

# LLM-Powered Predictive Inference with Online Text Time Series

Yingying Fan<sup>a,1,2</sup>, Jinchi Lv<sup>a,1,2</sup>, Ao Sun<sup>a,1</sup>, and Yurou Wang<sup>b,1</sup>

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Time series predictive inference is an important yet challenging task in economics and business, where existing approaches are often designed for low-frequency, survey-based data. With the recent advances of large language models (LLMs), there is growing potential to leverage high-frequency online text data for improved time series prediction, an area still largely unexplored. This paper proposes the LLM-TS, an LLM-based approach for time series predictive inference incorporating online text data. The LLM-TS is based on a joint time series framework that combines survey-based low-frequency data with LLM-generated high-frequency surrogates. The framework relies only on an error correlation assumption, combining a text-embedding-augmented ARX model for the observed gold-standard measurements with a VARX model for the LLM-generated surrogates. LLM-TS employs LLMs such as ChatGPT and the trained BERT models to construct LLM surrogates. Online text embeddings are extracted via LDA and BERT. We establish the asymptotic properties of the method and provide two forms of constructed prediction intervals. We also extend LLM-TS to incorporate deep learning backbones. To demonstrate the practical power of LLM-TS, we apply it to a critical real-world application: inflation forecast. We construct two large high-frequency online text data sets from the U.S. and China, and use LLMs to extract inflation-related signals from texts that reflect price dynamics. The finite-sample performance and practical advantages of LLM-TS are illustrated through extensive simulations and two noisy real data examples, highlighting its potential to improve time series prediction in economic applications.

Large language models | Inflation prediction | Online texts | Asymptotic distributions | Time series | Deep learning

Time series predictive inference is a longstanding problem with widespread applications in statistics, economics, business, health and medical sciences, genomics, biology, and other fields. Traditional time series prediction methods often rely on structural models (1–3) and laborious field-collected data. While these methods are fully interpretable, they face declining prediction accuracy and the high costs associated with large-scale data collection.

Emerging alternative data streams, particularly text data from news media and social platforms, demonstrate the viability of unstructured content as novel inputs for time series prediction (4–7). Such online data is cheaply collected, and the fact that large language models (LLMs) have capabilities in processing complex linguistic patterns suggests largely untapped potential for time series modeling and research (8, 9). However, the prediction advantages of LLMs are tempered by their black-box nature and limited interpretability regarding the sources of their predictive power.

The above dilemmas motivate a core research question: How can we effectively combine the prediction capabilities of LLMs using online text with the interpretability of established structural models using field-collected data? Our suggested framework addresses such challenge through an integrated forecasting system that harmoniously combines traditional econometric models with the LLM-powered prediction models. The method tackles two key obstacles: 1) the effective combination of limited high-accuracy, low-frequency official data and abundant high-frequency but less robust LLM-powered surrogates; and 2) the inherent conflict between complex machine learning and deep learning architectures, and the need for interpretable results. Through correlation modeling between traditional structure models and LLM-powered surrogates, our approach enhances the predictive inference accuracy while maintaining clear explanations for structural relationships. We name our new framework as the LLM-powered time series predictive inference (LLM-TS).

An illustrative application is inflation forecasting. Inflation is a key macroeconomic indicator that guides monetary policy and reflects overall economic conditions.

## Significance Statement

Large language models (LLMs) are increasingly used to build time-series indices. By scanning online texts, they provide high-frequency signals and low-cost experimentation for applications such as market simulations and policy evaluations. However, compared to structural models using carefully collected field data, LLM-based indices are often less stable and provide limited insight into their underlying mechanisms. This paper introduces a general forecasting framework that merges traditional econometric models with LLM outputs, converting low-quality LLM signals into a statistically coherent framework that yields reliable forecasts and valid prediction intervals. Using a decade of Wall Street Journal news articles and a collection of 159.5 million social media posts, we show that our framework accurately forecasts inflation with robustness across different information environments.

Author affiliations: <sup>a</sup>Data Sciences and Operations Department, Marshall School of Business, University of Southern California, Los Angeles, CA 90089; <sup>b</sup>Paula and Gregory Chow Institute for Studies in Economics, Xiamen University, China 361005

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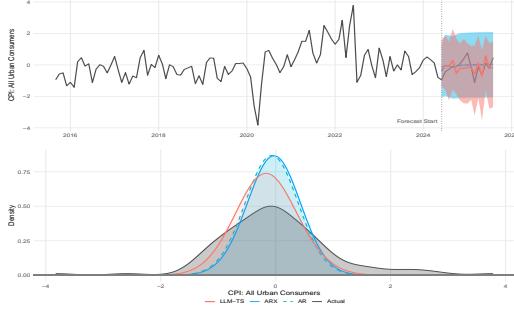
<sup>1</sup>Four authors contributed equally to this work.

<sup>2</sup>To whom correspondence should be addressed. E-mail: fanyingy@usc.edu or jinchilv@usc.edu



**Fig. 1.** A time-series comparison between the LLM-generated daily inflation index (Eq. (22)) and the CPI from September 2015 to August 2025 in the [U.S. economy](#). The blue curve represents the LLM-generated daily inflation index capturing high-frequency variations through analysis of unstructured Wall Street Journal news articles, and the red circles depict the normalized monthly urban CPI. To facilitate the comparison of trend changes, we apply a 30-day moving average smoothing to LLM-generated daily inflation index in the figure. Additionally, both the monthly survey-based CPI and the LLM-generated daily inflation index are standardized and shifted by subtracting 0.5, ensuring that their fluctuations are centered around zero.

It is closely linked to the interest rates, labor-market dynamics (such as wage growth and employment), production costs, and financial-market sentiment (10–13). Traditional forecasting approaches rely on structural autoregressive models, which regress current inflation indices on their own lags and a small set of macro variables (14, 15). A new and complementary information source arises from economic narratives (16) embedded in newspapers, blogs, and social-media discussions about price changes. These online texts provide real-time signals that are often absent from official statistics. In this paper, we examine two distinct economic environments: the U.S. and China. We collect a ten-year text data set from the Wall Street Journal (<https://www.wsj.com>) and a corpus of 159.5 million posts from Sina Weibo (<https://www.weibo.com>), the largest social-media platform in China. From these unstructured texts we use an LLM pipeline to construct a daily, text-based inflation index (Eq. (22)). Fig. 1 and *Fig. S5 in SI Appendix* compare our high-frequency inflation indices derived from the Wall Street Journal (WSJ) news articles and Weibo posts, respectively, with the official monthly Consumer Price Index (CPI), a standard measurement of inflation. The LLM-derived series tracks the same broad inflationary patterns as the gold-standard CPI while providing richer, higher-frequency signals; the daily index exhibits fine-grained volatility that converges toward the monthly survey-based benchmark.



**Fig. 2.** Top panel: The CPI forecasts and corresponding prediction intervals across different dates given by the AR, ARX (unemployment rate as exogenous), and LLM-TS models. Bottom panel: The kernel density plots of predicted CPI values given by different models compared to the actual CPI values. The LLM-TS model exploits the LLM-generated synthetic surrogates constructed using ChatGPT and the trained BERT models, and the LDA embeddings for online text embeddings, as discussed in Section [LLM-CPI](#). We adopt the period from January 2019 to [May 2024](#) as the training sample, and the period from [June 2024](#) to [August 2025](#) to evaluate the out-of-sample forecasts.

The proposed LLM-TS leverages the predictive power of LLMs and the large-scale data sources of online text time series to enhance the effectiveness of [predictive inference](#). Fig. 2 and *Fig. S10 in SI Appendix* illustrate the prediction power of the suggested LLM-TS method in forecasting high-frequency inflation dynamics using WSJ and Weibo text data, respectively. The black solid curve represents the standardized actual CPI values. Here, we consider three models: a classical autoregressive (AR) model relying on the historical CPI values (green dashed curve); an autoregressive model with the unemployment rate as an exogenous predictor (ARX), also referred to as Gordon's "triangle model" (17) (blue solid curve); and our suggested LLM-TS model that combines the prediction power of the structural ARX model and the LLM-powered surrogate model (red solid curve).

The results show clearly that the LLM-TS most closely tracks the actual CPI trajectory, successfully capturing turning points and underlying trends. More importantly, the prediction interval given by the LLM-TS model is substantially narrower than those given by the AR and ARX models, while still maintaining the desired nominal coverage rate (i.e., 95%). In contrast, the prediction intervals from traditional models are excessively wide and less informative. The advantages of the LLM-TS are more pronounced in the Chinese economic setting, which is considerably more challenging due to data scarcity. These findings highlight the effectiveness of the LLM-TS in delivering both accurate point forecasts and tight uncertainty quantification, especially in the presence of complex economic signals. While inflation forecasting serves as a concrete application in this work, we demonstrate later that it is broadly applicable and can be extended to more complex time-series prediction tasks, including those built upon state-of-the-art deep learning backbone models.

## Problem setup

We observe  $\{(y_t, \mathbf{z}_t), t = 1, \dots, T\}$ , where  $y_t \in \mathbb{R}$  is a continuous measurement of interest, and  $\mathbf{z}_t \in \mathbb{R}^d$  represents related field-collected data that may have predictive power for  $y_t$ . Additionally, we have a data set of online text time series  $\{\mathcal{D}_t, t = 1, \dots, T\}$ , where  $\mathcal{D}_t$  is an unstructured text data set collected during period  $t$ .

Our goal is to construct a  $(1 - \alpha)100\%$  prediction interval  $\text{PI}_{T+h}$  for  $h$ -step-ahead prediction, such that

$$\liminf_{T \rightarrow \infty} \mathbb{P}\{y_{T+h} \in \text{PI}_{T+h}\} \geq 1 - \alpha.$$

Traditional approaches rely on structural models, such as the AR or ARX models, where  $\mathbf{z}_t$  serves as exogenous variables. However, these methods often generate overly wide and noninformative prediction intervals  $\text{PI}_{T+h}$ .

