_T)$

also in this case. In fact, the constraint on the packet failure probability

$$\frac{\pi_{\kappa}(F)}{\pi_{\kappa}(1)}\rho_{T} \leq \sigma \tag{35}$$

can be rewritten as $\pi_{\underline{\kappa}}(F)\rho_T - \pi_{\underline{\kappa}}(1)\sigma \leq 0$.

Note that the structure of the cost function is significantly different with respect to the throughput case. In fact, under the hypothesis of Theorem 3, while $\mathcal{J}_{\mathrm{P}}(\underline{\kappa}+\underline{u}_{j}\delta)>\mathcal{J}_{\mathrm{P}}(\underline{\kappa}+\underline{u}_{r}\delta)$, in this case the equality holds, i.e., $\mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa}+\underline{u}_{j}\delta)=\mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa}+\underline{u}_{r}\delta)$. Therefore, the overall cost is insensitive to the state in which the secondary source increases the transmission probability. More formally, fix j, r, and δ , with $0< j < r \le T$ and $-\min(\kappa_{j},\kappa_{r}) \le \delta \le \min(1-\kappa_{j},1-\kappa_{r})$, then

$$\mathcal{J}_{P}^{fp}(\underline{\kappa} + \underline{u}_{i}\delta) = \mathcal{J}_{P}^{fp}(\underline{\kappa} + \underline{u}_{r}\delta). \tag{36}$$

Interestingly, while the cost in terms of failure probability is insensitive to the state in which the secondary source increases/decreases the transmission probability, the secondary source's throughput increases faster if the transmission probability is increased in the states with a small index. These considerations result in an overall behavior of the reward/cost tradeoff analogous to that resulting from a definition of the primary source's cost in terms of achieved throughput. Then, Theorems 3 and 4 hold for this definition of the cost. A detailed proof can be found in Appendix E.

Note that again transmission by the secondary source in state 0 does not have any effect on the cost of the primary source, while the reward of the secondary source increases as κ_0 is increased; thus, the optimal value for κ_0 is one.

As $\mathcal{J}_{P}(\underline{\kappa})$, also the average cost $\mathcal{J}_{P}^{fp}(\underline{\kappa})$ is a strictly increasing function of any variable κ_{θ} , with $\theta > 0$. Therefore, for this constraint also, the optimal policy lies in the set $\{\kappa:\Delta(0,\kappa)=\sigma\}\cup\{1\}$.

Since Theorems 3 and 4 hold, then it is possible to construct a sequence of randomized policies achieving an improved reward and converging to the optimal randomized policy defined in Theorem 5. Therefore, the optimal policy has the same structure of the optimal policy found for the previously considered case.

Other constraints, as well as other secondary source's performance metrics, may lead to a different optimal policy structure. For instance, the activity of the secondary source may be limited by a constraint on the average number of transmissions¹¹ of the packets of the primary source. For this metric, the average cost of the primary source is

$$\mathcal{J}_{\mathbf{P}}^{ntx}(\underline{\kappa}) = 1 + \sum_{t=1}^{T-1} \prod_{i=1}^{t} (\rho + (1-\rho)\lambda\kappa_i). \tag{37}$$

Observe that an increased transmission probability in state j increases all the terms of the sum with $t \ge j$.

Similarly to the throughput case, the cost increase associated to an increased transmission probability of the secondary source in state j is larger than the same increase in state r > j. However,

 $^{11}\mbox{This}$ performance metric is sometimes referred to as delay in the technical literature.

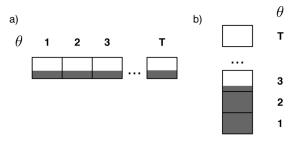


Fig. 6. Graphical representation of (a) the policy in which the secondary source accesses slots in which the primary source transmits with fixed probability, and (b) the optimal policy.

the difference between the average costs associated with the resulting policies may be larger than in the throughput case. Therefore, for some regions of the parameters, the secondary source may be forced to concentrate its transmissions in the last transmissions of the primary source's packets.

The optimization problem admits an analogous formulation, and it is possible to derive the structure of the optimal policy by following a logical procedure entirely similar to the one presented earlier.

V. DISCUSSION FOR THE GENERAL CASE

The structure shown earlier holds if $\nu^*=\nu$, i.e., if primary source's transmission does not alter the decoding probability at the secondary receiver. The reward collected by the secondary source associated with transmission in state 0 or state $\theta>0$ is then the same. This assumption may fit some configurations of the network and receiver capabilities, e.g., the secondary source is much closer to the secondary receiver than the primary source, or the secondary receiver can effectively decode and cancel the signal from the primary source.

However, in general, $\nu^* \ge \nu$. In the following, the case $\nu^* > \nu$ is discussed. This means that if $\kappa_0 = \kappa_t, t > 0$, the average reward of the secondary source in state 0 is larger than the reward in t. In fact, recalling that $\tilde{\omega}_S(\kappa_\theta, \theta)$ is the average reward collected by the secondary source in θ if the transmission probability is κ_θ , we have

$$\tilde{\omega}_S(\kappa_0, 0) = (1 - \nu)\kappa_0 = (1 - \nu)\kappa_t > (1 - \nu^*)\kappa_t = \tilde{\omega}_S(\kappa_t, t) \quad (38)$$

Note that the observations made earlier on the average cost of the primary source remain valid. The average cost is a monotonic increasing function of the transmission probabilities and for $0 < \delta < 1 - \max(\kappa_j, \kappa_T)$ and $0 < j < r \le T$, the following holds:

$$\mathcal{J}_{P}(\underline{\kappa} + \underline{u}_{j}\delta) > \mathcal{J}_{P}(\underline{\kappa} + \underline{u}_{r}\delta)$$
 (39)

for any κ .

Interference due to primary source's transmission at the secondary receiver makes state 0 more *desirable* to the secondary source. As observed earlier, interference increases the average number of transmissions of the primary source's packets. Therefore, the activity of the secondary source reduces the fraction of slots spent by the primary source in the idle state.

Depending on ν^* , ν , ρ^* , and ρ , an increased transmission in a state $\theta > 0$ may decrease the average throughput of the secondary source. Some insights can be extrapolated through the analysis of the case T=2, i.e., the primary receiver transmits the packets at most twice. The average reward of the secondary source is

$$W_{S}(\underline{\kappa}) = \frac{(1-\alpha)(1-\nu)\kappa_{0} + (1-\nu^{*})\alpha(\kappa_{1} + (\rho + (1-\rho)\lambda\kappa_{1})\kappa_{2})}{1 + \alpha(\rho + (1-\rho)\lambda\kappa_{1})}.$$
(40)

 $W_S(\underline{\kappa})$ is a monotonically increasing function of κ_0 , irrespective of κ_1 and κ_2 . In fact

$$\frac{\partial \mathcal{W}_{S}(\underline{\kappa})}{\partial \kappa_{0}} = \frac{(1-\alpha)(1-\nu)}{1+\alpha(\rho+(1-\rho)\lambda\kappa_{1})}$$
(41)

which is trivially positive for any admissible set of parameters. The secondary source's transmission in state 0 does not modify the transition probabilities of the Markov chain. Therefore, any increase of κ_0 corresponds to an increased average reward, and since it does not influence the cost, again it is optimal to set κ_0 =1.

Similar considerations apply to κ_2 , and, more generally, to transmission in state T. We obtain

$$\frac{\partial W_{S}(\underline{\kappa})}{\partial \kappa_{2}} = 1 - \nu^{*} - \frac{1 - \nu^{*}}{1 + \alpha(\rho + (1 - \rho)\lambda\kappa_{1})}$$
(42)

which is positive independently of κ_0 and κ_1 .

Transmission in state 1, instead, alters the transition probabilities, and increases the fraction of time spent by the primary source in state 2, while reducing the time spent in states 0 and 1. The total time spent in the absence of interference from the primary source, which is

$$\frac{1-\alpha}{1+\alpha(\rho+(1-\rho)\lambda\kappa_1)}\tag{43}$$

decreases as κ_1 is increased. If the secondary receiver incurs a high failure probability when decoding a signal interfered by the primary source, the average reward of the secondary source may suffer because of the larger average number of transmissions of the primary source's packets due to transmission in state 1. The derivative $\partial W_S(\underline{\kappa})/\partial \kappa_1$ is shown in (44) at the bottom of the next page, and is positive if ν^* is smaller than the threshold in (45), also shown at the bottom of the next page ¹².

There are, thus, regions of the parameters and transmission probability κ_2 , such that an increased transmission probability in state 1 results in a smaller average reward. Note that an upper bound for $\partial \mathcal{W}_S(\underline{\kappa})/\partial \kappa_1$ is obtained by setting $\kappa_2=1$. In fact, transmission in state 1 increases the steady-state probability of state 2, while transmission in the latter state does not modify the steady-state distribution. If it is easy to see that the threshold in (45), if computed with $\kappa_2=1$, becomes smaller than or equal to 1 for any admissible set of parameters. Therefore, there exists a region of parameters such that the derivative of the average reward with respect to κ_1 is negative. In this region, any throughput-optimal policy sets $\kappa_1=0$. The optimal policy may, therefore, have a different structure than the one shown earlier

 $^{^{12}\}kappa_0$ is set to one in the equations.

¹³In general, an upper bound is obtained by setting $\kappa_T = 1$.

for the case $\nu^* = \nu$. In particular, note that the optimal policy may not belong to the set of policies $\{\underline{\kappa} : \Delta(\underline{0},\underline{\kappa}) = \sigma\} \cup \underline{1}$, i.e., the optimal policy may provide a performance reduction to the primary source smaller than the maximum allowed.

