

STAT 7200

Introduction to Advanced Probability

Lecture 11

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- Probability Inequalities
- Almost Sure Convergence

“A First Look at Rigorous Probability Theory” (Jeffrey Rosenthal)
Sections 5.1 and 5.2

Markov and Chebychev's Inequalities

Theorem 1 (Markov's Inequality)

For non-negative random variable X , if $\alpha > 0$, then $P(X \geq \alpha) \leq E(X)/\alpha$.

- **Proof** Let $A = \{\omega : X(\omega) \geq \alpha\}$, then we have $X \geq \alpha 1_A$. By the order-preserving property of expectation, Markov's inequality follows immediately.
- Despite its simplicity, Markov's inequality can be quite useful in practice. For example, applying it to $f(X) = (X - E(X))^2$ gives us Chebychev's inequality:

Corollary 2 (Chebychev's Inequalities)

For any random variable Y with finite variance, for $\alpha \geq 0$, we have $P(|Y - \mu_Y| \geq \alpha) \leq \text{Var}(Y)/\alpha^2$.

The Cauchy-Schwarz Inequality

Theorem 3 (Cauchy-Schwarz Inequality)

For random variables X, Y such that $E(X^2) < \infty, E(Y^2) < \infty$, we have $E(|XY|) \leq \sqrt{E(X^2)E(Y^2)}$.

• Proof

$$\begin{aligned} 0 &\leq E \left[\left(\frac{|X|}{\sqrt{E(X^2)}} - \frac{|Y|}{\sqrt{E(Y^2)}} \right)^2 \right] \\ &= 1 + 1 - 2 \frac{E(|YX|)}{\sqrt{E(X^2)}\sqrt{E(Y^2)}} \\ &= 2 \left[1 - \frac{E(|YX|)}{\sqrt{E(X^2)E(Y^2)}} \right] \end{aligned}$$

The Cauchy-Schwarz Inequality

Theorem 4 (Cauchy-Schwarz Inequality)

For random variables X, Y such that $E(X^2) < \infty, E(Y^2) < \infty$, we have $E(|XY|) \leq \sqrt{E(X^2)E(Y^2)}$.

- Special case 1: correlation must always be between -1 and 1 . For r.v.s X and Y with finite variances, $\text{Cov}(X, Y) = E((X - EX)(Y - EY))$. Then $|\text{Cov}(X, Y)| \leq E(|X - EX||Y - EY|) \leq \sqrt{\text{Var}(X)\text{Var}(Y)}$, and $|\text{Corr}(X, Y)| \leq 1$.
- Special case 2: the Cramér-Rao inequality. For estimator $\hat{\theta}$ and score function U . $E[U] = 0$ and $\text{Var}[U] = I(\theta)$.
- Note: some resources define CS inequality as $|E(XY)| \leq \sqrt{E(X^2)E(Y^2)}$, but this isn't as tight, and we can prove it by using this CS inequality and Jensen's together.

Jensen's Inequality

Theorem 5 (Jensen's Inequality)

For random variable X with finite mean, and a convex function $\phi : \mathbb{R} \rightarrow \mathbb{R}$, we have $E(\phi(X)) \geq \phi(E(X))$.

- **Convex Function:** $\phi(x)$ is convex if for any $x, y \in \mathbb{R}$ and $0 < t < 1$, $t\phi(x) + (1 - t)\phi(y) \geq \phi(tx + (1 - t)y)$. For instance, x^2 , $|x|$, e^x are all examples of convex functions.
- One key property of convex function is that given any point $(x_0, \phi(x_0))$, you may find a straight line $g(x) = a + bx$ that is below $\phi(x)$ and also passes through $(x_0, \phi(x_0))$ (usually it is the tangent at x_0 if the derivative of $\phi(x)$ exists at $x = x_0$). That is, $g(x) \leq \phi(x)$ and $g(x_0) = \phi(x_0) = a + bx_0$.
- **Proof:** Apply the above property for $x_0 = E(X)$, then we have $E(\phi(X)) \geq E(g(X)) = a + bE(X) = g(E(X)) = \phi(E(X))$.