The key question we address is: How can the online text time series  $\{\mathcal{D}_t, t = 1, \dots, T\}$  be effectively leveraged in predictive inference to achieve both valid and informative predictions under mild assumptions? To tackle this problem, we adopt a two-step procedure. First, we use LLMs to form surrogate data  $\{(\mathbf{y}_t^S, \mathbf{x}_t), t = 1, \dots, T\}$  using the online text time series, where  $\mathbf{y}_t^S \in \mathbb{R}^K$  is an LLM-powered surrogate vector correlated with the target response  $y_t$ , and  $\mathbf{x}_t \in \mathbb{R}^p$  represents text embeddings from period  $t$ . This step is nontrivial, requiring the extraction of meaningful signals from highly noisy text data. We provide an information extraction process in this work (i.e., Algorithm 1).

Based on  $\{(\mathbf{y}_t^S, \mathbf{x}_t), t = 1, \dots, T\}$ , we propose the LLM-powered time series predictive inference (LLM-TS) framework, which integrates LLM-generated surrogates with conventional field-collected measurements. The framework combines a text-embedding-augmented autoregressive model for the observed gold-standard measurements with a vector autoregressive model for the LLM-generated surrogates. These two models are connected through their cross-sectional error correlation structure. Extending the methodology of McCaw et al. (18) to time-series settings, LLM-TS achieves improved time series prediction by reducing the model error of the text-embedding-augmented autoregressive model using the surrogate data.

The proposed LLM-TS framework provides tight prediction intervals with theoretical guarantees, as demonstrated through simulations and real-world data examples.

## Related work

**Prediction-powered inference and synthetic surrogate joint modeling.** Our study contributes to recent methodological developments in *prediction-powered inference* and *synthesized surrogate joint modeling*, which aim to incorporate LLM-generated predictions into formal statistical procedures. For example, Angelopoulos et al. (19) and Zrnic and Candès (20) proposed combining experimental and synthetic data through cross-validation to improve statistical efficiency, while McCaw et al. (18) stabilized raw LLM outputs by treating them as synthetic proxies within a joint-likelihood framework. These works focus primarily on estimation and hypothesis testing rather than time-series forecasting. A recent work of Bashari et al. (21) introduced the synthetic-powered conformal predictive inference. Their framework is highly general and yields strong performance when the surrogate and target nonconformity score distributions are well aligned. Our approach provides a distinct direction based on the joint residual modeling, which does not require such distributional alignment. A more detailed comparison between the two methods is provided in *Section 6 of SI Appendix*.

**Time series predictive inference.** Our work is also related to time series predictive inference, a long-standing problem in statistics and econometrics. Phillips (1) rigorously discussed the sampling distribution of forecasts for a first-order autoregressive model. Fuller and Hasza (2) and Stine (22) extended this discussion to general-order autoregressive models. Bootstrap methods for time series prediction were

explored in (23, 24). For an overview of traditional time series predictive inference, readers can refer to the works of (3, 25). However, for modern complex time series predictive inference, traditional structural approaches may exhibit reduced predictive power and provide less informative prediction intervals. A substantial body of work has focused on leveraging the predictive power of machine learning and deep learning in the time series field; see, e.g., (26–29). For a comprehensive review, refer to Petropoulos et al. (30). Our LLM-TS framework provides a general approach for combining the strengths of classical time-series methodologies with the predictive power of modern LLMs. It preserves the interpretability and provides tighter prediction intervals. Moreover, the core idea extends naturally to deep learning architectures.

**LLMs empowering economic and business research.** Our study contributes to recent advances in LLMs empowering economic and business research. Recent advancements in LLMs have shown potential to support economic and business research through their ability to generate synthetic data that approximates real-world patterns (31, 32). These models enable cost-effective experimentation in applications such as market simulations and policy evaluations, offering researchers a flexible tool for preliminary analysis. However, the reliability of LLM-driven insights remains uncertain due to challenges such as inherent biases, reliance on outdated training data, and limited adaptability to real-time events (33–35). These limitations highlight the importance of developing complementary methods to evaluate and refine the LLM outputs. Our work addresses these challenges by providing an effective approach that leverages LLM outputs, even when they are of low quality or unreliable, to develop a rigorous statistical framework for time series predictive inference. We further demonstrate the power of our methodology through simulations and complex real-world data examples.

**Inflation forecasting.** Our work also contributes to the literature on modern inflation forecasting, a long-studied challenge in economics (14, 17, 36). Recent advances in data availability have enabled new approaches to this problem. For instance, Medeiros et al. (37) demonstrated improved inflation predictions by applying machine learning methods to a broad set of macroeconomic indicators from Mccracken and Ng (38). Many theoretical and empirical studies have also examined the important role of economic narratives in economic fluctuations (5, 6, 39). More recently, Hong et al. (7) incorporated text data from Wall Street Journal articles to enhance forecasting accuracy. Building upon these developments, we explore an alternative approach that integrates both text and macroeconomic data sources while leveraging the power of LLMs for prediction and inference. Our results suggest that this combined method can offer significant improvements over existing approaches.

## LLM-TS

We proceed with assuming that the LLM-powered surrogates data  $\{(\mathbf{y}_t^S, \mathbf{x}_t), t = 1, \dots, T\}$  has already been obtained. The detailed process for obtaining it is deferred to Section *LLM-CPI*. A key ingredient of the suggested LLM-TS method is a joint time-series model integrating a target structural model and an LLM-powered surrogate model on the surrogate data.

373 We begin with introducing the target model on the  
 374 observed data set  $\{y_t, t = 1, \dots, T\}$ . For concreteness,  
 375 we showcase the idea using an autoregressive model with  
 376 exogenous variables of order  $q_1$ , referred to as ARX( $q_1$ ) model,  
 377 as the target structural model

$$379 \quad y_t = \sum_{l=1}^{q_1} \alpha_l y_{t-l} + \mathbf{z}_t^\top \boldsymbol{\theta} + \mathbf{x}_t^\top \boldsymbol{\beta} + \epsilon_t \quad [1]$$

382 for  $t = 1, \dots, T$ . Here,  $\boldsymbol{\alpha} = (\alpha_1, \dots, \alpha_{q_1})^\top \in \mathbb{R}^{q_1}$  denotes  
 383 the autoregressive coefficient vector,  $\boldsymbol{\theta} \in \mathbb{R}^d$  represents the  
 384 regression coefficient vector for the field-collected covariates,  
 385  $\boldsymbol{\beta} \in \mathbb{R}^p$  stands for the regression coefficient vector for the  
 386 text embedding features based on online time series, and  $\epsilon_t$   
 387 is the scalar model error.

388 While the LLM-generated surrogates  $\mathbf{y}_t^S$  may be related  
 389 to the target response, they should not be used directly  
 390 as predictors of the target response due to some intrinsic  
 391 limitations of theirs. First, the LLM-based predictions  
 392 are inherently stochastic, such that repeated executions or  
 393 slight prompt perturbations can yield noticeable different  
 394 results. When such high-variance outputs are used directly as  
 395 covariates, this variability can propagate into the forecasting  
 396 model, leading to unstable estimates and reduced reproducibility.  
 397 Second, because LLM surrogates and embeddings are  
 398 derived from the same textual sources, directly including  
 399  $\mathbf{y}_t^S$  could inflate variance through redundant information,  
 400 exacerbating collinearity and further inflating variance of  
 401 predictive inference.

402 Instead of the direct prediction approach, we incorporate  
 403 the LLM-generated surrogates into the target structural  
 404 model to reduce its error, thereby enhancing the target  
 405 model's predictive and inference performance. This idea is  
 406 inspired by the recent work of (18), who introduced a general  
 407 framework for integrating synthetic surrogates to empower  
 408 the testing procedure for genome-wide association studies.  
 409 Building upon this, we adapt and extend their approach to the  
 410 time series prediction framework, enabling the incorporation  
 411 of LLM-generated synthetic surrogates into the joint time-  
 412 series modeling. To formalize this idea, we choose the LLM-  
 413 powered surrogate model as the vector autoregressive model  
 414 with exogenous variables of order  $q_2$  (referred to as VARX( $q_2$ )  
 415 model hereafter)

$$417 \quad \mathbf{y}_t^S = \sum_{l=1}^{q_2} \mathbf{A}_l^S \mathbf{y}_{t-l}^S + \mathbf{B}^S \mathbf{x}_t + \epsilon_t^S, \quad [2]$$

420 where  $\mathbf{y}_t^S = (y_{t,1}^S, \dots, y_{t,K}^S)^\top \in \mathbb{R}^K$  contains the  $K$ -  
 421 dimensional LLM-generated surrogates at time  $t$ ,  $\mathbf{A}_l^S \in$   
 422  $\mathbb{R}^{K \times K}$  with  $l = 1, \dots, q_2$  denote the autoregressive coefficient  
 423 matrices,  $\mathbf{B}^S \in \mathbb{R}^{K \times p}$  represents the regression coefficient  
 424 matrix for the exogenous text embedding features, and  
 425  $\epsilon_t^S = (\epsilon_{t,1}^S, \dots, \epsilon_{t,K}^S)^\top$  is the model error vector. Here, we  
 426 assume  $q_2 \leq q_1$  without loss of generality. The LLM-powered  
 427 surrogate model (2) is not intended to capture the true data-  
 428 generating process of the LLM surrogates, but serves as a  
 429 working model for their temporal dynamics and relationships  
 430 with the text data. Indeed, the surrogate model can be a mis-  
 431 specified model for surrogate data  $\{(\mathbf{y}_t^S, \mathbf{x}_t), t = 1, \dots, T\}$ ,  
 432 while still achieving the goal of enhancing target model's  
 433 performance.

