

CSCI567 – Machine Learning

Assignment #1

By,

Suraj R

USC ID: 4979219097

Question 1:

1(a) Density Estimation:

We know that Beta Distribution is given by

$$f(x) = \left(\frac{x^{\alpha-1} (1-x)^{\beta-1} (\alpha+\beta-1)!}{(\alpha-1)(\beta-1)} \right) \quad (1)$$

And

$$\beta = 1$$

Substituting this in (1), we do get,

$$f(x) = \alpha x^{\alpha-1}$$

Likelihood function L is given by,

$$L = \prod_{i=1}^n f(x_i) = \prod_{i=1}^n \alpha x_i^{\alpha-1}$$

Then,

$$\ln(L) = \ln\left(\prod_{i=1}^n \alpha x_i^{\alpha-1}\right) = \sum_{i=1}^n (\alpha - 1) \ln(x_i)$$

$$n \ln(\alpha) + \sum_{i=1}^n (\alpha - 1) \ln(x_i)$$

$$\frac{\partial \ln L}{\partial \alpha} = \frac{n}{\alpha} + \sum_{i=1}^n \ln(x_i) = 0$$

Hence MLE of α is,

$$\hat{\alpha} = -\frac{n}{\sum_{i=1}^n \ln(x_i)}$$

Normal Distribution

Given that,

$$f(x) = \left[\frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \right]$$

Here $\mu = \theta$ and since the variance of a diagonal matrix is also θ it follows that $\sigma^2 = \theta$.

Substituting this in the above equation,

$$f(x) = \left[\frac{1}{\sqrt{2\pi\theta}} e^{-\frac{(x-\theta)^2}{2\theta}} \right]$$

The likelihood function as discussed in the lecture is given by,

$$L = \prod_{i=1}^n f(x_i) = \prod_{i=1}^n \left[\frac{1}{\sqrt{2\pi\theta}} e^{-\frac{(x-\theta)^2}{2\theta}} \right]$$

The log likelihood is given by,

$$\begin{aligned} \ln(L) &= \ln \left(\prod_{i=1}^n \left[\frac{1}{\sqrt{2\pi\theta}} e^{-\frac{(x-\theta)^2}{2\theta}} \right] \right) \\ &= -\frac{1}{2} \sum_{i=1}^n \ln(2\pi\theta) - \sum_{i=1}^n \left(\frac{x_i^2}{2\theta} \right) - \sum_{i=1}^n \frac{\theta}{2} + \sum_{i=1}^n x_i \\ &= -\frac{n}{2} \ln(2\pi\theta) - \sum_{i=1}^n \frac{x_i^2}{2\theta} - \frac{n\theta}{2} + \sum_{i=1}^n x_i \end{aligned}$$

Taking the partial derivative with respect to $\partial\theta$

$$\begin{aligned} \frac{\partial \ln L}{\partial \theta} &= -\frac{n}{2\theta} + \sum_{i=1}^n \frac{x_i^2}{2\theta} - \frac{n}{2} = 0 \\ n\theta^2 + n\theta - \sum_{i=1}^n x_i^2 &= 0 \end{aligned}$$

Since it's a quadratic equation, the roots of the equation is,

$$\hat{\theta} = -\frac{-n \pm \sqrt{n^2 - \sum_{i=1}^n 4nx_i^2}}{2n}$$

Hence the MLE of θ is as above.

1 (b)

Let x_1, x_2, \dots, x_n are the training inputs and y_1, y_2, \dots, y_n be the outputs. There are N I.i.d samples which are from the normal distribution having mean and variance

$$L = \frac{1}{\sqrt{2\pi\sigma^2}} \prod_{i=1}^N e^{\frac{-(y_i - \omega - \omega^t x_i)}{2\sigma^2}} \quad (1)$$

Taking the partial derivative with respect to ω_0 to find the estimate of parameter ω_0 and setting it to zero.

$$\sum_{i=1}^N \omega_0 = \sum_{i=1}^N y_i \sum_{i=1}^N \omega^t x_i$$

Therefore

$$\widehat{\omega}_0 = \frac{1}{n} \sum_{i=1}^N \omega^t x_i$$

Hence $\widehat{\omega}_0 = \bar{y} - \omega^t \bar{x}$. Substituting this in 1 and taking the partial derivative with respect to $\widehat{\omega}_0$ and setting it to zero,

$$0 + \frac{1}{\sigma^2} \sum_{i=1}^N (y_i - \bar{y}) - \omega^t (x_i - \bar{x}) = 0$$

$$\widehat{\omega} = \left[\sum_{i=1}^N (x_i - \bar{x}) - \omega^t (x_i - \bar{x}) \right] \left[\sum_{i=1}^N (y_i - \bar{y}) - \omega^t (y_i - \bar{y}) \right]$$

Also, $x_i^c = x_i - \bar{x}, y_i^c = y_i - \bar{y}$

Substituting this in the above equation,

$$\widehat{\omega} = \left[\sum_{i=1}^N (x_i^c)(x_i^c)^T \right]^{-1} \left[\sum_{i=1}^N (y_i^c)(y_i^c)^T \right]^{-1}$$

$$\widehat{\omega} = (X_c^T X_c)^{-1} X_c^T Y_c$$

Hence proved.

