

Stratified resampling ++

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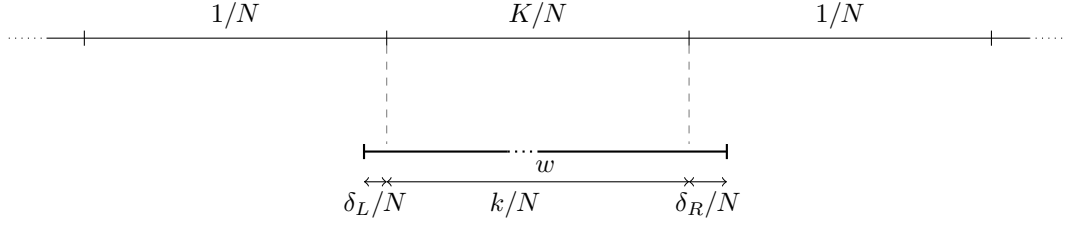
(A1) The conditional distribution of parental indices $a_t^{(1:N)}$ given offspring counts $\nu_t^{(1:N)}$ is uniform over all assignments such that $|\{j : a_t^{(j)} = i\}| = \nu_t^{(i)}$ for all i .

There are complex dependencies between the offspring counts, but we can still find some constraints on the distribution of each count conditional on the corresponding weight. Write the i^{th} weight in the form $w_t^{(i)} = (K + \delta)/N$, where $\delta \in [0, 1)$ and $K \in \{0, \dots, N - 1\}$. The distribution of $\nu_t^{(i)}$ depends not only on $w_t^{(i)}$ but also on where the i^{th} weight interval falls with respect to the length- $(1/N)$ intervals for inversion sampling. There are two cases to consider, which are illustrated in Figure 1. Note that Case (b) cannot happen if $K = 0$.

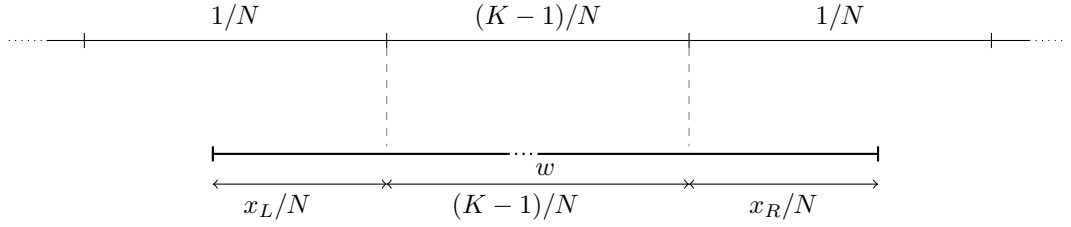
In any case $\nu_t^{(i)} \in \{k - 1, k, k + 1, k + 2\}$ almost surely. To define a probability distribution over these four values, we introduce the notation $p_j := \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + j \mid w_t^{(i)}]$, for $j = -1, 0, 1, 2$. Since the sample within each interval of length $1/N$ is uniform over that interval, we find the probabilities given in Table 1, in terms of δ and the other quantities $\delta_L, \delta_R \in [0, \delta]$ and $x_L, x_R \in [\delta, 1]$ defined in Figure 1. The probabilities do not depend on k , but of course the corresponding values of $\nu_t^{(i)}$ do. By definition $\delta_L + \delta_R = \delta$ and $x_L + x_R = 1 + \delta$.

| | Case (a) | Case (b) | L.B. | U.B. |
|----------|----------------------------------|-------------------------|--------------------|------------------|
| p_{-1} | 0 | $x_L x_R - \delta$ | 0 | $1/4$ |
| p_0 | $1 - \delta + \delta_L \delta_R$ | $1 + \delta - 2x_L x_R$ | $(1 - \delta)^2/2$ | $1 - 3\delta/4$ |
| p_1 | $\delta - 2\delta_L \delta_R$ | $x_L x_R$ | $\delta/2$ | $(1 + \delta)/2$ |
| p_2 | $\delta_L \delta_R$ | 0 | 0 | $1/4$ |

Table 1: Marginal probability distribution of $\nu_t^{(i)}$ conditional on $w_t^{(i)}$, in terms of δ and the quantities defined in Figure 1, along with upper and lower bounds on these in terms of δ only, which hold in both cases i.e. whenever $w_t^{(i)} = (K + \delta)/N$.



(a) The parent under consideration is automatically assigned K offspring, plus up to two more. ($\delta_L + \delta_R = \delta$.)



(b) This case can only occur when $K \geq 1$. The parent under consideration is automatically assigned $K - 1$ offspring, plus up to two more. ($x_L + x_R = 1 + \delta$.)

