Tool: Linear regression

The Analytics Edge:

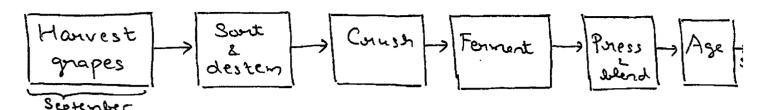
The price of mature wines may be predicted from data available when grapes are picked.

Using a simple leave on negression model with weather variables it is possible to predict which is traditionally done by wine experts and the predictions can often we improved.

Overview: Bondeaux is a sugion in Forance that is well known for making wines. The major sneason for the success is the excellent environment for growing vines in Bondeaux. Roughly 90% of the wines produced in Bondeaux are ned wines. Often these wines are necognized as some of the firest in the would.

Much of the wine in the oregion has been perioduced in the same way for hundreds of years yet there is significant differences in quality and price from year to year.

Onley Ashenfelter, a profession at Princeton developed a simple yet powerful analytics approach to Irelp predict the quality a price of Bondeaux wires.



Bondeaux wines taste when when they are older and hence there is incertive to store them till they come of age.

The younger wires are typically more unpleasant to

Key question: Can one predict how good a wine will be usen it matures?

This is useful since en primeur on wine futures gives people on oppositivity to buy wries couly a thus invest in it before it is bottled.

This is often based on sample of the wine much before it ages. Whe experts give scores (whe retry based on the tasting.

Example: The 1982 virtage of Chateau Latour was sold at 250 points a case enprimeur in 1988 and was valued at 9000 points in 2007.

Predictors Viritage (Year wire is made) Auchty of wine

Ashenfelter focused on the virtage as predictor for the quality of the wine, by averaging auction price across choten From the data, one can observe that:

- 1) Older the wine, greater is the value
- 2) However there is still significant variation in average prices that is weeplained.

To explain the quality of the wine better (as approximated by the price of the wine), he proposed the weather as a good predictor of quality. In Bondeaux, the weather significantly chapes from year to year for him to believe its to be a good predictor.

To study the analytics approach to this problem, we make use of the dataset from the websike www. I iquid asset. com. The data provided is:

TIME-SV marvest rain (Ang-Sep) violege

Time since virtage

WRAIN

Winter stain in months preceding

Virtage (OCt-March)

DEGREES

A verage de grees in °C over growing season (Apr-82p)

Analytics on Bosideaux wine-data: R

Read and bosic analysis of data:

wine & read. csv ("wine. csv")

Stor (wine)

While data frame consists of 38 Observations of 6 variables

Summary (wire)

VINT, LPRICE, WRAIN, DEGREES, vintage year, Jog of Price written summer temperature

HRAIN, TIME-SV Harvest rain Time since Vintage

18. na (wine)

1954 and 1956 wine prices not evallable in date set since they are oranely sold now Prices from 1981 to 1989 not in date

Plot (wire \$VINT, wine & LPRICE)

Scatter plat of vintage and price variable inductes negative relationsh but there is considerable variation Still that heeds to be captured Matrix of Scatterplats denown

pains (une) Split detaset

wine train & subset (wine, wine \$ VINT & 1978 A! is no (wine \$ LPRICE)

wine test < Subset (wine, wine \$VINT > 1978)

(excelleding 1954, 1956) and test set from 1939 to 1989.

Note that we do not have prices for all the points in the test set.

```
1 variable regression
```

model 1 & lm (LPRICE NVINT, data = winetrain)

(Alternative: model 1 < lm (wire train & LPRICE ~ wiretrain \$ VINT)
Summary (model 1)

In () is the basic model to fit a linear model to the data. Here LPRICE is fit with the VINT predictor The negression equation fit is:

LPRICE = 72,99 - 0.0378 VINT estimates intercept

Both estimated coefficients are significant at 0.01 level $R^2 = 0.2005$ Adjusted $R^2 = 0.1657$.

Plot (wine torain \$ VINT, whetrain \$LPRICE)

Obline (model 1). Plats the west fit line with a slope of -0.0378

model 1 \$ residuals

Sser & Sum (model 1 \$ residuals2)

Sun of squares of emons

SSEI & Sun ((winetrain & LPRICE - Mean (winetrain & LPRICE))^2)
Total sun of squares

1 - Ssel Ssti This gives R2 of 0.20048

(Also known as model R2)

abline (a = mean (wine train & LIRICE), & = 0)

Older the wine greater the value. However, there is still significent variation.

