

# Asymptotic Genealogies of Sequential Monte Carlo Algorithms

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26 April 2019

# State space models

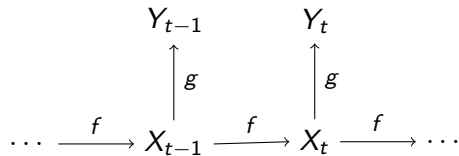
$$X_0, \dots, X_T \in \mathcal{X}$$

$$Y_0, \dots, Y_T \in \mathcal{Y}$$

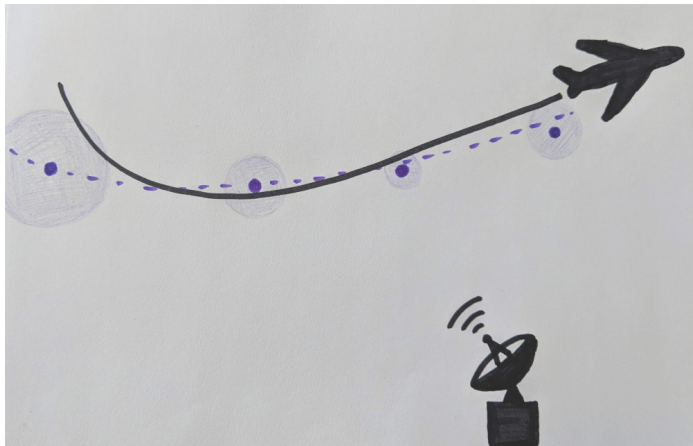
$$X_0 \sim \mu(\cdot)$$

$$X_{t+1} \mid (X_t = x_t) \sim f(\cdot \mid x_t)$$

$$Y_t \mid (X_t = x_t) \sim g(\cdot \mid x_t)$$



# Target tracking



use noisy radar data to infer position/trajectory of aircraft:

- $f$  models how aircraft moves
- $g$  models uncertainty in radar measurements

# Inference problems

**Filtering:** where is it now?  $p(x_t|y_{0:t})$

**Prediction:** where will it go next?  $p(x_{t+1}|y_{0:t})$

**Smoothing:** where has it been?  $p(x_{0:t}|y_{0:t})$

Smoothing is “harder” than filtering/prediction.

# Deterministic solutions

## Kalman filter

Under a linear Gaussian model:

$$X_0 \sim \mathcal{N}(0, \Sigma_0)$$

$$X_{t+1} \mid (X_t = x_t) \sim \mathcal{N}(Ax_t, \Sigma_x)$$

$$Y_t \mid (X_t = x_t) \sim \mathcal{N}(Bx_t, \Sigma_y)$$

we can recursively compute filtering distributions. Then a backward pass of RTS smoother provides smoothing distributions.

This class of models is rather restrictive.

## Extended Kalman filter

In non-linear Gaussian models, use a local linear approximation and apply Kalman filter.

This requires gradients and performs poorly in models that are very non-linear.

## Unscented Kalman filter

For highly non-linear Gaussian models, replace the transition step of Kalman filter by propagating a representative set of points through  $f$ .

## Deterministic solutions

- Kalman filter provides optimal filter for linear Gaussian models, but extended/unscented Kalman filter are no longer optimal.
- All of these methods require model to be Gaussian. Deterministic solutions are available for some other conjugate families, but these are still restrictive.
- Solutions are also available in the case that  $\mathcal{X}$  is finite (integrals become sums), but we will generally consider  $\mathcal{X}$  to be a continuous space.

Sequential Monte Carlo (SMC) provides general-purpose (stochastic) methods that do not require a tractable model. Only requires sampling from  $f(\cdot|x)$ , and pointwise evaluation of  $g(y|x)$  up to a normalising constant for each  $y$ .

# Sequential Monte Carlo

**Prior:**  $p(x_{0:t}) = \mu(x_0) \prod_{i=1}^t f(x_i|x_{i-1})$

**Likelihood:**  $p(y_{0:t}|x_{0:t}) = \prod_{i=0}^t g(y_i|x_i)$

**Posterior:**  $p(x_{0:t}|y_{0:t}) \propto \mu(x_0)g(y_0|x_0) \prod_{i=1}^t f(x_i|x_{i-1})g(y_i|x_i)$

- Represent posterior distribution at time  $t$  with  $N$  particles.
- Posterior factorises sequentially — avoid increase of dimension with  $T$ .

