

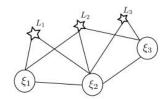
## 第四章作业提示

主讲人 会打篮球的猫





① 某时刻,SLAM 系统中相机和路标点的观测关系如下图所示,其中  $\xi$  表示相机姿态,L 表示观测到的路标点。当路标点 L 表示在世界坐标系下时,第 k 个路标被第 i 时刻的相机观测到,重投影误差为  $\mathbf{r}(\xi_i, L_k)$ 。另外,相邻相机之间存在运动约束,如 IMU 或者轮速计等约束。



- 1 请绘制上述系统的信息矩阵  $\Lambda$ .
- 2 请绘制相机  $\xi_1$  被 marg 以后的信息矩阵  $\Lambda'$ .

$$\mathbf{J}_{2} = \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}} = \begin{bmatrix} \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{1}} & \mathbf{0} & \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{3}} & \mathbf{0} & \mathbf{0} & \mathbf{0} \end{bmatrix}$$

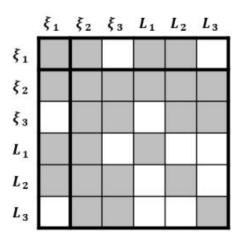
$$\mathbf{\Lambda}_{2} = \mathbf{J}_{2}^{\top} \mathbf{\Sigma}_{2}^{-1} \mathbf{J} = \begin{bmatrix} (\frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{1}})^{\top} \mathbf{\Sigma}_{2}^{-1} \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{1}} & \mathbf{0} & (\frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{1}})^{\top} \mathbf{\Sigma}_{2}^{-1} \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{3}} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ (\frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{3}})^{\top} \mathbf{\Sigma}_{2}^{-1} \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{1}} & \mathbf{0} & (\frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{3}})^{\top} \mathbf{\Sigma}_{2}^{-1} \frac{\partial \mathbf{r}_{13}}{\partial \boldsymbol{\xi}_{3}} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \end{bmatrix}$$

$$(46)$$

其实就是将系统所有约束对应的小信息矩阵累加的一个过程,只有该约束相连接的变量所对应位置才有值,其余为0。

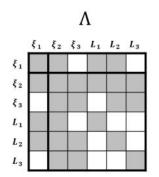


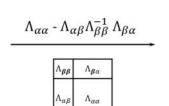
#### Λ

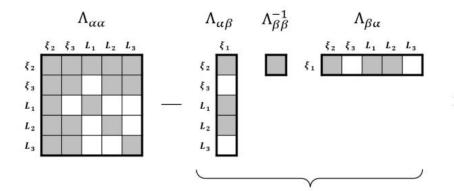


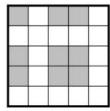
将 $\xi_1$ 进行边缘化,根据原来的系统,与 $\xi_1$ 相连的有 $\xi_1$ 、 $L_1$ 和 $L_2$ ,边缘化操作后,这三者之间便会两两相连。





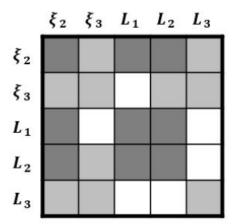








#### ۸′



可以发现,和 $\Lambda_{\alpha\alpha}$ 相比,增加了 $L_1$ 和 $L_2$ 之间的关联,信息矩阵更稠密了。



- ② 阅读《Relationship between the Hessian and Covariance Matrix for Gaussian Random Variables》. 证明信息矩阵和协方差的逆之间的关系。
  - · Hessian和协方差逆的关系

一个多元高斯随机变量的概率 密度函数的负对数函数的Hessian矩 阵,和其协方差矩阵的逆是相等的。

$$\mathbf{C}\delta\mathbf{x} = \delta\mathbf{z} \sim \mathcal{N}\left(\mathbf{0}, \mathbf{\Sigma}_{z}\right)$$

$$p\left(\mathbf{C}\delta\mathbf{x}\right) = \alpha \exp\left(\delta\mathbf{x}^{T}\mathbf{C}^{T}\mathbf{\Sigma}_{z}^{-1}\mathbf{C}\delta\mathbf{x}\right) = \alpha \exp\left(\delta\mathbf{x}^{T}\mathbf{\Sigma}_{x}^{-1}\delta\mathbf{x}\right)$$

