

The equity and efficiency effects of a relative GPA reward in college admissions

Tatiana Reyes*

November 30, 2022

[Click here for the latest version.](#)

Abstract

Many college admissions systems use a combination of GPA and standardized test scores to determine access to more selective programs. In this paper, I study the impacts of a 2013 reform in the Chilean admission system that sought to increase equity by introducing a third component, based on a student's GPA relative to the historical average at their high school. Simulating the admission mechanism with and without the relative GPA boost, I classify applicants into three groups: (i) those who gained access to more selective programs (pulled-up), (ii) those who lost access to more selective programs (pushed-down), and (iii) those whose admission was unaffected. Applying the same procedure in earlier years, I identify the same groups, facilitating a difference-in-differences design to estimate the impacts of the 2013 reform on enrollment, persistence, and graduation. Pulled-up students were able to persist in their newly accessed programs, resulting in more selective degree attainment with no effect on overall BA completion. Pushed-down students, who tended to come from better-educated/higher-income families, experienced comparable-sized reductions in the probability of graduating from selective programs, offset by gains in graduation from less selective programs. I conclude that the reform improved equity with little or no loss in efficiency.

*University of California, Berkeley. Email: tatiana_reyes@berkeley.edu. I am thankful to my advisors David Card, Jesse Rothstein and Christopher Walters for their support and guidance. I would also like to give special thanks to Hadar Avivi, Livia Alfonsi, Monica Saucedo, Damián Vergara, Pablo Muñoz, Harrison Wheeler, and Marina Dias for their thoughtful discussion of this paper at its earliest stage. This paper also benefited from discussion and suggestions from Sebastián Otero, Zach Bleemer, and the participants from the UC Berkeley Labor Seminar.

Introduction

The notion of higher education, especially at selective colleges, as a vehicle for upward social mobility makes the issue of access to these programs policy relevant (Autor, 2014; Chetty et al., 2017; Turner, 2020). Admission criteria at selective colleges typically rely on a combination of standardized tests and high school grades, but consistent evidence of test score disparities between students from different backgrounds raises concerns about the equity implications of these rules (J. M. Rothstein, 2004; Card & Rothstein, 2007; Zwick & Greif Green, 2007). Interventions such as top-percent programs and affirmative action policies are examples of systemic efforts to narrow admission gaps between students from different backgrounds.¹ However, there is wide disagreement on the effects of such interventions on the students they are designed to help, and on other students who are potentially harmed by the introduction of preferences for disadvantaged students.²

In this paper I use detailed student records, combined with the admissions formulas used by selective college programs in Chile, to evaluate the equity and efficiency impacts of a 2013 reform designed to improve access to the country’s most selective programs for students from disadvantaged high schools. Prior to the reform, students submitted ranked lists of selective college programs to a centralized system using a single offer deferred acceptance (DA) algorithm to rank students and allocate offers of admission.³ Each college program (e.g., Mechanical Engineering at University of Chile) used a combination of high school GPA and scores on a standardized test (the “PSU” test) to rank students. The 2013 reform introduced a third component, based on the difference between the student’s GPA and the historical mean GPA at her high school. This “GPA⁺” component

¹For California’s “Eligibility in the Local Context” see Bleemer (2021), and S. E. Black et al. (2020) for Texas Top Percent Policy. For Brazil’s affirmative action Otero et al. (2021) and Mello (2022) and Bagde et al. (2016) study an affirmative action policy in India.

²Dillon & Smith (2020) highlight this potential trade-off between equity and efficiency. In the case of California, Arcidiacono & Lovenheim (2016) find mixed evidence on the benefit of admission through affirmative action. On the other hand Bleemer (2022) presents evidences that support that the benefit of more selective university enrollment is greater for affirmative actions underrepresented minorities enrollees. Moreover, Bleemer (2021) is one of the first studies to attempt to quantify impacts on the winners and losers from a top percent plan, using a structural model of admissions for students in the University of California. His findings suggest that the gains for the pulled up group are larger in magnitude than the losses for the pushed down group.

³The DA is based on Gale & Shapley (1962) and described with detail in Rios et al. (2021)

was designed to boost the admission chances for students who performed much better than the average for students from the same high school, partly offsetting the lower average PSU scores and lower average GPA's at relatively disadvantaged schools in Chile.

The introduction of the this new component in the admissions formulas created three groups of students: (1) those who were admitted to a higher-ranked program under the new formula (a group I call the “pulled-up”), (2) those who lost access to the program they would have been admitted to in the absence of the reform, and were instead admitted to a lower-ranked program (a group I call the “pushed-down”), and (3) those whose admissions outcomes were unaffected. The available data allow me to identify all three groups in the first year of the new system (2013). I am also able to identify the same three groups who would have been present if the reform had been adopted in 2012. I then conduct a simple difference-in-differences (DD) analysis of enrollment, persistence, graduation, and post-graduation outcomes, treating the pulled-up and pushed-down students as separate treated groups and the unaffected students as a control group. In a robustness analysis, I show that the impacts from this DD approach are very similar to the effects implied by a regression discontinuity (RD) approach, focusing on students who narrowly win or lose access to their top-ranked program choices.

I find that, as intended, the pulled-up group included students from lower-income and less-educated families who attended mainly public schools. These students accepted their admissions offers at higher rates than in the previous year, and ended up in more selective programs with higher-scoring peers. Over the next 8 years I find that they graduated from significantly higher-ranked programs than their comparisons from the previous year, though their eventual rate of completing a bachelor's degree was nearly identical. Preliminary results for their first few years of labor market entry show, if anything, small increases in earnings. For pushed-down students, the results are largely symmetric (though of slightly smaller magnitude). These students, who tended to come from higher-income and better-educated families, were less likely to accept their admission offers than the comparison group from the previous year, and more likely to skip the year, retake the PSU test, and re-enter the admission pool the next year. They end

up graduating with a BA at the same rate as the previous cohort, but from less selective programs and colleges outside the selective system. Given the comparability of the impacts on the winners and losers from the reform, the evidence suggests that the GPA⁺ boost led to some improvement in the equity of the selective admissions system in Chile, with no change in efficiency.

The next section of the paper begins with an overview of the Chilean college system, which includes both selective institutions (which participate in the centralized admission system) and non-selective institutions (which charge relatively high tuition and use their own admissions rules).⁴ At the end of each year, after taking the PSU, students submit a rank order list (ROL) of preferences to the centralized admission system. Programs rank students based on GPA and PSU test scores, with different programs using different weights for the two components.⁵ The DA algorithm generates a single admission offer for each applicant: students who are unsatisfied with this choice can choose to take a year out of school, then retake the PSU and reapply the following year, or enroll in a program in the non-selective system. For each cohort of applicants, I have access to their rank order list of programs, and information on the ranking rules used by different programs. Using these data I am able to reproduce the admission offers for 99.9% of the applicants from cohorts before and after the reform. I also observe student enrollment (by college and program) and graduation outcomes for each student, including those in the selective and non-selective colleges.

To evaluate the 2013 GPA⁺ reform, I use data on the 2013 applicants, and on the numbers of students offered admission to each program, but adjust the ranking rules of each program to take out the GPA⁺ component. I then re-run the DA algorithm to generate admissions for the 2013 cohort in the absence of the reform. Comparing admission offers with and without the reform identifies the pulled-up and pushed-down students who win or lose access to a higher-ranked program, as well as the relatively large ($\sim 90\%$) of students whose admissions offers are the same. To measure the causal effects

⁴They use private admission requirements, which limits the knowledge of how students are ranked/selected if excess demand occurs.

⁵Some programs also have additional restrictions as minimum PSU scores and minimum application scores.

of the reform on the enrollment and graduation outcomes of the pulled-up and pushed down-students, I use the 2012 cohort of applicants, and compare their actual admission offers to those they would have received if the GPA^+ reform had been adopted one year earlier. This identifies potentially pulled up and pushed down groups in 2012. Since these groups were not exposed to the reform, but their ranked lists and PSU/GPA/ GPA^+ performance measures are very similar to those of the same groups in 2013, their outcomes form counterfactuals for the pulled-up and pushed-down groups in 2013 (after adjusting for economy-wide trends using the changes in outcomes of the unaffected groups using a DD approach).

This DD approach allows me to measure the separate effects of the reform on the winners and losers from the reform, and test whether the gains in outcomes for the winners are as large as the losses for the losers. However, its validity rests on two key assumptions. First, I have to assume that in the absence of the reform, the trend in the outcomes of interest for pulled-up, pushed-down, and unaffected groups would have evolved similarly - the so-called “parallel trends” assumption. I test this assumption by comparing 2011 and 2012 cohort. Following the same simulation strategy to classify students into the three relevant groups I estimate the same difference-in-differences specification for cohorts for which no reform was implemented (2011 and 2012). I find no significant difference between them when no reform is implemented.

A second key assumption is that the rank order list reported by students doesn’t change with the incorporation of GPA^+ in the application score. With no restrictions in the report of preferences, the dominant strategy for the DA algorithm is to report preferences truthfully (Gale & Shapley, 1962; Roth, 1982). While the Chilean system limits students to submitting just 10 choices, most students list fewer than 10, suggesting that most students had no incentive to change their lists in the presence of the GPA^+ boost (Haeringer & Klijn, 2009; Pathak & Sönmez, 2013). Nevertheless, some recent papers suggest that reported ranks, even under a DA system, depend on the probability of admission (Fack et al., 2019; Larroucau & Rios, 2018). If so, some students who received a relatively large GPA^+ boost may have changed their reported list of preferred

schools in 2013, relative to what they would have reported in 2012. To check that this behavior is not driving my results I estimate the same models in a sample that excludes students with “very high” boost scores.⁶ I find the same qualitative and quantitative results.

As a further validation exercise, I implement a regression discontinuity (RD) design, which does not rely on previous cohort comparisons. Specifically, I begin by estimating my DD models for the (relatively large) subset of pulled-up and pushed-down applicants who gain or lose access to their top-ranked program because of the GPA⁺ reform. The impacts of the reform on this group are very similar to the impacts on the overall groups. I then conduct an RD analysis using a sample of students whose admission scores (under the 2013 rules) are relatively close to the cutoff for their first-ranked choice, and using as a running variable their admission score as determined by that choice.⁷ I find that the impacts of passing the threshold for the first choice program are comparable in sign and magnitude to the DD estimates for the pulled-up and pushed-down groups. The estimates suggest that the winners from the 2013 reform experienced a significant gain in the selectivity of the program to which they were initially offered admission, and of the program from which they eventually graduate (which in most cases is the same), with no effect on BA completion in the 8 years after the application round. Likewise, the losers experienced a significant loss in the selectivity of the program to which they were initially offered admission, and of the program from which they eventually graduate, with again no effect on BA completion.

My results also align with the results from other equity admission interventions that find that access-oriented admission policies at selective universities can promote economic mobility without efficiency losses (Otero et al., 2021; Bleemer, 2021; S. E. Black et al., 2020). Consistently with the results reported in S. E. Black et al. (2020) for the Texas

⁶I compare the admission selectivity of the programs ranked in the first choice for students with the same boost before and after the boost was implemented. Only students with boost higher than 195 points list on average more selective programs after the reform. I define the checking sample as students with boost scores lower than 150 to be conservative.

⁷Programs have additional restrictions, like minimum test scores, that can make students above the cutoff not being offered admission into their first choice. These restrictions are easier to incorporate in the DD strategy.

Top Percent policy, I find similar graduation rates (inferred) for pulled-up students than for the average students pre-reform, suggesting that pulled-up students did not struggle more.

This paper contributes to the understanding of equity admission interventions and the effect of admission to more selective universities for students who would not normally have access to them (S. E. Black et al., 2020; Bleemer, 2021, 2022; Arcidiacono & Lovenheim, 2016; Arcidiacono et al., 2016; Otero et al., 2021; Mello, 2022; Bagde et al., 2016). I build on prior empirical research employing a difference-in-differences approach, and take advantage of the transparency of the admission criteria in order to precisely identify the treatment groups resulting from the admission reform. Unlike earlier studies, this admissions change affected the full spectrum of selective colleges, not just the access to a single institution. Thus, I study the effect in the entire population of applicants and on the entire college system (selective and non-selective institutions). Contrary to the mismatch hypothesis (Sowell, 1972), that states that low-test students targeted by access-oriented admission programs, like affirmative action, would be better off by attending programs where they match their peer characteristics, I find that the probability of graduation from the admission to a more selective program does not decrease.⁸ This paper also contribute to the early, but growing, literature that evaluate changes in the assignment mechanism, in this case which inputs are used, on the basis of students outcomes (Agarwal et al., 2020; Otero et al., 2021; Larroucau & Rios, 2020).

1 Related literature

There is a significant body of literature devoted to studying the returns to college, and more specifically, the returns for varying levels of quality or selectivity. In particular, Dale & Krueger (2002); D. A. Black & Smith (2004); Lindahl & Regnér (2005); Dale & Krueger (2014) highlight the difficulties of deriving causal estimates from observational data. Recently, numerous studies have used a regression discontinuity strategy to adjust

⁸Several papers study the mismatch hypothesis with varying results, see for example Sander & Taylor (2012); Arcidiacono & Lovenheim (2016); J. Rothstein & Yoon (2008); Bleemer (2022); Arcidiacono et al. (2011).

for selection bias and have shown that applicants at admissions thresholds gain from admission into selective institutions (e.g., Hoekstra (2009); Zimmerman (2014); Anelli (2020)).⁹ In addition, Cohodes & Goodman (2014); Goodman et al. (2015); Zimmerman (2014) present evidence from the United States that attending a selective university tends to increase graduation rates. In conclusion, the majority of evidence suggests that college quality has a beneficial impact on student performance, although this result is not universal.

However, these methods may be inadequate for evaluating the effectiveness of access-oriented policies. Students at the margin may differ from those who are targeted by the admission policies. Dale & Krueger (2014) provides evidence of heterogeneous returns to selective degrees in the United States - positive for underrepresented groups but zero on average - by analyzing the differences in outcomes for students with similar sets of admission offers but different enrollment decisions. Zimmerman (2019) and J. Hastings et al. (2009) document heterogeneous effect for the case of Chile in terms of field of study and family income.¹⁰ Additionally, treatment effects for those outside of the discontinuity may vary. I expand upon the research that employs differences-in-differences to examine the consequences beyond the admissions threshold.¹¹

My paper contributes to the literature by evaluating the effects of access-oriented policies to selective policies not only evaluating the effects on the targeted group of students but also on the displaced students. In this sense, my paper is most closely connected to S. E. Black et al. (2020), however, I take advantage of my setting to construct the treatment groups intuitively and transparently. The relationship between selectivity and outcomes for the two affected categories of students is needed to evaluate the efficiency impact on the entire system. It will be beneficial if institutions with a greater level of

⁹Another body of research focuses on the differential returns to fields of study; see for example Kirkeboen et al. (2016) and J. Hastings et al. (2009)

¹⁰Zimmerman (2019) argues that the greatest returns to top business program attendance in Chile apply only to students from high-income families. Compared to J. Hastings et al. (2009), my regression discontinuity analysis provides larger results. This discrepancy is expected since I only examine threshold crossing for the first choice, which results in greater effects than other threshold crossings. Prior studies averaged across all thresholds.

¹¹Otero et al. (2021) overcomes this challenge with a combination of admission thresholds and an exogenous score shifter.

selectivity have more and better learning materials. In contrast, it can be negative if students increase their likelihood of poor performance and school withdrawal. In the situation of differential impacts, student-resorting policies have the potential to generate both efficient and equitable benefits or costs. In order to evaluate efficiency I consider enrollment and medium-term outcomes such as dropout, college graduation, and earnings in the entire system. My results align with the results from other equity admission interventions that find that access-oriented admission policies at selective universities can promote economic mobility without efficiency losses (Otero et al., 2021; Bleemer, 2021; S. E. Black et al., 2020).

The research on the effects on graduation and earnings of access-oriented policies focuses mostly on affirmative action and Top N percent programs (Arcidiacono et al., 2011; Arcidiacono & Lovenheim, 2016; J. Rothstein & Yoon, 2008; Bleemer, 2022; Otero et al., 2021; Mello, 2022; Bagde et al., 2016; S. E. Black et al., 2020; Bleemer, 2021; Kapor et al., 2020). The "percent plans" implemented in Florida, California, and Texas ensured admission to the public university systems for students with high grades compared to their high school peers, independent of their standardized test scores. The Chilean reform is similar to these policies in that it increases the likelihood of admission for students with strong grades and is demographically blind. However, there are numerous significant distinctions. The Chilean reform did not ensure access but rather increased the likelihood. Related to this, another distinction is that the Chilean reform compares current students to prior students from the same school, whereas the percent plans compared students from the same cohort. A further advantage of the Chilean context is the transparency of admission rules. The majority of the college admission systems in which access-oriented policies have been studied have some arbitrary component or they are structured in such a way that students could behave strategically to take advantage of changes in the admission policies. Cullen et al. (2013); Estevan et al. (2017); Mello (2021) analyze the school switching behavior for the US and Brazil. Concha-Arriagada (2022) shows that this occurred in Chile during the second and third years following the implementation of the reform but the problem was quickly resolved in 2016.

Much of the research on affirmative action has centered on measuring academic mismatch. The mismatch hypothesis posits that graduation rates for minority students who attended selective post-secondary institutions would be lower than for those who attended colleges and universities where their academic credentials are better matched to the institutional average. However, results have not been conclusive (Loury & Garman, 1993; J. Rothstein & Yoon, 2008; Sander & Taylor, 2012; Dillon & Smith, 2017, 2020; Arcidiacono et al., 2011, 2014; Bleemer, 2022, 2021). My setting is ideal to evaluate this hypothesis. Similar to Bleemer (2021) for the case of California, I find that the benefits of more-selective enrollment are at least as large for high-GPA students whose low standardized test scores would have normally disqualified them from selective universities as they are for the higher-standardized test students admitted to those universities and that the graduation rate for the pulled-up students was roughly equivalent to the average for the non-affected students.

