```
PROBABILITY PROBLEMS
                           Fri Apr 1 05:00:15 EDT 2016
   pseudopi
       Throwing chalk.
       Boston Preliminary 2005.
   birthday
       Pseudo-random cannot be really random.
       Boston Preliminary 2008
   pagerank
        Whats popular these days?
       Boston Preliminary 2012
   betterbias
       The Better Bias Bureau improves biased dice.
       Boston Preliminary 2002.
    trends
       Luck or unluck tis common enough.
       Boston Preliminary 2015
   markov
       What is the future of my state?
       Boston Preliminary 2014
```

Pseudo-Random Computation of PI

One of the classic demonstrations of probability is the following. The professor draws a large square on the blackboard, and draws its inscribed circle. Then standing with her back to the board, she throws pieces of chalk at the square. After this she counts the number M of hits in the circle and the number N  $\geq$  M of hits in the square (including those in the circle), and demonstrates that M/N is about PI/4. This is because PI/4 is the area of the inscribed circle divided by the area of the square, and the probability of hitting any small part of the square is roughly identical to hitting any other small part of the square.

You have been asked to simulate the demonstration in the computer. The square is to be simulated by the unit square in the XY-plane, [0,1]x[0,1], which has (0,0) as its lower left corner and (1,1) as its upper right corner. To simulate throwing the chalk, two random integers X and Y are 'drawn uniformly' (see below for details) from the range 0...S-1, where S>0 is some integer. Then the coordinates where the chalk strikes are set at ((X+0.5)/S, (Y+0.5)/S). These are inside the square, so all our 'throws' count toward N. They are inside the circle, and count toward M, if and only if the chalk strikes at a distance of 0.5 or less from the center of the circle, (0.5, 0.5).

Thus if S = 100 and the first two random integers drawn are 37 and 69, the chalk point is (0.375, 0.695) which is distance 0.23 from (0.5, 0.5), and is therefore in the circle and counts toward both M and N.

Drawing Random Numbers

You are asked to draw pseudo-random numbers according to the equation:

RANDOM = ( RANDOM \* MULTIPLIER ) mod MODULUS

where RANDOM is the value of the pseudo-random number, the equation steps from the the last pseudo-random number to the next pseudo-random number, and MULTIPLIER and MODULUS are fixed values that determine the pseudo-random number sequence.

To get started, RANDOM is initialized to a value called SEED. The first pseudo-random number in the sequence is not SEED, but the first number after SEED in the sequence.

If MULTIPLIER and MODULUS have good values for this purpose, the resulting sequence of numbers appears when tested to be truly random and uniformly distributed in the range from 1 through MODULUS - 1. Uniformly distributed means all values in this range are equally probable. The choices

MULTIPLIER = 7\*\*5 = 16807 MODULUS = 2\*\*31 - 1 = 2147483647

are very good for this purpose.

For example, if MULTIPLIER and MODULUS are as just given, and the SEED is 374332679, then the first two random numbers are 1429733890 and 1342962947.

A remaining difficulty is how to convert uniformly distributed integers from 1 through MODULUS - 1 to uniformly distributed integers from 0 through S-1. An easy solution, which we will adopt, is to set

```
S = MODULUS - 1
```

and subtract 1 from each value of RANDOM. Thus 'a chalk throw' is simulated by executing

```
RANDOM = ( MULTIPLIER * RANDOM ) mod MODULUS
X = RANDOM - 1
X = (X + 0.5 ) / S
RANDOM = ( MULTIPLIER * RANDOM ) mod MODULUS
Y = RANDOM - 1
Y = (Y + 0.5 ) / S
```

to yield (X,Y) in the unit square.

Implementation of the above algorithm requires integers longer than 32 bits. In C or C++ you can use doubles and the fmod function. Or you can use 'long long's and the % operator. In JAVA you can use 'long's and the % operator. Remember, 'long's are only 32 bits in C and C++, but are 64 bits in JAVA. 'long long's are 64 bits in C and C++.

