Wishart Distribution & Hotelling's T2 test

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```
library(car) # scatterplotMatrix(...)
```

Loading required package: carData

Problem 1: Choose a 3-by-3 covariance matrix with non-zero covariances (the off-diagonal elements should not be 0). Also a choose a sample size n (e.g., n=100, 500, 1000, etc.). Using the covariance matrix you chose, simulate 1,000 data sets from a trivariate normal distribution.

1.1

Creating a covariance matrix:

```
set.seed(2) # Set the seed for reproducibility
m <- matrix(runif(9, -1, 1), nrow = 3, ncol = 3) # 9 random numbers between (-1 and 1)
Sigma <- crossprod(m) # Create a symmetric matrix</pre>
# Manipulating the values to give different types of correlations
Sigma[1, 2] \leftarrow Sigma[2, 1] \leftarrow 0.7
Sigma[1, 3] \leftarrow Sigma[3, 1] \leftarrow -0.1
Sigma[2, 3] <- Sigma[3, 2] <- 0.1
Sigma
##
                         [,2]
               [,1]
                                   [,3]
## [1,] 0.5825248 0.700000 -0.100000
## [2,] 0.7000000 2.015412 0.100000
## [3,] -0.1000000 0.100000 0.998936
round(eigen(Sigma)$values, 2) # Check all of the eigenvalues are positive
## [1] 2.30 1.02 0.27
#cor(Sigma)
```

Simulation Using the covariance matrix above, simulate 1,000 data sets from a trivariate normal distribution. The function returns an array simulation of size T where element, X_i , is a (nxp) Trivariate Normal Matrix,

```
library(MASS) # mvrnorm(...)
create_trivariate_norm <- function(T, n, p, Sigma) {
    simulation <- array(0, c(n, 3, T))
    for (t in 1:T) {
        X <- mvrnorm(n, rep(0, p), Sigma)
        simulation[, , t] <- X
    }
    return (simulation)
}</pre>
```

```
p <- 3
T <- 1000
n <-1000
simulation <- create_trivariate_norm(T=T, n=n, p=p, Sigma=Sigma)</pre>
head(simulation[, , 1:1]) # show head of 1 element
##
               [,1]
                            [,2]
                                       [,3]
## [1,] -0.53998012  0.05359939 -0.9002465
## [2,] 0.93330577 0.84812834 -1.3662784
## [3,] 0.49569373 1.19344562 2.2044041
## [4,] -0.89115617 -1.35083552 -0.2640324
## [5,] 0.76822026 0.95473478 -0.7041917
## [6,] -0.02500744 2.42210916 -1.1682155
Pairwise scatterplot of simulated data (simulation)
scatterplotMatrix(simulation[, , 1])
\alpha
0
4
                                                              -1 0 1 2 3
                    1
                         2
                                                       -3
cor(simulation[, , 1]) # correlation matrix
              [,1]
                        [,2]
## [1,] 1.0000000 0.6551686 -0.1087917
## [2,] 0.6551686 1.0000000 0.1137658
## [3,] -0.1087917 0.1137658 1.0000000
cov(simulation[, , 1]) # covariance matrix
##
               [,1]
                          [,2]
                                      [,3]
## [1,] 0.62589869 0.7689229 -0.08502275
## [2,] 0.76892291 2.2006727 0.16671582
## [3,] -0.08502275 0.1667158 0.97583089
```

- 1.1. Try to make sense of the covariance matrix by examining the pairwise scatter plots using the data you simulate.
 - The diagnol of the scatter plot matrix above plots the marginal distribution of the first element in sim1 and it can be observed that each marginal distribution of X_i is a normal distribution. The off-diagonol elements of the scatter plot matrix plots the joint distribution between random variables, X_i and X_j . Below the pairwise plots, is the correlation and covariance matrices respectively of the sample. It cna be observed that the plots align with the correlation values. For example, X_1 and X_2 have a positive correlation of 0.655. Their respective scatter plots have a diagonal line that slopes upward from left to right, suggesting a moderately positive linear relationship. Additionally X_1 and X_2 have higher covariance values relative to the other pairwise joint distributions, and so the data points are narrowly spread across the joint distribution. Conversely, X_1 and X_3 have low covariance and a low negative correlation. This is reflected by a high degree of spread in data points and a low negative slope.

1.2

Compute and store S, the sample covariance.