1 variable regression

IN(LPRICE N WRAIN, data=wine train)

R² = 0.018

In (LPRICE NDEGREES, data=wine train)

R² = 0.435

Un (LPRICE NHRAIN, data = wine train)

R² = 0.31

2 vanables

LOOK at the effects of DEGREES & HRAIN on price

Plot (whetrain & DEGREES, whetrain & HRAIN)

abline (h=mean (whetrain & HRAIN))

abline (v=mean (whetrain & DEGREES))

To add the dependent variable information,

plot (whietrain & DEGREES, whetrain GHRAIN,

col = if else (whietrain & LPRICE, mean (whitrain GLPRICE)

"red", "black"))

Figure indicates that hot and day summers produce wires that obtain Jugher perices while cooler Summers with more rain give lower priced wines.

Note that 1961 is the year when an extremely high quality wire was produced.

model 2 & In (LPRICE, DEGREES+ HRAIN, date = wine train) Sunnary (model 2)

LPRICE = -10.69 + 0.602 DEGREES - 0.0045 HRAIN.

A.ll there vanishes are extremely significant un this fit with $R^2 = 0.70$ and adjusted $R^2 = 0.68$. Note that one of the varieties here refers to the interest Complete regression

model 3 & Im (LPRICEN., data = wine train)
Summary (model 3)

Note that TIME_SV coefficients are not defined due to singularities (it perfectly correlates with VINT variable)

Drop the variable and redo degression

model 4 @ Im (LPRICE ~ VINT+ DEGREES + MRAIN+ WRAIN, data = wire rrain)

R2 = 0.8286 and adjusted R2 = 0.7943

Coefficients indicate that high quality wines convelte strongly in a positive manner with summer temperatures, negatively correlate with showeest main, convelate positively with virtage and writer main

Con (wire train)

High here refers of to magnitude of correlation Provides consolution
matrix of the variables

- High consolution Detween
independent variables is not
Good (multicollinearity)

(Typically > 0.7, <-0.7)

- High consolution Detween
dependent and independent
variables is good

Result indicates that by adding in weather variables, the R2 increases to 0.8 (80% of the variation can be explained) in comparison to R2 of 0.2 (20% of the variation can be explained) until the variation can be explained) until

Note that dropping the VINT variable decreases

R² to 0.75 and adjusted R² to 0.74

It seems reasonable to keep it.

Confirt (model 4)

at 2.5% and 92.5% lead

Confrit (modely, lord = 0.99) Provides confidence intend at 0.5 and 99.5%. Devel

Pure de chors

Sty (winetest)

wineprediction < predict (model 4, wine test)

This function predicts the outcome of the model 4 (from the linear regression) on the

values of the variables in the data frame

wine test Actual values Preduction -1.808 1980 -1.99

-1.724 -1.539

SSE4 (Sum (winetest & LPRICE[1:2) - wineprediction [1:2])?

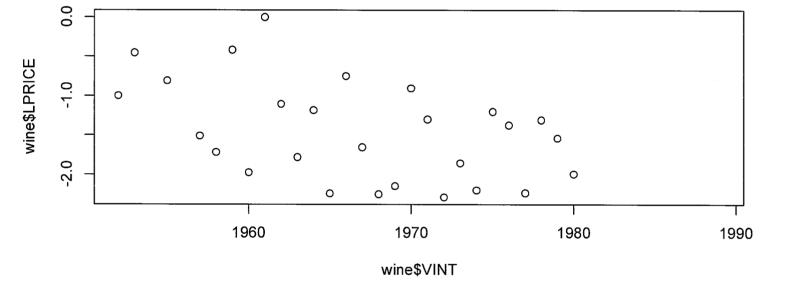
SSt 4 & Sum ((wine test & LPRICE[1:2] - mean (wine train)) 2 Note we should use training test mean to compute SST for cheeking effect of This gives on out of sample regresso 1-8504/8564 (or test R2 of 0.794)

