## - Set Up

```
{\tt from} \ \_{\tt future} \_ \ {\tt import} \ {\tt division}
import pandas as pd
import numpy as np
import seaborn as sns
import matplotlib.pyplot as plt
%matplotlib inline
import statsmodels.api as sm
import statsmodels.formula.api as smf
from statsmodels.compat import lzip
import scipy as sp
# upload data
from google.colab import files
uploaded= files.upload()
import io
d = pd.read_csv(io.BytesIO(uploaded['GSS.2006 (1).csv']))
     Choose Files GSS.2006 (1).csv
     • GSS.2006 (1).csv(text/csv) - 8232424 bytes, last modified: 2/17/2023 - 100% done
     Saving GSS.2006 (1).csv to GSS.2006 (1) (1).csv
```

### Variables

For my dependent variable I will explore mental health using the mntlhlth variable. This variable comes from the following question: Now thinking about your mental health, which includes stress, depression, and problems with emotions, for how many days during the past 30 days was your mental health not good?

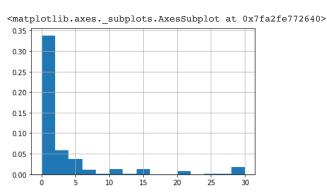
I will also note I am only using data from 2006.

As this is a count variable, a Poisson regression is the appropriate regression type to use.

### d.mntlhlth.describe()

count	1717.0	00000			
mean	3.0	00000			
std	6.6	31656			
min	0.0	00000			
25%	0.0	00000			
50%	0.0	00000			
75%	3.0	00000			
max	30.000000				
Name:	mntlhlth,	dtype:	float64		

#### d.mntlhlth.hist(bins=15, density=True)



```
d.mntlhlth.var()
```

44.64452214452214

The mean (3) is not equal to the variance (45), which suggess that a normal Poisson regression is not ideal, a negative binomial is more appropriate. Additionally, from the histogram, it appears that there is a large amount of "zero" responses, which suggests that a zero-inflated negative binomial would be most appropriate to use.

```
variables = ['mntlhlth', 'realinc','socfrend', 'sex']
#recode to make more interpretable
d = d[variables].copy()
d['socfrend'] = 7 - d.socfrend
d['inc10k'] = d.realinc/10000
d_noNA = d.dropna()
```

The independent variables I will explore are income and frequency of time spent with friends. I am mainly interested in how frequency of time spent with friends impacts mental health, particularly in light of the pandemic where socializing was drastically decreased and we saw a lot of mental health issues increase. As such, I hypothesize that the more frequently a person spends time with friends, the fewer days of poor mental health they will have.

I will include real income as a control variable as the effects of income on mental health are well known (the more money you have the less financial stress you have to endure, generally allowing for better mental health). I suspect that the more money a person makes the fewer poor mental health days they would experience.

I will lastly include sex of the respondent as another variable to act as a control, as there may be sex differences in mental health. Particularly women are more likely to be depressed and thus may have more poor mental health days.

The friend variable I am using is socfrend which comes from the following question:

Would you use this card and tell me which answer comes closest to how often you do the following things... C. Spend a social evening with friends who live outside the neighborhood?

Which had the following possible responses. -100 .i: Inapplicable

-99 .n: No answer

-98 .d: Do not Know/Cannot Choose

-97 .s: Skipped on Web

1 ALMOST DAILY

2 SEV TIMES A WEEK

3 SEV TIMES A MNTH

4 ONCE A MONTH

5 SEV TIMES A YEAR

6 ONCE A YEAR

7 NEVER

To make this variable easier to interpret, I recoded it to be 1 as being never and 7 as being almost daily so an increase in number corresponds with the increase of time spent with friends.

I will also exclude the various non-answers (-97 to -100) in my models by subsetting the data.

Additionally, to make income more interpretable I recoded it to be inc10k so it is for every increase in 10 thousand dollars instead of increase in 1 dollar.

### - OLS

```
from statsmodels.regression.linear_model import OLS
m1 = OLS.from_formula('mntlhlth ~ inc10k + socfrend + C(sex)', data =d, subset=d.socfrend>0).fit()
m1.summary()
```

```
OLS Regression Results
       Dep. Variable: mntlhlth
                                                        0.022
                                          R-squared:
           Model:
                       OLS
                                        Adj. R-squared: 0.019
          Method:
                       Least Squares
                                          F-statistic:
                                                        7.424
           Date:
                       Wed, 22 Feb 2023 Prob (F-statistic): 6.41e-05
           Time:
                       20:44:21
                                        Log-Likelihood: -3300.6
      No. Observations: 1007
                                                        6609
                                             AIC:
                                             BIC:
        Df Residuals: 1003
                                                        6629
         Df Model:
                       3
      Covariance Type: nonrobust
                 coef std err t P>Itl [0.025 0.975]
      Intercept 2.6786 0.620 4.318 0.000 1.461 3.896
      C(sex)[T.2] 0.4604 0.408 1.128 0.260 -0.340 1.261
       inc10k -0.2488 0.062 -3.982 0.000 -0.371 -0.126
      socfrend 0.2672 0.131 2.033 0.042 0.009 0.525
ml.predict(d[['inc10k', 'socfrend', 'sex']]).describe()
               1721.000000
     count
                   3,155762
     mean
     std
                   0.979095
     min
                  -0.649254
     25%
                   2.718843
     50%
                   3.284597
     75%
                   3.771407
                   5.002692
     max
     dtype: float64
```