435 We are now ready to introduce our joint LLM-TS model.  
 436 Following (18), we assume that the errors of both models  
 437 jointly follow a multivariate normal distribution

$$438 \quad (\epsilon_t, (\epsilon_t^S)^\top)^\top \sim N(\mathbf{0}^\top, \boldsymbol{\Sigma}), \quad [3]$$

440 with the error covariance matrix given by

$$441 \quad \boldsymbol{\Sigma} = \begin{bmatrix} \sigma_{TT}^2 & \boldsymbol{\Sigma}_{TS} \\ \boldsymbol{\Sigma}_{ST} & \boldsymbol{\Sigma}_{SS} \end{bmatrix}. \quad 442$$

443 We further assume that the errors are independent across  
 444 different periods, meaning that the autoregressive terms in  
 445 the models account for all cross-temporal correlations across  
 446 months. Here,  $\sigma_{TT}^2$  represents the variance of the target  
 447 model error,  $\boldsymbol{\Sigma}_{SS}$  denotes the covariance matrix of the LLM-  
 448 powered surrogate model errors, and  $\boldsymbol{\Sigma}_{TS}$  captures the cross-  
 449 covariance between the target model and the LLM-powered  
 450 surrogate model errors. Such formulation allows us to model  
 451 the correlation structures between the target model and LLM-  
 452 powered surrogate predictions. It is important to emphasize  
 453 that the target model and the LLM-powered surrogate model  
 454 do *not* share any model parameters; their only connection is  
 455 through the correlations of their model errors.

456 **Estimation based on the joint modeling.** In view of the joint  
 457 LLM-TS model (3), we can rewrite the random error of  
 458 the target model Eq. (1) using the properties of conditional  
 459 normal distribution as

$$460 \quad \epsilon_t = \boldsymbol{\gamma}^\top \boldsymbol{\epsilon}_t^S + e_t, \quad [4]$$

461 where  $\boldsymbol{\gamma} = \boldsymbol{\Sigma}_{SS}^{-1} \boldsymbol{\Sigma}_{ST}$  represents the regression coefficient  
 462 vector linking the LLM-generated surrogate model errors to the  
 463 target model error, and random error  $e_t \sim N(0, \sigma_e^2)$  with  
 464  $\sigma_e^2 = \sigma_{TT}^2 - \boldsymbol{\Sigma}_{TS} \boldsymbol{\Sigma}_{SS}^{-1} \boldsymbol{\Sigma}_{ST}$ . The decomposition in (4) allows  
 465 us to isolate the portion of variation in the target model error  
 466 that can be explained by the surrogate model errors, thereby  
 467 reducing the overall variance of the target model error and  
 468 ensuring more accurate prediction and inference.

469 By substituting Eq. (2) into decomposition Eq. (4), it  
 470 holds that

$$471 \quad \epsilon_t = \boldsymbol{\gamma}^\top \left( \mathbf{y}_t^S - \sum_{l=1}^{q_2} \mathbf{A}_l^S \mathbf{y}_{t-l}^S \right) - \boldsymbol{\gamma}^\top \mathbf{B}^S \mathbf{x}_t + e_t \quad [5]$$

$$472 \quad := \boldsymbol{\gamma}^\top \mathbf{D}(\mathbf{y}_t^S) - \boldsymbol{\gamma}^\top \mathbf{B}^S \mathbf{x}_t + e_t,$$

473 where  $\mathbf{D}(\mathbf{y}_t^S) := \mathbf{y}_t^S - \sum_{l=1}^{q_2} \mathbf{A}_l^S \mathbf{y}_{t-l}^S$  captures the error  
 474 component of the LLM predictions after accounting for their  
 475 autoregressive structure. Plugging expression (5) into the  
 476 target model Eq. (1), we can equivalently write the joint  
 477 LLM-TS model (3) as a joint LLM-powered ARX model  
 478 given by

$$479 \quad y_t = \sum_{l=1}^{q_1} \alpha_l y_{t-l} + \mathbf{z}_t^\top \boldsymbol{\theta} + \mathbf{x}_t^\top (\boldsymbol{\beta} - \boldsymbol{\gamma}^\top \mathbf{B}^S) + \boldsymbol{\gamma}^\top \mathbf{D}(\mathbf{y}_t^S) + e_t \quad [6]$$

$$480 \quad = \sum_{l=1}^{q_1} \alpha_l y_{t-l} + \mathbf{z}_t^\top \boldsymbol{\theta} + \mathbf{x}_t^\top \boldsymbol{\delta} + \boldsymbol{\gamma}^\top \mathbf{D}(\mathbf{y}_t^S) + e_t,$$

481 where  $\boldsymbol{\delta} = \boldsymbol{\beta} - \boldsymbol{\gamma}^\top \mathbf{B}^S$ . It is seen that the model error  $e_t$  in  
 482 (6) has a reduced variance compared to that in the target  
 483 model Eq. (1). The stronger the model error correlations, the

greater gains in the variance reduction, and thereby, better prediction and inference accuracy using Eq. (6). Throughout the rest of the paper, the joint LLM-TS model is implicitly referred to as model Eq. (6).

To estimate the parameters of the joint LLM-powered ARX model Eq. (6), we exploit a two-step approach. We first construct estimators  $(\hat{\mathbf{A}}_1^S, \dots, \hat{\mathbf{A}}_{q_2}^S, \hat{\mathbf{B}}^S)$  of the parameters in Eq. (2) by solving the optimization problem

$$\min_{\mathbf{A}_1^S, \dots, \mathbf{A}_{q_2}^S, \mathbf{B}^S} \frac{1}{T} \sum_{t=q_2+1}^T \left\| \mathbf{y}_t^S - \sum_{l=1}^{q_2} \mathbf{A}_l^S \mathbf{y}_{t-l}^S - \mathbf{B}^S \mathbf{x}_t \right\|^2. \quad [7]$$

Given such estimates, we further estimate the error component  $\mathbf{D}(\mathbf{y}_t^S)$  using the plug-in estimator

$$\hat{\mathbf{D}}(\mathbf{y}_t^S) = \mathbf{y}_t^S - \sum_{l=1}^{q_2} \hat{\mathbf{A}}_l^S \mathbf{y}_{t-l}^S. \quad [8]$$

We then form estimates  $(\hat{\boldsymbol{\alpha}}, \hat{\boldsymbol{\theta}}, \hat{\boldsymbol{\delta}}, \hat{\boldsymbol{\gamma}})$  of the parameters in the joint LLM-powered ARX model Eq. (6) by solving the optimization problem

$$\min_{\boldsymbol{\alpha}, \boldsymbol{\theta}, \boldsymbol{\delta}, \boldsymbol{\gamma}} \frac{1}{T} \sum_{t=q_2+1}^T \left( y_t - \sum_{l=1}^{q_1} \alpha_l y_{t-l} - \mathbf{z}_t^\top \boldsymbol{\theta} - \mathbf{x}_t^\top \boldsymbol{\delta} - \boldsymbol{\gamma}^\top \hat{\mathbf{D}}(\mathbf{y}_t^S) \right)^2. \quad [9]$$

With the estimates given in (7)–(9) above, we can construct the one-step-ahead forecast  $\hat{y}_{T+1}$  using the joint LLM-powered ARX model Eq. (6) as

$$\hat{y}_{T+1} = \sum_{l=1}^{q_1} \hat{\alpha}_l y_{T+1-l} + \mathbf{z}_{T+1}^\top \hat{\boldsymbol{\theta}} + \mathbf{x}_{T+1}^\top \hat{\boldsymbol{\delta}} + \hat{\boldsymbol{\gamma}}^\top \hat{\mathbf{D}}(\mathbf{y}_{T+1}^S), \quad [10]$$

where  $\hat{\mathbf{D}}(\mathbf{y}_{T+1}^S)$  is calculated using Eq. (8) with  $t = T+1$ . For the multi-step-ahead forecasts, we employ a rolling horizon approach. Specifically, the  $h$ -step-ahead forecast  $\hat{y}_{T+h}$  can be constructed as

$$\hat{y}_{T+h} = \sum_{l=1}^{q_1} \hat{\alpha}_l \hat{y}_{T+h-l} + \mathbf{z}_{T+h}^\top \hat{\boldsymbol{\theta}} + \mathbf{x}_{T+h}^\top \hat{\boldsymbol{\delta}} + \hat{\boldsymbol{\gamma}}^\top \hat{\mathbf{D}}(\mathbf{y}_{T+h}^S), \quad [11]$$

where  $\hat{y}_{T+h-1}, \hat{y}_{T+h-2}, \dots$  are iteratively computed based on the forecasts  $\hat{y}_t$  from previous time stamps with  $\hat{y}_t = y_t$  for  $t \leq T$ , and covariates  $\mathbf{x}_{T+h}$ ,  $\mathbf{z}_{T+h}$ , and  $\hat{\mathbf{D}}(\mathbf{y}_{T+h}^S)$  are used as the inputs.

Under mild regularity conditions, we can show that for each fixed  $h \geq 1$ ,

$$\hat{y}_{T+h} - y_{T+h} = \sum_{r=0}^{h-1} (\mathbf{A}^r)_{11} e_{T+h-r} + o_p(1), \quad [12]$$

where matrix  $\mathbf{A}$  is defined in *Section 1 of SI Appendix* and  $\mathbf{A}^0$  is defined as the identity matrix  $\mathbf{I}$ . The proof details are deferred to *Section 1 of SI Appendix*.

