## RIDGE REGRESSION

## Generalized Ridge regression

Consider the regression model

$$Y = X\beta_* + \varepsilon \,,$$

where  $X \in \mathbb{R}^{n \times d}$ ,  $\beta_*$  is an unknown vector in  $\mathbb{R}^d$  and  $\varepsilon \sim \mathcal{N}(0, \sigma^2 I_n)$ . Define the generalized Ridge estimator by:

$$\widehat{\beta} \in \operatorname{Argmin}_{\beta \in \mathbb{R}^d} \left\{ (Y - X\beta)^\top W (Y - X\beta) + (\beta - \beta_0)^\top \Delta (\beta - \beta_0) \right\},\,$$

where  $\beta_0 \in \mathbb{R}^d$ ,  $W \in \mathbb{R}^{n \times n}$  is a diagonal matrix with elements in [0,1],  $\Delta \in \mathbb{R}^{d \times d}$  is a definite-positive matrix.

- 1. Provide the expression of  $\widehat{\beta}$  when  $\beta_0 = 0$ ,  $W = I_n$  and  $\Delta = \lambda I_d$  where  $\lambda > 0$ .

  Proof in lecture notes.
- 2. Solve the optimization problem in the general case.

For all  $\beta \in \mathbb{R}^d$ , write

$$\mathcal{L}(\beta) = (Y - X\beta)^{\top} W(Y - X\beta) + (\beta - \beta_0)^{\top} \Delta(\beta - \beta_0).$$

Therefore, for all  $\beta \in \mathbb{R}^d$ ,

$$\nabla \mathcal{L}(\beta) = 2\left(\left(X^{\top}WX + \Delta\right)\beta - \Delta\beta_0 - X^{\top}WY\right).$$

Note that  $X^{\top}WX + \Delta$  is definite-positive so that  $\nabla \mathcal{L}(\beta) = 0$  has a unique solution given by

$$\widehat{\beta} = \left( X^{\top} W X + \Delta \right)^{-1} \left( \Delta \beta_0 + X^{\top} W Y \right).$$

3. Compute  $\mathbb{E}[\widehat{\beta}]$  and show that the estimator is unbiased when  $\beta_0 = \beta_*$ .

Assuming that the design is not random,

$$\mathbb{E}[\widehat{\beta}] = (X^{\top}WX + \Delta)^{-1} (\Delta\beta_0 + X^{\top}W\mathbb{E}[Y]).$$

This yields

$$\mathbb{E}[\widehat{\beta}] = \left(X^{\top}WX + \Delta\right)^{-1} \left(\Delta\beta_0 + X^{\top}WX\beta_*\right).$$

In the case where  $\beta_0 = \beta_*$ ,

$$\mathbb{E}[\widehat{\beta}] = (X^{\top}WX + \Delta)^{-1} (X^{\top}WX + \Delta) \beta_* = \beta_*$$

and the estimator is unbiased.

4. Compute  $\mathbb{V}[\widehat{\beta}]$  and the mean squared error  $\mathbb{E}[\|\widehat{\beta} - \beta_*\|_2^2]$  when  $\beta_0 = \beta_*$ .

By definition of  $\widehat{\beta}$ ,

$$\begin{split} \mathbb{V}[\widehat{\beta}] &= \left( \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{X} + \boldsymbol{\Delta} \right)^{-1} \boldsymbol{X}^\top \boldsymbol{W} \mathbb{V}[\boldsymbol{Y}] \boldsymbol{W}^\top \boldsymbol{X} \left( \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{X} + \boldsymbol{\Delta} \right)^{-1} \\ &= \sigma^2 \left( \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{X} + \boldsymbol{\Delta} \right)^{-1} \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{W}^\top \boldsymbol{X} \left( \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{X} + \boldsymbol{\Delta} \right)^{-1} \\ &= \sigma^2 \left( \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{X} + \boldsymbol{\Delta} \right)^{-1} \boldsymbol{X}^\top \boldsymbol{W}^2 \boldsymbol{X} \left( \boldsymbol{X}^\top \boldsymbol{W} \boldsymbol{X} + \boldsymbol{\Delta} \right)^{-1} \,. \end{split}$$

If  $\beta_0 = \beta_*$ , as the estimator is unbiased,

$$\begin{split} \mathbb{E}[\|\widehat{\beta} - \beta_*\|_2^2] &= \operatorname{Trace}\left(\mathbb{V}[\widehat{\beta}]\right) \\ &= \sigma^2 \operatorname{Trace}\left(\left(X^\top W X + \Delta\right)^{-1} X^\top W^2 X \left(X^\top W X + \Delta\right)^{-1}\right) \\ &= \sigma^2 \operatorname{Trace}\left(X^\top W^2 X \left(X^\top W X + \Delta\right)^{-2}\right) \,. \end{split}$$

5. Assume that  $W = I_n$ ,  $\beta_0 = 0$  and  $\Delta = V\Lambda V^{\top}$  where  $X = UDV^{\top}$  is a singular value decomposition of X and  $\Lambda$  is a diagonal matrix with positive diagonal components. Provide an expression of  $\hat{\beta}$  as a function of U, D, V,  $\Lambda$  and Y.

In the proposed setting,

$$\widehat{\beta} = (X^{\top}X + \Delta)^{-1} X^{\top}Y.$$

Let  $X = UDV^{\top}$  be a singular value decomposition of X and choose  $\Delta = V\Lambda V^{\top}$ . Then,

$$\begin{split} \widehat{\beta} &= \left( (UDV^\top)^\top UDV^\top + V\Lambda V^\top \right)^{-1} (UDV^\top)^\top Y \\ &= \left( VD^\top U^\top UDV^\top + V\Lambda V^\top \right)^{-1} VD^\top U^\top Y \\ &= V \left( D^\top D + \Lambda \right)^{-1} D^\top U^\top Y \,. \end{split}$$

Contrary to the classical Ridge estimator, this estimator shrinks values of  $\beta$  with a different penalty for each component thanks to the matrix  $\Lambda$ .