In general, if $\nu^* > \nu$, the mutual interaction between the activity of the secondary source and that of the primary source becomes more involved, and it is hard to provide a structure for the optimal policy. Intuitively, the larger ν^* , the smaller the transmission probabilities in states $\theta=1,2,\ldots,T$, as the secondary source may maximize its own throughput by preserving the steady-state probability of state 0. The same reason may force the secondary source to concentrate its transmissions in the states corresponding to the last transmissions of a primary source's packet.

Numerical results illustrating the aforementioned discussion are shown in Section VII.

VI. ONLINE APPROACHES: STATE OBSERVATION AND MODEL KNOWLEDGE

The resolution of the LP of (19) necessitates the knowledge of the transition probability kernel as well of the cost functions. However, it can be observed that a relatively small number of parameters (the failure probabilities ρ , ρ^* , ν , and ν^* , and the arrival probability α) determine the transition probability matrix and the cost functions. Therefore, the estimation of the statistics of the stochastic process and of the cost functions is faster than in a totally unstructured environment.

The realization of the policy requires the perfect identification of the state of the primary user. In the network considered herein, the estimation of the state within the state space can be obtained by combined channel sensing and packet header decoding. In fact, the secondary user can distinguish state 0 from any other transmission state $\Theta > 0$ by sensing the channel and detecting the presence of a signal. The header of the packets transmitted by the primary user contains their sequence number. Therefore, by decoding the header, the secondary user can count retransmissions of the same packet.

By decoding packet header and ACK/NACK feedback sent by the primary and secondary receivers, the secondary user can estimate the transition probability matrix and the cost functions, as well as identify the state of the primary user. Note that, as the decoding of packet headers and ACK/NACK is crucial to establish communications and instrumental for distributed access mechanisms, these packets are generally strongly encoded and available to all the neighbors of a node.

If the statistics of the Markov chain and the cost functions are unknown, under the assumption of idealized state observation, reinforcement learning algorithms [21] can be employed to iteratively converge to the optimal strategy based on a sample path of observations. The convergence rate of learning algorithms decreases as the state space gets larger. However, techniques which approximate the learned functions may speed up the learning rate [22].

In more complex network scenarios, the exact identification of the state of the network as well as the estimation of the statistics of the stochastic process which models its temporal evolution might be very challenging. As observed in some recent work which extends the framework presented herein to online learning [23], [24], the secondary user may get access to only some features of the state space. For instance, if the primary user stores packets in a buffer, the number of packets in the buffer is hidden to the secondary user. If the secondary user fails to decode the header of a primary user's packet, it may detect the presence of a signal but the retransmission index remains unknown. Another example of hidden state variable is the channel state of the primary links. Channel knowledge would increase the effectiveness of secondary user transmission. In fact, the secondary user can potentially reduce the impact of the generated interference by scheduling transmissions in those time slots in which the link between the primary transmitter and the primary receiver is very strong, and thus interference would not impair packet reception, or very weak, and thus the primary user packet would fail in any case. In the absence of channel state information, the secondary user bases its decision making on the average effect of actions over channel states, that is, the failure probability associated with idleness and transmission. Analogously, if a backoff mechanism is implemented by the primary users to regulate channel access, the secondary user may be unable to distinguish between idleness due to empty buffer or backoff.

In general, by observing the operations of the nodes, it is possible to acquire a significant amount of information about the state of the network. The amount of information collectible by the secondary user depends on the transmission, access, and networking protocols. For instance, the rigid access structure provided by time division multiple access (TDMA) provides more information to the observer than random access. In fact, an idle TDMA slot means that the assigned user has an empty buffer, whereas idleness in random access may be related to the access

If the statistics of the process and the state-observation map are known to the secondary user, then the secondary user can base its decision on a belief vector [25] collecting the maximum likelihood distribution of the real state of the system. Since a priori knowledge of statistics and state-observation map is unrealistic in general scenarios, the approach proposed in [24]

$$\frac{\partial \mathcal{W}_{S}(\underline{\kappa})}{\partial \kappa_{1}} = \frac{\alpha(1 - \lambda(1 - \kappa_{2} - \alpha - \nu + \alpha\nu)(1 - \rho) + \alpha\rho - \nu^{*}(1 + (1 - \rho)\lambda\kappa_{2} + \alpha\rho))}{(1 + \alpha(\rho + (1 - \rho)\lambda\kappa_{1}))^{2}}$$

$$\nu^{*} < \frac{1 - \lambda(-1 + \kappa_{2} + \alpha + \nu - \alpha\nu)(-1 + \rho) + \alpha\rho}{1 + (1 - \rho)\lambda\kappa_{2} + \alpha\rho} \tag{45}$$

$$\nu^* < \frac{1 - \lambda(-1 + \kappa_2 + \alpha + \nu - \alpha\nu)(-1 + \rho) + \alpha\rho}{1 + (1 - \rho)\lambda\kappa_2 + \alpha\rho} \tag{45}$$

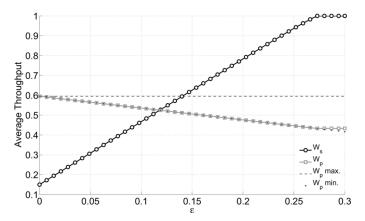


Fig. 7. Throughput as a function of the maximum fraction of throughput loss, where α =0.8, ρ =0.3, ν = ν *=0, and λ =0.3.

is to optimize the distribution of the states in the observation space based on the estimated cost functions, which collects all the possible observations.

According to this discussions, the framework presented in this paper opens many exciting new areas of investigation.

VII. NUMERICAL RESULTS

In this section, numerical results validating the findings and observations made throughout this paper are presented. We recall that α is the probability that the primary source transmits a fresh packet in a slot not allocated to packet retransmission; ρ is the probability that the primary receiver correctly decodes a packet sent by the primary source in a slot in which the secondary source is silent. The failure probability at the primary receiver if the secondary source transmits is $\rho^* = \rho + \lambda(1 - \rho)$, where λ is the failure probability increase. ν and ν^* are the failure probability at the secondary receiver if the primary source is silent and transmits, respectively. A failure probability increase is also defined for the secondary receiver by λ_S such that $\nu^* = \nu + \lambda_S(1 - \nu)$.

In Sections VII-A and VII-B, we present numerical results for $\nu^* = \nu$, i.e., the Z-channel, where the constraint is defined on throughput and failure probability, respectively. Section VII-C presents numerical results for the case $\nu^* > \nu$.

In all the following plots, the maximum number of transmissions of a primary source's packet is fixed to T=4.

A. Constraint on the Primary Source's Throughput, $\nu^* = \nu$

In this section, numerical results for the Z-channel network with a constraint on the throughput loss of the primary source are presented. The performance loss is parameterized through ϵ , defined as the maximum *fraction* of throughput loss of the primary source, i.e., the maximum throughput loss is $\sigma=W_P(\underline{0})\epsilon=(1-\mathcal{J}_P(\underline{0}))\epsilon$.

In Figs. 7 and 8, the throughput and the secondary source's transmission probability are depicted as a function of ϵ . In the picture, \mathcal{W}_{P} max and \mathcal{W}_{P} min correspond to the throughput achieved by the primary source when the secondary source is silent and the minimum throughput of the primary source according to the constraint.¹⁴

¹⁴In this and in the following section, the throughput of the secondary source is normalized to $(1-\nu)$.

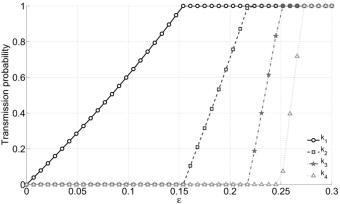


Fig. 8. Transmission probabilities as a function of the maximum fraction of throughput loss, where $\alpha = 0.8$, $\rho = 0.3$, $\nu = \nu^* = 0$, and $\lambda = 0.3$.

The throughput of the secondary source increases as ϵ is increased. A larger ϵ allows the secondary source to interfere more with the primary source. The throughput actually achieved by the primary source decreases according to the increased maximum performance loss allowed, and it can be observed that the policy of the secondary source lowers the throughput of the primary one as much as possible in order to maximize the secondary throughput. When the throughput of the secondary source is equal to one, corresponding to the former transmitting with probability one in every slot, the throughput of the primary source stops decreasing, as the secondary source cannot interfere more.

Fig. 8 shows that the policy of the secondary source follows the structure discussed earlier. Thus, with ϵ =0, the secondary source is allowed to transmit only in the slots where the primary is not accessing the channel (κ_0 =1 and κ_t =0, t>0). As ϵ increases, the transmission probability in state 1, i.e., κ_1 , increases until it reaches unity. Then, κ_2 starts to increase and so on until all the κ_t 's are set to unity.

The rate increase of the various κ_t 's is different. In particular, the rate increase of the κ_t 's corresponding to transmission in states with small indices is smaller than those corresponding to large indices. In fact, interference in the states corresponding to the first transmissions of a packet generates a larger primary source's throughput reduction than interference in the later transmissions. Conversely, the throughput of the primary source gets larger, and so does the maximum throughput loss.

Figs. 9 and 10 show the same quantities as a function of α , i.e., the arrival rate of new packets at the primary source. As expected, the throughput of the secondary source decreases as α increases. Fig. 11 depicts the average number of transmissions of the primary packets for the same parameters, with λ =0.3 and λ =0.9.

A larger α means that the primary source is accessing the channel more often. Therefore, the number of slots in which the secondary source can transmit while meeting the constraint on the throughput loss of the primary source decreases. However, there is another effect of a large α that needs to be considered besides the scarcity of empty slots (in which the secondary source transmits with probability one). In fact, if the probability that a fresh packet is transmitted in an idle slot by the primary source is small, an increased average number of transmissions for each

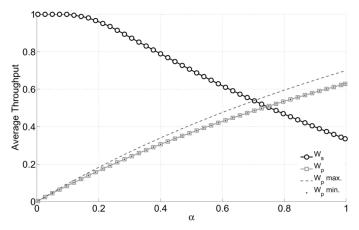


Fig. 9. Throughput as a function of the arrival rate α , where ϵ =0.1, ρ =0.3, ν = ν *=0, and λ =0.3.