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Input
```

For each of several test cases, one line containing four numbers in the order:

#### N MULTIPLIER MODULUS SEED

The numbers may be separated by spaces or tabs. All input numbers are positive integers below 2\*\*31 (but some products computed by intermediate computations will be larger).

Input ends with an end of file.

The simulation is to be done with RANDOM initialized to SEED (SEED is NOT the first pseudo-random number) and S = MODULUS - 1.

#### Output

----

For each test case one line containing five numbers in the order:

N MULTIPLIER MODULUS SEED PI\_ESTIMATE

where the first four numbers are copied from the input, and PI\_ESTIMATE equals 4\*M/N expressed as a decimal number with exactly 5 decimal places.

## Example Input

\_\_\_\_\_

100	16807	2147483647	374332679
1000	16807	2147483647	374332679
10000	16807	2147483647	374332679
100000	16807	2147483647	374332679
1000000	16807	2147483647	374332679

#### Example Output

\_\_\_\_\_

100 16807 2147483647 374332679 3.20000 1000 16807 2147483647 374332679 3.13600 10000 16807 2147483647 374332679 3.15960 100000 16807 2147483647 374332679 3.13888 1000000 16807 2147483647 374332679 3.14167 pseudopi.txt 02/22/16 04:50:18 3 of 3

File: pseudopi.txt

Author: Bob Walton <walton@seas.harvard.edu>

Date: Wed Oct 19 07:19:20 EDT 2005

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The Birthday Paradox

Suppose you pick m numbers at random, each between 1 and n. How large does m have to be before we can expect to see two numbers that are the same?

The answer is about the square root of 2n, which is surprisingly small. For example, if n=365, the number of days in the year, m=28 will do. One way to pick numbers from 1 through 365 is to pick people and take their birthdays. Thus if you have 28 people in a room, you can expect two to have the same birthday. For this reason the phenomenon we have just described is called the 'Birthday Paradox'.

The Birthday Paradox has some surprising applications. Suppose you have a pseudo-random number generator that uses the sequence

$$x(i+1) = A*x(i)^2 + B*x(i) + C$$
 (modulo n)  
 $x(0) = D$ 

for some constants n, A, B, C, and D to generate numbers x(0), x(1), x(2), ... which we hope act like a sequence of random numbers. Suppose they really are like random numbers. Then by the birthday paradox the sequence of numbers will start repeating itself after about m = square root of 2n numbers. That is, if we find the smallest m > 0 and i > 0 such that x(i+m) = x(i), then typical values of m and i are on the order of the square root of 2n. Note that given such values, by the nature of the above equation, x(j+m) = x(j) for all j >= i, and we say the sequence has a cycle of length m.

The theory here is not rigorous, because the sequence x(0), x(1), x(2), is not a rigorously random sequence. In this problem you will be given n, A, B, C, and D and asked to compute i and m (the start and length of first cycle) and print these and the square root of n. Just to see if the theory works most of the time.

Input

----

For each test case, one line containing

n A B C D

Here  $2 \le n \le 40,000$ ;  $0 \le A,B,C,D \le n$ . The input ends with an end of file.

Output

----

For each test case, one line containing

nABCDimr

where r = ceil ( sqrt ( 2 \* n ) ) as an integer (ceil is the ceiling function and sqrt the square root function), and each integer of the 8 integers is printed right adjusted in exactly 7 columns.

## Remark

\_\_\_\_

We limit n <= 40,000 in order to permit implementations to use 32 bit integers. Note, however, that A \* x \* x may not fit into 32 bits, though x \* x and A \* n will. If you want to use 64 bit integers ('long long' in C and C++ and 'long' in JAVA), you can compute with larger values of x.

If you use a vector of integers of length n the size of n will still be limited. But there are simple clever implementations that do not need a vector and use almost no memory for any size of n.

# Sample Input

-----

100 43 23 17 5 199 0 2 0 1 8191 5 2685 0 7 32749 0 1944 0 5

# Sample Output

-----

100	43	23	17	5	2	2	15
199	0	2	0	1	0	99	20
8191	5	2685	0	7	155	115	128
32749	0	1944	0	5	0	32748	256

#### Remark

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The pseudo-random number generators that are actually used pick A, B, C, and D so the shortest cycle is of length n or n-1. One has to be smart about picking A, B, C, and D. One old but usable set is

$$A = 0$$
,  $B = 7^5$ ,  $C = 0$ ,  $D > 0$ ,  $n = 2^31 - 1$ 

or in other words.

$$x(i+1) = 7^5 * x(i) modulo (2^31-1)$$

with D any non-zero value. Another set that you can test your program with is

$$A = 0$$
,  $B = 1944$ ,  $C = 0$ ,  $D > 0$ ,  $n = 2^15 - 19$ 

which should have i = 0 and m = n - 1, the maximum possible cycle length.

The irony is that to get a good random number generator, the sequence cannot really be random.

#### Remark

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Pollard's rho algorithm makes clever use of cases such as

$$A = 1$$
,  $B = 0$ ,  $C = n - 2$ 

to find factors of n for values of n up to 2^256 + 1.

birthday.txt 02/22/16 04:37:08 3 of 3

File: birthday.txt

Author: Bob Walton <walton@seas.harvard.edu>

Date: Wed Oct 15 09:18:40 EDT 2008

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# PageRank

Google uses a page ranking algorithm to determine the importance of each web page. The rank of a page is the probability that a particular random web surfer will be looking at the page at any given moment.

Suppose we represent the web as a graph with N nodes, each representing a page, and with a directed edge representing each link from one page to another.

The random surfer in question behaves as follows.

The surfer chooses a first page randomly from among the N pages.

To move from a current page to the next page, the surfer does the following. If the current page is the source of zero links, the surfer chooses the next page randomly from among the N pages of the web. Otherwise, with probability 1-ALPHA, the surfer does the same thing (as if the page sourced zero links), and with probability ALPHA, the surfer chooses a link sourced at the current page at random and follows that link.

When choosing from among a set of pages or links at random the surfer gives equal probability to each page or link that might be chosen. So to choose at random from among the N pages of the web, each page has probability 1/N of being chosen. And to choose at random from among Q links sourced at the current page, each link has probability 1/Q of being chosen.

You have been asked to compute the probability for each page that that page will be the K'th page visited by the surfer, for a given web graph and value of K.

### Input

----

For each of several test cases, first a line containing the name of the test case, and then a line containing

#### N ALPHA K

where N is the number of pages and K is the number of pages the surfer is to visit. After these two lines are N lines each containing a list of page numbers followed by a 0. Pages are numbered 1, 2, ..., N. The i'th of these lines contains the numbers of the pages targeted by links sourced at page i.

So in the sample input below, node 1 has two links to node 2, node 2 has links to nodes 1, 2, and 3, and node 3 has zero links. Note that a node may have several links to the same target node and may have links to itself.

 $1 \le N \le 100; 0 \le ALPHA \le 1; 1 \le K \le 10,000.$ 

Input ends with an end of file.

#### Output

\_\_\_\_\_

For each test case, one line containing the name of the test case (copied exactly from the input), followed by N lines each containing a page number i followed by the probability the K'th page surfed will be page i. The page numbers must be in increasing order. The probabilities must be output with exactly 6 decimal places.

