Weak convergence proof (in progress)

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Theorem 1. Let $\nu_t^{(1:N)}$ denote the offspring numbers in an interacting particle system satisfying the standing assumption and such that, for any N sufficiently large, $\mathbb{P}\{\tau_N(t)=\infty\}=0$ for all finite t. Suppose that there exists a deterministic sequence $(b_N)_{N\geq 1}$ such that $\lim_{N\to\infty}b_N=0$ and

$$\frac{1}{(N)_3} \sum_{i=1}^{N} \mathbb{E}_t \{ (\nu_t^{(i)})_3 \} \le b_N \frac{1}{(N)_2} \sum_{i=1}^{N} \mathbb{E}_t \{ (\nu_t^{(i)})_2 \}$$
 (1)

for all N, uniformly in $t \geq 1$. Then the rescaled genealogical process $(G_{\tau_N(t)}^{(n,N)})_{t\geq 0}$ converges weakly to Kingman's n-coalescent as $N \to \infty$.

Proof. Define $p_t := \max_{\xi \in E} \{1 - p_{\xi\xi}(t)\} = 1 - p_{\Delta\Delta}(t)$, where Δ denotes the trivial partition of $\{1, \ldots, n\}$ into singletons. For a proof that the maximum is attained at $\xi = \Delta$, see Lemma 1. Following Möhle (1999), we now construct the two-dimensional Markov process $(Z_t, S_t)_{t \in \mathbb{N}}$ with transition probabilities

$$\mathbb{P}(Z_{t} = j, S_{t} = \eta \mid Z_{t-1} = i, S_{t-1} = \xi) = \begin{cases}
1 - p_{t} & \text{if } j = i \text{ and } \xi = \eta \\
p_{\xi\xi}(t) + p_{t} - 1 & \text{if } j = i + 1 \text{ and } \xi = \eta \\
p_{\xi\eta}(t) & \text{if } j = i + 1 \text{ and } \xi \neq \eta \\
0 & \text{otherwise.}
\end{cases} \tag{2}$$

The construction is such that the marginal (S_t) has the same distribution as the genealogical process of interest, and (Z_t) has jumps at all the times (S_t) does plus some extra jumps. (The definition of p_t ensures that the probability in the second case is non-negative, attaining the value zero when $\xi = \Delta$.)

Denote by $0 = T_0^{(N)} < T_1^{(N)} < \dots$ the jump times of the rescaled process $(Z_{\tau_N(t)})_{t \ge 0}$, and $\omega_i^{(N)} := T_i^{(N)} - T_{i-1}^{(N)}$ the corresponding holding times $(i \in \mathbb{N})$.

Lemma 1.
$$\max_{\xi \in E} (1 - p_{\xi\xi}(t)) = 1 - p_{\Delta\Delta}(t)$$
.

Proof. Consider any $\xi \in E$ consisting of k blocks $(1 \le k \le n-1)$, and any $\xi' \in E$ consisting of k+1 blocks. From the definition of $p_{\xi\eta}(t)$ (Koskela et al., 2018, Equation (1)),

$$p_{\xi\xi}(t) = \frac{1}{(N)_k} \sum_{\substack{i_1,\dots,i_k \text{all distinct}}} \nu_t^{(i_1)} \cdots \nu_t^{(i_k)}. \tag{3}$$

Similarly,

$$p_{\xi'\xi'}(t) = \frac{1}{(N)_{k+1}} \sum_{\substack{i_1, \dots, i_k, i_{k+1} \\ \text{all distinct}}} \nu_t^{(i_1)} \cdots \nu_t^{(i_k)} \nu_t^{(i_{k+1})}$$

$$= \frac{1}{(N)_k (N-k)} \sum_{\substack{i_1, \dots, i_k \\ \text{all distinct}}} \left\{ \nu_t^{(i_1)} \cdots \nu_t^{(i_k)} \sum_{\substack{i_{k+1} = 1 \\ \text{also distinct}}}^N \nu_t^{(i_{k+1})} \right\}. \tag{4}$$

Discarding the zero summands,

$$p_{\xi'\xi'}(t) = \frac{1}{(N)_k(N-k)} \sum_{\substack{i_1,\dots,i_k \\ \text{all distinct:} \\ \nu_t^{(i_1)},\dots,\nu_t^{(i_k)} > 0}} \left\{ \nu_t^{(i_1)} \cdots \nu_t^{(i_k)} \sum_{\substack{i_{k+1}=1 \\ \text{also distinct}}}^N \nu_t^{(i_{k+1})} \right\}.$$
 (5)

