

Math - 680

HW - 2

(1)

(i) To show  $\rightarrow$  choose  $E(X-a)^2 = E(X-EX)^2$

Solution  $\rightarrow$

$$E(X-a)^2 = E(X \times (X - EX - a)) = E(X-EX)^2 + (EX-a)^2$$

OR  $\rightarrow E[(X-EX)^2 + 2(X-EX)(EX-a) + (EX-a)^2]$

$$E(X-EX)^2 \leq E(X-EX)^2 + (EX-a)^2$$

$\Rightarrow$

Which is minimum when  $a = EX$

[This was property of squared loss function]

(ii)

To show  $\rightarrow$  choose  $E|X-a| = E|X-m|$

where  $m$  is the median of  $X$ .

$$E|X-a| = \int_{-\infty}^{\infty} |X-a| f(x) dx \quad (\text{derivative})$$

$$= \int_{-\infty}^a (a-x) f(x) dx +$$

$$\int_a^{\infty} (x-a) f(x) dx$$

$$= \int_{-\infty}^a f(x) dx - \int_a^{\infty} f(x) dx$$

$$= 2F(a) - 1$$

$$2F(a) - 1 = 0$$

$$\Rightarrow F(a) = 1/2$$

So  $a$  is median

(2) (i) Posterior distribution of  $p$ .

$$P(p|h) \propto \pi(p) \propto \beta(p) L(h|p)$$

$$= \beta(\alpha, \beta) p^{\alpha-1} (1-p)^{\beta-1} \Rightarrow \prod_{i=1}^h \binom{h}{x_i} p^{x_i} (1-p)^{h-x_i}$$

$$= \frac{\beta(\alpha, \beta) p^{\alpha-1} (1-p)^{\beta-1} p^{\sum_{i=1}^h x_i} (1-p)^{h-\sum_{i=1}^h x_i}}{\beta(\alpha, \beta) p^{\alpha-1} (1-p)^{\beta-1}} \Rightarrow$$

$$\Rightarrow \sim \text{Beta}\left(\sum_{i=1}^h x_i + \alpha, h - \sum_{i=1}^h x_i + \beta\right).$$

(ii) To show  $\Rightarrow \hat{p}_{\text{Bayes}} = \frac{\sum_{i=1}^h x_i + \alpha}{\alpha + \beta + h}$

Bayesian estimation of  $p \Rightarrow \hat{p}_{\text{Bayes}}$

Posterior mean  $\Rightarrow E(p|h) \Rightarrow$

$$\Rightarrow \text{Which is equal to } \Rightarrow \frac{\sum_{i=1}^h x_i + \alpha}{\alpha + \beta + h}$$

Which can be further classified as

$$\Rightarrow \frac{h}{\alpha + \beta + h} \cdot \frac{\sum x_i}{h} + \frac{\alpha + \beta}{\alpha + \beta + h} \cdot \frac{\alpha}{\alpha + \beta}$$

$$\text{So } \hat{p}_{\text{Bayes}} \Rightarrow \frac{\sum x_i + \alpha}{\alpha + \beta + h}$$



(iii) To show  $\hat{p}_{\text{Bayes}}$  is given by

$$R(r, \hat{p}_{\text{Bayes}}) = \frac{nr(1-p)}{(\alpha + \beta + n)^2} + \left( \frac{np + \alpha}{\alpha + \beta + n} - r \right)^2$$

$$\begin{aligned} R(r, \hat{p}_{\text{Bayes}}) &= \text{Var}(p) + \text{Bias}^2(p) \\ &= \text{Var}\left(\frac{\sum_{i=1}^n x_i + \alpha}{\alpha + \beta + n}\right) + \left(\frac{1}{\alpha + \beta + n} E\left(\sum_{i=1}^n x_i\right) + \frac{\alpha}{\alpha + \beta + n} - p\right)^2 \end{aligned}$$

$$= \frac{np(1-p)}{(\alpha + \beta + n)^2} + \left( \frac{np + \alpha}{\alpha + \beta + n} - p \right)^2$$

$$\Rightarrow \frac{(\alpha(p-1) + \beta p)^2 + np(1-p)}{(\alpha + \beta + n)^2}$$

Risk function is given by

$$R(r, \hat{p}_{\text{Bayes}}) = \frac{\cancel{np} np(1-p) + (\alpha(p-1) + \beta p)^2}{(\alpha + \beta + n)^2}$$