```
Sample Input
-----
-- SAMPLE 1 --
3 0.00 2
2 2 0
3 2 1 0
-- SAMPLE 2 --
3 1.00 2
2 2 0
3 2 1 0
-- SAMPLE 3 --
3 0.85 3
2 2 0
3 2 1 0
Sample Output
-- SAMPLE 1 --
1 0.333333
2 0.333333
3 0.333333
-- SAMPLE 2 --
1 0.222222
2 0.555556
3 0.222222
-- SAMPLE 3 --
1 0.265648
2 0.468704
3 0.265648
```

pagerank.txt

### Reference:

"PageRank: Standing on the Shoulders of Giants", by Massimo Franceschet, Communications of the ACM, Vol 54, No 6, June 2011, pp 92-101.

File: pagerank.txt

Author: Bob Walton <walton@seas.harvard.edu>

Date: Sat Oct 6 03:26:45 EDT 2012

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Achieving Better Bias

The Better Bias Bureau takes M faced biased dice with known biases and makes from them N faced biased dice with desired biases. More specifically, if you have an M faced biased die with probability pm[j] of throwing face j, for j from 1 though M, but you really want instead an N faced biased die with probability pn[i] of throwing face i, for i from 1 through N, then good old B3 will give you a method of achieving your desires.

The method is this. You are to consider throwing your M faced die an infinite number of times and writing out the resulting faces, f1, f2, f3, ..., as an infinite fraction .f1f2f3... base M. The fraction will be a number, call it F, in the range from 0 through 1. For any number r, 0 <= r <= 1, we can compute the probability that F <= r.

Let rn[i] be defined so that

- probability F <= rn[i]</pre>
- = probability that an N faced die throw is k for some k <= i</pre>
- = sum of pn[k] for k from 1 through i

Then when you throw a value F with your M faced die, you should declare it to be face i of the N faced die if

This method produces a properly biased N faced die. Here one ignores the cases where F = rn[i] for some i, as the probability of this happening is 0.

The neat thing about all this is that you do not actually have to throw the M faced die an infinite number of times to decide what the value of i is. After some finite number of throws, depending on the value of F, you can stop, knowing the answer i.

Suppose you stop as soon as possible. What is the expected number of throws of the M faced die necessary to produce one throw of the N faced die?

Note:

----

For technical reasons, relating to the elimination of ambiguity in judging a contest, we use only one inequality if i=1 or i=N. To be a little more specific, the judge's test data has been adjusted so that for all cases that must actually be tested, either both the inequalities rn[i-1] < F and F < rn[i] will be true by a margin definitely larger than the accuracy of the computation, or one of these inequalities will be false by a margin definitely larger than the accuracy of the computation. But it is NOT possible to adjust the judge's data so the inequalities at the ends, rn[0] = 0.0 < F and F < rn[N] = 1.0, will be true or false by a margin definitely larger than the accuracy of the computation. So we suppress these two inequalities, which we can do without changing the validity of the method.

# Input

\_\_\_\_

For each of several test cases:

- \* M followed by pm[1], pm[2], ..., pm[M] in order.
- \* N followed by pn[1], pn[2], ..., pn[N] in order.

Here 2 <= M <= 20, 2 <= N <= 20, and for all j and i 0.01 <= pm[j] <= 0.99, 0.01 <= pn[i] <= 0.99. The sum of all the pm[j] equals 1.0, and the sum of all pn[i] equals 1.0.

Numbers are separated by whitespace, which may consist of any sequence of spaces, tabs, or newlines.

Input ends with a line containing just a single 0.

You should use double precision floating point arithmetic to do input and computation, as more than 6 significant figures of accuracy are necessary.

# Output

\_\_\_\_\_

For each test case, one line containing the expected number of throws of the M sided die needed to generate one throw of the N sided die. This number should have exactly 2 decimal places.