# Sequential Monte Carlo

## Algorithm

After initialisation, iterate these steps:

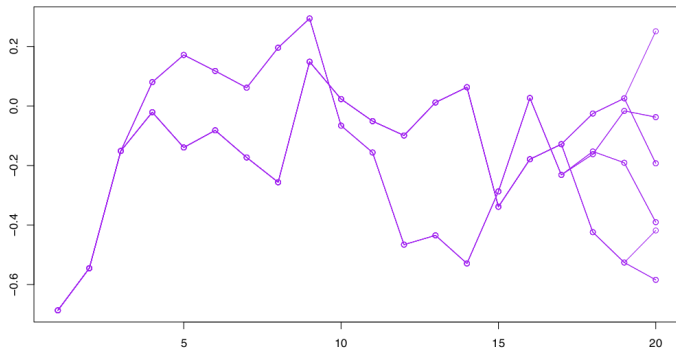
- **Propagate:** move the particles through the transition  $f$
- **Calculate weights:** weight each particle according to  $g$
- **Resample:** duplicate high-weight particles and kill off low-weight ones

Approximate posterior distribution  $p(x_{0:t}|y_{0:t})$  by the empirical measure of the particles:

$$\hat{p}(x_{0:t}|y_{0:t}) = \frac{1}{N} \sum_{i=1}^N \delta_{x_{0:t}^{(i)}}(x_{0:t})$$



# Ancestral degeneracy



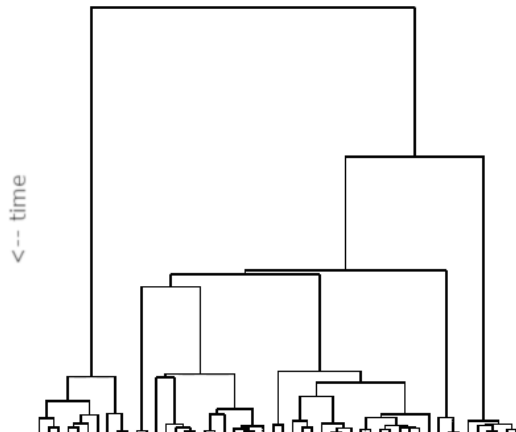
For smoothing we need a sample of trajectories.

Resampling means that trajectories of time  $T$  particles coalesce backwards in time.

# Kingman's coalescent

Looking backwards in time, each pair of lineages coalesces with unit rate.

This is the limiting coalescent process in many population models as  $N \rightarrow \infty$ .



# Coalescent for SMC

- Analysing the coalescent can help us to understand the performance of various SMC algorithms.
- The limiting coalescent of an SMC algorithm may depend on the resampling mechanism.
- For multinomial resampling, the limiting coalescent is the Kingman coalescent.
- What about other (more used) resampling schemes?

# Resampling

- The total number of particles  $N$  remains fixed.
- The particles after resampling are equally weighted.
- The resampling scheme is unbiased; that is, the expected number of offspring of each particle  $i$  is equal to  $Nw_t^{(i)}$ .

# Resampling

**Multinomial:**  $v_t^{(1:N)} \stackrel{d}{=} \text{Multinomial}(N, w_t^{(1:N)})$

**Residual:**  $v_t^{(1:N)} \stackrel{d}{=} \lfloor Nw_t^{(1:N)} \rfloor + \text{Multinomial}(R, r_t^{(1:N)} / R)$

$$r_t^{(i)} := (Nw_t^{(i)} - \lfloor Nw_t^{(i)} \rfloor)$$

$$R := \sum r_t^{(i)}$$

- Residual resampling yields lower Monte Carlo variance than multinomial resampling.
- Residual resampling is widely used by practitioners.
- Analysing the coalescent for residual resampling is a work in progress.

# Particle Gibbs

Hidden Markov model where transition depends on a parameter  $\theta$ :

$$\theta \sim \nu(\cdot)$$

$$X_0 \sim \mu(\cdot)$$

$$X_{t+1} \mid (X_t = x_t) \sim f_\theta(\cdot \mid x_t) \quad t = 0, \dots, T-1$$

$$Y_t \mid (X_t = x_t) \sim g(\cdot \mid x_t) \quad t = 0, \dots, T$$

**Gibbs sampler:** alternately sample from  $p(\theta \mid x_{0:t}, y_{0:t})$  and  $p(x_{0:t} \mid \theta, y_{0:t})$ .

- SMC is an appropriate method for sampling from  $p(x_{0:t} \mid \theta, y_{0:t})$
- To target the correct posterior distribution, need to use *conditional SMC*.

