### Outline

These lecture notes are based on the forthcoming book by Dudoit and van der Laan (2007).

Related articles and tech reports may be downloaded from Sandrine Dudoit's website

www.stat.berkeley.edu/~sandrine

and Mark van der Laan's website

www.stat.berkeley.edu/~laan.

# Outline: Part III. Applications to Genomics and Software Implementation

- Identification of Differentially Expressed and Co-Expressed Genes in High-Throughput Gene Expression Experiments [Chapter 9].
- Multiple Tests of Association with Biological Annotation Metadata [Chapter 10].
- HIV-1 Sequence Variation and Viral Replication Capacity [Chapter 11].
- Genetic Mapping of Complex Human Traits Using Single Nucleotide Polymorphisms: The ObeLinks Project [Chapter] [12].
- Software Implementation [Chapter 13].

### High-Throughput Gene Expression Experiments

- Identification of differentially expressed (DE) genes, i.e., genes whose expression measures are associated with possibly censored biological and clinical covariates and outcomes. Simultaneous test, for each gene, of the null hypothesis of no association between the expression measures and the covariates and outcomes.
- Identification of co-expressed (CE) genes, i.e., pairs (or sets) of genes with associated expression measures. Simultaneous test, for each gene pair, of the null hypothesis of no association (e.g., zero correlation) between their expression measures.

The multiple testing results can be used as a basis for further analyses, e.g., clustering and graph theoretical analyses.

Callow et al. (2000). Apo AI dataset. The Apo AI microarray experiment was carried out as part of a study of lipid metabolism and artherosclerosis susceptibility in mice.

Apolipoprotein AI (Apo AI) is a gene known to play a pivotal role in high-density lipoprotein (HDL) cholesterol metabolism and mice with the Apo AI gene knocked-out have very low HDL cholesterol levels.

The goal of the experiment was to identify differentially expressed genes in the livers of Apo AI knock-out mice compared to inbred control mice.

The treatment group consists of 8 inbred C57Bl/6 mice with the Apo AI gene knocked-out and the control group consists of 8 inbred C57Bl/6 mice.

For each of the 16 mice, target samples of complementary DNA (cDNA) were obtained from messenger RNA (mRNA) by reverse transcription and labeled using the red-fluorescent dye Cyanine 5 (Cy5).

The reference sample used in all hybridizations was prepared by pooling cDNA from the 8 control mice and was labeled with the green-fluorescent dye Cyanine 3 (Cy3).

Combined red- and green-labeled target cDNA samples were hybridized to microarrays with 6,384 spots (=  $4 \times 4 \times 19 \times 21$ ), including 257 probe sequences related to lipid metabolism.

Pre-processing. Each of the 16 hybridizations produced a pair of 16-bit TIFF images, which were processed by seeded region growing segmentation and morphological opening background adjustment (Yang et al. (2002); R package Spot, experimental.act.cmis.csiro.au/Spot/index.php).

The resulting fluorescence intensity measures were normalized by within-print-tip-group loess robust local regression (Dudoit et al. (2002); Dudoit and Yang (2003); Yang et al. (2001); Yang and Paquet (2005); Bioconductor R package marray, www.bioconductor.org).

Among the 6,384 spots on the microarray, only those 5,548 spots corresponding to actual cDNA sequences are retained for subsequent analyses. The other 836 spots are either blank or control spots.

Data. The data for each of the n = 16 mice consist of the following.

- Z, a binary covariate/genotype (1 for treatment vs. 0 for control).
- X = (X(m) : m = 1, ..., M), an M = 5,548-dimensional outcome/phenotype vector of microarray expression measures.

The R package ApoAl provides microarray data objects, at various levels of processing, for the Apo AI experiment (www.stat.berkeley.edu/~sandrine/MTBook).

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Multiple testing question. Identify differentially expressed genes by testing, for each of the M = 5,548 probes, whether there is a difference in mean expression measures between knock-out and control mice.

Parameters of interest. For the mth probe, the parameter of interest is the difference in mean expression measures  $\psi(m)$ between treatment and control mice, that is,

$$\psi(m) = E[X(m)|Z=1] - E[X(m)|Z=0], \qquad m = 1, \dots, M.$$

Null hypotheses. Consider two-sided tests of the M null hypotheses of no differences in mean expression measures vs. the alternative hypotheses of different mean expression measures,

$$H_0(m) = I(\psi(m) = 0)$$
 vs.  $H_1(m) = I(\psi(m) \neq 0)$ .

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Test statistics. Two-sample Welch t-statistics,

$$T_n(m) = \frac{\psi_n(m) - \psi_0(m)}{\sigma_n(m)} = \frac{\bar{X}_1(m) - \bar{X}_0(m) - 0}{\sqrt{\frac{\sigma_{0,n}^2(m)}{n_0(m)} + \frac{\sigma_{1,n}^2(m)}{n_1(m)}}},$$

where the null values  $\psi_0(m)$  are zero and  $n_k(m) = \sum_i \mathrm{I}(Z_i = k)$ ,  $\bar{X}_k(m) = \sum_i \mathrm{I}(Z_i = k) \, X_i(m) / n_k(m)$ , and  $\sigma_{k,n}^2(m) = \sum_i \mathrm{I}(Z_i = k) \, (X_i(m) - \bar{X}_k(m))^2 / (n_k(m) - 1)$  denote, respectively, the sample sizes, sample means, and sample variances for the expression measures of probe m in treatment (k = 1) and control (k = 0) mice (note that the sample sizes  $n_k(m)$  may differ across probes m due to missing data).

Test statistics null distribution. Non-parametric bootstrap estimator of the null shift and scale-transformed test statistics null distribution, B = 5,000 samples.

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### Multiple testing procedures.

