

# STAT 7200

## Introduction to Advanced Probability

### Lecture 19

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- Weak Convergence

# Equivalent Definitions of Weakly Convergence

## Theorem 1 (Equivalent Definitions of Weakly Convergence)

*The following statements are all equivalent definition of weak convergence:*

*(1)  $\{\mu_n\}$  converges weakly to  $\mu$ . (Original definition)*

*(2)  $\mu_n(A) \rightarrow \mu(A)$  for all measurable set  $A$  such that  $\mu(\partial A) = 0$ . ( $\partial A$  is defined as the boundary of set  $A$ )*

*(3)  $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$  for all  $x \in \mathbb{R}$  such that  $\mu(\{x\}) = 0$ . That is, the convergence of CDFs. (Note,  $\{x\}$  is the boundary of set  $(-\infty, x]$ .)*

*(4) (Skorohod's Theorem) there are random variable  $Y, Y_1, Y_2, \dots$  defined on the same probability triple, with  $\mathcal{L}(Y) = \mu$  and  $\mathcal{L}(Y_n) = \mu_n$  such that  $Y_n \rightarrow Y$  with probability 1 (This theorem connects the strongest type of convergence: convergence almost surely, with the weak convergence.)*

*(5)  $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$  for all bounded Borel-measurable functions  $f : \mathbb{R} \rightarrow \mathbb{R}$  such that  $\mu(D_f) = 0$ , where  $D_f$  is the set of discontinuous points of  $f$ . (The continuous condition of definition 1) is relaxed.)*

# Structure of Proof

- Our proof will follow the following structure:
- We have proved:  $(5) \Rightarrow (1)$ ,  $(5) \Rightarrow (2)$  and  $(2) \Rightarrow (3)$

## Proof: (1) $\Rightarrow$ (3)

- (1)  $\{\mu_n\}$  converges weakly to  $\mu$ :  $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$  for all bounded continuous functions  $f$ .

(3)  $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$  for all  $x \in \mathbb{R}$  such that  $\mu(\{x\}) = 0$ .

- **Strategy:** We can not apply (1) directly by setting  $f = 1_{(-\infty, x]}$  since  $1_{(-\infty, x]}$ , although bounded, is discontinuous at  $x$ . We may resolve this issue by constructing continuous approximation of  $1_{(-\infty, x]}$ .
- **Proof:** For any  $\varepsilon > 0$  (which is used to control how good the approximation is), define  $f(t) = 1$  for  $t \leq x$  and  $0$  for  $t \geq x + \varepsilon$ , but let  $f(t)$  be a linear function on  $(x, x + \varepsilon)$ .
  - As  $f$  is now continuous and  $1_{(-\infty, x]} \leq f \leq 1_{(-\infty, x + \varepsilon]}$ :

$$\limsup_n \mu_n((-\infty, x]) \leq \limsup_n \int f d\mu_n = \int f d\mu \leq \mu((-\infty, x + \varepsilon])$$

- Let  $\varepsilon \rightarrow 0$ . By the continuity of probability, we have  $\limsup_n \mu_n((-\infty, x]) \leq \mu((-\infty, x])$

## Proof: (1) $\Rightarrow$ (3): continued

- **Proof: continued** Similarly, define  $f(t) = 1$  for  $t \leq x - \varepsilon$  and 0 for  $t \geq x$ , but let  $f(t)$  be a linear function on  $(x - \varepsilon, x)$ . Then  $f$  is linear and  $1_{(-\infty, x - \varepsilon]} \leq f \leq 1_{(-\infty, x]}$ . And:

$$\liminf_n \mu_n((-\infty, x]) \geq \liminf_n \int f d\mu_n = \int f d\mu \geq \mu((-\infty, x - \varepsilon])$$

- Let  $\varepsilon \rightarrow 0$ ,  $\liminf_n \mu_n((-\infty, x]) \geq \mu((-\infty, x)) = \mu((-\infty, x])$ . The last equality holds since  $\mu(\{x\}) = 0$ .
- In summary:

$$\liminf_n \mu_n((-\infty, x]) \geq \mu((-\infty, x]) \geq \limsup_n \mu_n((-\infty, x])$$

- we then must have:

$$\lim_n \mu_n((-\infty, x]) = \mu((-\infty, x])$$

## Proof: (4) $\Rightarrow$ (5)

- (4) there are random variable  $Y, Y_1, Y_2, \dots$  defined on the same probability triple, with  $\mathcal{L}(Y) = \mu$  and  $\mathcal{L}(Y_n) = \mu_n$  such that  $Y_n \rightarrow Y$  with probability 1.

