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Contents

- The RLS Algorithm
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The least squares objective function is

$$\xi^{d}(k) = \sum_{i=0}^{k} \lambda^{k-i} \varepsilon^{2}(i)$$

$$= \sum_{i=0}^{k} \lambda^{k-i} \left[d(i) - \mathbf{x}^{T}(i) \mathbf{w}(k) \right]^{2}$$
(1)

 λ is a weighting factor with $0 \ll \lambda < 1$. This parameter is called forgetting factor.



By differentiating $\xi^d(k)$ with respect to $\mathbf{w}(k)$

$$\frac{\partial \xi^d(k)}{\partial \mathbf{w}(k)} = -2\sum_{i=0}^k \lambda^{k-i} \mathbf{x}(i) [d(i) - \mathbf{x}^T(i)\mathbf{w}(k)]$$
 (2)



The optimal vector $\mathbf{w}(k)$ is

$$-\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{T}(i) \mathbf{w}(k) + \sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) d(i) = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$



$$\mathbf{w}(k) = \left[\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{T}(i)\right]^{-1} \sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) d(i)$$
$$= \mathbf{R}_{D}^{-1}(k) \mathbf{p}_{D}(k)$$
(3)

where $\mathbf{w}(k)$ in (3) is the optimal coefficient vector, and $\mathbf{R}_D(k)$ and $\mathbf{p}_D(k)$ are called deterministic correlation matrix and deterministic cross-correlation vector respectively.



The computation of the inverse matrix is avoided through the use of the matrix inversion lemma, where

$$\mathbf{S}_{D}(k) = \mathbf{R}_{D}^{-1}(k) = \frac{1}{\lambda} \left[\mathbf{S}_{D}(k-1) - \frac{\mathbf{S}_{D}(k-1)\mathbf{x}(k)\mathbf{x}^{T}(k)\mathbf{S}_{D}(k-1)}{\lambda + \mathbf{x}^{T}(k)\mathbf{S}_{D}(k-1)\mathbf{x}(k)} \right]$$
(4)



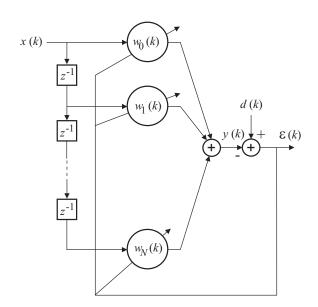


Figure 1: Adaptive FIR filter.



Algorithm 5.1 Conventional RLS Algorithm

Initialization

$$\mathbf{S}_D(-1) = \delta \mathbf{I}$$

where δ can be the inverse of the input-signal power estimate

$$\mathbf{p}_D(-1) = \mathbf{x}(-1) = [0 \ 0 \ \dots \ 0]^T$$



Do for $k \geq 0$:

$$\mathbf{S}_{D}(k) = \frac{1}{\lambda} [\mathbf{S}_{D}(k-1) - \frac{\mathbf{S}_{D}(k-1)\mathbf{X}(k)\mathbf{X}^{T}(k)\mathbf{S}_{D}(k-1)}{\lambda + \mathbf{X}^{T}(k)\mathbf{S}_{D}(k-1)\mathbf{X}(k)}]$$

$$\mathbf{p}_{D}(k) = \lambda \mathbf{p}_{D}(k-1) + d(k)\mathbf{x}(k)$$

$$\mathbf{w}(k) = \mathbf{S}_{D}(k)\mathbf{p}_{D}(k)$$

If necessary compute

$$y(k) = \mathbf{w}^T(k)\mathbf{x}(k)$$

$$\varepsilon(k) = d(k) - y(k)$$



Alternative RLS algorithm is obtained if (3) is rewritten as

$$[\sum_{i=0}^k \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^T(i)] \mathbf{w}(k) = \lambda [\sum_{i=0}^{k-1} \lambda^{k-i-1} \mathbf{x}(i) d(i)] + d(k) \mathbf{x}(k)$$

(5)



Considering that $\mathbf{R}_D(k-1)\mathbf{w}(k-1) = \mathbf{p}_D(k-1)$

$$[\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{T}(i)] \mathbf{w}(k) =$$

$$[\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{T}(i) - \mathbf{x}(k) \mathbf{x}^{T}(k)] \mathbf{w}(k-1) + \mathbf{x}(k) d(k)$$

(6)



Defining the *a priori* error as

$$e(k) = d(k) - \mathbf{x}^{T}(k)\mathbf{w}(k-1)$$
(7)

and substituting d(k) in equation (6) then

$$\mathbf{w}(k) = \mathbf{w}(k-1) + e(k)\mathbf{S}_D(k)\mathbf{x}(k)$$
(8)



Algorithm 5.2 Alternative RLS Algorithm

Initialization

$$\mathbf{S}_D(-1) = \delta \mathbf{I}$$

where δ can be the inverse of an estimate of the input signal power

$$\mathbf{x}(-1) = \mathbf{w}(-1) = [0 \ 0 \ \dots \ 0]^T$$



Do for $k \ge 0$

$$e(k) = d(k) - \mathbf{x}^{T}(k)\mathbf{w}(k-1)$$
 $\psi(k) = \mathbf{S}_{D}(k-1)\mathbf{x}(k)$

$$\mathbf{S}_D(k) = \frac{1}{\lambda} \left[\mathbf{S}_D(k-1) - \frac{\boldsymbol{\psi}(k) \boldsymbol{\psi}^T(k)}{\lambda + \boldsymbol{\psi}^T(k) \mathbf{X}(k)} \right]$$

$$\mathbf{w}(k) = \mathbf{w}(k-1) + e(k)\mathbf{S}_D(k)\dot{\mathbf{x}}(k)$$

If necessary compute

$$y(k) = \mathbf{w}^{T}(k)\mathbf{x}(k)$$
 $\varepsilon(k) = d(k) - y(k)$



No. of divisions is reduced using $\phi(k) = \frac{\psi^T(k)}{\lambda + \psi^T(k)\mathbf{X}(k)}$ that can be used as follows

$$\mathbf{S}_D(k) = \frac{1}{\lambda} [\mathbf{S}_D(k-1) - \psi(k)\boldsymbol{\phi}^T(k)]$$
 (9)

Properties Of The LS Solutions



Orthogonality Principle

Define the matrices $\mathbf{X}(k)$ and $\mathbf{d}(k)$ that contains all the information

$$\mathbf{X}(k) =$$

$$\mathbf{X}(k) = \begin{bmatrix} x(k) & \lambda^{1/2}x(k-1) & \cdots & \lambda^{k/2}x(0) \\ x(k-1) & \lambda^{1/2}x(k-2) & \cdots & 0 \\ \vdots & \vdots & & \vdots \\ x(k-N) & \lambda^{1/2}x(k-N-1) & \cdots & 0 \end{bmatrix}$$
$$= [\mathbf{x}(k) \ \lambda^{1/2}\mathbf{x}(k-1) \ \dots \ \lambda^{k/2}\mathbf{x}(0)]$$

$$= \left[\mathbf{x}(k) \ \lambda^{1/2} \mathbf{x}(k-1) \ \dots \ \lambda^{k/2} \mathbf{x}(0) \right] \tag{10}$$

$$\mathbf{d}(k) = [d(k) \ \lambda^{1/2} d(k-1) \ \dots \ \lambda^{k/2} d(0)]^T$$
 (11)

where $\mathbf{X}(k)$ is $(N+1) \times (k+1)$ and $\mathbf{d}(k)$ is $(k+1) \times (1)$.