- FWER control: Joint single-step maxT procedure (SS maxT) and minP procedure (SS minP), joint step-down maxT procedure (SD maxT) and minP procedure (SD minP), marginal single-step Bonferroni (1936) procedure (SS Bonferroni), marginal step-down Holm (1979) procedure (SD Holm), and marginal step-up Hochberg (1988) procedure (SU Hochberg).
- gFWER control: qFWER(k)-controlling augmentation multiple testing procedure, based on FWER-controlling joint single-step maxT procedure, for an allowed number  $k \in \{5, 10, 50, 100\}$  of false positives (gFWER(k) AMTP).
- TPPFP control: TPPFP(q)-controlling augmentation multiple testing procedure, based on FWER-controlling joint single-step maxT procedure, for an allowed proportion  $q \in \{0.05, 0.10, 0.25, 0.50\}$  of false positives (TPPFP(q) AMTP).
- FDR control: Marginal step-up Benjamini and Hochberg (1995) procedure (SU BH) and Benjamini and Yekutieli (2001) procedure (SU BY), TPPFP-based MTPs of van der Laan et al. (2004b) (TPPFP-based 1, TPPFP-based 2).

All multiple testing procedures single out 8 probes, out of 5,548 spotted probe sequences, as being differentially expressed between knock-out and control mice.

The negative t-statistics suggest that the probes are under-expressed in the Apo AI knock-out mice compared to the control mice.

The 8 most extreme probes actually correspond to only 4 distinct genes and 1 EST: ApoAl (2 copies), ApoCIII (2 copies), Steroldesaturase (2 copies), Catechol - O - methyltransferase (1 copy), and a novel EST (1 copy).

All changes were confirmed by real-time quantitative PCR (RT-PCR), as described in Callow et al. (2000).

Hyperlinked Supplementary Table 9.1.

The presence of ApoAl among the under-expressed genes is to be expected, as this is the gene that was knocked out in the treatment mice.

The ApoCIII gene, also associated with lipoprotein metabolism, is located very close to the ApoAl locus. Callow et al. (2000) showed that the down-regulation of ApoCIII is actually due to genetic polymorphism rather than lack of ApoAl. The presence of ApoAl and ApoCIII among the under-expressed genes thus provides a validation of the statistical methodology, if not a biologically novel finding.

Steroldesaturase is an enzyme that catalyzes one of the terminal steps in cholesterol synthesis.

Liver membrane-bound Catechol - O - methyltransferase was found to be a relevant factor in blood pressure regulation in rats (Tsunoda et al., 2003).

The novel EST shares sequence similarity to a family of ATPases.

The Apo AI experiment is rather unusual, in the sense that 8 spotted probe sequences clearly stand out from the remaining 5,540 probes as being differentially expressed.

Such a dichotomy in gene expression is seldom observed in other applications of the microarray technology.

For example, in many cancer microarray studies, genes tend to exhibit a continuum of change in expression measures and it is difficult to identify distinct groups of genes.

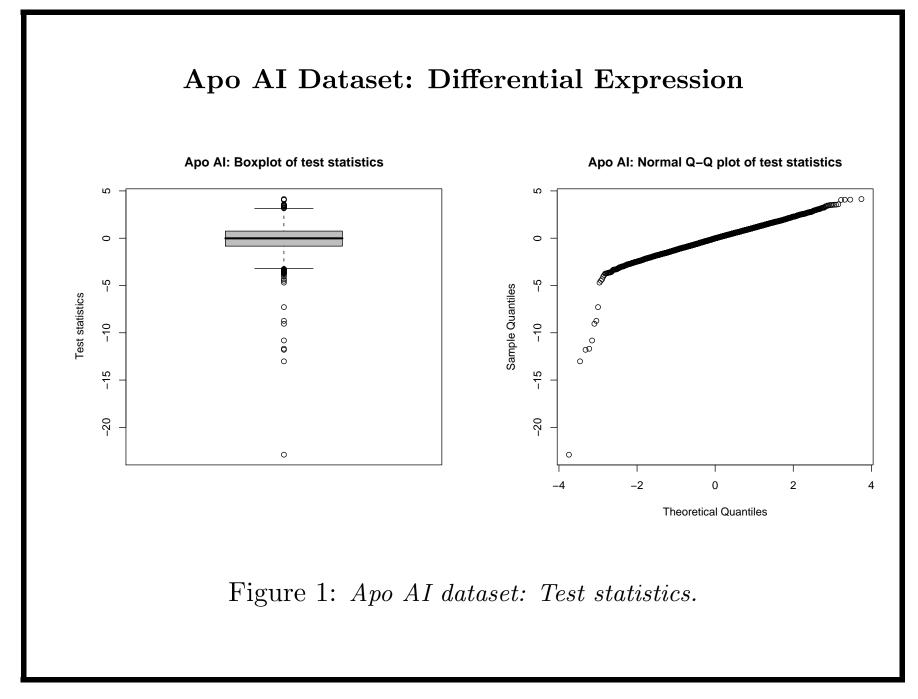
Differences in patterns of differential expression likely reflect the nature of the target samples under investigation.

The Apo AI experiment compares relatively pure cell samples (hepatocytes), from wild-type and knock-out mice with an otherwise identical genetic background. In contrast, human cancer microarray studies typically assay samples composed of a variety of cell types, from genetically diverse individuals.