(iv)  $\alpha + \beta = \sqrt{n/4}$

(a) for  $\hat{p}_{\text{Bayes}}$

$$\begin{aligned} &\Rightarrow \frac{\sum_{i=1}^n x_i + \alpha}{\alpha + \beta + n} \\ &= \frac{\sum_{i=1}^n x_i + \sqrt{n/4}}{\sqrt{n/4} + \sqrt{n/4} + n} \end{aligned}$$

$$z) \frac{\sum_{i=1}^n x_i - \sqrt{n/4}}{2(\sqrt{n/4}) + 1} \Rightarrow \boxed{\frac{\sum_{i=1}^n x_i + \sqrt{n/4}}{1 + \sqrt{n/4}}}$$

$$(b) R(p, p_{\text{Bayes}}) = \left( \frac{h p + \alpha}{\alpha + \beta + h} \right)^2 + \frac{h(p)(1-p)}{(\alpha + \beta + h)^2}$$

$$= \frac{(p\beta - p\alpha - \alpha)^2}{(\alpha + \beta + h)^2} + \frac{np - np^2}{(\alpha + \beta + n)}$$

$$\Rightarrow \frac{(p\alpha + p\beta - \alpha)^2 + np - np^2}{(\alpha + \beta + h)^2}$$

$$\text{Put } \alpha = \beta = \sqrt{n/4}$$

$$\Rightarrow \frac{(p(\sqrt{n/4}) + p(\sqrt{n/4}) - \sqrt{n/4})^2 + np - np^2}{(\sqrt{n/4} + \sqrt{n/4} + 1)^2}$$

[simplification  $\Rightarrow$ ]

$$\frac{(\sqrt{n/4})^2 - 2(\sqrt{n/4})(\sqrt{n/4}) + (\sqrt{n/4})^2 + np - np^2}{(2\sqrt{n/4} + 1)^2}$$

it has  $2\sqrt{n/4}$  so  $np^2$  comes  
to  $np - np^2$   
induct.

$$\Rightarrow \frac{n/4 - 2p + 4p^2 + np - np^2}{(2\sqrt{n/4} + 1)^2}$$



$$\Rightarrow \boxed{\frac{h}{9(9+\sqrt{h})^2}}$$

(3)

(i)  $C(0,1) = 3, \quad C(1,0) = 2$

Bayes rule  $\Rightarrow$

(From page 23)  
Lecture 4

$$F^*(x) = \begin{cases} 1 & \text{if } \frac{g_1(x)}{g_0(x)} > \frac{\pi_0(0,1)}{\pi_1(1,0)} \\ 0 & \text{if } \frac{g_1(x)}{g_0(x)} < \frac{\pi_1(0,1)}{\pi_0(1,0)} \end{cases}$$

For this problem bayes rule

$$F^*(x) = \begin{cases} 1 & \text{if } P(Y=1|X=x) = C(0,1) \\ 0 & \text{if } P(Y=0|X=x) = C(1,0) \end{cases}$$

(ii) Bayes decision Boundary

(From page 24)  
Lecture 4

$$\left\{ x: \frac{g_2(x)}{g_0(x)} = \frac{\pi_0}{\pi_2} \right\}$$

$$\left\{ x: \frac{g_1(x)}{g_0(x)} = \frac{3}{2} \right\} = \left\{ -1.582, 0.074 \right\}$$

(iii) By using Newton-Raphson method

$$\Rightarrow 0.65 \times \sup \left( -\frac{1}{2} (n-1)^2 \right) + 0.35 \sup \left\{ -\frac{(n-1)^2}{16} \right\} \frac{2.3}{2}$$

$$\Rightarrow -1.58, 0.27$$

Optimal classification region:

$$\Omega_1^* = (1.58, 0.27), \Omega_2^* = (-\infty, -1.58)$$

(4) Please refer to R code file

(5)

(i) For 0-1 loss, Bayes classifier:

$$X|Y=1 \sim N(\mu_1, \Sigma), \mu_1 = (2, 1)^T$$

$$X|Y=0 \sim N(\mu_0, \Sigma), \mu_0 = (1, 2)^T$$

This equation is satisfied by the decision boundary

$$P(Y=1|X=x) = P(Y=0|X=x)$$

$$\Rightarrow P(X=x|Y=1)P(Y=1) = P(X=x|Y=0)P(Y=0)$$

$$\Rightarrow \frac{P(X=x|Y=1)}{P(X=x|Y=0)} = \frac{P(Y=0)}{P(Y=1)}$$



taking log,

$$= \log \left[ \frac{P(X=1|Y=1)}{P(X=1|Y=0)} \right] = \log 1$$

$$\left( \log\left(\frac{e}{b}\right) = \log\left(\frac{e}{\frac{1}{b}}\right) = \log e - \log \frac{1}{b} \right) = \log P(X=1|Y=1) - \log(P(X=1|Y=0)) = 0$$

log of any no. = 0

Comparing in terms of  $\mu$ .

$$\Rightarrow -\frac{1}{2} (\mu - \mu_1)^T \Sigma^{-1} (\mu - \mu_1) + \frac{1}{2} (\mu - \mu_0)^T \Sigma^{-1} (\mu - \mu_0) = 0$$

$$\Rightarrow (\mu_1 + \mu_0)^T (\mu_1 - \mu_0) = 0$$

$$\Rightarrow \boxed{\mu_1 - \mu_2 = 0}$$

$$f^*(\mu) = \begin{cases} 1 & \text{if } \mu_1 > \mu_2 \\ 0 & \text{if } \mu_1 < \mu_2 \end{cases}$$

(ii) Mean up R code file

(iii) Bayes train error = 0.295

Bayes test error = 0.219