```
S.array <- array(apply(simulation, p, cov), c(p, p, T))
```

Compute and store the traces.

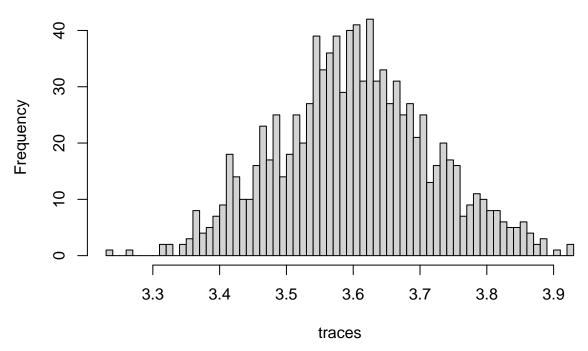
```
traces <- apply(S.array, p, function(x) sum(diag(x)))
traces[1:3]</pre>
```

[1] 3.802402 3.641897 3.725560

Distribution of the traces

```
hist(traces, breaks = 50)
```

Histogram of traces

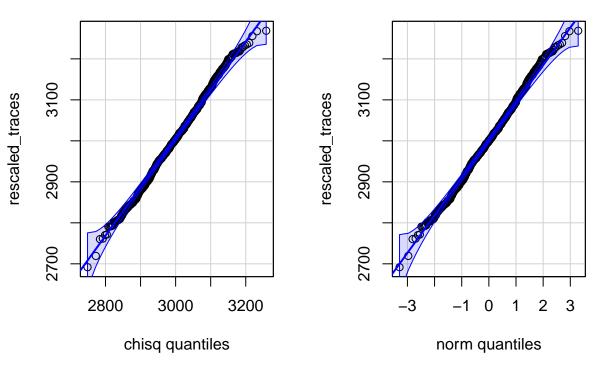


QQ Plots of χ^2 and \mathcal{N} Distributions on re-scaled data.

```
rescaled_traces <- (n-1)*traces/mean(diag(Sigma))
par(mfrow = c(1, 2))
qqPlot(rescaled_traces, distribution = "chisq", df=3*(n-1), id = F, main = "Chi2 QQ Plot")
qqPlot(rescaled_traces, id = F, main = "Normal QQ Plot")</pre>
```

Chi2 QQ Plot

Normal QQ Plot

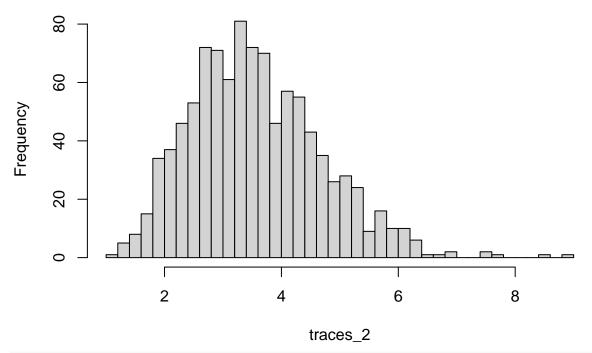


- 1.2 During the simulation, you will generate 1,000 Wishart distributed random matrices. Calculate the trace for each of them. Explain what distribution the traces should follow and examine their histogram.
 - We would expect that the (rescaled) trace would follow a χ^2 distribution or a linear combination of χ^2 , but the histogram looks like a bell-shaped distribution suggesting a \mathcal{N} distribution. This is because the Central Limit Theorem (CLT) effect kicks in when dealing with sums. The CLT tells us that the sum of random variables tend to be normally distributed as the number of elements in the sum increase. Thus, a χ^2 with very high degrees of freedom, (n=1000 above) tends to become a normal distribution.

To further test if CLT effect occurred, a simulation with a much smaller size (n=10) will be conducted to see if the trace plot and QQ-plots suggest a \mathcal{N} distribution.

```
p <- 3
T <- 1000
n <-10 # 100 instead of 10 this time
simulation_2 <- create_trivariate_norm(T=T, n=n, p=p, Sigma=Sigma)
S_2.array <- array(apply(simulation_2, p, cov), c(p, p, T))
traces_2 <- apply(S_2.array, p, function(x) sum(diag(x)))
hist(traces_2, breaks = 50)</pre>
```

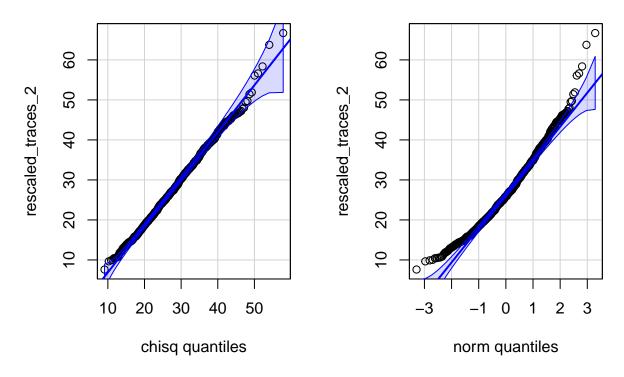
Histogram of traces_2