Figure 1: Cases for stratified resampling with a fixed weight $w = (K + \delta)/N$

Theorem 1. Let $\nu_t^{(1:N)}$ denote the offspring numbers in an interacting particle system satisfying (A1) such that, for any N sufficiently large, $\mathbb{P}[\tau_N(t) = \infty] = 0$ for all finite t . Suppose that there exists a deterministic sequence $(b_N)_{N \geq 1}$ such that $\lim_{N \rightarrow \infty} b_N = 0$ and

$$\frac{1}{(N)_3} \sum_{i=1}^N \mathbb{E}_t[(\nu_t^{(i)})_3] \leq b_N \frac{1}{(N)_2} \sum_{i=1}^N \mathbb{E}_t[(\nu_t^{(i)})_2] \quad (1)$$

for all N , uniformly in $t \geq 1$. Fix $n \leq N$ and consider a randomly chosen sample of n terminal particles. Then the resulting rescaled genealogical process $(G_{\tau_N(t)}^{(n,N)})_{t \geq 0}$ converges in the sense of finite-dimensional distributions to Kingman's n -coalescent as $N \rightarrow \infty$.

Corollary 2. Consider an SMC algorithm using stratified resampling, such that (A1) is satisfied. Assume that there exists a constant $a \in [1, \infty)$ such that for all x, x', t ,

$$\frac{1}{a} \leq g_t(x, x') \leq a. \quad (2)$$

Assume that $\mathbb{P}[\tau_N(t) = \infty] = 0$ for all finite t . Let $(G_t^{(n, N)})_{t \geq 0}$ denote the genealogy of a random sample of n terminal particles from the output of the algorithm when the total number of particles used is N . Then, for any fixed n , the time-scaled genealogy $(G_{\tau_N(t)}^{(n, N)})_{t \geq 0}$ converges to Kingman's n -coalescent as $N \rightarrow \infty$, in the sense of finite-dimensional distributions.

Proof. Recall that the sequence of σ -algebras

$$\mathcal{H}_t := \sigma(X_{t-1}^{(1:N)}, X_t^{(1:N)}, w_{t-1}^{(1:N)}, w_t^{(1:N)}) \quad (3)$$

are such that $\nu_t^{(1:N)}$ is conditionally independent of the filtration \mathcal{F}_{t-1} given \mathcal{H}_t . With stratified resampling, conditional on the weights each offspring count almost surely takes one of four values: $\nu_t^{(i)} \in \{\lfloor Nw_t^{(i)} \rfloor - 1, \lfloor Nw_t^{(i)} \rfloor, \lfloor Nw_t^{(i)} \rfloor + 1, \lfloor Nw_t^{(i)} \rfloor + 2\}$. Denote $p_j^{(i)} := \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + j \mid \mathcal{H}_t]$ for $j = -1, 0, 1, 2$. Now

$$\begin{aligned} \mathbb{E}[(\nu_t^{(i)})_2 \mid \mathcal{H}_t] &= p_{-1}^{(i)}(\lfloor Nw_t^{(i)} \rfloor - 1)_2 + p_0^{(i)}(\lfloor Nw_t^{(i)} \rfloor)_2 + p_1^{(i)}(\lfloor Nw_t^{(i)} \rfloor + 1)_2 \\ &\quad + p_2^{(i)}(\lfloor Nw_t^{(i)} \rfloor + 2)_2 \end{aligned}$$

and

$$\begin{aligned} \mathbb{E}[(\nu_t^{(i)})_3 \mid \mathcal{H}_t] &= p_{-1}^{(i)}(\lfloor Nw_t^{(i)} \rfloor - 1)_3 + p_0^{(i)}(\lfloor Nw_t^{(i)} \rfloor)_3 + p_1^{(i)}(\lfloor Nw_t^{(i)} \rfloor + 1)_3 \\ &\quad + p_2^{(i)}(\lfloor Nw_t^{(i)} \rfloor + 2)_3 \\ &= p_{-1}^{(i)}(\lfloor Nw_t^{(i)} \rfloor - 3)(\lfloor Nw_t^{(i)} \rfloor - 1)_2 + p_0^{(i)}(\lfloor Nw_t^{(i)} \rfloor - 2)(\lfloor Nw_t^{(i)} \rfloor)_2 \\ &\quad + p_1^{(i)}(\lfloor Nw_t^{(i)} \rfloor - 1)(\lfloor Nw_t^{(i)} \rfloor + 1)_2 + p_2^{(i)}\lfloor Nw_t^{(i)} \rfloor(\lfloor Nw_t^{(i)} \rfloor + 2)_2 \\ &\leq \lfloor Nw_t^{(i)} \rfloor \{p_0^{(i)}(\lfloor Nw_t^{(i)} \rfloor)_2 + p_1^{(i)}(\lfloor Nw_t^{(i)} \rfloor + 1)_2\} \\ &= \lfloor Nw_t^{(i)} \rfloor \mathbb{E}[(\nu_t^{(i)})_2 \mid \mathcal{H}_t] \\ &\leq a^2 \mathbb{E}[(\nu_t^{(i)})_2 \mid \mathcal{H}_t] \end{aligned}$$

