

Analysis of MK2 transcript levels in TCGA-LUAD (NSCLC) dataset

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Revisions (since rev2):	
1. We’ve added the concept of validating RNASeq counts using a secondary extraction process. Using TCGAbiolinks, we extracted normalized RSEM reads and compared these reads to those obtained via OncoLnc to ensure that the read counts were largely similar at a patient-to-patient level	

2. The original scatter plot showing differences in MK2 is restricted to only “Adenocarcinomas” within LUAD. We’ve applied that restriction to all the regression analyses as well
3. Some small clean-ups to the code
4. We’ve re-analyzed the proportion dead by tertiles of MK2 (instead of quantiles)
5. One advantage of downloading TCGA data is that we also happened to grab MK2 transcript levels from 50 samples with normal tissue controls, thus allowing us to calculate fold change in MK2 transcript levels between normal and tumor in matched samples.
6. R censoring is now performed at one year and two year. This is in part to harmonize analyses with the validation cohort, where censoring could only be done at 2 years.

Data retrieval and cleaning

First, we are downloading clinical data from TCGA-LUAD dataset using the TCGAbiolinks API. For the MK2 transcript data, I grabbed these data from oncoLnc.org using “MAPKAPK2” as the gene name and TCGA-LUAD as the cohort.

The code used to generate the TCGA manual reads is provided elsewhere.

Retrieve data from TCGA dataset, __join it with the MK2 transcript data, with additional joins for the manually extracted reads from TCGA and for Casp3 (from OncoLnc)

We retrieve the clinical data directly using GDC_query. For the MK2 transcript data, this was downloaded from OncoLnc.org using the following parameters: Gene-MAPKAPK2, and percentile: 50/50. We did this because this allowed us to get transcript levels on all available patients (since upper and lower 50% would include all patients). We then calculated the quantiles manually, so that in the future, if we needed to change the threshold, we could do so without re-downloading the OncoLnc data.

In this version, there is an additional step. Using TCGAbiolinks() we have manually extracted normalized RSEM reads from the LUAD dataset and compared it to what was given to us by OncoLnc

```
library(knitr)
opts_chunk$set(tidy.opts=list(width.cutoff=50), tidy=TRUE)
```

```
knitr::opts_chunk$set(echo = TRUE)
rm(list = ls())
library("SummarizedExperiment")
```

```
## Loading required package: MatrixGenerics
```

```
## Loading required package: matrixStats
```

```
##
```

```
## Attaching package: 'MatrixGenerics'
```

```
## The following objects are masked from 'package:matrixStats':
```

```
##
```

```
## colAlls, colAnyNAs, colAnys, colAvgsPerRowSet, colCollapse,
## colCounts, colCummaxs, colCummins, colCumprods, colCumsums,
## colDiffs, colIQRDiffs, colIQRs, colLogSumExps, colMadDiffs,
## colMads, colMaxs, colMeans2, colMedians, colMins, colOrderStats,
## colProds, colQuantiles, colRanges, colRanks, colSdDiffs, colSds,
## colSums2, colTabulates, colVarDiffs, colVars, colWeightedMads,
## colWeightedMeans, colWeightedMedians, colWeightedSds,
## colWeightedVars, rowAlls, rowAnyNAs, rowAnys, rowAvgsPerColSet,
```

```

##      rowCollapse, rowCounts, rowCummaxs, rowCummins, rowCumprods,
##      rowCumsums, rowDiffs, rowIQRDiffs, rowIQRs, rowLogSumExps,
##      rowMadDiffs, rowMads, rowMaxs, rowMeans2, rowMedians, rowMins,
##      rowOrderStats, rowProds, rowQuantiles, rowRanges, rowRanks,
##      rowSdDiffs, rowSds, rowSums2, rowTabulates, rowVarDiffs, rowVars,
##      rowWeightedMads, rowWeightedMeans, rowWeightedMedians,
##      rowWeightedSds, rowWeightedVars

## Loading required package: GenomicRanges

## Loading required package: stats4

## Loading required package: BiocGenerics

## Loading required package: parallel

##
## Attaching package: 'BiocGenerics'

## The following objects are masked from 'package:parallel':
##
##      clusterApply, clusterApplyLB, clusterCall, clusterEvalQ,
##      clusterExport, clusterMap, parApply, parCapply, parLapply,
##      parLapplyLB, parRapply, parSapply, parSapplyLB

## The following objects are masked from 'package:stats':
##
##      IQR, mad, sd, var, xtabs

## The following objects are masked from 'package:base':
##
##      anyDuplicated, append, as.data.frame, basename, cbind, colnames,
##      dirname, do.call, duplicated, eval, evalq, Filter, Find, get, grep,
##      grepl, intersect, is.unsorted, lapply, Map, mapply, match, mget,
##      order, paste, pmax, pmax.int, pmin, pmin.int, Position, rank,
##      rbind, Reduce, rownames, sapply, setdiff, sort, table, tapply,
##      union, unique, unsplit, which.max, which.min

## Loading required package: S4Vectors

##
## Attaching package: 'S4Vectors'

## The following objects are masked from 'package:base':
##
##      expand.grid, I, unname

## Loading required package: IRanges

## Loading required package: GenomeInfoDb

## Loading required package: Biobase

## Welcome to Bioconductor
##
##      Vignettes contain introductory material; view with
##      'browseVignettes()'. To cite Bioconductor, see
##      'citation("Biobase")', and for packages 'citation("pkgname)".

##
## Attaching package: 'Biobase'

```

```

## The following object is masked from 'package:MatrixGenerics':
##
##     rowMedians
## The following objects are masked from 'package:matrixStats':
##
##     anyMissing, rowMedians
library("dplyr")

##
## Attaching package: 'dplyr'
## The following object is masked from 'package:Biobase':
##
##     combine
## The following objects are masked from 'package:GenomicRanges':
##
##     intersect, setdiff, union
## The following object is masked from 'package:GenomeInfoDb':
##
##     intersect
## The following objects are masked from 'package:IRanges':
##
##     collapse, desc, intersect, setdiff, slice, union
## The following objects are masked from 'package:S4Vectors':
##
##     first, intersect, rename, setdiff, setequal, union
## The following objects are masked from 'package:BiocGenerics':
##
##     combine, intersect, setdiff, union
## The following object is masked from 'package:matrixStats':
##
##     count
## The following objects are masked from 'package:stats':
##
##     filter, lag
## The following objects are masked from 'package:base':
##
##     intersect, setdiff, setequal, union
library("DT")
library("TCGAbiolinks")
library(survival)
library(ggplot2)
library(ggpubr)
library(survminer)
library(knitr)
library("tidyr")

##
## Attaching package: 'tidyr'

```

```

## The following object is masked from 'package:S4Vectors':
##
##     expand
library("ggpmisc")

## Loading required package: ggpp
##
## Attaching package: 'ggpp'
## The following object is masked from 'package:ggplot2':
##
##     annotate
clinical <- GDCquery_clinic(project = "TCGA-LUAD",
  type = "clinical")
expr <- read.csv("../rawdata/LUAD.MAPKAPK2Exp.csv",
  header = TRUE)
colnames(expr) <- c("Patient", "MK2.Expression")
expr2 <- read.csv("../rawdata/LUAD.MAPKAPK2Exp.manual.csv",
  header = FALSE)
Casp3exp <- read.csv("../rawdata/LUAD.casp3Exp.2.csv")

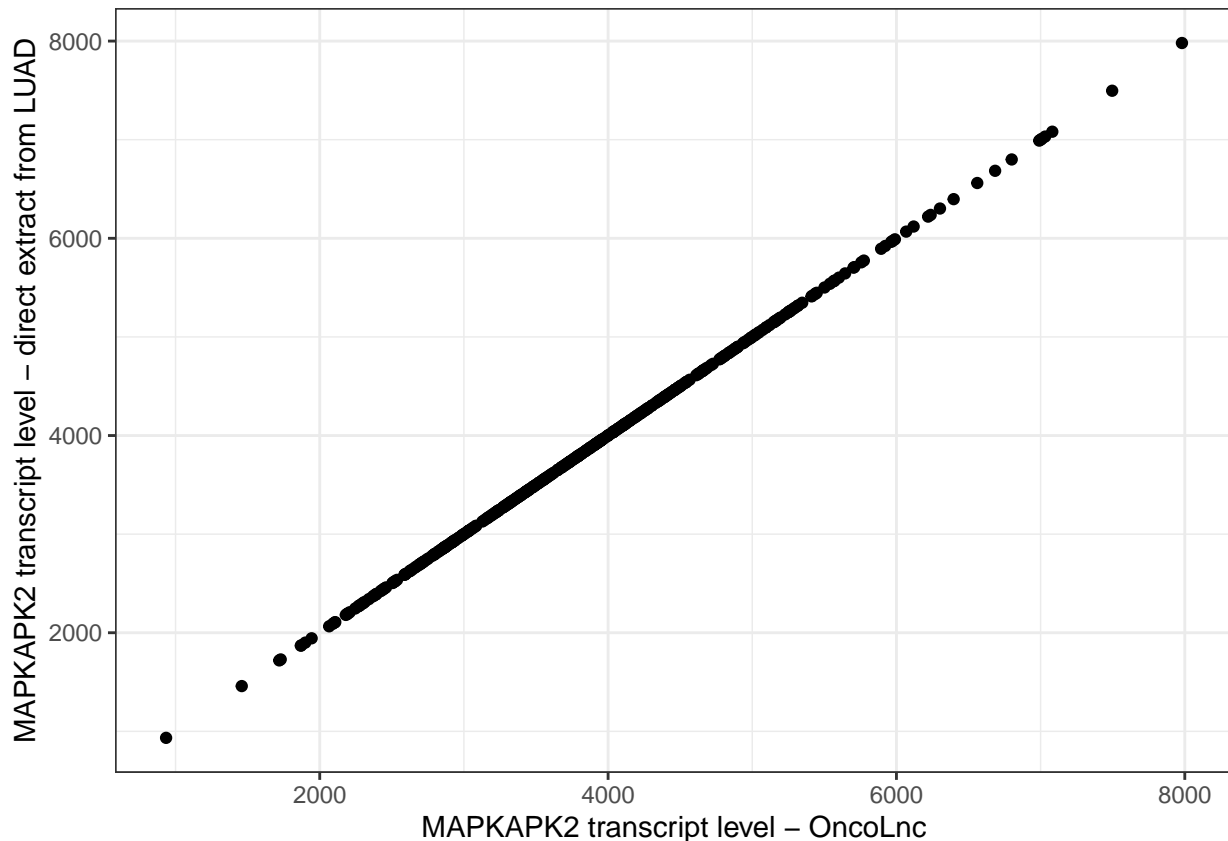
colnames(expr2) <- c("MK2.Expression.Manual", "Stage",
  "Patient")
expr2_normalcontrols <- expr2[grepl("11A", expr2$Patient),
  ]
# take out normal controls
expr2 <- expr2[!grepl("11A", expr2$Patient), ]

# remove non '01A' samples
expr3 <- expr2[grepl("01A", expr2$Patient), ]
expr3$Patient <- substring(expr3$Patient, 1, 12)
colnames(Casp3exp) <- c("Patient", "Casp3.Expression")
# the left_join. we need the clinical data, so we
# left_joined by TCGA patient id to the clinical
# dataset. This means that if there was a patient
# with clinical data without transcript levels,
# that would be a 'NA'.
clinical <- left_join(clinical, expr, c(submitter_id = "Patient"))
clinical <- left_join(clinical, expr3, c(submitter_id = "Patient"))
clinical <- left_join(clinical, Casp3exp, c(submitter_id = "Patient"))

ggplot(clinical, aes(x = MK2.Expression, y = MK2.Expression.Manual)) +
  geom_point() + theme_bw() + ylab("MAPKAPK2 transcript level - direct extract from LUAD") +
  xlab("MAPKAPK2 transcript level - OncoLnc")

## Warning: Removed 104 rows containing missing values (geom_point).

```



```
clinical <- clinical[!is.na(clinical$vital_status),
]
```

This shows that the correlation between MK2 Expression derived from OncoLnc and from TCGA is excellent. Thus, we feel confident that the MK2 reads data from OncoLnc can be used going forward. This makes it easier to grab mRNA transcript levels for other genes directly from OncoLnc rather than having to extract it from TCGA via GDC.

Clean the data, create a df for an LR and Cox PH analyses later

Now that we have left-joined the data, and checked to make sure that the join was done properly, we move on to prep the data for Cox PH work.

```
clinical$time <- mapply(function(x, y) if (is.na(x)) return(y) else return(x),
  clinical$days_to_death, clinical$days_to_last_follow_up)
clinical$age_at_diagnosis <- clinical$age_at_diagnosis/365
clinical$race <- as.factor(clinical$race)
clinical$smoking <- lapply(clinical$cigarettes_per_day,
  function(x) if (!is.na(x)) return(1) else return(0))
clinical$tumor_grade <- as.factor(clinical$tumor_grade)
clinical$vital_status <- as.factor(clinical$vital_status)
clinical$status <- lapply(clinical$vital_status, function(x) if (x ==
  "alive") return(0) else return(1))
clinical$smoking <- as.numeric(clinical$smoking)
```

```
# combine clinical stage
```

```

combineStage <- function(x) {
  if (!is.na(x)) {
    if (x == "Stage I" | x == "Stage IB" | x ==
        "Stage IA")
      return("Early Stage")
    if (x == "Stage II" | x == "Stage IIA" | x ==
        "Stage IIB")
      return("Early Stage")
    if (x == "Stage IIIA" | x == "Stage IIIB")
      return("Late Stage")
    if (x == "Stage IV")
      return("Late Stage")
  }
}

# combined histologic types

combinePath <- function(x) {
  if (!is.na(x)) {
    if (x == "Adenocarcinoma, NOS")
      return("Adenocarcinoma")
    if (x == "Mucinous adenocarcinoma")
      return("Adenocarcinoma")
    if (x == "Adenocarcinoma with mixed subtypes")
      return("Adenocarcinoma")
    if (x == "Papillary adenocarcinoma")
      return("Adenocarcinoma") else return("Non-Adenocarcinoma")
  }
}

# combine sub-stages into 1,2,3
clinical$ajcc_pathologic_stage_combined <- lapply(clinical$ajcc_pathologic_stage,
  combineStage)

clinical$ajcc_pathologic_stage_combined <- as.factor(as.character(clinical$ajcc_pathologic_stage_combined))

# combine histology types
clinical$primary_diagnosis_combined <- lapply(clinical$primary_diagnosis,
  combinePath)
clinical$primary_diagnosis_combined <- as.factor(unlist(clinical$primary_diagnosis_combined))

# get rid of NULL entries
clinical <- clinical[clinical$ajcc_pathologic_stage_combined !=
  "NULL", ]

# We already got rid of NULLs for
# ajcc_pathologic_stage. However, the factor
# levels still list 3 levels - early, late and
# NULL. To re-level, convert to character,

```

```

# re-convert to factor, and now only 2 levels.
clinical$ajcc_pathologic_stage_combined <- as.factor(as.character(clinical$ajcc_pathologic_stage_combined))

clinical <- clinical[!is.na(clinical$MK2.Expression),
]

# create censor variable for CoxPH
clinical$censor <- lapply(clinical$vital_status, function(x) if (x ==
  "Alive") return(0) else return(1))
clinical$censor <- as.numeric(clinical$censor)

# scale the MK2 variable
clinical$MK2.Expression <- clinical$MK2.Expression/1000
clinical$MK2.Expression.Manual <- clinical$MK2.Expression.Manual/1000
clinical$Casp3.Expression <- clinical$Casp3.Expression/1000
# create a simplified stage
combineStage.Simp <- function(x) {
  if (!is.na(x)) {
    if (x == "Stage I" | x == "Stage IB" | x ==
      "Stage IA")
      return("Stage 1")
    if (x == "Stage II" | x == "Stage IIA" | x ==
      "Stage IIB")
      return("Stage 2")
    if (x == "Stage IIIA" | x == "Stage IIIB")
      return("Stage 3")
    if (x == "Stage IV")
      return("Stage 4")
  }
}
clinical$ajcc_pathologic_stage_combined.2 <- lapply(clinical$ajcc_pathologic_stage,
  combineStage.Simp)
clinical$ajcc_pathologic_stage_combined.2 <- as.factor(unlist(clinical$ajcc_pathologic_stage_combined.2))

write.table(as.matrix(clinical), file = "../rawdata/tcga-luad-clinical.processed.csv",
  sep = ",", quote = FALSE, col.names = TRUE)

```

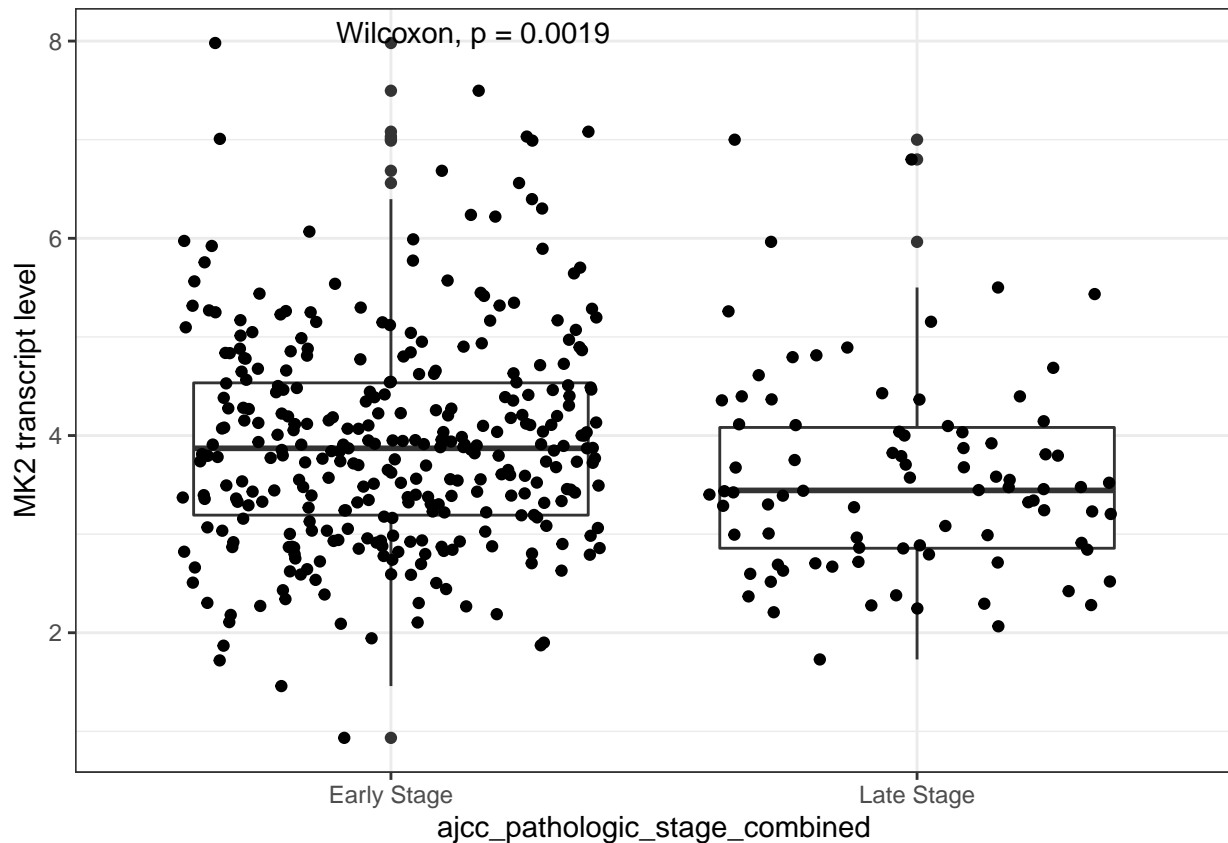
Exploratory graphs

Plot differences in transcript level by tumor histology and cancer stage

```

# plot MK2 differences for adenocarcinoma only
ggplot(clinical[clinical$primary_diagnosis_combined ==
  "Adenocarcinoma", ], aes(x = ajcc_pathologic_stage_combined,
  y = MK2.Expression, )) + geom_boxplot() + geom_point(position = position_jitter()) +
  stat_compare_means() + theme_bw() + ylab("MK2 transcript level")

