EC 320: Introduction to Econometrics

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Prologue

Housekeeping

Final

• This Friday

Analytical Homework 4

• Thursday or Friday? Up to class.

Homework 3

• Graded by tonight

Goal: Make quantitative statements about qualitative information.

• e.g., race, gender, being employed, living in Oregon, etc.

Approach: Construct binary variables.

- *a.k.a.* dummy variables or indicator variables.
- Value equals 1 if observation is in the category or 0 if otherwise.

Regression implications

- 1. Binary variables change the interpretation of the intercept.
- 2. Coefficients on binary variables have different interpretations than those on continuous variables.

Consider the relationship

$$\mathrm{Pay}_i = \beta_0 + \beta_1 \mathrm{School}_i + u_i$$

where

- ullet Pay_i is a continuous variable measuring an individual's pay
- School_i is a continuous variable that measures years of education

Interpretation

- β_0 : *y*-intercept, *i.e.*, Pay when School = 0
- β_1 : expected increase in Pay for a one-unit increase in School

Consider the relationship

$$\mathrm{Pay}_i = \beta_0 + \beta_1 \mathrm{School}_i + u_i$$

Derive the slope's interpretation:

$$egin{aligned} \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell+1] - \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell] \ &= \mathbb{E}[eta_0 + eta_1(\ell+1) + u] - \mathbb{E}[eta_0 + eta_1\ell + u] \ &= [eta_0 + eta_1(\ell+1)] - [eta_0 + eta_1\ell] \ &= eta_0 - eta_0 + eta_1\ell - eta_1\ell + eta_1 \ &= eta_1. \end{aligned}$$

The slope gives the expected increase in pay for an additional year of schooling.

Consider the relationship

$$\mathrm{Pay}_i = \beta_0 + \beta_1 \mathrm{School}_i + u_i$$

Alternative derivation

Differentiate the model with respect to schooling:

$$rac{d ext{Pay}}{d ext{School}} = eta_1$$

The slope gives the expected increase in pay for an additional year of schooling.

If we have multiple explanatory variables, e.g.,

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + \beta_2 \text{Ability}_i + u_i$$

then the interpretation changes slightly.

$$egin{aligned} \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell + 1 \wedge \operatorname{Ability} = lpha] - \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell \wedge \operatorname{Ability} = lpha] \\ &= \mathbb{E}[eta_0 + eta_1(\ell+1) + eta_2lpha + u] - \mathbb{E}[eta_0 + eta_1\ell + eta_2lpha + u] \\ &= [eta_0 + eta_1(\ell+1) + eta_2lpha] - [eta_0 + eta_1\ell + eta_2lpha] \\ &= eta_0 - eta_0 + eta_1\ell - eta_1\ell + eta_1 + eta_2lpha - eta_2lpha \\ &= eta_1 \end{aligned}$$

The slope gives the expected increase in pay for an additional year of schooling, holding ability constant.

If we have multiple explanatory variables, e.g.,

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + \beta_2 \text{Ability}_i + u_i$$

then the interpretation changes slightly.

Alternative derivation

Differentiate the model with respect to schooling:

$$rac{\partial ext{Pay}}{\partial ext{School}} = eta_1$$

The slope gives the expected increase in pay for an additional year of schooling, holding ability constant.

Consider the relationship

$$\mathrm{Pay}_i = eta_0 + eta_1 \mathrm{Female}_i + u_i$$

where Pay_i is a continuous variable measuring an individual's pay and Female_i is a binary variable equal to 1 when i is female.

Interpretation

 eta_0 is the expected Pay for males (*i.e.*, when $\mathrm{Female} = 0$):

$$egin{aligned} \mathbb{E}[\operatorname{Pay}|\operatorname{Male}] \ &= \mathbb{E}[eta_0 + eta_1 imes 0 + u_i] \ &= \mathbb{E}[eta_0 + 0 + u_i] \ &= eta_0 \end{aligned}$$

Consider the relationship

$$\mathrm{Pay}_i = eta_0 + eta_1 \mathrm{Female}_i + u_i$$

where Pay_i is a continuous variable measuring an individual's pay and Female_i is a binary variable equal to 1 when i is female.

Interpretation

 eta_1 is the expected difference in Pay between females and males:

$$egin{aligned} \mathbb{E}[ext{Pay}| ext{Female}] &- \mathbb{E}[ext{Pay}| ext{Male}] \ &= \mathbb{E}[eta_0 + eta_1 imes 1 + u_i] - \mathbb{E}[eta_0 + eta_1 imes 0 + u_i] \ &= \mathbb{E}[eta_0 + eta_1 + u_i] - \mathbb{E}[eta_0 + 0 + u_i] \ &= eta_0 + eta_1 - eta_0 \ &= eta_1 \end{aligned}$$

Consider the relationship

$$Pay_i = \beta_0 + \beta_1 Female_i + u_i$$

where Pay_i is a continuous variable measuring an individual's pay and Female_i is a binary variable equal to 1 when i is female.

