SOC542 Statistical Methods in Sociology II Binary outcomes I

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Course updates

- ► Homework 3 released today, due next Friday 4/4
 - Binary outcomes
 - Interaction terms
 - Marginal effects

Plan

- ► Linear probability model
- ► Logistic regression
- ► Probit regression

Binary outcomes

- A binary outcome variable y consists of two possible values, 0 or 1.
 - e.g. $y \sim Binomial(n, p)$ is a sequence of n observations, where $P(y_i = 1) = p$.
- ▶ We have already encountered binary independent variables, known as dummy variables, but now we want to put them on the left side of the regression equation.

Definition

► The linear probability model (LPM) is estimates the probability a binary dependent variable is equal to one as a linear function of predictors.

Specification

► The LPM is estimated using OLS:

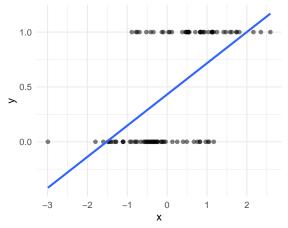
$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + u$$

► Thus,

$$P(y = 1|x_1, x_2, ...x_k) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + ... + \beta_k x_k + u$$

▶ The coefficient β_i represents the change in probability that y=1 associated with a unit-change in x_i , holding other regressors constant.

Fitting a line to a simulated binary outcome



Example: Diffusion of Microfinance¹

- Survey data from 75 villages in Karnataka, India
 - Focus only on women and 72 villages
 - Listwise deletion used to drop respondents missing key variables
 - N = 9064
- Dependent variable:
 - Membership in a micro-finance Self-Help Group (SHG), N = 3357
- Independent variables:
 - ► Age (continuous) and age squared
 - Caste (dummy, low/high)
- ► Fixed-effects:
 - Village (dummy)

¹Data from Banerjee, A., A. G. Chandrasekhar, E. Duflo, and M. O. Jackson. 2013. "The Diffusion of Microfinance." *Science* 341 (6144): 1236498–1236498. Link to paper. Harvard Dataverse link

Example: Predicting Self-Help Group membership²

²Based on a propensity score model in Davidson, Thomas, and Paromita Sanyal. 2017. "Associational Participation and Network Expansion: Microcredit Self-Help Groups and Poor Women's Social Ties in Rural India." Social Forces 95 (4): 1695–1724. https://doi.org/10.1093/sf/sox021.

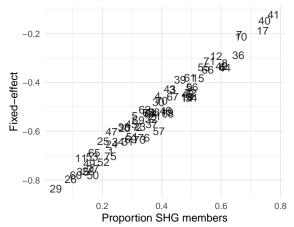
Correctly estimating fixed-effects and standard errors

Example: Predicting Self-Help Group membership

	Model 1	Model 2
(Intercept)	-0.702***	
	(0.057)	
Age	0.047***	0.047***
	(0.002)	(0.003)
Age2	-0.001***	-0.001***
	(0.000)	(0.000)
Caste (Lower)	0.060***	0.060**
	(0.011)	(0.019)
Num.Obs.	9064	9064
Std.Errors		by: village
	·	

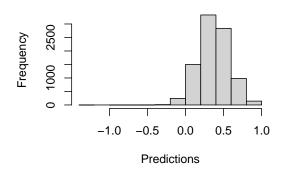
^{*} p <0.05, ** p <0.01, *** p <0.001 Village fixed-effects omitted.

Village fixed-effects



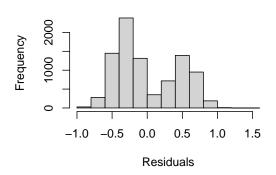
Predicted values

Histogram of Predictions

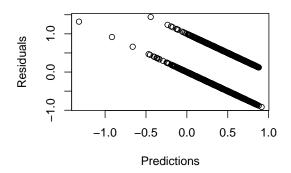


Residuals

Histogram of Residuals



Predicted values and residuals



Limitations

- Unrealistic predictions
 - Nothing constrains predictions to be probabilities bounded by [0,1] so the model can make unrealistic predictions
- Heteroskedastic errors
 - Requires the use of heteroskedasticity-robust standard errors³
- ► R² no longer reliable since most predicted values are not exactly 0 or 1
 - ▶ Under what circumstances could $R^2 = 1$ be achieved with a binary outcome?

³We will discuss this topic in more detail in Week 12.