```

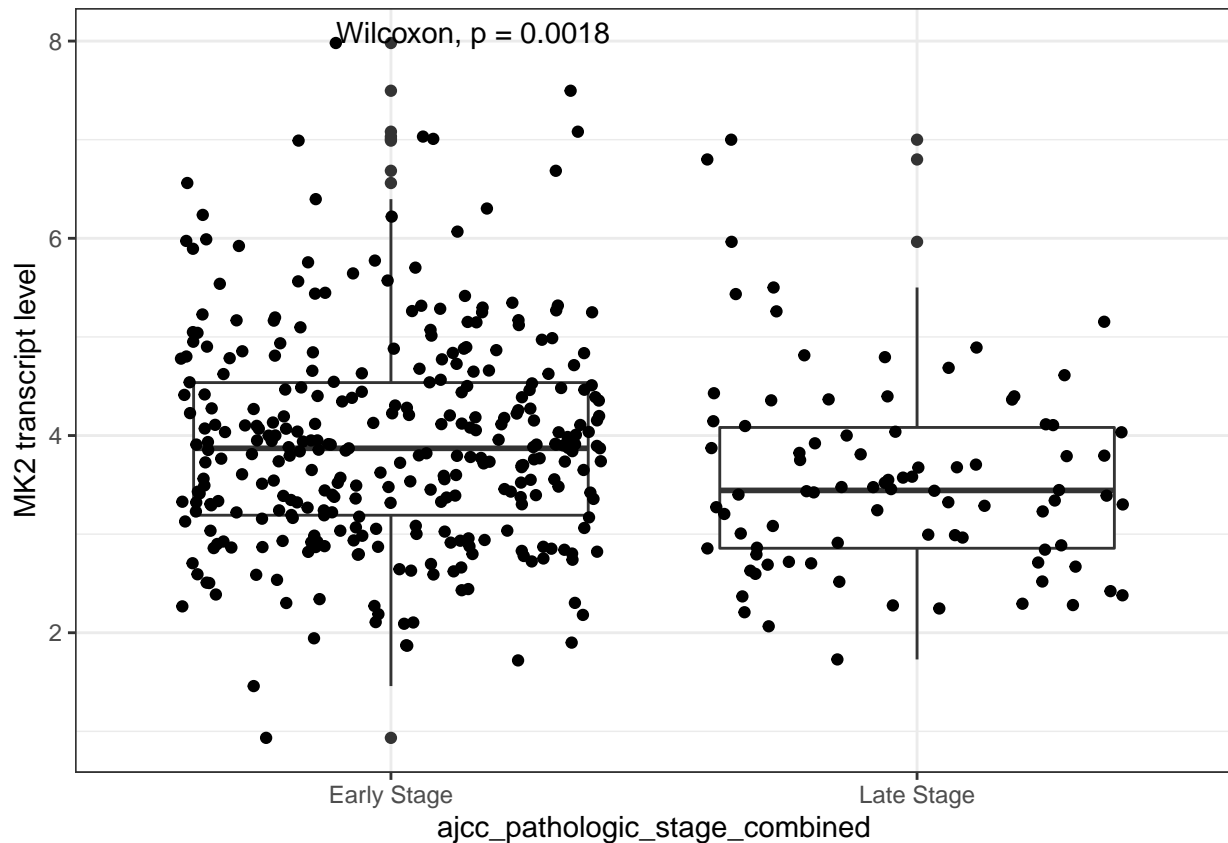



```
ggplot(clinical[clinical$primary_diagnosis_combined ==
  "Adenocarcinoma", ], aes(x = ajcc_pathologic_stage_combined,
  y = MK2.Expression.Manual, )) + geom_boxplot() +
  geom_point(position = position_jitter()) + stat_compare_means() +
  theme_bw() + ylab("MK2 transcript level")
```

```
## Warning: Removed 1 rows containing non-finite values (stat_boxplot).
```

```
## Warning: Removed 1 rows containing non-finite values (stat_compare_means).
```

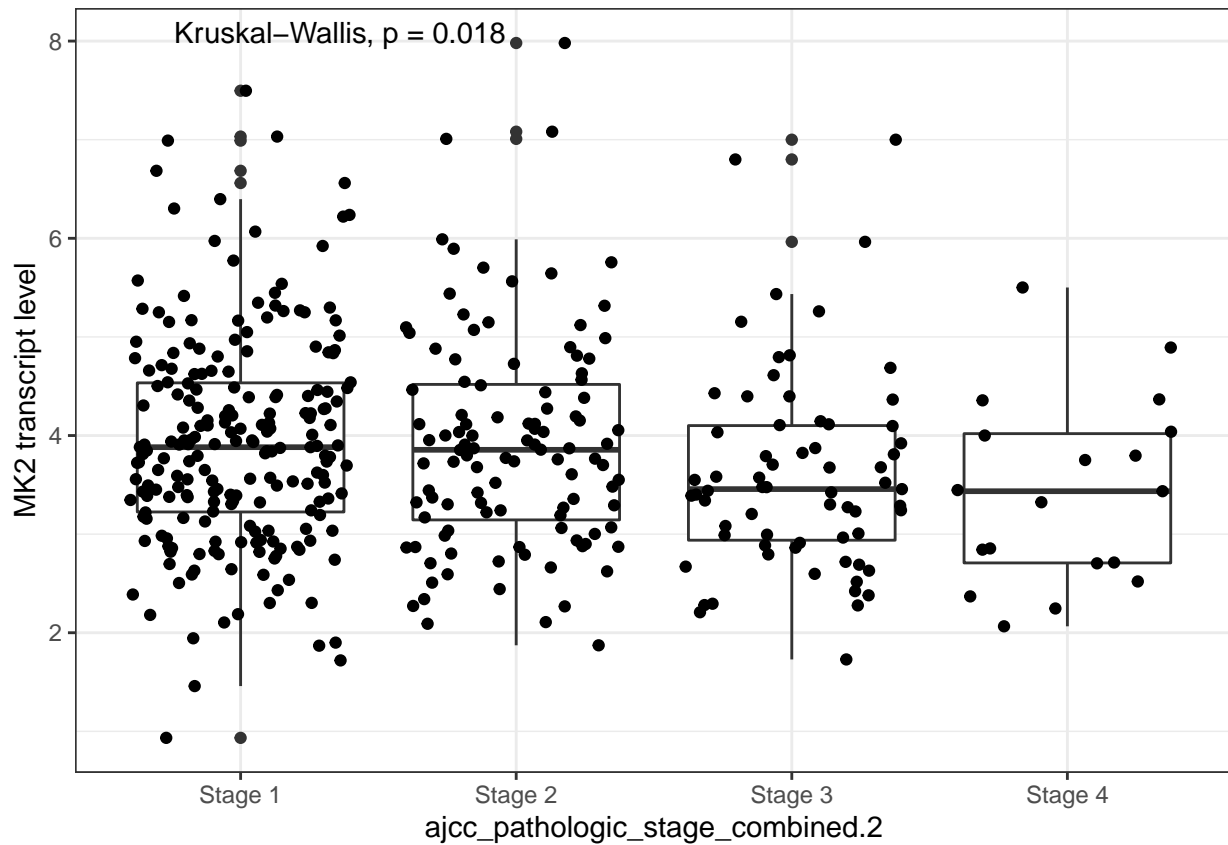
```
## Warning: Removed 1 rows containing missing values (geom_point).
```



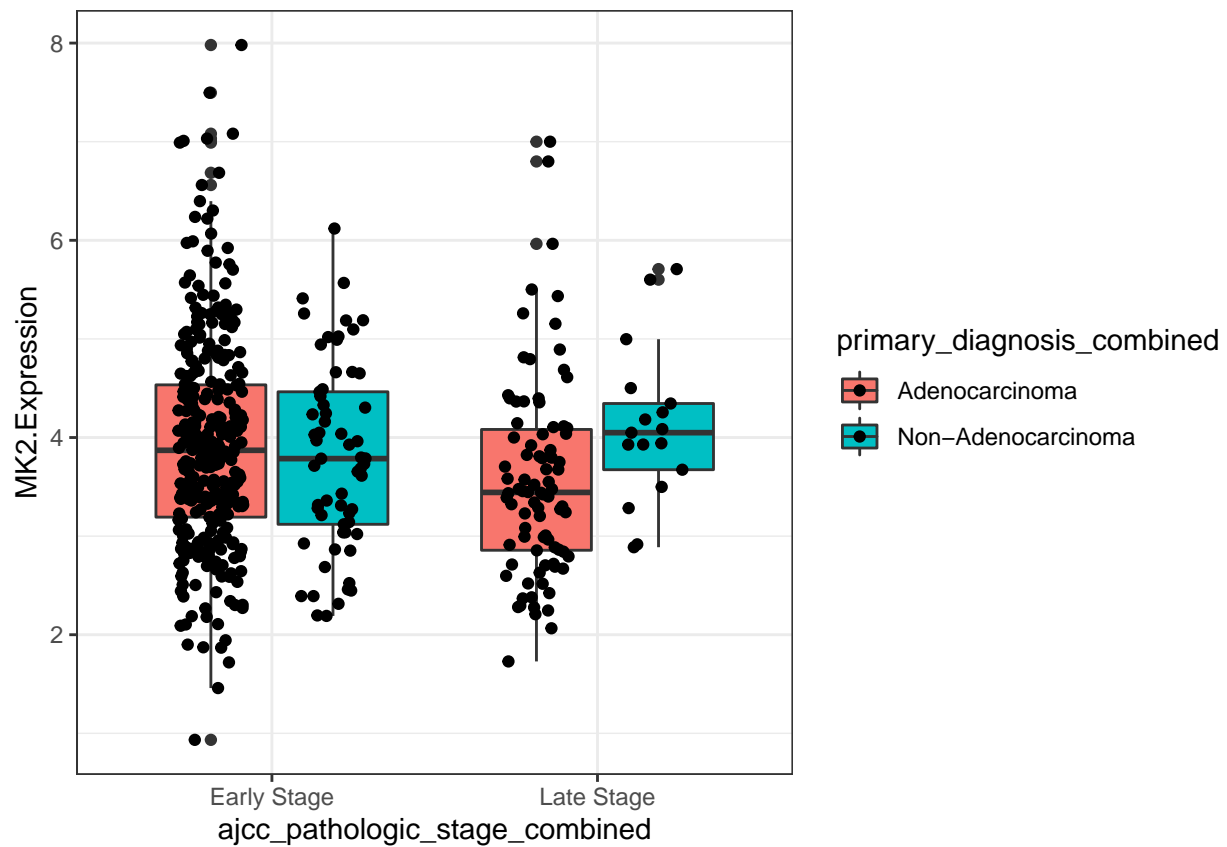
This is a key difference: MK2 is clearly associated with tumor stage. Early stage tumors have more MK2 transcript than late stage tumors. Thus, going forward, we will stratify our analyses by Tumor stage.

Here, we look at tumor differences more granularly (stage 1 vs. 2 vs. 3 vs. 4) and see a linear relationship - as stage increases, MK2 exp decreases. We also see that the LUAD cases that aren't really "AD" (non-Adenocarcinoma) don't share this relationship between Early/Late stage and MK2 transcript levels. So, when we prep the data for model work (below) we will restrict to Adenocarcinoma cases only

```
ggplot(clinical[clinical$primary_diagnosis_combined ==
  "Adenocarcinoma", ], aes(x = ajcc_pathologic_stage_combined.2,
  y = MK2.Expression, )) + geom_boxplot() + geom_point(position = position_jitter()) +
  stat_compare_means() + theme_bw() + ylab("MK2 transcript level")
```



```
ggplot(clinical, aes(x = ajcc_pathologic_stage_combined,
  y = MK2.Expression, fill = primary_diagnosis_combined)) +
  geom_boxplot() + geom_point(position = position_jitterdodge()) +
  theme_bw()
```

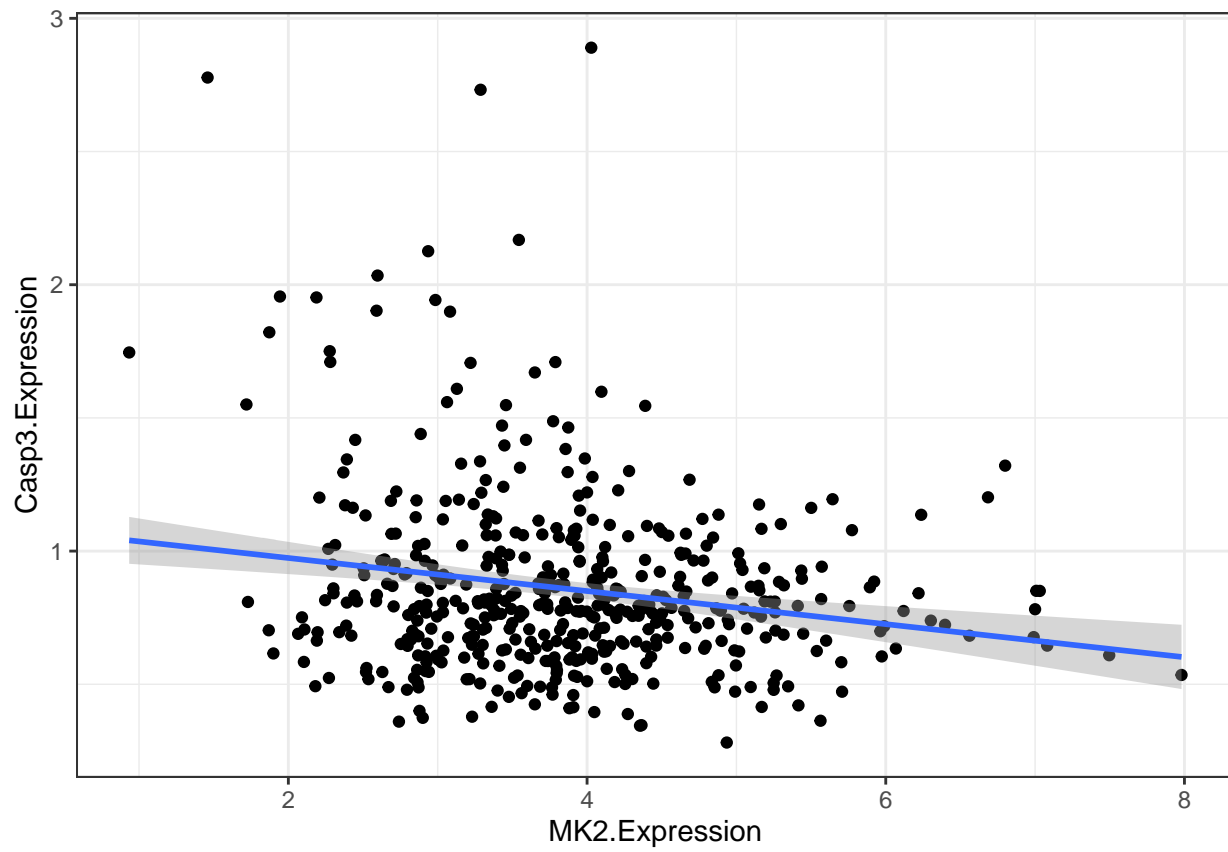


Explore the relationship between MK2 and Casp3

Here we look at the correlation between MK2 and Casp3 transcript levels.

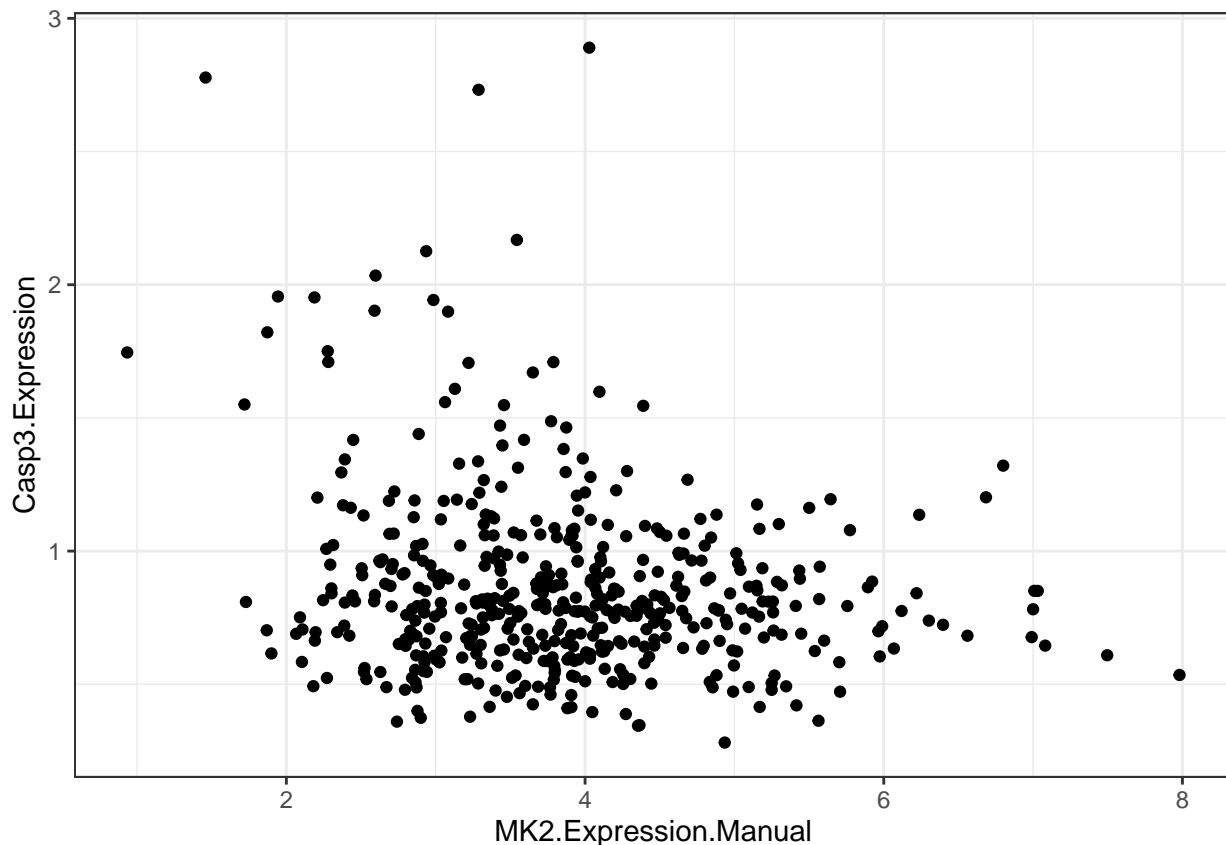
```
ggplot(clinical, aes(x = MK2.Expression, y = Casp3.Expression)) +  
  geom_point() + theme_bw() + geom_smooth(method = "lm")
```

```
## `geom_smooth()` using formula 'y ~ x'
```



```
ggplot(clinical, aes(x = MK2.Expression.Manual, y = Casp3.Expression)) +  
  geom_point() + theme_bw()
```

```
## Warning: Removed 2 rows containing missing values (geom_point).
```



We are technically able to “fit” a line through this data, but the relationship (i.e slope) is not particularly impressive. The raw data does not suggest any sort of linear relationship between Casp3 mRNA and MK2 mRNA levels.

