

1 Event History Analysis for psychological time-to-event data: A tutorial in R with examples
2 in Bayesian and frequentist workflows

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11

Abstract

12 Time-to-event data such as response times and saccade latencies form a cornerstone of
13 experimental psychology, and have had a widespread impact on our understanding of
14 human cognition. However, the orthodox method for analyzing such data – comparing
15 means between conditions – is known to conceal valuable information about the timeline of
16 psychological effects, such as their onset time and how they evolve with increasing waiting
17 time. The ability to reveal finer-grained, “temporal states” of cognitive processes can have
18 important consequences for theory development by qualitatively changing the key
19 inferences that are drawn from psychological data. Luckily, well-established analytical
20 approaches, such as event history analysis (EHA), are able to evaluate the detailed shape
21 of time-to-event distributions, and thus characterize the time course of psychological states.
22 One barrier to wider use of EHA, however, is that the analytical workflow is typically more
23 time-consuming and complex than orthodox approaches. To help achieve broader uptake of
24 EHA, in this paper we outline a set of tutorials that detail one distributional method
25 known as discrete-time EHA. We touch upon several key aspects of the workflow, such as
26 how to process raw data and specify regression models, and we also consider the
27 implications for experimental design. We finish the article by considering the benefits of
28 the approach for understanding psychological states, as well as the limitations and future
29 directions of this work. Finally, the project is written in R and freely available, which
30 means the approach can easily be adapted to other data sets.

31 *Keywords:* response times, event history analysis, Bayesian multilevel regression
32 models, experimental psychology, cognitive psychology

33 Word count: 10131 (body) + 1709 (references) + 3473 (body supplemental material)
34 + 393 (refs suppl. mat.)

35 **1. Introduction**

36 **1.1 Motivation and background context: Comparing means versus 37 distributional shapes**

38 In experimental psychology, it is standard practice to analyse response times (RTs),
39 saccade latencies, and fixation durations by calculating average performance across a series
40 of trials. Such comparisons between means have been the workhorse of experimental
41 psychology over the last century, and have had a substantial impact on theory development
42 as well as our understanding of the structure of cognition and brain function. Indeed, the
43 view that mean values are truth and variations around the mean are error is deeply
44 ingrained in experimental psychology (Bolger, Zee, Rossignac-Milon, & Hassin, 2019).

45 However, differences in mean RT conceal important pieces of information, such as when an
46 experimental effect starts, how it evolves with increasing waiting time, and whether its
47 onset is time-locked to other events (Panis, 2020; Panis, Moran, Wolkersdorfer, & Schmidt,
48 2020; Panis & Schmidt, 2016, 2022; Panis, Torfs, Gillebert, Wagemans, & Humphreys,
49 2017; Panis & Wagemans, 2009; Wolkersdorfer, Panis, & Schmidt, 2020). Such absolute
50 timing information is useful not only for the interpretation of experimental effects under
51 investigation, but also for cognitive psychophysiology and computational model selection
52 (Panis, Schmidt, Wolkersdorfer, & Schmidt, 2020).

53 As a simple illustration, Figure 1 summarises simulated single-subject data (200 trials
54 per condition) that shows how comparing means between two conditions can conceal the
55 shapes of the underlying RT and accuracy distributions. Indeed, compared to the
56 aggregation of data across trials (Figure 1A), a distributional approach offers the
57 possibility to reveal the time course of psychological states (Figure 1B). For example,
58 Figure 1B shows a first state (up to 400 ms after target onset) for which the early upswing
59 in hazard is equal for both conditions, and the emitted responses are always correct in
60 condition 1 and always incorrect in condition 2. In a second state (400 to 500 ms), hazard

- 61 is higher in condition 1, and conditional accuracies are close to .5 in both conditions. In a
62 third state (>500 ms), the effect disappears in hazard, and all conditional accuracies are
63 equal to 1 (see also Panis & Schmidt, 2016).

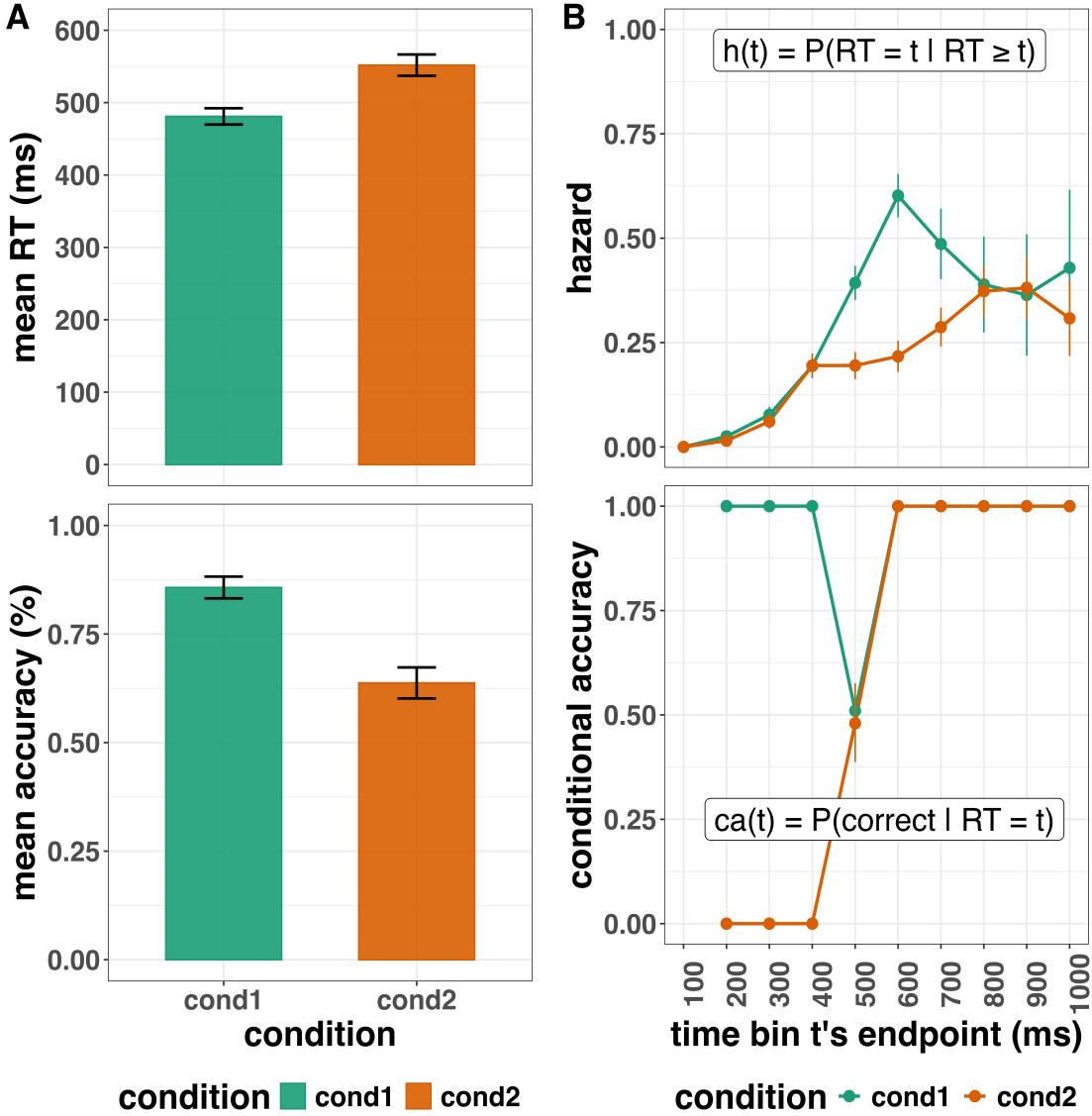


Figure 1. Simulated data showing mean performance versus distributional (EHA/SAT) analyses. (A) The mean RT (top) and overall accuracy (bottom) for two conditions are plotted. (B) The discrete-time hazard functions (top) and conditional accuracy functions (bottom) are plotted for the same data. The first second after target stimulus onset (time zero) is divided in ten bins of 100 ms. The first bin is (0,100], the last bin is (900,1000]. Note that the hazard and conditional accuracy estimates are plotted at the endpoint of each time bin. The definitions of discrete-time hazard and conditional accuracy are further explained in section 2.1.2. Error bars represent ± 1 standard error of the mean (A) or proportion (B).

64 Why does this matter for research in psychology? For many psychological questions,
65 the estimation of such “temporal states” information can be theoretically meaningful by
66 leading to more fine-grained understanding of psychological processes. Because EHA adds
67 a relatively under-used but ever-present dimension – the passage of time – to the theory
68 building toolkit, it provides one possible answer to the recent call for a temporal science of
69 behavior (Abney, Fausey, Suarez-Rivera, & Tamis-LeMonda, 2025).

70 **1.2 Aims**

71 Our ultimate aim in this paper is twofold: first, we want to convince readers of the
72 many benefits of using EHA when dealing with psychological RT data, and second, we
73 want to provide a set of practical tutorials, which provide step-by-step instructions on how
74 you actually perform a discrete-time EHA on RT data, as well as a complementary
75 discrete-time speed-accuracy tradeoff (SAT) analysis on timed accuracy data in case of
76 choice RT data (Figure 1B).

77 Even though EHA is a widely used statistical tool and there already exist many
78 excellent reviews (Allison, 1982; Blossfeld & Rohwer, 2002; Box-Steffensmeier, 2004;
79 Hosmer, Lemeshow, & May, 2011; Mills, 2011; Singer & Willett, 2003; Teachman, 1983)
80 and tutorials (Allison, 2010; Elmer, Van Duijn, Ram, & Bringmann, 2023; Landes,
81 Engelhardt, & Pelletier, 2020; Lougheed, Benson, Cole, & Ram, 2019; Stoolmiller, 2015;
82 Stoolmiller & Snyder, 2006), we are not aware of any tutorials that are aimed specifically at
83 psychological RT (+ accuracy) data, and which provide worked examples of the key data
84 processing and Bayesian multilevel regression modelling steps. Set within this context, our
85 overall aim is to introduce a set of tutorials, which explain **how** to do such analyses in the
86 context of experimental psychology, rather than repeat in any detail **why** you may do
87 them. Therefore, we hope that our tutorials will provide a pathway for research avenues in
88 experimental psychology that have the potential to benefit from using EHA in the future.

89 1.3 Structure

90 In what follows, the paper is organised in three main sections. In Section 2, we
91 provide a brief overview of EHA to orient the reader to the basic concepts that we will use
92 throughout the paper and why such an approach might be relevant for research in
93 experimental psychology. In Section 3, we outline a series of tutorials, which are written in
94 the R programming language and publicly available on our Github page
95 (https://github.com/sven-panis/Tutorial_Event_History_Analysis), along with all of the
96 other code and material associated with the project. The tutorials provide hands-on,
97 concrete examples of key parts of the analytical process, such as data wrangling, plotting
98 descriptive statistics, model fitting and planning future studies, so that others can apply
99 EHA to their own time-to-event data measured in RT tasks. In Section 4, we discuss the
100 strengths and weaknesses of the approach for researchers in experimental psychology.

101 **2. What is event history analysis and why is it relevant to research in**
102 **experimental psychology?**

103 **2.1 A brief introduction to event history analysis**

104 EHA is a class of statistical approaches to study the occurrence and timing of events,
105 such as disease onset, marriages, arrests, and job terminations (Allison, 2010). In this
106 section, we want to provide an intuition regarding how EHA works in general, as well as in
107 the context of experimental psychology. For those who want more detailed treatment of
108 EHA and/or regression equations, we refer the reader to several excellent textbooks on
109 these topics (Allison, 2010; Gelman, Hill, & Vehtari, 2020; Singer & Willett, 2003; Winter,
110 2019). We supply a few regression equations in section E of the Supplemental Material.

111 **2.1.1 Terminology and minimum requirements for EHA.** To avoid possible
112 confusion in terminology used, it is worth noting that EHA is known by various labels,
113 such as survival analysis, hazard analysis, duration analysis, failure-time analysis, and

¹¹⁴ transition analysis (Singer & Willett, 2003). In this paper, we choose to use the term EHA
¹¹⁵ throughout.

¹¹⁶ In terms of minimum requirements to apply EHA, one must be able to:

- ¹¹⁷ 1. define an event of interest that represents a qualitative change - a transition from one
¹¹⁸ discrete state to another - that can be situated in time (e.g., a button press, a
¹¹⁹ saccade onset, a fixation offset, etc.);
- ¹²⁰ 2. define time point zero in each trial (e.g., target stimulus onset, fixation onset, etc.);
- ¹²¹ 3. measure the passage of time between time point zero and event occurrence in discrete
¹²² or continuous time units in each trial.

