# High Dimensional Data Analysis With Dependency and Under Limited Memory

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Multivariate Spectral Density Estimation Under Weak Sparsity

Low Rank Tucker Approximation of a Tensor from Streaming Data

A Connection of These Two

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#### **Motivation**

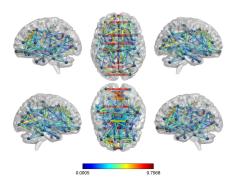


Figure: Interactions Between Regions in Brain

## Weakly & Strongly Stationary Time

## Definition (Weak Stationarity)

p-variate time series X is weakly stationary, if  $\mathbb{E}X_t = \mathbb{E}X_s$  for any t,s and  $\Gamma(\ell) := \mathbb{E}X_t X_{t-\ell}^{\top}$  only depends on the lag  $\ell$ .

## Definition (Strong Stationarity)

p-variate time series X is strongly stationary, if for any sequence  $t_1, \cdots, t_n, X_{t_1} \cdots X_{t_n}$  has the same distribution of  $X_{t_1+\tau} \cdots X_{t_n+\tau}$  for any integer  $\tau$ ,

#### **Gaussian Process**

### Definition (Gaussian Process)

*p*-variate time series X is Gaussian process if for any sequence  $t_1, \dots, t_n, X_{t_1} \dots X_{t_n}$  are jointly Gaussian distributed.

For Gaussian process, weak stationarity is equivalent to strong stationarity.

#### **Spectral Density**

Given a weakly stationary p-variate time series X, the spectral density at frequency  $\omega \in [-\pi,\pi)$  is defined

$$f(\omega) = \sum_{\ell=-\infty}^{\infty} \Gamma(\ell) e^{-i\omega\ell}$$

where  $\Gamma(\ell) = \mathbb{E} X_0 X_{-\ell}^{\top}$ .  $X_t$  is independent with  $X_s$ ,  $t \neq s$  iff  $f_{rs}(\omega) = 0$  for any  $\omega$ .

## Thresholding Estimator Under Weak Sparsity- A Example

Suppose that we have n observation of p-variate Gaussian distribution as follows.

$$y_i \overset{i.i.d}{\sim} \mathcal{N} \left( \mu, \begin{bmatrix} \sigma_1^2 & 0 & \cdots & 0 \\ 0 & \sigma_2^2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \\ 0 & 0 & 0 & \sigma_p^2 \end{bmatrix} \right),$$

$$i=1,\cdots,n$$
.

#### **A Example**

The maximum likelihood estimator for  $\mu_j$  is  $\bar{y}_j = \frac{1}{n} \sum_{i=1}^n y_{ij}$ . Does not Work Well Under Weak Sparsity

$$\mu \in \left\{ \mu \in \mathbb{R}^p, \sum_{j=1}^p |\mu_j|^q \leq c_0(p) \right\}.$$

for some  $0 \le q < 1$  and  $c_0(p)$  measures the weak sparsity.

#### **Solution: Thresholding**

Suppose  $\sigma_i \leq B$ , define element-wise thresholding operator

$$T_{\lambda}(x) = \begin{cases} x & |x| \ge \lambda \\ 0 & \text{else} \end{cases}$$

. Hard thresholding estimator  $T_{\lambda}(\bar{y}_j)$  can be shown asymptotically consistent under weak sparsity where we set

$$\lambda \propto Bc_0(p)\sqrt{\frac{\log p}{n}}$$

and assume  $\lambda \to 0$ .

## Two Key Ingredients for Thresholding

Two key ingredients under above example assuming weak sparsity. Cai & Liu 2011

► An element-wise concentration inequality :

$$\mathbb{P}(|\bar{y}_j - \mu_j| \ge \eta) \le 2 \exp(-n\eta^2/2\sigma_j^2).$$

 $ightharpoonup \sigma_j$  are uniformly bounded.

#### **Shortcomings for Hard Thresholding**

- $ightharpoonup \sigma_i$  may variate much
- ► B will appear in the thresholding value making convergence rate slow

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#### **Solution: Adaptive Thresholding**

Simply estimate  $\sigma_i$ , say with sample standard deviation:

$$\hat{\sigma}_j = \sqrt{1/(n-1)\sum_{i=1}^n (y_{ij} - \bar{y}_j)^2}$$

and replace  $B: \lambda_j \propto \hat{\sigma}_j c_0(p) \sqrt{\frac{\log p}{n}}.$ 

#### An Similar Example: Covariance Matrix

$$y_i \stackrel{i.i.d}{\sim} \mathcal{N}(0, \Sigma_{p \times p})$$

Goal: Estimate  $\Sigma$  assuming weak sparsity,  $\|\Sigma\|_1 \leq c_0(p)$  . Bickel & Levina 2008

#### An Similar Example: Covariance Matrix

Estimate the expectation of a vector of length  $p^2$ :

$$[(y_1y_1^{\top})_{rs}, 1 \leq r, s \leq p].$$

Remember Two In

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#### References

Bickel, P. J. & Levina, E. (2008). Covariance regularization by thresholding. *The Annals of Statistics*, 2577–2604.

Cai, T. & Liu, W. (2011). Adaptive thresholding for sparse covariance matrix estimation. Journal of the American Statistical Association, 106(494), 672–684.