# SOC542 Statistical Methods in Sociology II Probability and Bayesian Inference

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## **Plan**

- Probability review
- Bayes' theorem and its applications
- Comparing Bayesian and Frequentist approaches
- Bayesian estimation
- ► Lab: Bayesian regression in R

- $\triangleright$  P(A) refers to the probability of an event A
  - e.g. P(A) = 0.5 when referring to the probability of receiving a heads on a fair coin toss.
  - e.g.  $P(B) = \frac{1}{6}$  is the probability of rolling six with a fair die.
- ▶ In each case, we have a *random process* with a set of possible outcomes (e.g. heads or tails) referred to as the *sample space*.

## Simple probability

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  - P(A)P(A) = P(A) \* P(A) = 0.5 \* 0.5 = 0.25
- What is the probability of a sequence of N heads?
  ► P(A)<sup>N</sup>
- ▶ In this case, P(A) becomes vanishingly small as  $n \to \infty$  ▶  $0.5^{10} = 0.00098 = \frac{1}{1024}$

## Simple probability

▶ We can easily use simulations to verify our calculation. In this case, I use the rbinom function to simulate 1024 sequences of 10 tosses of a fair coin.

```
sims <- rbinom(1024, 10, 0.5)
print(length(sims[sims >= 10]))
## [1] 0
```

## Independence

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➤ The two events are independent of one another, so the joint probability is simply the product of the probabilities of the two events.

## Conditional probability and independence

- ▶ P(A) and P(B) are independent if and only if P(A|B) = P(A).
  - e.g. The number we rolled on the die has no effect on the outcome of the coin toss.

## Conditional probability and independence

► Consider a deck of 52 standard playing cards. What is the probability of randomly drawing an Ace?¹

 $<sup>^{1}</sup>$ This example is taken from Cunningham 2021, p. 17. It is an example of sampling without replacement.

## Conditional probability and independence

Consider a deck of 52 standard playing cards. What is the probability of randomly drawing an Ace?

$$P(Ace) = 4/52 = 1/13$$

Let's assume we pick an Ace and put it to the side. What's the probability we get another Ace?

## Conditional probability and independence

Consider a deck of 52 standard playing cards. What is the probability of randomly drawing an Ace?

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- Wrong answer:  $P(Ace_2) = \frac{4}{52} = \frac{1}{13}$ .

## Conditional probability and independence

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- ▶ Let's assume we pick an Ace and put it to the side. What's the probability we get another Ace?
- Wrong answer:  $P(Ace_2) = \frac{4}{52} = \frac{1}{13}$ .
- Correct answer:  $P(Ace_2) = P(Ace_2|Ace_1) = 3/51 = 0.059$ .
- This is an example of *conditional probability* since  $P(Ace_2|Ace_1) \neq P(Ace_1)$ .

## Conditional probability and independence

▶ We can express a conditional probability as:

$$P(A|B) = \frac{P(B,A)}{P(B)}$$

- ► The probability of *A* conditional on *B* is the **joint probability** of *A* and *B*, divided by the **marginal probability** of *B*.
- The denominator the sum of over possible joint probabilities of B and A,  $\sum_{A^*} P(B, A^*)$ .
  - ▶ The \* denotes that  $A^*$  may take multiple values.

## Conditional probability and independence

- ▶ If two events are independent, then P(A|B) = P(A).
- ► To reject independence, we need to show that  $P(A, B) \neq P(A)P(B)$

## Bayes' theorem

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

## Bayes' theorem

What's the probability it is going to rain given that we can see clouds?

$$P(Rain|Cloud) = \frac{P(Cloud|Rain)P(Rain)}{P(Cloud)}$$

## Bayes' theorem

- Let's say we live in England...
  - ▶ P(Cloud) = 0.7
  - P(Rain) = 0.3
  - ightharpoonup P(Cloud|Rain) = 1

$$P(Rain|Cloud) = \frac{P(Cloud|Rain)P(Rain)}{P(Cloud)} = \frac{1*0.3}{0.7} = \frac{0.3}{0.7} \approx 0.429$$

## Deriving Bayes' theorem

▶ Start with the definition of conditional probability:

$$P(A|B) = \frac{P(B,A)}{P(B)}$$

► Multiply each side by *P*(*B*):

$$P(A|B)P(B) = P(B,A)$$

▶ Analogously, if we start with P(B|A) we can get:

$$P(B|A)P(A) = P(B,A)$$

## **Deriving Bayes' theorem**

► The previous example shows that the following quantities are equal:

$$P(A|B)P(B) = P(B|A)P(A)$$

▶ Divide both sides by P(B) to get Bayes' theorem:

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

#### **COVID-19** tests

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$P(C19|+) = \frac{P(+|C19)P(C19)}{P(+)}$$

#### **COVID-19** tests

$$P(C19|+) = \frac{P(+|C19)P(C19)}{P(+)}$$

- ▶ P(C19|+): Probability you have COVID-19 given that you test positive.
- ▶ P(+|C19): Probability you test positive given that you have COVID-19.
- ► *P*(*C*19): Probability you have COVID-19 given population infection rates.
- $\triangleright$  P(+): Probability a test returns a positive result.