¹²³ These minimal requirements are fulfilled by the RT data obtained in single-button
¹²⁴ detection tasks, where the time-to-response is repeatedly measured in different trials in the
¹²⁵ same individual. In section A of the Supplemental Material we visualize this and other
¹²⁶ types of time-to-event data which are typically obtained in discrimination and bistable
¹²⁷ perception tasks.

¹²⁸ **2.1.2 Types of EHA.** There are different types of modeling approaches in EHA.
¹²⁹ For example, the definition of hazard and the type of models employed depend on whether
¹³⁰ one is using continuous or discrete time units. As a lab, and mainly for practical reasons,
¹³¹ we have much more experience using discrete-time EHA, and that is the approach that we
¹³² describe and focus on in this paper. This choice may seem counter-intuitive, given that RT
¹³³ is typically treated as a continuous variable. However, continuous forms of EHA require
¹³⁴ much more data to estimate the continuous-time hazard (rate) function well (Bloxom,
¹³⁵ 1984; Luce, 1991; Van Zandt, 2000). Thus, by trading a bit of temporal resolution for a
¹³⁶ lower number of trials, discrete-time methods seem ideal for dealing with typical
¹³⁷ psychological RT data sets for which there are less than ~200 trials per condition per
¹³⁸ participant (Panis, Schmidt, et al., 2020). Moreover, as indicated by Allison (2010),

139 learning discrete-time EHA methods first will help in learning continuous-time methods, so
140 it seems like a good starting point.

141 To apply discrete-time EHA, one divides the within-trial time in discrete, contiguous
142 time bins indexed by t (e.g., $t = 1:10$ time bins; Figure 1B). Then let RT be a discrete
143 random variable denoting the rank of the time bin in which a particular person's response
144 occurs in a particular trial (i.e., repeated measure). For example, a response in one trial
145 might occur at 546 ms and it would be in time bin 6 (any RTs from 501 ms to 600 ms).
146 One then calculates the sample-based estimate of the discrete-time hazard function of
147 event occurrence for each experimental condition (Figure 1B top). The discrete-time
148 hazard function gives you, for each time bin, the conditional probability that the event
149 occurs (sometime) in bin t , given that the event does not occur in previous bins. In other
150 words, it reflects the instantaneous risk that the event occurs in the current bin t , given
151 that it has not yet occurred in the past, i.e., in one of the prior bins ($t-1, t-2, \dots, 1$).

152 In the context of experimental psychology, it is often (but not always), the case that
153 responses can be classified as correct or incorrect. In those cases, one can also calculate the
154 conditional accuracy function (Figure 1B bottom). The conditional accuracy function gives
155 you for each time bin the conditional probability that a response is correct given that it is
156 emitted in time bin t (Allison, 2010; Kantowitz & Pachella, 2021; Wickelgren, 1977). The
157 conditional accuracy function is also known as the micro-level speed-accuracy tradeoff
158 (SAT) function. We refer to this extended (hazard + conditional accuracy) analysis for
159 choice RT data as EHA/SAT. The definitions of these and other discrete-time functions are
160 given in section B of the Supplemental Material.

161 **2.2 Benefits of event history analysis for research in experimental psychology**

162 Statisticians and mathematical psychologists recommend focusing on the hazard
163 function when analyzing time-to-event data for various reasons (Holden, Van Orden, &

¹⁶⁴ Turvey, 2009; Luce, 1991; Townsend, 1990). We do not cover these benefits in detail here,
¹⁶⁵ as these are more general topics that have been covered elsewhere in textbooks (see also
¹⁶⁶ section G of the Supplemental Material). Instead, here we focus on the benefits as we see
¹⁶⁷ them for common research programmes in experimental psychology.

¹⁶⁸ We highlight three benefits that we think are relevant to the domain of experimental
¹⁶⁹ psychology. First, as illustrated in Figure 1, compared to averaging data across trials,
¹⁷⁰ integrating results between hazard functions and their associated conditional accuracy
¹⁷¹ functions for choice RT data can be informative for understanding psychological processes,
¹⁷² in terms of inferences about the microgenesis and temporal organization of cognition and
¹⁷³ theoretical development. As such, the approach permits different kinds of questions to be
¹⁷⁴ asked, different inferences to be made, and it holds the potential to discriminate between
¹⁷⁵ theoretical accounts of psychological and/or brain-based processes. For example, what kind
¹⁷⁶ of theory or set of mechanisms could account for the shape of the functions and the
¹⁷⁷ temporally localized effects reported in Figure 1B (Panis & Schmidt, 2016)? Are there new
¹⁷⁸ auxiliary assumptions that computational models need to adopt (Panis, Moran, et al.,
¹⁷⁹ 2020)? Will the temporal effect patterns align nicely with EEG findings (Panis & Schmidt,
¹⁸⁰ 2022)? And are there new experiments that need to be performed to test the novel
¹⁸¹ predictions that follow from these analyses?

¹⁸² Second, compared to more conventional analytical approaches, EHA uses more of the
¹⁸³ data because it deals with missing data differently. It is conventional with RT data to either
¹⁸⁴ (a) use a response deadline and discard all trials without a response, or (b) wait in each
¹⁸⁵ trial until a response occurs and then apply data trimming techniques, i.e., discarding too
¹⁸⁶ short or too long RTs (and perhaps also erroneous responses) before calculating a mean RT
¹⁸⁷ (Berger & Kiefer, 2021). Discarding data can introduce biases, however. Rather than treat
¹⁸⁸ non-responses as missing data, EHA treats such trials as *right-censored* observations on the
¹⁸⁹ variable RT, because all we know is that RT is greater than some value. Right-censoring is
¹⁹⁰ a type of missing data problem and a nearly universal feature of survival data including RT

191 data. For example, if the censoring time was 1 second, then some trials result in observed
192 event times (those with a RT below 1 second), while the other trials result in response
193 times that are right-censored at 1 second. The fact that EHA can deal with
194 right-censoring, therefore, presents a analytical strength of the approach compared to many
195 common approaches in experimental psychology (ANOVA, linear regression, delta plots).

196 Third, the approach is generalisable and applicable to many tasks that are commonly
197 used in experimental psychology, such as detection, discrimination and bistable perception
198 tasks, and to a range of common experimental manipulations, such as
199 stimulus-onset-asynchrony (see section A of the Supplemental Material). The upshot is
200 that one general analytical approach, which holds several potential advantages, is widely
201 applicable to many substantive use-cases in the domain of experimental psychology,
202 irrespective of the analyst's current view on the nature of cognition (Barack & Krakauer,
203 2021).

204 **2.3 Implications for research design in experimental psychology**

205 Performing EHA in experimental psychology has implications for how experiments
206 are designed. More specifically, we consider three implications that researchers will need to
207 consider when using discrete-time EHA. First, one can use a response deadline in each trial
208 because EHA deals with right-censored observations.

209 Second, since the number of trials per condition are spread across bins, it is
210 important to have a relatively large number of trial repetitions per participant and per
211 condition. Accordingly, experimental designs using this approach typically focus on
212 factorial, within-subject designs, in which a large number of observations are made on a
213 relatively small number of participants (so-called small- N designs). This approach
214 emphasizes the precision and reproducibility of data patterns at the individual participant
215 level to increase the inferential validity of the design (Baker et al., 2021; Smith & Little,

²¹⁶ 2018). Note that because statistical power derives both from the number of participants
²¹⁷ and from the number of repeated measures per participant and condition, small- N designs
²¹⁸ can still achieve what are generally considered acceptable levels of statistical power, if they
²¹⁹ have a sufficient amount of data overall (Baker et al., 2021; Smith & Little, 2018).

²²⁰ Third, the width of each time bin will need to be determined. For instance, in Figure
²²¹ 1B we chose 100 ms in an arbitrary manner. In reality, however, bin width will need to be
²²² set by considering a number of factors simultaneously. The optimal bin width will depend
²²³ on (a) the length of the observation period in each trial, (b) the rarity of event occurrence,
²²⁴ (c) the number of repeated measures (or trials) per condition per participant, and (d) the
²²⁵ shape of the hazard function. Finding an appropriate bin width in a given user case before
²²⁶ fitting models will require testing a number of options, when calculating and plotting the
²²⁷ descriptive statistics (see section 3.1). The goal is to find the smallest bin width that is
²²⁸ supported by the amount of data available. Based on our experience, a bin width of 40 ms
²²⁹ is a good starting value when the number of repeated measures is 100 or less. Too small
²³⁰ bin widths will result in erratic hazard functions as many bins will have no events, and
²³¹ thus hazard estimates of zero. Interestingly, the time bins do not need to have the same
²³² width. For example, Panis (2020) used larger bins towards the end of the observation
²³³ period, as fewer events occurred there.

²³⁴ **3. Tutorials**

²³⁵ Tutorials 1a and 1b show how to calculate and plot the descriptive statistics of
²³⁶ EHA/SAT when there are one or two independent variables, respectively. Tutorials 2a and
²³⁷ 2b illustrate how to use Bayesian multilevel modeling to fit hazard and conditional
²³⁸ accuracy models, respectively. Tutorials 3a and 3b show how to implement, respectively,
²³⁹ multilevel models for hazard and conditional accuracy in the frequentist framework.
²⁴⁰ Tutorial 4 shows how to use simulation and power analysis for planning experiments.
²⁴¹ Additionally, to further simplify the process for other users, the first two tutorials rely on a

242 set of our own custom functions that make sub-processes easier to automate, such as data
243 wrangling and plotting functions (see section C of the Supplemental Material for a list of
244 the custom functions).

245 The content of the tutorials, in terms of EHA and multilevel regression modelling, is
246 mainly based on Allison (2010), Singer and Willett (2003), McElreath (2020), Heiss (2021),
247 Kurz (2023a), and Kurz (2023b). We used R (Version 4.5.1; R Core Team, 2024) and the
248 R-packages *bayesplot* (Version 1.13.0; Gabry, Simpson, Vehtari, Betancourt, & Gelman,
249 2019), *brms* (Version 2.22.0; Bürkner, 2017, 2018, 2021), *citr* (Version 0.3.2; Aust, 2019),
250 *cmdstanr* (Version 0.9.0.9000; Gabry, Češnovar, Johnson, & Brodner, 2024), *dplyr* (Version
251 1.1.4; Wickham, François, Henry, Müller, & Vaughan, 2023), *forcats* (Version 1.0.0;
252 Wickham, 2023a), *ggplot2* (Version 3.5.2; Wickham, 2016), *lme4* (Version 1.1.37; Bates,
253 Mächler, Bolker, & Walker, 2015), *lubridate* (Version 1.9.4; Grolemund & Wickham, 2011),
254 *Matrix* (Version 1.7.3; Bates, Maechler, & Jagan, 2024), *nlme* (Version 3.1.168; Pinheiro &
255 Bates, 2000), *papaja* (Version 0.1.3; Aust & Barth, 2024), *patchwork* (Version 1.3.0;
256 Pedersen, 2024), *purrr* (Version 1.0.4; Wickham & Henry, 2023), *RColorBrewer* (Version
257 1.1.3; Neuwirth, 2022), *Rcpp* (Eddelbuettel & Balamuta, 2018; Version 1.0.14; Eddelbuettel
258 & François, 2011), *readr* (Version 2.1.5; Wickham, Hester, & Bryan, 2024), *RJ-2021-048*
259 (Bengtsson, 2021, 2021), *rstan* (Version 2.32.7; Stan Development Team, 2024), *standist*
260 (Version 0.0.0.9000; Girard, 2024), *StanHeaders* (Version 2.32.10; Stan Development Team,
261 2020), *stringr* (Version 1.5.1; Wickham, 2023b), *tibble* (Version 3.3.0; Müller & Wickham,
262 2023), *tidybayes* (Version 3.0.7; Kay, 2024), *tidyverse* (Version 2.0.0; Wickham et al., 2019) and *tinylabels* (Version 0.2.5;
263 Barth, 2023) for all reported analyses.

265 **3.1 Tutorial 1a: Calculating descriptive statistics using a life table**

266 **3.1.1 Data wrangling aims.** Our data wrangling procedures serve two related
267 purposes. First, we want to summarise and visualise descriptive statistics using a life table.