Interpretation

 $\beta_0 + \beta_1$: is the expected Pay for females:

$$egin{aligned} \mathbb{E}[ext{Pay}| ext{Female}] \ &= \mathbb{E}[eta_0 + eta_1 imes 1 + u_i] \ &= \mathbb{E}[eta_0 + eta_1 + u_i] \ &= eta_0 + eta_1 \end{aligned}$$

Consider the relationship

$$Pay_i = \beta_0 + \beta_1 Female_i + u_i$$

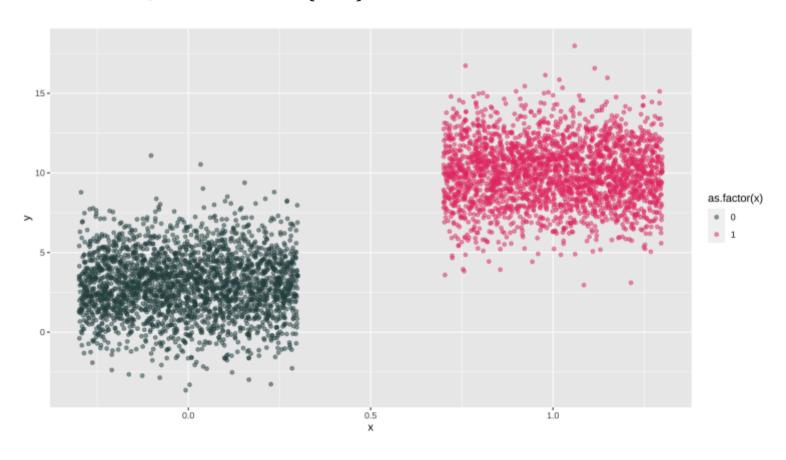
Interpretation

- β_0 : expected Pay for males (*i.e.*, when Female = 0)
- β_1 : expected difference in Pay between females and males
- $\beta_0 + \beta_1$: expected Pay for females
- Males are the reference group

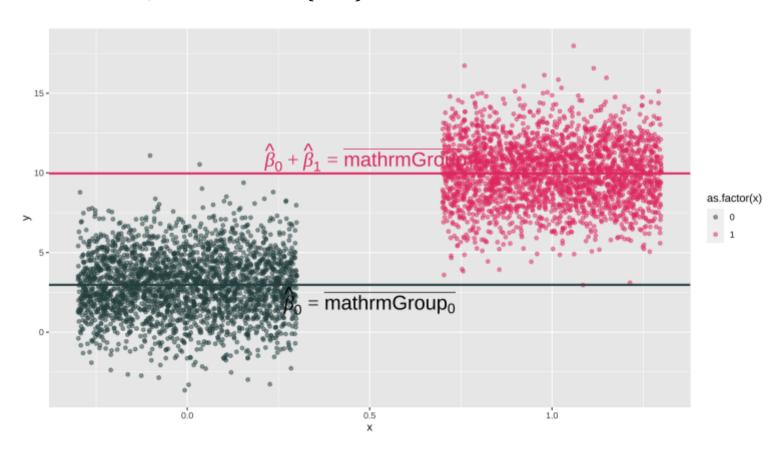
Note: If there are no other variables to condition on, then \hat{eta}_1 equals the difference in group means, *e.g.*, $ar{X}_{\mathrm{Female}} - ar{X}_{\mathrm{Male}}.$

Note₂: The *holding all other variables constant* interpretation also applies for categorical variables in multiple regression settings.

 $Y_i = eta_0 + eta_1 X_i + u_i$ for binary variable $X_i = \{0, \, 1\}$

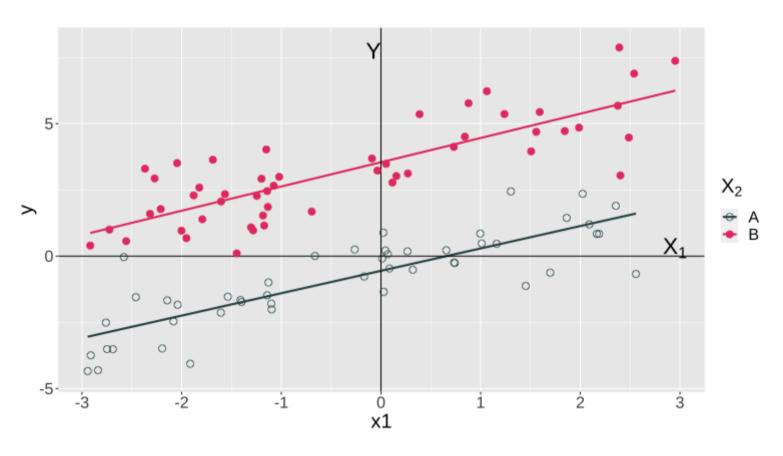


 $Y_i = eta_0 + eta_1 X_i + u_i$ for binary variable $X_i = \{0,\, 1\}$



Multiple Regression

Another way to think about it:



Question: Why not estimate $\mathrm{Pay}_i = \beta_0 + \beta_1 \mathrm{Female}_i + \beta_2 \mathrm{Male}_i + u_i$?

Answer: The intercept is a perfect linear combination of $Male_i$ and $Female_i$.

