

Toward a Unifying Information-Theoretic Framework for Re-identification Risk Quantification

Tayler Blake

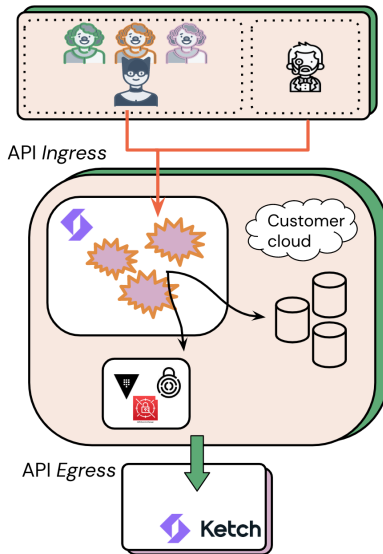
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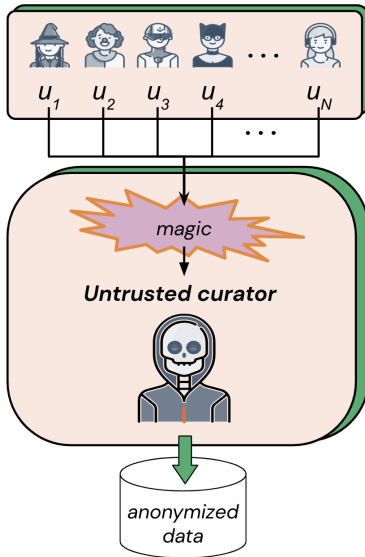
December 13, 2022

Introduction

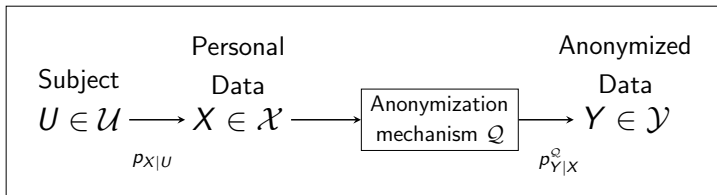
Re-identification probabilities: elusive requisites



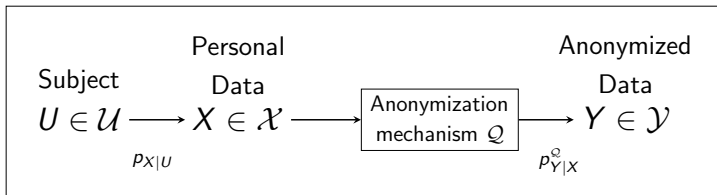
Re-identification probabilities: elusive requisites



The anonymization framework



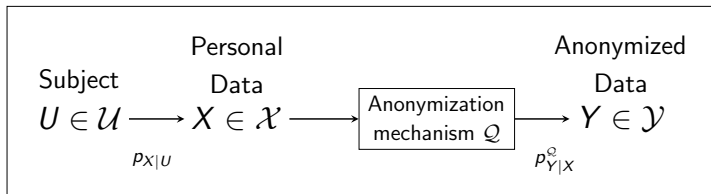
The anonymization framework



$$\mathcal{U} = \{u_1, \dots, u_N\}$$

The anonymization framework

$$\mathcal{X} = \{x_1, \dots, x_N\}$$



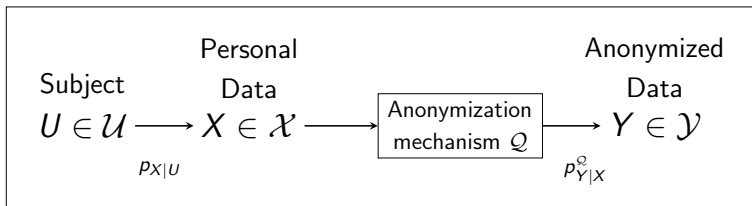
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The anonymization framework