# Note

3.00

You might find it interesting to consider at your leisure whether B3's method is optimal, i.e., achieves the minimal expected number of throws of the M faced die to generate one throw of the N faced die.

betterbias.txt 02/22/16 04:36:41 3 of 3

File: betterbias.txt

Author: Bob Walton <walton@deas.harvard.edu>

Date: Mon Oct 21 01:22:35 EDT 2002

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# Trends

Jack plays a solitary game in which he flips a coin and gives himself a point on heads but removes a point on tails. He starts with a score of zero, and keeps score as he flips coins. You would think that his score would not remain strictly positive for long, or strictly negative for long, but you would be wrong. Such long periods of being strictly positive or strictly negative are often mis-identified as trends that supposedly prove the coin is biased.

You have been asked to find the probability of a trend of length >= m in a sequence of n coin flips. Here a 'trend' is defined as a sequence of consecutive flips such that at the end of each flip Jack's score is strictly positive, or, at the end of each flip Jack's score is strictly negative. For comparison purposes, you are also asked to do this for both unbiased and biased coins.

Input

For each of several test cases, a line containing

n m p

where n and m are as above and p is the probability of heads.

 $0 < m,n \quad m \le n \quad (m**2)*n \le 50,000,000 \quad 0 \le p \le 1$ 

Input ends with an end of file.

### Output

-----

For each test case, a single line containing

n m p ptrend

where n, m, p are copied from the input line and ptrend is the probability of finding a trend of length >= m in a sequence of n flips. p and ptrend are to be printed with exactly 3 decimal places (even if p is input with fewer decimal places), whereas n and m are integers.

## Sample Input

-----

10 5 0.5

10 5 0.55

10 8 0.5

10 10 0.5

20 10 0.5

20 10 0.55

20 20 0.5

#### Sample Output

\_\_\_\_\_

10 5 0.500 0.703

10 5 0.550 0.712

10 8 0.500 0.410

10 10 0.500 0.246

20 10 0.500 0.666

20 10 0.550 0.688

20 20 0.500 0.176

File: trends.txt

Author: Bob Walton <walton@seas.harvard.edu>

Date: Thu Aug 13 06:26:42 EDT 2015

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Markov Recurrence

A finite Markov Chain is a finite set of N states S1, S2, ..., SN and a matrix p[i,j] of probabilities, 1 <= i,j <= N, such that if the 'system' is in state Si, the next state of the system will be Sj with probability p[i,j]. It is required that

$$0 \le p[i,j] \le 1$$
  
sum  $(1 \le j \le N) p[i,j] = 1$ 

Thus the 'transition' of the system from Si to Sj has probability p[i,j].

If a state sequence starts from Si it may or may not return to Si; such a return is called a 'recurrence'. Let f[i,t] be the probability that the sequence returns to Si for the FIRST time after the t'th transition (so f[i,1] == p[i,i]). Here t => 1 is thought of as 'time'. Then

$$f[i] = sum (1 <= t) f[i,t]$$

is the probability that the system returns to Si at some future time (the probability that Si ever recurs).

A state Si is said to be persistent if f[i] == 1, and to be transient if f[i] < 1.

A state Si is said to be no-return if f[i] == 0.

A state Si is said to be periodic if it is NOT no-return and there is an integer s > 1 such that f[i,t] == 0 if t mod s != 0. The largest such s is the period of Si.

A state Si that is NEITHER no-return or periodic is said to be aperiodic.

You have been asked to compute f[i] for the states Si of a Markov Chain and determine if Si is transient or persistent and if Si is no-return, periodic, or aperiodic. In the periodic case you are to determine the period.

## Input

----

For each of several test cases, first a line containing just the test case name, then a line containing just N, and then N lines containing the probabilities in the layout