#### COVID-19 tests

- ► Assume there is a 1% chance you have COVID-19.
- ► Assume a test has a false negative rate of 2%.
  - ▶ 98% of the time it correctly diagnoses COVID-19, 2% of the time it fails to detect it.
- ► Assume the same test has a false positive rate of 5%
  - ▶ 95% of the time it correctly rejects COVID-19 when a person is negative, 5% of the time it falsely diagnoses COVID-19.
- What is the probability you really have COVID-19 following a positive test?

COVID-19 tests: P(+|C19)

$$P(C19|+) = \frac{P(+|C19)P(C19)}{P(+)}$$

If we assume a false negative rate of 2%. Then the probability of a positive test given COVID-19 is P(+|C19) = 1 - 0.02 = 0.98.

COVID-19 tests: P(C19)

$$P(C19|+) = \frac{P(+|C19)P(C19)}{P(+)}$$

Assume 1% of the population has COVID-19, then P(C19) = 0.01.

## COVID-19 tests: P(+)

- ► To calculate the proportion of positive tests we need to count all the positive tests.
- ▶ We can thus reformulate Bayes rule as

$$\frac{P(B|A)P(A)}{P(B|A)P(A) + P(B|A*)P(A*)}$$

$$\frac{P(+|C19)P(C19)}{P(+|C19)P(C19) + P(+|C19-)P(C19-)}$$

## COVID-19 tests: P(+)

$$P(C19|+) = \frac{P(+|C19)P(C19)}{P(+)}$$

- We already know the first part of the denominator, P(+|C19)P(C19) = 0.98 \* 0.01.
- ► If the test has a false positive rate of 5%, P(+|C19-) = 0.05\*(1-0.01)
- ▶ Thus, we take the sum of these probabilities to get the marginal probability of a positive test:

$$P(+) = (0.98 * 0.01) + (0.05 * (1 - 0.01))$$

## COVID-19 tests: Calculating P(C19|+)

▶ If we plug the numbers into Bayes' theorem we get

$$P(C19|+) = \frac{0.98 * 0.01}{0.98 * 0.01 + 0.05 * 0.99}$$

▶ We can use R to do the calculation for us

## **Terminology**

#### **Posterior** $\propto$ **Likelihood** $\times$ **Prior**

- In the previous example,
  - $\triangleright$  P(C19|+) is the **posterior**.
  - ▶ P(+|C19) is the **likelihood of the data**.
  - $\triangleright$  P(C19) is the **prior**.
- ▶ The denominator P(+) is ensures the result is a probability. It is often described as the **evidence** or the **marginal likelihood**.

## **COVID-19** tests: Tabular explanation

- ► The four cells in the middle of the table represent the *joint* probabilities of two events.
- ► The row and column totals represent the *marginal probabilities* of each event.
- $\triangleright$   $\theta$  is used to denote the parameters we are estimating.

Test result	$\theta = C19+$	$\theta = C19-$	Marginal (Test)
+	P(+ C19)P(C19)	P(+ C19-)P(C19-)	$\sum_{\theta} P(+ \theta)P(\theta)$
-	P(- C19)P(C19)	P(- C19-)P(C19-)	$\sum_{\theta} P(- \theta)P(\theta)$
Marginal C19	P(C19+)	P(C19-)	1.0

#### **COVID-19** tests: Tabular explanation

- ▶ To calculate P(C19|+) we can take the *joint probability* of C19 and a positive test and divide it by the *marginal probability* of a positive test.
- We can get the relevant values directly from the table: 0.98 \* 0.01/0.06.