By using the matrices defined above

$$\mathbf{X}(k)\mathbf{X}^{T}(k)\mathbf{w}(k) = \mathbf{X}(k)\mathbf{d}(k) \tag{12}$$

Hence defining

$$\mathbf{y}(k) = \mathbf{X}^{T}(k)\mathbf{w}(k)$$

$$= [y(k) \ \lambda^{1/2}y(k-1) \ \cdots \ \lambda^{k/2}y(0)]^{T}$$
(13)

it follows from (12) that

$$\mathbf{X}(k)\mathbf{X}^{T}(k)\mathbf{w}(k) - \mathbf{X}(k)\mathbf{d}(k)$$

$$= \mathbf{X}(k)[\mathbf{y}(k) - \mathbf{d}(k)] = \mathbf{0}$$
(14)



This relation means that the weighted error vector given by

$$\varepsilon(k) = \begin{bmatrix} \varepsilon(k) \\ \lambda^{1/2} \varepsilon(k-1) \\ \vdots \\ \lambda^{k/2} \varepsilon(0) \end{bmatrix} = \mathbf{d}(k) - \mathbf{y}(k)$$
 (15)

is in the null space of $\mathbf{X}(k)$.



Example 1

Suppose $\lambda = 1$ and that

$$\mathbf{d}(1) = \begin{bmatrix} 0.5 \\ 1.5 \end{bmatrix} \quad \mathbf{X}(1) = \begin{bmatrix} 1 & -2 \end{bmatrix} \quad \mathbf{w}(1) = -\frac{2.5}{5}$$



The output of the adaptive filter is

$$\mathbf{y}(1) = \left[\begin{array}{c} -\frac{2.5}{5} \\ 1 \end{array} \right]$$

Note that

$$[1 - 2][\mathbf{y}(1) - \mathbf{d}(1)] = [1 - 2] \begin{bmatrix} -1 \\ -0.5 \end{bmatrix} = 0$$



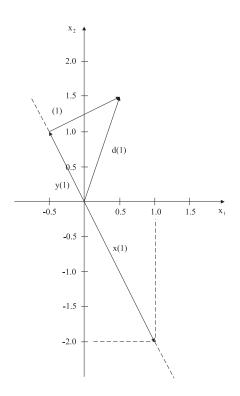


Figure 2: Geometric interpretation of least-square solution.



Relation Between LS and Wiener Solution

When $\lambda = 1$ the matrix $\frac{1}{k+1} \mathbf{R}_D(k)$ for large k is a consistent estimate of \mathbf{R} , if the process from which the input signal was taken is ergodic. The same is valid for $\frac{1}{k+1} \mathbf{p}_D(k)$ related to \mathbf{p} if the desired signal is ergodic.



In this case

$$\mathbf{R} = \lim_{k \to \infty} \frac{1}{k+1} \sum_{i=0}^{k} \mathbf{x}(i) \mathbf{x}^{T}(i) = \lim_{k \to \infty} \frac{1}{k+1} \mathbf{R}_{D}(k)$$
 (16)

and

$$\mathbf{p} = \lim_{k \to \infty} \frac{1}{k+1} \sum_{i=0}^{k} \mathbf{x}(i) d(i) = \lim_{k \to \infty} \frac{1}{k+1} \mathbf{p}_D(k)$$
 (17)

It can then be shown that

$$\mathbf{w}(k) = \mathbf{R}_D^{-1}(k)\mathbf{p}_D(k) = \mathbf{R}^{-1}\mathbf{p} = \mathbf{w}_o \tag{18}$$

for large k.



Influence of Initialization

The initialization of $\mathbf{S}_D(-1) = \delta \mathbf{I}$ causes a bias

$$\sum_{i=-1}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{T}(i) \mathbf{w}(k) = \left[\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{T}(i) + \frac{\lambda^{k+1}}{\delta} \mathbf{I} \right] \mathbf{w}(k)$$
$$= \mathbf{p}_{D}(k)$$
(19)



By multiplying $\mathbf{R}_D^{-1}(k)$ on both sides for $k \to \infty$

$$\mathbf{w}(k) + \frac{\lambda^{k+1}}{\delta} \mathbf{R}_D^{-1}(k) \mathbf{w}(k) = \mathbf{w}_o$$
 (20)

The bias is

$$\mathbf{w}(k) - \mathbf{w}_o \approx -\frac{\lambda^{k+1}}{\delta} \mathbf{S}_D(k) \mathbf{w}_o \tag{21}$$

For $\lambda < 1$ the bias tends to zero. For $\lambda = 1$, $\mathbf{S}_D(k)$ approaches a null matrix for large k.



Behavior of the Coefficient Vector

The input vectors $\mathbf{x}(k)$, for k = 0, 1, ..., are considered known, while the coefficients $w_i(k)$ for i = 0, 1, ..., N, are considered as stochastic processes. In this case for $k \geq N$.



$$E[\mathbf{w}(k)] = E\left\{ [\mathbf{X}(k)\mathbf{X}^{T}(k)]^{-1}\mathbf{X}(k)d(k) \right\}$$

$$= E\left[[\mathbf{X}(k)\mathbf{X}^{T}(k)]^{-1}[\mathbf{X}(k)\mathbf{X}^{T}(k)\mathbf{w}_{o} + \mathbf{n}(k)] \right]$$

$$= E[[\mathbf{X}(k)\mathbf{X}^{T}(k)]^{-1}\mathbf{X}(k)\mathbf{X}^{T}(k)\mathbf{w}_{o}] = \mathbf{w}_{o}$$
(22)

The equation shows that the estimate is unbiased when $\lambda \leq 1$.



