

Boundary effects in kernel regression

Xingyu Tao, Xuan Li, Sven Jacobs

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Research Module in Econometrics & Statistics

1. Introduction
2. Local linear regression
3. Boundary kernels
4. Simulation
5. Application
6. Conclusion

Introduction

Kernel regression

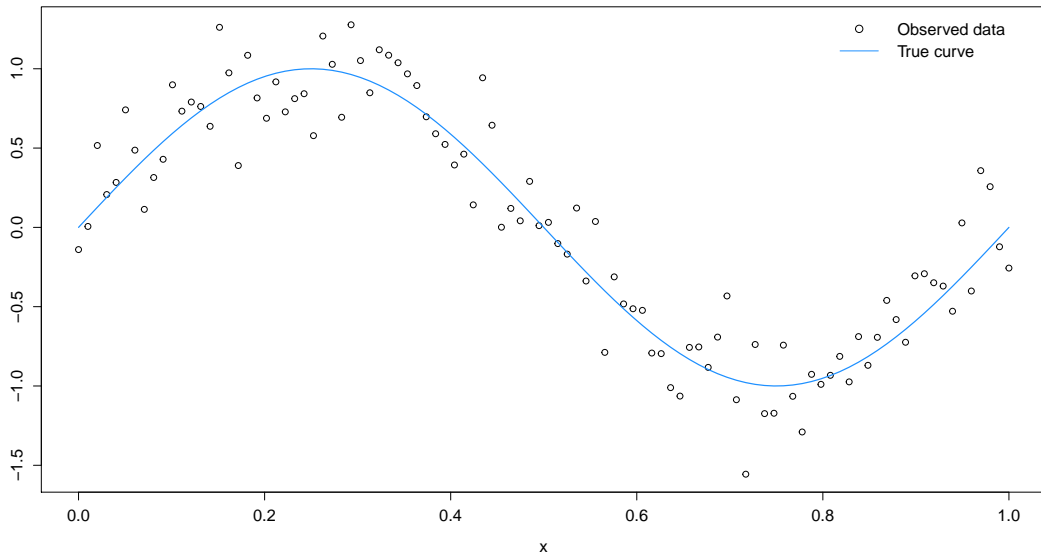
Setting

- iid data: $\{(X_i, Y_i)\}_{i=1, \dots, n}$
- Random design: $X \sim f_X, X_i \in [a, b] \subset \mathbb{R}$

$$Y_i = \underbrace{m(X_i)}_{\text{target}} + \epsilon_i$$

- $m(X_i) = \mathbb{E}[Y_i | X = X_i]$ (CEF)
- $\mathbb{E}[\epsilon_i | X = X_i] = 0, \text{Var}(\epsilon_i | X = X_i) = \sigma^2(X_i)$

Example



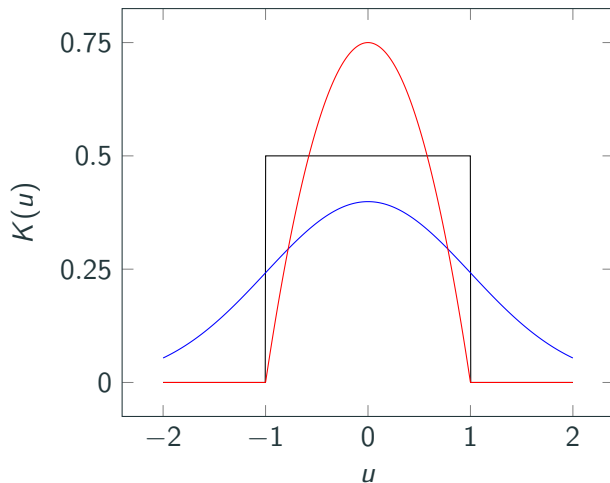
Nadaraya-Watson estimator

- Approximation of $m(x)$ by a weighted local average of the Y_i 's:

$$\hat{m}_{\text{NW}}(x) = \sum_{i=1}^n w(x, X_i, h) Y_i = \sum_{i=1}^n \frac{K\left(\frac{x-X_i}{h}\right)}{\sum_{j=1}^n K\left(\frac{x-X_j}{h}\right)} Y_i$$

- Kernel K :
 - Weight of Y_i smaller if $|x - X_i|$ larger
 - Most often: symmetric density
- Bandwidth h :
 - Smoothing parameter
 - Determines how fast weights decrease in $|x - X_i|$

Selection of kernels



Uniform

$$K(u) = \begin{cases} \frac{1}{2} & \text{if } |u| \leq 1 \\ 0 & \text{else} \end{cases}$$