268 A life table includes for each time bin, the risk set (i.e., the number of trials that are
269 event-free at the start of the bin), the number of observed events, and the estimates of the
270 discrete-time hazard probability $h(t)$, survival probability $S(t)$, probability mass $P(t)$,
271 possibly the conditional accuracy $ca(t)$, and their estimated standard errors (se). The
272 definitions of these quantities are provided in section B of the Supplemental Material.

273 Second, we want to produce two different data sets that can each be submitted to
274 different types of inferential modelling approaches. The two types of data structure we
275 label as ‘person-trial’ data and ‘person-trial-bin’ data. The ‘person-trial’ data (Table 1)
276 will be familiar to most researchers who record behavioural responses from participants, as
277 it represents the measured RT and accuracy per trial within an experiment. This data set
278 is used when fitting conditional accuracy models (Tutorials 2b and 3b).

Table 1
Data structure for ‘person-trial’ data

pid	trial	condition	rt	accuracy
1	1	congruent	373.49	1
1	2	incongruent	431.31	1
1	3	congruent	455.43	0
1	4	incongruent	622.41	1
1	5	incongruent	535.98	1
1	6	incongruent	540.08	1
1	7	congruent	511.07	1
1	8	incongruent	444.42	1
1	9	congruent	678.69	1
1	10	congruent	549.79	1

Note. The first 10 trials for participant 1 are shown. These data are simulated and for illustrative purposes only.

279 In contrast, the ‘person-trial-bin’ data (Table 2) has a different, more extended
 280 structure, which indicates in which bin a response occurred, if at all, in each trial.
 281 Therefore, the ‘person-trial-bin’ data generates a 0 in each bin until an event occurs and
 282 then it generates a 1 to signal an event has occurred in that bin. This data set is used
 283 when fitting discrete-time hazard models (Tutorials 2a and 3a). It is worth pointing out
 284 that there is no requirement for an event to occur at all (in any bin), as maybe there was
 285 no response on that trial or the event occurred after the time window of interest. Likewise,
 286 when the event occurs in bin 1 there would only be one row of data for that trial in the
 287 person-trial-bin data set.

Table 2
Data structure for ‘person-trial-bin’ data

pid	trial	condition	timebin	event
1	1	congruent	1	0
1	1	congruent	2	0
1	1	congruent	3	0
1	1	congruent	4	1
1	2	incongruent	1	0
1	2	incongruent	2	0
1	2	incongruent	3	0
1	2	incongruent	4	0
1	2	incongruent	5	1

Note. The first 2 trials for participant 1 from Table 1 are shown. The width of the time bins is 100 ms. These data are simulated and for illustrative purposes only.

288 **3.1.2 A real data wrangling example.** To illustrate how to quickly set up life
 289 tables for calculating the descriptive statistics (functions of discrete time), we use a
 290 published data set on masked response priming from Panis and Schmidt (2016), who were
 291 interested in the temporal dynamics of the effect of prime-target congruency in RT and
 292 accuracy data. In their first experiment, Panis and Schmidt (2016) presented a double
 293 arrow for 94 ms that pointed left or right as the target stimulus with an onset at time
 294 point zero in each trial. Participants had to indicate the direction in which the double
 295 arrow pointed using their corresponding index finger, within 800 ms after target onset.
 296 Response time and accuracy were recorded on each trial. Prime type (blank, congruent,

297 incongruent) and mask type were manipulated across trials (i.e., repeated measures of
 298 time-to-response). Here we focus for each participant on the subset of 220 trials in which
 299 no mask was presented. The 13-ms prime stimulus was a double arrow presented 187 ms
 300 before target onset in the congruent (same direction as target) and incongruent (opposite
 301 direction as target) prime conditions.

302 There are several data wrangling steps to be taken. First, we need to load the data
 303 before we (a) supply required column names, and (b) specify the factor condition with the
 304 correct levels and labels.

305 The required column names are as follows:

- 306 • “pid”, indicating unique participant IDs;
- 307 • “trial”, indicating each unique trial per participant;
- 308 • “condition”, a factor indicating the levels of the independent variable (1, 2, ...) and
 the corresponding labels;
- 310 • “rt”, indicating the response times in ms;
- 311 • “acc”, indicating the accuracies (1/0).

312 In the code of Tutorial 1a, this is accomplished as follows.

```
data_wr<-read_csv("../Tutorial_1_descriptive_stats/data/DataExp1_6subjects_wrangled.csv")
data_wr <- data_wr %>%
  rename(pid = vp, condition = prime_type, acc = respac, trial = TrialNr) %>%
  mutate(condition = condition + 1, # original levels were 0, 1, 2.
        condition = factor(condition,
                             levels=c(1,2,3),
                             labels=c("blank","congruent","incongruent")))
```

313 Next, we can set up the life tables and plot for each condition the discrete-time hazard
 314 function $h(t)$, survivor function $S(t)$, probability mass function $P(t)$, and conditional

accuracy function `ca(t)`. To do so using a functional programming approach, one has to nest the person-trial data within participants using the `group_nest()` function, and supply a user-defined censoring time and bin width to our custom function “`censor()`”, as follows.

```
data_nested <- data_wr %>% group_nest(pid)

data_final <- data_nested %>%
  # ! user input: censoring time, and bin width
  mutate(censored = map(data, censor, 600, 40)) %>%
  # create person-trial-bin data set
  mutate(ptb_data = map(censored, ptb)) %>%
  # create life tables without ca(t)
  mutate(lifetable = map(ptb_data, setup_lt)) %>%
  # calculate ca(t)
  mutate(condacc = map(censored, calc_ca)) %>%
  # create life tables with ca(t)
  mutate(lifetable_ca = map2(lifetable, condacc, join_lt_ca)) %>%
  # create plots
  mutate(plot = map2(.x = lifetable_ca, .y = pid, plot_eha,1))
```

Note that the censoring time (here: 600 ms) should be a multiple of the bin width (here: 40 ms). The censoring time should be a time point after which no informative responses are expected anymore, in case one waits for a response in each trial. In experiments that implement a response deadline in each trial the censoring time can equal that deadline time point. Trials with a RT larger than the censoring time, or trials in which no response is emitted during the data observation period, are treated as right-censored observations in EHA. In other words, these trials are not discarded, because they contain the information that the event did not occur before the censoring time. Removing such trials before calculating the mean event time would result in underestimation of the true mean.

The person-trial-bin oriented data set is created by our custom function `ptb()`, and it

329 has one row for each time bin (of each trial) that is at risk for event occurrence. The
330 variable “event” in the person-trial-bin oriented data set indicates whether a response
331 occurs (1) or not (0) for each bin. The next steps are to set up the life table using our
332 custom function setup_lt(), calculate the conditional accuracies using our custom function
333 calc_ca(), add the ca(t) estimates to the life table using our custom function join_lt_ca(),
334 and then plot the descriptive statistics using our custom function plot_eha(). One can now
335 inspect different aspects, including the life table for a particular condition of a particular
336 subject, and a plot of the different functions for a particular participant.

337 In general, it is important to visually inspect the functions first for each participant,
338 in order to identify individuals that may not be following task instructions (e.g., a flat
339 conditional accuracy function at .5 indicates that someone is just guessing), outlying
340 individuals, and/or different groups with qualitatively different behavior. Also, to select a
341 suited bin width for model fitting, one can test and compare various bin widths in the
342 censor function, and select the smallest one that is supported by the data. Too small bin
343 widths will result in erratic hazard functions because many bins will have estimates equal
344 to zero.

345 Table 3 shows the life table for condition “blank” (no prime stimulus presented) for
346 participant 6.

Table 3

The life table for the blank prime condition of participant 6.

bin	risk_set	events	hazard	se_haz	survival	se_surv	ca	se_ca
0	220	NA	NA	NA	1.00	0.00	NA	NA
40	220	0	0.00	0.00	1.00	0.00	NA	NA
80	220	0	0.00	0.00	1.00	0.00	NA	NA
120	220	0	0.00	0.00	1.00	0.00	NA	NA
160	220	0	0.00	0.00	1.00	0.00	NA	NA
200	220	0	0.00	0.00	1.00	0.00	NA	NA
240	220	0	0.00	0.00	1.00	0.00	NA	NA
280	220	7	0.03	0.01	0.97	0.01	0.29	0.17
320	213	13	0.06	0.02	0.91	0.02	0.77	0.12
360	200	26	0.13	0.02	0.79	0.03	0.92	0.05
400	174	40	0.23	0.03	0.61	0.03	1.00	0.00
440	134	48	0.36	0.04	0.39	0.03	0.98	0.02
480	86	37	0.43	0.05	0.22	0.03	1.00	0.00
520	49	32	0.65	0.07	0.08	0.02	1.00	0.00
560	17	9	0.53	0.12	0.04	0.01	1.00	0.00
600	8	4	0.50	0.18	0.02	0.01	1.00	0.00

Note. The column named “bin” indicates the endpoint of each time bin (in ms), and includes time point zero. For example the first bin is (0,40] with the starting point excluded and the endpoint included. At time point zero, no events can occur and therefore $h(t=0)$ and $ca(t=0)$ are undefined. $se =$ standard error. $ca =$ conditional accuracy. $NA =$ undefined.

Figure 4 displays the discrete-time hazard, survivor, conditional accuracy, and

348 probability mass functions for each prime condition for participant 6. By using
 349 discrete-time hazard functions of event occurrence – in combination with conditional
 350 accuracy functions for two-choice tasks – one can provide an unbiased, time-varying, and
 351 probabilistic description of the latency and accuracy of responses based on all trials of any
 352 RT data set.

Descriptive stats for subject 6

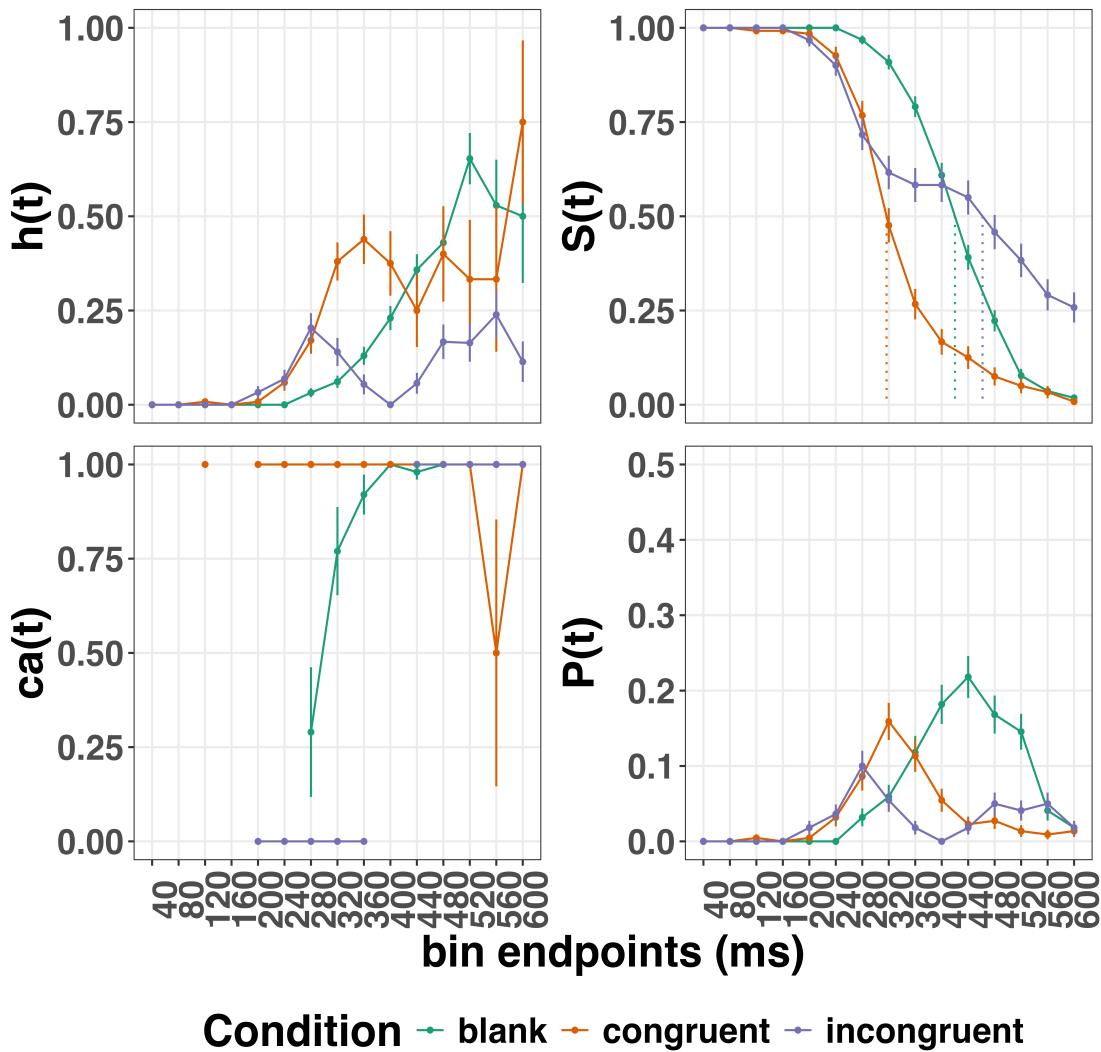


Figure 2. Estimated discrete-time hazard (h), survivor (S), conditional accuracy (ca) and probability mass (P) functions for participant 6. Vertical dotted lines indicate the estimated median RTs. Error bars represent ± 1 standard error of the respective proportion.