A closely connected literature evaluates the mismatch hypothesis in the particular subgroup of STEM programs Arcidiacono et al. (2016); Bleemer (2021). The STEM mismatch hypothesis holds that students admitted through access-oriented policies are less persistent in STEM fields than they would be at universities with fewer admission requirements. Contrary to what previous studies show (Arcidiacono & Lovenheim, 2016; Bleemer, 2022; Mountjoy & Hickman, 2021), the evidence for the Chilean case suggests that students pushed into more selective STEM programs by the reform have a higher probability of graduation. Even though the reform increases the probability of enrolling in a STEM program for pushed-up students, the majority of student applying to STEM programs in the pushed-up group had a fallback option of a STEM program, therefore, they were affected by getting access to better quality programs.

Lastly, this paper also relates to other studies interested in the same admission reform to answer different questions. The most related paper, Larroucau et al. (2015) evaluates the compositions of the students affected by the reform using the same simulation approach as this paper. Concha-Arriagada (2022) also relies on similar simulations to study the strategic behavior of students in 2015, after students learn about the construction of

the relative GPA boost and before the policy was fixed to address the strategic behavior. In a similar spirit, Fajnzylber et al. (2019) evaluate the effects of the reform in terms of the GPA inflation and learning effort. Finally, Larroucau & Rios (2018) use the variation from 2013 to 2014 in the weights associated with the relative GPA component to estimate models of preferences for program choices.

2 Context

The Chilean college admission system is an ideal setting to evaluate the effects of an access-oriented admission intervention like the 2013 reform. The reform introduced a new component based on the student’s relative GPA, designed to improve equity in the system. The transparency of the system, together with the availability of rich administrative data allows for the simulation of admission offers with and without the new GPA⁺ component even in years before the reform was implemented, facilitating the construction of meaningful counterfactuals for winners and losers of the reform.

2.1 Chilean College Admission System

The admission process to selective universities in Chile is a centralized score-based meritocracy, based solely on standardized admission test scores and the high school GPA score of the students. The assignment mechanism - that uses a deferred acceptance (DA) algorithm- generates a seemingly strategy-proof environment and can be replicated when admission preferences, program vacancies and applications scores are available. I discuss with detail these key characteristics to the implementation of my empirical strategy, particularly to the identification of the two treatment groups.

The college system and application procedure The Chilean college system has selective (public and private) and non-selective (private) colleges.¹² To enroll in a selective university students have to (i) graduate from high school, (ii) take the standardized

¹²In 2012 and 2013 the selective system was composed by 33 universities, which represented around 60% of college students.

admission test at the end of the academic year, and (iii) submit a rank ordered list of their preferences to the centralized admission system after learning about their test results. This process happens once-a-year and students can enroll only if they get an admission offer. To enroll in a non-selective college, students have to apply directly and follow the requirements of each institution.¹³

The admission process is organized around programs, instead of majors and universities. Programs have a highly fixed curriculum (which makes switching programs without going again through the application process hard and not common) with expected times for graduation between 4 to 7 years (5 being the mode). In most programs, students earn an academic degree after 4 years but they are required to attend a 5th year and pass a licensing exam to earn their professional degree and complete graduation. Programs provide the complete certification for most occupations, such as architecture, law, or medicine. This characteristic of the Chilean college system makes the relationship between college and labor market outcomes tighter compared to other settings.

The centralized admission process was established in the late 1960s in combination with an admission test (in the same spirit as the SAT) and a single-offer assignment mechanism based on a student-proposing deferred acceptance (DA) algorithm (Gale & Shapley, 1962; Abdulkadiroğlu, Pathak, & Roth, 2005, 2009). Its development and implementation in the country was led by Erika Grassau.¹⁴ New admission tests were redesigned at the beginning of 2000s and consist of a mandatory math and verbal exam, and one additional exam that could be science or history. Tests are taken simultaneously at a national level by the end of the academic year.¹⁵ After scores are published (tests and GPA scores), students can start their application - exclusively online through the Department of Evaluation, Measurement and Educational Registration (DEMRE for its acronym in Spanish) website and without any monetary cost - by submitting a list with

¹³In most of the cases colleges require the admission test score but don't set minimums for admission. Therefore, the restriction is a budgetary constrain.

¹⁴It is surprising the lack of recognition given to Erika Grassau and her team in charge of implementing that reform, considering how ahead of time it was when compared with the boom of the implementation of DA mechanisms in the last decade.

¹⁵The Chilean academic year normally goes from March to December, but it is shortened to November in the last high school year

no more than ten programs, ranked in strict order of preference (their Rank Order List - ROL).¹⁶ Once the application period is finished, students are assigned to programs with the DA algorithm.

Participation in the admission process is the only channel for students to enroll in any selective program.¹⁷ Because students with higher application scores are more likely to be offered admission to a program than a student with a lower application score, and selection can only be based on that, it is considered a score-based meritocratic system. A program is considered more selective than others if the application score of the last student admitted - the program cutoff score - is higher. The application score is a program-specific index that weights students' high school GPA and standardized test scores.

Deferred acceptance algorithm The Deferred Acceptance (DA) algorithm is the assignment procedure used to match students to programs, taking into consideration their preferences and the program vacancies.¹⁸ The algorithm can be described as follows: In the initial step, each student proposes to their most preferred program listed in their ROL. Programs provisionally accept students based on their application scores until they fill their total number of seats, rejecting the rest. In subsequent cycles, rejected students propose to their most-preferred program among those that have not previously rejected them, and programs reject provisionally accepted applicants with lower application scores. This process iterates until all students are assigned to a single program or all unassigned students have been rejected by every program they have ranked. See Rios et al. (2021) for a thorough description.

A studied theoretical characteristic of the DA mechanism is that it is strategy-proof, which makes reference to the fact that listing programs in order of true preferences is

¹⁶To help applicants in their decision-making, DEMRE distributes a directory that provides an overview of the university admission process, key dates, information about vacancies, extra requirements, and the application score formula for each program for each university. While waiting for their results students can access a simulation mode site with a help video that explicitly states “when selected in one of the preferences all the following ones are eliminated, therefore it is very important the strict order of preferences from higher to lower personal interest.”

¹⁷There are some special admission channels like switching students or students with disabilities but among those quotas admission score is always the selection criteria. This paper focuses on the regular admission channel.

¹⁸The variant of the student-proposing DA algorithm used by DEMRE establishes that all tied students for the last seat of a program must be admitted.

a weakly dominant strategy when students are allowed to rank every program, i.e. it cannot be manipulated by misrepresenting preferences (Dubins & Freedman, 1981; Roth, 1982). In the Chilean case, students are constrained to list only 10 choices, with extra conditions for some universities.¹⁹ Table 1 shows that 90% of applicants rank less than 10 programs with a mode of 3, in which case truthful reporting is a dominant strategy (Haeringer & Klijn, 2009; Pathak & Sönmez, 2013). Assumptions over the rank order list and details about the assignment mechanisms are used to simulate admissions with and without the relative GPA measure. Section 3.1 discuss this procedure.

2.2 Relative GPA Reform

The relative GPA reform created a grade-based measure that augment the admission criteria with a performance measure that takes account of between-school differences and boosts the admission chances for good students from schools with relatively low standardized test scores. The GPA^+ is based on the grades of a student relative to the historical distribution of GPAs at his or her high school and adds a positive boost to the GPA of students who score above the historical mean, with a maximum boost for those who score above the maximum past score at their school ($GPA^+ = GPA + \text{relative boost}$).

Equity concerns around college admission in the 1960s are what motivated the current admission system (meritocratic and transparent). Around the 2000s the admission test was changed in order to address socioeconomic differences in college admission but the socioeconomic gap in test scores persisted, even after controlling for income and parents' education. This evidence fueled a public debate that highlighted the need for a system able to identify high-ability students even when education conditions for them were not optimal to perform well in standardized test scores.

In the second half of 2012 academic year, the organization in charge of coordinating selective universities (CRUCH for its acronym in Spanish) informed the incorporation of

¹⁹Universidad de Chile and Pontificia Universidad Catolica de Chile limit the applications to their programs, in order to be valid, to the first 4 preferences. For details analysis on how this could affect the report of preferences see Lafortune et al. (2016)

a third element to calculate students' application scores in the 2013 admission process. The timing was such that students and programs had no scope for strategic responses, as students already have their GPA scores determined and universities have already made their capacity decisions.²⁰ Before the reform, application score (s_{ij}) for a student i to a program j was calculated as:

$$s_{ij} = \alpha_j \text{Tests Scores}_i + \beta_j \text{GPA}_i$$

The weights α_j and β_j were chosen by the programs under some minimum restrictions defined by the DEMRE such that $\alpha_j + \beta_j = 1$.²¹ After the reform was implemented, the GPA^+ measure was included in the formula

$$s'_{ij} = \alpha'_j \text{Tests Scores}_i + \beta'_j \text{GPA}_i + \gamma'_j \text{GPA}_i^+$$

with $\alpha'_j + \beta'_j + \gamma'_j = 1$. For its first year, γ'_j was fixed at a mandatory 10% for all the programs. From Figure 1 we can see that most of the programs opted for reducing the weight on β_j to allocate the 10% for the GPA^+ measure, therefore most of the variation observed in allocations comes from the introduction of the relative boost.

The proposed new component was designed to make more competitive the application of students that performed well at their high school by awarding them a boost to their GPA score if they perform above their school average ($\text{GPA}^+ = \text{GPA} + \text{relative boost}$). In Chile, grades are not fully curbed and they have an implicit reference to the minimum content expected by the national curriculum on each subject by year. Due to this, even the best student from a disadvantaged school that struggles to cover the minimum contents can have a very low GPA score. The GPA^+ component was designed such that with the boost, students that perform at the top of their school GPA distribution have a GPA^+ score that corresponds to that. By making the application score of good-performance students higher, the reform helped them access programs that would

²⁰The literal translation of the reform's name is "Ranking", which is misleading. Given that the score is assigned in relationship with the student's educational context rather than their class ranking, I will refer to it as relative GPA reform rather than Ranking reform.

²¹With a minimum 10% in each of the component.

have rejected them when their application score was lower.

Relative GPA measure in detail The relative GPA (GPA^+) measure is based on the GPA score of the student, but it is adjusted with a boost that depends on the historical average (\overline{GPA}) and the historical maximum high school GPA of their high school ($\max GPA$). The historical average and the historical maximum are constructed based on the high school GPAs of the students from the previous 3 cohorts at that school. It was chosen as a reference for the within-school measure to avoid within-classmates' competition. The formula to calculate the (GPA^+) score is the following

$$GPA_i^+ = \begin{cases} GPA_i & \text{if } GPA_i < \overline{GPA} \\ \overline{GPA} + \frac{850}{\max GPA} (GPA_i - \overline{GPA}) & \text{if } GPA_i \in [\overline{GPA}, \max GPA] \\ 850 & \text{if } GPA_i > \max GPA \end{cases}$$

Students with a GPA equal to or lower than the historical average at their schools have a relative GPA score equal to their GPA score. Students with a GPA bigger than the historical average but smaller than the historical maximum get their GPA score plus a boost that is determined by the slope of the line that connects the historical average GPA score with the historical maximum, which is for all schools the maximum possible score, 850.²² This implies that students in this range, from a school with a more spread out high school GPA distribution will have a smaller boost in terms of score points for each extra point in their GPA. Finally, students that perform above the historical maximum at their high school get the maximum possible score (850), even if the GPA is, measured in application points, very low.

In order to simulate the admission assignment under the the new mechanisms defined by the inclusion of the GPA^+ for cohorts previous to the implementation of the reform I construct the GPA^+ measure for the cohorts 2009 to 2012. According to the reform,

²²Figure 2 correspond to an example to represent the relationship between GPA, GPA^+ and the boost.

students who graduate from cohorts before 2009 or students who didn't attend a school had the relative GPA score equal to their GPA score.

2.3 Data

I focus my analysis on the entire universe of applicants to selective universities during the years 2012 (pre-reform) and 2013 (post-reform). For the first part of the empirical analysis, I construct a unique dataset that replicates college admission offers with and without the inclusion of the relative GPA measure in the admission process for the students in these two cohorts. This allows me to classify students into one of the three possible groups of analysis: pulled-up, pushed-down, or unaffected. To assess human capital acquisition, I add data on annual enrollment and graduation from selected and non-selective colleges for all the applicants to the 2012 and 2013 process. Finally, I add to the analysis information on employment and earnings on the private labor market up to 10 years following their application.

Admission process The relative GPA reform was implemented in the admission process of 2013. For that reason, my analysis focuses on the short and medium-long-term outcomes of all the students that participated in the admission process that year and the year before (2012). I use information from students in the 2011 cohort to validate my research design.²³

Administrative data at the student level from the admission process was shared upon request by DEMRE. It consists of socioeconomic and demographic information of applicants (gender, date of birth, self-reported family income, and parents' education), applications scores (tests scores, GPA, and relative GPA score), high school characteristics, application information (rank order list of program preferences listed in the application with their final status: valid/invalid, offer/no offer and waitlist), and enrollment informa-

²³Even though information for later cohorts is available I don't consider it in my analysis because my empirical strategy is sensitive to the strategic behavior observed during those years. After 2013, some students switched schools in their last year of high school to improve their GPA⁺ measurement. This potential for policy manipulation was fixed in the 2015 process.

tion (program, application score, and ranking of preference). This information is mainly used to simulate students' admission under a mechanism that uses two (test scores and GPA) or three (test scores, GPA and GPA⁺) inputs to calculate the application score.

The “new” mechanism incorporates the relative GPA measure (GPA⁺) into the application score formula. To compute the relative GPA measure for cohorts before the reform I use information from the national school records on high school performance for the entire population of high schoolers between 2002 and 2011 which is available online at the data platform of the Department of Education.²⁴ I compute the historical average and the historical maximum GPA at each school for each graduation cohort, and then the relative GPA score for students who graduated between 2008 to 2012 in the 2011 and 2012 admission process.²⁵ Figure A.I shows a binscatter graph with the boost score - i.e. the extra score relative to GPA- of the relative GPA score for students in application cohorts 2011 to 2013. The x-axis is the GPA score of the student minus the historical average high school GPA at the school of the student, therefore on the positive numbers we see the boost score in application points. Note that 2013 data is directly reported by DEMRE and 2011 and 2012 was calculated using the relative GPA score formula.

I also constructed a dataset with program characteristics like application score weights, application score restrictions, and the total number of seats from the public newsletter with the official information. Application score weights are required to calculate the application score under the two regimes. For each program, application scores under the status quo regime (s_{ij}) are calculated using weights from the 2012 process, and application scores under the GPA⁺ regime (s'_{ij}) are calculated with 2013 weights.²⁶

Enrollment and graduation outcomes To measure the effect of the reform on educational outcomes I track all the students that participate in the application processes of 2012 and 2013 using yearly information on enrollment and graduation provided publicly

²⁴<https://datosabiertos.mineduc.cl>

²⁵Students can participate in the admission process as many times as they want. The proportion of freshmen and older applicants is around 60% to 40% in each cohort.

²⁶Music, arts, and acting programs require an additional aptitude test, which score is not reported separately in the data. For those cases, the application score used for the alternative regime was the same as the one reported originally.

by the Department of Education. From the admission data I can observe who got an admission offer and to which program. I create variables to indicate if a student enrolls in their admission offer or if they enroll in a non-selective college instead. By using the enrollment file in the second year ($t = 2$) I can check if the student persisted at their admission offer, if they re-apply or switched to a different selective program, if they switched or persisted in a non-selective college, or if they dropped out of college.

Additionally, for each application cohort, I track graduation by 6th, 7th, and 8th years after application because yearly graduation files were available only up to 2020. I construct 3 graduation measures: (1) program graduation or graduation from their initial admission offer in 2012 or 2013, (2) graduation from some selective university to take into account that students that don't get their desired admission may switch or re-apply in the following years, and (3) graduation from a non-selective college which is always an alternative. Having access to data of the entire system allows me to measure the complete impact of the reform in the selective system - the one that DEMRE attempt to coordinate-, as well as the impact on the entire college system.

Labor market outcomes To study the effect on earnings of giving access to better programs to students that normally couldn't access them I use information from the Unemployment Insurance (UI) data. The UI data has information on all the dependent workers over 18 years old that participate in the private sector.²⁷ All the information is aggregated at the treatment group level. For pulled-up, pushed-down, and unaffected students I observe the fraction that was present in the labor market (participation) and bins for their monthly taxable income from 8 to 10 years after the admission process.

²⁷Data excludes: (i) workers subject to an apprenticeship contract; (ii) workers under 18 years of age; (iii) private home workers (until October 2020); (iv) pensioners; (v) independent or self-employed workers; and (vi) public sector workers. In a future version of the research, I will be able to include information on public sector workers and person-level data.

3 Empirical Strategy

The empirical strategy is divided in two parts. First, I simulate the admission mechanisms with and without the relative GPA measure. I classify students into 3 groups based on the admissions simulations: (i) pulled-up, students who gain access to more selective admissions when the third component is considered in the assignment mechanism, (ii) pushed-down, students who loss access to more selective programs with the new mechanism, and (iii) unaffected, students whose admission options are unaffected by the change in the mechanism. By simulating the admissions under the two mechanisms in earlier years, before the reform was implemented, I can identify the groups who would have been pulled-up and pushed-down in those years. This facilitate a difference-in-differences design to estimate the impact of the inclusion of the relative GPA on enrollment, graduation and earnings for the students affected by the reform.

3.1 Identification of treatment groups: pulled-up, pushed-down and unaffected

The inclusion of the relative GPA measure into the admission process enhanced the equity of the college admission system. Students with relatively low test scores but high GPA from low-educated and low-income families got admissions into more selective program when the third component (GPA^+) was considered. There is also a higher representation of females in the pulled-up group of students. Pushed-down students tend to be in higher proportions from private schools, males, and from highly educated and high income families. Looking at the impact in admission offers induced by the reform, most students affected had an admission one preference up or down with respect to the status-quo regime, they are move into or out of their 1st preference, and they get a new admission in the same field.

Simulation of the admission mechanism The relative GPA reform impacted the way that students were matched to the programs that they apply. Before its implementation the application score for a student i applying to a program j was calculated using

only 2 inputs: admission test scores e_i and GPA score g_i . With the implementation of the reform the new application score was calculated based on $s'_{ij}(e_i, g_i, c_i)$. Denote $\mu(\cdot)$ as the matching function defined by the mechanism that uses a Deferred Acceptance algorithm, the information from the pool of applicants, the application scores defined by the programs and the capacity restrictions of the program. The change in the inputs used by programs to evaluate students defines a new mechanism $\mu'(\cdot)$.