1 (c)

We know that,

$$\hat{f}(x) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h} K\left(\frac{x - X_i}{h}\right)$$

$$\begin{aligned}
E_{X_1, \dots, X_n}[\hat{f}(x)] &= \frac{1}{n} \sum_{i=1}^n E \left[\frac{1}{h} K \left(\frac{x - X_i}{h} \right) \right] \\
&= E \left[\frac{1}{h} K \left(\frac{x - X}{h} \right) \right] \\
&= \frac{1}{h} \int_{-\infty}^{\infty} K \left(\frac{x - t}{h} \right) f(t) dt \\
E_{X_1, \dots, X_n}[\hat{f}(x)] &= \frac{1}{h} \int_{-\infty}^{\infty} K \left(\frac{x - t}{h} \right) f(t) dt
\end{aligned}$$

Given that $z = \frac{x-t}{h} \Rightarrow t = -zh + x$

Applying this in the above equation and by using Taylor's theorem we get,

$$\begin{aligned}
E\hat{f}(x) &= \int K(z) f(x - hz) dz \\
&= \int K(z) \left[f(x) - hzf'(x) + h^2 z^2 \frac{f''(x)}{2} - h^3 z^3 \frac{f'''(x)}{3} + \dots \right] \\
&= \int f(x) K(z) dz - \int hzf'(x) K(z) dz + \int h^2 z^2 \frac{f''(x)}{2} K(z) dz + \dots
\end{aligned}$$

As we know that,

$$\int K(z) dz = 1, \int zK(z) = 0 \text{ and } \int z^2 K(z) = \sigma_K^2$$

$$= f(x) + h^2 \sigma_K^2 \frac{f''(x)}{2} + o(h^2)$$

$$\text{which is equal to, } E[\hat{f}(x)] = f(x) + h^2 \sigma_K^2 \frac{f''(x)}{2} + o(h^2) \quad (2)$$

In order to compute the bias, we have from equation 2 and the bias is given as below.

$$E[\hat{f}(x)] - f(x) = h^2 \sigma_K^2 \frac{f''(x)}{2} + o(h^2)$$

1. d

The kernel density estimation is given by,

$$\hat{f}(x) = \frac{1}{n} \sum_i^n K\left(\frac{x - X_i}{h}\right)$$

In MLE, there is only one global minimum which we find, but in kernel, there are many local minimum. So if we estimate h using MLE, MLE will give us only one value and many will be ignored. Hence MLE cannot be used to estimate h in $\hat{f}(x)$

Question 2.a

Based on the given data,

The Features are provided in x and y coordinates as below.

$$x = \{15, -7, -4, 29, 32, 37, 18, 40, -8, -11\}$$

$$y = \{49, 38, 47, 24, 36, 43, 9, -28, -19, 12\}$$

Using the given data, we compute the mean and variance of each feature:

For x:

$$\mu_x = \frac{1}{n} \sum_n x_n = \frac{15 - 7 - 4 + 29 + 32 + 37 + 18 + 40 - 8 - 11}{10} = 14.1$$

$$\sigma_x^2 = \frac{1}{(n-1)} \sum (x_n - \mu_x)^2$$

$$= \frac{1}{9} [(15 - 14.1)^2 + (-7 - 14.1)^2 + (-4 - 14.1)^2 + (29 - 14.1)^2 + (32 - 14.1)^2 + (37 - 14.1)^2 + (18 - 14.1)^2 + (40 - 14.1)^2 + (-8 - 14.1)^2 + (-11 - 14.1)^2]$$

$$\sigma_x^2 = 404.98$$

$$\sigma_x = 20.12$$

Similarly for y,

$$\mu_y = \frac{1}{n} \sum_n y_n = \frac{49 + 38 + 47 + 24 + 36 + 43 + 9 - 28 - 19 + 12}{10} = 21.1$$

$$\sigma_y^2 = \frac{1}{(n-1)} \sum (y_n - \mu_y)^2 = 743.65$$

$$\sigma_y = 27.26$$

Using the above values, we calculate the normalized and scaled co-ordinates.

$$x' = \left(\frac{x - \mu_x}{\sigma_x} \right) \text{ and } y' = \left(\frac{y - \mu_y}{\sigma_y} \right)$$

Mathematics	0.04	1.01
Mathematics	-1.04	0.61
Mathematics	-0.89	-0.92
Electrical Engineering	0.74	0.10
Electrical Engineering	0.88	0.54
Electrical Engineering	1.13	0.80
Computer Science	0.19	-0.44
Computer Science	1.28	-1.80
Computer Science	-1.09	-1.47
Computer Science	-1.24	-0.33

Given the coordinates of the student: $x = 9$ and $y = 18$

Using the above data, the normalized coordinates are: $x = -0.25$ and $y = -0.11$

We need to classify the given student based on the above information.