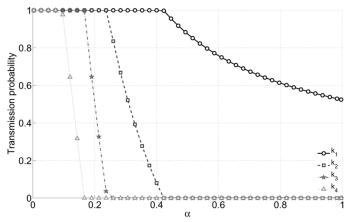


Fig. 10. Transmission probabilities as a function of the arrival rate α , where ϵ =0.1, ρ =0.3, ν = ν *=0, and λ =0.3.

packet has a smaller effect on the throughput of the primary source. The additional retransmissions forced by the interference are likely to substitute for slots in which the primary source would be idle anyway, and thus are the slots in which the primary source would incur the highest possible cost. On the other hand, if α is large, additional retransmissions are performed instead of new packet transmissions that collect an average cost smaller than that of an empty slot.

The relation between α and the interference generated by the secondary source to the primary receiver is illustrated in Figs. 10 and 11. Fig. 10 shows that the throughput trend in Fig. 9 does not only correspond to a smaller fraction of empty slots, but also that the policy in the states $\theta > 0$ is a function of the arrival rate. The fraction of slots in which the secondary source superposes its activity with that of the primary source, normalized by the fraction of slots in which the latter source transmits, decreases as α increases. The explanation for this behavior is illustrated previously. If α is small, the primary source is often idle, and the retransmissions induced by interference generate a smaller loss in the throughput of the primary source. In fact, if α is small, the secondary source is allowed to force more retransmissions (see Fig. 11). Note that, again, the policy follows the structure discussed in the previous section, where the κ_t 's sequentially turn off as the secondary source is forced to reduce the interference.

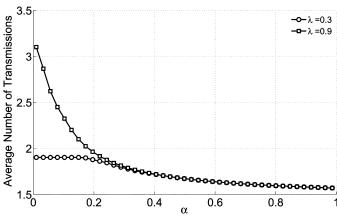


Fig. 11. Average number of transmissions as a function of α , where ϵ =0.1, ρ =0.3, ν = ν *=0, and λ =0.3.

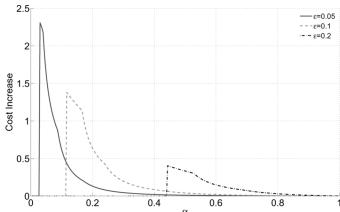


Fig. 12. Cost increase as a function of α , where ϵ =0.1, ρ =0.3, ν = ν *=0, and λ =0.3.

Finally, Fig. 12 shows $(\mathcal{J}_{S}(\underline{\widehat{\kappa}}_{hf}) - \mathcal{J}_{S}(\underline{\widehat{\kappa}}))/\mathcal{J}_{S}(\underline{\widehat{\kappa}})$ as a function of the arrival rate α , where $\mathcal{J}_{S}(\widehat{\underline{\kappa}}_{h,f})$ is the optimal cost for the horizontal flooding approach.¹⁵ The optimal transmission probabilities for the horizontal flooding approach are numerically found via an LP slightly more involved than that discussed herein. Thus, the curves represent the fraction of cost increase when horizontal flooding is adopted instead of vertical flooding. When α is sufficiently small, the cost increase is zero, as both approaches transmit in all states with probability one. As α increases, the secondary source is forced to reduce the fraction of time in which it transmits in both vertical and horizontal flooding. In the former case, the secondary source starts decreasing the transmission probability in state T, while in the latter, the transmission probability is reduced in all states t>0. However, as soon as $\underline{1}$ becomes inadmissible, in order to meet the constraint on the maximum throughput loss, the horizontal approach is forced to reduce the average transmission time of the secondary source much more quickly than the vertical approach. Then, as the arrival rate α is further increased, the cost increase diminishes, since the advantage due to the concentration of the interference in states with smaller indices vanishes. In fact, vertical flooding improves the delivery probability of a packet at the expense of a larger average number of transmissions.

¹⁵If $\mathcal{J}_{\mathrm{S}}(\widehat{\underline{\kappa}}_{h\,f}) = \mathcal{J}_{\mathrm{S}}(\widehat{\underline{\kappa}}) = 0$, the ratio is assumed to be equal to zero.

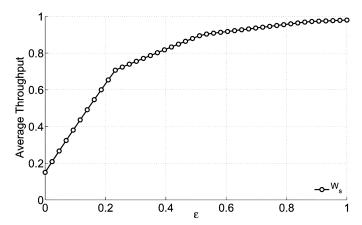


Fig. 13. Throughput of the secondary source as a function of the maximum performance loss ϵ , where α =0.8. ρ =0.3, ν = ν *=0, and λ =0.1.

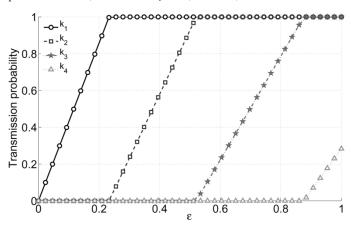


Fig. 14. Transmission probabilities as a function of the maximum performance loss ϵ , where α =0.8. ρ =0.3, ν = ν^* =0, and λ =0.1.

B. Constraint on the Primary Source's Failure Probability,

In this section, results for the optimization problem with a constraint defined on the failure probability of the primary source's packets are presented. ¹⁶ The activity of the secondary source increases the failure probability, and the maximum failure probability increase allowed is given by σ . This increase σ is again parameterized through ϵ , defined as the maximum relative failure probability increase, that is, $\sigma = \mathcal{J}_{S}^{\text{fp}}(\underline{0})(1+\epsilon)$.

Figs. 13–15 show the throughput of the secondary source, the secondary source transmission probability, and the failure probability as a function of ϵ , respectively. In Fig. 15, \mathcal{J}_{\min} is the failure probability associated with an idle secondary source, and \mathcal{J}_{\max} is the maximum failure probability according to the constraint.

Intuitively, the throughput as well as the overall fraction of slots in which the secondary source transmits increases as the maximum failure probability of the primary source's packets increases. The transmission strategy of the secondary source follows the structure discussed throughout this paper. As the constraint becomes less stringent, first transmission in state 0 is increased, then transmission in state 1 and so on, until the secondary source transmits with probability one in all states.

 16 We remark that by failure probability of a packet, we refer to the probability that all the T transmissions fail.

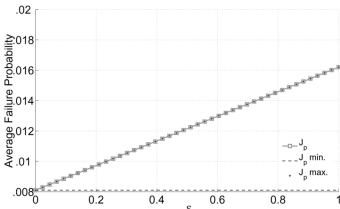


Fig. 15. Packet failure probability as a function of the maximum performance loss ϵ , where $\alpha = 0.8$, $\rho = 0.3$, $\nu = \nu^* = 0$, and $\lambda = 0.1$.

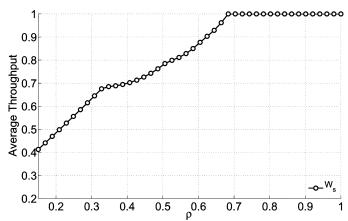


Fig. 16. Throughput of the secondary source as a function of the failure probability of the primary source in the absence of interference ρ , where α =0.8. ρ =0.3, ν = ν *=0, and λ =0.1.

Figs. 16–18 provide the same metrics of the previous figures as a function of the failure probability of the primary source's transmissions ρ . Note that the failure probability of the packets of the primary source if the secondary source is always idle is ρ^T .

The throughput of the secondary source as well as the transmission probabilities in the states $\theta > 0$ increase as ρ becomes larger. We observe the following.

- 1) the maximum $\sigma = \rho^T (1+\epsilon)$ polynomially increases with ρ . This means that the constraint becomes less stringent as ρ increases (see Fig. 18).
- 2) the primary source transmits in a larger fraction of slots as ρ increases, due to a larger average number of retransmissions. The secondary source has fewer empty slots in which to transmit without interfering with the primary source.

If ρ is small, and the secondary source keeps idle, the primary source transmits in a fraction of slots close to α . As ρ gets larger, the primary source increases the fraction of slots in which it transmits, because of the retransmissions. Nevertheless, the constraint becomes less stringent as ρ increases. In fact, the maximum allowed failure probability is $\sigma = \rho^T (1+\epsilon)$. Therefore, as ρ increases, the secondary source can increase its activity in states $\theta > 0$. The tradeoff between those two effects determines the optimal throughput achieved by the secondary

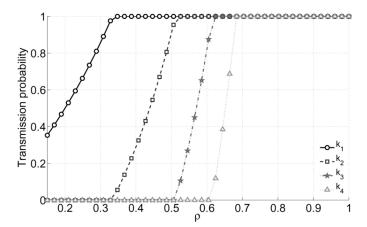


Fig. 17. Transmission probabilities as a function of the failure probability of the primary source in the absence of interference ρ , where α =0.8. ρ =0.3, ν = ν *=0, and λ =0.1.

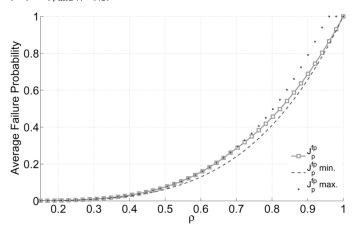


Fig. 18. Packet failure probability as a function of the failure probability of the primary source in the absence of interference ρ , where α =0.8. ρ =0.3, ν = ν *=0, and λ =0.1.

source. For the considered set of parameters, the increase of σ wins over the decrease of the number of empty slots.