A student i can be characterized by $\theta_i(\succ_i, e_i, g_i, c_i)$ composed of their rank order list (\succ_i) and their scores. In each application year, for some students the admission assignment under both mechanisms will differ, $\mu(\theta_i) \neq \mu'(\theta_i)$, and for others it won't $\mu(\theta_i) = \mu'(\theta_i)$. I classify the pool of applicants into 3 mutually exclusive groups:

- Pulled-Up: $PU_i = 1\{\mu(\theta_i) \prec \mu'(\theta_i)\}$ students who get access to a program ranked higher in their list with the new mechanism μ' than with the old mechanism μ .
- Pushed-Down: $PD_i = 1\{\mu(\theta_i) \succ \mu'(\theta_i)\}$ students who get access to a program ranked lower in their list with the new μ' than with the old mechanism μ .
- Unaffected: $C_i = 1\{\mu(\theta_i) = \mu'(\theta_i)\}$ corresponding to students with access to the same programs with and without the inclusion of the GPA⁺ measure.

Implementation of admission simulations For each student, in each application process, I start by computing their alternative application score. For students pre-reform this also includes computing the GPA⁺ score. For each program that the student listed, I use the weights from 2012 and 2013 to calculate the alternative application score (for students in the 2012 cohort I calculate s'_{ij} and for students in 2013 I compute s_{ij}).

I replicate the DA algorithm to simulate the admission assignment of students with the GPA⁺ measure for pre-reform students ($\hat{\mu}'(\theta_i)$), and without it for post-reform students ($\hat{\mu}(\theta_i)$). In order to test the quality of the replication I simulated the admission assignments using s'_{ij} for cohort 2013; I replicate 99.9% of the real assignment offers.

For each student in application cohort 2012 or 2013, I compare the simulated admission with the real admission offer and I classify them into the pulled-up (pushed-down)

group if the admission assignment with the GPA^+ measure was higher (lower) in the list than the assignment without it. Students are classified as unaffected if the admission program under both regimes is the same.

Simulation assumptions There are three main assumptions needed for the simulation to be valid as a counterfactual under the alternative mechanism.

Assumption 1 *The rank order list of preferences that the students submit would have been the same with and without the reform*

Assumption 1 has two components, one that refers to the stability of preference and one that refers to the reporting behavior. I assume that preferences are stable with respect to the reform, which means that the indirect utility associated with each program does not depend on the components and weights used by the programs to evaluate applicants.

In terms of reporting behavior, I use the traditional approach taken by the literature that establish that without restrictions on the number of applications, the dominant strategy with a Deferred Acceptance (DA) algorithm is truthful reporting (Gale & Shapley, 1962; Dubins & Freedman, 1981; Roth, 1982). As most of centralized admission system, the Chilean application system restrict the application list (up to 10 options), however, because more than 90% of the students list fewer than 10 options, the restrictions can be interpreted as not binding (Haeringer & Klijn, 2009; Abdulkadiroğlu & Sönmez, 2003; Abdulkadiroğlu, Pathak, Schellenberg, & Walters, 2020).

One possible concern rise from the recent literature on mechanisms design and their interest on using the information from the centralized admission systems to estimate school choice demands models (Agarwal et al., 2020; Fack et al., 2019; Larroucau & Rios, 2018). One way of rationalizing the fact that students don't fill up their application options relates to the idea that reporting behavior is based on students' feasible options. This behavior may violate assumption 1 if students that observe the boost (that potentially could increase the set of desirable options that they will be eligible for) reacted by adding more selective programs to the top of their list. This would create a problem in the identification of the treatment group if students get admitted to this added programs but similar students that didn't observed the boost (cohort of 2012) didn't get admitted

under the simulation (because they didn't list the new options).

To assess this potential threat I first compare the number of admission options listed in 2012 and 2013 by students with a boost (by adding a program to the top of the list, the total number could increase). Students that observe the boost in 2013 are not more likely to have longer application lists than students with the same calculated boost but who didn't observed it (cohort of 2012). Additionally, I check the selectivity of the most preferred program or top ranked program of students with a boost, in 2012 and 2013. Figure 3 show that the selectivity of the first option (measured as the application score of the last person admitted at that program) increased in 2013 only in the highest values of boost score distribution. In order to check for the sensitivity of the results I estimate the results without students with more than 150 points in their boost score (2% of the total sample and a conservative range compare to what is observed in the graph). As discussed in Section 8, results don't change qualitatively or quantitatively with this sample restriction.

Assumption 2 *The number of available seats per program each year would have been the same with or without the reform*

Assumption 3 *Standardized test scores and GPA scores would have been the same with and without the reform*

Assumption 2 and 3 are justified by the fact that the reform was announced in the last half of the academic year. At that point, universities have already made their capacity decisions and students' average GPA from the 4 year of high school was already determined, therefore there was no scope for strategic responses.²⁸

Characterization of treatment groups Table 2 shows the characteristics of the group of students identified as pulled-up, pushed-down and unaffected for cohorts of applicants in 2012 and 2013. Each year, pulled-up and pushed-down applicants account for approximately 4% of the applicant pool. From Table 2 we can see that the reform

²⁸After the first year, there is some evidence, at least anecdotal, about students switching schools in their last year in order to graduate from schools with very low maximum historical GPA in order to gain the maximum score from the GPA⁺ component. In 2015 this problem was addressed with a change in the policy, which established that the score was calculated relative to the GPA of the student and the school that they attended each year.

was able to impact the students that were targeted by it. Students in the pulled-up group have better GPA than those in the unaffected and pushed-down groups; yet, their exam scores are comparable to those in the unaffected group. Looking at pushed-down students, they have low GPA and high test scores. Moreover, pulled-up students are 3 times less likely to attend a private high school than a pushed-down student and looking at family characteristics, pulled-up students come from families with average income 30% lower than pushed-up students, and their parents are less educated.

Impact of the reform on admission offers Figure 4 presents the distribution of pulled-up and pushed-down students based on the number of positions moved in their rankings between the admission assignment with and without GPA⁺. If the most preferred program that the student could reach without the GPA⁺ measure was choice 3, but with the inclusion of the boost the student could get into their most preferred option (pulled-up students), then the student was moved 2 positions due to the reform. Figure 4 shows that the change in terms of preferences is similar for pulled-up and pushed-down groups and that most of the students affected by the reform were moved 1 position in the preference list.

A more detailed analysis of the distribution of rankings for admission is presented in Table 3. Each row presents the number of students with admission assignments in that ranking when the relative GPA measure is considered. Each column presents that total number of students with admission assignment in that preference choice when the GPA⁺ measure is not considered. Students assigned to the same program in both regimes are classified as unaffected and are presented in the table without background color (table diagonal). The percentage value in each cell correspond to the proportion of students in that group in that specific ranking combination. The main margins of treatment of the reform corresponds to movements into and out of students' 1st preference. The high percentage of students moved between no admission and 1st choice is not explained by a higher proportion of students with shorted rank order list but rather due to the a bigger proportion of students at the margin of the minimum requirements of not very demanded programs. More specifically, certain program establish complementary restrictions to

admission, as minimum application scores (taking all the components into consideration) or minimum test score averages. Students in this margin have twice higher proportion of their total rank order list as invalid due to this extra restrictions.

Finally, Tables 5 and 4 present the number of pulled-up and pushed-down student in each field with and without the inclusion of the GPA⁺ component, based on the fields of the admission and simulated admission. For both groups, in most of the cases, students move along their ranking but they stay in the same field (diagonal of the table).

3.2 Difference-in-differences design

I estimate the effect of the reform on human capital acquisition and earnings, on the group of pulled-up and pushed-down students. My difference-in-differences design compares the outcomes of students who apply in cohorts after the implementation of the reform - therefore affected by it - versus those in cohorts before the implementation of the reform. With the estimation of the effect of the reform on pulled-up and pushed-down students, I analyze the (outcome) efficiency impact of the reform on the system.

The parameters of interest to evaluate the effect of the inclusion of the relative GPA measure in the admission process can be expressed as the conditional average treatment effect for the group of students pulled-up and pushed-down.

$$\tau(PU) = \mathbb{E}[Y_i(\mu') - Y_i(\mu) | PU_i = 1]$$

$$\tau(PD) = \mathbb{E}[Y_i(\mu') - Y_i(\mu) | PD_i = 1]$$

In the potential outcome framework $Y_i = D_i Y_i(1) + (1 - D_i) \cdot Y_i(0)$ is the outcome of a student i , and $D_i = 1$ {when the relative GPA is used for admission assignment}. The observed outcomes is represented by $Y_i = 1\{t(i) = 2012\} \cdot Y_i(0) + 1\{t(i) = 2013\} \cdot Y_i(1)$. Assuming additive separability to capture any changes in time uncorrelated to the determinants of the outcomes with and without the inclusion of the GPA⁺ measure, I

estimate models of the form:

$$Y_i = \beta_1 PU_i + \beta_2 PD_i + \beta_3(PU_i \cdot Post_i) + \beta_4(PD_i \cdot Post_i) + \beta_5 Post_i + X_i' \Gamma + \varepsilon_i$$

where Y_i is the outcome variable of interest to evaluate the reform: enrollment, graduation and earnings. PU_i indicates if the student belong to the pulled-up group, PD_i indicates if the student belong to the pushed-down group, $Post_i$ is an indicator that takes the value of 1 if the students apply post reform. The omitted group are students that get access to the same programs under both regimes. X_i' is a vector of individual characteristics such as gender, family income, type of school, GPA and standardized test scores to control for possible changes in the composition characteristics of pulled-up and pushed-down students between 2012 and 2013.²⁹

Here β_3 and β_4 are the estimates of the parameter of interest to evaluate the reform. β_3 captures the effect on outcome Y_i of gaining access to the a more preferred, but also more selective program due to the inclusion of the GPA⁺ measure in the admission process. Likewise, β_4 captures the effect of losing access to more selective programs with the reform.³⁰

Identification assumption The key identification assumption is that the outcomes for these three groups of students would have evolved similarly for the cohorts 2012 and 2013 if the reform would have not been implemented. I cannot directly test that, however, I conduct a placebo exercise with data from the 2011 application cohort that present suggestive evidence in support of it.

Following the same procedure used for cohort 2012, I start by computing the boost score for each student in 2011, and application scores for each program in their rank order list. With that, and keeping constant the vacancies observed that year I re-run the DA algorithm using the three components application score. Using the simulated admission

²⁹Results are presented with and without controls. Most of the results are quantitatively and statistically unchanged.

³⁰The new admission program is more preferred by definition of the treatment group, but it has to be more selective because if it wasn't the case, that program would have been reached in the status quo scenario.

assignment I classify 2011 students into pulled-up, pushed-down and unaffected. Finally, I estimate the diff-in-diff specification but with the variable $Post_i$ indicating if the student was observed in the 2012 admission process.

Table 6 shows the estimates for this placebo exercise, which can be interpreted as the effect in enrollment and graduation for pulled-up and pushed-down students when no reform is implemented. As expected, there is no significant effect suggesting that when no reform is implemented these groups follow a similar trend. The estimates would be biased if the coefficients of interest reflect sample selection resulting from the impact of the reform on the composition of applicants. However, there is no change in the trend of total applicants, and no change in the probability of pulled-up students to reapply compared with the 2011 cohort. There also would be bias in the estimates if there were unexpected changes in 2013 in other determinants of outcomes that differentially affected the three groups. I am aware of no such change.

Notably, the intervention considered for this diff-in-diff evaluation occurred just once, so considerations regarding the calendar time of the comparison group observations, such as those stated by Goodman-Bacon (2021); Baker et al. (2022); De Chaisemartin & d’Haultfoeuille (2020), do not apply in this context.

4 Enrollment Results

The change in the admission mechanism due to the inclusion of the relative GPA measure had a large impact on initial enrollment for pulled-up and pushed-down students. However, this change fades out with time; 3 years after the implementation of the reform the changes in the probability of enrollment is zero for pulled-up and pushed-down group.

The difference-in-differences estimates in Table 7 show that, for pulled-up students there is a large effect in the probability of students choosing to enroll in their admission offer. After the reform, pulled-up students are 22 p.p. more likely to enroll in the selective program that they were admitted. This is a 40% effect on enrollment.³¹ For pushed-down students the probability of enrollment decreases by 16.7 p.p. The difference (in

³¹Table A.I presents the average enrollment rates in the selective system for the 3 groups.

absolute value) between the effect on enrollment for pulled-up and pushed-down students is significant, indicating that the inclusion of the GPA^+ measure improved the system in terms of identifying successful applicants, i.e., there is an increase in the total number of students that decide to enroll once admission is offered.

The total effect on enrollment uncover changes at two margins: the extensive margin - students that gain or lose the possibility of admission in the selective system - and the intensive margin - students that improve (worsen) their admission in the selective system, but that with or without the reform would have had some admission on the system. On the extensive margin, the reform changed the probability of a student of getting access to some selective program in pulled-up and pushed-down students by approximately 20%.

The total effect on initial enrollment is not fully driven by students at the extensive margin. To study the intensive margin, I restrict the sample to students that would have got some admission under the two regimes. Observing the admission offers under the two regimes allows me to correct for the potential selection bias of only observing enrollment if a student actually gets an offer.³² Columns 5 and 6 of Table 7 presents the results restricted to the group of student at the intensive margin. The estimates on initial enrollment for pulled-up students after the reform is smaller (17 p.p.) but still large. Compared with the pushed-down students (11 p.p.), I find evidence of higher intensity of preferences for pulled-up students, i.e., that the reaction, in terms of enrollment decision, from getting access to a program higher in the rank order list is stronger than the reaction from losing access to it, for the pushed-down group.

I summarize the changes in the programs that students attend using traditional measures of quality like selectivity and graduation rate. Table 8 shows how the characteristics of the peers and programs that students attend before and after the reform changed. Columns 1 and 2 show the diff-in-diff estimates of a regression in which the dependent variable is one of these average program characteristics before the reform. The first 3 rows show that pulled-up students attend more selective programs after the reform, in

³²All students in the pulled-up group got an admission offer in 2013 (if not they could not be better than without the GPA^+ measure), but not all pulled-up students got an admission offer in 2012 because the reform was still not implemented.

the sense that the average student at the program they enroll had higher test scores and GPAs than the average student at the programs they enroll before the reform was implemented. Graduation on time is an indicator of the probability that a student graduates in the number of years set by the program; after the reform pulled-up students enroll in programs where the average student is more likely to graduate on time. The results are symmetrical for pushed-down students.³³

Enrollment up to 4 years after the reform If students are unsatisfied with their initial admission offer students can enroll in a non-selective college or re-apply to the selective system the following year (normally after taking extra test preparation courses). Columns 3 and 4 in Table 7 show that in the first year, pushed-down students compensate for the decrease in the probability of enrolling in a selective program by enrolling in the non-selective system. However, the 3.9 p.p. increase in the probability of enrollment in the non-selective system does not offset completely the decrease in the probability of enrollment in the selective system. This means that the reform leads to some pushed-down students not enrolling in any university in the first year after high school.

Table 9 shows that pushed-down students are 7 p.p. more likely to reapply to the selective system after the reform was implemented. Table 10 presents the changes in enrollment at any program for the pulled-up and pushed-down group up to 4 years after the implementation of the reform using the same diff-in-diff specification. The initial difference in enrollment (even considering non-selective programs) generated by the reform is fully reversed in the second year for pushed-down students. Column 3 shows that 3 years after the implementation of the reform pulled-up students are still 1.5 p.p. more likely to be enrolled relative to before the implementation of the reform. This difference is fully offset 4 years after the reform.

The selectivity of the programs that students attend changed after the reform and a difference persisted throughout time. Table 11 presents the change in the average test score of the peers at the same program, up to 4 years after the implementation

³³The expected graduation time of the programs that pushed-down students enroll after the reform are on average 0.07 years shorter.

of the reform. The initial enrollment for pulled-up students is at significantly more selective programs after the reform, however students in the 2012 cohort seem to react the second year after the implementation of the reform by reapplying and enrolling at more selective programs (which makes the difference in the selectivity between the two cohort to go down). From columns 2, 3, and 4 we see that pulled-up students ended enrolled in programs with peers with on average 0.1 s.d. higher test scores after the reform was implemented. The behavioral response is similar for pushed-down students. After the reform, students are less likely to enroll and they enroll at less selective programs. However, when they reapply they are able to reach more selective programs (specially 2 years after the reform). Four years after the implementation of the reform there are no differences in the probability of enrollment for push-down students, however, the selectivity of the programs is lower than without the reform (peers have on average 0.08 s.d. smaller test scores).

5 Graduation Effects

I find that pulled-up students are 8.4 p.p. more likely to complete their initial admission program; pushed-down students have a comparable opposite effect (-8.2 p.p.). An alternate exercise designed to test for mismatch hypothesis confirms this preliminary evidence against it. Pulled-up and pushed-down students have no effect on the probability of college graduation when considering graduation from any program (and not just from the new programs granted admission as a result of the reform) and the probability for them to remain enrolled due to the delay enrollment.

Admission program completion There is a positive effect in the probability of complete their initial admission program for pulled-up students, with a comparable opposite effect for pushed-down students. Columns 1-3 of Table 12 present the results for graduation from the admission program at different points in time. Consistently, there is a large positive effect (8.4 p.p. increase by 8 years after the reform) of 36% on the likelihood of graduation from the admission program for pulled-up students. Column 4 also shows

that pulled-up students are more likely to graduate on time after the implementation of the reform. For pushed-down students the effects on graduation are similar in magnitude but with the opposite sign.

In essence, the reform enabled pulled-up students access to more selective programs which increased their likelihood of enrolling in and graduating from those programs. Putting the graduation effect for pulled-up students into perspective, the implied graduation rate for the marginal student admitted by the relative GPA is 38% (8.4/21.9). This does not differ much from the average graduation rate of unaffected students post-reform (40%) or from the pre-reform level of 39% percent. In addition, the impacts are qualitatively comparable to the findings of other equitable college admission programs, such as S. E. Black et al. (2020) and Bleemer (2021).