We need to calculate L_1 and L_2 using the formula which is mentioned as below.

$$L_1 = |x - x_n| + |y - y_n|$$

$$L_2 = \sqrt{|x - x_n|^2 + |y - y_n|^2}$$

ID	Class	X_n	Y_n	L_1 Distance	L_2 Distance
1	Mathematics	0.04	1.01	1.41	1.25
2	Mathematics	-1.04	0.61	1.51	0.52
3	Mathematics	-0.89	-0.92	1.45	0.66
4	Electrical Engineering	0.74	0.1	1.2	0.04
5	Electrical Engineering	0.88	0.54	1.78	0.42
6	Electrical Engineering	1.13	0.8	2.29	0.83
7	Computer Science	0.19	-0.44	0.77	0.11
8	Computer Science	1.28	-1.8	3.22	2.86
9	Computer Science	-1.09	-1.47	2.2	1.85
10	Computer Science	-1.24	-0.33	1.21	0.05

Using the above information, we now calculate

For $K=1$ and L_1

$nn(x)$ with L_1 is 0.19

Hence the given student belongs to Computer Science.

For $K=3$ and L_1

$nn(x)$ with L_1 and $K = 3$ Is Electrical Engineering = 1.2 Computer Science = {0.77, 1.21}

Hence the given student belongs to Computer Science.

For $K=1$ and L_2

$nn(x)$ with L_2 and $K = 1$ is Electrical Engineering = 0.04

Hence the given student belongs to Electrical Engineering.

For $K=3$ and L_2

$nn(x)$ with L_2 and $K = 3$ is Computer Science {0.05, 0.11} Electrical Engineering {0.04}

Hence the given student belongs to Computer Science.

We can see that when $k=3$, we got the same results as Computer Science irrespective of L_1 and L_2 . But when $k=1$, the neighbors changed based on L_1 and L_2 metrics.

Hence we can see the difference in the values.

Question 2.b

From the definition of Total Probability theorem, we have,

$$\sum_n K_c = K$$

$$p(x|y = c_i) = \frac{K_c}{N_c V}$$

$$p(y = c_i) = \frac{N_c}{N}$$

$$p(x) = \sum_{i=1}^n p(x|y = c_i) \times p(y = c_i)$$

$$\frac{K_{c_1}}{N_{c_1} V} \times \frac{N_{c_1}}{N} + \frac{K_{c_2}}{N_{c_2} V} \times \frac{N_{c_2}}{N} + \dots + \frac{K_{c_n}}{N_{c_n} V} \times \frac{N_{c_n}}{N}$$

$$= \frac{K_{c_1} + K_{c_2} + K_{c_3} + \dots + K_{c_n}}{NV} = \frac{K}{NV}$$

$$\Leftrightarrow p(x) = \frac{K}{NV}$$

Now calculate $p(Y = c|x)$

$$p(y = c|x) = \frac{p(x|y = c) \times p(y = c)}{p(x)}$$

$$= \left[\frac{\frac{K_c}{N_c V} \times \frac{N_c}{N}}{\frac{K}{N_c V}} \right] = \frac{K_c}{K}$$

$$\Leftrightarrow p(y = c|x) = \frac{K_c}{K}$$

Question 3:

Consider linear regression of the form,

$$y(x, \omega) = \omega_0 + \omega^T x$$

And the sum of squares of error function of the form,

$$E(\omega) = \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{y(\widetilde{x}_n^m, \omega) - t_n\}^2$$

Upon substituting,

$$\begin{aligned} &= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{(\omega_0 + \omega^T(x_n + \epsilon_m))\}^2 \\ &= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{\omega_0 + \omega^T \epsilon_m + \omega^T x_n - t_n\}^2 \\ &= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{(y(x_n, \omega) - t_n)\}^2 + \omega^T \epsilon_m \\ &= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{(y(x_n, \omega) - t_n)\}^2 + \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \omega^{T^2} \epsilon_m^2 \end{aligned}$$

Using the (a+b)² formula and expanding and equating the noise term to zero, we get

$$E(\omega) = \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{(y(x_n, \omega) - t_n)\}^2$$

3b the dropout noise corresponds to setting of $\widetilde{x}_{n,d}$ to 0 with the probability δ and $\frac{x_{n,d}}{(1-\delta)}$ with the probability $1 - \delta$. Hence from the definition of variance and expected value, we know that

$$E(\widetilde{x}_n) = 0 * (\delta) + \frac{x_n}{1 - \delta} * (1 - \delta)$$

$$\begin{aligned}
E(\widetilde{x}_n) &= x_n \\
Var[\omega^T \widetilde{x}_n] &= E[(\omega^T \widetilde{x}_n)^2] - E(\omega^T \widetilde{x}_n)^2 \\
Var[\omega^T \widetilde{x}_n] &= E[(\omega^T \widetilde{x}_n)^2] - x_n^2 \\
&= \sum_{d=1}^D \left[\omega_d^2 \frac{x_{n,d}^2}{(1-\delta)^2} (1-\delta) \right] - x_n^2 \\
&= \sum_{d=1}^D \left[\omega_d^2 \frac{x_{n,d}^2}{(1-\delta)^2} - x_{n,d}^2 \right]
\end{aligned}$$