C. Case $\nu^* > \nu$

In this section, illustrative results for the general case $\nu^* > \nu$ are shown. As discussed in Section V, in this case, the structure of the optimal policy depends on the parameters. In fact, due to the effect of the interference by the primary source at the secondary receiver, the secondary source may be forced to be silent in states $\theta > 0$ in order not to decrease the steady-state probability of the empty-slot state 0. In the following, the constraint is defined on the throughput loss of the primary source.

The throughput and the transmission probabilities as a function of λ_S are depicted in Figs. 19 and 20. We recall that $\lambda_S{\in}[0,1]$ determines how decoding at the secondary receiver is hampered by primary source's transmissions. The values $\lambda_S{=}0$ and $\lambda_S{=}1$ correspond to $\nu^*{=}\nu$ and $\nu^*{=}1$, respectively.

As a first observation, the throughput of the secondary source decreases as λ_S increases. In fact, the effect of interference both decreases the reward associated with transmission in the states $\theta > 0$ and forces the secondary source to reduce its overall activity. For the same reason, the throughput of the primary source increases and moves close to the maximum throughput, that is, the secondary source rarely interferes with the primary source.

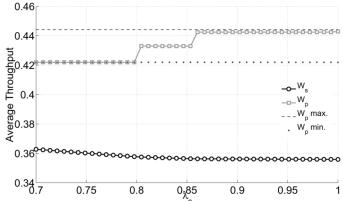


Fig. 19. Throughput as a function of λ_S , where α =0.5, λ =0.6, ρ = ν =0.2, and ϵ =0.05.

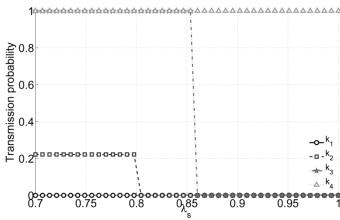


Fig. 20. Transmission probabilities as a function of λ_S , where α =0.5, λ =0.6, ρ = ν =0.2, and ϵ =0.05.

In fact, as ν^* gets closer to one, the secondary source reduces transmission, and interference, in states $0 < \theta < T$. This is done in order to avoid retransmissions, which would reduce the availability of white space. Interference in state T does not induce a higher probability of further retransmissions. Therefore, κ_T remains one as long as it is admissible according to the constraint. Note that κ_1 , i.e., the transmission probability in the state which has the largest impact on the average number of retransmissions, is set to 0. For this configuration of parameters, the policy takes the opposite form with respect to that described for the case $\nu^* = \nu$, i.e., the secondary source concentrates transmissions in the last states of the chain, in order to have a smaller impact on the number of transmissions of the primary source's packets.

VIII. CONCLUSION

In contrast to much prior work on cognitive networks, in this paper, we have investigated a scenario wherein the secondary source is allowed to superpose its transmissions over those of the primary source. The secondary source aims to maximize its own throughput, while guaranteeing a bounded performance loss for the primary source. We derived the optimal transmission policy for the secondary user when the primary user adopts a retransmission-based error control scheme. If the decoding probability at the secondary receiver is not increased by the primary source's transmissions, the resulting optimal strategy of the secondary user has a unique structure. In particular, the optimal

throughput is achieved by the secondary user by concentrating its interference to the primary user in the first transmissions of a packet. This is a first step toward a better understanding of interference control strategies in dynamic wireless networks.

APPENDIX A PROOF OF THEOREM 1

Proof: Theorem 1 states that if any component κ_{θ} of the vector $\underline{\kappa}$ is increased, with $\theta > 0$, then $\mathcal{J}_{P}(\underline{\kappa})$ increases. This corresponds to the intuitive fact that a larger transmission probability of the secondary source in any of the states in which the primary source transmits results in a smaller throughput achieved by the latter.

In order to prove this result, we show that $\partial \mathcal{J}_{P}(\underline{\kappa})/\partial \kappa_{\theta} > 0$, $\forall \theta > 0$.

Let us introduce the following notation

$$\mathcal{N}_{J_{\mathcal{P}}}(\underline{\kappa}) = (1 - \alpha) + \alpha \sum_{t=1}^{T} \prod_{i=1}^{t} (\rho + (1 - \rho)\lambda \kappa_{i})$$
 (46)

$$\mathcal{D}(\underline{\kappa}) = 1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} (\rho + (1-\rho)\lambda \kappa_i). \tag{47}$$

The cost is then $\mathcal{J}_{\mathrm{P}}(\underline{\kappa}) = \mathcal{N}_{\mathcal{J}_{\mathrm{P}}}(\underline{\kappa})/\mathcal{D}(\underline{\kappa})$. In the following, the obvious dependence of the aforementioned functions on $\underline{\kappa}$ is dropped from the notation.

The derivative of the cost of the primary source can be obtained through the well-known formula

$$\partial \mathcal{J}_{P}/\partial \kappa_{\theta} = \frac{\frac{\partial \mathcal{N}_{J_{P}}}{\partial \kappa_{\theta}} \mathcal{D} - \frac{\partial \mathcal{D}}{\partial \kappa_{\theta}} \mathcal{N}_{J_{P}}}{(\mathcal{D})^{2}}.$$
 (48)

In the previous equation, the denominator is always positive, and thus we focus on the numerator. We have

$$\frac{\partial \mathcal{N}_{J_{P}}}{\partial \kappa_{\theta}} = \alpha \sum_{t=\theta}^{T} (1-\rho)\lambda \prod_{i=1, i\neq \theta}^{t} (\rho + (1-\rho)\lambda \kappa_{i})$$
(49)

$$\frac{\partial \mathcal{D}}{\partial \kappa_{\theta}} = \alpha \sum_{t=\theta}^{T-1} (1-\rho) \lambda \prod_{i=1, i \neq \theta}^{t} (\rho + (1-\rho)\lambda \kappa_{i}). \tag{50}$$

Through simple algebraic manipulation, we obtain the expression in (51), shown at the bottom of the page. Since

$$1 - \alpha^2 \prod_{i=1}^{T} (\rho + (1 - \rho)\lambda \kappa_i) > 0$$
 (52)

then all the terms in (51) are strictly positive. Therefore, the derivative is strictly positive and the cost is a monotonically increasing function of any element κ_{θ} with $\theta > 0$

APPENDIX B PROOF OF THEOREM 2

Proof: Theorem 2 states that if the secondary source transmits with a higher probability in any state θ , i.e., κ_{θ} is increased, then the average throughput achieved by the secondary source increases. This may appear a trivial consideration. However, it must be observed that the transmission probabilities κ_{θ} , $0 < \theta < T$ influence the steady-state distribution of the Markov chain of the network and thus influence the average throughput of the secondary user.

A larger κ_{θ} results in a larger probability that the primary source fails the θ th transmission of a packet, and thus, a larger probability that the Markov process moves to states $\theta+1,\ldots,T$. As a consequence, the steady-state probabilities of the latter states increase, while those of states $0,\ldots,\theta$ decrease. Intuition suggests that, in some cases, a larger κ_{θ} may result in a smaller overall average transmission probability of the secondary source, i.e., $\sum_t \pi_{\underline{\kappa}}(t)\kappa_t$. Theorem 2 instead ensures that a larger value of any of the κ_{θ} always results in a larger $\sum_t \pi_{\underline{\kappa}}(t)\kappa_t$.

The proof of this theorem is analogous to that of Theorem 1. In particular, we show in the following that $\partial \mathcal{W}_{S}(\underline{\kappa})/\partial \kappa_{\theta} > 0$, $\forall \theta \in \{0, \dots, T\}$.

Recalling the definition of $\mathcal{D}(\underline{\kappa})$ given in the previous Theorem, we write $\mathcal{W}_{S}(\underline{\kappa}) = \mathcal{N}_{W_{S}}(\underline{\kappa})/\mathcal{D}(\underline{\kappa})$, where

(49)
$$\mathcal{N}_{W_{S}}(\underline{\kappa}) = (1 - \alpha) + \alpha \sum_{t=1}^{T} \kappa_{t} \prod_{i=1}^{t-1} (\rho + (1 - \rho)\lambda \kappa_{i}). \quad (53)$$

¹⁷Note that in the degenerate cases $\alpha = 0$, $\lambda = 0$, or $\rho = 1(51)$ is equal to zero.

$$\frac{\partial \mathcal{N}_{J_{P}}}{\partial \kappa_{\theta}} \mathcal{D} - \frac{\partial \mathcal{D}}{\partial \kappa_{\theta}} \mathcal{N}_{J_{P}} = \left(\alpha (1 - \rho) \lambda \prod_{i=1, i \neq \theta}^{T} (\rho + (1 - \rho) \lambda \kappa_{i}) \right) \left(\mathcal{D} + 1 \right) + \left(1 - \alpha^{2} \prod_{i=1}^{T} (\rho + (1 - \rho) \lambda \kappa_{i}) \right) \left(\sum_{t=\theta}^{T-1} (1 - \rho) \lambda \prod_{i=1, i \neq \theta}^{t} (\rho + (1 - \rho) \lambda \kappa_{i}) \right)$$
(51)

In the following, we drop in the notation the dependence between these functions and the policy. The derivative of the numerator is

$$\frac{\partial \mathcal{N}_{W_S}}{\partial \kappa_{\theta}} = \alpha \prod_{i=1}^{\theta-1} (\rho + (1-\rho)\lambda \kappa_i) + \\
+ \alpha \sum_{t=\theta+1}^{T} (1-\rho)\lambda \kappa_t \prod_{i=1, i\neq \theta}^{t-1} (\rho + (1-\rho)\lambda \kappa_i). \quad (54)$$

Again, the derivative can be written as

$$\partial W_{\rm S}/\partial \kappa_{\theta} = \frac{\frac{\partial \mathcal{N}_{W_{\rm S}}}{\partial \kappa_{\theta}} \mathcal{D} - \frac{\partial \mathcal{D}}{\partial \kappa_{\theta}} \mathcal{N}_{W_{\rm S}}}{(\mathcal{D})^2}.$$
 (55)

We first show the following Lemma:

Lemma 1: Consider $\mathcal{N}_{W_{\mathrm{S}}}$ and \mathcal{D} as previously defined, the following can be shown:

$$\partial \mathcal{N}_{W_{S}}/\partial \kappa_{\theta} > \partial \mathcal{D}/\partial \kappa_{\theta}$$
 (56)

with $0 < \theta < T$.

