

Stat 565

The PACF & Estimation for ARMA

Jan 26 2016

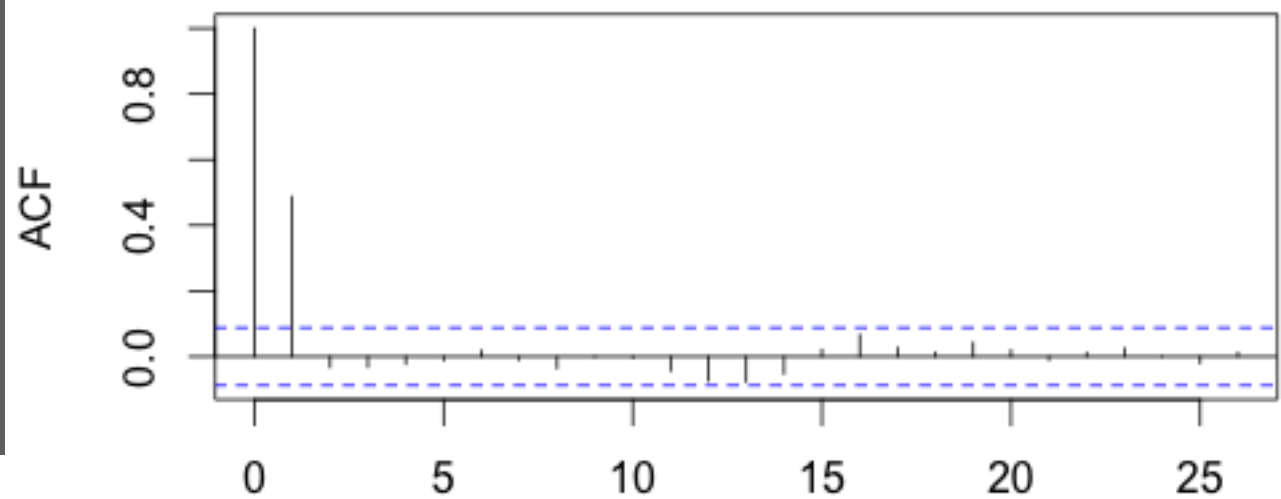
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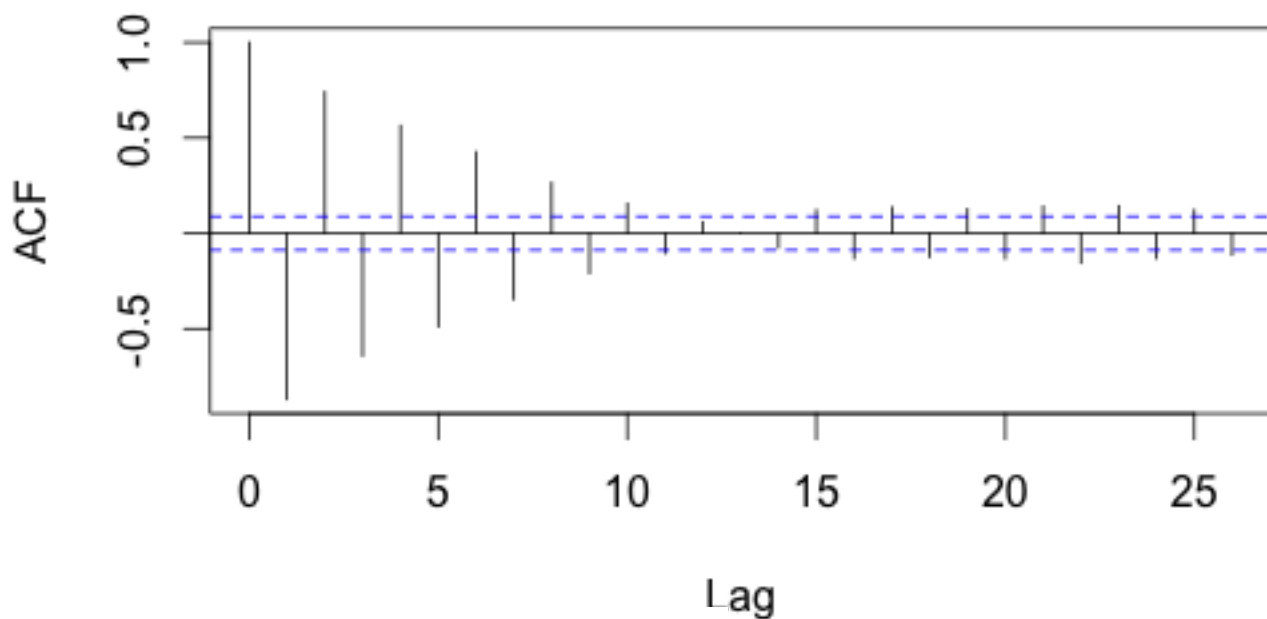
Which **basic** models might these simulated data come from?