Transient Behavior

Define

$$\Delta \mathbf{w}(k) = \mathbf{w}(k) - \mathbf{w}_o \tag{23}$$

Rewriting eq. (6) in matrix form it follows

$$\mathbf{R}_D(k)\mathbf{w}(k) = \lambda \mathbf{R}_D(k-1)\mathbf{w}(k-1) + \mathbf{x}(k)d(k)$$
 (24)



Defining the minimum output error as

$$e_o(k) = d(k) - \mathbf{w}_o^T \mathbf{x}(k) \tag{25}$$

and replacing d(k) in eq. (24)

$$\mathbf{R}_D(k)\Delta\mathbf{w}(k) = \lambda\mathbf{R}_D(k-1)\Delta\mathbf{w}(k-1) + \mathbf{x}(k)e_o(k)$$
 (26)



where it was used the following relation

$$\mathbf{R}_D(k) = \lambda \mathbf{R}_D(k-1) + \mathbf{x}(k)\mathbf{x}^T(k)$$
 (27)

The solution of eq. (26) is given by

$$\Delta \mathbf{w}(k) = \lambda^{k+1} \mathbf{R}_D^{-1}(k) \mathbf{R}_D(-1) \Delta \mathbf{w}(-1) + \mathbf{R}_D^{-1}(k) \sum_{i=0}^k \lambda^{k-i} \mathbf{x}(i) e_o(i)$$
(28)



By replacing $\mathbf{R}_D(-1)$ by $\frac{1}{\delta}$ I and taking the expected value

$$E[\Delta \mathbf{w}(k)] = \frac{\lambda^{k+1}}{\delta} E[\mathbf{S}_D(k)] \Delta \mathbf{w}(-1) + E[\mathbf{S}_D(k) \sum_{i=0}^k \lambda^{k-i} \mathbf{x}(i) e_o(i)]$$
(29)



Since $\mathbf{S}_D(k)$ is dependent on all past input vectors it can be considered independent of an individual $\mathbf{x}(i)$. Also, due to the orthogonality principle $e_o(i)$ can be considered uncorrelated to all elements of $\mathbf{x}(i)$. The last vector of eq. (29) cannot have large element values. The first vector of equation (29) has large element values only during the initial convergence, since as $k \to \infty$, $\lambda^{k+1} \to 0$ and $\mathbf{S}_D(k)$ has a nonincreasing behavior ,i.e. $\mathbf{R}_D(k)$ is assumed to remain positive definite as $k \to \infty$.



We conclude that the coefficients tend to \mathbf{w}_o almost independently

from any eigenvalue spread of **R**. In the simple case of $\lambda = 1$, consider the worst case which we replace $E[\mathbf{S}_D(k)]$ in (29) by

$$\mathbf{S}_{D_{\text{max}}} = \frac{\mathbf{q}_{\text{min}} \mathbf{q}_{\text{min}}^T}{\lambda_{\text{min}}(k+1)}$$
(30)

where \mathbf{q}_{\min} is the eigenvector of λ_{\min} . The value of the minimum eigenvalue affects the convergence only in the first few iterations.



Coefficient-Error-Vector Cov. Matrix

The coefficient-error-vector covariance matrix is given by

$$\operatorname{cov}\left[\Delta \mathbf{w}(k)\right] = E[(\mathbf{w}(k) - \mathbf{w}_o)(\mathbf{w}(k) - \mathbf{w}_o)^T]$$
$$= \sigma_n^2 E[\mathbf{S}_D(k)] \tag{31}$$

for $\lambda = 1$.



Proof

$$\mathbf{w}(k) - \mathbf{w}_o = \mathbf{S}_D(k)\mathbf{p}_D(k) - \mathbf{S}_D(k)\mathbf{S}_D^{-1}(k)\mathbf{w}_o$$

$$= (\mathbf{X}(k)\mathbf{X}^T(k))^{-1}\mathbf{X}(k)(\mathbf{d}(k) - \mathbf{X}^T(k)\mathbf{w}_o)$$

$$= (\mathbf{X}(k)\mathbf{X}^T(k))^{-1}\mathbf{X}(k)\mathbf{n}(k)$$

where
$$\mathbf{n}(k) = [n(k) \ \lambda^{1/2} n(k-1) \ \lambda n(k-2) \ \dots \ \lambda^{k/2} n(0)]^T$$
.



It then follows that

$$\operatorname{cov} \left[\Delta \mathbf{w}(k) \right] = E[(\mathbf{X}(k)\mathbf{X}^{T}(k))^{-1}\mathbf{X}(k)E[\mathbf{n}(k)\mathbf{n}^{T}(k)]\mathbf{X}^{T}(k)(\mathbf{X}(k)\mathbf{X}^{T}(k))^{-1}]$$
$$= \sigma_{n}^{2}E[\mathbf{S}_{D}(k)\mathbf{X}(k)\Lambda\mathbf{X}^{T}(k)\mathbf{S}_{D}(k)]$$

where $\mathbf{X}(k)$ was considered known and

$$\Lambda = \begin{bmatrix}
1 & & & & & \\
& \lambda & & & & \\
& & \lambda^2 & & \\
& 0 & & \ddots & \\
& & & \lambda^k
\end{bmatrix}$$



where for $\lambda = 1$,

$$\operatorname{cov} \left[\Delta \mathbf{w}(k) \right] = \sigma_n^2 \mathbf{S}_D(k) E[\mathbf{X}(k) \mathbf{X}^T(k) \mathbf{S}_D(k)]$$
$$= \sigma_n^2 E[\mathbf{S}_D(k) \mathbf{R}_D(k) \mathbf{S}_D(k)]$$
$$= \sigma_n^2 E[\mathbf{S}_D(k)]$$





Behavior of the Error Signal

With measurement noise, the error signal is given by

$$e(k) = d'(k) - \mathbf{w}^{T}(k-1)\mathbf{x}(k) + n(k)$$
(32)

then

$$E[e(k)] = E[d'(k)] - E[\mathbf{w}^T(k-1)]\mathbf{x}(k) + E[n(k)]$$

$$= E[d'(k)] - \mathbf{w}_o^T\mathbf{x}(k) + E[n(k)]$$

$$= E[n(k)]$$
(33)

assuming the filter order is suf. to model the desired signal.



If the noise signal has zero mean then

$$E[e(k)] = 0$$

The Minimum Mean Square Error

The minimum MSE in the presence of external uncorrelated noise is given by

$$\xi_{\min} = E[e^2(k)] = E[n^2(k)] = \sigma_n^2$$
 (34)

If the additive noise is correlated to the input and the desired signal, a more complicated expression for the MSE results.



When employing the a posteriori error the minimum MSE, denoted

by $\xi_{\min,p}$, differs from the corresponding value related to the *a priori* error. The following relation can be verified

$$\Delta \mathbf{w}(k) = \mathbf{S}_D(k)\mathbf{X}(k)\mathbf{n}(k) \tag{35}$$

When a measurement noise is present, the *a posteriori* error signal is given by

$$\varepsilon(k) = d'(k) - \mathbf{w}^{T}(k)\mathbf{x}(k) + n(k) = -\Delta\mathbf{w}^{T}(k)\mathbf{x}(k) + e_{o}(k)$$
 (36)



The expression for the MSE related to the *a posteriori* error is then given by

$$\xi(k) = E[\varepsilon^{2}(k)]$$

$$= E[e_{o}^{2}(k)] - 2E[\mathbf{x}^{T}(k)\Delta\mathbf{w}(k)e_{o}(k)]$$

$$+E[\Delta\mathbf{w}^{T}(k)\mathbf{x}(k)\mathbf{x}^{T}(k)\Delta\mathbf{w}(k)]$$
(37)