353 For example, for participant 6, the estimated hazard values in bin (240,280] are 0.03,

354 0.17, and 0.20 for the blank, congruent, and incongruent prime conditions, respectively. In

355 other words, when the waiting time has increased until *240 ms* after target onset, then the

356 conditional probability of response occurrence in the next 40 ms is more than five times

357 larger for both prime-present conditions, compared to the blank prime condition.

358 Furthermore, the estimated conditional accuracy values in bin (240,280] are 0.29, 1,

359 and 0 for the blank, congruent, and incongruent prime conditions, respectively. In other

360 words, if a response is emitted in bin (240,280], then the probability that it is correct is

361 estimated to be 0.29, 1, and 0 for the blank, congruent, and incongruent prime conditions,

362 respectively.

363 However, when the waiting time has increased until *400 ms* after target onset, then

364 the conditional probability of response occurrence in the next 40 ms is estimated to be

365 0.36, 0.25, and 0.06 for the blank, congruent, and incongruent prime conditions,

366 respectively. And when a response does occur in bin (400,440], then the probability that it

367 is correct is estimated to be 0.98, 1, and 1 for the blank, congruent, and incongruent prime

368 conditions, respectively.

369 These distributional results suggest that participant 6 is initially responding to the

370 prime even though (s)he was instructed to only respond to the target, that response

371 competition emerges in the incongruent prime condition around 300 ms, and that only

372 slower responses are fully controlled by the target stimulus. Qualitatively similar results

373 were obtained for the other five participants. When participants show qualitatively similar

374 distributional patterns, one might consider aggregating their data and plotting the

375 group-average distribution per condition (see Tutorial_1a.Rmd). More generally, these

376 results go against the (often implicit) assumption in research on priming that all observed

377 responses are primed responses to the target stimulus. Instead, the distributional data

378 show that fast responses are triggered exclusively by the prime stimulus, while only the

379 slower responses reflect primed responses to the target stimulus.

380 At this point, we have calculated and plotted the descriptive statistics for each type
381 of prime stimulus in EHA/SAT. As we will show in later Tutorials, statistical models for
382 hazard and conditional accuracy functions can be implemented as generalized linear mixed
383 regression models predicting event occurrence (1/0) and conditional accuracy (1/0) in each
384 bin of a selected time window for analysis. But first we consider calculating the descriptive
385 statistics for within-subject designs with two independent variables.

386 **3.2 Tutorial 1b: Generalising to a more complex design**

387 So far in this paper, we have used a simple experimental design, which involved one
388 condition with three levels. But psychological experiments are often more complex, with
389 crossed factorial designs and/or conditions with more than three levels. The purpose of
390 Tutorial 1b, therefore, is to provide a generalisation of the basic approach, which extends
391 to a more complicated design. We feel that this might be useful for researchers in
392 experimental psychology that typically use crossed factorial designs.

393 To this end, Tutorial 1b illustrates how to calculate and plot the descriptive statistics
394 for the full data set of Experiment 1 of Panis and Schmidt (2016), which includes two
395 independent variables: mask type and prime type. As we use the same functional
396 programming approach as in Tutorial 1a, we simply present the sample-based functions for
397 each participant as part of Tutorial_1b.Rmd for those that are interested.

398 **3.3 Tutorial 2a: Fitting Bayesian hazard models to interval-censored RT data**

399 In this third tutorial, we illustrate how to fit Bayesian multilevel regression models to
400 the RT data of the masked response priming data used in Tutorial 1a. Fitting (Bayesian or
401 non-Bayesian) regression models to time-to-event data is important when you want to
402 study how the shape of the hazard function depends on various predictors (Singer &

403 Willett, 2003).

404 In general, when fitting regression models, our lab adopts an estimation approach to
405 multilevel regression (Kruschke & Liddell, 2018; Winter, 2019), which is heavily influenced
406 by the Bayesian framework as suggested by Richard McElreath (Kurz, 2023b; McElreath,
407 2020). We also use a “keep it maximal” approach by specifying a full varying (or random)
408 effects structure (Barr, Levy, Scheepers, & Tily, 2013). This means that wherever possible
409 we include varying intercepts and slopes per participant. To make inferences, we use two
410 main approaches. We compare models of different complexity using information criteria
411 and cross-validation, to evaluate out-of-sample predictive accuracy (McElreath, 2020). We
412 also take the most complex model and evaluate key parameters of interest using point and
413 interval estimates.

414 **3.3.1 Hazard model considerations.** There are several analytic decisions one
415 has to make when fitting a discrete-time hazard model. First, because the first few bins
416 typically contain no responses, one has to select an analysis time window, i.e., a contiguous
417 set of bins for which there is data for each participant. Second, given that the dependent
418 variable (event occurrence) is binary, one has to select a link function (see section D of the
419 Supplemental Material). The cloglog link is preferred over the logit link when events can
420 occur in principle at any time point within a bin, which is the case for RT data (Singer &
421 Willett, 2003). Third, one has to choose whether to treat TIME (i.e., the time bin index t)
422 as a categorical or continuous predictor (see also section E of the Supplemental Material).
423 For example, when you want to know if cloglog-hazard is changing linearly or quadratically
424 over time, you should treat TIME as a continuous predictor. When you are only interested
425 in the effect of covariates on hazard, you can treat TIME as a categorical predictor (i.e., fit
426 an intercept for each bin), in which case you can choose between reference coding and
427 index coding. With reference coding, one defines the variable as a factor and selects one of
428 the k categories as the reference level. Brm() will then construct k-1 indicator variables
429 (see model M1d in Tutorial_2a.Rmd for an example). With index coding, one constructs

430 an index variable that contains integers that correspond to different categories (see models
 431 M0i and M1i below). As explained by McElreath (2020), the advantage of index coding is
 432 that the same prior can be assigned to each level of the index variable, so that each
 433 category has the same prior uncertainty.

434 In the case of a large- N design without repeated measurements, the parameters of a
 435 discrete-time hazard model can be estimated using standard logistic regression software
 436 after expanding the typical person-trial data set into a person-trial-bin data set (Allison,
 437 2010). When there is clustering in the data, as in the case of a small- N design with
 438 repeated measurements, the parameters of a discrete-time hazard model can be estimated
 439 using population-averaged methods (e.g., Generalized Estimating Equations), and Bayesian
 440 or frequentist generalized linear mixed models (Allison, 2010).

441 In general, there are three assumptions one can make or relax when adding
 442 experimental predictor variables and other covariates: The linearity assumption for
 443 continuous predictors (the effect of a 1 unit change is the same anywhere on the scale), the
 444 additivity assumption (predictors do not interact), and the proportionality assumption
 445 (predictors do not interact with TIME).

446 In tutorial_2a.Rmd we fit several Bayesian multilevel models (i.e., generalized linear
 447 mixed models) that differ in complexity to the person-trial-bin oriented data set that we
 448 created in Tutorial 1a. We decided to select the analysis time window (200,600] and the
 449 cloglog link. Below, we shortly discuss two of these models. The person-trial-bin data set is
 450 prepared as follows.

```
# read in the file we saved in tutorial 1a
ptb_data <- read_csv("Tutorial_1_descriptive_stats/data/inputfile_hazard_modeling.csv")

ptb_data <- ptb_data %>%
  # select analysis time range: (200,600] with 10 bins (time bin ranks 6 to 15)
  filter(period > 5) %>%
```

```

# define categorical predictor TIME as index variable named timebin
mutate(timebin = factor(period, levels = c(6:15)),
       # factor "condition" using reference coding, with "blank" as the reference level
       condition = factor(condition, labels = c("blank", "congruent", "incongruent")),
       # categorical predictor "prime" with index coding
       prime = ifelse(condition=="blank", 1, ifelse(condition=="congruent", 2, 3)),
       prime = factor(prime, levels = c(1,2,3)))

```

451 **3.3.2 Prior distributions.** To get the posterior distribution of each model

452 parameter given the data, we need to specify prior distributions for the model parameters
 453 which reflect our prior beliefs. In Tutorial_2a.Rmd we perform a few prior predictive
 454 checks to make sure our selected prior distributions reflect our prior beliefs (Gelman,
 455 Vehtari, et al., 2020).

456 The middle column of Supplementary Figure 3 (section F of the Supplemental
 457 Material) shows six examples of prior distributions for an intercept on the logit and/or
 458 cloglog scales. While a normal distribution with relatively large variance is often used as a
 459 weakly informative prior for continuous dependent variables, rows A and B of
 460 Supplementary Figure 3 show that specifying such distributions on the logit and cloglog
 461 scales actually leads to rather informative distributions on the original probability scale, as
 462 most mass is pushed to probabilities of 0 and 1.

463 **3.3.3 Model M0i: A null model with index coding.** When you do not want to
 464 make assumptions about the shape of the hazard function, or its shape is not smooth but
 465 irregular, then you can use a general specification of TIME, i.e., fit one grand intercept per
 466 time bin. In this first baseline or reference model, we use a general specification of TIME
 467 using index coding, and do not include experimental predictors. We call this model “M0i”.
 468 The other model (see section 3.3.4) extends model M0i by including our experimental
 469 predictor prime type.

470 Before we fit model M0i, we select the necessary columns from the data, and specify

471 our priors. In the code of Tutorial 2a, model M0i is specified as follows.

```
model_M0i <-
  brm(data = data_M0i,
       family = bernoulli(link="cloglog"),
       formula = event ~ 0 + timebin + (0 + timebin | pid),
       prior = priors_M0i,
       chains = 4, cores = 4,
       iter = 3000, warmup = 1000,
       control = list(adapt_delta = 0.999,
                      step_size = 0.04,
                      max_treedepth = 12),
       seed = 12, init = "0",
       file = "Tutorial_2_Bayesian/models/model_M0i")
```

472 After selecting the bernoulli family and the cloglog link, the model formula is
 473 specified. The specification “0 + …” removes the default intercept in brm(). The fixed
 474 effects include an intercept for each level of timebin. Each of these intercepts is allowed to
 475 vary across individuals (variable pid). We request 2000 samples from the posterior
 476 distribution for each of four chains. Estimating model M0i took about 30 minutes on a
 477 MacBook Pro (Sonoma 14.6.1 OS, 18GB Memory, M3 Pro Chip).

478 **3.3.4 Model M1i: Adding the effects of prime-target congruency.** Previous
 479 research has shown that psychological effects typically change over time (Panis, 2020;
 480 Panis, Moran, et al., 2020; Panis & Schmidt, 2022; Panis et al., 2017; Panis & Wagemans,
 481 2009). In the next model, therefore, we use index coding for both TIME (variable
 482 “timebin”) and the categorical predictor prime-target-congruency (variable “prime”), so
 483 that we get 30 grand intercepts, one for each combination of timebin level and prime level.
 484 Here is the model formula of this model that we call “M1i”.

```
event ~ 0 + timebin:prime + (0 + timebin:prime | pid)
```

485 Estimating model M1i took about 124 minutes.

486 **3.3.5 Compare the models.** There are two popular strategies to evaluate how
 487 well models will perform in predicting new data on average: Leave-One-Out (LOO)
 488 cross-validation and the Widely Applicable Information Criterion or WAIC (McElreath,
 489 2020). LOO-weights represent the optimal linear combination of models for predictive
 490 performance, with higher weights for models with better out-of-sample predictive
 491 performance. WAIC-weights represent the relative evidence for each model, with higher
 492 weights for models with a better fit while accounting for model complexity (Kurz, 2023a;
 493 McElreath, 2020).

```
model_weights(model_M0i, model_M1i, weights = "loo") %>% round(digits = 2) %>% format(nsmall=2)
```

494 ## model_M0i model_M1i

495 ## "0.00" "1.00"

```
model_weights(model_M0i, model_M1i, weights = "waic") %>% round(digits = 1) %>% format(nsmall=2)
```

496 ## model_M0i model_M1i

497 ## "0.00" "1.00"

498 Clearly, both the loo and waic weighting schemes assign a weight of 1 to model M1i,
 499 and a weight of 0 to model M0i.