The inner sum is

$$\sum_{\substack{i_{k+1}=1\\\text{also distinct}}}^{N} \nu_t^{(i_{k+1})} = \left\{ \sum_{i=1}^{N} \nu_t^{(i)} - \sum_{i \in \{i_1, \dots, i_k\}} \nu_t^{(i)} \right\} \le N - k$$
 (6)

since $\nu_t^{(i_1)}, \dots, \nu_t^{(i_k)}$ are all at least 1. Hence

$$p_{\xi'\xi'}(t) \le \frac{N-k}{(N)_k(N-k)} \sum_{\substack{i_1,\dots,i_k \\ \text{all distinct:} \\ \nu_t^{(i_1)},\dots,\nu_t^{(i_k)} > 0}} \nu_t^{(i_1)} \cdots \nu_t^{(i_k)} = p_{\xi\xi}(t). \tag{7}$$

Thus $p_{\xi\xi}(t)$ is decreasing in the number of blocks of ξ , and is therefore minimised by taking $\xi = \Delta$, which achieves the maximum n blocks. This choice in turn maximises $1 - p_{\xi\xi}(t)$, as required.

Lemma 2.

$$\lim_{N \to \infty} \mathbb{E} \left[\prod_{r=1}^{\tau_N(t)} (1 - p_r) \right] = e^{-\alpha t}$$
 (8)

where $\alpha := n(n-1)/2$.

Proof. The strategy is to find upper and lower bounds on $\mathbb{E}\left[\prod_{r=1}^{\tau_N(t)}(1-p_r)\right]$, both of which converge to $e^{-\alpha t}$.

Lower Bound

From Brown et al. (2020, Equation (14)), taking $\xi = \Delta$, we have

$$1 - p_t = p_{\Delta\Delta}(t) \ge 1 - \alpha(1 + O(N^{-1})) \left[\frac{B_n}{\alpha} D_N(t) + c_N(t) \right]$$
 (9)

where $B_n > 0$. Hence, by a multinomial expansion,

$$\prod_{r=1}^{\tau_{N}(t)} (1 - p_{r}) \ge \prod_{r=1}^{\tau_{N}(t)} \left\{ 1 - \alpha (1 + O(N^{-1})) \left[\frac{B_{n}}{\alpha} D_{N}(r) + c_{N}(r) \right] \right\}$$

$$= 1 + \sum_{k=1}^{\infty} \sum_{\substack{r_{1} < \dots < r_{k} \\ = 1}}^{\tau_{N}(t)} \prod_{j=1}^{k} \left\{ -\alpha (1 + O(N^{-1})) \left[\frac{B_{n}}{\alpha} D_{N}(r_{j}) + c_{N}(r_{j}) \right] \right\}$$

$$= 1 + \sum_{k=1}^{\infty} \left\{ -\alpha (1 + O(N^{-1})) \right\}^{k} \sum_{\substack{r_{1} < \dots < r_{k} \\ -1} < \dots < r_{k}}^{\tau_{N}(t)} \prod_{j=1}^{k} \left\{ \frac{B_{n}}{\alpha} D_{N}(r_{j}) + c_{N}(r_{j}) \right\}. \tag{10}$$

(The empty sum is defined to be zero and the empty product is defined to be one throughout.) Taking expectations,

$$\mathbb{E}\left[\prod_{r=1}^{\tau_N(t)} (1 - p_r)\right] \ge 1 + \sum_{k=1}^{\infty} \left\{-\alpha(1 + O(N^{-1}))\right\}^k \mathbb{E}\left[\sum_{\substack{r_1 < \dots < r_k \\ -1}}^{\tau_N(t)} \prod_{j=1}^k \left\{\frac{B_n}{\alpha} D_N(r_j) + c_N(r_j)\right\}\right]$$
(11)

(the infinite sum has only finitely many non-zero summands, since the inner sum is empty for $k > \tau_N(t)$, which justifies swapping the sum and expectation.) We want to show that the expectation on the right converges to $t^k/k!$, for reasons that will become clear. The strategy is to upper and lower bound this expectation by quantities that converge to $t^k/k!$.