Proof: If $\theta = T$, we have

$$\partial \mathcal{N}_{W_{S}}/\partial \kappa_{T} - \partial \mathcal{D}/\partial \kappa_{T} = \alpha \prod_{i=1}^{T-1} (\rho + (1-\rho)\lambda \kappa_{i})$$
 (57)

that is clearly positive.

Assume $\theta < T$. The expressions

$$1 + (1 - \rho)\lambda \kappa_{\theta+1} + \sum_{t=\theta+2}^{T} (1 - \rho)\lambda \kappa_t \prod_{i=\theta+1}^{t-1} (\rho + (1 - \rho)\lambda \kappa_i)$$
 (58)

and

$$(1-\rho)\lambda \left(1+\sum_{t=\theta+1}^{T-1}\prod_{i=\theta+1}^{t}(\rho+(1-\rho)\lambda\kappa_i)\right)$$
 (59)

are $\partial \mathcal{N}_{W_{\mathrm{S}}}/\partial \kappa_{\theta}$ and $\partial \mathcal{D}/\partial \kappa_{\theta}$ divided by

$$\alpha \prod_{i=1}^{\theta-1} (\rho + (1-\rho)\lambda \kappa_i),$$

respectively.

Equations (58) and (59) can be reorganized as

$$1 + (1 - \rho)\lambda \sum_{t=\theta+1}^{T} \rho^{t-\theta-1} \kappa_t + C_1 \tag{60}$$

and

$$(1 - \rho^{t-\theta-1})\lambda + (1 - \rho)\lambda^2 \sum_{t=\theta+1}^{T-1} (\rho^{t-\theta-1} - \rho^{t-\theta})\kappa_t + C_2 \quad (61)$$

respectively, where the constants C_1 and C_2 account for all the cross terms involving the multiplications of two or more variables κ_i . We do not provide here the expressions for C_1 and C_2 , as they are conceptually simple, but tedious. However, it is possible to show that $C_1 > C_2$.

Since $1 > (1 - \rho^{t-\theta-1})\bar{\lambda}$ and

$$(1-\rho)\lambda \sum_{t=\theta+1}^{T} \rho^{t-\theta-1} \kappa_t > (1-\rho)\lambda^2 \sum_{t=\theta+1}^{T-1} (\rho^{t-\theta-1} - \rho^{t-\theta}) \kappa_t$$
(62)

then $\partial \mathcal{N}_{W_{S}}/\partial \kappa_{\theta} > \partial \mathcal{D}/\partial \kappa_{\theta}$.

As for the proof of Theorem 1, since $(\mathcal{D})^2 > 0$, we focus on $\frac{\partial \mathcal{N}_{W_{\mathrm{S}}}}{\partial \kappa_{\theta}} \mathcal{D} - \frac{\partial \mathcal{D}}{\partial \kappa_{\theta}} \mathcal{N}_{W_{\mathrm{S}}}$. Due to Lemma 1, $\frac{\partial \mathcal{N}_{W_{\mathrm{S}}}}{\partial \kappa_{\theta}} > \frac{\partial \mathcal{D}}{\partial \kappa_{\theta}}$. Moreover, $\mathcal{D} \geq \mathcal{N}_{W_{\mathrm{S}}}$. In fact

$$W_{S}(\underline{\kappa}) = \frac{\mathcal{N}_{W_{S}}}{\mathcal{D}} \le 1 \tag{63}$$

Therefore

$$\frac{\partial \mathcal{N}_{W_{\mathrm{S}}}}{\partial \kappa_{\mathcal{A}}} \mathcal{D} > \frac{\partial \mathcal{D}}{\partial \kappa_{\mathcal{A}}} \mathcal{D} \ge \frac{\partial \mathcal{D}}{\partial \kappa_{\mathcal{A}}} \mathcal{N}_{W_{\mathrm{S}}}$$
 (64)

(56) and
$$\frac{\partial \mathcal{N}_{W_{\mathrm{S}}}}{\partial \kappa_{\theta}} \mathcal{D} - \frac{\partial \mathcal{D}}{\partial \kappa_{\theta}} \mathcal{N}_{W_{\mathrm{S}}} > 0.$$

APPENDIX C PROOF OF THEOREM 3

Proof: Theorem 3 states that starting from a policy $\underline{\kappa}$, for any pair of indices j and r with 0 < j < r such that $\kappa_j = \kappa_r$ and $\kappa_t = 0$, $r < t \le T$, if

$$\mathcal{J}_{P}(\underline{\kappa}') = \mathcal{J}_{P}(\underline{\kappa}'') \tag{65}$$

then

$$W_{S}(\underline{\kappa}') > W_{S}(\underline{\kappa}'')$$
 (66)

with $\underline{\kappa}' = \underline{\kappa} + \underline{u}_j \delta_j'$ and $\underline{\kappa}'' = \underline{\kappa} + \underline{u}_r \delta_r''$, $0 < \delta_j' \le 1 - \kappa_j$ and $0 < \delta_r'' \le 1 - \kappa_r$.

In words, if the policy obtained by increasing the jth element of $\underline{\kappa}$ by δ'_j and the policy obtained by increasing the rth element of $\underline{\kappa}$ by δ''_r incur the same primary source's average cost, then the former policy achieves a larger secondary source's average reward.

We briefly recall the notation introduced in the previous proofs. The average primary source's cost and secondary source's reward can be written, respectively, as

$$\mathcal{J}_{P}(\underline{\kappa}) = \frac{\mathcal{N}_{J_{P}}(\underline{\kappa})}{\mathcal{D}(\underline{\kappa})}$$
 (67)

$$W_{S}(\underline{\kappa}) = \frac{\mathcal{N}_{W_{S}}(\underline{\kappa})}{\mathcal{D}(\kappa)}$$
 (68)

where

$$\mathcal{N}_{J_{P}}(\underline{\kappa}) = (1 - \alpha) + \alpha \sum_{t=1}^{T} \prod_{i=1}^{t} (\rho + (1 - \rho)\lambda \kappa_{i}) > 0$$
 (69)

$$\mathcal{D}(\underline{\kappa}) = 1 + \alpha \sum_{t=1}^{T-1} \prod_{i=1}^{t} (\rho + (1-\rho)\lambda \kappa_i) > 0$$
 (70)

$$\mathcal{N}_{W_{\mathrm{S}}}(\underline{\kappa}) = (1 - \alpha) + \alpha \sum_{t=1}^{T} \kappa_t \prod_{i=1}^{t-1} (\rho + (1 - \rho)\lambda \kappa_i) > 0. \quad (71)$$

In the following, in order to simplify the notation, we drop the subscripts P and S and we refer to the primary source's cost and secondary source's reward when talking of cost and reward, respectively.

If the qth element of $\underline{\kappa}$, with q>0, is increased by δ , with $\delta \le 1-\kappa_q$, the average cost and reward can be written as

$$\mathcal{J}(\underline{\kappa} + \underline{u}_q \delta) = \frac{\mathcal{N}_J(\underline{\kappa}) + \Delta \mathcal{N}_J(q, \delta, \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta \mathcal{D}(q, \delta, \underline{\kappa})}$$
(72)

$$W(\underline{\kappa} + \underline{u}_q \delta) = \frac{\mathcal{N}_W(\underline{\kappa}) + \Delta \mathcal{N}_W(q, \delta, \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta \mathcal{D}(q, \delta, \underline{\kappa})}$$
(73)

where

There
$$\Delta \mathcal{N}_{J}(q, \delta, \underline{\kappa}) = \delta \left[\alpha(1-\rho)\lambda \sum_{t=q}^{T} \prod_{i=1, i \neq q}^{t} (\rho + (1-\rho)\lambda \kappa_{i}) \right]$$

$$\Delta \mathcal{D}(q, \delta, \underline{\kappa}) = \delta \left[\alpha(1-\rho)\lambda \sum_{t=q}^{T-1} \prod_{i=1, i \neq q}^{t} (\rho + (1-\rho)\lambda \kappa_{i}) \right]$$
(74)
$$(75)$$

$$\Delta \mathcal{N}_{W}(q, \delta, \underline{\kappa}) = \delta \left[\alpha \prod_{i=1}^{q-1} (\rho + (1-\rho)\lambda \kappa_{i}) + \alpha (1-\rho)\lambda \sum_{t=q+1}^{T} \kappa_{t} \prod_{i=1, i \neq q}^{t-1} (\rho + (1-\rho)\lambda \kappa_{i}) \right].$$
(76)

Thus, $\Delta \mathcal{N}_J(q, \delta, \underline{\kappa})$, $\Delta \mathcal{N}_W(q, \delta, \underline{\kappa})$, and $\Delta \mathcal{D}(q, \delta, \underline{\kappa})$ are linear functions of δ , and represent the increment of the numerator of the cost and reward, and of their denominator, corresponding to an increase δ of κ_q . Note that if δ is strictly positive and q > 0, 18 then $\Delta \mathcal{N}_J(q, \delta, \underline{\kappa})$, $\Delta \mathcal{D}(q, \delta, \underline{\kappa})$, and $\Delta \mathcal{N}_W(q, \delta, \underline{\kappa})$ are strictly positive.