In contrast, if we ignore the LLM-generated surrogates and rely solely on the traditional ARX model, the  $h$ -step-ahead forecast is then given by

$$\hat{y}_{T+h}^a = \sum_{l=1}^{q_1} \hat{\alpha}_l \hat{y}_{T+h-l}^a + \mathbf{z}_{T+h}^\top \hat{\boldsymbol{\theta}} + \mathbf{x}_{T+h}^\top \hat{\boldsymbol{\beta}}, \quad [13]$$

where  $\hat{y}_{T+h-1}^a, \hat{y}_{T+h-2}^a, \dots$  are iteratively computed via the ARX model prediction with  $\hat{y}_t^a = y_t$  for each  $t \leq T$ . The prediction error for this benchmark model is

$$\hat{y}_{T+h}^a - y_{T+h}^a = \sum_{r=0}^{h-1} (\mathbf{A}^r)_{11} e_{T+h-r} + o_p(1).$$

The efficiency gain of LLM-TS compared to the traditional ARX model can be quantified by the ratio of prediction error variances

$$\begin{aligned} \text{Efficiency} &= \frac{\text{Var} \left( \sum_{r=0}^{h-1} (\mathbf{A}^r)_{11} e_{T+h-r} \right)}{\text{Var} \left( \sum_{r=0}^{h-1} (\mathbf{A}^r)_{11} e_{T+h-r} \right)} \\ &= \frac{\sigma_{TT}^2}{\sigma_{TT}^2 - \Sigma_{TS} \Sigma_{SS}^{-1} \Sigma_{ST}}. \end{aligned} \quad [14]$$

Under the simplified assumptions of  $\Sigma_{SS} = \mathbf{I}$  and  $\Sigma_{ST} = \rho \mathbf{1}$  with  $\mathbf{I}$  and  $\mathbf{1}$  the identity matrix and the vector of ones, respectively, the efficiency gain in (14) reduces to

$$\text{Efficiency} = \frac{1}{1 - |S|\rho^2}. \quad [15]$$

Here, we require  $|S|\rho^2 < 1$  to guarantee the positive definiteness of the joint covariance matrix  $\Sigma$ . In light of (15), we see the practical benefits of incorporating the LLM-generated synthetic surrogates in the LLM-TS, and that the stronger the correlations between the target and the surrogates, the greater the efficiency gain. Moreover, under the residual correlation framework in Eq. (3), our joint modeling approach theoretically achieves uniformly lower prediction mean squared error (PMSE) than approaches that use directly the surrogate response  $\mathbf{y}_t^S$  as a predictor; see *Section 7 of SI Appendix* for a detailed theoretical comparison.

## Predictive inference via LLM-TS

We now introduce two ways of constructing the LLM-TS prediction intervals for time series *predictive inference*.

**Box-Jenkins prediction interval.** Based on Eq. (12), we can construct an asymptotic prediction interval for the  $h$ -step-ahead prediction  $\hat{y}_{T+h}$  once we obtain a consistent estimator of the error variance. One common approach to estimating such variance is through the sum of squared residuals (40)

$$\hat{\sigma}_e^2 := \frac{1}{T - q_1} \sum_{t=q_1+1}^T \hat{e}_t^2 \quad [16]$$

with  $\hat{e}_t = y_t - \sum_{l=1}^{q_1} \hat{\alpha}_l y_{t+1-l} - \mathbf{z}_t^\top \hat{\boldsymbol{\theta}} - \mathbf{x}_t^\top \hat{\boldsymbol{\delta}} - \hat{\boldsymbol{\gamma}}^\top \hat{\mathbf{D}}(\mathbf{y}_t^S)$ . Using the above error variance estimator, we can construct the Box-Jenkins (BJ) prediction interval (3) with confidence level  $1 - \alpha$  as

$$\text{PI}^{BJ} = (\hat{y}_{T+h}) [\hat{y}_{T+h} - \hat{z}_{\alpha/2}^h, \hat{y}_{T+h} + \hat{z}_{\alpha/2}^h], \quad [17]$$

where  $\hat{z}_{\alpha/2}^h = |z_{\alpha/2}| \sqrt{\sum_{r=0}^{h-1} (\hat{\mathbf{A}}^r)_{11}^2} \hat{\sigma}_e$ ,  $z_{\alpha/2}$  is the  $\alpha/2$  quantile of the standard normal distribution, and  $\alpha \in (0, 1)$ . The BJ prediction interval asymptotically covers the true value  $y_{T+h}$  if  $\hat{\sigma}_e^2$  is a consistent estimator of  $\sigma_e^2$ .

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**Bootstrap prediction interval.** We also suggest a residual-based bootstrap prediction interval for LLM-TS (41, 42). Given the estimated parameters  $(\hat{\alpha}, \hat{\theta}, \hat{\delta}, \hat{\gamma})$ , we compute the residuals as

$$\hat{e}_t = y_t - \sum_{l=1}^{q_1} \hat{\alpha}_l y_{t-l} - \mathbf{z}_t^\top \hat{\theta} - \mathbf{x}_t^\top \hat{\delta} - \hat{\gamma}^\top \hat{\mathbf{D}}(\mathbf{y}_t^S) \quad [18]$$

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for  $t = q_1 + 1, \dots, T$ . We generate the bootstrap residuals by drawing  $T+h$  samples with replacement from  $\{\hat{e}_t - \hat{\mu}, t = q_1 + 1, \dots, T\}$ , where  $\hat{\mu} = \sum_{t=q_1+1}^T \hat{e}_t / (T - q_1)$ . Denote by  $\{e_t^*, t = 1, \dots, T+h\}$  the bootstrap residuals. We then recursively calculate

$$y_t^* = \sum_{l=1}^{q_1} \hat{\alpha}_l y_{t-l}^* + \mathbf{z}_t^\top \hat{\theta} + \mathbf{x}_t^\top \hat{\delta} + \hat{\gamma}^\top \hat{\mathbf{D}}(\mathbf{y}_t^S) + e_t^* \quad [19]$$

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for  $t = q_1 + 1, \dots, T+h$ , with initial points  $\{y_t^* = e_t^*, t \leq q_1\}$ . We next refit the joint LLM-powered ARX model Eq. (6) using the bootstrap sample  $\{y_t^*, t = q_1 + 1, \dots, T\}$  and denote the refitted parameters as  $\{\hat{\alpha}^*, \hat{\theta}^*, \hat{\delta}^*, \hat{\gamma}^*\}$ . The  $h$ -step-ahead forecast for the bootstrap sample is calculated as

$$\hat{y}_{T+h}^* = \sum_{l=1}^{q_1} \hat{\alpha}_l^* \hat{y}_{T+h-l}^* + \mathbf{z}_{T+h}^\top \hat{\theta}^* + \mathbf{x}_{T+h}^\top \hat{\delta}^* + (\hat{\gamma}^*)^\top \hat{\mathbf{D}}(\mathbf{y}_{T+h}^S), \quad [20]$$

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where  $\hat{y}_{T+h-1}^*, \hat{y}_{T+h-2}^*, \dots$  are computed similarly with  $\hat{y}_t^*$  for each  $t \leq T$  being  $y_t^*$ . Using  $\hat{y}_{T+h}^*$  introduced above, the bootstrap residual is calculated as  $\hat{e}_{T+h}^* = y_{T+h}^* - \hat{y}_{T+h}^*$ .

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We repeat the bootstrap procedure (i.e., Eq. (19) and Eq. (20))  $B \geq 1$  times to obtain a sequence of bootstrap residuals  $\{\hat{e}_{T+h}^{*(b)}, b = 1, \dots, B\}$ . For each  $\alpha \in (0, 1)$ , denote the  $\alpha/2$  quantile and  $1 - \alpha/2$  quantile of the bootstrap residuals as  $\hat{q}_{\alpha/2}^h$  and  $\hat{q}_{1-\alpha/2}^h$ , respectively. Then we can construct the bootstrap prediction interval with confidence level  $1 - \alpha$  as

$$\text{PI}^{BOOT}(\hat{y}_{T+h}) = [\hat{y}_{T+h} + \hat{q}_{\alpha/2}^h, \hat{y}_{T+h} + \hat{q}_{1-\alpha/2}^h]. \quad [21]$$

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We provide rigorous theoretical guarantees for the asymptotic coverage rates of both the BJ and bootstrap prediction intervals in Theorems 2 and 3 of *Section 1 of SI Appendix*, respectively. These results ensure that our two proposed prediction intervals achieve the desired coverage level asymptotically under some mild regularity conditions.

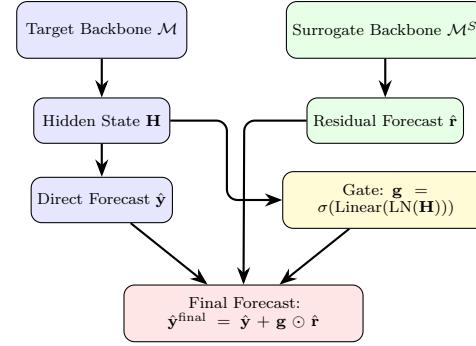
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**Extension of LLM-TS to deep learning models.** We emphasize that the linear structures in Eq. (1) and Eq. (2) are chosen primarily to align with the macroeconomic case studies considered in this paper. Macroeconomic indices are typically observed at a monthly frequency, resulting in short samples with relatively stable trend components. In such settings, linear specifications are effective and highly interpretable for short-horizon forecasting, and often perform competitively relative to more complex deep learning models; see *Tables S18 and S19 of SI Appendix* for empirical evidence.

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Our joint modeling framework is, however, not restricted to linear time series models. Its two core components: i) LLM-derived text embedding features and ii) the joint residual modeling extend naturally to modern deep learning architectures. Recent studies have shown that LLM-derived representations can serve as informative inputs for time-series

forecasting (29, 43); our case analysis further confirms that such representations remain effective even when embedded within classical structural models. We now introduce a deep-learning model extension of LLM-TS.



