# Exercise 3.1 MLE for the Bernoulli/ binomial model

$$\frac{d}{d\theta}p(D|\theta) = \frac{d}{d\theta}(\theta^{N_1}(1-\theta)^{N_0}) \tag{1}$$

$$= N_1 \theta^{N_1 - 1} (1 - \theta)^{N_0} - N_0 \theta^{N_1} (1 - \theta)^{N_0 1}$$
(2)

$$= \theta^{N_1 - 1} (1 - \theta)^{N_0 - 1} (N_1 (1 - \theta) - N_0 \theta) \tag{3}$$

$$= \theta^{N_1 - 1} (1 - \theta)^{N_0 - 1} (N_1 - N\theta) \tag{4}$$

$$\therefore \theta_{\text{MLE}} = \frac{N_1}{N} \tag{5}$$

### Exercise 3.2 Marginal likelihood for the Beta-Bernoulli model

$$p(D) = \frac{[(\alpha_1)\cdots(\alpha_1+N_1-1)][(\alpha_0)\cdots(\alpha_0+N_0-1)]}{(\alpha)\cdots(\alpha+N-1)}$$
(6)

$$= \frac{[(\alpha_1)\cdots(\alpha_1+N_1-1)][(\alpha_0)\cdots(\alpha_0+N_0-1)]}{(\alpha_1+\alpha_0)\cdots(\alpha_1+\alpha_0+N-1)}$$
(7)

$$= \frac{\Gamma(\alpha_1 + \alpha_0)}{\Gamma(\alpha_1 + \alpha_0 + N)} [(\alpha_1) \cdots (\alpha_1 + N_1 - 1)] [(\alpha_0) \cdots (\alpha_0 + N_0 - 1)]$$
(8)

$$= \frac{\Gamma(\alpha_1 + \alpha_0)}{\Gamma(\alpha_1 + \alpha_0 + N)} \frac{\Gamma(\alpha_1 + N_1)}{\Gamma(\alpha_1)} \frac{\Gamma(\alpha_0 + N_0)}{\Gamma(\alpha_0)}$$
(9)

$$= \frac{\Gamma(\alpha_1 + N_1)\Gamma(\alpha_0 + N_0)}{\Gamma(\alpha_1 + \alpha_0 + N)} \frac{\Gamma(\alpha_1 + \alpha_0)}{\Gamma(\alpha_1)\Gamma(\alpha_0)}$$
(10)

# Exercise 3.3 Posterior prdictive for Beta-Binomial model

$$Bb(1|\alpha'_1, \alpha'_0, 1) = \frac{B(1 + \alpha'_1, 1 - 1 + \alpha'_0)}{B(\alpha'_1, \alpha'_0)} \begin{pmatrix} 1\\1 \end{pmatrix}$$
(11)

$$=\frac{B(1+\alpha_1',\alpha_0')}{B(\alpha_1',\alpha_0')}\tag{12}$$

$$= \frac{\Gamma(\alpha_1' + \alpha_0')}{\Gamma(\alpha_1')\Gamma(\alpha_0')} \frac{\Gamma(1 + \alpha_1')\Gamma(\alpha_0')}{\Gamma(1 + \alpha_1' + \alpha_0')}$$
(13)

$$= \frac{\Gamma(\alpha_1' + \alpha_0')}{\Gamma(\alpha_1')\Gamma(\alpha_0')} \frac{\alpha_1'\Gamma(\alpha_1')\Gamma(\alpha_0')}{(\alpha_1' + \alpha_0')\Gamma(\alpha_1' + \alpha_0')}$$
(14)

$$=\frac{\alpha_1'}{\alpha_1'+\alpha_0'}\tag{15}$$

# Exercise 3.4 Beta updating from censored likelihood

$$p(\theta, X < 3) = p(\theta)p(X < 3|\theta) \tag{16}$$

$$= p(\theta)(\sum_{k=0}^{2} p(X = k|\theta))$$
 (17)

$$= p(\theta) \left( \sum_{k=0}^{2} \theta^{k} (1 - \theta)^{(5-k)} \right)$$
 (18)

$$= Beta(\theta|1,1)(\sum_{k=0}^{2} \theta^{k} (1-\theta)^{(5-k)})$$
(19)

$$=\sum_{k=0}^{2} \theta^{k} (1-\theta)^{(5-k)} \tag{20}$$

## Exercise 3.5 Uninformative prior for log-odds ratio

$$p(\theta) = p(\phi) \left| \frac{d\phi}{d\theta} \right| \tag{21}$$

$$= p(\phi)\theta^{-1}(1-\theta)^{-1} \tag{22}$$

$$\propto Beta(\theta|0,0) \ (\because p(\phi) \propto 1)$$
 (23)

#### Exercise 3.6 MLE for the Poisson distribution

$$D = \{x_1, x_2, \dots, x_N\} \tag{24}$$

$$p(D|\lambda) = \prod_{i=1}^{N} Poi(x_i|\lambda)$$
 (25)

$$= e^{-N\lambda} \frac{\lambda^{\sum_{i=1}^{N} x_i}}{\prod_{i=1}^{N} (x_i!)}$$
 (26)

$$\log p(D|\lambda) = -N\lambda + \sum_{i=1}^{N} x_i \log \lambda - \sum_{i=1}^{N} \log x_i!$$
(27)

$$\frac{\partial}{\partial \lambda} \log p(D|\lambda) = -N + \frac{1}{\lambda} \sum_{i=1}^{N} x_i$$
 (28)

$$\therefore \lambda_{MLE} = \frac{1}{N} \sum_{i=1}^{N} x_i \tag{29}$$

### Exercise 3.7 Bayesian analysis of the Poisson distribution

(a)