Hence

$$Var[\omega^T \widetilde{x}_n] = \frac{\delta}{1-\delta} \sum_{d=1}^D \omega_d^2 x_{n,d}^2$$

3.c By using the above results and the given E averaged over the noise distribution,

$$\begin{aligned}
E(\omega) &= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N \{y(\widetilde{x}_n^m, \omega) - t_n\}^2 \\
E(\omega) &= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N [(\omega_0 - t_n) + (\omega^T \widetilde{x}_n)]^2
\end{aligned}$$

Here let $a = (\omega_0 - t_n)$ and $b = (\omega^T \widetilde{x}_n)$ and using $(a+b)^2$ formula

$$= \frac{1}{2} \sum_{m=1}^M \sum_{n=1}^N (\omega_0 - t_n)^2 + (\omega^T \widetilde{x}_n)^2 + 2(\omega_0 - t_n)(\omega^T \widetilde{x}_n)$$

On simplifying,

$$= \frac{1}{2} \sum_{n=1}^N (M(\omega_0 - t_n)^2 + 2(\omega_0 - t_n)\omega^T \sum_{m=1}^M \widetilde{x}_n^m + \sum_{m=1}^M (\omega^T \widetilde{x}_n)^2)$$

Here,

$$\begin{aligned}
\sum_{m=1}^M \widetilde{x}_n^m &= Mx_n \text{ and } \sum_{m=1}^M (\omega^T \widetilde{x}_n)^2 = M(\omega^T \widetilde{x}_n)^2 \\
&= \frac{1}{2} \sum_{n=1}^N (M(\omega_0 - t_n)^2 + 2(\omega_0 - t_n)\omega^T Mx_n + M(\omega^T \widetilde{x}_n)^2)
\end{aligned}$$

On simplifying using the results of mean and variance, we get,

$$= \frac{M}{2} \left(\sum_{n=1}^N (y(x_n, \omega) - t_n)^2 \right) + \sum_{n=1}^N \frac{\delta}{1 - \delta} \sum_{d=1}^D \omega_d^2 x_{n,d}^2$$

From the above proof, it has been proved that minimizing E averaging over the noise distribution is equivalent to minimizing the sum of squares error for noise free input variables with the addition of weight decay regularization term.

Question 4:

4 a. We can use 3 entropy formula,

$$H[X] = - \sum_{k=1}^K P(x_k) \log P(x_k)$$

$$H[Y|X] = \sum_k P(x_k) H[Y|x_k]$$

$$IG = H[Y] - H[Y|X]$$

$$\text{Entropy: } H(T) = -\frac{9}{14} \log\left(\frac{9}{14}\right) - \frac{5}{14} \log\left(\frac{5}{14}\right) = 0.94$$

$$\begin{aligned} H(\text{temp}, T) &= \frac{4}{14} H(T_{\text{hot}}) + \frac{6}{14} H(T_{\text{mild}}) + \frac{4}{14} H(T_{\text{cool}}) \\ &= \frac{4}{14} \left(-\frac{2}{4} \log\left(\frac{2}{4}\right) - \frac{2}{4} \log\left(\frac{2}{4}\right) \right) + \frac{6}{14} \left(-\frac{4}{6} \log\left(\frac{4}{6}\right) - \frac{2}{6} \log\left(\frac{2}{6}\right) \right) \\ &\quad + \frac{4}{14} \left(-\frac{3}{4} \log\left(\frac{3}{4}\right) - \frac{1}{4} \log\left(\frac{1}{4}\right) \right) \\ &= \frac{4}{14} (1) + \frac{6}{14} (0.918) + \frac{4}{14} (0.81) \\ &= 0.911 \end{aligned}$$

$$\text{Gain}(\text{temperature}, T) = 0.940 - 0.911 = 0.029$$

$$\begin{aligned} H(\text{outlook}, T) &= \frac{5}{14} H(T_{\text{sunny}}) + \frac{4}{14} H(T_{\text{overcast}}) + \frac{5}{14} H(T_{\text{rain}}) \\ &= \frac{5}{14} \left(-\frac{3}{5} \log\left(\frac{3}{5}\right) - \frac{2}{5} \log\left(\frac{2}{5}\right) \right) + \frac{4}{14} \left(-\frac{4}{4} \log\left(\frac{4}{4}\right) \right) + \frac{5}{14} \left(-\frac{3}{5} \log\left(\frac{3}{5}\right) - \frac{2}{5} \log\left(\frac{2}{5}\right) \right) \\ &= 0.347 + 0 + 0.347 \\ &= 0.694 \end{aligned}$$

$$Gain(overcast, T) = 0.940 - 0.694 = 0.246$$

$$\begin{aligned} H(Humidity, T) &= \frac{7}{14}H(T_{normal}) + \frac{7}{14}H(T_{high}) \\ &= \frac{7}{14} \left(\left(-\frac{3}{7} \log\left(\frac{3}{7}\right) - \frac{4}{7} \log\left(\frac{4}{7}\right) \right) + \frac{7}{14} \left(-\frac{6}{7} \log\left(\frac{6}{7}\right) - \frac{1}{7} \log\left(\frac{1}{7}\right) \right) \right) \\ &= 0.493 + 0.296 = 0.789 \end{aligned}$$

