

STAT 7200

Introduction to Advanced Probability

Lecture 19

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- Weak Convergence

Equivalent Definitions of Weakly Convergence

Theorem 1 (Equivalent Definitions of Weakly Convergence)

The following statements are all equivalent definition of weak convergence:

(1) $\{\mu_n\}$ converges weakly to μ . (Original definition)

(2) $\mu_n(A) \rightarrow \mu(A)$ for all measurable set A such that $\mu(\partial A) = 0$. (∂A is defined as the boundary of set A)

(3) $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$ for all $x \in \mathbb{R}$ such that $\mu(\{x\}) = 0$. That is, the convergence of CDFs. (Note, $\{x\}$ is the boundary of set $(-\infty, x]$.)

(4) (Skorohod's Theorem) there are random variable Y, Y_1, Y_2, \dots defined on the same probability triple, with $\mathcal{L}(Y) = \mu$ and $\mathcal{L}(Y_n) = \mu_n$ such that $Y_n \rightarrow Y$ with probability 1 (This theorem connects the strongest type of convergence: convergence almost surely, with the weak convergence.)

(5) $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$ for all bounded Borel-measurable functions $f : \mathbb{R} \rightarrow \mathbb{R}$. such that $\mu(D_f) = 0$, where D_f is the set of discontinuous points of f . (The continuous condition of definition 1) is relaxed.)

Structure of Proof

- Our proof will follow the following structure:
- We have proved: $(5) \Rightarrow (1)$, $(5) \Rightarrow (2)$ and $(2) \Rightarrow (3)$

Proof: (1) \Rightarrow (3)

- (1) $\{\mu_n\}$ converges weakly to μ : $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$ for all bounded continuous functions f .

(3) $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$ for all $x \in \mathbb{R}$ such that $\mu(\{x\}) = 0$.

- **Strategy:** We can not apply (1) directly by setting $f = 1_{(-\infty, x]}$ since $1_{(-\infty, x]}$, although bounded, is discontinuous at x . We may resolve this issue by constructing continuous approximation of $1_{(-\infty, x]}$.
- **Proof:** For any $\varepsilon > 0$ (which is used to control how good the approximation is), define $f(t) = 1$ for $t \leq x$ and 0 for $t \geq x + \varepsilon$, but let $f(t)$ be a linear function on $(x, x + \varepsilon)$.
 - As f is now continuous and $1_{(-\infty, x]} \leq f \leq 1_{(-\infty, x + \varepsilon]}$:

$$\limsup_n \mu_n((-\infty, x]) \leq \limsup_n \int f d\mu_n = \int f d\mu \leq \mu((-\infty, x + \varepsilon])$$

- Let $\varepsilon \rightarrow 0$. By the continuity of probability, we have $\limsup_n \mu_n((-\infty, x]) \leq \mu((-\infty, x])$

Proof: (1) \Rightarrow (3): continued

- **Proof: continued** Similarly, define $f(t) = 1$ for $t \leq x - \varepsilon$ and 0 for $t \geq x$, but let $f(t)$ be a linear function on $(x - \varepsilon, x)$. Then f is linear and $1_{(-\infty, x - \varepsilon]} \leq f \leq 1_{(-\infty, x]}$. And:

$$\liminf_n \mu_n((-\infty, x]) \geq \liminf_n \int f d\mu_n = \int f d\mu \geq \mu((-\infty, x - \varepsilon])$$

- Let $\varepsilon \rightarrow 0$, $\liminf_n \mu_n((-\infty, x]) \geq \mu((-\infty, x)) = \mu((-\infty, x])$. The last equality holds since $\mu(\{x\}) = 0$.
- In summary:

$$\liminf_n \mu_n((-\infty, x]) \geq \mu((-\infty, x]) \geq \limsup_n \mu_n((-\infty, x])$$

- we then must have:

$$\lim_n \mu_n((-\infty, x]) = \mu((-\infty, x])$$

Proof: (4) \Rightarrow (5)

- (4) there are random variable Y, Y_1, Y_2, \dots defined on the same probability triple, with $\mathcal{L}(Y) = \mu$ and $\mathcal{L}(Y_n) = \mu_n$ such that $Y_n \rightarrow Y$ with probability 1.