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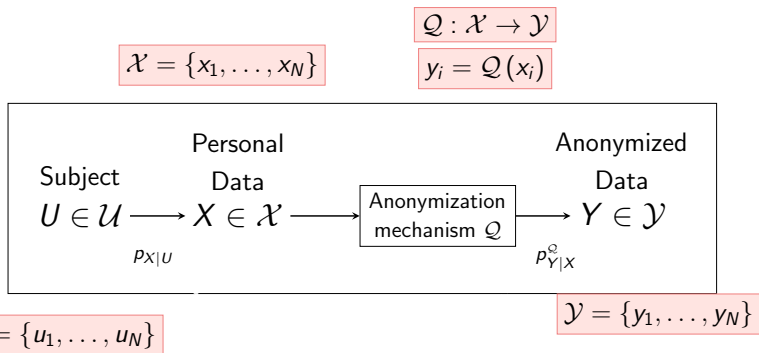
$$Q: \mathcal{X} \rightarrow \mathcal{Y}$$

$$y_i = Q(x_i)$$

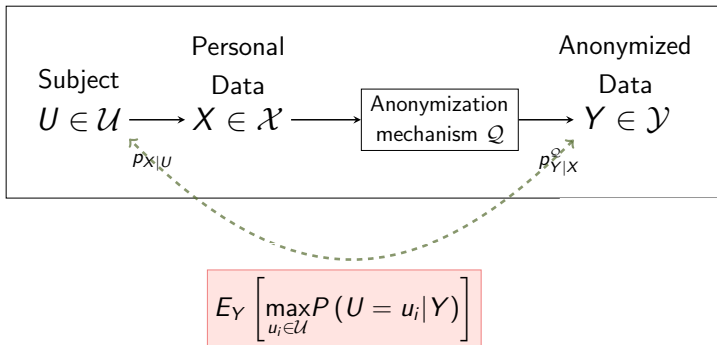


$$\mathcal{U} = \{u_1, \dots, u_N\}$$

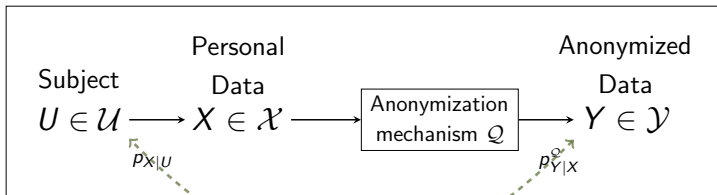
The anonymization framework



Evaluating re-identification risk



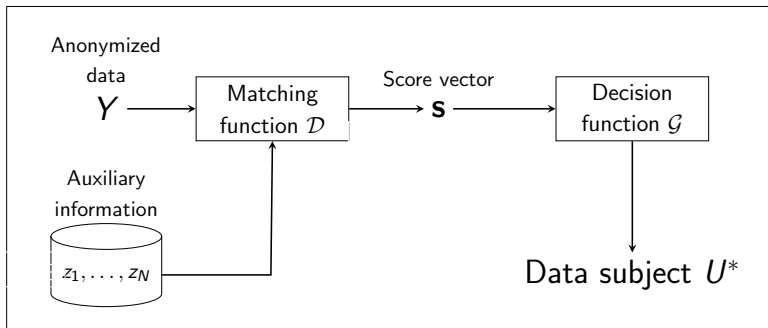
Evaluating re-identification risk



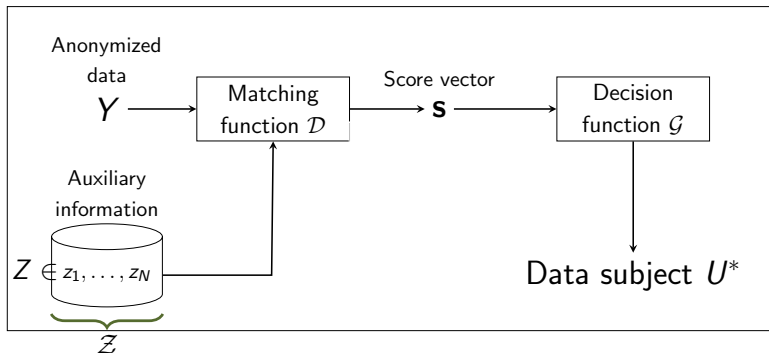
$$\text{PIE} = I(U, Y) = \sum_u \sum_y p_{U,Y}(u, y) \log \frac{p_{U,Y}(u, y)}{p_U(u) p_Y(y)}$$