$$Gain(humidity, T) = 0.151$$

$$\begin{aligned} H(wind, T) &= \frac{8}{14}H(T_{weak}) + \frac{6}{14}H(T_{strong}) \\ &= \frac{8}{14} \left(-\frac{2}{8} \log\left(\frac{2}{8}\right) - \frac{6}{8} \log\left(\frac{6}{8}\right) \right) + \frac{6}{14} \left(-\frac{3}{6} \log\left(\frac{3}{6}\right) - \frac{3}{6} \log\left(\frac{3}{6}\right) \right) \\ &= 0.464 + 0.429 = 0.893 \end{aligned}$$

$$Gain(wind, T) = 0.047$$

$$\begin{aligned} H(Temperature, Sunny) &= \frac{2}{5}H(T_{hot}) + \frac{2}{5}H(T_{mild}) + \frac{1}{5}H(T_{cool}) \\ &= 0.4 \end{aligned}$$

$$Gain(temperature, T) = 0.571$$

$$H(humidity, Sunny) = \frac{3}{5}H(T_{high}) + \frac{2}{5}H(T_{normal}) = 0$$

$$Gain(humidity, T) = 0.971$$

$$H(temperature, rain) = \frac{3}{5}H(T_{mild}) + \frac{2}{5}H(T_{cool}) = 0.951$$

$$Gain(temperature, T) = 0.02$$

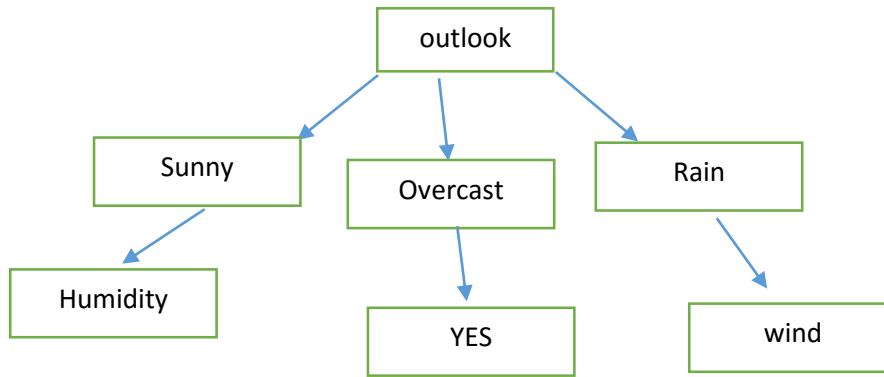
$$H(humidity, rain) = \frac{2}{5}H(T_{high}) + \frac{3}{5}H(T_{normal}) = 0.951$$

$$Gain(humidity, T) = 0.02$$

$$H(wind, rain) = \frac{3}{5}H(T_{weak}) + \frac{2}{5}H(T_{strong}) = 0$$

$$Gain(wind, T) = 0.951$$

Since Humidity and wind are having highest values at depth=2, we consider these attributes as the factors.