Strengths

- Despite the limitations, the LPM has grown in popularity as an alternative to logistic regression
 - Simple to estimate
 - Easy to interpret
 - Prediction range often plausible or irrelevant

Addressing the limitations

- ▶ We can address the limitations of the LPM by using a different functional form to ensure that predicted values are constrained to the [0,1] range
- ➤ To do this must extend the linear model by using a link function to map a linear model onto a non-linear outcome space.
- Such models are known as generalized linear models (GLM).

The logit function

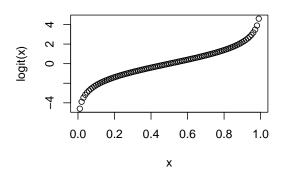
▶ The **logit** function takes values in the range [0,1] and maps them to the range $[-\infty,\infty]$

$$logit(x) = log(\frac{x}{1 - x})$$

The logit function

```
logit <- function(x) {return(log(x/(1-x)))}
logit(c(0, 0.01, 0.5, 0.99, 1))
## [1] -Inf -4.59512 0.00000 4.59512 Inf</pre>
```

The logit function



The inverse logit function

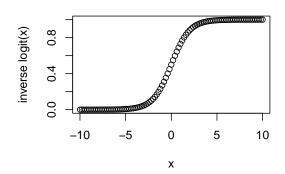
► The **inverse logit function** reverses this transformation, mapping values back to the [0,1] range:

$$logit^{-1}(x) = \frac{e^x}{1 + e^x} = \frac{1}{1 + e^{-x}}$$

The inverse logit function

```
invlogit <- function(x) {return((exp(1)^x)/(1 + exp(1)^x))}
invlogit(c(-2, -1, 0, 1, 10)) %>% round(3)
## [1] 0.119 0.269 0.500 0.731 1.000
```

The inverse logit function



We can write the following model for a binary outcome, where $p_i = P(y_i = 1 | x_{1i}, x_{2i}, ..., x_{ki})$:

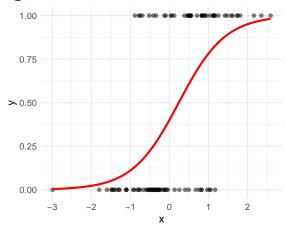
$$logit(p_i) = log(\frac{p_i}{1 - p_i}) = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \beta_k x_{ki}$$

▶ The model can be expressed in terms of p_i using the inverse-logit function (also known as the logistic function):

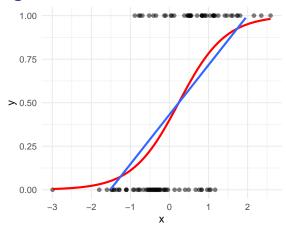
$$p_i = logit^{-1}(\beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + ... + \beta_k x_{ki})$$

$$p_{i} = \frac{1}{1 + e^{-(\beta_{0} + \beta_{1} x_{1i} + \beta_{2} x_{1i} + \dots + \beta_{k} x_{ki})}}$$

Fitting a logistic curve to simulated data



Fitting a logistic curve to simulated data



Error terms

- ▶ OLS regression has an error-term because we assume that the outcome is normally-distributed with two parameters, μ and σ^2 , i.e. $y \sim N(\mu, \sigma^2)$.
- ▶ Binary random variables follow a Bernoulli distribution,⁴ defined by a single parameter, p, the probability of a success. Thus, $y \sim Bernoulli(p)$.
- ▶ Therefore, there is *no error term in logistic regression*.

⁴Recall this is equivalent to Binomial(1, p).

Estimation

- There is no direct algebraic solution to obtain such estimates (unlike OLS regression, i.e. $(X^TX)^{-1}X^Ty$).
- ► In frequentist statistics, the parameters in a logistic regression (and most other GLMs) are estimated using Maximum Likelihood Estimation (MLE).
 - ► The MLE estimates have the same general properties as OLS estimates w.r.t standard errors, confidence intervals, and p-values.

Maximum Likelihood Estimation

For logistic regression, where β is a vector of coefficients and X is a matrix of predictors, the **likehood** function is written as

$$P(y|\beta, X) = \prod_{i=1}^{n} (logit^{-1}(X_{i}\beta))^{y_{i}} (1 - logit^{-1}(X_{i}\beta))^{1-y_{i}}$$

- ▶ The goal of MLE is to find the β that *maximizes* this function.
 - It finds the parameter values most likely to have produced the observed data.
- An *iterative* algorithm is used to find the parameters that maximize the likelihood function, typically using the logarithm of the likelihood function for computational efficiency.
 - Unlike OLS, models can sometimes fail to converge on an appropriate solution. This can be an issue when trying to fit complex models.