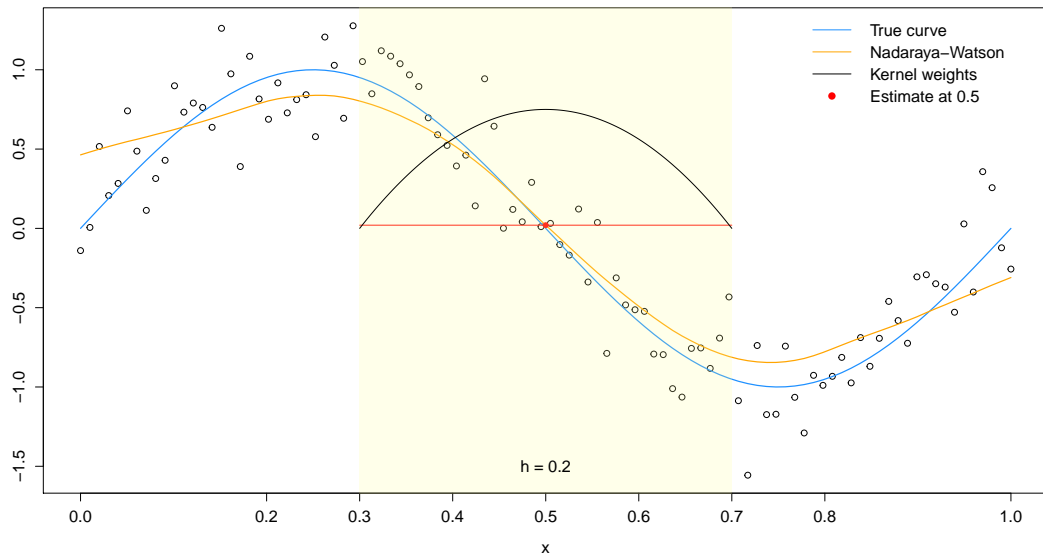
Gaussian

$$K(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}}$$

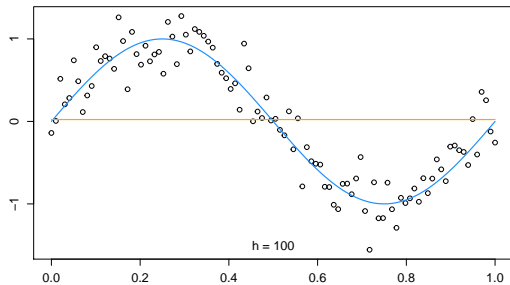
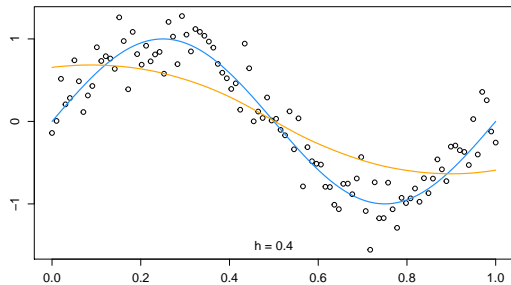
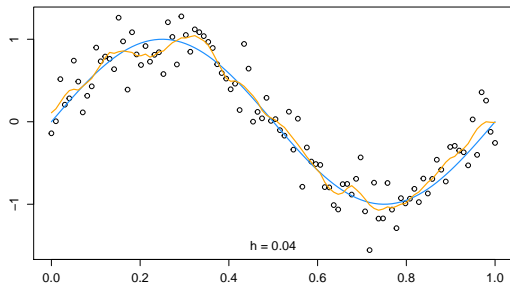
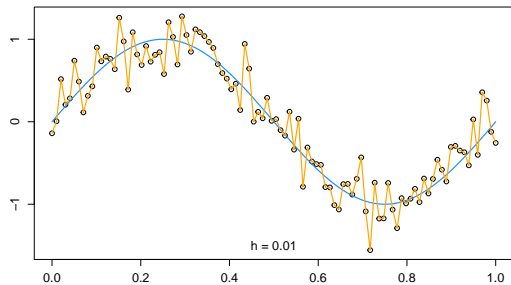
Epanechnikov

$$K(u) = \begin{cases} \frac{3}{4}(1 - u^2) & \text{if } |u| \leq 1 \\ 0 & \text{else} \end{cases}$$

Example continued: Nadaraya-Watson



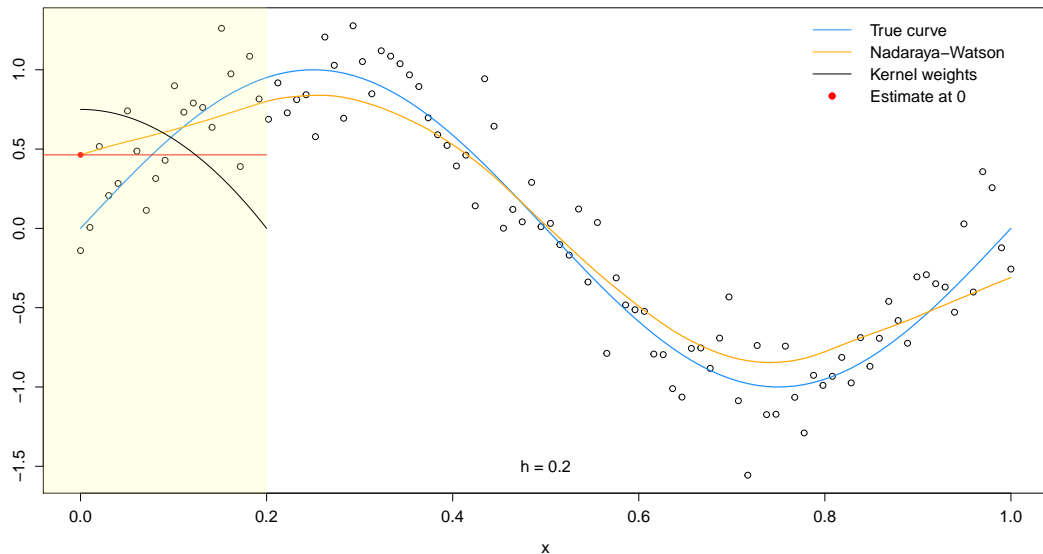
Role of bandwidth



Introduction

Motivation

Example continued: Boundary effects



Solution approaches

1. Local linear regression
 - Automatic boundary adaption

Solution approaches

1. Local linear regression
 - Automatic boundary adaption
2. Special boundary kernels
 - Adjustment of NW estimator near boundaries

