14.4: Bayesian Methods

Taylor

University of Virginia

Introduction

This is the main difference between frequentist statistics and Bayesian statistics:

Frequentist statistics assumes that parameter values are fixed.

Bayesian statistics assumes that parameter values are random.

To emphasize this difference in notation, instead of writing $f(x_1, \ldots, x_n; \theta)$, we will write

$$f(x_1,\ldots,x_n|\theta)$$

to emphasize that we are conditioning on a random variable θ .

↓□▶ ↓□▶ ↓ □▶ ↓ □▶ ↓ □ ♥ ♀ ○

Taylor (UVA) "14.4" 2 / :

Introduction

You may have seen Bayes' rule in terms of probabilities of simple events A and B as

$$P(B|A) = \frac{P(A|B)P(B)}{P(A)} = \frac{P(A|B)P(B)}{P(A|B)P(B) + P(A|B^c)P(B^c)}$$

Now we'll use Bayes' theorem for densities

$$h(\theta|x_1,\ldots,x_n) = \frac{f(x_1,\ldots,x_n|\theta)g(\theta)}{\int f(x_1,\ldots,x_n|\theta)g(\theta)d\theta}$$

Taylor (UVA) "14.4" 3 / 1-

Introduction

Here's our main formula again:

$$h(\theta|x_1,\ldots,x_n) = \frac{f(x_1,\ldots,x_n|\theta)g(\theta)}{\int f(x_1,\ldots,x_n|\theta)g(\theta)d\theta}$$

 $f(x_1, \ldots, x_n | \theta)$ is still called the likelihood. But $g(\theta)$ is called the **prior distribution**. The prior distribution is chosen to reflect prior knowledge we have about the parameter before we see our data.

After we see our data, we can compute the **posterior distribution** $h(\theta|x_1,\ldots,x_n)$. This is the main thing we're after here. This is the distribution of our unknown quantity after we take into account all possible information.

Taylor (UVA) "14.4" 4 / 14

Last thing before an example...

In the formula

$$h(\theta|x_1,\ldots,x_n) = \frac{f(x_1,\ldots,x_n|\theta)g(\theta)}{\int f(x_1,\ldots,x_n|\theta)g(\theta)d\theta}$$

The denominator $\int f(x_1,\ldots,x_n|\theta)g(\theta)d\theta$ is just a normalizing constant to make our density integrate to 1 (make sure you see why...it isn't a function in θ). Sometimes we don't care about it. Because of this, people doing Bayesian statistics will use "proportional to" symbol a lot like this:

$$h(\theta|\mathbf{x}) \propto f(\mathbf{x}|\theta)g(\theta)$$

This means that the posterior distribution is proportional to the pointwise product (in θ) between f and g.

Taylor (UVA) "14.4" 5 / 1

Example 14.7

Suppose we have data from a binomial distribution. The only parameter that we are uncertain about is p. Let's let p be a random variable now.

p needs to be in the interval (0,1). We might use the Beta distribution as a prior for p because its support is the interval (0,1). The parameters to this distribution will be α and β .

$$g(p; \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha - 1} (1 - p)^{\beta - 1}$$

Taylor (UVA) "14.4" 6 / 14

Remember how we said we can use the ∞ symbol? If we take the product $f(\mathbf{x}|\theta)g(\theta)$ and we can "recognize" the density, then we already know the normalizing constant. This happens in our first example (but not always). And instead of writing θ , we'll write our parameter as p.

Let $X|p \sim \text{Binomial}(n, p)$ and $p \sim \text{Beta}(\alpha, \beta)$.

$$h(\rho|\mathbf{x}) \propto f(\mathbf{x}|\rho)g(\rho)$$

$$= \left[\binom{n}{x} p^{x} (1-\rho)^{n-x} \right] \left[\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha-1} (1-\rho)^{\beta-1} \right]$$

$$\propto p^{x} (1-\rho)^{n-x} p^{\alpha-1} (1-\rho)^{\beta-1}$$

$$= p^{\alpha+x-1} (1-\rho)^{n+\beta-x-1}$$

And this looks like a beta distribution, so

$$h(p|\mathbf{x}) = \text{Beta}(\alpha + x, n + \beta - x)$$

Taylor (UVA) "14.4" 7 / 14

$$h(p|\mathbf{x}) = \mathsf{Beta}(\alpha + x, n + \beta - x)$$

Using quick formulas for a Beta distribution

$$E[p|x_1,...,x_n] = \frac{\alpha + x}{(\alpha + x) + (n + \beta - x)}$$

$$= \frac{\alpha + x}{\alpha + n + \beta}$$

$$= \frac{\alpha}{\alpha + n + \beta} \frac{\alpha + \beta}{\alpha + \beta} + \frac{x}{\alpha + n + \beta} \frac{n}{n}$$

$$= \frac{\alpha}{\alpha + \beta} w_1 + \frac{x}{n} w_2$$

Where $w_1 = \frac{\alpha + \beta}{\alpha + n + \beta}$ and $w_2 = \frac{n}{\alpha + n + \beta}$. So the posterior average is a weighted average of the prior and likelihood averages.