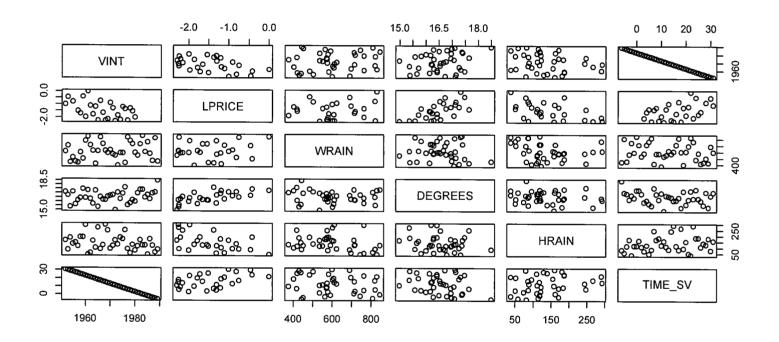
Test R2 Model R2 Variables 0.788 0.435 DEGREES

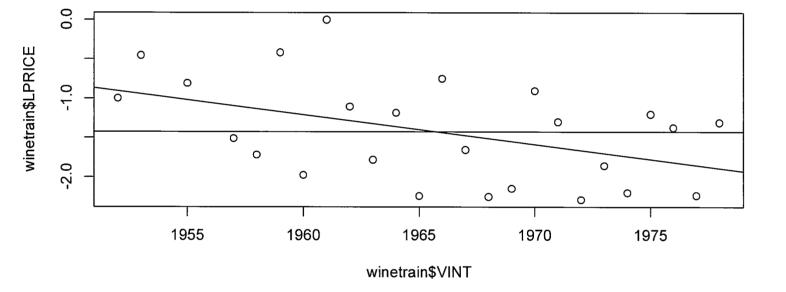
- 0.08 0.70 DEGREES, HRAIN

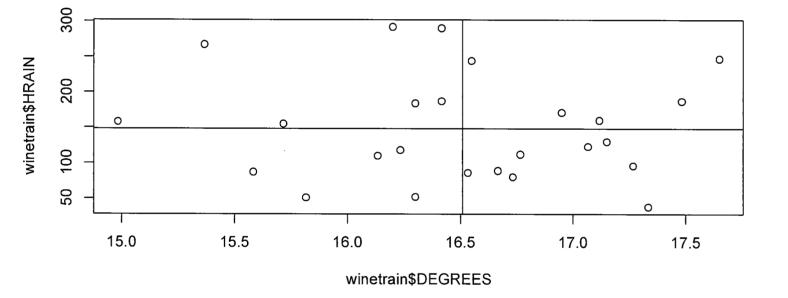
0.794 DEGREES, KRAIN, 0.82 WRAIN, VINT

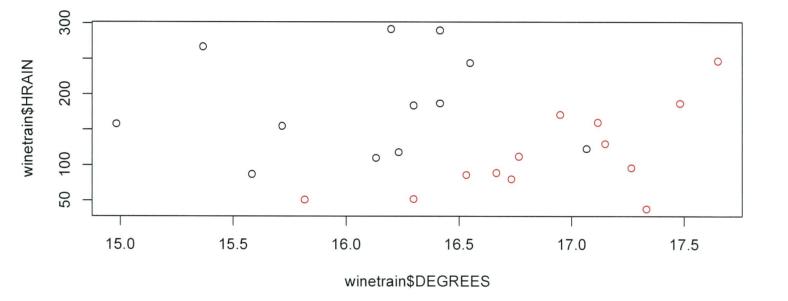
Better model RL does not men better test R2 which can also be negative.











The nesults from preduction indicate that 1989 wine would be of very ough quality.

I with the predictions of the Dest wine critics?

Ashenfelter prediction

Robert Parker prediction

1986 - mediocre dine to
below average growing
Season temperature a
above average honvest
vainfall

1986 - Very good d Sometimes exceptional

1989 - excellent vintage 1990 - even abetter vintage New York Times later
Said it was fentstic wine
At first Robert Parker
Said It would be
OR like 1985
what then said later
Its the vintage of
the century

2000-2001 Both prediction from the model L expert agreed that it was very shigh quality wine Data

Source: http://www.liquidasset.com Price of wire (from auchors), weather Information for the virtage

Years 1952 - 1989 Fairly small data set

Model

Linear oregression model to predict wine quality (represented by price) in terms of vintage, Summer temperature, wither main, honvest main

Decision

Develop a prediction on the quality of wine that is known only when it matures typically using weather variable information available at the time of making the wine.