Defining MK2 thresholds for “high” and “low” - Top 1/3 vs Bottom 2/3

Mechanistically, we’re more interested in whether high vs. low MK2 transcript levels impact outcome. For this, we first categorize MK2 transcript levels based on quantiles and only take the lower and upper third in our groups for comparison. The MK2_Expression_topQ variable essentially captures this. If the transcript is in the top 33% then we return 1, everything else (i.e the bottom 66%) returns 0.

```
# classify the MK2 expression variable as either
# top or bottom quantile, based on quantile
# definitions

bottom_quant <- 0.33
top_quant <- 0.66

clinical$MK2_Expression_bottomQ <- lapply(clinical$MK2.Expression,
  function(x) if (x < quantile(clinical$MK2.Expression,
    bottom_quant)) return(1) else return(0))

clinical$MK2_Expression_topQ <- lapply(clinical$MK2.Expression,
  function(x) if (x > quantile(clinical$MK2.Expression,
    top_quant)) return(1) else return(0))
```

Changes in proportion of people who died across quantiles of MK2

Prep data for using MK2 quantiles

The (poorly named) MK2_CoxPH_quantile df is the df we're using for all of our modeling purposes. This was originally used just for CoxPH, hence the name. We restrict this dataset to only those with follow-up available, and Adenocarcinomas only within LUAD. Additionally, we perform some more cleanup/create new variables as shown below.

```
# quantile and logistic regression

# this is the new data frame for this new round
# of calculations.
MK2_CoxPH_quantile <- clinical[!is.na(clinical$time) &
  clinical$primary_diagnosis_combined == "Adenocarcinoma",
  ]

# this provides the variable for R censoring at
# one year and the outcome variable for logit
# regression.
MK2_CoxPH_quantile$death_at_1year <- mapply(function(dead,
  time) if (dead == 1 & time < 366) return(1) else if (dead ==
  0 & time > 366) return(0) else if (dead == 0 &
  time < 366) return(0) else if (dead == 1 & time >
  366) return(0), MK2_CoxPH_quantile$censor, as.numeric(MK2_CoxPH_quantile$time))

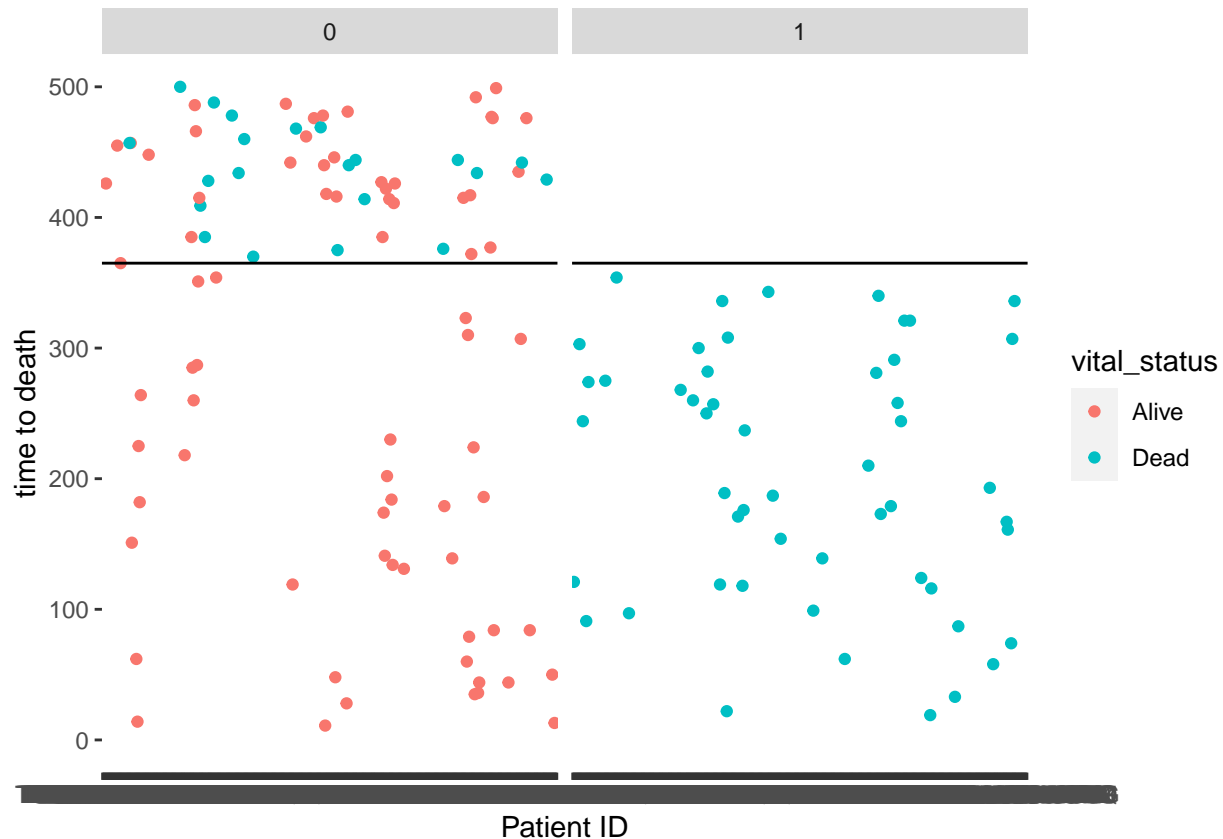
MK2_CoxPH_quantile$death_at_2year <- mapply(function(dead,
  time) if (dead == 1 & time < 366 * 2) return(1) else if (dead ==
  0 & time > 366 * 2) return(0) else if (dead ==
  0 & time < 366 * 2) return(0) else if (dead ==
  1 & time > 366 * 2) return(0), MK2_CoxPH_quantile$censor,
  as.numeric(MK2_CoxPH_quantile$time))

MK2_CoxPH_quantile$death_at_3year <- mapply(function(dead,
  time) if (dead == 1 & time < 366 * 3) return(1) else if (dead ==
  0 & time > 366 * 3) return(0) else if (dead ==
  0 & time < 366 * 3) return(0) else if (dead ==
  1 & time > 366 * 3) return(0), MK2_CoxPH_quantile$censor,
  as.numeric(MK2_CoxPH_quantile$time))

# this piece of code above is critical - so let's
# double check and make sure everything got done
# right

ggplot(MK2_CoxPH_quantile, aes(x = submitter_id, y = time,
  color = vital_status)) + geom_point() + facet_wrap(~death_at_1year) +
  ylim(0, 500) + xlab("Patient ID") + ylab("time to death") +
  geom_hline(yintercept = 365)
```

```
## Warning: Removed 258 rows containing missing values (geom_point).
```



Interpretation off the plot above. So, we have colored each patient's vital status by alive or dead, and we have faceted by whether they were assigned as "Death in one year" or not. So, if our code worked, all patients who died in less than one year (blue color; $y < 365$) should be classified as dead at one year ("1"). All patients who were alive in that time period should be classified as red and in the "0" column. After 365 days, everyone should have gotten a classification of "0" for death at one year, since they are all by definition still alive at 1 year, regardless of subsequent alive or dead status (color should mix, all colors in the "0" column). And that's what we see...

```
# lastly, going forward, we wish to show time in
# months rather than days, so we'll divide by 30
MK2_CoxPH_quantile$time <- MK2_CoxPH_quantile$time/30
```

Proportion of deaths by MK2 quantiles

```
# Now, moving on.

MK2_CoxPH_quantile$MK2_quantile <- ntile(MK2_CoxPH_quantile$MK2.Expression,
3)
MK2_CoxPH_quantile$Casp3_quantile <- ntile(MK2_CoxPH_quantile$Casp3.Expression,
3)
a <- aggregate(death_at_1year ~ MK2_quantile + ajcc_pathologic_stage_combined,
FUN = sum, data = MK2_CoxPH_quantile)
b <- aggregate(death_at_1year ~ MK2_quantile + ajcc_pathologic_stage_combined,
FUN = length, data = MK2_CoxPH_quantile)

quantile_df <- as.data.frame(a$MK2_quantile)
quantile_df$stage <- as.data.frame(a$ajcc_pathologic_stage_combined)
```



```

quantile_df$prop <- as.data.frame(unlist(a$death_at_1year/b$death_at_1year))
quantile_df$prop <- unlist(quantile_df$prop)
colnames(quantile_df) <- c("quantile", "stage", "prop")
p1 <- ggplot(quantile_df[quantile_df$stage == "Early Stage",
], aes(x = quantile, y = prop)) + geom_point() +
  geom_smooth(method = "lm", se = FALSE) + stat_poly_eq(formula = "y~x",
  aes(label = paste(..eq.label.., ..rr.label.., sep = "~~~")),
  parse = TRUE) + theme_bw() + ylim(0, 0.5)

p2 <- ggplot(quantile_df[quantile_df$stage == "Late Stage",
], aes(x = quantile, y = prop)) + geom_point() +
  geom_smooth(method = "lm", se = FALSE) + stat_poly_eq(formula = "y~x",
  aes(label = paste(..eq.label.., ..rr.label.., sep = "~~~")),
  parse = TRUE) + theme_bw() + ylim(0, 0.5)

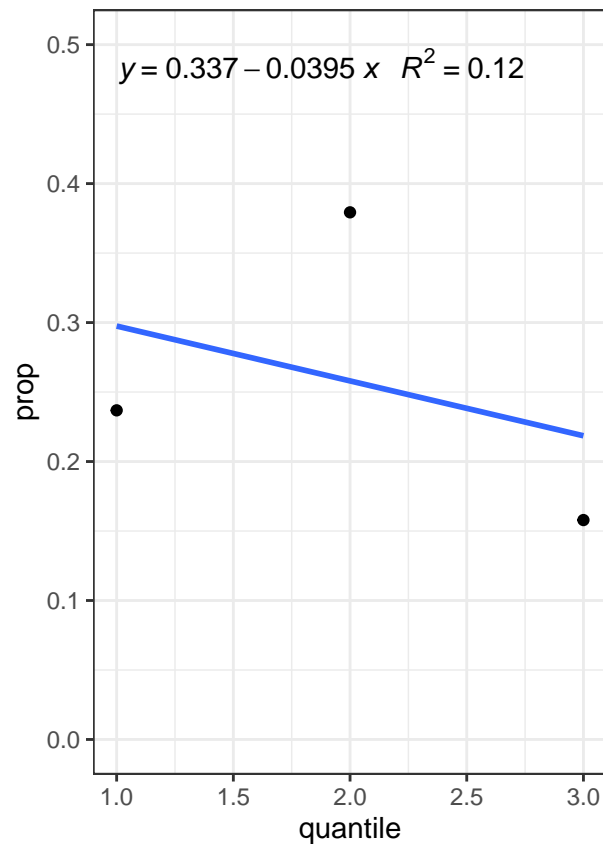
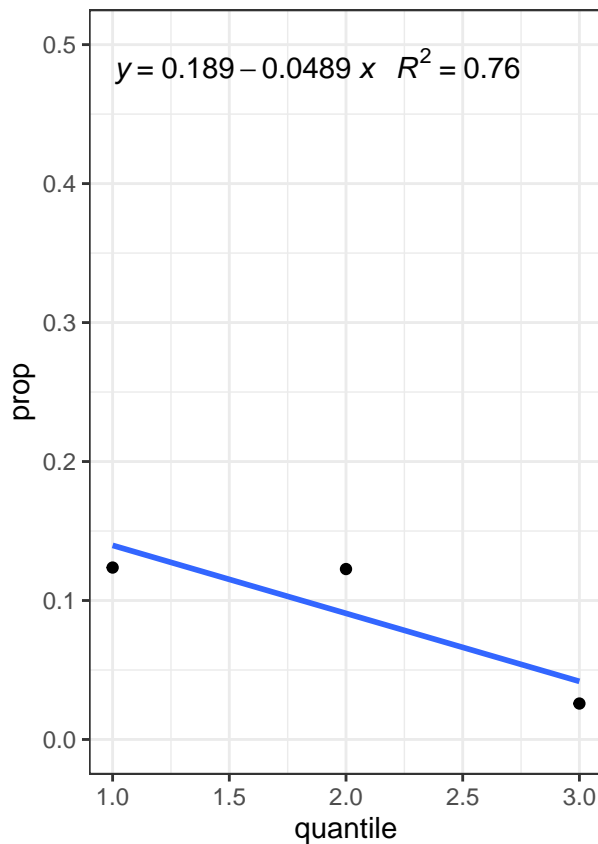
# This graph shows the relationship between MK2
# quantile and proportion of patients who died in
# Early and Late Stage cancers. The relationship
# between MK2 quantile and Proportion of those
# patients who died (denominator: all patients in
# that quantile)
gridExtra::grid.arrange(p1, p2, ncol = 2)

```

```

## `geom_smooth()` using formula 'y ~ x'
## `geom_smooth()` using formula 'y ~ x'

```



Logistic regression modeling death at 1 year ~ MK2+covariates

Data clean-up

```
MK2_CoxPH_quantile$death_at_1year <- as.factor(MK2_CoxPH_quantile$death_at_1year)
# logistic regression and ROC
MK2_CoxPH_quantile$ajcc_pathologic_stage_early_stage <- unlist(MK2_CoxPH_quantile$ajcc_pathologic_stage)
MK2_CoxPH_quantile$MK2_quantile <- as.factor(MK2_CoxPH_quantile$MK2_quantile)
MK2_CoxPH_quantile$MK2_Expression_topQ <- as.factor(unlist(MK2_CoxPH_quantile$MK2_Expression_topQ))
```

Logistic regression for death at 1 year, stratified by early/late, MK2 as a continuous variable

Univariate regression

So, here we run a univariate analysis, with MK2 as a continuous variable, stratified by stage. We see a signal in early but not late stage disease.

```
# LR stratified by Early/Late
summary(glm(death_at_1year ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage == "Early Stage", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined == "Early Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.7132  -0.4651  -0.4104  -0.3454   2.4882
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   -0.8740     0.7590  -1.152   0.2495
## MK2.Expression -0.3910     0.2041  -1.916   0.0554 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 189.71  on 318  degrees of freedom
## Residual deviance: 185.76  on 317  degrees of freedom
## AIC: 189.76
##
## Number of Fisher Scoring iterations: 5
summary(glm(death_at_1year ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage == "Late Stage", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined == "Late Stage", ])
##
##
```

```
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.8449  -0.8022  -0.7821   1.5805   1.6744
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   -0.69350    0.91543  -0.758    0.449
## MK2.Expression -0.08846    0.24994  -0.354    0.723
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 99.880  on 85  degrees of freedom
## Residual deviance: 99.753  on 84  degrees of freedom
## AIC: 103.75
##
## Number of Fisher Scoring iterations: 4
```

In this univariate regression, there is a suggestion of a negative association between MK2 (treated as a continuous variable) and death at one year, in the early stage patients. That is, higher MK2 = lower odds of death at one year. the CIs are imprecise, and thus the p values is 0.06. We do not see this in the late-stage patients.

Multivariate regression

Now, we add covariates to the univariate model above. Again, because there is a baseline imbalance in transcript level between early and late stage, we are stratifying rather than adjusting. MK2 here remains a continuous variable, for now.

```
summary(glm(death_at_1year ~ MK2.Expression + smoking +
  age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combi
  "Early Stage", ], family = binomial(link = "logit")))
```

```
##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression + smoking + age_at_diagnosis +
##      gender, family = binomial(link = "logit"), data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pat
##      "Early Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.7519  -0.4579  -0.3556  -0.2933   2.5860
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   -1.01796    1.63391  -0.623    0.5333
## MK2.Expression  -0.32422    0.21042  -1.541    0.1234
## smoking         0.13949    0.49344   0.283    0.7774
## age_at_diagnosis -0.01141    0.02121  -0.538    0.5906
## gendermale      0.82992    0.43657   1.901    0.0573 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 174.15  on 311  degrees of freedom
```

```
## Residual deviance: 166.81 on 307 degrees of freedom
## (7 observations deleted due to missingness)
## AIC: 176.81
##
## Number of Fisher Scoring iterations: 5
summary(glm(death_at_1year ~ MK2.Expression + smoking +
  age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combi
    "Late Stage", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression + smoking + age_at_diagnosis +
##     gender, family = binomial(link = "logit"), data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pat
##     "Late Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.1177  -0.8219  -0.7066   1.4406   1.9247
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   -2.64865    1.98157  -1.337   0.181
## MK2.Expression -0.06162    0.26183  -0.235   0.814
## smoking       -0.42478    0.51632  -0.823   0.411
## age_at_diagnosis 0.03072    0.02594   1.184   0.236
## gendermale      0.33048    0.49610   0.666   0.505
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 99.880 on 85 degrees of freedom
## Residual deviance: 97.263 on 81 degrees of freedom
## AIC: 107.26
##
## Number of Fisher Scoring iterations: 4
```

Even with adjustment, there is again a suggestion of a negative association between MK2 transcript level and death at one year, with a p value of 0.1. There is a strong positive association between male gender and death at one year.

Model 2 in paper: Logistic regression for death at 1 year, stratified by early/late, but using Top 1/3 MK2 expression (dichotomous MK2 variable)

Now, we are interested in using MK2 as a dichotomous variable - high (top 1/3) or low (bottom 2/3). Clinically, with an appropriate reference cohort, this might be more translatable. Thus, we modeled death at one year, as a logistic regression, but now with MK2 as a dichotomous variable. The univariate analysis was significant, so here we run the multi-variate regression.

Model 2 - Univariate Regression

```
# LR stratified by Early/Late
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pat
  "Early Stage", ], family = binomial(link = "logit")))

##
```

```
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##      "Early Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.5127  -0.5127  -0.5127  -0.2289   2.7037
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -1.9629     0.2136  -9.190 < 2e-16 ***
## MK2_Expression_topQ1 -1.6659     0.6227  -2.675  0.00747 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##    Null deviance: 189.71  on 318  degrees of freedom
## Residual deviance: 179.35  on 317  degrees of freedom
## AIC: 183.35
##
## Number of Fisher Scoring iterations: 6
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
  "Late Stage", ], family = binomial(link = "logit")))
##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##      "Late Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.8421  -0.8421  -0.8421   1.5550   1.9214
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -0.8544     0.2670  -3.200  0.00137 **
## MK2_Expression_topQ1 -0.8196     0.6835  -1.199  0.23047
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##    Null deviance: 99.88  on 85  degrees of freedom
## Residual deviance: 98.26  on 84  degrees of freedom
## AIC: 102.26
##
## Number of Fisher Scoring iterations: 4
```

We see highly significant association between high MK2 expression (as a dichotomous variable) and death at 1 year, in early but not late stage cancer.

Model 2: Multivariate regression

So, now we will run 2 models. MK2 as dichotomous variable, with covariates, stratified by stage with the outcome of death at 1 year, as well as death at 2 years.

Here we first look at Early Stage only

```
logreg_topq1_early <- glm(death_at_1year ~ MK2_Expression_topQ +
  smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_s
    "Early Stage", ], family = binomial(link = "logit"))

summary(logreg_topq1_early)
```

```
##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ + smoking +
##      age_at_diagnosis + gender, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
##      "Early Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.7059  -0.4897  -0.3689  -0.2019   2.8548
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -1.62809    1.47184  -1.106   0.2687
## MK2_Expression_topQ1 -1.50091    0.63011  -2.382   0.0172 *
## smoking           0.16785    0.49734   0.337   0.7357
## age_at_diagnosis  -0.01533    0.02089  -0.734   0.4629
## gendermale        0.78866    0.43959   1.794   0.0728 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 174.15  on 311  degrees of freedom
## Residual deviance: 161.58  on 307  degrees of freedom
##      (7 observations deleted due to missingness)
## AIC: 171.58
##
## Number of Fisher Scoring iterations: 6
```

```
exp(logreg_topq1_early$coefficients)
```

```
##              (Intercept) MK2_Expression_topQ1      smoking
##              0.1963034      0.2229281      1.1827597
##      age_at_diagnosis      gendermale
##              0.9847829      2.2004382
```

```
exp(confint.lm(logreg_topq1_early))
```

```
##              2.5 %    97.5 %
## (Intercept)    0.01084259 3.5540422
## MK2_Expression_topQ1 0.06451920 0.7702661
## smoking        0.44451696 3.1470578
```

```
## age_at_diagnosis      0.94512249 1.0261075
## gendermale           0.92651066 5.2259825

logreg_topq1_early_2y <- glm(death_at_2year ~ MK2_Expression_topQ +
  smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_s
    "Early Stage", ], family = binomial(link = "logit"))

summary(logreg_topq1_early_2y)

##
## Call:
## glm(formula = death_at_2year ~ MK2_Expression_topQ + smoking +
##      age_at_diagnosis + gender, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
##      "Early Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.8105  -0.6614  -0.5588  -0.4594   2.2780
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -2.184387   1.112186  -1.964   0.0495 *
## MK2_Expression_topQ1 -0.695725   0.345540  -2.013   0.0441 *
## smoking         0.373370   0.360243   1.036   0.3000
## age_at_diagnosis  0.005672   0.015531   0.365   0.7150
## gendermale      0.389860   0.304645   1.280   0.2006
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 287.50  on 311  degrees of freedom
## Residual deviance: 279.64  on 307  degrees of freedom
##      (7 observations deleted due to missingness)
## AIC: 289.64
##
## Number of Fisher Scoring iterations: 4

exp(logreg_topq1_early_2y$coefficients)

##              (Intercept) MK2_Expression_topQ1      smoking
##              0.1125467      0.4987126      1.4526216
##      age_at_diagnosis      gendermale
##              1.0056877      1.4767739

exp(confint.lm(logreg_topq1_early_2y))

##              2.5 %    97.5 %
## (Intercept)    0.01261513 1.0040923
## MK2_Expression_topQ1 0.25267507 0.9843244
## smoking         0.71498909 2.9512471
## age_at_diagnosis  0.97541719 1.0368977
## gendermale      0.81091195 2.6893934
```