Estimation in R

We can easily estimate this using the glm function in R. The family argument is used to select the appropriate model.

MLE convergence

We can inspect the output to verify that the model converge on the maximum-likelihood estimates.

```
logit.mle$converged # Has model converged?
## [1] TRUE
logit.mle$iter # How many MLE iterations used?
## [1] 5
```

Comparison with the LPM

	LPM	Logistic
Age	0.047***	0.300***
	(0.002)	(0.014)
Age2	-0.001***	-0.004***
	(0.000)	(0.000)
Caste (Lower)	0.060***	0.315***
	(0.011)	(0.054)
Num.Obs.	9064	9064
Log.Lik.	-5430.095	-5115.536

^{*} p <0.05, ** p <0.01, *** p <0.001 Village fixed-effects omitted.

Bayesian estimation

- Logistic regression can alternatively be estimated using Bayesian methods.
 - We can use the same MCMC approach as used for OLS regression.
- ▶ Bayesian estimation does not converge on a single maximum likelihood estimate of $\hat{\beta}$, but produces a posterior distribution, $\hat{\beta}_{posterior}$
- ▶ If a *uniform prior* is used, the posterior density is proportional to the likelihood function and the *mode* is equal to the maximum likelihood estimate.
 - Typically, we can do better by using more informative priors.

Bayesian estimation in R

Note: Bayesian estimation can be considerably slower than MLE. The model requires more iterations than the default to ensure convergence, likely due to the nested structure of the data (more informative priors would likely help).

Comparing Maximum Likelihood and Bayesian estimates⁵

	MLE	Bayes
(Intercept)	-7.066	-7.098
	[-7.782, -6.378]	[-7.778, -6.444]
Age	0.300	0.300
	[0.273, 0.328]	[0.274, 0.327]
Age2	-0.004	-0.004
	[-0.004, -0.003]	[-0.004, -0.003]
Caste (Lower)	0.315	0.320
	[0.209, 0.422]	[0.206, 0.420]
Num.Obs.	9064	9064
Log.Lik.	-5115.536	-5115.446
ELPD		-5192.1

Village fixed-effects omitted.

Interpretation: Terminology

- If an outcome occurs with probability p, the **odds** of the outcome are defined as $\frac{p}{1-p}$.
 - ▶ If p = 0.5, the odds = $\frac{0.5}{1 0.5} = 1$. If $p = \frac{2}{3}$, the odds ≈ 2 .
- The log odds is the natural logarithm of the odds
 - ► This is also known as the **logit** function, $log(\frac{p}{1-p})$
- ► An **odds ratio** is the ratio of two odds.
- ▶ Odds ratios can be used as a way to communicate changes in the probability scale:

$$OR(0.6, 0.8) = \frac{\frac{0.6}{1-0.6}}{\frac{0.8}{1-0.8}} = \frac{1.5}{4} = 0.375$$

Interpretation: Log-odds

- ▶ Since the outcome can be expressed as $log(\frac{p}{1-p})$, the coefficients in the regression output are log odds.
 - Where $\beta_{age} = 0.3$, a 1 year increase is associated with a 0.3 increase in the log-odds of SHG membership.
 - $eta_{caste}=0.315$, implying that belonging to a lower caste group versus a higher caste group changes the log-odds of SHG membership by 0.315.

Interpretation: Odds-ratios

- We can get the odds ratio by exponentiating the coefficients
 - OR(age) = exp(0.3) = 1.35. A 1 year increase in age is associated with a 34% increase in the odds of SHG membership.
 - ► $OR(caste) = \frac{Odds(SHG=1|lower-caste)}{Odds(SHG=1|higher-caste)} = exp(0.315) = 1.37$. This implies that low caste residents are more likely to belong to SHGs (a 37% increase in the odds of SHG membership).