A general adversarial re-identification framework

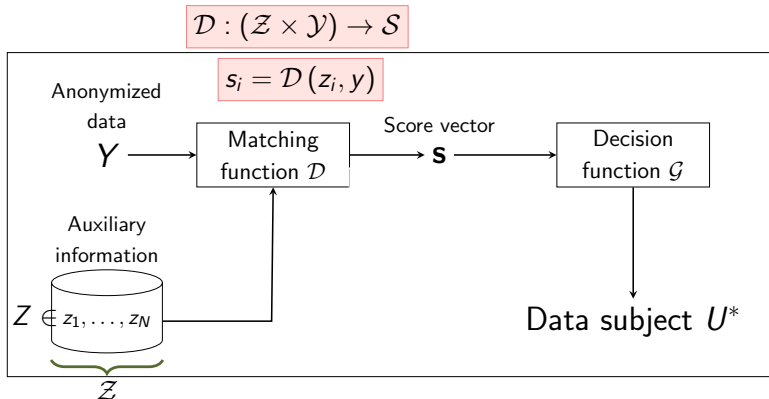
Re-identification framework



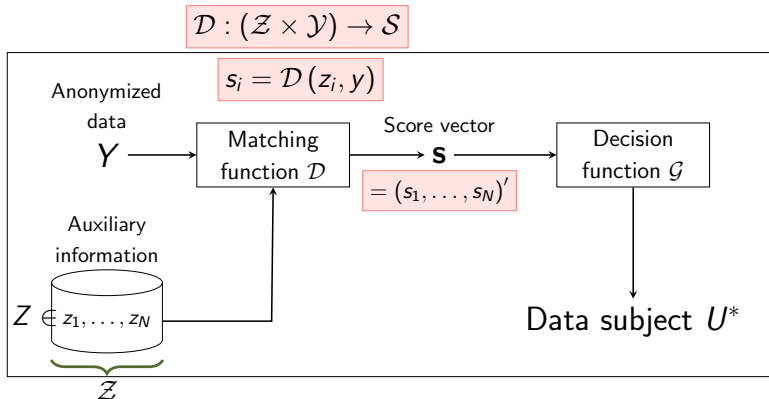
Re-identification framework



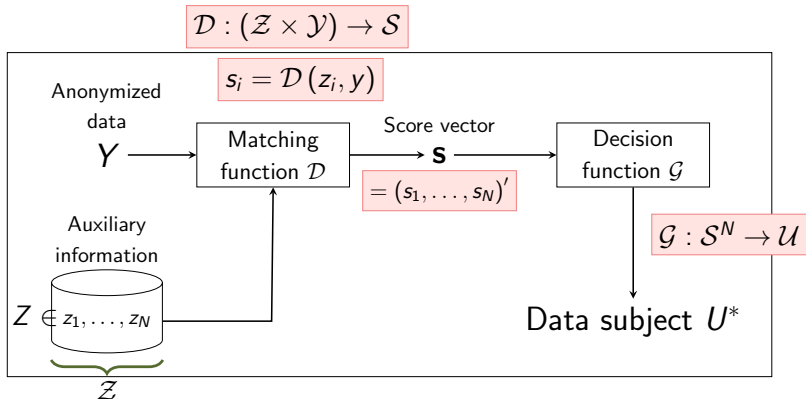
Re-identification framework



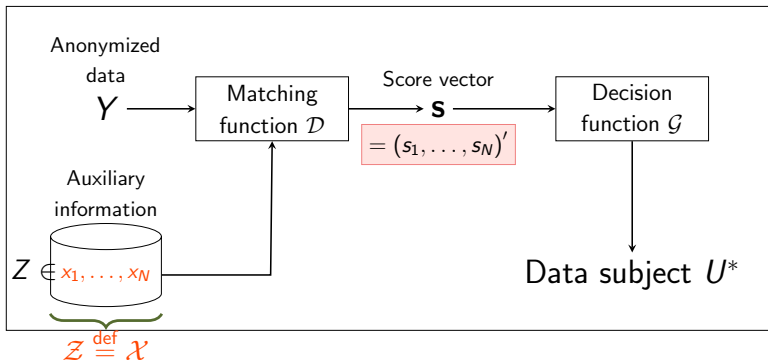
Re-identification framework



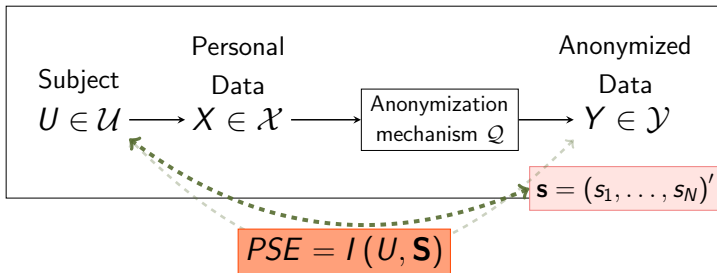
Re-identification framework



The maximum-knowledge intruder



Evaluating re-identification risk



Bounding posterior re-identification probabilities

Personal System Entropy (PSE)

The *personal system entropy* $PSE = I(U, \mathbf{S})$ is a lower bound for the PIE.

Proposition

For any matching function \mathcal{D} and any auxiliary information z_1, \dots, z_N ,

$$I(U, \mathbf{S}) \leq I(U, Y)$$

with equality if and only if U , \mathbf{S} , and Y form the Markov chain $U \rightarrow \mathbf{S} \rightarrow Y$ (i.e. \mathbf{S} is a sufficient statistic for U).

e.g. if $\mathbf{s} = (q_1(y), \dots, q_N(y))'$ where $q_i(y) = P(Y = y | U = u_i)$, then $PIE = PSE$.

Bayes re-identification error rate

Proposition (Bayes error, PIE and PSE)

$$\beta_{U|\mathbf{S}} \geq 1 - \frac{I(U, \mathbf{S}) + 1}{\log(1 - \beta_U)} \geq 1 - \frac{I(U, Y) + 1}{\log(1 - \beta_U)}.$$

where