Local linear regression

Definition

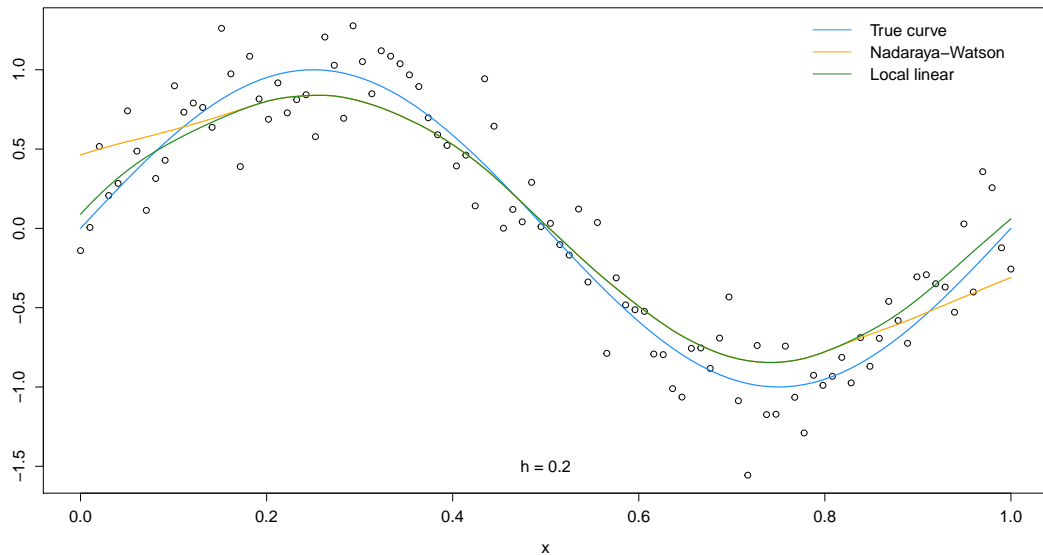
- Weighted local fitting of a line

$$\hat{m}_{\text{NW}}(x) = \arg \min_{\beta_0} \sum_{i=1}^n (Y_i - \beta_0)^2 K \left(\frac{x - X_i}{h} \right)$$

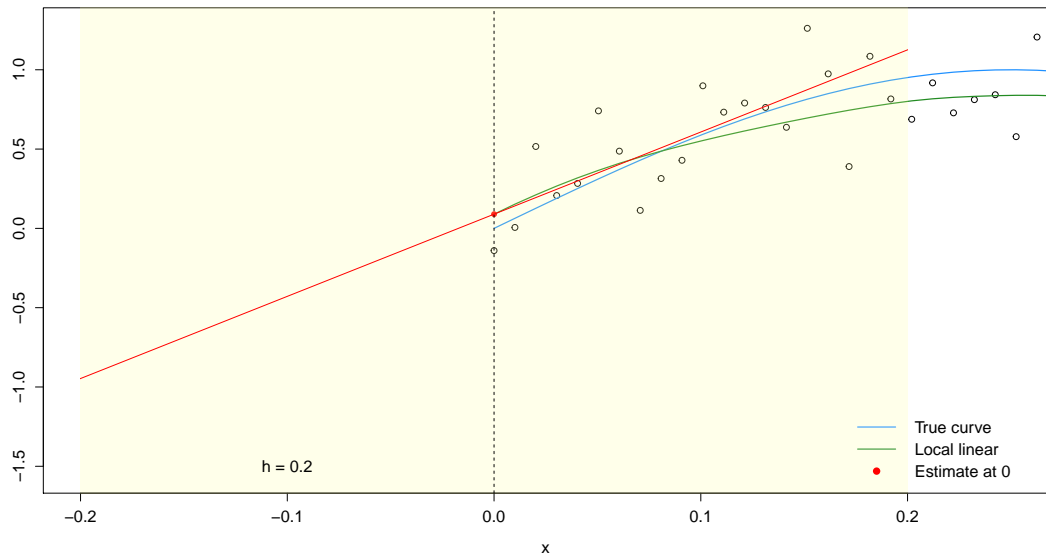
$$\begin{aligned} \hat{m}_{\text{LL}}(x) &= \arg \min_{\beta_0} \sum_{i=1}^n (Y_i - \beta_0 - \beta_1(X_i - x))^2 K \left(\frac{x - X_i}{h} \right) \\ &= \sum_{i=1}^n w(x, X_i, h) Y_i \end{aligned}$$

⇒ Weighted Least Squares (WLS)

Example continued: Local linear



Example continued: Local linear



Local linear regression

Theoretical comparison

Estimation in the interior: Assumptions

(A1) m is twice continuously differentiable

(A2) Regularity conditions on f , σ^2 and K

(A3) $h \rightarrow 0$, $nh \rightarrow \infty$ as $n \rightarrow \infty$

(A4) $x \in [a + h, b - h]$

Estimation in the interior: Asymptotics

$$\text{ABias}(\hat{m}_{\text{NW}}(x)|\mathbf{X}) = \frac{1}{2}\kappa_2(K)m''(x)h^2 + \kappa_2(K)\frac{m'(x)f'(x)}{f(x)}h^2 = \mathcal{O}(h^2)$$

$$\text{ABias}(\hat{m}_{\text{LL}}(x)|\mathbf{X}) = \frac{1}{2}\kappa_2(K)m''(x)h^2 = \mathcal{O}(h^2)$$

$$\text{AVar}(\hat{m}_{\text{NW}}(x)|\mathbf{X}) = \text{AVar}(\hat{m}_{\text{LL}}(x)|\mathbf{X}) = \frac{R(K)\sigma^2(x)}{f(x)}\frac{1}{nh} = \mathcal{O}\left(\frac{1}{nh}\right)$$

- Bias of LL typically smaller but of same order
- Asymptotic variance is the same

Estimation at the boundary: Assumptions

(A1) m is twice continuously differentiable

- $\lim_{x \downarrow a} m''(x) = m''(a)$ and $\lim_{x \uparrow b} m''(x) = m''(b)$

