SOC542 Statistical Methods in Sociology II Introduction and Review

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Plan

- Introductions
- Course outline
- Review
 - Notation
 - Statistics
- Lab: R, RStudio, and statistics refresher

Learning goals

- ► The use of statistical analysis in empirical social science research
- Multiple regression
 - Conceptual and statistical background
 - Implementation in R
 - Interpretation
 - Violations assumptions and robustness
 - Frequentist and Bayesian approaches to estimation
- Proficiency in data handling, analysis, and visualization using R

Structure

- 1. OLS regression (Weeks 1-4)
- 2. Non-linear variables and interactions (5-6)
- 3. Model checking and missing data (7)
- 4. Generalized linear models (8-11)
- 5. Clustered data (12)
- **6.** Causal inference with observational data (13)
- 7. Presentations (14)

- ► Theory-driven statistical modeling
 - Causal, scientific models to guide implementation of statistical analyses

- Multiple approaches to estimation and inference
 - Quantification of uncertainty via Bayesian inference
 - Reduced emphasis on p-values

- Data-driven pedagogy
 - Preference for code and simulations over mathematical formalism

- Robust inferences
 - Appropriately specified statistical models
 - Model checking and comparison

- ► Reproducibility and transparency
 - ▶ Emphasis on reproducible and transparent analytic procedures

- Scientific communication
 - Clear, precise summaries of statistical models using tables and visualizations

- ► Theory-driven statistical modeling
- Multiple approaches to estimation and inference
- Data-driven pedagogy
- Robust inferences
- Reproducibility and transparency
- Scientific communication

Assessment

- ► Homework assignments (40%)
 - ► Four assignments, each worth 10%
- ► Final paper (50%)
 - ▶ Phase 1: Research question, data, and methodology
 - Phase 2: Preliminary results
 - Phase 3: Submit final paper
- Presentations (10%)
 - Present results from final paper

Required readings

- Regression and Other Stories by Andrew Gelman, Jennifer Hill, and Aki Vehtari
 - Main textbook, covers applied OLS regression and generalized models in R
- ▶ Bayes Rules! An Introduction to Applied Bayesian Modeling by Johnson, Alicia A., Miles Q. Ott, Mine Dogucu. 2021.
 - More detailed introduction Bayesian inference with examples in R

Recommended readings

- Causal Inference: The Mixtape by Scott Cunningham
 - More formal econometric background and overview of causal approaches
- Statistical Rethinking, 2nd ed., by Richard McElreath
 - Supplementary textbook, provides additional material and deepens understanding of Bayesian inference and modern statistics
 - Recommended: McElreath's YouTube lecture series (2023 series currently underway)
- R for Data Science by Hadley Wickham and Garrett Grolemund.
 - ▶ A useful reference for data manipulation in R via the tidyverse
- Data Visualization: A Practical Introduction by Kieran Healy.
 - ▶ Great introduction to data visualization using R and ggplot

Vectors

- ► A vector is a sequence of numbers
 - e.g. We take the heights of everyone in the class and record them in vector v

$$v = \{6.1, 5.9, 5.7, 6.0, 6.2, ..., 5.9\}$$

- We typically arrange these vertically as columns in a data table.
- ▶ We can use *indexing* to reference specific elements of the vector
 - \triangleright v_i refers to arbitrary element i
 - \triangleright v_1 indexes the first element, 6.1, and so on

Summation

The uppercase letter Σ is used to denote a summation. We can use it here to take the sum of all the values in vector v, where k is the length of the vector.

$$\sum v_i = \sum_{i=1}^k v_i = v_1 + v_2 + ... + v_k$$

We can compute sums in R using the sum() function, where the thing we are summing over is included in the parentheses, e.g. sum(v).

Products

The uppercase letter Π is used to denote the product operation. We will encounter it far less frequently than summation.

$$\prod v_i = \prod_{i=1}^k v_i = v_1 * v_2 * ... * v_k$$

Matrices

- We often want to represent multiple vectors as a matrix.
- ► Let's say we also collected each students' age, we could represent the ages as a vector *u*.
 - v and u can be combined together in a matrix M (typically we use lowercase for vectors and constants and uppercase for matrices):

$$M = \begin{cases} 6.1 & 24 \\ 5.9 & 22 \\ 5.7 & 27 \\ 6.0 & 30 \\ 6.2 & 25 \\ \dots & \dots \\ 5.9 & 26 \end{cases}$$

Matrices

- We can index elements of a matrix in the following way
 - $ightharpoonup M_{i,j}$ refers to the i^{th} row of column j
 - e.g. $M_{4,2}$ indexes the age of the 4^{th} student

Vectors and matrices in R

```
v <- c(1,2,3)
print(sum(v))
## [1] 6
print(prod(v))
## [1] 6</pre>
```

Vectors and matrices in R

We can use cbind to combine vectors columnwise into a matrix.

```
u \leftarrow c(1,1,1)
M <- cbind(v,u)
print(M[3,1]) # M[row, column]
## v
## 3
print(M)
##
        v u
## [1,] 1 1
## [2,] 2 1
## [3,] 3 1
```

Vectors and matrices in R

We can *transpose* this matrix using t if we want to treat these columns as rows.