500 **3.3.6 Evaluating parameter estimates in model M1i.** To make causal
 501 inferences from the parameter estimates in model M1i, we first plot the densities of the
 502 draws from the posterior distributions of its population-level parameters in Figure 5,
 503 together with point (median) and interval estimates (80% and 95% credible intervals). A

504 credible interval is a range of values that contains a parameter's true value with a specified
 505 probability, given the observed data and model.

Posterior distributions for population-level effects in Model M1i

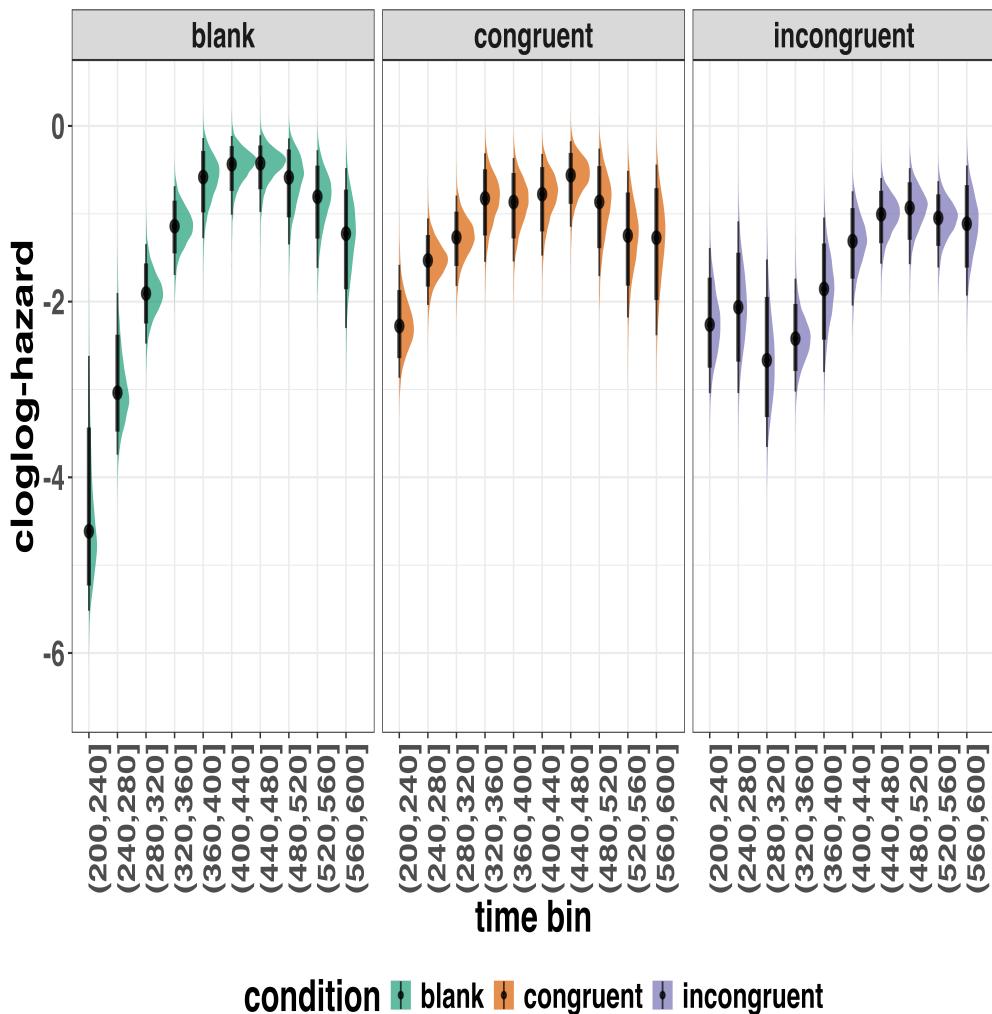


Figure 3. Medians and 80/95% credible intervals of the posterior distributions of the population-level parameters of model M1i.

506 Because the parameter estimates are on the cloglog-hazard scale, we can ease our
 507 interpretation by plotting the expected value of the posterior predictive distribution – the
 508 predicted hazard values – at the population level (Figure 6A), and for each participant in

the data set (Figure 6B).

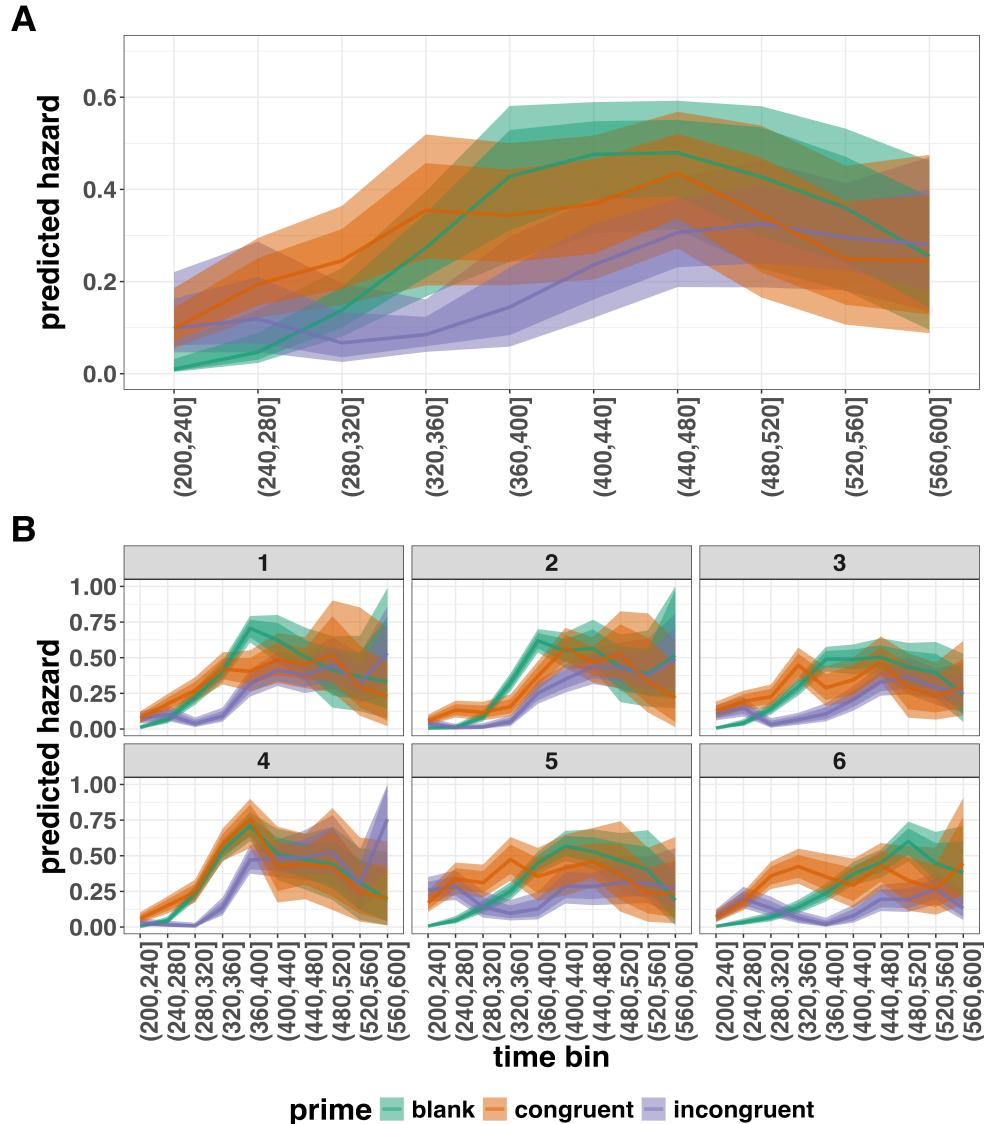


Figure 4. Point (median) and 80/95% credible interval summaries of the hazard estimates (expected values of the draws from the posterior predictive distributions) in each time bin at the population level (A), and for each participant (B).

As we are actually interested in the effects of congruent and incongruent primes, relative to the blank prime condition, we can construct two contrasts (congruent-blank, incongruent-blank), and plot the posterior distributions of these contrast effects, both at

⁵¹³ the population level (Figure 7A) and at the participant level (Figure 7B).

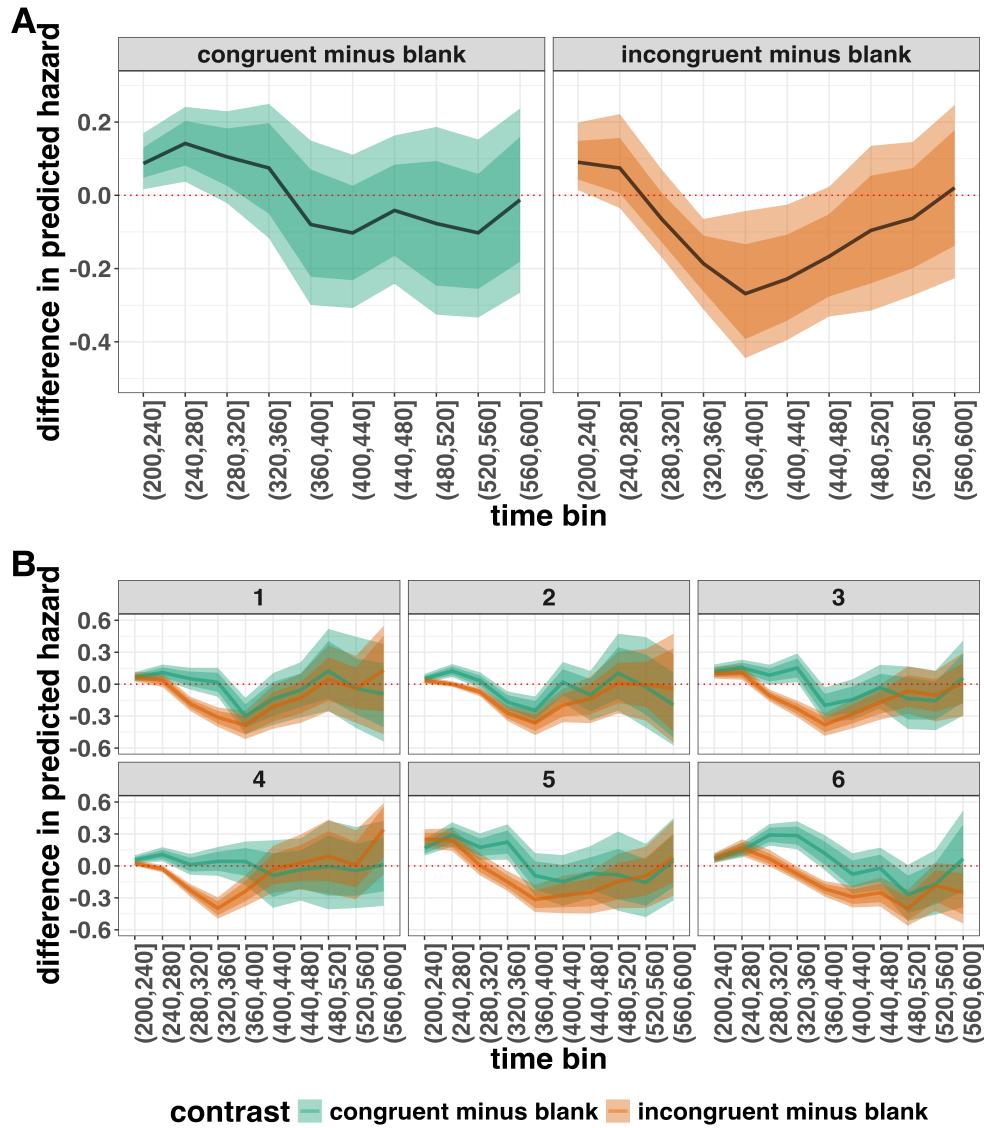


Figure 5. Point (mean) and 80/95% credible interval summaries of estimated differences in hazard in each time bin at the population level (A), and for each participant (B).