(A2) Regularity conditions on f , σ^2 and K

(A3) $h \rightarrow 0$, $nh \rightarrow \infty$ as $n \rightarrow \infty$

(A4) $x \in \{a, b\}$

Estimation at the boundary: Asymptotics

$$\text{ABias}(\hat{m}_{\text{NW}}(a)|\mathbf{X}) = 2 \int_0^\infty u K(u) du m'(a) h = \mathcal{O}(h)$$

$$\text{ABias}(\hat{m}_{\text{NW}}(b)|\mathbf{X}) = -2 \int_0^\infty u K(u) du m'(b) h = \mathcal{O}(h)$$

$$\text{ABias}(\hat{m}_{\text{LL}}(a)|\mathbf{X}) = \frac{1}{2} \kappa_2(\tilde{K}) m''(a) h^2 = \mathcal{O}(h^2)$$

$$\text{AVar}(\hat{m}_{\text{LL}}(a)|\mathbf{X}) = \frac{R(\tilde{K}) \sigma^2(a)}{f(a)} \frac{1}{nh} = \mathcal{O}\left(\frac{1}{nh}\right)$$

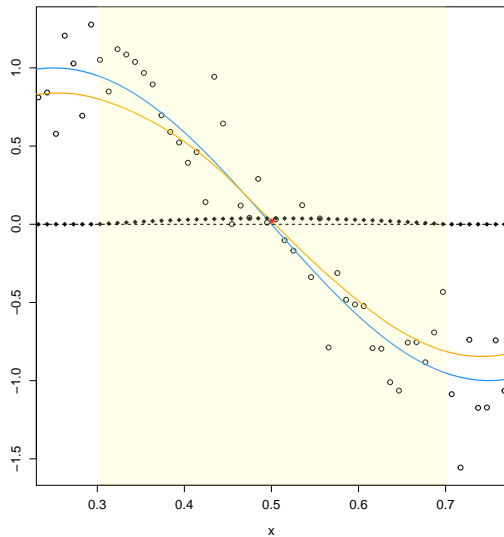
- ABias of NW is $\mathcal{O}(h)$ instead of $\mathcal{O}(h^2)$
- LL preserves the order (but constants differ)

Local linear regression

Graphical comparison

- NW and LL both linear smoothers: $\hat{m}(x) = \sum_{i=1}^n w(x, X_i, h) Y_i$
- $\{w(x, X_i, h)\}_{i=1}^n$: Effective kernel at x

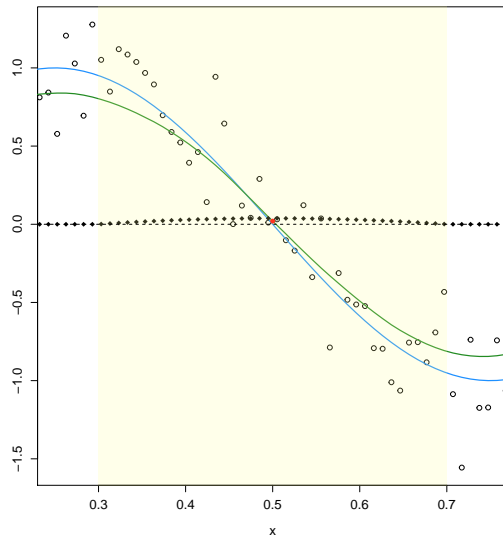
Effective kernel in the interior



— True curve

— NW

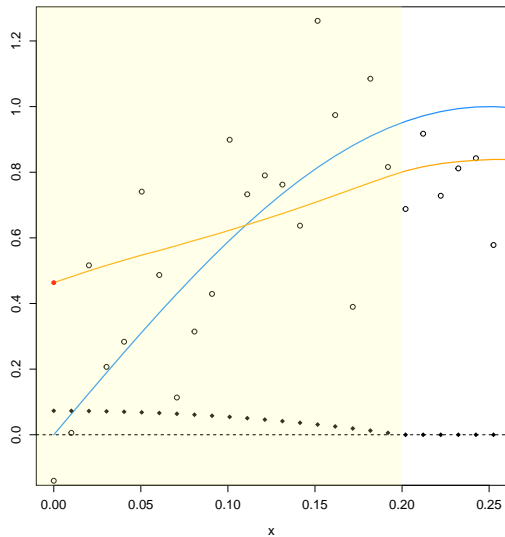
— LL



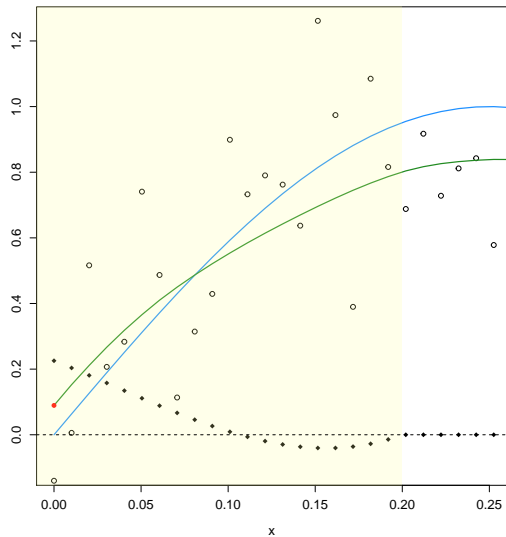
• Estimate at 0.5

• Effective kernel at 0.5

Effective kernel at the boundary



— True curve — NW — LL



• Estimate at 0 • Effective kernel at 0

Boundary kernels

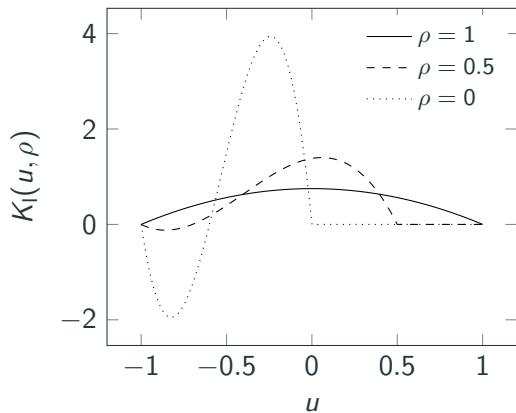
Idea

- Assume that $\text{supp}(K) = [-1, 1]$ and w.l.o.g. $X_i \in [0, 1]$
 \Rightarrow Boundary region: $[0, h) \cup (1 - h, 1]$
- Idea: Construct asymmetric kernel that restores the original moment properties
- For $x = \rho h$ a left boundary kernel $K_l(u, \rho)$ satisfies for $\rho \in [0, 1]$

