

PoliSci 4782 Political Analysis II

Categorical Outcome Models

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Roadmap

When having categorical outcomes, ask yourself two questions:

- Is the number of categories just equal to 2?
 - If so, a binary outcome model will do (logit or probit).
 - If not, we will need a categorical outcome model (which is also based either logit or probit)
- Is there any intrinsic ordering between different categories (so that we can rank those categories in an order, i.e. low-income, middle-income, high-income)?
 - ordered outcome models (ordinal logit or ordinal probit)
 - unordered outcome models (multinomial logit or multinomial probit)

Ordinal Logit/Probit for Ordered Categorical Outcome

Ordinal Logit/Probit

- It extends from logit/probit.
- In theory, we can combine multiple adjacent categories (since they are ordered) together to form a spectrum.
- We can model the probability of moving from one category to the next (upper) category, given a changeable baseline level (which is linked to the regression intercept).

Ordinal Logit

- Consider an ordered categorical outcome Y that has k possible categories.
- For brevity, those categories are denoted as $a, b, c, \dots k$
- We can set up the following logit models with changing cutpoint c_i :

$$Pr(y > a) = \text{logit}^{-1}(X\beta - c_1)$$

$$Pr(y > b) = \text{logit}^{-1}(X\beta - c_2)$$

$$Pr(y > c) = \text{logit}^{-1}(X\beta - c_3)$$

...

$$Pr(y > j) = \text{logit}^{-1}(X\beta - c_{k-1})$$

Cutpoints

- The parameters c_k (*cutpoints*) get larger as we move to a higher category: $c_1 < c_2 < \dots < c_{k-1}$.
- Cutpoints help locate the predicted categories by logit, but do not meddle in the cause-effect relationships between Y and \mathbf{X} .
- Thus, an ordinal logit is a logit with $k - 1$ intercepts indicating the cutpoints between categories.

Interpreting Ordinal Logit

$$Pr(Y_i \leq j) = \frac{\exp(\beta X_i - c_j)}{1 + \exp(\beta X_i - c_j)}$$

- β indicates how the corresponding variable affects the probability of changing from any given category to the next category, generally speaking (odds interpretation and divide-by-4 apply).
- The intercept c_j indicates the baseline probability that the categorical outcome is no greater than category $j + 1$ (while other explanatory variables are held at 0).
- If the outcome variable has k categories, we will have $k - 1$ intercepts.

Illustrative Example

nes96 dataset from faraway package on U.S. 1996 National Election Study:

- PID: party identification, ordered from strong Democrat to strong Republican (outcome variable)
- income: income level
- age: age group
- educ: education level

Ordinal Logit in R

```
28 library(MASS)
29 model1 <- polr(PID ~ age + as.numeric(educ) + as.numeric(income), nes96)
30 summary(model1)
31
32
```

33:1 lecture 9: categorical outcome

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```
Call:
polr(formula = PID ~ age + as.numeric(educ) + as.numeric(income),
      data = nes96)
```

Coefficients:

	Value	Std. Error	t value
age	0.001135	0.003579	0.317
as.numeric(educ)	0.047513	0.038780	1.225
as.numeric(income)	0.057229	0.010329	5.541

Intercepts:

	Value	Std. Error	t value
strDem weakDem	-0.1507	0.2812	-0.5359
weakDem indDem	0.8001	0.2812	2.8448
indDem indind	1.2811	0.2836	4.5168
indind indRep	1.4442	0.2846	5.0746
indRep weakRep	1.8759	0.2873	6.5294
weakRep strRep	2.7303	0.2942	9.2810

Residual Deviance: 3457.807

AIC: 3475.807

Interpretation


The **coefficient** of income ($\beta = 0.06$): the odds of moving from one category to the next (which applies to any two adjacent PIDs) increases by a factor of $\exp(0.06) = 1.06$ when income rises by one level.


Intercepts indicate the baseline probability for each category when independent variables are held at 0:

- `strDem|weakDem` (the first intercept) is -0.15 , so the probability of being a strong is $\text{invlogit}(-0.15) = 0.46$.
- `weakDem|indDem` is 0.80 , so the probability of being a strong Democrat **or** a weak Democrat is $\text{invlogit}(0.80) = 0.69$.
- ...

Interpretation

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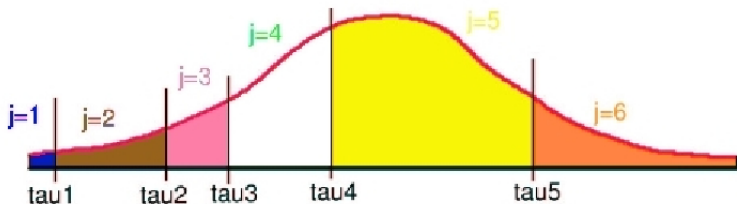
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- What is the baseline probability of being a weak Democrat?
- Answer: $\text{invlogit}(0.80) - \text{invlogit}(-0.15) = 0.23$

Ordinal Probit



- Intuition: imagine you throw a ball from your own yard ($j = 1$) forward and it eventually falls into one of k possible yard divided by $k - 1$ fences (“cutpoints”).
- The probit function $\Phi(X\beta)$ determines the (latent) distance which we don’t care; cutpoints (τ) help translate the latent value into an observed category J .
- The same coefficient interpretation approach as that for probit.

Multinomial Logit/Probit for Unordered Categorical Outcome

Unordered Categorical/Nominal Outcomes

- no superiority and inferiority between categories
- no meaningful distance between categories
- distinction between ordered and unordered categorical variable can be tricky and theory-dependent

Multiple Equation Models

- The new challenge in multinomial regression is that we are unable to put categories to a spectrum (since they are not ordered).
- The solution is to use multiple equations instead of one equation and to focus on one single category each time.
- This approach of using a system of equations is called “multiple equation model,” which has other important applications in issues such as exploring complex causation and dealing with missing data.

Multinomial Logit/Probit

- It is an extension of logit/probit.
- In multinomial logit, we pick one category as the reference group and do a logit on the the probability of observations changing from that category to each of the remaining categories one at a time.
- In multinomial probit, we also pick one category, use standard normal distribution to compare the utility value of that category to each of the remaining categories one at a time, and eventually model the probability that one prefers a given category over the reference category.
- Both multinomial logit and probit yield $k - 1$ sets of intercepts and coefficients, given the outcome has k categories.

Multinomial logit model

- We apply logistical regression to the odds ratio of being in one category vis-a-vis the reference category:

$$\eta_{ij} = X_i\beta_j = \log \frac{p_{ij}}{p_{i1}}, \quad j = 2, \dots, J$$

Note: we have individual index i (for our units) and category index j

- The inverse logit function serves as the mean function by which we can compute the predicted probability of moving from the reference category to the corresponding one:

$$p_{ij} = \frac{\exp(X_{ij}\beta_{ij})}{1 + \sum_{j=2}^J \exp(X_{ij}\beta_{ij})}$$

Illustrative Example

We continue to use `nes96` dataset for demo and treat `PID` as an unordered categorical variable (this is a judgment call)

Multinomial Logit in R