AR(1), MA(1), or white noise

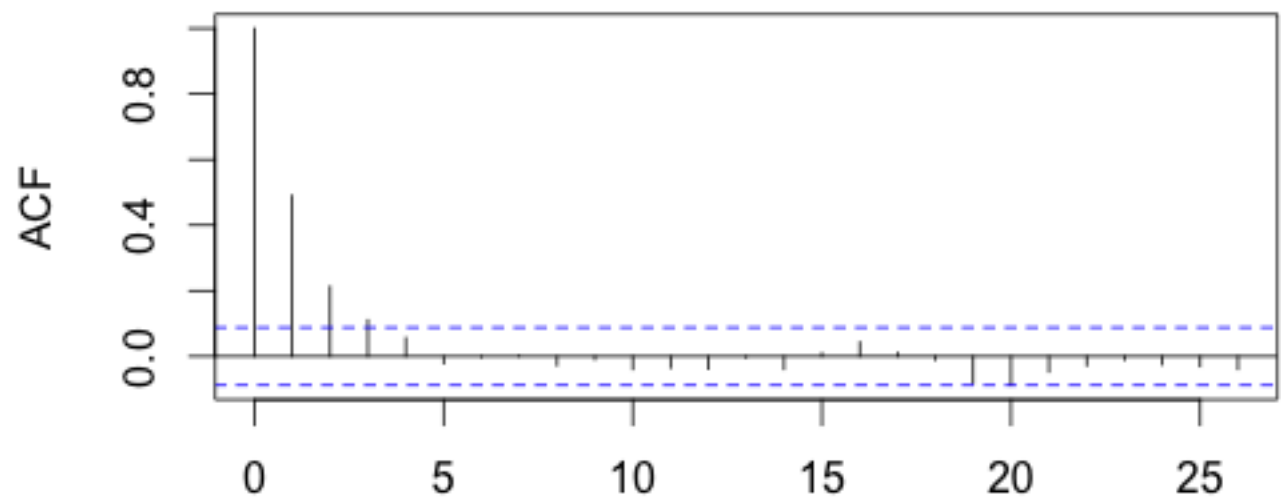
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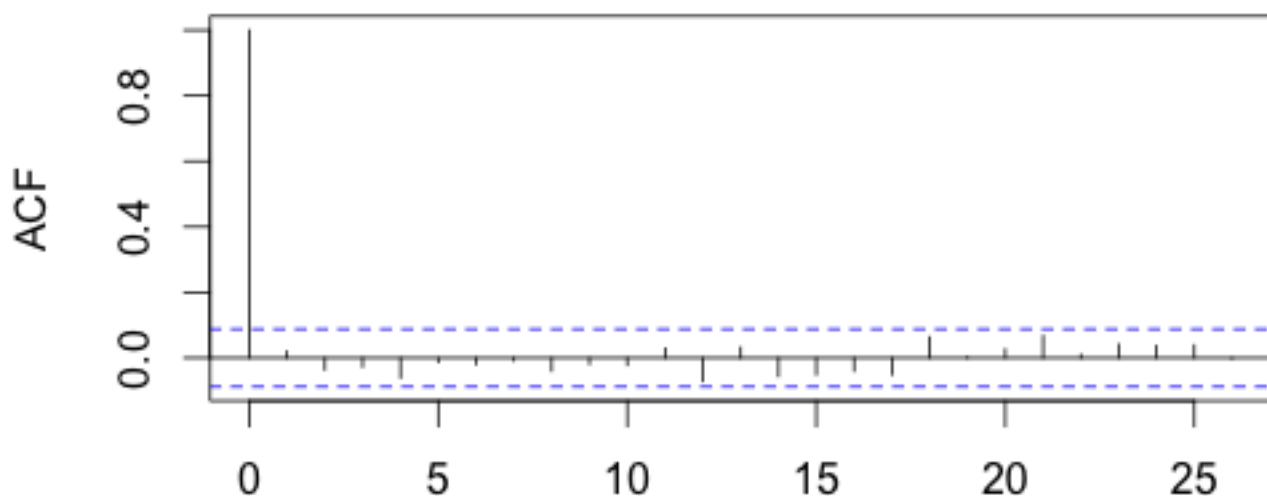