(5) $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$ for all bounded Borel-measurable functions $f : \mathbb{R} \rightarrow \mathbb{R}$. such that $\mu(D_f) = 0$, where D_f is the set of discontinuous points of f .

- **Proof:** Pick an appropriate f . First, we want to show that $P(f(Y_n) \rightarrow f(Y)) = 1$. Note that
 - ▶ $0 \leq P(Y_n(\omega) \rightarrow Y(\omega), D_f) \leq P(D_f) = 0$
 - ▶ $1 = P(Y_n \rightarrow Y) = P(Y_n \rightarrow Y, D_f) + P(Y_n \rightarrow Y, D_f^c) = P(Y_n \rightarrow Y, D_f^c)$
 - ▶ $\{\omega : f(Y_n) \rightarrow f(Y)\} \supseteq \{\omega : Y_n(\omega) \rightarrow Y(\omega)\} \cap \{\omega : Y(\omega) \in D_f^c\}$

so $f(Y_n) \rightarrow f(Y)$ wp1 by (4) and monotonicity of P .

- Because f is bounded, $f(Y)$ is integrable, so $E[f(Y_n)] \rightarrow E[f(Y)]$ by the dominated convergence theorem.

Proof: (3) \Rightarrow (4)

- (3) $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$ for all $x \in \mathbb{R}$ such that $\mu(\{x\}) = 0$.
(4) there are random variables Y, Y_1, Y_2, \dots defined on the same probability triple, with $\mathcal{L}(Y) = \mu$ and $\mathcal{L}(Y_n) = \mu_n$ such that $Y_n \rightarrow Y$ with probability 1
- **Strategy:** We will construct random variables with CDFs $F_n(x) = \mu_n((-\infty, x])$, $F(x) = \mu((-\infty, x])$, then we will show the convergence of these random variables using the fact that the corresponding CDFs converge.

Proof: (3) \Rightarrow (4): Probability Integral Transform

We can construct random variable with given CDF using probability integral transform theorem.

Useful facts:

- $F(q) \geq p \iff Q(p) \leq q$
- $F(q) < p \Rightarrow Q(p) > q$
- When the CDF is continuous and strictly increasing, the quantile function is the inverse of CDF.
- The quantile function $Q(p)$ is a non-decreasing function, same as the CDF.

Proof: (3) \Rightarrow (4): Probability Integral Transform

- For random variable U that follows uniform distribution, given any CDF $F(x)$, define quantile function $Q(p) = \inf\{x : F(x) \geq p\}$, then the random variable $Q(U)$ follows distribution with CDF $F(x)$.
- **Proof:**

$$P[Q(U) \leq x] = P[F(x) \geq U] = F(x)$$

Proof: (3) \Rightarrow (4): continued

- 1 Another fact we'll need, any nondecreasing function f has at most a countable number of discontinuities.
- 2 Call the set of discontinuities D_f . Suppose for starts that $f : [a, b] \rightarrow \mathbb{R}$ (e.g. a quantile function).
$$D_f = \bigcup_n \left\{ \frac{1}{n} \leq f(x^+) - f(x^-) \right\}.$$
- 3 $|\left\{ \frac{1}{n} \leq f(x^+) - f(x^-) \right\}| \leq n[f(b) - f(a)]$
- 4 Countable unions of finite sets are countable!

Proof: (3) \Rightarrow (4): continued

- 1 Remove the assumption of compact domain...suppose that $f : \mathbb{R} \rightarrow \mathbb{R}$ (e.g. a cdf).
- 2 Call the set of discontinuities D_f again. For any integer z , let D_{zf} be the set of discontinuities on $[z, z + 1]$.
- 3 Clearly $D_f = \bigcup_{z \in \mathbb{Z}} D_{zf}$.
- 4 Countable unions of countable sets are countable!