⁵¹⁴ The point estimates and quantile intervals can also be reported in a table (see
⁵¹⁵ Tutorial_2a.Rmd for details).

⁵¹⁶ **Example conclusions for M1i.** What can we conclude from model M1i about
⁵¹⁷ our research question, i.e., the temporal dynamics of the effect of prime-target congruency

518 on RT? In other words, in which of the 40-ms time bins between 200 and 600 ms after
519 target onset does changing the prime from blank to congruent or incongruent affect the
520 hazard of response occurrence (for a prime-target stimulus-onset-asynchrony of 187 ms)?

521 If we want to estimate the population-level effect of prime type on hazard, we can
522 base our conclusion on the credible Intervals (CrIs) in Figure 4A. The contrast “congruent
523 minus blank” was estimated to be 0.09 hazard units in bin (200,240] (95% CrI = [0.02,
524 0.17]), and 0.14 hazard units in bin (240,280]) (95% CrI = [0.04, 0.25]). For the other bins,
525 the 95% credible interval contained zero. The contrast “incongruent minus blank” was
526 estimated to be 0.09 hazard units in bin (200,240] (95% CrI = [0.01, 0.21]), -0.19 hazard
527 units in bin (320,360] (95% CrI = [-0.31, -0.06]), -0.27 hazard units in bin (360,400] (95%
528 CrI = [-0.45, -0.04]), and -0.23 hazard units in bin (400,440] (95% CrI = [-0.40, -0.03]). For
529 the other bins, the 95% credible interval contained zero.

530 There are thus two phases of performance for the average person between 200 and
531 600 ms after target onset. In the first phase, the addition of a congruent or incongruent
532 prime stimulus increases the hazard of response occurrence compared to blank prime trials
533 in the time period (200, 240]. In the second phase, only the incongruent prime decreases
534 the hazard of response occurrence compared to blank primes, in the time period (320,440].
535 The sign of the effect of incongruent primes on the hazard of response occurrence thus
536 depends on how much waiting time has passed since target onset.

537 If we want to focus more on inter-individual differences, we can study the
538 subject-specific hazard functions in Figure 4B. Note that three participants (1, 2, and 3)
539 show a negative difference for the contrast “congruent minus incongruent” in bin (360,400]
540 – subject 2 also in bin (320,360].

541 Interestingly, all subjects show a tendency in their mean difference (congruent minus
542 blank) to “dip” around that time (Figure 4B). Therefore, future modeling efforts could
543 incorporate the trial number into the model formula, in order to also study how the effects

544 of prime type on hazard change on the long experiment-wide time scale, next to the short
545 trial-wide time scale. In Tutorial_2a.Rmd we provide a number of model formulae that
546 should get you going.

547 **3.4 Tutorial 2b: Fitting Bayesian conditional accuracy models**

548 In this fourth tutorial, we illustrate how to fit a Bayesian multilevel regression model
549 to the timed accuracy data from the masked response priming data used in Tutorial 1a.
550 The general process is similar to Tutorial 2a, except that (a) we use the person-trial data,
551 (b) we use the symmetric logit link function, and (c) we change the priors (our prior belief
552 is that conditional accuracy values between 0 and 1 are equally likely). To keep the tutorial
553 short, we only fit one conditional accuracy model, which was based on model M1i from
554 Tutorial 2a and labelled M1i_ca.

555 To make inferences from the parameter estimates in model M1i_ca, we first plot the
556 densities of the draws from the posterior distributions of its population-level parameters in
557 Figure 8, together with point (median) and interval estimates (80% and 95% credible
558 intervals).

Posterior distributions for population-level effects in Model M1i_ca

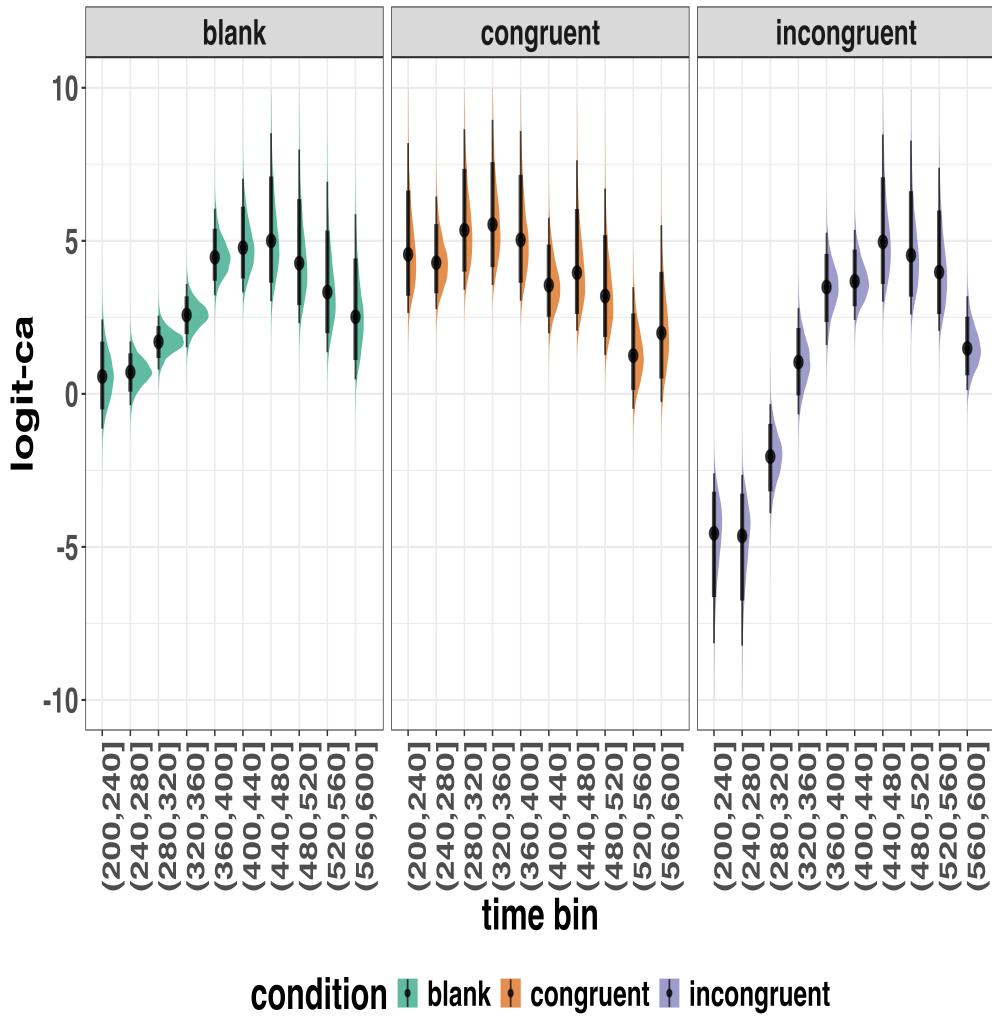


Figure 6. Medians and 80/95% credible intervals of the posterior distributions of the population-level parameters of model M1i_ca. ca = conditional accuracy.

Because the parameter estimates are on the logit-ca scale, we can ease our

interpretation by plotting the expected value of the posterior predictive distribution – the predicted conditional accuracies – at the population level (Figure 9A), and for each participant in the data set (Figure 9B).

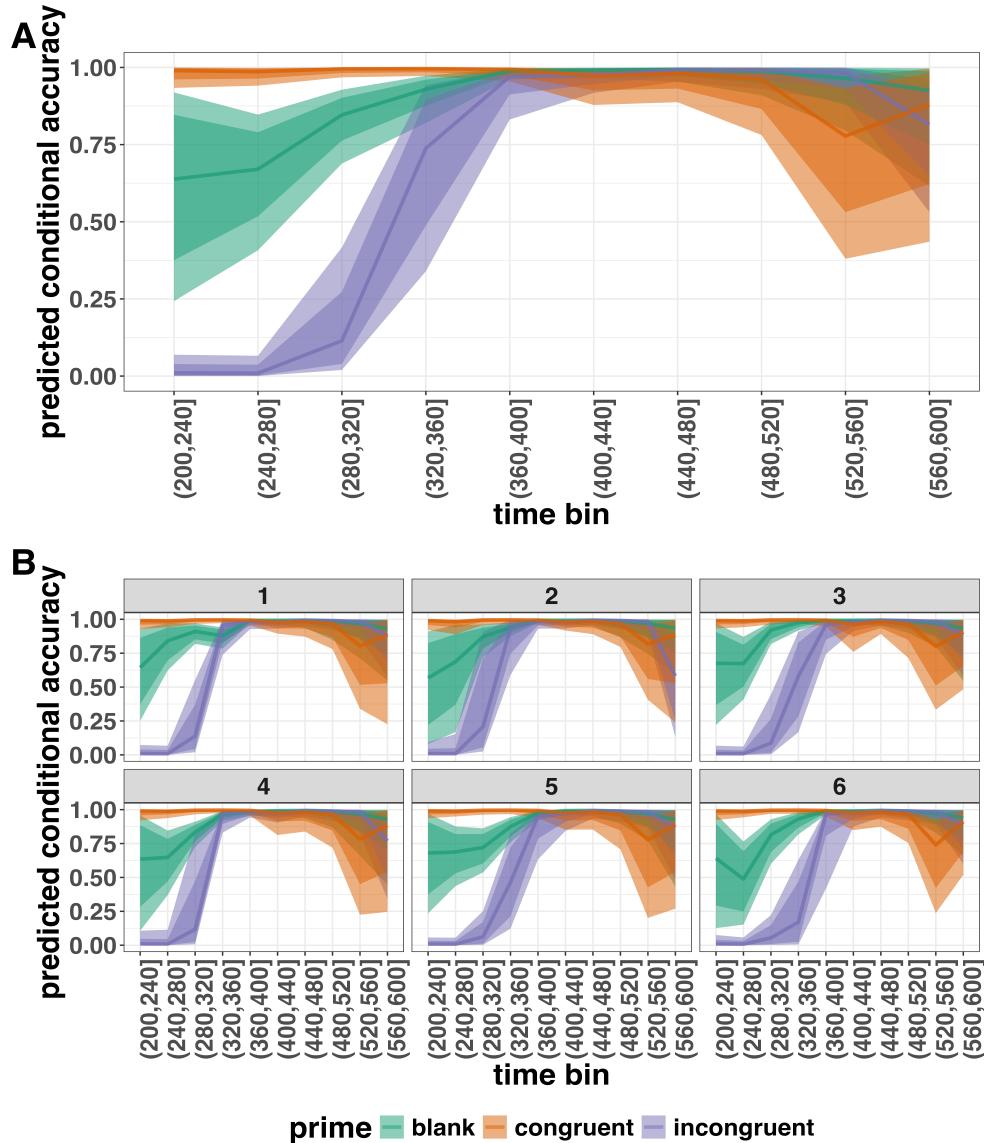


Figure 7. Point (median) and 80/95% credible interval summaries of the conditional accuracy estimates (expected values of the draws from the posterior predictive distributions) in each time bin at the population level (A), and for each participant (B).

563 As we are actually interested in the effects of congruent and incongruent primes,

564 relative to the blank prime condition, we can construct two contrasts (congruent-blank,
 565 incongruent-blank), and plot the posterior distributions of these contrast effects at the
 566 population level (Figure 10A) and for each participant (Figure 10B).