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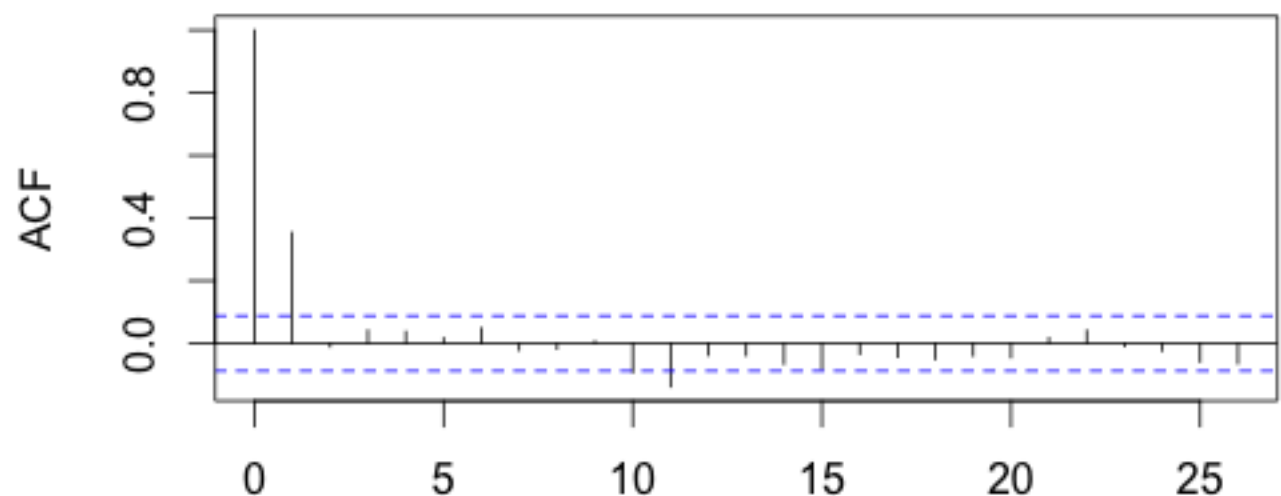
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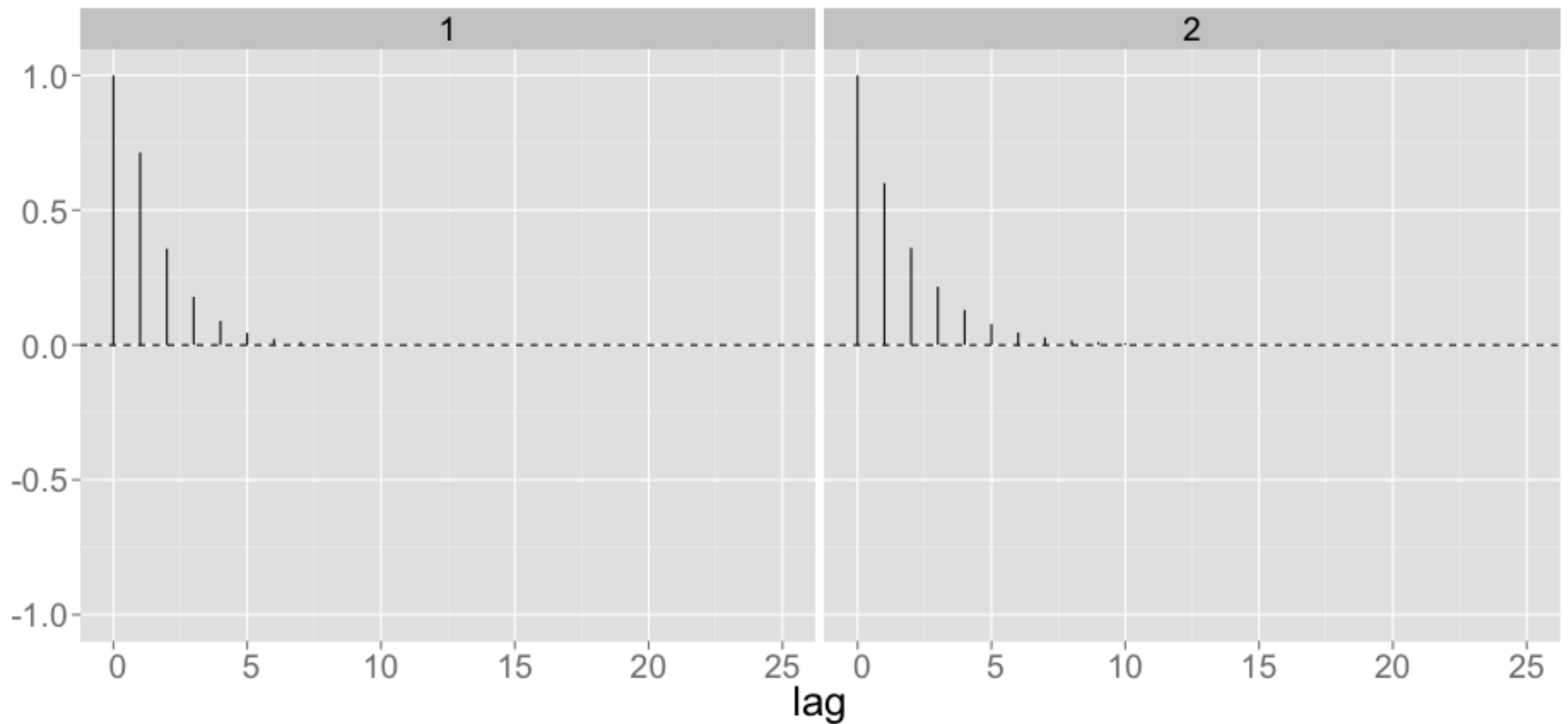
5