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**Fig. 3. LLM-TS with deep learning backbones.** The target backbone  $\mathcal{M}$  produces a direct forecast and a hidden representation  $\mathbf{H}$ . The surrogate backbone  $\mathcal{M}^S$  predicts a residual correction  $\hat{\mathbf{r}}$ . A gating mechanism computes horizon-wise weights  $\mathbf{g} \in [0, 1]^H$  and adaptively combines the two paths.

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The central innovation of our framework lies in the joint residual modeling strategy. In many applications, some surrogate series related to the target variable are available, but directly incorporating them as symmetric inputs can distort the intrinsic temporal lag structure of the target series and degrade forecasting accuracy (27, 28). We therefore impose strict *channel independence* in the target model: the target series is forecast *solely* from its own lagged history. At the same time, discarding surrogate information entirely may forfeit valuable leading or cross-series signals.

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Our joint residual modeling principle resolves such tension by allowing surrogate series to contribute *only* through residual corrections rather than as direct inputs. To generalize beyond the linear models, we replace the target and surrogate components with deep time-series backbones  $\mathcal{M}$  and  $\mathcal{M}^S$ , respectively. The target backbone produces an  $H$ -step direct forecast  $\hat{\mathbf{y}}$  along with a hidden representation  $\mathbf{H}$ , while the surrogate backbone predicts the corresponding residual correction  $\hat{\mathbf{r}}$ . Because deep learning models typically perform direct multi-step forecasting, surrogate information is most informative at short horizons, whereas long-horizon behavior is governed primarily by the intrinsic dynamics of the target series. To accommodate such heterogeneity, we introduce a *gating mechanism* (44) that regulates adaptively the contribution of surrogate residuals across different forecast horizons, as shown in Figure 3. The gate incorporates selectively the surrogate information only when it improves prediction accuracy, thereby ensuring stability and robustness while preserving the dominant role of the target backbone in long-horizon forecasting. A full deep learning extension of the LLM-TS framework, together with comparisons against state-of-the-art forecasting models on standard benchmarks, is provided in *Section 5 of SI Appendix*.

## Simulation examples

In this section, we present several simulation examples to examine the finite-sample performance of the LLM-TS method. The synthetic data sets used in these simulations

are derived from the *Wall Street Journal* (WSJ) corpus; see *Section 2 of SI Appendix* for simulation setting details.

As seen in *Table S1 in SI Appendix*, the LLM-TS model consistently achieves lower both relative root prediction mean squared error (rPMSE) and relative sign prediction error (rSign) than the benchmark models. Such performance advantage becomes even more pronounced as the model error correlation level  $\rho$  increases. Importantly, the improvement holds consistently across both short and long term horizons, demonstrating the advantages and robustness of the LLM-TS method in forecasting tasks compared to classical approaches.

*Table S2 in SI Appendix* reports the average coverage of the prediction intervals and their mean interval lengths (in parentheses) for each forecast horizon  $H$  and correlation level  $\rho$ . The AR model maintains high coverage across all cases. The two LLM-TS variants—based on the BJ and bootstrap intervals—match the nominal coverage in the short run and stay close in the long run, with only occasional slight undercoverage. Both LLM-TS variants produce much narrower prediction intervals than the AR model, demonstrating their ability to produce more accurate forecasts without losing coverage.

We also conduct additional simulations to test LLM-TS's robustness to different types of model misspecifications: the omitted key predictor, overfitting, and non-Gaussian error distribution (i.e., t-distribution). In all these cases, the LLM-TS remains robust by providing reliable predictive inference results; see *Section 2 of SI Appendix* for details.

## LLM-CPI

In this section, we validate the LLM-TS framework for inflation forecasting across two distinct economic environments: the U.S. and China. Our approach to constructing high-frequency, LLM-powered surrogate signals is readily generalizable to a broad class of time-series prediction problems whenever unstructured data is available as auxiliary information. When applied specifically to inflation forecasting, we refer to this framework as the LLM-CPI.

### LLM-based high-frequency inflation index construction.

**Online text time-series data collection.** To capture real-time inflation narratives across the distinct economic environments of the U.S. and China, we leverage two complementary information platforms that reflect both professional and grassroots perspectives. In the U.S. context, news coverage in *The Wall Street Journal* (WSJ) tracks systematically the monetary policy developments, price pressures, and macro-financial conditions, serving as a forward-looking signal of dynamics in a mature market economy. In contrast, reflecting China's unique information ecosystem, Weibo provides a critical real-time channel for public discourse, with 588 million monthly active users (45). As a platform studied widely for its role in shaping and amplifying public sentiment (46, 47), Weibo offers a granular view of consumer expectations and perceptions that differs fundamentally from traditional Western media outlets.

To operationalize this design, we collect the full corpus of WSJ news articles published between September 1, 2015 and August 31, 2025. For the Chinese market, we aggregate all Weibo posts containing predefined inflation-related keywords following (6) from January 1, 2019 to September 30, 2025.

See *Sections 3 and 4 of SI Appendix* for the data details. The resulting WSJ corpus comprises 422,444 news articles, while the Weibo data set contains approximately 159.5 million short posts. We should emphasize that the collected text data sets not only pertain to inflation-related contents, but also include background noise, including advertisements, e-commerce contents, and sales promotions.

**LLM-based inflation index construction.** Our text analysis begins with standard preprocessing steps to remove duplicate samples and non-textual elements such as emojis and special symbols. The primary challenge, however, lies in identifying inflation-related texts from a large volume of unrelated contents. Extracting such signals from large text collections is difficult for three major reasons. First, *contextual ambiguity* arises because commercial or promotional texts often resemble genuine economic discussions, making it difficult to distinguish informative content from noise. Second, *linguistic differences* across platforms add complexity, ranging from the formal financial language used in the *Wall Street Journal* to the informal and highly varied expressions found in user-generated contents on Weibo. Third, *semantic diversity* implies that inflation and price changes are described in many different ways across professional and grassroots discussions, which limits the effectiveness of simple filtering rules. As a result, directly applying unsupervised learning methods to identify inflation-related content has limited effectiveness.

To address these practical challenges, we develop a hierarchical LLM-based learning framework that is summarized in Algorithm 1. We first employ a chain-of-thought (CoT; 48) prompting strategy using GPT-4 (specifically `gpt-4-turbo-2024-04-09`) to annotate two random samples of 8,000 WSJ news articles and 30,000 Weibo posts, respectively. Leveraging the few-shot learning capabilities of GPT-4 (49), we implement a hierarchical annotation procedure in which documents are first classified into broad categories (e.g., Inflation, Lifestyle, and Entertainment), and those identified as inflation-related are subsequently assigned a continuous severity score; see *Sections 3 and 4 of SI Appendix* for prompt details.

Using these high-quality annotations, we further fine-tune a set of domain-specific BERT models. Specifically, we use *Category-BERT* for thematic classification and *CPI-BERT* for sentiment quantification across both text data sets. In addition, for corpora where commercial content is prevalent, we fine-tune an *Advertisement-BERT* model to filter promotional noise. Given the linguistic characteristics of the data, we adopt the `bert-base-cased` architecture for the WSJ corpus to preserve case-sensitive information in English text, and the `bert-base-chinese` architecture for processing Weibo posts to ensure accurate character-level tokenization. The fine-tuned BERT models exhibit strong performances on the held-out test sets, which are reported in *SI Appendix*; see *Figs. S13–S14* for the WSJ data and *Figs. S2–S4* for the Weibo data.

In the final deployment stage, we apply such cascaded framework sequentially to the full text data sets. The WSJ news articles are processed directly for thematic relevance, while for Weibo, the LLM pipeline first removes commercial content using *Advertisement-BERT* and then applies *Category-BERT* to identify inflation-related texts. Only texts classified as inflation-related are passed to the final scoring

869 **Algorithm 1** Constructing the LLM-generated high-  
 870 frequency index  
 871 1: **Input:** The preprocessed text data set.  
 872 2: **Output:** The LLM-generated high-frequency index.  
 873 3: **Step 1: Sampling**  
 874 4: Randomly sample a subset of text documents.  
 875 5: **Step 2: Annotation**  
 876 6: **for** each document in the sampled subset **do**  
 877 7:   Use *chain-of-thought prompting* with GPT-4 to anno-  
 878   tate the document:  
 879 8:     (a) If commercial content exists, determine if the  
 880       document is an advertisement;  
 881 9:     (b) Determine whether the document is relevant  
 882       to inflation;  
 883 10:    (c) If related to inflation, assess the continuous  
 884       degree of inflation it represents.  
 885 11: **Step 3: Fine-tuning**  
 886 12: Fine-tune the following BERT models using the annotated  
 887    data set:  
 888 13:   (a) **Advertisement-BERT:** Identify advertisement  
 889       content (when applicable);  
 890 14:   (b) **Category-BERT:** Identify whether non-  
 891       advertisement posts are related to inflation;  
 892 15:   (c) **CPI-BERT:** Estimate the continuous degree for  
 893       target-related documents.  
 894 16: **Step 4: Sequential prediction**  
 895 17: **if** commercial content exists in the corpus **then**  
 896 18:   Apply **Advertisement-BERT** to remove advertise-  
 897       ment documents.  
 898 19: Apply **Category-BERT** to filter out non-relevant docu-  
 899       ments.  
 900 20: Apply **CPI-BERT** *only* to the remaining target-related  
 901       documents.  
 902 21: Compute the LLM-generated daily high-frequency index  
 903       using Eq. (22).  
 904

905 stage. This filtering procedure yields 256,747 relevant WSJ  
 906 news articles (a 60.8% retention rate), and reduces the original  
 907 Weibo data set to 7.35 million inflation-relevant posts (a  
 908 4.6% retention rate) from 1.49 million unique users. The  
 909 retained documents are then processed by the *CPI-BERT*  
 910 regression model to generate fine-grained sentiment scores,  
 911  $\text{Score}_i \in [0, 1]$ . *Figs. S6 and S15 in SI Appendix* depict the  
 912 daily volumes of inflation-up and inflation-down discussions at  
 913 this final stage for both WSJ and Weibo data sets, respectively.  
 914 See *Sections 3 and 4 of SI Appendix* for more fine-tuning and  
 915 prediction details.