- Violates no perfect collinearity assumption.
- OLS can't estimate all three parameters simultaneously.
- Known as dummy variable trap.

Practical solution: Select a reference category and drop its indicator.

Dummy Variable *Trap?*

-168.

NA

Don't worry, R will bail you out if you include perfectly collinear indicators.

NA

Example

NA

10.9 -15.4 7.78e-52

NA

Thanks, R.

#> 2 black

#> 3 nonblack

Omitted variable bias (OVB) arises when we omit a variable that

- 1. Affects the outcome variable Y
- 2. Correlates with an explanatory variable X_i

Biases OLS estimator of β_j .

Example

Let's imagine a simple population model for the amount individual i gets paid

$$ext{Pay}_i = eta_0 + eta_1 ext{School}_i + eta_2 ext{Male}_i + u_i$$

where $School_i$ gives i's years of schooling and $Male_i$ denotes an indicator variable for whether individual i is male.

Interpretation

- β_1 : returns to an additional year of schooling (*ceteris paribus*)
- β_2 : premium for being male (*ceteris paribus*) If $\beta_2>0$, then there is discrimination against women.

Example, continued

From the population model

$$\mathrm{Pay}_i = eta_0 + eta_1 \mathrm{School}_i + eta_2 \mathrm{Male}_i + u_i$$

An analyst focuses on the relationship between pay and schooling, i.e.,

$$ext{Pay}_i = eta_0 + eta_1 ext{School}_i + (eta_2 ext{Male}_i + u_i)$$
 $ext{Pay}_i = eta_0 + eta_1 ext{School}_i + arepsilon_i$

where $arepsilon_i = eta_2 \mathrm{Male}_i + u_i$.

We assumed exogeneity to show that OLS is unbiasedness. But even if $\mathbb{E}[u|X]=0$, it is not necessarily true that $\mathbb{E}[arepsilon|X]=0$ (false if $eta_2
eq 0$).

Specifically, $\mathbb{E}[arepsilon|\mathrm{Male}=1]=eta_2+\mathbb{E}[u|\mathrm{Male}=1]
eq 0$. Now OLS is biased.

Let's try to see this result graphically.

The true population model:

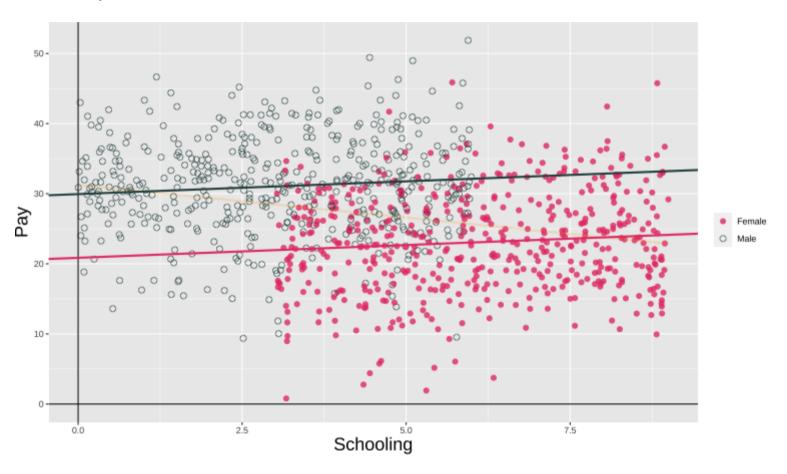
$$\mathrm{Pay}_i = 20 + 0.5 \times \mathrm{School}_i + 10 \times \mathrm{Male}_i + u_i$$

The regression model that suffers from omitted-variable bias:

$$\mathrm{Pay}_i = \hat{eta}_0 + \hat{eta}_1 imes \mathrm{School}_i + e_i$$

Finally, imagine that women, on average, receive more schooling than men.

Unbiased regression: $\widehat{\mathrm{Pay}}_i = 20.9 + 0.4 imes \mathrm{School}_i + 9.1 imes \mathrm{Male}_i$



Example: Weekly Wages

Q1: What is the reference category?

Q2: Interpret the coefficients.

Q3: Suppose you ran lm(wage ~ nonsouth, data = wage_data) instead. What is the coefficient estimate on nonsouth? What is the intercept estimate?

Example: Weekly Wages

Q1: What is the reference category?

Q2: Interpret the coefficients.

Q3: Suppose you ran lm(wage ~ south + nonblack, data = wage_data) instead. What is the coefficient estimate on nonblack? What is the coefficient estimate on south? What is the intercept estimate?

Example: Weekly Wages

129.

Answer to Q3:

#> 3 nonblack

```
lm(wage ~ south + nonblack, data = wage_data) %>% tidy()
#> # A tibble: 3 × 5
               estimate std.error statistic p.value
#>
    term
                                   <dbl>
                                           <dbl>
#>
    <chr>
                 <dbl>
                          <dbl>
#> 1 (Intercept)
                 518.
                          11.7 44.3 0
#> 2 south
              -98.6 9.84 -10.0 2.89e-23
```

11.4 11.3 3.43e-29