By replacing the expression (35) in equation (37) above

$$\xi(k) = E[e_o^2(k)] - 2E[\mathbf{x}^T(k)\mathbf{S}_D(k)\mathbf{X}(k)\mathbf{n}(k)e_o(k)]$$

$$+E[\Delta \mathbf{w}^T(k)\mathbf{x}(k)\mathbf{x}^T(k)\Delta \mathbf{w}(k)]$$

$$= \sigma_n^2 - 2E[\mathbf{x}^T(k)\mathbf{S}_D(k)\mathbf{X}(k)] \begin{bmatrix} \sigma_n^2 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

$$+E[\Delta \mathbf{w}^T(k)\mathbf{x}(k)\mathbf{x}^T(k)\Delta \mathbf{w}(k)]$$



$$= \sigma_n^2 - 2E[\mathbf{x}^T(k)\mathbf{S}_D(k)\mathbf{x}(k)]\sigma_n^2 + E[\Delta\mathbf{w}^T(k)\mathbf{x}(k)\mathbf{x}^T(k)\Delta\mathbf{w}(k)]$$

$$= \xi_{\min,p} + E[\Delta\mathbf{w}^T(k)\mathbf{x}(k)\mathbf{x}^T(k)\Delta\mathbf{w}(k)]$$
(38)

where in the second equality the additional noise is uncorrelated with the input signal and that $e_o(k) = n(k)$. This equality occurs when the adaptive filter has sufficient order to identify the unknown system.

 $\xi_{\min,p}$ related to the *a posteriori* error in eq. (38) is not the same as minimum MSE of the *a priori* error ξ_{\min} . The term $E[\Delta \mathbf{w}^T(k)\mathbf{x}(k)\mathbf{x}^T(k)\Delta \mathbf{w}(k)]$ in eq. (38) determines the excess MSE of the RLS algorithm.



It is possible to verify that the following expressions for $\xi_{\min,p}$ are accurate approximations

$$\xi_{\min,p} = \left\{1 - 2E[\mathbf{x}^{T}(k)\mathbf{S}_{D}(k)\mathbf{x}(k)]\right\}\sigma_{n}^{2}$$

$$= \left\{1 - 2\operatorname{tr}\left[E\left(\mathbf{S}_{D}(k)\mathbf{x}(k)\mathbf{x}^{T}(k)\right)\right]\right\}\sigma_{n}^{2}$$

$$= \left\{1 - 2\operatorname{tr}\left[\frac{1 - \lambda}{1 - \lambda^{k+1}}\mathbf{I}\right]\right\}\sigma_{n}^{2}$$

$$= \left\{1 - 2(N+1)\left[\frac{1 - \lambda}{1 - \lambda^{k+1}}\right]\right\}\sigma_{n}^{2}$$

$$= \left\{1 - 2(N+1)\left[\frac{1 - \lambda}{1 + \lambda + \lambda^{2} + \dots + \lambda^{k}}\right]\right\}\sigma_{n}^{2} \qquad (39)$$



Example 5.2

Repeat the equalization problem of Example 3.1 of Chapter 3 using the RLS algorithm.

- (a) Using $\lambda = 0.99$, run the algorithm and save the matrix $\mathbf{S}_D(k)$ at the iteration 500 and compare with the inverse of the input signal correlation matrix.
- (b) Plot the convergence path for the RLS algorithm on the MSE surface.



Solution:

(a) The inverse of matrix \mathbf{R} , as computed in the Example 3.1, is given by

$$\mathbf{R}^{-1} = 0.45106 \begin{bmatrix} 1.6873 & 0.7937 \\ 0.7937 & 1.6873 \end{bmatrix}$$
$$= \begin{bmatrix} 0.7611 & 0.3580 \\ 0.3580 & 0.7611 \end{bmatrix}$$



The initialization matrix $\mathbf{S}_D(-1)$ was a diagonal matrix with the diagonal elements equal to 0.1. The matrix $\mathbf{S}_D(k)$ at the 500th

iteration, obtained by averaging the results of 30 experiments, was

$$\mathbf{S}_D(500) = \begin{bmatrix} 0.0078 & 0.0037 \\ 0.0037 & 0.0078 \end{bmatrix}$$

Also, the obtained values of the deterministic cross-correlation vector was

$$\mathbf{p}_D(500) = \begin{bmatrix} 95.05 \\ 46.21 \end{bmatrix}$$



Now if we divide each element of the matrix \mathbf{R}^{-1} by

$$\frac{1-\lambda^{k+1}}{1-\lambda} = 99.34$$

The resulting matrix is

$$\frac{1}{99.34}\mathbf{R}^{-1} = \begin{bmatrix} 0.0077 & 0.0036\\ 0.0036 & 0.0077 \end{bmatrix}$$

as can be noted the values of the elements of the above matrix are close to the average values of the corresponding elements of matrix $\mathbf{S}_D(500)$.



Similarly, if we multiply the cross-correlation vector \mathbf{p} by 99.34, the

result is

$$99.34\mathbf{p} = \begin{bmatrix} 94.61\\47.31 \end{bmatrix}$$

The values of the elements of this vector are also close to the corresponding elements of $\mathbf{p}_D(500)$.



(b) The convergence path of the RLS algorithm on the MSE surface is depicted in Fig. 5.3. The reader should notice that the RLS algorithm approaches the minimum using large steps when the coefficients of the adaptive filter are far away from the optimum solution.



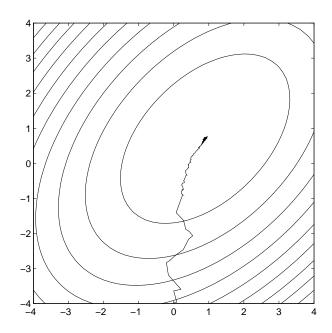


Figure 3: Convergence path of the RLS adaptive filter.



Excess MSE and Misadjustment

In the real implementation there is always an excess of MSE at the output caused by $\Delta \mathbf{w}(k) = \mathbf{w}(k) - \mathbf{w}_o$. The MSE for a given coefficient vector is

$$\xi(k) = \xi_{\min,p} + E[(\mathbf{w}(k) - \mathbf{w}_o)^T \mathbf{R}(\mathbf{w}(k) - \mathbf{w}_o)]$$

$$= \xi_{\min,p} + E[\Delta \mathbf{w}^T(k) \mathbf{R} \Delta \mathbf{w}(k)]$$
(40)



Now considering that $\Delta w_j(k)$ for $j = 0 \dots N$ are random variables

with zero mean and independent of $\mathbf{x}(k)$, then by employing eqs. (31) and (17), the ensemble average of the MSE can be calculated as

$$E[\xi(k)] = \xi_{min,p} + E[\Delta \mathbf{w}^{T}(k)\mathbf{R}\Delta \mathbf{w}(k)]$$
$$= \xi_{min,p} + E[\operatorname{tr}(\mathbf{R}\Delta \mathbf{w}(k)\Delta \mathbf{w}^{T}(k))]$$



$$= \xi_{min,p} + \operatorname{tr} \left(\mathbf{R} E[\Delta \mathbf{w}(k) \Delta \mathbf{w}^{T}(k)] \right)$$

$$= \xi_{min,p} + \sigma_{n}^{2} \operatorname{tr} \left(\mathbf{R} \mathbf{S}_{D}(k) \right)$$

$$= \sigma_{n}^{2} \left(1 + \operatorname{tr} \left(\mathbf{R} \frac{\mathbf{R}^{-1}}{k+1} \right) \right) for \ k \to \infty$$

$$= \sigma_{n}^{2} \left(1 + \frac{N+1}{k+1} \right)$$

where $\lambda = 1$ and $\xi_{min,p} = \xi_{min} = \sigma_n^2$. The minimum MSE is reached after the algorithm has operated a number of samples larger than the filter order.