Now we look at late stage

```
logreg_topq1_late <- glm(death_at_1year ~ MK2_Expression_topQ +
  smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_s
    "Late Stage", ], family = binomial(link = "logit"))

summary(logreg_topq1_late)
```

```
##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ + smoking +
##   age_at_diagnosis + gender, family = binomial(link = "logit"),
##   data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
##     "Late Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.1548  -0.8278  -0.6714   1.3897   2.1483
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -2.63173    1.80098  -1.461   0.144
## MK2_Expression_topQ1 -0.72657    0.69371  -1.047   0.295
## smoking        -0.38921    0.51949  -0.749   0.454
## age_at_diagnosis  0.02873    0.02663   1.079   0.281
## gendermale       0.33859    0.49884   0.679   0.497
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 99.880  on 85  degrees of freedom
## Residual deviance: 96.109  on 81  degrees of freedom
## AIC: 106.11
##
## Number of Fisher Scoring iterations: 4
```

```
exp(logreg_topq1_late$coefficients)
```

```
##           (Intercept) MK2_Expression_topQ1           smoking
##           0.07195393           0.48356425           0.67759368
##           age_at_diagnosis           gendermale
##           1.02914235           1.40296293
```

```
exp(confint.lm(logreg_topq1_late))
```

```
##              2.5 %   97.5 %
## (Intercept)  0.001998996 2.589984
## MK2_Expression_topQ1 0.121623019 1.922616
## smoking        0.241030494 1.904876
## age_at_diagnosis 0.976038424 1.085136
## gendermale      0.519985525 3.785307
```

```
logreg_topq1_late_2y <- glm(death_at_2year ~ MK2_Expression_topQ +
  smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_s
    "Late Stage", ], family = binomial(link = "logit"))

summary(logreg_topq1_late_2y)
```



```
##
## Call:
## glm(formula = death_at_2year ~ MK2_Expression_topQ + smoking +
##      age_at_diagnosis + gender, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##      "Late Stage", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.3149  -1.0925  -0.9229   1.2055   1.5937
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -1.09315    1.47675  -0.740    0.459
## MK2_Expression_topQ1 -0.37233    0.54314  -0.686    0.493
## smoking        -0.25976    0.46684  -0.556    0.578
## age_at_diagnosis    0.01838    0.02209   0.832    0.405
## gendermale       -0.02267    0.43926  -0.052    0.959
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 118.48  on 85  degrees of freedom
## Residual deviance: 116.76  on 81  degrees of freedom
## AIC: 126.76
##
## Number of Fisher Scoring iterations: 4
```

```
exp(logreg_topq1_late_2y$coefficients)
```

```
##              (Intercept) MK2_Expression_topQ1          smoking
##              0.3351602          0.6891295          0.7712360
##      age_at_diagnosis          gendermale
##              1.0185536          0.9775869
```

```
exp(confint.lm(logreg_topq1_late_2y))
```

```
##              2.5 %   97.5 %
## (Intercept)    0.01774918 6.328876
## MK2_Expression_topQ1 0.23386708 2.030639
## smoking        0.30464012 1.952484
## age_at_diagnosis 0.97476408 1.064310
## gendermale      0.40793462 2.342719
```

Model 2: MK2 multivariate model with Casp3

Now we run the multi-variate model again but with casp3 expression

```
summary(glm(death_at_2year ~ MK2_Expression_topQ +
  smoking + age_at_diagnosis + gender + Casp3.Expression,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Early Stage", ], family = binomial(link = "logit")))
```

```
##
## Call:
## glm(formula = death_at_2year ~ MK2_Expression_topQ + smoking +
##      age_at_diagnosis + gender + Casp3.Expression, family = binomial(link = "logit"),
```

```
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##      "Early Stage", ]
##
## Deviance Residuals:
##      Min        1Q    Median        3Q        Max
## -0.8350  -0.6598  -0.5608  -0.4615   2.2744
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -2.255784    1.179108  -1.913   0.0557 .
## MK2_Expression_topQ1 -0.688504    0.347914  -1.979   0.0478 *
## smoking         0.373941    0.360311   1.038   0.2994
## age_at_diagnosis  0.005639    0.015532   0.363   0.7166
## gendermale       0.392415    0.305005   1.287   0.1982
## Casp3.Expression  0.080910    0.443139   0.183   0.8551
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 287.50  on 311  degrees of freedom
## Residual deviance: 279.61  on 306  degrees of freedom
##      (7 observations deleted due to missingness)
## AIC: 291.61
##
## Number of Fisher Scoring iterations: 4
```

Addition of Casp3 does not change the effect of MK2_expression on death at 2 year in this model.

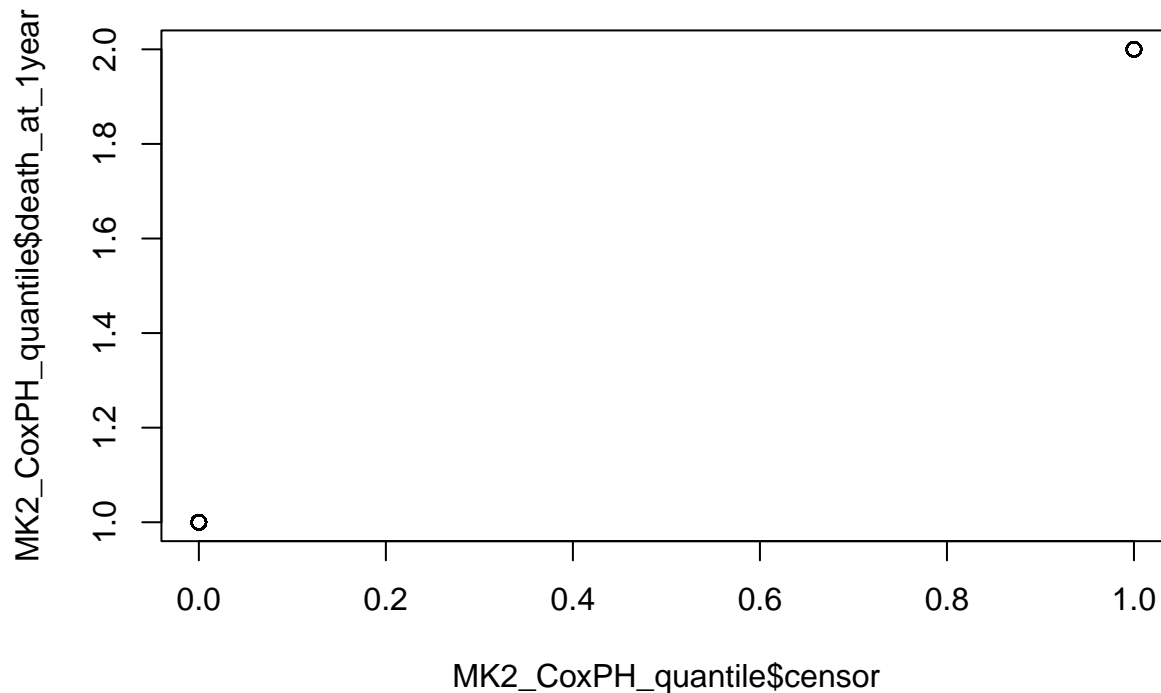
Cox Proportional Hazards modeling, with right censoring at 1 year.

Data Cleaning

```
# data cleaning prior to Cox PH model
MK2_CoxPH_quantile$MK2_Expression_topQ <- as.factor(unlist(MK2_CoxPH_quantile$MK2_Expression_topQ))
MK2_CoxPH_quantile$gender <- as.factor(MK2_CoxPH_quantile$gender)

# be careful here. as.numeric alone will return
# the factor level, not the value!
MK2_CoxPH_quantile$censor <- as.numeric(as.character(MK2_CoxPH_quantile$death_at_1year))
MK2_CoxPH_quantile$censor.2y <- as.numeric(as.character(MK2_CoxPH_quantile$death_at_2year))
MK2_CoxPH_quantile$censor.3y <- as.numeric(as.character(MK2_CoxPH_quantile$death_at_3year))

# there should be no discrepancy - only two dots
# should show up
plot(MK2_CoxPH_quantile$censor, MK2_CoxPH_quantile$death_at_1year)
```



Cox PH model R censored at one year, stratified by early/late, MK2 as a continuous variable

Model 1: MK2 only

```
# univariate Cox PH
summary(coxph(Surv(time, censor) ~ MK2.Expression,
  id = submitter_id, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Early Stage", ]))

## Call:
## coxph(formula = Surv(time, censor) ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined == "Early Stage", ], id = submitter_id)
##
##      n= 319, number of events= 28
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## MK2.Expression -0.3488    0.7055   0.1930 -1.808   0.0707 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##              exp(coef) exp(-coef) lower .95 upper .95
## MK2.Expression    0.7055      1.417   0.4833    1.03
##
## Concordance= 0.606 (se = 0.046 )
## Likelihood ratio test= 3.45 on 1 df,  p=0.06
## Wald test               = 3.27 on 1 df,  p=0.07
## Score (logrank) test = 3.25 on 1 df,  p=0.07

# Multi-variate Cox PH
MK2_coxph_model_tb_1y_early <- coxph(Surv(time, censor) ~
  MK2.Expression + gender + smoking + age_at_diagnosis,
```

```

    id = submitter_id, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
      "Early Stage", ])

MK2_coxph_model_tb_1y_late <- coxph(Surv(time, censor) ~
  MK2.Expression + gender + smoking + age_at_diagnosis,
  id = submitter_id, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
    "Late Stage", ])

# PH testing
cox.zph(MK2_coxph_model_tb_1y_early)

##               chisq df    p
## MK2.Expression   1.522  1 0.22
## gender           0.419  1 0.52
## smoking          3.276  1 0.07
## age_at_diagnosis 1.616  1 0.20
## GLOBAL           6.371  4 0.17
cox.zph(MK2_coxph_model_tb_1y_late)

##               chisq df    p
## MK2.Expression   0.3655  1 0.55
## gender           1.2832  1 0.26
## smoking          0.2681  1 0.60
## age_at_diagnosis 0.0349  1 0.85
## GLOBAL           1.9921  4 0.74

# results
summary(MK2_coxph_model_tb_1y_early)

## Call:
## coxph(formula = Surv(time, censor) ~ MK2.Expression + gender +
##       smoking + age_at_diagnosis, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
##       "Early Stage", ], id = submitter_id)
##
##      n= 312, number of events= 25
##      (7 observations deleted due to missingness)
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2.Expression  -0.27510   0.75949  0.19775 -1.391   0.1642
## gendermale       0.83411   2.30276  0.41878  1.992   0.0464 *
## smoking          0.06909   1.07153  0.47106  0.147   0.8834
## age_at_diagnosis -0.01332   0.98676  0.02050 -0.650   0.5158
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2.Expression      0.7595      1.3167    0.5155    1.119
## gendermale          2.3028      0.4343    1.0134    5.233
## smoking             1.0715      0.9332    0.4256    2.698
## age_at_diagnosis    0.9868      1.0134    0.9479    1.027
##
## Concordance= 0.644 (se = 0.058 )
## Likelihood ratio test= 7.33 on 4 df,  p=0.1

```

```
## Wald test          = 7.09  on 4 df,   p=0.1
## Score (logrank) test = 7.36  on 4 df,   p=0.1

summary(MK2_coxph_model_tb_1y_late)

## Call:
## coxph(formula = Surv(time, censor) ~ MK2.Expression + gender +
##       smoking + age_at_diagnosis, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##       "Late Stage", ], id = submitter_id)
##
## n= 86, number of events= 23
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2.Expression -0.07991   0.92320  0.21887 -0.365   0.715
## gendermale      0.20038   1.22187  0.41969  0.477   0.633
## smoking         -0.24635   0.78165  0.43026 -0.573   0.567
## age_at_diagnosis 0.02883   1.02925  0.02252  1.280   0.201
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2.Expression      0.9232      1.0832      0.6012      1.418
## gendermale           1.2219      0.8184      0.5368      2.781
## smoking              0.7817      1.2793      0.3363      1.817
## age_at_diagnosis     1.0293      0.9716      0.9848      1.076
##
## Concordance= 0.601 (se = 0.058 )
## Likelihood ratio test= 2.58  on 4 df,   p=0.6
## Wald test            = 2.39  on 4 df,   p=0.7
## Score (logrank) test = 2.43  on 4 df,   p=0.7
```

So here, we can see that in multi-variate Cox PH model, we see a signal for MK2 and improved survival, but only in the early stage tumors.

MK2 with casp 3

```
summary(coxph(Surv(time, censor) ~ MK2.Expression +
  gender + smoking + age_at_diagnosis + Casp3.Expression,
  id = submitter_id, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
  "Early Stage", ]))

## Call:
## coxph(formula = Surv(time, censor) ~ MK2.Expression + gender +
##       smoking + age_at_diagnosis + Casp3.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##       "Early Stage", ], id = submitter_id)
##
## n= 312, number of events= 25
## (7 observations deleted due to missingness)
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2.Expression -0.25541   0.77460  0.20031 -1.275   0.2023
## gendermale      0.84606   2.33045  0.41925  2.018   0.0436 *
## smoking         0.07316   1.07591  0.47159  0.155   0.8767
## age_at_diagnosis -0.01301   0.98707  0.02050 -0.635   0.5256
## Casp3.Expression 0.30076   1.35089  0.58307  0.516   0.6060
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2.Expression    0.7746    1.2910    0.5231    1.147
## gendermale        2.3305    0.4291    1.0247    5.300
## smoking           1.0759    0.9294    0.4269    2.711
## age_at_diagnosis  0.9871    1.0131    0.9482    1.028
## Casp3.Expression  1.3509    0.7403    0.4308    4.236
##
## Concordance= 0.646 (se = 0.059 )
## Likelihood ratio test= 7.58 on 5 df, p=0.2
## Wald test              = 7.38 on 5 df, p=0.2
## Score (logrank) test = 7.67 on 5 df, p=0.2
```

Similar to our logistic reg model, no real difference with Casp3 addition to the model.

Model 1 in paper: Cox PH model R censored at one year and two year, stratified by early late, MK2 as dichotomous (top 1/3 vs. bottom 2/3)

Model 1: MK2+other covariates, at 1 year and 2 year

```
summary(coxph(Surv(time, censor) ~ MK2_Expression_topQ +
  gender + smoking + age_at_diagnosis, id = submitter_id,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Early Stage", ]))
```

```
## Call:
## coxph(formula = Surv(time, censor) ~ MK2_Expression_topQ + gender +
##       smoking + age_at_diagnosis, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##       "Early Stage", ], id = submitter_id)
##
##      n= 312, number of events= 25
##      (7 observations deleted due to missingness)
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2_Expression_topQ1 -1.42504  0.24050  0.61622 -2.313  0.0207 *
## gendermale           0.81309  2.25486  0.41812  1.945  0.0518 .
## smoking              0.06489  1.06704  0.47247  0.137  0.8908
## age_at_diagnosis     -0.01696  0.98318  0.02001 -0.848  0.3966
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2_Expression_topQ1    0.2405    4.1580    0.07188    0.8047
## gendermale              2.2549    0.4435    0.99361    5.1171
## smoking                 1.0670    0.9372    0.42268    2.6937
## age_at_diagnosis        0.9832    1.0171    0.94536    1.0225
##
## Concordance= 0.701 (se = 0.053 )
## Likelihood ratio test= 12.76 on 4 df, p=0.01
## Wald test              = 10.18 on 4 df, p=0.04
## Score (logrank) test = 11.53 on 4 df, p=0.02
```

```
summary(coxph(Surv(time, censor) ~ MK2_Expression_topQ +
  gender + smoking + age_at_diagnosis, id = submitter_id,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
```

```

    "Late Stage", ]))

## Call:
## coxph(formula = Surv(time, censor) ~ MK2_Expression_topQ + gender +
##       smoking + age_at_diagnosis, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
##       "Late Stage", ], id = submitter_id)
##
## n= 86, number of events= 23
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2_Expression_topQ1 -0.71417  0.48960  0.62288 -1.147  0.252
## gendermale           0.19689  1.21761  0.41857  0.470  0.638
## smoking              -0.20205  0.81705  0.43065 -0.469  0.639
## age_at_diagnosis      0.02776  1.02815  0.02282  1.217  0.224
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2_Expression_topQ1  0.4896  2.0425  0.1444  1.660
## gendermale           1.2176  0.8213  0.5361  2.766
## smoking              0.8171  1.2239  0.3513  1.900
## age_at_diagnosis      1.0281  0.9726  0.9832  1.075
##
## Concordance= 0.626 (se = 0.058 )
## Likelihood ratio test= 4 on 4 df, p=0.4
## Wald test = 3.51 on 4 df, p=0.5
## Score (logrank) test = 3.6 on 4 df, p=0.5
summary(coxph(Surv(time, censor.2y) ~ MK2_Expression_topQ +
  gender + smoking + age_at_diagnosis, id = submitter_id,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
    "Early Stage", ]))

## Call:
## coxph(formula = Surv(time, censor.2y) ~ MK2_Expression_topQ +
##       gender + smoking + age_at_diagnosis, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined ==
##       "Early Stage", ], id = submitter_id)
##
## n= 312, number of events= 54
## (7 observations deleted due to missingness)
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2_Expression_topQ1 -0.643511  0.525444  0.319120 -2.017  0.0437 *
## gendermale           0.420566  1.522823  0.273765  1.536  0.1245
## smoking              0.309145  1.362260  0.330173  0.936  0.3491
## age_at_diagnosis      0.004461  1.004471  0.014624  0.305  0.7603
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2_Expression_topQ1  0.5254  1.9032  0.2811  0.9821
## gendermale           1.5228  0.6567  0.8905  2.6042
## smoking              1.3623  0.7341  0.7132  2.6020
## age_at_diagnosis      1.0045  0.9955  0.9761  1.0337
##
## Concordance= 0.624 (se = 0.038 )

```

```
## Likelihood ratio test= 8.26 on 4 df, p=0.08
## Wald test = 7.82 on 4 df, p=0.1
## Score (logrank) test = 8.05 on 4 df, p=0.09

summary(coxph(Surv(time, censor.2y) ~ MK2_Expression_topQ +
  gender + smoking + age_at_diagnosis, id = submitter_id,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Late Stage", ]))

## Call:
## coxph(formula = Surv(time, censor.2y) ~ MK2_Expression_topQ +
## gender + smoking + age_at_diagnosis, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
## "Late Stage", ], id = submitter_id)
##
## n= 86, number of events= 39
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## MK2_Expression_topQ1 -0.483550 0.616590 0.420764 -1.149 0.250
## gendermale          0.008068 1.008101 0.322004 0.025 0.980
## smoking             -0.103562 0.901620 0.339616 -0.305 0.760
## age_at_diagnosis     0.018713 1.018889 0.016820 1.113 0.266
##
##              exp(coef) exp(-coef) lower .95 upper .95
## MK2_Expression_topQ1 0.6166      1.6218    0.2703    1.407
## gendermale          1.0081      0.9920    0.5363    1.895
## smoking             0.9016      1.1091    0.4634    1.754
## age_at_diagnosis     1.0189      0.9815    0.9858    1.053
##
## Concordance= 0.604 (se = 0.048 )
## Likelihood ratio test= 3.2 on 4 df, p=0.5
## Wald test = 2.95 on 4 df, p=0.6
## Score (logrank) test = 2.99 on 4 df, p=0.6
```

Model 1: MK2 + other covariates + casp3, at 1 year and 2 year

```
summary(coxph(Surv(time, censor) ~ MK2_Expression_topQ +
  gender + smoking + age_at_diagnosis + Casp3.Expression,
  id = submitter_id, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Early Stage", ]))

## Call:
## coxph(formula = Surv(time, censor) ~ MK2_Expression_topQ + gender +
## smoking + age_at_diagnosis + Casp3.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
## "Early Stage", ], id = submitter_id)
##
## n= 312, number of events= 25
## (7 observations deleted due to missingness)
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## MK2_Expression_topQ1 -1.40534 0.24528 0.61767 -2.275 0.0229 *
## gendermale          0.82233 2.27580 0.41871 1.964 0.0495 *
## smoking             0.07424 1.07707 0.47357 0.157 0.8754
## age_at_diagnosis    -0.01664 0.98350 0.01999 -0.832 0.4051
## Casp3.Expression     0.30803 1.36074 0.56469 0.545 0.5854
## ---
```