ACF

One is AR(1), $\alpha_1 = 0.6$

The other is ARMA(1, 1), $\beta_1 = 0.5$, $\alpha_1 = 0.5$



Which is which?

Partial autocorrelation function

Basic idea: what is the correlation between x_t and x_{t+h} , after taking into account $x_{t+1}, x_{t+2}, \dots, x_{t+h-1}$?

Technically:

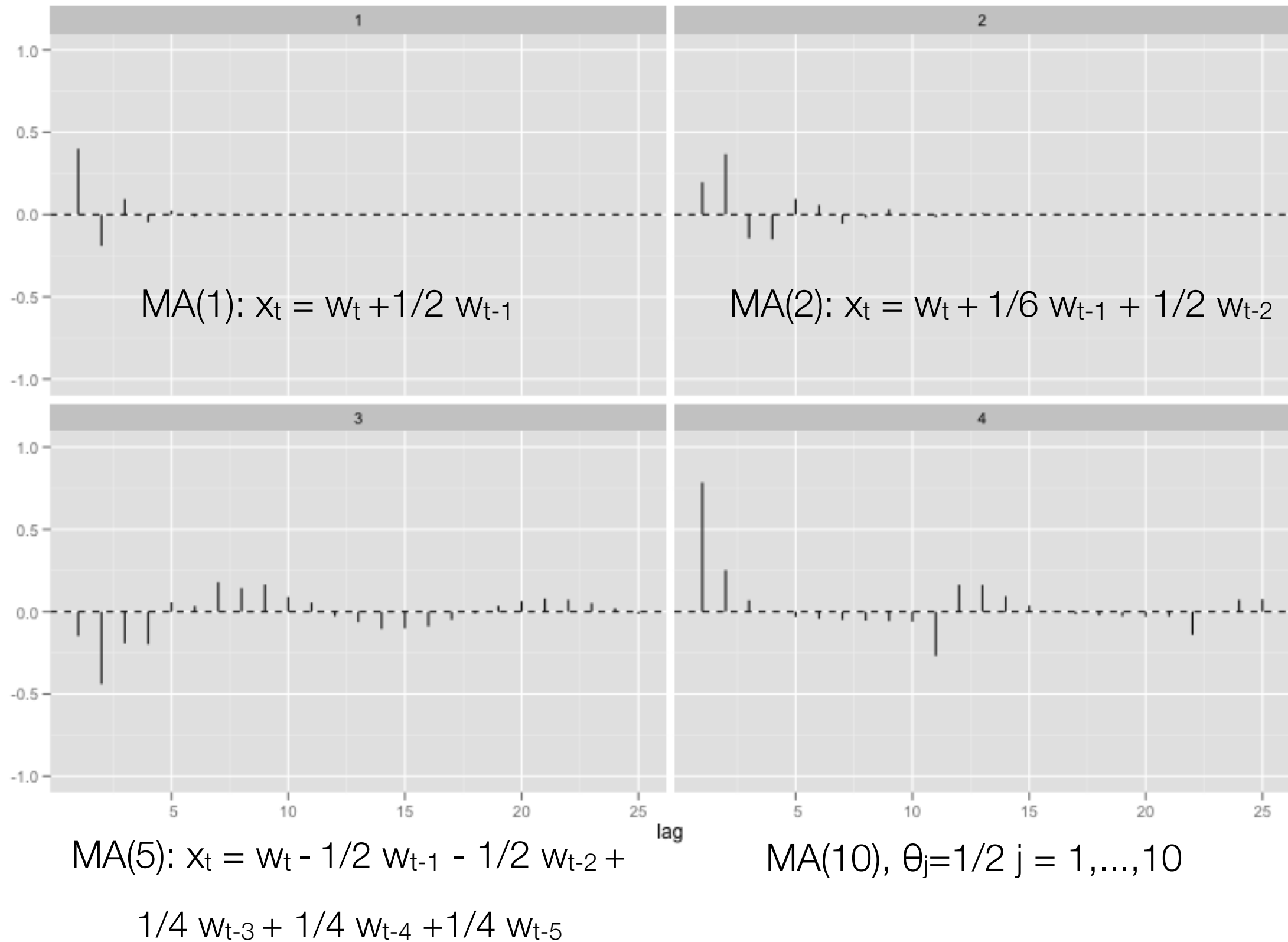
Regress x_t on $x_{t+1}, x_{t+2}, \dots, x_{t+h-1}$ to find the fitted value \hat{x}_t .

Regress x_{t+h} on $x_{t+1}, x_{t+2}, \dots, x_{t+h-1}$ to find the fitted value \hat{x}_{t+h} .

Find $\text{cor}(x_t - \hat{x}_t, x_{t+h} - \hat{x}_{t+h})$, call this $\text{PACF}(h) = \phi_{hh}$