4 b

We know that,

$$GiniIndex = \sum_{k=1}^K p_k(1 - p_k)$$

$$Cross\ Entropy = - \sum_{k=1}^k p_k \log(p_k)$$

$$\sum_{k=1}^K p_k(1 - p_k) \leq - \sum_{k=1}^k p_k \log(p_k)$$

$$\sum_{k=1}^k [p_k(1 - p_k + \log(p_k))] \leq 0$$

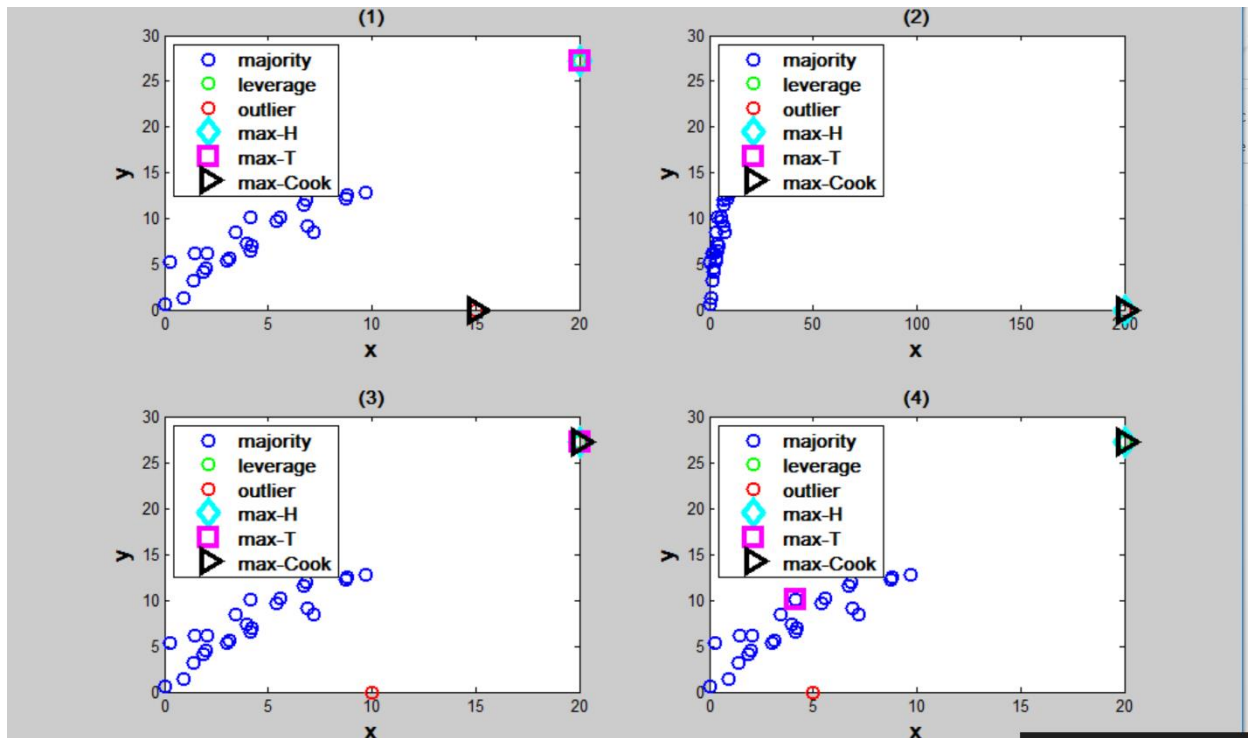
for each $k \in K$,

$$p_k(1 - p_k + \log(p_k)) \leq 0 \text{ because } 0 \leq p_k \leq 1$$

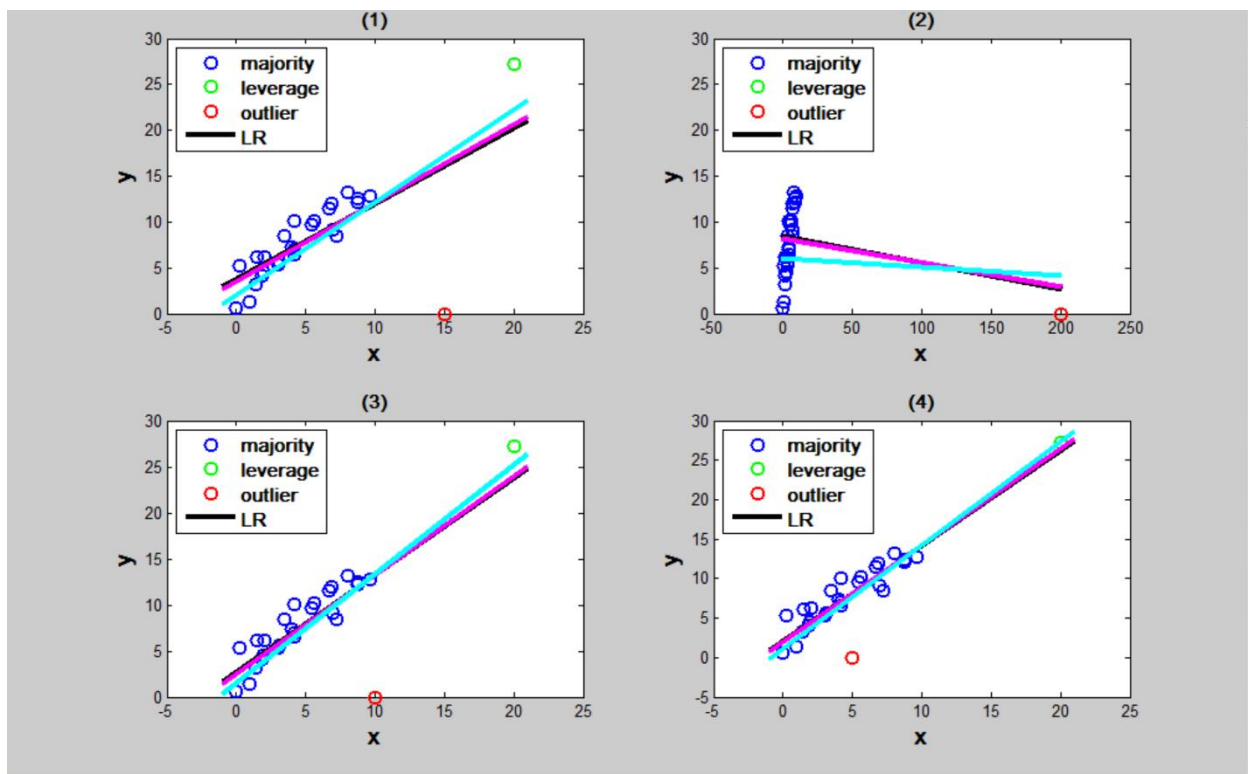
This term is less than $\log(p_k)$ because $1 - p_k \geq 0$. Therefore Gini index is less than or equal to cross entropy. $[\log(p_k) \leq 0 \forall 0 \leq p_k \leq 1]$