The last line uses the almost sure bound $w_t^{(i)} \leq a^2/N$ which follows from (2) along with the form of the weights in Algorithm ???. Note that some terms in the above expressions may be equal to zero when $w_t^{(i)}$ is small enough, but the bound always holds nonetheless. Since the above holds for all i , applying the tower rule we have

$$\frac{1}{(N)_3} \sum_{i=1}^N \mathbb{E}_t[(\nu_t^{(i)})_3] \leq \frac{a^2}{N-2} \frac{1}{(N)_2} \sum_{i=1}^N \mathbb{E}_t[(\nu_t^{(i)})_2]$$

satisfying (1) with $b_N := a^2/(N-2) \rightarrow 0$. The result then follows by applying Theorem 1. \blacksquare

Lemma 3. Consider an SMC algorithm using stratified resampling. Suppose that

$$\varepsilon \leq q_t(x, x') \leq \varepsilon^{-1}$$

uniformly in x, x' for some $\varepsilon \in (0, 1]$, and that there exist $\zeta > 0$ and $\delta \in (0, 1)$ such that

$$\mathbb{P}[\max_i w_t^{(i)} - \min_i w_t^{(i)} \geq 2\delta/N \mid \mathcal{F}_{t-1}] \geq \zeta$$

for infinitely many t . Then, for all $N > 1$, $\mathbb{P}[\tau_N(t) = \infty] = 0$ for all finite t .

Proof. It is sufficient **by a Borel-Cantelli argument, which is written somewhere else**—SB to prove that under the stated conditions

$$\sum_{r=0}^{\infty} \mathbb{P}[c_N(r) > 2/N^2 \mid \mathcal{F}_{r-1}] = \infty.$$

Firstly,

$$\begin{aligned} \mathbb{P}[c_N(t) \leq 2/N^2 \mid \mathcal{H}_t] &= \mathbb{P}[c_N(t) = 0 \mid \mathcal{H}_t] = \mathbb{P}[\nu_t^{(i)} = 1 \forall i \in \{1, \dots, N\} \mid \mathcal{H}_t] \\ &\leq \mathbb{P}[\nu_t^{(i^*)} = 1 \mid \mathcal{H}_t], \end{aligned} \tag{4}$$

where $i^* := \operatorname{argmax}_i \{w_t^{(i)}\}$ (but note that the inequality holds when i^* is taken to be any particular index). Define for any $k \in \mathbb{Z}$

$$p_k^{(i)} := \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + k \mid \mathcal{H}_t].$$

Since stratified resampling almost surely results in $\nu_t^{(i)} \in \{\lfloor Nw_t^{(i)} \rfloor - 1, \lfloor Nw_t^{(i)} \rfloor, \lfloor Nw_t^{(i)} \rfloor + 1, \lfloor Nw_t^{(i)} \rfloor + 2\}$ we have that $p_k^{(i)} \equiv 0$ for $k \notin \{-1, 0, 1, 2\}$, and

$$\sum_{k=-1}^2 p_k^{(i)} = \sum_{k=-1}^2 \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + k \mid w_t^{(1:N)}] = 1.$$

Up to a proportionality constant C ,

$$\begin{aligned} p_k^{(i)} &= C \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + k \mid w_t^{(1:N)}] \\ &\times \sum_{\substack{a_{1:N} \in \{1, \dots, N\}^N: \\ |\{j: a_j = i\}| = \lfloor Nw_t^{(i)} \rfloor + k}} \mathbb{P}[a_t^{(1:N)} = a_{1:N} \mid \nu_t^{(i)}, w_t^{(1:N)}] \prod_{j=1}^N q_{t-1}(X_t^{(a_j)}, X_{t-1}^{(j)}) \end{aligned}$$

for each $k \in \{-1, 0, 1, 2\}$. We can bound each probability above and below using the almost sure bounds on q_{t-1} stated in the Lemma:

$$C \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + k \mid w_t^{(1:N)}] \varepsilon^N \leq p_k^{(i)} \leq C \mathbb{P}[\nu_t^{(i)} = \lfloor Nw_t^{(i)} \rfloor + k \mid w_t^{(1:N)}] \varepsilon^{-N}$$

then eliminate the constant C by normalising, to obtain lower bounds

$$\begin{aligned} p_k^{(i)} &\geq \frac{C \mathbb{P}[\nu_t^{(i)} = \lfloor N w_t^{(i)} \rfloor + k \mid w_t^{(1:N)}] \varepsilon^N}{\sum_{j=-1}^2 C \mathbb{P}[\nu_t^{(i)} = \lfloor N w_t^{(i)} \rfloor + j \mid w_t^{(1:N)}] \varepsilon^{-N}} \\ &= \mathbb{P}[\nu_t^{(i)} = \lfloor N w_t^{(i)} \rfloor + k \mid w_t^{(1:N)}] \varepsilon^{2N}. \end{aligned} \quad (5)$$

Suppose that $\max_i w_t^{(i)} - \min_i w_t^{(i)} \geq 2\delta/N$. Then that at least one of $\{\max_i w_t^{(i)} \geq (1 + \delta)/N\}$ and $\{\min_i w_t^{(i)} \leq (1 - \delta)/N\}$ occurs. We will now examine each of these possibilities.