```
p[1,1] p[1,2] ... p[1,N]
p[2,1] p[2,2] ... p[2,N]
.....
p[N,1] p[N,2] ... p[N,N]
```

 $1 \le N \le 100$ ,  $0 \le p[i,j] \le 1$ , and for each i, sum p[i,j] over all j = 1.

Input probabilities may have many decimal places and the lines containing them may be long. Double precision floating point will suffice for input and computation.

Input ends with an end of file.

#### Output

\_\_\_\_\_

For each test case, the first a line containing an exact copy of the test case name input line, then N lines each with the format:

$$f[#] = #.### X Y$$

where # denotes a digit, X is either 'persistent' or 'transient', Y is either 'no-return', 'period #' or 'aperiodic'.

The only whitespace in the output are 4 single spaces, 2 surrounding the =, 1 before X, and 1 before Y. The N lines are in order of increasing f[#] index, i.e., f[1], f[2], ..., f[N].

The input will be such that a state i will be persistent if and only if f[i] >= 0.9995, and there will be no ambiguous cases.

WARNING: f[i] < 0.0005 does NOT mean the state is no-return. You must use a calculation not involving f[i] to determine whether a state is no-return (use your period calculation).

```
Sample Input
-----
-- SAMPLE 1 --
2
0 1
1 0
-- SAMPLE 2 --
2
0.5 0.5
0 1
-- SAMPLE 3 --
2
0 1
0 1
```

```
Sample Output
-----
-- SAMPLE 1 --
f[1] = 1.000 persistent period 2
f[2] = 1.000 persistent period 2
-- SAMPLE 2 --
f[1] = 0.500 transient aperiodic
f[2] = 1.000 persistent aperiodic
-- SAMPLE 3 --
f[1] = 0.000 transient no-return
f[2] = 1.000 persistent aperiodic
Note
----
```

For finite markov chains, persistent aperiodic states are called 'ergodic'. For infinite markov chains, persistent aperiodic states with finite expected time of return are called 'ergodic', but there are also persistent aperiodic states with infinite expected time of return which called 'null states'.

markov.txt 02/22/16 04:44:49 3 of 3

File: markov.txt

Author: Bob Walton <walton@seas.harvard.edu>

Date: Wed Oct 15 07:30:03 EDT 2014

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