916 To construct the LLM-generated daily inflation index,  
 917 we pair each document's continuous inflation score with its  
 918 publication date, forming a data set  $\{(\text{Score}_i, \text{Date}_i), i = 1, \dots, N\}$ , where  $\text{Date}_i$  denotes the posting date of the  $i$ th  
 919 document entry. The *LLM-generated daily inflation index* for  
 920 day  $d$  is defined as

$$\text{Inflation}_d = \frac{\sum_{i=1}^N \text{Score}_i \mathbb{I}(\text{Date}_i = d)}{\sum_{i=1}^N \mathbb{I}(\text{Date}_i = d)}, \quad [22]$$

921 where  $\mathbb{I}(\cdot)$  is the indicator function. Fig. 1 plots the LLM-  
 922 generated daily inflation index  $\text{Inflation}_d$  in (22) and the  
 923 monthly survey CPI for the WSJ data. See *Fig. S5 in SI*  
 924 *Appendix* for the corresponding plot for the Weibo data. Since

925 the LLM-generated daily inflation index  $\text{Inflation}_d$  highly  
 926 fluctuates, we smooth it into three ten-day periods within  
 927 each month in practical application. We denote the resulting  
 928 LLM-generated inflation index as  $\{y_{t,k}^S \in \mathbb{R}, t = 1, \dots, T, k = 1, \dots, K\}$ , where  $K = 3$  represents the three periods within  
 929 each month, and each  $y_{t,k}^S$  corresponds to a surrogate of the  
 930 CPI generated by LLMs (i.e., ChatGPT and the trained  
 931 BERT models) for the  $k$ th period of the  $t$ th month.

932 **Online text embeddings.** Our suggested LLM-CPI framework  
 933 incorporates two text embedding methods: the topic proba-  
 934 bility embeddings from the latent Dirichlet allocation (LDA)  
 935 (50), and the BERT embeddings extracted from the fine-tuned  
 936 CPI-BERT model architecture. Specifically, we implement  
 937 the LDA model to derive topic probability distributions  
 938 from the text data. Each document is represented as  
 939 a  $K$ -dimensional vector, where elements correspond to  
 940 posterior probabilities of memberships in  $K$  latent thematic  
 941 clusters. These document-topic distributions are temporally  
 942 aggregated monthly through averaging; see *Figs. S7 and S16*  
 943 in *SI Appendix* for the LDA topic results, as well as *Tables S22*  
 944 and *S23 in SI Appendix* for related hashtags from the Weibo  
 945 data. For the BERT embeddings, we extract 768-dimensional  
 946 vectors through mean pooling of the final hidden layer right  
 947 before the output layer of the fine-tuned CPI-BERT model  
 948 (i.e., a deep neural network), capturing semantic patterns in  
 949 individual documents (i.e., news articles or posts). These  
 950 document-level embeddings are averaged within each month  
 951 to create monthly LLM-based economic text features; see  
 952 *Sections 3 and 4 of SI Appendix* for the text embedding  
 953 details.

## 954 **LLM-CPI forecasting and inference results.**

955 **Low-frequency survey CPI and other indicators.** The target variable  
 956  $y_t$  in Eq. (1) is the observed monthly inflation rate. For  
 957 both the U.S. and Chinese economic analyses, we use the  
 958 urban CPI as the measure of inflation. Following standard  
 959 normalization, we obtain the series  $y_t$  for  $t = 1, \dots, T$ , where  
 960 the sign of  $y_t$  indicates whether inflation has increased or  
 961 decreased relative to the previous month. As an additional  
 962 macroeconomic control, we also collect the national urban  
 963 surveyed unemployment rate reported monthly by the NBSC  
 964 and standardize it in the same manner. See *Sections 3*  
 965 and *4 of SI Appendix* for details on data collection and  
 966 preprocessing. The predictors in the LLM-CPI model consist  
 967 of two components. The first component comprises text  
 968 embeddings extracted from the online texts, as described  
 969 above. The second component is a widely used inflation-  
 970 related macroeconomic control variable, standardized and  
 971 denoted as  $z_t$  for  $t = 1, \dots, T$ .

972 **Out-of-sample forecasting.** Based on the proposed LLM-TS  
 973 framework in Eq. (1), Eq. (2), and Eq. (6), we have the high-  
 974 frequency LLM-based inflation index  $y_{t,k}^S$  as our surrogates  
 975 in Eq. (2), the text embeddings  $\mathbf{x}_t$ , and the target variable  
 976  $y_t$  in Eq. (1). For easy reference, we will name the LLM-CPI  
 977 method with the LDA and BERT embeddings as LLM+LDA  
 978 and LLM+BERT, respectively.

979 We now assess the out-of-sample forecasting performance  
 980 of these two methods. To isolate the effect of the text  
 981 embedding features, we first exclude the macroeconomic  
 982 predictor  $z_t$  (i.e., the unemployment rate). We compare

993 the LLM+LDA and LLM+BERT models against three well-  
994 established inflation forecasting benchmarks: the AR model,  
995 RW model, and AVE model, as well as two deep learning  
996 models: the PatchTST (28) and Time-LLM (29). In addition,  
997 we compare to the direct text-based model using the LDA  
998 and BERT embeddings as covariates (without LLM-CPI for  
999 model error variance reduction). For simplicity, we refer to  
1000 the text-based prediction model with the LDA embeddings  
1001 as the LDA model, and that with the BERT embeddings as  
1002 the BERT model (with slight abuse of terminology). Detailed  
1003 model specifications are provided in *Section 3 of SI Appendix*.

1004 The out-of-sample forecasting period spans the  $H$  months  
1005 immediately preceding the end of the online text time series  
1006 up to its endpoint, yielding  $H$  forecast steps. Specifically,  
1007 the online text series ends on August 31, 2025 for the WSJ  
1008 corpus and on September 30, 2025 for the Weibo corpus. This  
1009 evaluation window is strictly independent of the samples used  
1010 for model fitting and model selection, and occurs strictly  
1011 after the release dates of the LLMs used for annotation.  
1012 These design choices ensure the integrity of the out-of-sample  
1013 assessment. Additional discussions on potential information  
1014 leakage considerations are provided in *Sections 3 and 4 of SI Appendix*. We choose the AR model as the baseline and  
1015 evaluate the performance using the relative root prediction  
1016 mean squared error ( $rPMSE^{AR}(H)$ ) and the relative sign  
1017 prediction error ( $rSign^{AR}(H)$ ) defined in Eq. (23) and Eq.  
1018 (24) of *SI Appendix*. For both performance measures, we now  
1019 have  $Q = 1$  due to a single observation for each month.

1020 Table 1 summarizes the results across different forecast  
1021 horizons  $H$  for the WSJ data; see *Table S10 of SI Appendix* for  
1022 the case of Weibo data. We outline the key findings below. 1) *Text  
1023 embeddings enhance prediction accuracy.* Incorporating  
1024 textual signals (without the LLM-CPI framework) reduces  
1025 consistently the forecast errors compared to the baselines.  
1026 Specifically, the LDA model achieves an average relative  
1027 prediction mean squared error ( $rPMSE^{AR}(H)$ ) of 0.996, out-  
1028 performing the RW (1.325), AVE (1.096), Time-LLM (1.811),  
1029 and PatchTST (1.157) benchmarks. The BERT-based model  
1030 performs even better, attaining an average  $rPMSE^{AR}(H)$  of  
1031 0.950. Both embedding-based models also exhibit strong  
1032 directional accuracy. The relative sign prediction error  
1033 ( $rSign^{AR}(H)$ ) is 0.889 for the LDA model and 0.806 for  
1034 the BERT model, both substantially lower than those of  
1035 all classical time series benchmark models and comparable  
1036 to the performance of deep learning-based approaches. 2) *The  
1037 LLM-CPI method improves substantially the prediction  
1038 performance.* Incorporating LLM-based surrogate signals  
1039 systematically improves performance across both point and  
1040 directional forecast metrics. For the LDA-based specification,  
1041 the LLM-CPI framework reduces the  $rPMSE^{AR}(H)$  from  
1042 0.996 to 0.916. The improvement is most pronounced for  
1043 the BERT-based specification. The LLM+BERT model  
1044 achieves the best overall performance, lowering  $rPMSE^{AR}(H)$   
1045 to 0.882. More importantly, it delivers a marked improvement  
1046 in directional accuracy: the relative sign prediction error  
1047 drops from 0.806 for the standalone BERT model to 0.639  
1048 under the LLM-CPI framework.

1049 **Inflation predictive inference.** We further evaluate the predictive  
1050 inference performance of LLM-CPI. We employ the Box-  
1051 Jenkins (BJ) procedure for constructing the prediction  
1052 intervals in LLM-CPI. We compare with both AR and SPI

1053 prediction intervals (21). The corresponding results for LLM-  
1054 CPI with the bootstrap prediction interval are presented in  
1055 *Sections 3 and 4 of SI Appendix*.

1056 Table 1 reports the prediction interval coverage rates and  
1057 average interval lengths across different forecast horizons  
1058  $H = 8$  to 15; see *Table S11 of SI Appendix* for the case  
1059 of Weibo data. The major findings are summarized below.