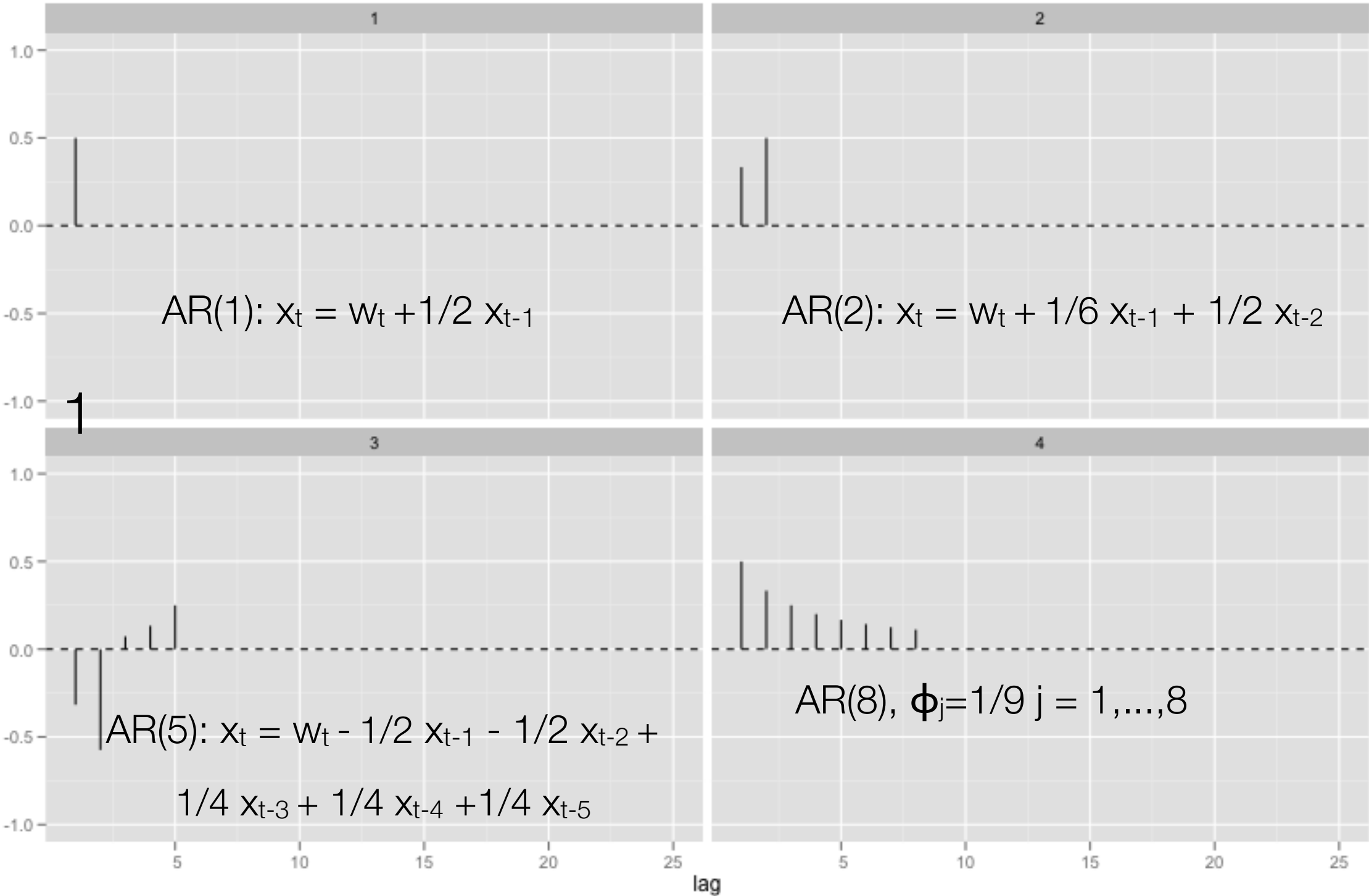
PACF

MA(q)



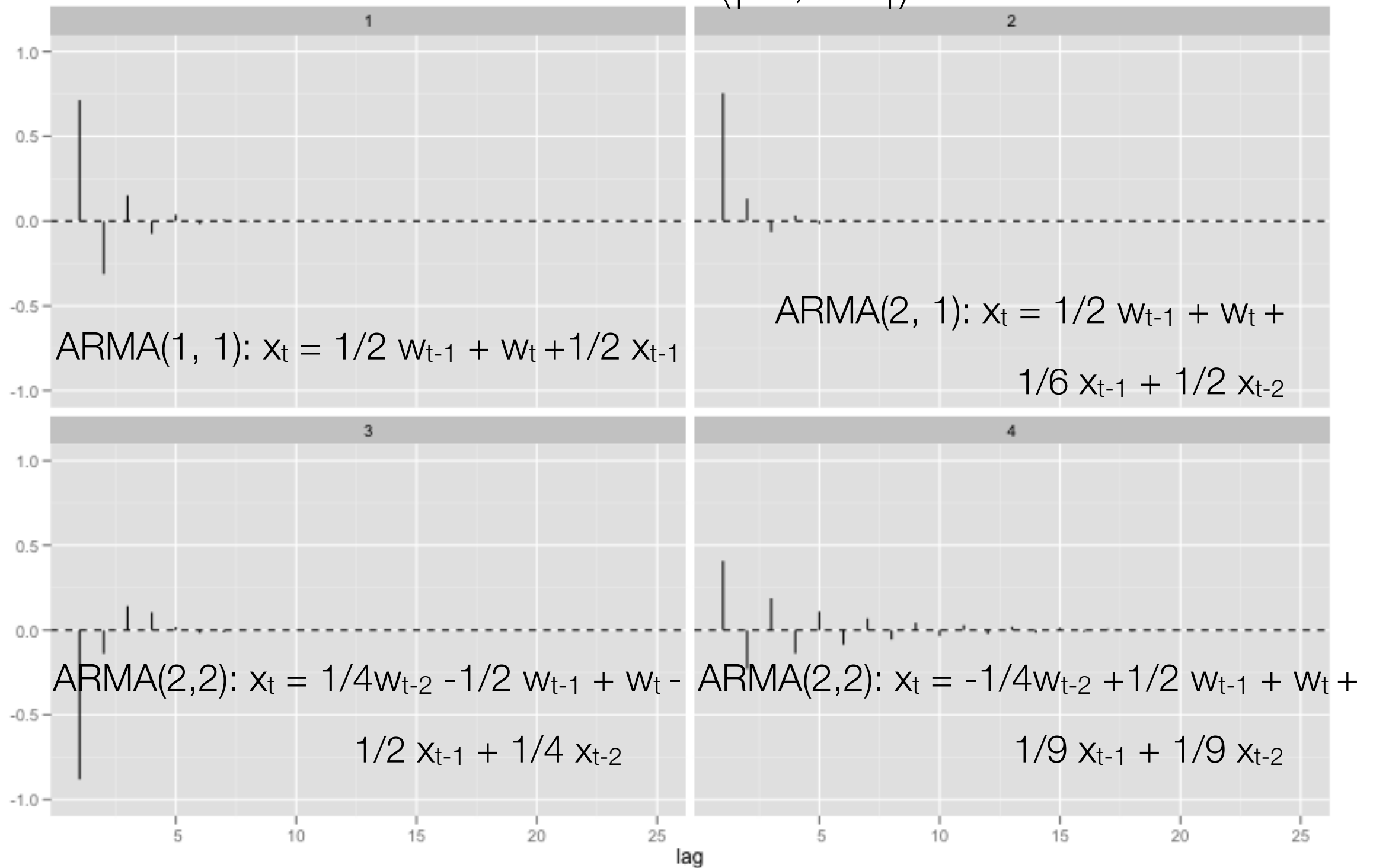
PACF

AR(p)