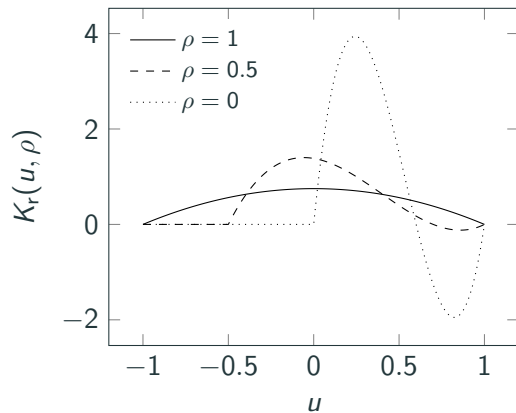
$$\int_{-1}^{\rho} u^j K_l(u, \rho) du = \begin{cases} 1 & j = 0 \\ 0 & j = 1 \\ \kappa_{2,\rho}(K_l) \neq 0 & j = 2 \end{cases}$$

- Note: Kernel shape changes with relative location of x to boundary!

Epanechnikov boundary kernels, Müller (1991)

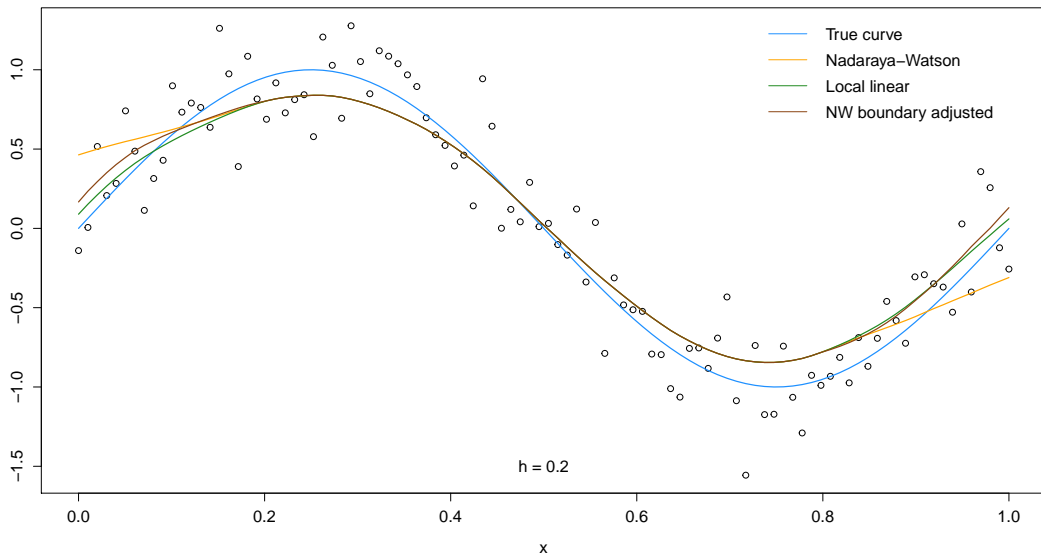


Left boundary kernels



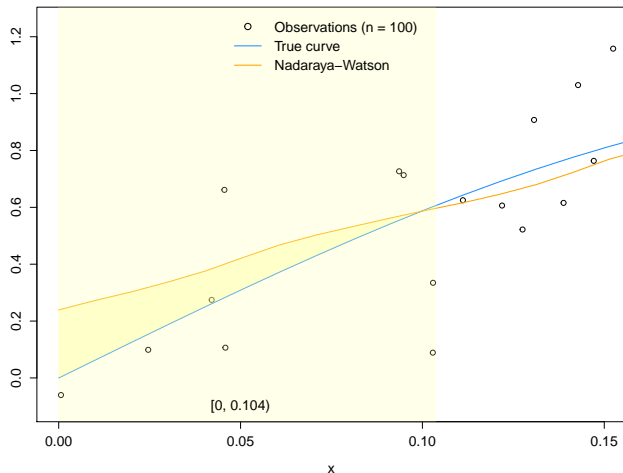
Right boundary kernels

Example continued: Boundary kernels



Simulation

Simulation set-up



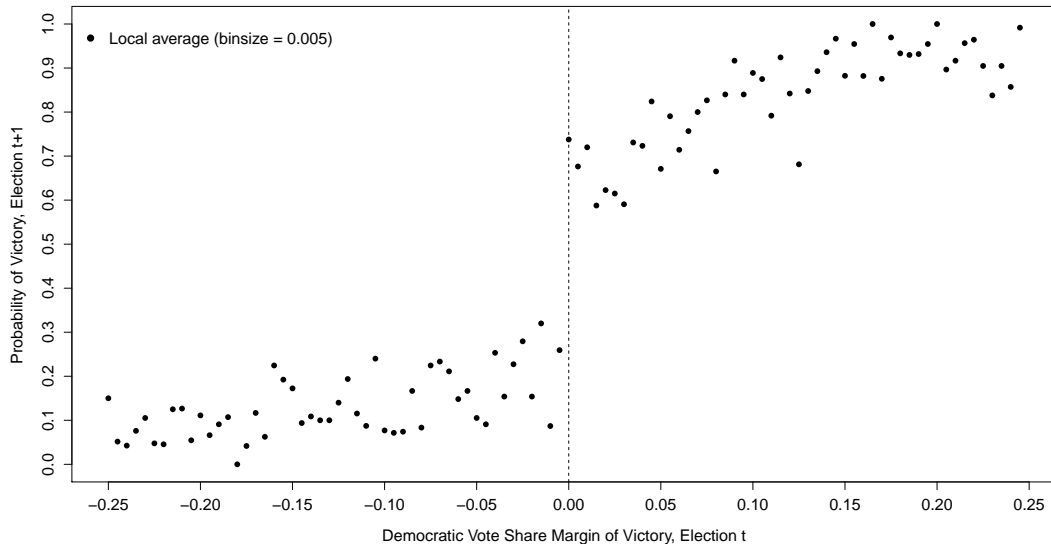
- $m(x) = \sin(2\pi x)$
- $X \sim \mathcal{U}(0, 1)$
- $\epsilon \sim \mathcal{N}(0, 0.25^2)$
- Target: Mean Integrated Squared Error (MISE) over boundary region
- Asymptotic optimal bandwidth: $h_{\text{AMISE}}(n)$
- 10000 Monte-Carlo repetitions
- Sample sizes from 50 to 5000