Value

Predictions are comparable to and con also sometimes Deat the prediction of qualitative expents using a simple model.

Sunnary of Junear regression output (R)

Res, Juals

(Provider a Summary of the residuals from the linear regression model
To access the residuals, use residuals (Ilabom) claubican additionals (Ilabom) charlished

Coefficients

Provides estimated coefficients,

Standard error of coefficients,

t value and Significance

value. To access coefficients,

Use coefficients (model 1) or

model 1 & coefficients, You can

access Std. Error & & y

Coefficients (summay (model 1)) [, "Std. En:

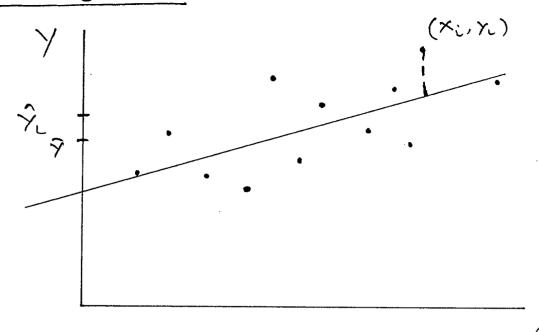
Residual Standard envion Priorides the averege arount
that was ponge will deviate
from true regression line
Priorides a neasure of lack
of fit of linear model to date

Multiple R2 and adjusted R2

R2 is a measure between Or 1 to indicate the amount of variability explained by regression Adjusted R2 accounts for number of predictor

Fotostutie L p-value

Test to see if atleast one of the predictors is non zero



P= 1 n= 12

Multiple linéar regression model

Data: Observations (Yi, Xi,, ..., Xip) for i=1,...,r

of absentations of absentations

P = Number of predictor variables

y = Dependent variable (outcome)

X, , , x p = Independent variables (productions)

Coefficients Bo, B,,.., Bp are chosen to minimise the sum of squared residuals:

min
$$\sum_{i=1}^{n} \left(y_i - \beta_0 - \beta_i \times_{i_1} - \dots - \beta_p \times_{ip} \right)^2$$

 $\beta_0, \dots, \beta_p = i=1$

Main points: Linear stagression

Solving the optimization problem gives using matrix notation

where
$$y = \begin{pmatrix} y_1 \\ y_n \end{pmatrix}$$
, $X = \begin{pmatrix} 1 \times 11 & \dots \times 1p \\ \vdots & \vdots & \vdots \\ 1 \times n_1 & \dots \times np \end{pmatrix}$, $P = \begin{pmatrix} P_0 \\ P_1 \\ \vdots \\ P_p \end{pmatrix}$

Optimel solution
$$\hat{\beta} = (x'x)^{-1}x'y$$
 of R computes coefficients

Futted values $\hat{y} = \times \hat{\beta}$

(2) The estimates have standard errors associated with them. This is board on the intuition that we develop a regression estimate using observed data to estimate the true population regression.

Assuming observations y are uncorrelated, have constant varionce of and X are non-random

$$Var(\hat{\beta}) = (x \times x)^{-1} \times x \times x^{-1}$$
 $Var(\mathbf{Y})$

$$= (x \times x)^{-1} \sigma^{2}$$
To estimate σ^{2} , use $\hat{\sigma}^{2} = \sum_{i=1}^{2} (x_{i} - \hat{y_{i}})^{2}$

$$= \sum_{i=1}^{2} (x_{i} - \hat{y_{i}})^{2}$$

The division by n-p-1 is to make the estimator unbiased with $E(\hat{\sigma}^2) = \sigma^2$

Standard error of the coefficients is equal to the square moof of the diagonal element of the matrix $(x'x)^{-1}r^2$.

Under null hypothesis that $\beta_i = 0$, t value = $\frac{\beta_i}{Stateror(\beta_i)}$ (Known on t-Statutic)

If t value is high (in absolute value), The null hypothesis will be neglected to claim that β_c is significant predictor in the model. This is evaluated as the P-value (P(>|t|)).