PACF

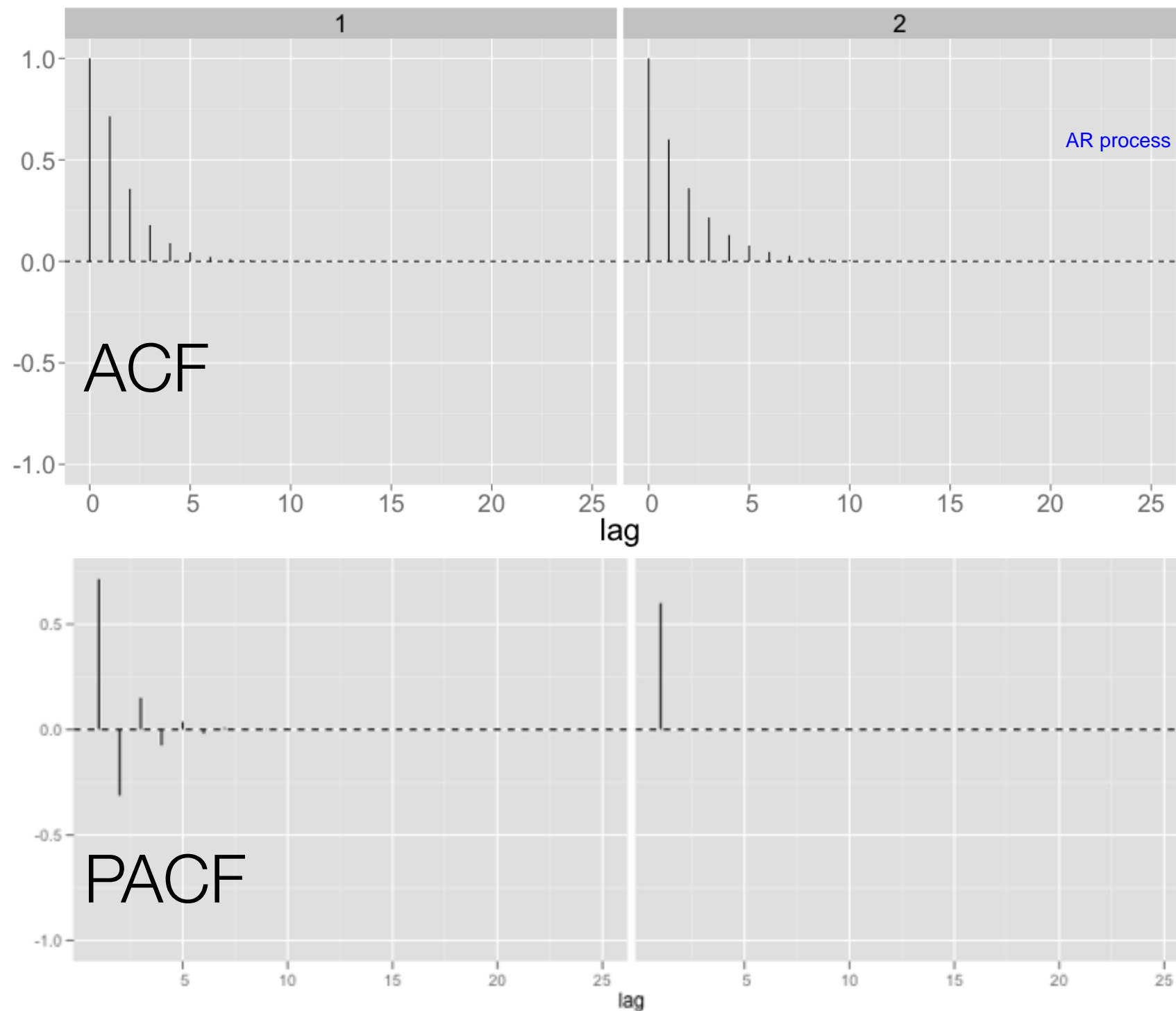
ARMA(p, q)



	MA(q)	AR(p)	ARMA(p, q)
ACF	zero lags > q	tails off	tails off
PACF	tails off	zero lags > p	tails off