1060 **1) Text embeddings improve prediction interval efficiency.**  
1061 Models that incorporate text embeddings produce shorter  
1062 prediction intervals than the autoregressive benchmark while  
1063 maintaining nominal coverage. As shown in the bottom panel  
1064 of Table 1, the LDA and BERT models (without the LLM-CPI  
1065 structure) achieve average interval lengths of 3.902 and 3.806,  
1066 respectively, both of which are tighter than the AR benchmark  
1067 (3.932). Importantly, both models retain a 100% coverage rate  
1068 across all horizons  $H$ . These results reinforce the conclusion  
1069 that textual signals contain valuable predictive information  
1070 that reduces uncertainty, with BERT embeddings yielding the  
1071 sharpest intervals among the standalone text-based models.  
1072 **2) LLM-CPI yields further gains in inference precision.**  
1073 The LLM-augmented models deliver substantial additional  
1074 reductions in interval length, reflecting markedly improved  
1075 inference efficiency. In particular, the LLM+LDA model  
1076 reduces the average interval length to 2.920, corresponding to  
1077 a reduction of more than 25% relative to the AR benchmark,  
1078 while maintaining a reliable average coverage rate of 0.912.  
1079 Similarly, the LLM+BERT model achieves an average interval  
1080 length of 3.731 with a coverage rate of 0.930. The SPI  
1081 method yields wider prediction intervals and exhibits lower  
1082 empirical coverage. Together, these results demonstrate  
1083 that the LLM-CPI framework produces tighter and more  
1084 informative prediction intervals by leveraging LLM-based  
1085 surrogate signals to denoise textual information and reduce  
1086 prediction variance.

1087 Additional empirical results that incorporate the unem-  
1088 ployment rate as a predictor are reported in *Tables S14*  
1089 and *S15 of SI Appendix* for the case of WSJ data, and  
1090 *Tables S10 and S11 of SI Appendix* for the case of Weibo  
1091 data. These results continue to show that the LLM-CPI  
1092 model outperforms consistently the benchmark methods.  
1093 Moreover, the corresponding analyses for the WSJ corpus and  
1094 Weibo corpus exhibit the same qualitative pattern, indicating  
1095 that the performance gains of the LLM-CPI framework  
1096 are robust and broadly applicable across different economic  
1097 environments. Together, these findings suggest a degree of  
1098 *universality* in the effectiveness of online text signals and LLM-  
1099 based surrogate information within the LLM-TS framework.

1100 **The impact of COVID-19 on inflation.** A fundamental question in  
1101 the evaluation of machine learning-based economic forecasting  
1102 models is whether observed performance gains reflect genuine  
1103 predictive content or instead arise from overfitting or spurious  
1104 correlations. To address such concern, we exploit the COVID-  
1105 19 pandemic, which began in early 2020 and constituted a  
1106 major global economic shock affecting both the U.S. and  
1107 Chinese economies, as a natural exogenous event to conduct  
1108 a structural robustness check of the LLM-CPI framework. To  
1109 address this concern, we partition the data into three regimes  
1110 for the U.S. analysis: the *pre-pandemic* period (October 2016  
1111 to January 2020), the *during-pandemic* period (February 2020  
1112 to May 2023), and the *post-pandemic* period (June 2023  
1113 to November 2023). We compare the performance of the  
1114 LLM-CPI framework across these three periods and find that  
1115 the framework maintains its superior performance even during  
1116 the most challenging period of the COVID-19 pandemic.

1117 **Table 1. The rPMSE<sup>AR</sup>( $H$ ), rSign<sup>AR</sup>( $H$ ), Coverage <sub>$m$</sub> ( $H$ ), and Length <sub>$m$</sub> ( $H$ ) results across different horizons  $H$ .**

Type	Method	8	9	10	11	12	13	14	15	Ave.
rPMSE <sup>AR</sup> ( $H$ )	RW	1.172	0.987	0.985	1.081	1.955	1.852	1.026	1.541	1.325
	AVE	1.102	1.095	1.076	1.033	1.194	0.964	1.057	1.246	1.096
	LDA	0.977	0.978	0.981	1.001	0.993	1.003	1.029	1.009	0.996
	BERT	0.919	0.916	0.919	0.954	0.968	0.984	0.969	0.972	0.950
	LLM+LDA	0.882	0.887	0.888	0.898	0.926	0.927	0.964	0.955	0.916
	LLM+BERT	0.827	0.826	0.828	0.877	0.929	0.934	0.925	0.912	0.882
rSign <sup>AR</sup> ( $H$ )	Time-LLM	1.202	2.638	1.904	1.507	1.237	2.334	1.380	2.284	1.811
	PatchTST	1.124	0.911	0.934	0.895	1.638	1.400	1.002	1.354	1.157
	RW	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	AVE	0.833	2.000	2.000	1.667	0.556	0.556	0.444	1.000	1.132
	LDA	0.667	1.000	1.000	1.333	0.444	0.556	0.778	1.333	0.889
	BERT	0.500	1.000	1.000	1.333	0.444	0.444	0.556	1.167	0.806
Coverage <sub><math>m</math></sub> ( $H$ ) (Length <sub><math>m</math></sub> ( $H$ ))	LLM+LDA (BJ)	0.667	1.000	1.000	1.333	0.667	0.444	0.556	1.167	0.854
	LLM+BERT (BJ)	0.500	1.000	1.000	0.667	0.333	0.333	0.444	0.833	0.639
	Time-LLM	0.500	2.667	1.667	2.000	0.556	0.667	0.667	1.167	1.236
	PatchTST	0.500	1.000	1.000	1.667	0.333	0.333	0.333	2.000	0.896
	BERT	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	(3.754)	(3.774)	(3.796)	(3.817)	(3.832)	(3.816)	(3.818)	(3.844)	(3.806)	
Coverage <sub><math>m</math></sub> ( $H$ ) (Length <sub><math>m</math></sub> ( $H$ ))	LDA	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	(3.853)	(3.875)	(3.897)	(3.916)	(3.928)	(3.916)	(3.901)	(3.926)	(3.902)	
	LLM+LDA (BJ)	0.875	0.889	0.900	0.909	1.000	0.923	0.929	0.867	0.912
	(2.864)	(2.960)	(3.052)	(3.151)	(3.189)	(2.852)	(2.614)	(2.681)	(2.920)	
	LLM+BERT (BJ)	0.875	0.889	0.900	1.000	0.917	1.000	0.929	0.933	0.930
	(3.420)	(3.560)	(3.750)	(3.985)	(4.031)	(3.765)	(3.411)	(3.927)	(3.731)	
Coverage <sub><math>m</math></sub> ( $H$ ) (Length <sub><math>m</math></sub> ( $H$ ))	SPI (LDA)	0.875	0.889	0.900	0.909	0.917	0.923	0.929	0.933	0.909
	(3.318)	(3.318)	(3.318)	(3.320)	(3.311)	(3.320)	(3.324)	(3.319)	(3.319)	
	SPI (BERT)	0.875	0.889	0.900	0.909	0.917	0.923	0.929	0.933	0.909
	(3.414)	(3.416)	(3.413)	(3.389)	(3.337)	(3.328)	(3.347)	(3.335)	(3.372)	

1148 Top panel: the relative cumulative PMSE values compared to the AR benchmark; smaller values indicate better performance. Middle panel: the  
1149 relative cumulative sign prediction error values compared to the AR benchmark. Bottom panel: the values in the parentheses are interval length.  
1150 All methods are without unemployment rate, and the BJ prediction interval is used.

1152 to August 2025)\*. This time segmentation enables us to  
1153 examine changes in text content, topic structures, and model  
1154 performance across markedly different economic regimes,  
1155 thereby providing a stringent test of the robustness of the  
1156 proposed approach.

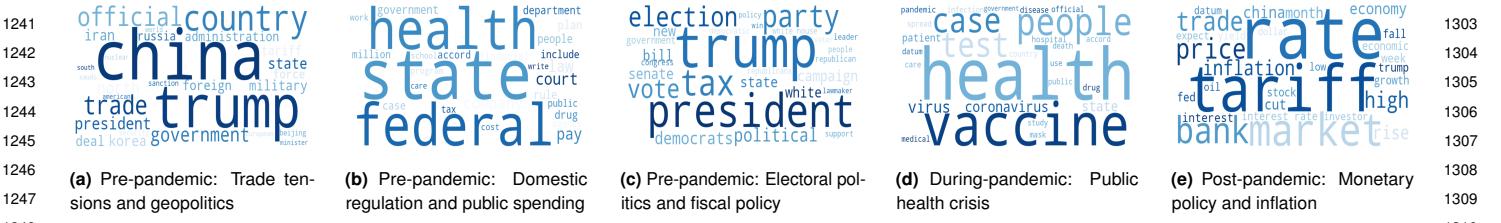
1157 For each period, we independently train the LDA model  
1158 to extract period-specific latent topics. To evaluate the  
1159 prediction performance, we designate the final six months of  
1160 each period as the testing sample, with the remaining (i.e.,  
1161 earlier) months used for training. We also apply the model  
1162 selection procedure on the training sample as discussed in  
1163 *Section 3 of SI Appendix* to reduce the dimensionality of text  
1164 embeddings. *Tables S12 and S16 in SI Appendix* report the  
1165 root prediction mean squared error (PMSE) and the sign  
1166 prediction error (Sign) for the corresponding periods in the  
1167 U.S. and Chinese economies, respectively<sup>†</sup>. These results  
1168 unveil that the LLM-CPI model commonly outperforms the  
1169 benchmark AR model in terms of both PMSE and Sign  
1170 over different periods **across both U.S. and Chinese economic**  
1171 **settings**.

1173 \*The timeline of the COVID-19 pandemic in the U.S. is available at [https://en.wikipedia.org/](https://en.wikipedia.org/wiki/COVID-19_pandemic_in_the_United_States)  
1174 [wiki/COVID-19\\_pandemic\\_in\\_the\\_United\\_States](https://en.wikipedia.org/wiki/COVID-19_pandemic_in_the_United_States). For China, given the presence of nationwide  
1175 lockdowns, we instead define the pre- and during-lockdown period from January 1, 2019 to  
1176 December 31, 2021, and the post-lockdown period from January 1, 2022 to December 31, 2023;  
1177 see [https://en.wikipedia.org/wiki/COVID-19\\_pandemic\\_in\\_mainland\\_China](https://en.wikipedia.org/wiki/COVID-19_pandemic_in_mainland_China).