Table 1: Apo AI dataset: FWER-controlling non-parametric bootstrap $based\ single-step\ maxT\ MTP.$ 

$\mathbf{Gene}$	${\bf Spot}$	Estimate	$t extsf{-statistic}$	${f Unadjusted}$	${\bf Adjusted}$
name	ID			p-value	$p extsf{-value}$
		$\psi_{n}(m)$	$T_{n}\left( m ight)$	$P_{0n}(m)$	$\widetilde{P}_{0n}(m)$
1 Apoa1	2149	-3.17	-22.89	0.0000	0.0014
ApoAl					
Apo AI,	lipid-Img				
1					
Sc5d	4139	-1.03	-13.02	0.0000	0.0136
Steroldesati	ırase				
EST, We	akly simi	lar to C-5 ST	EROL DESATU	URASE [Saccharomy	ces cerevisiae], lipid-UC
1					
Comt	5356	-1.86	-11.80	0.0000	0.0214
Catechol —	O — methy	ltransferase			
CATECH	HOL O-M	ETHYLTRAN	SFERASE, ME	MBRANE-BOUND	FORM, Brain-Img
1					
Apoa1	540	-3.05	-11.69	0.0000	0.0224
ApoAl					
EST. His	ghly simil	ar to APOLIP	OPROTEIN A	-I PRECURSOR [M <sub>1</sub>	us musculus], lipid-UG
				L	

Gene	${\bf Spot}$	Estimate	$t extsf{-statistic}$	${f Unadjusted}$	${f Adjusted}$	
name	ID			$p extsf{-} extbf{value}$	$p extsf{-} extbf{value}$	
		$\psi_{n}\left( m ight)$	$T_{n}\left( m ight)$	$P_{0n}(m)$	$\widetilde{P}_{0n}(m)$	
2 Apoc3	1739	-0.96	-10.81	0.0002	0.0322	
ApoCIII						
Apo CIII, lipid-Img						
1 EST	1496	-0.99	-9.05	0.0000	0.0606	
2 Apoc3 ApoCIII	2537	-1.02	-8.74	0.0002	0.0694	
ESTs, Highly similar to APOLIPOPROTEIN C-III PRECURSOR [Mus musculus], lipid-UG						
$^2$ Sc5d	4941	-0.97	-7.29	0.0002	0.1346	
Steroldesaturase						
similar to yeast sterol desaturase, lipid-Img						
$4_{\sf Casp7}$	954	-0.31	-4.70	0.0018	0.4840	
Caspase7						



Apo Al: Unadjusted and adjusted p-values vs. test statistics FWER-controlling non-parametric bootstrap-based single-step maxT MTP

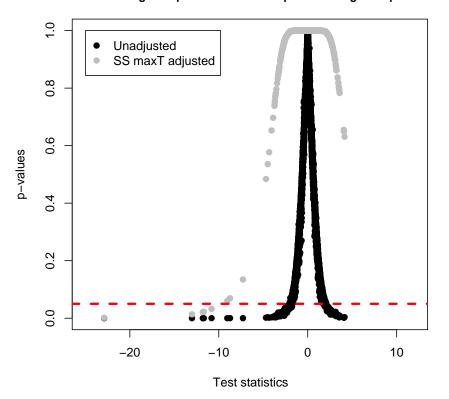


Figure 2: Apo AI dataset: FWER-controlling non-parametric bootstrapbased single-step maxT MTP, test statistics and p-values.

Apo Al: Test statistics and cut-offs for top 10 probes FWER-controlling non-parametric bootstrap-based single-step maxT MTP

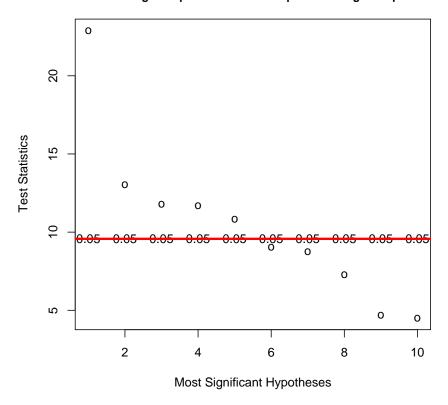


Figure 3: Apo AI dataset: FWER-controlling non-parametric bootstrapbased single-step maxT MTP, test statistics and cut-offs.

Apo AI: Parameter estimates and confidence regions for top 10 probes FWER-controlling non-parametric bootstrap-based single-step maxT MTP

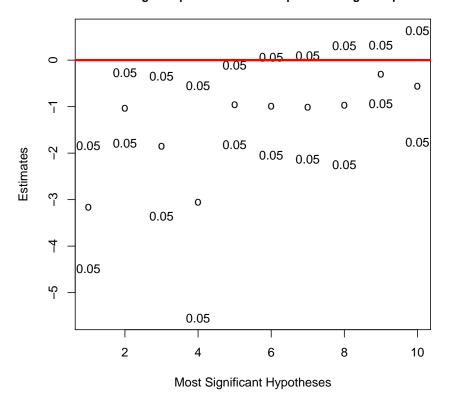


Figure 4: Apo AI dataset: FWER-controlling non-parametric bootstrapbased single-step maxT MTP, parameter estimates and confidence regions.

#### Apo Al: FWER-controlling non-parametric bootstrap-based MTPs

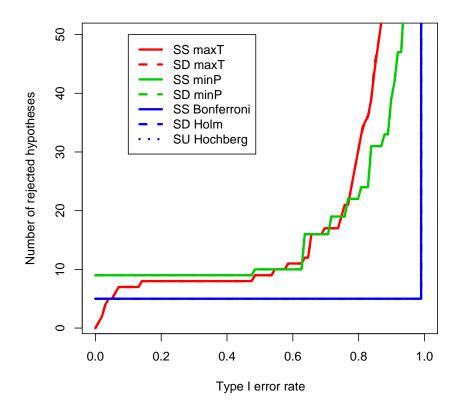
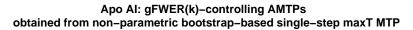


Figure 5: Apo AI dataset: FWER-controlling non-parametric bootstrapbased MTPs.



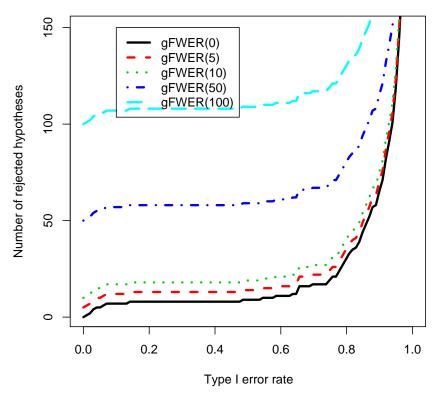
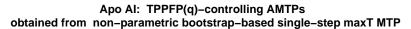


Figure 6: Apo AI dataset: gFWER-controlling non-parametric bootstrapbased AMTPs.