(5)  $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$  for all bounded Borel-measurable functions  $f : \mathbb{R} \rightarrow \mathbb{R}$ . such that  $\mu(D_f) = 0$ , where  $D_f$  is the set of discontinuous points of  $f$ .

- **Proof:** Pick an appropriate  $f$ . First, we want to show that  $P(f(Y_n) \rightarrow f(Y)) = 1$ . Note that
  - ▶  $0 \leq P(Y_n(\omega) \rightarrow Y(\omega), D_f) \leq P(D_f) = 0$
  - ▶  $1 = P(Y_n \rightarrow Y) = P(Y_n \rightarrow Y, D_f) + P(Y_n \rightarrow Y, D_f^c) = P(Y_n \rightarrow Y, D_f^c)$
  - ▶  $\{\omega : f(Y_n) \rightarrow f(Y)\} \supseteq \{\omega : Y_n(\omega) \rightarrow Y(\omega)\} \cap \{\omega : Y(\omega) \in D_f^c\}$

so  $f(Y_n) \rightarrow f(Y)$  wp1 by (4) and monotonicity of  $P$ .

- Because  $f$  is bounded,  $f(Y)$  is integrable, so  $E[f(Y_n)] \rightarrow E[f(Y)]$  by the dominated convergence theorem.

## Proof: (3) $\Rightarrow$ (4)

- (3)  $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$  for all  $x \in \mathbb{R}$  such that  $\mu(\{x\}) = 0$ .  
(4) there are random variables  $Y, Y_1, Y_2, \dots$  defined on the same probability triple, with  $\mathcal{L}(Y) = \mu$  and  $\mathcal{L}(Y_n) = \mu_n$  such that  $Y_n \rightarrow Y$  with probability 1
- **Strategy:** We will construct random variables with CDFs  $F_n(x) = \mu_n((-\infty, x])$ ,  $F(x) = \mu((-\infty, x])$ , then we will show the convergence of these random variables using the fact that the corresponding CDFs converge.



## Proof: (3) $\Rightarrow$ (4): Probability Integral Transform

- **Proof:** We can construct random variable with given CDF using probability integral transform theorem.
- This theorem states that, for random variable  $U$  that follows uniform distribution, given any CDF  $F(x)$ , define quantile function  $Q(p) = \inf\{x : F(x) \geq p\}$ , then the random variable  $Q(U)$  follows distribution with CDF  $F(x)$ .
- By definition  $F(q) \geq p \Rightarrow Q(p) \leq q$

$$P[Q(U) \leq x] = P[F(x) \geq U] = F(x)$$

- Other useful results include:
  - a)  $F(q) < p \Rightarrow Q(p) \geq q$
  - b) When the CDF is continuous and strictly increasing, the quantile function is the inverse of CDF.
  - c) The quantile function  $Q(p)$  is a non-decreasing function, same as the CDF.

## Proof: (3) $\Rightarrow$ (4): continued

- **Proof: continued** Let  $F_n(x) = \mu_n((-\infty, x])$ ,  $F(x) = \mu((-\infty, x])$ , and let  $(\Omega, \mathcal{F}, P)$  be a probability triple with the uniform measure over  $\Omega = [0, 1]$ , and  $Y_n(\omega) = \inf\{y : F_n(y) \geq \omega\}$ ,  $Y(\omega) = \inf\{y : F(y) \geq \omega\}$ . Then  $Y_n$  has CDF  $F_n(x)$  and  $Y$  has CDF  $F(x)$ .
  - Now we will show that  $Y_n(\omega) \rightarrow Y(\omega)$  if  $Y(\omega)$  is continuous at  $\omega$ .
  - Firstly, define  $Y(\omega) = y$ . Then  $y - \varepsilon < Y(\omega) < y + \varepsilon$  implies:

$$F(y - \varepsilon) < \omega < F(y + \varepsilon)$$

if  $Y(\omega)$  is continuous at  $\omega$ .