Consider a Gaussian random vector  $\theta$  with mean  $\theta^*$  and covariance matrix  $\Sigma_{\theta}$  so its joint probability density function (PDF) is given by:

$$p(\boldsymbol{\theta}) = (2\pi)^{-\frac{N_{\boldsymbol{\theta}}}{2}} |\boldsymbol{\Sigma}_{\boldsymbol{\theta}}|^{-\frac{1}{2}} \exp\left[-\frac{1}{2} (\boldsymbol{\theta} - \boldsymbol{\theta}^{\star})^T \boldsymbol{\Sigma}_{\boldsymbol{\theta}}^{-1} (\boldsymbol{\theta} - \boldsymbol{\theta}^{\star})\right]$$
(A.1)

The objective function can be defined as its negative logarithm:

$$J(\theta) \equiv -\ln p(\theta) = \frac{N_{\theta}}{2} \ln 2\pi + \frac{1}{2} \ln |\mathbf{\Sigma}_{\theta}| + \frac{1}{2} (\theta - \theta^{\star})^{T} \mathbf{\Sigma}_{\theta}^{-1} (\theta - \theta^{\star})$$
(A.2)

which is a quadratic function of the components in  $\theta$ . By taking partial differentiations with respect to  $\theta_l$  and  $\theta_{l'}$ , the (l, l') component of the Hessian matrix can be obtained:

$$\mathcal{H}^{(l,l')}(\boldsymbol{\theta}^{\star}) = \frac{\partial^2 J(\boldsymbol{\theta})}{\partial \theta_l \partial \theta_{l'}} \bigg|_{\boldsymbol{\theta} = \boldsymbol{\theta}^{\star}} = (\boldsymbol{\Sigma}_{\boldsymbol{\theta}}^{-1})^{(l,l')} \tag{A.3}$$

so the Hessian matrix is equal to the inverse of the covariance matrix:

$$\mathcal{H}(\boldsymbol{\theta}^{\star}) = \boldsymbol{\Sigma}_{\boldsymbol{\theta}}^{-1} \tag{A.4}$$

For Gaussian random variables, the second derivatives of the objective function are constant for all  $\theta$  because the objective function is a quadratic function of  $\theta$ . Therefore, the Hessian matrix can be computed without obtaining the mean vector  $\theta^*$ .



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We can then see it as an information. The covariance of score function above is the definition of Fisher Information. As we assume  $\theta$  is a vector, the Fisher Information is in a matrix form, called Fisher Information Matrix:

$$\mathrm{F} = \mathop{\mathbb{E}}_{p(x| heta)} ig[ 
abla \log p(x| heta) \, 
abla \log p(x| heta)^{\mathrm{T}} ig] \, .$$

However, usually our likelihood function is complicated and computing the expectation is intractable. We can approximate the expectation in  ${\bf F}$  using empirical distribution  $\hat{q}(x)$ , which is given by our training data  $X=\{x_1,x_2,\cdots,x_N\}$ . In this form,  ${\bf F}$  is called Empirical Fisher:

$$\mathrm{F} = rac{1}{N} \sum_{i=1}^{N} 
abla \log p(x_i| heta) \, 
abla \log p(x_i| heta)^{\mathrm{T}} \, .$$

 ${\tt Claim:}$  The negative expected Hessian of log likelihood is equal to the Fisher Information Matrix  ${\tt F.}$ 

 ${\it Proof.}$  The Hessian of the log likelihood is given by the Jacobian of its gradient:

$$egin{aligned} & \operatorname{H}_{\log p(x| heta)} = \operatorname{J}\left(rac{
abla p(x| heta)}{p(x| heta)}
ight) \ & = rac{\operatorname{H}_{p(x| heta)}p(x| heta) - 
abla p(x| heta)}{p(x| heta)} p(x| heta) \ & = rac{\operatorname{H}_{p(x| heta)}p(x| heta)}{p(x| heta)} - rac{
abla p(x| heta)}{p(x| heta)} 
abla p(x| heta) \frac{
abla p(x| heta)}{p(x| heta)} \cdot \frac{$$