The expected excess in the MSE is defined by

$$\Delta \xi(k) = E[\Delta \mathbf{w}^{T}(k) \mathbf{R} \Delta \mathbf{w}(k)] \tag{41}$$

For $\lambda < 1$. From equation (26) one can show that

$$\Delta \mathbf{w}(k) = \lambda \mathbf{R}_D^{-1}(k) \mathbf{R}_D(k-1) \Delta \mathbf{w}(k-1) + \mathbf{R}_D^{-1}(k) \mathbf{x}(k) e_o(k)$$
(42)



By applying equation (42) to (41) it follows that

$$E[\Delta \mathbf{w}^{T}(k)\mathbf{R}\Delta \mathbf{w}(k)] = \rho_1 + \rho_2 + \rho_3 + \rho_4 \tag{43}$$

where

$$\rho_1 = \lambda^2 E[\Delta \mathbf{w}^T (k-1) \mathbf{R} \Delta \mathbf{w} (k-1)] \tag{44}$$



$$\rho_2 \approx \lambda (1 - \lambda) E[\Delta \mathbf{w}^T(k - 1)] E[\mathbf{x}(k) e_o(k)]$$

$$= 0$$
(45)

Following a similar approach it can be shown that $\rho_3 = 0$.



3- Evaluation of ρ_4

$$\rho_4 = E[\mathbf{x}^T(k)\mathbf{R}_D^{-1}(k)\mathbf{R}\mathbf{R}_D^{-1}(k)\mathbf{R}\mathbf{R}^{-1}\mathbf{x}(k)e_o^2(k)]$$

$$\approx (1-\lambda)^2 E[\mathbf{x}^T(k)(\mathbf{I}-\Delta\mathbf{I}(k))^2\mathbf{R}^{-1}\mathbf{x}(k)]\xi_{\min}$$
(46)

Eq. (46) can be simplified to

$$\rho_4 = (1 - \lambda)^2 \operatorname{tr} \left\{ \mathbf{I} + E[\Delta \mathbf{I}^2(k)] \right\} \xi_{\min}$$
 (47)

where $tr[\cdot]$ means trace of $[\cdot]$.



By using (44), (45) and (47), it follows that

$$E[\Delta \mathbf{w}^{T}(k)\mathbf{R}\Delta \mathbf{w}(k)] = \lambda^{2} E[\Delta \mathbf{w}^{T}(k-1)\mathbf{R}\Delta \mathbf{w}(k-1)] + (1-\lambda)^{2} \operatorname{tr} \{\mathbf{I} + E[\Delta \mathbf{I}^{2}(k)]\}\xi_{\min}$$
(48)

Asymptotically, the solution of the equation above is

$$\xi_{\text{exc}} = \frac{1 - \lambda}{1 + \lambda} \text{tr} \left\{ \mathbf{I} + E[\Delta \mathbf{I}^2(k)] \right\} \xi_{\text{min}}$$
 (49)

The term $E[\Delta \mathbf{I}^2(k)]$ is dependent on fourth-order statistics of the input signal. Assume diagonal dominance of \mathbf{R} and $\Delta \mathbf{R}$, and use the definition of $\Delta \mathbf{I}(k)$



$$E[\Delta \mathbf{I}_{ii}^2(k)] = (1 - \lambda) \frac{E[\Delta r_{ii}^2(k)]}{[\sigma_x^2]^2}$$

$$(50)$$

where σ_x^2 is variance of x(k).

$$\Delta r_{ii}(k) = \lambda \Delta r_{ii}(k-1) + x(k-i)x(k-i) - r_{ii}$$
 (51)

and using the independence between $\Delta r_{ii}(k)$ and x(k) it follows asymptotically

$$E[\Delta r_{ii}^{2}(k)] = \frac{1}{1 - \lambda^{2}} \sigma_{x^{2}(k-i)}^{2} = \frac{1}{1 - \lambda^{2}} \sigma_{x^{2}}^{2}$$
 (52)



By substituting eq. (52) in (50)

$$E[\Delta \mathbf{I}_{ii}^2] = \frac{1 - \lambda}{1 + \lambda} \frac{\sigma_{x^2}^2}{\sigma_x^2} = \frac{1 - \lambda}{1 + \lambda} \mathcal{K}$$
 (53)

where \mathcal{K} is dependent on input signal statistics. For Gaussian signals $\mathcal{K}=2$.

Returning to our main objective, the excess of MSE is

$$\xi_{\text{exc}} = (N+1)\frac{1-\lambda}{1+\lambda}(1+\frac{1-\lambda}{1+\lambda}\mathcal{K})\xi_{\text{min}}$$
 (54)



If $\lambda \approx 1$ and \mathcal{K} is not very large then

$$\xi_{\text{exc}} = (N+1)\frac{1-\lambda}{1+\lambda}\xi_{\min} \tag{55}$$

The misadjustment can be deduced from (54)

$$M = (N+1)\frac{1-\lambda}{1+\lambda}(1+\frac{1-\lambda}{1+\lambda}\mathcal{K})$$
 (56)



Behavior in Nonstat. Environ.

If the input signal and/or the desired signal are nonstationary, the optimal coefficients are time variant $\mathbf{w}_o(k)$. That autocorrelation matrix $\mathbf{R}(k)$ and/or the crosscorrelation vector $\mathbf{p}(k)$ are time variant.



