## Problem 1

Formulation:

$$\rho = \chi_{1} \exp \left( A_{11} \left( \frac{A_{11} \times z}{A_{11} \times A_{21} \times z} \right)^{2} \right) \rho_{1}^{sat} + \chi_{2} \exp \left( A_{21} \left( \frac{A_{11} \times z}{A_{11} \times A_{11} \times z} \right)^{2} \right) \rho_{1}^{sat}$$

The goal is to fit A12 and A21 with knot squares

$$\min_{A_{12}, A_{21}} \sum_{i=1}^{N} \left( \gamma_{i} - \rho(x_{i}, A_{12}, A_{21}) \right)^{2}$$

subject to  $X_1 + X_2 = 1$ 

$$l_{\text{M}_{10}}(\rho_{\text{Sale}}) = \alpha_1 - \frac{\alpha_2}{T + \alpha_3}$$
,  $\alpha_1, \alpha_2, \alpha_3$  in Table

T: 20

All of the constraints can be solved and substituted into objective functions

Saturation Pressures

$$P_{1}^{\text{SAt}} = P_{\text{VATE}}^{\text{SAT}} = 10^{\left(8.07131 - \frac{1730.63}{20 + 233.426}\right)} = 17.4733$$

$$p_2^{3NC} = p_{1M-4ionne}^{5AC} = 10^{-1.93155} - \frac{1554.679}{20 + 240.337} = 28.8241$$

Binary

$$\begin{array}{ccc}
Min & \sum_{l=1}^{N} & \left( Y_{l} - \left[ X_{i,l} \exp \left( A_{il} \left( \frac{A_{il} \left( 1 - X_{i,k} \right)}{A_{in} X_{i,k} A_{il} \left( 1 - X_{i,k} \right)} \right)^{2} \right) \rho_{i,kl}^{skl} + \left( 1 - X_{i,l} \right) \exp \left( A_{il} \left( \frac{A_{in} X_{i,l}}{A_{in} X_{i,k} A_{i,l} \left( 1 - X_{i,k} \right)} \right)^{2} \right) \rho_{i,kl}^{skl}
\end{array}$$

$$f(a) = \begin{cases} y_1 - p(x_1, a) \\ y_1 - p(x_1, a) \end{cases}$$

The gradient

$$\frac{\partial f_1}{\partial a_1} \frac{\partial f_2}{\partial a_2} \dots \frac{\partial f_n}{\partial a_n} f(a)$$

$$\frac{\partial f_n}{\partial a_n} \dots \frac{\partial f_n}{\partial a_n} \qquad \frac{\partial f_n}{\partial a_n} \qquad \frac{\partial f_n}{\partial a_n} = \frac{\partial f_n}{\partial a_n} \qquad \frac{\partial f_n}{\partial a_n} = \frac{\partial$$

$$p(y; x, \theta) = \frac{1}{1 + exp(-y \theta^T X)}$$

Assuming X is iid, the liklihood is:

$$L(0,0) = \prod_{i=1}^{N} \frac{1}{1+\exp(-y0^{T}X)} = \frac{1}{\prod_{i=1}^{N} \left[1+\exp(-y0^{T}X)\right]}$$

where

Taking the negative log, to get the negative log likelihood function gives:

$$\mathcal{L}(0,0) = -\left(M(1) - \ln\left(\prod_{i=1}^{N} \left[1 + \exp\left(-\gamma_{i} \theta^{T} x_{i}\right)\right]\right)\right)$$

$$= \sum_{i=1}^{N} \ln\left(1 + \exp\left(-\gamma_{i} \theta^{T} x_{i}\right)\right)$$

To find the max likelihood of  $\theta$ , we need to minimize  $\ell(0,0)$ , so take derivative of  $\ell$  w.i.t.  $\theta$ 

$$\frac{\partial L}{\partial \sigma} = \sum_{i=1}^{N} \frac{1}{1 + \exp(-\gamma_i \sigma^T x_i)} \left( -\gamma_i \times_i \left( \exp(-\gamma_i \sigma^T x_i) \right) \right)$$

$$= \sum_{i=1}^{N} \frac{-\gamma_{i} \times_{i} \left( \exp\left(-\gamma_{i} \Theta^{T} \times_{i}\right) \right)}{1 + \exp\left(-\gamma_{i} \Theta^{T} \times_{i}\right)} = 0$$

variable substitute

$$= \bigcup_{i=1}^{N} \alpha \left( \frac{r}{1+r} \right) = \frac{\partial L}{\partial \theta}$$

$$\frac{96}{95} = \sum_{N}^{151} \alpha \frac{91}{9} \left(\frac{14N}{N}\right) \frac{96}{94}$$

$$= \sum_{i=1}^{N} \frac{(y_i \times i)^2 \exp(-y_i \theta^T \times i)}{\left(1 + \exp(-y_i \theta^T \times i)\right)^2}$$

$$\frac{\partial^{2} l}{\partial \theta^{2}} = \sum_{i=1}^{N} \alpha_{i} \frac{\partial}{\partial r} \left( \frac{r}{1+r} \right) = \frac{\partial k}{\partial \theta}$$

$$= \sum_{i=1}^{N} \alpha_{i} \frac{\partial}{\partial r} \left( \frac{r}{1+r} \right) \cdot \frac{\partial r}{\partial \theta}$$

$$= \sum_{i=1}^{N} \frac{(\gamma_{i} \times i)^{2}}{(1+exp(-\gamma_{i} \theta^{T} \times i))^{2}}$$

$$= \sum_{i=1}^{N} \frac{(\gamma_{i} \times i)^{2}}{(1+exp(-\gamma_{i} \theta^{T} \times i))^{2}}$$

$$\frac{\partial^2 \ell}{\partial \theta^2} = \sum_{i=1}^{N} \frac{(x_i)^2 \exp(-y_i \theta^T x_i)}{\left( + \exp(-y_i \theta^T x_i) \right)^2}$$

Xi is always positive, the exp function is always positive, therefore

Because the Hessian is positive definite, then L is strictly convex, so there is only one DMAX and it is unique.