Mismatch hypothesis The mismatch hypothesis establishes that applicants with lower test scores targeted by equitable admission policies would benefit from enrolling in less selective universities, where their academic qualifications more closely “match” those of their peers (Sowell, 1972). This hypothesis found empirical support on some of the mixed results from the research around affirmative action policies like (Arcidiacono & Lovenheim, 2016). However, the evidence presented so far for the relative GPA reform contradict this hypothesis; I interpret the fact that students in the pulled-up group enroll in more selective programs after the reform and increase their probability of graduation from those programs as evidence against the mismatch hypothesis.

Because the main specification doesn’t control for the tuple (specific pair of admission programs with and without the GPA^+ measure) of admission programs, one possible concern refers to the potential imbalances on the programs that student get admitted with and without the reform, between 2012 and 2013.³⁴ In order to control for that, I estimate an alternative specification that includes as a control the admission assignment without the reform. This way, I can ensure that all the variation captured by the diff-in-diff comes from pulled-up students with the same admission assignment without the

³⁴Students in the pulled-up group are by definition admitted to more selective programs post-reform, but this is relative to their own assignment.

reform and with admission to more selective programs after the reform.³⁵ Table 14 shows the result from this exercise. Contrary to the mismatch hypothesis, more selective admission increased the graduation probability for pulled-up students, with similar effect than the estimated before (9 p.p.).

STEM In recent year there has been an special interest around STEM degrees, and the focus around this topic for access-oriented policies has not been the exception (Loury & Garman, 1993; Holzer & Neumark, 2000; Arcidiacono et al., 2016). Arcidiacono et al. (2016) study major degrees for the case of California campuses when affirmative action policies were in place; their research states that a better matching of science students to universities by preparation level could increase minority science graduation.

I find that the effect on degree completion in STEM for STEM applicants is positive and significant (6 p.p.). Column 1 in Table 15 shows that the relative GPA reform increase the probability for pulled-up students to get admitted in a STEM program. Column 2 presents the effect for enrollment in a STEM program, conditional on student listing some STEM program in their application and column 3 also presents enrollment results but focusing on students at the intensive margin of treatment. The effect on enrollment (16.9 p.p.) compared to the effect on graduation in the same sample (6.1 p.p.) suggest that the implied graduation rate for the marginal student admitted by the reform is higher than the graduation rate in STEM degrees for the unaffected students in the entire system (36% vs 24%).

College completion There is no effect of the reform (pulled-up or pushed-down group) on human capital acquisition when it is measured as college completion and when the possibility for students to be still enrolled 8 years after the application process is considered. However, pulled-up students are more likely to earn degrees from selective programs after the reform.

Table A.III shows the average graduation from any program by 6, 7, and 8 years after

³⁵Remember that the definition on pulled-up group is based on the ranking of the preference, but if something was ranked higher and was less selective than the admission assignment without the GPA⁺ measure, then the algorithm would have assigned the student to that program pre-reform.

the implementation of the reform by treatment groups. Notice that graduation from any program captures some of the indirect effects of the reform in reapplications (therefore late enrollment in the selective system) and enrollment in the non-selective system. This could be one of the reasons why, even 8 years after the application process, there are still important changes in graduation rates relative to the previous year, suggesting that the lack of more graduation data limits the full analysis of the reform.

The difference-in-differences estimates for the effect of graduation from any program are presented in Table 16. There is no change in college completion for pulled-up and pushed-up students by 7 year after application due to the reform. However, there is a negative effect on graduation 8 years after the reform for pushed-down students, i.e., without the reform they are more likely to have completed some program. I interpret that results as a consequence of the behavioral response in enrollment for pushed-down students. As a consequence of their late enrollment after the reform (they are weaker candidates due to the introduction of the relative GPA and they take more attempts to enroll in the programs that they like) they are more likely to graduate late (even after 8 years from the implementation of the reform). Column 4 of Table 16 presents the result when the dependant variable indicates if the student graduate or is still enrolled 8 years after application. The null effect implies that pushed-down students are not acquiring less human capital after the reform.

Table 17 present the results divided by graduation from any selective program and Table 18 from any non-selective program. These results also suggest that changes in graduation at 8 years after application for pushed-down students are driven mostly by changes from selective enrollment, which requires a late enrollment if the student wants to enroll in a different program than the admission offered by the new mechanism after the inclusion of the relative GPA measure.

In summary, the reform made pulled-up students more likely to graduate from more selective programs, with no impact in college completion. For pushed-down students, the inclusion of the GPA⁺ made them less likely to graduate by 8 years after, however, this is not due to a decrease on the probability of college completion but due to a delayed

enrollment in selective programs, for some of the students that didn't enroll or didn't stay in the program admitted after the reform.

6 Heterogeneity Analysis

I first examine the effects of the reform dividing the group of pulled-up and pushed-down students into two groups based on changes of selectivity (measured as average of test scores) between the admission program with GPA⁺ and without GPA⁺. Table 19 shows in column 1 that the effects on initial enrollment are positive and larger for pulled-up students with smaller changes in selectivity relative to the group with larger changes in selectivity. Column 2 presents the effect on graduation from the initial admission; the effects are positive for pulled-up students and larger for students with a bigger change in selectivity. For graduation from any program, small increases in selectivity have a detrimental effect on students, but this effect appears to be driven by students taking longer than eight years from their participation in the admissions process to graduate. Results for pushed-down students follow a similar pattern across all the outcomes, students with bigger reduction in selectivity are less likely to enroll, graduate from the initial admission program and graduate 8 years from their participation in the admission process from any program; however, the effects are non-significant when the outcome of graduation or still enrolled is considered.

From Table 3 we observe that the main margin in which the reform affected students was increasing (decreasing) the admission of pulled-up (pushed-down) students into their 1st choice. Table 20 presents in columns 1 and 2 the main results for this sample, i.e., students moved to and from their 1st choice when the relative GPA was considered. On this sample the effects on for the initial admission program are bigger than in the entire population; however, when behavioral responses are considered there is no effect on human capital acquisition 8 years after the implementation of the reform.

Appendix B examines the differential effects by gender, income and boost score of student on the main outcomes of enrollment, graduation from admission, college completion

from any program, and graduation from a selective program. Table B.I shows differential effect of enrollment (15 p.p.) only for students with higher boost. In terms of graduation from admission Table B.II suggest that the main effect for pulled-up students is driven by the effect on females and students with high boost. There are no differential effect on college completion when any program is considered. I find some indications of variation of impacts across gender, family income, and boost score, but the overall picture is pretty consistent.

7 Alternative empirical approach

Using a regression discontinuity (RD) design that permits a direct test of the identification assumptions, and does not rely in previous cohorts, I evaluate the impact of getting access to the most desired program after the implementation of the reform.

I use a regression discontinuity design to estimate the effect on enrollment and graduation of threshold crossing the 1st preference's cutoff because, as shown in Table 3, the main margin of treatment of the reform is with respect to individuals moved to and from their first preference. I estimate the effects of crossing the admission cutoff for the most preferred program (δ) on enrollment, selectivity of the enrolled program, graduation from the admission program, and any graduation after 8 years using a standard regression discontinuity specifications of the form

$$Y_i = f(r_i) + \delta C_i + \eta_i$$

where Y_i is one of the outcomes listed above for individual i ; r_i is the difference between the admissions score assigned to i 's most preferred program and the admission cutoff score to that program or running variable; $f(r_i)$ is a smooth function (results presented in Appendix F for polynomials of degree 1 to 5) of the running variable (which can change on either side of the cutoff); C_i indicates if i 's application score is greater than the cutoff score (so i is admitted to the most preferred choice), and η_i is an error term. I estimate this equation using data from all the programs with excess of demand (for which the

cutoff is meaningful) on the whole range, with the exception of the linear specification, for which I limit the data to a small score window close to the cutoff.

Table 21 provides a summary of the principal results from the RD estimator employing a polynomial of order 3 and the diff-in-diff estimates at the 1st choice margin for pulled-up and pushed-down groups. The RD estimates are more similar to the results for pushed-down students (but smaller for graduation from the initial admission), with the same sign and order of magnitude.

students who could resemble pulled-up and pushed-down students. If compliers of the RD are a blend of these two groups, one would anticipate a smaller estimate for the RD.

In addition, I estimate the same RD model while limiting the sample to students with a boost score greater than 5 (the average boost score for pushed-down students at the margin of first and no admission) in an effort to recover the effects from a population that is more comparable to the pulled-up group of students. Table F.VI shows the enrollment results, while Table F.VI displays the graduation estimates. In both instances, the outcomes are greater and comparable to the diff-in-diff outcomes (19.5 p.p. for enrollment and 9.9 p.p for graduation).

The tables F.XII and F.XII provide the findings of an alternative exercise designed to quantify the effects on a subset of pulled-up students. This experiment focuses on students with admission at their 1st or 2nd preference with the relative GPA (treatment margins 1-2) and with simulated admission at their 2nd preference without the GPA⁺ measure. In this sample, the threshold crossing is only explained by the boost). By comparing pulled-up students with very comparable unaffected students, the sample restriction aims to determine the effect of threshold crossing for pulled-up students (non-crossing but similar - close to the margin). The small sample size resulting from the requisite makes the results unstable and imprecise; still, the sign and magnitude of the values for graduation from the admission offer fluctuate around the diff-in-diff estimate for the margin between first and second preference (7.7 p.p).

A potential threat to the regression discontinuity (RD) design is that people might try to sort themselves above the cutoff in order to receive an offer from their preferred

program. Figures in Appendix F show that there are no discontinuities around the cutoffs in the density of applicants and in the observed characteristics support assumption against that type of sorting. In addition, the McCrary (2008) test is negligible and fails to reject the null hypothesis of no sorting.

8 Robustness checks

I conduct a number of checks to verify the robustness of my conclusions. I check different samples (removing students with boost higher than 150 points or students attending programs over 6 years) and estimating my results clustering at the school-year level, and all of them support my main findings.

Changes in ROL due to the reform The key assumption for the identification of pulled-up and pushed-down groups is that the rank order list (ROL) of the application submitted by the applicants in each process would not change under a different assignment mechanism. Recent literature present evidence raising concerns over the inclusion of a more selective programs when the boost score is observed. By checking the selectivity of the first preference listed by students in 2012 and 2013 (measure as the cutoff score of that program) for students with the same boost we see some increase in the selectivity when boost is larger than 150.

As a robustness check I estimate the main results presented above but removing students with boost score higher than 150. The tables with the results for this case are presented in Appendix C. Results are not only qualitative but also quantitative similar for all the outcomes.

Sensitivity of the results to long programs Given the instability of graduation results even after 8 year of participation on the admission process, I restrict the analysis only to programs with expected graduation time of less than 6 year in Appendix D and to less than 7 years in Appendix D. Both sets of results present similar results in terms of magnitude and significance than the ones discussed previously.

Inference The previous results have been estimated using robust standard errors. Alternately, in Appendix E I present the main results allowing clustering at the school-year level. Nonetheless, any of the results take into consideration the potential error associated with the estimation of the pulled-up and pushed-down groups. Results presented in Appendix E are virtually equivalent to the results presented above.

9 Labor Market Outcomes

Finally, I study the labor market effects of the reform.³⁶ An important challenge refers to the long graduation times observed in the previous section, and the even longer span of time needed to account for the behavioral responses of reapplication to the selective system when students were not satisfied by their admission offer. Therefore, by studying earnings ten years after the implementation of the reform I am not able to fully capture the effect of the reform on earnings, limiting the analysis. Moreover, aggregated data - earnings with an indicator of group of treatment but without individual characteristics - only allows for very preliminary evidence at group level.

Figure 5 presents earnings histograms for pulled-up and pushed-down groups of students pre and post implementation of the relative GPA reform. In each case histograms are presented relative to the unaffected group. Even though at the moment I cannot calculate the diff-in-diff estimates, a preliminary review of the aggregated data confirms that pulled-up and pushed-down students do not do worse than before the implementation of the reform. Overall, in terms of outcome efficiency - graduation and earning-, the evidence confirms that the new assignment mechanism didn't make the system less efficient.

³⁶Up to this date, access to individual level data required to estimate the difference-in-differences specification used in the previous sections is under approval.

10 Conclusion

This paper studies the impact of providing students with access to more selective college alternatives. I use the variation on admission generated by the inclusion of a relative GPA measure motivated by equity concerns. I explore the effects of the reform on enrollment, graduation and earning for the two groups directly and indirectly affected by this change: (i) students who gain access to more selective programs (pulled-up) and (ii) students who lose access to more selective programs (pushed-down).

The transparency of the college admission process combined with the properties of the assignment mechanism and the richness of the data available allow me to cleanly identify the groups of students affected by the reform, one of the big challenges in the evaluation of admission reforms. By simulation of the admission offers with and without the inclusion of the relative GPA measure I identify the group of affected students. The replication of the admissions with the GPA^+ in the years before the reform helps me to identify the group of student that would have been affected. This simulation facilitate the implementation of a difference-in-difference design.

This empirical strategy compares the outcomes of students in the pulled-up and pushed-down groups before and after the implementation of the reform, therefore, before and after they get access to these more selective programs. The transitory variation on outcomes is controlled by the second difference with respect to the group of unaffected students.

I find that the incorporation of the relative GPA measure into the college admissions application score formula expanded the options available for students with significant less resources. As a result of the reform, pulled-up students became more likely to enroll in a selective program, and they chose to enroll in programs where their peers have higher test scores, GPA scores, and graduation rates. Contrary to the prediction of the mismatch hypothesis, reform-targeted applicants with lower test scores gained from enrolling in more selective options, boosting their likelihood of graduation by 8.4 percentage points.

For pushed-down students, I find that their likelihood of graduating from the admission program assigned by the new mechanism decreases by 8.2 p.p., but they are not

less likely to receive a bachelor's degree. There is however an impact in the timing of their enrollment that would be interesting to study with more details once more data on graduation and earning becomes available. Nevertheless, preliminary evidence confirms that there is no negative impact on earning for pushed-down students.

Collectively, the evidence presented above indicate that test-based meritocratic admission system can be improved by the inclusion of in-school performance metric, increasing admission equity without incurring an efficiency penalty.

References

(n.d.).

- Abdulkadiroğlu, A., Angrist, J. D., Dynarski, S. M., Kane, T. J., & Pathak, P. A. (2011). Accountability and flexibility in public schools: Evidence from boston’s charters and pilots. *The Quarterly Journal of Economics*, 126(2), 699–748.
- Abdulkadiroğlu, A., Pathak, P. A., & Roth, A. E. (2005). The new york city high school match. *American Economic Review*, 95(2), 364–367.
- Abdulkadiroğlu, A., Pathak, P. A., & Roth, A. E. (2009). Strategy-proofness versus efficiency in matching with indifference: Redesigning the nyc high school match. *American Economic Review*, 99(5), 1954–78.
- Abdulkadiroğlu, A., Pathak, P. A., Schellenberg, J., & Walters, C. R. (2020). Do parents value school effectiveness? *American Economic Review*, 110(5), 1502–39.
- Abdulkadiroğlu, A., & Sönmez, T. (2003). School choice: A mechanism design approach. *American economic review*, 93(3), 729–747.
- Agarwal, N., Hodgson, C., & Somaini, P. (2020). *Choices and outcomes in assignment mechanisms: The allocation of deceased donor kidneys* (Tech. Rep.). National Bureau of Economic Research.
- Anelli, M. (2020). The returns to elite university education: A quasi-experimental analysis. *Journal of the European Economic Association*, 18(6), 2824–2868.
- Angrist, J. D., & Rokkanen, M. (2015). Wanna get away? regression discontinuity estimation of exam school effects away from the cutoff. *Journal of the American Statistical Association*, 110(512), 1331–1344.
- Arcidiacono, P., Aucejo, E., Coate, P., & Hotz, V. J. (2014). Affirmative action and university fit: Evidence from proposition 209. *IZA Journal of Labor Economics*, 3(1), 1–29.
- Arcidiacono, P., Aucejo, E. M., Fang, H., & Spenner, K. I. (2011). Does affirmative action lead to mismatch? a new test and evidence. *Quantitative Economics*, 2(3), 303–333.
- Arcidiacono, P., Aucejo, E. M., & Hotz, V. J. (2016). University differences in the graduation of minorities in stem fields: Evidence from california. *American Economic Review*, 106(3), 525–62.
- Arcidiacono, P., & Lovenheim, M. (2016). Affirmative action and the quality-fit trade-off. *Journal of Economic Literature*, 54(1), 3–51.
- Autor, D. H. (2014). Skills, education, and the rise of earnings inequality among the “other 99 percent”. *Science*, 344(6186), 843–851.
- Bagde, S., Epple, D., & Taylor, L. (2016). Does affirmative action work? caste, gender, college quality, and academic success in india. *American Economic Review*, 106(6), 1495–1521.

- Baker, A. C., Larcker, D. F., & Wang, C. C. (2022). How much should we trust staggered difference-in-differences estimates? *Journal of Financial Economics*, 144(2), 370–395.
- Biró, P. (2012). University admission practices-hungary. *www. matching-in-practice. eu* accessed July, 21, 2014.
- Black, D. A., & Smith, J. A. (2004). How robust is the evidence on the effects of college quality? evidence from matching. *Journal of econometrics*, 121(1-2), 99–124.
- Black, D. A., & Smith, J. A. (2006). Estimating the returns to college quality with multiple proxies for quality. *Journal of labor Economics*, 24(3), 701–728.
- Black, S. E., Denning, J. T., & Rothstein, J. (2020). *Winners and losers? the effect of gaining and losing access to selective colleges on education and labor market outcomes* (Tech. Rep.). National Bureau of Economic Research.
- Bleemer, Z. (2021). Top percent policies and the return to postsecondary selectivity, by zachary bleemer, cshe 1.21.
- Bleemer, Z. (2022). Affirmative action, mismatch, and economic mobility after california’s proposition 209. *The Quarterly Journal of Economics*, 137(1), 115–160.
- Card, D., & Rothstein, J. (2007). Racial segregation and the black–white test score gap. *Journal of Public Economics*, 91(11-12), 2158–2184.
- Chen, L. (2012). University admission practices–ireland. *Mip country profile*, 8.
- Chetty, R., Friedman, J. N., Saez, E., Turner, N., & Yagan, D. (2017). *Mobility report cards: The role of colleges in intergenerational mobility* (Tech. Rep.). national bureau of economic research.
- Cohodes, S. R., & Goodman, J. S. (2014). Merit aid, college quality, and college completion: Massachusetts’ adams scholarship as an in-kind subsidy. *American Economic Journal: Applied Economics*, 6(4), 251–85.
- Concha-Arriagada, C. (2022). *Should i stay, or should i go? strategic responses to improve college admission chances* (Tech. Rep.). Working paper.
- Cullen, J. B., Jacob, B. A., & Levitt, S. (2006). The effect of school choice on participants: Evidence from randomized lotteries. *Econometrica*, 74(5), 1191–1230.
- Cullen, J. B., Long, M. C., & Reback, R. (2013). Jockeying for position: Strategic high school choice under texas’ top ten percent plan. *Journal of Public Economics*, 97, 32–48.
- Dale, S. B., & Krueger, A. B. (2002). Estimating the payoff to attending a more selective college: An application of selection on observables and unobservables. *The Quarterly Journal of Economics*, 117(4), 1491–1527.
- Dale, S. B., & Krueger, A. B. (2014). Estimating the effects of college characteristics over the career using administrative earnings data. *Journal of human resources*, 49(2), 323–358.
- De Chaisemartin, C., & d’Haultfoeuille, X. (2020). Two-way fixed effects estimators with heterogeneous treatment effects. *American Economic Review*, 110(9), 2964–96.