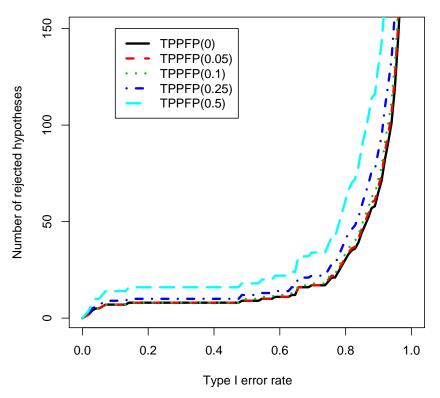


Figure 7: Apo AI dataset: TPPFP-controlling non-parametric bootstrapbased AMTPs.

#### Apo Al: FDR-controlling non-parametric bootstrap-based MTPs

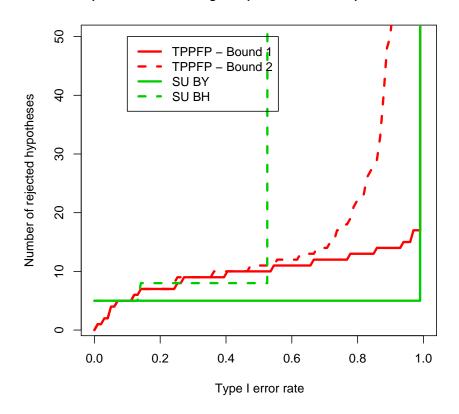


Figure 8: Apo AI dataset: FDR-controlling non-parametric bootstrapbased MTPs.

In addition to passing genetic messages from DNA to the protein-making machinery of the cell, ribonucleic acids (RNA) serve many other cellular functions.

microRNAs (miRNA) are small, non-coding RNAs involved in gene regulation and developmental timing (microrna.sanger.ac.uk).

By binding to messenger RNA (mRNA), miRNAs regulate gene expression post-transcriptionally and affect the abundance of a wide range of proteins, in diverse biological processes.

By now, hundreds of miRNAs have been identified, in various multicellular organisms, including the fruitfly *Drosophila* melanogaster and humans, and many are evolutionary conserved.

Although the biological functions of miRNAs are still largely unknown, miRNAs are predicted to regulate up to 30% of all protein-coding genes.

Each mammalian miRNA is believed to regulate approximately 200 genes and many genes have several target sites for one or several different miRNAs.

The large number of miRNA genes, their diverse expression patterns, and the abundance of miRNA targets, suggest the involvement of miRNAs in a variety of diseases, including cancers and viruses.

More than half of the known human miRNA genes are located in genomic regions related to cancers, such as, fragile sites, minimal regions of loss of heterozigosity, minimal regions of amplification, and common breakpoint regions.

miRNAs have also been implicated in several mammalian viruses, such as, the Epstein-Barr virus and the human immunodeficiency virus (HIV).

Lu et al. (2005). Cancer microRNA dataset. Lu et al. (2005) used a bead-based flow cytometric profiling method to measure the levels of 217 known human miRNAs in n = 186 cell samples derived from cancerous and non-cancerous tissues.

These authors found that predictors based on miRNA expression measures are better able to distinguish developmental lineage, differentiation state, and cancer state, than the best corresponding predictors based on genome-wide mRNA expression measures from the same cells.

Pre-processing. log<sub>2</sub>-transform; exclude cell lines; exclude any miRNA with expression measures below a detection threshold of  $\log_2 32 = 5$  in more than half of the n = 186 samples.

Data. The data for each of the n = 186 samples consist of the following.

- Y, a binary cancer status outcome/phenotype (1 for cancerous vs. 0 for non-cancerous).
- X = (X(j) : j = 1, ..., J), a J = 155-dimensional covariate vector of real-valued miRNA expression measures.
- W, a 19-dimensional tissue type indicator vector confounding variable.

The data are available at www.broad.mit.edu/cancer/pub/miGCM.

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### Goals.

• Identify differentially expressed miRNAs, i.e., miRNAs whose expression measures are associated with cancer status (cancerous vs. non-cancerous).

Our approach is based on tests for regression coefficients in logistic models that relate cancer status Y to miRNA expression measures X(j), while adjusting for the confounding variable tissue type W.

• Identify co-expressed miRNAs, i.e., pairs (groups) of miRNAs with correlated expression profiles across tissue samples.

Logistic regression model. For the jth miRNA, fit a logistic regression model that includes expression measure X(j) and tissue type W as covariates,

$$logit(E[Y|X(j), W]) = \alpha(j) + \beta(j)X(j) + \gamma(j)W, \qquad j = 1, \dots, J,$$

where logit(z) = log(z/(1-z)) is the logit function,

 $\alpha(j)$  a baseline effect parameter,

 $\beta(j)$  a main effect parameter for the expression measure X(j) of the jth miRNA, and

 $\gamma(j)$  a miRNA-specific 19-dimensional parameter vector adjusting for tissue type W.

Parameters of interest. For the jth miRNA, the parameter of interest is the logistic regression coefficient  $\beta(j)$  for the expression measure X(j).

Null hypotheses. Consider two-sided tests of the J null hypotheses of no association between the expression measures X(j) and cancer status Y,

$$H_0(j) = I(\beta(j) = 0)$$
 vs.  $H_1(j) = I(\beta(j) \neq 0)$ .

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Test statistics. t-statistics for logistic regression coefficients,

$$T_n(j) = \frac{\beta_n(j) - \beta_0(j)}{\sigma_n(j)},$$

where the null values  $\beta_0(j)$  are zero and  $\beta_n(j)$  are logistic regression parameter estimators with estimated standard errors  $\sigma_n(j)$  (as implemented in the function glm from the R package stats, with the call glm(Y ~ X(j) + W, family="binomial"), using the binomial family and iteratively reweighted least squares (IWLS)).