First the lower bound. From Koskela et al. (2018, Equation (8)),

$$\sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\} \ge \sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k c_N(r_j)$$

$$\ge \frac{1}{k!} \left(\sum_{s=1}^{\tau_N(t)} c_N(s) \right)^k - \frac{1}{k!} \binom{k}{2} \left(\sum_{s=1}^{\tau_N(t)} c_N(s)^2 \right) \left(\sum_{s=1}^{\tau_N(t)} c_N(s) \right)^{k-2}$$

$$\ge \frac{1}{k!} t^k - \frac{1}{k!} \binom{k}{2} (t+1)^{k-2} \left(\sum_{s=1}^{\tau_N(t)} c_N(s)^2 \right) \tag{12}$$

by the definition of τ_N . Then

$$\mathbb{E}\left[\sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\} \right] \ge \frac{1}{k!} t^k - \frac{1}{k!} \binom{k}{2} (t+1)^{k-2} \mathbb{E}\left[\sum_{s=1}^{\tau_N(t)} c_N(s)^2\right] \longrightarrow \frac{1}{k!} t^k \tag{13}$$

as $N \to \infty$ using Brown et al. (2020, Equation (5)), which is shown to hold via lemmata 1 and 3 therein. Now for the upper bound. We start with a multinomial expansion and some manipulations of the sums:

$$\sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\} = \frac{1}{k!} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\}$$

$$= \frac{1}{k!} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}} \left(\frac{B_n}{\alpha} \right)^{k-|\mathcal{I}|} \left\{ \prod_{i \in \mathcal{I}} c_N(r_i) \right\} \left\{ \prod_{j \notin \mathcal{I}} D_N(r_j) \right\}$$

$$= \frac{1}{k!} \sum_{\mathcal{I} \subseteq \{1,\dots,k\}}^{k} \left(\frac{B_n}{\alpha} \right)^{k-|\mathcal{I}|} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i \in \mathcal{I}} c_N(r_i) \right\} \left\{ \prod_{j \in \mathcal{I}} D_N(r_j) \right\}$$

$$= \frac{1}{k!} \sum_{l=0}^{k} \binom{k}{l} \left(\frac{B_n}{\alpha} \right)^{k-l} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^{l} c_N(r_i) \right\} \left\{ \prod_{j=l+1}^{k} D_N(r_j) \right\}$$

$$= \frac{1}{k!} \sum_{l=0}^{r_N(t)} \binom{k}{l} \left(\frac{B_n}{\alpha} \right)^{k-l} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^{l} c_N(r_i) \right\} \left\{ \prod_{j=l+1}^{k} D_N(r_j) \right\}. (14)$$

Then, using that $D_N(s) \le c_N(s)$ for all s (see Lemma 3 for a proof), along with the definition of τ_N ,

$$\frac{1}{k!} \sum_{\substack{r_1 \neq \cdots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^k c_N(r_i) \right\} + \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(\frac{B_n}{\alpha} \right)^{k-I} \sum_{\substack{r_1 \neq \cdots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^I c_N(r_i) \right\} \left\{ \prod_{j=I+1}^k D_N(r_j) \right\} \\
\leq \frac{1}{k!} \left\{ \sum_{r=1}^{\tau_N(t)} c_N(r) \right\}^k + \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(\frac{B_n}{\alpha} \right)^{k-I} \sum_{\substack{r_1 \neq \cdots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^{k-1} c_N(r_i) \right\} \left\{ D_N(r_k) \right\} \\
\leq \frac{1}{k!} \left\{ t + c_N(\tau_N(t)) \right\}^k + \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(\frac{B_n}{\alpha} \right)^{k-I} \left\{ \sum_{r_1 \neq \cdots \neq r_{k-1}}^{\tau_N(t)} \prod_{i=1}^{k-1} c_N(r_i) \right\} \left\{ \sum_{r_k = 1}^{\tau_N(t)} D_N(r_k) \right\} \\
\leq \frac{1}{k!} \left\{ t + c_N(\tau_N(t)) \right\}^k + \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(\frac{B_n}{\alpha} \right)^{k-I} \left(\sum_{r=1}^{\tau_N(t)} c_N(r) \right)^{k-1} \left(\sum_{r=1}^{\tau_N(t)} D_N(r) \right) \\
\leq \frac{1}{k!} \left\{ t + c_N(\tau_N(t)) \right\}^k + \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(\frac{B_n}{\alpha} \right)^{k-I} \left(t + 1 \right)^{k-1} \left(\sum_{r=1}^{\tau_N(t)} D_N(r) \right). \tag{15}$$

