

# Lista 1

## Modelagem com Apoio Computacional

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Carregamento dos dados:

```
fatigue_df <- data.frame(  
  work_MJ_m3 = c(  
    11.5, 13.0, 14.3, 15.6, 16.0, 17.3, 19.3, 21.1, 21.5, 22.6, 22.6, 24.0,  
    24.0, 24.6, 25.2, 25.5, 26.3, 27.9, 28.3, 28.4, 28.6, 30.9, 31.9, 34.5,  
    40.1, 40.1, 43.0, 44.1, 46.5, 47.3, 48.7, 52.9, 56.6, 59.9, 60.2,  
    60.3, 60.5, 62.1, 62.8, 66.5, 67.0, 67.1, 67.9, 68.8, 75.4, 100.5 ),  
  life_cycles = c(  
    3280, 5046, 1563, 4707, 977, 2834, 2266, 2208, 1040, 700, 1583, 482,  
    804, 1093, 1125, 884, 1300, 852, 580, 1066, 1114, 386, 745, 736,  
    750, 316, 456, 552, 355, 242, 190, 127, 185, 255, 195,  
    283, 212, 327, 373, 125, 187, 135, 245, 137, 200, 190))  
  
head(fatigue_df)
```

	work_MJ_m3	life_cycles
1	11.5	3280
2	13.0	5046
3	14.3	1563
4	15.6	4707
5	16.0	977
6	17.3	2834

Função para ajuste do modelo de moda (log-BS):

```

mle_mode_bs <- function(x, t) {
  x <- as.matrix(x)
  t <- as.matrix(t)

  fit_initial_log <- lm.fit(x, log(t))
  beta_initial_log_mode <- c(fit_initial_log$coef)

  k <- length(beta_initial_log_mode)
  n <- length(t)

  mu_initial_log <- x %*% beta_initial_log_mode
  alpha_initial_approx <- sqrt((4 / n) *
    sum((sinh((log(t) - mu_initial_log) / 2)) ^ 2))
  phi_initial <- alpha_initial_approx^2

  phi_initial <- max(0.01, min(0.99, phi_initial))

  thetaStar <- c(beta_initial_log_mode, phi_initial)

  loglik_mode <- function(par) {
    log_mode_Y <- x %*% par[1:k]
    phi_param <- par[k+1]
    if (phi_param <= 0 || phi_param >= 1) {
      return(NA)}

    alpha_orig <- sqrt(phi_param)
    beta_orig <- exp(log_mode_Y) / (1 - phi_param)
    terms <- log(t)

    l_i <- -log(alpha_orig) - 0.5 * log(2 * pi) - 0.5 * terms +
      log(sqrt(t / beta_orig) + sqrt(beta_orig / t)) -
      (1 / (2 * alpha_orig^2)) * (t / beta_orig + beta_orig / t - 2)

    if (any(!is.finite(l_i))) {
      return(.Machine$double.xmax)}

    return(-sum(l_i))}

  est <- optim(
    par = thetaStar,
    fn = loglik_mode,
    method = "BFGS",

```

```

    hessian = TRUE,
    control = list(maxit = 2000, reltol = 1e-12))

if (est$conv != 0) {
  warning("FUNCTION DID NOT CONVERGE!")
}

coef <- (est$par)[1:k]
phi_est <- est$par[k + 1]

mode_hat_log <- x %*% coef
mode_hat <- exp(mode_hat_log)

SHess = solve(est$hessian)
SE = sqrt(diag(SHess))

tval = est$par / SE
matcoef = cbind(est$par, SE, tval, 2 * (1 - pnorm(abs(tval))))

AIC <- 2 * est$value + 2 * (k + 1)
BIC <- 2 * est$value + (k + 1) * log(n)

result <- list(
  phiHat = phi_est,
  betaHat_log_mode = coef,
  modeHat = mode_hat,
  AIC = AIC,
  BIC = BIC,
  matcoef = matcoef
)

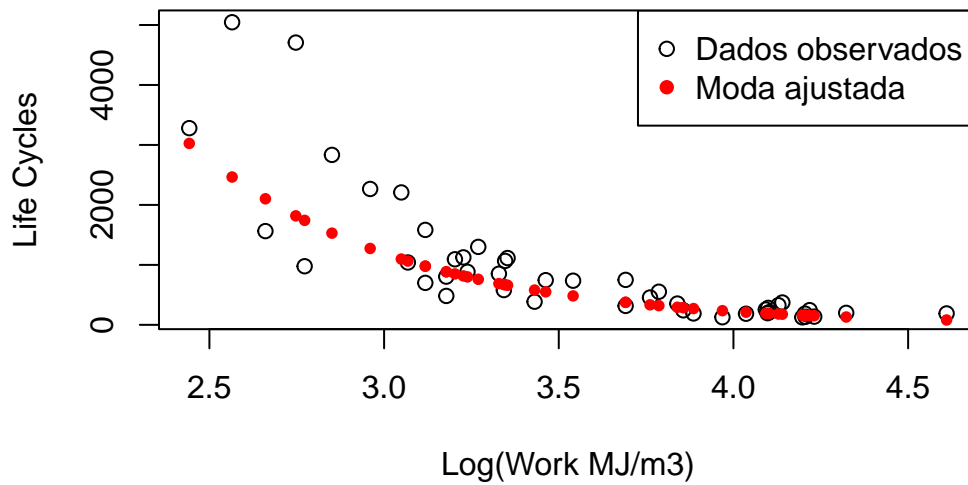
return(result)
}

```

Estimativas do modelo com base na moda:

