

Lista 1

Modelagem com Apoio Computacional

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Carregamento dos dados

```
fatigue_df <- data.frame(work_MJ_m3 = c(11.5,13.0,14.3,15.6,16.0,17.3,19.3,
21.1,21.5,22.6,22.6,24.0,24.0,24.6,25.2,25.5,26.3,27.9,28.3,28.4,28.6,30.9,
31.9,34.5,40.1,40.1,43.0,44.1,46.5,47.3,48.7,52.9,56.6,59.9,60.2,60.3,60.5,
62.1,62.8,66.5,67.0,67.1,67.9,68.8,75.4,100.5),life_cycles = c(3280,5046,
1563,4707,977,2834,2266,2208,1040,700,1583,482, 804,1093,1125,884,1300,
852,580,1066,1114,386,745,736, 750,316,456,552,355,242,190,127,185,255,
195, 283,212,327,373,125,187,135,245,137,200,190))

head(fatigue_df)
```

	work_MJ_m3	life_cycles
1	11.5	3280
2	13.0	5046
3	14.3	1563
4	15.6	4707
5	16.0	977
6	17.3	2834

Ajuste do modelo com base na moda

```
mle_moda_bs <- function(x, t) {
  x <- as.matrix(x)
  t <- as.matrix(t)
```

```

fit <- lm.fit(x, log(t))
beta <- c(fit$coef)

k <- length(beta)
n <- length(t)

mu_init <- x %*% beta
alpha <- sqrt((4/n)*sum((sinh((log(t) - mu_init)/2))^2))

phi_init <- alpha^2 #chute inicial

thetaStar <- c(beta, phi_init)

loglik_moda <- function(par) {
  log_moda_Y <- x %*% par[1:k]
  phi_param <- par[k+1]
  if (phi_param <= 0 || phi_param >= 1) {
    return(NA)}

  alpha_orig <- sqrt(phi_param)
  beta_orig <- exp(log_moda_Y) / (1 - phi_param)
  terms <- log(t)

  #pdf da distribuição log-BS
  z <- (sqrt(t/beta_orig) - sqrt(beta_orig/t))/alpha_orig
  pdf_bs <- (1/(alpha_orig*sqrt(t)))*(sqrt(t/beta_orig) +
sqrt(beta_orig/t))*dnorm(z)
  pdf_bs[pdf_bs <= 0] <- .Machine$double.xmin
  #substitui valor menor ou igual a 0 por um número muito pequeno

  #log-verossimilhança
  l_i <- log(pdf_bs)
  if (any(!is.finite(l_i))) {
    return(.Machine$double.xmax)}
  #retorna um valor muito grande se algum log for infinito

  #verossimilhança total (loglik-plus)
  return(-sum(l_i))}

est <- optim(
  par = thetaStar,
  fn = loglik_moda,

```

```

method = "BFGS",
hessian = TRUE,
control = list(maxit = 2000, reltol = 1e-12))

if (est$conv != 0) {
  warning("FUNCTION DID NOT CONVERGE!")}

coef <- (est$par)[1:k]
phi_est <- est$par[k + 1]

moda_hat_log <- x %*% coef
moda_hat <- exp(moda_hat_log)

SHess = solve(est$hessian)
SE = sqrt(diag(SHess))

tval = est$par / SE
matcoef = cbind(est$par, SE, tval, 2 * (1 - pnorm(abs(tval))))

AIC <- 2 * est$value + 2 * (k + 1)
BIC <- 2 * est$value + (k + 1) * log(n)

result <- list(
  phiHat = phi_est,
  betaHat_log_moda = coef,
  modaHat = moda_hat,
  AIC = AIC,
  BIC = BIC,
  matcoef = matcoef)

return(result)}