As expected, the OLS is not appropriate as it gave negative predicted values, which do not make sense given the variable mntlhealth is a count that can only go from 0-30. Additionally, it only predicts a maximum of 5 days of poor mental health where there was a large skew above 5 days in the dataset, indicating that there could be a model that would give more accurate predictions.

The value of r and adjusted r-squared are also very low at 0.019 and 0.017 respectively, which suggests it may not be a very good predictive model

#### → Poisson

```
m2 = smf.poisson('mntlhlth ~ inc10k + socfrend + C(sex)', data =d, subset=d.socfrend>0).fit()
m2.summary()
     Optimization terminated successfully.
               Current function value: 4.783848
               Iterations 6
                      Poisson Regression Results
       Dep. Variable: mntlhlth
                                     No. Observations: 1007
          Model:
                     Poisson
                                       Df Residuals: 1003
         Method:
                     MLE
                                         Df Model:
                                                      3
           Date:
                     Wed, 22 Feb 2023 Pseudo R-squ.: 0.03653
           Time:
                     20:44:59
                                      Log-Likelihood: -4817.3
        converged:
                     True
                                          LL-Null:
                                                      -5000.0
      Covariance Type: nonrobust
                                       LLR p-value:
                                                      7.254e-79
                 coef std err z P>|z| [0.025 0.975]
      Intercept 1.0800 0.058 18.561 0.000 0.966 1.194
     C(sex)[T.2] 0.1502 0.037 4.034 0.000 0.077 0.223
       inc10k -0.1190 0.008 -14.843 0.000 -0.135 -0.103
      socfrend 0.0836 0.011 7.391 0.000 0.061 0.106
```

According to this model, with every increase in frequency of time spent with friends, is a 0.08 higher expected log-count of days with poor mental health, net of other factors. This is a significant increase, but is not what I would have expected.

For every 10,000 dollar increase in income, there is a 0.1190 lower expected log-count of days spent with poor mental health. This is also significant and aligns with my previous hypothesis.

Females compared to males have a 0.1502 higher expected log-count of days with more mental health, net all other factors. This aligns with my earlier hypothesis and is a significant increase.

However, this does not address the overdispersion of the dependent variable. As such, a negative binomial should be run instead.

### Negative Binomial

```
from statsmodels.discrete.discrete_model import NegativeBinomial
nb1 = NegativeBinomial.from_formula('mntlhlth ~ inc10k + socfrend + C(sex)', data =d, subset=d.socfrend>0).fit()
nb1.summary()
    Optimization terminated successfully.
             Current function value: 1.760224
             Iterations: 14
             Function evaluations: 21
             Gradient evaluations: 21
                NegativeBinomial Regression Results
      Dep. Variable: mntlhlth
                                 No. Observations: 1007
         Model: NegativeBinomial Df Residuals: 1003
        Method: MLE
                                    Df Model:
                   Date:
         Time:
                   20:45:07 Log-Likelihood: -1772.5
       converged: True
                                     LL-Null:
                                                -1781.3
                                   LLR p-value: 0.0005528
     Covariance Type: nonrobust
               coef std err z P>|z| [0.025 0.975]
     Intercept 0.9747 0.237 4.109 0.000 0.510 1.440
     C(sex)[T.2] 0.1486 0.162 0.919 0.358 -0.168 0.465
      inc10k -0.0875 0.022 -3.937 0.000 -0.131 -0.044
     socfrend 0.0836 0.052 1.621 0.105 -0.018 0.185
       alpha 6.0533 0.406 14.926 0.000 5.258 6.848
```

As the alpha parameter is statistically significant, this suggest that the Poisson was not the most appropriate model, and we are correct in using the negative binomial model instead.

According to this model, with every increase in frequency of time spent with friends, is a 0.0836 higher expected log-count of days with poor mental health, net of other factors. This is no longer a significant increase however.

For every 10,000 dollar increase in income, there is a 0.0875 lower expected log-count of days spent with poor mental health. This is significant and aligns with my previous hypothesis.