Almost Sure Convergence

- **Pointwise Convergence** Suppose we have random variables Z, Z_1, Z_2, \dots on probability triple (Ω, \mathcal{F}, P) such that for each $\omega \in \Omega$, $\lim_{n \rightarrow \infty} Z_n(\omega) = Z(\omega)$. Then we may say that $\{Z_n\}$ converges to Z pointwise,
- This type of convergence is usually unnecessarily strong in probability theory. So a slightly weaker version is more popular and useful: we would only require $Z_n(\omega)$ converges to $Z(\omega)$ with probability one. That is, $P(\{\omega \in \Omega : \lim_{n \rightarrow \infty} Z_n(\omega) = Z(\omega)\}) = 1$.
- **Almost Sure Convergence** We say that $\{Z_n\}$ converges to Z almost surely (or a.s., or with probability 1), if the above condition holds. And we usually denote it as $P(Z_n \rightarrow Z) = 1$.
- **Example** Consider the uniform measure (Ω, \mathcal{F}, P) on $[0, 1]$. Define $Z_n(\omega) = 1_{[0, \frac{1}{2^n}]}(\omega)$, then for each $\omega > 0$, we have $\lim_n Z_n(\omega) = 0$ and $\lim_n Z_n(0) = 1$. So $P(\lim_{n \rightarrow \infty} Z_n(\omega) = 0) = P((0, 1]) = 1$. That is, $\{Z_n\}$ converges to $Z = 0$ almost surely (but not pointwise).

Almost Sure Convergence: A lemma

Lemma 6 (5.2.1)

For r.v.s Z, Z_1, Z_2, \dots such that for each $\varepsilon > 0$, $P(|Z_n - Z| \geq \varepsilon \text{ i.o.}) = 0$. Then $P(Z_n \rightarrow Z) = 1$. The converse is also true.

- **Proof:** Pick $\omega \in \Omega$. $Z_n(\omega) \rightarrow Z(\omega)$ iff $\forall \epsilon > 0$, $\exists N$ such that $n \geq N$ implies $|Z_n(\omega) - Z(\omega)| < \epsilon$.
- $Z_n(\omega) \rightarrow Z(\omega)$ iff $\omega \in \bigcup_{N=1}^{\infty} \bigcap_{n \geq N} \{|Z_n(\omega) - Z(\omega)| < \epsilon\}$
- $Z_n(\omega) \rightarrow Z(\omega)$ iff $\omega \in \liminf_{n \rightarrow \infty} \{|Z_n(\omega) - Z(\omega)| < \epsilon\} := \{|Z_n(\omega) - Z(\omega)| < \epsilon \text{ a.a.}\}$
- $P(Z_n \rightarrow Z) = P(\{\omega \in \Omega : |Z_n(\omega) - Z(\omega)| < \epsilon \text{ a.a.}\}) = 1 - P(|Z_n - Z| \geq \epsilon \text{ i.o.})$

Almost Sure Convergence and Borel-Cantelli Lemma

- According to the Borel-Cantelli Lemma: for a sequence of event A_1, A_2, \dots , $\sum_n P(A_n) < \infty$ implies $P(A_n \text{ i.o.}) = 0$. Combining this with the lemma we obtained in previous slide:

Lemma 7

For r.v.s Z, Z_1, Z_2, \dots such that for any $\varepsilon > 0$, $\sum_n P(|Z_n - Z| \geq \varepsilon) < \infty$. Then $P(Z_n \rightarrow Z) = 1$.

- **Example** Let Z_1, Z_2, \dots be random variables such that $P(Z_n = 1) = \frac{1}{2^n}$ and $P(Z_n = 0) = 1 - \frac{1}{2^n}$. Then for $1 > \varepsilon > 0$, $P(|Z_n| \geq \varepsilon) = \frac{1}{2^n}$ and $\sum_n P(|Z_n| \geq \varepsilon) = 1 < \infty$, so we must have $Z_n \rightarrow 0$ almost surely.
- **Note** The converse of this lemma is not necessarily true. For instance, consider Z_1, Z_2, \dots defined on $([0, 1], \mathcal{M}, \lambda)$, $Z_n(\omega) = 1_{[0, \frac{1}{n}]}(\omega)$.