Simulation results

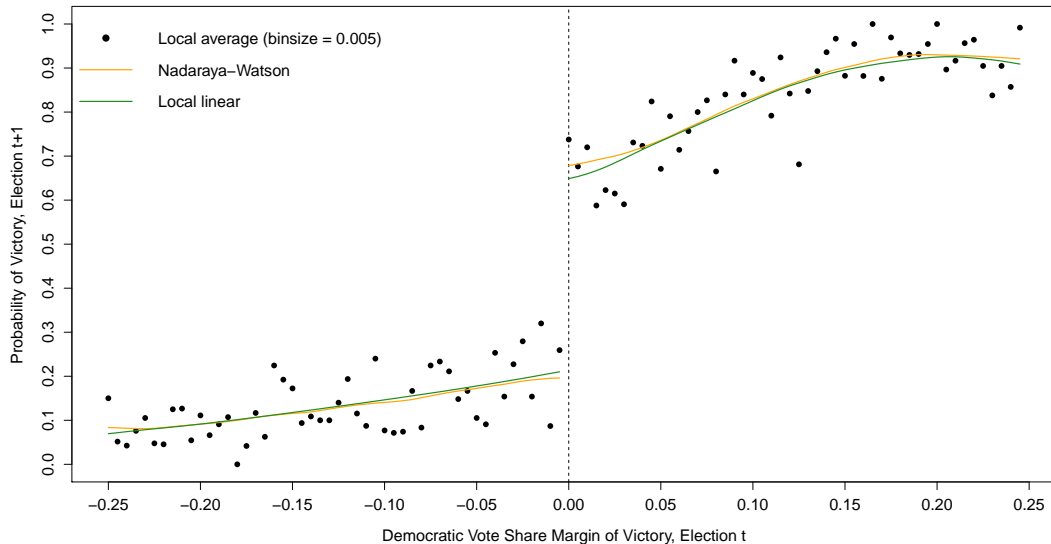
n	Boundary region	Change in MISE to NW [%]	
		Local linear	NW boundary adjusted
50	[0, 0.119)	789.29	1578.07
100	[0, 0.104)	−41.90	1143.28
250	[0, 0.086)	−65.13	1982.67
500	[0, 0.075)	−72.98	14.92
1000	[0, 0.065)	−78.80	−63.95
2500	[0, 0.055)	−85.12	−77.80
5000	[0, 0.047)	−88.50	−83.76

Application

Regression discontinuity design (RDD), Lee (2008)



Regression discontinuity design (RDD), Lee (2008)



Conclusion

- Often in econometrics the boundaries are of great interest
- NW with potentially severe boundary bias → Adjustment!
- Boundary kernels with asymptotic correction, but complicated and impractical
- LL regression as a simple and intuitive method to automatically reduce boundary effects
 - Plus favorable asymptotic properties in interior

⇒ Recommendation: Local linear regression

Contact

Slides and codes are hosted on GitHub.

For further questions contact us via email.



`github.com/svjaco`



`s6xitaoo(at)uni-bonn.de` (Xingyu)



`s6xuliii(at)uni-bonn.de` (Xuan)



`s.jacobs(at)uni-bonn.de` (Sven)

Appendix

Local polynomial regression: Definition

$$\begin{aligned}\hat{m}_{\text{LP}}(x) &= \arg \min_{\beta_0} \sum_{i=1}^n \left(Y_i - \sum_{j=0}^p \beta_j (X_i - x)^j \right)^2 K \left(\frac{x - X_i}{h} \right) \\ &= \sum_{i=1}^n w(x, X_i, h) Y_i\end{aligned}$$

Local polynomial regression: Motivation

- Minimization of RSS

$$\sum_{i=1}^n (Y_i - \hat{m}(X_i))^2$$

- p -th order Taylor expansion:

