

(5.4,5.5) ETS Experiments and Hierarchical Normal Models

Weihang Ren

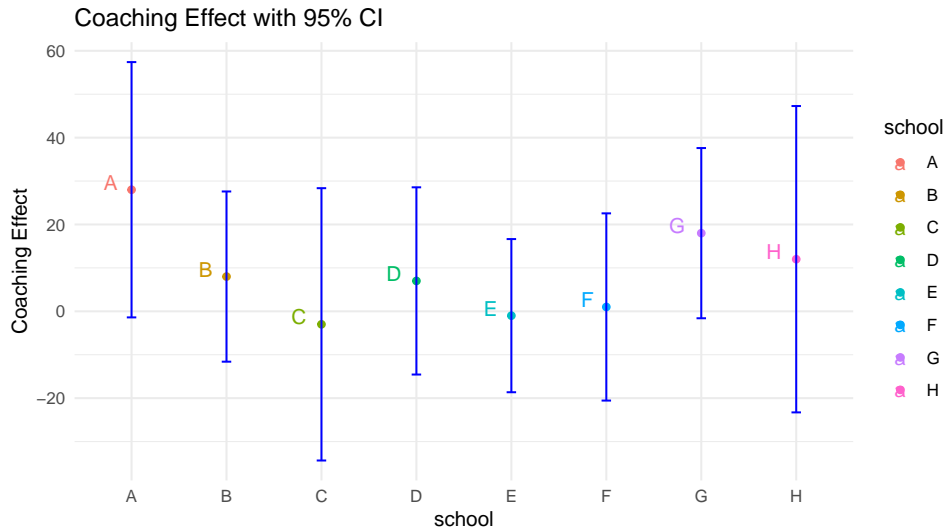
The ETS Dataset

- ▶ ETS performed a study to analyze the effects of coaching programs on test scores (SAT Verbal).
- ▶ Data was collected from eight high schools.
- ▶ Data:
 - ▶ The estimated coaching effect $y_j, j = 1, 2, \dots, 8$,
 - ▶ And its sampling variance σ_j^2 .
 - ▶ Approximated normal sampling distribution.

High School	A	B	C	D	E	F	G	H
Coaching Effect	28	8	-3	7	-1	1	18	12
SE of C.E.	15	10	16	11	9	11	10	18

- ▶ Question: How effective are SAT-V prep courses?

The ETS Dataset



The ETS dataset : Data Structure

- ▶ J independent experiments (schools).
- ▶ Experiment j is to estimate θ_j with n_j data points y_{ij} (Coaching Effect).

$$y_{ij}|\theta_j \sim N(\theta_j, \sigma^2), i = 1, \dots, n_j, j = 1, \dots, J$$

σ^2 is assumed to be known.

- ▶ Denote $\bar{y}_{.j} = \frac{1}{n_j} \sum_{i=1}^{n_j} y_{ij}$, then

$$\bar{y}_{.j}|\theta_j \sim N(\theta_j, \sigma_j^2), \sigma_j^2 = \sigma^2/n_j$$

Nonhierarchical Approaches

Separate Estimates

Pooled Estimates

Nonhierarchical Approaches

Separate Estimates

- Consider each estimate θ_j separately.

$$\bar{y}_{\cdot j} = \frac{1}{n_j} \sum_{i=1}^{n_j} y_{ij}$$

Pooled Estimates

- Consider a single common effect

$$\bar{y}_{..} = \frac{1}{\sum_{j=1}^J n_j} \sum_{i,j} y_{ij} = \frac{\sum_{j=1}^J \frac{1}{\sigma_j^2} \bar{y}_{\cdot j}}{\sum_{j=1}^J \frac{1}{\sigma_j^2}}$$

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	df	SS	MS	$E(MS \sigma^2, \tau)$
Between Groups	$J - 1$	$\sum_i \sum_j (\bar{y}_{\cdot j} - \bar{y}_{..})^2$	$SS/(J - 1)$	$n\tau^2 + \sigma^2$
Within Groups	$J(n - 1)$	$\sum_i \sum_j (y_{ij} - \bar{y}_{\cdot j})^2$	$SS/J(n - 1)$	σ^2
Total	$Jn - 1$	$\sum_i \sum_j (y_{ij} - \bar{y}_{..})^2$	$SS/(Jn - 1)$	