Recall that

$$\mathbf{w}(k) = \mathbf{w}(k-1) + \mathbf{S}_D(k)\mathbf{x}(k)(d(k) - \mathbf{x}^T(k)\mathbf{w}(k-1))$$
(57)

and

$$d(k) = \mathbf{x}^{T}(k)\mathbf{w}_{o}(k-1) + e'_{o}(k)$$

$$\tag{58}$$



one can replace eq. (58) in (57) then

$$\mathbf{w}(k) = \mathbf{w}(k-1) + \mathbf{S}_D(k)\mathbf{x}(k)\mathbf{x}^T(k)[\mathbf{w}_o(k-1) - \mathbf{w}(k-1)] + \mathbf{S}_D(k)\mathbf{x}(k)e'_o(k)$$
(59)

Consider that $\mathbf{x}(k)$ and $e'_o(k)$ are approximately orthogonal and $\mathbf{w}(k-1)$ is independent of $\mathbf{x}(k)$, then

$$E[\mathbf{w}(k)] = E[\mathbf{w}(k-1)] + E[\mathbf{S}_D(k)\mathbf{x}(k)\mathbf{x}^T(k)]$$

$$(\mathbf{w}_o(k-1) - E[\mathbf{w}(k-1)])$$
(60)



It is needed to compute $E[\mathbf{S}_D(k)\mathbf{x}(k)\mathbf{x}^T(k)]$, for nonstationary

environment. In this case

$$\mathbf{R}_{D}(k) = \sum_{l=0}^{k} \lambda^{k-l} \mathbf{R}(l) + \Delta \mathbf{R}(k)$$
 (61)

since
$$E[\mathbf{R}_D(k)] = \sum_{l=0}^k \lambda^{k-l} \mathbf{R}(l)$$
.

If the environment is varying in a slower pace than the memory of the adaptive RLS algorithm then

$$\mathbf{R}_D(k) \approx \frac{1}{1-\lambda} \mathbf{R}(k) + \Delta \mathbf{R}(k) \tag{62}$$



Considering that $(1 - \lambda)||\mathbf{R}^{-1}(k)\Delta\mathbf{R}(k)|| < 1$ then

$$\mathbf{S}_D(k) \approx (1 - \lambda)\mathbf{R}^{-1}(k) - (1 - \lambda)^2\mathbf{R}^{-1}(k)\Delta\mathbf{R}(k)\mathbf{R}^{-1}(k)$$
(63)

it then follows that

$$E[\mathbf{w}(k)] = E[\mathbf{w}(k-1)] + \{(1-\lambda)E[\mathbf{R}^{-1}(k)\mathbf{x}(k)\mathbf{x}^{T}(k)]$$

$$+(1-\lambda)^{2}E[\mathbf{R}^{-1}(k)\Delta\mathbf{R}(k)\mathbf{R}^{-1}(k)\mathbf{x}(k)\mathbf{x}^{T}(k)]\}$$

$$(\mathbf{w}_{o}(k-1) - E[\mathbf{w}(k-1)])$$

$$\approx E[\mathbf{w}(k-1)] + (1-\lambda)(\mathbf{w}_{o}(k-1) - E[\mathbf{w}(k-1)])$$
(64)

where $\Delta \mathbf{R}(k)$ is independent of $\mathbf{x}(k)$ and has zero expected value.



Now defining the lag error vector in the coefficients as

$$\mathbf{l}_{\mathbf{W}}(k) = E[\mathbf{w}(k)] - \mathbf{w}_o(k) \tag{65}$$

from equation (64) it can be concluded that

$$\mathbf{l}_{\mathbf{W}}(k) = \lambda \mathbf{l}_{\mathbf{W}}(k-1) - \mathbf{w}_o(k) + \mathbf{w}_o(k-1)$$
 (66)

This eq. is equivalent to pass the optimal value $\mathbf{w}_o(k)$ through a first order filter

$$L_i(z) = -\frac{z-1}{z-\lambda} W_{oi}(z) \tag{67}$$



The filter transient response converges with a time constant

$$\tau = \frac{1}{1 - \lambda} \tag{68}$$

To calculate the MSE suppose that the coefficients values are a first order Markov process

$$\mathbf{w}_o(k) = \lambda_{\mathbf{W}} \mathbf{w}_o(k-1) + \mathbf{n}_{\mathbf{W}}(k) \tag{69}$$

where the elements of $\mathbf{n}_{\mathbf{W}}(k)$ are zero mean Gaussian noise processes with variance $\sigma_{\mathbf{W}}^2$, and $\lambda_{\mathbf{W}} < 1$.



 $\lambda < \lambda_{\mathbf{W}} < 1$, since the optimal coefficients must vary slower than the filter tracking speed, i.e. $\frac{1}{1-\lambda} << \frac{1}{1-\lambda_{\mathbf{W}}}$.

The excess of MSE due to lag is

$$\xi_{\text{lag}} = E[\mathbf{l}_{\mathbf{W}}^{T}(k)\mathbf{R}\mathbf{l}_{\mathbf{W}}(k)]$$

$$= E[\text{tr}(\mathbf{R}\mathbf{l}_{\mathbf{W}}(k)\mathbf{l}_{\mathbf{W}}^{T}(k))]$$

$$= \text{tr}(\mathbf{R}E[\mathbf{l}_{\mathbf{W}}(k)\mathbf{l}_{\mathbf{W}}^{T}(k)])$$
(70)



If $\lambda_{\mathbf{W}} \approx 1$ then from (67) and (69)

$$\xi_{\text{lag}} \approx \operatorname{tr}\left[\mathbf{R} \cdot \frac{\sigma_{\mathbf{w}}^{2}}{1 - \lambda^{2}} \mathbf{I}\right] = \frac{\sigma_{\mathbf{w}}^{2}}{1 - \lambda^{2}} \operatorname{tr} \mathbf{R}$$

$$= \frac{(N+1)\sigma_{\mathbf{w}}^{2} \sigma_{x}^{2}}{1 - \lambda^{2}}$$
(71)



For general $\lambda_{\mathbf{W}}$ it can be shown that

$$\xi_{\text{lag}} \approx \frac{(N+1)\sigma_{\mathbf{w}}^{2}\sigma_{x}^{2}}{\lambda_{\mathbf{w}}(1+\lambda^{2}) - \lambda(1+\lambda_{\mathbf{w}}^{2})} \left(\frac{1-\lambda}{1+\lambda} - \frac{1-\lambda_{\mathbf{w}}}{1+\lambda_{\mathbf{w}}}\right)$$
(72)

If $\lambda = 1$ and $\lambda_{\mathbf{W}} \approx 1$, the MSE due to lag tends to infinity indicating that the RLS cannot track any change in the environment.



Since $\lambda \approx 1$ we can rewrite (72) as

$$\xi_{\text{lag}} \approx (N+1) \frac{\sigma_{\mathbf{w}}^2}{2(1-\lambda)} \sigma_x^2 \tag{73}$$

The total excess MSE accounting for the lag and finite memory is

$$\xi_{\text{total}} \approx (N+1) \left[\frac{1-\lambda}{1+\lambda} \xi_{\text{min}} + \frac{\sigma_{\mathbf{w}}^2 \sigma_x^2}{2(1-\lambda)} \right]$$
 (74)



By differentiating with respect to λ , setting the result to zero, an optimum value for λ can be found

$$\lambda_{\text{opt}} = \frac{1 - \frac{\sigma_{\mathbf{W}}\sigma_x}{2\sigma_n}}{1 + \frac{\sigma_{\mathbf{W}}\sigma_x}{2\sigma_n}} \tag{75}$$

It can be used when it falls in an acceptable range of values for λ .