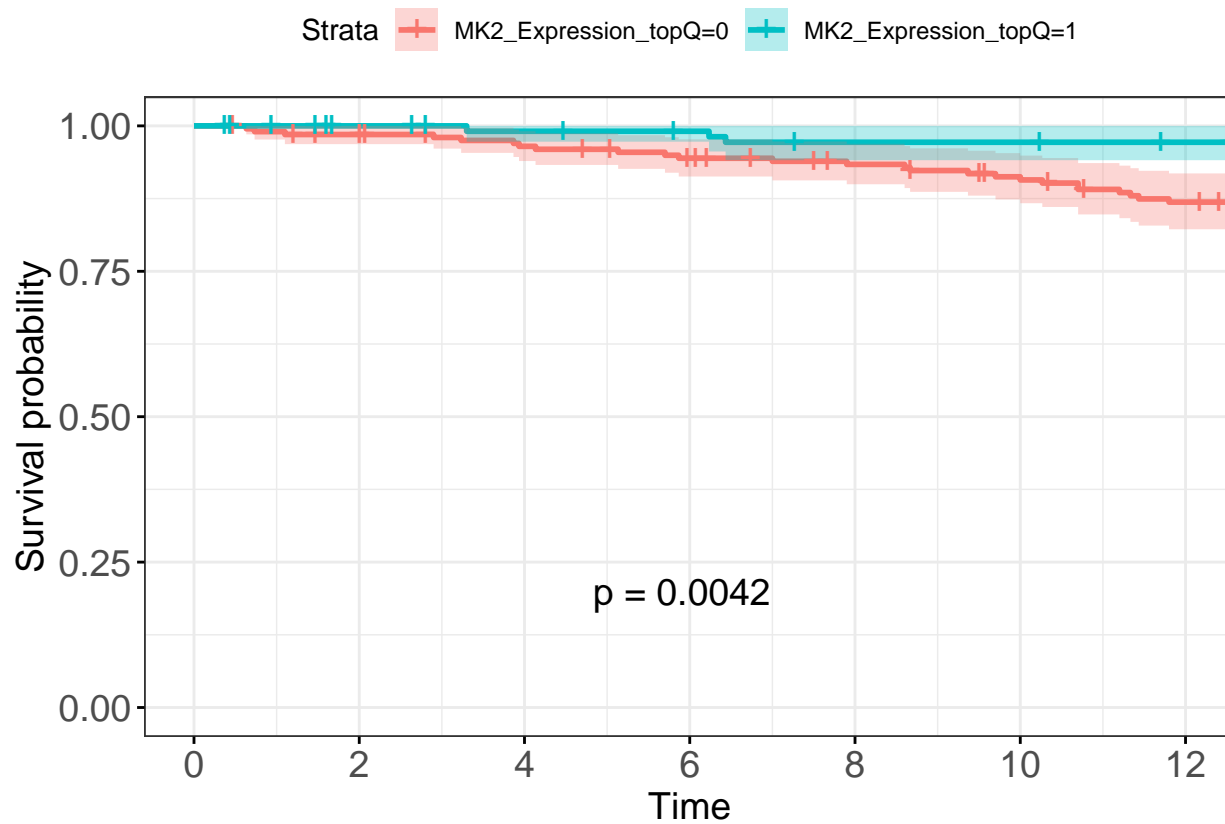
```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##               exp(coef) exp(-coef) lower .95 upper .95
## MK2_Expression_topQ1    0.2453    4.0769    0.0731    0.8231
## gendermale              2.2758    0.4394    1.0017    5.1706
## smoking                 1.0771    0.9284    0.4257    2.7248
## age_at_diagnosis        0.9835    1.0168    0.9457    1.0228
## Casp3.Expression        1.3607    0.7349    0.4499    4.1157
##
## Concordance= 0.707  (se = 0.054 )
## Likelihood ratio test= 13.04  on 5 df,   p=0.02
## Wald test              = 10.5   on 5 df,   p=0.06
## Score (logrank) test = 11.81  on 5 df,   p=0.04
```

So, the same pattern here - we don't really see an impact of adding casp3 to the multivariable Cox PH model.

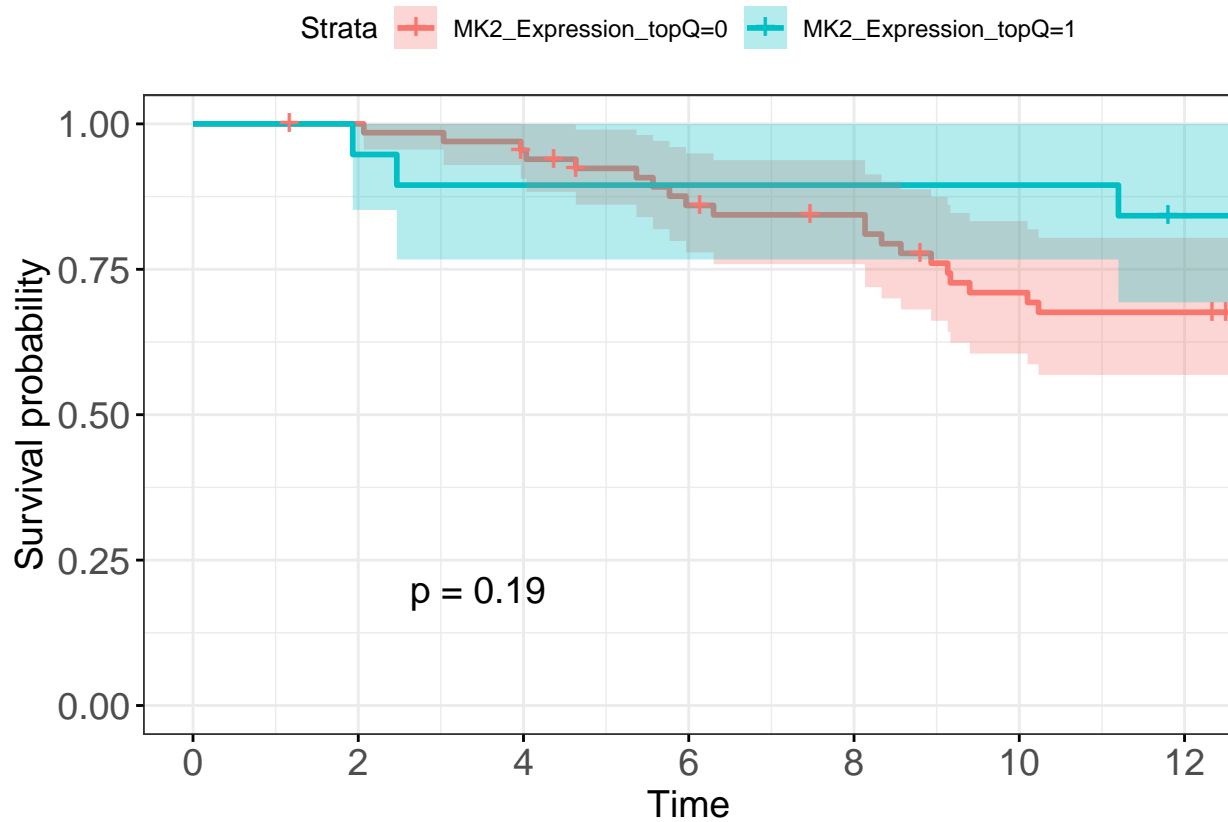
Model Graphics

KM curve - here, we're going to stratify MK2 by "hi" vs. "low" - Model 1

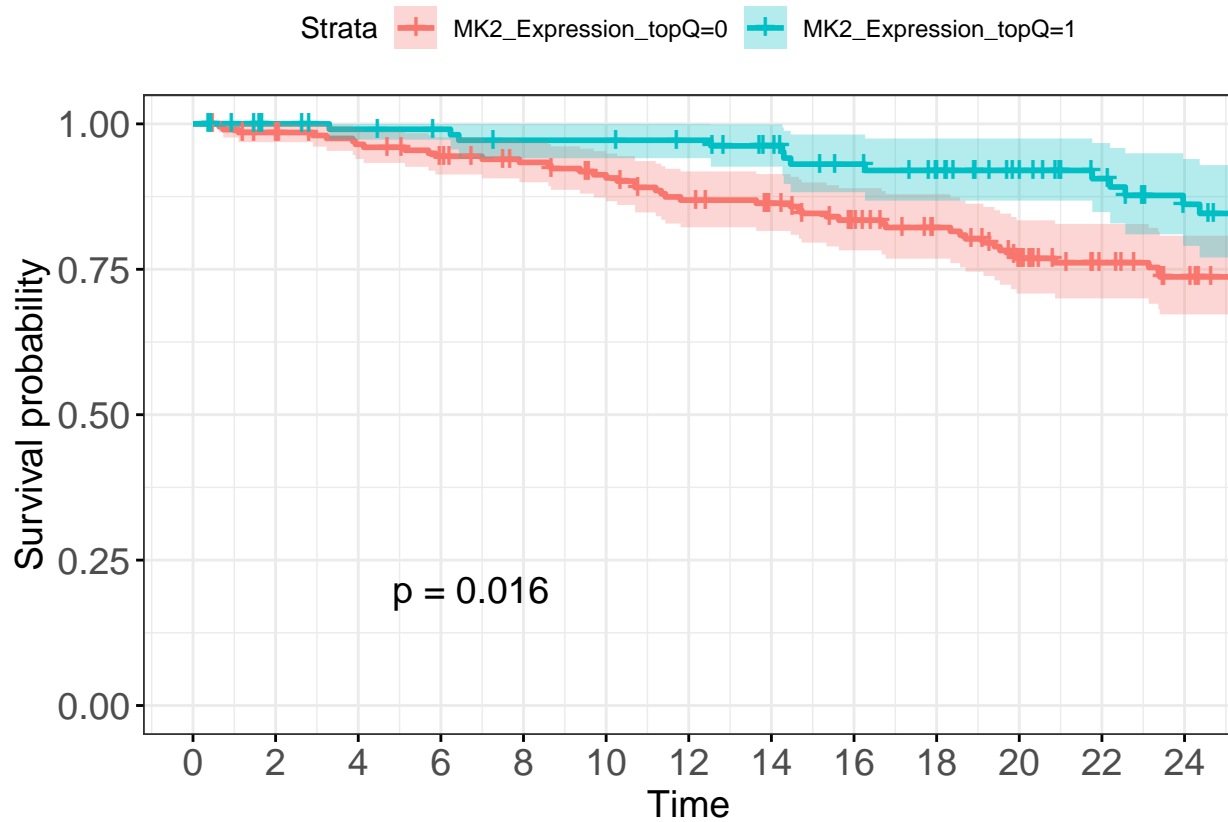
```
library(survminer)
library(survival)
library(ggsci)
km_early <- survfit(Surv(time, censor) ~ MK2_Expression_topQ,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Early Stage", ], type = "kaplan-meier")
ggsurvplot(km_early, xlim = c(0, 12), conf.int = TRUE,
  break.x.by = 2, pval = TRUE, font.y = 14, font.x = 14,
  font.tickslab = 14, ggtheme = theme_bw())
```



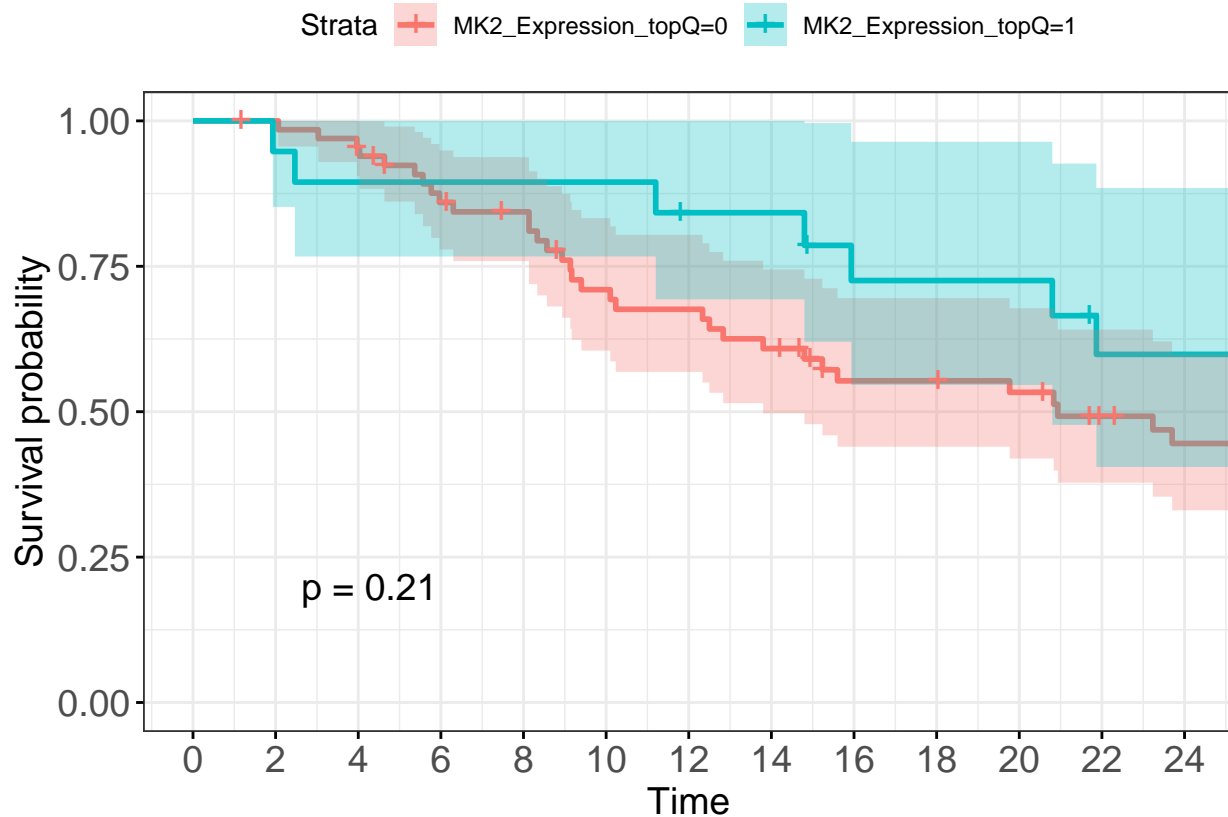
```
km_late <- survfit(Surv(time, censor) ~ MK2_Expression_topQ,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Late Stage", ], type = "kaplan-meier")
ggsurvplot(km_late, xlim = c(0, 12), conf.int = TRUE,
  break.x.by = 2, pval = TRUE, font.y = 14, font.x = 14,
  font.tickslab = 14, ggtheme = theme_bw())
```



```
km_early.2y <- survfit(Surv(time, censor.2y) ~ MK2_Expression_topQ,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Early Stage", ], type = "kaplan-meier")
ggsurvplot(km_early.2y, xlim = c(0, 24), conf.int = TRUE,
  break.x.by = 2, pval = TRUE, font.y = 14, font.x = 14,
  font.tickslab = 14, ggtheme = theme_bw())
```



```
km_late.2y <- survfit(Surv(time, censor.2y) ~ MK2_Expression_topQ,
  data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
    "Late Stage", ], type = "kaplan-meier")
ggsurvplot(km_late.2y, xlim = c(0, 24), conf.int = TRUE,
  break.x.by = 2, pval = TRUE, font.y = 14, font.x = 14,
  font.tickslab = 14, ggtheme = theme_bw())
```



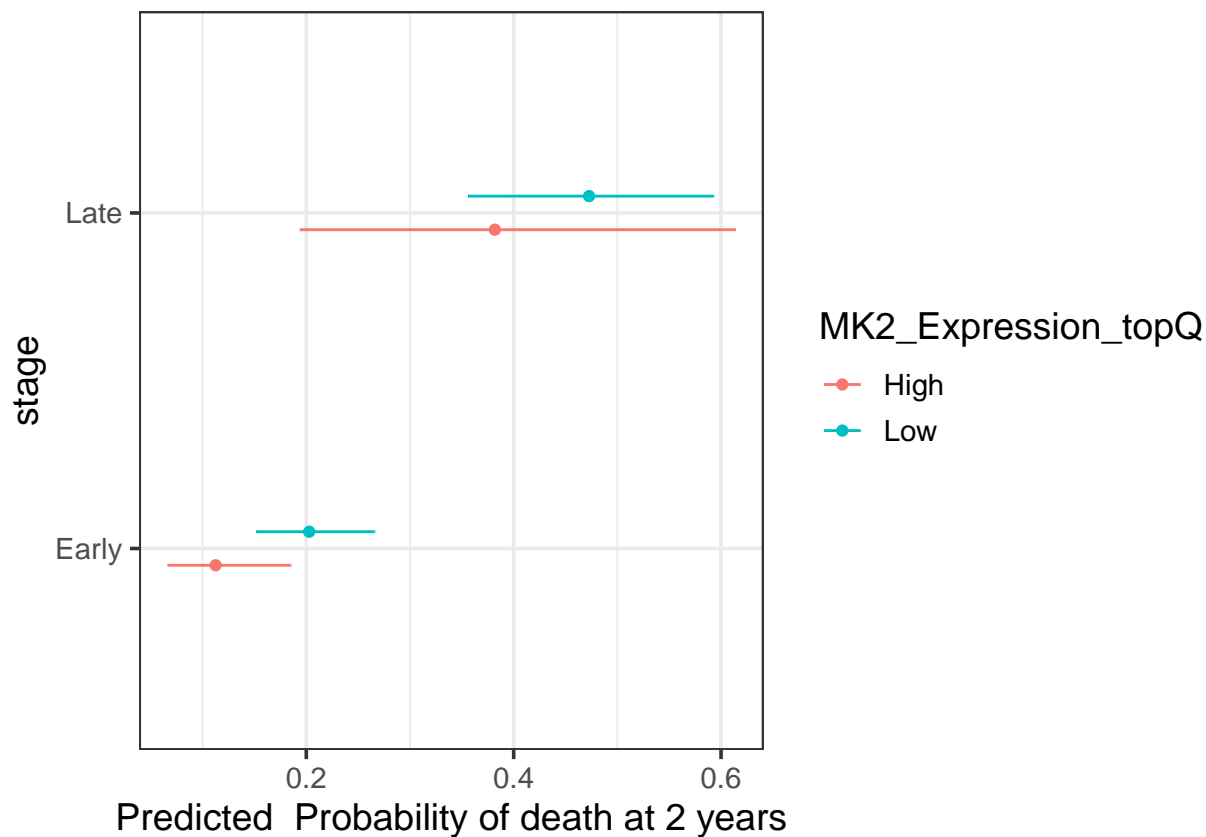
Predicted probabilities graphs for Logistic Regression (Model 2)

```
library(effects)

## Loading required package: carData
## lattice theme set by effectsTheme()
## See ?effectsTheme for details.

library(ggthemes)
MK2efx_early <- Effect("MK2_Expression_topQ", logreg_topq1_early_2y)
MK2efx_late <- Effect("MK2_Expression_topQ", logreg_topq1_late_2y)

MK2efx_early <- as.data.frame(MK2efx_early)
MK2efx_late <- as.data.frame(MK2efx_late)
MK2efx_early$stage <- "Early"
MK2efx_late$stage <- "Late"
MK2efx <- rbind(MK2efx_early, MK2efx_late)
MK2efx$MK2_Expression_topQ <- lapply(MK2efx$MK2_Expression_topQ,
  function(x) if (x == 0) return("Low") else return("High"))
MK2efx$MK2_Expression_topQ <- as.factor(unlist(MK2efx$MK2_Expression_topQ))
ggplot(MK2efx, aes(x = stage, y = fit, group = MK2_Expression_topQ,
  color = MK2_Expression_topQ, ymin = lower, ymax = upper)) +
  geom_point(position = position_dodge(width = 0.2)) +
  geom_errorbar(width = 0, position = position_dodge(width = 0.2)) +
  theme_bw(base_size = 14) + ylab("Predicted Probability of death at 2 years") +
  coord_flip()
```



Demographics / Table 1

```
library(tableone)
listVars <- c("age_at_diagnosis", "gender", "smoking",
             "ajcc_pathologic_stage", "ajcc_pathologic_n", "prior_malignancy",
             "primary_diagnosis", "time", "race", "MK2.Expression")
catVars <- c("gender", "smoking", "ajcc_pathologic_stage",
            "ajcc_pathologic_n", "prior_malignancy", "primary_diagnosis",
            "race")

# Table 1 where we stratify by Early/Late Stage
table1 <- CreateTableOne(listVars, clinical, catVars,
                        strata = c("ajcc_pathologic_stage_combined"), addOverall = TRUE)
clinical$MK2_Expression_topQ_t1 <- unlist(clinical$MK2_Expression_topQ)
kable(print(table1))
```

		Stratified by ajcc_pathologic_stage_combined
		Overall
##	n	483
##	age_at_diagnosis (mean (SD))	65.77 (9.97)
##	gender = male (%)	219 (45.3)
##	smoking = 1 (%)	331 (68.5)
##	ajcc_pathologic_stage (%)	
##	Stage I	5 (1.0)
##	Stage IA	129 (26.7)
##	Stage IB	128 (26.5)
##	Stage II	1 (0.2)
##	Stage IIA	50 (10.4)