According to the hypothesis of the theorem $\mathcal{J}(\underline{\kappa}+\underline{u}_j\delta_j')=\mathcal{J}(\underline{\kappa}+\underline{u}_r\delta_r'')$. This equality can be rewritten as

$$\frac{\mathcal{N}_{J}(\underline{\kappa}) + \Delta \mathcal{N}_{J}(j, \delta'_{j}, \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta \mathcal{D}(j, \delta'_{j}, \underline{\kappa})} = \frac{\mathcal{N}_{J}(\underline{\kappa}) + \Delta \mathcal{N}_{J}(r, \delta''_{r}, \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta \mathcal{D}(r, \delta''_{r}, \underline{\kappa})}.$$
 (77)

The increases of the numerators and denominators can be rewritten as

$$\Delta \mathcal{D}(r, \delta, \underline{\kappa}) = \delta A \tag{78}$$

$$\Delta \mathcal{N}_J(r, \delta, \underline{\kappa}) = \delta B \tag{79}$$

$$\Delta \mathcal{N}_W(r, \delta, \kappa) = \delta G. \tag{80}$$

Note that, since $\kappa_j = \kappa_r$, the difference between the increase of the denominator when κ_j or κ_r are increased by δ is a constant C equal to

$$\Delta \mathcal{D}(j, \delta, \underline{\kappa}) - \Delta \mathcal{D}(r, \delta, \underline{\kappa}) =$$

$$= \delta \alpha (1 - \rho) \lambda \sum_{t=j}^{r-1} \prod_{i=1, i \neq j}^{t} (\rho + (1 - \rho) \lambda \kappa_i)$$

$$= \delta C > 0. \tag{81}$$

Analogously, the difference between the numerators of the cost and reward increases are

$$\Delta \mathcal{N}_{J}(j, \delta, \underline{\kappa}) - \Delta \mathcal{N}_{J}(r, \delta, \underline{\kappa}) = \Delta \mathcal{D}(j, \delta, \underline{\kappa}) - \Delta \mathcal{D}(r, \delta, \underline{\kappa})$$

$$= \delta C \tag{82}$$

and

$$\Delta \mathcal{N}_{W}(j, \delta, \underline{\kappa}) - \Delta \mathcal{N}_{W}(r, \delta, \underline{\kappa}) =$$

$$= \delta \alpha \left[\prod_{i=1}^{j-1} (\rho + (1-\rho)\lambda \kappa_{i}) - \prod_{i=1}^{r-1} (\rho + (1-\rho)\lambda \kappa_{i}) \right] +$$

$$+ \delta \alpha (1-\rho)\lambda \sum_{t=j+1}^{r} \kappa_{t} \prod_{i=1, i \neq j}^{t-1} (\rho + (1-\rho)\lambda \kappa_{i})$$

$$= \delta F > 0$$
(83)

respectively.

According to (78)–(83), and omitting the dependence of the quantities on $\underline{\kappa}$, we rewrite $\mathcal{J}(\underline{\kappa}+\underline{u}_j\delta'_j)$, $\mathcal{J}(\underline{\kappa}+\underline{u}_r\delta''_r)$, $\mathcal{W}(\underline{\kappa}+\underline{u}_j\delta'_j)$, and $\mathcal{W}(\underline{\kappa}+\underline{u}_r\delta''_r)$ as shown in (84)–(87), respectively, shown at the bottom of the page.

$$\mathcal{J}(\underline{\kappa} + \underline{u}_{j}\delta'_{j}) = \frac{\mathcal{N}_{J}(\underline{\kappa}) + \Delta\mathcal{N}_{J}(j, \delta'_{j}, \underline{\kappa}) + \Delta\mathcal{N}_{J}(r, \delta'_{j}, \underline{\kappa}) - \Delta\mathcal{N}_{J}(r, \delta'_{j}, \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta\mathcal{D}(j, \delta'_{j}, \underline{\kappa}) + \Delta\mathcal{D}(r, \delta'_{j}, \underline{\kappa}) - \Delta\mathcal{D}(r, \delta'_{j}, \underline{\kappa})} = \frac{\mathcal{N}_{J} + (B + C) \delta'_{j}}{\mathcal{D} + (A + C) \delta'_{j}}$$
(84)

$$\mathcal{J}(\underline{\kappa} + \underline{u}_r \delta_r'') = \frac{\mathcal{N}_J(\underline{\kappa}) + \Delta \mathcal{N}_J(r, \delta_r'', \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta \mathcal{D}(r, \delta_r'', \underline{\kappa})} = \frac{\mathcal{N}_J + B \, \delta_r''}{\mathcal{D} + A \, \delta_r''}$$
(85)

$$\mathcal{W}(\underline{\kappa} + \underline{u}_{j}\delta'_{j}) = \frac{\mathcal{N}_{W}(\underline{\kappa}) + \Delta\mathcal{N}_{W}(j,\delta'_{j},\underline{\kappa}) + \Delta\mathcal{N}_{W}(r,\delta'_{j},\underline{\kappa}) - \Delta\mathcal{N}_{W}(r,\delta'_{j},\underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta\mathcal{D}(j,\delta'_{j},\underline{\kappa}) + \Delta\mathcal{D}(r,\delta'_{j},\underline{\kappa}) - \Delta\mathcal{D}(r,\delta'_{j},\underline{\kappa})} = \frac{\mathcal{N}_{W}(\underline{\kappa}) + (G+F)\delta'_{j}}{\mathcal{D} + (A+C)\delta'_{j}}$$
(86)

$$\mathcal{W}(\underline{\kappa} + \underline{u}_r \delta_r'') = \frac{\mathcal{N}_W(\underline{\kappa}) + \Delta \mathcal{N}_W(r, \delta_r'', \underline{\kappa})}{\mathcal{D}(\underline{\kappa}) + \Delta \mathcal{D}(r, \delta_r'', \underline{\kappa})} = \frac{\mathcal{N}_W + G \, \delta_r''}{\mathcal{D} + A \, \delta_r''}$$
(87)

 $^{^{18}}$ Together with the assumptions $\lambda,\,\alpha,\,1\!-\!\rho\!>\!\!0$.

Note that

$$\mathcal{J}(\underline{\kappa} + \underline{u}_{j}\delta) = \frac{\mathcal{N}_{J} + (B + C) \delta}{\mathcal{D} + (A + C) \delta} > \frac{\mathcal{N}_{J} + B \delta}{\mathcal{D} + A \delta} = \mathcal{J}(\underline{\kappa} + \underline{u}_{r}\delta) \quad (88)$$

for any δ , with $0 < \delta \le \min(1 - \kappa_i, 1 - \kappa_r)$.

Choose δ_r'' , with $0 < \delta_r'' \le 1 - \kappa_r$, and denote the cost of policy $\underline{\kappa} + \underline{u}_r \delta_r''$ with $Z = \mathcal{J}(\underline{\kappa} + \underline{u}_r \delta_r'')$. Observe that, due to the monotonicity of the cost function, then $\mathcal{J}(\underline{\kappa}) < Z \le 1$.

Since the cost function is continuous with respect to any element of the policy vector and $\kappa_i = \kappa_r$ by assumption, then there always exists δ'_i such that $\mathcal{J}(\underline{\kappa} + \underline{u}_i \delta'_i) = \mathcal{J}(\underline{\kappa} + \underline{u}_r \delta''_r) = Z$, with $0<\delta_j^{'}<\delta_r^{''}\le 1-\kappa_r=1-\kappa_j$. The values for δ_j' and δ_r'' can be readily found to be

$$\delta_j' = \frac{\mathcal{D} Z - \mathcal{N}_J}{B + C - (A + C) Z} \tag{89}$$

$$\delta_r'' = \frac{\mathcal{D} Z - \mathcal{N}_J}{B - A Z}.$$
(90)

In order to complete the proof, the following inequality needs to be proved:

$$\mathcal{W}(\underline{\kappa} + \underline{u}_{j}\delta'_{j}) - \mathcal{W}(\underline{\kappa} + \underline{u}_{r}\delta''_{r}) =
= \frac{\mathcal{N}_{W} + (G+F)\delta'_{j}}{\mathcal{D} + (A+C)\delta'_{j}} - \frac{\mathcal{N}_{W} + G\delta''_{r}}{\mathcal{D} + A\delta''_{r}} > 0.$$
(91)

By substituting (89) and (90) into (91), we obtain (92), shown at the bottom of the page. Note that

$$\frac{\mathcal{D}Z - \mathcal{N}_J}{(B\mathcal{D} - A\mathcal{N}_J)((B+C)\mathcal{D} - (A+C)\mathcal{N}_J)} > 0$$
 (93)

In fact, recalling that $\kappa_j = \kappa_r$ by hypothesis, we have

$$B - A = \alpha (1 - \rho) \lambda \left(\prod_{i=1, i \neq r}^{T} (\rho + (1 - \rho)\lambda) \kappa_i \right)$$
$$= \alpha (1 - \rho) \lambda \left(\prod_{i=1, i \neq j}^{T} (\rho + (1 - \rho)\lambda) \kappa_i \right) > 0 \quad (94)$$

and since $\mathcal{D} \geq \mathcal{N}_{J}$, 19 then

$$(B\mathcal{D} - A\mathcal{N}_J)((B+C)\mathcal{D} - (A+C)\mathcal{N}_J) > 0 \tag{95}$$

Moreover, due to Theorem 1, we have for $\delta_i' > 0$

$$Z = \mathcal{J}(\underline{\kappa} + \underline{u}_j \delta'_j) > \mathcal{J}(\underline{\kappa}) = \frac{\mathcal{N}_J}{\mathcal{D}}.$$
 (96)

¹⁹Because $\frac{\mathcal{N}_J}{\mathcal{D}} = \mathcal{J}(\underline{\kappa}) \leq 1$.