- Deming, D. J. (2011). Better schools, less crime? *The Quarterly Journal of Economics*, 126(4), 2063–2115.
- Deming, D. J., Hastings, J. S., Kane, T. J., & Staiger, D. O. (2014). School choice, school quality, and postsecondary attainment. *American Economic Review*, 104(3), 991–1013.
- Dillon, E. W., & Smith, J. A. (2017). Determinants of the match between student ability and college quality. *Journal of Labor Economics*, 35(1), 45–66.
- Dillon, E. W., & Smith, J. A. (2020). The consequences of academic match between students and colleges. *Journal of Human Resources*, 55(3), 767–808.
- Dubins, L. E., & Freedman, D. A. (1981). Machiavelli and the gale-shapley algorithm. *The American Mathematical Monthly*, 88(7), 485–494.
- Estevan, F., Gall, T., Legros, P., & Newman, A. F. (2017). The top-ten way to integrate high schools.
- Fack, G., Grenet, J., & He, Y. (2019). Beyond truth-telling: Preference estimation with centralized school choice and college admissions. *American Economic Review*, 109(4), 1486–1529.
- Fajnzylber, E., Lara, B., & León, T. (2019). Increased learning or gpa inflation? evidence from gpa-based university admission in chile. *Economics of Education Review*, 72, 147–165.
- Gale, D., & Shapley, L. S. (1962). College admissions and the stability of marriage. *The American Mathematical Monthly*, 69(1), 9–15.
- Goodman, J., Hurwitz, M., Smith, J., & Fox, J. (2015). The relationship between siblings’ college choices: Evidence from one million sat-taking families. *Economics of Education Review*, 48, 75–85.
- Goodman-Bacon, A. (2021). Difference-in-differences with variation in treatment timing. *Journal of Econometrics*, 225(2), 254–277.
- Haeringer, G., & Klijn, F. (2009). Constrained school choice. *Journal of Economic theory*, 144(5), 1921–1947.
- Hastings, J., Kane, T. J., & Staiger, D. O. (2009). Heterogeneous preferences and the efficacy of public school choice. *NBER Working Paper*, 2145, 1–46.
- Hastings, J. S., Neilson, C. A., & Zimmerman, S. D. (2013). *Are some degrees worth more than others? evidence from college admission cutoffs in chile* (Tech. Rep.). National Bureau of Economic Research.
- Hoekstra, M. (2009). The effect of attending the flagship state university on earnings: A discontinuity-based approach. *The review of economics and statistics*, 91(4), 717–724.
- Holzer, H., & Neumark, D. (2000). Assessing affirmative action. *Journal of Economic literature*, 38(3), 483–568.
- Kapor, A., Karnani, M., & Neilson, C. (2022). *Aftermarket frictions and the cost of off-platform options in centralized assignment mechanisms* (Tech. Rep.). National Bureau

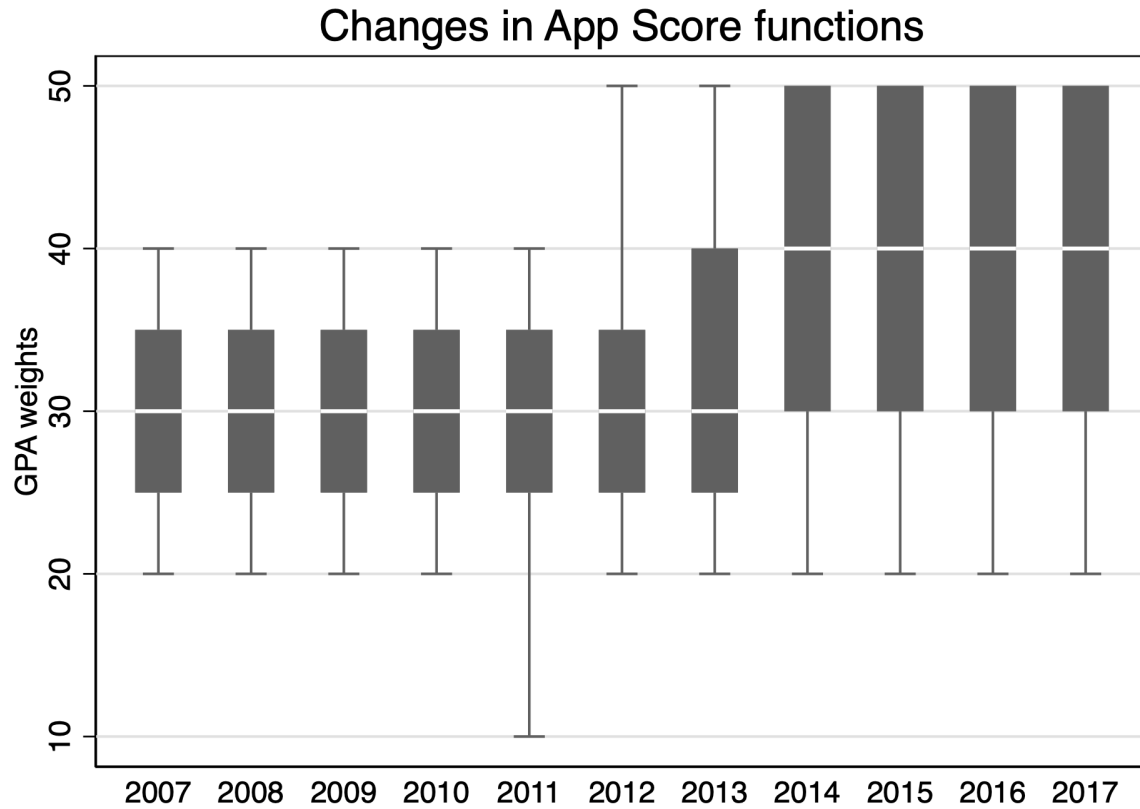
of Economic Research.

- Kapor, A., et al. (2020). *Distributional effects of race-blind affirmative action* (Tech. Rep.).
- Kirkeboen, L. J., Leuven, E., & Mogstad, M. (2016). Field of study, earnings, and self-selection. *The Quarterly Journal of Economics*, 131(3), 1057–1111.
- Krueger, A., Rothstein, J., & Turner, S. (2006). Race, Income, and College in 25 Years: Evaluating Justice O'Connor's Conjecture. *American Law and Economics Review*, 8(2), 282-311. Retrieved from <https://ideas.repec.org/a/oup/amlawe/v8y2006i2p282-311.html>
- Lafortune, J., Figueroa, N., & Saenz, A. (2016). *Do you like me enough? the impact of restricting preferences ranking in a university matching process* (Tech. Rep.). Working Paper.
- Larroucau, T., & Rios, I. (2018). Do “short-list” students report truthfully? strategic behavior in the chilean college admissions problem. *Preprint, submitted September, 1*(10.13140).
- Larroucau, T., & Rios, I. (2020). *Dynamic college admissions and the determinants of students' college retention*. unpublished manuscript, University of Pennsylvania.
- Larroucau, T., Ríos, I., & Mizala, A. (2015). Efecto de la incorporación del ranking de notas en el proceso de admisión a las universidades chilenas. *Pensamiento Educativo, Revista de Investigación Latinoamericana (PEL)*, 52(1), 95–118.
- Lindahl, L., & Regnér, H. (2005). College choice and subsequent earnings: Results using swedish sibling data. *Scandinavian Journal of Economics*, 107(3), 437–457.
- Loury, L. D., & Garman, D. (1993). Affirmative action in higher education. *The American Economic Review*, 83(2), 99–103.
- Luffade, M. (2017). The value of information in centralized school choice systems (job market paper).
- Mello, U. (2021). Affirmative action and the choice of schools.
- Mello, U. (2022). Centralized admissions, affirmative action, and access of low-income students to higher education. *American Economic Journal: Economic Policy*, 14(3), 166–97.
- Mora, R., & Romero-Medina, A. (2001). Understanding preference formation in a matching market.
- Mountjoy, J. (2022). Community colleges and upward mobility. *American Economic Review*, 112(8), 2580–2630.
- Mountjoy, J., & Hickman, B. R. (2021). *The returns to college (s): Relative value-added and match effects in higher education* (Tech. Rep.). National Bureau of Economic Research.
- Otero, S., Barahona, N., & Dobbin, C. (2021). *Affirmative action in centralized college admission systems: Evidence from brazil* (Tech. Rep.). Working paper.

- Pathak, P. A., & Sönmez, T. (2013). School admissions reform in chicago and england: Comparing mechanisms by their vulnerability to manipulation. *American Economic Review*, 103(1), 80–106.
- Prakhov, I., & Yudkevich, M. (2019). University admission in russia: Do the wealthier benefit from standardized exams? *International Journal of Educational Development*, 65, 98–105.
- Rios, I., Larroucau, T., Parra, G., & Cominetti, R. (2021). Improving the chilean college admissions system. *Operations Research*, 69(4), 1186–1205.
- Roth, A. E. (1982). The economics of matching: Stability and incentives. *Mathematics of operations research*, 7(4), 617–628.
- Rothstein, J., & Yoon, A. H. (2008). *Affirmative action in law school admissions: What do racial preferences do?* (Tech. Rep.). National Bureau of Economic Research.
- Rothstein, J. M. (2004). College performance predictions and the sat. *Journal of Econometrics*, 121(1-2), 297–317.
- Sander, R., & Taylor, S. (2012). *Mismatch: How affirmative action hurts students it's intended to help, and why universities won't admit it*. Basic Books.
- Saygin, P. O. (2016). Gender differences in preferences for taking risk in college applications. *Economics of Education Review*, 52, 120–133.
- Sowell, T. (1972). *Black education: Myths and tragedies*. David McKay.
- Turner, N. (2020). Income segregation and intergenerational mobility across colleges in the united states.
- Zimmerman, S. D. (2014). The returns to college admission for academically marginal students. *Journal of Labor Economics*, 32(4), 711–754.
- Zimmerman, S. D. (2019). Elite colleges and upward mobility to top jobs and top incomes. *American Economic Review*, 109(1), 1–47.
- Zwick, R., & Greif Green, J. (2007). New perspectives on the correlation of sat scores, high school grades, and socioeconomic factors. *Journal of Educational Measurement*, 44(1), 23–45.

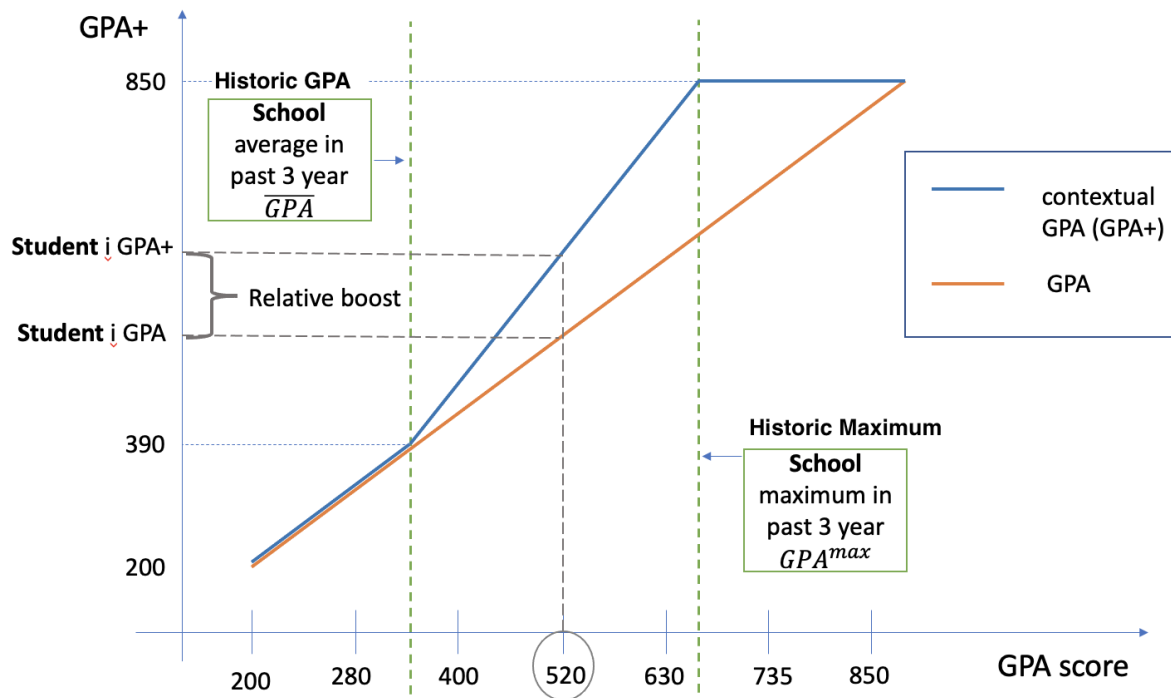
Figures and Tables

Figure 1: Weights of GPA components ($\text{GPA} + \text{GPA}^+$) in application scores by year



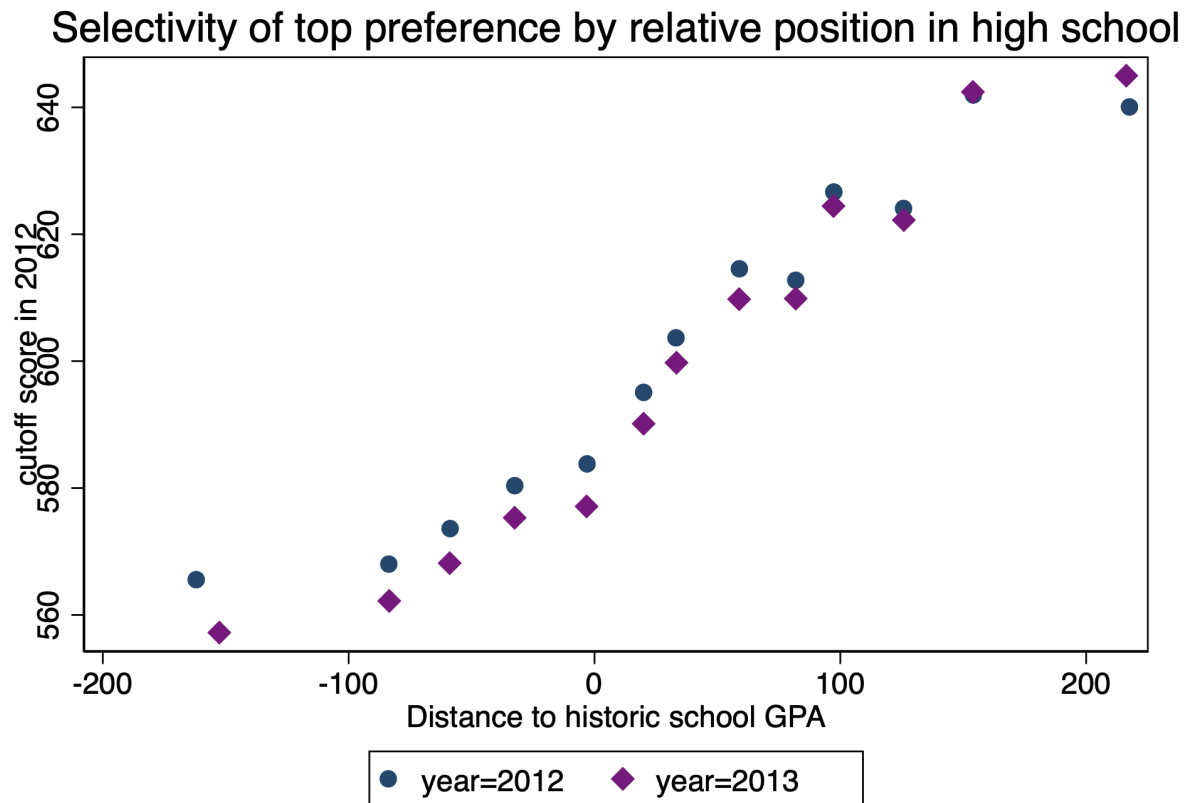
Notes: This figure shows the whisker plots for the distribution of the weights of the GPA components assigned by programs in the application score formula. The middle box represents 50% of the data, the white line corresponds to the median weight and the maximum and minimum values are displayed with vertical lines (“whiskers”).