Taylor (UVA) "14.4" 8 / 14

Earlier when we were doing frequentist statistics, we did point estimation, confidence intervals, and hypothesis testing. We only used the likelihood function.

Now we have a probability distribution for our parameter. A few things worth mentioning:

- we can calculate the mode of this distribution (kind of like MLE)
- 2 we can calculate the mean of this distribution (kind of like MLE too)
- answer questions like "What's the probability our parameter was less than a half"
- we don't have to be careful with the distinction between probability and likelihood
- we can predict new data like this $p(x_{new}|x_{old}) = \int p(x_{new}|\theta)p(\theta|x_{old})d\theta$
- we took into account a subjective prior distribution to reflect what we already know about p

Find a Bayesian *credible interval* for *p*.

$$h(p|\mathbf{x}) = \mathsf{Beta}(\alpha + x, n + \beta - x)$$

Answer...just take appropriate quantiles from the posterior distribution.

Taylor (UVA) "14.4" 10 / 14

Example 14.8

$$X_1,\ldots,X_n \stackrel{iid}{\sim} \mathcal{N}(\mu,\sigma_1^2)$$
 and $\mu \sim \mathcal{N}(\mu_0,\sigma_0^2)$. Check that we have $f(x_1,\ldots,x_n|\mu) = (2\pi\sigma_1^2)^{-n/2} \exp\left[-\frac{\sum (x_1-\mu)^2}{2\sigma_1^2}\right]$

$$h(\mu|x_1, \dots, x_n) \propto f(x_1, \dots, x_n|\mu)g(\mu)$$

$$= (2\pi\sigma_1^2)^{-n/2} \exp\left[-\frac{\sum (x_1 - \mu)^2}{2\sigma_1^2}\right] \times$$

$$(2\pi\sigma_0^2)^{-1/2} \exp\left[-\frac{(\mu - \mu_0)^2}{2\sigma_0^2}\right]$$

$$\propto \exp\left[-\frac{1}{2}\left\{\frac{\sum (x_1 - \mu)^2}{\sigma_1^2} + \frac{(\mu - \mu_0)^2}{\sigma_0^2}\right\}\right]$$

$$= \exp\left[-\frac{1}{2}\left\{\frac{\sum x_1^2 - 2\mu \sum x_i + n\mu^2}{\sigma_1^2} + \frac{(\mu - \mu_0)^2}{\sigma_0^2}\right\}\right]$$

↓□▶ ↓□▶ ↓ □▶ ↓ □▶ ↓ □ ♥ ♀ ○

Taylor (UVA) "14.4" 11 / 14

$$\dots = \exp\left[-\frac{1}{2} \left\{ \frac{\sum x_1^2 - 2\mu \sum x_i + n\mu^2}{\sigma_1^2} + \frac{(\mu - \mu_0)^2}{\sigma_0^2} \right\} \right]$$

$$= \exp\left[-\frac{1}{2} \left\{ \frac{\sum x_1^2 - 2\mu \sum x_i + n\mu^2}{\sigma_1^2} + \frac{\mu^2 - 2\mu\mu_0 + \mu_0^2}{\sigma_0^2} \right\} \right]$$

$$\propto \exp\left[-\frac{1}{2} \left\{ \frac{-2\mu \sum x_i + n\mu^2}{\sigma_1^2} + \frac{\mu^2 - 2\mu\mu_0}{\sigma_0^2} \right\} \right]$$

$$= \exp\left[-\frac{1}{2} \left\{ \mu^2 \left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right) - 2\mu \left(\frac{\sum x_i}{\sigma_1^2} + \frac{\mu_0}{\sigma_0^2}\right) \right\} \right]$$

$$= \exp\left[-\frac{1}{2} \left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right) \left\{ \mu^2 - 2\mu \left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right)^{-1} \left(\frac{\sum x_i}{\sigma_1^2} + \frac{\mu_0}{\sigma_0^2}\right) \right\} \right]$$

$$\propto \exp\left[-\frac{1}{2} \left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right) \left\{ \mu - \left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right)^{-1} \left(\frac{\sum x_i}{\sigma_1^2} + \frac{\mu_0}{\sigma_0^2}\right) \right\}^2 \right]$$

Taylor (UVA) "14.4" 12 / 14

So $h(\mu|x_1,\ldots,x_n)$ is *still* normally distributed with mean

$$\left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right)^{-1} \left(\frac{\sum x_i}{\sigma_1^2} + \frac{\mu_0}{\sigma_0^2}\right)$$

and variance

$$\left(\frac{n}{\sigma_1^2} + \frac{1}{\sigma_0^2}\right)^{-1}$$

Taylor (UVA) "14.4" 13 / 14

Summary of Examples

In the examples we did, the posterior distribution is from the sample family as the prior distribution (but it usually has different parameters). This is a very special case; we say here that the prior is **conjugate** to the data distribution. Said differently, the prior is a **conjugate** prior for the likelihood.

Taylor (UVA) "14.4" 14 / 14