Revisiting valid instruments: A Bayesian approach

Jonathan Sweeney

University of Hawaii at Manoa

Applied Micro Workshop, February 28, 2017

IV estimation can be categorized into three uses:

- Simultaneous equation modeling (Wright 1928)
- ► Measurement error (Friedman 1957)
- ► Omitted variables/Causal inference (Angrist 1990)

Why not OLS?

▶ Variable of interest is correlated with the error term.

$$y = \alpha_1 + x\beta_1 + \epsilon$$

Why not OLS?

▶ Variable of interest is correlated with the error term.

$$E[x\epsilon] \neq 0$$

$$\hat{\beta}_1 = (X'X)^{-1}X'y = (X'X)^{-1}X'(X\beta_1 + \epsilon) = \beta_1 + (X'X)^{-1}X'\epsilon$$

Simulation

```
# Generate instrument data
n <- 1000
z <- rnorm(n)
# Specify error terms with for 1st and 2nd stage
e \leftarrow rnorm(n, sd = 1)
u \leftarrow rnorm(n, sd = 1)
# Generate 1st and 2nd stage outcomes
x < -1 + 2 * z + 0.5 * 11 + e
y < -3 + 1.5 * x + u
d <- data.frame(z, x, y)</pre>
```

Simulation

```
# OLS results
ols <- lm(y ~ x)
summary(ols)</pre>
```

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 2.93101 0.03328 88.07 <2e-16 ***
x 1.60188 0.01345 119.13 <2e-16 ***
```

Correct the bias with an IV.

Introduce an instrument, z, strongly correlated with x but not correlated with ϵ .

$$\hat{\beta}_{1,IV} = (Z'X)^{-1}Z'y = (Z'X)^{-1}Z'X\beta_1 + (Z'X)^{-1}Z'\epsilon$$

Simulation

```
# IV results
iv <- ivreg(y ~ x | z, data = d)
summary(iv)</pre>
```

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.02611 0.03497 86.52 <2e-16 ***
x 1.50226 0.01583 94.87 <2e-16 ***
```

Bayesian IV estimation

- Performs well even with a weakly identified model (i.e. weak instruments)
- Accomodates finite samples

Many thanks to Danilo Freire (King College London) for sharing his code.

Bayesian IV estimation: Bayes' Rule

$$p(\theta|y) \propto p(\theta)p(y|\theta)$$

- Observe data y
- ightharpoonup Define model parameters θ
- ▶ Specify likelihood function $p(y|\theta)$
- Specify prior densities $p(\theta)$

Simulation (R/Stan code is on sweenejo.github.io)

$$y = \alpha_1 + \beta_1 x + u$$
$$x = \alpha_2 + \beta_2 z + e$$

- ▶ Use same simulated data as before.
- ▶ Define model parameters $\theta = (\alpha_1, \alpha_2, \beta_1, \beta_2, \sigma_1, \sigma_2, \rho)$

Simulation

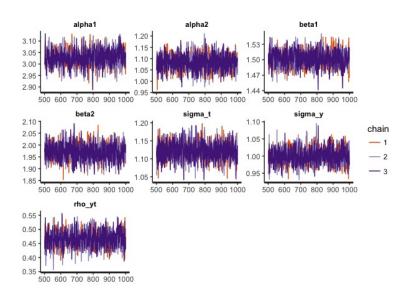
$$\begin{bmatrix} u \\ e \end{bmatrix} = \begin{bmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{bmatrix}$$

▶ Note, this is similar to what Altonji et al. (2005) do.

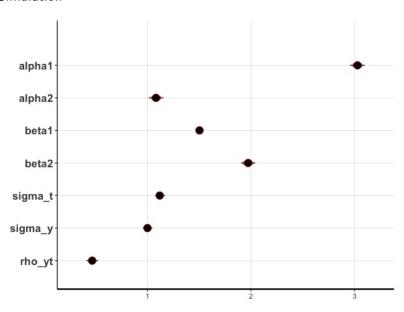
Simulation

- Likelihood function will be a multivariate normal.
- Specify uniform priors for the variance parameters, and flat normals for the others.