Females compared to males have a 0.1486 higher expected log-count of days with more mental health, net all other factors, though this is not a significant difference.

Both sex and time spent with friends were not significant, which is counter to what I previously assumed.

```
np.exp(nb1.params)
```

```
Intercept 2.688801
C(sex)[T.2] 1.163486
inc10k 0.909371
socfrend 1.089861
dtype: float64
```

Another way to interpret this model is to look at the incident rate ratios above.

Females have an incident rate ratio (for poor mental health days) that is approximately 1.2 times higher, net of frequency of time spent with friends and income.

With every increase in 10,000 dollars of income, the incident rate ratio is approximately 0.9 times higher, net of all other factors.

With every marginal increase in frequency of time spent with friends, the incident rate ratio is approximately 1.1 times higher, net of all other factors.

# Zero-inflated Negative Binomial

```
znb1 = nb1 = sm.GLM.from_formula('mntlhlth ~ inc10k + socfrend + C(sex)', data =d, subset=d.socfrend>0, family=sm.families.Negaticznb1.summary()
```

```
Generalized Linear Model Regression Results
 Dep. Variable: mntlhlth
                                 No. Observations: 1007
                GLM
                                   Df Residuals:
                                                  1003
    Model:
 Model Family: NegativeBinomial
                                     Df Model:
                                                  3
 Link Function: log
                                      Scale:
                                                  1.0000
    Method:
                IRLS
                                  Log-Likelihood: -2204.6
     Date:
                Wed, 22 Feb 2023
                                   Deviance:
                                                  2406.7
     Time:
                22:05:51
                                   Pearson chi2: 3.61e+03
 No. Iterations: 9
Covariance Type: nonrobust
           coef std err z P>Izl [0.025 0.975]
Intercept 0.9891 0.114 8.685 0.000 0.766 1.212
C(sex)[T.2] 0.1514 0.074 2.039 0.041 0.006 0.297
```

According to this model, with every increase in frequency of time spent with friends, is a 0.0860 higher expected log-count of days with poor mental health, net of other factors. This is now significant again in this model, which aligns with my hypothesis.

For every 10,000 dollar increase in income, there is a 0.0950 lower expected log-count of days spent with poor mental health. This is significant and aligns with my previous hypothesis.

Females compared to males have a 0.1514 higher expected log-count of days with more mental health, net all other factors. This is also significant and aligns with my previous assumptions.

# Comparing Models

	OLS	Poisson	Negative binomial	Zero-inflated NB
count	1007.000000	1007.000000	1007.000000	1007.000000
mean	3.972498	3.200709	3.944721	3.944721
std	5.040007	5.915379	5.044245	5.044245
min	0.015604	0.024472	0.028221	0.028221
25%	1.886836	0.874289	1.881954	1.881954
50%	2.843877	1.137268	2.729740	2.729740
75%	3.571564	1.722337	3.490840	3.490840
max	27.423365	29.111711	27.495331	27.495331

The average absolute deviations are lowest for Poisson, and much more behaved at the 3rd quartile of deviations, which again suggests that Poisson is the right route to go. However, as discussed above the alpha value suggests that the negative binomial model is the better model considering the over-dispersed nature of the dependent variable.

Additionally, since the negative binomial and the zero-inflated negative binomial have the same absolute deviations, this suggest that the zero-inflated negative binomial is not giving any additional information from the negative binomial. This suggests that the negative binomial is the best predictive model.

### Predicted Values

Females (3.2) have more poor mental health days than males (2.8), with all else at means. This is as expected.

For males with income set at the mean, those who spend the most time with friends (3.6)have more poor mental health days than those who spend the least amount of time with friends (2.1). This is a difference of 1.5 days. As you increase in time spent with friends, all the predicted values also increase, which is contrary to my hypothesis, but it is not a significant increase according to the model.

For females with income set at the mean, those who spend the most time with friends (4.2)have more poor mental health days than those who spend the least amount of time with friends (2.5). This is a difference of around 2 days. As you increase in time spent with friends, all the predicted values also increase, which is contrary to my hypothesis, but it is not a significant increase according to the model.

#### Conclusion

Due to the over-dispersed nature of the dependent variable, confirmed by a significant alpha value of the negative binomial model, the negative binomial model is the best model to use as the zero-negative binomial model did not have a lower absolute deviation.

Since the negative binomial model is the most appropriate, the coefficients from that model are the most trustworthy. In that model time spent with friends was not a significant predictor of number of poor mental health days. This is surprising as I thought that more time with friends would improve your mental health, resulting in fewer poor mental health days. This is not as big of a factor as I thought it would be when controlling for income and sex.

As predicted, higher income was significantly associated with fewer poor mental health days, net of all factors. Females also had more poor mental health days than males, as expected.

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