3 Quality of fit:

a) Residual Standard error, sesidual sun of squares sun of square errors and Jotel sun of squares SSE = $\sum_{i=1}^{\infty} (\gamma_i - \bar{\gamma}_i)^2$ where $\bar{\gamma}_i = \bar{\beta}_0 + \hat{\beta}_i \times i^{\dagger}_i + i \bar{\beta}_i \times i^{\dagger}_i$

Residual Standard = $\sqrt{\frac{SSE}{N-P-1}}$ | Measure of lack of the europe of the property of the standard of the

Note that it is possible for models with more variables to have higher tresidual standard covor if the decrese in SSE is Small relative to increse in p.

1) R2 and adjusted R2

 $R^{2} \left(\text{coefficient of determination} \right) = \frac{SSR(Sun.t squareo from regression)}{SST(Total Sun.t squareo)}$ $\left(\begin{array}{c} SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SST = SSR+SSE \end{array} \right) = \frac{1 - SSE_{eq} Sun.t squareo}{SST} \\ SSE = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SSR = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SST = \frac{2}{L_{eq}^{2}} \left(\begin{array}{c} SSE_{eq} \end{array} \right) \\ SS$

R² measures the proportion of the total variation in y (dependent variable) that is accounted for by variation in negressons (dependent var). This provides information on the goodness of the fit of the model.

Regnession is a homeontal =) (x has no explanatory power)

Regnession line penfectly

Fits all points in a straight (x has penfect)

line

R2 = 1

(x has penfect explantory power)

All values of Yi lie in =) R2 cannot be carreted.

As we increase the number of variables in the model, R^2 will increase.

It will never decrease in this case and hence one needs to be careful of overfitting the date The adjusted R2 Statistic peralises the R2 Statistic as more variables are added in the fit.

The adjusted R² can be negative and its value will always be less than on equal to R².

The adjisted R2 increases when a new explanatory variable is added only if this variable increases the R2 more than would be expected by Chance. The adjisted R2 is prunarily useful in the feature selection steps of model building.

Adjusted
$$R^2 = 1 - \left(1 - R^2\right) \left(\frac{m-1}{m-p-1}\right)$$

c) F-statistic

Ho: B= B2= ... = Bp = 0

H.: Atlest one Bili non zero

$$F-Stestistic = \frac{(SST-SSE)/p}{(SSE)/(m-p-1)}$$

When strone is no relationship doctured response 2 predictors, F^2 is expected to be close to 1.

If H, is true, we expect F to the greater than 1.