- The reason: if  $F(y + \varepsilon) = \omega$ , for any  $\delta > 0$ ,  $F(y + \varepsilon) < \omega + \delta$ , then  $Y(\omega + \delta) \geq y + \varepsilon = Y(\omega) + \varepsilon$ . This indicates that there is a jump of at least size  $\varepsilon$  of  $Y(\omega)$  at  $\omega$ , which is a contradiction to the continuity of  $Y$  at  $\omega$ .
- Thus,  $F(y - \varepsilon) < \omega < F(y + \varepsilon)$

## Proof: (3) $\Rightarrow$ (4): continued

- **Proof: continued** In the previous slide, if  $Y$  is continuous at  $\omega$ , and  $Y(\omega) = y$ , then  $F(y - \varepsilon) < \omega < F(y + \varepsilon)$  for all  $\varepsilon > 0$ .
- Now for a particular  $\varepsilon$ , we can always find  $0 < \varepsilon' < \varepsilon$ , such that  $\mu(y - \varepsilon') = \mu(y + \varepsilon') = 0$ . ( $P(x) > 0$  only for at most countably many  $\{x\}$ ). Then  $F_n(y - \varepsilon') \rightarrow F(y - \varepsilon')$  and  $F_n(y + \varepsilon') \rightarrow F(y + \varepsilon')$ . Thus, for large enough  $n$ , we have:

$$F_n(y - \varepsilon') < \omega < F_n(y + \varepsilon')$$

- Since  $\omega < F_n(y + \varepsilon')$ ,  $\omega \leq F_n(y + \varepsilon')$ , then:

$$Y_n(\omega) \leq y + \varepsilon' = Y(\omega) + \varepsilon'.$$

- Since  $F_n(y - \varepsilon') < \omega$ , then  $y - \varepsilon' < Y_n(\omega)$ , which implies a weaker inequality:

$$Y_n(\omega) \geq y - \varepsilon' = Y(\omega) - \varepsilon'.$$

- In summary,  $|Y_n(\omega) - Y(\omega)| \leq \varepsilon' < \varepsilon$  for large enough  $n$ . Thus,  $Y_n(\omega) \rightarrow Y(\omega)$  when  $Y$  is continuous at  $\omega$

## Proof: (3) $\Rightarrow$ (4): continued

- **Proof: continued** Finally, we need to establish the fact that  $Y$  is continuous with probability 1. Or equivalently,  $D_Y$ , the set of the discontinuous points of  $Y$ , has probability 0. This statement can be justified based on : a)  $Y$  is defined on a uniform measure over  $[0, 1]$ , b)  $D_Y$  is at most countable.
- To show  $D_f$  is at most countable, let us first create a partition of  $\mathbb{R} = \bigcup_{z \in \mathbb{Z}} (z, z + 1]$ , and define  $\Omega_z = \{\omega : z < Y(\omega) \leq z + 1\} = Y^{-1}((z, z + 1])$ . Since  $Y$  is a non-decreasing function, each  $\Omega_z$  should be an interval on  $[0, 1]$  and  $\{\Omega_z\}$  forms a partition of  $[0, 1]$ .
- Then Let  $D_f^z =$  Discontinuous points in  $\Omega_z$ . Clearly  $D_f = \bigcup_z D_f^z$ .
- Next define  $D_f^z \supseteq D_f^z(m) = \{\text{jumps that have size} \geq m^{-1}\}$ .
- Clearly  $|D_f^z| \leq m$  and  $D_f = \bigcup_z \bigcup_m D_f^z(m)$ . Therefore the number of discontinuous points is at most countable.