```

Estimativas do modelo

```

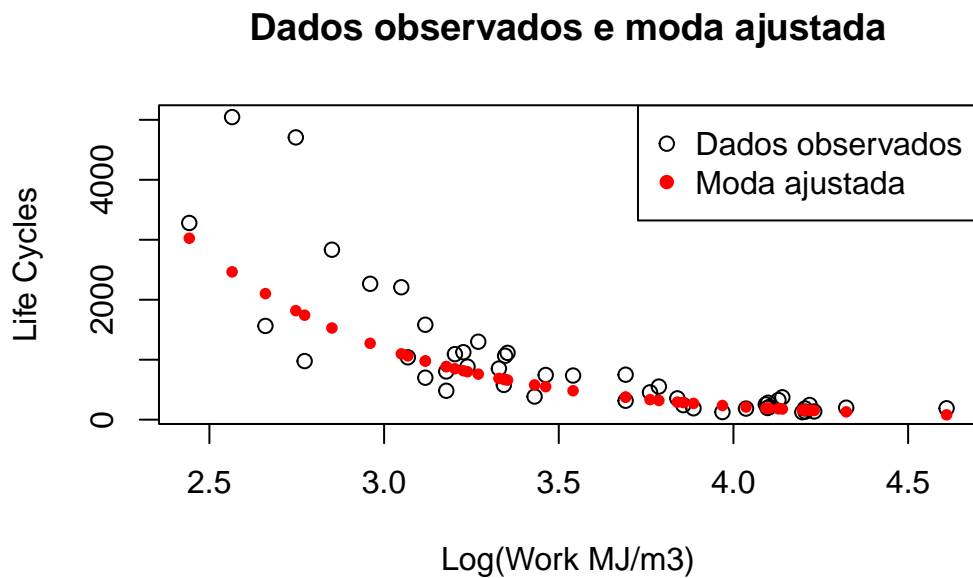
Zx <- model.matrix(~ log(fatigue_df$work_MJ_m3))
Zy <- fatigue_df$life_cycles

fit_moda <- mle_moda_bs(x = Zx , t = Zy)
fit_moda$matcoef

```

		SE	tval
(Intercept)	12.095313	0.39169573	30.879358 0.00000e+00
log(fatigue_df\$work_MJ_m3)	-1.670763	0.10844480	-15.406574 0.00000e+00
	0.168396	0.03510972	4.796278 1.61641e-06

```
plot(log(fatigue_df$work_MJ_m3), Zy,
     main = "Dados observados e moda ajustada",
     xlab = "Log(Work MJ/m3)", ylab = "Life Cycles")
points(log(fatigue_df$work_MJ_m3), fit_moda$modaHat,
       col = "red", pch = 16, cex = 0.8)
legend("topright",
      legend = c("Dados observados", "Moda ajustada"),
      col = c("black", "red"), pch = c(1, 16))
```