One is AR(1), $\alpha_1 = 0.6$

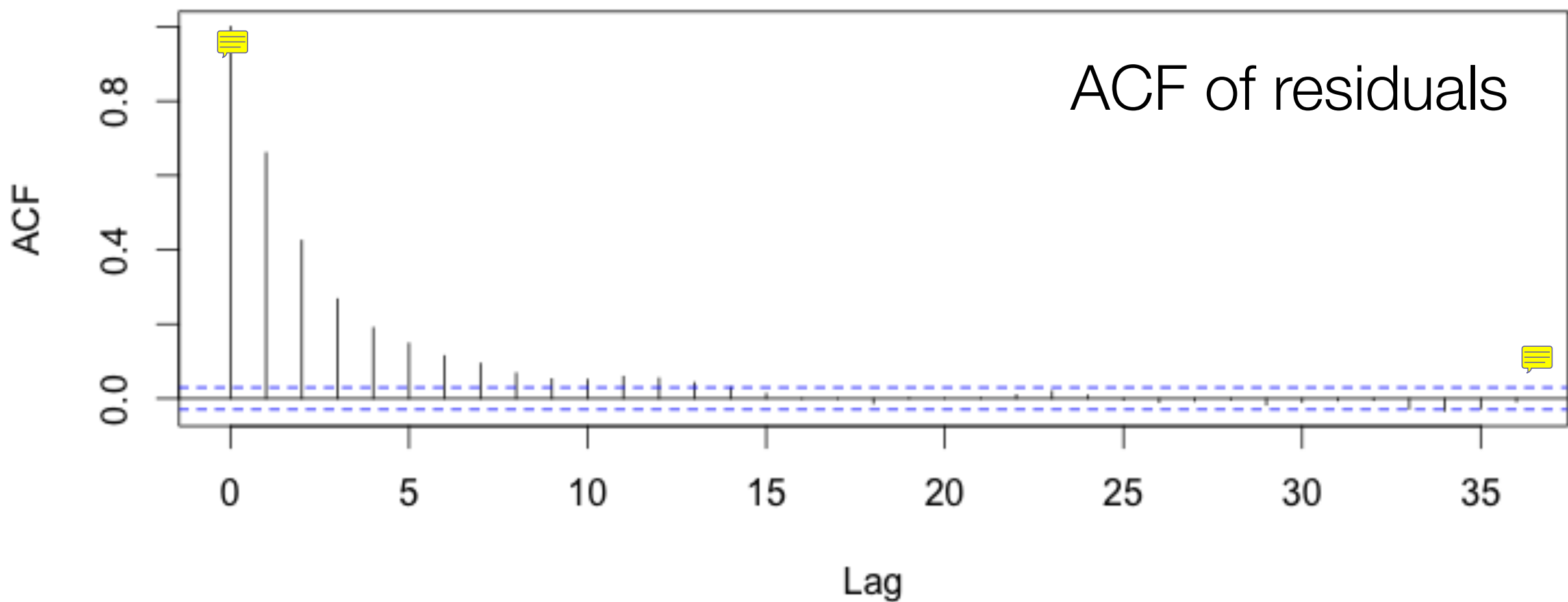
The other is ARMA(1, 1), $\beta_1 = 0.5$, $\alpha_1 = 0.5$



Which is which?

Corvallis temperature

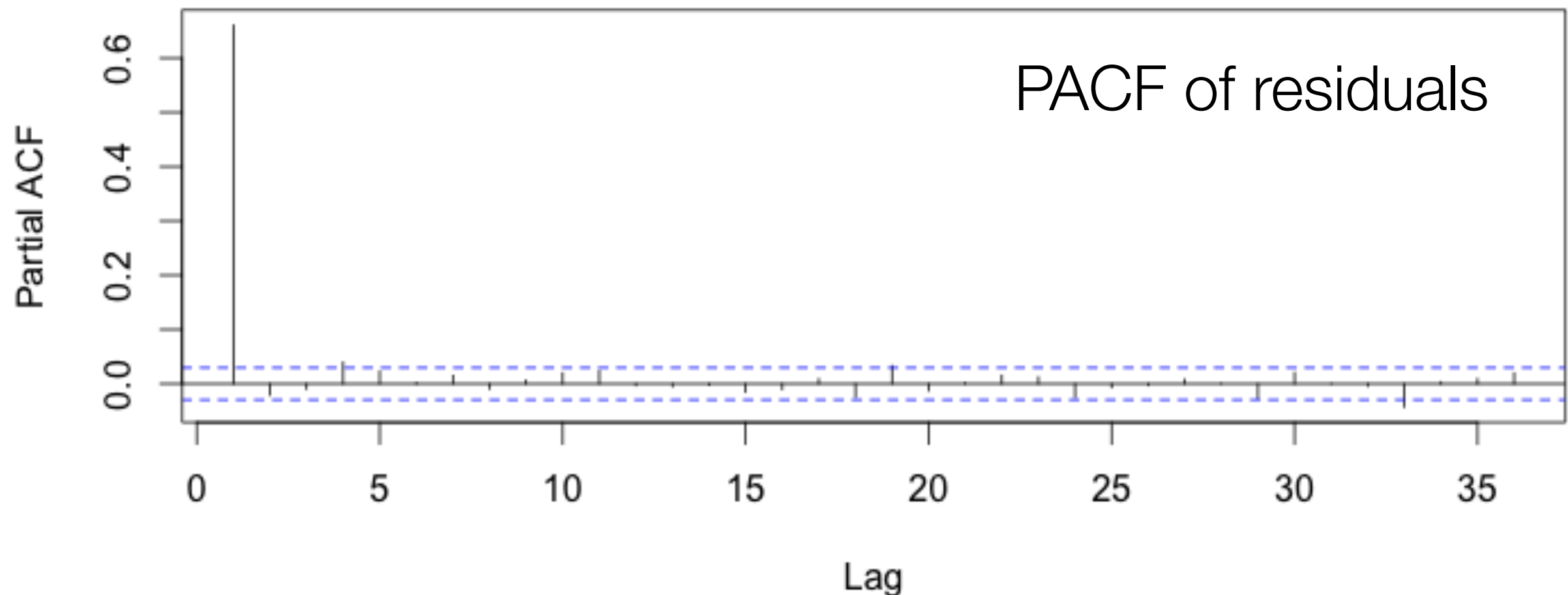
Series corv\$residual



Corvallis temperature

```
pacf(corv$residual, na.action = na.pass)
```

Series corv\$residual



AR(1) looks like a good model

How do we fit it? I.e. how do we estimate α_1 ?

$$\text{ARMA}(p, q)$$

Assume we know the order of our process, i.e. we know p and q .

