Multivariate Statistical Analysis

Homework 2

Lucas Fellmeth, Helen Kafka, Sven Bergmann

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Problem 1

(a)

We have $X \in \mathbb{R}^p$, $\mathbb{E}[X] = \mu$, $\operatorname{Cov}(X) = \Sigma$, A is a $p \times p$ constant matrix and $\operatorname{tr}(Avv^\top) = v^\top Av$. Because A is a constant $p \times p$ matrix, A is symmetric. Because of this symmetry, it follows that it has a Cholesky decomposition as $A = C^\top C$.

Let y = CX.

Then

$$\mathbb{E}[X^{\top}AX] = \mathbb{E}[X^{\top}C^{\top}CX]$$

$$= \mathbb{E}[(CX)^{\top}CX]$$

$$= \mathbb{E}[y^{\top}y]$$

$$= \sum_{i} \mathbb{E}[y_{i}^{2}]$$

$$= \sum_{i} \operatorname{Var}(y_{i}) + \mathbb{E}[y_{i}]^{2}$$

$$= \operatorname{tr}(\Sigma_{y}) + \mu_{y}^{\top}\mu_{y}$$

where $\Sigma_y = \mathbb{E}[(y - \mathbb{E}[y])(y - \mathbb{E}[y])^{\top}] = C\Sigma C^{\top}$ and $\mu_y = C \cdot \mu$.

$$\implies \mathbb{E}[X^{\top}AX] = \operatorname{tr}(C\Sigma C^{\top}) + \mu^{\top} \underbrace{C^{\top}C}_{=A} \mu$$
$$= \operatorname{tr}(\Sigma C^{\top}C) + \mu^{\top}A\mu$$
$$= \operatorname{tr}(\Sigma A) + \mu^{\top}A\mu$$

(b)

 X_1, \ldots, X_n uncorrelated:

$$\mathrm{Cov}(X_i,X_j)=0$$
 for $i\neq j$ and $\mathrm{Cov}(X_i,X_i)=\mathrm{Var}(X_i)=\sigma^2$

$$\Longrightarrow \Sigma = \begin{pmatrix} \sigma^2 & 0 & 0 & \dots & 0 \\ 0 & \sigma^2 & 0 & \dots & 0 \\ 0 & 0 & \ddots & \dots & \vdots \\ \vdots & \vdots & \dots & \ddots & \vdots \\ 0 & 0 & \dots & \dots & \sigma^2 \end{pmatrix}, J = \begin{pmatrix} 1 & \dots & 1 \\ \vdots & \ddots & \vdots \\ 1 & \dots & 1 \end{pmatrix} \in \mathbb{R}^{p \times p}.$$

$$A = I - \frac{1}{p}J = \begin{pmatrix} 1 - \frac{1}{p} & -\frac{1}{p} & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{1}{p} & 1 - \frac{1}{p} & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{1}{p} & -\frac{1}{p} & \ddots & \dots & \vdots \\ \vdots & \vdots & \dots & \ddots & \vdots \\ -\frac{1}{p} & -\frac{1}{p} & \dots & \dots & 1 - \frac{1}{p} \end{pmatrix}$$

$$A\Sigma = \begin{pmatrix} (1 - \frac{1}{p})\sigma^2 & -\frac{\sigma^2}{p} & \dots & -\frac{\sigma^2}{p} \\ -\frac{\sigma^2}{p} & (1 - \frac{1}{p})\sigma^2 & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{\sigma^2}{p} & -\frac{\sigma^2}{p} & \ddots & \dots & \vdots \\ \vdots & \vdots & \dots & \ddots & \vdots \\ -\frac{\sigma^2}{p} & -\frac{\sigma^2}{p} & \dots & \dots & (1 - \frac{1}{p})\sigma^2 \end{pmatrix} \in \mathbb{R}^{p \times p}.$$

$$\mathbb{E}[X^{\top}AX] \stackrel{a)}{=} \operatorname{tr}(A\Sigma) + \mu^{\top}A\mu$$