Random variables

- The value of a random variable depends on the outcome of random events
 - e.g. x is a random variable representing the outcome of a series of coin tosses

Probability distributions

- Random variables are drawn from probability distributions
 - The probability of tossing a coin and getting a head is a defined by the Bernoulli distribution
 - The number of heads in a sequence of coin tosses is defined by the binomial distribution
 - The height of a randomly chosen adult male is drawn from a normal distribution

Probability distributions

- Distributions are defined by parameters that modify their shape and scale.
 - ► A Bernoulli distribution has a single parameter, p
 - A binomial distribution has two parameters, *n* and *p*
 - A normal distribution has two parameters, a mean μ and a standard deviation of σ
 - Often we refer to this using the shorthand $N(\mu, \sigma)$

Probability distributions in R

A random variable drawn from a binomial distribution can take a value of 0 or 1, where p is the probability of a 1. p=0.5 is equivalent to a fair coin toss. The Bernoulli distribution is a special case where p=1.

```
x <- rbinom(n=1, 1, 0.5)
print(x)
## [1] 0</pre>
```

Probability distributions in R

In this case, I make 10 draws from a binomial distribution, simulating ten tosses of a fair coin.

```
x <- rbinom(10, 1, 0.5)
print(x)
```

[1] 0 1 0 0 1 1 0 0 1 0

Probability distributions in R

A Normal, or Gaussian, distribution is a continuous distribution with two parameters. The standard normal distribution has a mean $\mu=0$ and standard deviation $\sigma=1$. Here is a single random draw:

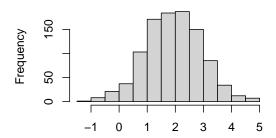
```
x <- rnorm(1, mean = 0, sd = 1)
print(x)
```

[1] 1.477438

Probability distributions in R

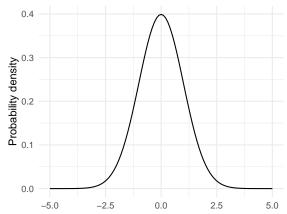
In this case, we can make 1000 draws from a normal distribution with a mean of 2 and a standard deviation of 1. Since there are a lot of values it is best to plot them using a histogram.

Histogram of x



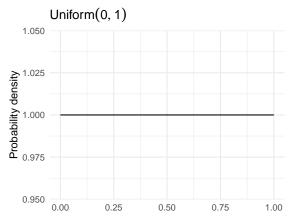
Probability distributions in R

A *probability density function* shows the probability of observing particular random draws from a given distribution.



Probability distributions in R

In a uniform distribution, every value in the interval [a, b] is equally likely.



Expected values

- ▶ The expectation of a random variable is denoted by $\mathbb{E}[x]$. It is the long-run average of the random variable over many repeated trials.
- ▶ The expected value of a constant x = c is c. E.g. $\mathbb{E}[2] = 2$
- ▶ The expected value of a random variable is its mean.
 - e.g. If x represents a vector of values drawn from a normal distribution, our best guess as to the value of any one realization x_i is μ .
- We can use expectations notation to represent relationships between variables, e.g. $\mathbb{E}[y|x]$

Expectations as weighted averages

Discrete case, where p_i is the probability of observing a particular value of x

$$\mathbb{E}[x] = \sum_{i=1}^k x_i p_i$$

▶ In the continuous case, where f(x) is a probability density function, the expectation is an integral over the possible values of x.

$$\mathbb{E}[x] = \int_{-\infty}^{\infty} x f(x) dx$$

Populations and samples

- Classical statistics is based upon the assumption that our observations x are drawn from an underlying population.
- ► In a simple random sample, we draw *n* instances of a random variable from the population
 - ► These draws are assumed to be *independent* and *identically* distributed (IID)
- ► For example, a hypothetical population might be adults residing in the United States and a sample would be a randomly sampled subset of these adults.

