KERNELS

Warm-up

Let \mathcal{H} be a RKHS associated with a positive definite kernel $k: X \times X \to \mathbb{R}$.

1. Prove that for all $(x, y) \in X \times X$,

$$|f(x) - f(y)| \le ||f||_{\mathcal{H}} ||k(x, \cdot) - k(y, \cdot)||_{\mathcal{H}}.$$

The proof follows from Cauchy-Schwarz inequality as, for all $(x, y) \in X^2$,

$$|f(x) - f(y)| = |\langle f, k(x, \cdot) \rangle_{\mathcal{H}} - \langle f, k(x, \cdot) \rangle_{\mathcal{H}}| = |\langle f, k(x, \cdot) - k(y, \cdot) \rangle_{\mathcal{H}}|.$$

2. Prove that the kernel k associated with \mathcal{H} is unique, i.e. if \widetilde{k} is another potitive definite kernel satisfying the RKHS properties for \mathcal{H} , then $k = \widetilde{k}$.

Write, for all $x \in X$,

$$||k(x,\cdot)-\widetilde{k}(x,\cdot)||_{\mathcal{U}}^2 = \langle k(x,\cdot)-\widetilde{k}(x,\cdot),k(x,\cdot)-\widetilde{k}(x,\cdot)\rangle = k(x,x)-\widetilde{k}(x,x)+\widetilde{k}(x,x)-k(x,x) = 0.$$

3. Prove that for all $x \in X$, the function defined on \mathcal{H} by $\delta_x : f \mapsto f(x)$ is continuous.

Kernel Ridge regression

Let \mathcal{H} be a RKHS on \mathcal{X} with kernel k. We consider the regression model $Y_i = f^*(X_i) + \xi_i$, $i \in \{1, ..., n\}$, with ξ_i , $1 \le i \le n$, independent centered noise with finite variance. The unknown function f^* is estimated by the solution \widehat{f} of the convex minimization problem

$$\widehat{f} = \operatorname*{argmin}_{f \in \mathcal{H}} \left\{ \frac{1}{n} \sum_{i=1}^{n} (Y_i - f(X_i))^2 + \frac{\lambda}{n} ||f||_{\mathcal{H}}^2 \right\},$$

with $\lambda > 0$.

Solving Kernel ridge regression

1. Check that $\widehat{f}: x \mapsto \sum_{j=1}^n \widehat{\beta}_j k(X_j, x)$ where $\widehat{\beta} = (\widehat{\beta}_1, \dots, \widehat{\beta}_n)^{\top}$ is solution to

$$\widehat{\beta} = \operatorname*{argmin}_{\beta \in \mathbb{R}^n} \left\{ \|Y - K\beta\|^2 + \lambda \beta^\top K\beta \right\}$$

with K defined by $K = (k(X_i, X_j))_{1 \le i,j \le n}$. Comment on this result.

There exists β such that, for all x,

$$\widehat{f}(x) = \sum_{j=1}^{n} \beta_j k(X_j, x).$$

This yields

$$\frac{1}{n}\sum_{i=1}^{n}(Y_{i}-f(X_{i}))^{2}+\frac{\lambda}{n}\|f\|_{\mathcal{H}}^{2}=\frac{1}{n}\sum_{i=1}^{n}(Y_{i}-\sum_{j=1}^{n}\beta_{j}k(X_{j},X_{i}))^{2}+\frac{\lambda}{n}\langle\sum_{j=1}^{n}\beta_{j}k(X_{j},\cdot),\sum_{i=1}^{n}\beta_{i}k(X_{i},\cdot)\rangle,$$

which gives the result, since

$$\langle \sum_{j=1}^{n} \beta_j k(X_j, \cdot), \sum_{i=1}^{n} \beta_i k(X_i, \cdot) \rangle = \sum_{i,j=1}^{n} \beta_i \beta_j k(X_i, X_j) = \beta^{\top} K \beta.$$

2. Assume that K is non-singular. Give an explicit expression for $\widehat{\beta}$.

Write, for all β ,

$$L(\beta) = \|Y - K\beta\|_2^2 + \lambda \beta^\top K\beta.$$

The gradient of L is then given by

$$\nabla L(\beta) = -2K^{\top}(Y - K\beta) + \lambda(K\beta + K^{\top}\beta)$$
$$= -2K(Y - K\beta) + 2\lambda K\beta.$$

The minimum $\widehat{\beta}$ of L satisfies

$$\Leftrightarrow -2K(Y - K\widehat{\beta}) + 2\lambda K\widehat{\beta} = 0$$

$$\Leftrightarrow \widehat{\beta} = (K + \lambda I)^{-1}Y.$$

Bias and variance

We assume that $f^* \in \mathcal{H}$ and we write

$$f_V^*: x \mapsto \sum_{i=1}^n \beta_i^* k(X_i, x)$$

for the projection of f^* onto the linear span $V = \operatorname{span}\{k(X_i,.): i=1,\ldots,n\}$, with respect to the Hilbert norm $\|\cdot\|_{\mathcal{H}}$. We write $K = \sum_{i=1}^n \lambda_i u_i u_i^{\top}$ for an eigenvalue decomposition of K.