Question 5.1

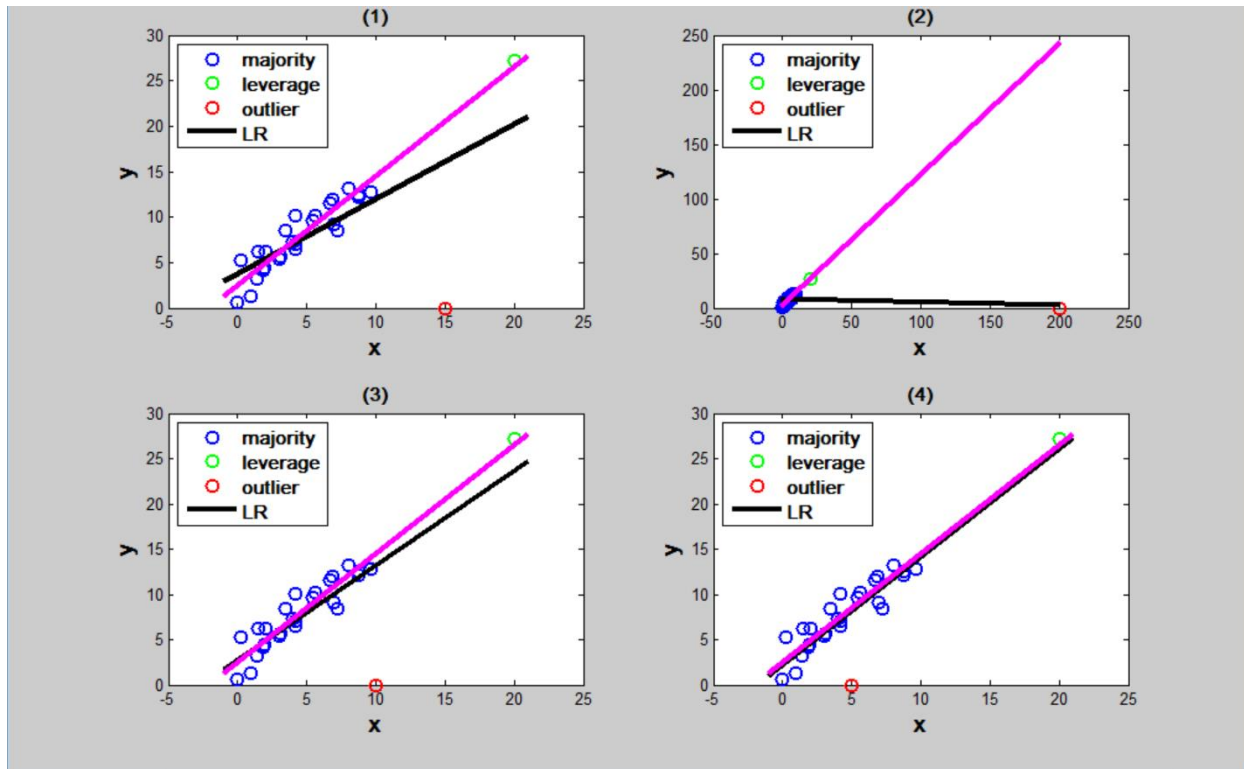
Linear regression model for the dataset data1



Linear regression model for the dataset data2



Linear regression model for the dataset data3 and data4



- The weight decay parameter doesn't significantly reduce the influence of outlier sample. Larger the value of lambda greater will be the penetration hence variance will decrease but bias increases.
- We could significantly reduce if we use Gaussian distribution because Laplace distribution has sharp peak when compared to Gaussian. Also Laplace distribution is quadratic and it grows faster than Gaussian.

Question 5 a.

Probability density function for Laplace is,

$$f(x|\mu, b) = \frac{1}{2b} e^{\left(-\frac{(x-\mu)}{b}\right)}$$

If Laplace distribution is used to model the noise,

$$f(noise|\mu, b) = \frac{1}{2b} e^{\left(-\frac{noise}{b}\right)}$$

On taking log likelihood,

$$Lf(noise|\mu, b) = \ln \pi_{i=1}^n \frac{1}{2b} e^{\left(-\frac{noise}{b}\right)}$$

$$= \ln \sum_{i=1}^n \frac{1}{2b} e^{\left(-\frac{\text{noise}}{b}\right)}$$

$$= -n \ln 2b - \sum_{i=1}^n \frac{|y_n - \omega^T x|}{b}$$

Thus inorder to maximize the left hand side, the right hand side should be minimized. Hence the objective function is $\sum_{i=1}^n \frac{|y_n - \omega^T x|}{b}$

Question 5.2

Leave one out strategy

- Test and Validation datasets using Decision Tree and *gini index* split criterion.