```
rescaled_traces_2 <- (n-1)*traces_2/mean(diag(Sigma))
par(mfrow = c(1, 2))
qqPlot(rescaled_traces_2, distribution = "chisq", df=3*(n-1), id = F, main = "Chi2 QQ Plot")
qqPlot(rescaled_traces_2, id = F, main = "Normal QQ Plot")</pre>
```

Chi2 QQ Plot

Normal QQ Plot



The simulation above used the same covariance matrix as the original simulation however only 10 random variables were used to generate a multivariate normal distribution for each t_1 sample. The histogram of traces is no longer a bell-shape distribution which suggests not a normal distribution. Additionally the rescaled traces follow the χ^2 distribution in QQ-plot but do not follow a $\mathcal N$ distribution in the QQ-plot so well. This simulation is evidence that the first simulation results were affected by the Central Limit Theorem.

Problem 2 starts on next page.

Problem 2: Find a good data example to conduct a two-sample Hotelling's T^2 test. Do not use the data example discussed in this course. Please (1) include visualizations as exploratory methods and (2) make conclusion in the context of the data example.

```
conclusion in the context of the data example.
\# X = random population 1, Y = random population 2
Hotelling.T2.2sample=function(X, Y) {
  n = dim(X)[1]; m=dim(Y)[1]; p=dim(X)[2]
  if(p!= dim(Y)[2])
    return ("Error: the dimensions of X and Y are not the same")
  X.bar=colMeans(X); Y.bar=colMeans(Y)
  X.S=cov(X); Y.S=cov(Y)
  pooled.S=((n-1)*X.S+(m-1)*Y.S)/(m+n-2)
  T2=t(X.bar-Y.bar)%*%solve((1/n+1/m)*pooled.S)%*%(X.bar-Y.bar)
  p.value=1-pf(T2/((n+m-2)*p/(n+m-1-p)),p,n+m-1-p)
  return (list(X.bar=X.bar, Y.bar=Y.bar, T2=T2, p.value=p.value))
}
# dual function Hotelling.T2
Hotelling.T2=function(X, Y=NULL, mu0=NULL) {
  if(is.null(Y) && is.null(mu0))
    return("Error: mu0 is not specified")
  if(!is.null(X) && !is.null(mu0))
    obj=Hotelling.T2.1sample(X, mu0)
  if(!is.null(X) && !is.null(Y))
    obj=Hotelling.T2.2sample(X,Y)
  return (obj)
}
Using built-in R dataset, mtcars.
data("mtcars")
head(mtcars)
```

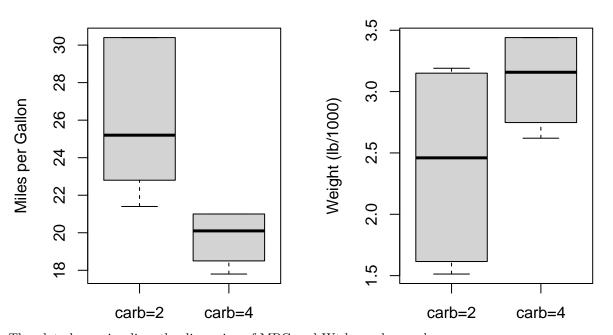
```
wt qsec vs am gear carb
##
                     mpg cyl disp hp drat
## Mazda RX4
                     21.0
                            6 160 110 3.90 2.620 16.46
## Mazda RX4 Wag
                            6 160 110 3.90 2.875 17.02 0
                                                                 4
                                                                      4
                     21.0
                                                           1
## Datsun 710
                     22.8 4 108 93 3.85 2.320 18.61 1 1
                                                                      1
## Hornet 4 Drive
                     21.4
                            6 258 110 3.08 3.215 19.44 1 0
                                                                 3
                                                                      1
                            8 360 175 3.15 3.440 17.02
                                                                      2
## Hornet Sportabout 18.7
                                                         0
                                                            0
                                                                 3
## Valiant
                            6 225 105 2.76 3.460 20.22 1 0
                     18.1
                                                                      1
# subset mtcars to only include samples with 4 or 6 cylinders
subset_mtcars <- mtcars[mtcars$cyl %in% c(4,6),]</pre>
# carb = Number of carburetors
# create subset X with samples of carb = 2, keeping only mpg and wt columns
X <- subset_mtcars[subset_mtcars$carb == 2, c("mpg", "wt")]</pre>
# create subset Y with samples of carb = 4, keeping only mpg and wt columns
Y <- subset_mtcars[subset_mtcars$carb == 4, c("mpg", "wt")]
n1 <- nrow(X); n2 <- nrow(Y)</pre>
n1;n2
## [1] 6
## [1] 4
```

2.1 Explatory Data Analysis on differences between X and Y.

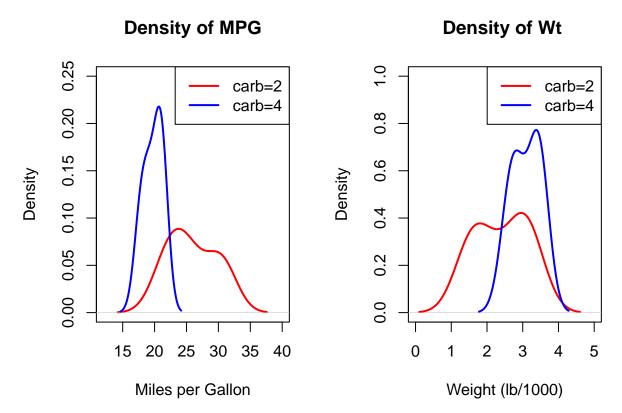
For the 2 groups, mpg (miles per gallon) and wt (weight - lb/1000), will be explored below.

MPG Distribution

Wt Distribution



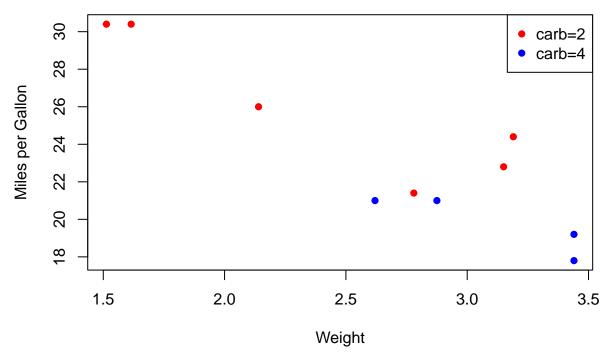
The plot above visualizes the dispersion of MPG and Wt by each sample group.



It can be observed from the marginal density plots that the variables appear to be approximately marginally normally distributed due to bell-shaped curves.