Interpretation: Intuition

Why do we get an odds ratio and not an odds when we exponentiate a log odds?

$$\begin{split} exp(\beta_{caste}) &= exp(\beta_{caste_{low}}^* - \beta_{caste_{high}}^*) \\ &= exp(log(odds(p|caste_{low})) - log(odds(p|caste_{high}))) \\ &= \frac{exp(log(odds(p|caste_{low}))}{exp(log(odds(p|caste_{high}))))} \\ &= \frac{odds(p|caste_{low})}{odds(p|caste_{high})} \end{split}$$

Interpretation: The divide-by-4 rule

- ► The divide-by-4 rule provides a quick way to assess the effects of predictors in a logistic regression on the probability scale:
 - The logistic curve is steepest at the center, where $\beta_0 + \beta X = 0$ and $logit^{-1}(\beta_0 + \beta X) = 0.5$. The slope (or the derivative of the logistic function) is maximized.
 - At this point, $\frac{\beta e^0}{(1+e^0)^2} = \frac{\beta}{(1+1)^2} = \beta/4$.
- ▶ $\beta/4$ is the *maximum difference* in P(y=1) corresponding to a unit change in x.
 - ► This provides an simple approximation for the *maximum effect* of a predictor.
- ▶ For example, $\beta_{Age}/4 = 0.3/4 = 0.075$. Thus, the maximum effect of a 1-year results in a maximum 7.5% increase in the probability of SHG membership.

Confidence intervals

- ➤ The standard formula for calculating confidence intervals assumes normality. This assumption is violated by logistic regression so standard (Wald) confidence intervals are incorrect.⁶
- Instead, confidence intervals for GLMs are calculated by using information from the likelihood function using the profile likelihood approach.
 - ▶ Note: Profile intervals are considerably slower to compute than Wald intervals.
- ► For Bayesian models, we can construct credible intervals using the posterior distribution in the same fashion as OLS models.

 $^{^6}$ In practice, the two approaches often produce very similar results, as the following example shows.

Confidence intervals

The confint function in R allows us to calculate the correct confidence intervals (in this case for the effect of age). confint.default provides standard Wald confidence intervals. The intervals are exponentiated to get odds ratios.

```
round(exp(confint.default(logit.mle)[2,]),4)

## 2.5 % 97.5 %

## 1.3130 1.3873

round(exp(confint(logit.mle))[2,],4)

## 2.5 % 97.5 %

## 1.3133 1.3877
```

Model fit: Log-likelihood

► The log-likelihood of a model is defined as

$$\sum_{i=1}^{n} log(p_i)y_i + log(1-p_i)(1-y_i)$$

- ▶ If $y_i = 1$, we add $log(p_i)$, otherwise we add $log(1 p_i)$.
- It is always negative.⁷ A higher score indicates a better fit
 But like R², adding more variables tends to increase the score.
- ▶ A related measure known as **deviance** is simply −2 times the log-likelihood.

⁷Recall $log_e(1) = 0$.

Model fit: Log-likelihood

We can calculate the log-likelihood using the formula above or the logLik function:

```
y <- data$shg
pred_probs <- predict(logit.mle, type = "response")
sum(log(pred_probs)*y + log(1 - pred_probs)*(1-y))
## [1] -5115.536
print(logLik(logit.mle))
## 'log Lik.' -5115.536 (df=75)
```

Model fit: Log-likelihood

	Village	Caste	Age	Full	+Religion
(Intercept)	-1.421	-1.470	-6.981	-7.066	-7.053
	(0.243)	(0.244)	(0.357)	(0.358)	(0.381)
Caste (Lower)		0.286		0.315	0.316
		(0.052)		(0.054)	(0.055)
Age			0.298	0.300	0.300
			(0.014)	(0.014)	(0.014)
Age2			-0.004	-0.004	-0.004
			(0.000)	(0.000)	(0.000)
Hindu					-0.014
					(0.139)
Log.Lik.	-5431.917	-5417.034	-5132.504	-5115.536	-5115.531

Model fit: Log-likelihood (Bayesian)

We can use the same formula to calculate the log-likelihood for a Bayesian model.⁸ Do you notice any problems with such an approach?

```
pred_probs.bayes <- predict(logit.bayes, type = "response")</pre>
sum(log(pred probs.bayes)*v + log(1 - pred probs.bayes)*(1-v))
## [1] -5115.612
```

⁸The standard 'logLik' function does not work for Bayesian models.

Model fit: Log-likelihood (Bayesian)

To incorporate the uncertainty in the posterior distribution, we can take the average log-likelihood of each point over all posterior samples S, known as the **log-pointwise-predictive-density**. 9

^{## [1] -5115.438}

⁹See McElreath p. 210

Model fit: Log-likelihood (Bayesian)

► The estimates using predict and posterior_predict are almost identical, but you should always use the full posterior distribution when computing any summary statistics from Bayesian models.