How do we estimate the β s, α s, and σ^2 ?

Your turn

Look for...

Name three common approaches
to finding an estimate:

Hints:

M_{ethod} of M_{oment}

L_{east} S_{quare}

M_{aximum} L_{ikelihood}

The default in R
Has nice properties

$$M \text{ ————— } M \text{ ————— }$$

We have the theoretical ACF in terms of α , β and σ .

Set the theoretical ACF equal to our sample ACF and solve for the parameters.

So, we don't get too confused, let $\rho(h)$ denote the theoretical ACF, and $r(h)$ the sample ACF.

Corvallis temperature

Assume the residuals can be modelled by an AR(1).

$$\rho(1) = \phi \qquad \phi = \alpha, \text{ notation slip}$$
$$r(1) = \hat{\phi} = 0.6607$$

$$\text{AR}(p)$$

For $\text{AR}(p)$ processes we write down a recursion,

$$\rho(h) = \phi_1 \rho(h-1) + \dots + \phi_p \rho(h-p), \quad h = 1, \dots, p$$

$$\sigma^2 = \gamma(0) (1 - \phi_1 \rho(1) - \dots - \phi_p \rho(p))$$

The Yule-Walker equations

Derive Yule-Walker eqns

Yule-Walker Estimates

$$r(1) = \hat{\phi}_1 + \hat{\phi}_2 r(1) + \dots + \hat{\phi}_p r(p)$$

$$r(2) = \hat{\phi}_1 r(1) + \hat{\phi}_2 + \dots + \hat{\phi}_p r(p-2)$$

$$\vdots$$

$$r(p) = \hat{\phi}_1 r(p-1) + \hat{\phi}_2 r(p-2) + \dots + \hat{\phi}_p$$

A set of p equations in p unknowns, solve for $\hat{\phi}_1$ to $\hat{\phi}_p$.

$MA(q)$ and $ARMA(p,q)$

The method of moments approach gets complicated. End up with non-linear equations to be solved numerically.

The method of moments estimators have bad properties for MA and $ARMA$ processes anyway, so we'll leave it here.

Your turn

Remember linear regression?

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

We can find estimates for β_0, β_1 by minimizing

$$\sum_{i=1}^n \left(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i \right)^2$$

L _____ S _____

Consider the AR(1) process

$$X_t = \alpha_1 X_{t-1} + W_t$$

Define the residual, $e_t = x_t - \hat{\alpha}_1 x_{t-1}$

We could consider finding the $\hat{\alpha}_1$ that minimises the sum of squared residuals,

$$\sum_{t=2}^n (x_t - \hat{\alpha}_1 x_{t-1})^2 = \sum_{t=2}^n e_t^2$$

↑
since we don't see x_0

called
conditional
least squares

L_____ S_____ for MA

In general we can always define these residuals, but for MA and ARMA processes they are recursive. For example, MA(1)

$$e_1 = x_1 + \beta_1 e_0 \quad \text{assume } e_0 = 0$$

$$e_2 = x_2 + \beta_1 e_1$$

$$\vdots$$

$$e_n = x_n + \beta_1 e_{n-1}$$

L_____ S_____ in ARMA

For a general ARMA(p,q):

$$e_t = x_t - \alpha_1 x_{t-1} - \dots - \alpha_p x_{t-p} \\ + \beta_1 e_{t-1} + \dots + \beta_q e_{t-q}$$

And we have to set $e_p = \dots = e_{p+1-q} = 0$,
and sum starting at $t = p$, to avoid the x_t
we haven't observed

The minimization is done numerically

M _____ L _____

Assume a distribution for the white noise (usually Gaussian), then write the joint density function of our data as a function of the parameters, the likelihood,

$$L(\beta, \theta, \sigma^2) = f(x_1, x_2, \dots x_n; \beta, \theta, \sigma^2)$$

Find the parameters that maximise the likelihood.

For the non-statisticans: The joint density, f , tells us the probability of our data given certain parameter values. The likelihood, L , tells us how *likely* certain parameters are given our data (f and L are the same function, we just switch what we consider to be the variable). We estimate the parameters by choosing the most likely parameters given the data we saw.

M_____L_____

Assuming our white noise is Gaussian then the ARMA(p, q), x_t , process is also Gaussian and the likelihood is

$$x_t \sim \text{Normal}_n(\mathbf{0}, \Sigma)$$

where $\Sigma_{ij} = \text{Cov}(x_i, x_j) = \gamma(|i - j|)$

It's complicated, but there are general algorithms for maximizing it.

Maximum Likelihood

The way the function `arima` in R does it by default.

Nice asymptotic properties, deals with missing data easily.

Always lower variance than method of moments.

Fitting in R

```
arima(ts, order = c(p, d, q))
```

↑
ignore for now

Thursday

I'm out of town.

Chris will lead lecture.

Bring laptops!