Test statistics null distribution. Non-parametric bootstrap estimator of the null shift and scale-transformed test statistics null distribution, B = 5,000 samples.

Multiple testing procedures. FWER-controlling single-step maxT procedure.

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Table 2: Cancer miRNA dataset, differential expression: Tests for logistic regression coefficients. Number of differentially expressed miRNAs (out of J = 155) for different nominal FWER levels  $\alpha$ .

Nominal FWER, $\alpha$	Number of miRNAs, $R_n$
0.05	90~(58%)
0.01	53 (34%)

All 90 miRNAs that are significantly differentially expressed at level  $\alpha = 0.05$  have negative test statistics  $(T_n(j) < -3.8)$ , suggesting under-expression in cancerous compared to non-cancerous tissues.

Table 3: Cancer miRNA dataset, differential expression: Tests for logistic regression coefficients. 53 most significantly differentially expressed miRNAs between cancerous and non-cancerous tissues (bootstrap-based single-step maxT adjusted p-values < 0.01).

Name	miRNA target sequence	Adjusted $p$ -value	Test statistic
hsa — miR — 98	UGAGGUAGUAAGUUGUAUUGUU	0.0038	-4.88
hsa — miR — 28	AAGGAGCUCACAGUCUAUUGAG	0.0038	-4.79
hsa-miR-196	UAGGUAGUUUCAUGUUGUUGG	0.0038	-4.79
hsa — miR — 30a	CUUUCAGUCGGAUGUUUGCAGC	0.0038	-4.78
hsa — miR — 30e	UGUAAACAUCCUUGACUGGA	0.0038	-4.78
hsa — miR — 99a#	AACCCGUAGAUCCGAUCUUGUG	0.0038	-4.77
hsa — miR — 335	UCAAGAGCAAUAACGAAAAAUGU	0.0038	-4.72
hsa — let — 7e	UGAGGUAGGAGGUUGUAUAGU	0.0038	-4.69
hsa $-$ miR $-$ 23b#	AUCACAUUGCCAGGGAUUACCAC	0.0038	-4.67
hsa - miR - 214	ACAGCAGGCACAGACAGGCAG	0.0038	-4.67
hsa — miR — 99b	CACCCGUAGAACCGACCUUGCG	0.0038	-4.67
hsa — miR — 30c	UGUAAACAUCCUACACUCUCAGC	0.0038	-4.66
hsa - miR - 30b	UGUAAACAUCCUACACUCAGC	0.0038	-4.66
hsa — miR — 338	UCCAGCAUCAGUGAUUUUGUUGA	0.0038	-4.65

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Name	miRNA target sequence	Adjusted $p$ -value	Test stat stic
hsa — miR — 103	AGCAGCAUUGUACAGGGCUAUGA	0.0038	-4.64
hsa - miR - 185	UGGAGAGAAAGGCAGUUC	0.0038	-4.63
hsa $-$ miR $-$ 151 $\ast$	UCGAGGAGCUCACAGUCUAGUA	0.0038	-4.62
hsa $-$ miR $-$ 100 $\#$	AACCCGUAGAUCCGAACUUGUG	0.0038	-4.61
$hsa - miR - 20\_(sub\_1)$	UAAAGUGCUUAUAGUGCAGGUAG	0.0038	-4.61
hsa - miR - 129*	AAGCCCUUACCCCAAAAAGCAU	0.0038	-4.60
hsa $-$ miR $-$ 22 $\#$	AAGCUGCCAGUUGAAGAACUGU	0.0038	-4.60
hsa $-$ let $-$ 7d $\#$	AGAGGUAGUAGGUUGCAUAGU	0.0038	-4.58
hsa - miR - 107	AGCAGCAUUGUACAGGGCUAUCA	0.0038	-4.58
rno - miR - 352	AGAGUAGUAGGUUGCAUAGUA	0.0038	-4.58
hsa - miR - 197	UUCACCACCUUCUCCACCCAGC	0.0038	-4.57
hsa - miR - 32	UAUUGCACAUUACUAAGUUGC	0.0038	-4.57
hsa - miR - 342	UCUCACACAGAAAUCGCACCCGUC	0.0038	-4.56
hsa - miR - 324 - 5p	CGCAUCCCCUAGGGCAUUGGUGU	0.0038	-4.51
hsa - miR - 128b	UCACAGUGAACCGGUCUCUUUC	0.0038	-4.51
hsa - miR - 126*	CAUUAUUACUUUUGGUACGCG	0.0038	-4.50
hsa - miR - 19b	UGUGCAAAUCCAUGCAAAACUGA	0.0038	-4.49
$hsa - miR - 151\_(sub\_1)$	ACUAGACUGAGGCUCCUUGAGG	0.0038	-4.49
hsa — miR — 199a*	UACAGUAGUCUGCACAUUGGUU	0.0038	-4.48
hsa — let — 7i	UGAGGUAGUUUGUGCU	0.0038	-4.48
hsa - miR - 10b	UACCCUGUAGAACCGAAUUUGU	0.0038	-4.47
miR - 292 - 3p	AAGUGCCGCCAGGUUUUGAGUGU	0.0040	-4.46