##	Stage IIB	67 (13.9)
##	Stage IIIA	69 (14.3)
##	Stage IIIB	10 (2.1)
##	Stage IV	24 (5.0)
##	ajcc_pathologic_n (%)	
##	N0	313 (64.8)
##	N1	90 (18.6)
##	N2	68 (14.1)
##	N3	2 (0.4)
##	NX	10 (2.1)
##	prior_malignancy = yes (%)	79 (16.4)
##	primary_diagnosis (%)	
##	Acinar cell carcinoma	22 (4.6)
##	Adenocarcinoma with mixed subtypes	103 (21.3)
##	Adenocarcinoma, NOS	289 (59.8)
##	Bronchio-alveolar carcinoma, mucinous	5 (1.0)
##	Bronchiolo-alveolar adenocarcinoma, NOS	3 (0.6)
##	Bronchiolo-alveolar carcinoma, non-mucinous	17 (3.5)
##	Clear cell adenocarcinoma, NOS	2 (0.4)
##	Micropapillary carcinoma, NOS	2 (0.4)
##	Mucinous adenocarcinoma	13 (2.7)
##	Papillary adenocarcinoma, NOS	21 (4.3)
##	Signet ring cell carcinoma	1 (0.2)
##	Solid carcinoma, NOS	5 (1.0)
##	time (mean (SD))	879.55 (873.37)
##	race (%)	
##	american indian or alaska native	1 (0.2)
##	asian	8 (1.7)
##	black or african american	50 (10.4)
##	not reported	44 (9.1)
##	white	380 (78.7)
##	MK2.Expression (mean (SD))	3.86 (1.05)
##		Stratified by ajcc_pathologic_stage_combined
##		Early Stage
##	n	380
##	age_at_diagnosis (mean (SD))	66.01 (9.71)
##	gender = male (%)	171 (45.0)
##	smoking = 1 (%)	259 (68.2)
##	ajcc_pathologic_stage (%)	
##	Stage I	5 (1.3)
##	Stage IA	129 (33.9)
##	Stage IB	128 (33.7)
##	Stage II	1 (0.3)
##	Stage IIA	50 (13.2)
##	Stage IIB	67 (17.6)
##	Stage IIIA	0 (0.0)
##	Stage IIIB	0 (0.0)
##	Stage IV	0 (0.0)
##	ajcc_pathologic_n (%)	
##	N0	297 (78.2)
##	N1	77 (20.3)
##	N2	0 (0.0)
##	N3	0 (0.0)
##	NX	6 (1.6)

##	prior_malignancy = yes (%)	65 (17.1)		
##	primary_diagnosis (%)			
##	Acinar cell carcinoma	16 (4.2)		
##	Adenocarcinoma with mixed subtypes	77 (20.3)		
##	Adenocarcinoma, NOS	230 (60.5)		
##	Bronchio-alveolar carcinoma, mucinous	5 (1.3)		
##	Bronchiolo-alveolar adenocarcinoma, NOS	3 (0.8)		
##	Bronchiolo-alveolar carcinoma, non-mucinous	13 (3.4)		
##	Clear cell adenocarcinoma, NOS	1 (0.3)		
##	Micropapillary carcinoma, NOS	2 (0.5)		
##	Mucinous adenocarcinoma	12 (3.2)		
##	Papillary adenocarcinoma, NOS	15 (3.9)		
##	Signet ring cell carcinoma	1 (0.3)		
##	Solid carcinoma, NOS	5 (1.3)		
##	time (mean (SD))	929.04 (921.78)		
##	race (%)			
##	american indian or alaska native	1 (0.3)		
##	asian	6 (1.6)		
##	black or african american	41 (10.8)		
##	not reported	30 (7.9)		
##	white	302 (79.5)		
##	MK2.Expression (mean (SD))	3.91 (1.06)		
##		Stratified by ajcc_pathologic_stage_combined		
##		Late Stage	p	test
##	n	103		
##	age_at_diagnosis (mean (SD))	64.89 (10.86)	0.314	
##	gender = male (%)	48 (46.6)	0.859	
##	smoking = 1 (%)	72 (69.9)	0.827	
##	ajcc_pathologic_stage (%)		<0.001	
##	Stage I	0 (0.0)		
##	Stage IA	0 (0.0)		
##	Stage IB	0 (0.0)		
##	Stage II	0 (0.0)		
##	Stage IIA	0 (0.0)		
##	Stage IIB	0 (0.0)		
##	Stage IIIA	69 (67.0)		
##	Stage IIIB	10 (9.7)		
##	Stage IV	24 (23.3)		
##	ajcc_pathologic_n (%)		<0.001	
##	N0	16 (15.5)		
##	N1	13 (12.6)		
##	N2	68 (66.0)		
##	N3	2 (1.9)		
##	NX	4 (3.9)		
##	prior_malignancy = yes (%)	14 (13.6)	0.481	
##	primary_diagnosis (%)		0.621	
##	Acinar cell carcinoma	6 (5.8)		
##	Adenocarcinoma with mixed subtypes	26 (25.2)		
##	Adenocarcinoma, NOS	59 (57.3)		
##	Bronchio-alveolar carcinoma, mucinous	0 (0.0)		
##	Bronchiolo-alveolar adenocarcinoma, NOS	0 (0.0)		
##	Bronchiolo-alveolar carcinoma, non-mucinous	4 (3.9)		
##	Clear cell adenocarcinoma, NOS	1 (1.0)		
##	Micropapillary carcinoma, NOS	0 (0.0)		

##	Mucinous adenocarcinoma	1 (1.0)	
##	Papillary adenocarcinoma, NOS	6 (5.8)	
##	Signet ring cell carcinoma	0 (0.0)	
##	Solid carcinoma, NOS	0 (0.0)	
##	time (mean (SD))	696.95 (636.03)	0.017
##	race (%)		0.449
##	american indian or alaska native	0 (0.0)	
##	asian	2 (1.9)	
##	black or african american	9 (8.7)	
##	not reported	14 (13.6)	
##	white	78 (75.7)	
##	MK2.Expression (mean (SD))	3.66 (0.99)	0.032

	Overall	Early Stage	Late Stage	p	test
n	483	380	103		
age_at_diagnosis (mean (SD))	65.77 (9.97)	66.01 (9.71)	64.89 (10.86)	0.314	
gender = male (%)	219 (45.3)	171 (45.0)	48 (46.6)	0.859	
smoking = 1 (%)	331 (68.5)	259 (68.2)	72 (69.9)	0.827	
ajcc_pathologic_stage (%)				<0.001	
Stage I	5 (1.0)	5 (1.3)	0 (0.0)		
Stage IA	129 (26.7)	129 (33.9)	0 (0.0)		
Stage IB	128 (26.5)	128 (33.7)	0 (0.0)		
Stage II	1 (0.2)	1 (0.3)	0 (0.0)		
Stage IIA	50 (10.4)	50 (13.2)	0 (0.0)		
Stage IIB	67 (13.9)	67 (17.6)	0 (0.0)		
Stage IIIA	69 (14.3)	0 (0.0)	69 (67.0)		
Stage IIIB	10 (2.1)	0 (0.0)	10 (9.7)		
Stage IV	24 (5.0)	0 (0.0)	24 (23.3)		
ajcc_pathologic_n (%)				<0.001	
N0	313 (64.8)	297 (78.2)	16 (15.5)		
N1	90 (18.6)	77 (20.3)	13 (12.6)		
N2	68 (14.1)	0 (0.0)	68 (66.0)		
N3	2 (0.4)	0 (0.0)	2 (1.9)		
NX	10 (2.1)	6 (1.6)	4 (3.9)		
prior_malignancy = yes (%)	79 (16.4)	65 (17.1)	14 (13.6)	0.481	
primary_diagnosis (%)				0.621	
Acinar cell carcinoma	22 (4.6)	16 (4.2)	6 (5.8)		
Adenocarcinoma with mixed subtypes	103 (21.3)	77 (20.3)	26 (25.2)		
Adenocarcinoma, NOS	289 (59.8)	230 (60.5)	59 (57.3)		
Bronchio-alveolar carcinoma, mucinous	5 (1.0)	5 (1.3)	0 (0.0)		
Bronchiolo-alveolar adenocarcinoma, NOS	3 (0.6)	3 (0.8)	0 (0.0)		
Bronchiolo-alveolar carcinoma, non-mucinous	17 (3.5)	13 (3.4)	4 (3.9)		
Clear cell adenocarcinoma, NOS	2 (0.4)	1 (0.3)	1 (1.0)		
Micropapillary carcinoma, NOS	2 (0.4)	2 (0.5)	0 (0.0)		
Mucinous adenocarcinoma	13 (2.7)	12 (3.2)	1 (1.0)		
Papillary adenocarcinoma, NOS	21 (4.3)	15 (3.9)	6 (5.8)		
Signet ring cell carcinoma	1 (0.2)	1 (0.3)	0 (0.0)		
Solid carcinoma, NOS	5 (1.0)	5 (1.3)	0 (0.0)		
time (mean (SD))	879.55 (873.37)	929.04 (921.78)	696.95 (636.03)	0.017	
race (%)				0.449	
american indian or alaska native	1 (0.2)	1 (0.3)	0 (0.0)		
asian	8 (1.7)	6 (1.6)	2 (1.9)		
black or african american	50 (10.4)	41 (10.8)	9 (8.7)		
not reported	44 (9.1)	30 (7.9)	14 (13.6)		

	Overall	Early Stage	Late Stage	p	test
white	380 (78.7)	302 (79.5)	78 (75.7)		
MK2.Expression (mean (SD))	3.86 (1.05)	3.91 (1.06)	3.66 (0.99)	0.032	

Table 1 where we stratify by MK2 'hi' vs. 'low'

```
table1b <- CreateTableOne(listVars, clinical, catVars,
  strata = c("MK2_Expression_topQ_t1"), addOverall = TRUE)
kable(print(table1b))
```

```
##                               Stratified by MK2_Expression_topQ_t1
##                               Overall
##   n                               483
##   age_at_diagnosis (mean (SD))    65.77 (9.97)
##   gender = male (%)               219 (45.3)
##   smoking = 1 (%)                 331 (68.5)
##   ajcc_pathologic_stage (%)
##     Stage I                       5 ( 1.0)
##     Stage IA                     129 (26.7)
##     Stage IB                     128 (26.5)
##     Stage II                      1 ( 0.2)
##     Stage IIA                     50 (10.4)
##     Stage IIB                     67 (13.9)
##     Stage IIIA                    69 (14.3)
##     Stage IIIB                    10 ( 2.1)
##     Stage IV                     24 ( 5.0)
##   ajcc_pathologic_n (%)
##     N0                           313 (64.8)
##     N1                           90 (18.6)
##     N2                           68 (14.1)
##     N3                           2 ( 0.4)
##     NX                           10 ( 2.1)
##   prior_malignancy = yes (%)      79 (16.4)
##   primary_diagnosis (%)
##     Acinar cell carcinoma         22 ( 4.6)
##     Adenocarcinoma with mixed subtypes 103 (21.3)
##     Adenocarcinoma, NOS          289 (59.8)
##     Bronchio-alveolar carcinoma, mucinous 5 ( 1.0)
##     Bronchiolo-alveolar adenocarcinoma, NOS 3 ( 0.6)
##     Bronchiolo-alveolar carcinoma, non-mucinous 17 ( 3.5)
##     Clear cell adenocarcinoma, NOS 2 ( 0.4)
##     Micropapillary carcinoma, NOS 2 ( 0.4)
##     Mucinous adenocarcinoma      13 ( 2.7)
##     Papillary adenocarcinoma, NOS 21 ( 4.3)
##     Signet ring cell carcinoma   1 ( 0.2)
##     Solid carcinoma, NOS         5 ( 1.0)
##   time (mean (SD))                879.55 (873.37)
##   race (%)
##     american indian or alaska native 1 ( 0.2)
##     asian                             8 ( 1.7)
##     black or african american       50 (10.4)
##     not reported                    44 ( 9.1)
##     white                          380 (78.7)
##   MK2.Expression (mean (SD))      3.86 (1.05)
```

##		Stratified by MK2_Expression_topQ_t1	
##		0	
##	n	319	
##	age_at_diagnosis (mean (SD))	66.11 (9.96)	
##	gender = male (%)	153 (48.0)	
##	smoking = 1 (%)	213 (66.8)	
##	ajcc_pathologic_stage (%)		
##	Stage I	3 (0.9)	
##	Stage IA	81 (25.4)	
##	Stage IB	80 (25.1)	
##	Stage II	0 (0.0)	
##	Stage IIA	31 (9.7)	
##	Stage IIB	47 (14.7)	
##	Stage IIIA	53 (16.6)	
##	Stage IIIB	6 (1.9)	
##	Stage IV	18 (5.6)	
##	ajcc_pathologic_n (%)		
##	N0	201 (63.0)	
##	N1	62 (19.4)	
##	N2	50 (15.7)	
##	N3	0 (0.0)	
##	NX	6 (1.9)	
##	prior_malignancy = yes (%)	50 (15.7)	
##	primary_diagnosis (%)		
##	Acinar cell carcinoma	15 (4.7)	
##	Adenocarcinoma with mixed subtypes	74 (23.2)	
##	Adenocarcinoma, NOS	187 (58.6)	
##	Bronchio-alveolar carcinoma, mucinous	5 (1.6)	
##	Bronchiolo-alveolar adenocarcinoma, NOS	2 (0.6)	
##	Bronchiolo-alveolar carcinoma, non-mucinous	10 (3.1)	
##	Clear cell adenocarcinoma, NOS	2 (0.6)	
##	Micropapillary carcinoma, NOS	0 (0.0)	
##	Mucinous adenocarcinoma	9 (2.8)	
##	Papillary adenocarcinoma, NOS	11 (3.4)	
##	Signet ring cell carcinoma	1 (0.3)	
##	Solid carcinoma, NOS	3 (0.9)	
##	time (mean (SD))	869.21 (953.82)	
##	race (%)		
##	american indian or alaska native	0 (0.0)	
##	asian	6 (1.9)	
##	black or african american	35 (11.0)	
##	not reported	36 (11.3)	
##	white	242 (75.9)	
##	MK2.Expression (mean (SD))	3.27 (0.60)	
##		Stratified by MK2_Expression_topQ_t1	
##		1	p test
##	n	164	
##	age_at_diagnosis (mean (SD))	65.12 (9.99)	0.307
##	gender = male (%)	66 (40.2)	0.129
##	smoking = 1 (%)	118 (72.0)	0.290
##	ajcc_pathologic_stage (%)		0.359
##	Stage I	2 (1.2)	
##	Stage IA	48 (29.3)	
##	Stage IB	48 (29.3)	

```

##      Stage II                      1 ( 0.6)
##      Stage IIA                     19 (11.6)
##      Stage IIB                      20 (12.2)
##      Stage IIIA                     16 ( 9.8)
##      Stage IIIB                      4 ( 2.4)
##      Stage IV                       6 ( 3.7)
##  ajcc_pathologic_n (%)              0.162
##      N0                           112 (68.3)
##      N1                            28 (17.1)
##      N2                            18 (11.0)
##      N3                             2 ( 1.2)
##      NX                             4 ( 2.4)
##  prior_malignancy = yes (%)         0.663
##  primary_diagnosis (%)              0.356
##      Acinar cell carcinoma          7 ( 4.3)
##      Adenocarcinoma with mixed subtypes 29 (17.7)
##      Adenocarcinoma, NOS            102 (62.2)
##      Bronchio-alveolar carcinoma, mucinous 0 ( 0.0)
##      Bronchiolo-alveolar adenocarcinoma, NOS 1 ( 0.6)
##      Bronchiolo-alveolar carcinoma, non-mucinous 7 ( 4.3)
##      Clear cell adenocarcinoma, NOS    0 ( 0.0)
##      Micropapillary carcinoma, NOS     2 ( 1.2)
##      Mucinous adenocarcinoma          4 ( 2.4)
##      Papillary adenocarcinoma, NOS     10 ( 6.1)
##      Signet ring cell carcinoma        0 ( 0.0)
##      Solid carcinoma, NOS             2 ( 1.2)
##  time (mean (SD))                   899.66 (692.84) 0.717
##  race (%)                           0.078
##      american indian or alaska native 1 ( 0.6)
##      asian                           2 ( 1.2)
##      black or african american        15 ( 9.1)
##      not reported                     8 ( 4.9)
##      white                           138 (84.1)
##  MK2.Expression (mean (SD))          5.01 (0.75) <0.001

```

	Overall	0	1	p	test
n	483	319	164		
age_at_diagnosis (mean (SD))	65.77 (9.97)	66.11 (9.96)	65.12 (9.99)	0.307	
gender = male (%)	219 (45.3)	153 (48.0)	66 (40.2)	0.129	
smoking = 1 (%)	331 (68.5)	213 (66.8)	118 (72.0)	0.290	
ajcc_pathologic_stage (%)				0.359	
Stage I	5 (1.0)	3 (0.9)	2 (1.2)		
Stage IA	129 (26.7)	81 (25.4)	48 (29.3)		
Stage IB	128 (26.5)	80 (25.1)	48 (29.3)		
Stage II	1 (0.2)	0 (0.0)	1 (0.6)		
Stage IIA	50 (10.4)	31 (9.7)	19 (11.6)		
Stage IIB	67 (13.9)	47 (14.7)	20 (12.2)		
Stage IIIA	69 (14.3)	53 (16.6)	16 (9.8)		
Stage IIIB	10 (2.1)	6 (1.9)	4 (2.4)		
Stage IV	24 (5.0)	18 (5.6)	6 (3.7)		
ajcc_pathologic_n (%)				0.162	
N0	313 (64.8)	201 (63.0)	112 (68.3)		
N1	90 (18.6)	62 (19.4)	28 (17.1)		
N2	68 (14.1)	50 (15.7)	18 (11.0)		

	Overall	0	1	p	test
N3	2 (0.4)	0 (0.0)	2 (1.2)		
NX	10 (2.1)	6 (1.9)	4 (2.4)		
prior_malignancy = yes (%)	79 (16.4)	50 (15.7)	29 (17.7)	0.663	
primary_diagnosis (%)				0.356	
Acinar cell carcinoma	22 (4.6)	15 (4.7)	7 (4.3)		
Adenocarcinoma with mixed subtypes	103 (21.3)	74 (23.2)	29 (17.7)		
Adenocarcinoma, NOS	289 (59.8)	187 (58.6)	102 (62.2)		
Bronchio-alveolar carcinoma, mucinous	5 (1.0)	5 (1.6)	0 (0.0)		
Bronchiolo-alveolar adenocarcinoma, NOS	3 (0.6)	2 (0.6)	1 (0.6)		
Bronchiolo-alveolar carcinoma, non-mucinous	17 (3.5)	10 (3.1)	7 (4.3)		
Clear cell adenocarcinoma, NOS	2 (0.4)	2 (0.6)	0 (0.0)		
Micropapillary carcinoma, NOS	2 (0.4)	0 (0.0)	2 (1.2)		
Mucinous adenocarcinoma	13 (2.7)	9 (2.8)	4 (2.4)		
Papillary adenocarcinoma, NOS	21 (4.3)	11 (3.4)	10 (6.1)		
Signet ring cell carcinoma	1 (0.2)	1 (0.3)	0 (0.0)		
Solid carcinoma, NOS	5 (1.0)	3 (0.9)	2 (1.2)		
time (mean (SD))	879.55 (873.37)	869.21 (953.82)	899.66 (692.84)	0.717	
race (%)				0.078	
american indian or alaska native	1 (0.2)	0 (0.0)	1 (0.6)		
asian	8 (1.7)	6 (1.9)	2 (1.2)		
black or african american	50 (10.4)	35 (11.0)	15 (9.1)		
not reported	44 (9.1)	36 (11.3)	8 (4.9)		
white	380 (78.7)	242 (75.9)	138 (84.1)		
MK2.Expression (mean (SD))	3.86 (1.05)	3.27 (0.60)	5.01 (0.75)	<0.001	

Additional models and analysis

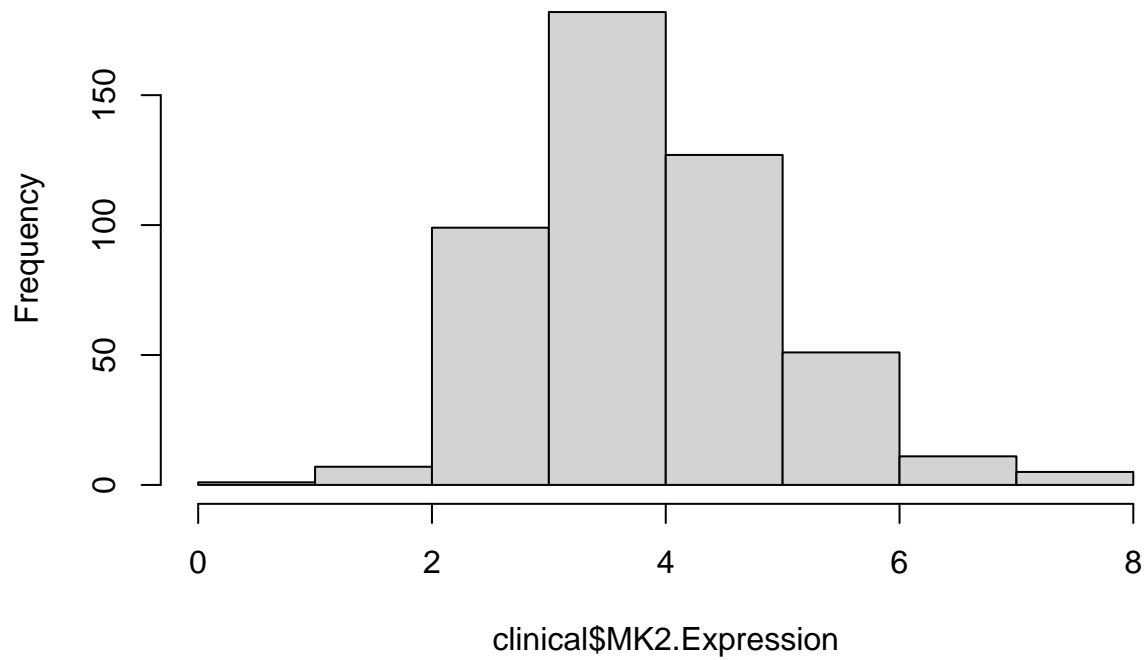
Continuous MK2 expression, no R censoring: Survival ~ expression (continuous) + clinical stage (stage 1, stage 2,...)