Figure 2



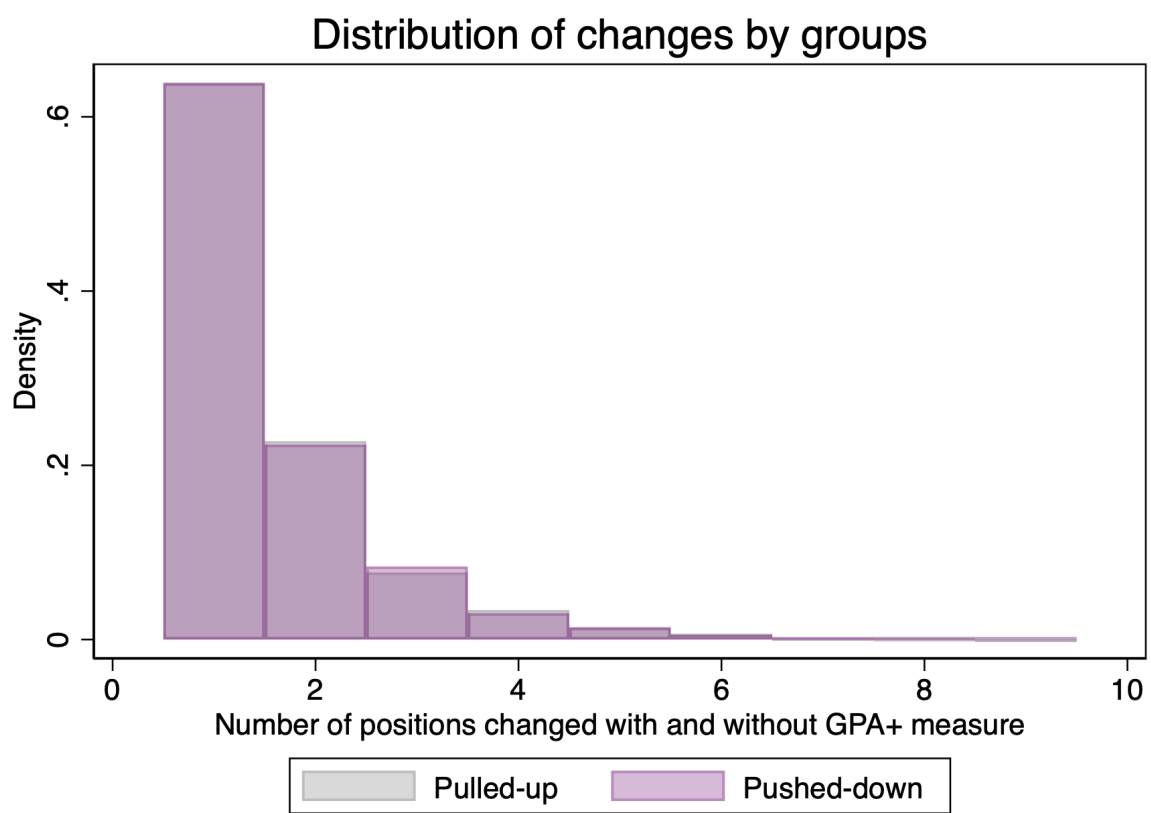
Notes: exemplary figure to show how GPA^+ depends on school averages and how it relates to the GPA score. Boost is obtained from the difference between GPA^+ score and GPA.

Figure 3



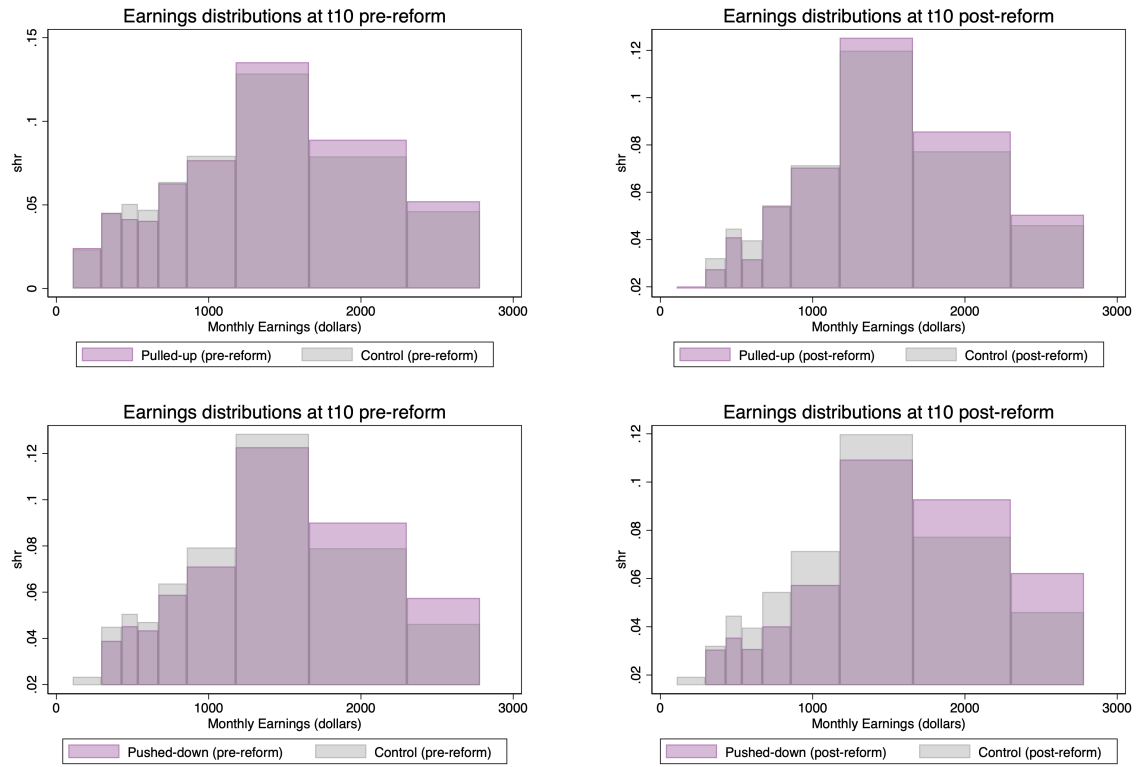
Notes: binscatter of the selectivity of the 1st preference by boost. Selectivity measure as the cutoff (application score of the last person admitted in the programs, measured pre-reform) of the program listed 1st. The x-axis have the GPA^+ measure, but centered around the average score of the school. By centered at the school average we have that positive values correspond to the boost score.

Figure 4



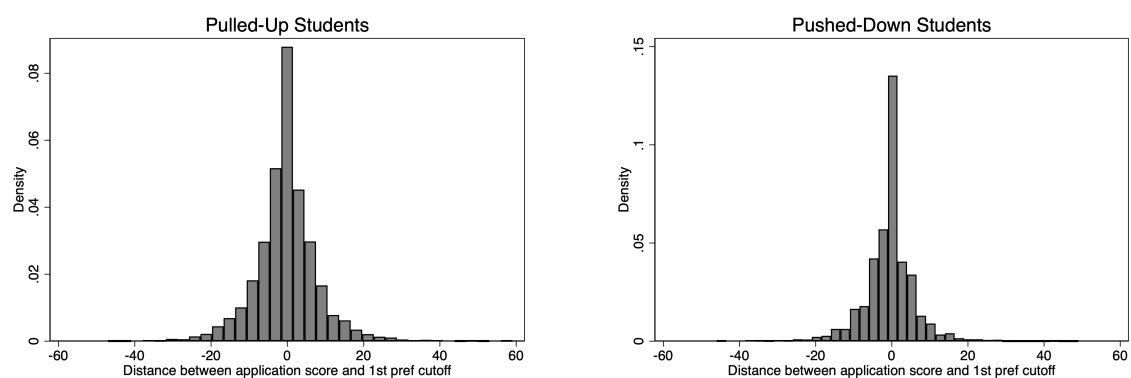
Notes: distribution of pulled-up and pushed-down students based on the number of positions moved in their ranking between admission with and without GPA⁺.

Figure 5: Earnings distribution



Notes: Earnings distribution for pulled-up and pushed-down groups, relative to unaffected, 10 years after application. Figures on the left show earnings distribution for students in cohort 2012 (pre-reform) and figures on the right show earnings distribution for students in cohort 2013 (post-reform).

Figure 6: Distribution by distance between application score and cutoff for 1st preference



Notes: Histogram for pulled-up and pushed-down groups for students who win or lose their first preference when the relative GPA is considered.

Table 1: Distribution of students reporting rankings by year

Ranking		2012	2013
1	Total N	116,336	118,208
	Only 1 (%)	0.07	0.07
	Up to 1 (%)	0.07	0.07
2	Total N	108,715	110,264
	Only 2 (%)	0.09	0.10
	Up to 2 (%)	0.16	0.17
3	Total N	98,166	98,245
	Only 3 (%)	0.17	0.20
	Up to 3 (%)	0.32	0.37
4	Total N	78,828	74,152
	Only 4 (%)	0.16	0.17
	Up to 4 (%)	0.48	0.55
5	Total N	60,420	53,693
	Only 5 (%)	0.14	0.14
	Up to 5 (%)	0.62	0.68
6	Total N	44,322	37,403
	Only 6 (%)	0.10	0.09
	Up to 6 (%)	0.72	0.78
7	Total N	32,720	26,182
	Only 7 (%)	0.07	0.07
	Up to 7 (%)	0.79	0.84
8	Total N	24,208	18,477
	Only 8 (%)	0.06	0.05
	Up to 8 (%)	0.85	0.89
9	Total N	17,041	12,572
	Only 9 (%)	0.04	0.03
	Up to 1 (%)	0.89	0.92
10	Total N	12,582	9,167
	Only 10 (%)	0.11	0.08
	Up to 10 (%)	1.00	1.00

Notes: The table shows the total number of students reporting each ranking, the percentage of students reporting a total of each ranking, and the percentage of students reporting each ranking or less options. In 2011 the maximum number of choices was increase and students were nudge to take advantage of that and list 10 options.

Table 2: Summary Statistics for Groups of Interest

	Unaffected		Pulled-up		Pushed-down	
	2012	2013	2012	2013	2012	2013
N	108,167	109,440	3,753	4,515	4,416	4,253
Female (%)	53	52	62	60	41	40
Public School (%)	28	27	29	29	26	25
Voucher School (%)	53	54	60	60	47	47
Private School (%)	19	18	10	11	27	28
Family Inc (\$/mo)	689	714	573	594	809	869
Father with HS (%)	67	67	64	61	74	75
Mother with HS (%)	73	73	69	70	78	79
Father with College (%)	26	26	20	19	34	35
Mother with College (%)	21	21	16	16	27	29
Capital City (%)	39	39	46	46	54	53
Std Math	0.68	0.65	0.74	0.65	1.05	1.13
Std Verbal	0.66	0.65	0.70	0.62	1.01	1.04
Std GPA	0.75	0.73	1.40	1.28	0.42	0.58
Boost score	21	22	60	57	6	8

Notes: This table shows the summary statistics for the groups of interest, the year before and after the reform.

Table 3: Distribution of students by ranking with and without GPA⁺

With GPA ⁺		Without GPA ⁺									
Ranking	1	2	3	4	5	6	7	8	9	10	NA
1	48,434	1,166	415	139	59	31	16	2	7	0	397
%	0.44	0.26	0.09	0.03	0.01	0.01	0.00	0.00	0.00	0.00	0.09
2	1,149	19,404	530	187	72	36	12	5	2	3	262
%	0.27	0.18	0.12	0.04	0.02	0.01	0.00	0.00	0.00	0.00	0.06
3	347	590	10,500	240	113	23	11	5	3	2	202
%	0.08	0.14	0.10	0.05	0.03	0.01	0.00	0.00	0.00	0.00	0.04
4	124	184	279	4,303	96	50	18	6	0	1	92
%	0.03	0.04	0.07	0.04	0.02	0.01	0.00	0.00	0.00	0.00	0.02
5	37	72	124	101	2,288	47	22	5	5	1	58
%	0.01	0.02	0.03	0.02	0.02	0.01	0.00	0.00	0.00	0.00	0.01
6	397	28	36	46	59	1,158	21	13	5	1	45
%	0.01	0.01	0.01	0.01	0.01	0.01	0.00	0.00	0.00	0.00	0.01
7	9	12	12	16	22	21	651	12	8	3	26
%	0.00	0.00	0.00	0.00	0.01	0.00	0.01	0.00	0.00	0.00	0.01
8	2	3	10	5	8	7	14	338	5	4	11
%	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
9	2	1	3	4	3	3	5	14	193	8	7
%	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
10	1	1	0	5	0	1	3		2	123	5
%	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
NA	269	220	141	94	62	29	13	15	8	9	22,048
%	0.06	0.05	0.03	0.02	0.01	0.01	0.00	0.00	0.00	0.00	0.20

Notes: this table presents the number of students in 2013 with admission at different ranking of their rank order list with and without the GPA⁺. Values on green correspond to pushed-down cases and values on blue correspond to pulled-up students. Column 1 - row 1 (and all the diagonal) shows the number of students that with admission in their top choice under both regimes, therefore, they are classify as unaffected. The percentage value under the total number of students represent the proportion of students in that treatment group that have that combination of rankings.

Table 4: Distribution of pulled-up students by fields with and without GPA⁺

With GPA ⁺		Without GPA ⁺								
Ranking	MedOdon	Health	Sci	Engi	Tech	Business	Art	SocSci	Law	Educ
MedOdon	126	54	7	10	2	2	0	3	4	1
Health	3	425	51	21	34	21	2	32	2	44
Sci	0	19	57	23	25	3	0	5	0	10
Engi	1	7	23	359	111	51	2	5	2	2
Tech	0	9	19	79	266	20	20	7	1	10
Business	0	1	12	19	29	224	3	8	3	6
Art	0	0	1	0	7	0	18	7	0	4
SocSci	0	8	8	3	17	32	15	200	18	47
Law	0	0	0	1	4	11	0	29	63	9
Educ	0	4	4	6	13	9	6	29	4	194

Notes: Total number of pulled-up student in 2013 in each field combination based on the field of the program that they get admitted with the GPA⁺ and the field of the program that they get admitted without the GPA⁺.

Table 5: Distribution of pushed-down students by fields with and without GPA⁺

With GPA ⁺	Without GPA ⁺									
Ranking	MedOdon	Health	Sci	Engi	Tech	Business	Art	SocSci	Law	Educ
MedOdon	126	54	7	10	2	2	0	3	4	1
Health	3	425	51	21	34	21	2	32	2	44
Sci	0	19	57	23	25	3	0	5	0	10
Engi	1	7	23	359	111	51	2	5	2	2
Tech	0	9	19	79	266	20	20	7	1	10
Business	0	1	12	19	29	224	3	8	3	6
Art	0	0	1	0	7	0	18	7	0	4
SocSci	0	8	8	3	17	32	15	200	18	47
Law	0	0	0	1	4	11	0	29	63	9
Educ	0	4	4	6	13	9	6	29	4	194

Notes: Total number of pushed-down student in each field combination based on the field of the program that they get admitted with the GPA⁺ and the field of the program that they get admitted without the GPA⁺.

Table 6: Difference-in-differences estimates for 2012 and 2011

	(1)	(2)	(3)	(4)
	Enroll	Enroll	Grad by 8yr	Grad by 8yr
Pulled-Up	0.001	-0.015	0.006	-0.009
	(0.011)	(0.010)	(0.011)	(0.011)
Pushed-Down	-0.010	-0.002	0.007	0.005
	(0.009)	(0.009)	(0.011)	(0.010)
Observations	211,872	211,872	211,872	211,872
Controls		✓		✓

Robust standard errors in parentheses

Notes: columns 1 and 3 have the estimates from the difference-in-difference without controls and columns 2 and 4 have the estimates for the same outcomes but controlling by individual characteristics.

Table 7: Diff-in-diff estimates for enrollment

	(1) Enrollment	(2) Enrollment	(3) Non-Select	(4) Non-Select	(5) Enrollment	(6) Enrollment
Pulled-Up x after	0.199*** (0.011)	0.219*** (0.010)	-0.049*** (0.007)	-0.057*** (0.007)	0.165*** (0.0113)	0.175*** (0.0111)
Pushed-Down x after	-0.136*** (0.010)	-0.167*** (0.010)	0.023*** (0.007)	0.039*** (0.006)	-0.0947*** (0.00987)	-0.110*** (0.00965)
Obs.	234,544	234,544	234,544	234,544	186,734	186,734
Controls		✓		✓		✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.000	0.000	0.008	0.048	0.000	0.000

Robust standard errors in parentheses

Notes: columns 1 and 2 have estimates when the outcome is enrollment at the admission program. Column 3 and 4 have estimates for an indicator if the student enroll in a non-selective program. Columns 2 and 4 control for standardized test scores, GPA, family income, region, type of high school and gender. Column 5 and 6 restrict the sample to students with some admission offer under both regimes to capture enrollment effect on students at the intensive margin.

Table 8: Changes in peer characteristics at chosen programs

Program Charact.	Diff-in-Diff			Pre-Reform (\bar{x})	
	Pulled-up	Pushed-down	Control	Pulled-up	Pushed-down
Math (std)	0.264*** (0.008)	-0.212*** (0.007)	1.104 [0.610]	1.262 [0.603]	1.272 [0.611]
Verbal (std)	0.235*** (0.008)	-0.230*** (0.008)	1.114 [0.567]	1.240 [0.537]	1.261 [0.530]
GPA (std)	0.280*** (0.009)	-0.288*** (0.008)	1.165 [0.575]	1.317 [0.519]	1.276 [0.544]
Grad on time	0.044*** (0.007)	-0.040*** (0.006)	0.389 [0.263]	0.384 [0.271]	0.374 [0.267]
E(grad time)	0.026 (0.027)	-0.065** (0.027)	5.110 [0.725]	5.163 [0.779]	5.161 [0.820]

Robust standard errors in parentheses. Standard deviation in square brackets.

Notes: Columns 1 and 2 show the results for the main diff-in-diff specification for the outcome 5 different outcomes: (i) average math score of students enrolled at the chosen program pre-reform, (ii) average verbal score of students enrolled at the chosen program pre-reform, (iii) average GPA score of the students enrolled at the chosen program pre-reform, (iv) probability of graduation on time by the students enrolled at the chosen program pre-reform, (v) expected graduation time based on the class structure at the chosen program. Columns 3-5 show the averages and standard deviation of these variables for the 3 groups of interest, pre-reform.

Table 9: Effect on re-application by second year

	(1) Reapplication	(2) Reapplication
P-Up x after	-0.0387*** (0.0091)	-0.0374*** (0.0091)
P-Down x after	0.0739*** (0.0086)	0.0717*** (0.0086)
Obs.	234,544	234,544
Controls		✓

Robust standard errors in parentheses

Notes: diff-in-diff estimates using an indicator if the student participate on the application process on the second year.

Table 10: Effect on total enrollment up to 4 years after the reform

	(1) Enroll at t=1	(2) Enroll at t=2	(3) Enroll at t=3	(4) Enroll at t=4
P-Up x after	0.062*** (0.008)	0.018** (0.008)	0.015* (0.008)	0.003 (0.009)
P-Down x after	-0.077*** (0.007)	-0.008 (0.007)	0.005 (0.008)	0.008 (0.008)
Obs.	234,544	234,544	234,544	234,544
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: diff-in-diff estimates for an indicator is the student is enrolled in some program at different points in time. Column 1 shows the results for the same year of applications; column 2 for two years after the application process; column 3 for three years after and column 4 for four years after.