Therefore

$$\mathcal{D}Z - \mathcal{N}_J > 0. \tag{97}$$

The theorem is then proved if the following inequality holds:

$$\mathcal{D}(BF-CG)+(CG-AF)\mathcal{N}_J+(A-B)C\mathcal{N}_W)=$$

$$(\mathcal{D}-\mathcal{N}_J)(FB-CG)+(B-A)(\mathcal{N}_JF-\mathcal{N}_WC)>0. (98)$$

Define

$$X = \alpha \prod_{i=1}^{T} (\rho + (1 - \rho)\lambda \kappa_i) > 0$$
 (99)

Then

$$\mathcal{D} - \mathcal{N}_J = \alpha \left(1 - \prod_{i=1}^T (\rho + (1 - \rho) \lambda \kappa_i) \right) = \alpha - X > 0 \quad (100)$$

and

$$B - A = \frac{(1 - \rho)\lambda}{\rho + (1 - \rho)\lambda\kappa_i} X > 0. \tag{101}$$

In the second term of (98), \mathcal{N}_I can be split into two terms

$$\mathcal{N}_{J} = (1 - \alpha) + \alpha \sum_{t=1}^{r-1} \prod_{i=1}^{t} (\rho + (1 - \rho)\lambda \kappa_{i}) + \alpha \sum_{t=r}^{r-1} \prod_{i=1}^{t} (\rho + (1 - \rho)\lambda \kappa_{i})$$
(102)

so that the second summation corresponds to the summation in $B^{20}(\mathcal{D}-\mathcal{N}_J)F$ $B+(B-A)\mathcal{N}_JF$ in (98) can be simplified as shown in (103), at the bottom of the next page.

Analogously, \mathcal{N}_W can be rewritten as

$$\mathcal{N}_W = (1 - \alpha) + \alpha \sum_{t=1}^r \kappa_t \prod_{i=1}^{t-1} (\rho + (1 - \rho)\lambda \kappa_i) + \alpha \sum_{t=r+1}^T \kappa_t \prod_{i=1}^{t-1} (\rho + (1 - \rho)\lambda \kappa_i)$$

$$(104)$$

so that one summation corresponds to the summation in G. The term $-(\mathcal{D}-\mathcal{N}_J)G$ $C-(B-A)\mathcal{N}_WC$ in (98) can be rewritten as reported in (105), shown at the bottom of the next page.

Equation (98) is the sum of (103) and (105). By hypothesis $\kappa_t = 0, \forall t > r$, and thus G = 0. The term $-\alpha G C$ in (105) is then

²⁰In B, the summation has the additional term corresponding to t=T and the products do not have the term i=j.

$$\frac{(\mathcal{D}Z - \mathcal{N}_J)(\mathcal{D}(BF - CG) + (CG - AF)\mathcal{N}_J + (A - B)C\mathcal{N}_W)}{(B\mathcal{D} - A\mathcal{N}_I)((B + C)\mathcal{D} - (A + C)\mathcal{N}_I)} > 0 \tag{92}$$

equal to zero. Equation (106), shown at the bottom of the page, reorganizes the sum of (103) and (105).

The first term of (106)

$$\alpha F \left(B - X \alpha (1 - \rho) \lambda \prod_{i=1, i \neq j}^{T} (\rho + (1 - \rho) \lambda \kappa_i) \right)$$
 (107)

is positive. In fact, $X \le 1$ and

$$B - \alpha(1 - \rho)\lambda \prod_{i=1, i \neq j}^{T} (\rho + (1 - \rho)\lambda \kappa_i)$$

$$= \alpha(1 - \rho)\lambda \sum_{t=r}^{T-1} \prod_{i=1, i \neq j}^{t} (\rho + (1 - \rho)\lambda \kappa_i) > 0. (108)$$

Moreover, it can be shown that F>C. The proof is analogous to that of Lemma 1, and is not reported herein. As a consequence, the second term of (106) is positive. All the other terms are trivially positive.

The inequality is then proved, as well as the Theorem.

APPENDIX D PROOF OF THEOREM 4

Proof: Theorem 4 states that starting from a policy $\underline{\kappa}$, for any pair of indices j and r, with 0 < j < r, such that $\kappa_j = \kappa_r$ and $\kappa_t = 0$, $r < t \le T$, if

$$\mathcal{J}_{P}(\kappa') = \mathcal{J}_{P}(\kappa'') \tag{109}$$

$$(\mathcal{D}-\mathcal{N}_{J})F B+(B-A)\mathcal{N}_{J}F = \alpha F B - X F B + \frac{(1-\rho)\lambda}{\rho + (1-\rho)\lambda\kappa_{j}}\mathcal{N}_{J}$$

$$= \alpha F B + (1-\alpha)\frac{(1-\rho)\lambda}{\rho + (1-\rho)\lambda\kappa_{j}}X F - X F\alpha(1-\rho)\lambda X F \prod_{i=1,i\neq j}^{T} (\rho + (1-\rho)\lambda\kappa_{i})$$

$$+ X F\alpha(1-\rho)\lambda \sum_{t=1}^{J-1} \prod_{i=1,i\neq j}^{t} (\rho + (1-\rho)\lambda\kappa_{i}) + X F C.$$
(103)

$$-(\mathcal{D}-\mathcal{N}_{J})GC - (B-A)\mathcal{N}_{W}C = -\alpha C G + X C G - \frac{(1-\rho)\lambda}{\rho + (1-\rho)\lambda\kappa_{j}} X C \mathcal{N}_{W}$$

$$= -\alpha C G - (1-\alpha)\frac{(1-\rho)\lambda}{\rho + (1-\rho)\lambda\kappa_{j}} X C + X C \alpha \prod_{i=1}^{r-1} (\rho + (1-\rho)\lambda\kappa_{i}) +$$

$$-X C \alpha (1-\rho)\lambda \sum_{t=1}^{r} \kappa_{t} \prod_{i=1, i\neq j}^{t-1} (\rho + (1-\rho)\lambda\kappa_{i})$$

$$= -\alpha C G - \frac{(1-\rho)\lambda}{\rho + (1-\rho)\lambda\kappa_{j}} X (1-\alpha) C + X C\alpha (1-\rho)\lambda \sum_{t=1}^{j} \kappa_{t} \prod_{i=1, i\neq j}^{t-1} (\rho + (1-\rho)\lambda\kappa_{i}) +$$

$$-X C F + X C \alpha \prod_{i=1}^{j-1} (\rho + (1-\rho)\lambda\kappa_{i}). \tag{105}$$

$$\mathcal{D}(BF-CG) + (CG-AF)\mathcal{N}_J + (A-B)C\mathcal{N}_W = \alpha F\left(B-X\alpha(1-\rho)\lambda \prod_{i=1,i\neq j}^T (\rho + (1-\rho)\lambda\kappa_i)\right) +$$

$$(1-\alpha)\frac{(1-\rho)\lambda}{\rho + (1-\rho)\lambda\kappa_j} X(F-C) + X F \alpha(1-\rho)\lambda \sum_{t=1}^{j-1} \prod_{i=1,i\neq j}^t (\rho + (1-\rho)\lambda\kappa_i) +$$

$$X C \alpha \prod_{i=1}^{j-1} (\rho + (1-\rho)\lambda\kappa_i) + X C \alpha(1-\rho)\lambda \sum_{t=1}^{j} \kappa_t \prod_{i=1,i\neq j}^{t-1} (\rho + (1-\rho)\lambda\kappa_i)$$

$$(106)$$

then

$$W_{\rm S}(\underline{\kappa}') < W_{\rm S}(\underline{\kappa}'')$$
 (110)

with $\underline{\kappa}' = \underline{\kappa} - \underline{u}_j \delta_j'$ and $\underline{\kappa}'' = \underline{\kappa} - \underline{u}_r \delta_r''$, $0 < \delta_j' \le \kappa_j$ and $0 < \delta_r'' \le \kappa_r$. The proof is similar to that provided in the previous Appendix.

As done in the previous proof, we fix δ''_r , with $0 < \delta''_r \le 1 - \kappa_r$, and we denote the cost associated with the policy obtained by decreasing κ_r by δ''_r with $Z = \mathcal{J}(\underline{\kappa} - \underline{u}_r \delta''_r)$. Note that $0 \le Z < \mathcal{J}_P(\underline{\kappa})$. Through considerations entirely analogous to those provided in Appendix C, it can be shown that there exists δ'_j such that $\mathcal{J}_P(\underline{\kappa} - \underline{u}_j \delta'_j) = \mathcal{J}_P(\underline{\kappa} - \underline{u}_r \delta''_r) = Z$, with $0 < \delta'_j < \delta''_r \le \kappa_r = \kappa_j$. The corresponding values of δ'_j and δ''_r can be readily found to be

$$\delta_j' = \frac{\mathcal{N}_J - \mathcal{D}Z}{B + C - (A + C)Z} \tag{111}$$

$$\delta_r'' = \frac{\mathcal{N}_J - \mathcal{D}Z}{B - A Z}.\tag{112}$$

Observe that the aforementioned values of δ'_j and δ''_r are the opposites of those in (89) and (90).