```
# scatterplot of mpg vs. wt for X and Y datasets
plot(X$wt, X$mpg, col = "red", pch = 16,
    main = "Joint Density (MPG vs Weight)",
    xlab = "Weight", ylab = "Miles per Gallon",
    xlim = range(c(X$wt, Y$wt)), ylim = range(c(X$mpg, Y$mpg)))
points(Y$wt, Y$mpg, col = "blue", pch = 16)
legend("topright", legend = c("carb=2", "carb=4"),
    col = c("red", "blue"), pch = 16)
```

Joint Density (MPG vs Weight)



It is difficult to assess if the variables, MPG and Wt, are jointly normally distributed because there are few samples. The ellispe-like shape and lack of tight clusters does support the notion that there is little evidence to say the features are not jointly normally distributed.

Calculating mean differences among features between sample groups, sample variances, and pooled variance.

```
p=2
# calculate mean of each sample's features
mean1=matrix(colMeans(X, p, 1))
mean2=matrix(colMeans(Y, p, 1))
mean1; mean2
##
          [,1]
## [1,] 25.900
## [2,] 2.398
##
            [,1]
## [1,] 19.75000
## [2,] 3.09375
# calculate difference in means between samples' features
mean.diff = mean1-mean2
mean.diff
            [,1]
## [1,] 6.15000
## [2,] -0.69575
# calculate each sample variance
S1=cov(X); S2=cov(Y);
# pooled variance
Sp=((n1-1)*S1+(n2-1)*S2)/(n1+n2-2)
Sp
```

```
## mpg wt
## mpg 9.981250 -1.8126562
## wt -1.812656 0.4142048
```

Performing the Hotelling's T^2 test on X and Y.

Hotelling.T2(X, Y)

```
## $X.bar
##
      mpg
               wt
## 25.900
           2.398
##
## $Y.bar
##
                   wt
        mpg
## 19.75000
             3.09375
##
## $T2
##
             [,1]
## [1,] 14.10104
##
## $p.value
##
               [,1]
## [1,] 0.02853453
```

2.2

The p-value is 0.0285. At a significance level of 0.05, we can conclude that the difference between the mean vectors (mpg, wt) of the two groups, 2 carburetor vehicles and 4 carburetor vehicles, is statistically significant. It should be noted that the p-value is close to exceeding the significance level which means the evidence supporting the rejection of the null is not very strong, but still strong enough to support the claim at a significance level of 0.05.