$$\begin{split} &= \sum_{i=1}^{p} (1 - \frac{1}{p})\sigma^{2} + (\mu \dots \mu)^{\top} \begin{pmatrix} 1 - \frac{1}{p} & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{1}{p} & 1 - \frac{1}{p} & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{1}{p} & -\frac{1}{p} & \ddots & \dots & \vdots \\ \vdots & \vdots & \dots & \ddots & \vdots \\ -\frac{1}{p} & -\frac{1}{p} & \dots & \dots & 1 - \frac{1}{p} \end{pmatrix} \begin{pmatrix} \mu \\ \vdots \\ \mu \end{pmatrix} \\ &= p \cdot (1 - \frac{1}{p}) \cdot \sigma^{2} + (\mu \cdot (1 - \frac{1}{p}) + (p - 1) \cdot (-\frac{1}{p}) \cdot \mu \dots \mu \cdot (1 - \frac{1}{p}) + (p - 1) \cdot (-\frac{1}{p}) \cdot \mu) \begin{pmatrix} \mu \\ \vdots \\ \mu \end{pmatrix} \\ &= (p - 1) \cdot \sigma^{2} + (0 \dots 0) \cdot \begin{pmatrix} \mu \\ \vdots \\ \mu \end{pmatrix} \\ &= (p - 1) \cdot \sigma^{2} \end{split}$$

(c)

 $Cov(X_i, X_j) = \rho \sigma^2, i \neq j, Cov(X_i, X_i) = \sigma^2.$

$$\Sigma = \text{Cov}(X) = \begin{pmatrix} \sigma^2 & \rho \sigma^2 & \rho \sigma^2 & \dots & \rho \sigma^2 \\ \rho \sigma^2 & \sigma^2 & \rho \sigma^2 & \dots & \rho \sigma^2 \\ \rho \sigma^2 & \rho \sigma^2 & \ddots & \dots & \vdots \\ \vdots & \vdots & \dots & \ddots & \rho \sigma^2 \\ \rho \sigma^2 & \rho \sigma^2 & \dots & \rho \sigma^2 & \sigma^2 \end{pmatrix}$$

$$A = I - \frac{1}{p}J = \begin{pmatrix} 1 - \frac{1}{p} & -\frac{1}{p} & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{1}{p} & 1 - \frac{1}{p} & -\frac{1}{p} & \dots & -\frac{1}{p} \\ -\frac{1}{p} & -\frac{1}{p} & \ddots & \dots & \vdots \\ \vdots & \vdots & \dots & \ddots & \vdots \\ -\frac{1}{p} & -\frac{1}{p} & \dots & \dots & 1 - \frac{1}{p} \end{pmatrix} \in \mathbb{R}^{p \times p}.$$

$$\begin{split} \operatorname{tr}(A\Sigma) &= \sum_{i=1}^{p} (1 - \frac{1}{p}) \cdot \sigma^2 + (p-1) \cdot (-\frac{1}{p}) \cdot \rho \sigma^2 \\ &= p \cdot (1 - \frac{1}{p}) \cdot \sigma^2 + (p-1) \cdot (-\frac{1}{p}) \cdot \rho \sigma^2 \\ &= (p-1) \cdot \sigma^2 + (1-p)\rho \cdot \sigma^2 \\ &= (p-1) \cdot \sigma^2 + (\rho-p \cdot \rho) \cdot \sigma^2 \\ &= (p-1+\rho-p \cdot \rho) \cdot \sigma^2 \\ \mathbb{E}[X^\top AX] &= \operatorname{tr}(A\Sigma) + \underbrace{\mu^\top A \mu}_{=0 \text{ (as shown above)}} \\ &= (p-1+\rho-p \cdot \rho) \cdot \sigma^2 \end{split}$$

Problem 2

(a)

 $X \sim \mathcal{N}(\mu, \Sigma)$. We have to prove that $Z = \Sigma^{-\frac{1}{2}}(X - \mu) \sim \mathcal{N}(0, I)$.

$$\mathbb{E}[Z] = \mathbb{E}[\Sigma^{-\frac{1}{2}}(X - \mu)] = \Sigma^{-\frac{1}{2}}(\mathbb{E}[X] - \mu) = \Sigma^{-\frac{1}{2}}(\mu - \mu) = 0,$$

$$\operatorname{Var}(Z) = \operatorname{Var}(\Sigma^{-\frac{1}{2}}(X - \mu)) = (\Sigma^{-\frac{1}{2}})^{2}\operatorname{Var}(X - \mu)$$

$$= \Sigma^{-1} \cdot \operatorname{Var}(X) = \Sigma^{-1} \cdot \Sigma = I.$$

$$\implies Z \sim \mathcal{N}(\mu, \Sigma).$$

(b)

$$Z = \Sigma^{-\frac{1}{2}}(X - \mu)$$

= $\Sigma^{-\frac{1}{2}}(X - 0)$
= $\Sigma^{-\frac{1}{2}} \cdot X$.

```
Sigma <- matrix(data = c(1, -2, 0, -2, 5, 0, 0, 0, 2), nrow = 3, ncol = 3)
print(Sigma)
```

```
## [,1] [,2] [,3]
## [1,] 1 -2 0
## [2,] -2 5 0
## [3,] 0 0 2
```