Means

► The population mean value of a random variable is defined in the following manner

$$\mu_{\mathsf{x}} = \frac{1}{\mathsf{N}} \Sigma_{i=1}^{\mathsf{N}} \mathsf{x}_i = \frac{\Sigma_{i=1}^{\mathsf{N}} \mathsf{x}_i}{\mathsf{N}}$$

▶ The sample mean \bar{x} is defined by the following equation. n is lowercase to denote $n \subseteq N$

$$\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i = \frac{\sum_{i=1}^{n} x_i}{n}$$

Variance

Variance is the average of the squared deviations from the mean. For the population it is defined as

$$\sigma_x^2 = \frac{1}{N} \Sigma (x_i - \mu)^2$$

▶ The sample variance includes a degrees of freedom correction.

$$\sigma_{\bar{x}}^2 = \frac{1}{n-1} \Sigma (x_i - \bar{x})^2$$

Standard deviation

Variance is difficult to interpret. The standard deviation is a scaled measure of variance, typically denoted using σ . In the population, it is equal to

$$\sigma_{x} = \sqrt{\sigma_{x}^{2}} = \sqrt{\frac{1}{N}\Sigma(x_{i} - \mu)^{2}}$$

► The sample standard deviation is thus

$$\sigma_{\bar{x}} = \sqrt{\sigma_{\bar{x}}^2} = \sqrt{\frac{1}{n-1}\Sigma(x_i - \bar{x})^2}$$

The sampling distribution of the mean

- ► The mean of an IID random sample is distributed according to a sampling distribution
- ▶ It has the following properties:

$$\mathbb{E}[\bar{x}] = \mu_{x}$$
 $var(\bar{x}) = \sigma_{\bar{x}}^{2} = \frac{\sigma_{x}^{2}}{N}$

Simulating the sampling distribution

We can simulate the sample mean of 1000 normal distributions with the same population parameters and compare the sample means to the expectations.

```
mu <- 100; sigma <- 10; n <- 100 # population parameters
sims <- replicate(1000, mean(rnorm(n, mu, sigma)))
print(round(mean(sims),2)) # E[mu] = 100
## [1] 99.99
print(round(var(sims),2)) # E[sigma^2/n] = (10^2)/100 = 1
## [1] 0.9</pre>
```

The Law of Large Numbers

- ▶ When sample size is large, \bar{x} is close to μ_x with a high probability
- ► A large IID sample can therefore approximate the sampling distribution
 - ▶ Such samples are *asymptotic* because approximations become exact in the limit as $n \to \infty$
- ▶ Under such conditions, \bar{x} is considered a *consistent* estimator for μ_{x}

The Law of Large Numbers

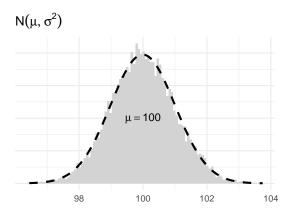
The sample mean from a large IID sample closely approximates the population mean.

```
large.sample <- rnorm(1e6, mu, sigma)</pre>
print(mean(large.sample))
```

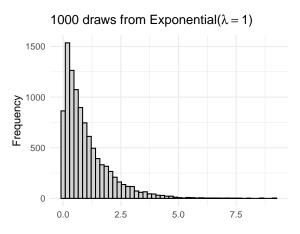
[1] 99.99705

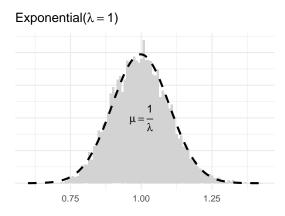
The Central Limit Theorem

- ▶ CLT: The distribution of \bar{x} is well approximated by a normal distribution when n is large
 - ▶ This is approximately true even if *x* are not normally distributed



Distribution of 10,000 sample means, where n = 100





Distribution of 10,000 sample means, where n = 100.

Standard error of the sample mean

► The standard error of the sample mean is defined as the sample standard deviation divided by the square root of *n*.

$$SE_{\bar{x}} = \frac{\sigma_{\bar{x}}}{\sqrt{n}}$$

- ▶ The standard error is used to communicate *uncertainty* since we cannot observe the true population mean μ but only the sample mean \bar{x} .
- ► Theoretically, the standard error is the standard deviation of the sampling distribution.