Test result	$\theta = C19+$	$\theta = C19-$	Marginal (Test)
+	0.98*0.01	0.05*(1-0.01)	0.06
_	(1-0.98)*0.01	(1-0.05)*(1-0.01)	0.94
Marginal C19	0.01	(1-0.01)	1.0

## **Changing our priors**

▶ Let's change our prior to assume 10% COVID-19 prevalence in the population (perhaps this is a more reasonable assumption at the moment...)

```
(0.98*0.1) / ((0.98*0.1) + (0.05*0.9))
```

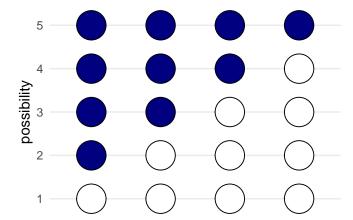
## [1] 0.6853147

- Now we get a much higher posterior probability.
- ► We could easily alter the calculation by incorporating other prior information, e.g. symptoms, exposure

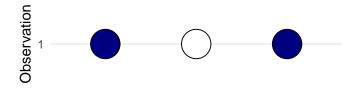
### McElreath's marble counting example

- Consider a bag containing four marbles
- The marbles can be white or blue
- We draw a sample of marbles from the bag (with replacement)

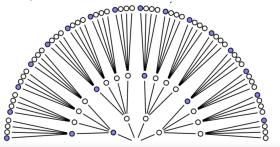
### Conjecture: Five possibilities



A sample from the bag produces



### Sampling and possibilities



McElreath 2020, Fig. 2.2 (p. 22)

### Counting the possibilities

<u> </u>	\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\
Conjecture	Ways to produce [B,W,B]
[W,W,W,W]	$0 \times 4 \times 0 = 0$
[B,W,W,W]	$1 \times 3 \times 1 = 3$
[B,B,W,W]	$2 \times 2 \times 2 = 8$
[B,B,B,W]	$3 \times 1 \times 3 = 9$
[B,B,B,B]	$4 \times 0 \times 4 = 0$

### From counts to probability

Conjecture	Propoportion B	Ways [B,W,B]	Plausibility
[W,W,W,W]	0.00	0	0.00
B,W,W,W]	0.25	3	0.15
[B,B,W,W]	0.50	8	0.40
[B,B,B,W]	0.75	9	0.45
[B,B,B,B]	1.00	0	0.00

### **Summary**

- We enumerated the set of plausible data generating processes p
- ▶ We counted the ways we could produce the data given each value of *p*. This is known as the *likelihood*.
- ▶ We normalized these counts to get *posterior* probabilities, which indicate the relative plausibility of each option *p*.
- The most plausible value is the one that has the most ways of generating the data.

#### **Incorporating prior information**

Now let's say we pick another marble and it's blue. We can use the prior information to update our counts.

Conjecture	Ways to produce [B]	Prior counts	New counts
[W,W,W,W]	0	0	$0 \times 0 = 0$
[B,W,W,W]	1	3	$3 \times 1 = 3$
[B,B,W,W]	2	8	$8 \times 2 = 16$
[B,B,B,W]	3	9	$9 \times 3 = 27$
[B,B,B,B]	4	0	$0 \times 4 = 0$

#### Bayes' theorem and data analysis

▶ In a general sense, we can think about Bayesian inference as calculating the posterior distribution in the following way:

$$\textit{Posterior} = \frac{\textit{Probability of the data} * \textit{Prior}}{\textit{Average probability of the data}}$$

## **Bayesian inference**

"Bayesian inference is reallocation of credibility across possibilities" - John Kruscke<sup>2</sup>

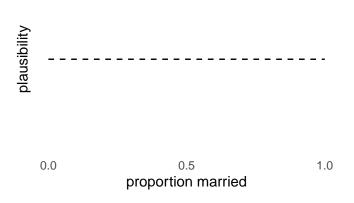
<sup>&</sup>lt;sup>2</sup>Chapter 2 of Kruschke's 2015 book *Doing Bayesian Data Analysis* provides an outline of his argument and is available online.

## Bayesian inference for a continuous parameter

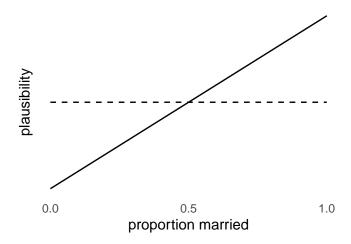
#### Estimating the marriage rate

- Assume a demographer is interested in estimating the marriage rate in the populuation.
- The demographer starts out with a "flat" prior
  - ► The marriage rate could be anywhere from 0 (nobody is married) to 1 (everybody is married).
- ► The demographer samples people at random and asks them their marital status.