Table 11: Effect on peers test scores at enrollment up to 4 years after the reform

	(1) Selectivity at t=1	(2) Selectivity at t=2	(3) Selectivity at t=3	(4) Selectivity at t=4
P-Up x after	0.160*** (0.008)	0.093*** (0.009)	0.117*** (0.010)	0.107*** (0.012)
P-Down x after	-0.120*** (0.008)	-0.113*** (0.008)	-0.084*** (0.010)	-0.082*** (0.011)
Obs.	187,534	190,703	181,081	175,037
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: diff-in-diff estimates for average test score (standardized) of students at program chosen by applicants at different points in time. Column 1 through 4 show the results from 1st to 4th year since the moment of application.

Table 12: Effect on graduation from initial admission program

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.042*** (0.008)	0.072*** (0.009)	0.084*** (0.010)	0.043*** (0.010)
P-Down x after	-0.033*** (0.007)	-0.060*** (0.009)	-0.082*** (0.009)	-0.039*** (0.010)
Obs.	234,544	234,544	234,544	234,544
Controls	✓	✓	✓	✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.397	0.366	0.917	0.742

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table 13: Effect on graduation from initial admission program conditional on some admission offer with both mechanism

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.037*** (0.009)	0.065*** (0.011)	0.078*** (0.011)	0.034*** (0.012)
P-Down x after	-0.020** (0.009)	-0.040*** (0.010)	-0.062*** (0.011)	-0.017 (0.011)
Obs.	186,734	186,734	186,734	186,734
Controls	✓	✓	✓	✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.195	0.100	0.305	0.295

Robust standard errors in parentheses

Notes: Diff-in-diff results for the sample of students with some admission with and without the inclusion of the GPA⁺ measure. Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admitted program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table 14: Mismatch effect exercise

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.047*** (0.007)	0.078*** (0.009)	0.091*** (0.010)	0.049*** (0.010)
Obs.	234,529	234,529	234,529	234,529
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: This table shows the effect on graduation from admission into a more selective program. The diff-in-diff specification controls by the admission program without the relative GPA reform in order to ensure that estimation uses only variation from students with admission to more selective programs after the reform, and not from potential changes in the compositions of admission programs between 2012 and 2013.

Table 15: Effect on STEM applicants

	(1) Admission	(2) Enrollment	(3) Enrollment	(4) Grad by 8yr	(5) Grad or enroll by 8 yr
P-Up x after	0.061*** (0.010)	0.216*** (0.014)	0.169*** (0.015)	0.061*** (0.014)	0.052*** (0.016)
P-Down x after	-0.030*** (0.010)	-0.166*** (0.013)	-0.120*** (0.012)	-0.047*** (0.014)	-0.035** (0.016)
Obs.	234,544	110,791	97,350	97,350	97,350
Controls	✓	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Column 1 shows the coefficient for the indicator of admission offer in STEM using the main diff-in-diff specification. Column 2 shows the effect on enrollment for STEM applicants. Column 3 restrict the sample of column 2 only to students that have some admission offer with and without GPA. Column 4 present the effects on graduation for the same sample than column 3. Finally, column 5 present the results conditional on enrollment.

Table 16: Effects on graduation from any program

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad or enroll by 8 yr
P-Up x after	-0.003 (0.009)	0.005 (0.011)	-0.008 (0.011)	0.006 (0.010)
P-Down x after	-0.012 (0.009)	-0.015 (0.010)	-0.032*** (0.011)	-0.006 (0.010)
Obs.	234,544	234,544	234,544	234,544
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: diff-in-diff estimates for the indicator if the student graduate from some program by 6, 7 or 8 years. Column 4 show the results when the dependent variable takes the value of 1 if the student graduate or if the student is enrolled in some program 8 years after application.

Table 17: Effect on graduation from a selective program

	(1)	(2)	(3)	(4)
	Grad by 6yr	Grad by 7yr	Grad by 8yr	Grad or enroll by 8yr
P-Up x after	0.019** (0.009)	0.030*** (0.010)	0.019* (0.011)	0.024** (0.010)
P-Down x after	-0.022*** (0.008)	-0.028*** (0.010)	-0.046*** (0.010)	-0.030*** (0.010)
Obs.	234,544	234,544	234,544	234,544
Controls	✓ Selective	✓ Selective	✓ Selective	✓ Selective

Robust standard errors in parentheses

Notes: Columns 1-3 show the results for graduation from a selective program by 6, 7 or 8 years after application. Columns 4-5 show the same results for non-selective programs.

Table 18: Effects on graduation from a non-selective program

	(1)	(2)	(3)	(4)
	Grad by 6yr	Grad by 7yr	Grad by 8yr	Grad or enroll by 8yr
P-Up x after	-0.022*** (0.004)	-0.025*** (0.005)	-0.026*** (0.005)	-0.019*** (0.006)
P-Down x after	0.009*** (0.004)	0.013*** (0.004)	0.014*** (0.004)	0.022*** (0.005)
Obs.	234,544	234,544	234,544	234,544
Controls	✓ Non-Selective	Non-Selective	✓ Non-Selective	✓ Non-Selective

Robust standard errors in parentheses

Notes: Columns 1-3 show the results for graduation from a selective program by 6, 7 or 8 years after application. Columns 4-5 show the same results for non-selective programs.

Table 19: Differential effect for students with big and small changes in selectivity

	(1) Enrollment	(2) Grad by 8yr	(3) Grad by 8yr from any	(4) Grad or enroll by 8 yr
Small Pulled-Up x after	0.195*** (0.015)	0.069*** (0.016)	-0.034** (0.017)	0.005 (0.015)
Big Pulled-Up x after	0.151*** (0.016)	0.080*** (0.016)	0.007 (0.017)	-0.003 (0.015)
Small Pushed-Down x after	-0.108*** (0.013)	-0.040** (0.016)	0.000 (0.017)	0.007 (0.015)
Big Pushed-Down x after	-0.112*** (0.014)	-0.085*** (0.016)	-0.036** (0.017)	-0.014 (0.015)
Obs.	186,734	186,734	186,734	186,734
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Based on how much the average test score of the peers (selectivity of the programs) changed between the simulated program and the admission program, the pulled-up and pushed-down groups are split into big and small changes in selectivity. The sample contains only students who have some admission offer in both regimes. Column 1 shows how the reform affected enrollment for the four subgroups. Column 2 shows the change in the probability of program completion. Column 3 shows the effects of graduating from any program 8 years after admission. Column 4 shows the effects of graduating or still being in school 8 years after application, which takes into account the fact that students in the selective system may switch programs, which will cause them to graduate from college later.

Table 20: Effects on students move into or out of their 1st preference with and without GPA⁺

	(1) Enroll	(2) Select	(3) Grad by 8yr	(4) Grad (any) by 8yr	(5) Grad or enroll by 8 yr
P-Up x after	0.281*** (0.014)	0.283*** (0.007)	0.112*** (0.014)	-0.005 (0.015)	-0.002 (0.014)
P-Down x after	-0.179*** (0.013)	-0.228*** (0.007)	-0.101*** (0.015)	-0.026* (0.016)	0.003 (0.014)
Obs.	225,618	225,531	225,618	225,618	225,618
Controls	✓	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Column 1 presents the results for enrollment at their admission offer. Column 2 presents the estimates for the change in the average test scores at the program of enrollment. Column 3 to 5 show the estimates for graduation from the admission offer, from any program and for an indicator of graduate or still enroll 8 years after the application process.

Table 21: Difference-in-differences and RD comparison

	(1) RD	(2) DD Pulled-Up	(3) DD Pushed-Down
Enrollment	0.178*** (0.006)	0.281*** (0.014)	-0.179*** (0.013)
Selectivity	0.16*** (0.005)	0.157*** (0.010)	-0.168*** (0.010)
Grad from Admission	0.067*** (0.007)	0.112*** (0.014)	-0.101*** (.015)
Grad from Any	-0.004 (0.007)	-0.005 (0.015)	-0.026* (0.016)
Graduation or Enroll	-0.004 (0.007)	-0.002 (0.014)	0.003 (0.014)
Observations	118,205	225,618	225,618
Controls	✓	✓	✓

Robust standard errors in parentheses

Notes: Column 1 presents the results for the RD specification. It compares people with very similar application scores for their 1st preference in 2013. Column 2 presents the results for pulled-up group with the diff-in-diff specification. It compares students who are above the threshold for their 1st choice to their “past cohort selves” who did not get their 1st choice. Their “past cohort selves” have the same score, presumably not too far below the threshold, but did not get treated. Column 3 presents the results for pushed-down group with the diff-in-diff specification. It compares people who are below the 2013 threshold to their “past cohort selves” who did get their 1st choice.

The equity and efficiency effects of a relative GPA reward in college admissions

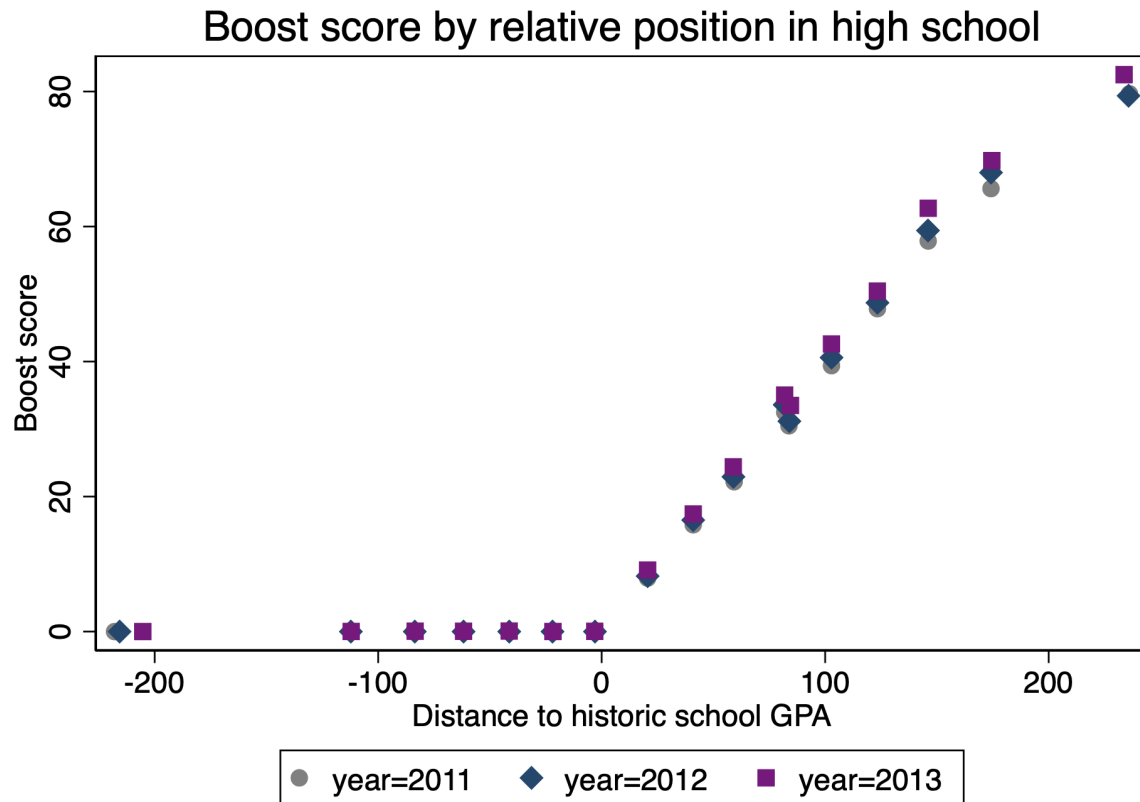
Appendix

Tatiana Reyes - UC Berkeley¹

¹Email: tatiana_reyes@berkeley.edu. This version: November 30, 2022.

A Appendix

Figure A.I



Notes: boost score for cohort 2011, 2012 and 2013. For 2013 GPA^+ (and the inferred boost) was provided on the application data. For 2011 and 2012 boost was calculated according to the GPA^+ formula using education records of the universe of high school students graduated between 2008 and 2012.

Table A.I: Enrollment rates at admission program by groups, before and after the reform

Total	Unaffected	Pulled-Up	Pushed-Down
Enrollment Pre-Reform (2012)	0.80	0.83	0.91
Enrollment Reform (2013)	0.79	0.87	0.85
Difference	-0.01	0.04	-0.06

Program	Unaffected	Pulled-Up	Pushed-Down
Enrollment Pre-Reform (2012)	0.60	0.53	0.78
Enrollment Reform (2013)	0.62	0.75	0.66
Difference	0.02	0.22	-0.12

Non-selective	Unaffected	Pulled-Up	Pushed-Down
Enrollment Pre-Reform (2012)	0.13	0.12	0.08
Enrollment Reform (2013)	0.10	0.05	0.08
Difference	-0.03	-0.07	0.00

Notes: averages for a variable that indicates if the student choose to enroll in the admission assignment. The difference by group, between after and before the reform is shown in the 3rd row.

Table A.II: Effect on graduation from admission program conditional on enrollment

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	-0.012 (0.012)	-0.009 (0.013)	-0.010 (0.014)	-0.000 (0.014)
P-Down x after	-0.005 (0.010)	-0.018 (0.012)	-0.034*** (0.012)	0.001 (0.012)
Obs.	144,540	144,540	144,540	144,540
Controls	✓	✓	✓	✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.302	0.126	0.020	0.989

Robust standard errors in parentheses

Notes: Diff-in-diff results for the sample of students that enroll in 1st year. Columns 1-3 show estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table A.III: Graduation averages from any program by groups, before and after the reform

	Unaffected	Pulled-Up	Pushed-Down
Grad by 6yr Pre-Reform (2012)	0.22	0.24	0.21
Grad by 6yr Reform (2013)	0.21	0.22	0.19
Difference	-0.01	-0.02	-0.02
	Unaffected	Pulled-Up	Pushed-Down
Grad by 7yr Pre-Reform (2012)	0.36	0.40	0.35
Grad by 7yr Reform (2013)	0.34	0.37	0.34
Difference	-0.02	-0.03	-0.01
	Unaffected	Pulled-Up	Pushed-Down
Grad by 8yr Pre-Reform (2012)	0.46	0.51	0.47
Grad by 8yr Reform (2013)	0.42	0.45	0.42
Difference	-0.04	-0.06	-0.05

Notes: averages for a variable that indicates if the student graduates from some program (selective or non-selective). The difference by group, between after and before the reform is shown in the 3rd row.

B Heterogeneity Analysis

Table B.I: Heterogeneity: Effects on enrollment by gender, family income and boost

	(1) Enrollment	(2) Enrollment	(3) Enrollment
P-Up x after	0.193*** (0.017)	0.192*** (0.014)	0.080*** (0.029)
P-Down x after	-0.148*** (0.013)	-0.145*** (0.011)	-0.148*** (0.012)
P-Up x after x Characteristic	0.009 (0.022)	0.020 (0.021)	0.155*** (0.031)
P-Down x after x Characteristic	0.027 (0.020)	0.033 (0.022)	0.012 (0.020)
After x Characteristic	-0.005 (0.004)	-0.002 (0.004)	-0.023*** (0.004)
Obs.	234,544	234,544	234,544
Controls	✓	✓	✓
Characteristic	Female	Low Income	Boost
PU Fraction	61%	45%	85%
PD Fraction	41%	30%	32%

Robust standard errors in parentheses

Notes: main diff-in-diff specification for enrollment fully interacted with (i) female indicator, (ii) low income indicator, and (iii) boost indicator.

Table B.II: Heterogeneity: Effects on graduation from same program by gender, family income and boost

	(1) Grad by 8yr	(2) Grad by 8yr	(3) Grad by 8yr
P-Up x after	0.042*** (0.014)	0.082*** (0.013)	0.025 (0.022)
P-Down x after	-0.051*** (0.012)	-0.058*** (0.012)	-0.055*** (0.011)
P-Up x after x Characteristic	0.047** (0.019)	-0.025 (0.020)	0.065*** (0.025)
P-Down x after x Characteristic	-0.020 (0.020)	-0.013 (0.021)	-0.049** (0.021)
After x Characteristic	0.010*** (0.004)	0.006* (0.004)	-0.015*** (0.004)
Obs.	234,544	234,544	234,544
Controls	✓	✓	✓
Characteristic	Female	Low Income	Boost
PU Fraction	61%	45%	85%
PD Fraction	41%	30%	32%

Robust standard errors in parentheses

Notes: main diff-in-diff specification for graduation from assigned program fully interacted with (i) female indicator, (ii) low income indicator, and (iii) boost indicator.

Table B.III: Heterogeneity: Effects on graduation from any program by gender, family income and boost

	(1) Grad by 8yr	(2) Grad by 8yr	(3) Grad by 8yr
P-Up x after	-0.023 (0.017)	-0.008 (0.015)	0.006 (0.028)
P-Down x after	-0.004 (0.014)	-0.002 (0.013)	-0.008 (0.013)
P-Up x after x Characteristic	0.003 (0.023)	-0.030 (0.023)	-0.018 (0.031)
P-Down x after x Characteristic	-0.008 (0.022)	-0.026 (0.024)	-0.042* (0.023)
After x Characteristic	0.003 (0.004)	0.011*** (0.004)	-0.015*** (0.004)
Obs.	234,544	234,544	234,544
Controls	✓	✓	✓
Characteristic	Female	Low Income	Boost
PU Fraction	61%	45%	85%
PD Fraction	41%	30%	32%

Robust standard errors in parentheses

Notes: main diff-in-diff specification for graduation from any program fully interacted with (i) female indicator, (ii) low income indicator, and (iii) boost indicator.