Verificação de ajuste do modelo

```
S_bs_moda <- function(y, mu, phi) {
  alpha <- sqrt(phi)
  beta <- mu / (1 - phi)
  xi <- (sqrt(y/beta) - sqrt(beta/y)) / alpha
```

```

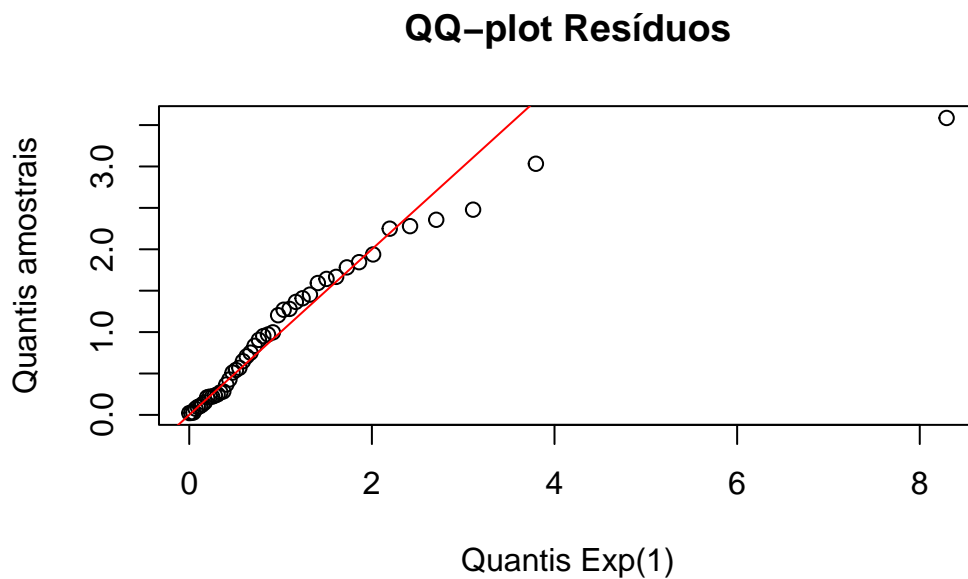
S <- 1 - pnorm(xi)
return(S)}

residuos_bs_moda <- function(y, mu_hat, phi_hat) {
  S <- S_bs_moda(y, mu_hat, phi_hat)
  r_cox <- -log(S)
  r_quant <- qnorm(S)
  list(coxsnell = r_cox, quantile = r_quant)}

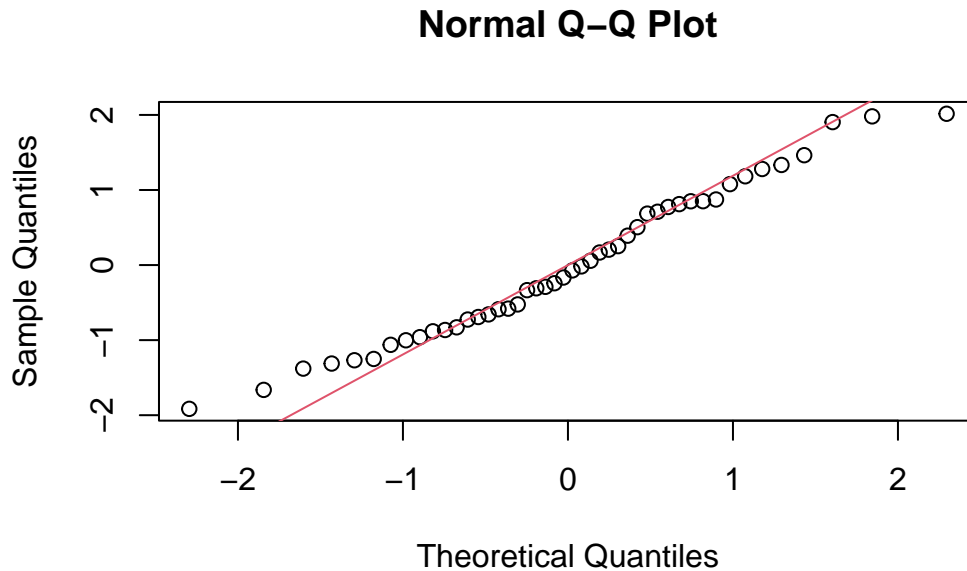
res_moda <- residuos_bs_moda(
  y = fatigue_df$life_cycles,
  mu_hat = fit_moda$modaHat,
  phi_hat = fit_moda$phiHat)

# QQ plot
a <- ppoints(2000)
QGG <- qexp(a)
qqplot(QGG, res_moda$coxsnell,
  xlab = "Quantis Exp(1)", ylab = "Quantis amostrais",
  main = "QQ-plot Resíduos")
abline(0,1,col="red")

```



```
# QQ plot quantílicos
qqnorm(res_moda$quantile); qqline(res_moda$quantile, col=2)
```



Os dois gráficos são bem similares aos gráficos para o modelo baseado na média. Indicando também adequação do modelo com base na moda.

Comparação dos modelos

Moda

		SE	tval	
(Intercept)	12.095313	0.39169573	30.879358	0.00000e+00
log(fatigue_df\$work_MJ_m3)	-1.670763	0.10844480	-15.406574	0.00000e+00
	0.168396	0.03510972	4.796278	1.61641e-06

Média

		se.coef	tval	
(Intercept)	12.2797340	0.3893978	31.535187	0
log(fatigue_df\$work_MJ_m3)	-1.6707690	0.1084439	-15.406766	0
	0.4103574	0.0427819	9.591846	0