Taking expectations,

$$\lim_{N \to \infty} \mathbb{E} \left[\sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\} \right] \le \frac{1}{k!} \lim_{N \to \infty} \mathbb{E} \left[\left\{ t + c_N(\tau_N(t)) \right\}^k \right]$$

$$+ \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(\frac{B_n}{\alpha} \right)^{k-I} (t+1)^{k-1} \lim_{N \to \infty} \mathbb{E} \left[\left(\sum_{r=1}^{\tau_N(t)} D_N(r) \right) \right]$$

$$= \frac{1}{k!} t^k.$$

$$(16)$$

The limit follows from Brown et al. (2020, Equations (3),(4)) along with the fact that, since $c_N(s) \in [0,1]$ for all s, $\mathbb{E}[c_N(s)^k] \leq \mathbb{E}[c_N(s)]$ for all $k \geq 1$, and the expansion

$$\mathbb{E}\left[\frac{1}{k!}\left\{t + c_N(\tau_N(t))\right\}^k\right] = \mathbb{E}\left[\frac{1}{k!}\sum_{i=0}^k \binom{k}{i} t^i c_N(\tau_N(t))^{k-i}\right] = \frac{1}{k!}\left\{t^k + kt^{k-1}\mathbb{E}[c_N(\tau_N(t))] + \dots\right\} \longrightarrow \frac{1}{k!}t^k. \quad (17)$$

Combining these upper and lower limits, we conclude that

$$\lim_{N \to \infty} \mathbb{E} \left[\sum_{\substack{r_1 < \dots < r_k \\ =1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\} \right] = \frac{1}{k!} t^k$$
 (18)

and thus

$$1 + \sum_{k=1}^{\infty} \left\{ -\alpha (1 + O(N^{-1})) \right\}^k \mathbb{E} \left[\sum_{\substack{r_1 < \dots < r_k \\ -1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ \frac{B_n}{\alpha} D_N(r_j) + c_N(r_j) \right\} \right] \longrightarrow 1 + \sum_{k=1}^{\infty} (-\alpha)^k \frac{1}{k!} t^k = e^{-\alpha t}$$
 (19)

as $N \to \infty$.

Upper Bound

From Koskela et al. (2018, Lemma 1 Case 1), taking $\xi = \Delta$, we have

$$1 - p_t = p_{\Delta\Delta}(t) \le 1 - \alpha(1 + O(N^{-1})) \left[c_N(t) - \binom{n-1}{2} D_N(t) \right]. \tag{20}$$

A multinomial expansion as before yields

$$\prod_{r=1}^{\tau_N(t)} (1 - p_r) \le 1 + \sum_{k=1}^{\infty} \left\{ -\alpha (1 + O(N^{-1})) \right\}^k \sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ c_N(r_j) - \binom{n-1}{2} D_N(r_j) \right\}. \tag{21}$$

Analogously to (13), we can write

$$\sum_{\substack{r_1 < \dots < r_k \\ = 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ c_N(r_j) - \binom{n-1}{2} D_N(r_j) \right\} = \frac{1}{k!} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^k c_N(r_i) \right\} \\
+ \frac{1}{k!} \sum_{I=0}^{k-1} \binom{k}{I} \left(-\binom{n-1}{2} \right)^{k-I} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{otherwise}}} \left\{ \prod_{i=1}^I c_N(r_i) \right\} \left\{ \prod_{j=I+1}^k D_N(r_j) \right\}.$$
(22)

We start by dealing with the second term:

$$\frac{1}{k!} \sum_{I=0}^{k-1} {k \choose I} \left(-\binom{n-1}{2} \right)^{k-I} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}} \left\{ \prod_{i=1}^{I} c_N(r_i) \right\} \left\{ \prod_{j=I+1}^{k} D_N(r_j) \right\}$$

$$= \frac{1}{k!} \sum_{\substack{I=0: \\ k-I \text{ even}}}^{k-1} {k \choose I} {n-1 \choose 2}^{k-I} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}} \left\{ \prod_{i=1}^{I} c_N(r_i) \right\} \left\{ \prod_{j=I+1}^{k} D_N(r_j) \right\}$$

$$- \frac{1}{k!} \sum_{\substack{I=0: \\ k-I \text{ odd}}}^{k-1} {k \choose I} {n-1 \choose 2}^{k-I} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{all distinct}}} \left\{ \prod_{i=1}^{I} c_N(r_i) \right\} \left\{ \prod_{j=I+1}^{k} D_N(r_j) \right\}. \quad (23)$$

This is lower bounded by

$$0 - \frac{1}{k!} \sum_{\substack{I=0\\(k-I) \text{ odd}}}^{k-1} \binom{k}{I} \binom{n-1}{2}^{k-I} (t+1)^{k-1} \left(\sum_{s=1}^{\tau_N(t)} D_N(s) \right)$$
 (24)

using that $c_N(r)$, $D_N(r) \ge 0$ for all r to bound the even terms below, and arguments from (14) to bound the odd terms above. The same arguments lead to the upper bound

$$\frac{1}{k!} \sum_{\substack{I=0\\(k-I)\text{odd}}}^{k-1} \binom{k}{I} \binom{n-1}{2}^{k-I} (t+1)^{k-1} \left(\sum_{s=1}^{\tau_N(t)} D_N(s) \right) - 0.$$
 (25)