rangs
Parker
Robert

The contract of the contract o					L			Statement of the later of the l			State Control of the	-	State State St	214	100				
	Austria Plesting & Gruner Veltliner	FX	168	8 118 871 8		881 85C	95R 82C 98	200	87C				25 80 80	Z			Z	Z	TN T
	Alsace	TN TN	79I 90T	791 87R	82R	91R 90R	90R	87R	90R		THE STATE OF	86R	82C	75C	82C	-	84C	90R	2C 90C
	Bordeaux: St Julien/Pavillac/St Estephe	87E		87E 95T 8	NAMES		87T	196 T	85C	790	98E 90E		94T 90		800		858	84R	91 821
	Bordeaux: Margaux	88E	97E	186		89E 94T	86T	B8	85C	-			8 90T 86	100000	868		87R	77R	8E 83I
	Bordeaux: Graves	89E	91E	87E	88		94T	100		-		89R	89E	79R	888		888	710	9T 86R
	Bordeaux: Pomerol	888	98E 96E	106		90E 95T	196	85E	89T	100000	96R 96R	1		R 65C 9	OR 96R		868		4R 87
The contract	Bordeaux: St. Emilion	_	93E 92E	88E 99T	Resident.	90E 96T	88R 96T 86	87T	86T	_	98R 88R	No.	38E	269	948				5R 83
Part	Bordeaux: Barsac/Sauternes	91E	97E	88R 96R	95R	95R	4,550	-			96R 90R	96R 70R	94R	70C	75R		75R	878	DT 86
The contribution of the	Burgundy: Cate de Nuits red	93E 91E	95E 88I	89I 98T				T68	72C		93R 85C	846	9 2 2 9	78R	75C		277	298	90
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The control of the co	Russingly white	92E 91E	90g			386		926	77.0		100	828	820	208	586		e e	SEC	
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The continuation of the co	Champagne		z z				¥ 200	- 1	7.50			STREET.	F Z	Z	94R		로 10 6	94R	38 930
The continuation of the co	Languedoc and Roussillon		91R 87R 92F	90i 88R 8			r Z	r Z					Z	Z			F	F	TN TI
The control co	Loire Valley white	Z	881 901		2C 82R 96R	District of	84C	91R	87C		Security Sec		870	289	84C		830	i Z	TN TI
The control of the co	Rhone: Cote Rotie/Hermitage	NT 88E	187 T8I	92E 89T	968	89R 87E	STATE OF THE PARTY		88C	78R 92R	92R 92T		84C	75E	856				36 84
The continuation of the co	Rhone: Chateauneuf du Pape	NT 888	868	92R 95T	106	96T 98E	90E 98E 82		36C 85C	78C 65C	95R 94T	88R 60C	NO SHOW	72C	7C 70C	88C 77C	88C 97		96 820
The continuation of the co	Germany: Mosel-Saar-Ruwer	Z	951 821	156	Maria	91R 76C	92T	91T	94R	87C 88R	96R 91C	92R 84C	85C 88	70R		86R 681		92C	8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8
The state of the s	Germany: Rhine Rieshing	NT 94R	931 88R	198		95R 69C	ST PROPERTY.	R 91T 86C	87R 88R	85C 87R	96R 90C	90R 84R	84C 86	C 70R 8	7C 76C	87R 701	86R 70		2R 98C
The state of the s	Campania: Taurasi	L.	92T	92T 92R	_	91R 92R	93R 93R N		TN TN		TN TN			Z	TN T		L N	TN	TN T
The control of the co	Friulis Collin Whites	100	94R 911	930 900	856	88C 90C	90C 85C N1	Z					Z	Z	Z			Z	TN
The control of the co	Diadovent Rashavacca	100	901 90E	90E 93T	69	96 T 90E	90E 92T 93	B 97T 87C	900	100	98R 97T	90	168	980	9		89B 07		Med
The control of the co	Treminal C. Darbaresto	1-2	100 IOS	T o		1 04E	OFT GOT			-		9	ş	1					action in-
The state of the s	Pledmont; Balolo		978	100		į		1			100					uno uno			1
The control of the co	frentho-Alto Adige Whites							I N					2	ž	2	ž	ž	ž	ž
The control of the co	Tuscany: Bolgheri		ave azi	HE S	2	Sharp I	94K 88K 94	Nes Nes	Z	Z	N N N N N N N N N N N N N N N N N N N	NO N	N N	Z	Z	- Z	Z	E Z	Z
The continuity of the continui	Tuscany: Brunello di Montalcino	r Z	93T 91I	97T 91E	<u>.</u>	95T	95T 88R 95	R 88R 93R	NT 90R	TN TN	94R NT	94R NV	NN NN	N TN	2R 92R	LN LN	E Z	FN	E
	Tuscany: Chianti Classico	90R	92E 90R	96T 91R	921	93R	94R 88R 95	R 88R 89R	H 89R	TN TN	90R NT	92R NT	NV 93	R 60C	IT SOR	FN FN	Z FZ	LN LN	A NT