1. Check that

$$K\widehat{\beta} = \sum_{i=1}^{n} \frac{\lambda_i}{\lambda_i + \lambda} \langle Y, u_i \rangle u_i \quad \text{with} \quad Y = (Y_1, \dots, Y_n)^{\top}.$$

Since $(u_i)_{1 \leq i \leq n}$ is an orthonormal basis of \mathbb{R}^n ,

$$K\widehat{\beta} = \sum_{i=1}^{n} \langle K\widehat{\beta}, u_i \rangle u_i$$

$$= \sum_{i=1}^{n} \langle K(K + \lambda I)^{-1} Y, u_i \rangle u_i$$

$$= \sum_{i=1}^{n} \langle Y, (K + \lambda I)^{-1} K u_i \rangle u_i$$

$$= \sum_{i=1}^{n} \frac{\lambda_i}{\lambda + \lambda_i} \langle Y, u_i \rangle u_i.$$

2. Check that

$$\|\mathbb{E}[K\widehat{\beta}] - K\beta^*\|_2^2 = \sum_{i=1}^n \left(\frac{\lambda \lambda_i}{\lambda_i + \lambda}\right)^2 \langle \beta^*, u_i \rangle^2.$$

First, note that, for all $1 \le i \le n$,

$$\langle \mathbb{E}[Y], u_i \rangle = \langle K\beta^*, u_i \rangle = \langle \beta^*, Ku_i \rangle = \lambda_i \langle \beta^*, u_i \rangle.$$

Consequently,

$$\|\mathbb{E}[K\widehat{\beta}] - K\beta^*\|^2 = \left\| \sum_{i=1}^n \frac{\lambda_i}{\lambda_i + \lambda} \langle \mathbb{E}[Y], u_i \rangle u_i - \sum_{i=1}^n \langle K\beta^*, u_i \rangle u_i \right\|_2^2$$

$$= \left\| \sum_{i=1}^n \left(\frac{\lambda_i^2}{\lambda_i + \lambda} - \lambda_i \right) \langle \beta^*, u_i \rangle u_i \right\|_2^2$$

$$= \sum_{i=1}^n \left(\frac{\lambda \lambda_i}{\lambda_i + \lambda} \right)^2 \langle \beta^*, u_i \rangle^2.$$

3. We assume henceforth that the $\varepsilon_i = Y_i - f^*(X_i)$, i = 1, ..., n, have a covariance $\mathbb{V}[\varepsilon] = \sigma^2 I_n$. Check that the covariance matrix of $K\widehat{\beta}$ is equal to

$$\mathbb{V}[K\widehat{\beta}] = \sum_{i=1}^{n} \left(\frac{\lambda_i \sigma}{\lambda_i + \lambda}\right)^2 u_i u_i^{\top}.$$

Since $\widehat{\beta} = (K + \lambda I)^{-1} y$,

$$\begin{aligned} \mathbb{V}[K\widehat{\beta}] &= K \mathbb{V}[(K+\lambda I)^{-1}Y]K^{\top} \\ &= K(K+\lambda I)^{-1} \mathbb{V}[Y](K+\lambda I)^{-1}K \\ &= \sigma^{2}K^{2}(K+\lambda I)^{-2} \\ &= \sum_{i=1}^{n} \left(\frac{\lambda_{i}\sigma}{\lambda_{i}+\lambda}\right)^{2} u_{i}u_{i}^{\top}, \end{aligned}$$

using the eigenvector decomposition of K.

4. We define $||f||_n^2 := \frac{1}{n} \sum_{i=1}^n f(X_i)^2$. Prove that

$$\mathbb{E}\left[\|\widehat{f} - f^*\|_n^2\right] = \frac{1}{n} \sum_{i=1}^n \left(\frac{\lambda_i}{\lambda + \lambda_i}\right)^2 \left(\lambda^2 \langle \beta^*, u_i \rangle^2 + \sigma^2\right).$$