We can always write the maximum weight as $w_t^{(i^*)} = \frac{1+\gamma}{N}$ for some $\gamma \geq 0$. Then, using (4),

$$\mathbb{P}[c_N(t) > 2/N^2 \mid \mathcal{H}_t] \geq 1 - \mathbb{P}[\nu_t^{(i^*)} = 1 \mid \mathcal{H}_t] = \begin{cases} 0 & \text{if } \gamma = 0 \\ 1 - p_0^{(i^*)} & \text{if } \gamma \in (0, 1) \\ 1 - p_{-1}^{(i^*)} & \text{if } \gamma \in [1, 2) \\ 1 & \text{if } \gamma \geq 2. \end{cases}$$

If $\gamma \in (0, 1)$ then

$$1 - p_0^{(i^*)} \geq \frac{3\gamma\varepsilon^{2N}}{4}$$

using (5) and Table 1 (p_0 , U.B.). Similarly, if $\gamma \in [1, 2)$ then by Table 1 (p_{-1} , U.B.),

$$1 - p_{-1}^{(i^*)} \geq \left(1 - \frac{1}{4}\right) \varepsilon^{2N} \geq \frac{3\varepsilon^{2N}}{4}.$$

So overall, under the constraint $\max_i w_t^{(i)} \geq (1 + \delta)/N$, we have

$$\mathbb{P}[c_N(t) > 2/N^2 \mid \mathcal{H}_t] \geq \min_{\gamma \geq \delta} \left\{ \frac{3\gamma\varepsilon^{2N}}{4} \mathbb{1}_{\{\gamma \in (0, 1)\}} + \frac{3\varepsilon^{2N}}{4} \mathbb{1}_{\{\gamma \in [1, 2)\}} + \mathbb{1}_{\{\gamma \geq 2\}} \right\} = \frac{3\delta\varepsilon^{2N}}{4}.$$

Now for the minimum weight. Let $j^* := \arg\min_i \{w_t^{(i)}\}$ and write $w_t^{(j^*)} = \frac{1-\gamma}{N}$, for some $\gamma \in [0, 1]$. Then we have

$$\mathbb{P}[c_N(t) > 2/N^2 \mid \mathcal{H}_t] \geq 1 - \mathbb{P}[\nu_t^{(j^*)} = 1 \mid \mathcal{H}_t] = \begin{cases} 1 - p_1^{(j^*)} & \text{if } \gamma \in (0, 1) \\ 0 & \text{if } \gamma = 0. \end{cases}$$

If $\gamma \in (0, 1]$ then

$$1 - p_1^{(j^*)} \geq \left(1 - \frac{1 + (1 - \gamma)}{2}\right) \varepsilon^{2N} = \frac{\gamma\varepsilon^{2N}}{2},$$

again using Table 1 (p_1 , U.B.). Therefore, under the constraint $\min_i w_t^{(i)} \leq (1 - \delta)/N$, we have

$$\mathbb{P}[c_N(t) > 2/N^2 \mid \mathcal{H}_t] \geq \min_{\gamma \geq \delta} \left\{ \frac{\gamma\varepsilon^{2N}}{2} \right\} = \frac{\delta\varepsilon^{2N}}{2}.$$

Combining both cases, we find for arbitrary r

$$\mathbb{P}[c_N(r) > 2/N^2 \mid \mathcal{H}_r] \geq \frac{\delta\varepsilon^{2N}}{2} \mathbb{1}_{\{\max_i w_r^{(i)} - \min_i w_r^{(i)} \geq 2\delta/N\}}$$

so

$$\begin{aligned}\mathbb{P}[c_N(r) > 2/N^2 \mid \mathcal{F}_{r-1}] &\geq \frac{\delta\varepsilon^{2N}}{2} \mathbb{P}[\max_i w_r^{(i)} - \min_i w_r^{(i)} \geq 2\delta/N \mid \mathcal{F}_{r-1}] \\ &\geq \zeta \frac{\delta\varepsilon^{2N}}{2} > 0\end{aligned}$$

for infinitely many r . Hence

$$\sum_{r=0}^{\infty} \mathbb{P}[c_N(r) > 2/N^2 \mid \mathcal{F}_{r-1}] = \infty$$

as required. ■