$$m(X_i) \approx \sum_{j=0}^p \frac{m^{(j)}(x)}{j!} (X_i - x)^j$$

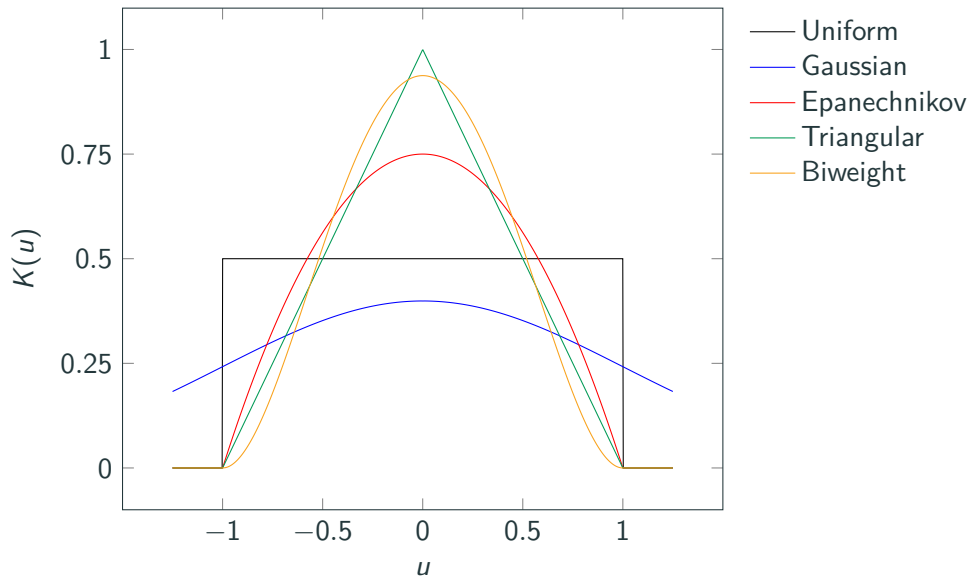
Local polynomial regression: Motivation

$$\sum_{i=1}^n \left(Y_i - \sum_{j=0}^p \underbrace{\frac{m^{(j)}(x)}{j!}}_{\equiv \beta_j} (X_i - x)^j \right)^2$$

- Weighting of observation (X_i, Y_i) using a kernel:

$$\hat{m}_{\text{LP}}(x) = \arg \min_{\beta_0} \sum_{i=1}^n \left(Y_i - \sum_{j=0}^p \beta_j (X_i - x)^j \right)^2 K \left(\frac{x - X_i}{h} \right)$$

Kernels: Prominent examples

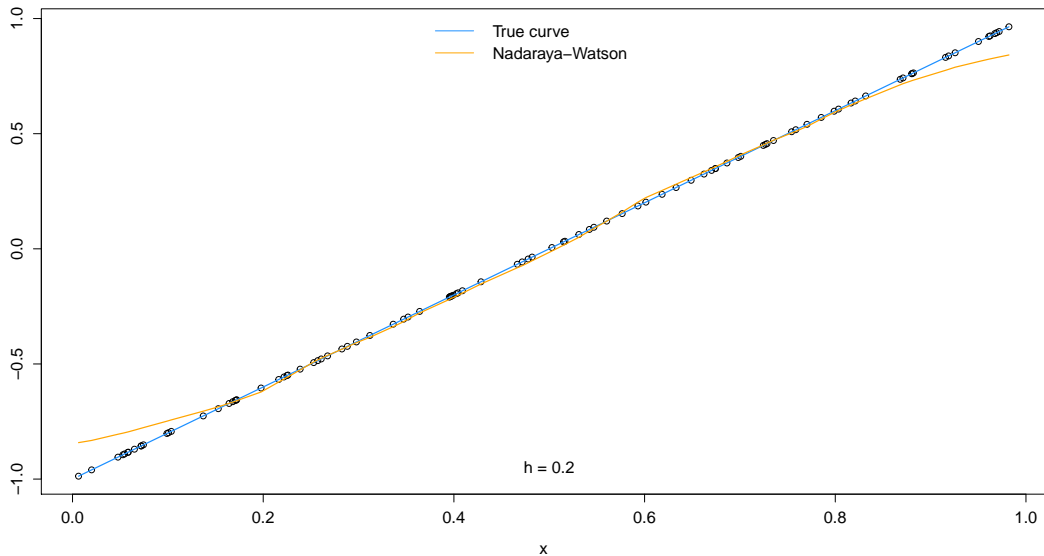


Kernels: Efficiency

Kernel	Function	$R(K)$	$\kappa_2(K)$	Efficiency
Epanechnikov	$K_E(u) = \frac{3}{4}(1 - u^2)$	$\frac{3}{5}$	$\frac{1}{5}$	100.0%
Biweight	$K_B(u) = \frac{15}{16}(1 - u^2)^2$	$\frac{5}{7}$	$\frac{1}{7}$	99.39%
Triangular	$K_T(u) = 1 - u $	$\frac{2}{3}$	$\frac{1}{6}$	98.59%
Gaussian	$K_G(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}}$	$\frac{1}{2\sqrt{\pi}}$	1	95.12%
Uniform	$K_U(u) = \frac{1}{2}$	$\frac{1}{2}$	$\frac{1}{3}$	92.95%

$$\text{Eff}(K) \equiv \frac{\kappa_2(K_E)^{\frac{1}{2}} R(K_E)}{\kappa_2(K)^{\frac{1}{2}} R(K)}$$

Example II: Boundary effects



Others: Assumptions Section 2

- (A1) m is twice continuously differentiable
- (A2) f is continuously differentiable and $f(x) > 0$
- (A3) σ^2 is continuous and $\sigma^2(x) > 0$
- (A4) K is a symmetric and bounded pdf with $\kappa_2(K) \equiv \int u^2 K(u) du < \infty$ (finite variance) and $R(K) \equiv \int K(u)^2 du < \infty$ (square integrability)
- (A5) $h \rightarrow 0, nh \rightarrow \infty$ as $n \rightarrow \infty$
- (A6) $x \in [a + h, b - h] / x \in \{a, b\}$

Effective kernels: Insights

- LL regression automatically modifies the kernel to correct the bias exactly to first order when there is asymmetry in the smoothing window

$$\begin{aligned} E[\hat{m}(x)|\mathbf{X}] &= \sum_{i=1}^n w_i(x) m(X_i) \\ &= m(x) \underbrace{\sum_{i=1}^n w_i(x)}_{=1} + m'(x) \underbrace{\sum_{i=1}^n w_i(x)(X_i - x)}_{=0 \text{ for LL}} + R \end{aligned}$$

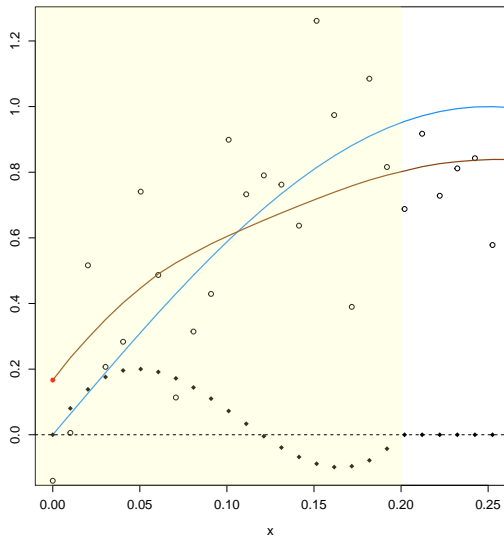
Others: Epanechnikov boundary kernel, Müller (1991, Table 1)

