

Warm-up

Let \mathcal{H} be a RKHS associated with a positive definite kernel $k : \mathbf{X} \times \mathbf{X} \rightarrow \mathbb{R}$.

1. Prove that for all $(x, y) \in \mathbf{X} \times \mathbf{X}$,

$$|f(x) - f(y)| \leq \|f\|_{\mathcal{H}} \|k(x, \cdot) - k(y, \cdot)\|_{\mathcal{H}}.$$

The proof follows from Cauchy-Schwarz inequality as, for all $(x, y) \in \mathbf{X}^2$,

$$|f(x) - f(y)| = |\langle f, k(x, \cdot) \rangle_{\mathcal{H}} - \langle f, k(y, \cdot) \rangle_{\mathcal{H}}| = |\langle f, k(x, \cdot) - k(y, \cdot) \rangle_{\mathcal{H}}|.$$

2. Prove that the kernel k associated with \mathcal{H} is unique, i.e. if \tilde{k} is another positive definite kernel satisfying the RKHS properties for \mathcal{H} , then $k = \tilde{k}$.

Write, for all $x \in \mathbf{X}$,

$$\|k(x, \cdot) - \tilde{k}(x, \cdot)\|_{\mathcal{H}}^2 = \langle k(x, \cdot) - \tilde{k}(x, \cdot), k(x, \cdot) - \tilde{k}(x, \cdot) \rangle = k(x, x) - \tilde{k}(x, x) + \tilde{k}(x, x) - k(x, x) = 0.$$

3. Prove that for all $x \in \mathbf{X}$, the function defined on \mathcal{H} by $\delta_x : f \mapsto f(x)$ is continuous.

Kernel Ridge regression

Let \mathcal{H} be a RKHS on \mathcal{X} with kernel k . We consider the regression model $Y_i = f^*(X_i) + \xi_i$, $i \in \{1, \dots, n\}$, with ξ_i , $1 \leq i \leq n$, independent centered noise with finite variance. The unknown function f^* is estimated by the solution \hat{f} of the convex minimization problem

$$\hat{f} = \operatorname{argmin}_{f \in \mathcal{H}} \left\{ \frac{1}{n} \sum_{i=1}^n (Y_i - f(X_i))^2 + \frac{\lambda}{n} \|f\|_{\mathcal{H}}^2 \right\},$$

with $\lambda > 0$.

Solving Kernel ridge regression

1. Check that $\hat{f} : x \mapsto \sum_{j=1}^n \hat{\beta}_j k(X_j, x)$ where $\hat{\beta} = (\hat{\beta}_1, \dots, \hat{\beta}_n)^\top$ is solution to

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^n} \{ \|Y - K\beta\|^2 + \lambda \beta^\top K \beta \}$$

with K defined by $K = (k(X_i, X_j))_{1 \leq i, j \leq n}$. Comment on this result.

There exists β such that, for all x ,

$$\hat{f}(x) = \sum_{j=1}^n \beta_j k(X_j, x).$$

This yields

$$\frac{1}{n} \sum_{i=1}^n (Y_i - f(X_i))^2 + \frac{\lambda}{n} \|f\|_{\mathcal{H}}^2 = \frac{1}{n} \sum_{i=1}^n (Y_i - \sum_{j=1}^n \beta_j k(X_j, X_i))^2 + \frac{\lambda}{n} \langle \sum_{j=1}^n \beta_j k(X_j, \cdot), \sum_{i=1}^n \beta_i k(X_i, \cdot) \rangle,$$

which gives the result, since

$$\langle \sum_{j=1}^n \beta_j k(X_j, \cdot), \sum_{i=1}^n \beta_i k(X_i, \cdot) \rangle = \sum_{i,j=1}^n \beta_i \beta_j k(X_i, X_j) = \beta^\top K \beta.$$

2. Assume that K is non-singular. Give an explicit expression for $\hat{\beta}$.

Write, for all β ,

$$L(\beta) = \|Y - K\beta\|_2^2 + \lambda \beta^\top K \beta.$$

The gradient of L is then given by

$$\begin{aligned} \nabla L(\beta) &= -2K^\top (Y - K\beta) + \lambda(K\beta + K^\top \beta) \\ &= -2K(Y - K\beta) + 2\lambda K\beta. \end{aligned}$$

The minimum $\hat{\beta}$ of L satisfies

$$\begin{aligned} &\Leftrightarrow -2K(Y - K\hat{\beta}) + 2\lambda K\hat{\beta} = 0 \\ &\Leftrightarrow \hat{\beta} = (K + \lambda I)^{-1} Y. \end{aligned}$$

Bias and variance

We assume that $f^* \in \mathcal{H}$ and we write

$$f_V^* : x \mapsto \sum_{i=1}^n \beta_i^* k(X_i, x)$$

for the projection of f^* onto the linear span $V = \text{span}\{k(X_i, \cdot) : i = 1, \dots, n\}$, with respect to the Hilbert norm $\|\cdot\|_{\mathcal{H}}$. We write $K = \sum_{i=1}^n \lambda_i u_i u_i^\top$ for an eigenvalue decomposition of K .

