

Previous lecture: Chernoff bound.

Remark: For Chernoff bound, X is a sum of independent Poisson trials.

Question: What about dependent r.v.s?

Let us revisit Balls and Bins problem. Recall that the r.v.s related to bin loads are dependent.

3.2 Martingales.

A sequence of random variables X_0, X_1, \dots , is said to be a martingale sequence if for all $i > 0$,

$$E[X_i | X_0, \dots, X_{i-1}] = X_{i-1}. \quad E[X_i | Y_0, Y_1, \dots, Y_{i-1}] = X_{i-1}.$$

$\hookrightarrow E[X_i | X_{i-1}] = X_{i-1}.$

Example: A gambler w/ initial capital x_0 and capital after i^{th} bet X_i play a fair game s.t. expected gain/loss from each bet is zero. Then X_0, X_1, \dots is a martingale.

Conditional distribution and expectations.

Let X be a r.v. and E be any event that occurs w. nonzero probability. The conditional density function of X given E is given by $\Pr[X=x | E]$. Further E can be an event that another r.v. Y takes a particular value.

We denote the joint density function of X and Y by $p(x, y)$ and

$$\Pr[X=x | Y=y] = \frac{p(x, y)}{\Pr[Y=y]} = \frac{p(x, y)}{\sum_x p(x, y)}$$

Conditional expectation of X given $Y=y$ is

$$\mathbb{E}[X | Y=y] = \frac{\sum_x x \cdot p(x, y)}{\sum_x p(x, y)} \quad \begin{matrix} \mathbb{E}[X | Y=y] \\ = \sum x \cdot \Pr[X=x | Y=y] \end{matrix}$$

Remark: $\mathbb{E}[X | Y]$ is a r.v. that is defined to be equal to $f(Y)$ and $f(y) = \mathbb{E}[X | Y=y]$.

Lemma: $\mathbb{E}[\mathbb{E}[X | Y]] = \mathbb{E}[X]$. $\rightarrow \mathbb{E}[x_i] = \mathbb{E}[\mathbb{E}[x_i | x_0, \dots, x_{i-1}]] = \mathbb{E}[x_{i-1}] = \dots = \mathbb{E}[x_0]$.

Azuma's inequality:

Let x_0, x_1, \dots be a martingale sequence s.t for every k , $|x_k - x_{k-1}| \leq C_k$ where C_k may depend on k .

Then for all $t \geq 0$ and any $\lambda > 0$,

$$\Pr[|X_t - x_0| \geq \lambda] \leq 2 \cdot \exp\left(-\frac{\lambda^2}{2 \sum_{k=1}^t C_k^2}\right).$$

If all C_k are equal and are independent of k ,

$$\Pr[|X_t - x_0| \geq \lambda C \sqrt{t}] \leq 2 \cdot \exp\left(-\frac{\lambda^2}{2}\right).$$

$\mathbb{E}[Z] = n \cdot \left(1 - \frac{1}{n}\right)^m$.

Concentration bound for the no. of empty bins:

Let Z be the no. of empty bins after m balls are thrown into n bins. Let x_1, x_2, \dots denote the random choices made by balls $1, 2, \dots$.

Let $Z_t = \mathbb{E}[Z | x_1, \dots, x_t]$ denote the conditional expectation of r.v Z at time t .

Observation: $Z_0 = \mathbb{E}[Z]$ and $Z_m = Z$.

$$\mathbb{E}[Z] = n \left(1 - \frac{1}{n}\right)^m \sim n \cdot e^{-m/n}.$$

Claim: Z_0, Z_1, \dots form a martingale.

That is $\mathbb{E}[Z_t | x_1, \dots, x_t] = Z_{t-1}$ for all $t > 0$.

Let $Z(Y, t)$ be defined as the expectation of Z given that Y bins are empty at time t . (x_1, \dots, x_t give the value of Y at time t . Let us denote it by Y_t).

$$\begin{aligned} Z(Y, t) &= \mathbb{E}[Z | Y \text{ bins are empty at } t] \\ &= Y \left(1 - \frac{1}{n}\right)^{m-t}. \end{aligned}$$

$$Z_{t-1} = Z(Y_{t-1}, t-1) = Y_{t-1} \cdot \left(1 - \frac{1}{n}\right)^{m-t+1}.$$

Let us look at how Z_t can be computed. Firstly $Y_t = Y_{t-1}$ with a prob of $1 - \frac{Y_{t-1}}{n}$ and $Y_t = Y_{t-1} - 1$ w/p

$$\frac{Y_{t-1}}{n}$$

t^{th} ball \rightarrow w.p $1 - \frac{Y_{t-1}}{n}$ goes to nonempty bin
 \rightarrow w.p $\frac{Y_{t-1}}{n}$ goes to empty bin.

Case when $Y_t = Y_{t-1}$:

$$1 - \frac{Y_{t-1}}{n} \quad Z_t = Z(Y_t, t) = Z(Y_{t-1}, t) = Y_{t-1} \left(1 - \frac{1}{n}\right)^{m-t}$$

$$\rightarrow Z_t - Z_{t-1}$$

Case when $Y_t = Y_{t-1} - 1$:

$$\frac{Y_{t-1}}{n} \quad Z_t = Z(Y_t, t) = (Y_{t-1} - 1) \left(1 - \frac{1}{n}\right)^{m-t}$$

$$\rightarrow Z_t - Z_{t-1}$$

$$E[Z_t | Y_{t-1}] = \left(1 - \frac{Y_{t-1}}{n}\right) \cdot Y_{t-1} \cdot \left(1 - \frac{1}{n}\right)^{m-t}$$

$$+ \left(\frac{Y_{t-1}}{n}\right) (Y_{t-1} - 1) \left(1 - \frac{1}{n}\right)^{m-t}$$

$$= Y_{t-1} \left(1 - \frac{1}{n}\right)^{m-t} \left[1 - \frac{Y_{t-1}}{n} + \frac{Y_{t-1}}{n} - \frac{1}{n}\right]$$

$$= Y_{t-1} \left(1 - \frac{1}{n}\right)^{m-t+1} = Z_{t-1}$$

Thus, Z_0, Z_1, \dots form a martingale.

Claim: $|Z_t - Z_{t-1}| \leq \left(1 - \frac{1}{n}\right)^{m-t}$ ✓

From this, we get that $C_t = \left(1 - \frac{1}{n}\right)^{m-t}$. Using Azuma's ineq. we get the following:

$$Pr[|z - \mathbb{E}[z]| \geq \lambda] = Pr[|z_m - z_0| \geq \lambda]$$

$$\leq 2 \cdot \exp\left(-\frac{\lambda^2}{2 \sum_{t=1}^m C_t^2}\right) \cdot \} *$$

w.p. $(1 - *)$ r.v. z is concentrated in the λ -radius of $\mathbb{E}[z]$.

Reference:

1. Kamath, Motwani, Palem, Spirakis "Tail Bounds for Occupancy and Sat. threshold conj". 1994.
2. Section 4.3 of Motwani-Raghavan.