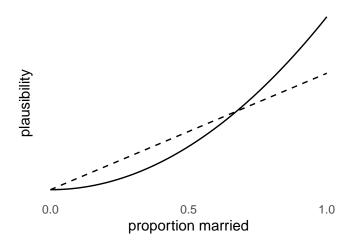
## Assume zero knowledge with a flat (uniform) prior



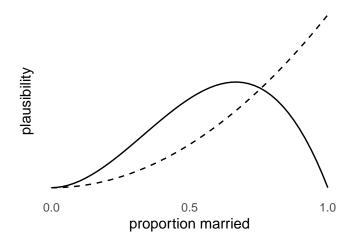
### First observation: Married



### Second observation: Married



## Third observation: Single



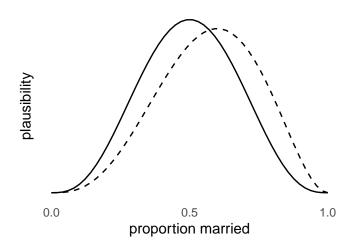
### Fourth observation: Married



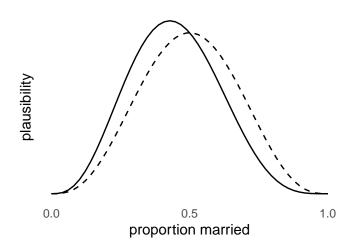
## Fifth observation: Single



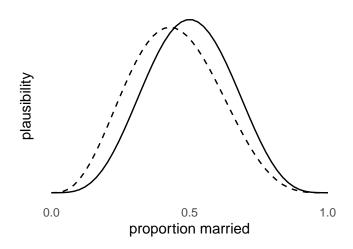
## Sixth observation: Single



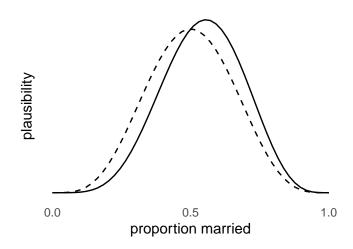
## Seventh observation: Single



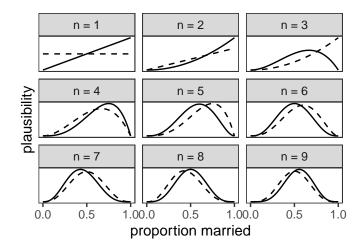
## Eighth observation: Married



## Nineth observation: Single



### **Overview**



## **Bayesian Updating**

- This example demonstrates the concept of Bayesian updating
   We use new information to update our beliefs
- ► Each time we update we use the previous **posterior** as the new **prior**!

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## **Bayesian Updating**

- This example demonstrates the concept of Bayesian updating
   We use new information to update our beliefs
- ► Each time we update we use the previous **posterior** as the new **prior**!
- ▶ Most of the time we use all our data at once to get the final posterior rather than iteratively updating.
- ▶ Bayesian updating is order invariant: we will get the same result regardless of the way observations are ordered.

## Formalizing a model

- ► The previous calculations are an example of the *binomial* distribution
  - $\triangleright$  Recall the distribution has two parameters N and p
- ► The goal of this analysis is to produce an estimate of the parameter *p*.
- We can thus write down a model to describe our analysis of marriage:

Marriage 
$$\sim$$
 Binomial(N, p)  
p  $\sim$  Uniform(0, 1)

► The goal of this analysis is to produce an estimate of the parameter *p*. In this case, we started with a flat, uniform prior.

Thomas Bayes (1701-1761)



Source: Wikipedia.

Pierre-Simon Laplace (1749-1827)



Source: Wikipedia.

Ronald Fisher (1890-1962)



Source: Wikipedia.

### **Historical developments**

- Frequentist (or "Fisherian") statistics dominated for most of the 20th century.
- Bayesian inference critiqued as too subjective and difficult to implement for complex problems.
- Reversal over the past couple of decades as critiques of Bayesian approach debunked, cheap compute power makes it tractable, and key tenets of Frequentist statistics are questioned (e.g. controversy over p-hacking<sup>3</sup>).
- ► The Bayesian approach is now mainstream in statistics and much of the natural sciences, but the social sciences have been slower to adopt.<sup>4</sup>

<sup>&</sup>lt;sup>3</sup>See Imbens 2021 reading from Week 1.