H. S.	Sielly: Etna	92T 91T	88R 90I	92R 93I	880			TN TN T		TN TN	TN TN	TN FN	Z FZ	FZ	LN L	LN LN	Z Z	E Z	TN
	Veneto: Valpolicella (Amarone)	94R NT	94E 90R	93R 90R		STATE OF THE PARTY.	89C 89C NI		TN TN		TN	TN TN		FN		TN TN	Z F	FZ.	TN
Handing the part of the part o	Portugal dry wines	Z	90T	84E 90T	89R		FN.	r N						L				FZ	F
A continue and the cont	Jortugal vintage port	TN TN	94T 90T	TN TN	90T	92T	N	N	PERM	ALC: NO.			N	N	86R	2.000	9-30	N	N N
Statistic continues and statis	lioja	92T 88T	106	85R 92R	871	94E 86E	82C	858	90R			Married World		AL STATE		92C 75C		8 SR 8	140
Leat Channet Shoupers	Abera del Duero	92T 89T	94T 91T	86R 93T		95E 87C	88T	R 92R 90R	90R 87R	82R 74R			77C	BGR	87C	and the last	FN	FN.	FN F
Teat-Channels and the contribution of the cont	eodes	9 E	91R 90R 89R	TN		TN	LN.	F.N						8,8			F	12	
Total Control											100								1
Total Control	North Coast Cabernet Sauvignon	187 396	941 941	15 III		96T 78C		106	95E	941			90R	92R	86R		80R 92	8 90T 8	3R 70C
Constitution of the consti	North Epast Chardonnay	94E 88I	90 R 88I	871		90C 87C	89R	87C	88G	-			90C	980	856	86C 88C	83C 86	8 208 E)C 82C
Althorithministry Mily and the control of the contr	North Coast Zinfandel	92R 74I	851 781	791 781		830	87R 86C 85	E 89C 87R	92C	90C 91C	Service .	Estan	87C	88 R	800	82C 82C	836 86		200
Ministration between the state of the state	North Coast Pinot Hon	93R 84II	90R 86I	871 90E		-	90E 89R 90	E 88R 88R	92R 88R	88R 86R	86R 85C	87C 86C	840 86		5C 84C	836 850	80C 84	FN	FN P
Hands believe the believe that the belie	Central Eoast	93R 87E	87R 88I	92E 92R	92E	92R	90 R	Ľ.						E.	P.Z	FN FN	r F	E Z	FN
14. 1 1. 1 1. 1 1. 1 1. 1 1. 1 1. 1 1.	Oregon: Willamette Valley Pinot Nour	TN TN	86R 94T	85T	88	86E	168	83C	92R	87C	30C 86C		85C	65C	84C	386 866	LN LN	LN.	LN L
143. 143. 143. 143. 143. 143. 143. 143.	Washington Cabernet Sauvignor/Sylah	94T		2,000		89R	106	188	90R	850	1200		78R	72C	780	-	- Z	E	Z Z
Anticological state of the control o	Donantino	# F		876		F.N	FN	FN	F	μN			FN	H	12		FN	12	1 1
1. Data state and all all all all all all all all all al	enough cons	Sales Sylicity									8					1			- -
No. 24. 25. 25. 25. 25. 25. 25. 25. 25. 25. 25	South: Barossa/McLaten Yale	94T 79I	89T 85I	94R SGR		386	95E	30E	90R	87R 89R	88R 88C	85C 87C	90R 86	E NT	83C	85C 88C	F Z	ŁN .	FN
14. Table 1. Seed 1. S	West	91T 90T	T68	82I 91R		88E	90R	HN					FN	FN	E Z	TN TN	FN	L N	TN T
Acide Anison Mark							100												
N S47 S48	vevr South Wales	170 10)		84K 8/K	Z.			z						84 4	Z	Z	Z	Z	Z
Fig. 1. Set 1. S	fictoria/Tasmania	94T 78I	858	93R 91R	Z		F.	Z								_	LN FN	E N	TN
Additional No. 1791 888 918 898 801 918 908 826 826 87 87 87 87 87 87 87 87 87 87 87 87 87	4114	87T 92T		90R	Z	Z	Z	Z						N	F		F.N	L N	I.N
NT 91E 891 91R 93R 38R 38R 38R 38R 38R 38R 38R 38R 38R 3		791	89R 841 94B	i i	100	174	aus	808		6		100	2	1	702	-	FIN		1
NI N	iew Cealand										-		760	2 1	2			ž	
	south Africa	85 B B B B B B B B B B B B B B B B B B B	93R 86R 90R	92R 85C N		Z	Z	Z						Z	HZ			LN	LN

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Rating Ranges:

| So | 100 | Extraordinary, | 90 | 86 | Outstanding, | 80 | 89 | Rove Average to Excellent | 170 | 173 | Average | 60 | 160 | Below Average, | 6 | Appailing | Manufacture, may be too old, | E | Early maturing and accessible, | NV = Vimage not declared, | I = Inregalar, even among the best wines, | NT = Not yet sufficiently fasted to rate, | R = Ready to drink, | T = Still tamp, youthful, or slow to mature.

80 - 89 Above Average to Excellent.