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Name	miRNA target sequence	Adjusted $p$ -value	Test statistic
hsa — miR — 136	ACUCCAUUUGUUUUGAUGAUGGA	0.0042	-4.45
mmu - miR - 10b	CCCUGUAGAACCGAAUUUGUGU	0.0042	-4.45
hsa - let - 7f	UGAGGUAGUAGAUUGUAUAGUU	0.0042	-4.44
hsa - miR - 302	UAAGUGCUUCCAUGUUUUGGUGA	0.0042	-4.43
mmu $-$ let $-$ 7g	UGAGGUAGUAGUUUGUACAGU	0.0042	-4.43
hsa - miR - 10a	UACCCUGUAGAUCCGAAUUUGUG	0.0042	-4.42
hsa — miR — 34b	AGGCAGUGUCAUUAGCUGAUUG	0.0042	-4.42
hsa — miR — 92	UAUUGCACUUGUCCCGGCCUGU	0.0042	-4.42
hsa - miR - 101	UACAGUACUGUGAUAACUGAAG	0.0044	-4.38
hsa - miR - 16	UAGCAGCACGUAAAUAUUGGCG	0.0046	-4.37
mmu — miR — 339	UCCCUGUCCUCCAGGAGCUCA	0.0046	-4.37
hsa - miR - 19a	UGUGCAAAUCUAUGCAAAACUGA	0.0046	-4.37
hsa - miR - 152	UCAGUGCAUGACAGAACUUGG	0.0052	-4.35
hsa — miR — 23a	AUCACAUUGCCAGGGAUUUCC	0.0052	-4.34
hsa - miR - 186	CAAAGAAUUCUCCUUUUGGGCUU	0.0072	-4.30
rno - miR - 343	UCUCCCUCCGUGUGCCCAGU	0.0096	-4.29
hsa - miR - 140	AGUGGUUUUACCCUAUGGUAG	0.0096	-4.28

# Located in minimal deleted regions, minimal amplified regions, and breakpoint regions involved in human cancers (Calin et al., 2004).

Our findings are in agreement with the original publication of Lu et al. (2005), the main distinctions being that the single-step maxT procedure takes into account the joint distribution of the test statistics and the logistic regression model allows adjusting for the confounding variable tissue type when comparing expression measures between cancerous and non-cancerous tissues.

Five of the highly significant miRNAs are located in minimal deleted regions, minimal amplified regions, and breakpoint regions involved in human cancers (Calin et al., 2004). Specifically, hsa - let - 7d and hsa - miR - 23b have been associated with urothelial cancer; hsa - miR - 22 with hepatocellular cancer; hsa - miR - 99a with lung cancer; hsa - miR - 100 with breast, cervical, lung, and ovarian cancers.

It would be of interest, as a follow-up analysis, to examine the target sequences of the differentially expressed miRNAs for the potential identification of common motifs.

Parameters of interest.  $M = J(J-1)/2 = 155 \times 154/2 = 11,935$  correlation coefficients for the expression measures of distinct pairs (j, j') of miRNAs,

$$\rho(j, j') = \text{Cor}[X(j), X(j')], \qquad j = 1, \dots, J - 1, \ j' = j + 1, \dots, J.$$

Null hypotheses. Consider two-sided tests of the M null hypotheses of no correlation in expression measures between pairs of miRNAs,

$$H_0(j, j') = I(\rho(j, j') = 0)$$
 vs.  $H_1(j, j') = I(\rho(j, j') \neq 0)$ .

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Test statistics. Difference statistics,

$$T_n(j, j') = \sqrt{n}(\rho_n(j, j') - \rho_0(j, j')),$$

where the null values  $\rho_0(j,j')$  are zero and  $\rho_n(j,j')$  are empirical correlation coefficients.

Test statistics null distribution. Non-parametric bootstrap estimator of the null shift and scale-transformed test statistics null distribution, B = 5,000 samples.

Multiple testing procedures. FWER-controlling single-step maxT procedure.

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Table 4: Cancer miRNA dataset, co-expression: Tests for correlation coefficients. Number of co-expressed miRNA pairs (out of M=11,935) for different nominal FWER levels  $\alpha$ .

Nominal FWER, $\alpha$	Number of miRNA pairs, $R_n$
0.05	8,916 (75%)
0.01	7,479~(63%)

Correlation coefficients found to be significantly different from zero at nominal FWER level  $\alpha = 0.05$  range from 0.26 to 0.99, with median value 0.55.

Only 8% of all pairwise correlation coefficients are negative and none significantly so.

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Table 5: Cancer miRNA dataset, co-expression: Tests for correlation coefficients. Twenty most significantly co-expressed pairs of miRNAs (bootstrap-based single-step maxT MTP).

Names		Correlation coefficient
hsa — miR — 106a#	hsa $-$ miR $-$ 17 $-$ 5p#	0.99
mmu - miR - 200b	hsa - miR - 200b	0.99
mmu - miR - 200b	hsa - miR - 200c	0.99
hsa $-$ miR $-$ 107 $\dagger$	hsa - miR - 103	0.99
hsa - miR - 200b	hsa - miR - 200c	0.99
hsa $-$ miR $-$ 145 $\ddagger$	$hsa - miR - 143 \ddagger$	0.98
$hsa - miR - 199a \_(sub \_1)$	mmu - miR - 199b	0.98
hsa $-$ miR $-$ 17 $-$ 5p	$hsa - miR - 20\_(sub\_1)$	0.97
hsa $-$ miR $-$ 19a $\#$	hsa $-$ miR $-$ 19b#	0.97
hsa — miR — 29a	hsa $-$ miR $-$ 30a $*$	0.97
hsa $-$ miR $-$ 181a	hsa - miR - 181c	0.97
$hsa - miR - 199a\_(sub\_1)$	hsa $-$ miR $-$ 199a $\ast$	0.97
hsa — miR — 29b_(sub_2)	hsa — miR — 29c	0.97

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Names		Correlation coefficient
hsa — miR — 199a*	mmu — miR — 199b	0.96
hsa — miR — 200a	hsa $-$ miR $-$ 141	0.96
$hsa - miR - 20\_(sub\_1) \#$	mmu — miR — $106a$	0.96
hsa $-$ miR $-$ 106a	$hsa - miR - 20\_(sub\_1) \#$	0.96
hsa — miR — 200a	hsa — miR — 200a	0.96
hsa — miR — 23b	hsa — miR — 23a	0.96
hsa $-$ miR $-$ 10a	hsa-miR-10b	0.96

Several pairs are composed of miRNAs in the same family (e.g., hsa - miR - 10a and hsa - miR - 10b).

# Up-regulated by the proto-oncogene c-MYC (O'Donnell et al., 2005).

- † Increases cell growth in lung carcinomas (Cheng et al., 2005).
- ‡ Expressed at lower levels in cancerous and pre-cancerous tissues compared to normal colon tissues (Michael et al., 2003).