MinLeaf	Training	Test	Validation
1	94.88679	98.53488	97.87234
2	95.88679	98.06977	96.34043
3	95.88679	98.13953	96.74468
4	95.96226	97.13953	96.74468
5	95.81132	97.27907	96.74468
6	95.73585	97.27907	95.21277
7	95.73585	96.27907	95.21277
8	95.28302	96.27907	94.21277
9	96.92453	96.81395	93.55319
10	96.84906	96.81395	93.55319

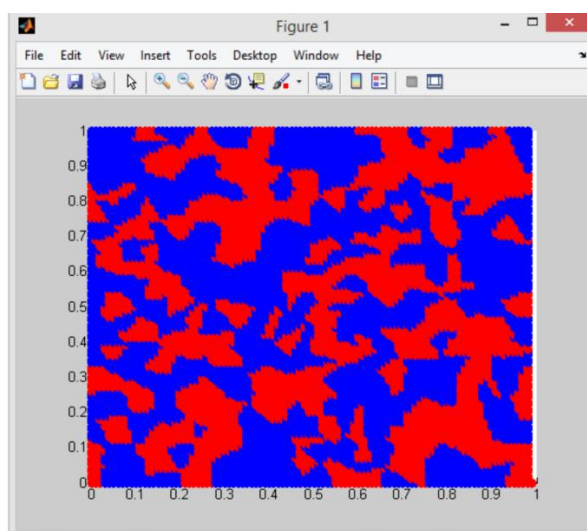
Test and Validation datasets using Decision Tree and *cross – entropy* split criterion.

MinLeaf	Training	Test	Validation
1	97.26415	98.06977	96.340426
2	97.26415	98.67442	95.808511
3	97.26415	98.67442	95.744681
4	97.26415	96.67442	94.744681

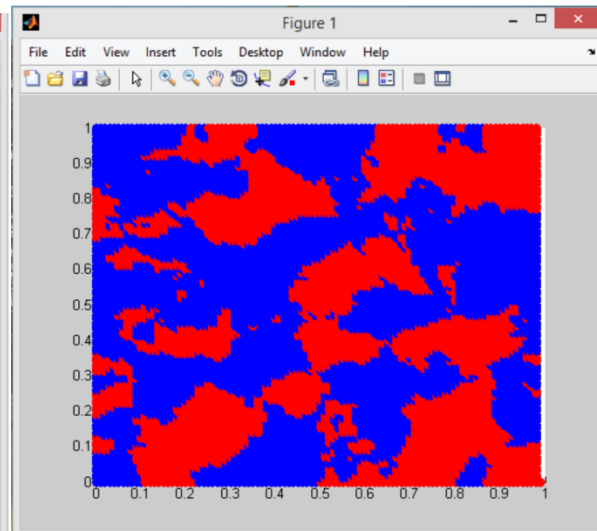
5	98.81132	96.2093	96.744681
6	98.73585	96.81395	96.212766
7	98.73585	96.81395	97.212766
8	96.73585	94.81395	93.212766
9	97.83019	94.81395	92.553191
10	97.75472	96.81395	94.553191

- Test and Validation datasets using k nearest neighbor approach and values of k as: 1,3,5, 15

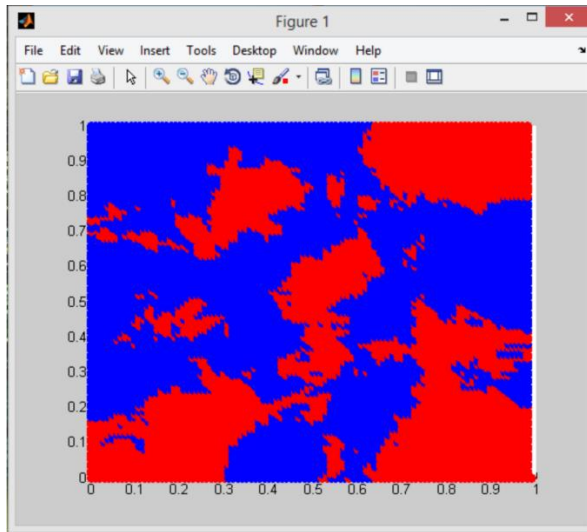
K	Training	Test	Validation
1	54.33962	55.74419	63.70213
3	91.81132	63.93023	63.70213
5	93.09434	59.13953	63.70213
7	92.26415	56.53488	63.70213
9	94.09434	60.74419	63.70213
11	91.39623	60.74419	63.70213
13	93.4717	60.74419	63.70213
15	94.01887	61.74419	63.70213



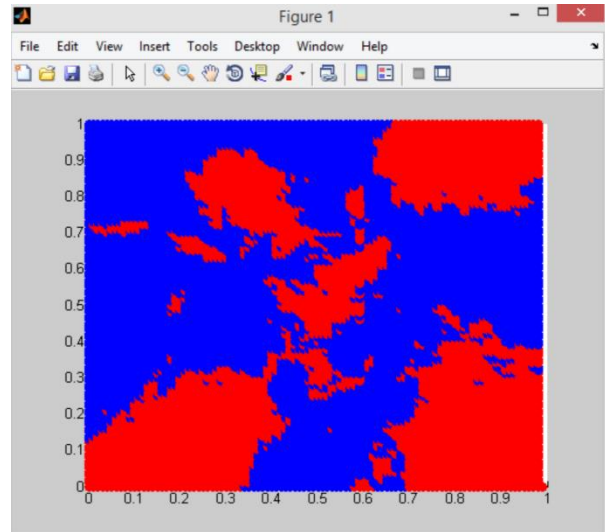
K=1



K=5



K=15



k=25

Collaboration: Neel shah