Model fit: R^2 and fraction correctly predicted

- $ightharpoonup R^2$ is no longer a useful measure of fit for binary outcomes.
- ▶ A simple fit measure is the fraction of cases correctly predicted, where $\hat{y_i} = 1$ if $p_i > 0.5$ and $\hat{y_i} = 0$ if $p_i \leq 0.5$, but this approach throws out information about the predicted probabilities.

Model fit: Pseudo-R²

- ► There are several different approaches to construct **pseudo**- R^2 statistics for logistic regression. These measures approximate an R^2 by ranging between 0 and 1, but are not equivalent. ¹⁰
- ▶ One of the more common variants is McFadden's pseudo- R^2 :

$$R^2 = 1 - \frac{LL(M_{full})}{LL(M_{intercept})}$$

▶ This is the standard formula for R^2 , where the log-likelihood of an intercept-only model as the total sum of squares and the fully parameterized model is the sum of squared errors.

¹⁰See this blog post for a discussion of several different measures.

Model fit: McFadden's pseudo- R^2

Model fit: Bayesian held-out likelihood

- For Bayesian models, we can use the loo function to calculate a held-out likelihood score using the entire posterior distribution.
 - The elpd_loo score provides an approximation of the LOO-CV expected log-pointwise predictive density.
- ► This will be slightly more conservative (lower) than the in-sample log-likelihood.

Model fit: Bayesian held-out likelihood

```
print(loo(logit.bayes))
##
## Computed from 2000 by 9064 log-likelihood matrix.
##
##
        Estimate SE
## elpd_loo -5192.1 43.9
## p loo 76.7 1.0
## looic 10384.2 87.7
## -----
## MCSE of elpd_loo is 0.2.
## MCSE and ESS estimates assume MCMC draws (r_eff in [0.0, 2.9]).
##
## All Pareto k estimates are good (k < 0.7).
## See help('pareto-k-diagnostic') for details.
```

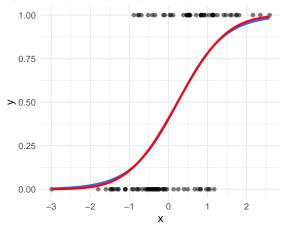
The probit regression model is similar to logistic regression but uses a cumulative normal function Φ instead of the inverse logistic function:

$$P(y = 1|X) = \Phi(X\beta)$$

- Historically, probit has been preferred in some cases for computational reasons, but the two models tend to produce similar results.¹¹
- Like logistic regression, probit regression can be estimated using MLE or Bayesian approaches.

 $^{^{11}\}mbox{See}$ this blog post for an example of a comparison using Monte Carlo simulation.

Probit (red) and logistic (blue) curves



Estimation

The only change to the model is the link function.

Comparison with logistic regression

	Logistic	Probit			
(Intercept)	-7.066***	-4.064***			
	(0.358)	(0.202)			
Age	0.300***	0.169***			
	(0.014)	(800.0)			
Age2	-0.004***	-0.002***			
	(0.000)	(0.000)			
Caste (Lower)	0.315***	0.190***			
	(0.054)	(0.032)			
Num.Obs.	9064	9064			
Log.Lik.	-5115.536	-5123.682			
F	17.330	19.151			
* - <0.05 ** - <0.01 *** - <0.001					

^{*} p <0.05, ** p <0.01, *** p <0.001

Interpretation

- ► The coefficients in a probit regression are more difficult to interpret than OLS or logistic regression.
- ▶ In general, positive coefficients indicate increases in the predicted probability of the outcome, while negative coefficients indicate decreases.¹²

¹²See the Stata blog for further discussion.

Summary

- ► LPM
 - Estimate using OLS
 - ▶ Easy to estimate and interpret, but can make bad predictions
- Logistic regression
 - Logistic function used to apply linear model to non-linear outcome
 - Interpret log-odds and odds-ratios
 - Estimate using MLE or Bayes
- Probit regression
 - Cumulative normal distribution as link
 - Similar fit to logistic but more difficult to interpret

Next week

- Logistic regression continued
 - ▶ Predictions and marginal effects
 - Interaction terms

Lab

► Regression models for binary outcomes