```
67 library(nnet)
68
69 model3 <- multinom(PID ~ as.numeric(age) + as.numeric(educ) + as.numeric(income), nes96)
70
71 summary(model3)
72
```

73:1 (Untitled) ↕

R Script

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```
converged
> summary(model3)
Call:
multinom(formula = PID ~ as.numeric(age) + as.numeric(educ) +
  as.numeric(income), data = nes96)
```

Coefficients:

	(Intercept)	as.numeric(age)	as.numeric(educ)	as.numeric(income)
weakDem	0.7021256	-2.209823e-02	0.04643949	0.001542891
indDem	-1.0150463	-1.918221e-02	0.12621262	0.045812593
indind	-1.6613264	-9.281564e-03	-0.10361522	0.055683638
indRep	-2.0857629	1.825893e-06	0.03347729	0.073664606
weakRep	-1.1482475	-9.190595e-03	0.03984032	0.070180544
strRep	-1.9303261	-1.849700e-03	0.10130247	0.087050320

Std. Errors:

	(Intercept)	as.numeric(age)	as.numeric(educ)	as.numeric(income)
weakDem	0.4817966	0.006315216	0.07119764	0.01722851
indDem	0.5890586	0.007658225	0.08266329	0.02176761
indind	0.8931024	0.011048023	0.12428004	0.03316717
indRep	0.6513791	0.007700525	0.08538329	0.02422957
weakRep	0.5478319	0.006727317	0.07436736	0.02039010
strRep	0.5456234	0.006460797	0.07145470	0.02036117

```
Residual Deviance: 3428.426
AIC: 3476.426
```

Interpreting Intercepts

- Estimated intercepts indicate the baseline probabilities that individuals switch from the reference category to the corresponding category, while other explanatory variables are held at 0.
- "0" intercept indicates the baseline probability of being in the reference category, while other explanatory variables are held at 0.

```
> intercept <- c(0, coef(model3)[ ,1])
> intercept
```

	weakDem	indDem	indind	indRep	weakRep	strRep
0.0000000	0.7021256	-1.0150463	-1.6613264	-2.0857629	-1.1482475	-1.9303261

```
> exp(intercept)/sum(exp(intercept))
```

	weakDem	indDem	indind	indRep	weakRep	strRep
0.24056876	0.48547684	0.08717867	0.04568087	0.02988160	0.07630652	0.03490674

```
>
```

Interpreting Coefficients

- Almost the same as to logit and ordinal logit (odds interpretation and divide-by-4 apply).
- Coefficients represent the log-odds of moving from the baseline category to a given category given a unit change in the corresponding explanatory variable.
- Exponentiated coefficients thus tell us how odds changes in the multiplicative way if the corresponding explanatory variable increases by one unit.

Independence of Irrelevant Alternatives (IIA)

- Theoretically, multinomial logit assumes IIA, which means that our preference between categories A and B is independent to any other category (the red-bus/blue-bus problem).
- This is often time an unrealistic assumption.
 - think about the fact that adding a third candidate often splits the votes
- If IIA is clearly violated, we need multinomial probit, which is theoretically less demanding but computationally more challenging

Multinomial Probit

- It can be run in R with the help of `mlogit` package or `MNP` package.
- It may take a long time to run and end up with unstable, weakly identified results, due to technical challenges.
- Coefficient interpretation is almost the same as to probit and ordinal probit (change in the z-score for the probability of changing from the reference category to the corresponding category).

Coming Up

- No lab or lab assignment this week (instructional break)
- Lecture and lab on duration outcomes next week