By Brown et al. (2020, Equation (4)), both bounds have vanishing expectation as $N \to \infty$. We are left with the first term in (18), which is upper bounded by

$$\frac{1}{k!} \sum_{\substack{r_1 \neq \dots \neq r_k \\ \text{old distance}}}^{\tau_N(t)} \left\{ \prod_{i=1}^k c_N(r_i) \right\} \leq \frac{1}{k!} \left(\sum_{s=1}^{\tau_N(t)} c_N(s) \right)^k \leq \frac{1}{k!} \{ t + c_N(\tau_N(t)) \}^k$$
 (26)

the expectation of which converges to $t^k/k!$ as in (16). We use Koskela et al. (2018, Equation (8)) to construct a lower bound:

$$\frac{1}{k!} \sum_{\substack{r_1 \neq \dots \neq r_k \text{all distinct}}}^{\tau_N(t)} \left\{ \prod_{i=1}^k c_N(r_i) \right\} \ge \frac{1}{k!} \left(\sum_{s=1}^{\tau_N(t)} c_N(s) \right)^k - \frac{1}{k!} \binom{k}{2} \left(\sum_{s=1}^{\tau_N(t)} c_N(s)^2 \right) \left(\sum_{s=1}^{\tau_N(t)} c_N(s) \right)^{k-2} \\
\ge \frac{1}{k!} t^k - \frac{1}{k!} \binom{k}{2} (t+1)^{k-2} \left(\sum_{s=1}^{\tau_N(t)} c_N(s)^2 \right) \tag{27}$$

The expectation of this bound also converges to $t^k/k!$, using Brown et al. (2020, Equation (5)). We can now conclude that

$$\lim_{N \to \infty} \mathbb{E} \left[\sum_{\substack{r_1 < \dots < r_k \\ 1}}^{\tau_N(t)} \prod_{j=1}^k \left\{ c_N(r_j) - \binom{n-1}{2} D_N(r_j) \right\} \right] = \frac{1}{k!} t^k$$
 (28)

and thus

$$1 + \sum_{k=1}^{\infty} \left\{ -\alpha (1 + O(N^{-1})) \right\}^{k} \mathbb{E} \left[\sum_{\substack{r_{1} < \dots < r_{k} \\ = 1}}^{\tau_{N}(t)} \prod_{j=1}^{k} \left\{ c_{N}(r_{j}) - \binom{n-1}{2} D_{N}(r_{j}) \right\} \right] \longrightarrow 1 + \sum_{k=1}^{\infty} (-\alpha)^{k} \frac{1}{k!} t^{k} = e^{-\alpha t} \quad (29)^{k} \left[\sum_{\substack{r_{1} < \dots < r_{k} \\ = 1}}^{\tau_{N}(t)} \prod_{j=1}^{k} \left\{ c_{N}(r_{j}) - \binom{n-1}{2} D_{N}(r_{j}) \right\} \right]$$

as $N \to \infty$.

We now have upper and lower bounds on $\lim_{N\to\infty} \mathbb{E}\left[\prod_{r=1}^{\tau_N(t)}(1-p_r)\right]$, both of which are equal to $e^{-\alpha t}$, and the result follows.

Lemma 3. $D_N(t) \leq c_N(t)$ for all t.

Proof.

$$D_{N}(t) := \frac{1}{(N)_{2}} \sum_{i=1}^{N} (\nu_{t}^{(i)})_{2} \frac{1}{N} \left\{ \nu_{t}^{(i)} + \frac{1}{N} \sum_{j \neq i}^{N} (\nu_{t}^{(j)})^{2} \right\} \leq \frac{1}{(N)_{2}} \sum_{i=1}^{N} (\nu_{t}^{(i)})_{2} \frac{1}{N} \left\{ \nu_{t}^{(i)} + \frac{1}{N} \sum_{j \neq i}^{N} N \nu_{t}^{(j)} \right\}$$

$$= \frac{1}{(N)_{2}} \sum_{i=1}^{N} (\nu_{t}^{(i)})_{2} \frac{1}{N} \left\{ \sum_{j=1}^{N} \nu_{t}^{(j)} \right\} \leq \frac{1}{(N)_{2}} \sum_{i=1}^{N} (\nu_{t}^{(i)})_{2} = c_{N}(t). \tag{30}$$

References

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