Several of the identified pairs are composed of miRNAs in the same family (e.g., hsa - miR - 10a and hsa - miR - 10b). The two most significantly correlated miRNAs are a pair of paralogs, hsa - miR - 17 - 5p (chromosome 17) and hsa - miR - 106a(chromosome X), which belong to miRNA clusters believed to be up-regulated by the proto-oncogene c-MYC (O'Donnell et al., 2005). hsa - miR - 19a, hsa - miR - 19b, and hsa - miR - 20 are also members of these paralogous miRNA clusters.

hsa - miR - 107 has been shown to increase cell growth in lung carcinomas (Cheng et al., 2005).

hsa - miR - 143 and hsa - miR - 145, located within 1.7 kb on human chromosome 5, were found to be expressed at lower levels in cancerous and pre-cancerous tissues compared to normal colon tissues (Michael et al., 2003).

It would be of interest to investigate the biological and medical implications of the identified clusters of co-expressed miRNAs.

Segal et al. (2004). HIV-1 dataset. Studying genomic sequence variation for the human immunodeficiency virus type 1 (HIV-1) could potentially give important insight into genotype-phenotype associations for the acquired immune deficiency syndrome (AIDS).

- Phenotype. Replication capacity (RC) of HIV-1, which reflects the severity of the disease.
- Genotypes. Codons/amino acids in the protease and reverse transcriptase regions of the viral strand.

Goal. Relate HIV-1 protein sequence variation to viral replication capacity.

The protease (PR) enzyme affects the reproductive cycle of the virus by breaking protein peptide bonds during replication.

The reverse transcriptase (RT) enzyme synthesizes double-stranded DNA from the virus' single-stranded RNA genome, thereby facilitating integration into the host's chromosome.

Because the PR and RT regions are essential to viral replication, many antiretrovirals (protease inhibitors and reverse transcriptase inhibitors) have been developed to target these specific genomic locations.

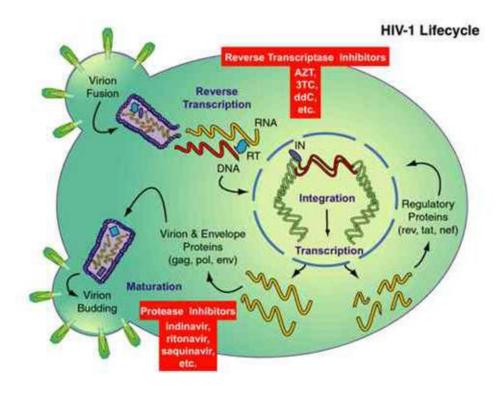


Figure 11: HIV-1 lifecycle. Diagram of the HIV-1 lifecycle and modes of action of protease and reverse transcriptase inhibitors.

Data. The HIV-1 dataset comprises n=317 records, linking viral replication capacity with PR and RT sequence data, from individuals participating in studies at the San Francisco General Hospital and the Gladstone Institute of Virology and Immunology.

The data for each of the n = 317 patients consist of the following.

- Y, a continuous replication capacity outcome/phenotype.
- X = (X(m): m = 1, ..., M), an M = 96 + 186 = 282-dimensional covariate vector of binary codon genotypes in the PR (pr4-pr99) and RT (rt38-rt223) HIV-1 regions. Codons are recoded as binary covariates, with value of 0 corresponding to the wild-type codon, i.e., the most common codon among the n = 317 patients, and value of 1 for mutant codons, i.e., all other codons.

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Multiple testing question. Test for each of the M=282 codon positions whether viral replication capacity Y is associated with the corresponding binary codon genotype  $X(m) \in \{0, 1\}$ .

Parameters of interest. For the mth codon position, the parameter of interest is the difference in mean replication capacity  $\psi(m)$  for viruses with mutant and wild-type codons, that is,

$$\psi(m) = E[Y|X(m) = 1] - E[Y|X(m) = 0], \qquad m = 1, \dots, M.$$

Null hypotheses. Consider two-sided tests of the M null hypotheses of no differences in mean RC vs. the alternative hypotheses of different mean RC,

$$H_0(m) = I(\psi(m) = 0)$$
 vs.  $H_1(m) = I(\psi(m) \neq 0)$ .

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Test statistics. Two-sample pooled-variance t-statistics,

$$T_n(m) = \frac{\psi_n(m) - \psi_0(m)}{\sigma_n(m)} = \frac{\bar{Y}_{1,n}(m) - \bar{Y}_{0,n}(m) - 0}{\sigma_{p,n}(m)\sqrt{\frac{1}{n_0(m)} + \frac{1}{n_1(m)}}},$$

$$\sigma_{p,n}^2(m) = \frac{(n_0(m) - 1)\sigma_{0,n}^2(m) + (n_1(m) - 1)\sigma_{1,n}^2(m)}{n_0(m) + n_1(m) - 2},$$

where the null values  $\psi_0(m)$  are zero,  $n_k(m) = \sum_i I(X_i(m) = k)$  denotes the number of patients with codon genotype

$$X(m) = k \in \{0, 1\}$$
 at position  $m$ , and

$$\bar{Y}_{k,n}(m) = \sum_{i} I(X_i(m) = k) Y_i/n_k(m)$$
 and

$$\sigma_{k,n}^2(m) = \sum_i I(X_i(m) = k) (Y_i - \bar{Y}_{k,n}(m))^2 / (n_k(m) - 1)$$
 denote,

respectively, the sample means and sample variances for the RC of patients with codon genotype  $X(m) = k \in \{0, 1\}$  at position m.

The pooled-variance estimators are denoted by  $\sigma_{p,n}^2(m)$ .

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Test statistics null distribution. Non-parametric bootstrap estimator of the null shift and scale-transformed test statistics null distribution, B = 7,500 samples.

Multiple testing procedures.