In this model, we use MK2 expression as a continuous variable - in other words, don't split into low vs. high (for that, see below). The 'MK.Expression' variable is the MK2 transcript level. Here, we're going to use the full follow-up time without any R censoring.

```
# COX PH model
coxph_model_all <- coxph(Surv(time, censor.2y) ~ gender +
  smoking + MK2.Expression, id = submitter_id, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_path
    "Early Stage", ])

# examine MK2 expression values
hist(clinical$MK2.Expression)
```

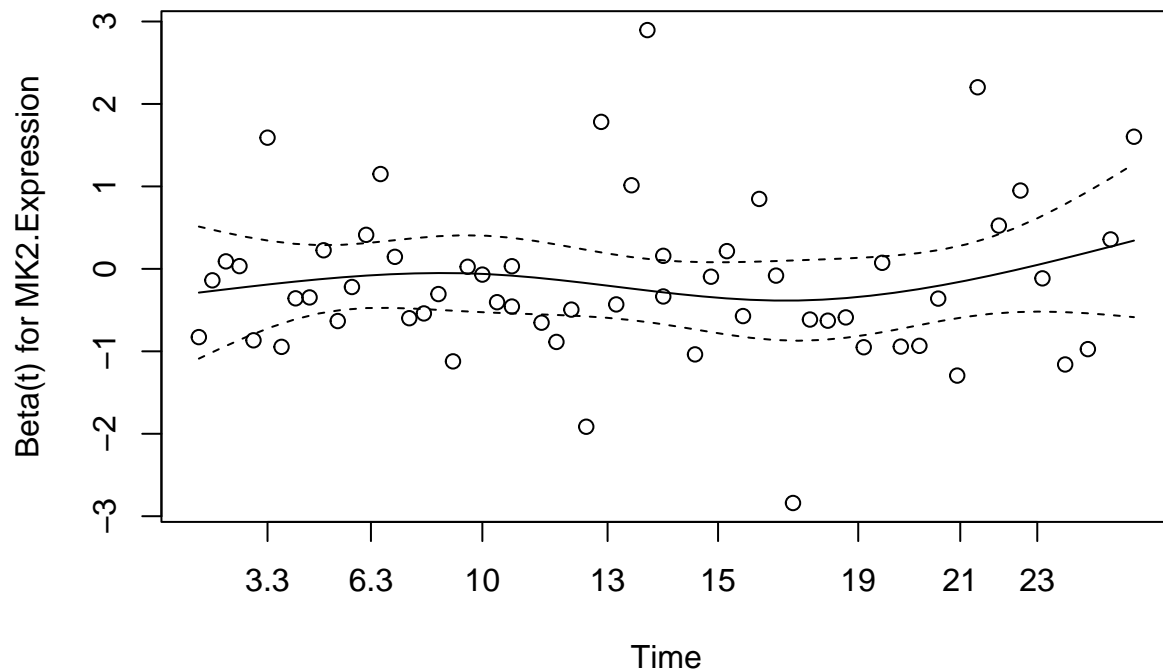
Histogram of clinical\$MK2.Expression



```
cox.zph(coxph_model_all)
```

```
##           chisq df    p
## gender      2.021  1 0.16
## smoking     1.142  1 0.29
## MK2.Expression 0.142  1 0.71
## GLOBAL      3.305  3 0.35
```

```
plot(cox.zph(coxph_model_all)[3])
```



```
summary(coxph_model_all)
```

```
## Call:
## coxph(formula = Surv(time, censor.2y) ~ gender + smoking + MK2.Expression,
##       data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined ==
##       "Early Stage", ], id = submitter_id)
##
##      n= 319, number of events= 58
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## gendermale      0.54980  1.73290  0.26657  2.063  0.0392 *
## smoking         0.04049  1.04132  0.29428  0.138  0.8906
## MK2.Expression -0.16220   0.85027  0.12556 -1.292  0.1964
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##              exp(coef) exp(-coef) lower .95 upper .95
## gendermale      1.7329    0.5771    1.0277    2.922
## smoking         1.0413    0.9603    0.5849    1.854
## MK2.Expression   0.8503    1.1761    0.6648    1.088
##
## Concordance= 0.605 (se = 0.037 )
## Likelihood ratio test= 6.57 on 3 df,  p=0.09
## Wald test            = 6.5 on 3 df,  p=0.09
## Score (logrank) test = 6.64 on 3 df,  p=0.08
```

Logistic regression for death at 1 year, stratified by specific stage

As shown above, LR using MK2 as a continuous variable but tumor stage dichotomized as early/late produced some interesting results. We wondered if we were “overlumping” stage by combining Stage I and II into “Early Stage”. Here we split stages into 1-4. Of course, now the problem was low numbers. These models did not yield any additional information, so were not considered further.

```
# LR stratified by individual stage
```

```
summary(glm(death_at_1year ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 == "Stage 1", ], family = binomial(link = "logit")))
```

```
##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression, family = binomial(link = "logit"),
##     data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##     "Stage 1", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.4456  -0.3181  -0.2899  -0.2583   2.6770
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   -1.9805     1.2743  -1.554   0.120
## MK2.Expression -0.2988     0.3375  -0.885   0.376
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 74.394 on 210 degrees of freedom
```

```

## Residual deviance: 73.576  on 209  degrees of freedom
## AIC: 77.576
##
## Number of Fisher Scoring iterations: 6
summary(glm(death_at_1year ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic
"Stage 2", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 2", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.9166  -0.6776  -0.5733  -0.4215   2.2349
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    0.2439     1.0268   0.238  0.8123
## MK2.Expression -0.4773     0.2768  -1.724  0.0846 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 100.47  on 107  degrees of freedom
## Residual deviance:  97.09  on 106  degrees of freedom
## AIC: 101.09
##
## Number of Fisher Scoring iterations: 4
summary(glm(death_at_1year ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic
"Stage 3", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 3", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.0027  -0.8448  -0.7861   1.4666   1.7445
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    0.05054     1.05892   0.048  0.962
## MK2.Expression -0.27542     0.29351  -0.938  0.348
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 79.905  on 66  degrees of freedom
## Residual deviance: 78.958  on 65  degrees of freedom

```



```

## AIC: 82.958
##
## Number of Fisher Scoring iterations: 4
summary(glm(death_at_1year ~ MK2.Expression, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic
  "Stage 4", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 4", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.1913  -0.6872  -0.5152  -0.4273   1.8134
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -3.8831     2.4529  -1.583   0.113
## MK2.Expression   0.7118     0.6355   1.120   0.263
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 19.557  on 18  degrees of freedom
## Residual deviance: 18.216  on 17  degrees of freedom
## AIC: 22.216
##
## Number of Fisher Scoring iterations: 4
summary(glm(death_at_1year ~ MK2.Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pat
  "Stage 1", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2.Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 1", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.3550  -0.3550  -0.3550  -0.1586   2.9604
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -2.7327     0.3649  -7.490 6.91e-14 ***
## MK2.Expression_topQ1 -1.6367     1.0704  -1.529   0.126
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 74.394  on 210  degrees of freedom
## Residual deviance: 70.985  on 209  degrees of freedom
## AIC: 74.985

```

```
##
## Number of Fisher Scoring iterations: 7
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_patl
  "Stage 2", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 2", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.7339  -0.7339  -0.7339  -0.3381   2.4043
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -1.1741     0.2775  -4.231 2.33e-05 ***
## MK2_Expression_topQ1 -1.6591     0.7787  -2.131  0.0331 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 100.474  on 107  degrees of freedom
## Residual deviance:  94.152  on 106  degrees of freedom
## AIC: 98.152
##
## Number of Fisher Scoring iterations: 5
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_patl
  "Stage 3", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 3", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.8898  -0.8898  -0.8898   1.4953   2.0074
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -0.7221     0.2956  -2.443  0.0146 *
## MK2_Expression_topQ1 -1.1497     0.8151  -1.411  0.1584
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 79.905  on 66  degrees of freedom
## Residual deviance: 77.506  on 65  degrees of freedom
```

```

## AIC: 81.506
##
## Number of Fisher Scoring iterations: 4
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pat
  "Stage 4", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 4", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.7585  -0.6681  -0.6681  -0.6681   1.7941
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -1.3863     0.6455  -2.148  0.0317 *
## MK2_Expression_topQ1  0.2877     1.3229   0.217  0.8278
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 19.557  on 18  degrees of freedom
## Residual deviance: 19.511  on 17  degrees of freedom
## AIC: 23.511
##
## Number of Fisher Scoring iterations: 4
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pat
  "Stage 1", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_pathologic_stage_combined.2 ==
##      "Stage 1", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.3550  -0.3550  -0.3550  -0.1586   2.9604
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -2.7327     0.3649  -7.490 6.91e-14 ***
## MK2_Expression_topQ1 -1.6367     1.0704  -1.529  0.126
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 74.394  on 210  degrees of freedom

```

```

## Residual deviance: 70.985  on 209  degrees of freedom
## AIC: 74.985
##
## Number of Fisher Scoring iterations: 7
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pat
"Stage 2", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined.2 ==
##      "Stage 2", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.7339  -0.7339  -0.7339  -0.3381   2.4043
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -1.1741     0.2775  -4.231 2.33e-05 ***
## MK2_Expression_topQ1 -1.6591     0.7787  -2.131  0.0331 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 100.474  on 107  degrees of freedom
## Residual deviance:  94.152  on 106  degrees of freedom
## AIC: 98.152
##
## Number of Fisher Scoring iterations: 5
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pat
"Stage 3", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
##      data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined.2 ==
##      "Stage 3", ])
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.8898  -0.8898  -0.8898   1.4953   2.0074
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      -0.7221     0.2956  -2.443  0.0146 *
## MK2_Expression_topQ1 -1.1497     0.8151  -1.411  0.1584
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##

```

```
## Null deviance: 79.905 on 66 degrees of freedom
## Residual deviance: 77.506 on 65 degrees of freedom
## AIC: 81.506
##
## Number of Fisher Scoring iterations: 4
summary(glm(death_at_1year ~ MK2_Expression_topQ, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pat
"Stage 4", ], family = binomial(link = "logit")))

##
## Call:
## glm(formula = death_at_1year ~ MK2_Expression_topQ, family = binomial(link = "logit"),
## data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic_stage_combined.2 ==
## "Stage 4", ])
##
## Deviance Residuals:
## Min 1Q Median 3Q Max
## -0.7585 -0.6681 -0.6681 -0.6681 1.7941
##
## Coefficients:
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.3863 0.6455 -2.148 0.0317 *
## MK2_Expression_topQ1 0.2877 1.3229 0.217 0.8278
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 19.557 on 18 degrees of freedom
## Residual deviance: 19.511 on 17 degrees of freedom
## AIC: 23.511
##
## Number of Fisher Scoring iterations: 4
```

Jackknife-type analysis for effects due to long period of patient recruitment

It appears that patients were recruited across many years. Chemo regimens change with time. To Figure out whether this plays a role, we re-estimate the odds ratio for high MK2 levels in the multivariate Linear Regression Model (Model 2) multiple times, each time removing only the patients recruited in a particular year

```
MK2_tvary <- MK2_CoxPH_quantile
years <- as.data.frame(table(as.factor(as.character(clinical$year_of_diagnosis))))
MK2_tvary$year_of_diagnosis <- as.factor(as.character(MK2_tvary$year_of_diagnosis))

returnPointEst <- function(x) {
  modelresult <- glm(death_at_1year ~ MK2_Expression_topQ +
    smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$ajcc_pathologic
    "Early Stage", ], family = binomial(link = "logit"))
  return(exp(modelresult$coefficients[1]))
}

years$estimate <- lapply(as.character(years$Var1),
  returnPointEst)
```

```

years$estimate <- lapply(years$estimate, function(x) return(x[[1]]))

returnPointEstlci <- function(x) {
  modelresult <- glm(death_at_1year ~ MK2_Expression_topQ +
    smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_patholog
    "Early Stage", ], family = binomial(link = "logit"))
  return(exp(confint.lm(modelresult)[2]))
}

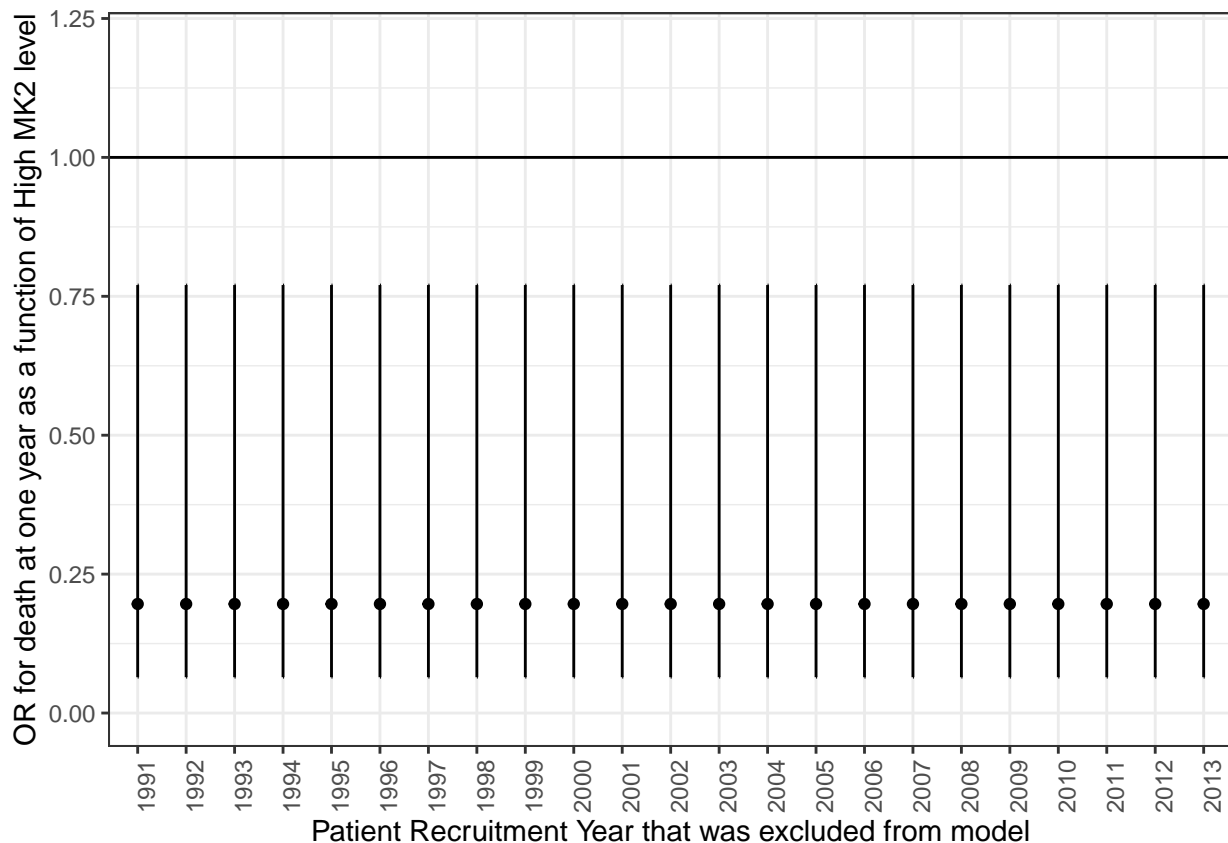
returnPointEstuci <- function(x) {
  modelresult <- glm(death_at_1year ~ MK2_Expression_topQ +
    smoking + age_at_diagnosis + gender, data = MK2_CoxPH_quantile[MK2_CoxPH_quantile$aajcc_patholog
    "Early Stage", ], family = binomial(link = "logit"))
  return(exp(confint.lm(modelresult)[7]))
}

years$estimate_lci <- lapply(as.character(years$Var1),
  returnPointEstlci)
years$estimate_uci <- lapply(as.character(years$Var1),
  returnPointEstuci)

years$estimate <- as.numeric(years$estimate)
years$estimate_lci <- as.numeric(years$estimate_lci)
years$estimate_uci <- as.numeric(years$estimate_uci)

ggplot(years, aes(x = Var1, y = estimate)) + geom_point() +
  geom_errorbar(aes(ymin = estimate_lci, ymax = estimate_uci),
    width = 0) + theme_bw() + ylim(0, 1.2) + geom_hline(yintercept = 1) +
  ylab("OR for death at one year as a function of High MK2 level") +
  xlab("Patient Recruitment Year that was excluded from model") +
  theme(axis.text.x = element_text(angle = 90, hjust = 1))