In the complex data case the objective function is

$$\xi^{d}(k) = \sum_{i=0}^{k} \lambda^{k-i} |\varepsilon(i)|^{2} = \sum_{i=0}^{k} \lambda^{k-i} |d(i) - \mathbf{w}^{H}(i)\mathbf{x}(k)|^{2}$$
$$= \sum_{i=0}^{k} \lambda^{k-i} \left[d(i) - \mathbf{w}^{H}(i)\mathbf{x}(k)\right] \left[d^{*}(i) - \mathbf{w}^{T}(i)\mathbf{x}^{*}(k)\right]$$
(76)

Differentiating $\xi^d(k)$ with respect to the complex coefficient $\mathbf{w}(k)$ leads to

$$\frac{\partial \xi^d(k)}{\partial \mathbf{w}(k)} = -\sum_{i=0}^k \lambda^{k-i} \mathbf{x}^*(i) [d(i) - \mathbf{w}^H(i) \mathbf{x}(k)]$$
 (77)

The optimal vector is computed by equating the equation above to zero that is

$$-\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{H}(i) \mathbf{w}(k) + \sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) d^{*}(i) = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

leading to the following expression

$$\mathbf{w}(k) = \left[\sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) \mathbf{x}^{H}(i)\right]^{-1} \sum_{i=0}^{k} \lambda^{k-i} \mathbf{x}(i) d^{*}(i)$$
$$= \mathbf{R}_{D}^{-1}(k) \mathbf{p}_{D}(k)$$
(78)



The matrix inversion lemma to the case of complex data is given by

$$\mathbf{S}_{D}(k) = \mathbf{R}_{D}^{-1}(k) = \frac{1}{\lambda} \left[\mathbf{S}_{D}(k-1) - \frac{\mathbf{S}_{D}(k-1)\mathbf{x}(k)\mathbf{x}^{H}(k)\mathbf{S}_{D}(k-1)}{\lambda + \mathbf{x}^{H}(k)\mathbf{S}_{D}(k-1)\mathbf{x}(k)} \right]$$
(79)



Algorithm 5.3 Conventional Complex RLS Algorithm

Initialization

$$\mathbf{S}_D(-1) = \delta \mathbf{I}$$

where δ can be the inverse of the input signal power estimate

$$\mathbf{p}_D(-1) = \mathbf{x}(-1) = [0 \ 0 \dots 0]^T$$



Do for $k \ge 0$:

$$\mathbf{S}_{D}(k) = \frac{1}{\lambda} [\mathbf{S}_{D}(k-1) - \frac{\mathbf{S}_{D}(k-1)\mathbf{X}(k)\mathbf{X}^{H}(k)\mathbf{S}_{D}(k-1)}{\lambda + \mathbf{X}^{H}(k)\mathbf{S}_{D}(k-1)\mathbf{X}(k)}]$$

$$\mathbf{p}_{D}(k) = \lambda \mathbf{p}_{D}(k-1) + d^{*}(k)\mathbf{x}(k)$$

$$\mathbf{w}(k) = \mathbf{S}_{D}(k)\mathbf{p}_{D}(k)$$

If necessary compute

$$y(k) = \mathbf{w}^{H}(k)\mathbf{x}(k)$$
$$\varepsilon(k) = d(k) - y(k)$$



An alternative complex RLS algorithm has an updating equation described by

$$\mathbf{w}(k) = \mathbf{w}(k-1) + e^*(k)\mathbf{S}_D(k)\mathbf{x}(k)$$
(80)

where

$$e(k) = d(k) - \mathbf{w}^{H}(k-1)\mathbf{x}(k)$$
(81)



Algorithm 5.4 Alternative Complex RLS Algorithm

Initialization

$$\mathbf{S}_D(-1) = \delta \mathbf{I}$$

where δ can be the inverse of an estimate of the input signal power

$$\mathbf{x}(-1) = \mathbf{w}(-1) = [0 \ 0 \dots 0]^T$$



Do for $k \geq 0$

$$e(k) = d(k) - \mathbf{w}^{H}(k-1)\mathbf{x}(k)$$
$$\boldsymbol{\psi}(k) = \mathbf{S}_{D}(k-1)\mathbf{x}(k)$$



$$\mathbf{S}_{D}(k) = \frac{1}{\lambda} [\mathbf{S}_{D}(k-1) - \frac{\boldsymbol{\psi}(k)\boldsymbol{\psi}^{H}(k)}{\lambda + \boldsymbol{\psi}^{H}(k)\mathbf{X}(k)}]$$
$$\mathbf{w}(k) = \mathbf{w}(k-1) + e^{*}(k)\mathbf{S}_{D}(k)\mathbf{x}(k)$$

If necessary compute

$$y(k) = \mathbf{w}^{H}(k)\mathbf{x}(k)$$
$$\varepsilon(k) = d(k) - y(k)$$



Example 5.3

Assume that an adaptive filter of sufficient order is employed to identify an unknown system of order N, and produces a misadjustment of 10%. Assume the input signal is a white Gaussian noise with unit variance and $\sigma_n^2 = 0.001$.

- (a) Compute the value of λ required by the RLS algorithm in order to achieve the desired result when N=9.
- (b) For values in the range $0.9 < \lambda < 0.99$, which orders should the adaptive filters have?



Solution:

(a) The desired misadjustment expression is

$$M = 0.1 = (N+1)\frac{1-\lambda}{1+\lambda} \left(1 + \frac{1-\lambda}{1+\lambda}\mathcal{K}\right) = 10a(1+2a)$$

where $a = \frac{1-\lambda}{1+\lambda}$ and $\mathcal{K} = 2$. By solving this equation we obtain

$$a = \frac{-\frac{1}{2} \pm \sqrt{\frac{1}{4} + 0.02}}{2}$$

where the valid solution is

$$a = \frac{1}{4} \left(-1 + \sqrt{1 + 0.08} \right) = 0.0098076$$



Then solving for λ

$$\lambda = \frac{1 - a}{1 + a} = 0.980507$$

By employing the simplest expression we obtain

$$\lambda = \frac{1 - \frac{M}{(N+1)}}{1 + \frac{M}{(N+1)}} = \frac{1 - 10^{-2}}{1 + 10^{-2}} = 0.98$$

where M is the misadjustment.



Solution:

(b) Since

$$\frac{1}{N+1} = \frac{1}{M} \frac{1-\lambda}{1+\lambda} \left(1 + \frac{1-\lambda}{1+\lambda} \mathcal{K} \right) = 10a(1+2a)$$

for $\lambda = 0.90$, a = 0.052631578

$$\frac{1}{N+1} = 0.5817$$

so that N = 0.7190 and as a result only one coefficient can be employed in the adaptive filter.



For $\lambda = 0.99$, a = 0.005025125,

$$\frac{1}{N+1} = 0.05075$$

so that N=18.7 and as a result 19 coefficients can be employed in the adaptive filter.