$$p(\lambda|D) \propto p(\lambda)p(D|\lambda)$$
 (30)

$$\propto \lambda^{a-1} e^{-\lambda b} e^{-N\lambda} \frac{\lambda^{\sum_{i=1}^{N} x_i}}{\prod_{i=1}^{N} (x_i!)}$$
(31)

$$= e^{-\lambda(N+b)} \frac{\lambda^{\sum_{i=1}^{N} x_i}}{\prod_{i=1}^{N} (x_i!)}$$
(32)

$$\propto Ga(\lambda|a + \sum_{i=1}^{N} x_i, b + N) \tag{33}$$

(b)

$$\frac{a + \sum_{i=1}^{N} x_i}{b + N} \to \frac{1}{N} \sum_{i=1}^{N} x_i \tag{34}$$

$$=\lambda_{MLE} \tag{35}$$

### Exercise 3.8 MLE for the uniform distribution

(a)

$$D = \{x_1, \dots, x_N\} \tag{36}$$

$$p(D|a) = \prod_{i=1}^{N} \frac{1}{2a} I(x_i \in [-a, a])$$
(37)

If  $\forall i - a \leq x_i \leq a$ , then  $p(D|a) = \frac{1}{(2a)^n}$ . More smaller a, more larger p(D|a).  $\hat{a} = \max\{|x_1|, \ldots, |x_N|\}$ 

(b)

$$p(x_{n+1}|\hat{a}) = \frac{1}{2\hat{a}}I(x_{n+1} \in [-\hat{a}, \hat{a}])$$
(38)

$$= \begin{cases} 0 & (x_{n+1} \notin [-\hat{a}, \hat{a}]) \\ \frac{1}{2\hat{a}} & (x_{n+1} \in [-\hat{a}, \hat{a}]) \end{cases}$$
(39)

(c)

If we use MLE approach, the probability between  $-\hat{a}$  and  $\hat{a}$  is 0. Bayesian approach with introducing a wide range prior is better.

# Exercise 3.9 Bayesian analysis of the uniform distribution

$$p(\theta|D) = \frac{p(D,\theta)}{p(D)} \tag{40}$$

$$= \begin{cases} \frac{(N+K)b^N}{K} p(D,\theta) & (m \le b) \\ \frac{(N+K)m^{N+K}}{Kb^K} p(D,\theta) & (m > b) \end{cases}$$

$$\tag{41}$$

$$p(\theta|D) = \frac{p(D,\theta)}{p(D)}$$

$$= \begin{cases} \frac{(N+K)b^{N}}{K} p(D,\theta) & (m \le b) \\ \frac{(N+K)m^{N+K}}{Kb^{K}} p(D,\theta) & (m > b) \end{cases}$$

$$= \begin{cases} \frac{(N+K)b^{N+K}}{Kb^{K}} p(D,\theta) & (m > b) \\ \frac{(N+K)m^{N+K}}{\theta^{N+K+1}} I(\theta \ge \max(D,b)) & (m \le b) \\ \frac{(N+K)m^{N+K}}{\theta^{N+K+1}} I(\theta \ge \max(D,b)) & (m > b) \end{cases}$$

$$= (N+K) \{ \max(D,b) \}^{N+K} \theta^{-(N+K+1)} I(\theta \ge \max(D,b))$$

$$= Poynto(\theta|N+K) \max(D,b)$$

$$= Poynto(\theta|N+K) \max(D,b)$$

$$= (A4)$$

$$= (N+K)\{\max(D,b)\}^{N+K}\theta^{-(N+K+1)}I(\theta \ge \max(D,b))$$
(43)

$$= \operatorname{Pareto}(\theta|N+K, \max(D, b)) \tag{44}$$

### Exercise 3.10 Taxicab (tramcar) problem

(a)

$$p(\theta|\{100\}) = \operatorname{Pareto}(\theta|1, 100) \tag{45}$$

$$=100\theta^{-2}I(\theta \ge 100) \tag{46}$$

(b)

mean

not exist

mode

100

median

$$\int_{100}^{x} 100\theta^{-2} I(\theta \ge 100) d\theta = \frac{1}{2}$$
 (47)

$$100 \int_{100}^{x} \theta^{-2} d\theta = \frac{1}{2} \tag{48}$$

$$-100\left[\frac{1}{\theta}\right]_{100}^{x} = \frac{1}{2} \tag{49}$$

$$\frac{1}{x} - \frac{1}{100} = -\frac{1}{200}$$

$$x = 200$$
(50)

$$x = 200 \tag{51}$$

(c)

$$p(D'|D,\alpha) = \int_{\theta} p(D'|\theta)p(\theta|D,\alpha)d\theta$$
 (52)

$$= \int_{\theta} U(x|0,\theta) \operatorname{Pareto}(\theta|1,m) d\theta \tag{53}$$

$$= \int_{\theta} \frac{1}{\theta} I[0 \le x \le \theta] m \theta^{-2} I(\theta \ge m) d\theta \tag{54}$$

$$= m \int_{\theta} \theta^{-3} I[0 \le x \le \theta] I[\theta \ge m] d\theta \tag{55}$$

$$= m\left[-\frac{1}{2}\theta^{-2}\right]_{\max(x,m)}^{\infty} \tag{56}$$

$$=\frac{m}{2\{\max(x,m)\}^2}\tag{57}$$

(d)

$$p(x = 100|D, \alpha) = \frac{1}{200} \tag{58}$$

$$p(x = 50|D, \alpha) = \frac{1}{200} \tag{59}$$

$$p(x=150|D,\alpha) = \frac{100}{2 \cdot 150^2} = \frac{1}{450}$$
 (60)