```



Normal vs. tumor tissue

One of the benefits of downloading data from the TCGA is that it includes the “11” barcode, which is normal tissue

```
expr2_normalcontrols$Patient <- substring(expr2_normalcontrols$Patient,
1, 12)
expr2_normalcontrols$sample.type = "Normal Tissue"
expr2_t <- expr3
expr2_t$sample.type = "Tumor"

exp2_paired <- left_join(expr2_normalcontrols, expr2_t,
by = c("Patient"))
exp2_paired <- exp2_paired[!is.na(exp2_paired$MK2.Expression.Manual.y),
]
colnames(exp2_paired) <- c("MK2.Expression.Normal",
"Stage.1", "Patient", "sample.type.1", "MK2.Expression.Tumor",
"Stage.2", "sample.type.2")

# These data are plotted elsewhere. (See graphpad
# file)

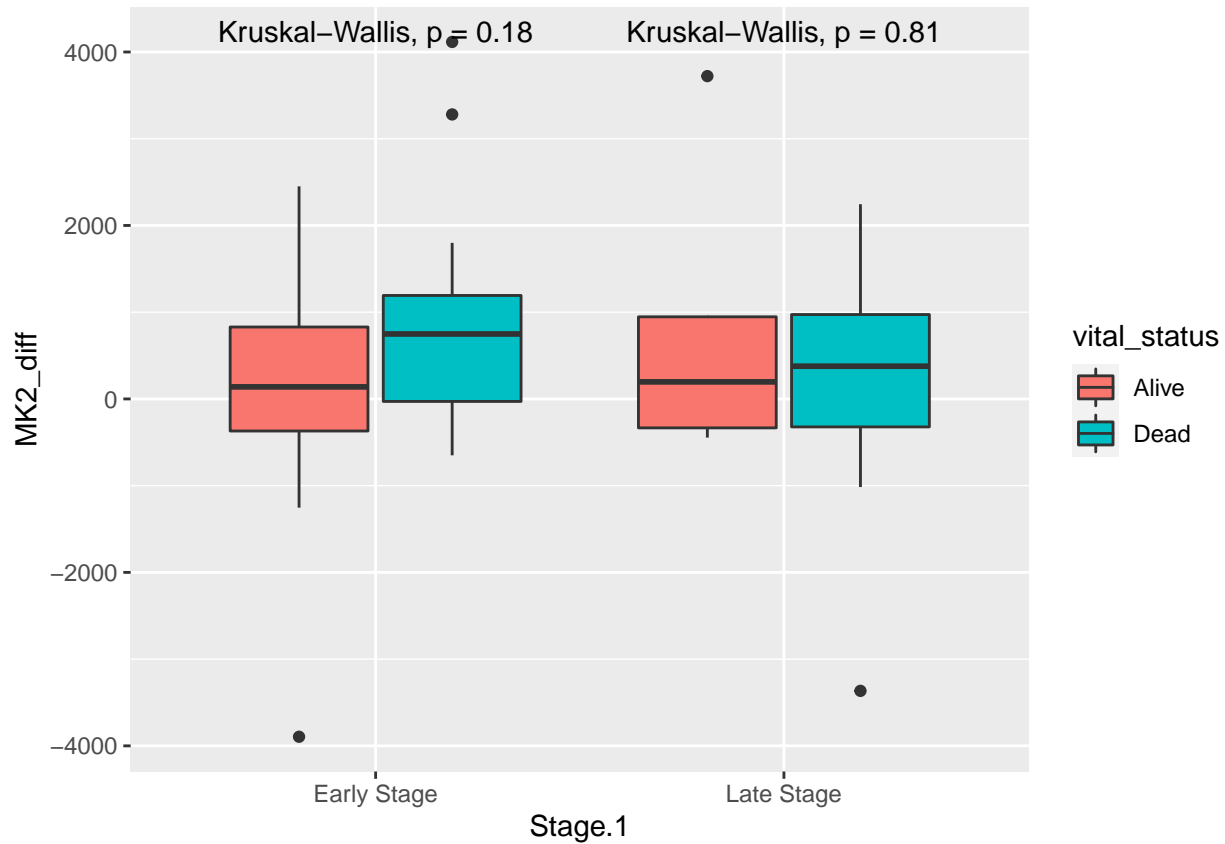
# look at FC in MK2 as a predictor of survival

exp2_paired$MK2_diff <- exp2_paired$MK2.Expression.Tumor -
exp2_paired$MK2.Expression.Normal
exp2_paired$MK2_FC <- exp2_paired$MK2.Expression.Tumor/exp2_paired$MK2.Expression.Normal
exp2_paired$Patient <- substring(exp2_paired$Patient,
```

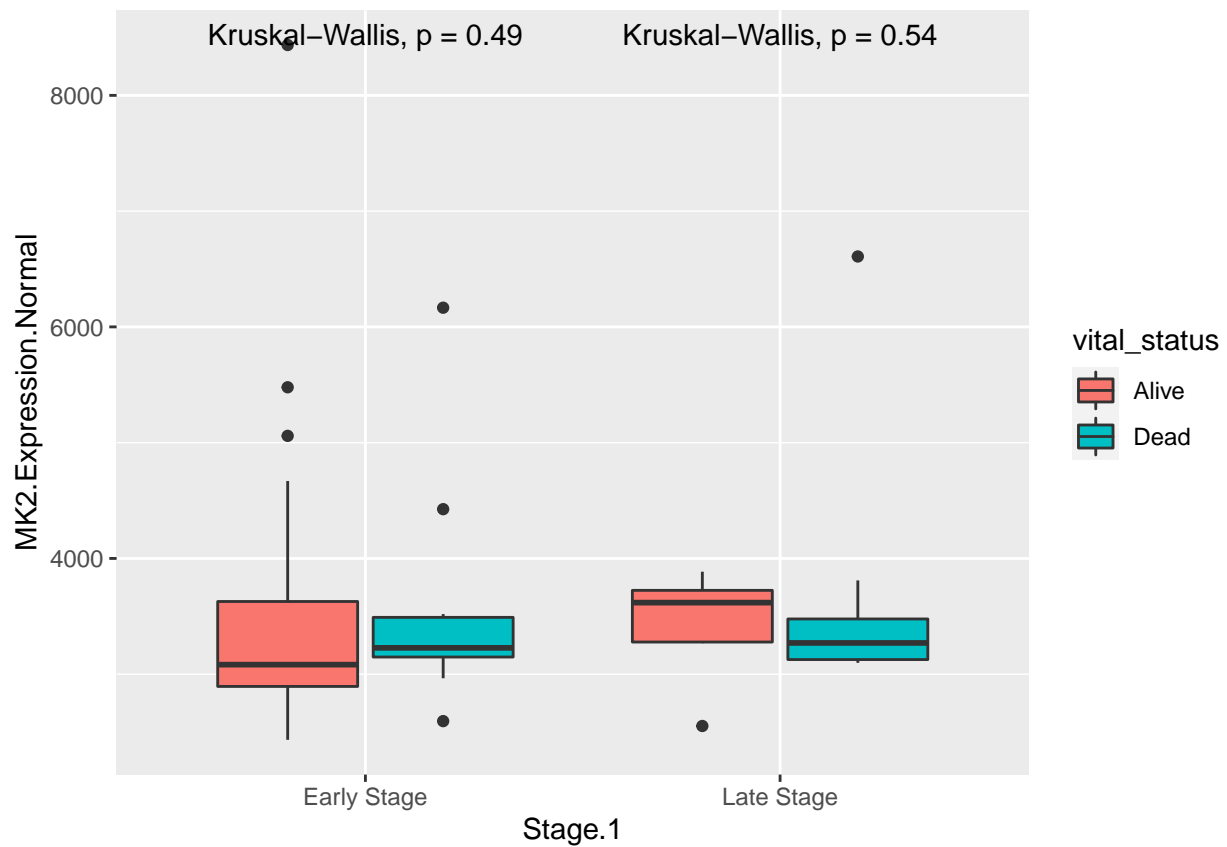
```
1, 12)
```

```
exp2_paired <- left_join(exp2_paired, clinical, by = c(Patient = "submitter_id"))
```

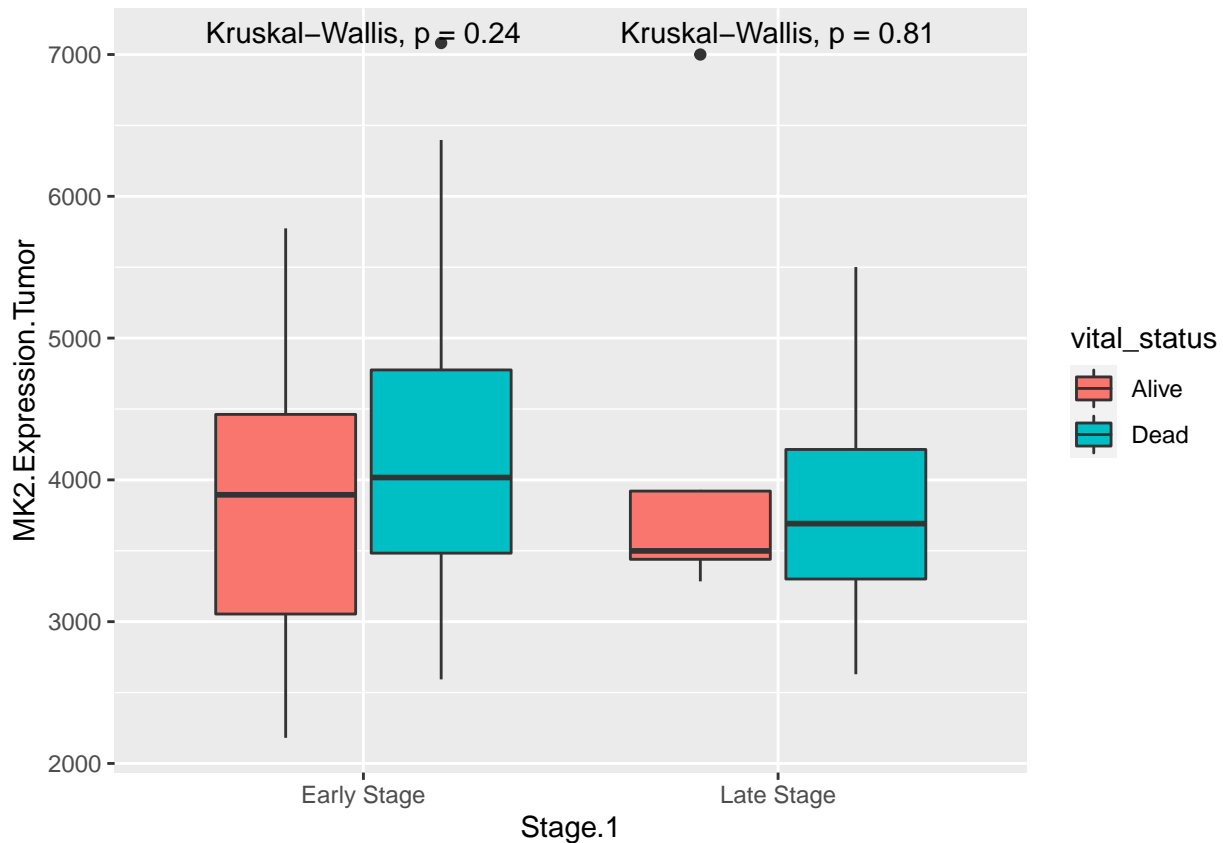
```
ggplot(exp2_paired, aes(x = Stage.1, y = MK2_diff,  
  fill = vital_status)) + geom_boxplot() + stat_compare_means()
```



```
ggplot(exp2_paired, aes(x = Stage.1, y = MK2.Expression.Normal,  
  fill = vital_status)) + geom_boxplot() + stat_compare_means()
```

```
ggplot(exp2_paired, aes(x = Stage.1, y = MK2.Expression.Tumor,
  fill = vital_status)) + geom_boxplot() + stat_compare_means()
```



```
summary(coxph(Surv(time, censor) ~ MK2_FC, id = Patient,
  data = exp2_paired))
```

```
## Call:
## coxph(formula = Surv(time, censor) ~ MK2_FC, data = exp2_paired,
##       id = Patient)
##
## n= 56, number of events= 26
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## MK2_FC 0.07524    1.07814  0.52357 0.144   0.886
##
##      exp(coef) exp(-coef) lower .95 upper .95
## MK2_FC      1.078      0.9275  0.3864    3.008
##
## Concordance= 0.461 (se = 0.057 )
## Likelihood ratio test= 0.02 on 1 df,  p=0.9
## Wald test               = 0.02 on 1 df,  p=0.9
## Score (logrank) test = 0.02 on 1 df,  p=0.9
```

Hsp27 and MK2

Here, we explore correlations between Hsp27 and MK2

```
HSP27Exp <- read.csv("../rawdata/LUAD.Hsp27Exp.csv")
HSP27Exp <- left_join(HSP27Exp, expr, by = c("Patient"))
```

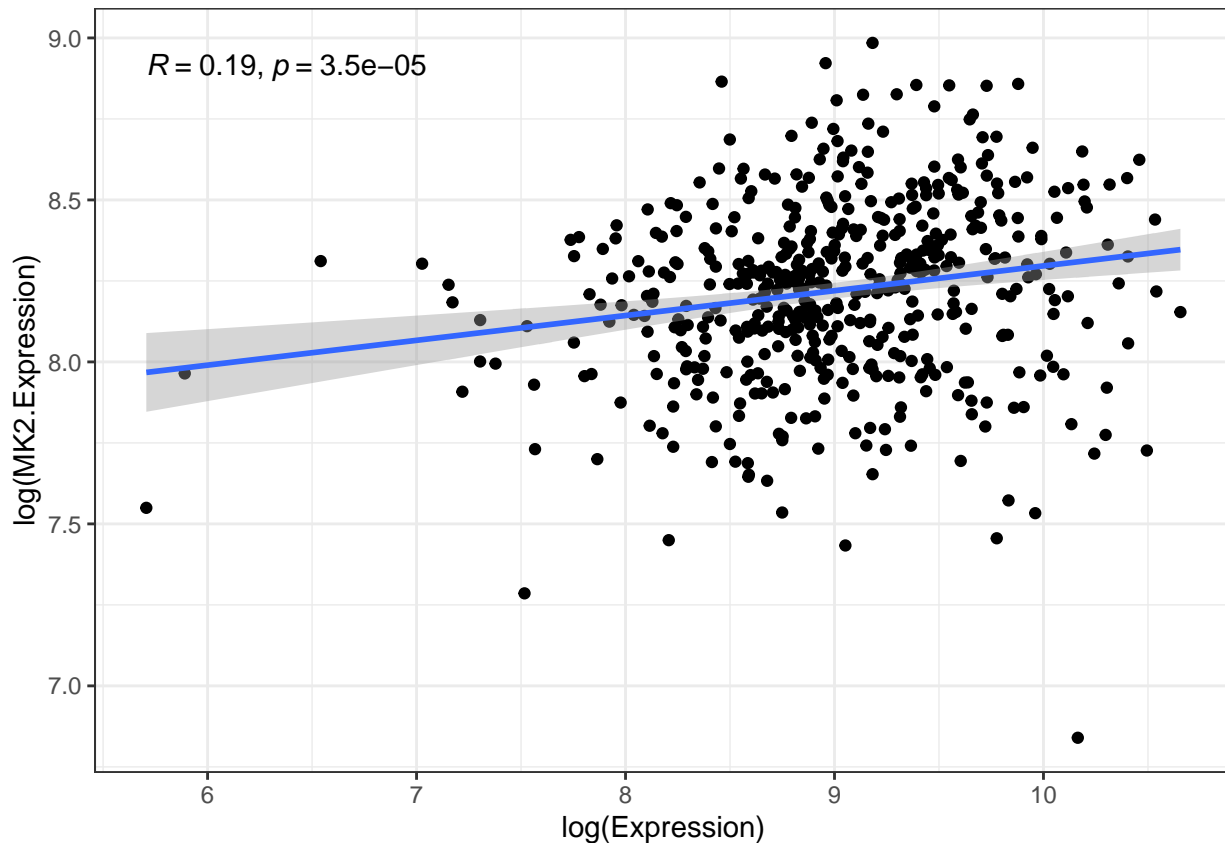
```
library(ggpubr)
ggplot(HSP27Exp, aes(x = log(Expression), y = log(MK2.Expression))) +
  geom_point() + geom_smooth(method = "lm") + theme_bw() +
  stat_cor()
```

```
## `geom_smooth()` using formula 'y ~ x'
```

```
## Warning: Removed 1 rows containing non-finite values (stat_smooth).
```

```
## Warning: Removed 1 rows containing non-finite values (stat_cor).
```

```
## Warning: Removed 1 rows containing missing values (geom_point).
```



```
# System Info
```

```
sessionInfo()
```

```
## R version 4.1.0 (2021-05-18)
## Platform: x86_64-pc-linux-gnu (64-bit)
## Running under: Ubuntu 20.04.2 LTS
##
## Matrix products: default
## BLAS: /usr/lib/x86_64-linux-gnu/blas/libblas.so.3.9.0
## LAPACK: /usr/lib/x86_64-linux-gnu/lapack/liblapack.so.3.9.0
##
## locale:
## [1] LC_CTYPE=en_US.UTF-8 LC_NUMERIC=C
## [3] LC_TIME=en_US.UTF-8 LC_COLLATE=en_US.UTF-8
## [5] LC_MONETARY=en_US.UTF-8 LC_MESSAGES=en_US.UTF-8
## [7] LC_PAPER=en_US.UTF-8 LC_NAME=C
```

```

## [9] LC_ADDRESS=C LC_TELEPHONE=C
## [11] LC_MEASUREMENT=en_US.UTF-8 LC_IDENTIFICATION=C
##
## attached base packages:
## [1] parallel stats4 stats graphics grDevices utils datasets
## [8] methods base
##
## other attached packages:
## [1] tableone_0.13.0 ggthemes_4.2.4
## [3] effects_4.2-0 carData_3.0-4
## [5] ggsci_2.9 ggpmisc_0.4.0
## [7] ggpp_0.4.0 tidyr_1.1.3
## [9] survminer_0.4.9 ggpubr_0.4.0
## [11] ggplot2_3.3.5 survival_3.2-11
## [13] TCGAbiolinks_2.20.0 DT_0.18
## [15] dplyr_1.0.7 SummarizedExperiment_1.22.0
## [17] Biobase_2.52.0 GenomicRanges_1.44.0
## [19] GenomeInfoDb_1.28.1 IRanges_2.26.0
## [21] S4Vectors_0.30.0 BiocGenerics_0.38.0
## [23] MatrixGenerics_1.4.3 matrixStats_0.60.1
## [25] knitr_1.33
##
## loaded via a namespace (and not attached):
## [1] readxl_1.3.1 backports_1.2.1
## [3] BiocFileCache_2.0.0 plyr_1.8.6
## [5] splines_4.1.0 digest_0.6.27
## [7] htmltools_0.5.2 fansi_0.5.0
## [9] magrittr_2.0.1 memoise_2.0.0
## [11] openxlsx_4.2.4 Biostrings_2.60.1
## [13] readr_1.4.0 R.utils_2.10.1
## [15] prettyunits_1.1.1 colorspace_2.0-2
## [17] blob_1.2.2 rvest_1.0.0
## [19] rappdirs_0.3.3 mitools_2.4
## [21] haven_2.4.1 xfun_0.25
## [23] crayon_1.4.1 RCurl_1.98-1.3
## [25] jsonlite_1.7.2 lme4_1.1-27.1
## [27] zoo_1.8-9 glue_1.4.2
## [29] gtable_0.3.0 zlibbioc_1.38.0
## [31] XVector_0.32.0 MatrixModels_0.5-0
## [33] DelayedArray_0.18.0 car_3.0-11
## [35] abind_1.4-5 SparseM_1.81
## [37] scales_1.1.1 DBI_1.1.1
## [39] rstatix_0.7.0 Rcpp_1.0.7
## [41] xtable_1.8-4 progress_1.2.2
## [43] proxy_0.4-26 foreign_0.8-81
## [45] bit_4.0.4 km.ci_0.5-2
## [47] survey_4.0 htmlwidgets_1.5.3
## [49] httr_1.4.2 ellipsis_0.3.2
## [51] pkgconfig_2.0.3 XML_3.99-0.7
## [53] R.methodsS3_1.8.1 farver_2.1.0
## [55] nnet_7.3-16 dbplyr_2.1.1
## [57] utf8_1.2.2 tidyselect_1.1.1
## [59] labeling_0.4.2 rlang_0.4.11
## [61] polynom_1.4-0 AnnotationDbi_1.54.1

```

## [63] munsell_0.5.0	cellranger_1.1.0
## [65] tools_4.1.0	cachem_1.0.6
## [67] downloader_0.4	generics_0.1.0
## [69] RSQLite_2.2.8	broom_0.7.8
## [71] evaluate_0.14	stringr_1.4.0
## [73] fastmap_1.1.0	yaml_2.2.1
## [75] bit64_4.0.5	zip_2.2.0
## [77] survMisc_0.5.5	purrr_0.3.4
## [79] KEGGREST_1.32.0	nlme_3.1-152
## [81] quantreg_5.86	formatR_1.11
## [83] R.oo_1.24.0	xml2_1.3.2
## [85] biomaRt_2.48.2	compiler_4.1.0
## [87] filelock_1.0.2	curl_4.3.2
## [89] png_0.1-7	e1071_1.7-7
## [91] ggsignif_0.6.2	tibble_3.1.4
## [93] stringi_1.7.4	highr_0.9
## [95] TCGAbiolinksGUI.data_1.12.0	forcats_0.5.1
## [97] lattice_0.20-44	Matrix_1.3-4
## [99] nloptr_1.2.2.2	KMsurv_0.1-5
## [101] vctrs_0.3.8	pillar_1.6.2
## [103] lifecycle_1.0.0	estimability_1.3
## [105] data.table_1.14.0	bitops_1.0-7
## [107] insight_0.14.2	conquer_1.0.2
## [109] R6_2.5.1	gridExtra_2.3
## [111] rio_0.5.27	boot_1.3-28
## [113] MASS_7.3-54	assertthat_0.2.1
## [115] withr_2.4.2	GenomeInfoDbData_1.2.6
## [117] mgcv_1.8-36	hms_1.1.0
## [119] labelled_2.8.0	grid_4.1.0
## [121] class_7.3-19	minqa_1.2.4
## [123] rmarkdown_2.11	