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where the second line is a result of applying quotient rule of derivative. Taking expectation wrt. our model, we have:

$$\begin{split} \underset{p(x|\theta)}{\mathbb{E}} \left[ \mathbf{H}_{\log p(x|\theta)} \right] &= \underset{p(x|\theta)}{\mathbb{E}} \left[ \frac{\mathbf{H}_{p(x|\theta)}}{p(x|\theta)} - \left( \frac{\nabla p(x|\theta)}{p(x|\theta)} \right) \left( \frac{\nabla p(x|\theta)}{p(x|\theta)} \right)^{\mathrm{T}} \right] \\ &= \underset{p(x|\theta)}{\mathbb{E}} \left[ \frac{\mathbf{H}_{p(x|\theta)}}{p(x|\theta)} \right] - \underset{p(x|\theta)}{\mathbb{E}} \left[ \left( \frac{\nabla p(x|\theta)}{p(x|\theta)} \right) \left( \frac{\nabla p(x|\theta)}{p(x|\theta)} \right)^{\mathrm{T}} \right] \\ &= \int \frac{\mathbf{H}_{p(x|\theta)}}{p(x|\theta)} p(x|\theta) \, \mathrm{d}x - \underset{p(x|\theta)}{\mathbb{E}} \left[ \nabla \log p(x|\theta) \, \nabla \log p(x|\theta)^{\mathrm{T}} \right] \\ &= \mathbf{H}_{\int p(x|\theta) \, \mathrm{d}x} - \mathbf{F} \\ &= \mathbf{H}_{1} - \mathbf{F} \\ &= -\mathbf{F} \, . \end{split}$$

Thus we have  $\mathbf{F} = -\mathbb{E}_{p(x|\theta)} \big[ \mathbf{H}_{\log p(x|\theta)} \big].$ 

Indeed knowing this result, we can see the role of  ${\bf F}$  as a measure of curvature of the log likelihood function.

似然分布的概率密度函数的负 对数函数的Hessian矩阵的期望等于 信息矩阵。



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综上所述,负对数最大似然估计的 Hessian 矩阵就是信息矩阵,同时也等于协方差的逆。所以平时往往就直接将协方差矩阵的逆当做信息矩阵。

#### 第三题



③ 请补充作业代码中单目 Bundle Adjustment 信息矩阵的计算,并输出正确的结果。正确的结果为:奇异值最后 7 维接近于 0,表明零空间的维度为 7.

```
H.block(i*6,i*6,6,6) += jacobian_Ti.transpose() * jacobian_Ti;

/// 请补充完整作业信息矩阵块的计算

H.block(poseNums*6+j*3, i*6, 3, 6) += jacobian_Pj.transpose() * jacobian_Ti;

H.block(i*6, poseNums*6+j*3, 6, 3) += jacobian_Ti.transpose() * jacobian_Pj;

H.block(poseNums*6+j*3, poseNums*6+j*3, 3, 3) += jacobian_Pj.transpose() * jacobian_Pj;
```

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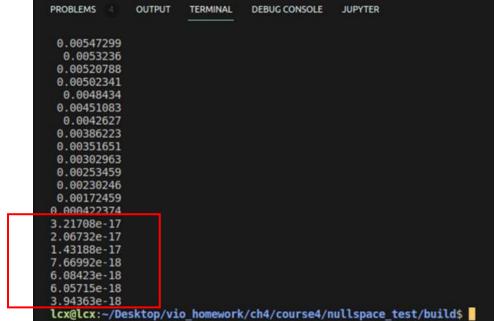
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度为 7.



在单目视觉SLAM中,信息 矩阵存在一个7维零空间,对 应7自由度的不客观量。



# 感谢各位聆听 Thanks for Listening