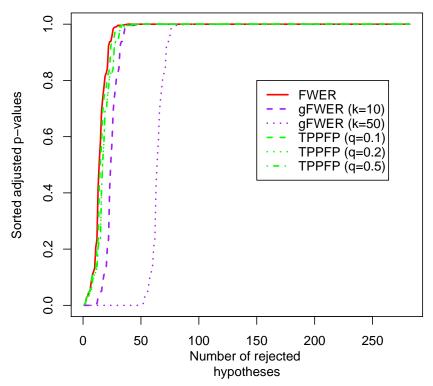
- 1. FWER-controlling single-step maxT procedure (SS maxT).
- 2. gFWER(k)-controlling augmentation procedure, based on FWER-controlling single-step maxT procedure, for an allowed number  $k \in \{10, 50\}$  of false positives (gFWER(k) AMTP).
- 3. TPPFP(q)-controlling augmentation procedure, based on FWER-controlling single-step maxT procedure, for an allowed proportion  $q \in \{0.10, 0.20, 0.50\}$  of false positives (TPPFP(q) AMTP).

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Table 6: HIV-1 dataset. t-statistics and sorted adjusted p-values for FWER-controlling single-step maxT procedure, gFWER(k)-controlling augmentation procedure (k = 5), and TPPFP(q)-controlling augmentation procedure (q = 0.10).

Codon	t-statistic	${\bf Adjusted} \ \ p{\bf -values}$		
position		SS maxT	gFWER(k) AMTP	TPPFP(q) AMTP
pr32	-9.755	0.0001	0	0.0001
pr47	-9.579	0.0013	0	0.0013
pr34	-8.843	0.0087	0	0.0087
pr55	-8.150	0.0104	0	0.0104
pr90	-6.237	0.0396	0	0.0396
rt184	-6.162	0.0431	0.0001	0.0431
pr43	-6.118	$\underline{0.0444}$	0.0013	$\underline{0.0444}$
pr54	-5.539	0.0780	0.0087	0.0780
rt41	-5.225	0.0978	0.0104	0.0978
pr46	-5.224	0.0980	0.0396	0.0978
pr82	-4.521	0.1678	0.0431	0.0980
rt215	-4.479	0.1740	$\underline{0.0444}$	0.1678
rt121	-4.070	0.2380	0.0780	0.1740

#### HIV-1 Sequence Variation and Viral Replication Capacity HIV: M=282 Codon Positions.



HIV-1 dataset. Sorted adjusted p-values for FWER-Figure 12: controlling single-step maxT procedure, gFWER(k)-controlling augmentation procedure  $(k \in \{10, 50\})$ , and TPPFP(q)-controlling augmentation procedure  $(q \in \{0.10, 0.20, 0.50\})$ .

The 13 codon positions with the smallest adjusted p-values all have negative t-statistics, suggesting that mutant codons (recoded as 1) are associated with decreased viral replication capacity.

The specific mutations observed in the present study are consistent with those found in the literature.

Protease: Vpr32I, Mpr46I, Ipr54V/L/T, Vpr82A/T/F/S, and Lpr90M increase the resistance of HIV-1 to various protease inhibitors.

Reverse transcriptase: Mrt41L, Mrt184V/I, and Trt215Y/F are related to azidothymidine (AZT) resistance.

The multiple testing procedures developed in Dudoit and van der Laan (2007) and related articles are implemented in the R package multtest, released as part of the Bioconductor Project, an open-source software project for the analysis of biomedical and genomic data.

Please consult the package documentation (e.g., helpfiles, manuals) and the book chapters by Dudoit and van der Laan (2007, Section 13.1) and Pollard et al. (2005) for details.

Bioconductor R package: multtest.

Authors: Katherine S. Pollard, Yongchao Ge, and Sandrine Dudoit.

URL: www.bioconductor.org.

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Test statistics. t-statistics for tests of regression coefficients in linear models and Cox proportional hazards survival models; F-statistics for tests of equality of means in one-way and two-way designs.

Weighted and robust rank-based versions of the above test statistics are implemented.

Test statistics null distribution. Bootstrap null shift and scale-transformed; permutation (Chapter 2 in Dudoit and van der Laan, 2007).

#### Multiple testing procedures.

- FWER control: single-step Bonferroni (1936); step-down Holm (1979); step-up Hochberg (1988); single-step maxT and minP (Chapter 4 in Dudoit and van der Laan, 2007; Dudoit et al., 2004; Pollard and van der Laan, 2004); step-down maxT and minP (Chapter 5 in Dudoit and van der Laan, 2007; van der Laan et al., 2004a).
- gFWER and TPPFP control: augmentation multiple testing procedures (Chapter 6 in Dudoit and van der Laan, 2007; van der Laan et al., 2004b).
- FDR control: step-up Benjamini and Hochberg (1995) and Benjamini and Yekutieli (2001); TPPFP-based (Chapter 6 in Dudoit and van der Laan, 2007; van der Laan et al., 2004b).

- Numerical summaries. Parameter estimates; test statistics; unadjusted and adjusted p-values; test statistic cut-offs; parameter confidence regions; estimated null distribution.
- Graphical summaries. Type I error rate vs. # rejections; # rejections vs. adjusted p-values; adjusted p-values vs. test statistics ("volcano" plots).
- Software design.
  - Function closure. Allow uniform data input for all MTPs and facilitate the extension of the package's functionality, by implementing, for example, new types of test statistics.
  - Class/method object-oriented programming. Represent and operate on the results of multiple testing procedures.

## Software Implementation: SAS Macros

SAS macros are available to compute the following components of a MTP (Birkner et al., 2005):

- *t*-statistics;
- non-parametric bootstrap estimates of the null shift and scale-transformed test statistics null distribution;
- adjusted p-values for the FWER-controlling single-step maxT procedure;
- adjusted p-values for the gFWER- and TPPFP-controlling augmentation procedures.

Author: M. D. Birkner.

URL: www.stat.berkeley.edu/~sandrine/MTBook.

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