Proof: (3) \Rightarrow (4): continued

- **Proof: continued** Let $F_n(x) = \mu_n((-\infty, x])$, $F(x) = \mu((-\infty, x])$, and let (Ω, \mathcal{F}, P) be a probability triple with the uniform measure over $\Omega = [0, 1]$, and $Y_n(\omega) = \inf\{y : F_n(y) \geq \omega\}$, $Y(\omega) = \inf\{y : F(y) \geq \omega\}$. Then Y_n has CDF $F_n(x)$ and Y has CDF $F(x)$.
 - Now we will show that $\{\omega : Y \text{ is continuous at } \omega\} \subseteq \{\omega : Y_n(\omega) \rightarrow Y(\omega)\}$
 - Roadmap: assume cty of Y , get inequality for F , get inequality for F_n , get inequality for Y_n .

Proof: (3) \Rightarrow (4): continued

- **Proof: continued** Define $Y(\omega) = y$. Then $y - \varepsilon < y < y + \varepsilon$ implies:

$$F(y - \varepsilon) < \omega < F(y + \varepsilon)$$

if $Y(\omega)$ is continuous at ω . Why?

- The weak inequalities are obviously true. if $F(y + \varepsilon) = \omega$, then for any $\delta > 0$, $F(y + \varepsilon) < \omega + \delta$, then $Y(\omega + \delta) \geq y + \varepsilon = Y(\omega) + \varepsilon$. This indicates that there is a jump of at least size ε of $Y(\omega)$ at ω , which is a contradiction to the continuity of Y at ω .
- The other inequality is easier. It's just the contrapositive of $F(q) \geq p \Rightarrow Q(p) \leq q$.

Proof: (3) \Rightarrow (4): continued

- **Proof: continued** In the previous slide, if Y is continuous at ω , and $Y(\omega) = y$, then $F(y - \varepsilon) < \omega < F(y + \varepsilon)$ for all $\varepsilon > 0$.
- Now, in addition to looking at F , we'll look at the F_n s. These converge to F , but only at F 's continuity points.
- For a particular ε , we can always find $0 < \varepsilon' < \varepsilon$, such that $\mu(y - \varepsilon') = \mu(y + \varepsilon') = 0$. ($\mu(x) > 0$ only for at most countably many x). Then $F_n(y - \varepsilon') \rightarrow F(y - \varepsilon')$ and $F_n(y + \varepsilon') \rightarrow F(y + \varepsilon')$. Thus, for large enough n , we have:

$$F_n(y - \varepsilon') < \omega < F_n(y + \varepsilon')$$

- Since $\omega < F_n(y + \varepsilon')$, $\omega \leq F_n(y + \varepsilon')$, then:
$$Y_n(\omega) \leq y + \varepsilon' = Y(\omega) + \varepsilon'.$$
- Since $F_n(y - \varepsilon') < \omega$, then :
$$Y_n(\omega) \geq y - \varepsilon' = Y(\omega) - \varepsilon'.$$

- So $|Y_n(\omega) - Y(\omega)| \leq \varepsilon' < \varepsilon$ for large enough n when Y is cts at ω

Proof: (3) \Rightarrow (4): continued

- ① **Proof: continued** We showed $\{\omega : Y \text{ is continuous at } \omega\} \subseteq \{\omega : Y_n(\omega) \rightarrow Y(\omega)\}.$
- ② By monotonicity $P(Y \text{ is continuous at } \omega) \leq P(Y_n(\omega) \rightarrow Y(\omega)).$
- ③ We need to establish the fact that Y is continuous with probability 1. Or equivalently, D_Y , the set of the discontinuous points of Y , has probability 0.
- ④ We showed discontinuities of $Y(\omega)$, call it D_Y , is at most a countable. So $P(D_Y) = 0.$