Using the simplest expression for M the results are almost the same since

$$N = M \frac{1+\lambda}{1-\lambda} - 1$$

for $\lambda = 0.90$, N = 0.9 meaning that only an adaptive filter with one coefficient would be able to a achieve the desired misadjustment for this value of λ . For $\lambda = 0.99$, N = 18.9 meaning that adaptive filters up to order 18 would be able to a achieve the desired misadjustment for this value of λ .



System Identification Simulation

The conventional RLS algorithm was employed in the identification of the system discussed in the previous Chapter. The forgetting factor was chosen $\lambda = 0.99$.



The learning curves of the MS *a priori* error are depicted in Fig. 4,

for different values of the eigenvalue spread. Also the measured misadjustment in each example is given in Table 1. From these results we conclude that the RLS algorithm is insensitive to the eigenvalue spread. On Table 1 is given the misadjustment predicted by theory. As can be seen the analytical results are quite accurate.

$$M = (N+1)\frac{1-\lambda}{1+\lambda}(1+\frac{1-\lambda}{1+\lambda}\mathcal{K}) \tag{82}$$



Table 2 summarizes the results. Note the close agreement between

the measurement results and those predicted by the equations

$$E[||\Delta \mathbf{w}(k)_Q||^2] \approx \frac{(1-\lambda)(N+1)}{2\lambda} \frac{\sigma_n^2 + \sigma_e^2}{\sigma_x^2} + \frac{(N+1)\sigma_{\mathbf{w}}^2}{2\lambda(1-\lambda)}$$
(83)

$$\xi(k)_Q \approx \xi_{min}^2 + \sigma_e^2 + \frac{(N+1)\sigma_{\mathbf{W}}^2 \sigma_x^2}{2\lambda(1-\lambda)}$$
(84)

For the simulations with 12 and 10 bits, the discrepancy between the measured and theoretical estimates of $E[||\Delta \mathbf{w}(k)_Q||^2]$ are caused by the freezing of some tap coefficients.



The optimal value of λ was too small, $\lambda = 0.99$ was used. The measured excess of MSE was 0.0254, the theoretical value predicted was 0.0418. The theoretical result is not as accurate as before, due to a number of approximations used in the analysis. However, the equation below provides a good indication of what is expected in the practical implementation.

$$\xi_{\text{exc}} \approx (N+1)\frac{1-\lambda}{1+\lambda}(1+\frac{1-\lambda}{1+\lambda}\mathcal{K})\xi_{\text{min}} + \frac{(N+1)\sigma_{\mathbf{W}}^{2}\sigma_{x}^{2}}{\lambda_{\mathbf{W}}(1+\lambda^{2}) - \lambda(1+\lambda_{\mathbf{W}}^{2})}(\frac{1-\lambda}{1+\lambda} - \frac{1-\lambda_{\mathbf{W}}}{1+\lambda})$$
(85)



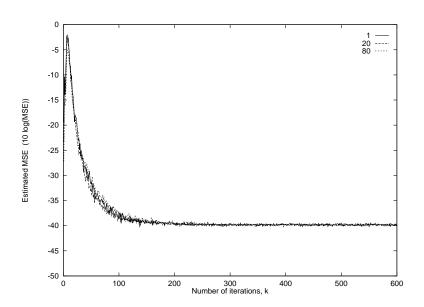


Figure 4: Learning curves for RLS algorithm for eigenvalue spreads: 1, 20, and 80; $\lambda = 0.99$.



Table 1: Evaluation of the RLS Algorithm

	Misadjustment		
$rac{\lambda_{max}}{\lambda_{min}}$	Experiment	Theory	
1	0.04211	0.04020	
20	0.04211	0.04020	
80	0.04547	0.04020	



Table 2: Finite Precision Implementation of the RLS Algorithm

	$\xi(k)_Q$		$E[\Delta \mathbf{w}(k)_Q ^2]$	
No. of bits	Experiment	Theory	Experiment	Theory
16	$1.566 \ 10^{-3}$	$1.500\ 10^{-3}$	$6.013 \ 10^{-5}$	6.061 10 ⁻⁵
12	$1.522 \ 10^{-3}$	$1.502\ 10^{-3}$	$3.128 \ 10^{-5}$	$6.261 \ 10^{-5}$
10	$1.566\ 10^{-3}$	$1.532 \ 10^{-3}$	$6.979 \ 10^{-5}$	$9.272 \ 10^{-5}$



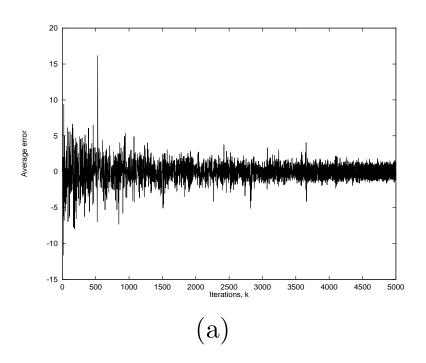
Signal Enhancement Simulations

LMS and RLS Algorithms

We solved the same signal enhancement problem with the conventional RLS and LMS algorithms.

An appropriate value for μ in the LMS case is 0.0204, whereas $\lambda = 1.0$ was used for the RLS. The learning curves for the algorithms are shown in Fig. 5, we see the faster convergence of the RLS. Plotting the output errors after convergence we noted the larger variance of the MSE for the RLS. This result is due to the small signal to noise ratio. Fig. 6 depicts the output error for the RLS.







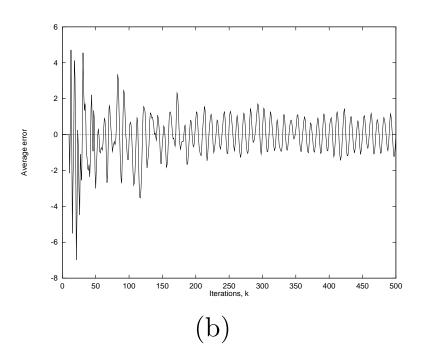
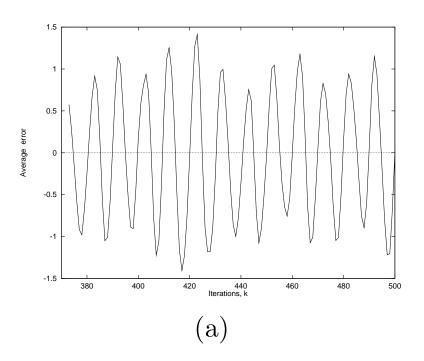


Figure 5: Learning curves for the (a) LMS and (b) RLS algorithms.







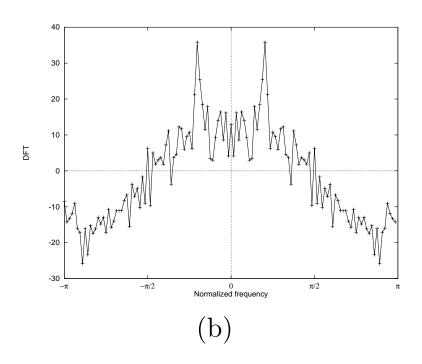


Figure 6: (a) Output error for the RLS algorithm and (b) DFT of the output error.