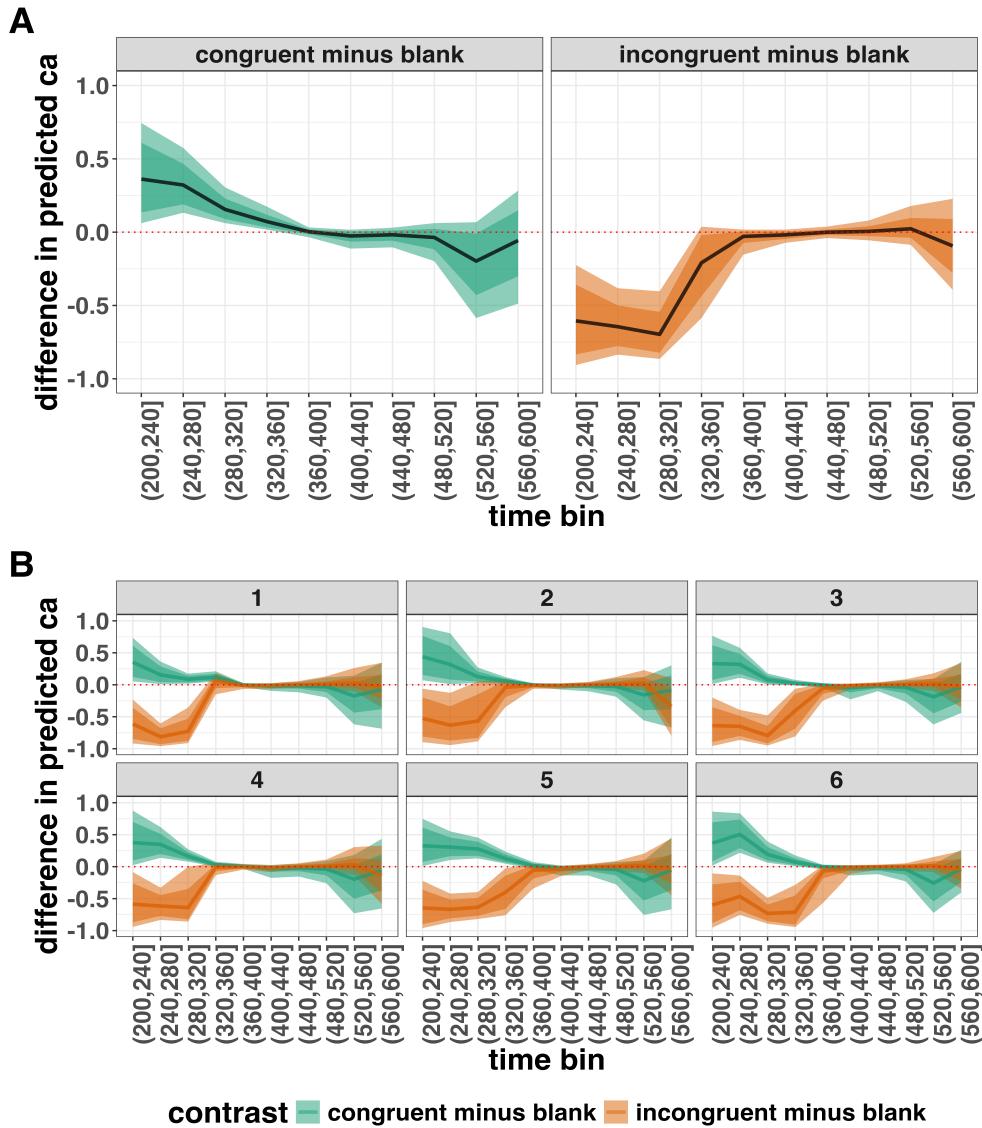


Figure 8. Point (mean) and 80/95% credible interval summaries of estimated differences in conditional accuracy in each time bin at the population level (A), and for each participant (B).

Based on Figure 10A we see that on the population level congruent primes have a

positive effect on the conditional accuracy of emitted responses in time bins (200,240],

(240,280], (280,320], and (320,360], relative to the estimates in the baseline condition

(blank prime; red dashed lines in Figure 10A). Incongruent primes have a negative effect on

571 the conditional accuracy of emitted responses in the first time bins, relative to the
572 estimates in the baseline condition.

573 Before we move to our Tutorial on planning experiments, we also provide code to fit
574 hazard and conditional accuracy models in the frequentist framework (see
575 Tutorial_3a.Rmd and Tutorial_3b.Rmd). However, because multilevel generalized linear
576 regression models often do not converge with complex random-coefficient structures, we do
577 not discuss them here.

578 3.5 Tutorial 4: Planning

579 In the final tutorial, we look at planning a future experiment, which uses EHA.

580 **3.5.1 Background.** The general approach to planning that we adopt here involves
581 simulating reasonably structured data to help guide what you might be able to expect from
582 your data once you collect it (Gelman, Vehtari, et al., 2020). The basic structure and code
583 follows the examples outlined by Solomon Kurz in his ‘power’ blog posts
584 (<https://solomonkurz.netlify.app/blog/bayesian-power-analysis-part-i/>) and Lisa
585 DeBruine’s R package faux{} (<https://debruine.github.io/faux/>) as well as these related
586 papers (DeBruine & Barr, 2021; Pargent, Koch, Kleine, Lermer, & Gaube, 2024).

587 **3.5.2 Basic workflow.** The basic workflow is as follows:

- 588 1. Fit a regression model to existing data.
- 589 2. Use the regression model parameters to simulate new data.
- 590 3. Write a function to create 1000s of datasets and vary parameters of interest (e.g.,
591 sample size, trial count, effect size).
- 592 4. Summarise the simulated data to estimate likely power or precision of the research
593 design options.

594 Ideally, in the above workflow, we would also fit a model to each dataset and
595 summarise the model output, rather than the raw data. However, when each model takes

596 several hours to build, and we may want to simulate many 1000s of datasets, it can be
597 computationally demanding for desktop machines. So, for ease, here we just use the raw
598 simulated datasets to guide future expectations.

599 In the below, we only provide a high-level summary of the process and let readers
600 dive into the details within the tutorial should they feel so inclined.

601 **3.5.3 Fit a regression model and simulate one dataset.** We again use the
602 data from Panis and Schmidt (2016) to provide a worked example. We fit an index coding
603 model on a subset of time bins (six time bins in total) and for two prime conditions
604 (congruent and incongruent). We chose to focus on a subsample of the data to ease the
605 computational burden. We also used a full varying effects structure, with the model
606 formula as follows:

```
event ~ 0 + timebin:prime + (0 + timebin:prime | pid)
```

607 We then took parameters from this model and used them to create a single dataset
608 with 200 trials per condition for 10 individual participants. The raw data and the
609 simulated data are plotted in Figure 12 and show quite close correspondence, which is
610 re-assuring. But, this is only one dataset. What we really want to do is simulate many
611 datasets and vary parameters of interest, which is what we turn to in the next section.

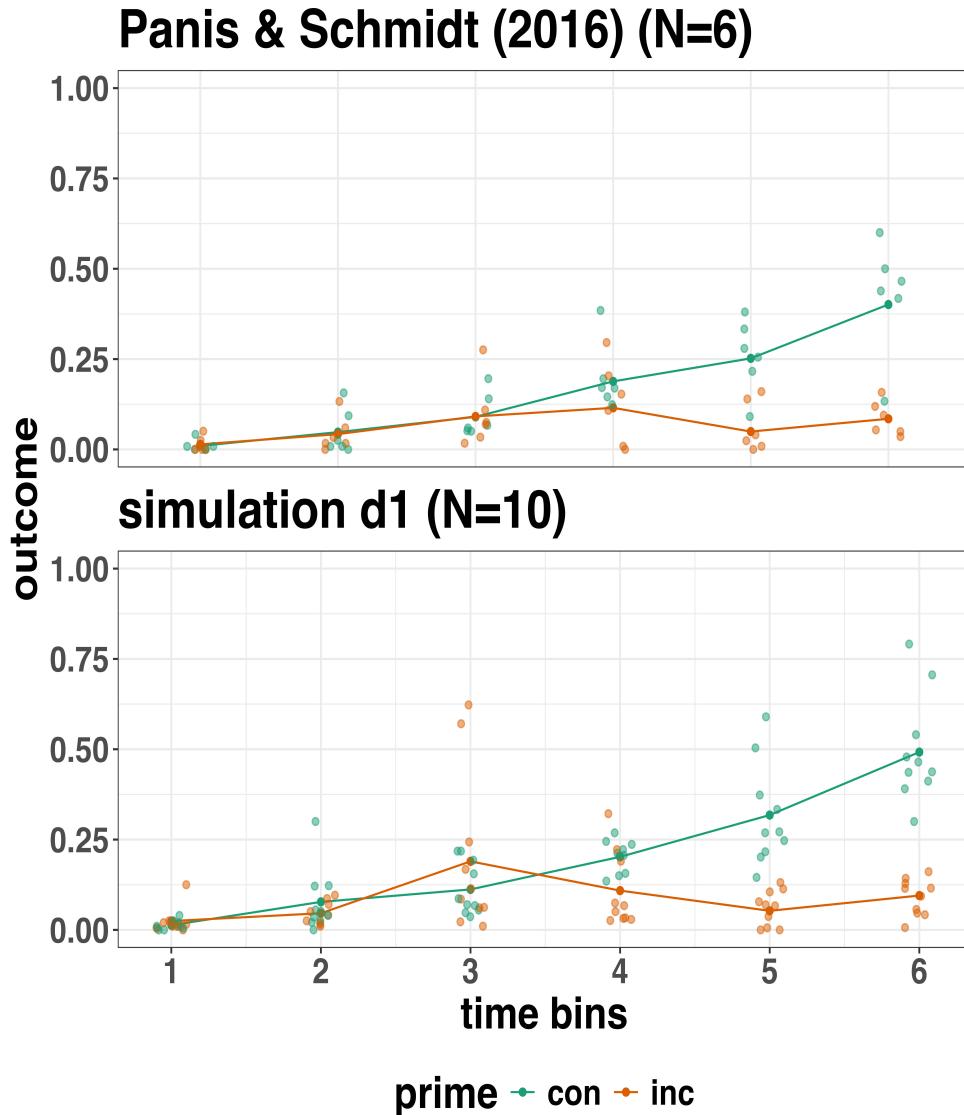


Figure 9. Raw data from Panis and Schmidt (2016) and simulated data from 10 participants.

3.5.4 Simulate and summarise data across a range of parameter values.

612 Here we use the same data simulation process as used above, but instead of simulating one
 613 dataset, we simulate 1000 datasets per variation in parameter values. Specifically, in
 614 Simulation 1, we vary the number of trials per condition (100, 200, and 400), as well as the
 615 effect size in bin 6. We focus on bin 6 only, in terms of varying the effect size, just to make
 616 things simpler and easier to understand. The effect size observed in bin 6 in this subsample
 617

of data was a 79% reduction in hazard value from the congruent prime (0.401 hazard value) to the incongruent prime condition (0.085 hazard value). In other words, a hazard ratio of 0.21 (e.g., $0.085/0.401 = 0.21$). As a starting point, we chose three effect sizes, which covered a fairly broad range of hazard ratios (0.25, 0.5, 0.75), which correspond to a 75%, 50% and 25% reduction in hazard value as a function of prime condition.

Summary results from Simulation 1 are shown in Figure 13A. Figure 13A depicts statistical “power” as calculated by the percentage of lower-bound 95% confidence intervals that exclude zero when the difference between prime condition is calculated (congruent - incongruent). In other words, what fraction of the simulated datasets generated an effect of prime that excludes the criterion mark of zero. We are aware that “power” is not part of a Bayesian analytical workflow, but we choose to include it here, as it is familiar to most researchers in experimental psychology.

The results of Simulation 1 show that if we were targeting an effect size similar to the one reported in the original study, then testing 10 participants and collecting 100 trials per condition would be enough to provide over 95% power. However, we could not be as confident about smaller effects, such as a hazard ratio of 50% or 25%. From this simulation, we can see that somewhere between an effect size of a 50% and 75% reduction in hazard value, power increases to a range that most researchers would consider acceptable (i.e., >95% power). To probe this space a little further, we decided to run a second simulation, which varied different parameters.

In Simulation 2, we varied the effect size between a different range of values (0.5, 0.4, 0.3), which correspond to a 50%, 60% and 70% reduction in hazard value as a function of prime condition. In addition, we varied the number of participants per experiment between 10, 15, and 20 participants. Given that trial count per condition made little difference to power in Simulation 1, we fixed trial count at 200 trials per condition in Simulation 2. Summary results from Simulation 2 are shown in Figure 13B. A summary of these power

⁶⁴⁴ calculations might be as follows (trial count = 200 per condition in all cases):

- ⁶⁴⁵ • For a 70% reduction (0.3 hazard ratio), N=10 would give nearly 100% power.
- ⁶⁴⁶ • For a 60% reduction (0.4 hazard ratio), N=10 would give nearly 90% power.
- ⁶⁴⁷ • For a 50% reduction (0.5 hazard ratio), N=15 would give over 80% power.