[1] 0.01

Standard error as standard deviation of the sampling distribution

We can show that a standard error of a large random sample is a good approximate of the standard deviation of the sampling distribution.

```
N <- 10000
s.dist <- replicate(1000, mean(rnorm(N)))
print(round(sd(s.dist),3))
## [1] 0.01
x <- rnorm(N)
print(round(sd(x)/sqrt(N),3))</pre>
```

Estimating the standard error of the mean in R

```
mu <- 10 # population mean
sigma2 <- 1 # population standard deviation
N <- 100
x <- rnorm(N, mu, sigma2)
print(mean(x)) # sample mean
## [1] 9.862993
print(sd(x)/sqrt(N)) # sample SE
## [1] 0.09972622</pre>
```

Confidence intervals

► The standard error is used to define a confidence interval. By convention, a 95% confidence interval has lower and upper bounds of

$$[\bar{x} - 1.96SE_{\bar{x}}, \bar{x} + 1.96SE_{\bar{x}}]$$

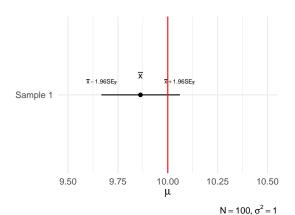
When n is large, the 95% confidence interval around \bar{x} contains the true value μ_x in 95% of all possible random samples.

Confidence intervals in R

We can use the results of the previous simulation to compute confidence intervals.

```
xbar <- mean(x)
SE <- sd(x)/sqrt(N)
lower <- xbar-1.96*SE
print(lower)
## [1] 9.667529
upper <- xbar+1.96*SE
print(upper)
## [1] 10.05846</pre>
```

Confidence intervals in R



One-sample t-tests

- We can use a one-sample t-test to assess whether there is a statistically significant difference between our sample mean and a null hypothesis.
- ▶ The *test statistic t* is obtained using the following equation:

$$t = \frac{\bar{x} - \mu}{\frac{\sigma_{\bar{x}}}{\sqrt{n}}}$$

p-values and statistical significance

- ▶ In classical statistics, we use a test statistic to calculate a p-value.
 - This represents the probability of drawing a test statistic at least as large as the observed test statistic, assuming the null hypothesis is correct
- ➤ Smaller p-values indicate that a result is less likely to be due to chance
 - By convention, p < 0.05 is considered the threshold for statistical significance.

One-sample t-tests in R

```
t.test(x, mu=10)
##
##
    One Sample t-test
##
## data: x
## t = -1.3738, df = 99, p-value = 0.1726
## alternative hypothesis: true mean is not equal to 10
## 95 percent confidence interval:
## 9.665114 10.060871
## sample estimates:
## mean of x
## 9.862993
```

One-sample t-tests in R

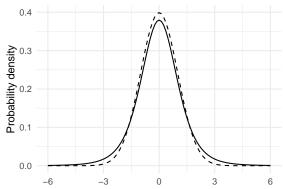
We can calculate the t-statistic and directly look up the corresponding p-value using a Student t distribution with N-1 degrees of freedom. 1

```
t <- (xbar-mu)/(sd(x)/sqrt(N))
print(t)
## [1] -1.373835

p <- pt(abs(round(t,5)), df = N-1, lower.tail = FALSE)
print(round(p*2,4)) # p*2 = two-tailed p-value
## [1] 0.1726</pre>
```

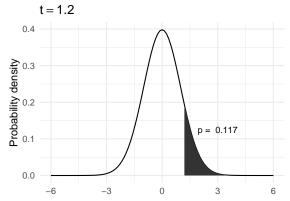
 $^{^{1}}$ Rounding ensures that the results are equivalent to the 't.test' function used above.

The t-test and the Student t distribution



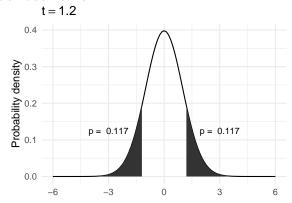
Student t distribution with 5 d.f. N(0,1) dashed line

Area under one tail



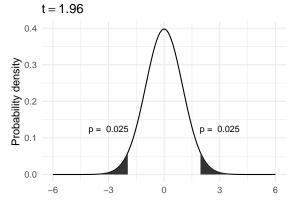
Student t distribution with 99 degrees of freedom.

Area under both tails



Student t distribution with 99 degrees of freedom.

Area under both tails



Student t distribution with 999 degrees of freedom.

Random sampling and p-values

Another way to show the same result is to draw random variables from a Student t distribution and calculate the proportion that fall above the chosen significance threshold.

[1] 0.05

We can expect to see a t-statistic where $|t| \ge 1.96$ approximately 5% of the time purely due to chance.