Table B.IV: Heterogeneity: Effects on graduation from a selective program by gender, family income and boost

	(1) Grad by 8yr	(2) Grad by 8yr	(3) Grad by 8yr
P-Up x after	-0.015 (0.017)	0.006 (0.015)	0.038 (0.026)
P-Down x after	-0.006 (0.014)	-0.007 (0.013)	-0.022* (0.013)
P-Up x after x Characteristic	0.026 (0.022)	-0.010 (0.022)	-0.025 (0.029)
P-Down x after x Characteristic	-0.022 (0.022)	-0.035 (0.023)	-0.031 (0.023)
After x Characteristic	0.018*** (0.004)	0.021*** (0.004)	-0.021*** (0.004)
Obs.	234,544	234,544	234,544
Controls	✓	✓	✓
Characteristic	Female	Low Income	Boost
PU Fraction	61%	45%	85%
PD Fraction	41%	30%	32%

Robust standard errors in parentheses

Notes: main diff-in-diff specification for graduation from a selective program fully interacted with (i) female indicator, (ii) low income indicator, and (iii) boost indicator.

Table B.V: Heterogeneity: Effects on graduation or enroll after 8 years from a selective program by gender, family income and boost

	(1) Grad or enroll by 8 yr	(2) Grad or enroll by 8 yr	(3) Grad or enroll by 8 yr
P-Up x after	-0.015 (0.017)	-0.001 (0.013)	-0.030 (0.029)
P-Down x after	0.028** (0.013)	0.014 (0.012)	0.013 (0.013)
P-Up x after x Characteristic	0.005 (0.021)	-0.021 (0.021)	0.037 (0.031)
P-Down x after x Characteristic	-0.014 (0.020)	0.020 (0.022)	-0.013 (0.020)
After x Characteristic	0.010** (0.004)	0.013*** (0.004)	-0.007* (0.004)
Obs.	234,544	234,544	234,544
Controls	✓	✓	✓
Characteristic	Female	Low Income	Boost
PU Fraction	61%	45%	85%
PD Fraction	41%	30%	32%

Robust standard errors in parentheses

Notes: main diff-in-diff specification for graduation or enrollment after 8 year from a selective program fully interacted with (i) female indicator, (ii) low income indicator, and (iii) boost indicator.

C Main results from Section 8, boost sensitivity

Table C.I: Diff-in-diff estimates for enrollment

	(1) Enrollment	(2) Enrollment	(3) Non-Select	(4) Non-Select
P-Up x after	0.197*** (0.011)	0.219*** (0.011)	-0.047*** (0.007)	-0.056*** (0.007)
P-Down x after	-0.136*** (0.010)	-0.167*** (0.010)	0.023*** (0.006)	0.039*** (0.006)
Obs.	233,789	233,789	233,789	233,789
Controls		✓		✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.000	0.000	0.014	0.064

Robust standard errors in parentheses

Notes: columns 1 and 2 have estimates when the outcome is enrollment at the admission program. Column 3 and 4 have estimates for an indicator if the student enroll in a non-selective program. Columns 2 and 4 control for standardized test scores, GPA, family income, region, type of high school and gender.

Table C.II: Diff-in-diff estimates for graduation from GPA⁺ program on sample without boost > 150

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.041*** (0.008)	0.070*** (0.009)	0.083*** (0.010)	0.041*** (0.010)
P-Down x after	-0.033*** (0.007)	-0.060*** (0.009)	-0.082*** (0.009)	-0.038*** (0.010)
Obs.	233,789	233,789	233,789	233,789
Controls	✓	✓	✓	✓
Test	0	0	0	0
p-value	0.483	0.418	0.968	0.852

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table C.III: Diff-in-diff estimates for any graduation on sample without boost > 150

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad or enroll by 8 yr
P-Up x after	-0.005 (0.010)	0.005 (0.011)	-0.007 (0.011)	0.007 (0.010)
P-Down x after	-0.012 (0.009)	-0.014 (0.010)	-0.032*** (0.011)	-0.006 (0.010)
Obs.	233,789	233,789	233,789	233,789
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

D Main results from Section 8, long programs

Table D.I: Diff-in-diff estimates for enrollment on sample without long programs

	(1)	(2)	(3)	(4)
	Enrollment	Enrollment	Non-Select	Non-Select
P-Up x after	0.210*** (0.014)	0.238*** (0.013)	-0.067*** (0.010)	-0.078*** (0.009)
P-Down x after	-0.141*** (0.013)	-0.177*** (0.013)	0.025*** (0.009)	0.043*** (0.009)
Obs.	178,760	178,760	178,760	178,760
Controls		✓		✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.000	0.001	0.002	0.009

Robust standard errors in parentheses

Notes: results for sample without students in programs with 6 or 7 expected year. Columns 1 and 2 have estimates when the outcome is enrollment at the admission program. Column 3 and 4 have estimates for an indicator if the student enroll in a non-selective program. Columns 2 and 4 control for standardized test scores, GPA, family income, region, type of high school and gender.

Table D.II: Diff-in-diff estimates for graduation from GPA⁺ program on sample without long programs

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.065*** (0.011)	0.084*** (0.012)	0.098*** (0.012)	0.061*** (0.013)
P-Down x after	-0.051*** (0.011)	-0.077*** (0.012)	-0.087*** (0.012)	-0.061*** (0.013)
Obs.	178,760	178,760	178,760	178,760
Controls	✓	✓	✓	✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.367	0.667	0.552	0.981

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table D.III: Diff-in-diff estimates for any graduation on sample without long programs

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad or enroll by 8 yr
P-Up x after	0.008 (0.013)	0.007 (0.014)	0.008 (0.014)	0.009 (0.013)
P-Down x after	-0.030** (0.012)	-0.034** (0.013)	-0.040*** (0.014)	-0.008 (0.013)
Obs.	178,760	178,760	178,760	178,760
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Main results from Section 8, extra long programs

Table D.IV: Diff-in-diff estimates for enrollment on sample without extra long programs

	(1) Enrollment	(2) Enrollment	(3) Non-Select	(4) Non-Select
P-Up x after	0.195*** (0.011)	0.216*** (0.011)	-0.052*** (0.007)	-0.060*** (0.007)
P-Down x after	-0.136*** (0.010)	-0.167*** (0.010)	0.023*** (0.007)	0.038*** (0.006)
Obs.	228,741	228,741	228,741	228,741
Controls		✓		✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.000	0.001	0.004	0.029

Robust standard errors in parentheses

Notes: results for sample without students in programs with 6 or 7 expected year. Columns 1 and 2 show estimates when the outcome is enrollment at the admission program. Column 3 and 4 have estimates for an indicator if the student enroll in a non-selective program. Columns 2 and 4 control for standardized test scores, GPA, family income, region, type of high school and gender.

Table D.V: Diff-in-diff estimates for graduation from GPA⁺ program on sample without extra long programs

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.047*** (0.008)	0.066*** (0.009)	0.080*** (0.010)	0.042*** (0.010)
P-Down x after	-0.034*** (0.008)	-0.055*** (0.009)	-0.072*** (0.010)	-0.035*** (0.010)
Obs.	228,741	228,741	228,741	228,741
Controls	✓	✓	✓	✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.272	0.411	0.547	0.623

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table D.VI: Diff-in-diff estimates for any graduation on sample without extra long programs

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad or enroll by 8 yr
P-Up x after	0.002 (0.010)	0.001 (0.011)	-0.003 (0.011)	0.005 (0.010)
P-Down x after	-0.014 (0.009)	-0.011 (0.010)	-0.023** (0.011)	-0.005 (0.010)
Obs.	228,741	228,741	228,741	228,741
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

E Main results with clustered standard errors at school-year level

Table E.I: Diff-in-diff estimates for enrollment at admission program with school-year cluster standard errors

	(1) Enrollment	(2) Enrollment	(3) Non-Select	(4) Non-Select
P-Up x after	0.199*** (0.012)	0.219*** (0.011)	-0.049*** (0.007)	-0.057*** (0.007)
P-Down x after	-0.136*** (0.011)	-0.167*** (0.011)	0.023*** (0.007)	0.039*** (0.007)
Obs.	234,544	234,544	234,544	234,544
Controls		✓		✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.000	0.001	0.014	0.063

Robust standard errors in parentheses

Notes: Columns 1 and 2 show estimates when the outcome is enrollment at the admission program. Column 3 and 4 have estimates for an indicator if the student enroll in a non-selective program. Columns 2 and 4 control for standardized test scores, GPA, family income, region, type of high school and gender.

Table E.II: Diff-in-diff estimates for graduation at admission program with school-year cluster standard errors

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad on time
P-Up x after	0.042*** (0.008)	0.072*** (0.009)	0.084*** (0.010)	0.043*** (0.011)
P-Down x after	-0.033*** (0.008)	-0.060*** (0.009)	-0.082*** (0.010)	-0.039*** (0.010)
Obs.	234,544	234,544	234,544	234,544
Controls	✓	✓	✓	✓
Test	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$	$\beta_3 = -\beta_4$
p-value	0.412	0.376	0.918	0.751

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table E.III: Diff-in-diff estimates for college graduation (any program) with school-year cluster standard errors

	(1) Grad by 6yr	(2) Grad by 7yr	(3) Grad by 8yr	(4) Grad or enroll by 8 yr
P-Up x after	-0.003 (0.010)	0.005 (0.011)	-0.008 (0.011)	0.006 (0.010)
P-Down x after	-0.012 (0.009)	-0.015 (0.010)	-0.032*** (0.011)	-0.006 (0.010)
Obs.	234,544	234,544	234,544	234,544
Controls	✓	✓	✓	✓

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

Table E.IV: Diff-in-diff estimates for graduation with school-year cluster standard errors

	(1)	(2)	(3)	(4)	(5)	(6)
	Grad by 6yr	Grad by 7yr	Grad by 8yr	Grad by 6yr	Grad by 7yr	Grad by 8yr
P-Up x after	0.019** (0.009)	0.030*** (0.011)	0.019* (0.011)	-0.022*** (0.004)	-0.025*** (0.005)	-0.026*** (0.005)
P-Down x after	-0.022*** (0.008)	-0.028*** (0.010)	-0.046*** (0.010)	0.009** (0.004)	0.013*** (0.004)	0.014*** (0.005)
Obs.	234,544	234,544	234,544	234,544	234,544	234,544
Controls	✓ Selective	✓ Selective	✓ Selective	✓ Non-Selective	Non-Selective	✓ Non-Selective

Robust standard errors in parentheses

Notes: Columns 1-3 show diff-in-diff estimates for an indicator if the student graduate from the admission program by 6, 7 and 8 years after application. Column 4 show the results for the outcome of graduation on time.

F Regression Discontinuity Results

Table F.I: RD estimates on enrollment for crossing threshold for 1st preference

	(1) Enroll	(2) Enroll	(3) Enroll	(4) Enroll	(5) Enroll
RD estimator	0.170*** (0.009)	0.185*** (0.008)	0.184*** (0.007)	0.178*** (0.005)	0.136*** (0.012)
Obs.	90,205	90,205	90,205	90,205	15,623
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2013 sample of students at program with excess of demand. The outcome variable indicates if the students enrolled at the admission program. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.II: RD estimates on peers' selectivity for crossing threshold for 1st preference

	(1) Selectivity	(2) Selectivity	(3) Selectivity	(4) Selectivity	(5) Selectivity
RD estimator	0.197*** (0.005)	0.211*** (0.005)	0.217*** (0.004)	0.217*** (0.003)	0.166*** (0.007)
Obs.	84,770	84,770	84,770	84,770	15,011
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2013 sample of students at program with excess of demand. The outcome variable correspond to the average test score of the students enrolled at the same program. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.III: RD estimates on enrollment for crossing threshold for 1st preference

	(1) Grad by 8yr	(2) Grad by 8yr	(3) Grad by 8yr	(4) Grad by 8yr	(5) Grad by 8yr
RD estimator	0.065*** (0.009)	0.068*** (0.008)	0.067*** (0.007)	0.063*** (0.006)	0.071*** (0.014)
Obs.	90,205	90,205	90,205	90,205	15,623
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2013 sample of students at program with excess of demand. The outcome variable indicates if the students graduate from the admission program 8 years after the admission process. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.IV: RD estimates on enrollment for crossing threshold for 1st preference

	(1)	(2)	(3)	(4)	(5)
	Grad by 8yr	Grad by 8yr	Grad by 8yr	Grad by 8yr	Grad by 8yr
RD estimator	-0.006 (0.010)	-0.006 (0.009)	-0.012 (0.008)	-0.021*** (0.006)	0.002 (0.015)
Obs.	90,205	90,205	90,205	90,205	15,623
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2013 sample of students at program with excess of demand. The outcome variable indicates if the students graduate from any program 8 years after the admission process. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.V: RD estimate on enrollment for crossing threshold for 1st preference for boost students

	(1)	(2)	(3)	(4)	(5)
	enroll	enroll	enroll	enroll	enroll
RD estimator	0.199*** (0.012)	0.199*** (0.011)	0.195*** (0.009)	0.182*** (0.007)	0.193*** (0.013)
Obs.	47,584	47,584	47,584	47,584	15,490
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2012 sample of students at program with excess of demand and a boost score larger than 5. The outcome variable indicates if the students graduate from any program 8 years after the admission process. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.VI: RD estimate on graduation from the admission offer for crossing threshold for 1st preference for boost students

	(1)	(2)	(3)	(4)	(5)
	Grad by 8yr	Grad by 8yr	Grad by 8yr	Grad by 8yr	Grad by 8yr
RD estimator	0.096*** (0.014)	0.095*** (0.012)	0.099*** (0.010)	0.082*** (0.008)	0.095*** (0.015)
Obs.	47,584	47,584	47,584	47,584	15,490
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2012 sample of students at program with excess of demand and a boost score larger than 5. The outcome variable indicates if the students graduate from any program 8 years after the admission process. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.VII: RD estimate on college graduation for crossing threshold for 1st preference for boost students

	(1)	(2)	(3)	(4)	(5)
	Grad in 8yr	Grad in 8yr	Grad in 8yr	Grad in 8yr	Grad in 8yr
RD estimator	0.001 (0.014)	-0.011 (0.013)	-0.011 (0.011)	-0.027*** (0.009)	-0.001 (0.015)
Obs.	47,584	47,584	47,584	47,584	15,490
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: 2012 sample of students at program with excess of demand and a boost score larger than 5. The outcome variable indicates if the students graduate from any program 8 years after the admission process. The running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20).

Table F.VIII: Diff-in-diff estimates on enrollment at admission offer

	(1) Enroll	(2) Enroll	(3) Enroll	(4) Enroll
P-Up x after	0.189*** (0.017)	0.094*** (0.028)	0.276*** (0.030)	0.542*** (0.036)
P-Down x after	-0.113*** (0.015)	-0.088*** (0.023)	-0.117*** (0.026)	-0.498*** (0.039)
Obs.	137,400	61,188	116,834	140,033
Controls	✓	✓	✓	✓
	Moved 1-2	Moved 2-3	Moved 1-3	Moved 1-0

Robust standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

Notes: diff-in-diff estimates using the sample of students for whom the admissions with and without the inclusion of the GPA⁺ measurement moves them between the respective margins.

Table F.IX: Diff-in-diff estimates of peers performance at admission offer

	(1) Selectivity	(2) Selectivity	(3) Selectivity	(4) Selectivity
P-Up x after	0.250*** (0.009)	0.231*** (0.013)	0.254*** (0.015)	0.443*** (0.011)
P-Down x after	-0.256*** (0.010)	-0.221*** (0.012)	-0.216*** (0.016)	-0.128*** (0.013)
Obs.	137,341	61,163	116,783	139,990
Controls	✓	✓	✓	✓
	Moved 1-2	Moved 2-3	Moved 1-3	Moved 1-0

Robust standard errors in parentheses

Notes: diff-in-diff estimates using the sample of students for whom the admissions with and without the inclusion of the GPA⁺ measurement moves them between the respective margins. Selectivity is measured as the average test score of the students at the admission program; for students without any admission selectivity is measure as the average test score of students without any admission.

Table F.X: Diff-in-diff estimates for graduation from admission offer

	(1) Grad by 8yr	(2) Grad by 8yr	(3) Grad by 8yr	(4) Grad by 8yr
P-Up x after	0.077*** (0.019)	0.059** (0.029)	0.160*** (0.033)	0.181*** (0.029)
P-Down x after	-0.073*** (0.020)	-0.042 (0.027)	-0.089** (0.035)	-0.240*** (0.030)
Obs.	137,400	61,188	116,834	140,033
Controls	✓ Moved 1-2	✓ Moved 2-3	✓ Moved 1-3	✓ Moved 1-0

Robust standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

Notes: diff-in-diff estimates using the sample of students for whom the admissions with and without the inclusion of the GPA⁺ measurement moves them between the respective margins.

Table F.XI: Diff-in-diff estimates for college graduation

	(1) Grad by 8yr	(2) Grad by 8yr	(3) Grad by 8yr	(4) Grad by 8yr
Pulled-Up x after	-0.028 (0.020)	0.028 (0.031)	0.033 (0.036)	0.010 (0.038)
Pushed-Down x after	-0.019 (0.021)	-0.005 (0.029)	-0.024 (0.037)	-0.111*** (0.037)
Obs.	137,400	61,188	116,834	140,033
Controls	✓ Moved 1-2	✓ Moved 2-3	✓ Moved 1-3	✓ Moved 1-0

Robust standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

Notes: diff-in-diff estimates using the sample of students for whom the admissions with and without the inclusion of the GPA⁺ measurement moves them between the respective margins.

Table F.XII: RD estimates on enrollment for crossing threshold for 1st preference for pulled-up students

	(1) Enroll	(2) Enroll	(3) Enroll	(4) Enroll	(5) Enroll
RD estimator	0.137*** (0.027)	0.141*** (0.023)	0.122*** (0.020)	0.117*** (0.017)	0.156*** (0.024)
Obs.	18,375	18,375	18,375	18,375	4,588
Controls	✓	✓	✓	✓	✓
Poly	5	4	3	2	1

Robust standard errors in parentheses

Notes: Sample of analysis only consider students with admission at their 1st or 2nd preference (margin of treatment 1-2) and with the admission simulated without the reform at 2nd preference (therefore threshold crossing mostly due to the boost) in 2013. The sample restriction attempt to capture the effect of threshold crossing for pulled-up students by comparing them with very similar controls (non-crossing but similar score). Running variable is the application score in the most preferred choice minus the cutoff score (ex-post defined by the application of the last person admitted at that program based on the vacancies) at that program. Columns 1-4 use the entire sample. Column 5 estimate the effect with a bandwidth of 20 point (optimal bandwidth range between 10 and 20). Specification allow for different slope in each side of the cutoff.