$$\beta_U = 1 - \max_{u_i \in \mathcal{U}} P(U = u_i) = \text{Bayes error probability before observing } \mathbf{S}$$

$$\beta_{U|\mathbf{S}} = 1 - E_{\mathbf{S}} \left[\max_{u_i \in \mathcal{U}} P(U = u_i | \mathbf{S}) \right] = \text{Bayes error probability after observing } \mathbf{S}$$

Bayes re-identification error rate

Then!

$$1 - \beta_{U|S} \leq \underbrace{\frac{I(U, \mathbf{S}) + 1}{\log(1 - \beta_U)}}_{\substack{\text{we can estimate this} \\ \text{using the adorable} \\ \text{result on the next slide}}} \leq \frac{I(U, Y) + 1}{\log(1 - \beta_U)}$$

When p_U is Uniform, this becomes

$$1 - \beta_{U|S} \leq \frac{I(U, \mathbf{S}) + 1}{\log N} \leq \frac{I(U, Y) + 1}{\log N}$$

Estimation of the PSE

Let $\mathcal{D}(X_i, Y_i) \sim \underset{\substack{\text{genuine} \\ \text{score} \\ \text{distribution}}}{f_G}$, $\mathcal{D}(X_i, Y_j) \sim \underset{\substack{\text{imposter} \\ \text{score} \\ \text{distribution}}}{f_I}$ for $i \neq j$.

Theorem (1)

Let f_G be a one-dimensional genuine score distribution and f_I be a one-dimensional imposter score distribution. Then

$$I(U, \mathbf{S}) \rightarrow D(f_G || f_I) \text{ as } N \rightarrow \infty,$$

where $D(f_G || f_I)$ denotes the Kullback-Leibler divergence of f_G from f_I .

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