Testing for differences in means

- ▶ Often we want to know whether the mean values of two random variables, μ_X and μ_Y , are different from one another.
- ▶ We can test for this by computing calculating the difference between the sample means and the uncertainty about that difference
- ▶ The equation for the two-sample t-test is:

$$t = \frac{\mu_{x} - \mu_{y}}{SE(\mu_{x} - \mu_{y})}$$

where

$$SE(\mu_x - \mu_y) = \sqrt{\frac{\sigma_x^2}{n_x} + \frac{\sigma_y^2}{n_y}}$$

Testing for differences in means

```
n < -1000
x \leftarrow rnorm(n, mean = 12)
y \leftarrow rnorm(n, mean = 11)
xbar <- mean(x)
ybar <- mean(y)</pre>
varx \leftarrow (1/(n-1))*sum((x-xbar)^2)
vary \leftarrow (1/(n-1))*sum((y-ybar)^2)
SE <- sqrt((varx/n)+(vary/n))
t <- (xbar-ybar)/SE
print(round(t,3))
## [1] 20.831
```

Testing for differences in means

```
t.test(x, y)
##
##
   Welch Two Sample t-test
##
## data: x and y
## t = 20.831, df = 1994.7, p-value < 2.2e-16
## alternative hypothesis: true difference in means is not
## 95 percent confidence interval:
## 0.8753134 1.0572567
## sample estimates:
## mean of x mean of y
## 11.97403 11.00774
```

Type I and Type II errors

- ► A Type I (false positive) error occurs when we incorrectly reject the null hypothesis
 - e.g. In the one-sample test, we reject the hypothesis that $\bar{x}=10$ given $\mu=10$
- ► A Type II (false negative) error occurs when we incorrectly fail to reject the null hypothesis
 - e.g. If we did not have sufficient statistical power for the test above, we might fail to reject the null that $\bar{x} = \bar{y}$
- ▶ In classical statistics, if our chosen significance level is p < 0.05 then we expect to see such errors approximately 5% of the time

Sign (S) and magnitude (M) errors

- Gelman, Hill and Vehtari draw attention to two kinds of errors that are often overlooked but can lead to profoundly incorrect inferences:
 - ► Sign errors: the sign of a relationship is incorrect
 - e.g. We observe a positive relationship between x and y when the true relationship is negative
 - Magnitude errors: the observed effect is severely over- or under-estimated
 - e.g. We think $\bar{x} = 10$ but $\bar{x} = 0.1$

Expected mean and variance of two random variables

▶ The expected mean of the sum of two random variables is

$$E[x + y] = E[x] + E[y] = \mu_x + \mu_y$$

► The expected variance is the sum of the variances plus twice their covariance

$$var(x + y) = var(x) + var(y) + 2cov(x, y)$$

If x and y are independent then cov(x, y) = 0 and var(x + y) = var(x) + var(y)

Covariance

- ► Covariance is the a measure of the joint variability of two random variables
- ▶ The expectation of the covariance between *x* and *y* is

$$cov(x, y) = E[xy] - E[x]E[y]$$

For a population, the covariance is

$$cov(x,y) = \frac{1}{N} \Sigma(x_i - \mu_x)(y_i - \mu_y)$$

► Sample covariance is defined as

$$cov(x,y) = \frac{1}{n-1}\Sigma(x_i - \bar{x})(y_i - \bar{y})$$

Correlation

► Correlation is a scaled version of covariance. We divide the covariance by the product of the standard deviations.

$$\rho(x,y) = \frac{\frac{1}{n-1}\Sigma(x_i - \bar{x})(y_i - \bar{y})}{\sigma_x \sigma_y} = \frac{cov(x,y)}{\sigma_x \sigma_y}$$

▶ The Greek letter ρ is typically used to refer to correlation. The correlation coefficient ranges from -1 to 1.

Why R?

- ► Free and open-source
- Versatile
 - Statistical computing (lm, glm, 'rstanarm")
 - Data manipulation (tidyverse)
 - Visualization (ggplot)
 - General purpose programming
- Active developer community
 - Lots of packages with regular updates
 - New statistical and econometric packages

Why R?



Source: Kieran Healey

RStudio

Overview

- RStudio is an Integrated Development Environment for programming in R
 - ▶ Run code in the console or in scripts
 - Easy to view data, objects in memory, plots
 - Easy to create output such as papers or slides
 - Integration with Github

RMarkdown

Overview

- RMarkdown is an interactive coding environment
 - RMarkdown documents can combine text, LaTeX code, R code, and any output.
 - ► These slides are rendered using RMarkdown
 - You will be using RMarkdown for your homework assignments and hopefully your papers

Next lecture

Ordinary Least Squares regression with two variables

Lab 1

- RStudio and RMarkdown
- Manipulating random variables
- Intro to tidyverse and ggplot for data manipulation and visualization