<sup>&</sup>lt;sup>4</sup>See Scott and Bartlett 2019

#### Theoretical foundations

- Frequentist
  - Long-run probabilities
  - Sampling distributions
- Bayesian
  - Probability theory

#### Sample size

- Frequentist
  - Properties of estimators depend on minimal sample size
- Bayesian
  - ► No minimum sample size
  - But larger samples allow for more precise estimates

#### **Point estimates**

- Frequentist
  - Models produce point estimates
- Bayesian
  - No singular point estimates
    - Many different summaries of the posterior distribution are possible (e.g. mean, median, mode)

#### **P-values**

- Frequentist
  - p-values used to communicate statistical significance
- Bayesian
  - Critique: p-values are based on arbitrary distributional assumptions
  - Uncertainty is captured by entire posterior distribution
  - ▶ Bayes' Factor is a Bayesian version of a p-value<sup>5</sup>

See Kruschke and Liddell 2018

#### **Confidence intervals**

- Frequentist
  - Confidence intervals defined using test statistics and conventions
  - Assumption that a parameter is fixed and that interval is derived from a sample
- Bayesian
  - Critique: Frequentist conventions are arbitrary
  - Assumption that a parameter has a distribution
  - Credible intervals or compatibility intervals can be used to summarize the posterior distribution

#### Confidence intervals: Interpretation of a 95% interval

- Frequentist
  - Over many repeat samples, 95% of calculated confidence intervals would contain the true value of the parameter
- Bayesian (assume an interval over 95% of the posterior distribution)
  - There is a 95% probability that the estimated parameter lies within the defined range, given the model and the data.
  - "What the interval indicates is a range of parameter values compatible with the model and the data." McElreath, p. 54.

#### **Bayesian Estimation**

- ► Three methods for estimating the posterior distribution
  - Analytical calculations
  - Grid and quadratic approximation
  - Markov Chain Monte Carlo

#### **Analytical calculations**

- ► For simple problems we can use calculus to provide an analytical solution for the posterior distribution
- But this approach does not scale well beyond simple problems like the marriage example

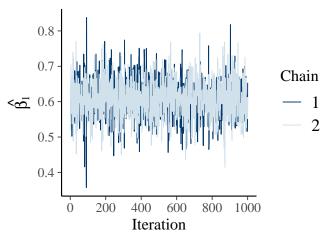
#### **Grid and quadratic approximation**

- Grid approximation (see McElreath 2.4.3)
  - We can approximate continuous spaces by using grids
    - ▶ But the method also scales very poorly to complex examples
- Quadratic approximation (see McElreath 2.4.4)
  - ▶ A more robust approach that involves using distributions to approximate the posterior
  - Flexible for many regression problems but also has trouble scaling

### Markov Chain Monte Carlo (MCMC)

- Use simulation to draw samples from the posterior distribution
  - A computationally intensive approach
  - ► Samples provide an approximation for complex spaces
  - More efficient for complex models than quadratic approximation
- ▶ MCMC has led to major advances in Bayesian methods since the 1990s (see McElreath 2.4.5).

### Samples from a Markov Chain



#### Stan and Hamiltonian Monte Carlo

- Stan is a programming language developed for statistical computing
- It implements Hamiltonian Monte Carlo (HMC) sampling
  - A variant of MCMC methods based on Hamiltonian physics
  - Approximates the posterior by "flicking" a particle and observing its movement
- ► HMC is highly effective at solving even complex problems<sup>6</sup>
  - ► It provides lots of useful diagnostics making it easier to debug than early MCMC approaches
  - Greater flexibility as it not require conjugacy

<sup>&</sup>lt;sup>6</sup>See McElreath Chapter 9 and Betancourt 2018 for a more advanced conceptual overview.

## **Bayesian Regression**

- ► Regression coefficients are the *unknown* parameters that we want to estimate given a model and the observed data.
- We can formalize these assumptions by writing down a model that looks something like this:

$$y_i \sim ext{Normal}(\mu_i, \sigma)$$
 $\mu_i = eta_0 + eta_1 x_i$ 
 $eta_0 \sim ext{Normal}(0, 10)$ 
 $eta_1 \sim ext{Normal}(0, 1)$ 
 $\sigma \sim ext{Uniform}(0, 1)$ 

## **Bayesian Regression**

- In this case, we make the assumption that  $y_i$  is normally distributed and that we can express its mean in terms of x (recall that  $E[y|x] = \beta_0 + \beta_1 x_i$ )
- ► Each *parameter* in the model has a *prior* distribution. We specify these before we have seen any data
- After estimating a model using the data we get the *posterior* distribution for each parameter

## **Bayesian Regression**

- We will be using stan\_glm to estimate these kinds of models using HMC
- ► The *posterior distributions* of the parameters are then analyzed to make inferences about the relationship between *x* and *y*
- ► We can also use the posterior to generate new data consistent with the model

### **Final remarks**

"All models are wrong, but some are useful" - George Box<sup>7</sup>

 $<sup>^{7}\</sup>mathrm{This}$  aphorism is attributed to statistician George Box. See Wikipedia for further discussion.