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- For school A, effect is 7.9 with se 4.2

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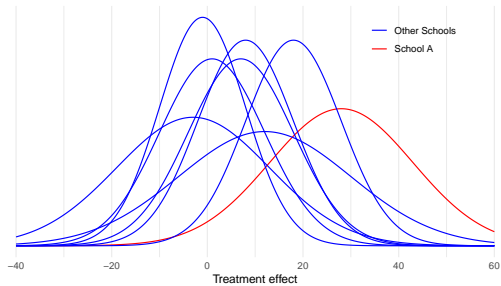
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- $P(\theta_1 > 28.4) = \frac{1}{2}?$

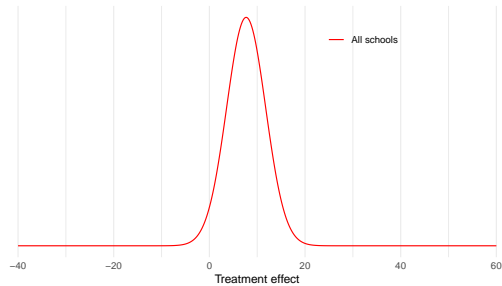
- For school A, effect is 7.9 with se 4.2
- $P(\theta_1 < 7.9) = \frac{1}{2}?$

The ETS Dataset

Separate model



Pooled model



Hierarchical Approaches

$$\hat{\theta}_j = \lambda_j \bar{y}_{\cdot j} + (1 - \lambda_j) \bar{y}_{\cdot \cdot}$$

- ▶ The separate estimate is posterior mean if J values θ_j have independent uniform prior density on $(-\infty, \infty)$.

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- ▶ θ_j 's are conditionally independent given (μ, τ) .
- ▶ Assign non-informative uniform hyperprior to μ given τ . We can afford to be vague because J experiments are highly informative about μ .

Hierarchical Approaches: Joint and Conditional Posterior Distribution

- Joint Posterior Distribution:

$$\begin{aligned} p(\theta, \mu, \tau | y) &\propto p(\mu, \tau) p(\theta | \mu, \tau) p(y | \theta) \\ &\propto p(\mu, \tau) \prod_{j=1}^J N(\theta_j | \mu, \tau^2) \prod_{j=1}^J N(\bar{y}_{\cdot j} | \theta_j, \sigma_j^2) \end{aligned}$$

- Conditional Posterior Distribution: θ_j are conditionally independent, so the conditional posterior has J components. Each of them is a normal so

$$\bar{y}_{\cdot j} | \theta_j \sim N(\theta_j, \sigma_j^2); \theta_j | (\mu, \tau) \sim N(\mu, \tau^2) \implies \theta_j | \mu, \tau, y \sim N(\hat{\theta}_j, V_j)$$

where

$$\hat{\theta}_j = \frac{\frac{1}{\sigma_j^2} \bar{y}_{\cdot j} + \frac{1}{\tau^2} \mu}{\frac{1}{\sigma_j^2} + \frac{1}{\tau^2}} \text{ and } V_j = \frac{1}{\frac{1}{\sigma_j^2} + \frac{1}{\tau^2}}$$

Hierarchical Approaches: Marginal Posterior Distribution

Fully Bayesian Treatment for $p(\mu, \tau|y)$:

1. Brute Force:

$$p(\mu, \tau|y) = \int p(\theta, \mu, \tau|y) d\theta$$

2. Analytic Solution:

$$p(\mu, \tau|y) = \frac{p(\theta, \mu, \tau|y)}{p(\theta|y)}$$

3. There is a different solution for Hierarchical Normal Model

$$p(\mu, \tau|y) \propto p(\mu, \tau)p(y|\mu, \tau)$$

- $p(y|\mu, \tau)$ could be written in closed form: $\bar{y}_{\cdot j}|\mu, \tau \sim N(\mu, \sigma_j^2 + \tau^2)$

$$E(e^{it\bar{y}_{\cdot j}}|\mu, \sigma) = E(E(e^{it\bar{y}_{\cdot j}}|\theta_j)|\mu, \sigma) = E(e^{it\theta_j}|\mu, \sigma)e^{-\frac{1}{2}\sigma_j^2} = e^{it\mu - \frac{1}{2}(\sigma_j^2 + \tau^2)}$$