Simulation (R/Stan code is on sweenejo.github.io)



Simulation



Bayesian IV estimation

- Performs well even with a weakly identified model (i.e. weak instruments)
- Accomodates finite samples
- Can a Bayesian approach help us address questions of instrument validity?

Yes!

An important advantage of our Bayesian analysis is that neither the exclusion restriction nor the monotonicity assumption is essential, and consequently violations of these assumptions are easily addressed (Imbens and Rubin 1997)

Let's replicate Imbens and Rubin (1997) to see how useful this approach is.

 Use data from randomized community trial of impact of vitamin A on children's survival (Sommer and Zeger 1991)

Type (<i>C</i>)	Assignment (Z)	Vitamin supplements (D)	Survival (Y)	Number of units
Complier or never-taker	0	0	0	74
Complier or never-taker	0	0	1	11,514
Never-taker	1	0	0	34
Never-taker	1	0	1	2,385
Complier	1	1	0	12
Complier	1	1	1	9,663

$$Y \leftarrow C$$

$$C \leftarrow Z$$

- Exclusion restriction: Treatment assignment (Z) is not directly correlated with potential outcomes (Y).
- Strict monotonicity: We can ignore defiers.

Let's relax the exclusion restriction. There may be direct correlation between Z and Y (selection).

- ▶ Define a set of model parameters
- ▶ Probability of survival $\eta_{C,Z}$
- lacktriangle Probability of being a complier ω

Does taking vitamin A increase childrens' survival?

$$CACE = \eta_{C,1} - \eta_{C,0}$$

Specify posterior distribution, which is a product of 5 distributions:

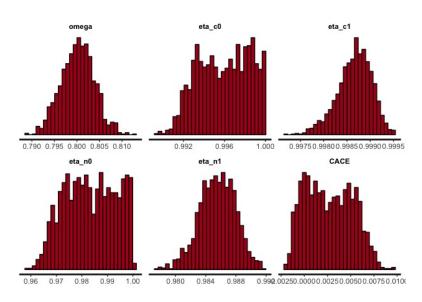
$$p(\omega|C,Z,D,Y) \propto p(\omega)\omega^{N_C}(1-\omega)^{N_N}$$

$$p(\eta_{C,Z}|C,Z,D,Y) \propto p(\eta_{C,Z}) \prod_{i \in C,Z} \eta_{C,Z}^{Y_i} (1-\eta_{C,Z})^{1-Y_i}$$

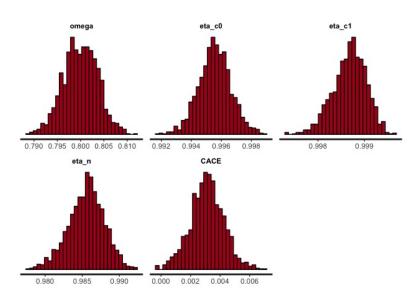
- ▶ Note: We do not observe the compliance status of the control group. We'll treat this as discrete missing data, and marginalize it out of the distribution.
- \blacktriangleright This means we will sample from ω to specify compliance status of missing observations.

We'll specify flat beta distributions (beta(1,1)) as priors for all parameters.

Estimates relaxing the exclusion restriction.



Estimates with the exclusion restriction. ($\eta_N = \eta_{N,0} = \eta_{N,1}$)



- ➤ Supplementing with vitamin A increases childrens' survival by ~0.003%, or saves ~3 children for every 1000.
- ▶ Without the exclusion restriction the distribution of CACE has a flat top over a wide range, although the mean is still ~0.003%.

The Bayesian likelihood approach allows us to relax the fundamental IV assumptions, including the exclusion restriction and strict monotonicity.

- Instead of relying on a handwaving defense of our instruments validity, we can evaluate parameter sensitivity with and without valid instrument assumptions.
- ► The next steps are to extend this approach to models with continuous parameters.