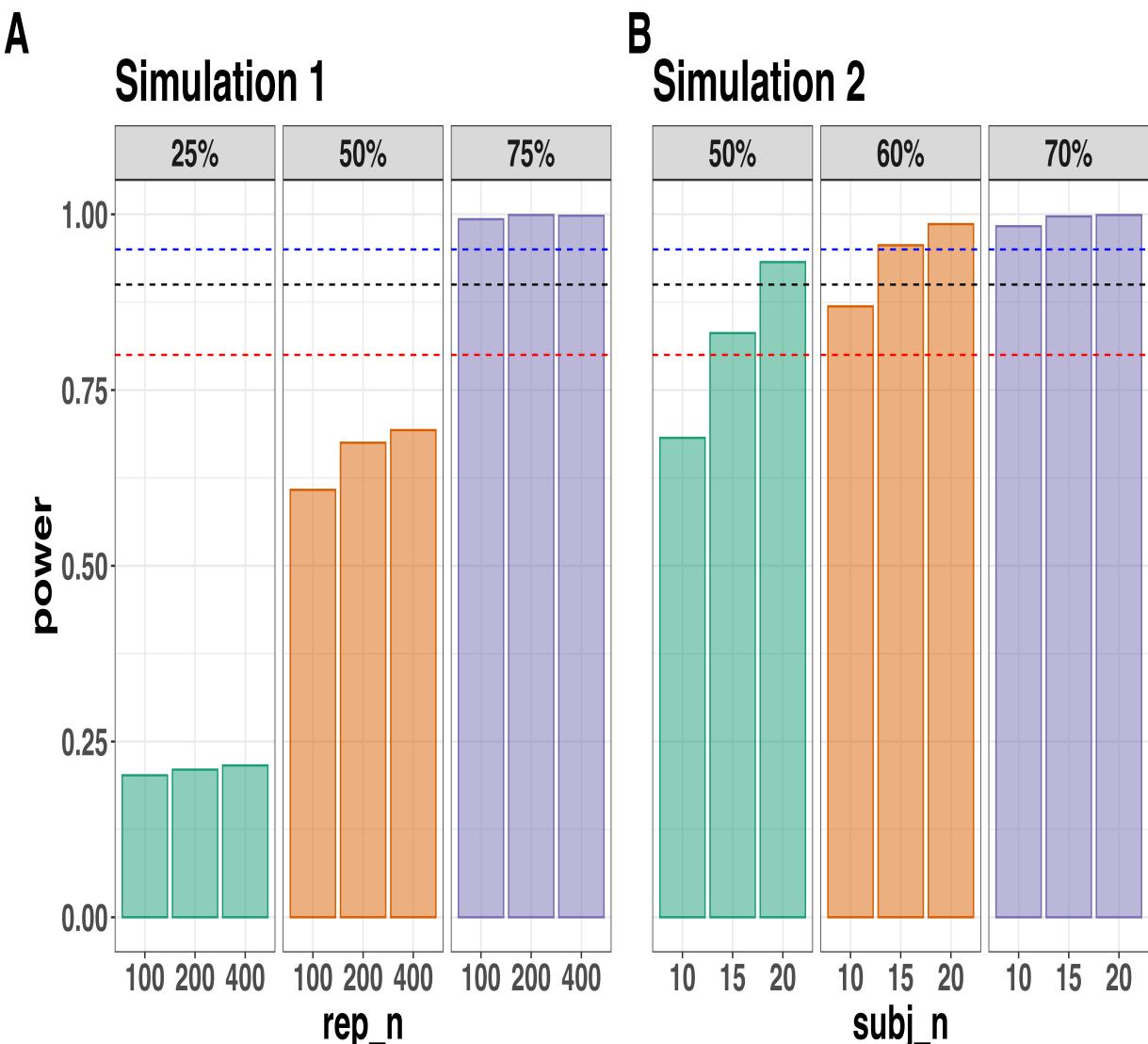


Figure 10. Statistical power across data Simulation 1 (A) and Simulation 2 (B). Power was calculated as the percentage of lower-bound 95% confidence intervals that exclude zero when the difference between prime condition is calculated (congruent - incongruent). In Simulation 1, the effect size was varied between a 25%, 50% and 75% reduction in hazard value, whereas the trial count was varied between 100, 200 and 400 trials per condition (the number of participants was fixed at N=10). In Simulation 2, the effect size was varied between a 50%, 60% and 70% reduction in hazard value, whereas the number of participants was varied between N=10, 15 and 20 (the number of trials per condition was fixed at 200). The dashed lines represent 80% (red), 90% (black) and 95% (blue) power. Abbreviations: rep_n = the number of trials per experimental condition; subj_n = the number of participants per simulated experiment.

648 **3.5.5 Planning decisions.** Now that we have summarised our simulated data,

649 what planning decisions could we make about a future study? More concretely, how many
650 trials per condition should we collect and how many participants should we test? Like
651 almost always when planning future studies, the answer depends on your objectives, as well
652 as the available resources (Lakens, 2022). There is no straightforward and clear-cut answer.

653 Some considerations might be as follows:

- 654 • How much power or precision are you looking to obtain in this particular study?
- 655 • Are you running multiple studies that have some form of replication built in?
- 656 • What level of resources do you have at your disposal, such as time, money and
657 personnel?
- 658 • How easy or difficult is it to obtain the specific type of sample?

659 If we were running this kind of study in our lab, what would we do? We might pick a

660 hazard ratio of 0.4 or 0.5 as a target effect size since this is much smaller than that
661 observed previously (Panis & Schmidt, 2016). Then we might pick the corresponding
662 combination of trial count per condition (e.g., 200) and participant sample size (e.g., N=10
663 or N=15) that takes you over the 80% power mark. If we wanted to maximise power based
664 on these simulations, and we had the time and resources available, then we would test
665 N=20 participants, which would provide >90% power for an effect size of 0.5.

666 **But**, and this is an important “but”, unless there are unavoidable reasons, no matter

667 what planning choices we made based on these data simulations, we would not solely rely
668 on data collected from one single study. Instead, we would run a follow-up experiment that
669 replicates and extends the initial result. By doing so, we would aim to avoid the Cult of
670 the Isolated Single Study (Nelder, 1999; Tong, 2019), and thus reduce the reliance on any
671 one type of planning tool, such as a power analysis. Then, we would look for common
672 patterns across two or more experiments, rather than trying to make the case that a single
673 study on its own has sufficient evidential value to hit some criterion mark.

674

4. Discussion

675 This main motivation for writing this paper is the observation that EHA and SAT
676 analysis remain under-used in psychological research. As a consequence, the field of
677 psychological research is not taking full advantage of the many benefits EHA/SAT provides
678 compared to more conventional analyses. By providing a freely available set of tutorials,
679 which provide step-by-step guidelines and ready-to-use R code, we hope that researchers
680 will feel more comfortable using EHA/SAT in the future. Indeed, we hope that our
681 tutorials may help to overcome a barrier to entry with EHA/SAT, which is that such
682 approaches require more analytical complexity compared to standard approaches. While
683 we have focused here on within-subject, factorial, small- N designs, it is important to realize
684 that EHA/SAT can be applied to other designs as well (large- N designs with only one
685 measurement per subject, between-subject designs, etc.). As such, the general workflow
686 and associated code can be modified and applied more broadly to other contexts and
687 research questions. In the following, we discuss the main use-cases, issues relating to model
688 complexity and interpretability, as well as limitations of the approach.

689 **4.1 What are the main use-cases of EHA for understanding cognition and brain
690 function?**

691 For those researchers, like ourselves, who are primarily interested in understanding
692 human cognitive and brain systems, we consider two broadly-defined, main use-cases of
693 EHA. First, as we hope to have made clear by this point, EHA is one way to investigating
694 a “temporal states” approach to cognitive processes, by tracking behavior as a function of
695 step-wise increases in absolute waiting time. EHA thus provides a way to uncover the
696 microgenesis of cognitive effects, by revealing when cognitive states may start and stop,
697 how states are replaced with others, as well as what they may be tied to or interact with.
698 Therefore, if your research questions concern **when psychological states occur, and**

699 how they are temporally organized, our EHA tutorials could be useful tools for you to
700 use.

701 Second, even if you are not primarily interested in studying the temporal
702 organization of cognitive states, EHA could still be a useful tool to consider using, in order
703 to qualify inferences that are being made based on comparisons between means. Given that
704 distinctly different inferences can be made from the same data based on whether one
705 computes a mean across trials or a RT distribution of events (Figure 1), it may be
706 important for researchers to supplement comparisons between means with EHA. Therefore,
707 if you have a lot of right-censored observations in your RT data set, and/or your research
708 question concerns whether and when responses occur, and whether and when experimental
709 manipulations affect the instantaneous risk of response occurrence, then EHA should be
710 your method of choice.

711 4.2 Model complexity versus interpretability

712 Hazard and conditional accuracy models can quickly become very complex when
713 adding more than one time scale, due to the many possible higher-order interactions. For
714 example, some of the models discussed in Tutorial 2a, which we did not focus on in the
715 main text, contain two time scales as covariates: the passage of time on the within-trial
716 time scale, and the passage of time on the across-trial (or within-experiment) time scale.
717 However, when trials are presented in blocks, and blocks of trials within sessions, and when
718 the experiment comprises a number of sessions, then four time scales can be defined
719 (within-trial, within-block, within-session, and within-experiment). From a theoretical
720 perspective, adding more than one time scale – and their interactions – can be important
721 to capture plasticity and other learning effects that may play out on such longer time
722 scales, and that are probably present in each experiment in general (Schöner & Spencer,
723 2016). From a practical perspective, therefore, some choices need to be made to balance
724 the amount of data that is being collected per participant, condition and across the varying

725 timescales. As one example, if there are several times cales of relevance, then it might be
726 prudent for interpretational purposes to limit the number of experimental predictor
727 variables (conditions). This is of course where planning and data simulation efforts would
728 be important to provide a guide to experimental design choices (see Tutorial 4 and section
729 2.3).

730 **4.3 Limitations**

731 Compared to the orthodox method – comparing means between conditions – the
732 most important limitation of multilevel hazard and conditional accuracy modeling is that it
733 might take a long time to estimate the parameters using Bayesian methods or the model
734 might have to be simplified significantly to use frequentist methods.

735 Another issue is that you need a relatively large number of trials per condition to
736 estimate the discrete-time hazard function with relatively high temporal resolution (e.g., 20
737 ms), which is required when testing predictions of process models of cognition. Indeed, in
738 general, there is a trade-off between the number of trials per condition and the temporal
739 resolution (i.e., bin width) of the discrete-time hazard function. Therefore, we recommend
740 researchers to collect as many trials as possible per experimental condition, given the
741 available resources and considering the participant experience (e.g., fatigue and boredom).

742 For instance, if the maximum session length deemed reasonable is between 1 and 2 hours,
743 what is the maximum number of trials per condition that you could reasonably collect?

744 After consideration, it might be worth conducting multiple testing sessions per participant
745 and/or reducing the number of experimental conditions. There is a user-friendly online tool
746 for calculating statistical power as a function of the number of trials as well as the number
747 of participants, and this might be worth consulting to guide the research design process
748 (Baker et al., 2021). Finally, if you have a lot of repeated measurements per condition per
749 participant, you can of course also try continuous-time methods (Allison, 2010; Elmer et
750 al., 2023).

751

5. Conclusions

752 Estimating the temporal distributions of RT and accuracy provide a rich source of
753 information on the time course of cognitive processing, which have been largely
754 undervalued in the history of experimental psychology and cognitive neuroscience. We
755 hope that by providing a set of hands-on, step-by-step tutorials, which come with
756 custom-built and freely available code, researchers will feel more comfortable embracing
757 EHA and investigating the shape of empirical hazard functions and the temporal profile of
758 cognitive states. On a broader level, we think that wider adoption of such approaches will
759 have a meaningful impact on the inferences drawn from data, as well as the development of
760 theories regarding the structure of cognition.

761

Author contributions

762 Conceptualization: S. Panis and R. Ramsey; Software: S. Panis and R. Ramsey;
763 Writing - Original Draft Preparation: S. Panis; Writing - Review & Editing: S. Panis and
764 R. Ramsey; Supervision: R. Ramsey.

765

Conflicts of Interest

766 The author(s) declare that there were no conflicts of interest with respect to the
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768

Prior versions

769 All of the submitted manuscript and Supplemental Material was previously posted to
770 a preprint archive: <https://doi.org/10.31234/osf.io/57bh6>

771

Supplemental Material

772

Disclosures**773 Data, materials, and online resources**

774 Link to public archive:
775 https://github.com/sven-panis/Tutorial_Event_History_Analysis
776 Supplemental Material: Panis_Ramsey_suppl_material.pdf

777 Ethical approval

778 Ethical approval was not required for this tutorial in which we reanalyze existing
779 data sets.

780

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