By substituting the aforementioned equations in

$$W_{S}(\underline{\kappa}') - W_{S}(\underline{\kappa}'') = \frac{N_{W} - (G + F) \delta'_{j}}{\mathcal{D} - (A + C) \delta'_{i}} - \frac{N_{W} - G \delta''_{r}}{\mathcal{D} - A \delta''_{r}}$$
(113)

the same fraction as in (92) is obtained, and according to (95) and (98)

$$\frac{(\mathcal{D}(BF-CG)+(CG-AF)\mathcal{N}_J+(A-B)C\mathcal{N}_W)}{(B\mathcal{D}-A\mathcal{N}_J)((B+C)\mathcal{D}-(A+C)\mathcal{N}_J)} > 0. (114)$$

Differently from the proof of Theorem 3, since $Z < \mathcal{J}_P(\underline{\kappa}) = \mathcal{N}_J/\mathcal{D}$, then

$$\mathcal{D}Z - \mathcal{N}_J < 0. \tag{115}$$

Therefore

$$W_{S}(\kappa') - W_{S}(\kappa'') < 0. \tag{116}$$

and the theorem is proved.

APPENDIX E

PROOF OF THEOREMS 3 AND 4 FOR CONSTRAINT ON THE FAILURE PROBABILITY

Proof: As discussed in Section IV-B, the cost increase/decrease induced by an increase/decrease of the transmission probability in state θ >0 does not depend on θ .

In this case, a transmission probability increased by δ , 21 in state j or r results in the same increase of the average cost. More formally, fix j, r, and δ , with $0 < j < r \le T$, $-\min(\kappa_j, \kappa_r) \le \delta \le \min(1 - \kappa_j, 1 - \kappa_r)$, and $\delta \ne 0$; then, $\mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa} + \underline{u}_j \delta) = \mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa} + \underline{u}_r \delta)$.

 $^{21}\mathrm{A}$ decreased transmission probability corresponds to a negative δ in the following.

Therefore, starting from a policy $\underline{\kappa}$, and defining the policies $\underline{\kappa}' = \underline{\kappa} + \underline{u}_j \delta'_j$ and $\underline{\kappa}'' = \underline{\kappa} + \underline{u}_r \delta''_r$, then

$$\mathcal{J}_{P}^{fp}(\underline{\kappa}') = \mathcal{J}_{P}^{fp}(\underline{\kappa}'') = Z \tag{117}$$

with $0 \le Z \le 1$, $Z \ne \mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa})$, only if $\delta_j' = \delta_r'' = \delta(Z)$, where $\delta(Z)$ is a function of Z. Note that $\delta(Z) > 0$ if $Z > \mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa})$, and $\delta(Z) < 0$ otherwise.

According to the notation introduced in Appendix C, the difference between the rewards achieved with policies $\underline{\kappa}'$ and $\underline{\kappa}''$ is

$$\mathcal{W}_{S}(\underline{\kappa}') - \mathcal{W}_{S}(\underline{\kappa}'') = \frac{\mathcal{N}_{W} + (G+F) \delta(Z)}{\mathcal{D} + (A+C) \delta(Z)} - \frac{\mathcal{N}_{W} + G \delta(Z)}{\mathcal{D} + A \delta(Z)}$$
(118)

which is larger than zero if the following holds:

$$\delta(Z)\bigg(F\left(\mathcal{D} + A\delta(Z)\right) - C(\mathcal{N}_W + G\delta(Z))\bigg) > 0. \quad (119)$$

Since F > C, as previously stated, and

$$\mathcal{W}(\underline{\kappa}'') = \mathcal{W}(\underline{\kappa} + \underline{u}_r \delta(Z)) = \frac{\mathcal{N}_W + G\delta(Z)}{\mathcal{D} + A\delta(Z)} \le 1$$
 (120)

by construction, then

$$F\left(\mathcal{D} + A\delta(Z)\right) - C(\mathcal{N}_W + G\delta(Z)) > 0. \tag{121}$$

Finally, if $Z>\mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa})$, then $\delta(Z)>0$ and $\mathcal{W}(\underline{\kappa}')>\mathcal{W}(\underline{\kappa}'')$. If $Z<\mathcal{J}_{\mathrm{P}}^{\mathrm{fp}}(\underline{\kappa})$, then $\delta(Z)<0$ and $\mathcal{W}(\underline{\kappa}')<\mathcal{W}(\underline{\kappa}'')$.

REFERENCES

- Q. Zhao, L. Tong, A. Swami, and Y. Chen, "Decentralized cognitive MAC for opportunistic spectrum access in ad hoc networks: A POMPD framework," *IEEE J. Sel. Areas Commun.*, vol. 25, no. 3, pp. 589–600, Apr. 2007.
- [2] O. Simeone, Y. Bar-Ness, and U. Spagnolini, "Stable throughput of cognitive radios with and without relaying capability," *IEEE Trans. Wireless Commun.*, vol. 55, no. 12, pp. 2351–2360, Dec. 2007.
- [3] S. Geirhofer, L. Tong, and B. M. Sadler, "Dynamic spectrum access in the time domain: Modeling and exploiting white space," *IEEE Commun. Mag.*, vol. 45, no. 5, pp. 66–87, May 2007.
- [4] R. Urgaonkar and M. J. Neely, "Opportunistic scheduling with reliability guarantees in cognitive radio networks," in *Proc. 27th IEEE Conf. Comput. Commun.*, Phoenix, AZ, Apr. 2008, pp. 1301–1309.
- [5] W. Zhang and U. Mitra, "A spectrum-shaping perspective on cognitive radio: Uncoded primary transmission case," in *Proc. IEEE Int. Symp. Inf. Theory*, Toronto, ON, Canada, Jul. 2008, pp. 1338–1342.
- [6] L. Cao and H. Zheng, "Distributed rule-regulated spectrum sharing," IEEE J. Sel. Areas Commun., vol. 26, no. 1, pp. 130–145, Jan. 2008.
- [7] H. Su and X. Zhang, "Cross-layer based opportunistic MAC protocols for QoS provisioning over cognitive radio wireless networks," *IEEE J. Sel. Areas Commun.*, vol. 26, no. 1, pp. 118–129, Jan. 2008.
- [8] Y. Xing, C. N. Mathur, M. A. Haleem, R. Chandramouli, and K. P. Subbalakshmi, "Dynamic spectrum access with QoS and interference temperature constraints," *IEEE Trans. Mobile Comput.*, vol. 6, no. 4, pp. 423–433, Apr. 2007.
- [9] J. Mitola, Cognitive Radio: An Integrated Agent Architecture for Software-Defined Radio. Stockholm, Sweden: Royal Inst. Technol., 2000.
- [10] R. Zhang, "Optimal power control over fading cognitive radio channel by exploiting primary user CSI," in *Proc. IEEE Global Telecommun. Conf.*, Nov.–Dec. 30–4, 2008, pp. 1–5.

- [11] P. Popovski, H. Yomo, K. Nishimori, R. D. Taranto, and R. Prasad, "Opportunistic interference cancellation in cognitive radio systems," in *Proc. 2nd IEEE Int. Symp. New Front. Dyn. Spectrum Access Netw.*, Apr. 2007, pp. 472–475.
- [12] M. Gastpar, "On capacity under receive and spatial spectrum-sharing constraints," *IEEE Trans. Inf. Theory*, vol. 53, no. 2, pp. 471–487, Feb. 2007.
- [13] M. Levorato, U. Mitra, and M. Zorzi, "On optimal control of wireless networks with multiuser detection, hybrid ARQ and distortion constraints," in *Proc. 28th IEEE Conf. Comput. Commun.*, Rio de Janeiro, Brazil, Apr. 2009, pp. 1926–1934.
- [14] R. A. Tannious and A. Noratinia, "Cognitive radio protocols based on exploiting hybrid arq retransmissions," *IEEE Trans. Wireless Commun.*, vol. 9, no. 9, pp. 2833–2841, Sep. 2010.
- [15] K. Eswaran, M. Gastpar, and K. Ramchandran, "Bits through ARQs: Spectrum sharing with a primary packet system," in *Proc. IEEE Int. Symp. Inf. Theory*, Jun. 2007, pp. 2171–2175.
- [16] R. Zhang, "On active learning and supervised transmission of spectrum sharing based cognitive radios by exploiting hidden primary radio feedback," *IEEE Trans. Commun.*, vol. 58, no. 10, pp. 2960–2970, Oct. 2010.
- [17] T. Cover and J. Thomas, *Elements of Information Theory*. New York: Wiley. 1991.
- [18] D. P. Bertsekas, Dynamic Programming and Optimal Control, 2nd ed. Belmont, MA: Athena Scientific, 2001, vol. 2.
- [19] K. W. Ross, "Randomized and past-dependent policies for Markov decision processes with multiple constraints," *Oper. Res.*, vol. 37, no. 3, pp. 474–477, May–Jun. 1989.
- [20] S. P. Meyn and R. L. Tweedie, Markov Chains and Stochastic Stability. London, U.K.: Springer-Verlag, 1993.
- [21] S. Mahadevan, "Average reward reinforcement learning: Foundations, algorithms, and empirical results," *Mach. Learn.*, vol. 22, no. 1, pp. 159–195, 1996.
- [22] F. Fu and M. Van Der Schaar, "Structure-aware stochastic control for transmission scheduling," Arxiv Preprint ArXiv:1003.2471 2010.
- [23] S. Firouzabadi, M. Levorato, D. ONeill, and A. Goldsmith, "Learning interference strategies in cognitive arq networks," in *Proc. IEEE Global Telecommun. Conf.*, Nov. 2010, pp. 1–6.
- [24] M. Levorato, S. Firouzabadi, and A. Goldsmith, "Cognitive interference networks with partial and noisy observations: A learning framework," in *Proc. IEEE Global Telecommun. Conf.*, pp. 1–6.
- [25] L. Kaelbling, M. Littman, and A. Cassandra, "Planning and acting in partially observable stochastic domains," *J. Artif. Intell.*, vol. 101, no. 1–2, pp. 99–134, 1998.

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