1178 †We do not use the relative errors here since the baseline AR model occasionally yields zero error  
1179 when the forecast horizon  $H$  is small.

1200 Fig. 4 displays topic visualizations for the WSJ corpus  
1201 across the three periods. The changes in dominant topics  
1202 and keywords reveal a clear shift in textual narratives  
1203 across these regimes. The pre-pandemic topics capture  
1204 policy- and geopolitics-driven cost pressures, regulatory  
1205 and fiscal influences on prices, and political uncertainty  
1206 affecting expectations, all of which play a central role in  
1207 shaping inflation dynamics and help explain their strong  
1208 predictive power for the U.S. inflation. During the pandemic,  
1209 the topic reflects the evolution of the COVID-19 public  
1210 health crisis, emphasizing infection spread, medical capacity,  
1211 vaccination efforts, and government responses that shaped  
1212 economic activity during the pandemic. The post-pandemic  
1213 macroeconomic environment—characterized by interest rate  
1214 adjustments, inflation concerns, trade and tariff dynamics,  
1215 and financial market responses involving banks, investors, and  
1216 asset prices—captures policy-driven and market-mediated  
1217 inflation dynamics in the post-pandemic period, making it  
1218 highly informative for inflation forecasting.

1219 For the Weibo data, which reflects grassroots public  
1220 opinion, the text signals provide more diverse information.  
1221 The selected topics are illustrated in *Figs. S11 and S12 in*  
1222 *SI Appendix*. Beyond topic modeling, the Weibo platform  
1223 also allows for deeper analysis by exploiting user-generated  
1224 hashtags. We select the 10 most frequently occurring  
1225 hashtags in posts to represent the human-readable meaning  
1226 of the topics.



**Fig. 4.** The word cloud visualizations of the LDA topics extracted from the WSJ corpus.

of each topic word, see *Section 7 of SI Appendix* for more details. *Table S24 in SI Appendix* presents the most-discussed original hashtags corresponding to our selected topic during the lockdown period. It is seen that the lockdown-period topics emphasize crisis-related themes such as debt defaults, bankruptcies, and city lockdowns, reflecting the immediate economic fallout. *Tables S25–S28 in SI Appendix* list the four selected topics and related hashtags during the post-lockdown period. These LDA-identified topics highlight how inflationary pressures in China arise from the interaction of input cost shocks, financial market adjustments, household behavior, and external risks.

**Robustness analysis of LLM-TS in the U.S. economy.** The preceding analysis demonstrates that the LLM-TS performs well in predicting urban CPI in both the U.S. and Chinese economies. A natural next question is whether LLM-TS remains effective when applied to other related macroeconomic variables. We address this question by extending LLM-TS to a broad set of macroeconomic price and financial indices. To this end, we analyze a comprehensive panel of U.S. macroeconomic indicators from the FRED-MD database (38). Detailed descriptions of the selected indices are provided in *Table S13 of SI Appendix*.

As shown in *Table S17 of SI Appendix*, the LLM-TS delivers substantial accuracy gains across the vast majority of series. Relative to the AR benchmark, LLM-TS improves the relative prediction mean squared error (rPMSE) in 17 out of 23 cases, and improves the sign accuracy in 16 out of 23 cases.

*Table S19 of SI Appendix* further compares LLM-TS with deep learning-based forecasting methods. The LLM-TS outperforms substantially these alternatives. This result is consistent with the nature of the data: deep forecasting models typically require large-scale time-series corpora to fully exploit their expressive capacity, whereas monthly macroeconomic series contain only a few dozen observations, placing these methods in a few-shot learning regime. Moreover, the CPI and related macroeconomic indices exhibit relatively stable and low-complexity temporal dynamics, for which classical AR-type models and the joint residual modeling employed by LLM-TS are already well suited. In such settings, prediction difficulty arises primarily from data scarcity rather than model flexibility, limiting the advantages of high-capacity neural network models.

Table 2 summarizes the predictive inference results across different U.S. macroeconomic price indices. Overall, LLM-TS produces shorter prediction intervals than both the AR benchmark and SPI while maintaining nominal coverage, achieving the best overall performance for 20 out of 23 series. Relative to the AR benchmark, these results show that

**Table 2. The Coverage<sub>m</sub>(H) and interval length results across different price indices.**

Target	AR	LLM-TS (BJ)	SPI
WPSFD49207	1.00 (4.031)	1.00 (3.105)	1.00 (4.361)
WPSFD49502	1.00 (4.031)	1.00 (3.085)	1.00 (4.302)
PPICMM	0.92 (3.751)	0.58 (2.546)	0.75 (3.330)
CPIAUCSL	1.00 (3.997)	1.00 (2.865)	1.00 (3.268)
CPIAPPSL	0.92 (3.949)	0.83 (2.659)	0.92 (3.428)
CPITRNSL	1.00 (4.024)	1.00 (2.455)	1.00 (3.369)
CPIMEDSL	1.00 (3.959)	1.00 (3.189)	0.92 (3.311)
CUSR0000SAC	1.00 (4.033)	1.00 (2.335)	1.00 (3.975)
CUSR0000SAD	1.00 (3.928)	0.83 (1.474)	1.00 (2.485)
CUSR0000SAS	1.00 (3.914)	1.00 (3.056)	1.00 (2.858)
CPIULFSL	1.00 (4.061)	1.00 (2.739)	1.00 (3.210)
CUSR0000SA0L2	1.00 (4.033)	1.00 (2.889)	1.00 (3.384)
CUSR0000SA0L5	1.00 (3.996)	1.00 (2.823)	1.00 (3.057)
PCEPI	1.00 (3.988)	1.00 (2.704)	1.00 (3.347)
DNDGRG3M086SBEA	1.00 (4.032)	1.00 (2.319)	1.00 (3.577)
DSERRG3M086SBEA	1.00 (3.806)	1.00 (3.882)	0.92 (3.048)
DDURRG3M086SBEA	1.00 (3.937)	0.75 (2.469)	1.00 (3.223)
OILPRICEX	1.00 (4.071)	1.00 (2.277)	1.00 (3.351)
WPSID61	1.00 (3.996)	1.00 (3.818)	1.00 (3.545)
WPSID62	1.00 (4.015)	0.83 (2.381)	1.00 (4.104)
S&P 500	0.83 (3.889)	1.00 (2.950)	0.75 (3.357)
S&P div yield	1.00 (3.927)	1.00 (2.850)	0.75 (2.899)
VIXCLSX	1.00 (3.993)	0.92 (2.794)	0.83 (2.862)
<b>Count (smallest length)</b>	0	20	3

The values in the parentheses are interval length, and coverage near the nominal level of 0.95 with smaller interval length is preferred.

incorporating text embeddings and joint residual modeling in the LLM-TS framework yields robust and systematic gains in predictive inference. The gains are particularly pronounced for the headline CPI, core inflation components, PCE inflation, and apparel and transportation prices, as well as key financial market variables, including the S&P 500, dividend yield, and the VIX. These series are especially sensitive to public narratives, market sentiment, and forward-looking economic expectations that are reflected prominently in financial news coverage.

Compared to SPI, the improvements can be partly attributed to the Box–Jenkins framework, which exploits effectively the approximate normality of macroeconomic index innovations while leveraging LLM-based surrogate information.

## Discussions

We have investigated in this paper the problem of time series forecasting. Motivated by the recent developments in LLMs, our suggested framework of LLM-TS **predictive inference** is rooted on the joint time series model of both observed low-frequency field data and LLM-generated high-

frequency text surrogates. We have applied the proposed framework to inflation forecasting and inference, exploiting the correlations between low-frequency, survey-based inflation measurements and high-frequency, LLM-generated inflation signals constructed from two large-scale online text data sets. The model conditions on lagged monthly inflation indices, lagged LLM-generated daily inflation surrogates, macroeconomic covariates, and online text embeddings. Supported by theoretical guarantees, LLM-TS is shown to deliver accurate point forecasts and tight inflation predictive inference results in both U.S. and Chinese economic environments, thanks to the power of LLMs such as ChatGPT and the trained BERT models as well as text embeddings via LDA and BERT. We have further generalized the LLM-TS framework to deep learning-based time-series models, demonstrating that the underlying joint residual modeling principle is broadly applicable and not restricted to linear specifications.

It would be of interest to incorporate time-series surrogates generated by different LLM tools within the joint modeling framework. Exploring more advanced text-embedding archi-

1365 lectures to extract informative textual features is another  
 1366 promising direction. In addition, extending conformal  
 1367 predictive inference to deep learning-based LLM-TS models  
 1368 would further enhance uncertainty quantification in nonlinear  
 1369 forecasting settings. These directions lie beyond the scope of  
 1370 the current paper and are left for future research.  
 1371

## Data availability

The analytic code is fully documented in executable R  
 1372 Markdown files and publicly available at [https://github.com/  
 1373 suntiansheng/LLM-CPI-prediction-and-inference](https://github.com/suntiansheng/LLM-CPI-prediction-and-inference). The under-  
 1374 lying data sets include the proprietary Wall Street Journal  
 1375 content and Weibo posts accessed under the platform terms.  
 1376 Because the WSJ terms of use may treat certain derived data  
 1377 sets (e.g., embeddings) as reproductions, creating potential  
 1378 legal uncertainty, we do not share the WSJ data. We release  
 1379 only non-expressive derived data sets from Weibo to support  
 1380 reproducibility.  
 1381

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