Hierarchical Approaches: Posterior distribution of μ given τ . $p(\mu|\tau, y)$

- From previous page

$$p(\mu, \tau|y) \propto p(\tau)p(\mu|\tau) \prod_{j=1}^J N(\bar{y}_{\cdot j}|\mu, \sigma_j^2 + \tau^2)$$

- Given τ fixed, and $p(\mu|\tau) \propto 1$, the $\log p(\mu, \tau|y)$ is quadratic in μ . This implies $p(\mu|\tau, y)$ must be normal.

$$\mu|\tau, y \sim N(\hat{\mu}, V_{\mu})$$

where

$$\hat{\mu} = \frac{\sum_{j=1}^J \frac{1}{\sigma_j^2 + \tau^2} \bar{y}_{\cdot j}}{\sum_{j=1}^J \frac{1}{\sigma_j^2 + \tau^2}} \text{ and } V_{\mu} = \frac{1}{\sum_{j=1}^J \frac{1}{\sigma_j^2 + \tau^2}}$$

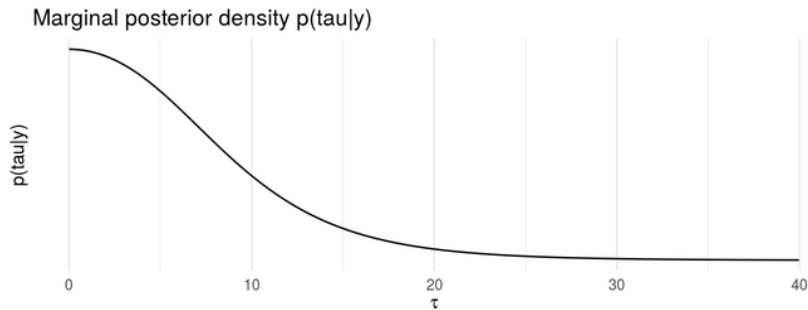
Hierarchical Approaches: Posterior distribution of τ . $p(\tau|y)$

$$\begin{aligned} p(\tau|y) &= \frac{p(\mu, \tau|y)}{p(\mu|\tau, y)} \propto \frac{p(\tau)p(\mu|\tau)p(y|\mu, \tau)}{p(\mu|\tau, y)} \\ &\propto \frac{p(\tau)p(\mu|\tau) \prod_{j=1}^J N(\bar{y}_{\cdot j}|\mu, \sigma_j^2 + \tau^2)}{N(\hat{\mu}, V_\mu)} \\ &\propto p(\tau) \frac{\prod_j (\sigma_j^2 + \tau^2)^{-1/2} \exp \left[-\frac{1}{2(\sigma_j^2 + \tau^2)} (\bar{y}_{\cdot j} - \mu)^2 \right]}{V_\mu^{-1/2} \exp \left[-\frac{1}{2V_\mu} (\mu - \hat{\mu})^2 \right]} \\ &\propto p(\tau) \frac{\prod_j (\sigma_j^2 + \tau^2)^{-1/2} \exp \left[-\frac{1}{2(\sigma_j^2 + \tau^2)} (\mu - \hat{\mu} + \hat{\mu} - \bar{y}_{\cdot j})^2 \right]}{V_\mu^{-1/2} \exp \left[-\frac{1}{2V_\mu} (\mu - \hat{\mu})^2 \right]} \\ &\propto p(\tau) V_\mu^{1/2} \prod_{i=1}^J (\sigma_j^2 + \tau^2)^{-1/2} \exp \left(-\frac{(\bar{y}_{\cdot j} - \hat{\mu})^2}{2(\sigma_j^2 + \tau^2)} \right) \end{aligned}$$