1. Check that

$$K\hat{\beta} = \sum_{i=1}^n \frac{\lambda_i}{\lambda_i + \lambda} \langle Y, u_i \rangle u_i \quad \text{with} \quad Y = (Y_1, \dots, Y_n)^\top.$$

Since $(u_i)_{1 \leq i \leq n}$ is an orthonormal basis of \mathbb{R}^n ,

$$\begin{aligned} K\hat{\beta} &= \sum_{i=1}^n \langle K\hat{\beta}, u_i \rangle u_i \\ &= \sum_{i=1}^n \langle K(K + \lambda I)^{-1} Y, u_i \rangle u_i \\ &= \sum_{i=1}^n \langle Y, (K + \lambda I)^{-1} K u_i \rangle u_i \\ &= \sum_{i=1}^n \frac{\lambda_i}{\lambda + \lambda_i} \langle Y, u_i \rangle u_i. \end{aligned}$$

2. Check that

$$\|\mathbb{E}[K\hat{\beta}] - K\beta^*\|_2^2 = \sum_{i=1}^n \left(\frac{\lambda\lambda_i}{\lambda_i + \lambda} \right)^2 \langle \beta^*, u_i \rangle^2.$$

First, note that, for all $1 \leq i \leq n$,

$$\langle \mathbb{E}[Y], u_i \rangle = \langle K\beta^*, u_i \rangle = \langle \beta^*, Ku_i \rangle = \lambda_i \langle \beta^*, u_i \rangle.$$

Consequently,

$$\begin{aligned} \|\mathbb{E}[K\hat{\beta}] - K\beta^*\|_2^2 &= \left\| \sum_{i=1}^n \frac{\lambda_i}{\lambda_i + \lambda} \langle \mathbb{E}[Y], u_i \rangle u_i - \sum_{i=1}^n \langle K\beta^*, u_i \rangle u_i \right\|_2^2 \\ &= \left\| \sum_{i=1}^n \left(\frac{\lambda_i^2}{\lambda_i + \lambda} - \lambda_i \right) \langle \beta^*, u_i \rangle u_i \right\|_2^2 \\ &= \sum_{i=1}^n \left(\frac{\lambda\lambda_i}{\lambda_i + \lambda} \right)^2 \langle \beta^*, u_i \rangle^2. \end{aligned}$$

3. We assume henceforth that the $\varepsilon_i = Y_i - f^*(X_i)$, $i = 1, \dots, n$, have a covariance $\mathbb{V}[\varepsilon] = \sigma^2 I_n$. Check that the covariance matrix of $K\hat{\beta}$ is equal to

$$\mathbb{V}[K\hat{\beta}] = \sum_{i=1}^n \left(\frac{\lambda_i \sigma}{\lambda_i + \lambda} \right)^2 u_i u_i^\top.$$

Since $\hat{\beta} = (K + \lambda I)^{-1}y$,

$$\begin{aligned} \mathbb{V}[K\hat{\beta}] &= K \mathbb{V}[(K + \lambda I)^{-1}Y] K^\top \\ &= K(K + \lambda I)^{-1} \mathbb{V}[Y] (K + \lambda I)^{-1} K \\ &= \sigma^2 K^2 (K + \lambda I)^{-2} \\ &= \sum_{i=1}^n \left(\frac{\lambda_i \sigma}{\lambda_i + \lambda} \right)^2 u_i u_i^\top, \end{aligned}$$

using the eigenvector decomposition of K .

4. We define $\|f\|_n^2 := \frac{1}{n} \sum_{i=1}^n f(X_i)^2$. Prove that

$$\mathbb{E} \left[\|\hat{f} - f^*\|_n^2 \right] = \frac{1}{n} \sum_{i=1}^n \left(\frac{\lambda_i}{\lambda + \lambda_i} \right)^2 (\lambda^2 \langle \beta^*, u_i \rangle^2 + \sigma^2).$$