		SE	tval
(Intercept)	12.095313	0.39169573	30.879358 0.00000e+00
log(fatigue_df\$work_MJ_m3)	-1.670763	0.10844480	-15.406574 0.00000e+00
	0.168396	0.03510972	4.796278 1.61641e-06

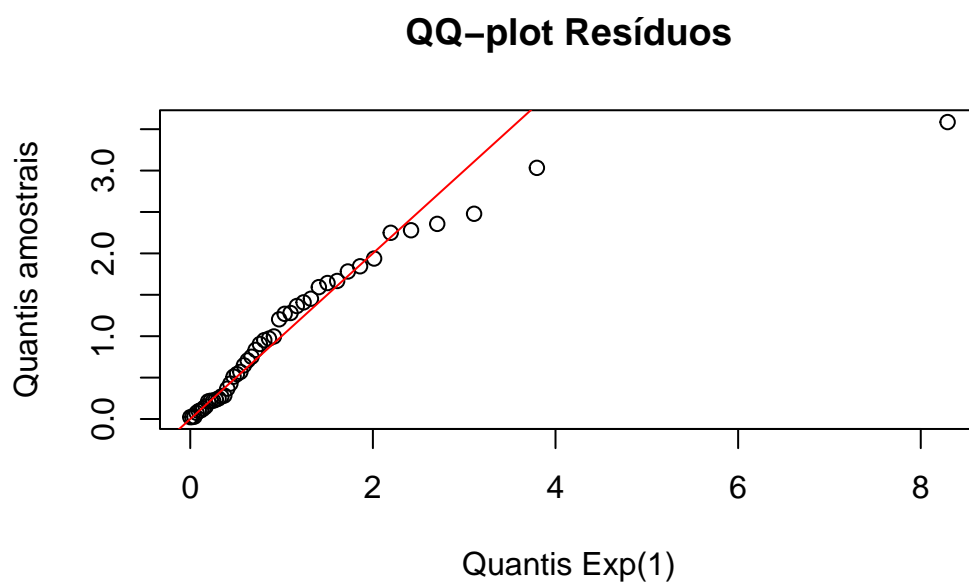
## Dados observados e moda ajustada



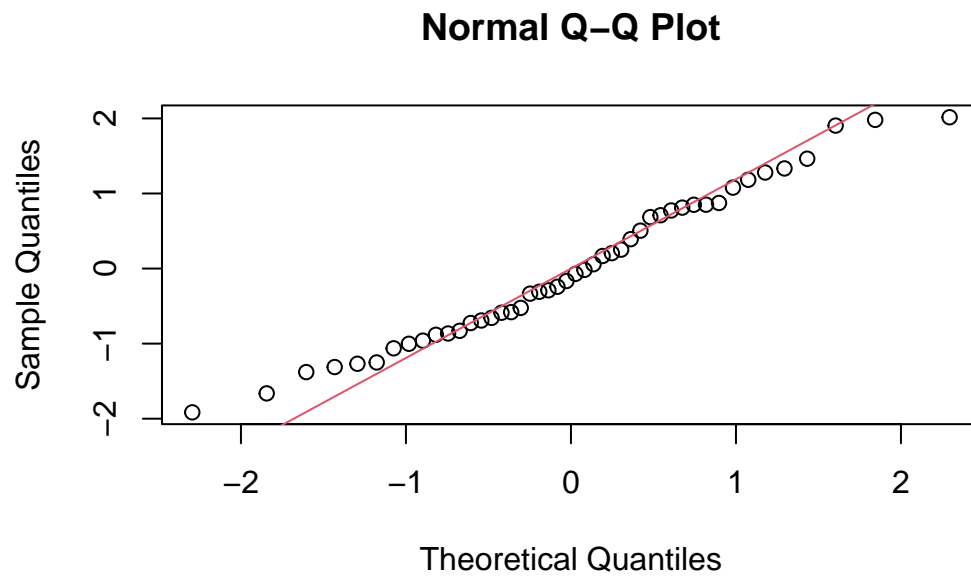
Análise do ajuste do modelo com base na moda:

```
S_bs_mode <- function(y, mu, phi) {  
  alpha <- sqrt(phi)  
  beta <- mu / (1 - phi)  
  xi <- (sqrt(y/beta) - sqrt(beta/y)) / alpha  
  S <- 1 - pnorm(xi)  
  return(S)  
}  
  
residuos_bs_mode <- function(y, mu_hat, phi_hat) {  
  S <- S_bs_mode(y, mu_hat, phi_hat)  
  r_cox <- -log(S)  
  r_quant <- qnorm(S)  
  list(coxsnell = r_cox, quantile = r_quant)  
}  
  
res_mode <- residuos_bs_mode(  
  y = fatigue_df$life_cycles,  
  mu_hat = fit_mode_bs$modeHat,  
  phi_hat = fit_mode_bs$phiHat  
)
```

```
# QQ plot
a <- ppoints(2000)
QGG <- qexp(a)
qqplot(QGG, res_mode$coxsnell,
       xlab = "Quantis Exp(1)", ylab = "Quantis amostrais",
       main = "QQ-plot Resíduos")
abline(0,1,col="red")
```



```
# QQ plot Quantílicos
qqnorm(res_mode$quantile); qqline(res_mode$quantile, col=2)
```



Ambos gráficos são bem similares aos gráficos para o modelo baseado na média.

**Comparação com o modelo com base na média:**

		se.coef	tval	
(Intercept)	12.2797340	0.3893978	31.535187	0
log(fatigue_df\$work_MJ_m3)	-1.6707690	0.1084439	-15.406766	0
	0.4103574	0.0427819	9.591846	0