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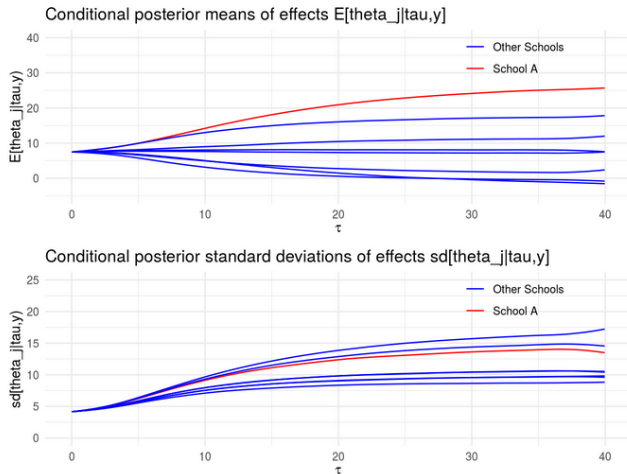
$$\begin{aligned} & - \sum_j \frac{1}{(\sigma_j^2 + \tau^2)} (\mu - \hat{\mu} + \hat{\mu} - \bar{y}_{\cdot j})^2 + \frac{1}{V_\mu} (\mu - \hat{\mu})^2 \\ &= \overbrace{- \sum_j \frac{(\hat{\mu} - \bar{y}_{\cdot j})^2}{(\sigma_j^2 + \tau^2)}}^{\text{keep}} - \sum_j \frac{(\mu - \hat{\mu})^2}{(\sigma_j^2 + \tau^2)} + \sum_j \frac{2(\hat{\mu} - \bar{y}_{\cdot j})(\mu - \hat{\mu})}{(\sigma_j^2 + \tau^2)} + \frac{1}{V_\mu} (\mu - \hat{\mu})^2 \\ &= \overbrace{- \sum_j \frac{(\hat{\mu} - \bar{y}_{\cdot j})^2}{(\sigma_j^2 + \tau^2)}}^{\text{keep}} + 2(\mu - \hat{\mu}) \sum_j \frac{(\hat{\mu} - \bar{y}_{\cdot j})}{(\sigma_j^2 + \tau^2)} \\ &= \overbrace{- \sum_j \frac{(\hat{\mu} - \bar{y}_{\cdot j})^2}{(\sigma_j^2 + \tau^2)}}^{\text{keep}} + 2(\mu - \hat{\mu})(\hat{\mu} V_\mu^{-1} - \hat{\mu} V_\mu^{-1}) = - \sum_j \frac{(\hat{\mu} - \bar{y}_{\cdot j})^2}{(\sigma_j^2 + \tau^2)} \end{aligned}$$

The ETS Dataset: Evaluate $p(\tau|y)$ on a Grid



- ▶ Values of τ near zero are most plausible;
- ▶ Values of τ larger than 10 are less than half as likely as $\tau = 0$, and $P(\tau > 25) \approx 0$.

The ETS Dataset: $p(\theta_j|\tau, y) = \int p(\theta_j|\mu, \tau, y)p(\mu|\tau, y)d\mu$



- ▶ For most of the likely values of τ , the estimated effects are relatively close together;
- ▶ As τ increases, the population distribution allows the eight effects to be more different from each other, and hence the posterior uncertainty in each individual θ_j increases

The ETS Dataset: Conclusion

- ▶ v.s. Pooled
 - ▶ Too much pulling together of the estimates in the eight schools; - τ is on the boundary of its parameter space.
- ▶ v.s. Seperate
 - ▶ Ordering of the effects in the eight schools is essentially the same as would be obtained by the eight separate estimates.
 - ▶ Bayesian probability that the effect in school A is as large as 28 points is less than 10%, which is substantially less than the 50% probability based on the separate estimate for school A.
- ▶ Hierarchical model is flexible enough to adapt to the data, thereby providing posterior inferences that account for the partial pooling as well as the uncertainty in the hyperparameters.