Calculate the matrix square root:

```
Sigma_sqrt <- expm::sqrtm(Sigma)
print(Sigma_sqrt)</pre>
```

```
## [,1] [,2] [,3]
## [1,] 0.7071068 -0.7071068 0.000000
## [2,] -0.7071068 2.1213203 0.000000
## [3,] 0.0000000 0.0000000 1.414214
```

Check if the matrix square root times the matrix square root equals A:

```
print(Sigma_sqrt %*% Sigma_sqrt)
```

```
## [,1] [,2] [,3]
## [1,] 1 -2 0
## [2,] -2 5 0
## [3,] 0 0 2
```

So it follows:

$$Z = \begin{pmatrix} 0.7071068 \cdot x_1 + (-0.7071068) \cdot x_2 \\ (-0.7071068) \cdot x_1 + 2.1213203 \cdot x_2 \\ 1.414214 \cdot x_3 \end{pmatrix}$$

Problem 3

 $X = (X_1, X_2)$ with joint pdf:

$$f(x_1, x_2) = \begin{cases} 2\varphi(X_1)\varphi(X_2), & X_1 \cdot X_2 > 0, \\ 0, & \text{otherwise} \end{cases}$$

where

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} exp(\frac{-x}{2})$$

$$\implies f(x_1, x_2) = 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot I_{\{x_1 \cdot x_2 > 0\}}(x_1, x_2)$$

Case 1 $(x_1, x_2 > 0)$:

$$\begin{split} f_{X_1}(x_1) &= \int_0^\infty 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) dx_2 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \int_0^\infty \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) dx_2 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot \frac{1}{2} \cdot \underbrace{\int_{-\infty}^\infty \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) dx_2}_{=1 \text{ (pdf of } \mathcal{N}(0,1) \text{ over its whole support)}} \\ &= \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot I_{\{0,\infty\}}(x_1). \\ f_{X_2}(x_2) &= \int_0^\infty 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) dx_1 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \int_0^\infty \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) dx_1 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot \frac{1}{2} \cdot \underbrace{\int_{-\infty}^\infty \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) dx_1}_{=1 \text{ (pdf of } \mathcal{N}(0,1) \text{ over its whole support)}} \\ &= \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot I_{\{0,\infty\}}(x_2). \end{split}$$

Case 2 $(x_1, x_2 < 0)$:

$$\begin{split} f_{X_1}(x_1) &= \int_0^\infty 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) dx_2 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \int_{-\infty}^0 \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) dx_2 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot \frac{1}{2} \cdot \underbrace{\int_{-\infty}^\infty \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) dx_2}_{=1 \text{ (pdf of } \mathcal{N}(0,1) \text{ over its whole support)}} \\ &= \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot I_{\{-\infty,0\}}(x_1). \\ f_{X_2}(x_2) &= \int_0^\infty 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) dx_1 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \int_{-\infty}^0 \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) dx_1 \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot \frac{1}{2} \cdot \underbrace{\int_{-\infty}^\infty \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) dx_1}_{=1 \text{ (pdf of } \mathcal{N}(0,1) \text{ over its whole support)}} \\ &= \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot I_{\{-\infty,0\}}(x_2). \end{split}$$

Then

$$f_{X_1}(x_1) = \frac{1}{\sqrt{2\pi}} exp(\frac{-x_1}{2}) \cdot I_{\{-\infty,\infty\}}(x_1),$$

$$f_{X_2}(x_2) = \frac{1}{\sqrt{2\pi}} exp(\frac{-x_2}{2}) \cdot I_{\{-\infty,\infty\}}(x_2).$$

Thus, $f_{X_1}(x_1) \sim \mathcal{N}(0,1)$ and $f_{X_2}(x_2) \sim \mathcal{N}(0,1)$, but X is not multivariate Normal since the support of $f(X_1, X_2)$ is not \mathbb{R}^2 .