For $0 \leq \rho \leq 1$:

$$K_l(u, \rho) = 6(1+u)(\rho-u) \frac{1}{(1+\rho)^3} \left\{ 1 + 5 \left(\frac{1-\rho}{1+\rho} \right)^2 + 10 \frac{1-\rho}{(1+\rho)^2} u \right\}, u \in [-1, \rho]$$

$$K_r(u, \rho) = K_l(-u, \rho), u \in [-\rho, 1]$$

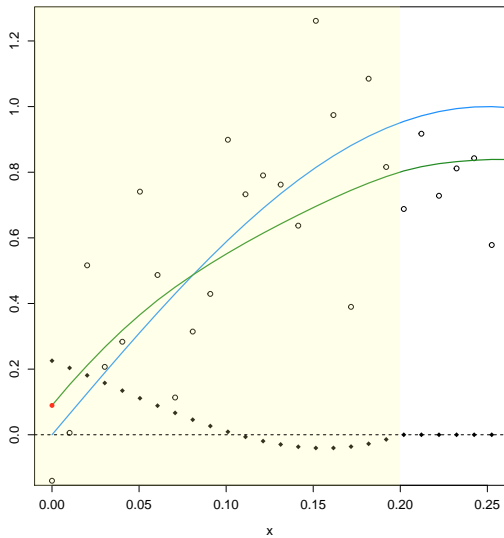
Others: Effective kernel at the boundary



— True curve

— NW boundary adjusted

— LL



• Estimate at 0

• Effective kernel at 0

Others: Optimal bandwidths

- Minimization of (weighted) AMISE:

$$h_{\text{AMISE}} \equiv \arg \min_{h>0} \text{AMISE}(\hat{m}|\mathbf{X})$$

$$h_{\text{AMISE}}^{\text{LL}}(n) = \left\{ \frac{R(K) \int \sigma^2(x) \, dx}{\kappa_2(K)^2 \int (m''(x))^2 f(x) \, dx} \right\}^{1/5} \cdot n^{-1/5}$$

$$h_{\text{AMISE}}^{\text{NW}}(n) = \left\{ \frac{R(K) \int \sigma^2(x) \, dx}{\kappa_2(K)^2 4 \int \left[\frac{1}{2} m''(x) + \frac{f'(x)m'(x)}{f(x)} \right]^2 f(x) \, dx} \right\}^{1/5} \cdot n^{-1/5}$$

Simulation: Extended results

n	Boundary region	Change in MISE to NW [%]	
		Local linear	NW boundary adjusted
50	[0, 0.119)	789.29	1578.07
100	[0, 0.104)	-41.90	1143.28
250	[0, 0.086)	-65.13	1982.67
500	[0, 0.075)	-72.98	14.92
1000	[0, 0.065)	-78.80	-63.95
2500	[0, 0.055)	-85.12	-77.80
5000	[0, 0.047)	-88.50	-83.76
10000	[0, 0.041)	-91.14	-88.26

Note: For $n = 10000$ the number of Monte-Carlo repetitions is 1000 instead of 10000.

Simulation: Robustness check h_{AMISE}

CV-optimal bandwidths for NW and LL compared to the asymptotically optimal bandwidths

n	$h_{\text{CV}}^{\text{NW}}$	$h_{\text{CV}}^{\text{LL}}$	h_{AMISE}	Change to h_{AMISE} [%]	
				Nadaraya-Watson	Local linear
50	0.097	0.128	0.119	−18.59	6.99
100	0.080	0.105	0.104	−22.96	1.25
250	0.062	0.085	0.086	−28.44	−1.55

Note: CV-optimal bandwidths are computed as a mean over 1000 repetitions.

Simulation: Robustness check MISE

n	Boundary region	Change in MISE to NW [%]	
		Local linear	NW boundary adjusted
Results for cross-validated bandwidths			
50	[0, 0.097)	703.01	3373.15
100	[0, 0.080)	−20.87	4561.98
250	[0, 0.062)	−41.54	443.39
Results for asymptotically optimal bandwidths			
50	[0, 0.119)	789.29	1578.07
100	[0, 0.104)	−41.90	1143.28
250	[0, 0.086)	−65.13	1982.67

Others: Cross Validation (CV)

- Leave One Out Cross Validation (LOOCV):

$$CV(h) \equiv \frac{1}{n} \sum_{i=1}^n (Y_i - \hat{m}_{-i}(X_i, h))^2$$

$$\hat{h}_{CV} \equiv \arg \min_{h>0} CV(h)$$

- \hat{h}_{CV} is essentially an unbiased estimator of h_{MISE}

Extensions

- Local polynomial regression (e.g. cubic)
- Multivariate regression
 - Boundary issues more severe
- Estimation of other functionals (e.g. derivatives)

Introductory literature

Monographs

- Fan, J. and I. Gijbels (1996). *Local Polynomial Modelling and its Applications*. Monographs on Statistics and Applied Probability 66. Boca Raton: Chapman & Hall/CRC.
- Wand, M. and M. Jones (1995). *Kernel Smoothing*. Monographs on Statistics and Applied Probability 60. Boca Raton: Chapman & Hall/CRC.

Papers

- Fan, J. (1992). "Design-adaptive nonparametric regression". *Journal of the American Statistical Association* 87 (420), pp. 998–1004.
- Müller, H.-G. (1991). "Smooth optimum kernel estimators near endpoints". *Biometrika* 78 (3), pp. 521–530.